Larger Numbers Can Impede Adaptation in Asexual Populations despite Entailing Greater Genetic Variation

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Abstract

Periodic bottlenecks play a major role in shaping the adaptive dynamics of natural and laboratory populations of asexual microbes. Here we study how they affect the 'Extent of Adaptation' (*EoA*), in such populations. *EoA,* the average fitness gain relative to the ancestor, is the quantity of interest in a large number of microbial experimental-evolution studies which assume that for any given bottleneck size (N_0) and number of generations between bottlenecks (*g*), the harmonic mean size *(HM=N_0g)* will predict the ensuing evolutionary dynamics. However, there are no theoretical or empirical validations for *HM* being a good predictor of *EoA*. Using experimental-evolution with *Escherichia coli* and individual-based simulations, we show that *HM* fails to predict *EoA* (i.e., higher N_0g does not lead to higher *EoA*). This is because although higher *g* allows populations to arrive at superior benefits by entailing increased variation, it also reduces the efficacy of selection, which lowers *EoA*. We show that *EoA* can be maximized in evolution experiments by either maximizing N_0 and/or minimizing *g*. We also conjecture that N_0/*g* is a better predictor of *EoA* than N_0g. Our results call for a re-evaluation of the role of population size in predicting fitness trajectories. They also aid in predicting adaptation in asexual populations, which has important evolutionary, epidemiological and economic implications.

Introduction

Population size is a key demographic parameter that affects several ecological and evolutionary processes including the rate of adaptation (Desai et al. 2007; Desai and Fisher 2007; Gerrish and Lenski 1998; Lanfear et al. 2014; Samani and Bell 2010; Wilke 2004), efficacy of selection (Petit and Barbadilla 2009), organismal complexity (LaBar and Adami 2016), fitness decline (Katju et al. 2015), repeatability of evolution (Lachapelle et al. 2015; Szendro et al. 2013; Vogwill et al. 2016), etc. Interestingly though, what constitutes a useful measure of population size for predicting evolutionary outcomes often depends on the ecological/evolutionary question being addressed and the population-genetics quantity in question (Charlesworth 2009). Consequently, it is crucial to use the relevant measure of population size while constructing or empirically validating any evolutionary theory.

Experimental evolution using asexual microbes has been one of the key tools in validating several tenets of evolutionary theory ((Kassen 2014), reviewed in (Kawecki et al. 2012)). Most such studies deal with populations that face regular and periodic bottlenecks during their propagation (Kawecki et al. 2012). The absolute population size keeps changing regularly because of these periodic bottlenecks. Therefore, in order to make predictions and claims based on population size in such experiments, it is important to define a proper measure of population size depending upon the question of interest (Charlesworth 2009; Kawecki et al. 2012; Lanfear et al. 2014; Wang et al. 2016).

Previous theoretical studies have shown that the harmonic mean of population size *(HM)* over time acts as the measure of population size that can explain and predict the fixation probabilities of beneficial mutations in such systems (Patwa and Wahl 2008; Wahl and Gerrish 2001). Specifically,

if a population grows from size N_0 to N_f via binary fissions within a growth phase, and is diluted

back periodically to N_0 by random sampling at the end of the growth phase, then the relevant

population-size measure for fixation probabilities is given by $N_e \approx N_0*log_2(N_f/N_0) = N_0g$, where g

refers to the number of generations between successive bottlenecks and N_0g is the harmonic mean

size (Lenski et al. 1991). From an evolutionary perspective, periodic bottlenecks play two opposite

roles in such experiments. On the one hand, harsher bottlenecks (entailed by higher g) reduce the

probability that a given beneficial mutation would fix due to sampling errors during the bottleneck.

On the other hand, higher values of g also imply an increase in N_f, which causes an increase in

mutational opportunities (binary fissions) during exponential growth (Wahl et al. 2002). This is

expected to increase the total supply of mutations that would survive drift, which in turn should

increase the raw material available for evolution. It has been predicted that exponential growth

between N_0 and N_f influences fixation probabilities more than the elimination by sampling

(Heffernan and Wahl 2002). More nuanced and complex measures of population size (Campos

and Wahl 2009, 2010) also suggest that adaptation rates in terms of fixation probabilities would

have a positive relationship with N_0 and g, the two population size parameters amenable to

experimental manipulation.

Unfortunately, most experimental evolution studies with serially bottlenecked asexual populations

do not focus on the fixation probabilities of beneficial mutants. Instead, they are interested in the

average amount of fitness gained with respect to the ancestor at a given time (we call this quantity

the extent of adaptation, *EoA*) during the course of evolution (De Visser and Rozen 2005; Desai

et al. 2007; Lachapelle et al. 2015; Lenski et al. 1991; Rozen et al. 2008; Samani and Bell 2010).

Several experimental studies, dealing with quantities akin to *EoA* for periodically bottlenecked

asexual populations, have used *HM ($=N_0g$)* for quantifying the evolutionarily relevant (i.e.

predictive of the magnitude of evolutionary response) population size (De Visser and Rozen 2005;

Desai et al. 2007; Lachapelle et al. 2015; Lenski et al. 1991; Rozen et al. 2008; Samani and Bell

2010). However, to the best of our knowledge, there is no theoretical basis or empirical

justifications (Y. Raynes et al. 2014) for this usage of *HM*. Here we use a combination of agent-

based simulations and long-term evolutionary experiments using *Escherichia coli* to investigate

the interplay of N_0 and *g* in shaping the *EoA* of asexual populations. Since *HM* has been widely

used by experimentalists in the context of *EoA*-like quantities, we begin by testing the suitability

of *HM* as a predictor of *EoA*. We show that populations with similar values of N_0g can have

markedly different *EoA* trajectories, and this result applies to both real (bacterial) as well as

simulated populations. Secondly, we demonstrate that although increasing the value of *g* (making

the periodic bottleneck harsher) promotes adaptation through an increased supply of variation, it

also reduces the efficacy of selection which impedes adaptation by restricting the spread of large-

effect beneficial mutations. When these two opposing aspects of bottlenecks are considered

together, counterintuitively, *EoA* turns out to have a negative relationship with *g*. Thirdly, we show

that populations with similar *HM* can not only have different fitness trajectories but can also differ

markedly in terms of how frequency-distribution of fitness amongst individuals changes during

adaptation. Finally, we show that, for a given mutation rate, N_0/g can be a better predictor of *EoA*

trajectories, *i.e.*, populations with similar N_0/g have similar fitness trajectories and populations

with higher N_0/g adapt faster. Our findings thus introduce a new way of thinking about the

relationship between population size and adaptive trajectories.

Our approach differs from previous studies in two important ways. First, unlike many studies

(Campos and Wahl 2009, 2010; Heffernan and Wahl 2002; Wahl and Gerrish 2001) we focus on

how *EoA* (and not long-term fixation probabilities) is shaped by bottleneck size (N_0) and bottleneck

ratios (N_0/N_f). This makes our study directly relevant to a rich body of microbial experimental

evolution literature ((De Visser and Rozen 2005; Desai et al. 2007; Lachapelle et al. 2015; Lenski

et al. 1991; Rozen et al. 2008; Samani and Bell 2010), reviewed in (Kawecki et al. 2012)). Second,

many previous theoretical studies on periodically bottlenecked systems (where $N_f = N_0 2^g$), assume

that the culture volume (and therefore N_f) is a constant, and then go on to explore what value of N_0

or g leads to the minimum loss of variation during bottlenecks and/or in the long run (Campos and

Wahl 2009, 2010; Heffernan and Wahl 2002; Wahl et al. 2002; Wahl and Gerrish 2001; Wahl and

Zhu 2015). In our simulations, we remove this restriction and seek to compare loss of variation in

those cases where both N_0 and N_f can be different (e.g. between a population grown in 50 ml of

medium versus one grown in (say) 1 ml of medium). Clearly, it is possible to have two populations

with very different N_0 and N_f values that can nevertheless have similar values of $N_0 g$. One of the

questions that we investigate is whether such populations have similar fitness trajectories or not.

Thus, our results make it possible to compare the expected *EoA* across experimental studies that

employ similar environments but different culture volumes, which is a rather common scenario in

experimental evolution studies (Lachapelle et al. 2015; Y. Raynes et al. 2014; Yevgeniy Raynes

et al. 2012; Rozen et al. 2008; Samani and Bell 2010).

Methods

Experimental evolution, assays and statistical analysis

Here we present a brief description of the experimental protocol, relegating the details to Supplementary Methods. Our primary aim was to investigate if a commonly used measures of population size in experimental evolution, namely harmonic mean (N_0g), could predict *EoA* trajectories. We also wanted to see if populations with similar values of N_f have similar *EoA*. To this end, we experimentally evolved three different population regimens (LL, SL, and SS) in Nutrient broth containing a sub-lethal cocktail of three antibiotics (Norfloxacin, Rifampicin and Streptomycin) for ~380 generations in batch culture. Each regimen consisted of 8 independently evolving replicate populations, all of which were started from a single *Escherichia coli* MG 1655 colony. The three population regimens were propagated at different bottleneck sizes: LL faced lenient bottlenecks (1/10), whereas SS ($1/10^4$) and SL ($1/10^6$) experienced much harsher bottlenecks. LL and SL were grown at larger culture volumes (100 ml, culture in flasks) than SS (1.5 ml, culture in 24 well-plates). Thus, in terms of N_f, LL = SL >> SS but in terms of N_0g LL >> SL = SS (see Table S1 for the values of these parameters). The *EoA*-trajectories of the three population types were reconstructed by assaying the maximum population-wide growth rates (R) and carrying capacities (K) of each replicate at different time-points during evolution following standard protocols (Karve et al. 2015, 2016), the details of which can be found in Supplementary Methods. K of a population was defined as the maximum OD value attained over a period of twenty four hours (the highest value in the sigmoidal growth curve) (Karve et al. 2016; Novak et al. 2006). R was estimated as the maximum slope of the growth curve over a running window of four OD readings (each window spanning one hour) (Karve et al. 2015, 2016; Vogwill et al. 2016).

