Seed size and its rate of evolution correlate with species diversification across

angiosperms

Short title: Seed size and angiosperm diversification

J. Igea[1,4], E. F. Miller[1,2], A. S. T. Papadopulos[3], A. J. Tanentzap[1]

Author affiliations:

[1] Department of Plant Sciences, University of Cambridge, Downing St, Cambridge,

CB2 3EA, UK

[2] Department of Zoology, University of Cambridge, Downing St, Cambridge, CB2

3EJ, UK

[3] Jodrell Laboratory, Royal Botanic Gardens, Kew, Richmond, TW9 3AB, UK.

[4]*Corresponding author:*

Dr. Javier Igea

Department of Plant Sciences, University of Cambridge, Downing St, Cambridge,

CB2 3EA, UK

ji247@cam.ac.uk

Abstract

Species diversity varies greatly across the different taxonomic groups that comprise the Tree of Life (ToL). This imbalance is particularly conspicuous within angiosperms, but is largely unexplained. Seed mass is one trait that may help clarify why some lineages diversify more than others because it confers adaptation to different environments, which can subsequently influence speciation and extinction. The rate at which seed mass changes across the angiosperm phylogeny may also be linked to diversification by increasing reproductive isolation and allowing access to novel ecological niches. However, the magnitude and direction of the association between seed mass and diversification has not been assessed across the angiosperm phylogeny. Here, we show that absolute seed size and the rate of change in seed size are both associated with variation in diversification rates. Based on the largest available angiosperm phylogenetic tree, we found that smaller-seeded plants had higher rates of diversification, possibly due to improved colonisation potential. The rate of phenotypic change in seed size was also strongly positively correlated with speciation rates, providing rare, large-scale evidence that rapid morphological change is associated with species divergence. Our study now reveals that variation in morphological traits and, importantly, the rate at which they evolve can contribute to explaining the extremely uneven distribution of diversity across the ToL.

Introduction

Angiosperms are one of the most species-rich clades on Earth and have dominated terrestrial plant communities since the Late Cretaceous [1]. The astounding diversity

of flowering plants is distributed extremely unevenly across the ToL. Each of the five

most species-rich angiosperm families contains >10,000 species while more than 200

families contain <100 species each [2]. An enduring pursuit in evolutionary biology is

to explain this uneven distribution of biodiversity, not only in angiosperms, but also

across the whole ToL [3].

Biological traits offer one way to explain disparity in species diversification if they

confer adaptation to different environments. Seed mass is one such trait that is

particularly important for angiosperms because it integrates across many

characteristics of an individual's life history strategy [4]. Along with adult plant size,

seed mass affects survival, reproductive lifespan and dispersal [5]. These life history

characteristics contribute to fitness and adaptation, which are the ultimate

determinants of whether lineages diversify or go extinct [6]. In support of this idea,

seed mass has been shown to correlate negatively with diversification in the

Polygonaceae [7], but this has not been investigated across large taxonomic scales. As

seed mass varies over ten orders of magnitude in angiosperms, from the minute 1 μg

seeds of some orchids to the >18 kg seeds of the sea coconut (*Lodoicea maldivica*),

this huge variation may coincide with variation in species diversity. Generalising the

direction and magnitude of a link between seed mass and diversification across

taxonomic scales has, however, proved difficult. Some life history characteristics

encapsulated by seed mass are expected to promote speciation or extinction, while

others may simultaneously counteract such effects [8].

The rate of change in key biological traits, such as seed size, can be as important in

driving macroevolutionary dynamics as the absolute values of the traits themselves

[9]. This is because phenotypic divergence may cause reproductive isolation that

results in speciation [10]. Nevertheless, few empirical studies have detected a

correlation between rates of phenotypic evolution and lineage diversification

([9,11,12] but see [13]). A correlation between the two may be expected where a trait

can change more rapidly in some species than others in response to selective pressures

(i.e. high "evolvability" [14]). This rapid change may enable greater access to new

ecological niches or quicker establishment of reproductive isolation, thereby

increasing the rate of speciation (λ) [15]. In the case of seed mass, the ability to switch

rapidly from small seeds with high dispersal ability to larger seeds with lower

dispersal ability might promote cycles of rapid colonisation and isolation or permit

adaptation to new dispersal vectors in novel environments. Rapid evolution of new

phenotypes may also allow individuals to escape harsh environmental conditions and

competitive interactions [16], thereby decreasing extinction rates (μ). The overall

outcome of these processes on net diversification ($r = \lambda$-μ) will ultimately depend

upon which of these rates responds more strongly to phenotypic change.

Here, we show that both seed mass and its phenotypic rate of evolution correlate with

speciation and extinction across the angiosperm ToL. Our approach combined the

most comprehensive phylogenetic timetree available [17] with an unparalleled dataset

of seed mass measurements from over 30,000 angiosperm species [18]. We estimated

rates of speciation, extinction and seed size evolution across the phylogeny using

Bayesian, method-of-moments, and maximum likelihood analyses – each with

different assumptions regarding rate variation through time. We then tested whether

there were any links between rates of diversification and both seed size and its rate of

evolution. Additionally, we examined whether these links were consistent across

different methodologies and timescales.

Results

Our results point to a strong association between angiosperm diversification and rates

of seed size evolution irrespective of analytical method or timescale, with weaker

evidence for a link between macroevolutionary dynamics and absolute seed size. In

the first instance, we calculated rates of speciation (λ), extinction (μ) and seed size

evolution using Bayesian Analysis of Macroevolutionary Mixtures (BAMM) [19].

