bioRxiv preprint doi: https://doi.org/10.1101/059691; this version posted April 10, 2017. The copyright holder for this preprint (which was not certified by peer review) is the author/funder, who has granted bioRxiv a license to display the preprint in perpetuity. It is made available under aCC-BY-NC-ND 4.0 International license.

NAMMCO JCNB/SWG/2017-JWG

Meta population modelling of narwhals in East Canada and West Greenland - 2017

Lars Witting
Greenland Institute of Natural Resources,
P. O. Box 570, DK-3900 Nuuk, Greenland. Email: lawi@natur.gl

ABSTRACT
This paper combines a catch allocation model for narwhals in East Canada and West Greenland with Bayesian population modelling of the eight summer aggregations of narwhals in the region. The catch allocation model allocates the catches in different hunting areas and seasons to the different summer aggregations, and the population models analyse the impact of these catches on the population dynamics of the eight narwhal aggregations.

The population models run from 1970, and the catch allocation model needs population trajectories from 1970 to the present in order to estimate the catches taken from the different summer aggregations during this period. In an initial run it uses linear transitions between the available abundance estimates; but more elaborate population trajectories are estimated by the fit of the population models to the abundance data. The two models are therefore run in an iterative manner until the catch histories that are estimated by the allocation model, and the abundance trajectories that are estimated by the population models, converge between runs.

Given a converged model and potential future catch options for the different hunts, the model estimates the probabilities of fulfilling management options for the eight summer aggregations of narwhals.

INTRODUCTION

This paper develops a meta population dynamic model for eight summer aggregations of narwhals in East Canada and West Greenland. It combines a catch allocation model that was developed at the 2014 meeting of the JWG (JWG 2015) with eight population dynamic models, in the aim of assessing the population dynamic implications of hunting.

One of the main challenges in this assessment is the divergence between the population structure of narwhals and the geographical and yearly structure of the hunt. Narwhals return to specific summering grounds and their population structure is best described by the geographical distinctiveness of the different summering grounds; a view that has identified eight populations of narwhal in the Baffin Bay region in East Canada and West Greenland (JWG 2015).

These populations, however, are hunted not only on the summering grounds, but also on their spring and fall migrations in other areas, as well as on the wintering grounds; and these latter areas are typically shared to some degree by individuals from different summer ground.

This lead to the development of a catch allocation model that uses satellite tracking,

bioRxiv preprint doi: https://doi.org/10.1101/059691; this version posted April 10, 2017. The copyright holder for this preprint (which was not certified by peer review) is the author/funder, who has granted bioRxiv a license to display the preprint in perpetuity. It is made available under aCC-BY-NC-ND 4.0 International license.

Hunt	Season	Smith	Jones	Inglefield	Melville	Somerset	Admiralty	Eclipse	Baffin
Etah	Spring	1	$0/n$	0	0	0	0	0	0
Qaanaaq	Summer	0	0	1	0	0	0	0	0
Grise Fjord	Spring	$0/n$	1	$0/n$	0	$0/n$	0	0	0
Grise Fjord	Summer	0	1	0	0	0	0	0	0
Grise Fjord	Fall	$0/n$	1	$0/n$	0	$0/n$	0	0	0
Upernavik	Summer	0	0	0	1	0	0	0	0
Ummannaaq	Fall	$0/n$	$0/n$	$0/n$	1/9	1	0/42	0/26	$0/n$
Disko Bay	Winter	$0/n$	$0/n$	$0/n$	1/7	$0/n$	1/42	1/6	$0/n$
CCA	Spring	0	0	0	0	1	0/4	0/5	0
CCA	Summer	0	0	0	0	1	0	0	0
CCA	Fall	0	0	0	0	1	7/42	1/26	0
Arctic Bay	Spring	0	0	0	0	1	1	1/5	0
Arctic Bay	Summer	0	0	0	0	0	1	0	0
Arctic Bay	Fall	0	0	0	0	$0/n$	1	6/26	0
Pond Inlet	Spring	0	$0/n$	$0/n$	0	2/2	4/4	1	$0/n$
Pond Inlet	Summer	0	0	0	0	0	0	1	0
Pond Inlet	Fall	0	$0/n$	$0/n$	0	0/14	4/42	1	$0/n$
BIC	Spring	0	$0/n$	$0/n$	0	0/2	0/4	0/6	1
BIC	Summer	0	0	0	0	0	0	0	1
BIC	Fall	0	$0/n$	$0/n$	0	0/5	10/42	16/26	1
BIS	Spring	0	0	0	0	0/2	0/4	0/6	n/n
BIS	Summer	0	0	0	0	0	0	0	1
BIS	Fall	0	0	0	0	0/5	0/42	2/26	n/n
BIS	Winter	0	0	0	0	0/2	0/42	1/6	n/n

Table 1: The proportional availability matrix ($\mathbf{P}$) of narwhals from summer aggregations to hunting regions [x/y: available (x) over total (y)]. Black numbers (defined zeros and hunts) are fixed, blue (partial hunts) and red (probable zeros and hunts) are beta(α, β) distributions ($\alpha = x+1$; $\beta = y - x + 1$); red for sensitivity by changes in n.

expert judgment, and abundance estimates to allocate the catches from 24 seasonal and geographically separated hunts between the eight populations of narwhals (JWG 2015). I use this model to produce historical catch histories for the eight summer aggregations, and this allows me to construct population dynamic models that assess the implications of the hunt on the eight summer aggregations. The catch allocation and population dynamics models are then integrated to provide a tool that can assess the sustainability of potential future catches in the 24 hunts, as defined by the impacts of the hunts on the eight summer aggregations of narwhals.

CATCH ALLOCATION MODEL

The catch allocation model was developed by JWG (2015), and it allocates catches taken in different hunting regions and seasons to the eight summer aggregations of narwhals in East Canada and West Greenland.

The model uses a proportional availability matrix (Table 2 & 1) to describe the availability of the narwhals from the different summer aggregations to the hunts in the different regions and seasons, and it uses a catch matrix (Tables 10 and 11) to describe the annual total removals (catches plus loss) in the different hunts.

To allocate the catches from the different hunts to the different summer aggregations, the model needs an additional matrix that describes the abundance in the different stocks per year. These abundance estimates are needed to estimate the relative availability of the different stocks to the different hunts, so that the catches can be allocated to the summer

bioRxiv preprint doi: https://doi.org/10.1101/059691; this version posted April 10, 2017. The copyright holder for this preprint (which was not certified by peer review) is the author/funder, who has granted bioRxiv a license to display the preprint in perpetuity. It is made available under aCC-BY-NC-ND 4.0 International license.

Hunt	Season	Smith	Jones	Inglefield	Melville	Somerset	Admiralty	Eclipse	Baffin
Etah	Spring	1.00	0.00	0.00	0.00	0.00	0.00	0.00	0.00
Qaanaaq	Summer	0.00	0.00	1.00	0.00	0.00	0.00	0.00	0.00
Grise Fjord	Spring	0.00	1.00	0.00	0.00	0.00	0.00	0.00	0.00
Grise Fjord	Summer	0.00	1.00	0.00	0.00	0.00	0.00	0.00	0.00
Grise Fjord	Fall	0.00	1.00	0.00	0.00	0.00	0.00	0.00	0.00
Upernavik	Summer	0.00	0.00	0.00	1.00	0.00	0.00	0.00	0.00
Ummannaaq	Fall	0.00	0.00	0.00	0.11	1.00	0.00	0.00	0.00
Disko Bay	Winter	0.00	0.00	0.00	0.14	0.00	0.02	0.17	0.00
CCA	Spring	0.00	0.00	0.00	0.00	1.00	0.00	0.00	0.00
CCA	Summer	0.00	0.00	0.00	0.00	1.00	0.00	0.00	0.00
CCA	Fall	0.00	0.00	0.00	0.00	1.00	0.17	0.04	0.00
Arctic Bay	Spring	0.00	0.00	0.00	0.00	1.00	1.00	0.2	0.00
Arctic Bay	Summer	0.00	0.00	0.00	0.00	0.00	1.00	0.00	0.00
Arctic Bay	Fall	0.00	0.00	0.00	0.00	0.00	1.00	0.23	0.00
Pond Inlet	Spring	0.00	0.00	0.00	0.00	1.00	1.00	1.00	0.00
Pond Inlet	Summer	0.00	0.00	0.00	0.00	0.00	0.00	1.00	0.00
Pond Inlet	Fall	0.00	0.00	0.00	0.00	0.00	0.1	1.00	0.00
BIC	Spring	0.00	0.00	0.00	0.00	0.00	0.00	0.00	1.00
BIC	Summer	0.00	0.00	0.00	0.00	0.00	0.00	0.00	1.00
BIC	Fall	0.00	0.00	0.00	0.00	0.00	0.24	0.62	1.00
BIS	Spring	0.00	0.00	0.00	0.00	0.00	0.00	0.00	1.00
BIS	Summer	0.00	0.00	0.00	0.00	0.00	0.00	0.00	1.00
BIS	Fall	0.00	0.00	0.00	0.00	0.00	0.00	0.08	1.00
BIS	Winter	0.00	0.00	0.00	0.00	0.00	0.00	0.17	1.00

Table 2: The proportional availability matrix ($\mathbf{P}$) of narwhals from summer aggregations to hunting regions. Black numbers are fixed, blue and red are point estimates of beta distributions; red for sensitivity only.

aggregations.

Allocation matrix

The allocation model was developed in the form of a 24 rows by 8 columns allocation matrix. The eight columns are the individual summer aggregations, and the 24 rows represent hunts divided by 10 regions with some of the hunts divided by season. For each summer aggregation and hunt there is a cell in the matrix, and the matrix is devised so that when multiplied by a vector of removals, the resulting vector will determine the total removals from each summer aggregation.

Each cell of the allocation matrix, $\mathbf{A}$, has a value

$$\mathbf{A}_{i,j,t} = \frac{\mathbf{P}_{i,j} N_{i,t}}{\sum_i \mathbf{P}_{i,j} N_{i,t}} \tag{1}$$

where, $\mathbf{A}_{i,j,t}$ is the proportion of the jth hunt that is assigned to the ith summer aggregation in year t, $\mathbf{P}_{i,j}$ is the proportional availability of the ith summer aggregation to the jth hunt, and $N_{i,t}$ is the abundance of the ith summer aggregation in year t.

This model assumes that for each summer aggregation there is a proportion between zero and one, $\mathbf{P}_{i,j}$, that is available to hunters during the hunting period on the hunting grounds. Each individual that is available is then at equal risk of being taken in the hunt. The sum of the $\mathbf{A}_{i,j,t}$ should be 1 for each row of $\mathbf{A}_t$.