To analyze the data, we performed separate repeated measures (RM) ANOVA for K and R.

"Regimen-type" (LL/SL/SS) was treated as the categorical factor, and TIME (9 time-points) as the

repeated measures factor. We also included an interaction of Regimen-type and TIME in our model

to determine if the fitness trajectories of the regimes were significantly different from each other.

Furthermore, to compare the fitness values at each time point, we used a nested-design ANOVA

with "regimen-type" (SS, SL or LL, fixed factor) and "replicate-line" (1-8, random factor, nested

in regimen-type). We used Holm-Šidàk correction (Abdi 2010) for controlling the family-wise

error rates. For all ANOVAs where there was a significant effect of "Regimen-type" after the

Holm-Šidàk correction, we used Tukey's HSD to compare pairwise differences between LL, SL,

and SS (See Supplementary Methods for details and rationale).

Simulations of microbial evolution

Any difference between the three regimens in our experiment can, in principle, be due to some

idiosyncratic properties of the experimental organism (*E. coli*) or potential differences between the

selection environments in flasks and plates. In order to account for that possibility and enhance the

generalizability of our results, we used an Individual Based Model (IBM) to simulate bacterial

growth under resource-limited conditions (Wahl et al. 2002). Except for differences in the amount

of resources, our model contained no other parameters specific to *E. coli* or related to differences

in culture conditions. Thus, in terms of differences between the *EoA* of the regimens, if the model

output matched the empirical observations then our results were likely to be applicable for other

asexual systems. Treating our experiment as a case-study, we used our model to investigate if our

results were generalizable.

Our simulations start with a nearly clonal distribution of fitness effects. In our model, an individual

bacterium was characterized by three principal parameters: efficiency, threshold, and body-mass.

The simulation (coded in the C programming language) began with a fixed amount of resources

available in the environment, utilized by the bacteria for growth. A typical individual was

represented by an array that specified three principal parameters: (1) Bodymass, (2) Efficiency,

and (3) Threshold. Efficiency and Threshold were the only two evolvable parameters. Bacteria

consumed resources in an iterative and density-dependent manner. The parameter $Bodymass_i$ of

the ith individual represented how big the particular individual was during a given iteration. Its

efficiency (K_eff_i) specified how much food it assimilated per iteration. If *population size/ K_eff_i*

< 1, *10(1 - (population size/ K_eff_i))* units were added to $Bodymass_i$. Otherwise, $Bodymass_i$

remained unchanged. $Bodymass_i$ increased with cumulative assimilation. When $Bodymass_i$

becomes greater than or equal to $thres_i$ (its threshold parameter), the individual *i* underwent binary

fission and divided into two equally sized daughter individuals. Each fission event had a fixed

probability of giving rise to mutations based on a mutation rate that remained constant for all

individuals in the population. K_eff_i and $thres_i$ mutate independently, and were the only two

parameters that could undergo mutation. The mutated value was drawn from a static normal

distribution with the frequency of deleterious mutations being much higher than that of beneficial

mutations, which is in line with experimental observations (See Table S3; Kassen and Bataillon,

2006; Eyre-Walker and Keightley, 2007). The distribution of mutational effects remained fixed

throughout the simulation (Kassen and Bataillon, 2006) due to which, *EoA* was expected to

eventually approach a plateau. When the population ran out of resources (once the amount of body-

mass accumulated per unit time by the population went below a pre-decided threshold so that the

sigmoidal curve reached a plateau), it was sampled according to the sampling ratio being studied.

The above process was repeated for 400 generations, where each generation represented two-fold

growth in population size (see Supplementary Methods for a more detailed description of the

model). We also checked if our model met several intuitive theoretical predictions that had not

been coded directly (see Supplementary data (SD1); also see Figs. S1, S2, and S3).

Density-dependent growth, clonal interference (CI), the presence of deleterious mutations, the

presence of variable fitness effects of mutations, etc. are some key features that are instrumental

in shaping the adaptive dynamics of periodically bottlenecked asexual populations (Patwa and

Wahl 2008; Sniegowski and Gerrish 2010). Unfortunately, the complex interactions of so many

features are difficult to capture in analytical models (Sniegowski and Gerrish 2010). Consequently,

previous theoretical studies have been forced to make simplifying assumptions like the absence of

deleterious mutations (Desai and Fisher 2007; Wahl and Gerrish 2001), constancy of beneficial

mutational effects (Desai and Fisher 2007), constancy of N_f (Campos and Wahl 2009, 2010; Wahl

and Gerrish 2001; Wahl and Zhu 2015), the presence of discrete generations (Campos and Wahl

2009, 2010; Desai and Fisher 2007), etc. (see Table S2 for details). Our IBM avoids these

simplifying assumptions, which might explain why some of the features captured by our model

have not been reported earlier. Moreover, our study is in the context of *EoA*, while most of the

earlier studies have investigated fixation probabilities.

Results

HM **failed to predict and explain the** *EoA* **trajectories of experimental populations**

RM ANOVA on all three regimens indicated a significant Regimen-type $\times$ TIME interaction for

both K ($F_{16, 168} = 5.72$; $p < 0.000001$) and R ($F_{16, 168} = 7.306$; $p < 0.000001$). However, in principle,

this interaction could be driven by the fact that the LL populations had a much larger increase in

K and R compared to the SL and SS populations. Since our primary interest was to check whether

the SL and SS populations differed in terms of these two fitness measures, we repeated the RM

ANOVA for only these two regimens and again found a significant Regimen-type $\times$ TIME

interaction for both K ($F_{8,\,112} = 2.070$; $p = 0.0446$) and R ($F_{8,\,112} = 3.594$; $p = 0.000948$). Since the

interaction term was significant, we chose not to interpret the main effects of Regimen-Type or

TIME.

Individual ANOVAs showed that the *EoA* of SS was greater than that of SL at 5/6 and 4/5 time-

points which had significant difference in terms of K (Fig. 1a) and R (Fig. 1b). The *p*-values and

the *F*-values (with corresponding *df*) for each time-point for K and R are presented in Tables S4

and S5 respectively. Thus, particularly during the last two thirds of the evolution experiment, the

EoA of SS was consistently higher than that of SL. The effect sizes (Cohen's *d* (Cohen 1988)) of

EoA differences between SL and SS were found to be either medium or large (with the majority

being large effects; see Table S6) for several points on the *EoA* trajectory. Thus, similar *HM* can

give rise to fairly different adaptive trajectories. This observation is consistent with recent

empirical findings that question the validity of harmonic mean as an evolutionarily relevant

population size (Y. Raynes et al. 2014). Surprisingly, SS had a larger overall *EoA* than SL despite

having lower N_f. Interestingly, despite having similar N_f, LL typically had much larger *EoA* than

SL, which is explainable by the fact that the latter regimen suffered more severe bottlenecks. This

shows that similar N_f does not lead to similar *EoA* if the bottleneck ratios are different.

In summary, *HM* failed to predict the *EoA* trajectories of our experimental populations as, in spite

of having similar values of N_0g, the SL and SS regimens had markedly different adaptive *(EoA)*

trajectories for K (Fig. 1a) as well as R (Fig. 1b).

Fig. 1. Experimental *EoA* trajectories in terms of carrying capacity and maximum growth
rate. (**a**) *EoA* of carrying capacity *(K)*. (**b**) *EoA* of maximum growth rate *(R)*. Data points show
mean ± SEM for 8 replicates. * refers to cases when all three pairwise differences (LL-SL, LL-SS,
and SL-SS) are significant (Tukey post hoc $p < 0.05$). # refers to significant difference across LL-
SL and LL-SS, but not SL-SS (See Tables S4 and S5). SS and SL have markedly different adaptive
trajectories despite having similar harmonic mean population sizes.

Simulations also revealed that *HM* fails to predict adaptive trajectories.

We simulated evolution in populations with similar *HM* but with different values of N_0 and g, such

that the product (N_0g) remained constant. If *HM* ($=N_0g$) were a good predictor of how much a

population is expected to adapt, then these three treatments were expected to show similar *EoA*.

This was not found to be the case for both *K* (Fig. 2a) and *R* (Fig. 2b), which was consistent with

our experimental observations of *EoA* trends in SL and SS (Fig. 1; also see Supplementary

Methods and Fig. S4). XX', SS', and SL' were also found to be remarkably different in terms of

the adaptive increase in average efficiency of individuals (Fig. S4a). Interestingly, populations

with similar *HM* were also found to differ in terms of the frequency distributions of the efficiency

parameters amongst their constituent individuals (Fig. S5). To elucidate why N_0g could not explain

EoA trajectories, we determined how *EoA* varied with N_0 and g, independently.

