TITLE: Conspecific sperm precedence is reinforced but sexual selection weakened in

sympatric populations of *Drosophila*

AUTHORS: Dean M. Castillo[1,2] and Leonie C. Moyle[1]

AUTHOR AFFILIATIONS: [1]Department of Biology, 1001 East Third Street, Indiana

University, Bloomington, Indiana 47405

CORRESPONDING AUTHOR [2] Department of Molecular Biology and Genetics, 526

Campus Rd, Cornell University, Ithaca, New York 14853 dmc79@cornell.edu (607) 254-

5208

KEYWORDS: speciation, sexual selection, sperm competition

13
14
15 **Word Count:** 4,161 (not including abstract or methods)   **Figure Count:** 6

SUMMARY: Sexual selection is well recognized as a driver of reproductive isolation

between lineages. However, selection for increased reproductive isolation could

reciprocally change the outcomes of sexual selection, when these processes share a

genetic basis. Direct selection for reproductive isolation occurs in the context of

'reinforcement', where selection acts to increase prezygotic barriers to reduce the cost of

heterospecific matings. Many studies of reinforcement focus on premating reproductive

barriers, however postmating traits-such as conspecific sperm precedence (CSP)-can also

respond to reinforcing selection. We tested whether i) CSP responded to reinforcing

selection, and ii) this response in sympatric populations altered intraspecific sperm

competition (ISC) and the strength of sexual selection, with the sister species *Drosophila*

pseudoobscura and *D. persimilis*. We used sperm competition experiments to evaluate

differences in CSP and ISC between two sympatric and two allopatric populations of *D.*

pseudoobscura. Using multiple genotypes for each population allowed us to estimate not

only patterns of phenotype divergence, but also the opportunity for sexual selection

within each population. Consistent with a pattern of reinforcement, the sympatric

populations had higher mean CSP. Moreover, ISC was altered in sympatric populations,

where we observed decreased average offensive sperm competitive ability against

conspecific males, allowing less opportunity for sexual selection to operate within these

populations. These data demonstrate that strong reinforcing selection for reproductive

isolation can have consequences for sexual selection and sexual interactions within

species, in these important postmating sperm competition traits.

Introduction

47   When closely-related species comes into contact, the presence of heterospecifics

can influence sexual interactions and therefore alter patterns of selection on reproductive

traits. In cases where these species have the potential to interbreed, selection can favor

divergence in sexual traits to avoid costs of heterospecific mating, a type of reproductive

character displacement commonly called reinforcement [1-3]. The frequency at which

reinforcement contributes to speciation is still under debate [3-4] although several recent

examples provide strong evidence for reinforcement acting on mating traits [5-10].

Regardless, the mating trait changes that evolve in response to reinforcement can have

collateral effects on intraspecific sexual dynamics [6]. This can in turn alter the

magnitude and efficacy of sexual selection specifically within populations exposed to

heterospecifics. These potential reciprocal interactions between sexual selection and

reproductive isolation remain relatively untested [6-7], but can have important

consequences for how we interpret evolution of sexual traits and interactions. For

example, patterns of reproductive trait evolution in rapid radiations, where sexual

selection is thought to be the primary driver, may be misinterpreted if they do not take

into account species interactions.

63   For reinforcement and sexual selection to reciprocally affect the evolution of

sexual traits, these traits must be involved in both processes and share a genetic basis.

Currently the best example of a shared genetic basis for sexual selection and reproductive

isolation comes from *Drosophila* sperm competition genes, several of which have been

shown to mediate both sexual selection through intraspecies sperm competition (ISC) and

reproductive isolation via conspecific sperm precedence (CSP) [11]. Conspecific sperm

precedence occurs when a female mates with both heterospecific and conspecific males

yet most of the progeny are sired by the conspecific male; this precedence can occur

either through competitive mechanisms (including male sperm competition and cryptic

female choice) or non-competitive mechanisms (resulting mainly from gametic

incompatibilities). CSP has proven to be a strong reproductive isolation barrier among

species in *Drosophila* [4,12-13] and in many other plant and animal species [4, and

references therein]. Although ubiquitous, CSP can be overlooked as a reproductive

isolating barrier because it involves inconspicuous phenotypes that are not readily

observed in the field [14]. Moreover, although reinforcement studies have

overwhelmingly focused on pre-mating traits, postcopulatory prezygotic traits including

CSP can also be the target of reinforcement [15-17]. Previous empirical studies have been

equivocal about whether heterospecific interactions and reinforcement select for

increased CSP specifically in sympatry, with no single study simultaneously estimating

and comparing levels of CSP in allopatric and sympatric populations [13, 18-25]

While reinforcing selection (acting on CSP) and sexual selection (acting on ISC)

could interact to influence evolutionary change in post-copulatory traits, the outcomes of

this interaction clearly will depend upon whether these forces act in concert or in

opposition. When sexual selection and reinforcing selection act in concert, trait evolution

can proceed faster than otherwise expected, but the direction of trait evolution remains

unchanged. In contrast, the potential feedback between sexual selection and reproductive

isolation can generate complex evolutionary outcomes when these forces act at cross-

purposes. For example, sperm competition is shaped by sexual conflict between males

and females (i.e. antagonistic pleiotropy [26-28]) and genotype-genotype interactions

(male-male [29-30] and male-female: [31-33]. Both are expected to maintain high

variance in the affected traits and, indeed, sperm competition genes are often highly

variable both in terms of molecular and phenotypic variation [30, 33-34]. In contrast,

under models of speciation by sexual selection—where isolation is generated by strong

disruptive selection between populations and directional selection within a population

[35-36]—genetic variance of traits that act as barriers to reproduction is expected to be

reduced and the overall trait mean shifted. The net effect of selection imposed by

intrapopulation sexual interactions and by reinforcement can together produce phenotypic

and genetic variation in sperm competition traits/genes that is different from the optimal

variation when sexual selection acts alone.

102   One way these potentially antagonistic optima could play out is when

reinforcement-mediated changes in the mean and variance of sperm competition traits

alter the opportunity for sexual selection among conspecifics [7]. Sperm competition

contributes to variance in reproductive success because male genotypes that can

disproportionately sire offspring increase their fitness compared to the fitness of rival

male conspecifics [37-38]. Strong sperm competition leads to greater opportunity for

sexual selection because there is greater variance in reproductive success compared to

scenarios where males have equal probability of siring offspring. This generates two

alternative predictions of the possible effects of reinforcement on sexual selection. First,

the response to strong directional selection from reinforcement on sperm competition

traits could lead to greater siring ability in intrapopulation sperm competition, increasing

variance in reproductive success and opportunity for sexual selection. Alternatively,

strong directional selection could reduce phenotypic variation so that competitive ability

is equalized among males, thus reducing the opportunity for sexual selection.

The strategy we used to evaluate the interaction between selection for increased

reproductive isolation (i.e. reinforcement) and sexual selection acting on sperm

competition genes was to estimate variation between genotypes in CSP and ISC in

parallel. Both CSP and ISC are measures of postcopulatory offensive sperm competition,

estimated by allowing females to mate sequentially with two different male genotypes

and scoring the paternity of the resulting progeny. Here our focus was on second-male or

'offensive' siring success. This is typically referred to as 'P2' and captures the ability of

the second mated male to sire offspring by displacing or disabling the sperm of the first

male. For our experiments the first male was either heterospecific (to estimate CSP) or

conspecific (to estimate ISC) tester male. By comparing the relative competitive success

of replicate male lines against a common set of either heterospecific and conspecific male

tester genotypes, we could estimate post-copulatory CSP and ISC in parallel in the same

experiment. Using this design we also estimated which genotype effects (male genotype,

female genotype, or the interaction) might shape CSP and ISC. Females experience the

most cost of heterospecific matings [39-41] and could control CSP via cryptic female

choice [42], thus we would expect strong female genotype effects on CSP. This contrasts

with previous studies of ISC where both male and female genetic effects, and their

interaction were significant effects [31,33]. Unlike ISC the phenotypic and genetic

variance for CSP has not been empirically explored and their similarity to ISC is

currently unknown.

In this study, we examine evidence for reinforcement of CSP among populations

of *Drosophila pseudoobscura* that are allopatric or sympatric with their closely related

sister species *D. persimilis*, and evaluate the potential consequences of these

heterospecific interactions for ISC and sexual selection within *D. pseudoobscura*

populations. One of the first clear empirical demonstrations of reinforcement on pre-

mating isolation was described in this species pair [43]. This finding suggests that

heterospecific interactions and matings are frequent and sustained over evolutionary time

and can act as a substantial selective agent on reproductive traits in this system. Here we

determine whether there is evidence that heterospecific interactions have selected for

increased CSP, by comparing this barrier among populations of *D. pseudoobscura* that

are allopatric or sympatric with *D. persimilis*. A pattern of stronger CSP specifically in

sympatry is consistent with reinforcement; moreover, because postcopulatory traits are

less likely to be directly affected by environmental conditions, this pattern is unlikely to

be explained by alternative phenomena, such as ecological selection, that could also

explain character displacement in sympatry (see Discussion). Using a consistent design

across all populations we could also estimate premating reproductive isolation in the

same experiment, and compare its strength in sympatry and allopatry. Second, we

evaluate whether selection for strong CSP in sympatry has affected ISC, and thereby

post-copulatory sexual selection, as might occur when CSP and ISC have shared genetic

architecture. Throughout, we test for differences in trait variation across a set of distinct

genotypes which allows us to specifically evaluate which sex is playing a more critical

role in determining variation in heterospecific and conspecific postcopulatory

interactions.

