- 1 Title: Unsupervised characterization of dynamic functional connectivity reveals
- 2 age-associated differences in temporal stability and connectivity states during rest
- 3 and task
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14 Abstract

Understanding brain functions as an outcome of underlying neuro-cognitive network mechanisms in 15 rest and task requires accurate spatiotemporal characterization of the relevant functional brain 16 17 networks. Recent endeavours of the Neuroimaging community to develop the notion of dynamic 18 functional connectivity is a step in this direction. A key goal is to detect what are the important events in time that delimits how one functional brain network defined by known patterns of correlated brain 19 20 activity transitions into a "new" network. Such characterization can also lead to more accurate 21 conceptual realization of brain states, thereby, defined in terms of time-resolved correlations. 22 Nonetheless, identifying the canonical temporal window over which dynamic functional connectivity 23 is operational is currently based on an ad-hoc selection of sliding windows that can certainly lead to 24 spurious results. Here, we introduce a data-driven unsupervised approach to characterize the high 25 dimensional dynamic functional connectivity into dynamics of lower dimensional patterns. The 26 whole-brain dynamic functional connectivity states bearing functional significance for task or rest 27 can be explored through the temporal correlations, both short and long range. The present study 28 investigates the stability of such short- and long-range temporal correlations to explore the dynamic 29 network mechanisms across resting state, movie viewing and sensorimotor action tasks requiring 30 varied degrees of attention. As an outcome of applying our methods to the fMRI data of a healthy 31 ageing cohort we could quantify whole-brain temporal dynamics which indicates naturalistic movie 32 watching task is closer to resting state than the sensorimotor task. Our analysis also revealed an 33 overall trend of highest short range temporal network stability in the sensorimotor task, followed by 34 naturalistic movie watching task and resting state that remains similar in both young and old adults. 35 However, the stability of neurocognitive networks in the resting state in young adults is higher than 36 their older counterparts. Thus, healthy ageing related differences in quantification of network stability 37 along task and rest provides a blueprint of how our approach can be used for cohort studies of mental 38 health and neurological disorders.

40 1. Introduction

41 Functional connectivity (FC) - most simplistically computed using the pairwise Pearson correlations 42 between brain regions using blood oxygen level dependent (BOLD) fMRI has proven to be a powerful 43 tool for studying the functional organization of the brain (Friston, et al. 1993). FC sheds light on the 44 functional coupling and connectedness between proximal and distal brain regions subserving crucial 45 role towards the neuronal processing of a task (Aertsen, Gerstein, Habib, & Palm, 1989) (Friston, 46 Frith, Liddle, & Frackowiak, 1993). However, emergence of superior computing prowess has allowed 47 us to critique the inferences drawn from time-averaged, static FC usually computed by collapsing the 48 functional dynamics across time (Ciric, Nomi, Uddin, & Satpute, 2017) (Mash, et al., 2019). 49 Recently, dynamic functional connectivity (dFC) has emerged as a major topic in the resting-state 50 BOLD fMRI literature (Hutchison et al., 2013). The more refined measure of Dynamic functional 51 connectivity (dFC) is commonly computed using the sliding window framework, which estimates 52 dFC by computing average FC over small windows of time, and subsequently sliding the window 53 over the entire duration of the BOLD time series (Hutchison & et al, 2013). Although, the sliding 54 window approach has been the most common, simple, and intuitive analysis strategy for estimating 55 dFC (Kudela, Harezlak, & Lindquist, 2017) (Preti, Bolton, & Van De Ville, 2017), the method suffers from prominent drawbacks. Arbitrary choice of window length, inherent variation present in the 56 57 estimate that can be confused with the empirical time-varying nature of FC, equal weighting of all 58 observations within the window leading to spurious fluctuations being magnified - all add to the woes 59 of sliding window based approach (Lindquist, Xu, Nebel, & Caffo, 2014) (Hindriks, et al., 2016) 60 (Preti, Bolton, & Van De Ville, 2017). Over the years, many meaningful extensions have been 61 suggested to improve sliding window approach. Independent component analysis (ICA) was used to 62 decompose windowed BOLD timecourses (Kiviniemi, et al., 2011). Several graph theoretical 63 summary measures such as assortivity, modularity, efficiency offer promising avenues to extract 64 information from dFC (Bullmore & Sporns, 2009). In addition, clustering algorithms such as K-

65 means clustering (Damaraju, et al., 2014) (Allen, et al., 2014), hidden Markov models (HMM) (Vidaurre, Smith, & Woolrich, 2017), temporal ICA (TICA) (Yaesoubi, Miller, & Calhoun, 2015) 66 allows to identify clustering-derived recurring connectivity patterns or dFC states. Several conceptual 67 68 alternative strategies such as wavelet transform coherence (Chang & Glover, 2010), a time/frequency 69 analysis strategy with an observation window for the frequency content of the time courses; and frame-wise analysis of the BOLD timecourses (Cabral, et al., 2017), which allows information to be 70 71 retrieved from the observation of single frames and yield temporally subsequent co-activation maps 72 (Liu, Chang, & Duyn, 2013); have been suggested (see (Preti, Bolton, & Van De Ville, 2017) for a 73 review).

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75 In spite of inherent limitations, dFC captures the fluctuations of FC, which contain meaningful 76 information on minute temporal scale (Hutchison & et al, 2013). While, accounting these fluctuations 77 are critical for understanding complex behaviour, nonetheless, stable representation of information 78 of neural activity and corresponding stability of FC patterns over time is critical for survival (Li, Lu, 79 & Yan, 2019). In other words, in an axiomatic sense, there must exist a temporally stable connectivity 80 pattern that corresponds to one or multiple functional states of the brain. Subsequent transition 81 between successive functional brain states can be characterized by estimating the dFC patterns. The 82 two most widely applied dynamic measures in brain/behaviour analyses that are constructed from 83 dFC time courses based on either applying sliding window or frame wise analysis are connection 84 variability (CV) and connectivity states (CS). Stable representation of information processing will 85 reflect the robustness of recurring patterns of CS concatenated across subjects to influence myriad of 86 behaviour (Bolton, Morgenroth, Preti, & Van De Ville, 2020). Identifying both connectivity states 87 and their stability in brain dynamics requires delimiting the dFC evolution and dFC stability with 88 measures of optimality which may or may not be crucially linked with the underlying subject specific 89 structural connectivity (Surampudi, et al., 2019). The goal of this article is to develop an unsupervised

approach to characterize optimal dFC states and go beyond what has been proposed in the existing
literature based on widely applied sliding window or frame wise based approach (Bolton,
Morgenroth, Preti, & Van De Ville, 2020).

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Previous studies exploring temporal dynamics of FC have tried to investigate the stability by 94 95 calculating the correlation between FC matrices computed from successive temporal windows 96 (Hansen, Battaglia, Spiegler, Deco, & Jirsa, 2015), characterizing CV of the functional connectivity 97 profile of a given region across time (Zhang & et al, 2016) (Guo, Zhao, Tao, Liu, & Palaniyappan, 98 2017), by estimating voxel level dFC maps using Kendall's coefficient of concordance with time 99 windows as raters (Li, Lu, & Yan, 2019), by estimating the standard deviation of global modularity 100 averaged across all timepoints and all participants (Hilger, Fukushima, Sporns, & Fiebach, 2019). FC 101 stability has been shown to increase with motor learning (Yu, Song, Huang, Song, & Liu, 2020), 102 decrease in patients of schizophrenia and their siblings (Guo, Zhao, Tao, Liu, & Palaniyappan, 2017), 103 was significantly higher in patients with major depressive disorder (Demirtas, et al., 2016). These 104 studies emphasise the neurobiological significance of the stability of FC.