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Fig. 2. Simulations: Adaption in three populations with similar harmonic mean size. Data
points show mean *EoA* $\pm$ SEM for 8 replicates. (**a**) Adaptation in terms of normalized carrying
capacity *(K)*. (**b**) Adaptation in terms of normalized maximum growth rate *(R)*. XX', SS' and SL'
had similar harmonic mean sizes and represent lenient, medium and harsh bottlenecks with $N_0 \approx$
3.6×10^3, 1.8×10^3, 9×10^2 and bottleneck ratio of 1/10; $1/10^2$, $1/10^4$ respectively. These simulations
suggest that populations with similar harmonic mean size can have markedly different *EoA*
trajectories.

***EoA* varied positively with N_0 but negatively with *g*.**

If N_0g were a good measure of the population size that has a positive relationship with *EoA*, then

increasing either N_0 or *g* or both should lead to greater *EoA*. We tested this intuitive prediction via

simulations using several combinations of N_0 and *g*, spanning four orders of magnitude for both

N_0 and the sampling ratio (N_0/N_f). Although *EoA* was found to increase with greater N_0 (Fig. 3a;

also see Fig. S6), the relationship between *EoA* and *g* turned out to be negative (Fig. 3b; also see

Fig. S7) which was reflected in terms of both individual-level (Fig. 3b—in terms of efficiency)

and population-level (Fig. S8—in terms of *R*) fitness parameters. The latter result implied that

larger values of N_f impeded adaptation in populations when the population size during the

bottleneck *(N_0)* was held constant. The nature (sign) of this relationship between *EoA* and *g* was

found to be robust to changes in mutation rate over a 100-fold range in our simulations (Fig. S10).

Fig. 3. Simulations: The relationship of *EoA* (expressed in terms of efficiency) with N_0 and *g*.
Data points show mean ± SEM; 8 replicates. The populations shown in **(a)** had the same bottleneck
ratio ($1/10^2$) but different bottleneck sizes (= N_0). *EoA* varies positively with N_0. On the other hand,
the populations shown in **(b)** had identical bottleneck size (= N_0) but different bottleneck ratios
(reflected by different values of *g*. Bottleneck ratios: BN1: 1/10 (*g* = 3.32); BN2: $1/10^2$ (*g* = 6.64);
BN3: $1/10^3$ (*g* = 9.96); BN4:$1/10^4$ (*g* = 13.28). *EoA* varies negatively with *g*. Also see Figs. S6, S7
and S8.

A negative relationship between *EoA* and *g* is particularly surprising because, in populations with

similar N_0, increase in *g* is expected to lead to an increase in the available variation. All else being

equal, this should have led to greater *EoA*. Since that was not the case, we went on to check if these

slowly adapting populations (with similar N_0 but higher *g*) were limited, qualitatively and/or

quantitatively, by the availability of variation.

The quantitative availability of beneficial traits could not explain why *EoA* varied negatively

with *g*.

Consider SM1 and SM4, treatment regimens which had similar starting population size (N_0) after

the first bottleneck but had *g* values of 3.32 and 13.28 respectively (SM refers to sampling ratio,

expressed in terms of log(10) (see Figs. 4 and 5). SM1 grew to a final size of $10N_0$ in one growth

phase (*i.e.*, before bottleneck), while SM4 grew to $10^4 N_0$. Consequently, SM1 faced a periodic

bottleneck of 1/10 whereas SM4 was sampled $1/10^4$. Since SM4 experienced approximately 278

times more fission events than SM1 per evolutionary generation, the former was expected to

undergo more mutations and consequently show more variation. Moreover, SM4 was also

expected to arrive at very large-effect benefits that were so rare that the probability of SM1

stumbling upon them was vanishingly low due to its lower mutational supply. As expected,

compared to SM1, SM4 had a greater within-population coefficient of variation in terms of

efficiency values (Fig. 4) and therefore was not limited by the supply of variation. To better

understand the contributions of phenotypes of different magnitudes to *EoA*, we classified the

phenotypes into 50 discrete static classes. We found that SM4 also had a continuous access to

highly fit genotypes (Fig. S11a) that were inaccessible to SM1 throughout the simulations. On the

basis of these observations, *EoA* can be expected to vary positively with *g* and thus SM4 was

expected to be fitter than SM1 at a given point of time in general. However, counterintuitively,

SM4 had a consistently lower *EoA* than SM1 (Fig. 4). Evidently, harsher periodic sampling

impeded adaptation despite resulting in increased substrate for selection. We also found that

although higher N_f allowed SM4 to arrive at extremely rare mutations with very large benefits,

these mutations failed to survive the harsh periodic bottlenecks by rising to large enough

frequencies (Fig. S12a). In other words, SM4 typically wasted the best mutation explored by it but

SM1 almost always conserved it. This explains why arriving at these rare mutations with very

large benefits did not make SM4 adapt more than SM1 in a sustained manner. However, this does

not explain why the *EoA* in SM4 was consistently lower than that of SM1.

Fig. 4. Simulations: Trajectories of efficiency in terms of across-population mean and within-
population coefficient of variation. The within-populations coefficient of variation (CV) was
computed for each replicate population across its constituent individuals using discrete frequency
distributions. The error bars represent SEM (8 replicates). Both SM1 and SM4 had similar
bottleneck size ($N_0 \approx 900$). SM1 experienced a periodic bottleneck of 1/10 whereas SM4
experienced a periodic bottleneck of $1/10^4$. SM4 had a consistently lower *EoA* than SM1 despite
having consistently more variation.

Fig. 5. **The efficacy of selection in SM1 was more than that in SM4. (a)** SM1 consistently experienced a much lower genetic load than SM1 (the error bars represent SEM (8 replicates)). **(b)** The lagging chunk was the major contributor to the Extent of Adaptation (*EoA*) in SM4 but not in SM1 (the error bars represent SEM (8 replicates)). This also means that the contribution of the nose to the *EoA* (which equals (1 – contribution of the lagging chunk)) in SM1 was much more than that of the lagging chunk. **(c) and (d)** Schematic representations of the distribution of efficiency across individuals during adaptation during the initial phases of evolution (before generation 80). Due to the high efficacy of selection in SM1, the majority of individuals were found in the nose **(c)**. On the other hand, a relatively low efficacy of selection due to harsher bottlenecks in SM4 resulted in most individuals being found in the lagging chunk (please refer to the text for more details) **(d)**. **(e) and (f)** During the later phases of evolution (around generation 360), the contributions of the nose to the overall *EoA* became relatively similar in SM1 and SM4.

The negative relationship between *EoA* and *g* can be explained in terms of the efficacy of selection

The efficacy of selection in eliminating deleterious mutations and spreading beneficial ones is an important factor that influences the increase of *EoA*. We quantified the inefficacy of selection in increasing *EoA* using the Genetic Load (*GL*), which was defined as: *GL = (Best Efficiency – Average Efficiency)/ Best Efficiency* (Crow 1958; Rice 2004). The term "Best Efficiency" refers to the highest efficiency value that succeeds in surviving the bottleneck. As discussed earlier, the magnitudes of the highest efficiency values explored by SM4 populations are much greater than those explored by SM1 (Fig. S11a). However, these high-fitness phenotypes of SM4 typically have such low frequencies that they almost always fail to survive bottlenecks and thus do not contribute significantly to the *EoA* (Fig. S12a). Therefore, we defined GL only in terms of the phenotypes that survived the bottlenecks. We found that the Best Efficiency (after bottlenecks) for SM1 was very similar to that of SM4 (Fig. S12b). We note here that the phenotypes that are fitter than the wild type but less fit than the best phenotype also contribute to the GL. Thus, consistently higher GL entails lower contribution of the best phenotype to the *EoA*. Furthermore, if these best phenotypes (with respect to which GL is defined) are similar across populations being compared, consistently lower contribution of the best phenotype to *EoA* would in turn entail slower rise of *EoA*.

SM4 consistently experienced a heavier genetic load than SM1, particularly during the initial phases of evolution (Fig. 5a). This genetic load was constituted largely by phenotypes that are fitter than the wild type ancestor but less fit than the best phenotype (Fig. S13). We labelled the top five occupied fitness classes as the "nose" (*sensu* Desai and Fisher (2007)) and all the classes inferior to the nose as the "lagging chunk." During the early phases of evolution, the relative

contribution of the lagging chunk to the *EoA* was much higher in SM4 than in SM1 (Fig. 5b). In

other words, the nose accounted for most of the *EoA* in SM1 but not in SM4 (shown schematically

in Figs. 5c and 5d). Thus, compared to SM1, the best phenotype of a typical SM4 population

needed to outcompete many more phenotypes (present in sizable frequencies) that were superior

to the wild-type but inferior to itself. This suggests that the efficacy of selection was higher in SM1

than in SM4, which in turn explains the faster increase in *EoA* in the former. As selection

proceeded, the GL of SM4 reduced greatly by generation 360 (Fig. 5a). This resulted in similar

contributions of the respective noses to the overall *EoA* in SM1 and SM4 (Figs. 5e and 5f).

The above observations suggest that during the early phases of evolution, populations with higher

g (here SM4) can face greater impediment (genetic load), which translates into a reduced *EoA*.

N_0/g is a better predictor of *EoA* than N_0g

Since our simulations suggested that the rate of adaptation is positively related to N_0 and negatively

related to *g*, we went on to test if N_0/g is a better predictor of adaptive trajectories than N_0g. This

turned out to be the case, not only in our simulations (Fig. 6; also see Figs. 2 and S14), but also for

our experiments. The N_0/g values of LL, SS and SL populations were approximately 3.01×10^9,

1.13×10^4, and 5.02×10^3, respectively, which led to a predicted *EoA* trend of LL>SS>SL, which

was observed for both *K* and *R* in the experiments (Fig. 1).