RESULTS

No difference between allopatric and sympatric populations in premating isolation

Because our experimental assessment of CSP involved first mating with a

heterospecific *D. persimilis* male, we were able to estimate the magnitude of premating

isolation in each *D. pseudoobscura* population in our experiment. We did not find

evidence for a pattern consistent with reinforcement of premating isolation mediated by

female mate preference. The average probability of heterospecific matings ranged from

46-52% between populations, and did not differ between allopatric and sympatric

populations (χ^2 test of independence: χ^2=1.185, df=1, *P*=0.2763; Wald's Test: χ^2=1.9,

df=4, *P*=0.75; Table 1). In pairwise tests between each allopatric and sympatric

population we also failed to reject the null hypothesis. Though we did not detect a signal

of reinforcement there was ample genetic variance in heterospecific mating rate between

female genotypes available for selection within each population (Fig. 1; Supplemental

Table 1). Only in one of the populations (Lamoille, which is allopatric) did the identity of

the *D. persimilis* tester line affect variation in premating isolation (Supplemental Table

1).

Reinforcement acts on conspecific sperm precedence

Unlike premating isolation, we observed a pattern consistent with reinforcement for

conspecific sperm precedence (CSP). Specifically, in sympatry we find both greater

average CSP (*t*=-6.5898, df=210.92, *P*<0.001; Wilcox *W*=4427.5, *P*<0.001) and less

phenotypic variation in this trait (Levene-type test χ^2=22.82, *P*<0.0001) when data were

pooled by geographic region (allopatry versus sympatry) (Table 1; Fig. 2A). These

differences in both the average and variance of CSP were also observed in pairwise tests

between individual allopatric and sympatric populations (Supplemental Table 2).

Reinforcement has collateral effects on intrapopulation sperm competition

ISC also differed between allopatric and sympatric populations, in both mean and

variance (Table 1; Fig 2B). First, mean offensive ability for ISC was significantly lower

in sympatric populations (t=3.738, df=246.55,P=0.0002; Wilcox's W=10280, P=0.0004).

This contrasts with the observed increase in offensive CSP in sympatric populations.

Second, there was more variation in ISC in the sympatric populations compared to the

allopatric populations (Leven-type test χ^2=5.74, P=0.0172). Given the differences in ISC

and CSP across populations, we used the mean CSP and ISC phenotype for each male x

female genotype combination within a population (i.e., each cell within the diallel

crossing design) to examine the pattern of relationship between the two phenotypes

across the four populations. We observed a significant negative relationship between CSP

and ISC (Pearson's r=-0.31, P=0.01; Fig. 5). Since each male or female genotype is

represented in multiple combinations we controlled for non-independence using a linear

mixed effect model, and confirmed that the negative slope of the relationship was

significant as indicated by a confidence interval that did not overlap zero (Profiled CI = -

0.451, -0.028).

Female genotype effects contribute to CSP and male x female genotype effects explain

both CSP and ISC

Of male, female, and male x female genotype effects that could contribute to explaining

the variance in CSP, we found that three out of the four populations had a significant

female genotype effect on CSP (Table 2; Fig 3), and all populations had a significant

male x female genotype interaction effect. The *D. persimilis* tester male line was also

significant in three out of four populations. There was no consistent pattern among

populations in which effect had the largest intraclass correlation (i.e. which explained the

largest proportion of variance; see Methods); in some populations the female genotype

effect had the largest intraclass correlation, while in others the male x female genotype

interaction had the largest intraclass correlation (Table 2). In contrast, for ISC in all four

populations we only observed significant male x female genotype interaction and a

significant effect of the first-tester male genotype (Table 3; Fig. 4). In every case, the

male x female genotype effect had a larger intraclass correlation (usually two to three

times greater) than the identity of the specific tester male genotype within each *D.*

pseudoobscura population.

The opportunity for sexual selection is decreased in sympatry

Our design allowed us to describe the reproductive success of males in terms of offensive

(second male) and defensive (first-tester male) success. We found that the sympatric

populations had significantly lower variance for reproductive success compared to the

allopatric populations (Figure 6; Supplemental Table 4). The variance in reproductive

success across all male genotypes (both offensive and defensive) in the allopatric

Lamoille population was significantly greater than both sympatric populations (Mt. St

Helena F=1.96, Bootstrap P=0.003; Sierra F=2.08, Bootstrap P=0.008), as was the

variance in reproductive success in the allopatric Zion population compared to the

sympatric populations (Mt. St Helena F=2.65, Bootstrap P=0.003; Sierra F=2.83,

Bootstrap P=0.004).This reduced variance in reproductive success in sympatry is a

product of lower offensive sperm competition values in sympatry, that result in equalized

differences in the siring success between offensive and defensive males.

DISCUSSION

Interactions with heterospecifics have the potential to drive divergent sexual selection

and the evolution of reproductive isolation, via reproductive character displacement and

reinforcement [6-7,44]. Using *D. pseudoobscura* and *D. persimilis*, we assessed whether

there was evidence for reinforcement of species barriers in sympatry via elevated female

preference or conspecific sperm precedence, traits that are known to contribute to

reproductive isolation across numerous taxa [2]. Premating isolation is historically

considered to be a strong barrier to isolation between these species, and one that

reinforcing selection has acted on [43], but we saw no evidence for reproductive

character displacement for this trait. In contrast we saw a clear signal of increased CSP in

sympatric populations, consistent with a pattern of reinforcement. Specifically, the

average CSP was higher, and the overall level of phenotypic variation was lower, in

sympatric populations, a pattern consistent with recent or recurrent directional selection

acting on CSP in these populations. We further asked whether reinforcement could have

collateral effects on intraspecific sperm competition and sexual selection, given that these

two traits are mechanistically and genetically linked [11,45]. We found that sympatric

populations also had lower ISC ability (lower offensive ability) than allopatric

populations, consistent with weakened sexual selection in sympatry.

Our results indicate that CSP can strongly contribute to reproductive isolation in

response to reinforcing selection. While CSP is known to be a barrier to gene flow in

Drosophila [12-13] and other taxa [2], its overall importance in nature has been difficult

to ascertain [14,16]. Moreover, previous studies of reinforcement sometimes qualitatively

describe variation in the target premating traits, but trait variance is typically not

quantified [5,9,17] even though models of speciation by sexual selection predict that

strong divergent selection will erode phenotypic variation in selected traits [46-47]. Our

observations of both increased mean CSP and reduced variation specifically in sympatry

provide compelling support for the inference that CSP has responded to strong selection

imposed by heterospecific interactions, and underscores the important role that CSP can

play in maintaining species boundaries.

The pattern of reproductive character displacement that we observed for CSP is

consistent with reinforcement, but other factors have been proposed to account for

reproductive character displacement including differential fusion [48] or ecological

differences that have collateral effects on mating traits [44,49]. Differential fusion

predicts that strong reproductive isolation evolves between species in allopatry and

merely prevents species collapse upon secondary contact, so that sympatric species

incidentally appear to have stronger isolation [50-51]. If differential fusion operates at the

deme/lineage level within a population we would expect the sympatric CSP values to be a

subset of allopatric CSP values [17]. This is not the case, however, because the sympatric

values of CSP are systematically higher than in allopatry (Figure 2). Regardless, the

differential persistence of demes/lineages with strong CSP in sympatry would

nevertheless be consistent with selection from standing variation leading to reinforcement

[52,53]. Similarly, several lines of evidence argue that systematic ecological differences

between allopatry and sympatry are unlikely to explain our observed postcopulatory

differences. Although both sympatric populations are located in California, they are

ecological distinct (collected from two different mountain ranges) in terms of numerous

ecological factors [54]. Indeed, habitat variation between sympatric populations of *D.*

pseudoobscura has led to differences in inversion frequencies maintained by ecological

forces that differ in these locations [55]. Moreover, the ecological differences across the

whole range of *D. pseudoobscura* are largely continuous, rather than uniquely

differentiating regions of allopatry and sympatry/co-occurrence with *D. persimilis*. Given

the ecological diversity between populations we do not expect a consistent direction of

natural selection acting on either the sympatric or allopatric populations. Arguably more

important, there are no established mechanisms whereby external ecological factors are

expected to have a direct effect on the strength of sperm competition consistent with our

observed pattern. Indirect effects of diet and nutrition can affect sperm competition

outcomes [56-57], but should not persist in the lab environment. Moreover, if ecological

mechanisms existed there is no reason to expect they would act in the specific direction

we observed here. Given this, while the ecological alternative to reinforcement might be

plausible for some premating phenotypes, it is unlikely to explain the postcopulatory

phenotypes that we examine here.