BAMM models rate heterogeneity through time and lineages, and accounts for

incomplete taxon sampling. We used a phylogenetic tree that contained 29,703

angiosperm species for the speciation/extinction analysis. The tree was subset to

13,577 species with seed size data for phenotypic evolution analysis. As expected,

given the high degree of taxonomic imbalance observed in the angiosperm phylogeny,

we found strong support for more than 500 shifts in the rates of diversification. There

was also marked heterogeneity in the rates of seed size evolution (Fig. 1), which

varied over three orders of magnitude (Fig. S1).

Figure 1. Seed mass and macroevolutionary rates estimated with BAMM across

the angiosperm tree of life. Phylogenetic tree of 13,577 species of flowering plants

with seed mass, rate of seed mass change, and speciation (λ), extinction (μ) and net

diversification (r) rates estimated by BAMM. Seed mass and rate data were

standardised to Z-scores so that variation could be directly compared. λ, μ and r were

calculated with a larger, 29,703-species tree.

We then estimated whether shifts in macroevolutionary dynamics (λ, μ and r)

estimated with BAMM were significantly correlated with absolute seed size and tip-specific rates of seed size evolution by comparing the empirical correlations to a null distribution generated using STructured Rate Permutations on Phylogenies (STRAPP), which is robust to phylogenetic pseudoreplication (see Methods for details) [20]. We were able to link major differences in diversity across angiosperm clades with both the present rate of phenotypic evolution and the absolute value of trait itself. Specifically, increased speciation was associated with a faster rate of seed size evolution (Spearman's $\rho = 0.55$, p-value < 0.0001; Fig. 2a). Increased extinction rates were similarly associated with higher evolvability ($\rho = 0.44$, p-value < 0.0001; Fig 2b), but given the weaker effect, the net outcome of λ-μ was that diversification rates were positively correlated with phenotypic change ($\rho = 0.49$, p-value < 0.0001; Fig. 2c). We also identified an association between seed size and both speciation ($\rho = -0.17$, p-value $= 0.003$; Fig. 2d), and extinction rates ($\rho = -0.17$, p-value $= 0.003$, Fig. 2e). As the correlations with speciation and extinction were in the same direction and of comparable magnitude, and estimates of extinction rates were relatively variable (Fig. 2e), net diversification rates did not change with seed size ($\rho = -0.12$, p-value $= 0.077$; Fig. 2f). Generally, the observed correlations arose from many phenotypically fast-evolving clades distributed across the phylogeny (Fig. S1) and were robust to prior choice in the BAMM analyses (Fig. S2).

Figure 2. Seed mass and seed mass rate of evolution correlate with macroevolutionary dynamics estimated with BAMM. Spearman correlations were calculated between speciation (λ), extinction (μ), and net diversification (r) and each of a) present-day rate of seed mass change and b) seed mass. Coloured lines are correlations for one sample of the BAMM posterior distribution, bold line is the

median. The insets show the density plots of the absolute difference between the

observed and null correlation calculated across 1,000 structured permutations of the

evolutionary rates on the phylogenetic tree (myr = million years).

Given recent discussion on the reliability of BAMM for estimating diversification

rates ([21], but see [22]), we tested the robustness of our results by using alternative

methodologies to infer macroevolutionary dynamics across clades at different

timescales. Ten, 2 million year-wide time slices from the present up to 20 million

years ago were defined. These time slices were used to identify the most inclusive

monophyletic clades of ≥4 species in which we had estimated a ≥70% probability of

recovering the correct crown age node of the clade (see Methods). For each resulting

clade in each time slice, we calculated diversification rates using a method-of-

moments estimator, which assumes rates are constant over time [23]. We also fitted a

series of time-dependent diversification models to each clade with RPANDA, which

uses a maximum likelihood approach to estimate speciation and extinction and allows

for incomplete taxon sampling [24] (see Table S1 for a summary of the best fitting

models for each time slice). Rates of seed size evolution were estimated within each

clade that also had ≥4 species with seed size data by fitting both Brownian motion

(BM) and early burst (EB) or accelerating decelerating [25] models of trait evolution.

Mirroring the BAMM results, we found a positive correlation between the rate of seed

size evolution and speciation rates that was consistent across time slices (Fig. 3a, Fig.

S3a). As expected given the weaker association between seed size and speciation

found in our BAMM analyses, correlations were generally weaker and non-significant

(Fig. 3b), except for one of the time slices (Fig. S3b).

Figure 3. Seed mass and seed mass rate of evolution correlate with speciation in the clade-based analysis. Correlation of (a) rate of seed mass evolution and (b) seed mass with speciation rate (λ) estimated using RPANDA in the clade-based analysis. The strength of correlations is shown as PGLS slopes and were calculated using mean clade-level seed mass across 10 time slices in our species-level phylogenetic tree. The size of the circles represents the number of clades in each time slice plotted at the median age of the time slice. Colour indicates the significance of the slope. A detailed representation of the results in each time slice is given in figures S13 and S14. Correlations calculated with speciation rates obtained with the method-of-moments estimator are given in Fig. S3.

We also found limited evidence that other traits that co-vary with seed size (Tables S2, S3) better explained our results. If our results were explained by seed size being a proxy of some other phenotypic trait that ultimately influenced speciation, we would expect this other phenotypic trait to be strongly correlated with seed size. We would also expect a significantly stronger correlation of speciation with both this trait and its rate of evolution when compared to seed size and its rate of evolution. By comparing the effects of genome size, life cycle, plant height and woodiness across a subset of 1,007 species in our dataset, we found that only the distinction between woody and herbaceous and annual and perennial species were more strongly correlated with macroevolutionary dynamics than absolute seed size. The rate of phenotypic evolution of all the continuous traits (genome size, seed size and plant height) was strongly and similarly correlated with speciation (Fig. S4, Table S4). Importantly, however, neither the rate of evolution nor the absolute values of both genome size and plant height were more strongly associated with speciation than seed size or its rate of evolution

(Fig. S4). These results therefore suggest that the correlation between

macroevolutionary dynamics and both seed size and its rate of evolution is not simply

mediated by other phenotypic traits.