Fig. 6. Simulations: *EoA* trajectories in terms of *K*. Populations with similar N_0/g (LBbar and HB) match more closely in terms of mean adaptive trajectories than populations with similar $N_0 g$ (LB and HB). LB: $N_0 \approx 3600$, bottleneck ratio: 1/10; HB: $N_0 \approx 900$, bottleneck ratio: $1/10^4$; LBbar: $N_0 \approx 225$, bottleneck ratio: 1/10.

Discussion

Overview of the main results and what they suggest

Most experimental studies with periodically bottlenecked asexual populations have used the harmonic mean *(HM)* as the measure of population size (De Visser and Rozen 2005; Desai et al. 2007; Lachapelle et al. 2015; Lenski et al. 1991; Rozen et al. 2008; Samani and Bell 2010; Vogwill et al. 2016) to investigate quantities akin to the extent of adaptation *(EoA)*. Desai et al. (2007) stated that the enhancement in mean population fitness with respect to time (a quantity equivalent to *EoA*) depends upon the harmonic mean of the population size in such populations (Desai et al. 2007). However, there has been no empirical or theoretical test for the validity of *HM* as a predictor of *EoA*. Therefore, as a starting-point, we performed evolutionary experiments on *E. coli*

populations to test if the harmonic mean of population size (HM $(=N_0g)$) can predict EoA. Our

experiments revealed that N_0g does not predict EoA (Fig. 1). This observation could be interpreted

in two ways. Either there was something wrong with HM in terms of predicting EoA, or there were

some idiosyncratic properties of our experimental system (e.g. different kinds of containers) that

masks the relationship between HM and EoA. Apart from their different numbers (whose effect

we study here) and the fact that the LL/SL treatments are grown in flasks while the SS treatment

is grown in tissue culture plates, there are no differences between the three treatments, and hence

the corresponding selection pressures. Since they are grown with continuous shaking, aeration is

unlikely to be a significant issue. To account for the possibility that some idiosyncrasies of our

experiments were responsible for our results, and to test if the results of our experimental case-

study were generalizable, we simulated the adaptive evolution of asexual populations that grow

via fission. For this purpose, we used a very generic Individual-Based Model (IBM) that did not

contain any $E.$ $coli$ specific functions or parameters. The idea here was that if the outcomes of the

simulations matched the experiments, we could be reasonably confident that the experimental

results are not due to some peculiarities of the $E.$ $coli$ system or experimental protocols. The

simulations also revealed no association between N_0g and EoA (Fig. 2) which strengthened the

first interpretation that N_0g is not a good predictor of EoA.

It must be added here that conventionally, HM has been treated as an evolutionarily relevant

measure of population size only in terms of neutral mutations (Charlesworth 2009; Kimura 1983).

However, at least in terms of fixation probabilities of beneficial mutations, it has been shown that

population size measures similar to HM can act as the relevant measure of population size (Campos

and Wahl 2009, 2010; Heffernan and Wahl 2002).

To investigate why N_0g is an inappropriate measure for predicting *EoA*, we used our IBM to test

how *EoA* varied with N_0 and g independently, and found the counter-intuitive result that *EoA* varies

negatively with g (Figs. 3b and S7). To explain this result, we probed the composition of our

simulated populations as they evolved (Figs. 4 and 5). We found that g plays a dual role in terms

of determining *EoA*. Higher values of g positively affect *EoA* by increasing the supply of variation,

but negatively affect *EoA* by decreasing the efficacy of selection, as reflected by a consistently

greater genetic load. We found that this second effect of g on *EoA* overshadows the first, something

that is underappreciated in the empirical literature. Since N_0 and g have positive and negative

relationships respectively with *EoA*, intuition suggests that a good predictor of *EoA* should also do

the same. One such expression (of the many possible, taking into account that in most evolutionary

experiments $N_0 >> g$) is N_0/g. N_0/g indeed turns out to be a better predictor of *EoA* in our

simulations than N_0g (Figs. 6 and S14). We show below how both measures, (i.e. N_0g and N_0/g)

could lead to similar predictions about *EoA* under certain circumstances, and why is it important

to consider the cases when this correspondence breaks down.

The rest of the discussion elaborates the various insights mentioned above (and some more) and

their consequences.

Periodic bottlenecks lead to increased variation but reduced adaptation

The growth of many natural asexual populations is punctuated by episodic bottlenecks caused by,

for example, abrupt dissociation from hosts or spread of infections across hosts (reviewed in (Abel

et al. 2015)), etc. Moreover, periodic sampling during sub-culturing is a common feature of most

asexual populations propagated during experimental evolution studies (Kawecki et al. 2012;

Lenski et al. 1991). Therefore, it is important to appreciate the complex role played by periodic

bottlenecks in the evolutionary dynamics of asexual populations.

Most experimental evolution studies with asexual microbes are started with either genetically

uniform/clonal replicate populations or a mixed inoculum of relatively small number of genotypes.

In such populations, de *novo* beneficial mutations are the principal basis of adaptation (Barrick et

al. 2009; Kawecki et al. 2012). That is why populations that experience greater number of binary

fissions per generation are expected to generate more *de novo* beneficial variation and thus, to have

a higher extent of adaptation (*EoA*). Now, the number of binary fissions per generation is given by

$N_0(2^g-1)/g$. This quantity varies positively with the number of generations before a bottleneck (*g*)

and also with the size of the population at the bottleneck (N_0). Thus, all else being equal, *HM*

(≈$N_0 g$) is expected to be a good predictor of *EoA*.

However, the above line of reasoning disregards the fact that there can be a significant loss of

variation during periodic bottlenecks. As *g* increases, N_0 represents a smaller fraction of the final

population size (N_f) before bottleneck, which in turn increases the chances of loss of variation. For

example, assume that there are two bacterial populations that have the same value of N_0 (=10^2) but

g values of 3.32 and 13.28, leading to N_f values of 10^3 and 10^6 respectively (Lenski et al. 1991).

For a given value of N_0, increasing the value of *g* decreases the probability that a new beneficial

mutation would survive the bottlenecks (Wahl et al. 2002; Wahl and Zhu 2015). All else being

equal, this should reduce *EoA*.

Thus, increasing *g* has opposite effects on supply and survival of mutations in a population. Several

theoretical studies have investigated which of these two effects is more important for adaptive

evolution in asexual populations. For example, it has been suggested that increasing *g* increases

the probability of fixation of a beneficial mutation (Heffernan and Wahl 2002). This implies that

the positive relationship between g and mutational supply can overcome the negative effect of

increasing g on adaptation. Other theoretical studies have also shown a positive relationship

between adaptively relevant population size and the product N_0g (Campos and Wahl 2009, 2010).

Unfortunately, this rich body of theoretical predictions are not in the context of quantities (like

EoA) that are experimentally tractable, which was one of the motivations behind this study.

Our experiments (Fig. 1) and simulations (Fig. 2) showed that populations with similar values of

N_0g can have very different adaptive trajectories, suggesting that N_0g is not a good predictor of

EoA. Moreover, our simulations predicted the relationship between *EoA* and g to be negative (Figs.

3b and S7) and not positive. These two results disagree with a rather large body of existing

literature, as outlined above. One way by which this can happen is if our model incorporates some

atypical assumptions which lead to the observed counter-intuitive results. However, if that were to

be the case, then one would also expect our model to show other unintuitive results. Therefore, we

first investigated whether various other predictions of our model matched those from the literature.

Our model was able to replicate several intuitive theoretical predictions that had not been coded

directly. See Supplementary Data (SD2, also see the associated Figs. S1, S3, S6, and S10).

Therefore, it is reasonable to state that our model was generic and a good descriptor of evolving

bacterial populations.

***EoA* varies negatively with *g* because higher *g* makes selection less effective**

In order to explain why *EoA* varies negatively with g, we simulated populations with similar values

of N_0 (*i.e.*, bottleneck size) but different degrees of harshness of the bottlenecks, namely SM1

(lenient bottleneck (= 1/10), $g =3.32$) and SM4 (harsh bottleneck (= $1/10^4$), $g = 13.28$) (Figs. 5 and

6).

Our results demonstrate that higher g decreases the efficacy of selection in terms of spreading

beneficial mutations and purging deleterious ones (Fig. 5, also see Figs. S11 and S12). As shown

in Fig. 5, very high-efficiency classes rise to very high frequencies in SM1 populations by

generation 80. However, such classes fail to do so in SM4 populations. Owing to lenient

bottlenecks (lower g), selection operates so effectively in SM1 that its best efficiency class quickly

converges with the modal class (Fig. S11b). This is also reflected by the proximity of the mean

class with the modal class in SM1 (Fig. S11c). Thus, once a high-fitness class arises in an SM1

population, its rapid spread results in a steep increase in the population's *EoA*. However, despite

having the same bottleneck size (= N_0) as SM1, SM4 populations exhibit a much slower rise in

their *EoA*. This happens due to two reasons. As opposed to SM1, high-fitness genotypes in SM4

need to rise to much higher frequencies to survive the harsh periodic bottlenecks. This results in

the removal (due to sampling) of several high-fitness classes from SM4 during the bottleneck (Fig.

S12a). More importantly, the higher mutational supply rate of SM4 increases the genetic load (Fig.

5), which ultimately results in a much slower rise in the *EoA* of SM4.