Our second major inference is that the response to reinforcing selection observed

in CSP has had a collateral effect on the magnitude of offensive ISC and the opportunity

for sexual selection in sympatric populations. The decrease in the opportunity for sexual

selection in sympatry appears to be the result of a negative genetic correlation between

CSP and ISC, as well as reduced variance in post-copulatory fitness based on ISC

estimates. Sperm competition strongly contributes to sexual selection in *D.*

pseudoobscura where multiple mating is frequent in wild caught females [58], and male

mating success, including sperm competition, is a major component of selection in

natural populations [59]. The observed reduction in offensive sperm competition differs

from both of our *a priori* expectations. One *a priori* hypothesis was that selection for

increased CSP in sympatry would select for increased offensive sperm competitive ability

among conspecifics, if offensive ability were a general trait that acted regardless of

whether the competitor was a conspecific or heterospecific male. In contrast, we observed

that ISC, was lower for sympatric populations compared to allopatric populations; that is,

average offensive ability was closer to 0.5, indicating a greater equalization in sperm

competitive ability among competing males. Our other *a priori* expectation was that

strong directional selection would alter sexual selection by reducing phenotypic variation.

However, the reduced phenotypic variation seen for CSP in sympatry was not mirrored

by reduced phenotypic variation for ISC. This observation is also inconsistent with an

alternative explanation-that selection for weaker ISC in sympatry indirectly increased

CSP. This alternative is more generally implausible as it requires that there has been

selection specifically to reduce ISC, solely in sympatry. Instead, we infer that selection

for stronger CSP in sympatry has reduced mean ISC in sympatric populations via a

negative genetic correlation between these two sperm competitive phenotypes.

For reinforcing selection to influence and interfere with sexual selection, the

selection favoring increased CSP must outweigh selection acting to maximize ISC. One

way CSP could have a larger effect on fitness than ISC is via a higher selective premium

specifically for females. Weaker CSP results in substantial fitness deficits for females

because of reproductive investment in low or no fitness hybrids, whereas weaker ISC

likely has a comparatively marginal effect on female fitness outcomes. Regardless, the

strength of reinforcing selection on CSP depends on the frequency of heterospecific

matings. Several lines of evidence suggest that heterospecific mating rates are common

between these species. First, from our data we observe a large range in the frequency

with which *D. pseudoobscura* females accept *D. persimilis* males in no-choice

experiments, with some genotypes on average accepting *D. persimils* 90% of the time.

Second, while no estimates for heterospecific mating rate exist from natural populations,

rare F1 progeny have been identified from wild collections [60]. Third, genetic evidence

suggests there has been post-speciation gene flow (i.e., evidence of movement of alleles

between species) between *D. pseudoobscura* and *D. persimils* [61-62]. Notably, these

estimates of realized gene flow will systematically underestimate the rate of

heterospecific matings, because they will only capture events that result in F1 progeny

that themselves then successfully reproduced; for example, given the presence of strong

CSP, many heterospecific matings may never produce hybrid progeny.

We were able to test the hypothesis that females face more costs of hybridization

[39-41] and that choice manifests as female control of sperm use patterns [63-65] by

contrasting the genotype effects (male, female, and male x female genotype effects)

between CSP and ISC. We observed significant male x female genotype interactions for

all populations for both CSP and ISC but, interestingly, only saw significant female

genotype effects for CSP. Significant female genotype effects for CSP suggest that

cryptic female choice may be operating similarly to premating isolation mechanisms

where females are observed to be the more "choosy" sex and female effects control the

level of reproductive isolation more so than male effects [66].

Strong female genotype effects on CSP are also consistent with the current

knowledge of postcopulatory sexual selection in the obscura group. Both *D.*

pseudoobscura and *D. persimilis* produce two sperm morphs: longer fertilizing eusperm

and shorter non-fertilizing parasperm. In *D. pseudoobscura*, the female reproductive tract

is spermicidal and higher proportions of parasperm help protect eusperm from these

negative effects [67]. Females in sympatric populations may have evolved more effective

spermicide against heterospecific males at a cost of spermicidal effectiveness with

conspecific males. In this case reproductive isolation would be mediated by cryptic

female choice and heterospecific male-female compatibility. This hypothesis may also be

consistent with our finding that the *D. persimilis* male genotype contributed significantly

to observed variation in CSP.

Reinforcement acting on CSP suggests that other prezygotic barriers that act

before CSP are not strong enough to limit the efficacy of selection on CSP in our

sympatric populations [14,16]. Indeed, our analysis of premating isolation (propensity to

mate with a heterospecific in the first mating) indicated that this potential barrier was

equally strong in sympatry and allopatry. This is interesting because one of the first

studies demonstrating reinforcement on premating barriers used the *Drosophila*

pseudoobscura and *D. persimilis* sister pair [43], although subsequent studies have found

more variable patterns [68-70; but see 71]. Our observation of a strong response in CSP

also suggests that the populations of *D. pseudoobscura* and *D. persimilis* we examined

are not strongly isolated by non-competitive (gametic) isolation, in agreement with

inferences from other studies of this specific species pair [70,72]. Though we lack data on

CSP from earlier collections in this species pair, our observations here might suggest that

the relative contribution of barriers to reproduction has changed in sympatry over time,

from premating isolation to CSP. Both gene flow between sympatry and allopatry, or a

cost to female premating preferences, might explain this shift over time. Depending on

the levels of gene flow among sympatric and allopatric populations, strong premating

isolation in sympatry could be lost due to "swamping effects" of allopatric gene flow [73]

or could lead to greater species wide reproductive isolation [74-75]. Our data suggest that

it's unlikely that gene flow from sympatry into allopatry created greater reproductive

isolation in allopatry (thereby reducing the signal of reinforcement) because the average

allopatric premating isolation in our experiment is similar to previous reports [43]. This

suggests that reduced premating isolation has emerged in sympatry, but it is difficult to

disentangle the effects of gene flow from the cost of female choice as causes of this

reduction. Both processes could contribute to the large variance we see for female

preference in sympatry compared to the more uniform level of premating isolation in

allopatry (Fig 1). The probability that strong female preference have been lost in

sympatry also depends on the frequency of this trait and any associated costs of

choosiness. When *D. pseudoobscura* stocks are kept in the absence of heterospecific

interactions female preference against heterospecifics decreases with longer periods of

experimental allopatry, suggesting that it may be costly to maintain this trait [76]. In

either case, the reduction in the strength of premating isolation in sympatry suggests that

this barrier to reproduction may only generate transient patterns of reinforcement.

Overall, our data suggest that strong reinforcing selection for reproductive

isolation can have consequences for sexual selection and sexual interactions, in these

important postmating sperm competition traits. The direction of this interaction provides

an interesting inversion to standard expectations about the connection between sexual

selection and speciation. Sexual selection is often thought of as a driver of sexual

characteristics whose evolutionary divergence then contributes to reproductive isolation.

But a direct genetic connection between these processes implies reproductive isolation

also has the reciprocal potential to shape sexual selection [77]. Based on our observations

of higher mean but lower variance in CSP in sympatry, a negative correlation between

CSP and ISC, and reduced variance in reproductive success via ISC among sympatric

conspecific males, we infer that strong selection for reproductive isolation within

populations exposed to heterospecific species has reduced the efficacy of sexual selection

in these populations, a collateral effect of reinforcing selection that has not previously

been demonstrated.

ACKNOWLEDGEMENTS

We would like to thank E. Walburn and J. Roesener for their assistance with crosses and

scoring progeny, J. Powers and the IU Light Microscopy Imaging Center for assistance

with the Leica microscope, M. Noor, A. Hish, and N. Phadnis for providing strains used

in this experiment, and Donn Castillo for help with collecting strains. Collections were

completed with assistance from IU Biology Department travel awards to DMC. Research

was supported by Indiana University Dept. of Biology funding to LCM and an American

Society of Naturalists student research award to DMC. DMC was supported by a

President's Diversity Initiative Dissertation Fellowship from the Indiana University

Graduate School.

LITERATURE CITED

1. Dobzhansky, T. 1951. Genetics and the Origin of Species. Columbia University Press

New York NY.

2. Howard, D.J . 1993. Reinforcement: origins, dynamics, and the fate of an evolutionary

hypothesis *in* R.G. Harrison, ed. Hybrid Zones and Evolutionary Process. Oxford

University Press, Oxford, pp. 46–69.

3. Servedio, M. R., and M. A. F. Noor. 2003. The role of reinforcement in speciation: Theory

and data. Ann. Rev. Ecol. Evol. 34: 339-364.

4. Howard, D. J. 1999. Conspecific sperm and pollen precedence and speciation. Ann. Rev. Ecol.

Evol. 30: 109-132.

5. Jaenike, J., K.A. Dyer, C. Cornish, M.S. Minhas. 2006. Asymmetrical reinforcement

and *Wolbachia* infection in *Drosophila*. Plos Biol. 4:1852-1862.

6. Higgie, M., and M.W. Blows. 2008. The evolution of reproductive character

displacement conflicts with how sexual selection operates within a species.

Evolution 62:1192-1203.

7. Higgie, M., and M.W. Blows. 2007. Are traits that experience reinforcement also

under sexual selection? Am. Nat. 170:409-420.

8. Porretta, D., and S. Urbanelli. 2012. Evolution of premating reproductive isolation

among conspecific populations of the sea rock-pool beetle *Ochthebius urbanelliae*

driven by reinforcing natural selection. Evolution 66:1284-1295.