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106 Quantifying the temporal stability of dFC patterns is of immediate concern to studies investigating 107 the relationship between resting state and task-related brain dynamics (Li, Lu, & Yan, 2019). 108 Spontaneous brain activity during rest is not random and show specific spatio-temporal organization 109 in state space (Deco, Jirsa, & Mcintosh, 2011). From a dynamical systems point of view, the resultant 110 emerging resting state functional connectivity of the brain networks, quantitatively, fits best with the 111 experimentally observed functional connectivity when the brain network operates at the edge of 112 instability. Under these near critical conditions, the slow fluctuating (< 0.1 Hz) resting state BOLD 113 networks emerge as structured noise fluctuations around a stable low firing activity equilibrium state 114 in the presence of latent "ghost" multi stable attractors (Deco & Jirsa, 2012). Recent work has further

115 demonstrated that during spontaneous resting state activity the ghost attractors makes frequent 116 excursion to functionally and behavioural relevant phase locking states in a low dimensional state 117 space (Vohryzek, Deco, Cessac, Kringelbach, & Cabral, 2020). Brain resides in a specific attractor 118 state defined by a certain FC pattern according to the cognitive demands of the task (Fedorenko & 119 Thompson-Schill, 2014) (Pillai & Jirsa, 2017). An overall increase in FC stability has been reported in the presence of the task (Gonzalez-Castillo & Bandettini, 2018). Subsequently, temporal stability 120 121 of FC guides the stability of a functional state. Thus, we tested the following hypothesis, unsupervised 122 dFC characterization will reveal task specific dFC stability patterns that are local in time, whereas for 123 the resting state dFC patterns, these functional states are composed of non-local correlations in time. 124 Prior studies have also explored changes in temporal stability of functional architecture in resting 125 state of healthy control and patients with psychiatric disorders, and different battery of tasks. Zhang 126 and colleagues showed disorder specific (ADHD, schizophrenia, autism spectrum disorder) 127 variability modifications in functional architecture of DMN, visual and subcortical regions of the 128 brain (Zhang & et al, 2016). Increased functional stability in high-order visual regions during 129 naturalistic movie watching task were identified (Li, Lu, & Yan, 2019), but these studies are limited 130 to stability of FC of a given region. The second test for an unsupervised dFC characterization 131 algorithm will be application to a specific neuroscience problem, such as investigation of lifespan 132 trajectories in healthy ageing. Although previous studies have explored the association between 133 dynamic functional connectivity and age (Viviano, Raz, Yuan, & Damoiseaux, 2017) (Chen, et al., 134 2017) (Xia, et al., 2018), how the stability of functional architecture modifies across age remains an 135 open question.

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137 The aim of the present study is three-fold: 1) to precisely characterise the stability of whole-brain 138 dFC patterns 2) to demonstrate that dFC patterns are locally stable during task 3) identify the dFC 139 patterns during task and rest for a cross-sectional population with age range over human adult lifespan

140 (18-88 years). This manuscript is organized as follows. First, we estimate BOLD phase coherence over time (Glerean, Salmi, Lahnakoski, Jääskeläinen, & Sams, 2012) which was used as a measure 141 of dFC for rest and task. Next, we proceed with unsupervised characterization of dFC subspaces 142 143 involved in task and rest. Subsequently, the temporal stability of dFC subspaces were computed using 144 two different measures - angular separation and the Mahalanobis distance (Mahalanobis, 1930) (Shen, 145 Kim, & Wang, 2010). Finally, the temporal stability of dFC was analysed to draw critical insights 146 about age associated differences to task and rest using a large human cohort of healthy ageing (Shafto, 147 et al., 2014).

149 **2. Methods**

- 150 2.1 Data sources and participants
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The data were collected as part of stage 2 of the Cambridge Centre for Ageing and Neuroscience 152 153 (CamCAN) project (available at http://www.mrc-cbu.cam.ac.uk/datasets/camcan) (Taylor, et al., 154 2017) (Shafto, et al., 2014). The CamCAN is a large-scale multimodal, cross-sectional, population-155 based study. The database includes raw and pre-processed structural magnetic resonance imaging 156 (MRI), resting state and active tasks using functional MRI (fMRI) and Magnetoencephalogram 157 (MEG), behavioural scores, demographic and neuropsychological data. From 3000 participants of 158 stage 1, a subset of approximately 700 participants who were cognitively healthy (MMSE score >25), 159 with no past or current treatment for drug abuse or usage, met hearing threshold greater than 35 dB 160 at 1000 Hz in both ears, had at least a corrected near vision of 20/100 with both eyes and could speak 161 English language (native English speaker or bilingual English from birth) were eligible for MRI 162 scanning. They were home interviewed and recruited to stage 2. The study was in compliance with 163 the Helsinki Declaration and was approved by the Cambridgeshire 2 Research Ethics Committee. The 164 fMRI data from resting state and task periods (naturalistic movie watching and sensorimotor task) 165 was used in the present study.

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167 2.2 Data acquisition and experimental paradigm

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The fMRI data were collected at MRC Cognition and Brain Sciences Unit, on a 3T Siemens TIM Trio scanner with a 32-channel head coil, the head movement was restricted with the aid of memory foam cushions. For the tasks, the instructions and visual stimuli were back projected onto the screen, auditory stimuli were presented via MR-compatible Etymotics headphones and manual responses from the participants made with the right hand were recorded using an MR-compatible button box

174 (Taylor, et al., 2017). The fMRI data for eyes-closed resting state and sensorimotor task were acquired 175 using Echo-Planar Imaging (EPI) sequence, consisted of 261 volumes, each volume with 32 axial slices (slice thickness 3.7mm, interslice gap 20%) acquired in descending order, TR 1970 ms, TE 30 176 177 ms, voxel-size 3 mm 3 mm 4.44 mm. The duration of both the scans was 8 min 40s. The fMRI data 178 for the naturalistic movie watching task were acquired using multi-echo EPI sequence, consisting of 179 193 volumes of 32 axial slices each (slice thickness 3.7mm, interslice gap 20%) acquired in 180 descending order, TR 2470 ms, TE [9,4,21.2,33,45,57] ms, voxel-size 3 mm 3 mm 4.44 mm. The 181 duration of the scan was 8 min 13s.

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183 The task-induced BOLD data from the naturalistic movie watching task was acquired from 184 participants, who watched 8 minutes of narrative preserved, condensed, black and white version of 185 Alfred Hitchcock's television drama "Bang! You're Dead". The participants were not aware of the 186 title of the movie but were instructed to pay attention to the movie. In the sensorimotor task, the trials 187 consisted of a binaural tone simulation at either 300, 600, or 1200 Hz and bilateral black and white 188 checkerboard. The participants were asked to button press with their right index finger if they hear or 189 see any stimuli. More details about the task paradigm have been presented here (Shafto, et al., 2014) 190 (Taylor, et al., 2017).

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192 2.3 Data pre-processing

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Pre-processed data was provided by Cam-CAN research consortium. Mean regional BOLD time series were estimated in 116 parcellated brain areas of Anatomical Automatic Labelling atlas (AAL) (Tzourio-Mazoyer, et al., 2002) (available at <u>http://www.gin.cnrs.fr/tools/aal</u>). We selected 50 participants, 25 were young adults (48% female; mean age = 24.1 ± 3.33 years) randomly selected from age range 18-28, and remaining 25 were old adults (52% female; mean age = 63.8 ± 2.63 years)

199	randomly selected over age range 60-68 years. Each participant's BOLD time series in the resting
200	state, naturalistic movie watching and sensorimotor tasks were extracted.