Discussion

Our study supports the ideas that variation in seed mass and, particularly, its rate of

evolution can help explain disparity in diversification across the angiosperm

phylogeny by playing a central role in plant life history. As we show with our clade-

based analysis, our results are repeatable across many methodologies, varying

timescales and are not restricted to a particular taxonomic rank (Figs 2, 3).

The robust association of high rates of phenotypic change with lineage diversification

has recently been observed in other taxonomic groups [9,26], but never across the

whole of the angiosperm ToL as we find here. Accelerated morphological evolution

may allow radiating lineages to occupy more complex adaptive landscapes [27].

Species with greater rate of change in their seed mass (i.e., higher evolvability) could

also shift between adaptive peaks or develop reproductive barriers more rapidly.

Alternatively, the theory of punctuated equilibria [28], whereby morphological

changes can arise from the speciation process, might also explain the connection of

phenotypic evolution with species divergence. However, current methods do not allow

us to distinguish whether speciation is responding to morphological change or vice

versa when reconstructing 250 million years of evolutionary history [9].

Seed mass itself will also co-vary with dispersal ability and environmental tolerance

in ways that can change speciation. For example, we found that smaller-seeded genera

had faster speciation rates. This may be because smaller-seeded genera generally

disperse over larger distances [29], which can promote speciation by creating isolated

populations [30]. However, the relationship between dispersal and speciation is highly

context dependent. The permeability of a landscape to dispersal determines the

dispersal distances that may promote species divergence. For example, long distance

dispersal may be needed for isolation to occur in continuous habitats but be less

effective in highly fragmented landscapes [31]. Therefore, the weak correlation that

we observed between seed size and diversification might reflect contradicting patterns

operating in different angiosperm clades. Dispersal syndromes may also modify the

effect of seed size on speciation. For instance, species with larger seeds are generally

associated with biotic dispersal that distributes seeds over greater distances than wind

or gravity dispersal [5]. However, broad-scale predictions for the effects of dispersal

syndromes on diversification may be inaccurate, since the former depend on

landscape connectivity [8] and can sometimes be inconsistent, e.g. a wind-dispersed

seed might be transported by an animal. Detailed contextual data will be necessary to

expand upon the mechanisms underlying our findings in specific regions and clades.

Although seed mass is associated with other traits that can affect diversification, there

is little evidence that these better explain our observed correlations or that seed size is

a mere proxy for one of these other traits. For example, genome size positively

correlates with seed mass [32], and faster rates of genome size evolution have been

linked to increased speciation in angiosperms [12]. Shorter, smaller-seeded plants also

tend to have faster life cycles, which may accelerate mutation rates [33,34] and

promote diversification [35]. But unlike other traits [12], both absolute seed size as

well as its rate of change were correlated with speciation. Thus, although other traits

surely influence diversification [36], we argue that our results generally reflect the

role of seed size as a trait that integrates across multiple aspects of life history

characteristics in ways that can predictably influence plant macroevolutionary

dynamics (Fig. S5).

Our analyses build upon the largest available phylogenetic timetree for angiosperms

in ways that do not consider topological and branch length uncertainty. Similar to

other megaphylogenies, the low sampling fraction (approximately 10% of described

flowering plants) and limited number of phylogenetic markers (a maximum of seven),

which were employed in constructing our phylogeny [17], may affect the inference of

macroevolutionary estimates [37,38]. In any case, we believe our results, particularly

the strong correlation between the rate of seed size evolution and speciation rates may

reflect general patterns on how biodiversity is generated across angiosperms. Detailed

studies in well-sampled clades can expand upon our findings and reveal different

relationships operating in particular groups of organisms.

The approach applied here can help to unravel the processes responsible for

generating large-scale asymmetries in biodiversity. It also offers the potential to test

how widely-varying traits and their rate of morphological evolution influence other

aspects of the evolution and adaptation of flowering plants (e.g. [17]). Clade-specific

exceptions arising from local interactions with non-focal traits [39] and specific

spatio-temporal contexts will undoubtedly interact with broad-scale

macroevolutionary patterns and may modulate the effects of seed mass on

diversification. Regardless, our results show that seed size and its rate of evolution

correlate with speciation and extinction across the flowering plants. This finding may

help to explain why some clades are much more species-rich than others and points to

the role of rapid morphological evolution in generating greater levels of diversity.

Materials and methods

Seed mass and phylogenetic dataset

Seed mass data for 31,932 species were obtained from the Royal Botanic Gardens

Kew Seed Information Database [18]. Species names were standardised with The

Plant List (TPL) nomenclature [2] and cleaned using the *Taxonstand* R package [40].

Further processing was performed with the *taxonlookup* R package [41], which is a

complete genus-family-order mapping for vascular plants that draws from TPL, the

Angiosperm Phylogeny website [42] and a higher-level manually-curated taxonomic

lookup [17].

We used the most comprehensive phylogenetic tree for land plants [17,43] that

comprises 31,389 species. Taxonomic information for our phylogenetic tree was run

through *Taxonstand* and *taxonlookup* as described above to prune it down to

angiosperms. The final phylogenetic tree contained 29,703 angiosperm species

belonging to 353 plant families (following APG IV, Fig.S6).