To set up the proportional availability matrix ($\mathbf{P}$; Table 2 & 1) the JWG reviewed each cell so that each cell in the matrix was given one of five designations:

bioRxiv preprint doi: https://doi.org/10.1101/059691; this version posted April 10, 2017. The copyright holder for this preprint (which was not certified by peer review) is the author/funder, who has granted bioRxiv a license to display the preprint in perpetuity. It is made available under aCC-BY-NC-ND 4.0 International license.

Year	Smith	Jones	Inglefield	Melville	Somerset	Admiralty	Eclipse	Baffin
1975	-	-	-	-	-	28260; 0.22	-	-
1981	-	-	-	-	32520; 0.1	-	-	-
1985	-	-	-	-	-	16400; 0.43	-	-
1986	-	-	8710; 0.25	-	-	-	-	-
1996	-	-	-	-	45360; 0.35	-	-	-
2001	-	-	-	-	-	-	-	-
2002	-	-	-	-	35810; 0.43	-	-	-
2003	-	-	-	-	-	5360; 0.5	-	10070; 0.31
2004	-	-	-	-	-	-	20230; 0.36	-
2007	-	-	8370; 0.25	6020; 0.86	-	-	-	-
2009	-	-	-	-	-	-	-	-
2010	-	-	-	-	-	18050; 0.22	-	-
2012	-	-	-	2980; 0.39	-	-	-	-
2013	16360; 0.65	12690; 0.33	-	-	49770; 0.2	35040; 0.42	10490; 0.24	17560; 0.35
2014	-	-	-	3090; 0.5	-	-	-	-

Table 3: Abundance estimates with CVs for summer aggregations of narwhal.

Defined zero: This designated cells that represented improbable situations such as a summer harvest that was not at a summering ground (e.g. a narwhal harvested in summer in Resolute could not come from the Smith Sound summer aggregation and would be assigned to Somerset Island stock) and to hunts in areas that could not have originated in a particular summering ground based on known movements.

Probable zero: This designated cells in which a summering aggregation was unlikely to be hunted but proximity during the hunting season, or a presumed migration route did not rule out possible catches and designated cells with no tag data.

Partial hunt: This designated cells with tag data showing a portion of the summering aggregation was available or not to hunters.

Probable hunt: This designated cells in which a summering aggregation was likely to be fully available to a hunt, based on its geographical proximity to a summering ground or migration route, but for which there was no quantitative evidence such as tag data.

Defined hunt: This designates cells representing hunts on summering grounds or known wintering areas of stocks.

Each cell in $\mathbf{P}$ was then parameterised as a beta(α, β) distribution, where $\alpha = x + 1$ and $\beta = y - x + 1$ would depend on the designation of the hunt. Defined zeros where obtained by $x = 0$ and $y = 9999$, defined hunts by $x = 9999$ and $y = 9999$, probable zeros by $x = 0$ and $y = n$, probable hunts by $x = y = n$, and for a partial hunt that was parameterised by the seasonal and geographical distribution of y sattelited trackes individuals, x would be the number of the individuals that migrated to the hunting ground. The simulations in this paper, however, set $n = 9999$ which implies that they do not distinguish between probable and defined designations.

The generation of catch histories

The abundance matrix in the initial run of the model is constructed as linear transitions between the abundance estimates in the abundance estimate matrix of Table 3. In subsequent runs, the abundance matrix is given by the abundance trajectories that the previous

bioRxiv preprint doi: https://doi.org/10.1101/059691; this version posted April 10, 2017. The copyright holder for this preprint (which was not certified by peer review) is the author/funder, who has granted bioRxiv a license to display the preprint in perpetuity. It is made available under aCC-BY-NC-ND 4.0 International license.

runs of the population dynamic models are estimating for the different summer aggregations of narwhals, given the catch histories that were estimated by the previous run of the allocation model. This iterative running of the two models was then conducted three to five times to ensure convergence of the catch histories and abundance trajectories.

In either case, for a given year t, we would have abundance estimates with cv's for all summer aggregations. Distributions of catch histories for the summer aggregations during these iterations were generated by a large number of runs of the allocation model over the entire catch history with yearly random draws of each abundance estimate (log-normal based on point estimate and cv) and a $\mathbf{P}_t$ matrix drawn from the underlying beta distributions.

POPULATION DYNAMIC MODELS

Separate population models with density regulated growth were constructed for each of the eight summer aggregations of narwhals. All the models were based on an age and sex structured Bayesian modelling framework that have been used in earlier applications for walrus (Witting and Born 2005, 2013), large cetaceans (Witting 2013), beluga (Heide-Jørgensen et al. In press), and narwhal (e.g., Witting and Heide-Jørgensen 2012a,b,c).

Some of the summering aggregations have only been surveyed once, and a full age-structured model is clearly over parameterized for these cases if the main purpose was parameter estimation by maximum likelihood. Yet, the main purpose here is instead to use a Bayesian framework to integrate prior knowledge on the life history biology of narwhals with survey estimates of abundance for the construction of realistic population dynamic models.

Witting (2009) used the case of beluga in West Greenland to analyse for influence of model uncertainty in the construction of realistic population models in Bayesian assessments of density regulated growth. Assessments were made for one age-structured and four structurally different discrete models, with all assessments being based on the same data. All models gave very similar estimates of current abundance and current production levels, making the choice of model basically a matter of taste.

I have chosen an age and sex structured framework that allows the model to be constructed directly from our prior knowledge of the life history of narwhals. This allows also for a later inclusion of sex structured catches, and age-structured catch data, as it has been done in earlier assessments of Greenland narwhals (Witting and Heide-Jørgensen 2012a,b,c). Age-structured data, however, were not included here, as I wanted to keep the population dynamic models relatively simple (and fast to simulate) in this first run of the meta model.

Let $x = 15$ be the maximum lumped age-class of the model. Let the number $N^{m/f}_{a,t+1}$ of males (m) and females (f) in age-classes $0 < a < x$ in year $t + 1$ be

$$N^{m/f}_{a+1,t+1} = p^{m/f}_a N^{m/f}_{a,t} - c^{m/f}_{a,t} \tag{2}$$

and the number of animals in age-class x be

$$N^{m/f}_{x,t+1} = p^{m/f}_x N^{m/f}_{x,t} + p^{m/f}_{x-1} N^{m/f}_{x-1,t} - c^{m/f}_{x,t} - c^{m/f}_{x-1,t} \tag{3}$$

where $p^{m/f}_a$ is the age specific survival rate of males/females, and $c^{m/f}_{a,t}$ is the age specific catch of males/females in year t. The age and gender (g) dependent survival rates $p^g_a = p\tilde{p}^g_a$

bioRxiv preprint doi: https://doi.org/10.1101/059691; this version posted April 10, 2017. The copyright holder for this preprint (which was not certified by peer review) is the author/funder, who has granted bioRxiv a license to display the preprint in perpetuity. It is made available under aCC-BY-NC-ND 4.0 International license.

are given as a product between a survival scalar p and a relative ($0 < \tilde{p}_a^g \leq 1$) survival rate, with relative survival being one for males and females older than one year of age. The age and gender specific catches $c_{a,t}^{m/f} = c_t^{m/f} \tilde{c}_a^{m/f}$ in year t is given as a product between the total catch of males/females ($c_t^{m/f}$), as specified by the catch history, and an age-specific catch selectivity ($\tilde{c}_a^{m/f}$) that is uniform except that no animals are taken from age-class zero.

The number of females and males in age-class zero is $N_{0,t}^f = \vartheta N_{0,t}$ and $N_{0,t}^m = (1 - \vartheta) N_{0,t}$, where ϑ is the fraction of females at birth, and

$$N_{0,t} = \sum_{a=a_m}^{x} B_{a,t} \tag{4}$$

where a_m is the age of the first reproductive event and $B_{a,t}$, the number of births from females in age class a, is

$$B_{a,t} = b_{a,t}\, \tilde{b}_a\, M_{a,t}^f \tag{5}$$

where $b_{a,t}$ is the birth rate in year t for age-class a females should they be at their age-specific reproductive peak, $0 < \tilde{b}_a \leq 1$ is the relative age-specific birth rate (1 for all mature females), and $M_{a,t}^f$ is the number of mature females in age-class a in year t, defined as

$$M_{a,t}^f = \begin{cases} 0 & \text{if } a < a_m \\ N_{a,t}^f & \text{if } a \geq a_m \end{cases} \tag{6}$$

Let $b_{a,t}$ be

$$b_{a,t} = \quad b^* + [b_{max} - b^*][1 - (\hat{N}_t / \hat{N}^*)^\gamma] \quad \text{for density regulated growth} \tag{7}$$

where b^* is the birth rate at population dynamic equilibrium (assuming zero catch and equilibrium denoted by $*$), b_{max} is the maximal birth rate, γ is the density dependence parameter, and the abundance component that imposes density dependence is the one-plus component

$$\hat{N}_t = \sum_{a=1}^{x} N_t^f + N_t^m \tag{8}$$

Given a stable age-structure and no catch, let, for a traditional model of exponential or density regulated growth, λ be a constant defined by $\hat{N}_{t+1} = \lambda \hat{N}_t$. The sustainable yield is then $\text{sy} = \hat{N}(\lambda - 1)$, and for the density regulated model there is an optimum $\partial \text{sy} / \partial \hat{N} = 0$; the maximum sustainable yield (msy) at $\hat{N}_{\text{msy}}$, also known by the maximum sustainable yield rate ($\text{msyr} = \text{msy}/\hat{N}_{\text{msy}}$) and the maximum sustainable yield level ($\text{msyl} = \hat{N}_{\text{msy}}/\hat{N}^*$).

STATISTICAL METHODS

The assessment model was fitted to data by projecting the population under the influence of the historical catches. A Bayesian statistical method (e.g, Berger 1985; Press 1989) was used, and posterior estimates of model parameters and other management related outputs were calculated. This implied an integration of the product between a prior distribution for each parameter and a likelihood function that links the probability of the data to the different parameterisations of the model.

bioRxiv preprint doi: https://doi.org/10.1101/059691; this version posted April 10, 2017. The copyright holder for this preprint (which was not certified by peer review) is the author/funder, who has granted bioRxiv a license to display the preprint in perpetuity. It is made available under aCC-BY-NC-ND 4.0 International license.

