

1 **Phylodynamics of H5N1 Highly Pathogenic Avian Influenza in** 2 **Europe, 2005–2010: Potential for molecular surveillance of new** 3 **outbreaks**

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14 **Abstract:** Previous Bayesian phylogeographic studies of H5N1 highly pathogenic avian
15 influenza viruses (HPAIVs) explored the origin and spread of the epidemic from China in-
16 to Russia, indicating that HPAIV circulated in Russia prior to its detection there in 2005. In
17 this study, we extend this research to explore the evolution and spread of HPAIV within
18 Europe during the 2005–2010 epidemic, using all available sequences of the HA and NA
19 gene regions that were collected in Europe and Russia during the outbreak. We use dis-
20 crete-trait phylodynamic models within a Bayesian statistical framework to explore the
21 evolution of HPAIV. Our results indicate that the genetic diversity and effective population
22 size of HPAIV peaked between mid-2005 and early 2006, followed by drastic decline in
23 2007, which coincides with the end of the epidemic in Europe. Our results also suggest that
24 domestic birds were the most likely source of the spread of the virus from Russia into Eu-
25 rope. Additionally, estimates of viral dispersal routes indicate that Russia, Romania, and
26 Germany were key epicenters of these outbreaks. Our study quantifies the dynamics of a
27 major European HPAIV pandemic and substantiates the ability of phylodynamic models to
28 improve molecular surveillance of novel AIVs.

29 **Keywords:** H5N1; Highly Pathogenic Avian Influenza; phylodynamic models; Bayesian
30 inference; phylogeography; Europe; Russia; Surveillance

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33 1. Introduction

34 The epidemiology of the H5N1 highly pathogenic avian influenza viruses (HPAIVs) is character-
35 ized both by the far-reaching economic impact for commercial poultry production, and also by the risk
36 of spread into mammalian species, including humans. Transmission into humans has mainly been at-
37 tributed to direct contact of susceptible individuals with infected poultry, including activities associat-
38 ed with the slaughter and dressing of poultry for human consumption [1]. The maintenance and spread
39 of avian influenza (AI) infection in domesticated poultry are mainly associated with the establishment
40 of the virus in free-range (or backyard) poultry and its subsequent introduction to commercial animals
41 via live-bird markets. Furthermore, wild birds are the natural reservoirs of the virus, and the movement
42 of wild birds infected with HPAIV is known to play an important role in the spread, circulation, and
43 maintenance of the virus through either direct or indirect contact with poultry [2–4].

44 The ancestral strain of H5N1 HPAIV, referred to as A/goose/Guangdong/1/96, emerged in com-
45 mercial domesticated geese in the Guangdong province of China in 1996 [5]. Since then, the strain has
46 undergone numerous genetic changes that have given rise to several distinct lineages, also referred to
47 as clades [6]. To date, ten H5N1 HPAIV clades have been identified based on comparisons of 859 he-
48 magglutinin (HA) gene sequences [7]. These viruses are now widespread in countries throughout Asia,
49 Europe, and Africa. Between 2004 and 2005, a new H5N1 HPAIV strain (sub-clade 2.2) emerged that
50 caused severe outbreaks in wild birds and poultry in northwestern China [6, 8]. In Europe, HPAI H5N1
51 was first detected at the end of 2005 in both poultry (in Russia, Romania, and Turkey) and wild birds
52 (in Croatia) [2, 9–11]. In January 2006, the first human infection outside Southeast Asia was reported
53 in Turkey. In February 2006, the first occurrence of HPAI H5N1 in Northern Europe was identified
54 from a wild mute swan (*Cygnus olor*) in Germany in the southwestern part of the Baltic Sea [11,12].
55 Between 2006 and 2010, the epidemic spread through 13 European Union (EU) member states located
56 eastward of a line extending from south-eastern Sweden to south-western Italy. Both poultry and wild-
57 bird populations were infected in Austria, Bulgaria, Czech Republic, Denmark, France, Greece, Hun-
58 gary, Poland, Slovak Republic, Slovenia, Switzerland, and the UK [10, 13]. Phylogenetic analysis of
59 HPAIVs isolated from European outbreaks distinguished at least two different sub-lineages of the 2005
60 Qinghai strain, which confirmed that the virus spread from southeastern Asia into Europe [10-13].

61 Since 2005, surveillance of avian influenza viruses in poultry and wild birds has become compulso-
62 ry in the EU. Member states have conducted extensive molecular surveillance for AI, with most of the
63 studies focused on genetic analysis of partially or fully sequenced AI viral genomes [14–18]. Molecu-
64 lar epidemiological studies of H5N1 HPAIV sequence data isolated between 2005–2010 in EU mem-
65 ber states, for example, were focused on the molecular characterization, genetic classification, identifi-
66 cation of genetic shifts and drifts, and the degree of relatedness to previously isolated viruses. Phylog-
67 enies of H5N1 HPAIV isolated in the EU were estimated using traditional parsimony, neighbor-
68 joining, and maximum-likelihood methods. However, those methods typically ignore various sources
69 of uncertainty, including uncertainty associated with estimates of the phylogenetic relationships, diver-
70 gence times, and history of geographic dispersal [19]. Furthermore, epidemiological studies that ex-
71 plored risk factors and the spatial and temporal evolution of H5N1 HPAIV in the EU between 2005–
72 2010 were mainly conducted separately from phylogenetic studies [20–24]. Phylodynamics is an
73 emerging field that aims to characterize the joint evolutionary and epidemical behavior of rapidly

74 evolving infectious diseases using tools borrowed from the field of phylogenetics [25]. The probabilis-
75 tic models used in the phylodynamic approach are mostly pursued in a Bayesian statistical framework
76 [26]. This approach treats parameters of the phylodynamic model as random variables, such that each
77 parameter is described by a specified prior probability distribution (and a corresponding inferred poste-
78 rior probability distribution). Accordingly, the Bayesian approach provides a natural way to estimate
79 (and accommodate) uncertainty in the phylodynamic model parameters, including the virus phylogeny,
80 divergence times, and history of geographic spread [27].

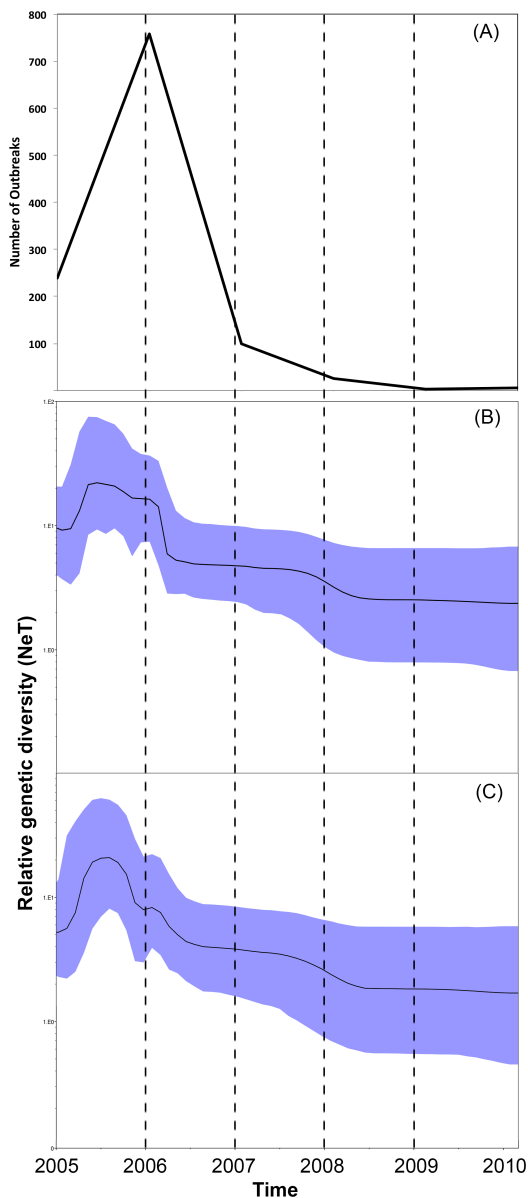
81 Recent evidence suggests that Bayesian methods may be more accurate in assessing HPAIV evolu-
82 tion than traditional phylogenetic methods. For example, Lemey *et al.* (2009) [27] adopted a Bayesian
83 phylodynamic approach to re-analyze a H5N1 HPAIV dataset initially studied using conventional
84 methods [28, 29]. Results of the Bayesian analysis did not support the earlier conclusion regarding the
85 epidemiological link between Guangdong and Indonesia, but instead suggested that the ancestral
86 strains of the Indonesian outbreaks originated from the Hunan province in south central China [27].
87 Furthermore, these studies demonstrated the potential of Bayesian phylogeographic methods to esti-
88 mate the posterior probability of each geographic location at any point in the phylogenetic tree, and the
89 ability for phylogenetic parameters to be estimated from different genomic segments of the H5N1
90 HPAIV genome without assuming that the sequences shared an identical phylogenetic history [27].

