

1 **Seed size and its rate of evolution correlate with species diversification across**
2 **angiosperms**

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4 Short title: Seed size and angiosperm diversification

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23

24 **Abstract**

25

26 Species diversity varies greatly across the different taxonomic groups that comprise
27 the Tree of Life (ToL). This imbalance is particularly conspicuous within
28 angiosperms, but is largely unexplained. Seed mass is one trait that may help clarify
29 why some lineages diversify more than others because it confers adaptation to
30 different environments, which can subsequently influence speciation and extinction.
31 The rate at which seed mass changes across the phylogeny may also be linked to
32 diversification by increasing reproductive isolation and allowing access to novel
33 ecological niches. However, the magnitude and direction of the association between
34 seed mass and diversification has not been assessed across the angiosperm phylogeny.
35 Here, we show that absolute seed size and the rate of change in seed size are both
36 associated with variation in diversification rates. Based on the largest available
37 angiosperm phylogenetic tree, we found that smaller-seeded plants had higher rates of
38 diversification, possibly due to improved colonisation potential. The rate of
39 phenotypic change in seed size was also strongly positively correlated with speciation
40 rates, providing rare, large-scale evidence that rapid morphological change is
41 associated with species divergence. Our study now reveals that variation in
42 morphological traits and, importantly, the rate at which they evolve can contribute to
43 explaining the extremely uneven distribution of diversity across the ToL.

44

45

46 **Introduction**

47 Angiosperms are one of the most species-rich clades on Earth and have dominated
48 terrestrial plant communities since the Late Cretaceous [1]. The astounding diversity

49 of flowering plants is distributed extremely unevenly across the ToL. Each of the five
50 most species-rich angiosperm families contains >10,000 species while more than 200
51 families contain <100 species each [2]. An enduring pursuit in evolutionary biology is
52 to explain this uneven distribution of biodiversity, not only in angiosperms, but also
53 across the whole ToL [3].

54

55 Biological traits offer one way to explain disparity in species diversification if they
56 confer adaptation to different environments. Seed mass is one such trait that is
57 particularly important for angiosperms because it integrates across many
58 characteristics of an individual's life history strategy [4]. Along with adult plant size,
59 seed mass affects survival, reproductive lifespan and dispersal [5]. These life history
60 characteristics contribute to fitness and adaptation, which are the ultimate
61 determinants of whether lineages diversify or go extinct [6]. In support of this idea,
62 seed mass has been shown to correlate negatively with diversification in the
63 Polygoneaceae [7], but this has not been investigated across large taxonomic scales.
64 As seed mass varies over ten orders of magnitude in angiosperms, from the minute 1
65 μg seeds of some orchids to the >18 kg seeds of the sea coconut (*Lodoicea*
66 *maldivica*), this huge variation may coincide with variation in species diversity.
67 Generalising the direction and magnitude of a link between seed mass and
68 diversification across taxonomic scales has, however, proved difficult. Some life
69 history characteristics encapsulated by seed mass are expected to promote speciation
70 or extinction, while others may simultaneously counteract such effects [8].

71

72 The rate of change in key biological traits, such as seed size, can be as important in
73 driving macroevolutionary dynamics as the absolute values of the traits themselves

74 [9]. This is because phenotypic divergence may cause reproductive isolation that
75 results in speciation [10]. Nevertheless, few empirical studies have detected a
76 correlation between rates of phenotypic evolution and lineage diversification
77 [9,11,12]. A correlation between the two may be expected where a trait can change
78 more rapidly in some species than others in response to selective pressures (i.e. high
79 “evolvability” [13]). This rapid change may enable greater access to new ecological
80 niches or quicker establishment of reproductive isolation, thereby increasing the rate
81 of speciation (λ) [14]. In the case of seed mass, the ability to switch rapidly from
82 small seeds with high dispersal ability to larger seeds with lower dispersal ability
83 might promote cycles of rapid colonisation and isolation or permit adaptation to new
84 dispersal vectors in novel environments. Rapid evolution of new phenotypes may also
85 allow individuals to escape harsh environmental conditions and competitive
86 interactions [15], thereby decreasing extinction rates (μ). The net outcome of these
87 processes on net diversification ($r = \lambda - \mu$) will ultimately depend upon which of these
88 rates responds more strongly to phenotypic change.

89

90 Here, we show that both seed mass and its phenotypic rate of evolution correlate with
91 speciation and extinction across the angiosperm ToL. Our approach combined the
92 most comprehensive phylogenetic timetree available [16] with an unparalleled dataset
93 of seed mass measurements from over 30,000 angiosperm species [17]. We estimated
94 rates of speciation, extinction and seed size evolution across the phylogeny using
95 Bayesian, method-of-moments, and maximum likelihood analyses – each with
96 different assumptions regarding rate variation through time. We then tested whether
97 there was any links between rates of diversification and both seed size and its rate of
98 evolution. Additionally, we examined whether this link was consistent across different

99 methodologies, timescales and taxonomic categories.