Evolution of carrying capacity can feedback into adaptive trajectories

Both our experiments and simulations showed that carrying capacity (K) can evolve during

adaptation in asexual microbes (Figs. 1a and 2a respectively), which is consistent with previous

results (Novak et al. 2006). Unfortunately, most models of asexual adaptation do not take into

account such adaptive changes in the carrying capacity (Campos and Wahl 2009, 2010; Gerrish

and Lenski 1998; Wahl and Gerrish 2001). Most evolution experiments keep the bottleneck ratio

(represented by g) constant (Kawecki et al. 2012; Lenski et al. 1991). This constancy of g ensures

that any evolutionary change in carrying capacity would also change N_0. In other words, if K

increases, a constant value of g throughout evolution would ensure an increase in N_0. Since higher

values of N_0 accelerate adaptation (Fig. 3a), the regularity of bottlenecks introduces a positive

feedback during evolution if K increases adaptively. Stated differently, a larger value of N_0 would

make a population evolve higher K, which in turn would increase the next N_0, and so on. We think

that this aspect of fitness should not be omitted from theoretical models of how microbes evolve,

particularly under resource-limited conditions, which are a common feature of experimental

evolution protocols (Kawecki et al. 2012; Lenski et al. 1991).

N_0/g is a better predictor of EoA than N_0g

As shown in Figs. 3b, S7 and 4, when selection is at work, EoA decreases with increasing g. This

suggests that a population size measure which is an increasing function of N_0 but a decreasing

function of g can be a better predictor of EoA than the conventional formula (HM). For example,

as shown in Fig. 6, we found that N_0/g is a better predictor of EoA than HM ($= N_0g$). Admittedly,

it is not possible to reason from this that the expression N_0/g will always be a good predictor of

EoA, and we make no such claims. We simply submit this expression as a potential candidate for

this purpose and hope that future theoretical work will be able to validate this empirically derived

quantity.

Most theoretical studies assume that the final population size attained in their study systems (N_f)

is constant (Campos and Wahl 2009, 2010; Gerrish and Lenski 1998; Heffernan and Wahl 2002;

Wahl et al. 2002; Wahl and Gerrish 2001). Interestingly, if the experimental populations that are

being compared have similar values of N_f (Desai et al. 2007; Y. Raynes et al. 2014; Vogwill et al.

2016), then the populations with larger values of N_0g will typically also have larger values of any

quantity that is an increasing function of N_0 but a decreasing function of g. This is because of two

reasons. First, if N_f is held constant, since $N_f = N_0 2^g$, increasing N_0 necessarily decreases g. Second,

in most empirical studies, $N_0 >> g$. Consequently, if N_f is assumed to be the same the populations

being compared, any prediction based on the relative values of $N_0 g$ will typically be similar to

predictions based on N_0/g (Fig. S15). However, whenever N_f is not held constant (e.g., Figs. 2, 3,

4, 6 and S14, and studies like (Lachapelle et al. 2015; Ramsayer et al. 2013; Y. Raynes et al. 2014;

Rozen et al. 2002; Samani and Bell 2010), N_0/g predicts EoA much better than $N_0 g$. The above

observations can explain why $N_0 g$ has been widely used across several empirical studies despite

failing to capture the effects of g on EoA accurately.

Implications of our results

At very long time-scales, the high-fitness mutations accessible only to SM4 (but not to SM1) may

end up surviving a harsh periodic bottleneck. A post-facto analysis of our SM4 simulations shows

that mutations of this kind rise to a frequency between 10^{-7} and 10^{-8} in a typical growth phase just

prior to bottlenecks in SM4. Since N_0 is close to 10^3 in these populations, the above high-quality

mutations would survive one bottleneck in every 10^4 to 10^5 growth phases which roughly amounts

to 1.3×10^5 to 1.3×10^6 generations. However, to this date, there are no reported experimental

evolution studies over this long a time-span. Therefore, we conclude that the observation that

increasing g decreases EoA should be relevant for the time-scales most commonly employed in

experimental evolution studies.

Our results can be used to compare the extents of adaptation in independent evolution experiments

with similar environments but dissimilar demographic properties (differences in terms of N_0 and/or

g and/or N_f). Such studies, which compare populations evolving in similar environments but with

dissimilar demographic properties, are reasonably common in the field of experimental evolution

(Desai et al. 2007; Lachapelle et al. 2015; Y. Raynes et al. 2014, 2014; Rozen et al. 2002; Samani

and Bell 2010; Vogwill et al. 2016).

Our study shows that in serially bottle-necked asexual populations, the destructive aspect of

bottlenecks (reduction in efficacy of selection by harsher bottlenecks) can overshadow their

constructive aspect (increase in supply of variation in harsher bottlenecks). This calls for a change

in perspective about periodic bottlenecks and a substantial re-evaluation of the role of population

size as a predictor of adaptive evolution.

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Supplementary information

Supplementary Methods

Experimental Evolution

Selection regimens: We propagated three distinct regimens (LL, SL, and SS) of *Escherichia coli* MG 1655 populations for more than 380 generations in Nutrient Broth with a fixed concentration of an antibiotic cocktail containing a mixture of three antibiotics at sub-lethal concentrations. 8 independently evolving replicate lines each of LL (large *HM* and large N_f, culture volume: 100ml (flasks)), SL (small *HM* but large N_f, culture volume: 100ml (flasks)), and SS (small *HM* and small N_f, culture volume: 1.5 ml (well plates)) were derived from a single *Escherichia coli* K-12 MG1655 colony and propagated during the course of our evolution-experiment. It is possible that the differences in culture containers or absolute numbers at the stationary phase could have exposed these populations to somewhat different selection pressures. However, our simulations had no such differences coded into them and yet yielded identical results. Thus, it is highly unlikely that the empirically observed differences between the SS, SL and LL populations arose due to differences in culture conditions. All the cultures were propagated at 37^OC, and were shaken continuously at 150 rpm. The three population regimens experienced different numbers of evolutionary generations (g) between periodic bottlenecks (*i.e.*, before they were sub-cultured). SS and SL had similar *HM (i.e.*, N_0g) albeit obtained through different combinations of N_0 (SS>SL) and g (SL>SS) such that the N_f of SL was approximately 73 times larger than that of SS. The N_f of SL was similar to that of LL, while the harmonic mean size of LL was > 16,500 times larger than that of SL and SS (Table S1). The bottleneck ratio for random sampling at the end of each growth phase was 1/10 for LL, $1/10^4$ for SS, and $1/10^6$ for SL. 1 ml cryostocks belonging to each of the twenty four independently evolving populations were stored periodically. This scheme automatically gave rise to the following rank-distribution in terms of the quantity N_0/g: SL<SS<LL.

Culture environment for experimental evolution:

24 independent bacterial populations *Escherichia coli* K-12 MG1655 were grown in Nutrient Broth with a fixed concentration of an antibiotic cocktail containing a mixture of three antibiotics at sub-lethal concentrations:

       1. Norfloxacin (0.015 μg/ml)
       2. Rifampicin (6 μg/ml)
       3. Streptomycin (0.1 μg/ml)

Composition of Nutrient Broth (Himedia Laboratories Pvt. Ltd.):

- Peptic digest of animal tissue (5 g/l)
- Sodium chloride (5 g/l)
- Beef extract (1.50 g/l)
- Yeast extract (1.50 g/l)

The following table summarizes the numerical properties of the three population regimes used in our study:

Regime type	Starter population size (N_0)	Final population size (N_f)	Dilution during bottleneck	No. of generations per dilution (g)	Harmonic mean size (HM)	Culture volume
SS	1.5x	15000x	1: 10^4	≈13.28	**~20x**	**1.5ml**
SL	x	10^6x	1:10^6	≈19.93	**~20x**	**100ml**
LL	10^5x	10^6x	1:10	≈3.32	**~3.32*10^5x**	**100ml**

Table S1. A summary of the experimental populations. x ≈ 10^5 in our experiments.

Fitness assays: To reconstruct the evolutionary trajectories of our experimental bacterial populations, we measured bacterial growth using an automated multi-well plate reader (Synergy HT, BIOTEK ® Winooski, VT, USA). Bacterial growth was measured in the same medium (containing the antibiotic cocktail) that the populations experienced during evolution. We used OD at 600 nm as a proxy for bacterial growth. Bacteria from the cryostocks belonging to each of the 24 populations were grown in 96 well plates. Each cryostock-derived population was assayed in three measurement-replicate wells in a 96 well plate. Each well contained 180 μl growth medium (Nutrient Broth with the antibiotic cocktail) containing 1:10^4 diluted cryostock. The plate was incubated at 37°C, and shaken continuously by the plate-reader throughout the growth assay. OD readings taken every 20 minutes during this incubation resulted in sigmoidal growth curves. Fitness measurements were done using cryostocks belonging to multiple time-points in order to reconstruct evolutionary trajectories. While reconstructing fitness trajectories, it was made sure that every 96 well-plate contained populations belonging to similar time-points (in terms of number of generations) during the course of evolution. We used the carrying capacity (*K*) and maximum population-wide growth rate (*R*) as the measure of fitness (Novak et al. 2006)). *K* of a population was defined as the maximum OD value attained over a period of twenty four hours (the highest value in the sigmoidal growth curve) (Karve et al. 2016; Novak et al. 2006). *R* was estimated as the maximum slope of the growth curve over a running window of four OD readings (each window spanning one hour) (Karve et al. 2015, 2016; Vogwill et al. 2016). Fitness measurements were done using cryostocks belonging to multiple time-points in order to reconstruct evolutionary trajectories.