91 Here, we extend the landmark study of Lemey *et al.* (2009) to estimate the history of HPAIV dis-
92 persal into and within Europe after its initial detection in Russia in 2005. Our data comprise 277 H5N1
93 HPAIV hemagglutinin (HA) and neuraminidase (NA) gene sequences (referred to as isolates) collected
94 between 2005 and 2010 in Europe and Russia. We adopt a discrete-trait phylodynamic model to esti-
95 mate both the history of viral migration between geographic areas, and also to infer the movement of
96 the virus among host species (wild versus domesticated birds). We also estimate temporal changes in
97 viral population size, and discuss the potential of phylodynamic methods to improve AI surveillance.
98 Our study provides quantitative estimates of the mechanisms that led to the rapid spread of HPAIV
99 throughout Europe in 2005–2010 and further illustrates the potential of this approach to improve the
100 prevention and control of novel emerging AIVs.

101 **2. Results and Discussion**

102 Our analyses of the viral population dynamics for both HA and NA genes reveal a strong increase in
103 genetic diversity (and, therefore, effective population size) in mid-2005 followed by a second increase
104 in early 2006 (Figure 1B and C). This coincides with the highest frequency of reported cases of H5N1
105 infections during the epidemic, when several independent introductions of H5N1 viruses belonging to
106 sub-clade 2.2 have been confirmed (Figure 1A)[15, 16, 30, 31]. The relatively high viral genetic diver-
107 sity during the early stage of the epidemic is consistent with the introduction of the virus into a novel
108 region, with a naïve population, where it was exposed to selection pressure imposed by environmental
109 and demographic factors [10, 23, 31]. We also inferred a third (and relatively minor) change in viral
110 genetic diversity during mid-2007, which coincides with the entry of the virus into new EU countries.
111 From 2005 to 2007 conditions seem to have been relatively favorable for the virus to perpetuate in the
112 susceptible population and, given the genetic diversity, reassortment may have occurred during this
113 time frame. The viral genetic diversity is then inferred to decline toward 2007 and to stabilize until

114 2010. The inferred decrease of viral genetic diversity coincides with the decline of the epidemic, sug-
115 gesting that prevailing epidemiological factors, such as host density or weather conditions, in 2007–
116 2010 were not sufficient to favor the establishment of the newly emerging H5N1 strain in the popula-
117 tion. In contrast to typical AIV epidemics (which exhibit seasonal variation in viral genetic diversity)
118 the genetic diversity of the H5N1 virus did not exhibit seasonal variation in genetic diversity during
119 the 2005–2010 period of the epidemic in Europe and Russia, which suggests that transmission of the
120 virus was influenced by factors different from those typically observed in AIV epidemics.



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123 **Figure 1.** (A) Temporal distribution of H5N1 HPAI outbreaks (per year) in domesticated
124 poultry and wild birds in Europe and Russia from between 2005 and 2010. (B) and (C)
125 Bayesian Skyline plots (BSP) illustrating temporal changes in the relative genetic diversity
126 of H5N1 HPAIV isolates from outbreaks in Europe and the Russia between May, 2005 and
127 June, 2010 estimated from the HA and NA gene sequences. Line plots summarize esti-
128 mates of the effective population size (N_eT), a measure of genetic diversity, for HA (above)
129 and NA (below) gene regions; the shaded regions correspond to the 95% HPD.

130 Our analyses indicate that domestic birds were the most probable ancestral host type for both genes.
131 Domestic birds are inferred to be the most probable host type along many branches of the HA and NA
132 phylogenies, suggesting that the evolution of the H5N1 HPAIV virus largely occurred in association
133 with domestic-bird populations. By contrast, the large proportion of wild-bird isolates (Figure 2) and
134 the inferred high transition rates between host types (Table 1) suggest that wild birds also played an
135 important role in viral dispersal between host populations and geographic areas. This finding is con-
136 sistent with a number of alternative hypotheses, including scenarios where (1) domestic birds were the
137 source of the spread of H5N1 HPAIV into Europe, or (2) where the HPAI strains were first transmitted
138 from non-EU wild birds to non-EU domestic birds (with little or no genetic change), and later transmit-
139 ted from non-EU domestic birds to EU domestic birds (Figure 2). The first scenario may be explained
140 by the illegal introduction of domestic birds into Europe, which has been implicated in the spread of
141 other diseases within some European countries [34]. The second scenario assumes very low selective
142 pressure on the virus within—and rapid transmission of the virus among—wild-bird populations,
143 which would explain the high mortality in wild-bird populations and the reported prevalence of the
144 virus along migratory routes [33]. These two seemingly contradictory scenarios may be reconciled as
145 follows: H5N1 HPAIV was initially transmitted from non-EU wild birds to domestic poultry as a less
146 virulent and non-reassortant strain of the virus. Subsequently, the virulent strain of H5N1 HPAIV that
147 caused the EU epidemic emerged as a reassortant strain from poultry, as suggested elsewhere [32]. The
148 high density of poultry populations may have resulted in elevated contact and transmission rates, ulti-
149 mately enhancing opportunities for viral evolution in poultry compared to wild birds. Additionally, the
150 three sequences isolated from non-avian species were obtained from hosts that prey on birds (two from
151 cats and one from a marten), which is both consistent with a wild-bird origin of infection (Table 1) and
152 with the results of previous studies [11, 35].

153 Our divergence-time estimates (Figure 2) suggest that H5N1 HPAIV originated and experienced re-
154 assortment among hosts between 1995–1996 (Figure S1). These results are consistent with the frequent
155 reports of highly reassorted H5N1 HPAIVs introductions into eastern and southeastern Asia since the
156 emergence of the ancestral strain in 1996 (A/goose/ Guangdong /1/96) [5, 36].

157 **Table 1.** Bayes factor (BF) tests for non-zero transition rates between host types for each
158 gene region (HA, NA). BF values > 6 indicate significant rates of exchange between host
159 types.

Gene	BF	From	To
HA	11049.7	Wild	Other
HA	11049.7	Other	Wild
HA	171.4	Domestic	Wild
NA	11049.7	Wild	Other
NA	11049.7	Other	Wild
NA	183.0	Domestic	Wild

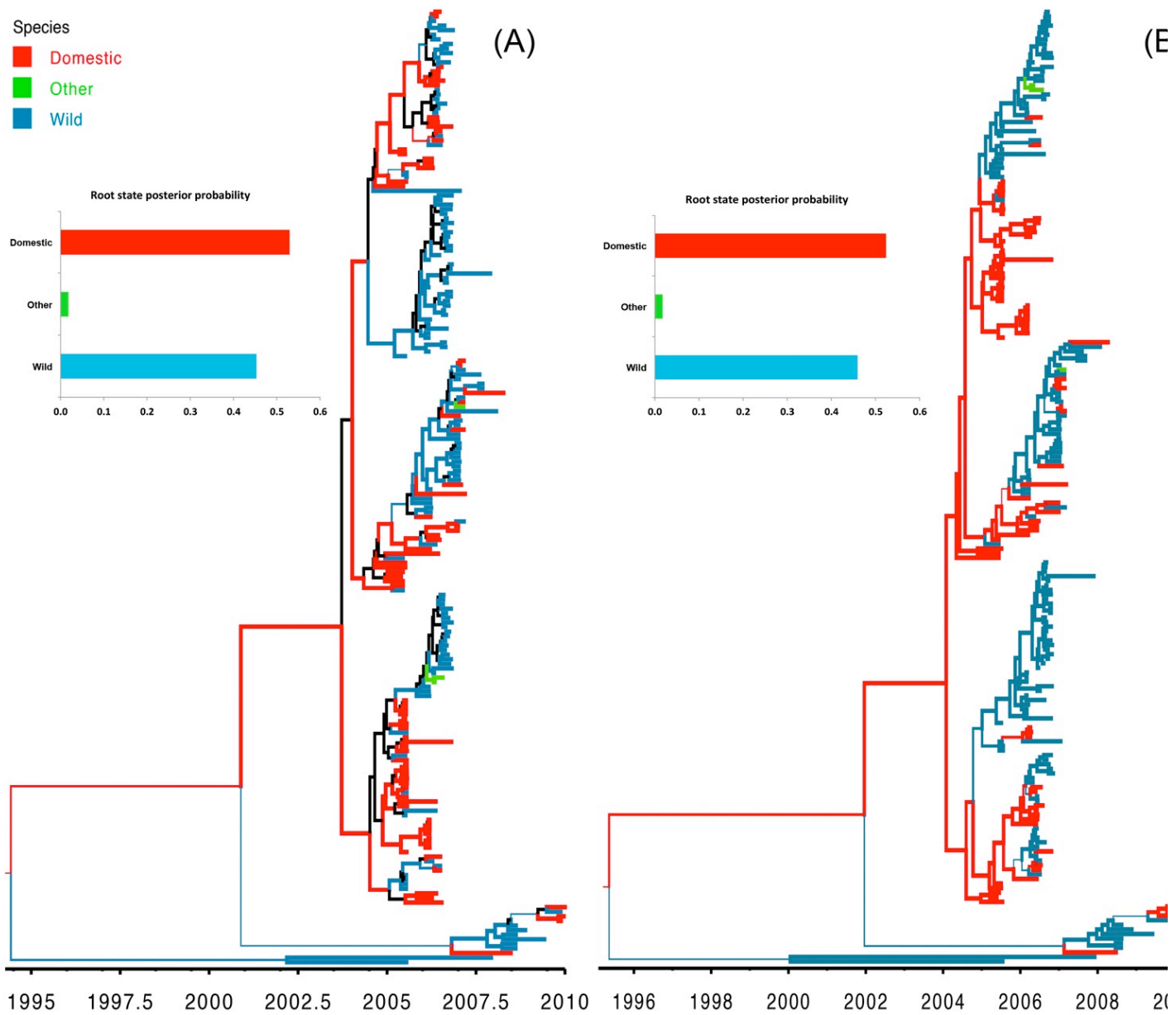


Figure 2. Maximum clade credibility (MCC) trees for H5N1 HPAIV hemagglutinin (HA) and neuraminidase (NA) gene regions (A and B, respectively). Branch lengths are rendered proportional to absolute time (see timescales), the branches are colored according to the most probable host type (wild birds, domestic birds, or other), and the branch thickness is proportional to the posterior probability of the inferred host state. Branches where the host state is uncertain (where the posterior probability of any hosts <0.5) are colored black. The posterior probabilities for the ancestral host states are shown in the upper left panel for each tree.