100

101 **Results**

102 Our results point to a strong association between angiosperm diversification and rates
103 of seed size evolution irrespective of analytical method or taxonomic scale, with
104 weaker evidence for a link between macroevolutionary dynamics and absolute seed
105 size. In the first instance, we calculated rates of speciation (λ), extinction (μ) and seed
106 size evolution using Bayesian Analysis of Macroevolutionary Mixtures (BAMM)
107 [18]. BAMM models rate heterogeneity through time and lineages, and accounts for
108 incomplete taxon sampling. We used a genus-level phylogenetic tree for this
109 particular analysis to maximise overlap between the phylogenetic and the seed mass
110 datasets and make computation feasible. The resulting tree contained 4105 genera. As
111 expected, given the high degree of taxonomic imbalance observed in the angiosperm
112 phylogeny, we found strong support for more than 100 shifts in the rates of
113 diversification (Fig. S1). There was also marked heterogeneity in the rates of seed size
114 evolution (Fig. 1), which varied over two orders of magnitude (Fig. S2).

115

116 We then estimated whether shifts in macroevolutionary dynamics (λ , μ and r)
117 estimated with BAMM were significantly correlated with absolute mean genus seed
118 size and rates of seed size evolution by comparing the empirical correlations to a null
119 distribution generated using STructured Rate Permutations on Phylogenies
120 (STRAPP), which is robust to phylogenetic pseudoreplication (see Methods for
121 details) [19]. We were able to link major differences in diversity across angiosperm
122 clades with both the rate of phenotypic evolution and the absolute value of trait itself.
123 Specifically, increased speciation was associated with a faster rate of seed size

124 evolution (Spearman's $\rho = 0.44$, p-value < 0.0001; Fig. 2a). Increased extinction rates
125 were similarly associated with higher evolvability ($\rho = 0.35$, p-value = 0.003), but
126 given the weaker effect, the net outcome of $\lambda - \mu$ was that diversification rates were
127 positively correlated with phenotypic change ($\rho = 0.42$, p-value < 0.0001; Fig. 2b).
128 We also identified an association between seed size and both speciation and extinction
129 ($\rho = -0.22$, p-value = 0.016 and $\rho = -0.20$, p-value = 0.045, respectively; Fig. 2b).
130 Again, as the effect on speciation was greater than that on extinction, diversification
131 rates increased with decreasing seed size ($\rho = -0.19$, p-value = 0.049; Fig. 2b).
132 Generally, the observed correlations arose from many phenotypically fast-evolving
133 clades distributed across the phylogeny (Fig. S2) and were robust to prior choice in
134 the BAMM analyses (Fig. S1).

135

136 We also obtained estimates of net diversification rates with other methods to account
137 for intra-generic variation in seed size and test the robustness of our findings. The
138 first method used a moment estimator that assumed diversification rates were constant
139 over time within genera that had a known crown age and number of species [20].
140 Monophyletic genera were then selected from a phylogenetic tree containing 8448
141 species present in our initial seed mass dataset so that they each contained ≥ 4 species
142 with seed size data and more than 30% of their described species, thereby ensuring
143 that the minimum probability of recovering the correct crown age node of any genus
144 was 70% (see Methods). A rate of seed size evolution was then estimated for the 70
145 resulting genera by fitting a Brownian Motion (BM) model of trait evolution.
146 Mirroring the BAMM analysis, diversification rates were strongly correlated with the
147 rate of seed size evolution (PGLS slope = 0.157, p-value = 0.004; df = 69), and
148 weakly negatively correlated with mean genus seed size (PGLS slope = -0.064, p-

149 value = 0.039; df = 69). Our results were similar when we did not limit the analysis
150 based on available seed size data (PGLS slope = -0.103, p-value = 0.004, df = 127; no
151 analysis was performed for trait evolution with these criteria because of considerable
152 missing data).