Diversification and phenotypic evolution analyses

Speciation, extinction and net diversification rates were estimated across the 29,703

species tree using BAMM version 2.5.0 [19]. BAMM models shifts in

macroevolutionary regimes across a phylogenetic tree using reversible-jump Markov

chain Monte Carlo (rjMCMC) sampling. The size of the tree precluded a single

analysis from readily converging. Therefore, we divided our initial tree into clades of

≤6000 species. This resulted in six monophyletic clades and one additional clade that

contained the backbone of the tree (i.e., one representative of the six monophyletic

clades) plus the remaining, unassigned species (Fig. S7). We then ran BAMM

speciation/extinction analyses for each of the seven clades (six monophyletic clades

plus the backbone set). Initial prior settings were calculated with the *setBAMMpriors*

function in *BAMMtools* [44], and the *expectedNumberOfShifts* parameter was set at

50. Following [12], we incorporated non-random incomplete sampling information by

calculating the proportion of species sampled inside each family and estimated the

backbone sampling as the overall proportion of sampled species. *Taxonlookup* was

used as a reference for these calculations. Analyses were run for 150 million

generations and convergence was verified by plotting chain traces and ensuring that

the effective sample sizes of all relevant parameters exceeded 200. The first 100

million generations were discarded as burn-in. The resulting event files for all seven

analyses were combined into a single event file, effectively generating one BAMM

result set that was then analysed following standard procedure. This clade-level

analysis allowed for diversification rate estimation with the most complete dataset that

was available. The resulting rate estimates results were strongly correlated with

estimates from a single speciation/extinction analysis using the 13,577 species that

were present in our seed size database (Fig. S8; $r = 0.92$, p-value < 0.001).

Rates of seed size evolution were also estimated with BAMM using the phylogenetic

tree of the 13,577 species that were present in our seed size database. Initial priors

were calculated as above and analyses were run for 300 million generations. The

initial 30 million generations were discarded as burn-in. We analysed BAMM prior

sensitivity following recent concerns ([45], but see [22]) by re-running both the

diversification and the trait evolution analyses for the 13,577 species dataset with

different settings for the *expectedNumberOfShifts* parameter of either 25, 50, 100 or

250. These analyses confirmed a low prior sensitivity for the posterior of the number

of expected rate shifts (Fig. S2).

Finally, we obtained clade-based measures of diversification and seed size evolution

across the species-level tree. Clades were defined as non-overlapping monophyletic

groupings of four or more species and their ages were defined using a series of 2-

million year wide time slices from the present up to 20 myr ago (see Fig. S13 for the

number of clades in each time slice). For each clade, we estimated its sampling

fraction by weighting the genus-specific sampling fractions (i.e., the number of

congeneric species in the 13,577 species tree divided by the total number of described

species for that genus) with the number of species from each genus present in the

clade. A minimum clade-specific sampling fraction of 0.3 was used for inclusion in

our analyses, ensuring that the crown sampling probability for each clade was at least

0.7 (the actual median clade crown sampling probabilities for the time slices ranged

between 0.90 and 0.96). Crown ages for the selected clades were then used to estimate

net diversification rates using the method-of-moments estimator [23]. Following

standard practice [46,47], we assumed three values of relative extinction fraction, ε =

0, 0.5 and 0.9. Different values did not affect our conclusions (results not shown);

therefore we present the results of the intermediate extinction fraction (ε = 0.5). We

also used RPANDA to fit a series diversification models that estimated time-

dependent rates for each clade. We fitted six different models of diversification: (i)

pure birth model with constant λ (speciation rate); (ii) pure birth model with

exponential λ; (iii) birth-death model with constant λ and μ (extinction); (iv) birth-

death model with exponential λ and constant μ; (v) birth-death model with constant λ

and exponential μ; and (vi) birth-death model with exponential λ and exponential μ.

We then used AIC-based model selection to select the best fitting model and obtain

the corresponding macroevolutionary parameters. Finally, we estimated the rates of

seed size evolution by fitting BM and EB models of evolution to the seed size data

within each clade using *fitContinuous* from the *geiger* R package [48]. We performed

AIC-based model selection to find the best fitting model of trait evolution. More than

99.7% of the clades showed a higher support for the BM model. Additionally, to

account for possible biases when analysing clades with many non-congeneric species,

we confirmed the results of our clade-based analysis by considering clades that only

contained congeneric species (Fig. S9).

Correlation of diversification and trait evolution

All rate variables were log-transformed for the correlation analyses. Following [20]

we treated seed mass evolutionary rates at the tips of the tree as character states and

used STRAPP to test for multiple associations between these present-day rates and

BAMM-estimated diversification dynamics. We similarly analysed the correlation of

absolute seed mass and speciation/extinction dynamics. STRAPP compares the

correlation between a focal trait and a macroevolutionary parameter (λ, μ or r) to a

null distribution of correlations. The null correlations are generated by permuting the

evolutionary rates in the tips of the phylogenetic tree while maintaining the location of

rate shift events in the phylogeny. In each case, we calculated the absolute difference

between the observed correlation of the macroevolutionary rate and the trait state and

the null correlation obtained by the structured permutations across 5000 samples from

the BAMM posterior (for an example of the observed correlation in one of the

samples in the posterior, see Fig. S10). The reported p-value was the proportion of

replicates where the null correlation coefficient was greater than the observed

correlation. We found a low type I error associated with our STRAPP correlation

analysis (p-value = 0.094, Fig. S11). We also investigated the correlation between the

mean speciation and mean phenotypic rates across all the branches in the tree using an

ordinary least squares regression (Fig. S12). We note however, that this regression

does not correct for phylogenetic dependence in the branch estimates, despite

suggesting something about patterns across the whole evolutionary history of the

angiosperms.

For the clade-based analyses, we estimated the correlation between speciation rates

(measured as λ at present time for the RPANDA analyses) and each of seed size rate

of evolution and phylogenetically-corrected mean seed size (i.e., trait value at the root

node of the clade under a BM). Correlations were estimated using phylogenetic

generalised least square (PGLS) as implemented in the R package *caper* [49] (Fig.

S13, Fig. S14). Similar results as presented in the main text were obtained when

analysing net diversification instead of speciation rates (Fig. S15). Finally, we

similarly analysed the correlation between speciation and each of seed size and its rate

of change when selecting only clades consisting of congeneric species. Again, this

analysis resulted in a similar pattern as the one presented in the main text (Fig. S9).