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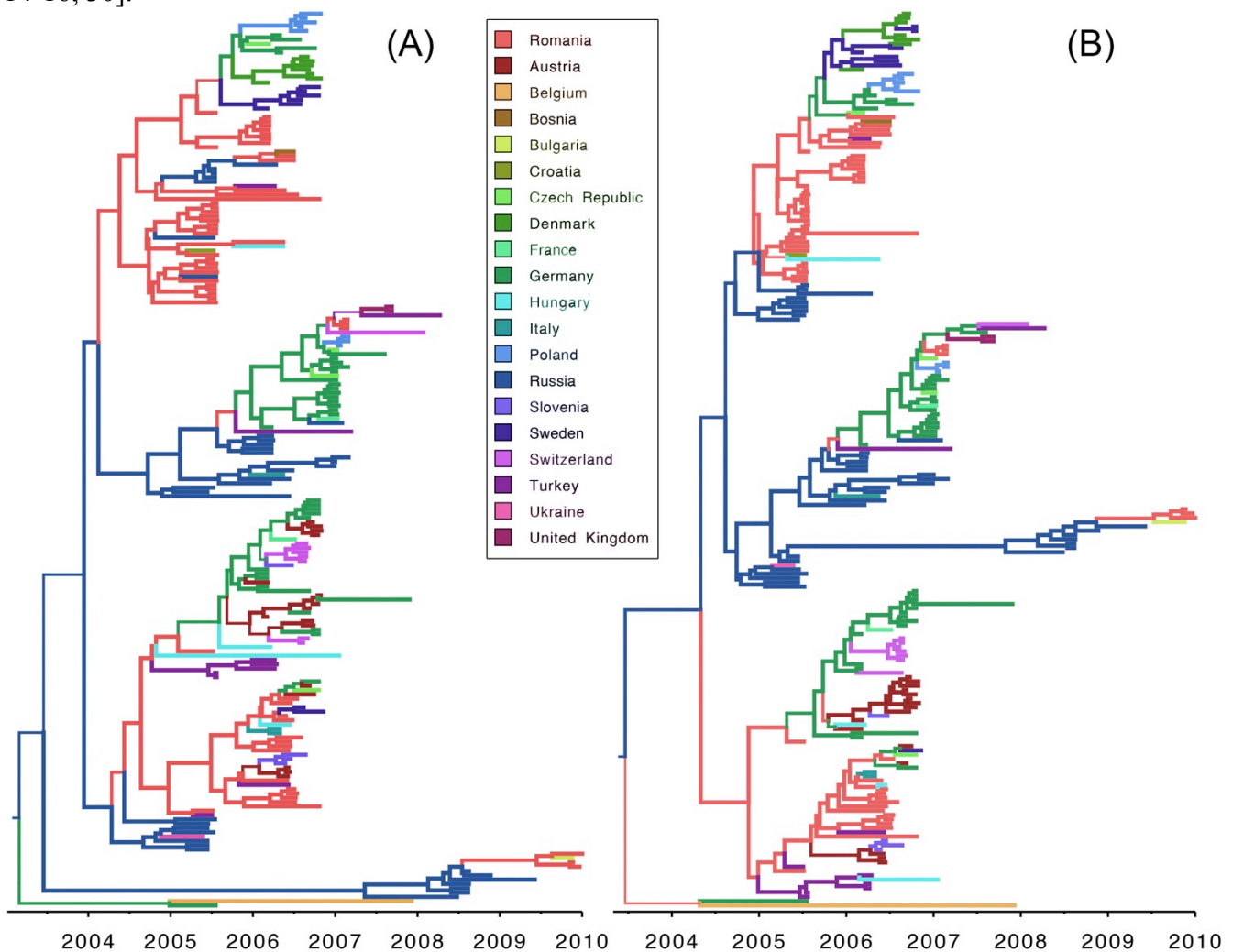
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170 Results of our discrete phylodynamic analyses indicate that Russia, Romania, and Germany were
171 important epicenters of the H5N1 HPAIV epidemic in Europe. Specifically, our results suggest that the
172 virus first originated and accumulated in Russia, with subsequent significant dispersal rates out of Rus-
173 sia into Romania (Figures 3–5, S2). Later, significant dispersal rates occurred between Romania and
174 Germany. In fact, exchange between Russia-Romania and Romania-Germany represent the most sig-
175 nificant dispersal rates of the virus during the course of the epidemic (Figure 5, Table S3). Finally, in
176 late 2005 and early 2006 the virus spread independently from Romania and Germany, respectively into
177 other areas of the EU (Figure S3); a scenario that is consistent with the results of previous studies [1,
178 14-16, 30].



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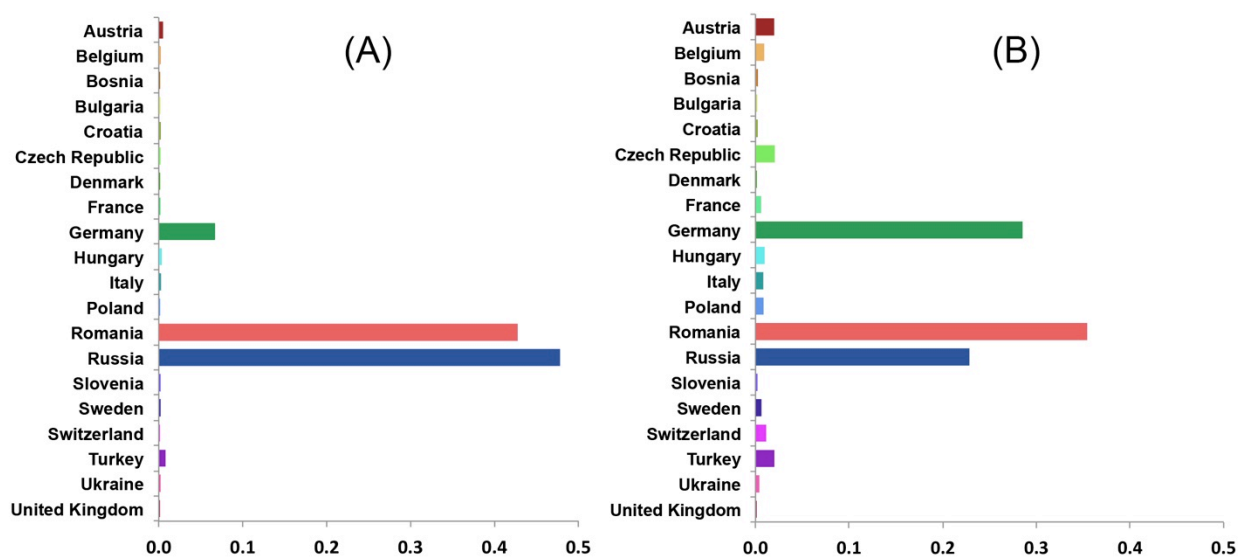
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Figure 3. Maximum clade credibility (MCC) tree for H5N1 HPAIV hemagglutinin (HA) and neuraminidase (NA) gene regions. Branch lengths are rendered proportional to absolute time (see timescales), the branches are colored according to the most probable geographic location, and the branch thickness is proportional to the posterior probability of the inferred area (see inset legends). Color gradients along branches reflect the inferred changes in the posterior probability of the corresponding areas.