201 2.4 Data analysis

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203 2.4.1 Characterization of dynamic functional connectivity

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205 Time-resolved dynamic functional connectivity (dFC) was estimated, for each individual, using 206 BOLD phase coherence (Figure 1A) (Glerean, Salmi, Lahnakoski, Jääskeläinen, & Sams, 2012) 207 (Ponce-Alvarez, et al., 2015) (Deco & Kringelbach, 2016) (Cabral, et al., 2017), which resulted in a 208 matrix with size NxNxT, where N=116 is the number of brain regions defined by AAL atlas, T is the 209 total number of time points (T=261 for resting state and Sensorimotor task, T=193 for naturalistic 210 movie watching task). We chose BOLD phase coherence instead of computing correlation over a 211 sliding window to calculate dFC, because BOLD phase coherence is an instantaneous measure with 212 maximum temporal resolution (Glerean, Salmi, Lahnakoski, Jääskeläinen, & Sams, 2012). BOLD 213 phase coherence does not require time-windowed averaging, that generates biased estimates if the 214 window length is short and reduces temporal resolution if the window length is longer (Glerean, 215 Salmi, Lahnakoski, Jääskeläinen, & Sams, 2012).

216

First, the instantaneous phases $\theta(n, t)$ of the BOLD time series for all the brain regions, *n*, was computed using Hilbert transform. The real-valued modulated BOLD signal s(t) is expressed as an analytical signal in the complex plane as:

220

221
$$z(t) = z_r(t) + jz_i(t) = s(t) + j HT[s(t)]$$
 (1)

223 Where, (HT [*]) represents the Hilbert transform. The instantaneous phase θ (*t*) is computed as 224 follows:

225

226
$$\theta(t) = \angle z(t) = \arctan \frac{z_i(t)}{z_r(t)} = \arctan \frac{HT[s(t)]}{s(t)}$$
 (2)

227

Given the phases of the BOLD time series, phase coherence i.e., dFC(n, p, t) for brain regions, *n* and *p* at time *t* is computed as:

(3)

230

231
$$dFC(n, p, t) = \cos(\theta(n, t) - \theta(p, t))$$

232

when, the phases of BOLD signals, $\theta(n,t)$, $\theta(p,t)$ of the brain regions n,p are synchronized, dFC(n,p,t) (ranges from -1 to 1) is close 1, when the phases from the BOLD signals of brain regions n,p are orthogonal dFC(n,p,t) is close to 0. Since the phases are undirected, dFC(n,p,t) is symmetric along the diagonal.

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In addition to this, to check for reliability, we compute dFC using a sliding-window approach (Hutchison & et al, 2013) with non-overlapping, gaussian windows, varying the window length (10, 20, 30 time points) (Supplementary information - **S 1**, **S 2**, **S 3**).

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242 2.4.2 Extracting Dominant dynamic functional connectivity

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Principal component analysis (PCA) was applied to participant-wise dFC(n, p, t) matrix of size NXN representing the FC between nth and pth brain area for each time point. PCA is an unsupervised, multivariate dimension reduction method that decomposes the data into a set of orthogonal principal components or leading eigenvectors sorted by their contribution to the overall variance (Friston, 1993). Thus, dFC (n, p, t) or simply dFC_t can be expressed as

$$dFC_t = V^T SV \tag{4}$$

250 where, matrix V of size NXN are set of eigenvectors, with each column of V of size 1XN

251 representing orthogonal principal component, and **S** the diagonal matrix $\begin{pmatrix} \lambda_1 & \cdots & 0 \\ \vdots & \ddots & \vdots \\ 0 & \cdots & \lambda_N \end{pmatrix}$, such that

252 $\lambda_1 > \lambda_2 \dots > \lambda_N$

253 If k is the number of principal components chosen to represent dFC, the corresponding subspace

254 D(n, k, t) or D_t , representative of dominant dFC pattern, can be expressed as

$$255 D = \tilde{V}^T \tilde{S} \tilde{V} (5)$$

256 where, \tilde{V}^T is a dimensionally reduced matrix of size N X k, \tilde{S} is a diagonal matrix $\begin{pmatrix} \lambda_1 & \cdots & 0 \\ \vdots & \ddots & \vdots \\ 0 & \cdots & \lambda_k \end{pmatrix}$.

In this study, we chose k = 3 because for all participants at least 99% variance in dFC matrix is captured by the 3 leading eigenvectors (S 4). The dimension of dFC(n, p, t) has been reduced to D(n, k, t).

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262 2.4.3 Computation of stability of dynamic functional architecture

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We seek to characterize the temporal stability of the dominant subspace D(n, p, t) (or referred to as simply D_t) by estimating how similar they are across time t. To estimate the similarity between dominant dFC configurations, we introduce two types of distance measures successive dFC subspaces, 1) angular distance 2) Normalised Euclidean distance (**Figure 1B**). We define angular distance as the principal angle between the dFC subspaces from different time points, given by thefollowing equation:

270

271
$$\phi(t_x, t_y) = \angle (\boldsymbol{D}_{t_x}, \boldsymbol{D}_{t_y})$$
 (6)

272

273 Where, each entry in the time X time *temporal stability matrix*, $\phi(t_x, t_y)$ is the principal angle 274 between the two N X k dimensional subspaces at t_x and t_y (Banerjee, Pillai, Sperling, Smith, & 275 Horwitz, 2012) (Björck & Golub, 1973). The principal angle ranges between 0 (low angular distance) 276 to $\pi/2$ (high angular distance).

For each individual, we calculate the angular distance between dominant dFC subspaces at t_x and t_y , by estimating the principal angle between them. The low principal angle between dominant dFC subspaces means that their dFC configurations are very similar. On the contrary, the high principal angle between dominant dFC subspaces means that their dFC configurations are dissimilar.

281

We define the normalised Euclidean distance between dominant dFC subspaces by the Mahalanobis distance. Mahalanobis distance measures the distance between points in space 1 from space 2 with the following equation:

285

286
$$M^{2} = \left(\boldsymbol{D}_{t_{x}} - \boldsymbol{D}_{t_{y}}\right)^{T} C^{-1} (\boldsymbol{D}_{t_{x}} - \boldsymbol{D}_{t_{y}})$$
(7)

where M^2 is the distance between each entry of D_{t_x} and D_{t_y} Subsequently, for each individual, we estimate the time X time *temporal stability matrix*, where each entry is the Mahalanobis distance (*M* ranges between 0.5 to 2.5), averaged across all brain parcels. Low *M* means that dominant dFC subspaces are similar, high *M* means that the dFC subspaces are dissimilar.

292 2.4.4 Quantifying complexity of temporal stability matrices

293 Entropy:

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To evaluate the informational content of temporal stability matrices we evaluated the entropy, for all three categories, rest, movie viewing and sensorimotor task in young and old adults. Entropy is defined by the following equation:

298

$$E = -\sum p \log(p)$$

300 where p contains the normalised histogram counts returned from 'imhist.m'. 'imhist.m' calculates 301 the histogram of temporal stability matrices and returns histogram counts.

(8)

302

Entropy in temporal stability matrices for empirical data and surrogate data were compared by generating random time series using MATLAB function 'randn.m' and down sampling them at 0.1 Hz mimicking BOLD activity of each subject. The D matrices with same dimensions as the empirical data were computed. Welch's corrected t-tests revealed significant differences between the entropy of surrogate and empirical temporal stability matrices of rest (p=0.000464) (S 5).