Statistics:

Repeated measures ANOVA were performed independently for each of the two growth parameters (*K* and *R*). "Regimen-type" (LL/SL/SS) was treated as the categorical factor, and TIME (9 time-points) as the repeated measures factor. We also include the interaction of the two factors (Regimen-type and TIME) in the ANOVA model. We did not perform post-hoc tests in the RM ANOVA analysis as it involved

very large number of comparisons that were irrelevant for the present question and thus reduced statistical power. Instead, we analyzed the two growth parameters (K and R) independently at each time point in the *EoA* trajectory (Fig 1A and 1B) using nested-design ANOVAs with population regimen-type (SS, SL or LL) as the fixed factor and replicate-line (1-8, nested in population-type) as the random factor. For each of these ANOVAs, we further corrected the *p*-value of the main-effect of "Regimen-type" using the Holm-Šidàk correction (Abdi 2010) to control the family-wise error rate. The means from all those ANOVAs which showed a significant (< 0.05) *p*-value after the Holm-Šidàk correction, were further subjected to Tukey's HSD to identify which pair-wise differences were significant.

Algorithm for the individual based model used in this study:

Our model simulates the growth of individual bacteria in density-dependent resource-limited conditions. Each bacterium is represented by an array which has the following components:

1. Determinant of efficiency (K_eff): determines how much food can be assimilated per unit time
2. Determinant for threshold (thres): how much food needs to be assimilated in order to divide
3. Bodymass: how big is the bacterium (where is it along its cell cycle) at any given time

At the beginning of the simulation, two global scaling quantities, Food_Proxy and Body_Proxy are declared for the whole population. As the names suggest, Food_Proxy acts as a proxy for the amount of available resources initially, while Body_Proxy (=250) is a proxy for bodymass of the ancestor. Each simulation run is started with 100 individuals and each individual is allotted K_eff$_i$ value given as Eff$_i$ * Food_Proxy. Here, Eff$_i$ is a random number picked from a uniform distribution U(0.95-1.05) and K_eff$_i$ determines how much food would be consumed in a density–dependent manner and when its food consumption would stop (as per the conditions given below). Similarly, the parameter for the threshold for each of the 100 starting individuals is assigned as a random number picked from a uniform distribution between 0.95*(Body_Proxy) and 1.05*(Body_Proxy).

Each bacterium has the same initial biomass (an arbitrarily small quantity, 10 units in this case).

Time is implicitly defined in our code and each iteration signifies one unit of time.

In each iteration, each bacterium "grows and divides" according to the following rules:

> *If for bacterium i, (population size/ K_eff$_i$) $\geq$ 1, it doesn't eat anything: its bodymass remains the same as the earlier iteration.*

> *If for bacterium i (population size/ K_eff$_i$) < 1, it eats $10*(1 - (population size/ K_eff_i))$ units of food in this iteration: its bodymass increases by $10*(1- (population size/ K_eff_i)$ units).*

> *If at the end of this iteration, bodymass$_i$ > thres$_i$, bacterium i divides into two equal parts. A small thermodynamic cost (constant for all individuals) is deducted so that the sum of the bodymass of the daughter cells is exactly 1 unit less than the bodymass of the mother cell at the time of division.*

If the bacterium divides, there is a 1 in 100 chance for each of the daughter cells that it
mutates. If a mutation occurs, the new parameter for efficiency is drawn from an already defined
normal distribution that is used throughout the simulation. The same applies to threshold.
(Threshold and efficiency mutate independently in each bacterium.)

The total size of the population is saved at the end of every iteration. The total amount of
food consumed during each iteration is also computed.

The above description (italics) represents all the processes that happen within an iteration.

The process is repeated (and the population grows) until the following conditions are fulfilled:

1. The number of iterations is greater than 2000.
2. The amount of food consumed during each iteration < 0.08* Food_Proxy

If the above conditions are met simultaneously, food consumption is stopped, a defined fraction
of individuals are sampled randomly and the whole process is started with this sample population
(this represents bottlenecking). The above process is continued for q bottlenecks. The bottleneck
ratio and the number q are predefined, depending upon the type of population being studied. This
gives rise to q sigmoidal growth curves. Two quantities are extracted from each sigmoidal curve:

1. Carrying capacity (K, the maximum size of the population in each growth phase)
2. Maximum linear growth rate (R, the maximum slope of population growth over 100
iterations). Straight lines were fit on overlapping moving windows of 100 iterations on the
entire time-series of population size values within each growth-phase. The maximum
value of the slope observed within the entire time-series of population size values
(sigmoidal curve) was taken to be the maximum growth rate (R).

Time series of carrying capacities and maximum growth rates are computed using the series of q
sigmoidal curves.

In each simulation used in this study, the carrying capacity of the first growth phase was $\approx$
1.8*Food_Proxy. The value of Food_Proxy was adjusted in such a way that it gave rise to the
desired value of the carrying capacity of the first growth phase in a simulation. The carrying
capacities corresponding to the subsequent growth phases was an emergent result of Darwinian
evolution in the simulations. A summary of the key differences between our model and previous
theoretical studies is given in Table S2.

Property	Wahl and Gerrish (2001)	Desai and Fisher (2007)	Campos and Wahl (2009)	Campos and Wahl (2010)	Wahl and Zhu (2015)	Our study
Direct analysis of bottlenecks	Yes	No	Yes	Yes	Yes	**Yes**
Direct predictions about *EoA*	No	Yes	No	No	No	**Yes**
Deleterious mutations present	No	No	Yes	Yes	Yes	**Yes (Majority)**
Clonal interference allowed	No	Yes	Yes	Yes	No	**Yes**
Size(s) of distinct beneficial mutations	Not applicable	Constant	Variable	Constant*	Not applicable	**Variable**
Variable N_f across populations being compared	No	Not applicable	No	No	No	**Yes**
Evolution in resource limited conditions	No	No	No	No	Yes	**Yes**
Overlapping generations	Yes	No	No	No	Yes	**Yes**

Table S2. A summary of the key differences between our model and previous theoretical studies.

*While assuming that distinct beneficial mutations can have different effects, Campos and Wahl
(2010) arrive at an expression of N_e that itself depends upon the size of such effects (s_b). They
then state that "Clearly an effective population size that depends on s_b is unsatisfactory." Then
they assume that "all beneficial mutations have the same effect" to arrive at a much simpler
expression.

The following simulation settings were used in our study:

Population type	Food_proxy	Bottleneck ratio	Number of bottlenecks (400 generations)
XX'	$2*10^4$	1/10	120
SS'	10^5	$1/10^2$	60
SL'	$5*10^6$	$1/10^4$	30
BN1	$5*10^3$	1/10	120
BN2	$5*10^4$	$1/10^2$	60
BN3	$5*10^5$	$1/10^3$	40
BN4	$5*10^6$	$1/10^4$	30
SM1	$5*10^3$	1/10	120
SM4	$5*10^6$	$1/10^4$	30
HB	$5*10^6$	$1/10^4$	30
LB	$2*10^4$	1/10	120

Population type	Food_proxy	Bottleneck ratio	Number of bottlenecks (400 generations)
LBbar	$1.25*10^3$	1/10	120
MBbar	$2.5*10^4$	$1/10^2$	60
SNFBN1	$5*10^4$	1/10	120
SNFBN2	$5*10^4$	$1/10^2$	60
SNFBN4	$5*10^4$	$1/10^4$	30
Fig. S6a (small bottleneck size)	$5*10^4$	1/10	120
Fig. S6a (large bottleneck size)	$5*10^6$	1/10	120
Fig. S6b (small bottleneck size)	$5*10^4$	$1/10^3$	40
Fig. S6B (large bottleneck size)	$5*10^6$	$1/10^3$	40
Fig. S6c (small bottleneck size)	$5*10^5$	$1/10^4$	30
Fig. S6c (large bottleneck size)	$5*10^6$	$1/10^4$	30
Fig. S7a (lenient bottleneck)	$5*10^2$	1/10	120
Fig. S7a (harsh bottleneck)	$5*10^6$	$1/10^5$	24
Fig. S7b (lenient bottleneck)	$5*10^4$	1/10	120
Fig. S7b (harsh bottleneck)	$5*10^6$	$1/10^3$	40
Fig. S7c (lenient bottleneck)	$5*10^5$	1/10	120
Fig. S7c (harsh bottleneck)	$5*10^7$	$1/10^3$	40
Fig. S7d (lenient bottleneck)	$5*10^6$	1/10	120
Fig. S7c (harsh bottleneck)	$5*10^7$	$1/10^2$	60

We checked if our simulations met several other theoretical expectations from the extant literature.
As expected, despite following the same distribution for mutations, large populations showed
curves of diminishing returns while adapting, whereas very small populations showed stepwise
increase in fitness with long periods of stasis (Fig. S1). This happens because very small
populations (but not large ones) need to wait for beneficial mutations to arise (Sniegowski and
Gerrish 2010). Moreover, the extent of adaptation showed a positive but saturating relationship
with an unambiguous measure of absolute population size in our simulations (Fig. 3a and S7a).
The simulations also revealed a non-monotonous relationship between fitness and mutation rate
(Orr 2000) (Fig. S7).