153

154 The second method that we used to account for intra-generic variation involved
155 estimating macroevolutionary dynamics and phenotypic evolution across sets of
156 clades in our species-level tree that were defined at different timescales. A series of
157 ten, 2 million year-wide time slices from the present up to 20 million years ago was
158 used to identify the most inclusive monophyletic clades of ≥ 4 species. The minimum
159 intra-clade sampling fraction was set at 30% as above. For each resulting clade in
160 each time interval, diversification rates and seed size evolution were estimated as in
161 our previous non-BAMM analyses. We also fitted a series of time-dependent
162 diversification models to each clade with RPANDA, which uses a maximum
163 likelihood approach to estimate speciation and extinction and allows for incomplete
164 taxon sampling [21]. By measuring the correlation of macroevolutionary rates and
165 morphological rates across sets of similarly aged clades we effectively eliminated any
166 spurious correlations caused by both rates being positively correlated with time.
167 Again, we found a positive correlation between the rate of seed size evolution and
168 speciation rates that was consistent across time slices (Fig. 3a, Fig. S3a). As expected
169 given the weak association between seed size and speciation found in our other
170 analyses, a significant negative correlation was only recovered in some of the
171 analysed time slices (Fig 3b, Fig. S3b).

172

173 **Discussion**

174 Our study supports the ideas that variation in seed mass and, particularly, its rate of
175 evolution can help explain disparity in diversification across the angiosperm
176 phylogeny by playing a central role in plant life history. As we show with our clade-
177 based analysis, our results are repeatable across many methodologies, varying
178 timescales and do not ascribe to a particular taxonomic category (Figs 2, 3). We also
179 showed that the effect of intra-generic variation in seed mass in the BAMM analysis
180 was negligible (S1 Appendix).

181

182 The robust association of high rates of phenotypic change with lineage diversification
183 has recently been observed in other taxonomic groups [9,12], but never across the
184 whole of the angiosperm ToL as we find here. Accelerated morphological evolution
185 may allow radiating lineages to occupy more complex adaptive landscapes [22].
186 Similarly, species with greater rate of change in their seed mass (i.e., higher
187 evolvability) could shift between adaptive peaks or develop reproductive barriers
188 more rapidly. Alternatively, the theory of punctuated equilibria [23], whereby
189 morphological changes can arise from the speciation process, might also explain the
190 connection of phenotypic evolution with species divergence. However, current
191 methods do not allow us to distinguish whether speciation is responding to
192 morphological change or vice versa when reconstructing 250 million years of
193 evolutionary history [9].

194

195 Seed mass itself will also covary with dispersal ability and environmental tolerance in
196 ways that can change speciation and extinction. For example, we found that smaller-
197 seeded genera had faster speciation rates. This may be because smaller-seeded genera
198 generally disperse over larger distances [24], which can promote speciation by

199 creating isolated populations [25]. However, the relationship between dispersal and
200 speciation is highly context dependent. The permeability of a landscape to dispersal
201 determines the values of dispersal that may promote species divergence [26]. The
202 weak correlation that we observed between seed size and diversification might
203 therefore reflect contradicting patterns of different strength operating at different
204 clades in the angiosperm phylogeny. Dispersal syndromes may also modify the effect
205 of seed size on speciation. For instance, species with larger seeds are generally
206 associated with biotic dispersal that distributes seeds over greater distances than wind
207 or gravity dispersal [5]. However, broad-scale predictions for the effects of dispersal
208 syndromes on diversification may be inaccurate, since the former depend on
209 landscape connectivity [8] and can sometimes be inconsistent, e.g. a wind-dispersed
210 seed might be transported by an animal. We also found that smaller-seeded genera had
211 higher rates of extinction, possibly due to smaller nutritional reserves that constrain
212 establishment, environmental tolerance and access to limiting resources [27]. Detailed
213 contextual data will be necessary to expand upon the mechanisms underlying our
214 findings in specific regions and clades.

215

216 Although seed mass is associated with other traits that can affect diversification, there
217 is little evidence that these better explain our observed correlations or that seed size is
218 a mere proxy for one of these other traits. For example, genome size positively
219 correlates with seed mass [28], and faster rates of genome size evolution have been
220 linked to increased speciation in angiosperms [29]. Shorter, smaller-seeded plants also
221 tend to have faster life cycles, which may accelerate mutation rates [30] and promote
222 diversification [31]. By comparing the effects of genome size and life cycle across a
223 subset of >900 genera in our dataset, we found that only the distinction between

224 strictly annual versus perennial genera influenced macroevolutionary dynamics in a
225 similar way as seed size (S1 Appendix, Fig. S4). But unlike other traits [29], both
226 absolute seed size as well as its rate of change were correlated with speciation and
227 extinction. Thus, although other traits surely influence diversification [32], we argue
228 that our results generally reflect the role of seed size as a trait that integrates across
229 multiple aspects of life history characteristics in ways that can predictably influence
230 plant macroevolutionary dynamics (Fig. S5).