Diversification dynamics and other phenotypic traits

Seed mass is central to a network of inter-correlated traits associated with plant life

history that can impact diversification. Some of these traits are genome size or plant

C-value (measured as picograms of DNA per haploid nucleus), plant height, life cycle

and woodiness. We compared the correlation between macroevolutionary parameters

(λ, μ, r) and each of seed mass, seed mass rate of evolution, C-value (i.e., genome

size), plant height, life cycle and woodiness across a dataset of 1,007 angiosperm

species for which all phenotypic traits could be assembled. Genome content and life

cycle data were downloaded from the Plant DNA C-values database [50]. Woodiness

data were obtained from [17,51]. Plant height data was obtained from the TRY

database [52]. The rates of phenotypic evolution for the continuous traits (seed size,

plant height and C-value) were calculated as the phenotypic rate inferred at the tips of

the tree in our main BAMM analysis (see above). Surprisingly, mean seed mass did

not differ between the 214 strictly annual and 793 perennial plants when accounting

for phylogenetic relationships using a phylogenetic ANOVA (Fig. S16, phylANOVA:

p-value = 0.308, significance assessed with 1,000 random simulations with *phytools*

[53]). In this reduced dataset, we ran STRAPP correlations for each focal trait with the

diversification parameters calculated from our main BAMM analysis. We then

calculated the absolute differences in the observed and the null correlations between

the macroevolutionary parameters and seed mass, C-value, "annuality" (a binary

variable specifying whether the species was strictly annual or not), plant height,

woodiness (a binary variable specifying whether the species was herbaceous or

woody) and the rates of seed size evolution, genome size evolution and plant height

evolution. As expected, all phenotypic traits and their rates of evolution were

correlated with each other (Fig S5).

Code availability

Scripts used to carry out the analysis described in the paper and generate the figures

will be deposited in Github (https://github.com/javierigea/XXXX) upon acceptance.

Acknowledgments

We thank V. Soria-Carrasco for help with analyses and D. Rabosky for useful advice on the BAMM analysis. D. A. Coomes, A. J. Helmstetter, T. Jucker and W. G. Lee kindly commented on an earlier draft. J.I. and A.J.T. thank the Gatsby Charitable Foundation, Wellcome Trust and Newton Trust for funding. E.F.M was funded by the BBSRC DTP at the University of Cambridge.

Author contributions

J.I and A.S.T.P conceived the study. J.I. and E.F.M. performed the analysis. J.I. and A.J.T interpreted the analysis and wrote the manuscript. All authors edited the manuscript.

References

1. Crane PR, Friis EM, Pedersen KR. The origin and early diversification of angiosperms. Nature. 1995;374: 27–33. doi:10.1038/374027a0

2. The Plant List. Version 1.1 [Internet]. 2013 [cited 20 Jun 2003]. Available: http://www.theplantlist.org/

3. Davies T, Barraclough T. The diversification of flowering plants through time and space: Key innovations, climate and chance. Reconstructing the Tree of Life: Taxonomy and Systematics of Species Rich Taxa. 2007. pp. 149–163. doi:DOI: 10.1201/9781420009538.ch10

4. Moles AT, Ackerly DD, Webb CO, Tweddle JC, Dickie JB, Westoby M. A brief history of seed size. Science (80-). 2005;307: 576–580.

doi:10.1126/science.1104863

5.   Moles A, Leishman M. The seedling as part of a plant's life history strategy. In:

Leck MA, Parker VT, Simpson RL, editors. Seedling ecology and evolution.

Cambridge: Cambridge University Press; 2008. pp. 215–235.

6.   Salguero-Gómez R, Jones OR, Jongejans E, Blomberg SP, Hodgson DJ,

Mbeau-Ache C, et al. Fast-slow continuum and reproductive strategies

structure plant life-history variation worldwide. Proc Natl Acad Sci U S A.

2016;113: 230–5. doi:10.1073/pnas.1506215112

7.   Kostikova A, Salamin N, Pearman PB. The role of climatic tolerances and seed

traits in reduced extinction rates of temperate polygonaceae. Evolution (N Y).

2014;68: 1856–1870. doi:10.1111/evo.12400

8.   Givnish TJ. Ecology of plant speciation. Taxon. 2010;59: 1326–1366.

9.   Rabosky DL, Santini F, Eastman J, Smith SA, Sidlauskas B, Chang J, et al.

Rates of speciation and morphological evolution are correlated across the

largest vertebrate radiation. Nat Commun. Nature Publishing Group; 2013;4:

1958. doi:10.1038/ncomms2958

10.   Coyne JA, Orr HA. Speciation. Sinauer Associates; 2004.

11.   Adams DC, Berns CM, Kozak KH, Wiens JJ. Are rates of species

diversification correlated with rates of morphological evolution? Proc R Soc B

Biol Sci. 2009;276: 2729–38. doi:10.1098/rspb.2009.0543

12.   Puttick MN, Clark J, Donoghue PCJ. Size is not everything: rates of genome

size evolution, not C-value, correlate with speciation in angiosperms. Proc Biol

Sci. 2015;282: 20152289-. doi:10.1098/rspb.2015.2289

13.   Cantalapiedra JL, Prado JL, Hernández Fernández M, Alberdi MT. Decoupled

ecomorphological evolution and diversification in Neogene-Quaternary horses.

470   Science (80-). 2017;355. Available:

471   http://science.sciencemag.org/content/355/6325/627/tab-pdf

472   14.   Pigliucci M. Is evolvability evolvable? Nat Rev Genet. Nature Publishing

473   Group; 2008;9: 75–82. doi:10.1038/nrg2278

474   15.   Ricklefs RE, Renner SS. Species richness within families of flowering plants.