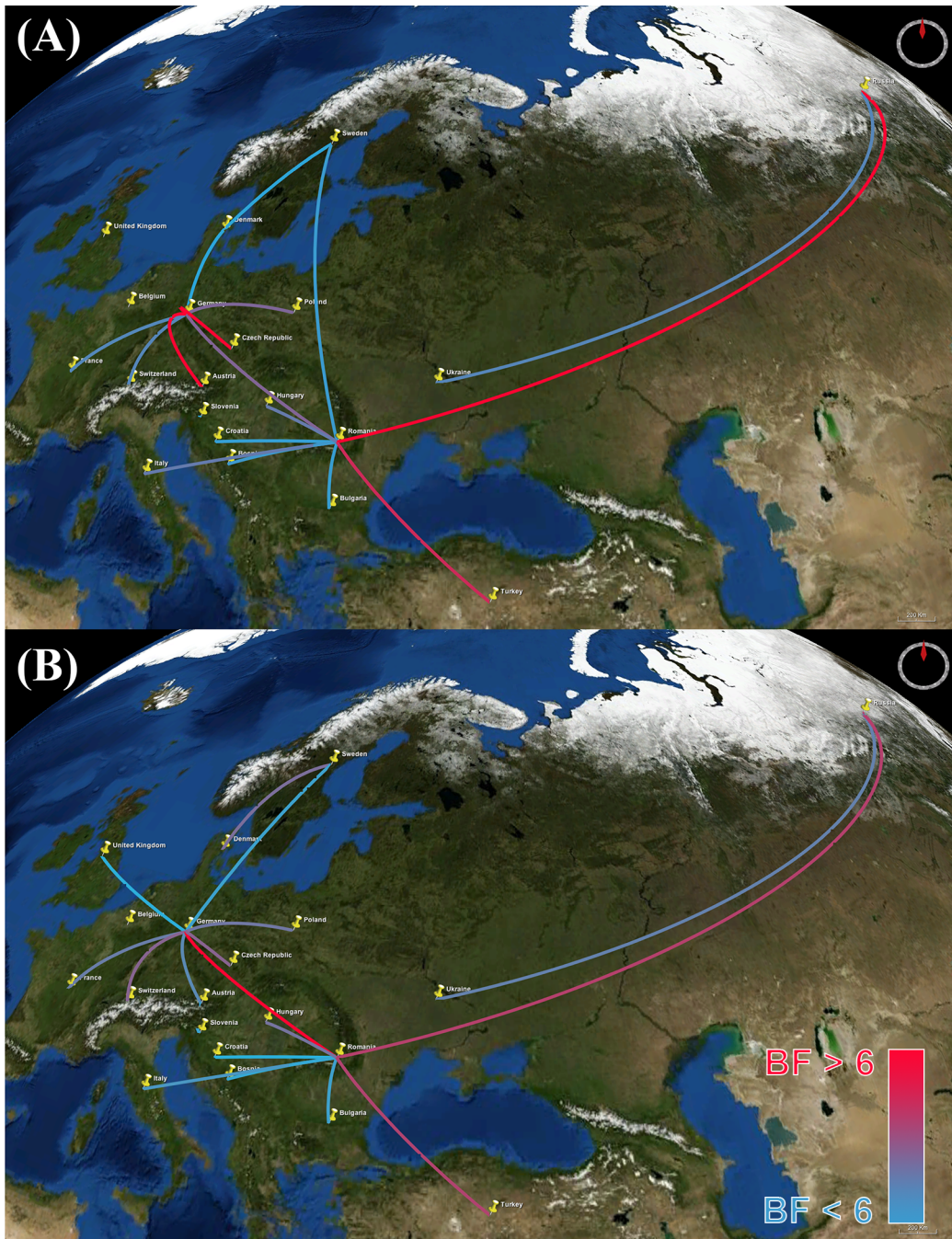


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Figure 4. Posterior probabilities of ancestral areas of H5N1 HPAIV hemagglutinin (HA; left) and neuraminidase (NA; right) gene regions collected in Europe and the Russia.



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191 **Figure 5.** Bayes factor (BF) test for significant non-zero dispersal rates in H5N1 HPAIV.
192 Only rates supported by a BF greater than six are indicated. The color gradient of lines cor-
193 respond to the probability of the inferred dispersal routes; blue lines and red lines indicate
194 relatively weak and strong support, respectively. The maps are based on satellite pictures
195 available in Google Earth (<http://earth.google.com>). A: HA gene. B: NA gene. The maps
196 are based on satellite images sourced from the NASA World Wind Java SDK
197 (<http://worldwind.arc.nasa.gov/java/>).

198 The inferred geographic expansion of the virus is consistent with our estimates of the viral effective
199 population size; the geographic expanse and the genetic diversity of the virus increased between 2005
200 and 2006, and subsequently declined between 2007 and 2010 (Figure 1 and S3). The geographic ex-
201 pansion of the virus entailed two major waves. The first wave began in early 2005 from Russia, inten-
202 sified in Romania, and resulted in the spread of the virus into Germany, Ukraine, and Turkey. The se-

203 cond wave began in early 2006, also commencing in Russia, intensified in both Romania and Germa-
204 ny, and resulted in the spread of the virus into Switzerland, Czech Republic, Denmark, and Sweden
205 [14, 15, 17, 30, 37]. Additionally, a minor wave in 2007 emerged and intensified within Germany, and
206 resulted in the spread of the virus into France, Hungary, and Poland [15].

207 The results of our phylodynamic analyses generally appear to be robust (Table 2). The significance
208 (p-value <0.001) of the observed association-index (AI) values and their 95% credible intervals for
209 both genes indicate a strong relationship between the inferred phylogeny of the H5N1 virus and the
210 discrete traits (geographic area and host type), and also indicate that both host type and geography
211 played important roles in the transmission of the virus. By contrast, the Kullback-Leibler (KL) values
212 for the discrete host-type models were fairly small for both genes (suggesting relatively poor model
213 fit), whereas the KL values for the geographic models were large, indicating good fit the between
214 model and the geographic data [27].

215 **Table 2.** Kullback-Leibler divergence and Association Index statistics assessing the fit of
216 the viral data to the discrete phylodynamic models.

Tree	Kullback-Leibler	Association Index
<i>Host Type</i>		
HA	0.71	6.01 (4.97,7.06)*
NA	0.72	7.22 (5.97,8.44)*
<i>Location (by Countries)</i>		
HA	3.24	9.1(7.87,10.3)*
NA	1.68	10.45(9.29,11.58)*

217 * Statistically significant (p -value < 0.001).

218 Although generally robust, the results of this study are based on analyses that necessarily entailed
219 several compromises, including: (1) imprecise geographic information on the isolation of H5N1 se-
220 quences; and; (2) incomplete and possibly biased sampling of H5N1 isolates. First, precise lati-
221 tude/longitude coordinates were unavailable for 65.8% of the sequences used in this study; instead
222 only the country of isolation was recorded. In these cases, we used the centroid of the country as a
223 proxy for the location, which is apt to depart substantially from the true location in most cases, which
224 may bias our inferences of the phylogeographic history of the H5N1 virus. This sampling issue pre-
225 cludes our use of the continuous phylodynamic model described by Lemey et al [38], as these models
226 require more precise information on the locations of the viral samples. Unfortunately, this information
227 is typically not recorded for publically available viral sequence data in GenBank or disease databases.
228 The greatest obstacle for applying continuous phylodynamic models to the H5N1 system is therefore
229 mainly an issue of data confidentiality. The impact of incomplete spatial metadata on the performance
230 of phylogeographic models has been discussed elsewhere [39, 40]. Inferences under the phylodynamic
231 models—like all statistical inference methods—assume that we have either a complete or random
232 sample of sequence data. In the present case, this requires that the H5N1 sequences were collected
233 randomly with respect to time (between ~1990 and 2010), geographic location, and host type. Myriad
234 practical considerations almost certainly result in very strongly biased samples, and the impact of these
235 departures from random sampling on the estimates are difficult to quantify. Although certainly incom-
236 plete, and almost certainly non-random, our study is based on all available sequence data for the HA

237 and NA gene regions associated with the H5N1 epidemic in the EU during 2005–2010, and therefore
238 reflects our best understanding based of the available data.

239 Phylodynamic analysis has not yet been widely embraced as a resource by public health agencies to
240 support the design of disease-surveillance strategies. For example, most of the published literature on
241 avian influenza—including disease summary reports and peer-reviewed scientific studies—upon which
242 global disease agencies base their decision-making process primarily include numerical summaries of
243 reported cases in each country, risk maps based purely on spatial and temporal distribution of the dis-
244 ease, and conventional phylogenetic analysis of sequenced isolates from these outbreaks [13–18, 22–
245 24, 42]. The 2009 AI global pandemic demonstrates the ability of phylodynamic analyses to provide
246 novel insights for decisions regarding animal and public health [42]. Our phylodynamic analyses of
247 an H5N1 HPAIV sequence dataset and associated metadata provides insights on the origin of the out-
248 break in Europe (*e.g.*, the ancestral location or host type) and the temporal and the spatial progression
249 of this epidemic. These inferences could be used to inform prevention and control measures to block
250 the virus at the source (*e.g.*, high-risk geographical areas or bird species), which can limit the spread of
251 the virus both to the domestic poultry population or to uninfected geographic areas. Furthermore, the
252 ability to infer and predict routes of viral transmission has clear implications for informing the selec-
253 tion of appropriate strains for vaccine production to more effectively control new reassortant strains of
254 AIV in future epidemics. Accordingly, incorporating phylodynamic analyses as a standard tool for the
255 molecular surveillance of AI might support the development of more effective (and cost effective) pol-
256 icy decisions for the control of this virus in high-risk regions.

257 **3. Conclusions**

258 The purpose of this study was to illustrate the potential of a Bayesian phylodynamic approach to in-
259 form AI surveillance and control programs and associated decision-making processes. Our results indi-
260 cate that the peak in viral genetic diversity between 2005–2006 coincided with its geographic spread
261 from Russia into Romania and Germany, and throughout the rest of the EU. This expansion was fol-
262 lowed by a drastic decline in viral genetic diversity, which corresponds with the epidemic decline in
263 most EU member states. Our findings corroborate previous conclusions that wild birds were the most
264 important vectors for the spread of the disease. However, our analyses indicate that domestic (rather
265 than wild) birds were the source of the epidemic, and that viral reassortment most likely occurred in
266 domestic species. Finally, our results provide further demonstration of the potential of the Bayesian
267 phylodynamic approach to effectively study the dynamics of disease origin and spread, which promis-
268 es to maximize the utility of molecular sequence data for the surveillance and control of emerging dis-
269 eases. These results will contribute to the design of surveillance and control strategies for H5N1
270 HPAIV in Europe and provide a methodological framework for molecular epidemiological modeling
271 of emerging novel HPAIVs at regional, national, and international scales. This work can also assist
272 global and European disease agencies in their effort to improve the efficiency of their surveillance sys-
273 tems by sharing and improving their genomic databases in the course of emergence of new AIVs.