308

309 Frobenius norm:

Frobenius norm was used to measure the differences between the temporal stability matrices computed for rest and the task conditions. Frobenius norm, also called the Euclidean norm of a matrix, is defined as the square root of the sum of the absolute squares of its elements. Here, we calculate Frobenius norm between temporal stability matrices with the following equation:

315
$$||x_F|| = \sqrt{\sum_{i=1}^T \sum_{j=1}^T |a_{ij} - b_{ij}|^2}$$
 (9)

where a_{ij} and b_{ij} are the entries in the temporal dynamic matrices of rest and any of the task conditions respectively (movie watching or sensorimotor). x_F is also computed between the two tasks.

318

319 Stochastic characterization of dFC

320 The temporal variation of two measures, principal angle and Mahalanobis distance between the dominant *dFC* subspaces essentially capture the degree of temporal variation in functional network. 321 Principal angular values close to $\frac{\pi}{2}$ or high Mahalanobis distance at a specific time point reflects the 322 reorganization of the functional state itself, whereas angular values closer to zero or low Mahalanobis 323 distance indicates minor deviation from previous time. To understand the underlying stochastic 324 325 characteristics of these measures, we use auto-regressive (AR) models where present values of $\phi(t)/M(t)$ are modelled as a linear weighted sum of values from past $\phi(t-1), \phi(t-2) \dots \phi(t-1)$ 326 $i)/M(t-1), M(t-2) \dots M(t-i)$. The AR (ρ) process, X_t $(\phi(t) \text{ or } M(t))$ is given by the 327 328 following equation:

329
$$X_t = c + \sum_{i=1}^{\rho} \varphi_i X_{t-i} + \varepsilon_t$$
(10)

330 where $\varphi_1 \dots \dots \dots \varphi_{\rho}$ are parameters of the model, *c* is a constant, ε_t is white noise and ρ is the lag 331 term or model order. The simplest AR process is AR (0) is essentially a white noise process. In AR 332 (1), the current value is dependent only on its immediately preceding value, and hence captures a 333 Markovian process. Optimal model of an AR process can be computed using the Akaike information 334 criterion (AIC) which is expressed as

$$335 \qquad AIC(\rho) = -2L + 2\rho \tag{11}$$

where *L* is the likelihood function computed by summing up over the mean squared error for an AR model of order ρ (Wagenmaker & Farrell, 2004) (H.Akaike, 1974).Optimal model order can be selected at a value of ρ where AIC is minimum. We varied the model order (ρ) from 0 to 100 and use the first minimal AIC value to select the best AR (ρ), model. If the model order is found to be greater than 1, the underlying process is considered non-Markovian.

341 3 Results

342 3.1 Dynamic functional connectivity (dFC) patterns during rest, continuous naturalistic movie 343 watching, and discrete sensorimotor task.

We computed the *dFC* from parcellated BOLD time series of resting state, naturalistic movie watching task where the participants watched and listened to an excerpt from Alfred Hitchcock's "Bang! You're Dead", and a sensorimotor task where participants responded by a button press to either a visual or an auditory stimulus from the Cam-CAN dataset (details in Methods). **Figure 2A** represents dFC obtained using BOLD phase coherence connectivity in resting state. We report the results of the analysis on young adults (age range 18-28) in this section.

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344

353 Dominant dFC subspaces were obtained by applying the unsupervised approach of Principal 354 Component Analysis (PCA) to BOLD time series at each time point, and then reconstructing 355 either the task or rest as the dynamics of a reduced dimensional dFC subspace. To demonstrate, 356 that the unsupervised characterization of dFC patterns indeed capture the functional brain 357 network organization, we computed the differences between the temporal stability matrices of 358 rest and the two task conditions; first using the measure of principal angles and second using 359 the measure of Mahalanobis distance. Thereafter, other measures of complexity and temporal 360 variability were tested.

361

362 3.1.1 Using angular distance to characterize temporal stability matrices

First, we calculate the principal angles among the dominant dFC subspaces generated across all time points. This resulted in time X time temporal stability matrix, averaged across all subjects, where each entry in the matrix is the angle between dominant dFC subspaces at t_x and t_y , as shown in **Figure 2BFigure 2**. We consider a dominant dFC configuration to be 367 stable if the subsequent subspaces are similar in configuration, i.e., less "angular distant" for 368 extended duration of time points. Results shown in Figure 2B indicate that the resting state has a global spread of shorter-lived, repeated patterns of stability than both tasks. On the 369 370 contrary, both the task cohorts, passive movie watching, and sensorimotor task, showed a local 371 spread of, longer-lived stability patterns suggesting that local temporal stability of functionally 372 connected networks are higher in the task than in resting state. To quantify these observations, 373 we calculated the entropy of temporal stability matrices of each category. The plots in Figure 374 4A, which represent entropy of temporal dynamic matrices of three categories, report resting 375 state to have the highest entropy, followed by movie watching task and sensorimotor task. The 376 distribution was parametric (normality check was done with Jarque-Bera test and verified with 377 D'Agostino-Pearson omnibus test), paired two-sample t-tests and effect size analysis using 378 Cohen's d, revealed significant differences (at 95% significance level) in entropy values 379 between resting state and movie watching task (p=0.0026 d=0.8), and resting state and 380 sensorimotor task (p=0.001, d=1.009). However, difference in movie watching task and 381 sensorimotor task were not significant (p=0.4907 ns). Further, to analyse how similar temporal 382 stability matrices across rest and tasks are, we calculate the Frobenius norm as shown in 383 Figure 4B. The results reveal a shorter Frobenius norm between the temporal dynamic 384 matrices of the resting state and movie watching task, than the resting state and sensorimotor 385 task.

386

387 *3.1.2* Using Mahalanobis distance to characterize temporal stability matrices

Alternatively, we evaluate the temporal stability of *dFC*, by estimating Mahalanobis distance, that resulted in a time X time temporal stability matrix. Each entry of this matrix is the Mahalanobis distance between dominant dFC subspaces (**Figure 3A**). Results, as shown in **Figure 3B** and **Figure 3C**, reveal global, shorter-lived repeated patterns of temporal stability 392 in resting state and local, longer-lived temporal stability patterns in both the tasks. The entropy 393 results (**Figure 4A**) reveal high entropy in the resting state, followed by movie watching task 394 and sensorimotor task. The distribution was non-parametric (normality check was done with 395 Jarque-Bera test and D'Agostino-Pearson omnibus test), we employed Wilcoxon matched 396 paired test to compute statistical significance between the entropy of temporal stability 397 matrices of each category, although the results did not reveal statistical significance, the trend 398 in entropy is similar to the trend in angular distance metric. We repeated the Frobenius norm 399 analysis, which produced similar results as the angular distance metric, as shown in Figure 400 **4B.**

401

402 **3.2 Unsupervised characterization of** *dFC* across healthy ageing

- 403 Next, we have included two cohorts, young and old adults from the Cam-CAN dataset and 404 carried out unsupervised characterisation of dFC using participant's resting state, movie 405 watching, and sensorimotor task data to identify age associated alterations in temporal 406 stability of dominant dFC subspaces.
- 407

408 3.2.1 Using principal angle to quantify temporal stability differences in dFC between young and 409 elderly

The time X time temporal stability matrix was computed for the aged cohort (age range 60-68) and compared with that of younger cohort computed in the section 3.1. A global spread of shorter duration of temporal stability patterns was observed in resting state and local spread of longer duration temporal stability patterns was observed in the task, in both young and old adults. Further, entropy analysis revealed (**Figure 4A**) a similar trend of peak entropy in resting state, followed by movie watching task and sensorimotor task in both young and old cohorts. The distribution was parametric (normality check was done with Jarque-Bera test and D'Agostino-Pearson omnibus test), paired two-sample t-test revealed significant differences in entropy values between resting state and movie watching task (p=0.000435, d=0.971), movie watching task, and sensorimotor task (p=0.0438, d=0.370), resting state and sensorimotor task (p=0.000567, d=1.319) of the older cohort (P values of young adults are reported in the previous section). The Frobenius norm analysis as shown in (**Figure 4B**) also revealed a similar trend in young and old adults i.e., shorter Frobenius norm between resting state and movie watching task than resting state and sensorimotor task