M	N_0	N^*	p	p_0	b	a_m	ϑ	γ	c_h	β_a
smith	-	$2{,}80^U$	${}^{2,2}_{.95,1}{}^b$	$.5{,}1^u$	.33	${}^{2,2}_{7,15}{}^b$	.5	$2{,}4^u$	f	-
jones	-	$2{,}60^U$	${}^{2,2}_{.95,1}{}^b$	$.5{,}1^u$	.33	${}^{2,2}_{7,15}{}^b$	.5	$2{,}4^u$	f	-
ingle	$1{,}25^U$	$3{,}30^U$	${}^{2,2}_{.95,1}{}^b$	$.5{,}1^u$	.33	${}^{2,2}_{7,15}{}^b$	.5	$2{,}4^u$	f	$.01{,}1^U$
melvi	$.8{,}20^U$	$3{,}30^U$	${}^{2,2}_{.95,1}{}^b$	$.5{,}1^u$	.33	${}^{2,2}_{7,15}{}^b$	.5	$2{,}4^u$	f	-
somer	$5{,}60^U$	$25{,}90^U$	${}^{2,2}_{.95,1}{}^b$	$.5{,}1^u$	.33	${}^{2,2}_{7,15}{}^b$	.5	$2{,}4^u$	f	-
admir	-	$10{,}40^U$	${}^{2,2}_{.95,1}{}^b$	$.5{,}1^u$	.33	${}^{2,2}_{7,15}{}^b$	.5	$2{,}4^u$	f	-
eclip	-	$5{,}50^U$	${}^{2,2}_{.95,1}{}^b$	$.5{,}1^u$	.33	${}^{2,2}_{7,15}{}^b$	.5	$2{,}4^u$	f	-
baffi	-	$3{,}60^U$	${}^{2,2}_{.95,1}{}^b$	$.5{,}1^u$	.33	${}^{2,2}_{7,15}{}^b$	.5	$2{,}4^u$	f	-

Table 4: **Prior distributions** for the different models (M). The list of parameters: N_0 is the initial abundance, N^* the population dynamic equilibrium abundance, p the yearly survival, p_0 the first year survival, b the birth rate, a_m the age of the first reproductive event, ϑ the female fraction at birth, γ the density regulation, c_h the catch history, and β_i the abundance estimate bias (i: data reference). Abundance is given in thousands. The prior probability distribution is given by superscripts; p: fixed value, u: uniform (min,max), U: log uniform (min,max), b: beta ${}^{a,b}_{i,x}$) with i=min and x=max, and f: file distribution.

Prior distributions

All models had the same priors on the biological parameters (see Table 4), and they were all initiated in 1970. All the summer aggregations with only one or two abundance estimates available (Smith, Jones, Eclipse, Baffin) and Admiralty seems to have had a very low exploitation rate in the beginning of the period, so for these I assumed that the population was close to the carrying capacity in 1970. For the remaining aggregations (Inglefield, Melville and Somerset), with a somewhat larger early exploitation, I assumed that the abundance in 1970 was lower than the carrying capacity.

The catch histories in a run of the population models were estimated by the allocation model over the complete catch history starting in 1970. These distributions were then used as catch priors in the population models, with specific catch histories being drawn from the prior for each iteration of a population model.

These catch priors were constructed from the distribution of possible total removals that was estimated for 2011 by 200 random draws of the allocation model (Figure 1), together with two complete catch histories, a minimum catch history (c_{min}, represented by the 1th percentile of this distribution over time) and a maximum catch history (c_{max}, represented by the 99th percentile). These percentiles were estimated from 1000 random draws of the allocation model over the entire catch history. The 2011 distribution was then rescaled to run from zero to one, with a value (x) drawn at random from the distribution for each parameterisation of a population model, with the catch history calculated as $c_t = c_{min,t} + x(c_{max,t} - c_{min,t})$.

Bayesian integration

The Bayesian integration was obtained by the sampling-importance-resampling routine (Jeffreys 1961; Berger 1985; Rubin 1988), where n_s random parameterisations θ_i ($1 \leq i \leq n_1$) are sampled from an importance function $h(\theta)$. This function is a probability distri-

bioRxiv preprint doi: https://doi.org/10.1101/059691; this version posted April 10, 2017. The copyright holder for this preprint (which was not certified by peer review) is the author/funder, who has granted bioRxiv a license to display the preprint in perpetuity. It is made available under aCC-BY-NC-ND 4.0 International license.

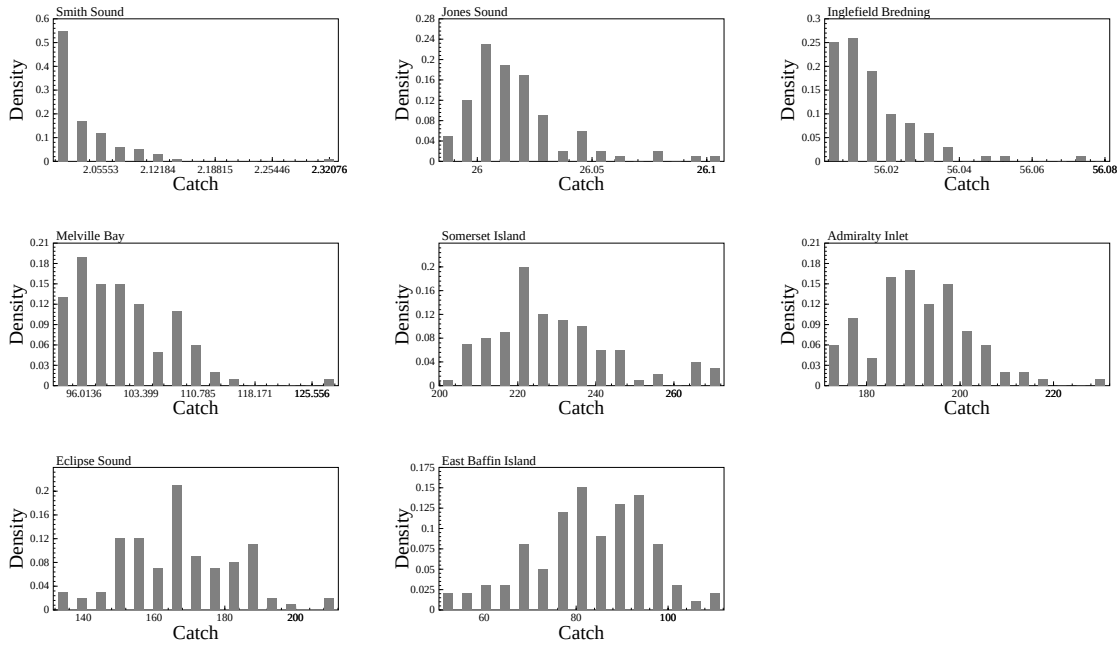

Figure 1: Catch distributions per summer aggregation, estimates for year 2011.

bution function from which a large number, n_s, of independent and identically distributed draws of θ can be taken. $h(\theta)$ shall generally be as close as possible to the posterior, however, the tails of $h(\theta)$ must be no thinner (less dense) than the tails of the posterior (Oh and Berger 1992). For each drawn parameter set θ_i the population was projected from the first year with a harvest estimate to the present. For each draw an importance weight, or ratio, was then calculated

$$w(\theta_i) = \frac{L(\theta_i)p(\theta_i)}{h(\theta_i)} \tag{9}$$

where $L(\theta_i)$ is the likelihood given the data, and $h(\theta_i)$ and $p(\theta_i)$ are the importance and prior functions evaluated at θ_i. In the present study the importance function is set to the joint prior, so that the importance weight is given simply by the likelihood. The n_s parameter sets were then re-sampled n_r times with replacement, with the sampling probability of the ith parameter set being

$$q_i = \frac{w(\theta_i)}{\sum_{j=1}^{n_s} w(\theta_j)} \tag{10}$$

This generates a random sample of the posterior distribution of size n_r.

The method of de la Mare (1986) was used to calculate the likelihood L under the assumption that observation errors are log-normally distributed (Buckland 1992)

$$L = \prod_i \prod_t \exp\left(-\frac{[\ln(\hat{N}_{i,t}/\beta_i N_t)]^2}{2\mathrm{CV}_{i,t}^2}\right) / \mathrm{CV}_{i,t} \tag{11}$$

where $\hat{N}_{i,t}$ is the point estimate of the ith set of abundance data in year t, $\mathrm{CV}_{i,t}$ is the coefficient of variation of the estimate, N_t is the simulated abundance, and β_i a bias term with is set to one for absolute abundance estimates.

bioRxiv preprint doi: https://doi.org/10.1101/059691; this version posted April 10, 2017. The copyright holder for this preprint (which was not certified by peer review) is the author/funder, who has granted bioRxiv a license to display the preprint in perpetuity. It is made available under aCC-BY-NC-ND 4.0 International license.

Probabilities of meeting objectives

In order to assess the sustainability of future catches on the eight populations of narwhals, I used the Bayesian framework to estimate the probabilities that an assumed management objective would be fulfilled for potential future catches from each population.

Given future annual catches c in the period 2015 to 2020, I applied the objective

$$ob = \begin{cases} N_{2020} > N_{2015} & \text{if } N_{2015} < N^* \text{ msyl} \\ c \leq 0.9 \text{ msy} & \text{if } N_{2015} \geq N^* \text{ msyl} \end{cases} \tag{12}$$

Given the population dynamic model and the data, the probability of meeting the objective ob is straightforwardly calculated from the Bayesian statistical method. For each parameterisation θ_i of the random sample of the posterior distribution of size n_2, we have perfect knowledge of the status of the population so that it can be determined if Eq. (12) is true or false. Hence, the probability $p(ob)$ of meeting the objective is

$$\begin{aligned} p(ob) &= \sum_{i=1}^{n_2} g(\theta_i, c)/n_2 \\ g(\theta_i, c) &= \begin{cases} 1 & \text{if } ob \text{ is truet} \\ 0 & \text{if } ob \text{ is false} \end{cases} \end{aligned} \tag{13}$$

with the sum given over the complete random sample of the posterior distribution.

While the sustainability of the hunt has to be identified at the population level, recommendations on the sustainability of potential future hunts should preferably be addressed in relation to hunting grounds. To achieve this, for a given a set of potential future catches for each hunt, I used the allocation model to calculate the distributions of future catches for the different populations, with these distributions reflecting the uncertainty in the allocation of catches between the populations. Then, by having these distributions, I could for the catches of each percentile of these distributions calculate the probability that the assumed management objective would be fulfilled for the different populations. This allows the set of potential catches for the different hunts to be adjusted until the probabilities of fulfilling management objectives would be above agreed threshold levels for all populations.

RESULTS

The convergence of the catch and abundance trajectories over the different iterations of the allocation and population dynamic models is shown in Figure 2.