Table S3. Distributions for parameters used in simulations:

Distribution used for	Distribution for efficiency parameter	Distribution for threshold parameter
Starting the simulation	Uniform random (0.95,1.05)	Uniform random (237.5,262.5)
Mutation	Normal random (0.9,0.22)	Normal random (1.1,0.22)

847 **Supplementary Data:**

848 **SD1. Statistical analysis of empirical results**

Generation	ANOVA $F_{(2,48)}$	ANOVA p	Holm-Šidàk corrected p	Tukey p LL-SL	Tukey p LL-SS	Tukey p SL-SS
40	36.75	1.4E-7	1E-6	0.0001	0.0001	0.0001
80	13.319	0.0002	0.0011	0.0001	0.0001	0.7942
120	14.365	0.0001	0.0008	0.0001	0.0001	0.0001
160	17.282	3.7E-5	0.0003	0.0001	0.0001	0.0001
200	2.894	0.0776	0.0776	-	-	-
240	9.359	0.0012	0.0050	0.0001	0.0001	0.0001
280	12.110	0.0003	0.0016	0.0001	0.0001	0.0001
320	6.769	0.0054	0.0161	0.0001	0.0001	0.0001
~ 390	3.33	0.0556	0.1081	-	-	-

Table S4. A summary of statistical analysis of carrying capacity (K) measurements in
empirical populations. The values in red represent statistically significant difference ($p<0.05$).
The p-values corresponding to nine independent ANOVAs (corresponding to nine different time
points) were subjected to Holm-Šidàk correction. Post-hoc (Tukey) comparisons were done only
in cases where the ANOVA p-values were less than 0.05 after Holm-Šidàk correction. These post–
hoc comparisons were done for three pairwise differences (LL-SL, LL-SS, and SL-SS) at each
time point. Holm-Šidàk correction was not done on Tukey p-values. The p-values are reported to
four decimal places.

Generation	ANOVA $F_{(2,48)}$	ANOVA p	Holm-Šidàk corrected p	Tukey p LL-SL	Tukey p LL-SS	Tukey p SL-SS
40	56.631	3.5E-9	3.1E-8	0.0001	0.0001	0.0001
80	5.996	0.0087	0.0344	0.0001	0.0001	0.3229
120	12.027	0.0003	0.0023	0.0001	0.0001	0.0001
160	12.287	0.0003	0.0023	0.0001	0.0001	0.0001
200	3.093	0.0665	0.1286	-	-	-
240	8.757	0.0017	0.0103	0.0001	0.0001	0.0001
280	7.785	0.0030	0.0147	0.0001	0.0001	0.0001
320	4.096	0.0315	0.0915	-	-	-
~ 390	0.925	0.4122	0.4122	-	-	-

Table S5. A summary of statistical analysis of maximum growth rate (R) measurements in
empirical populations. The values in red represent statistically significant difference ($p<0.05$).
The values in red represent statistically significant difference ($p<0.05$). The p-values
corresponding to nine independent ANOVAs (corresponding to nine different time points) were
subjected to Holm-Šidàk correction. Post-hoc (Tukey) comparisons were done only in cases where
the ANOVA p-value was less than 0.05 after Holm-Šidàk correction. These post–hoc comparisons
were done for three pairwise differences (LL-SL, LL-SS, and SL-SS) at each time point. Holm-
Šidàk correction was not done on Tukey p-values. The p-values are reported to four decimal places.

Effect sizes for the differences in *EoA* of SL and SS in terms of *K* and *R*

Fitness measure	Generation	Cohen's *d*	Inference about Effect Size
K	40	1.112	Large effect
	120	0.527	Medium effect
	160	0.786	Medium effect
	240	0.606	Medium effect
	280	0.514	Medium effect
	320	1.000	Large effect
R	40	1.172	Large effect
	120	0.770	Medium effect
	160	1.024	Large effect
	240	1.037	Large effect
	280	0.836	Large effect

Table S6. Analysis of the differences in *EoA* of populations with similar *HM* in terms of effect sizes. Cohen's d (Cohen 1988) was used to determine the effect sizes of the differences in the *EoA* of SS and SL. $0.2 < d < 0.5$ was interpreted as small effect, $0.5 < d < 0.8$ as medium effect, and $d > 0.8$ as large effect. The majority of differences between SL and SS were found to be of large effect size. The analysis was performed only at time points when the post-hoc (Tukey) p – values corresponding to SL-SS were < 0.05.

SD2. Results from individual-based simulations that match intuitive theoretical predictions that had not been coded directly

Firstly, as expected (Elena et al. 1996; Sniegowski and Gerrish 2010), very small populations showed discontinuous staircase-like (stepwise) trajectories of fitness increase whereas large populations showed smooth adaptive trajectories (Fig. S1). Secondly, *EoA* trajectories showed diminishing returns with time despite never hitting the explicitly coded wall of adaptive limit (Figs. 2, 3, 4, 5, S2a, S3, S4, S6, S7, and S8) (Lenski et al. 1991; Tenaillon et al. 2016). Thirdly, as expected, we found a non-monotonous relationship between *EoA* and mutation rate (Fig. S8c) (Orr 2000). Fourthly, *EoA* showed a positive but saturating relationship with N_0 (which is an unambiguous measure of absolute population size) (Fig. 3a) (Gerrish and Lenski 1998; Sniegowski and Gerrish 2010). All this was highly unlikely if our model incorporated unrealistic or atypical assumptions. Furthermore, for numerically similar populations (i.e. populations with similar N_0 and *g*) and identical time-frames, the results from our simulations were a very close match to the results of our experiments in terms of both *K* and *R* trajectories (Fig. S3). This again suggests that the IBM captures at least some features of the *EoA* of our *E. coli* populations. Finally, the *EoA* rank predictions generated for the three experimental populations based on our model agreed well with the empirical data (LL>SS>SL, Fig. 1).

Qualitative differences in *EoA* trajectory-shapes brought about by large differences in population size

Fig. S1. Qualitative differences in adaptive trajectories corresponding to populations with a large different in their sizes. Stepwise increase in fitness (with long periods of stasis) occurred in typically small populations such as the one shown in (**a**) as compared to smooth curves of diminishing returns in typically large populations such as the one shown in (**b**) (See the ordinates for absolute ranges of N_f during adaptation). The population shown in (**a**) experienced a periodic bottleneck of 1/10 while he population shown in (**b**) was bottlenecked $1/10^4$ periodically.

As expected (Sniegowski and Gerrish 2010), very small populations showed staircase-like (stepwise) trajectories of fitness increase.

Changes in standard deviation with sample size

Fig. S2. Increasing the number of replicate simulations from 8 to 20 did not result in increase in variation across replicates. (**a**) Increasing the number of independent replicate simulations from 8 to 20 didn't result in qualitative changes (ranks) of three populations with similar harmonic mean size but different N_0/g (mean ± SEM; N=20) (compare with Fig. 2a in the Main-text). (**b**) This increase in replicate number also didn't result in major changes in the standard deviation in carrying capacity during the course of adaptation. XX': $N_0 \approx 3.6*10^3$, bottleneck ratio: 1/10; SS': $N_0 \approx 1.8*10^3$, bottleneck ratio: $1/10^2$; SL': $N_0 \approx 9*10^2$, bottleneck ratio: $1/10^4$.

Since our simulations are agent-based (and consequently take a very long time to run), we decided to operate on a sample size of 8 replicates per population type throughout our study.

Agreement between experiments and simulations

Fig. S3. Agreement between experiments and simulations in terms of adaptive dynamics over identical time-scales in numerically similar populations. (**a**) Carrying capacity *(K)* versus bottleneck number *(BN)* (**b**) Maximum growth rate *(R)* versus bottleneck number *(BN)*. Data points

represent mean ± SD over 8 replicates. Each data point corresponds to the respective measure of
fitness *(K or R)* derived from the sample taken after *BN* bottlenecks. Range of population size: N_0
$\approx 10^{4.5}$; $N_f \approx 10^{8.5}$; bottleneck ratio = $1/10^4$. Each bottleneck corresponds to approximately 13.28
generations. Both carrying capacity and maximum growth rate are normalized with the ancestral
values.

The results of our experiments and simulations agree well in terms of the range and dynamics of
adaptation over identical time-scales in numerically similar populations. This applies to both
measures of population-level fitness: carrying capacity *(K)* and maximum growth rate *(R)*.

Adaptation in terms of efficiency and threshold:

**Fig. S4. Adaptation in three populations with similar *HM* in terms of measures of fitness at
the level of individuals** (**a**) Adaptive increase in averge individual efficiency within populations
with similar harmonic mean size (mean ± SEM; 8 replicates). (**b**) Adaptive decrease in average
individual threshold in populations with similar harmonic mean size (mean ± SEM; 8 replicates).
Threshold evolved so quickly that its adaptive decrease did not refelct the differnce observed in
EoA trjectories for *K* and *R* (Fig. 2 (Main text)). XX': $N_0 \approx 3.6*10^3$, bottleneck ratio: 1/10; SS': N_0
$\approx 1.8*10^3$, bottleneck ratio: $1/10^2$; SL': $N_0 \approx 9*10^2$, bottleneck ratio: $1/10^4$.