475   Evolution (N Y). [Society for the Study of Evolution, Wiley]; 1994;48: 1619–

476   1636. doi:10.2307/2410252

477   16.   Rees M, Westoby M. Game-theoretical evolution of seed mass in multi-species

478   ecological models. Oikos. 1997;78: 116–126. doi:10.2307/3545807

479   17.   Zanne AE, Tank DC, Cornwell WK, Eastman JM, Smith SA, FitzJohn RG, et

480   al. Three keys to the radiation of angiosperms into freezing environments.

481   Nature. 2014;506: 89–92. doi:10.1038/nature12872

482   18.   Kew Seed Information Database. In: Royal Botanic Gardens Kew [Internet].

483   2016 [cited 1 Mar 2016]. Available: http://data.kew.org/sid/

484   19.   Rabosky DL. Automatic detection of key innovations, rate shifts, and diversity-

485   dependence on phylogenetic trees. PLoS One. Public Library of Science;

486   2014;9: e89543. doi:10.1371/journal.pone.0089543

487   20.   Rabosky DL, Huang H. A robust semi-parametric test for detecting trait-

488   dependent diversification. Syst Biol. Oxford University Press; 2016;65: 181–

489   93. doi:10.1093/sysbio/syv066

490   21.   Moore BR, Höhna S, May MR, Rannala B, Huelsenbeck JP. Critically

491   evaluating the theory and performance of Bayesian analysis of

492   macroevolutionary mixtures. Proc Natl Acad Sci. 2016;113.

493   doi:10.1073/pnas.1518659113

22.   Rabosky DL, Mitchell JS, Chang J. Is BAMM flawed? Theoretical and

practical concerns in the analysis of multi-rate diversification models. Syst

Biol. 2017; preprint. doi:10.1093/sysbio/syx037

23.   Magallón S, Sanderson MJ. Absolute diversification rates in angiosperm

clades. Evolution (N Y). 2001;55: 1762–1780. doi:10.1111/j.0014-

3820.2001.tb00826.x

24.   Morlon H, Lewitus E, Condamine FL, Manceau M, Clavel J, Drury J.

RPANDA: An R package for macroevolutionary analyses on phylogenetic

trees. Fitzjohn R, editor. Methods Ecol Evol. 2016;7: 589–597.

doi:10.1111/2041-210X.12526

25.   Blomberg SP, Garland T, Ives AR. Testing for phylogenetic signal in

comparative data: behavioral traits are more labile. Evolution (N Y). Blackwell

Publishing Ltd; 2003;57: 717–745. doi:10.1111/j.0014-3820.2003.tb00285.x

26.   Price SL, Etienne RS, Powell S. Tightly congruent bursts of lineage and

phenotypic diversification identified in a continental ant radiation. Evolution.

2016; doi:10.1111/evo.12894

27.   Lovette IJ, Bermingham E, Ricklefs RE. Clade-specific morphological

diversification and adaptive radiation in Hawaiian songbirds. Proc Biol Sci.

2002;269: 37–42. doi:10.1098/rspb.2001.1789

28.   Eldredge N, Gould SJ. Punctuated equilibria: an alternative to phyletic

gradualism. In: Schopf T, editor. Models in Paleobiology. San Francisco:

Freeman Cooper; 1972. pp. 82–115.

29.   Coomes DA, Grubb PJ. Colonization, tolerance, competition and seed-size

variation within functional groups. Trends Ecol Evol. 2003;18: 283–291.

doi:10.1016/S0169-5347(03)00072-7

30.   Kisel Y, Barraclough TG. Speciation has a spatial scale that depends on levels

of gene flow. Am Nat. 2010;175: 316–34. doi:10.1086/650369

31. Claramunt S, Derryberry EP, Remsen J V, Brumfield RT. High dispersal ability

inhibits speciation in a continental radiation of passerine birds. Proc Biol Sci.

2012;279: 1567–74. doi:10.1098/rspb.2011.1922

32. Beaulieu JM, Moles AT, Leitch IJ, Bennett MD, Dickie JB, Knight CA.

Correlated evolution of genome size and seed mass. New Phytol. 2007;173:

422–37. doi:10.1111/j.1469-8137.2006.01919.x

33. Smith SA, Donoghue MJ. Rates of Molecular Evolution Are Linked to Life

History in Flowering Plants. Science (80-). 2008;322: 86–89.

doi:10.1126/science.1163197

34. Lanfear R, Ho SYW, Love D, Bromham L. Mutation rate is linked to

diversification in birds. Proc Natl Acad Sci U S A. 2010;107: 20423–8.

doi:10.1073/pnas.1007888107

35. Lanfear R, Ho SYW, Jonathan Davies T, Moles AT, Aarssen L, Swenson NG, et

al. Taller plants have lower rates of molecular evolution. Nat Commun. Nature

Publishing Group, a division of Macmillan Publishers Limited. All Rights

Reserved.; 2013;4: 1879. doi:10.1038/ncomms2836

36. Beaulieu JM, O'Meara BC. Detecting hidden diversification shifts in models of

trait-dependent speciation and extinction. Syst Biol. 2016; syw022-.

doi:10.1093/sysbio/syw022

37. Hinchliff CE, Smith SA. Some limitations of public sequence data for

phylogenetic inference (in plants). Joly S, editor. PLoS One. Public Library of

Science; 2014;9: e98986. doi:10.1371/journal.pone.0098986

38. Title PO, Rabosky DL. Do Macrophylogenies Yield Stable Macroevolutionary

Inferences? An Example from Squamate Reptiles. Syst Biol. 2016;94: syw102.

doi:10.1093/sysbio/syw102

39.   Donoghue MJ, Sanderson MJ. Confluence, synnovation, and depauperons in

plant diversification. New Phytol. 2015;207: 260–274. doi:10.1111/nph.13367

40.   Cayuela L, Granzow-de la Cerda Í, Albuquerque FS, Golicher DJ. taxonstand:

An r package for species names standardisation in vegetation databases.