9. Dyer, K.A., B.E. White, J.L. Sztepanacz, E.R. Bewick, and H.D. Rundle. 2014.

Reproductive character displacement of epicuticular compounds and their

contribution to mate choice in *Drosophila subquinaria* and *Drosophila recens*.

Evolution 68:1163-1175.

10. Kozak, G., M.G. Rolan, C. Rankhorn, A. Falater, E.L. Berdan, and R.C. Fuller. 2015.

Behavioral isolation due to cascade reinforcement in *Lucania* killifish. Am. Nat.

185:491-506.

11.Castillo, D.M., and L.C. Moyle. 2014. Intraspecific sperm competition genes enforce

post-mating species barriers in *Drosophila*. Proc. Roy. Soc. B281.

12. Price, C.S.C. 1997. Conspecific sperm precedence in *Drosophila*. Nature 388:663-666.

13. Chang, A.S. 2004. Conspecific sperm precedence in sister species of *Drosophila* with

overlapping ranges. Evolution 58:781-789

14. Marshall, J. L., M. L. Arnold and D. J. Howard, 2002 Reinforcement: the road not taken.

Trends Ecol. Evol.17: 558-563.

15. Servedio, M.R. 2001. Beyond reinforcement: the evolution of premating isolation by direct

selection on preferences and postmating, prezygotic incompatibilities. Evolution

55:1909–1920.

16. Lorch, P.D., and M.R. Servedio. 2007. The evolution of conspecific gamete

precedence and its effect on reinforcement. J. Evol. Biol.20:937-949.

17. Matute, D. R. 2010. Reinforcement of gametic isolation in Drosophila. PLoS Biol. 8: 3.

18. Hewitt, G.M., P. Mason, and R. Nichols. 1989. Sperm precedence and homogamy

across a hybrid zone in the alpine grasshopper *Podisma pedestris*. Heredity

62:343–353

19. Carney, S.E., M.B. Cruzan, and M.L. Arnold. 1994. Reproductive interactions

between hybridizing irises: analyses of pollen tube growth and fertilization

success. Am. J. Bot.81:1169–1175.

20. Metz, E.C. R. E. Kane, H. Yanagimachi and S. R. Palumbi. 1994. Fertilization

between closely related sea urchins is blocked by incompatibilities during sperm-

egg attachment and early stages of fusion. Biol. Bull. 187:23–34

21. Rieseberg, L.H. 1995. Interspecific pollen competition as a reproductive barrier

between sympatric species of *Helianthus* (Asteraceae). Am . J . Bot. 82, 515–519

22. Klips, R.A. 1999. Pollen competition as a reproductive isolating mechanism between

two sympatric *Hibiscus* species (Malvaceae). Am. J. Bot.86:269–272.

23. Williams, J .H., Jr., W.E. Friedman, and M.L. Arnold. 1999. Developmental selection

within the angiosperm style: using gamete DNA to visualize interspecific pollen

competition. Proc. Natl. Acad. Sci. U.S.A. 96:9201–9206.

24. Geyer, J.B., and S. R. Palumbi. 2005. Conspecific sperm precedence in two species or

tropical sea urchins. Evolution 59:97-105.

25. Peterson, M.A., E.L. Larson, M. Brassil, K.J. Buckingham, D Juarez, et al. 2011.

Cryptic gametic interactions confer both conspecific and heterospecific

advantages in the *Chrysochus* (Coloptera: Chrysomleidae) hybrid zone. Genetica

139:663-676.

26. Prout, T., and A.G. Clark. 1996. Polymorphism in genes that influence sperm

displacement. Genetics 144:401-408.

27. Civetta, A., and A.G. Clark. 2000. Correlated effects of sperm competition and

postmating female mortality. Proc. Natl. Acad. Sci. U.S.A. 97:13162-13165.

28. Fiumera, A.C., B L. Dumont and A.G. Clark. 2006. Natural variation in male-induced

'cost-of-mating' and allele-specific association with male reproductive genes in

Drosophila melanogaster. Phil. Trans. Roy. Soc. B 361: 355–361.

29. Clark, A.G. 2002. Sperm competition and the maintenance of polymorphism.

Heredity 88:148-153.

30. Zhang, R., A.G. Clark, and A.C. Fiumera. 2012. Natural genetic variation in male

reproductive genes contributes to non-transitivity of sperm competitive ability in

Drosophila melanogaster. Mol. Ecol. 22:1400-1415.

31. Clark, A. G., D. J. Begun, and T. Prout. 1999. Female x male interactions in

Drosophila sperm competition. Science 283:217-220.

32. Bjork, A., W.T. Starmer, D.M. Higginson, C. J. Rhodes, and S. Pitnick. 2007.

Complex interactions with females and rival males limit the evolution of sperm

offence and defence. Proc. Roy. Soc. B 274: 1779–1788.

33. Chow, C. Y., M. F. Wolfner, and A. G. Clark. 2010. The genetic basis for male x

female interactions underlying variation in reproductive phenotypes of

Drosophila. Genetics 186:1355-1365.

34. Wong, A., M.C. Turchin, M.F. Wolfner, and C.F. Aquadro. 2008 Evidence for

positive selection on *Drosophila melanogaster* seminal fluid protease

homologs.Mol. Biol. Evol. 25:497-506.

35. Panhuis, T.M., R. Butlin, M. Zuk and T. Tregenza, 2001 Sexual selection and speciation.

Trends Ecol. Evol.16: 364-371.

36. Kirkpatrick, M., and V. Ravigne. 2002 Speciation by natural and sexual selection:

Models and experiments. Am. Nat. 159:S22-S35.

37. Levitan, D.R. 2008. Gamete traits influence the variance in reproductive success, the

intensity of sexual selection, and the outcome of sexual conflict among

congeneric sea urchins. Evolution 62:1305-1316.

38. Pischedda, A. and W. R. Rice. 2012. Partiotioning sexual selection into its mating

success and fertilization success components. Proc. Natl. Acad. Sci.

U.S.A.109:2049-2053.

39. Trivers, R.L. 1972. Parental investment and sexual selection *in* B. Campbell, ed.,

Sexual selection and the Descent of Mane. Aldine, Chicago pp. 137-179.

40. Saetre, G.-P., M. Král, and S. Bures. Differential species recognition abilities of

males and females in a flycatcher hybrid zone. J. Avian Biol. 28:259-263.

41. Bonduriansky, R. 2011. Sexual selection and conflict as engines of ecological

diversification. Am. Nat. 178:729-745.

42. Manier, M.K., S. Lupold, J.M. Belote, W.T. Starmer, K.S. Berben *et al.*, 2013 Postcopulatory

sexual selection generates speciation phenotypes in *Drosophila*. Curr. Biol. 23: 1853-

1862.

43. Noor, M.A.F. 1995. Speciation driven by natural selection in *Drosophila*. Nature 375: 674-

675.

44. Pfennig KS, Pfennig DW. 2009. Character displacement: ecological and reproductive

responses to a common evolutionary problem. Q. Rev. Biol. 84:253-276.

45. Civetta, A., and S. Finn. 2014. Do candidate genes mediating conspecific sperm

precedence affect sperm competitive ability within species? A test case in

Drosophila G3 4:1701-1707.

46. Lande, R. 1981. Models of speciation by sexual selection on polygenic traits.

Proc. Natl. Acad. Sci. U.S.A. 78: 3721-3725.

47. Kirkpatrick, M. 1982. Sexual selection and the evolution of female choice. Evolution

36:1-12.

48. Templeton, A.R. 1981. Mechanisms of speciation- a population genetic apporach.

Ann. Rev. Ecol. Evol. 12:23-48.

49. Schluter, D. 2000. Ecological character displacement in adaptive radiation. Am. Nat.

156: S4–S16.

50. Noor, M. A. F. 1999. Reinforcement and other consequences of sympatry. Heredity,

83: 503-508.

51. Yukilevich, R. 2012. Asymmetrical patterns of speciation uniquely support

reinforcement in *Drosophila*. Evolution 66:1430-1446.

52. Ortiz-Barrientos, D., B. A. Counterman, and M.A.F. Noor. 2004. The genetics of

speciation by reinforcement. PLoS Biol. 2:e416.

53. Orr, H. A. & Betancourt, A. J. 2001. Haldane's sieve and adaptation from the

standing genetic variation. Genetics 157: 875–884

54. Baldwin, B G., D H. Goldman, D. J. Keil, R. Patterson, T. J. Rosatti, and D.

H. Wilken. 2012. The Jepson manual: vascular plants of California, second

edition. University of California Press, Berkeley, CA.

55. Schaeffer, S.W. 2008. Selection in heterogenous environments maintains the gene

arrangement polymorphism of *Drosophila pseudoobscura*. Evolution 62:3082-

3099.

56. Clark, S.C.A., N.P. Sharp, L. Rowe, A.F. Agrawal. 2012. Relative effectiveness of

mating success and sperm competition at eliminating deleterious mutations in

Drosophila melanogaster. PLoS One 7: e37351.

57. Zajitschek, F., S. Zajitschek, and M. Manier. 2017. High-protein paternal diet confers

an advantage to sons in sperm competition. Biol. Lett. 13: 20160914.

58. Anderson, W.A. 1974. Frequent multiple insemination in a natural population of

Drosphila pseudobscura. Am. Nat. 108:709-711.