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276 **4. Material and Methods**

277 *4.1 Sequence data*

278 Our dataset comprised complete (or nearly complete) H5N1 HPAIV haemagglutinin (HA) and neu-
279 raminidase (NA) gene sequences isolated in Europe and the Russia between May, 2005 and June,
280 2010, including information on the date, location, and host from which the sequences were isolated.
281 These data were obtained from the Global Initiative on Sharing All influenza data (GISAID) public
282 database (<http://platform.gisaid.org>), with the exception of sequences from Romania, which have been
283 described elsewhere [46]. Laboratory recombinants or highly cultured sequences were excluded from
284 the dataset. The length of the HA and NA gene regions ranged from 1621–1803 and 1211–1458 bp,
285 respectively. Our dataset comprised a total of 347 sequences, of which 277 (79.8%) were from Europe
286 and Russia, and 70 (20.2%) were from Asian and African countries that were included to root the phy-
287 logenetic tree (Table S1). Furthermore, we cross-matched each retrieved sequence with the NCBI In-
288 fluenza Virus Resource (IVR) (<http://www.ncbi.nlm.nih.gov/genomes/FLU/FLU.html>) database to
289 obtain information on the date, location, and host of origin of the sample. Because both IVR and
290 GISAID databases do not report the exact longitude and latitude (and occasionally the actual date of
291 collection), we consulted the metadata field for each sequence isolate to approximate the location and
292 date of collection, as well as to specify the host species. We identified metadata for each sequence by
293 comparing strain names between databases (*e.g.*, A/swan/Germany/6/2007), and, when available, we
294 cross-matched information with the corresponding publication [11, 47-52]. We specified the location
295 for each sample (latitude-longitude) using the centroid of the country from which samples were isolat-
296 ed. We converted the collection date for each sequence into fractional years (decimal days) in order to
297 estimate divergence times. If only the year of isolation was available (1.2% of the sequences), we spec-
298 ified the ages as the mid-point of the corresponding year. Similarly, if only the month of isolation was
299 available (10.4% of the sequences), we recorded the date as the mid-point of the corresponding month.
300 Finally, we classified host species as either *wild birds*, *domestic birds*, or *other* (non-avian hosts). Ta-
301 ble S2 summarizes the profile of 277 H5N1 HPAIV sequences isolated from reported outbreaks in
302 Europe and the Russia between May 2005, and June 2010. Finally, an epidemic curve was plotted us-
303 ing approximately 1130 outbreak reports retrieved from the Food and Agriculture Organization of the
304 United Nation (FAO) Global Animal Disease Information System EMPRES-i ([http://empres-](http://empres-i.fao.org/)
305 [i.fao.org/](http://empres-i.fao.org/)). The outbreak data included dates and locations of detected cases in domesticated poultry
306 and wild birds in Europe and Russia between 2005 and 2010.

307 *4.2 Preliminary phylogenetic analysis*

308 We aligned the H5N1 HPAIV HA and NA gene regions using MUSCLE version 3.8 [43]. The re-
309 sulting alignments were adjusted manually to ensure that these protein-coding gene regions remained
310 in frame (assessed by amino-acid translation using Mesquite version 3.01) [44]. We removed all se-
311 quences (25%) with 100% nucleotide identity and both gene regions were found to be free of homolo-
312 gous recombination using Recombination Detection Program version 3 (RDP3) [45]. For each gene
313 region, we selected a mixed substitution model (assignment of data partitions to substitution models)
314 by first defining six data subsets (the codon positions of each gene region), and then selecting among

315 mixed-substitution model using the Bayesian Information Criterion (BIC) implemented in Partition-
316 Finder v 1.1 [46]. Finally, we assessed the degree of topological (in)congruence between the two gene
317 regions by performing maximum-likelihood estimates of the phylogeny for the individual gene regions
318 under the selected mixed-substitution model; these analyses entailed 100 non-parametric bootstrap
319 replicate searches of each gene region using RAxML version 8 [47].

320 *4.3 Divergence-time estimation*

321 We estimated divergence times for the H5N1 HPAIV sequence dataset using the (relaxed) molecu-
322 lar-clock models implemented in BEAST v 1.8 [48]. We used the mixed-substitution model for each
323 gene region identified in the preliminary analyses detailed above. We assessed the fit of the sequence
324 data to three branch-rate prior models: (1) a strict molecular clock model (which assumes that substitu-
325 tion rates are stochastically constant across branches of the tree); (2) the uncorrelated exponential re-
326 laxed clock (UCED) model, which assumes that substitution rates on adjacent branches are sampled
327 from a shared exponential distribution, and; (3) the uncorrelated lognormal relaxed clock (UCLN)
328 model, which assumes that substitution rates on adjacent branches are drawn from a shared lognormal
329 distribution. We assessed the fit of the sequence data to these three branch-rate models by comparing
330 their corresponding marginal likelihoods, which were estimated using Tracer version 1.6 [48, 49].
331 Bayes factor (BF) comparisons indicated that the UCED branch-rate model provided the best fit for
332 both HA and NA gene regions (Bayes factor >25 for the log marginal likelihood). Estimation of diver-
333 gence times also requires a node-age model; we assumed a Bayesian skyline coalescent tree prior,
334 which allowed us to estimate changes in the effective population size through time [50]. Isolation dates
335 of the sequences were used to calibrate divergence-time estimates.

336 We approximated the joint posterior probability density of these model parameters using the Mar-
337 kov Chain Monte Carlo (MCMC) algorithms implemented in BEAST v 1.8. Each MCMC simulation
338 was run for 1×10^8 cycles, and thinned by sampling every 10,000 cycle. We performed four replicate
339 MCMC simulations to aid in assessing performance of the simulations. We assessed MCMC conver-
340 gence by comparing the parameter estimates from independent analyses for each parameter, ensuring
341 that the estimates of the marginal posterior probability densities were effectively identical for the four
342 independent chains. We also assessed convergence by the estimating effective sample sizes (ESS) for
343 each parameter using Tracer, and assessed mixing based on the acceptance ratios. Those evaluations
344 suggested that the MCMC algorithms provided a reliable approximation of the posterior probability
345 density, and suggested that the first 10% of the samples (the “burn-in”) should be discarded. We sum-
346 marized posterior results as a maximum clade credibility (MCC) tree using TreeAnnotator. We then
347 generated a Bayesian skyride plot to infer the population demographics of H5N1 HPAIV HA and NA
348 gene segments in Europe and the Russia. We plotted the inferred effective population size of the virus
349 between 2005–2010 in terms of relative genetic diversity ($N_e T$), where N_e is the effective population
350 size and T the generation time [51].

351 *4.4 Estimation of geographic history under the discrete phylodynamic model*

352 We incorporated information on the geographic locations where sequences were isolated to describe
353 the spatial dynamics of viral epidemics following a procedure described by Lemey *et al.* [27]. Briefly,

354 the approach jointly estimates phylogeny and history of discrete traits in a Bayesian statistical frame-
355 work, which enables inference of the timing of viral dispersal patterns while accommodating phyloge-
356 netic uncertainty. Here, the discrete states are geographic areas, and the goal is to estimate the history
357 of viral migration (transitions) between these areas through time. This is achieved using a model aver-
358 aging approach (based on Bayesian stochastic search variable selection; BSSVS) to describe the spatial
359 evolution of H5N1 HPAIV epidemic. We modeled the geographic transition of the H5N1 HPAIV be-
360 tween areas as discrete states under a continuous-time Markov model comprising a number of non-
361 zero transition rates that were identified by means of BSSVS. We assessed the fit of the H5N1 HPAIV
362 data to a number of candidate discrete phylogeographic models (Table S4), including both symmetric
363 and asymmetric discrete-trait models with irreversible and reversible transitions, respectively (using a
364 mean-one exponential prior for the rate parameters). Accordingly, our analyses were based on a com-
365 posite phylogenetic model that comprised: (1) the previously selected mixed-substitution model; (2)
366 the strict, UCLN, UCED branch-rate models to describe variation in substitution rate across branches;
367 (3) the coalescent Gaussian Markov Random field (GMRF) Bayesian Skyride model as a prior on the
368 node times in the tree, and; (4) the symmetric and asymmetric discrete-state phylodynamic models to
369 describe the history of viral migration between discrete geographic areas. Accordingly, we explored a
370 total of six candidate composite phylodynamic models (the symmetric and asymmetric variants of the
371 geographic model and the three branch-rate models).

