Contribution of *de novo* non-coding mutations to autism and identification of risk genes from whole-genome sequencing of affected families

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Abstract

Analyzing de novo mutations (DNMs) in protein-coding genes from whole-exome sequencing (WES) data has emerged as a powerful tool for mapping risk genes of autism spectrum disorder (ASD). The impact of non-coding mutations in ASD, however, has been largely unknown. This represents a large gap in our understanding of the genetics of ASD, as the majority of GWAS hits for a range of disorders fall into non-coding regions. To address this question, we performed a meta-analysis of DNMs using whole-genome sequencing (WGS) data from more than 300 individuals with ASD. We found that DNMs are enriched within brain transcriptional regulatory elements near genes involved in neuropsychiatric disorders. In these genes and in evolutionarily constrained genes, we also found an excess of DNMs that are predicted to affect pre-mRNA splicing. Collectively, we estimate that non-coding mutations explain at least one third of the ASD genetic risk attributable to DNMs. By combining information of non-coding DNMs with published WES data, we identified three new ASD risk genes at a false discovery rate (FDR) < 0.1, and 11 at a FDR < 0.3. A number of these genes are known to regulate critical processes in neural development and have been associated with other neuropsychiatric disorders. Taken together, our results demonstrate the pathogenic contribution of non-coding DNMs in ASD etiology and highlight some promising ASD risk genes. The analytic tools we provided in this study, for estimating contribution of non-coding mutations to disease risk and for mapping susceptibility genes using both coding and regulatory mutations, are applicable to any WGS studies on DNMs.

Introduction

Autism spectrum disorder (ASD) is a neurodevelopmental disorder with onset in early childhood[1,2]. Understanding the genetics of ASD would help understand the etiology of the disorder and may lead to the development of more effective diagnostics and therapies[3]. Nonetheless, ascribing ASD genetic risk to specific genomic loci has proved challenging. GWAS have only yielded a handful of small-effect variants that have not replicated between studies[4,5]. In contrast, screening for *de novo* mutations (DNMs) with potentially larger effect sizes has emerged as an effective approach for identifying ASD risk loci. Several large whole-exome sequencing (WES) studies have been performed on ASD parent-child trios or quartets[6–9], based on the principle that genes harboring an excess of deleterious DNMs are likely risk genes. These studies have identified approximately 70 ASD risk genes.

Compared with coding mutations, the contribution of noncoding mutations to ASD etiology is largely unknown. For a wide range of diseases and traits, the majority of GWAS hits fall into noncoding regions[10]. Non-coding variants may alter the activity of cis-regulatory elements or affect the splicing of pre-mRNAs, which may in turn alter gene dosage and isoform usage[11–15]. Beyond implicating specific disease-associated loci, knowledge of non-coding variants can also offer clues as to which cell types are relevant to disease etiology and how trans-acting genes may contribute to disease risk[12,16]. In the case of autism, disruption of chromatin remodeling, histone modification, and pre-mRNA splicing have been implicated in the disease[8,17–22]. Taken together, these results support the role of transcriptional mis-regulation in ASD etiology. Nevertheless, with few exceptions[23], most existing ASD studies focused on coding sequences, and little is known about cis-regulatory variants predisposing autism risk.

In this study, we explore the contribution of non-coding mutations to ASD by combining data from multiple whole-genome sequencing (WGS) studies in autism trios. We developed a computational framework to analyze whole-genome DNM data (Fig. 1). To assess the extent non-coding mutations contribute to autism risk, we compared the rates of DNMs within putative regulatory sequences in ASD subjects vs. controls. We employed multiple annotations to identify variants with possible regulatory effects, such as epigenomic modification patterns in brain and predicted

deleteriousness from computational methods[24–26]. Based on these burden analyses, we estimated

the effect sizes of different types of regulatory mutations, measured as relative risks. Combining

these estimates with the frequencies of mutations, we estimated the contribution of different types

of mutations to the variation of autism risk (liability) across individuals. Next, we aimed to identify

specific genes and regulatory elements underlying autism. We hypothesize that an ASD risk gene

may be disrupted in multiple ways: by coding mutations changing protein function, or by

regulatory mutations affecting gene expression level, or by splicing mutations. This motivates a

strategy of combining coding and non-coding mutations within/near a gene to assess its role in

autism. To optimally combine information from multiple mutational categories, we used a method

we developed previously, TADA (*Transmission and de novo association*)[27]. TADA assesses the

enrichment of DNMs in a gene relative to random expectation, weighing each mutation based on

its likely effect size (Fig. 1).