59. Anderson, W.A., L. Levine, O. Olvera, J.R. Powell, M.E de la Rosa, V. M. Salceda,

M. I. Gaso, and J. Guzman. 1979. Evidence for selection by male mating success

in natural populations of *Drosophila pseudoobscura*. Proc. Natl. Acad. Sci.

U.S.A. 76:1519-1523.

60. Dobzhansky, T. 1973. Is there gene exchange between *Drosophila pseudoobscura*

and *D. persimilis* in their natural habitats? Am. Nat. 107:312-314.

61. Powell, J.R. 1983. Interspecific cytoplasmic gene flow in the absence of nuclear gene

flow: evidence from *Drosophila*. Proc. Natl. Acad. Sci. U.S.A. 80:492-495.

62. Kulathinal, R. J., L. S. Stevison, and M. A. F. Noor. 2009. The genomics of

speciation in *Drosophila*: Diversity, divergence and introgression on a genome-

wide scale. PLoS Genetics, 5: e1000550.

63. Eberhard, W.G. 1996. Female control: sexual selection by crypitc female choice.

Princeton University Press, Princeton, NJ.

64. Manier, M.K., J.M. Belote, S. Lüpold, K.S. Berben, O. Ala-Honkola, W.F. Collins, and S.

567 Pitnick. 2013a. Rapid diversification of sperm precedence traits and processes among

568 three sibling *Drosophila* species. Evolution 67: 2348-2362

569 65. Tyler, F., X.A. Harrison, A. Bretman, T. Veen, R. Rodriguez-Munoz, and T.

570 Trezenga. 2013. Multiple post-mating barriers to hybridization in field crickets.

571 Molec. Ecol. 22:1640-1649.

572 66. Andersson, M.M., and L.W. Simmons. 2006. Sexual selection and mate choice. Trends in

573 Ecology & Evolution 21:296-302.

574 67. Holman, L., and R.R. Snook. 2008. A sterile sperm caste protects brother fertile sperm from

575 female-mediated death in *Drosophila pseudoobscura*. Current Biology 18:292-296.

576 68. Anderson, W.W., and Y.K. Kim, 2005 Sexual isolation between sympatric and allopatric

577 populations of *Drosophila pseudoobscura* and *D. persimilis*. Behavior Genetics 35: 305-

578 312.

579 69. Anderson, W.W., and Y.K. Kim, 2006 A further analysis of sexual isolation between

580 sympatric and allopatric populations of *Drosophila pseudoobscura* and *D. persimilis* -

581 Rejoinder to Noor and Ortiz-Barrientos. Behavior Genetics 36: 328-330.

582 70. Davis, J.S., D.M. Castillo, and L.C. Moyle. 2017. Remating responses are shaped by

583 male post- copulatory manipulation but not reinforcement in *D. pseudoobscura*.

584 Ecology and Evolution 7:507-515.

585 71. Noor, M. A. F. and D. Ortiz-Barrientos. 2006. Simulating natural conditions in the

586 laboratory: A re-examination of sexual isolation between sympatric and allopatric

587 populations of *Drosophila pseudoobscura* and *D. persimilis*. Behavior Genetics,

588 36: 322-327.

589 72. Lorch, P.D., and M.R. Servedio 2005. Postmating-prezygotic isolation is not an

important source of selection for reinforcement within and between species in

Drosophila pseudoobscura and *D. persimilis*. Evolution 59:1039-1045.

73. Sanderson, N. 1989. Can gene flow prevent reinforcement? Evolution 43: 1223-1235.

74. Coyne, J.A., and H. A. Orr. 2004. Speciation. Sinauer Associates, Inc.

75. Ortiz-Barrientos, D., A. Grealy, and P. Nosil, 2009. The genetics and ecology of

reinforcement: implications for the evolution of prezyogtic isolation in sympatry

and beyod. The Year in Evolutionary Biology 1168:156-182.

76. Myers, E.M., and W. A. Frankino. 2012. Time in a bottle: the evolutionary fate of

species discrimination in sibling *Drosophila* species. PLoS One 2:e31759

77. Servedio, M.R., and R. Burger. 2014. The counterintuitive role of sexual selection in

species maintenance and speciation. Proceedings of the National Academy of

Science of the USA 111:8113-8118.

78. Boorman, E., and G. A. Parker. 1976. Sperm (ejaculate) competition in *Drosophila*

melanogaster, and reproductive value of females to males in relation to female

age and mating status. Ecological Entomology 1:145–155.

79. Lupold, S., S. Pitnick, K.S. Berben, C.S. Blengini, J.M. Belote, and M.K. Manier. 2013.

Female mediation of competitive fertilization success in *Drosophila melanogaster*. Proc.

Natl. Acad. Sci. U.S.A.110: 10693-10698.

80. Dixon, S.M., J.A. Coyne, and M.A.F. Noor. 2003. The evolution of conspecific sperm

precedence in *Drosophila*. Molec. Ecol. 12:1179–1184.

81. Markow TA, O'Grady PM. 2005 Drosophila: A guide to species identification and

use. Academic Press, Waltham Mass.

82. McGaugh, S.E., C.S.S. Heil, B. Manzano-Winkler, L. Loewe, S. Goldstein, T.L.

Himmel, M.A.F. Noor. 2012. Recombination modulates how selection affects

linked sites in *Drosophila*. PLoS Biol., 10: e1001423

83. Ortiz-Barrientos, D., and M.A.F. Noor. 2005. Evidence for a one-allele assortative mating

locus. Science 310:1467.

84. Holtzman, S., D. Miller, R. Eisman, H. Kuwayama, T. Niimi, and T. Kaufman. 2010

Transgenic tools for members of the genus *Drosophila* with sequenced genomes.

Fly 4:349-362.

85. Lesnoff, M., and R. Lancelot. 2012. aod: Analysis of Overdispersed Data. R package

version 1.3.

86. Hui, W., Y.R. Gel, and J.L. Gastwirth. 2008. lawstat: an R package for law, public

policy and biostatistics. J. Stat. Soft, 28.

87. Brown, M.B., and A. B. Forsythe. 1974. Robust tests for equality of variances.

J. Am. Stat. Ass. 69:364-367.

88. Warton, D.I., and F.K.C. Hui. 2011. The arcsine is asinine: the analysis of proportions

in ecology. Ecology 92:3-10.

89. Fox, J. 2008. Applied regression analysis and generalized linear models 2nd Ed. Sage

Publications, Inc. Thousand Oaks, CA.

90. Menard, S.W. 2010. Logistic regression: from introductory to advanced concepts and

applications. Sage Publications, Inc. Thousand Oaks, CA.

91. Halekoh, U., and S. Højsgaar. 2014. A Kenward-Roger approximation and parametric

bootstrap methods for tests in linear mixed models - The R Package pbkrtest. J.

Stat. Soft. 59:1-30.

92. Eldridge, S.M., O.C. Ukoumunne, and J.B. Carlin. 2009. The intra-cluster correlation

coefficient in cluster randomized trials: a review of definitions. Int. Stat. Rev.

77:378-394.

93. Wade, M.J. 1979. Sexual selection and variance in reproductive success. Am. Nat.

114:742-747.

94. Shuster, S.M., W.R. Briggs, and P.A. Dennis. 2013. How multiple mating by females

affects sexual selection. Phil. Trans. Roy. Soc. B :20120046.

95. Efron, B., and R.J. Tibshirani. 1993. An introduction to the bootstrap. Chapman and

Hall, Inc. London, UK.

96. Davison, A.C., and C.V. Hinkley. 1997. Bootstrap methods and their application.

Cambridge University Press, Cambridge, UK.

97. Schenker, N., and J.F. Gentleman. 2001. On judging the significance of differences

by examining the overlap between confidence intervals. Am. Stat. 55:182-186.

Table 1. The average levels of reproductive isolation for each *D. pseudoobscura*

population measured from two barriers to reproduction: female preference (proportion of

females that did not mate with heterospecifics) and conspecific sperm precedence (CSP).

Higher values indicate stronger reproductive isolation. Interpopulation sperm precedence

(ISC) is included for comparison. The mean and variance estimates for CSP and ISC are

based on 64 replicates per populations A = allopatric; S = sympatric

	Fem. Pref.	CSP		ISC	
Population	Proportion (n)	Mean	Variance	Mean	Variance
Lamoille (A)	0.481 (179)	0.75	0.054	0.76	0.028
Zion (A)	0.476 (145)	0.77	0.041	0.80	0.047
Mt. St. Helena (S)	0.540 (200)	0.90	0.017	0.79	0.057
Sierra (S)	0.505 (222)	0.92	0.018	0.68	0.052

655
656

657
658

Table 2. The genotype effects that predict CSP. The maximum likelihood estimate (ML

est.) and intraclass correlation (ICC) are reported as point estimates from the full model.

The *P*-value for each term was calculated by comparing the observed Likelihood ratio

test statistic (LR) to the distribution generated by parametric bootstrap. Data were

bootstrap sampled according to the null hypothesis where the random effect of interest is

not included. The full and reduced models are then fit to each bootstrap sample to

determine the distribution for the LR test statistic. A = allopatric; S = sympatric. Bold

indicates significance at *P*<0.05. Italics indicates marginal significance *P*<0.06.