Methods Ecol Evol. 2012;3: 1078–1083. doi:10.1111/j.2041-

210X.2012.00232.x

41.   Pennell MW, FitzJohn RG, Cornwell WK. A simple approach for maximizing

the overlap of phylogenetic and comparative data. Methods Ecol Evol. 2015;

n/a-n/a. doi:10.1111/2041-210X.12517

42.   Stevens P. Angiosperm Phylogeny website. In: Angiosperm Phylogeny

Website. Version 12 [Internet]. 2012. Available:

http://www.mobot.org/MOBOT/research/APweb/

43.   Qian H, Jin Y. An updated megaphylogeny of plants, a tool for generating plant

phylogenies and an analysis of phylogenetic community structure. J Plant Ecol.

2015; rtv047. doi:10.1093/jpe/rtv047

44.   Rabosky DL, Grundler M, Anderson C, Title P, Shi JJ, Brown JW, et al.

BAMMtools: an R package for the analysis of evolutionary dynamics on

phylogenetic trees. Kembel S, editor. Methods Ecol Evol. 2014;5: 701–707.

doi:10.1111/2041-210X.12199

45.   Moore BR, Höhna S, May MR, Rannala B, Huelsenbeck JP. Critically

evaluating the theory and performance of Bayesian analysis of

macroevolutionary mixtures. Proc Natl Acad Sci. National Academy of

Sciences; 2016;113: 9569–9574. doi:10.1073/pnas.1518659113

46.   Wiens JJ. Explaining large-scale patterns of vertebrate diversity. Biol Lett. The

Royal Society; 2015;11: 20150506. doi:10.1098/rsbl.2015.0506

47.   Kozak KH, Wiens JJ. Testing the Relationships between Diversification,

Species Richness, and Trait Evolution. Syst Biol. Oxford University Press;

2016;65: 975–988. doi:10.1093/sysbio/syw029

48.   Pennell MW, Eastman JM, Slater GJ, Brown JW, Uyeda JC, Fitzjohn RG, et al.

Geiger v2.0: An expanded suite of methods for fitting macroevolutionary

models to phylogenetic trees. Bioinformatics. Oxford University Press;

2014;30: 2216–2218. doi:10.1093/bioinformatics/btu181

49.   Orme D. The caper package: comparative analysis of phylogenetics and

evolution in R. R Packag version. 2013;5.

50.   Bennett M, Leitch IJ. Plant DNA C-values database (release 6.0, Dec 2012)

[Internet]. 2012. Available: http://www.kew.org/cvalues/

51.   Zanne AE, Tank DC, Cornwell WK, Eastman JM, Smith SA, FitzJohn RG, et

al. Data from: Three keys to the radiation of angiosperms into freezing

environments. 2014; doi:10.5061/DRYAD.63Q27.2

52.   Kattge J, Diaz S, Lavorel S, Prentice IC, Leadley P, Bonisch G, et al. TRY - a

global database of plant traits. Glob Chang Biol. Blackwell Publishing Ltd;

2011;17: 2905–2935. doi:10.1111/j.1365-2486.2011.02451.x

53.   Revell LJ. phytools: An R package for phylogenetic comparative biology (and

other things). Methods Ecol Evol. Blackwell Publishing Ltd; 2012;3: 217–223.

doi:10.1111/j.2041-210X.2011.00169.x

54.   Winkler DW, Hall C. Do big plants have big seeds? Am Nat. 1989;130: 526–

543. doi:10.2307/2678832

Supporting Information captions

Fig. S1. Phylogenetic tree of 13,577 angiosperm species with branch colours indicating the rate of seed mass evolution estimated with BAMM. Branches were scaled by speciation rate as determined by a BAMM analysis on a larger 19,703 tree.

Fig. S2. Prior and posterior distribution of the number of rate shifts in BAMM for a) the speciation/extinction and b) phenotypic evolution analyses for *expectedNumberOfShifts* = 25, 50 and 100 and 250. The analyses in the main text were carried out with *expectedNumberOfShifts* = 50 for both speciation/extinction and phenotypic evolution analyses.

Fig. S3. Seed mass and its rate of evolution are associated with speciation in the clade-based analyses. (a) PGLS slope of the relationship between speciation rate (λ) from the method-of-moments estimator and the rate of seed mass evolution across 10 time slices. Circles are scaled to the number of clades in each time slice while colour indicates the significance of the slope. (b) PGLS slope of the relationship between speciation rate and mean clade seed mass. For a detailed representation of the results in each time slice, see figure S14 and S15.

Fig. S4. STRAPP correlations of diversification and phenotypic traits for 1,007 angiosperm species. The distribution of the absolute difference in the observed correlation minus the null correlation is plotted for each trait. The coloured dotted lines indicate the mean of that distribution, and the black dotted line indicates 0; a

distribution with mean = 0 would show no association between a focal trait and

macroevolutionary dynamics. STRAPP correlation of seed size (shown in red), C-

value (shown in yellow), life cycle (shown in light green), woodiness (shown in dark

green), height (shown in blue), seed size rate (shown in dark blue), C-value rate

(shown in purple); and height rate (shown in pink) with a) speciation rate (λ), b)

extinction rate (μ), and c) net diversification rate (r).

Fig. S5. Proposed effects of seed mass and other life history traits on diversification

(solid lines). Dashed lines indicate correlations between life history traits. Numbers

indicate reference where the link is proposed.

Fig. S6. Phylogenetic tree of 353 angiosperm families with representatives in our

analyses. The red bars indicate the levels of sampling for each family.