The sampling statistics of the last run of the Bayesian population models are shown in Table 5. The estimated trajectories of the eight summer aggregations are shown in Figure 3, and the posterior parameter estimates in Table 6, with plots of the posterior and realised prior distributions given in Figures 5 to 12. The final estimates of the catch histories per summer aggregation are shown in Figure 4.

Table 7 list the estimated total allowable takes for the different summer aggregations that will meet the management objective of Eq. (12) with probabilities from 0.5 to 0.95.

But management should define the total allowable takes for the different hunts (region and season), as these cannot generally be allocated directly to the different summer aggregation. Hence, Table 8 define possible total allowable takes for the different hunts,

bioRxiv preprint doi: https://doi.org/10.1101/059691; this version posted April 10, 2017. The copyright holder for this preprint (which was not certified by peer review) is the author/funder, who has granted bioRxiv a license to display the preprint in perpetuity. It is made available under aCC-BY-NC-ND 4.0 International license.

M	n_S	n_R	Unique	Max
smith	500	5	4964	2
jones	500	5	4926	2
ingle	2000	5	3672	11
melvi	500	5	4794	3
somer	500	5	4747	4
admir	500	5	4915	3
eclip	500	5	4907	2
baffi	500	5	4899	3

Table 5: **Sampling statistics** for the different models (M). The number of parameter sets in the sample (n_S) and the resample (n_R), the number of unique parameter sets in the resample, and the maximum number of occurrences of a unique parameter set in the resample. n_S and n_R are given in thousands.

with Table 9 giving the associated estimates of the probabilities that these takes from 2015 to 2020 will allow the management objective to be fulfilled for the different summer aggregations. These latter probability estimates have 90% confidence limits that reflect the uncertainty of the summer aggregation origin of the animals taken in the different hunts. The C0 option in Table 8 is the average take over the five-year period from 2009 to 2013.

REFERENCES

Berger, J. O. 1985. *Statistical decision theory and Bayesian analysis.* Second ed. Springer-Verlag, New York.

Buckland, S. T. 1992. Proposal for standard presentation of abundance estimates. *Report of the International Whaling Commission* 42:235.

De la Mare, W. K. 1986. Fitting population models to time series of abundance data. *Report of the International Whaling Commission* 36:399–418.

Heide-Jørgensen, M. P., R. G. Hansen, S. Fossette, N. H. Nielsen, D. L. Borchers, H. Stern and L. Witting 2016. Rebuilding beluga stocks in West Greenland. *Animal Conservation* :DOI: 10.1111/acv.12315.

Jeffreys, H. 1961. *Theory of probability.* 3rd edition ed. Clarendon Press, Oxford.

JWG 2015. Report of the Joint Meeting of the Narwhal Catch Allocation Sub-Group. *NAMMCO* Annual Report 2015:281–304.

Oh, M. S., and J. O. Berger 1992. Adaptive importance sampling in Monte Carlo integration.. *Journal of Statistics and Computer Simulation* 41:143–168.

Press, S. J. 1989. *Bayesian statistics: principles, models, and applications.* John Wiley, New York.

Rubin, D. B. 1988. Using the SIR algorithm to simulate posterior distributions. pp. 395–402, *In*: J. M. Bernardo, M. H. DeGroot, D. V. Lindley, and A. M. Smith (eds.) *Bayesian Statistics 3: Proceedings of the Third Valencia International Meeting, 1–5 June 1987.* Clarendon Press, Oxford.

Witting, L. 2009. Model uncertainty on assessments of West Greenland beluga and narwhals. *NAMMCO/SC/16-JCNB/SWG/2009-JWG/12* .

Witting, L. 2013. Selection-delayed population dynamics in baleen whales and beyond. *Population Ecology* 55:377–401.

bioRxiv preprint doi: https://doi.org/10.1101/059691; this version posted April 10, 2017. The copyright holder for this preprint (which was not certified by peer review) is the author/funder, who has granted bioRxiv a license to display the preprint in perpetuity. It is made available under aCC-BY-NC-ND 4.0 International license.

Witting, L., and E. W. Born 2005. An assessment of Greenland walrus populations. *ICES Journal of Marine Science* 62:266–284.

Witting, L., and E. W. Born 2013. Population dynamics of walrus in Greenland. *NAMMCO Scientific Publications* 9:http://dx.doi.org/10.7557/3.2612.

Witting, L., and M. P. Heide-Jørgensen 2012a. Assessment of East Greenland narwhals. *NAMMCO/SC/19-JCNB/SWG/2012-JWG/12* .

Witting, L., and M. P. Heide-Jørgensen 2012b. Assessment of West Greenland narwhal (excluding Uummannaq). *NAMMCO/SC/19-JCNB/SWG/2012-JWG/10* .

Witting, L., and M. P. Heide-Jørgensen 2012c. Estimation of dynamics for narwhals in Uummannaq, West Greenland. *NAMMCO/SC/19-JCNB/SWG/2012-JWG/11* .

bioRxiv preprint doi: https://doi.org/10.1101/059691; this version posted April 10, 2017. The copyright holder for this preprint (which was not certified by peer review) is the author/funder, who has granted bioRxiv a license to display the preprint in perpetuity. It is made available under aCC-BY-NC-ND 4.0 International license.

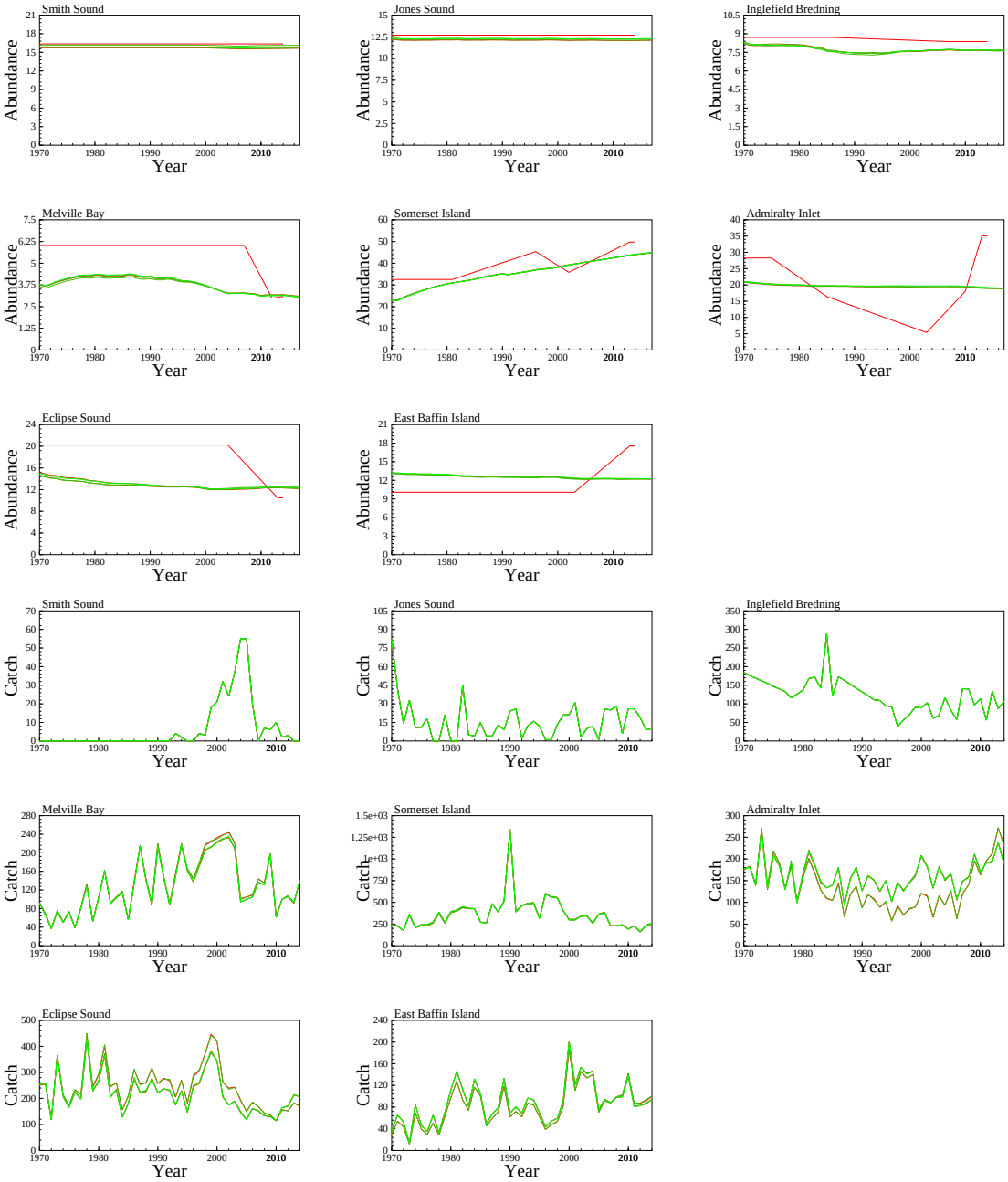

Figure 2: The convergence of the abundance trajectories and catch histories as a function of the number of iterations of the complete meta aggregation model, with iteration number increasing with colour transitions from clear red to clear green. Abundance is given in thousands.

bioRxiv preprint doi: https://doi.org/10.1101/059691; this version posted April 10, 2017. The copyright holder for this preprint (which was not certified by peer review) is the author/funder, who has granted bioRxiv a license to display the preprint in perpetuity. It is made available under aCC-BY-NC-ND 4.0 International license.