Using this computational framework, we found modest enrichment of *de novo* single nucleotide

variations (SNVs) in individuals with ASD in brain-active enhancers near genes with likely roles

in psychiatric disorders, as well as in positions affecting RNA splicing. Our conservative estimates

suggest that regulatory non-coding mutations contribute to about a third of *de novo* autism risk (i.e.

autism risk attributable to all DNMs). Although the sample size of WGS is only a tenth of that of

WES, we still identified three new ASD risk genes at a FDR < 0.1 and 11 at a FDR < 0.3. Multiple

lines of evidence support the possible roles of these genes in ASD.

Results

<u>WGS data of ASD and control families</u>

We combined DNM data from five WGS studies of ASD trios/quartets, including four published[28–

31] and one unpublished (Wu et al. See Methods for details), with a total of 300 affected subjects

(Supplementary Table 1). As controls, we used recently published whole-genome DNM data from

a large cohort of 700 unaffected trios[32]. Mutation data is limited to *de novo* SNVs. The validation

rate of *de novo* SNVs based on Sanger sequencing ranges from 85% to 94% in the six studies.

Largely in line with published mutation rates[29,33,34], the number of DNMs per subject in ASD

studies ranges from 57 to 63. The rate of DNMs in controls is lower (39), which could be due to

differences in sequencing depth, SNV calling procedures, or other factors. In subsequent analysis

of a particular mutation type (e.g. mutations within enhancers), we accounted for the differences

in the total mutation rates between cases and controls by comparing the relative mutation rates

with neutral mutations (e.g. synonymous mutations) serving as background (see Methods). Despite

the differences in absolute mutation rates, the mutational patterns are similar between cases and

controls (see Section 2.1-2.3 of Supplementary Methods, also Supplementary Figs. 1, 2, and 3).

We were able to replicate the findings of earlier WES studies using the coding mutations. We

found that nonsynonymous mutations from WGS data are about 1.2 fold enriched in individuals

with ASD versus controls (Fig. 2a), similar to previous numbers[6,35,36]. The burden (i.e. fold of

enrichment) is higher for nonsynonymous mutations within known ASD genes, genes likely

involved in neuropsychiatric disorders (dubbed "neuropsychiatric genes"), and genes intolerant to

mutations (Fig. 2a, See Methods for gene-set definitions). The burden of highly expressed brain

genes (from Brainspan: http://www.brainspan.org) is also higher than genes with lower expression

(Fig. 2a).

<u>Enrichment of functional *de novo* non-coding mutations in children with ASD</u>

To establish the importance of non-coding sequences in autism, we analyzed DNMs in putative

regulatory sequences of the brain. Enriched DNMs in these regions comparing with random

expectation would suggest that at least some of these DNMs are causative mutations. We used

fetal and adult brain H3K27ac regions as markers of regulatory activity[37,38]. We limited our

analysis to H3K27ac regions within 10 kb of transcription start sites (TSSs) of protein-coding

genes, including promoters and 5'UTR (promoters are defined as sequences within 1 kb upstream

of TSSs). We found 1.1 fold enrichment of *de novo* SNVs in promoters with H3K27ac marks

(active promoters), compared to de novo synonymous mutations though the results are not

statistically significant (Fig. 2b). In contrast, we found no enrichment of *de novo* SNVs in

promoter/5'UTR regions without H3K27ac marks (Fig. 2b). We also saw no enrichment of *de*

novo SNVs across all H3K27ac regions within 10 kb of TSSs (referred to as regulatory SNVs

hereafter) (Fig. 2b). Noncoding variants that disrupt or create transcription factor binding motifs

have been frequently associated with regulatory effects and complex phenotypes[39,40]. Indeed, we

found a significant ASD burden for motif-changing SNVs (see Methods for definition) in

H3K27ac enhancers ($P = 0.00019$, one-sided Fisher's exact tests were used in all burden analyses).

Notably, the commonly used annotations for non-coding variants, including GERP++[25], PhyloP[41]

and CADD[24] and Eigen[42], were not significantly associated with mutational burden (Fig. 2b and

Supplementary Table 2).

We hypothesized that regulatory sequences near ASD risk genes harbor excess noncoding

mutations, as suggested by an earlier study[23]. We found a 1.8-fold enrichment of regulatory SNVs

near ASD genes ($P = 0.069$) and a 1.2-fold enrichment near neuropsychiatric genes ($P = 0.032$)

(Fig. 2c). In contrast, we observed no enrichment of regulatory SNVs near genes that are unlikely

to be ASD risk genes (nonASD genes). Additionally, we found that SNVs falling into more distal

H3K27ac regions (10 kb - 50 kb to TSSs) are not enriched in known ASD or neuropsychiatric

genes (Fig. 2c).

Because of the concern of potential batch effects between ASD cases and controls, we performed

simulation-based alternative burden analyses, without using control data. The procedure is based

on reshuffling of DNMs in cases across enhancers according to estimated mutation rates. We found

that the burdens of motif-changing SNVs and of regulatory SNVs near neuropsychiatric genes are

similar to what we found in Figs. 2b and 2c ($P = 0.025$ and 0.0017, respectively, Supplementary

Table 3).