231

232 The approach applied here can help to unravel the processes responsible for
233 generating large-scale asymmetries in biodiversity. It also offers the potential to test
234 how widely-varying traits and their rate of morphological evolution influence other
235 aspects of the evolution and adaptation of flowering plants (e.g. [16]). Clade-specific
236 exceptions arising from local interactions with non-focal traits [33] and specific
237 spatio-temporal contexts will undoubtedly interact with broad-scale
238 macroevolutionary patterns and may modulate the effects of seed mass on
239 diversification. Regardless, our results show that seed size and its rate of evolution
240 correlate with speciation and extinction across the flowering plants. This finding may
241 help to explain why some clades are much more species-rich than others and points to
242 a key role of rapid morphological evolution in generating greater levels of diversity.

243

244 **Materials and methods**

245 *Seed mass and phylogenetic dataset*

246 Seed mass data for 31,932 species were obtained from the Royal Botanic Gardens
247 Kew Seed Information Database [17]. Species names were standardised with The
248 Plant List (TPL) nomenclature [2] and cleaned using the *Taxonstand* R package [34].

249 Further processing at the genus-level was performed with the *taxonlookup* R package
250 [35], which is a complete genus-family-order mapping for vascular plants that draws
251 from TPL, the Angiosperm Phylogeny website [36] and a higher-level manually-
252 curated taxonomic lookup [16]. Mean seed mass values were subsequently calculated
253 for a total of 4763 genera.

254

255 We used the most comprehensive phylogenetic tree for land plants [16,37] that
256 comprises 31,389 species. Taxonomic information for our phylogenetic tree was run
257 through *Taxonstand* and *taxonlookup* as described above to make it as comparable as
258 possible to the seed mass dataset. Monophyly of the genera in the tree was assessed
259 using the *Monophy* package [38]. The initial estimate was 16%, but we removed taxa
260 classified as outliers by *Monophy* (i.e., taxa that lay outside of an established “core
261 clade” for each genus) and this resulted in only 8% of the genera not being
262 monophyletic. The *Phyndr* [35] package was then used to generate a genus-level tree
263 with as much overlap as possible between the phylogeny and the trait data. The final
264 genus-level phylogenetic tree included representatives from 303 plant families (Fig.
265 S6). We also obtained a species-level tree with 8448 representatives from the initial
266 phylogenetic tree that were also present in the Kew Seed Information Database.

267

268 *Diversification and phenotypic evolution analyses*

269 First, speciation, extinction and net diversification rates and rates of seed size
270 evolution were estimated separately on the genus-level phylogeny using BAMM
271 version 2.5.0 [18]. BAMM models shifts in macroevolutionary regimes across a
272 phylogenetic tree using reversible-jump Markov chain Monte Carlo (rjMCMC)
273 sampling. Initial prior settings were calculated with the *setBAMMpriors* function in

274 *BAMMtools* [39], and the *expectedNumberOfShifts* parameter was set at 25 for both
275 the speciation/extinction and trait evolution analyses, respectively. We incorporated
276 non-random incomplete sampling information following BAMM protocols by
277 calculating the proportion of genera sampled inside each family and estimated the
278 backbone sampling as the overall proportion of sampled genera. *Taxonlookup* was
279 used as a reference for these calculations. All BAMM analyses were run for 50
280 million generations. We verified convergence by plotting chain traces and ensuring
281 that the effective sample sizes of all relevant parameters exceeded 200. The first 10
282 million generations were discarded as burn-in. The *plotPrior* function in *BAMMtools*
283 was used to check the difference between model priors and posteriors. We also
284 analysed BAMM prior sensitivity by re-running both the diversification and the trait
285 evolution analyses with different settings for the *expectedNumberOfShifts* parameter
286 of either 10, 25, 50 or 100 (Fig. S1).

287

288 Second, estimates of net diversification rates were obtained for a curated set of
289 angiosperm genera using the species-level phylogeny. Only monophyletic genera with
290 four or more species in the tree were considered for further analysis. A minimum
291 threshold value for genus specific sampling fraction was set at 0.3. This ensured that
292 the minimum probability of recovering the true root node of any genera was 0.7
293 (“crown sampling probability”, following [40,41]). Crown ages for the selected
294 genera were then calculated and used to estimate net diversification rates using the
295 method-of-moments estimator [20]. The total number of species in each genus was
296 obtained from *taxonlookup*. Following standard practice, we assumed three values of
297 relative extinction fraction, $\epsilon = 0, 0.5$ and 0.9 . Different values did not affect our
298 conclusions; therefore we present the results of the intermediate extinction fraction (ϵ

299 = 0.5). The rate of seed size evolution was estimated with by fitting a Brownian
300 Motion (BM) model of evolution to the seed size data within each genus using
301 *fitContinuous* from the package *geiger* [42].