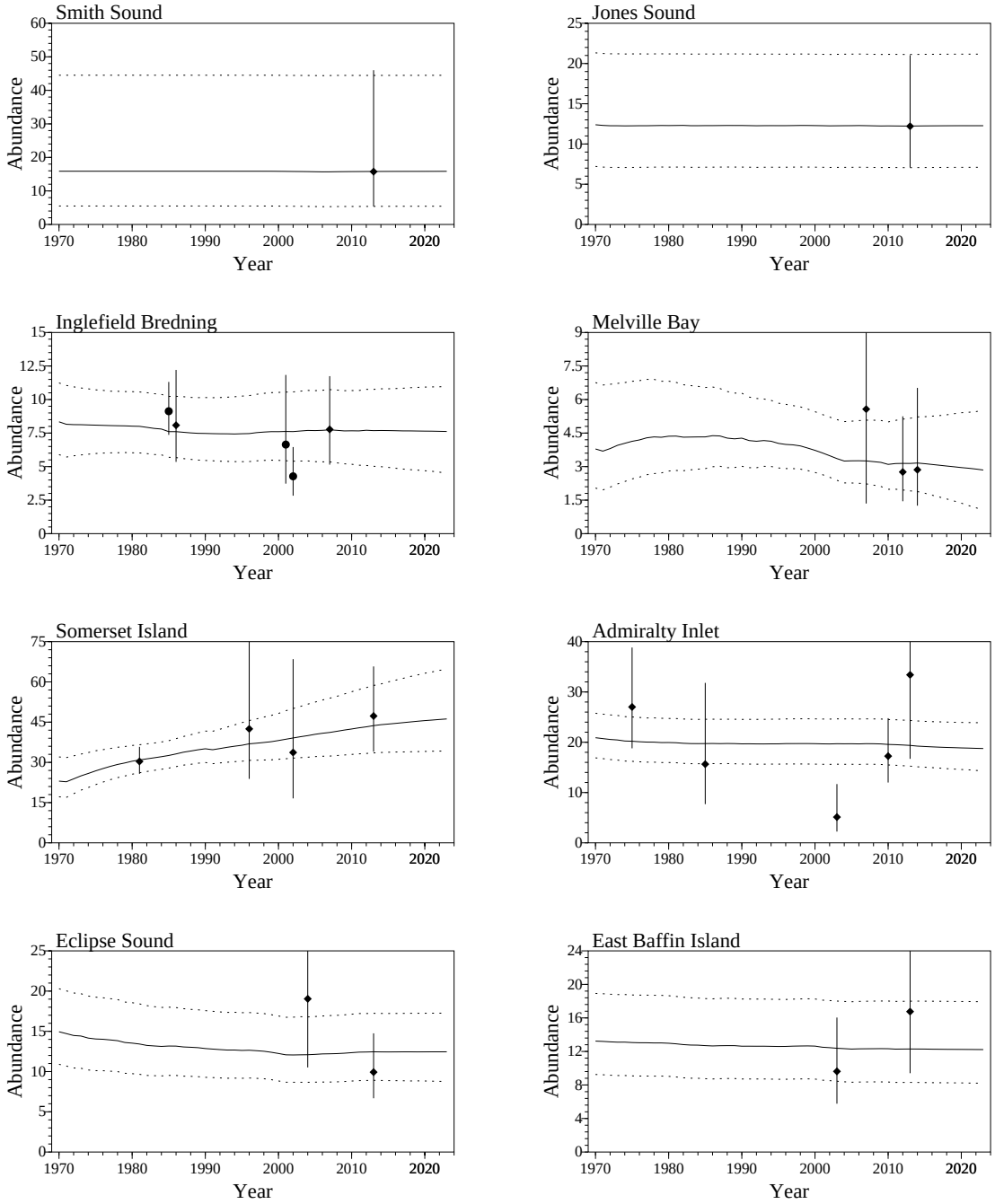

Figure 3: The trajectories of the different narwhal aggregations. Points with bars are the abundance estimates with 90% CI, solid curves the median, and dotted curves the 90% CI, of the estimated models. Abundance is given in thousands.

bioRxiv preprint doi: https://doi.org/10.1101/059691; this version posted April 10, 2017. The copyright holder for this preprint (which was not certified by peer review) is the author/funder, who has granted bioRxiv a license to display the preprint in perpetuity. It is made available under aCC-BY-NC-ND 4.0 International license.

M		N_0	N^*	r	msyr	p	p_0	a_m	γ	msyl	c_h
	$x_{.5}$	-	16	.035	.028	.97	.75	11	3	.67	.05
	$x_{.05}$	-	5.5	.013	.01	.96	.53	8.1	2.1	.62	.0062
smith	$x_{.95}$	-	45	.058	.045	.99	.97	14	3.9	.7	.29
	$x_{.5}$	-	12	.035	.027	.97	.75	11	3	.67	.29
	$x_{.05}$	-	7.2	.013	.0097	.96	.52	8.1	2.1	.62	.17
jones	$x_{.95}$	-	21	.058	.046	.99	.97	14	3.9	.7	.6
	$x_{.5}$	8.3	11	.023	.018	.97	.66	12	2.9	.65	.12
	$x_{.05}$	5.9	7.9	.0055	.0042	.95	.51	8.3	2.1	.6	.033
ingle	$x_{.95}$	11	25	.056	.044	.99	.95	14	3.9	.7	.48
	$x_{.5}$	3.8	7.6	.037	.029	.97	.77	11	3	.67	.16
	$x_{.05}$	2	4.4	.015	.012	.96	.53	8	2.1	.62	.026
melvi	$x_{.95}$	6.8	25	.06	.047	.99	.98	14	3.9	.71	.49
	$x_{.5}$	23	51	.038	.029	.97	.77	11	3	.67	.36
	$x_{.05}$	17	36	.022	.017	.96	.54	8	2.1	.62	.075
somer	$x_{.95}$	32	83	.058	.045	.99	.98	14	3.9	.71	.82
	$x_{.5}$	-	21	.035	.027	.97	.74	11	3	.66	.37
	$x_{.05}$	-	17	.012	.0092	.96	.52	8	2.1	.61	.086
admir	$x_{.95}$	-	26	.057	.045	.99	.97	14	3.9	.7	.7
	$x_{.5}$	-	15	.036	.028	.97	.76	11	3	.67	.38
	$x_{.05}$	-	11	.013	.01	.96	.53	8	2.1	.62	.077
eclip	$x_{.95}$	-	20	.06	.046	.99	.98	14	3.9	.71	.75
	$x_{.5}$	-	13	.036	.028	.97	.76	11	3	.67	.52
	$x_{.05}$	-	9.2	.015	.011	.96	.53	8	2.1	.62	.2
baffi	$x_{.95}$	-	19	.058	.045	.99	.98	14	3.9	.7	.77

M	N_t	d_t	$\dot{r}_t$	β_a
	16	1	.00029	-
	5.4	.98	6e-5	-
smith	44	1	.0011	-
	12	.99	.0011	-
	7.1	.97	.00058	-
jones	21	1	.0019	-
	7.7	.7	.012	.32
	4.8	.25	.0047	.23
ingle	11	.95	.02	.46
	3.1	.39	.032	-
	1.6	.11	.014	-
melvi	5.3	.78	.049	-
	45	.92	.0096	-
	34	.59	.0052	-
somer	61	.98	.022	-
	19	.92	.0089	-
	15	.8	.0051	-
admir	24	.96	.013	-
	12	.87	.013	-
	8.9	.63	.0079	-
eclip	17	.96	.02	-
	12	.94	.0076	-
	8.3	.83	.0046	-
baffi	18	.98	.012	-

Table 6: **Parameter estimates** for the different models (M). Estimates are given by the median ($x_{.5}$) and the 90% credibility interval ($x_{.05}$ - $x_{.95}$) of the posterior distributions. Abundance is given in thousands.

bioRxiv preprint doi: https://doi.org/10.1101/059691; this version posted April 10, 2017. The copyright holder for this preprint (which was not certified by peer review) is the author/funder, who has granted bioRxiv a license to display the preprint in perpetuity. It is made available under aCC-BY-NC-ND 4.0 International license.

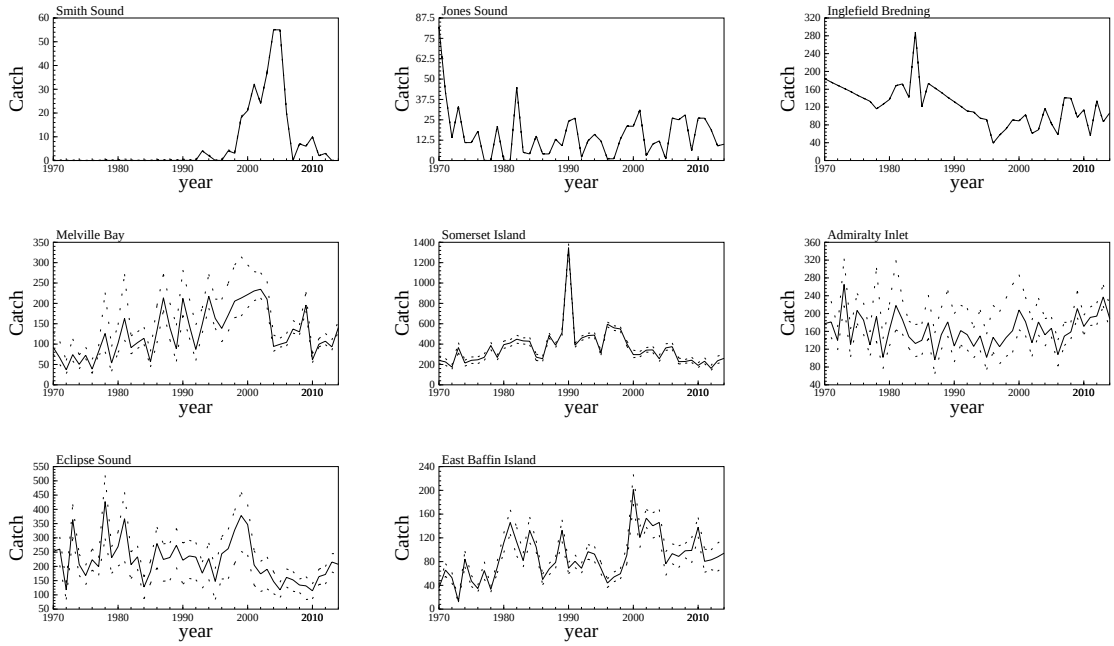

Figure 4: Yearly catches per summer aggregation with 90% confidence intervals.

P	smith	jones	ingle	melvi	somer	admir	eclip	baffi
F	0.50	0.50	0.50	0.50	0.50	0.50	0.50	0.50
0.50	250	196	117	110	930	328	241	218
0.55	225	182	109	103	896	308	230	206
0.60	200	169	101	96	863	288	217	194
0.65	180	156	91	88	828	268	206	181
0.70	159	144	81	79	791	247	194	170
0.75	141	130	71	70	754	226	181	158
0.80	121	116	60	60	714	206	168	144
0.85	101	101	47	49	671	181	150	127
0.90	80	84	29	37	614	152	129	109
0.95	57	61	13	20	538	111	101	83

Table 7: **Catch objective trade-off per stock.** The total annual removals per stock that meet given probabilities (P) of meeting management objectives. The simulated period is from 2017 to 2022, and F is the assumed fraction of females in the catch.

bioRxiv preprint doi: https://doi.org/10.1101/059691; this version posted April 10, 2017. The copyright holder for this preprint (which was not certified by peer review) is the author/funder, who has granted bioRxiv a license to display the preprint in perpetuity. It is made available under aCC-BY-NC-ND 4.0 International license.