424 3.2.2 Using Mahalanobis to quantify temporal stability of dFC between young and elderly

425 Mahalanobis distance between dominant dFC subspaces showed patterns similar to principal 426 angle in young and elderly. Further, we calculate entropy as shown in **Figure 4A**, of temporal 427 stability matrices of each category, in both young and old adults. The results indicate peak 428 entropy in resting state, followed by movie watching task and sensorimotor task, a similar 429 trend as the angular distance metric. In the elderly, the distribution was non-parametric 430 (normality check was done with Jarque-Bera test and D'Agostino-Pearson omnibus test). 431 Wilcoxon matched paired test revealed statistical significance between the entropy of 432 temporal stability matrices of movie watching task and sensorimotor task (p=0.0074, 433 d=0.379). Frobenius norm analysis as shown in **Figure 4B** revealed a shorter Frobenius norm 434 between resting state and movie watching task than resting state and sensorimotor task.

The entropy analysis between young and elderly in resting state and tasks is shown in **Figure** 436 **4A(inset)**. The analysis indicates entropy of resting state in older adults was higher than their 437 younger counterparts, in both angular distance and Mahalanobis distance metric but statistical 438 tests (independent t-test for angular distance metric and Wilcoxon rank-sum test for 439 Mahalanobis distance metric) did not reveal any statistical significance.

441 **3.3 Stochastic characterization of** *dFC*

442 We examined the stochastic structure of dFC evolution by investigating the principal angle 443 $\phi(t)$ and Mahalanobis distance M(t) as functions of time. $\phi(t)$ and M(t) are modelled as 444 auto-regressive or AR (ρ) process. The optimal model order was taken to be at the value which 445 yields lowest Akaike information criterion (AIC). The results from this analysis shown in 446 447 **Figure 5A** and **Figure 5B** reveal the best fit model that explains $\phi(t)$ has a model order $\rho \ge 1$ 448 6 i.e., the results suggest $\phi(t)$ of resting state, movie watching task and sensorimotor task, in 449 both young and old adults, is neither random ($\rho \neq 0$) nor markovian ($\rho \neq 1$) in nature, and is 450 dependent on at least 6 immediately preceding values of ϕ . For M(t), as shown in Figure 5C 451 and **Figure 5D** both resting state and tasks have the optimum model order $\rho \ge 6$, suggesting 452 M(t) is neither random ($\rho \neq 0$) nor markovian ($\rho \neq 1$) in both young and old adults.

453 **4 Discussion**

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455 The functional architecture of the brain is dynamic and changes on a minute temporal scale during 456 resting state and task (Gonzalez-Castillo J., et al., 2015) (Hutchison & et al, 2013) (Gonzalez-Castillo 457 & Bandettini, 2018) (Bolton, Morgenroth, Preti, & Van De Ville, 2020). While previous studies have 458 explored flexibility (Zhang & et al, 2016) (Yin, et al., 2016) and temporal variability (Zhang & et al, 459 2016) (Li, Lu, & Yan, 2019) of the functional architecture of a specific region, we propose a novel 460 unsupervised method, that captures the stability of whole-brain functional architecture on a minute 461 temporal scale. First, we apply the data-driven unsupervised approach to characterize the high 462 dimensional dynamic functional connectivity into lower dimensional patterns by identifying 463 temporally similar dominant FC configurations. Subsequently, using two different measures -464 principal angle and Mahalanobis distance applied on dFCs extracted across time, we capture the 465 stability of dFC through the temporal stability matrices that could be used to draw critical insights 466 about underlying functional brain states. For empirical validation, we explored modifications in 467 temporal stability matrices of whole-brain FC during a continuous, naturalistic movie watching task 468 and discrete, goal oriented sensorimotor task and showed that, in contrast to resting state, stability 469 increased during the task (stability was highest in the sensorimotor task, followed by naturalistic 470 movie watching task and resting state). Next, we explored ageing specific modulations in temporal 471 stability matrices of dFC patterns between resting state and task and showed increased stability in the 472 task in both young and old adults. Finally, we examined the stochastic properties of temporal stability 473 matrices using an auto-regressive modelling, and showed dominant whole-brain FC configurations 474 are neither random nor Markovian. We discuss the implications of these key results in the following 475 subsections.

477 **4.1 Stochastic properties of dynamic functional connectivity**

478 Studies describing brain dynamics have clustered recurring connectivity patterns into states, using 479 clustering algorithms like K-means clustering (Allen, et al., 2014) (Cabral, et al., 2017) (Damaraju, 480 et al., 2014), HMM (Cabral, et al., 2017) (Vidaurre, Smith, & Woolrich, 2017) (Vidaurre, et al., 2016) 481 (Quinn, et al., 2018), suggestive of stability of functional architecture of the brain. Yet, most of the 482 studies hypothesize a fixed number of discrete recurrent connectivity patterns or states with varying 483 temporal fractional occupancy. The homogenous states are essentially clustered ignoring their 484 temporal order and index. Studies have shown clustering time series requires ignoring some data and 485 few attempts at clustering time series have shown to be objectively incorrect in some cases 486 (Rakthanmanon, Keogh, Lonardi, & Evans, 2011) (Rahman, Damaraju, Saha, Plis, & Calhoun, 2020). 487 Rahman and colleagues (Rahman, Damaraju, Saha, Plis, & Calhoun, 2020) have proposed a novel 488 framework, relying on the concept of shapelets, 'statelets'- a high dimensional state-shape 489 representation of temporal dynamics of functional connectivity, instead of clustering. Another set of 490 prior studies have explored the other side of stability – flexibility, which characterises heterogenous 491 connectivity between a specific region and others over time (Yin, et al., 2016) (Harlalka, Bapi, Vinod, 492 & Roy, 2019) and temporal variability (Zhang & et al, 2016) (Li, Lu, & Yan, 2019) of functional 493 architecture in resting state (Li, Lu, & Yan, 2019), naturalistic movie watching task (Li, Lu, & Yan, 494 2019) and in disease (Zhang & et al, 2016). But these studies are restricted to temporal variability and 495 flexibility of the functional architecture of a specific region. Our main contribution in this study is an 496 unsupervised, data-driven approach to characterise the stability of whole-brain functional 497 connectivity patterns. A recent study (Faghiri, et al., 2020) has proposed a new method, where they 498 calculate the gradients of timeseries pair and use their weighted average of shared trajectory (WAST) 499 as a new estimator of dFC. This method defines a subspace on the raw BOLD fMRI timeseries where 500 as our approach estimated dFC with BOLD phase coherence and defined dominant whole-brain FC 501 patterns as dominant dFC subspaces with PCA and characterised temporally similar dominant whole502 brain FC patterns with two alternative measures, angular distance and verifying the same with 503 Mahalanobis distance (**Figure 1B**). The central idea is if the dominant FC configurations are similar 504 for extended time points, then they are considered to be stable.

505 Viduarre and colleagues (Vidaurre, Smith, & Woolrich, 2017) have shown dynamic switching 506 between brain networks and time spent visiting distinct brain networks are not random. Subsequently, 507 another study has shown that the switching dynamics of functional brain states in the resting state 508 follows AR model of order 1, or in other words a Markovian process fully explains the dFC evolution 509 when correlation was computed using a sliding window approach (liégeois, Laumann, Snyder, Zhou, 510 & Yeo, 2017). By constructing the unsupervised temporal stability matrices from two alternative 511 approaches - principal angle, $\phi(t)$ and Mahalanobis distance, M(t), we reveal that dFC evolution is 512 neither random nor Markovian (Figure 5A and Figure 5B) (Figure 5C and Figure 5D).