Lamoille (A)				
Effect	ML est.	LR	*P*-Value	Intraclass Corr.
Female	**0.4024**	**8.10**	**0.0067**	0.096
Male	0.0000	0.00	0.7509	0.00
M x F	**0.1154**	**3.52**	**0.0383**	0.027
D. persimilis	**0.3413**	**37.49**	**0.0013**	0.082
Zion (A)				
Effect	ML est.	LR	*P*-Value	Intraclass Corr.
Female	*0.2683*	*2.72*	*0.05632*	0.067
Male	0.0000	0.00	0.4190	0.00
M x F	**0.3315**	**16.30**	**0.00238**	0.0833
D. persimilis	**0.0865**	**6.44**	**0.0068**	0.0217
Mt St. Helena(S)				
Effect	ML est.	LR	*P*-Value	Intraclass Corr.
Female	**0.8408**	**5.77**	**0.0068**	0.188
Male	0.0000	0.00	0.9891	0.000
M x F	**0.3266**	**8.76**	**0.0026**	0.0737
D. persimilis	0.0000	0.00	0.9851	0.000
Sierra (S)				
Effect	ML est.	LR	*P*-Value	Intraclass Corr.
Female	0.3287	0.72	0.1760	0.071
Male	0.1529	0.27	0.2673	0.033
M x F	**0.5975**	**7.28**	**0.0046**	0.129
D. persimilis	**0.2487**	**8.16**	**0.0012**	0.053

Table 3. The genotype effects that predict ISC. The maximum likelihood estimate (ML

est.) and intraclass correlation (ICC) are reported as point estimates from the full model.

The *P*-value for each term was calculated by comparing the observed Likelihood ratio

test statistic (LR) to the distribution generated by parametric bootstrap. Data were

bootstrap sampled according to the null hypothesis where the random effect of interest is

not included. The full and reduced models are then fit to each bootstrap sample to

determine the distribution for the LR test statistic. A = allopatric; S = sympatric. Bold

indicates significance at *P*<0.05.

Lamoille (A)				
Effect	ML est.	LR	*P*-Value	Intraclass Corr.
Female	0.0668	0.825	0.2131	0.018
Male	0.0000	0.000	0.5037	0.000
M x F	**0.2098**	**29.93**	**0.0023**	0.057
GFP male	**0.0879**	**23.88**	**0.0010**	0.024
Zion (A)				
Effect	ML est.	LR	*P*-Value	Intraclass Corr.
Female	0.3003	3.202	0.0647	0.074
Male	0.0405	0.170	0.3721	0.010
M x F	**0.3056**	**22.47**	**0.0022**	0.076
GFP male	**0.0835**	**12.21**	**0.0011**	0.020
Mt. St. Helena (S)				
Effect	ML est.	LR	*P*-Value	Intraclass Corr.
Female	0.0000	0.000	1.0000	0.00
Male	0.0184	0.096	0.4120	0.005
M x F	**0.2195**	**52.44**	**0.0019**	0.060
GFP male	**0.0825**	**35.24**	**0.0010**	0.022
Sierra (S)				
Effect	ML est.	LR	*P*-Value	Intraclass Corr.
Female	0.0000	0.000	0.3744	0.00
Male	0.0000	0.000	1.0000	0.00
M x F	**0.4139**	**70.85**	**0.0021**	0.111
GFP male	0.0077	0.902	0.0886	0.002

677
678
679

Figure 1. Prezygotic reproductive isolation via female mating preference does not show a

pattern consistent with reinforcement. Reproductive isolation is measured by the

proportion of females that did not mate with heterospecifics in individual no-choice trials.

Significant variation among *D. pseudoobscura* female genotypes in female preference

occurs in each population (Supplemental Table 1) . Each point is the mean reproductive

isolation for each isofemale line tested against each of four *D. persimilis* tester males.

Error bars represent ± one standard error.

689
690

Figure 2. The phenotypic distributions of CSP (panel A) is consistent with a pattern of reinforcement. The distribution of ISC (panel B) shows a shift in ISC in the opposite direction compared to CSP for sympatric populations. The red line in each distribution represents the mean value. Significant differences determined by Welch's t-test and Wilcox tests between the allopatric and sympatric populations is denoted by *.

698
699
Figure 3. Conspecific sperm precedence (CSP) for all male-female genotype

combinations in each population demonstrating a significant effect of female genotype

and male-female genotype interaction on the outcome of CSP. A) Lamoille-Allopatry, B)

Zion-Allopatry, C) Mt. Dt. Helena-Sympatry, and D) Sierra-Sympatry. Each point

represents a specific male-female genotype combination. Error bars are ± one standard

error. Female genotypes are ordered by mean CSP. Each color represents a single male

genotype for each population. Colors were re-used between each population panel, but

actual second male genotypes were unique to each population.

709
710
Figure 4. Intrapopulation sperm competition (ISC) for all male-female genotype

combination in each population demonstrating a significant male-female genotype

interaction on the outcome of ISC. A) Lamoille-Allopatry, B) Zion-Allopatry, C) Mt. Dt.

Helena-Sympatry, and D) Sierra-Sympatry. Each point represents a specific male-female

genotype combination. Error bars are ± one standard error. Female genotypes are

ordered by mean ISC. Each color represents a single male genotype for each population.

Colors were re-used between each population panel, but actual second male genotypes

were unique to each population.

Figure 5. The negative correlation between intrapopulation sperm competition (ISC) and

conspecific sperm precedence (CSP) across all four populations with each point

representing a male-female genotype combination. Blue points are from allopatric

populations and green points are from sympatric populations.

Figure 6. The variance in reproductive success across populations calculated in the

framework that combines offensive and defensive males. Each point represents the

estimate for the variance in fitness for each population. The error bars are confidence

intervals generated from the empirical bootstrap distribution. Significance, denoted by *,

was assessed in pairwise comparisons between allopatric and sympatric populations using

empirical bootstrap hypothesis testing.

MATERIALS AND METHODS

Wild type fly stocks

All stocks were reared on standard media prepared by the Bloomington Drosophila Stock

Center, and were kept at room temperature (~22C). We used a set of isofemale lines

collected from four natural populations in the summers of 2013 and 2014. Allopatric D.

pseudoobscura were collected at Zion National Park, UT (kindly provided by N. Phadnis)

and Lamoille Canyon, NV (collected by D. Castillo). Sympatric D. pseudoobscura and D.

persimilis were collected at two sites: Mt. St. Helena, CA (D. pseudoobscura collected by

A. Hish/M. Noor and D. Castillo, and D. persimilis collected by D. Castillo); and, near

Meadow Vista and Forest Hill, CA (called here 'Sierra'; D. pseudoobscura and D.

persimilis collected by D. Castillo). For both sympatric populations, both species were

present in field collections and can be considered truly co-occurring/sympatric.

Conspecific sperm competition assay

Sperm competition assays generally involve mating an individual female sequentially

with two distinct male genotypes. In all experimental crosses between species, females

were paired first with a D. persimilis male and second with a D. pseudoobscura male; that

is, the assays are evaluating the "offensive" sperm competitive ability of conspecific

males to displace heterospecific sperm (equivalent to 'P2', or second male siring ability;

[78]). We focused on "offensive" sperm competition because D. pseudoobscura females

do not remate with D. persimilis males if they have first mated with a conspecific,

therefore we cannot evaluate "defensive" sperm competition in this cross. In this

experiment we partitioned the variance in CSP due to male genotype, female genotype,

and the male x female genotype interaction using a "diallel-like" crossing design, which

is commonly used for this purpose [31,33; Supplemental Fig 1). A diallel cross is a

mating scheme commonly used to estimate the genetic effects, additive genetic variance,

and heritability, of quantitative traits by crossing all parental genotypes in all possible

combinations [79]. Our design is "diallel-like" because we did not use progeny from the

diallel to estimate heritability. We completed separate CSP experiments for each of our

four D. pseudoobscura collection locations (Sympatric= Sierra and Mt. St. Helena,

Allopatric= Zion and Lamoille). For each population we used a 4x4x4 design: four D.

pseudoobscura female genotypes from that population, four D. persimilis genotypes as

first males ("tester males"), and four D. pseudoobscura male genotypes as second males

from the same population as females. Each 4x4x4 combination was replicated once (n=64

unique cross combinations for each population). If CSP is important for reproductive

isolation in sympatry it should be consistently strong across multiple heterospecific

genotypes. Accordingly, rather than rely on a single D. persimilis genotype, we aimed to

use multiple wild-collected D. persimilis tester male lines for our experiments. Of these,

two D. persimilis lines were collected at the same time and in the same traps as the D.

pseudoobscura strains at the Sierra location and another two at the Mt. St. Helena

location.

Virgin individuals were collected and aged 7 days prior to the initiation of an

experimental block. One day before mating, D. persimilis tester males were isolated

individually [80]. The following day, females were individually added (without

anesthesia) to a vial containing a tester male and were co-housed for 24 hours, after

which time the tester male was removed. We kept females housed individually in these

vials for 7 days before second mating (similar to [80]). After 7 days we inspected all vials

for the presence of larvae to determine if females had mated with the first D. persimilis

tester males. This was used to evaluate evidence for differences in successful first

matings (pre-mating isolation) among allopatric and sympatric populations, rather than

observing matings directly, as there is high variance in time to copulation in this

heterospecific pairing [70]. Only females that had mated (i.e. had produced larvae within

7 days) were retained for the remainder of the CSP experiment.