Hunt	Season	C0	C1
Etah	Spring	4	5
Qaanaaq	Summer	98	98
Grise Fjord	Spring	7	9
Grise Fjord	Summer	11	15
Grise Fjord	Fall	0	0
Upernavik	Summer	100	70
Ummannaaq	Fall	86	154
Disko Bay	Winter	73	97
CCA	Spring	4	6
CCA	Summer	74	118
CCA	Fall	2	3
Arctic Bay	Spring	31	41
Arctic Bay	Summer	141	188
Arctic Bay	Fall	0	0
Pond Inlet	Spring	58	77
Pond Inlet	Summer	55	73
Pond Inlet	Fall	4	5
BIC	Spring	12	11
BIC	Summer	100	91
BIC	Fall	44	40
BIS	Spring	5	5
BIS	Summer	9	8
BIS	Fall	12	11
BIS	Winter	0	0

Table 8: Catch option examples (C#) of maximum yearly removal per hunting region.

	Smith	Jones	Inglefield	Melville	Somerset	Admiralty	Eclipse	Baffin
C0	$4\,{}^{4}_{4}$	$18\,{}^{18}_{18}$	$98\,{}^{98}_{98}$	$110\,{}^{133}_{102}$	$223\,{}^{246}_{200}$	$213\,{}^{260}_{180}$	$113\,{}^{141}_{84}$	$146\,{}^{158}_{133}$
P0	$1.00\,{}^{1.00}_{1.00}$	$1.00\,{}^{1.00}_{1.00}$	$0.62\,{}^{0.62}_{0.62}$	$0.4\,{}^{0.45}_{0.24}$	$1.00\,{}^{1.00}_{1.00}$	$0.78\,{}^{0.85}_{0.67}$	$0.93\,{}^{0.96}_{0.88}$	$0.79\,{}^{0.83}_{0.75}$
C1	$5\,{}^{5}_{5}$	$24\,{}^{24}_{24}$	$98\,{}^{98}_{98}$	$84\,{}^{114}_{73}$	$349\,{}^{376}_{320}$	$280\,{}^{338}_{236}$	$147\,{}^{183}_{109}$	$133\,{}^{144}_{121}$
P1	$1.00\,{}^{1.00}_{1.00}$	$0.99\,{}^{0.99}_{0.99}$	$0.62\,{}^{0.62}_{0.62}$	$0.57\,{}^{0.64}_{0.37}$	$1.00\,{}^{1.00}_{0.99}$	$0.62\,{}^{0.72}_{0.47}$	$0.86\,{}^{0.94}_{0.74}$	$0.84\,{}^{0.87}_{0.8}$

Table 9: Examples of future annual removals (C#) per summer aggregation, with associated probabilities (P#) of fulfilling management objectives. The different removals follow from the catch options in Table 8, and the 90% confidence intervals of the estimates are given by the sub and super scripts.

bioRxiv preprint doi: https://doi.org/10.1101/059691; this version posted April 10, 2017. The copyright holder for this preprint (which was not certified by peer review) is the author/funder, who has granted bioRxiv a license to display the preprint in perpetuity. It is made available under aCC-BY-NC-ND 4.0 International license.

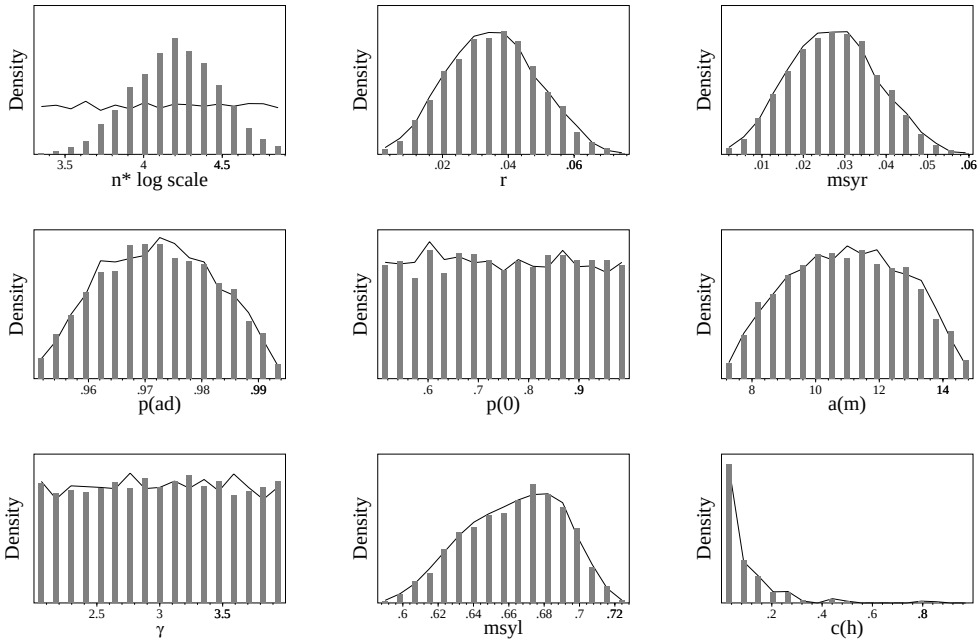

Figure 5: **Smith Sound** Realised prior (curve) and posterior (bars) distributions.

bioRxiv preprint doi: https://doi.org/10.1101/059691; this version posted April 10, 2017. The copyright holder for this preprint (which was not certified by peer review) is the author/funder, who has granted bioRxiv a license to display the preprint in perpetuity. It is made available under aCC-BY-NC-ND 4.0 International license.

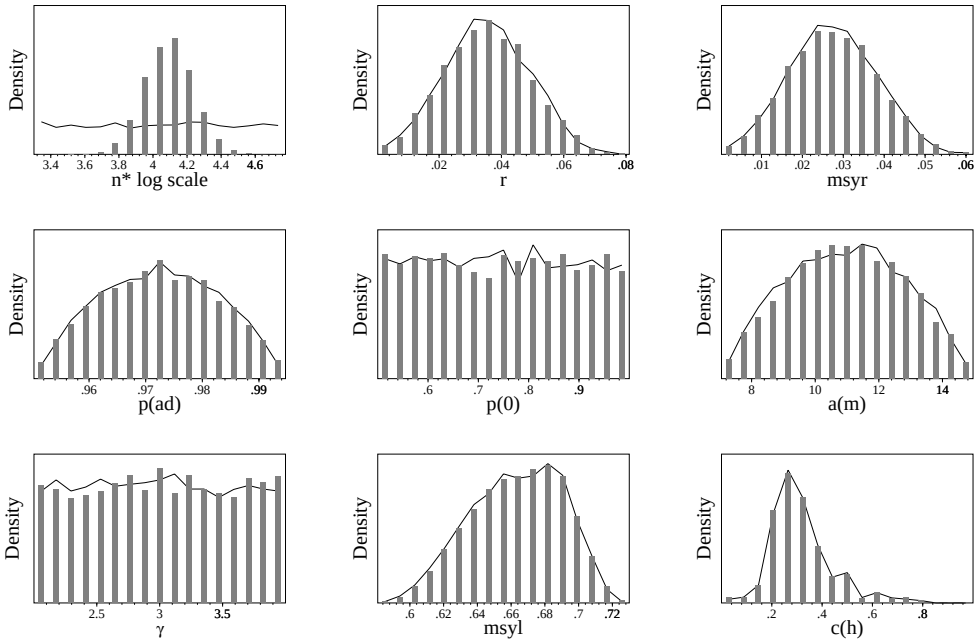

Figure 6: **Jones Sound** Realised prior (curve) and posterior (bars) distributions.

bioRxiv preprint doi: https://doi.org/10.1101/059691; this version posted April 10, 2017. The copyright holder for this preprint (which was not certified by peer review) is the author/funder, who has granted bioRxiv a license to display the preprint in perpetuity. It is made available under aCC-BY-NC-ND 4.0 International license.

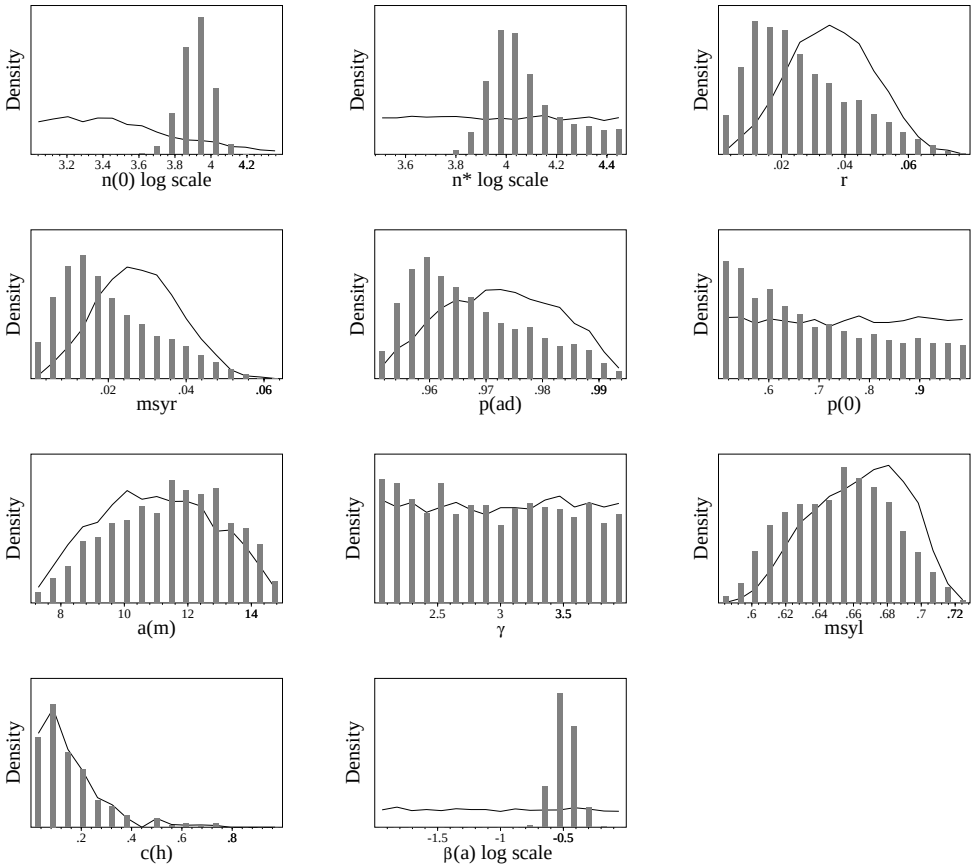

Figure 7: **Inglefield Bredning** Realised prior (curve) and posterior (bars) distributions.

bioRxiv preprint doi: https://doi.org/10.1101/059691; this version posted April 10, 2017. The copyright holder for this preprint (which was not certified by peer review) is the author/funder, who has granted bioRxiv a license to display the preprint in perpetuity. It is made available under aCC-BY-NC-ND 4.0 International license.