To show that the enrichment of regulatory SNVs is due to regulatory functions of H3K27ac

sequences, rather than to some generic properties of enhancers (such as higher GC content[43]), we

compared the burdens of SNVs in enhancers with strong activities in the developing cortex

(denoted as "brain-specific enhancers") vs. enhancers with strong activities in other tissues ("non-

brain specific enhancers"). The numbers of such enhancers are small, so we extended the distance

cutoff to 1 Mb around TSSs. Within brain-specific enhancers, the burdens of regulatory SNVs near

known ASD genes, neuropsychiatric genes, and intolerant genes are all higher than baseline

(though not significant because of the small mutational counts, Fig. 2d). In contrast, the
corresponding burdens are all below 1 for enhancers specifically active in other tissues (Fig. 2d).

Non-coding mutations may also become pathogenic by affecting pre-mRNA splicing. Using the
predictions from a recent paper[26], we found significant burdens of SNVs likely affecting RNA
splicing (splicing SNVs) in neuropsychiatric genes (Fig. 2e, $P = 0.011$) and in genes intolerant to
mutations (Fig. 2e, $P = 0.018$). In contrast, the SNVs that are not predicted to affect splicing show
no such patterns ("non-splicing" mutations in Fig. 2e).

In summary, we found multiple lines of evidence supporting the roles of regulatory and splicing
mutations in autism (see Supplementary Table 4 for detailed results of all burden analyses). The
specificity of our findings (e.g. we found burden in brain-specific enhancers but not enhancers
specific to other tissues) and the robustness of our results to different ways of performing burden
analysis argue that the results are due to true biological signals, rather than batch effects from
heterogeneous samples. We summarized the evidence supporting this claim in Section 2 of
Supplementary Methods.

Partition of ASD risk into coding and non-coding DNMs

We next assessed the relative contribution of non-coding versus coding DNMs to ASD liability.
We considered four classes of mutations: loss-of-function (nonsense and frameshift indels in
coding sequences), probably damaging missense (predicted by PolyPhen-2, denoted as Mis3),
regulatory SNVs, and splicing SNVs. The contribution of any type of mutation is measured as
Liability Variance Explained (LVE), taking into account both the mutational target sizes and the
average relative risks of mutations (see Methods). Note that our relative risks are defined on
mutations of causative genes. We used a strategy previously developed to translate results of
previous burden analysis into average relative risks of causal genes.[27] For coding mutations (LoF
and mis3), we used values estimated previously[8]. For regulatory and splicing SNVs, we used the
burden results from ASD genes with TADA q-value < 0.5[35] (see Methods and Section 3 of
Supplementary Methods for details and possible caveats).

Our relative risk estimates of regulatory SNVs and splicing SNVs are significantly lower than

those of coding SNVs, but regulatory SNVs are much more frequent (Table 1). Each class of

mutation explains only a small fraction of estimated total ASD genetic risk (Table 1), which is

broadly consistent with the conclusion of an earlier study[44]. Considering only the risk due to *de*

novo mutations, we found that non-coding SNVs (including regulatory SNVs and splicing SNVs)

explain 31% of the *de novo* risk (Fig. 2f). We argue that this estimate is very conservative . First,

we considered only enhancers within 10 kb of TSSs, which constitute about 36% of all enhancers

in our data. Second, our dataset contains only regulatory sequences active very early in

development (5, 7, and 12 post-conception week) or in the adult brain. Many enhancers active in

other related developmental periods are not covered in our data. Third, we only considered SNVs

in this study. *De novo* indels and CNVs, which are likely to be much more deleterious, were

ignored.

<u>Identification of ASD risk genes using both coding and noncoding mutations</u>

We performed integrative risk gene mapping using both coding and non-coding mutations. The

challenge posed by combining different types of mutations is that these mutations could have very

different damaging effects (e.g., nonsense mutations are generally more deleterious than other

mutations). Summarizing a gene's mutational burden in a manner agnostic to each mutation's

likely consequence would compromise the power to detect risk genes. We adopted TADA[27], to

address this issue. Briefly, TADA assesses the evidence of risk on each gene by comparing the

observed number of mutations in each mutational class (e.g. LoF or missense) with the expected

number. The evidence of each class is weighted by the average relative risk of that class, and the

evidence of all classes is combined into a single Bayes factor for the gene, reflecting the strength

of evidence (Fig. 1). In earlier studies, we have shown that TADA is much more powerful for

mapping risk genes than simple counting of mutation numbers in genes[8,35]. In this study, we

considered four classes of DNMs: LoF and mis3 coding mutations, regulatory SNVs within 10 kb

of TSS and splicing SNVs. For each class of mutations, we derived its average relative risk from

mutational burden analysis (Table 1), and estimated the mutation rate of each class for each gene

using a genome-wide mutation rate model[45]. Once we computed the results of all genes from

TADA, we corrected for multiple testing using Bayesian FDR[46].