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514 **4.2** Temporal stability of task related dynamic functional connectivity is higher than rest.

516 A key finding of our study indicates a global spread of shorter-lived, repeated patterns of stability 517 between dominant FC configurations in resting state and local spread of longer-lived repeated patterns 518 of stability in the task (in both continuous, naturalistic movie watching task and discrete goal oriented 519 sensorimotor task) (Figure 2B and Figure 3B). The resting state is shown to be a multistable 520 stationary state-regime at equilibrium (Deco & Jirsa, 2012). Ghosh and colleagues (Ghosh, Rho, 521 McIntosh, Kötter, & Jirsa, 2008) have demonstrated that resting state networks operate close to 522 instability and explore these states, before committing to one of these states. Deco and Jirsa (Deco 523 & Jirsa, 2012) have proposed that a repertoire of multistable states exists in resting state, that are 524 functionally meaningful and inherently supported by the neuroanatomical connectivity, and can be 525 rapidly activated even in the absence of any task. We speculate that in resting state the global spread 526 of shorter-lived repeated patterns of stability between dominant FC configurations is associated with 527 the exploration of multistable dynamic repertoire of states. On the contrary during a task (continuous 528 or discrete), the repertoire of multistable states are limited, as only task specific, cognitively relevant 529 brain networks are explored. The brain visits task specific stable states for duration that a putative 530 stimulus triggered cognitive process demands. This is associated with the local spread of longer-lived 531 temporal similarities between dominant functional connectivity subspaces in a task.

532 Our entropy results indicate the stability of functional connectivity architecture was highest in the 533 discrete, goal-oriented sensorimotor task, followed by continuous naturalistic movie watching task 534 and resting state (Figure 4A). This is in line with previous studies which report an increase in overall 535 stability of FC with the largest increase in between network connections (Elton & Gao, 2015) 536 (Gonzalez-Castillo & Bandettini, 2018), increase in stability of hemispheric homotopic connections 537 during a task (Gonzalez-Castillo J., Hoy, Handwerker, & Bandettini, 2014). Such increased stability 538 of FC during a task is hypothesised to be associated with cognitive constraints during a task 539 (Gonzalez-Castillo & Bandettini, 2018). Frobenius distance analysis results reveal the temporal 540 stability matrices of functional connectivity during continuous, naturalistic movie watching task was 541 closer to resting state than discrete, goal oriented sensorimotor task (Figure 4B). Considering our 542 Frobenius distance analysis, we hypothesized stability of functional connectivity architecture should 543 be highest in the sensorimotor task, followed by the naturalistic movie watching task, which was 544 validated by our entropy results. Our findings thus provide evidence of increased temporal stability 545 of whole-brain functional connectivity in task, highest in the discrete, goal-oriented task, followed by 546 continuous, naturalistic movie watching task and then resting-state, using a novel unsupervised 547 approach of characterising the stability of functional connectivity architecture.

548

549 4.3 Ageing introduces temporal variability in evolution of dynamic functional connectivity in 550 both rest and task

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552 Evidence from prior studies reveals the complexity of FC dynamics remains similar for all 553 participants irrespective of age. An earlier study (Viviano, Raz, Yuan, & Damoiseaux, 2017) found 554 no association between age and rate of switching between the FC states for resting brain. Our results 555 (Figure 2B and Figure 3B) indicate an overall trend of global spread of shorter-lived repeated 556 patterns of stability between dominant FC configurations in resting state and local spread of longer-557 lived repeated patterns of stability in the task was similar in both young and old adults. Our study also 558 revealed the highest stability of functional connectivity in the discrete, goal-oriented sensorimotor 559 task, followed by continuous, naturalistic movie watching task and resting state, a trend similar in both young and old adults (Figure 4A). Interestingly, McIntosh and colleagues (McIntosh, et al., 560 561 2010) have reported BOLD signal variability of hub-region decreases with age, suggestive of increase 562 in stability of hub regions with age. Our results, which contrasted the stability of functional 563 architecture in young and old adults (Figure 4A (inset)), found increased stability of functional 564 architecture in young adults in resting state. The neural noise hypothesis suggests the age-related 565 cognitive decline could be explained as a consequence of the increase in the noisy baseline activity of the brain (Voytek, et al., 2015) (Davis, et al., 2009). In accordance to this hypothesis, the decrease 566 567 in stability of the functional architecture of the brain in older adults can be explained with an increase 568 in neural noise with age. An important point to note, regardless of age associated changes in the 569 stability of functional architecture, our results did not reveal statistically significant differences. 570 Therefore, although there are differences in stability of functional architecture with age, their 571 magnitude may be modest.

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573 **3.4 Limitations and Future directions**

An important caveat of the current study was due to parcellation atlas used in the Cam-CAN dataset. The AAL atlas parcellates the brain regions into 116 structural parcels and few parcels span multiple functional regions. For future studies, for a more refined spatial profile of temporal stability of functional architecture, using a finely parcellated brain atlas is recommended. Researchers have shown stability of functional architecture is modified in patients of Schizophrenia, ADHD and ASD (Zhang & et al, 2016) (Guo, Zhao, Tao, Liu, & Palaniyappan, 2017). Hence, we can extrapolate that the temporal stability of functional architecture can provide a richer information to discoverbiomarkers for neurological and mental disorders.

3.5 Conclusion

584

585 In summary, the current study introduces a data-driven unsupervised approach to characterise the 586 temporal stability of functional architecture. When applied to a putative lifespan ageing data, the 587 whole-brain temporal dynamics of naturalistic movie watching task was found to be closer to resting 588 state than during sensorimotor task. Further, the study revealed peak temporal stability in 589 sensorimotor task, followed by naturalistic movie watching task and resting state, a trend similar in 590 both young and elderly. The temporal stability of functional architecture of the resting state was also 591 found to be higher in young adults than their older counterparts. The quantification of differences in 592 network stability associated with healthy ageing provides evidence for the potency of the temporal 593 stability measure to act as biomarker for multiple neurological disorders.

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610 **Declaration of competing interest**

- 611 The authors declare no conflicts of interest
- 612

613 **Ethics statement**

- 614 CamCAN dataset was collected in compliance with the Helsinki Declaration, and has been approved
- 615 by the local ethics committee, Cambridgeshire 2 Research Ethics Committee (reference:
- 616 10/H0308/50)
- 617

618 **5 References**

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849 **Figure captions:**

Figure 1: (**A**). The schematic diagram shows how the temporal stability of dynamic functional connectivity subspaces (dFC) are computed. Dominant dFC subspace, at each time point, is estimated using the first three principal components of dFC(t), that was computed using the measure of BOLD phase coherence. The similarity between dFC subspaces are calculated using principal angle (Angular distance) and Mahalanobis distance (Euclidean distance). If the dominant dFC subspaces are similar for extended timepoints, then they are considered to be stable. (**B**). A flowchart representation of the method

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858 Figure 2: (A) dFC matrices estimated using BOLD phase coherence.

(B) Time X Time temporal stability matrix of resting state, naturalistic movie watching task and discrete, sensorimotor task for young and old adults. Each entry in the matrix is the principal angle $\phi(t_x, t_y)$ between dominant dFC subspaces at t_x and t_y . The principal angle ranges between 0 (low angular distance) to $\pi/2$ (high angular distance). Resting state, in both young and old adults, has shorter-lived, global spread of patterns of temporal stability. On the contrary, both the tasks have a longer-lived, local spread of patterns of stability (indicated by arrows and rectangular boxes).