372 We assessed the relative fit of the six candidate phylodynamic models to the HA and NA sequence
373 datasets using Bayes factors. To this end, we first estimated the marginal likelihood for each of the six
374 candidate phylodynamic models from the resulting posterior samples using the posterior simulation-
375 based analogue of Akaike's information criterion (AICm) [52]. Bael et al. [53] have recently shown
376 the AICm to provide more reliable estimates of the marginal likelihood than the harmonic-mean esti-
377 mator. We then used the marginal-likelihood estimates to compute Bayes factors to select among the
378 candidate models. The UCED branch-rate model was again preferred for both HA and NA gene seg-
379 ments (BF > 25). Bayes factor (BF) comparisons indicated that the symmetric UCED branch-rate
380 model with irreversible transitions provided the best fit for both HA and NA gene regions (Bayes fac-
381 tor >25 for the log marginal likelihood). The posterior probability distribution of trees sampled under
382 the preferred model was summarized as an MCC consensus tree with posterior probabilities of geo-
383 graphic areas plotted at interior nodes using FigTree version 1.4 [54]. We used SPREAD version 1.0.6
384 [55] to assess the strength of transition rates between discrete geographic areas, using a BF cutoff = 6
385 to identify non-zero transition rates. Finally, we generated a keyhole markup language (KML) file to
386 visualize the geographic migration of the virus.

387 *4.5 Exploring the evolution of H5N1 HPAIV host infection*

388 We modeled the evolutionary movement of H5N1 HPAIV within and between host types (domestic
389 birds, wild birds, other). To this end, we again adopted a discrete-trait model, where the discrete states
390 in this case correspond to the host type, where the objective is to infer the history of H5N1 HPAIV
391 migration between hosts through time. The number of non-zero transition rates in the model was again
392 estimated using BSSVS. The relative strength of transition rates (e.g., wild → domestic) was estimated
393 using Bayes factors. Following [64], we estimated the ancestral states (host type) at internal nodes of

394 the tree under a composite phylogenetic model that included: (1) the previously selected mixed-
395 substitution model to describe evolution of nucleotides over the tree with branch lengths; (2) the pre-
396 ferred uncorrelated lognormal model to describe the variation of substitution rates across branches of
397 the tree [56]; (3) a constant population-size coalescent model to describe the temporal distribution of
398 node heights in the phylogeny, and; (4) the asymmetric discrete-state phylodynamic model to describe
399 the history of viral migration between host types as a continuous-time Markov process (with a number
400 of non-zero transition rates determined by the BSSVS procedure, using a uniform prior with mean = 1
401 for the rate parameters). We summarized the posterior probability distribution of parameter estimates
402 as an MCC consensus tree with the posterior probability of host state mapped to interior nodes of the
403 tree using FigTree. Additionally, we assessed the strength of transition rates between states (hosts) by
404 means of Bayes factors using SPREAD with a cutoff of six to identify non-zero transition rates. Use of
405 the asymmetric discrete trait model allowed us to assess the strength of directionality between states
406 (e.g., wild → domestic and domestic → wild bird).

407 *4.6 Assessing uncertainty in discrete-trait mappings and association statistics*

408 We used the Kullback-Leibler (KL) divergence statistic [57] to accommodate phylogenetic uncer-
409 tainty in the discrete-trait estimates (for host type and geographic location). The KL divergence
410 measures the departure between prior probability distribution and the corresponding posterior probabili-
411 ty distribution for a given parameter. The premise is simple: the posterior probability distribution in-
412 ferred for a given parameter is the updated version of prior probability distribution—it is updated by
413 the information in the data via the likelihood function. Accordingly, if the data contain little infor-
414 mation regarding the value of a parameter, its posterior probability distribution will closely resemble
415 the corresponding prior probability distribution, and the KL divergence between these two probability
416 distributions will therefore be small [27]. We calculated the KL divergence for each selected posterior
417 distribution of trees using a function written by Razavi [58] implemented in Matlab v 2013a [59]. The
418 Association Index (AI) was calculated using Bayesian Tip-Significance Testing (BaTS) software ver-
419 sion 1.0 to test the hypothesis that a taxon with a given trait (host type or geographic location) are
420 more likely to share traits with adjoining taxa than that expected by chance [60].

421 **Acknowledgments**

422 We are grateful to Matthew L. Scotch for assistance with Matlab. This research was supported by
423 National Science Foundation (NSF) grants DEB-0842181, DEB-0919529, DBI-1356737, and DEB-
424 1457835 awarded to B.R.M..

425 **Author Contributions**

426 M.A. and B.R.M. conceived and designed the study. M.A. and A.P. analyzed the data. B.R.M. pro-
427 vided advise regarding the analyses. M.A., B.R.M. and A.P. wrote the paper.

428 **Conflicts of Interest**

429 The authors declare no conflict of interest.

430 **References and Notes**

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595 **Supplementary Materials**

596 **Table S1** Bayes factor test for non-zero rates for location (country) state model for both
 597 HA and NA genes. Large BF values indicate most likely routes of dispersal between coun-
 598 try's location states.

Gene	BF	Between	
HA	77823.6	Russia	Romania
HA	77823.6	Germany	Austria
HA	77823.6	Germany	Czech Republic
HA	8639.4	Turkey	Romania
HA	317.0	Romania	Germany
HA	274.4	Poland	Germany
HA	90.6	Romania	Hungary
HA	87.4	Romania	Italy
HA	48.5	Ukraine	Russia
HA	46.4	Switzerland	Germany
HA	43.4	Germany	France
HA	16.0	Romania	Bosnia
HA	13.9	Sweden	Romania
HA	11.9	Romania	Bulgaria
HA	11.1	Romania	Croatia
HA	9.8	Slovenia	Austria
HA	7.7	Sweden	Denmark
HA	7.0	Germany	Denmark
NA	77823.6	Romania	Germany
NA	2984.9	Russia	Romania
NA	2502.1	Turkey	Romania
NA	441.2	Germany	Czech Republic
NA	409.8	Switzerland	Germany
NA	240.0	Sweden	Denmark
NA	190.9	Romania	Hungary
NA	149.2	Poland	Germany
NA	108.2	Germany	France
NA	87.0	Ukraine	Russia
NA	57.9	Germany	Austria
NA	27.5	Romania	Italy
NA	16.3	Romania	Bulgaria
NA	15.4	Slovenia	Austria
NA	15.3	Sweden	Germany
NA	11.1	Romania	Bosnia
NA	7.6	United Kingdom	Germany
NA	6.9	Romania	Croatia

599 Bayes factors (BF) > 6 indicate significant non-zero dispersal

600

601 **Table S2** GISAID database Accession numbers and isolate names matched with other pub-
 602 lic databases used in this study for both HA and NA gene segments, with the exception of
 603 sequences isolated from Romania.

Epi_Flu #	Isolate_Name
EPI_ISL_7012	A/chicken/Afghanistan/1573_7/2006
EPI_ISL_66178	A/coot/Voecklabruck/1585/2006
EPI_ISL_66194	A/duck/Leibnitz/243/2006
EPI_ISL_66195	A/duck/Mellach/335/2006
EPI_ISL_66196	A/duck/Schaerding/1398/2006
EPI_ISL_66199	A/duck/Wels/2025/2006
EPI_ISL_66200	A/duck/Wien/1836/2006
EPI_ISL_66201	A/egret/Wien/1977/2006
EPI_ISL_66202	A/goose/Wien/1966/2006
EPI_ISL_66219	A/swan/Krems/2354/2006
EPI_ISL_66220	A/swan/Krems/2547/2006
EPI_ISL_66221	A/swan/Krems/2675/2006
EPI_ISL_66222	A/swan/Laakirchen/2703/2006
EPI_ISL_66223	A/swan/Mellach/215/2006
EPI_ISL_66225	A/swan/Perg/1358/2006
EPI_ISL_66226	A/swan/Schaerding/1499/2006
EPI_ISL_66227	A/swan/Schaerding/1806/2006
EPI_ISL_66228	A/swan/Schwechat/2538/2006
EPI_ISL_66229	A/swan/Voecklabruck/1484/2006
EPI_ISL_66230	A/swan/Wien/1410/2006
EPI_ISL_66231	A/swan/Wien/2323/2006
EPI_ISL_74836	A/Anas platyrhynchos/Belgium/09-762-P1/2008
EPI_ISL_64083	A/Cygnus olor/BIH/1/2006
EPI_ISL_105998	A/common buzzard/Bulgaria/38WB/2010
EPI_ISL_1254	A/Goose/Guangdong/1/96
EPI_ISL_1909	A/goose/Guangdong/3/1997
EPI_ISL_10006	A/chicken/Hebei/718/2001
EPI_ISL_15275	A/duck/Guangxi/27/2003
EPI_ISL_4533	A/duck/Guangdong/173/2004
EPI_ISL_64893	A/Mallard/Huadong/hn/2005
EPI_ISL_78004	A/chicken/Xinjiang/27/2006
EPI_ISL_75197	A/great_crested_grebe/Qinghai/1/2009
EPI_ISL_6386	A/cygnus olor/Croatia/1/2005
EPI_ISL_15685	A/Cygnus olor/Czech Republic/6111/06
EPI_ISL_15689	A/Cygnus olor/Czech Republic/10814/06
EPI_ISL_15690	A/turkey/Czech Republic/10309-3/07
EPI_ISL_15691	A/chicken/Czech Republic/11242-38/07
EPI_ISL_15692	A/Cygnus olor/Czech Republic/10732/07
EPI_ISL_15693	A/Cygnus olor/Czech Republic/10662/06
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EPI_ISL_27768	A/domestic_duck/Germany/R1779/2007
EPI_ISL_27797	A/great crested grebe/Germany/R1406/07