Multiple measures of fitness in our study revealed that harmonic mean is not a good predictor of
adaptive trajectories because populations with similar harmonic mean size can have markedly
different adaptive trajectories (Fig. S4 and Fig. 2 (Main-text)). Identical trends were observed
when such populations (XX', SS', and SL') were compared in terms of two different measures of
population level fitness (Fig. 2 (Main-text)). In terms of fitness at the level of individuals,
efficiency showed the same trend as *R* and *K* (Fig. S4a). However,the adaptive trajectories
corresponding to XX', SS', and SL' were almost identical when expressed in terms of threshold.
Threshold evolved (decreased) so quickly and to such a large extent in almost all population types
that we simulated in this study (regardless of their *HM*) that most populations had similar
trajectories of threshold decrease (also see Fig. S7). Consequently, despite threshold being an
important determinant of fitness, adaptive differences amongst populations were best expressed

and explained in terms of trajectories of increase in efficiency and not in terms of decrease in
threshold. The trends shown by adaptive trajectories of efficiency increase were identical to those
shown by adaptive trajectories of *K* and *R*. Due to the above reasons, we focussed on population-
wide trait distributions only in terms of efficiency.

Populations with similar HM show remarkable differences in distributions of efficiency:

Fig. S5. The distributions of efficiency across constituent individuals during adaptation in populations with similar HM. The individuals of each simulated population (8 replicate populations each of XX' and SL') were classified into to a discrete frequency distribution of their efficiency values (50 static classes) just prior to the bottleneck. Higher class indices correspond to higher efficiencies. The best phenotype (in terms of fitness) explored by SL' was consistently fitter than the best phenotype explored by LB (**a**). The modal phenotype quickly converged with the

best available phenotype in most XX' populations but failed to do so in all SL' populations (**b**).
The mean phenotype in XX' approached the best phenotype very closely (**b** and **c**). However, there
was a consistently larger gap between the best phenotype and the modal phenotype in SL' (**b**) and
an even larger one between its best and mean phenotype (**b** and **c**). Our simulations revealed that
populations with similar harmonic mean size can differ appreciably from each other not only in
terms of their adaptive trajectories but also in terms of how the distribution of fitness amongst their
constituent individuals changes during adaptation.

EoA **shows an increasing relationship with** N_0**:**

Fig. S6. Simulations: *EoA* **trajectories of populations with similar bottleneck ratio but**
different bottleneck sizes (N_0**).** Data points show mean ± SEM; 8 replicates. **(a)** Bottleneck ratio
= 1/10; Small bottleneck size (N_0) ≈ 9000; Large bottleneck size (N_0) ≈ 900000. **(b)** Bottleneck

ratio = $1/10^3$; Small bottleneck size (N_0) ≈ 90; Large bottleneck size (N_0) ≈ 9000. (c) Bottleneck
ratio = $1/10^4$; Small bottleneck size (N_0) ≈ 90; Large bottleneck size (N_0) ≈ 900.

EoA **shows a decreasing relationship with *g*:**

Fig. S7. Simulations: *EoA* **trajectories of populations with similar bottleneck size (N_0) but**
different bottleneck ratios. Data points show mean ± SEM; 8 replicates. **(a)** N_0 ≈ 90; Lenient
bottleneck = 1/10 (*g* = 3.32); Harsh bottleneck = $1/10^5$ (*g* = 16.61) **(b)** N_0 ≈ 9000; Lenient
bottleneck = 1/10 (*g* = 3.32); Harsh bottleneck = $1/10^3$ (*g* = 9.96) **(c)** N_0 ≈ 90000; Lenient bottleneck
= 1/10 (*g* = 3.32); Harsh bottleneck = $1/10^3$ (*g* = 9.96) **(d)** N_0 ≈ 900000; Lenient bottleneck = 1/10
(*g* = 3.32); Harsh bottleneck = $1/10^2$ (*g* = 6.64).

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987 **Fig. S8. Simulations: *EoA* trajectories (in terms of *R*) of populations with similar bottleneck**
988 **size (N_0) but different bottleneck ratios.** Data points show mean ± SEM; 8 replicates. All the
989 population regimens shown here had $N_0 \approx 900$. Bottleneck ratios: BN1: 1/10; BN2: $1/10^2$;
990 BN4:$1/10^4$. Larger values of *g* lead to reduced *EoA* for a given number of generations.

993 **Threshold evolved (decreased) quickly and convergently in almost all populations**

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Fig. S9. Rapid and convergent reduction in threshold. Threshold evolved (decreased) so
quickly and convergently in most populations that we simulated that the effects The data points
show mean ± SEM (N=8). The populations shown in **(a)** were bottlenecked $1/10^2$ periodically.
The populations shown in **(b)** had $N_0 \approx 900$. Bottleneck ratios: BN1: 1/10; BN2: $1/10^2$; BN3:
$1/10^3$; BN4:$1/10^4$.

The relationship between *EoA* and *g* was robust to changes in mutation rates:

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Fig. S10. The relationship between *EoA* and *g* at three different mutation rates. (**a**) Adaptive
increase in normalized carrying capacity in BN1 populations at three mutation rates. (**b**) Adaptive
increase in normalized carrying capacity in BN4 populations at three mutation rates. (**c**)
Normalized carrying capacity in BN1 and BN4 at generation 200 at three μ values. *EoA* exhibited
a non-monotonous relationship with μ in both BN1 and BN4 populations, which is in line with
theoretical expectations (Orr 2000). The negative dependence of *EoA* on *g* was robust to changes
in mutation rate *(μ)* over a 100-fold range. We found that the relationship between *EoA* and μ can
be influenced by bottleneck ratio. This is in agreement with recent empirical findings (Raynes et
al. 2014; Raynes and Sniegowski 2014). The data points show mean ± SEM (8 replicates). Both
BN1 and BN4 had similar bottleneck size ($N_0 \approx 900$). BN1 experienced a periodic bottleneck of
1/10 whereas BN4 experienced a periodic bottleneck of $1/10^4$.

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Fig. S11. Simulations: Distributions of phenotypic effects across individuals during adaptation. The individuals of each simulated population (8 replicates each of SM1 and SM4) were classified into to a discrete frequency distribution of their efficiency values (50 static classes) prior to bottlenecks. Higher class indices correspond to higher efficiencies. (**a**) The best phenotype (in terms of efficiency) explored by SM4 was consistently fitter than the best phenotype explored by SM1. The modal phenotype quickly converged to the best available phenotype in all but one SM1 populations but failed to do so in all SM4 populations (**b**). The mean phenotype in SM1 approached the best phenotype very closely (**b and c**). However, there was a consistently larger gap between the best phenotype and the modal phenotype in SM4 (**c**) and an even larger one between its best and mean phenotype (**b and c**).

Locations of the best class before and after bottleneck:

Fig. S12. Locations of the best class before and after bottleneck. (a) Differences in the locations of the best class in the distribution of the efficeincy parameter before and after bottlenck in populations with similar N_0. The individuals of each simulated population (8 replicate populations each of SM1 and SM4) were classified into to a discrete frequency distribution of their efficiency values (50 static classes). While the best class of SM1 could survive the bottleneck in most cases (black circles), the best class of SM4 invariably failed to survive its harsh bottleneck (grey triangles). (b) The locations of the best class after bottlenecks were remarkably similar across SM1 and SM4. The data points show mean ± SEM (8 replicates; some error bars are smaller than the data symbols.

Thus, SM4 populations will typically waste their best mutations (as the latter will be lost dring
bottlenecks). However, SM1 populations will typically retain their best mutations across
bottlenecks.

Fig. S13. The fraction of beneficial variation with respect to the ancestor (in terms of
efficiency) in SM1 and SM4. The data points show mean ± SEM (8 replicates; some error bars
are smaller than the data symbols. Beneficial variation is refers to the variation by the individuals
that are fitter than the wild-type (ancestor).

N_0/g **can predict** *EoA* **better than** N_0g**:**

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Fig. S14. N_0/g **was a better predictor of** *EoA* **than** N_0g**.** Adaptive trajectories in populations with
similar N_0/g expressed in terms of individual-level fitness parameters: efficiency (**a**), threshold (**b**).

The data points show mean ± SD (8 replicates) in (a) and (b). Populations with similar N_0/g (LBbar
and HB) match more closely in terms of mean adaptive trajectories than populations with similar
N_0g (LB and HB). LBbar: $N_0 \approx 225$; bottleneck ratio = 1/10; MBbar: $N_0 \approx 450$; bottleneck ratio =
$1/10^2$; $N_0 \approx 900$; bottleneck ratio = $1/10^4$. HB: $N_0 \approx 900$, bottleneck ratio: $1/10^4$;

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1075   Populations with similar N_0/g had remarkably similar adaptive trajectories in terms of both
1076   efficiency and threshold. These populations had similar adaptive trajectories despite differing in
1077   terms of the intensity of the periodic botleneck over a 1000-fold range. While N_0*g failed to predict
1078   adaptive trajectories over this bottleneck range (Fig. 2 (Main text)), N_0/g could act as a much better
1079   predictor of adaptive trajectories.

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Fig. S15. *EoA* in populations with similar initial N_f, but different N_0g. When initial N_f is
constant, populations with higher N_0g also have higher N_0/g, and thus show higher *EoA*. However,
when N_f is not constant (e.g., see Figs. 2, 3b, 4, and 5 of the main text), higher N_0g does not lead
to higher *EoA*. SNFBN1: $N_0 \approx 9000$; bottleneck ratio = 1/10; $N_0g \approx 29900$; $N_0/g \approx 2711$. SNFBN2:
$N_0 \approx 900$; bottleneck ratio = $1/10^2$; $N_0g \approx 5975$; $N_0/g \approx 136$. SNFBN4: $N_0 \approx 9$; bottleneck ratio =
$1/10^4$; $N_0g \approx 120$; $N_0/g \approx 0.68$.

The above graph shows that both N_0/g and N_0g can successfully predict *EoA* if the populations
under consideration have similar initial N_f.

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