302

303 Finally, we obtained clade-based measures of diversification and seed size evolution
304 across the species-level tree. Clade-specific sampling fractions were calculated by
305 weighting the genus-specific sampling fractions (i.e., the number of congeneric
306 species in the 8448 species tree divided by the total number of described species for
307 that genus) by the number of species from each genus present in the clade. These
308 weighted sampling fractions were used to estimate the total clade species richness. A
309 minimum threshold of 0.3 was used for the clade-specific sampling fraction, again
310 ensuring that the crown sampling probability for each clade was at least 0.7 (the
311 actual median clade crown sampling probabilities for the time slices ranged between
312 0.84 and 0.95). We estimated net diversification rates with the method-of-moments
313 estimator as described above. We also used RPANDA to fit a series diversification
314 models that estimated time-dependent rates to each clade. We fitted six different
315 models of diversification: (i) pure birth model with constant λ (speciation rate); (ii)
316 pure birth model with exponential λ ; (iii) birth-death model with constant λ and μ
317 (extinction); (iv) birth-death model with constant λ and exponential μ ; (v) birth-death
318 model with exponential λ and constant μ ; and (vi) birth-death model with exponential
319 λ and exponential μ . We then used AIC-based model selection to select the best fitting
320 model and obtain the corresponding macroevolutionary parameters. Additionally, to
321 account for possible biases when analysing clades with many non-congeneric species,
322 we confirmed the results of our clade-based analysis but only considered clades that
323 only contained congeneric species (Fig. S7).

324

325 *Correlation of diversification and trait evolution*

326 All rate variables were log-transformed for the correlation analyses. We used
327 STRAPP to test for multiple associations between BAMM-estimated
328 macroevolutionary dynamics and each of seed mass (using genus mean values as
329 character states) and seed mass rate of evolution (using seed mass evolutionary rates
330 at the genus-level tips of the phylogeny as character states). STRAPP compares the
331 correlation between a focal trait and a macroevolutionary parameter (λ , μ or r) to a
332 null distribution of correlations. The null correlations are generated by permuting the
333 evolutionary rates in the tips of the phylogenetic tree while maintaining the location
334 of rate shift events in the phylogeny. In each case, we calculated the absolute
335 difference between the observed correlation of the macroevolutionary rate and the
336 trait state and the null correlation obtained by the structured permutations across 5000
337 samples from the BAMM posterior. The reported p-value was the proportion of
338 replicates where the null correlation coefficient was greater than the observed
339 correlation. We found a low type I error associated with our STRAPP correlation
340 analysis (p-value = 0.11, Fig. S8).

341

342 For both the curated genera dataset and the clade based analyses, we estimated the
343 relationship between speciation rates with seed size rate of evolution and mean genus
344 seed size with phylogenetic generalised least square (PGLS) as implemented in the R
345 package *caper* [43] (Fig. S9, Fig. S10). Similar results as presented in the main text
346 were obtained when analysing net diversification instead of speciation rates (Fig.
347 S11). Finally, we similarly analysed the correlation between speciation and each of
348 seed size and its rate of change, this time selecting only clades consisting of

349 congeneric species. Again, this analysis resulted in a similar pattern as the one
350 presented in the main text (Fig. S7).

351

352 *Code availability*

353 Scripts used to carry out the analysis described in the paper and generate the figures
354 will be deposited in Github upon acceptance.

355

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362

363 **Author contributions**

364 J.I and A.S.T.P conceived the study. J.I. and E.F.M. performed the analysis. J.I. and
365 A.J.T interpreted the analysis and wrote the manuscript. All authors edited the
366 manuscript.