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EPI_ISL_68261	A/wild duck/Bavaria/12/2007
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EPI_ISL_30650	A/chicken/East_Java/UT6031/2007
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EPI_ISL_21034	A/cygnus olor/Italy/742/2006
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EPI_ISL_11391	A/chicken/Tula/4/05
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EPI_ISL_11385	A/chicken/Suzdalka/06/05
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EPI_ISL_10052	A/Cygnus olor/Astrakhan/Ast05_2_6/2005
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EPI_ISL_64332	A/chicken/Reshoty/02/2006
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EPI_ISL_10435	A/grebe/Tyva/Tyv06_2/06
EPI_ISL_64356	A/duck/Omsk/1822/2006
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EPI_ISL_63440	A/chicken/Domodovo/MK/2007
EPI_ISL_11823	A/chicken/Moscow/2/2007
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EPI_ISL_80265	A/great_crested_grebe/Tyva/22/2010
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EPI_ISL_12829	A/common pochard/Switzerland/V558/2006
EPI_ISL_12830	A/common pochard/Switzerland/V592/2006
EPI_ISL_12831	A/common pochard/Switzerland/V762/2006
EPI_ISL_10107	A/turkey/Turkey/1/2005
EPI_ISL_83693	A/Pigeon/Turkey-Agri/09rs2841-8/2005
EPI_ISL_83692	A/Pigeon/Turkey-Kars/09rs2841-7/2005
EPI_ISL_83700	A/Chicken/Turkey-Edirne/09rs2843-5/2008
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EPI_ISL_103350	A/swan/England/AV380-2326/2008
EPI_ISL_99850	A/Common Buzzard/Berlin/1/2006
EPI_ISL_6924	A/chicken/Egypt/2253_1/2006
EPI_ISL_6925	A/turkey/Egypt/2253_2/2006
EPI_ISL_6384	A/duck/Ivory Coast/1787_18/2006
EPI_ISL_2932	A/chicken/Nigeria/SO494/2006
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EPI_ISL_66867	A/chicken/Nigeria/08RS848-121/2007
EPI_ISL_292	A/whooper swan/Mongolia/3/05
EPI_ISL_293	A/whooper swan/Mongolia/4/05
EPI_ISL_372	A/whooper swan/Mongolia/2/06
EPI_ISL_548	A/common goldeneye/Mongolia/12/2006
EPI_ISL_9940	A/migratory duck/Jiangxi/2300/2005
EPI_ISL_9939	A/migratory duck/Jiangxi/2295/2005
EPI_ISL_9938	A/migratory duck/Jiangxi/2136/2005
EPI_ISL_9937	A/migratory duck/Jiangxi/1701/2005
EPI_ISL_9936	A/migratory duck/Jiangxi/1657/2005
EPI_ISL_9935	A/migratory duck/Jiangxi/1653/2005
EPI_ISL_3903	A/duck/Zhejiang/11/2000
EPI_ISL_3900	A/duck/Shanghai/35/2002
EPI_ISL_3891	A/duck/Fujian/19/2000
EPI_ISL_68660	A/chicken/Hong Kong/786-2/1997
EPI_ISL_966	A/Chicken/Hong Kong/728/97
EPI_ISL_62720	A/bar-headed goose/Mongolia/X54/2009
EPI_ISL_62738	A/whooper swan/Mongolia/8/2009
EPI_ISL_76273	A/whooper swan/Mongolia/1/2010
EPI_ISL_76635	A/whooper swan/Mongolia/21/2010
EPI_ISL_93807	A/duck/Hokkaido/WZ101/2010
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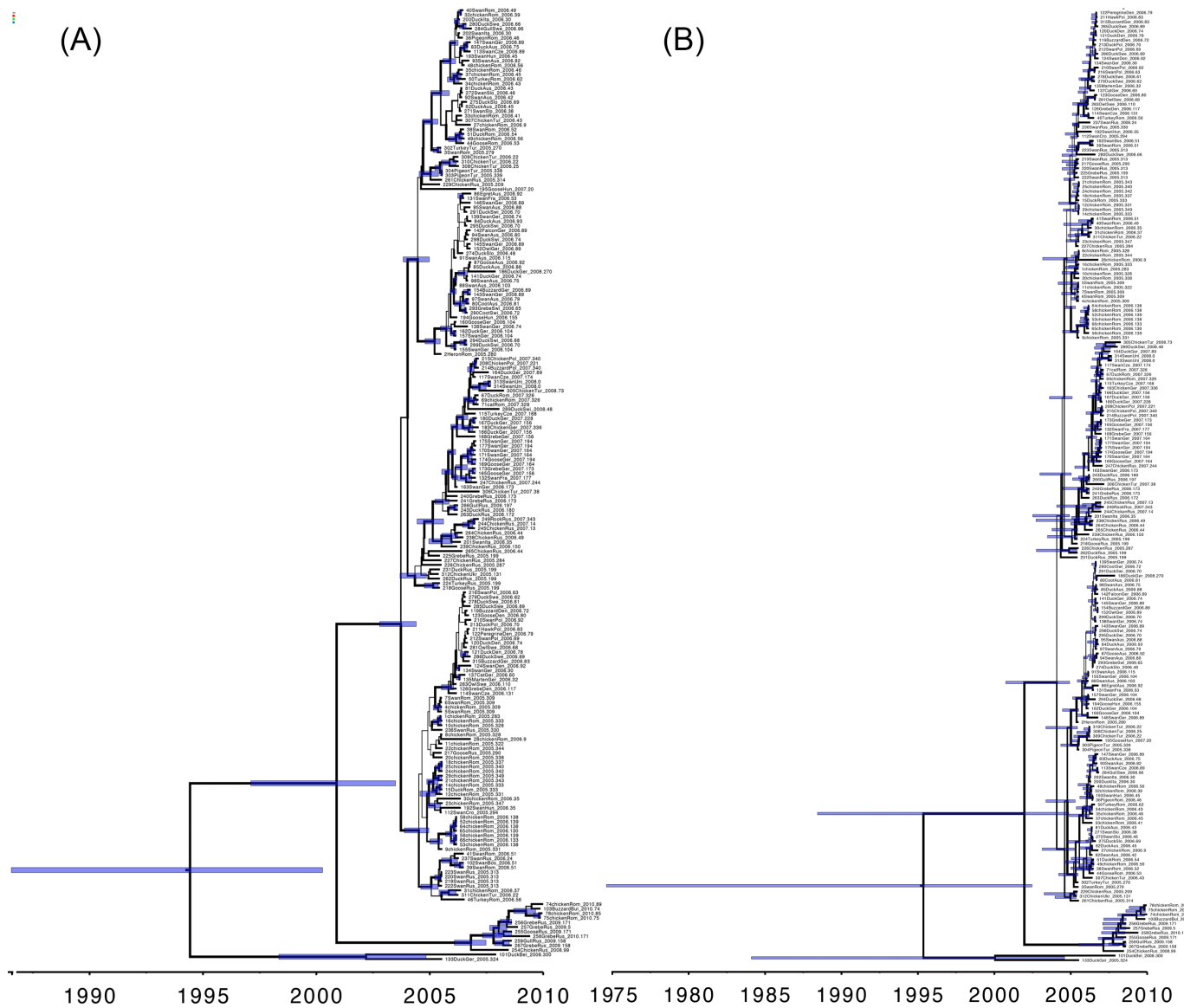
605 **Table S3** Summary profile of H5N1 HPAIV sequences isolated from reported outbreaks in
 606 Europe and the Russia between May 2005 and June 2010 ($N = 277$).

Country	N.s eq	Collection date	Host type
Austria	20	February 2006-April 2006	Wild
Belgium	1	November 2008	Wild
Bosnia	1	February 2006	Wild
Bulgaria	1	March 2010	Wild
Croatia	1	October 2005	Wild
Czech Republic	6	March 2006-July 2007	Wild & domestic
Denmark	8	March 2006-April 2006	Wild
France	2	February 2006-June 2007	Wild
Germany	57	November 2005-October 2008	Wild, domestic, & Other
Hungary	5	February 2006-January 2007	Wild
Italy	3	February 2006	Wild
Poland	9	March 2006-December 2007	Wild & domestic
Romania	78	October 2005-March 2010	Wild, domestic, & Other
Russia	51	July 2005-January 2010	Wild & domestic
Slovakia	2	February 2006	Wild
Slovenia	5	February 2006-March 2006	Wild
Sweden	11	March 2006-April 2006	Wild
Switzerland	13	February 2006-February 2008	Wild
Ukraine	1	May 2005	Domestic
United King- dom	2	January 2008	Wild

609 **Table S4** Candidate phylogeographic models for HA and NA gene segments explored for relative fit.
610