865

Figure 3: Mahalanobis distance between dominant dFC subspaces.

867 (A) Mahalanobis distance is a pairwise Euclidean distance between the ROIs of dominant dFC 868 subspace at t_x , with the whole-brain dominant dFC subspace at t_y .

(B) Time X Time temporal stability matrix of resting state, naturalistic movie watching task, and sensorimotor task, where each entry in the matrix is Mahalanobis $(M^2(t_x, t_y))$ distance between the dominant dFC subspaces. Mahalanobis distance between dominant dFC subspaces is low when the dFC configurations are similar. 873 (C) The profile of temporal stability estimated with Mahalanobis distance across the brain regions at874 different instances of time.

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Figure 4: (A) Plots representing entropy of temporal stability matrices of resting state (rest), naturalistic movie watching task (movie) and sensorimotor task (SMT), for Angular distance and Mahalanobis distance metric, in both young (magenta) and old adults (blue). Statistically significant differences are indicated using * ($\mathcal{P} \le 0.05$), ** ($\mathcal{P} \le 0.01$), *** ($\mathcal{P} \le 0.01$), *** ($\mathcal{P} \le 0.001$), **** ($\mathcal{P} \le 0.0001$), ns (not significant).

(inset) Distribution of entropy computed from temporal stability matrices of the resting state,
naturalistic movie watching, and sensorimotor task, each contrasted between young (magenta) and
old adults (blue) represented as a violin plot

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(B) Plots representing distribution of Frobenius distance between temporal stability matrices of
resting state, naturalistic movie watching (yellow) and resting state, sensorimotor task (pink) for
Angular distance, and Mahalanobis distance metric, in both young and old adults. The violin plots
reveal a shorter Frobenius norm between resting state and movie watching task than resting state and
sensorimotor task in both young and old adults.

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Figure 5: (A) Stochastic modelling of principal angle, ϕ (*t*) as autoregressive, AR (ρ) process. The model order (ρ) was varied from 0 to 100. The plot represents Akaike information criterion (AIC) values corresponding to the model order. Inset shows the first minima of the AIC value and its corresponding model order.

895 (B) Table shows first minimal AIC value and its corresponding model order of $\phi(t)$ for all the 896 categories

(C) Stochastic modelling of Mahalanobis distance, M (t) as AR (ρ) process. The model order (ρ)
was varied from 0 to 100. The plot represents AIC values corresponding to the model order. Inset
shows the first minima of the AIC value and its corresponding model order.
(D) Table shows first minimal AIC value and its corresponding model order of M (t) for all the
categories

905 Supplementary Figures:

906

907 **S** 1: Temporal stability matrices of the resting state, naturalistic movie watching task and 908 sensorimotor task, where each entry is the principal angle $\phi(t_x, t_y)$ between dominant dFC subspaces 909 at t_x and t_y , for young and old adults. For validation of the results where dFC was estimated using 910 BOLD phase coherence, we calculated dFC using sliding window approach with (window length) 911 WL = 10 time points.

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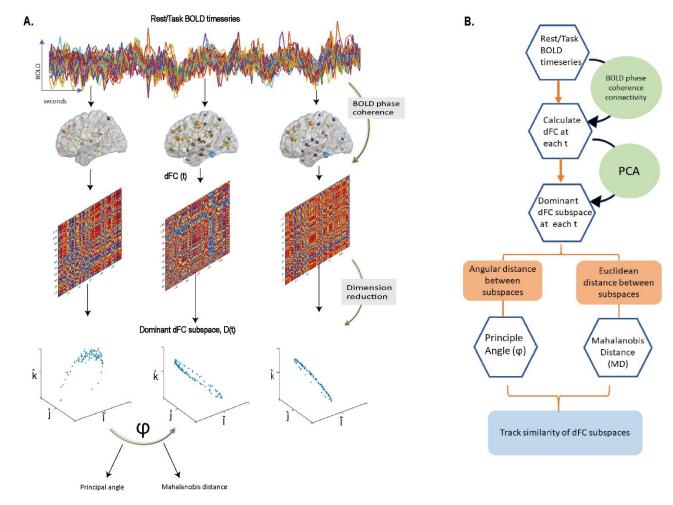
913 S 2: Temporal stability matrices of the resting state, naturalistic movie watching task, and
914 sensorimotor task, for both young and old adults. dFC was estimated using sliding window approach
915 with (window length) WL= 20 time points.

916

917 S 3: Temporal stability matrices of the resting state, naturalistic movie watching task, and
918 sensorimotor task, for both young and old adults. dFC was estimated using sliding window approach
919 with (window length) WL= 30 time points

- 921 **S** 4: The Plot represents the variance explained by all 116 principal components of the input dFC
- matrix for all categories. The first three principal components explain almost 99% of the variance of
- 923 the input matrix.
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- 925 S 5: Temporal stability matrices representing the temporal landscape of randomised BOLD signals,
- 926 resting state, naturalistic movie watching task, and sensorimotor task.