For the second mating, each individual female was paired with one of the four D.

pseudoobscura male genotypes from her own population to determine the strength of

CSP. These second males were also isolated one day before the introduction of the

female. Seven days after mating with the first male, females were transferred, without

anesthesia, to the vial containing the second male. Individual pairs were co-housed for 24

hours and the male was removed on the second morning. The female was kept for five

days (transferring after 2 days to avoid overcrowding of larvae). All progeny produced in

the five-day window after the second mating were collected; from these progeny a

maximum of 10 males and 10 females, randomly chosen from the total group of progeny,

were used to score CSP (P2) as described below.

Intrapopulation sperm competition assay

The design for intrapopulation sperm competition (ISC) assay mirrored the experimental

design for CSP except that, rather than a D. persimilis tester male, the first male was a D.

pseudoobscura tester male derived from the same population as the D. pseudobscura

female and second male genotypes in the trial. For each population we used a 4x2x4

design: four D. pseudoobscura female genotypes, two D. pseudoobscura GFP genotypes

as first males and 4 D. pseudoobscura male genotypes as second males. The same female

x second male genotypes were used in ISC and CSP experiments. Each combination was

replicated twice (n=64 for each population, with 32 unique cross combinations). This

allowed us to have a total sample size per population that matched the CSP experiment

(64 replicates per population, 256 replicates across all populations).

The details of the mating scheme (virgin collection, aging of individuals, isolation of

individuals, etc.) are identical to the CSP experiment. We did not observe matings

directly, but the average refractory period for D. pseudoobscura is 4 days [81], so we are

confident that on average only a single mating occurred in the 24 hour co-housing

timeframe. Each individual female was randomly assigned one of the two D.

pseudoobscura first male (tester) genotypes to determine the strength of P2 (second male

siring ability) by our four focal second male genotypes, against these tester male

genotypes. The female was kept for five days after the second mating (transferring after 2

days to avoid overcrowding of larvae). All progeny produced in the five-day window

after the second mating were collected and scored.

Generating visibly-marked tester males for quantifying CSP and ISC

To allow efficient progeny scoring, paternity was scored with the aid of visible markers

in both CSP and ISC experiments. This required us to generate marked male tester lines

with wild-caught D. persimilis (for CSP tester males) and D. pseudoobscura (for ISC

tester males) lines from each study population.

For CSP, to introduce a visible marker into wild-type wild-collected D. persimilis males

from our sympatric sites, we introgressed an X-linked marker ("short" or sh) from a D.

pseudoobscura line, into four of our collected D. persimilis genotypes (Supplemental Fig

2). These four D. persimilis tester males originated from isofmale lines collected at the

Sierra and Mt St. Helena locations and were used to evaluate the mean strength and

variation in CSP for all four D. pseudoobscura populations in the CSP experiment. We

first crossed these D. persimilis sh mutant males to females from each of the four wild-

type D. persimilis isofemale lines (keeping each tester genotype separate throughout this

process). This produced F1 daughters heterozygous for the sh allele, that were

backcrossed to wild type males from the same wildtype isofemale line. From the BC1

progeny we retained sh males, and these were backcrossed to the original D. persimilis

isofemale line to generate BC2s (Supplemental Fig 2). This process of alternating males

and females for each backcross generation within each D. persimilis isofemale line was

completed until the BC12. The alternation of male/female during backcrossing was

necessary because recombination only occurs in females, but to retain the marker we had

to select for sh males every second generation. After the BC12, the progeny within each

BC isofemale line were interbred to create males and females homozygous for the sh

allele. We did not directly evaluate how much of the sh line genome was introgressed in

each case, however, D. pseudoobscura and relatives have a much higher recombination

rate than D. melanogaster [82], and previous introgression lines between these species

have eliminated unwanted regions after 4 generations of backcrossing [83].

For ISC experiments, the marked tester males were created by introgressing a green

fluorescent protein marker (GFP) into 2 wild type D. pseudoobscura strains per location

(therefore 8 strains in total, using wild-collected isofemale lines that were not used as

female or second male genotypes for the ISC experiments). The original GFP strain was

obtained from the UCSD stock center (14011-0121.166) the creation of which is

described in Holtzman et al. [84]. We chose this marker because it is dominant [11]. We

chromosomally mapped the GFP insertion of the original GFP strain to the second

chromosome using a multiply marked (MM) strain, which contains visible recessive

markers on all of the major chromosomes (y;gl;or;inc kindly provided by N. Phadnis,

University of Utah). This mapping was completed in order to ensure the GFP insertion

was not on the 3rd chromosome which, in D. pseudoobscura, contains large inversions

that would have inhibited recombination of the marker into the wild-type backgrounds of

our D. pseudoobscura isofemale lines.

The original GFP line was created in a stock that carried the X-linked white mutation. To

eliminate the white allele from the population, in the parental cross we crossed the WT

line with the GFP carrying male, and then used only F1 males with wild-type X

chromosomes (no white mutation) to backcross in this initial generation. For the

remaining eight backcross generations, we used females to allow recombination. We then

chose 10 sibling pairs for each genotype to ensure the GFP marker was homozygous.

These sub-lines were inbred for two generations. In the second generation we testcrossed

the founder pair of individuals of each sub-line to ensure they were homozygous for the

GFP marker. We recovered 2-4 lines that were homozygous for the GFP marker for each

genotype. We then combined inbred lines that had originated from the same isofemale

genotype to reduce any potential effects of inbreeding depression that might have arisen

during marker introgression.

Scoring conspecific sperm precedence

Hybrid male progeny from D. pseudoobscura x D. persimilis crosses are sterile (there are

no motile sperm, observable by dissecting the testes). We used this sterility phenotype to

differentiate the male progeny of heterospecific versus conspecific males and therefore to

score CSP. For a given replicate we collected and dissected 10 male progeny that were

produced after the second mating. Each male was dissected individually in PBS buffer,

and its testes moved to a slide that had 1ul of PBS buffer. A cover slip was placed over

the slide and the testes were squashed, releasing sperm into the buffer. The slides were

examined under an EVOS FL microscope for the presence of motile sperm. If no motile

sperm were present, the male was scored as hybrid.

Because female hybrids are fertile in these crosses, the sh allele was used to

differentiate the female progeny of heterospecific versus conspecific males and therefore

to score CSP from female offspring. Since the sh allele is recessive we could not score F1

females directly, but instead scored their offspring for the presence of the sh allele. If an

F1 female was hybrid (and carrying the sh allele from the D. persimilis male) we would

expect half of her sons and half of her daughters to have the sh phenotype. We previously

confirmed that the half segregation held for known hybrid progeny. For each cross, ten

F1 females (that could be hybrid or purebred) were housed individually with a D.

pseudoobscura male that also carried the sh allele (UCSD stock center Dpse co;sh 14011-

0121.13). We chose a D. pseudoobscura male for these crosses to increase the number of

progeny to score since D. pseudoobscura females (and therefore any purebred female

progeny in our experiment) exhibit premating isolation with D. persimilis males; hybrid

females do not demonstrate a mating preference. After a week the parental individuals

were cleared from the vials and the vials were retained to score progeny. As progeny

eclosed they were scored for the presence of sh allele. Any F1 female that produced sh

progeny was considered hybrid. We required each F1 female to produce at least 10

progeny to be used in scoring CSP.

Our measure of CSP was then the number of purebred progeny out of the total

number of F1 individuals scored for a particular cross. If all progeny produced in a cross

were scored as hybrid, we did not use this replicate in our analyses because we could not

ensure that a second mating had taken place. Note that the frequency of this failure to

remate following a first mating does not differ between populations [70]. Every CSP

estimate was based on at least 10 scored progeny and, for the majority of the crosses, we

scored close to 20 individuals. In addition, to ensure that CSP estimated here does not

simply reflect stronger fecundity stimulation by conspecific males, in a pilot experiment

we determined that there was no difference in progeny production in heterospecific vs.

conspecific matings for any of the allopatric or sympatric populations, consistent with

previous work [70,72]. There was also no correlation between the total number of

progeny scored for CSP and the magnitude of CSP, and the number of progeny scored

did not differ between populations. These observations suggest that there are no

postzygotic survivorship barriers in hybrids between these species that would

systematically differ between sympatric and allopatric populations, confounding our

estimate of CSP.

Scoring intrapopulation sperm competition

We scored all progeny that eclosed in the five days after the second mating for the

presence/absence of the GFP phenotype. Our measure of sperm competition (P2) for ISC

was then the number of wild-type (non-GFP) progeny out of the total number of progeny

scored for a particular cross. If all progeny produced in a cross were GFP, we did not use

this replicate because we could not ensure that a second mating had taken place. (As with

CSP, the proportion of females that did not remate was not significantly different

between populations). Individuals were scored as they eclosed, using a Leica M205FA

Stereo Microscope that has an Hg fluorescent lamp attached and GFP filter. Individuals

were anesthetized and the ocelli were examined for GFP signal as described in Castillo

and Moyle [11].