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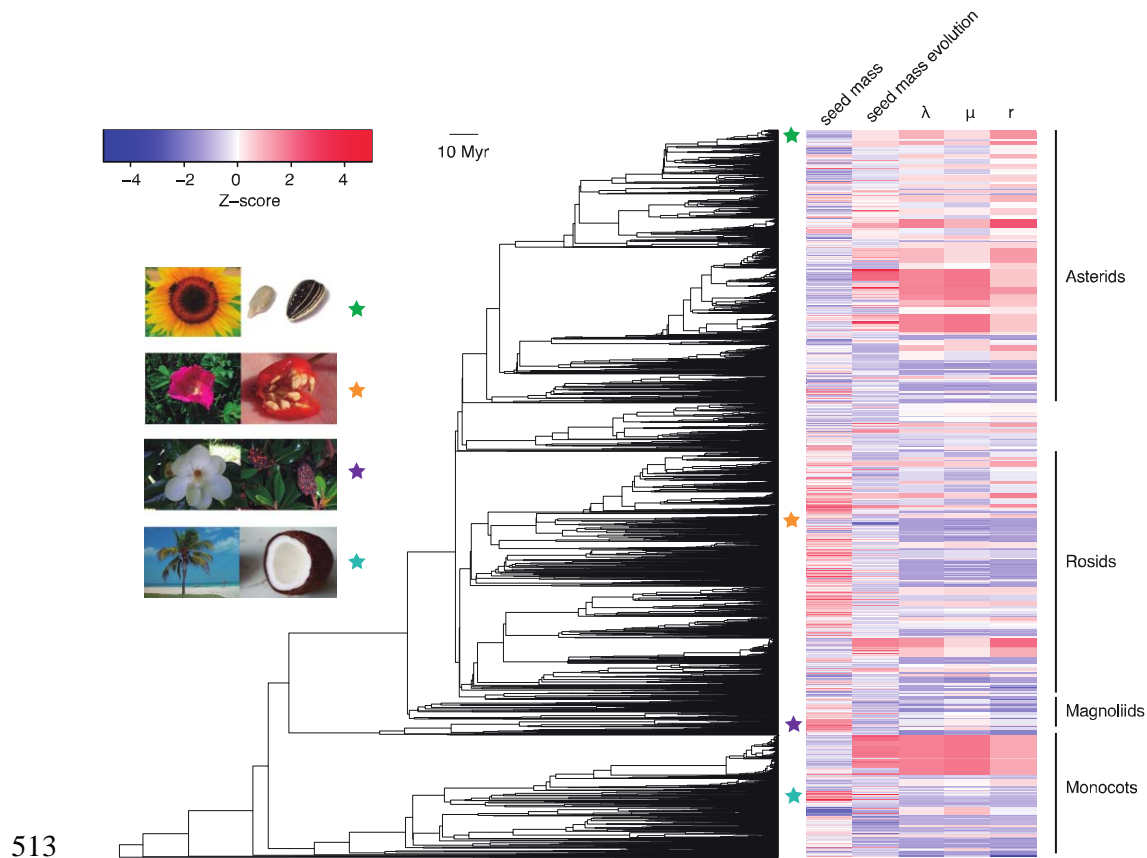
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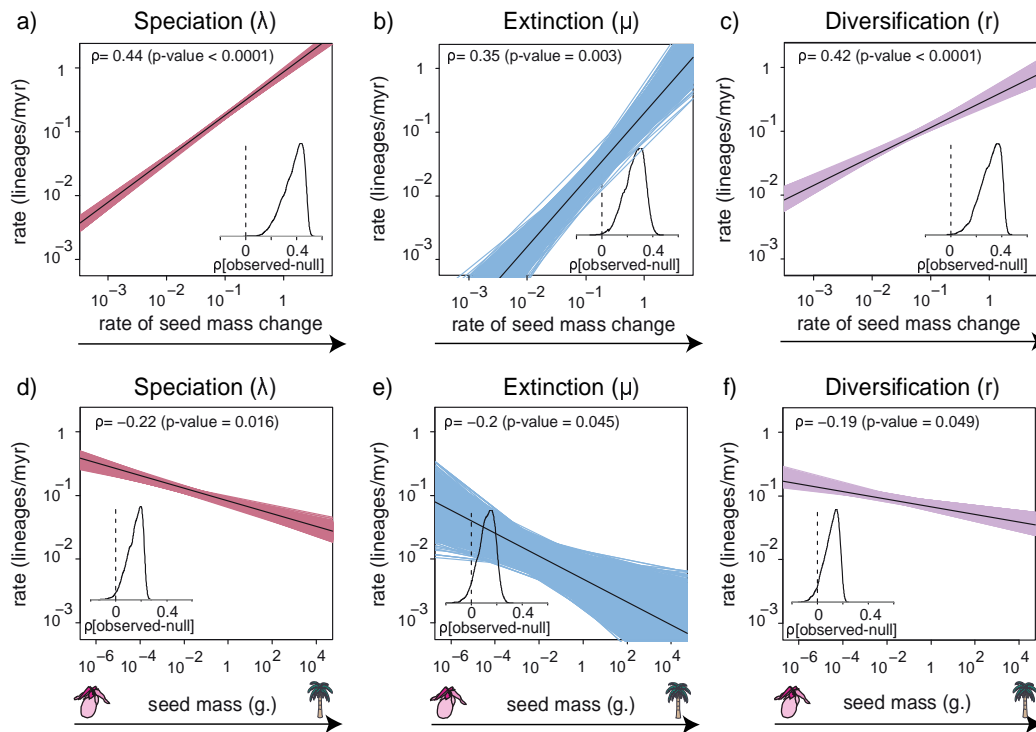
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514 **Figure 1. Seed mass and macroevolutionary rates estimated with BAMM across**
515 **the angiosperm tree of life.** Phylogenetic tree of 4105 genera of plants with seed
516 mass, rate of seed mass change, and speciation (λ), extinction (μ) and net
517 diversification (r) rates estimated by BAMM. Seed mass and rate data were
518 standardised to Z-scores so that variation could be directly compared.