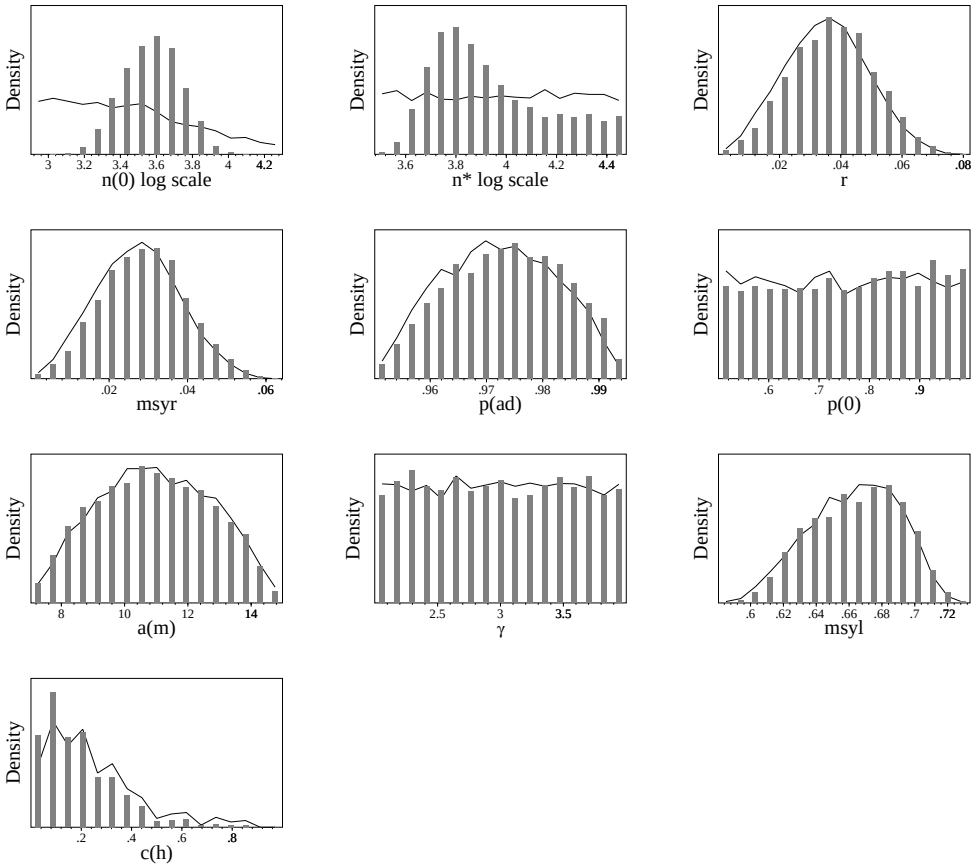

Figure 8: **Melville Bay** Realised prior (curve) and posterior (bars) distributions.

bioRxiv preprint doi: https://doi.org/10.1101/059691; this version posted April 10, 2017. The copyright holder for this preprint (which was not certified by peer review) is the author/funder, who has granted bioRxiv a license to display the preprint in perpetuity. It is made available under aCC-BY-NC-ND 4.0 International license.

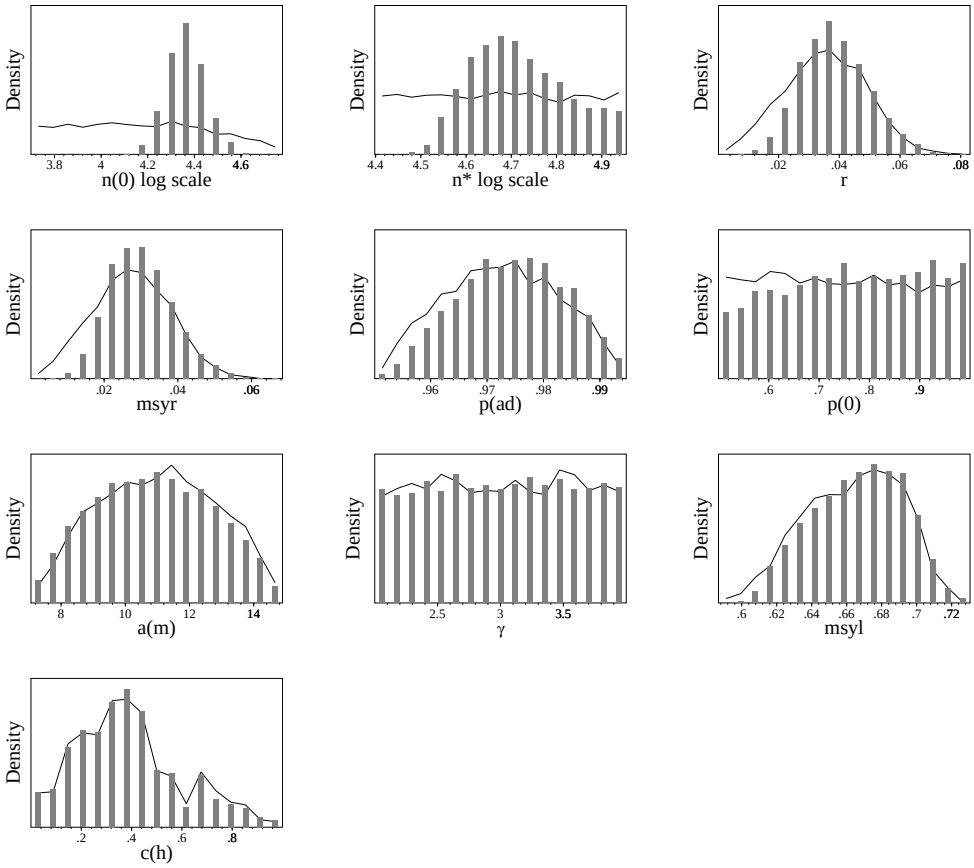

Figure 9: **Somerset Island** Realised prior (curve) and posterior (bars) distributions.

bioRxiv preprint doi: https://doi.org/10.1101/059691; this version posted April 10, 2017. The copyright holder for this preprint (which was not certified by peer review) is the author/funder, who has granted bioRxiv a license to display the preprint in perpetuity. It is made available under aCC-BY-NC-ND 4.0 International license.

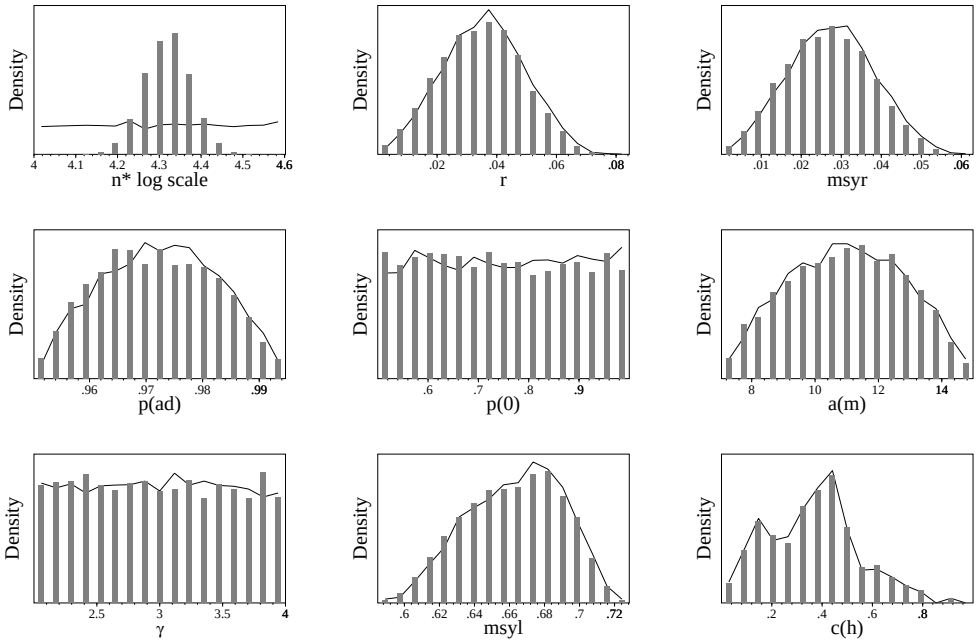

Figure 10: **Admiralty Inlet** Realised prior (curve) and posterior (bars) distributions.

bioRxiv preprint doi: https://doi.org/10.1101/059691; this version posted April 10, 2017. The copyright holder for this preprint (which was not certified by peer review) is the author/funder, who has granted bioRxiv a license to display the preprint in perpetuity. It is made available under aCC-BY-NC-ND 4.0 International license.

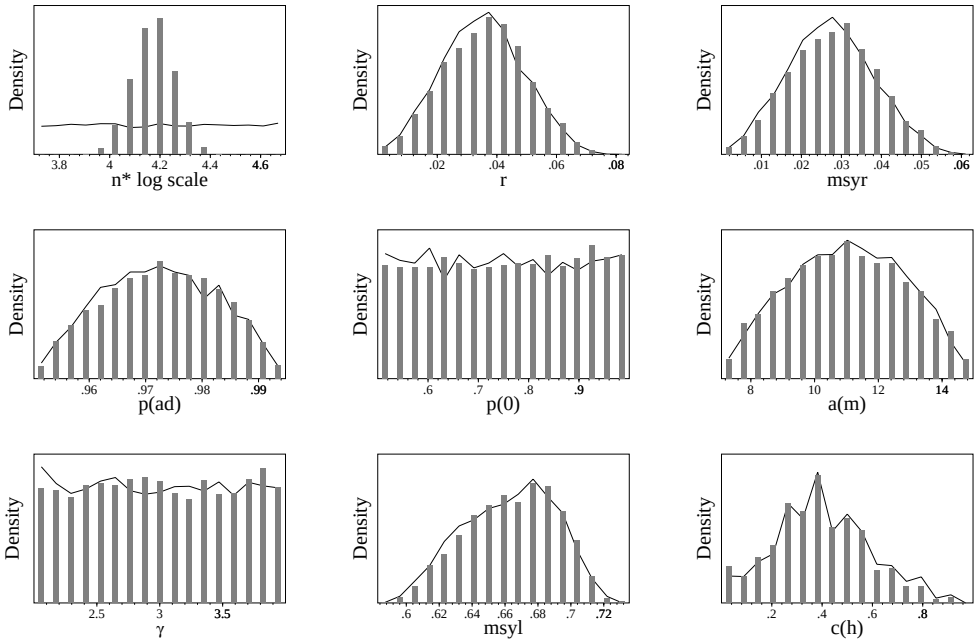

Figure 11: **Eclipse Sound** Realised prior (curve) and posterior (bars) distributions.

bioRxiv preprint doi: https://doi.org/10.1101/059691; this version posted April 10, 2017. The copyright holder for this preprint (which was not certified by peer review) is the author/funder, who has granted bioRxiv a license to display the preprint in perpetuity. It is made available under aCC-BY-NC-ND 4.0 International license.