Figure 1



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Figure 2

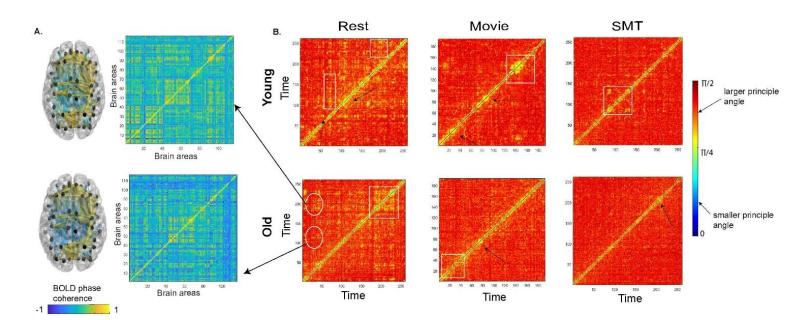


Figure 3

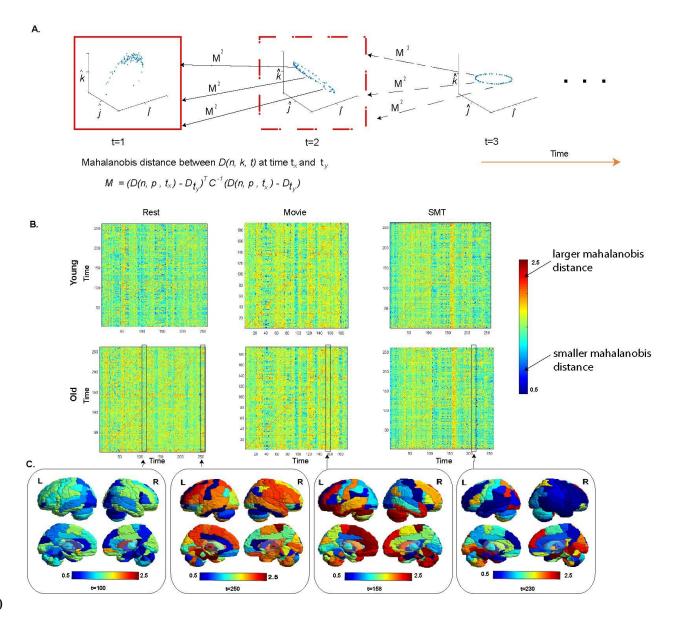
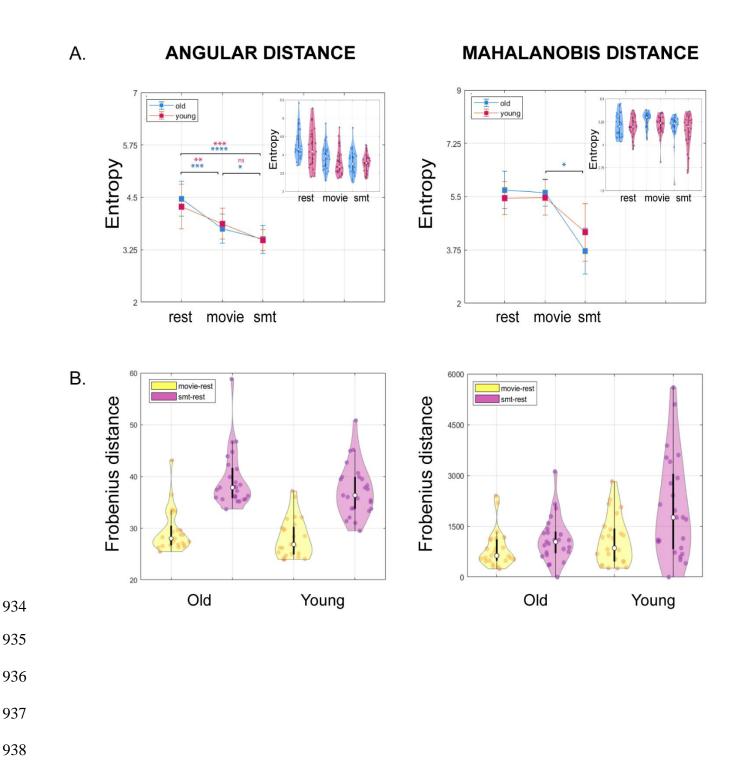
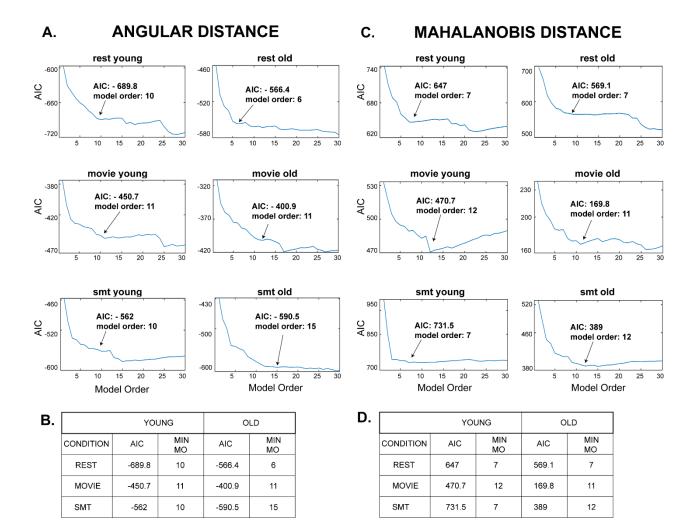


Figure 4



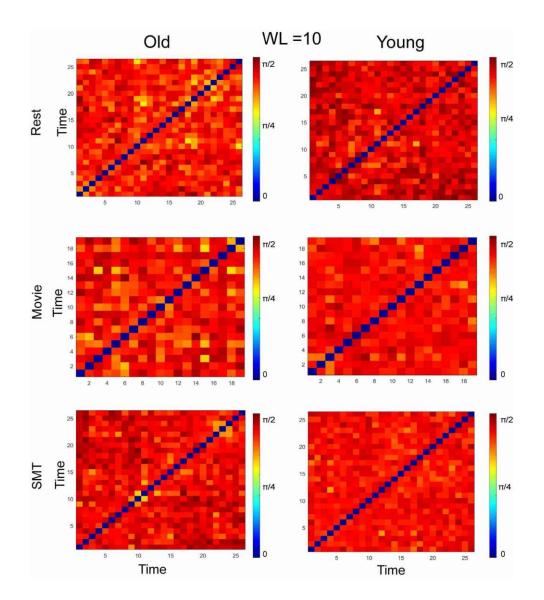
939 Figure 5

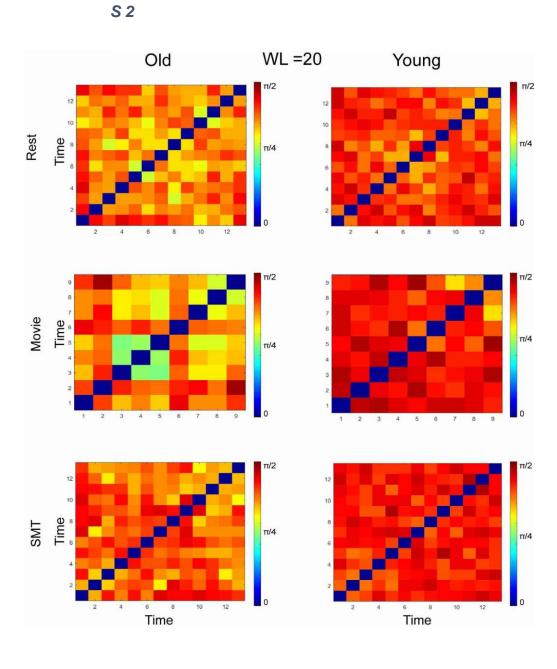


940941 Supplementary figures

S 1

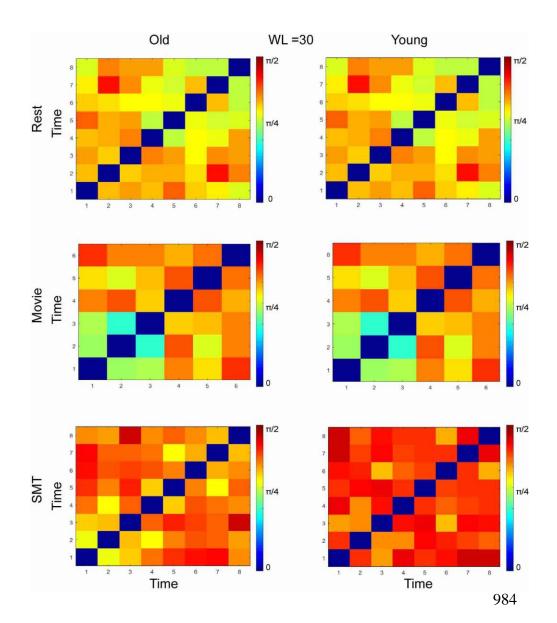
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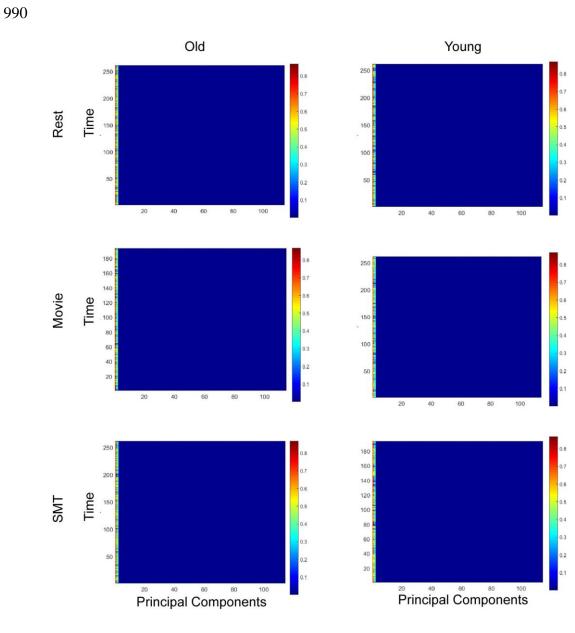


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