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Figure 2. Seed mass and seed mass rate of evolution correlate with

537 **macroevolutionary dynamics estimated with BMM.** Spearman correlations were

538 calculated between speciation (λ), extinction (μ), and net diversification (r) and each

539 of a) rate of seed mass change and b) seed mass. Coloured lines are correlations for

540 one sample of the BMM posterior distribution, bold line is the median. The insets

541 show the density plots of the absolute difference between the observed and null

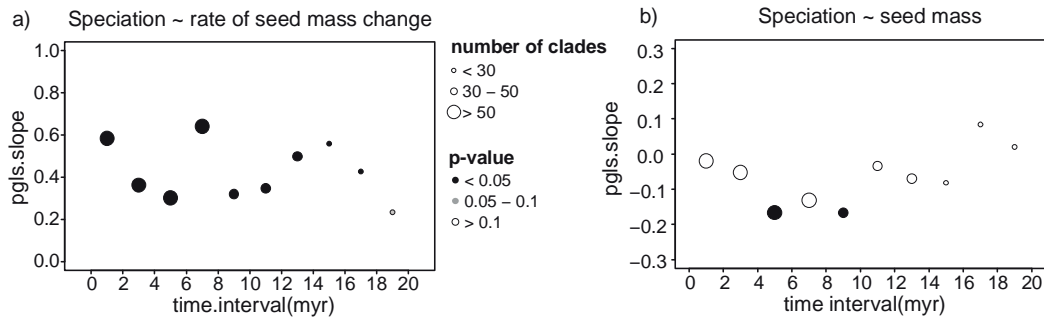
542 correlation calculated across 5000 structured permutations of the evolutionary rates

543 on the phylogenetic tree (myr = million years).

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549 **Figure 3. Seed mass and seed mass rate of evolution correlate with speciation in**

550 **the clade-based analysis.** Correlation of (a) rate of seed mass evolution and (b) seed

551 mass with speciation rate (λ) estimated using RPANDA in the clade-based analysis.

552 The strength of correlations is shown as PGLS slopes and were calculated using mean

553 clade-level seed mass across 10 time slices in the species-level phylogenetic tree. The

554 size of the circles represents the number of clades in each time slice while the colour

555 indicates the significance of the slope. A detailed representation of the results in each

556 time slice is given in figures S9 and S10.

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580 **Supporting Information captions**

581

582 **Fig. S1.** Prior and posterior distribution of the number of rate shifts in BAMM for a)
583 the speciation/extinction and b) phenotypic evolution analyses for
584 *expectedNumberOfShifts* = 10, 25, 50 and 100. The analyses in the main text were
585 carried out with *expectedNumberOfShifts* = 25 for both speciation/extinction and
586 phenotypic evolution analyses.

587

588 **Fig. S2.** Phylogenetic tree of 4105 angiosperm genera with branches scaled by
589 diversification rate as determined by BAMM. The colours of the branches indicate the
590 rate of seed mass evolution, also estimated with BAMM.

591

592 **Fig. S3.** Seed mass and its rate of evolution are associated with speciation in the
593 clade-based analyses. (a) PGLS slope of the relationship between speciation rate (λ)
594 from the method-of-moments estimator and the rate of seed mass evolution across
595 10 time slices in the species level phylogenetic tree. Circles are scaled to the number
596 of clades in each time slice while colour indicates the significance of the slope. (b)
597 PGLS slope of the relationship between speciation rate and mean clade seed mass.
598 For a detailed representation of the results in each time slice, see figure S12 and S13.

599

600 **Fig. S4.** STRAPP correlations of diversification and phenotypic traits for 938
601 angiosperm genera. The distribution of the absolute difference in the observed
602 correlation minus the null correlation is plotted for each trait. The coloured dotted
603 lines indicate the mean of that distribution, and the black dotted line indicates 0; a
604 distribution with mean = 0 would show no association between a focal trait and

605 speciation. STRAPP correlation of seed mass (shown in blue), C-value (shown in red)
606 and life cycle (shown in green) with a) speciation rate (λ), b) extinction rate (μ), and
607 c) net diversification rate (r).

608

609 **Fig. S5.** Proposed effects of seed mass and other life history traits on diversification
610 (solid lines). Dashed lines indicate correlations between life history traits. Numbers
611 indicate reference where the link is proposed.