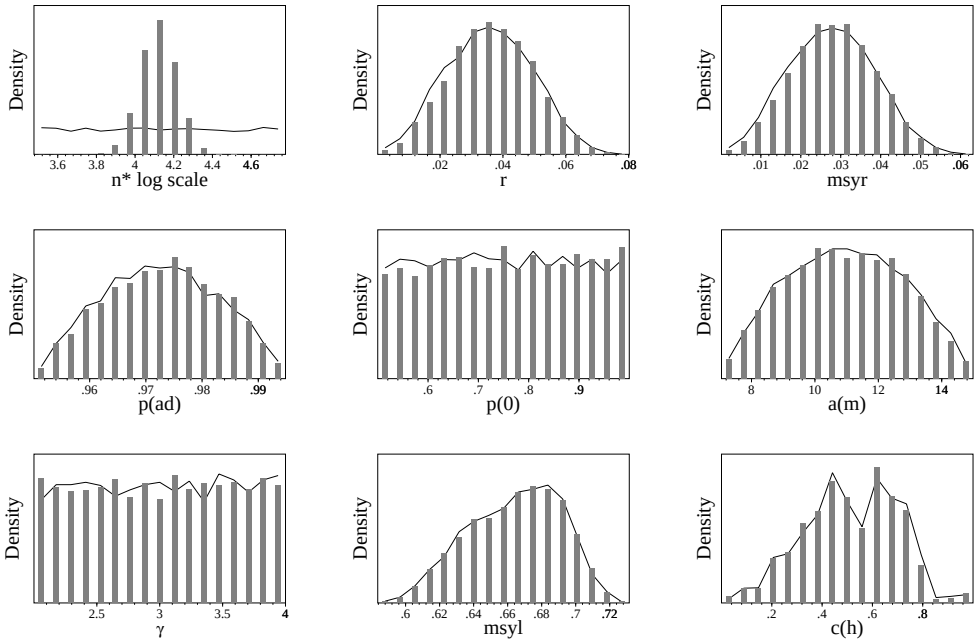

Figure 12: **East Baffin Island** Realised prior (curve) and posterior (bars) distributions.

bioRxiv preprint doi: https://doi.org/10.1101/059691; this version posted April 10, 2017. The copyright holder for this preprint (which was not certified by peer review) is the author/funder, who has granted bioRxiv a license to display the preprint in perpetuity. It is made available under aCC-BY-NC-ND 4.0 International license.

Year	E_s	Q_s	Up_s	Um_f	D_w
1970	0	184	70	86	129
1971	0	176	45	60	134
1972	0	169	24	35	78
1973	0	162	53	83	120
1974	0	155	35	61	83
1975	0	147	62	14	66
1976	0	140	25	35	74
1977	0	133	71	147	40
1978	0	116	64	238	342
1979	0	126	25	172	134
1980	0	137	70	190	163
1981	0	168	95	182	348
1982	0	172	68	211	99
1983	0	142	83	213	88
1984	0	288	92	273	87
1985	0	121	39	51	88
1986	0	173	93	126	203
1987	0	163	167	434	203
1988	0	153	98	294	203
1989	0	142	43	374	203
1990	0	132	146	1325	203
1991	0	122	104	290	203
1992	0	111	43	374	203
1993	4	109	117	391	134
1994	2	95	173	386	203
1995	0	92	130	207	163
1996	0	39	89	527	224
1997	4	57	113	495	272
1998	3	71	147	447	295
1999	18	91	150	329	335
2000	21	89	177	138	255
2001	32	103	198	124	182
2002	24	61	204	234	163
2003	37	69	182	226	157
2004	55	117	78	87	99
2005	55	83	89	209	51
2006	20	58	92	94	73
2007	0	141	123	87	86
2008	7	140	120	113	61
2009	6	97	177	118	116
2010	10	114	52	55	59
2011	2	56	91	100	52
2012	3	134	96	55	72
2013	0	87	82	101	66
2014	0	107	130	90	65

Table 10: Estimated total removal per hunting region in Greenland per year. E_s:Etah (Spring). Q_s:Qaanaaq (Summer). Up_s:Upernavik (Summer). Um_f:Ummannaaq (Fall). D_w:Disko Bay (Winter).

bioRxiv preprint doi: https://doi.org/10.1101/059691; this version posted April 10, 2017. The copyright holder for this preprint (which was not certified by peer review) is the author/funder, who has granted bioRxiv a license to display the preprint in perpetuity. It is made available under aCC-BY-NC-ND 4.0 International license.

Year	G_s	G_s	G_f	C_s	C_s	C_f	A_s	A_s	A_f	P_s	P_s	P_f	B_s	B_s	B_f	S_s	S_s	S_f	S_w
1970	11	37	34	4	30	18	112	22	21	145	75	30	7	12	30	8	2	1	0
1971	6	19	17	4	32	21	112	23	21	141	74	30	10	20	43	26	4	3	0
1972	2	7	5	4	35	20	121	26	23	44	22	8	6	14	28	27	5	4	0
1973	3	18	12	4	62	30	183	37	33	264	136	60	3	6	11	0	1	0	0
1974	2	6	3	3	32	18	66	12	10	139	63	29	13	26	56	34	5	4	0
1975	2	6	3	6	36	19	220	35	30	107	46	25	4	12	17	34	4	4	0
1976	4	9	5	7	33	17	147	27	22	171	78	40	5	12	18	16	2	2	0
1977	0	0	0	0	13	36	23	22	19	144	69	34	0	14	111	6	0	0	0
1978	0	0	0	5	14	5	83	16	12	198	100	48	6	15	29	3	0	1	0
1979	0	2	19	0	4	15	42	6	8	118	97	0	25	6	23	23	27	3	0
1980	0	0	0	1	44	1	160	0	18	121	66	34	14	50	75	40	0	0	0
1981	0	0	0	0	20	86	113	34	20	101	58	29	15	52	77	53	8	23	0
1982	0	45	0	4	73	0	99	31	19	188	52	0	0	9	103	56	11	27	0
1983	0	3	2	0	64	29	141	18	14	81	71	78	38	42	36	4	2	0	0
1984	2	2	0	0	0	0	164	5	0	94	8	8	15	60	66	68	0	0	0
1985	7	6	2	11	18	8	183	0	0	141	27	62	8	78	5	36	2	0	0
1986	2	2	0	4	10	5	87	45	29	113	83	31	2	7	11	22	8	26	0
1987	2	2	0	4	11	6	22	11	7	57	44	17	9	35	65	0	0	0	0
1988	6	6	1	4	12	8	75	39	24	55	49	16	17	44	56	1	0	2	0
1989	4	4	1	7	21	11	86	46	27	74	74	25	15	47	74	28	9	34	0
1990	9	12	3	4	17	8	39	28	16	32	39	14	8	35	61	3	1	2	0
1991	10	13	3	4	18	15	65	49	26	46	55	21	10	40	66	3	2	5	0
1992	1	1	0	3	23	7	56	41	29	48	54	19	8	38	53	2	1	3	0
1993	5	6	1	5	30	10	46	36	22	38	43	15	9	43	65	12	5	15	0
1994	6	7	3	4	28	12	54	42	25	44	47	21	8	36	58	16	6	21	0
1995	5	5	2	3	26	12	33	9	16	90	0	0	8	34	61	3	1	4	0
1996	0	1	0	1	13	7	59	43	20	40	57	26	3	14	28	8	3	14	0
1997	0	1	0	2	21	12	40	29	13	28	43	21	5	26	57	1	0	1	0
1998	4	7	2	4	47	17	54	41	18	21	106	2	5	29	57	2	1	2	0
1999	8	11	2	3	12	8	49	45	16	17	106	39	1	29	79	14	4	25	0
2000	8	11	2	2	38	12	66	64	23	69	79	58	18	79	153	9	44	0	0
2001	12	16	3	6	54	37	67	71	24	27	32	21	13	53	108	11	1	14	0
2002	3	0	0	0	37	21	22	11	63	48	29	0	0	98	73	9	0	30	0
2003	0	10	0	4	32	2	60	84	15	32	40	10	12	73	105	36	0	1	0
2004	0	3	9	0	13	59	81	50	21	27	14	39	34	71	94	12	0	21	0
2005	1	0	0	0	43	26	83	93	1	26	25	20	14	10	133	0	0	6	0
2006	0	26	0	1	73	74	170	3	3	25	56	20	5	45	111	0	0	1	0
2007	4	21	0	0	44	12	90	72	5	8	35	32	10	31	120	4	1	0	0
2008	0	23	5	0	45	8	65	78	35	16	58	9	2	52	65	0	4	22	0
2009	5	1	0	3	22	46	23	150	1	24	6	21	9	25	93	10	0	40	0
2010	10	16	0	2	48	14	51	89	32	20	15	37	17	77	76	14	1	20	0
2011	14	10	2	8	51	4	38	112	26	45	81	3	7	23	125	2	0	4	0
2012	2	17	0	0	47	23	4	65	100	25	63	23	9	31	98	0	1	10	0
2013	5	0	4	1	33	23	43	167	4	30	82	58	11	9	143	3	1	18	0
2014	1	0	9	0	45	32	63	46	81	33	63	59	16	22	140	1	0	11	0

Table 11: Estimated total removal per hunting region in Canada per year. G_s:Grise Fjord (Spring). G_s:Grise Fjord (Summer). G_f:Grise Fjord (Fall). C_s:CCA (Spring). C_s:CCA (Summer). C_f:CCA (Fall). A_s:Arctic Bay (Spring). A_s:Arctic Bay (Summer). A_f:Arctic Bay (Fall). P_s:Pond Inlet (Spring). P_s:Pond Inlet (Summer). P_f:Pond Inlet (Fall). B_s:BIC (Spring). B_s:BIC (Summer). B_f:BIC (Fall). S_s:BIS (Spring). S_s:BIS (Summer). S_f:BIS (Fall). S_w:BIS (Winter).