Figure S7. Angiosperm phylogenetic tree collapsed to monophyletic clades of ≤6000

species. The name of one representative species per clade is shown, and the numbers

in parentheses indicate the number of species included in each clade. The BAMM

analyses were carried out for six monophyletic clades (shown in red, yellow, green,

blue, dark blue and pink) and one "backbone" analysis with the remaining clades

(shown in grey) and one representative of each of the six monophyletic clades.

Fig. S8. Comparison of the speciation rates at the tip of the tree obtained with the

complete Zanne tree (29,703 species) and the seed size filtered tree (13,577 species).

The dotted line represents the 1:1 reference line.

Figure S9. Correlation of speciation with seed mass and seed mass rate of evolution

in the clade-based analysis only considering congeneric species. (a) PGLS slope of the

relationship of speciation rate - estimated with the method-of-moments estimator -

with mean clade seed mass across 10 time slices. The size of the circles represents the

number of clades in each time slice while the colour indicates the significance of the

slope. (b) PGLS slope of the relationship of speciation rate and the rate of seed mass

evolution.

Fig. S10. Correlation between speciation rate and rate of seed size evolution in a

random sample of the BAMM posterior. The dotted line represents the Spearman

correlation ($\rho = 0.47$, p-value < 0.001)

Fig. S11. Type I error analysis. We estimated the type I error rate of our analysis by

simulating neutral traits on the angiosperm phylogenetic tree. We performed 1,000

simulations and then ran 1,000 STRAPP tests with each simulated dataset. We

estimated the corresponding p-values for the association between traits and

diversification and calculated the type I error as the proportion of datasets where a

significant association (p-value < 0.05) was detected.

Fig. S12. Correlation of mean speciation and mean phenotypic rates across all the

branches of the angiosperm phylogenetic tree. The dotted line is the ordinary least

squares regression ($R^2 = 0.31$, p-value < 0.001)

Fig. S13. Correlations between clade rate of seed size evolution and speciation rate

(estimated with RPANDA) across time slices from: (a) 0 to 2 million years (myr); (b)

2 to 4 myr; (c) 4 to 6 myr; (d) 6 to 8 myr; (e) 8 to 10 myr; (f) 10 to 12 myr; (g) 12 to

14 myr; (h) 14 to 16 myr; (i) 16 to 18 myr; and (j) 18 to 20 myr. The degrees of

freedom (df) are equivalent to the number of clades minus one.

Fig. S14. Correlations between mean clade seed mass and speciation rate (estimated

with RPANDA) across time slices from: (a) 0 to 2 million years (myr); (b) 2 to 4 myr;

(c) 4 to 6 myr; (d) 6 to 8 myr; (e) 8 to 10 myr; (f) 10 to 12 myr; (g) 12 to 14 myr; (h)

14 to 16 myr; (i) 16 to 18 myr; and (j) 18 to 20 myr. The degrees of freedom (df) are

equivalent to the number of clades minus one.

Fig. S15. Correlation of (a) rate of seed mass evolution and (b) seed mass with net

diversification rate (r) estimated using RPANDA in the clade-based analysis. The

strength of correlations is shown as PGLS slopes and was calculated using mean

clade-level seed mass across 10 time slices. The size of the circles represents the

number of clades in each time slice while the colour indicates the significance of the

slope.

Fig. S16. Mean genus seed mass of strict annual (n = 214) and perennial (n = 793)

genera. No significant difference between the means of the two groups was found

when accounting for phylogeny (phylANOVA: p-value = 0.308, significance assessed

with 1,000 random simulations).

Fig. S17. Correlations between clade rate of seed size evolution and speciation rate

(estimated with the method-of-moments estimator) across time slices from: (a) 0 to 2

million years (myr); (b) 2 to 4 myr; (c) 4 to 6 myr; (d) 6 to 8 myr; (e) 8 to 10 myr; (f)

10 to 12 myr; (g) 12 to 14 myr; (h) 14 to 16 myr; (i) 16 to 18 myr; and (j) 18 to 20

myr.

Fig. S18. Correlations between mean clade seed mass and speciation rate (estimated

with the method-of-moments estimator) across time slices from: (a) 0 to 2 million

years (myr); (b) 2 to 4 myr; (c) 4 to 6 myr; (d) 6 to 8 myr; (e) 8 to 10 myr; (f) 10 to 12

myr; (g) 12 to 14 myr; (h) 14 to 16 myr; (i) 16 to 18 myr; and (j) 18 to 20 myr.

Table S1. RPANDA diversification models for the clade-based analyses. For each 2-

million year (myr) time slice, we counted the number of clades where the best-fitting

model was either i) birth-death model with constant λ (speciation) and μ (extinction)

(lambda.cst.mu.cst); pure birth model with constant λ (lambda.cst.mu0); pure birth

model with exponential λ (lambda.exp.mu.0); birth-death model with exponential λ

and constant μ (lambda.exp.mu.cst); birth-death model with exponential λ and

exponential μ (lambda.exp.mu.exp); or birth-death model with constant λ and

exponential μ (lambda.cst.mu.exp).

Table S2. Correlations of seed size and other phenotypic traits. Trait values were

obtained from a 1,007 species tree where all species had data for seed size, C-value

and plant height. The values are the slopes of the PGLS regressions and asterisks

denote statistically significant correlations (p-value < 0.05).

Table S3. Correlations of seed size rate of evolution and other phenotypic rates of

evolution. Rate values were obtained from a 1,007 species tree where all species had

data for seed size, C-value and plant height. The values are the slopes of the PGLS

regressions and asterisks denote statistically significant correlations (p-value < 0.05).

Table S4. STRAPP correlations (*rho*) for 1,007 species of angiosperms with seed size,

genome size (i.e., C-value), life cycle, height, woodiness data and rates of seed size,

C-value and height evolution. Significant correlation are shown in bold and p-values

are shown in parentheses.