612

613 **Fig. S6.** Phylogenetic tree of 303 angiosperm families with representatives in this
614 study. The red bars indicate the levels of sampling for each family.

615

616 **Figure S7.** Correlation of speciation with seed mass and seed mass rate of evolution
617 in the clade-based analysis only considering congeneric species. (a) PGLS slope of
618 the relationship of speciation rate - estimated with the method-of-moments estimator -
619 with mean clade seed mass across 10 time slices in the species level phylogenetic tree.
620 The size of the circles represents the number of clades in each time slice while the
621 colour indicates the significance of the slope. (b) PGLS slope of the relationship of
622 speciation rate and the rate of seed mass evolution.

623

624 **Fig. S8.** Type I error analysis. We estimated the type I error rate of our analysis by
625 simulating neutral traits on the angiosperm phylogenetic tree. We performed 1000
626 simulations and then ran 1000 STRAPP tests with each simulated dataset. We
627 estimated the corresponding p-values for the association between traits and
628 diversification and calculated the type I error as the proportion of datasets where a
629 significant association (p-value < 0.05) was detected.

630

631 **Fig. S9.** Correlations between clade rate of seed size evolution and speciation rate
632 (estimated with RPANDA) across time slices from: (a) 0 to 2 million years (myr); (b)
633 2 to 4 myr; (c) 4 to 6 myr; (d) 6 to 8 myr; (e) 8 to 10 myr; (f) 10 to 12 myr; (g) 12 to
634 14 myr; (h) 14 to 16 myr; (i) 16 to 18 myr; and (j) 18 to 20 myr.

635

636 **Fig. S10.** Correlations between mean clade seed mass and speciation rate (estimated
637 with RPANDA) across time slices from: (a) 0 to 2 million years (myr); (b) 2 to 4 myr;
638 (c) 4 to 6 myr; (d) 6 to 8 myr; (e) 8 to 10 myr; (f) 10 to 12 myr; (g) 12 to 14 myr; (h)
639 14 to 16 myr; (i) 16 to 18 myr; and (j) 18 to 20 myr.

640

641 **Figure S11.** Correlation of (a) rate of seed mass evolution and (b) seed mass with net
642 diversification rate (r) estimated using RPANDA in the clade-based analysis. The
643 strength of correlations is shown as PGLS slopes and were calculated using mean
644 clade-level seed mass across 10 time slices in the species-level phylogenetic tree. The
645 size of the circles represents the number of clades in each time slice while the colour
646 indicates the significance of the slope.

647

648 **Fig. S12.** Correlations between clade rate of seed size evolution and speciation rate
649 (estimated with the method-of-moments estimator) across time slices from: (a) 0 to 2
650 million years (myr); (b) 2 to 4 myr; (c) 4 to 6 myr; (d) 6 to 8 myr; (e) 8 to 10 myr; (f)
651 10 to 12 myr; (g) 12 to 14 myr; (h) 14 to 16 myr; (i) 16 to 18 myr; and (j) 18 to 20
652 myr.

653

654 **Fig. S13.** Correlations between mean clade seed mass and speciation rate (estimated

655 with the method-of-moments estimator) across time slices from: (a) 0 to 2 million
656 years (myr); (b) 2 to 4 myr; (c) 4 to 6 myr; (d) 6 to 8 myr; (e) 8 to 10 myr; (f) 10 to 12
657 myr; (g) 12 to 14 myr; (h) 14 to 16 myr; (i) 16 to 18 myr; and (j) 18 to 20 myr.

658

659 **Fig. S14.** Mean genus seed mass of strict annual ($n = 106$) and perennial ($n = 832$)
660 genera. No significant difference between the means of the two groups was found
661 when phylogeny was accounted for (PGLS, $F_{1,937} < 0.01$, $p\text{-value} = 0.972$).

662

663 **Fig. S15.** Mean genus seed mass and coefficient of variation for 2613 angiosperm
664 genera. The two variables were not significantly correlated (PGLS test, $F_{1,131} = 0.67$,
665 $p\text{-value} = 0.416$).

666

667 **Fig. S16.** Correlation of coefficient of variation in seed mass with macroevolutionary
668 dynamics. STRAPP correlations with speciation (λ), extinction (μ) and net
669 diversification (r) and the genus-level coefficient of variation in seed mass for 2613
670 angiosperm genera (i.e. each with seed mass data for >1 species in the Kew SID).
671 Correlations and insets calculated as in Fig. 2.

672

673 **Table S1.** STRAPP correlations for 938 genera of angiosperms with seed mass,
674 genome size (i.e., C-value) and life cycle data. Macroevolutionary parameters were
675 obtained from an analysis with 4105 genera (see Main text).

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