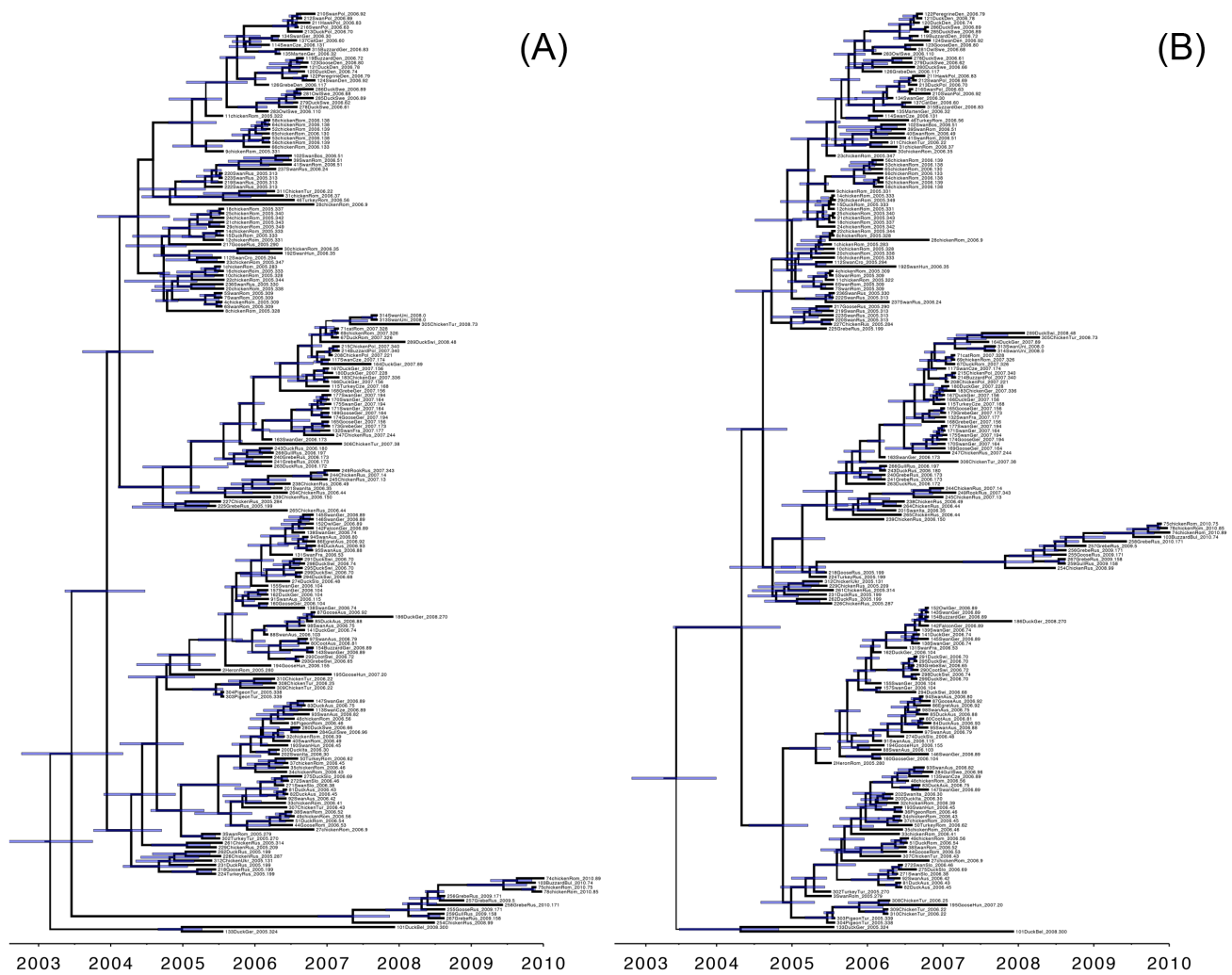
Model	Sites	Discrete-traits	Branch rates	Trees
1			Strict clock	
2		Symmetric reversible	UCLN	
3	GTR+GAMMA		UCED	Coalescent: GMRF
4			Strict clock	Bayesian Skyride
5		Asymmetric irreversible	UCLN	
6			UCED	

611
612



613
614 **Figure S1.** Maximum clade credibility (MCC) trees for H5N1 HPAIV hemagglutinin (HA)
615 and neuraminidase (NA) gene regions (A and B, respectively). Branch lengths are ren-
616 dered proportional to absolute time (see timescales), and the branch thickness is propor-
617 tional to the posterior probability of the inferred host state. Nodes correspond to median
618 ages and the blue horizontal bars at nodes represent the corresponding 95% HPDs for di-
619 vergence-time estimates.

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623

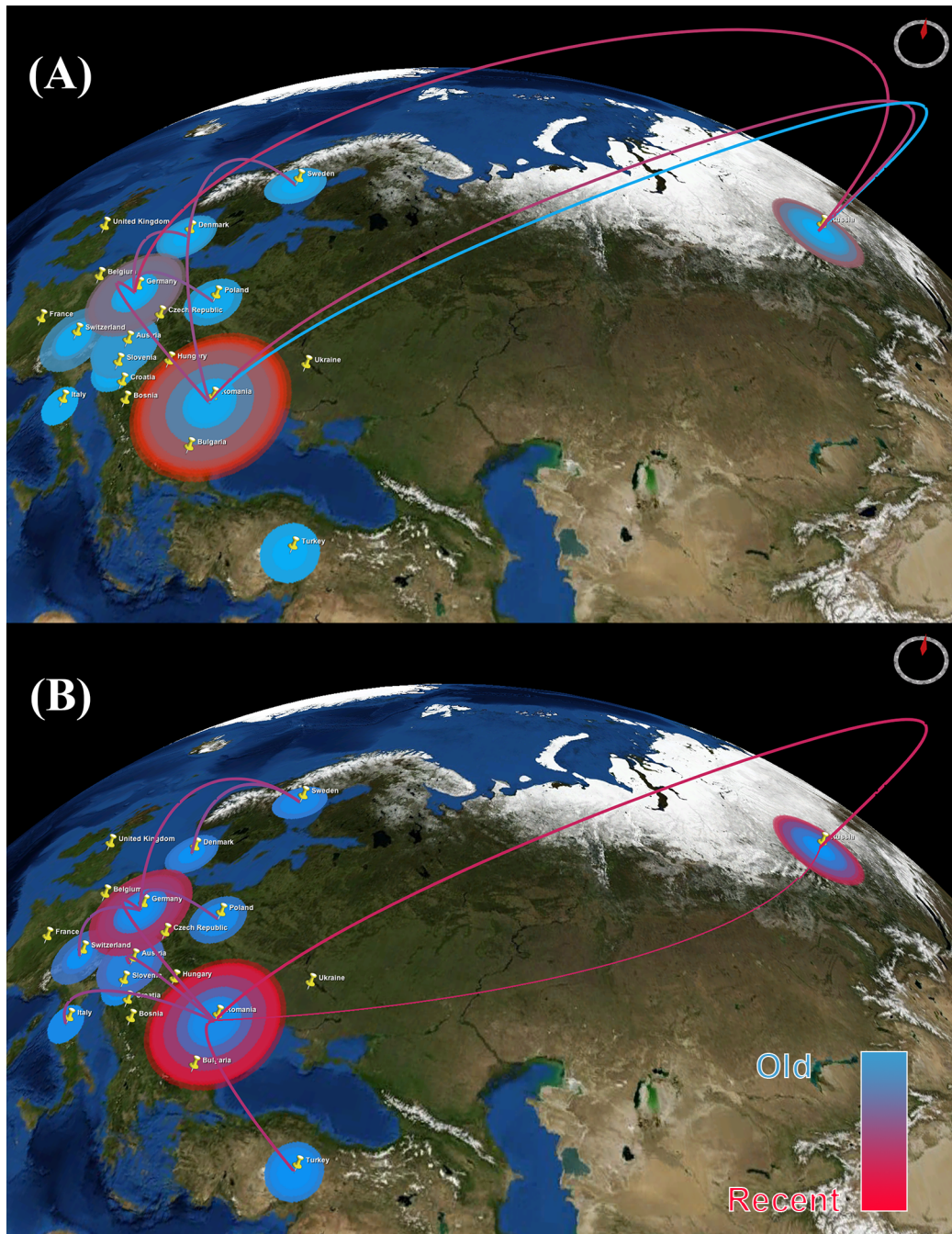
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626

627

Figure S2. Maximum clade credibility (MCC) trees for H5N1 HPAIV hemagglutinin (HA) and neuraminidase (NA) gene regions (A and B, respectively). Branch lengths are rendered proportional to absolute time (see timescales), and the branch thickness is proportional to the posterior probability of the inferred location state. Nodes correspond to median ages and the blue horizontal bars at nodes represent the corresponding 95% HPDs for divergence-time estimates.



628

629 **Figure S3. Temporal dynamics of the spatial diffusion of H5N1 HPAIV.** The figure
630 consists of snapshots of the dispersal pattern of avian influenza between May 2005 and
631 June 2010. Lines between locations represent branches in the MCC tree along which spa-
632 tial transmission occurs. The diameters of circles are proportional to square root of the
633 number of MCC branches maintaining a particular location state at each time-point. The
634 blue and red color gradients reflect the relative age of the transitions for each gene (older-
635 recent, respectively). A: HA gene. B: NA gene. The maps are based on satellite pictures
636 available in NASA World Wind (<http://worldwind.arc.nasa.gov/java/>).

637

638 **File S1 Temporal dynamics of the spatial diffusion of H5N1 HPAIV based on the HA gene.** The
639 KML file demonstrates the dispersal pattern of avian influenza between May 2005 and June 2010.
640 Lines between locations represent branches in the MCC tree along which spatial transmission occurs.
641 The diameters of circles are proportional to square root of the number of MCC branches maintaining a
642 particular location state at each time-point. The blue and red color gradients reflect the relative age of
643 the transitions for HA gene (older-recent, respectively).