Statistical analyses

All analyses were completed in R v 3.01.

Differences in the probability of first mating with heterospecifics

We evaluated evidence for a pattern consistent with reinforcement acting on first mating

(simple prezygotic isolation) in two ways. First, we used a $\chi2$ test of independence to test

the null hypothesis that the mating rate with heterospecifics was the same for alternative

geographic scenarios (allopatric vs. sympatric), after combining both allopatric and both

sympatric populations for this single comparison (pairwise tests among individual

populations gave the same result). Second, because $\chi2$ tests might lack power, and since

mating events can be coded as a binary variable (0 for did not mate, 1 for successful

mating), we used a logistic regression model with all four populations represented by a

categorical variable using the glmer function. We then tested whether there were any

differences in heterospecific mating between populations by conducting a Wald's test

(using the wald.test function from the aod package; [85]).

To evaluate whether there was significant variation within each population (i.e.,

among isofemale line genotypes) in the probability of mating with a heterospecific, we

used logistic regression. We first fit a full model where the probability of mating with a

heterospecific depended on the isofemale line, the D. persimilis tester line, and the male x

female genotype interaction, and tested significance of these effects using a Wald's test.

Because there was no significant interaction for any population, we fit a reduced model

that only contained the effects of isofemale line and D. persimilis tester line without the

interaction, and report these models in the results.

Differences in mean and variance of CSP and ISC between populations

We evaluated evidence for a pattern in CSP consistent with reinforcement, by evaluating

whether the allopatric and sympatric populations had a mean difference in CSP or

whether they differed in variance. For analyses of mean differences, we pooled the two

allopatric populations because there was no significant difference in mean CSP between

them (Allopatry $t = -0.45064$, df $= 123.62$, $P = 0.653$) and pooled the two sympatric

populations for the same reason (Sympatry $t = -0.86678$, df $= 125.87$, $P = 0.3877$). We

tested the hypothesis that the mean CSP differed between geographic scenarios using a

Welch's t-test that accounts for unequal variances between samples, and (given that the

data are not normally distributed) we also confirmed these results with a Wilcoxon

ranked sum test. To evaluate differences in variance, we again pooled the allopatric and

sympatric populations because the variance was equivalent between allopatric

populations ($\chi2 = 0.031899$, $P = 0.8585$), and between sympatric populations ($\chi2=$

0.80562, $P = 0.3711$). We compared the total phenotypic variation between geographical

classes of population with a Levene-type test implemented in the lawstat package in R

[86]. The specific test we used in the lawstat is a Kruskal-Wallis modified Brown-

Forsythe Levene-type test. The Brown-Forsythe test is based on the absolute deviations

from the median, which retains statistical power for many types of non-normal data [87].

Kruskal-Wallis tests are rank-based tests. We used the Kruskal-Wallis modification

because the variance in proportion data derived from binomial data does not accurately

reflect variance in the original data [88].

Using the same statistical approach as for CSP, we tested for differences in the mean and

variance between sympatric and allopatric populations for ISC, again pooling the

individual allopatric and sympatric populations as they were not significantly different

from one another for either measure (Allopatric mean t=-1.136, df=118,66, P=0.2593;

Sympatric mean t=0.191, df=125.72, P=0.8488; Allopatric variance $\chi 2$=0.949, P=0.3316;

Sympatric variance $\chi 2$=0.0796, P=0.7782). Note that, although we report results from

tests with these pooled data in the main text, we also observed significant differences in

pairwise tests between individual allopatric and sympatric populations, for both average

and variance measures of CSP and ISC (Supplemental Tables 2 and 3).

Genetic variation and genotype effects on CSP and ISC

Within each population we assessed whether female, male, or female x male

genotype predicted variation in the strength of CSP and ISC. While this can be tested

using a two-way ANOVA with interaction, the assumptions of ANOVA, including

normally distributed residuals and heterogeneity in the distribution of the residuals, are

typically violated by binomial data such as our sperm competition data [89-90]. We

instead chose binomial regression, as this more naturally models our count/binomial data.

The model is of the form

988
$$logit(p_{ijk}) = \mu + \alpha_i + \beta_j + (\alpha\beta)_{ij} + \varepsilon_{ijk}$$

The variable α is a categorical variable with four levels that represents male genotype.

The variable β is also a categorical variable with four levels that represents female

genotype. The variable (αβ) represents the male x female genotype interactions. Since we

were interested in partitioning the variance and estimating the variance components (σ_α^2,

$\sigma_\beta^2, \sigma_{\alpha\beta}^2$) we assumed that each variable was a random variable. To test the significance

of each variance component, we used a binomial regression in a mixed modeling

framework with parametric bootstrap [91]. In this bootstrap procedure, data are simulated

from the null model which lacks the random effect of interest. Then the full and reduced

models are fit to the simulated data to determine the bootstrap distribution of the

Likelihood Ratio test statistic. To the model above we also included a random effect of

tester male (*D. persimilis* for CSP and GFP *D. pseudoobscura* strain for ISC). To provide

an assessment of the relative importance of each variable we calculated the intraclass

correlation for each coefficient; a high correlation indicates that the variable explains

much of the variance in the data. The ICC for the female effect, for example, would be:

1003
$$ICC_F = \frac{\sigma_F^2}{\sigma_F^2 + \sigma_M^2 + \sigma_{MF}^2 + \sigma_T^2 + \frac{\pi^2}{3}}$$

Where *F* represents female variance, *M* represents male variance, *MF* represents the

interaction, and *T* represents the identity of the tester male. The $\frac{\pi^2}{3}$ replaces the residual

variance for the binomial model with logit link function. In the case of binomial

regression the ICC values are on the log scale, and there is no convenient transformation

to proportion scale [92], so they are presented here as a relative measure of variance

explained.

1010

Quantifying sexual selection and variance in male reproductive success

To evaluate whether the intensity/opportunity for sexual selection differs among

populations we require an estimate of variance in male reproductive success [93]. In a

natural population most males can gain fitness through offensive (P1) and defensive (P2)

sperm competition, so the best estimate for variance in reproductive success would be

total progeny produced. In our experiment we did not score lifetime progeny production,

and specific male genotypes were either used as offensive or defensive males only. As

such we estimated male fitness as the proportion of progeny sired, taking into

consideration that we had two distinct classes of males—tester first (defensive) males and

second (offensive) males--that may differ in their frequency and variance in fitness in the

experiment. Following Shuster et al. [94] we define total variance in male reproductive

success as the sum of within and between male class variance

$$V_{total} = (f_{P1})(V_{P1}) + (f_{P2})(V_{P2}) + (\bar{X}_{P2} - \bar{X}_{P1})^2(f_{P2})(f_{P1})$$

The two terms on the left hand of the equation represent the within class variance (for

example, V_{P1} is the variance in sperm competitive success between tester males and f_{P1} is

the frequency of tester males used in the experiment). The last term represents the

between class variance.

We were interested in reproductive variance at the level of male genotype so we averaged

biological replicates to generate mean fitness values for each individual genotype. We

used empirical bootstrap confidence intervals to estimate error that may have been a

product of averaging over replicates [95-96]. For the bootstrap procedure we sampled 16

data points, with replacement, from the 16 original empirical replicates for each genotype

(32 for defensive males). We then averaged these data points and calculated V_{total} as

described above. We completed 1000 bootstrap replicates for each population. We

constructed the 95% confidence interval using the bootstrap difference $\delta^* = V_{total} -$

V_{total}^* where * represents each bootstrap replicate. The interval is then $[V_{total} -$

$\delta^*_{0.05}, V_{total} - \delta^*_{0.95}]$.

The confidence intervals for the Zion population did not overlap with the confidence

intervals for either sympatric population and can be considered significantly different at

the 0.05 level (Supplemental Table 4). The Lamoille population confidence intervals

overlapped with the sympatric populations, but overlap in confidence intervals does not

mean parameters are not statistically different [97]. This is because confidence intervals

calculated for independent parameters cannot replace a comparative test of the

differences between two parameters. Therefore, we conducted bootstrap hypothesis

testing [95-96] to determine whether differences in V_{total} between populations were

significant, specifically by calculating bootstrap F statistics. The F statistic is a ratio of

any two variance parameters, for example $F = V_{total,pop1} \Big/ V_{total,pop2}$. We compared the

V_{total} in pairwise comparisons following standard bootstrap methods, where bootstrap

samples are generated under the null hypothesis, and then this distribution is compared to

the empirically observed statistic. For our scenario, the null hypothesis was that there was

no differences in V_{total} between populations. Therefore we sampled, with replacement,

offensive and defensive genotypes after pooling data from both populations. We then

randomly assigned each value to one of the two populations. This generated a bootstrap

replicate with approximately equal variance between the populations. We then could

calculate $F = {V_{total,pop1}}\big/{V_{total,pop2}}$ for each replicate. The bootstrap p-value is then

calculated by comparing the bootstrap statistic (F^*) to the observed statistics (F) using

$p^*(F) = \frac{1}{B}\sum_{j=1}^{B} I(F^* > F)$. I() is an indicator function that is equal to 1 when the

argument is true (boostrap statistic > observed statistic), and 0 when false. B is the

number of bootstrap replicates (1000 per population comparison).