TADA analysis with only coding mutations from a previous WES study (~3,500 samples)

identified 58 ASD risk genes at FDR < 0.1[35]. Adding regulatory SNVs and splicing SNVs from

our WGS data added three new ASD risk genes at FDR < 0.1, and 11 new genes at FDR < 0.3.

These predictions are robust to TADA parameters within a reasonable range (Supplementary Table

5). Each of the three genes at FDR < 0.1 (Table 2) has at least one LoF or Mis3 mutation, and the

evidence for these genes is strengthened by the presence of regulatory SNVs. We found extensive

evidence supporting the plausibility of these genes as ASD risk genes. *APBB1* is the target of

neuronal-RNA binding protein FMRP, whose loss of function causes Fragile X syndrome and

autistic features. All of the three genes have been identified as haploinsufficient genes[47,48] (Table

2, hypergeometric test, $P = 0.00079$) and are highly expressed in the brain (top 25% of all genes,

$P = 0.015$). The genes also tend to be evolutionarily constrained using either RIVS or another

metric based on frequency of LoF variants in ExAC (Table 2). Evolutionary constraint in the

human population has been shown to be a strong predictor of autism genes[45].

We performed multiple network analyses to further establish the link of new candidate genes to

autism. DAWN[49,50] is a recently developed method that predicts autism risk genes by virtue of the

genes' association with known ASD genes in co-expression networks of early developing brain.

We found two out of three new genes (*JUP* and *ARHGAP5*) have DAWN q-value < 0.05 in at

least one of the two critical developmental spatial-temporal windows for ASD[51] (Table 2,

enrichment $P = 0.00041$). Using GeneMania[52], we found additional evidence supporting the

connections between high-confidence ASD genes and our candidate genes. Between the two gene

sets, there are 75 co-expression links in GeneMania (Fig. 3a), significantly higher than chance

expectation *($P = 0.014$)*. To better understand the functions and biological processes that the

candidate genes are involved in, we obtained the top 30 neighboring genes of each candidate using

GeneMania (Supplementary Table 6). Gene Ontology (GO) enrichment analysis of the top 30

neighbors of the whole set of new ASD genes identified multiple biological processes related to

neuro-development, such as "negative regulation of neurogenesis" and "forebrain development"

(Table 3).

We next expanded the analysis to the 11 new genes at FDR < 0.3. We observed significant

enrichment of multiple gene annotations (Supplementary Table 7), including haploinsufficient

genes ($P = 0.014$), FMRP targets ($P = 0.010$), Gene Ontology "Brain" ($P = 0.024$), constrained

genes ($P = 0.0038$ using RVIS and $P = 0.025$ using variant frequency in ExAC), high expression

level in the early developing brain from BrainSpan ($P = 0.011$) and genes significantly co-

expressed with known ASD genes from DAWN analysis ($P = 0.00013$). The pathogenic roles of

these genes in ASD are further supported by previous studies (Supplementary Table 8).

We next explore the mechanisms by which non-coding mutations may contribute to ASD. We

found an interesting example in *PLXNB1* (a new risk gene with q = 0.22). This gene has important

functions in neurodevelopment (see Discussions). In the promoter region of *PLXNB1*, an SNV

disrupts the binding motif of the transcription factor RFX2 (Fig. 3b). Notably, in the motif-

matching region, multiple bases including the one hit by the DNM, are highly conserved based on

GERP scores (Fig. 3b). RFX2 is a transcription factor highly expressed in the tegmental region of

the the fetal mid-brain, a region linked to cognitive abilities[53] and social reward processing[54,55].

An intronic variant in *RFX2* has been associated with schizophrenia and bipolar disorder[56].

Additionally, the RFX2 binding motif is enriched in regulatory regions of several co-expression

gene modules (Modules 1, 5, 7, and 15)[37] that are active in the early developing brain. Taken

together, our analyses offer a testable hypothesis: the mutation in *PLXNB1* promoter changes the

affinity of RFX2, and as a result, perturbs the expression of *PLXNB1* and hence increases the risk

of ASD.

<u>Enhancers with recurrent *de novo* mutations</u>

Our previous analyses were performed at the gene level, and only considered enhancers within 10

kb of TSSs. We expanded these analyses to identify specific H3K27 enhancers with roles in autism

across the genome, based on the notion that multiple DNMs in a single enhancer are unlikely to

occur by chance. We found 25 enhancers with two or more SNVs in ASD cases, significantly

higher than random expectation (Fig. 3c, $P = 0.0025$). Since the expected number of recurrent

enhancers by chance is about 12, we estimated that half of these enhancers are likely to be causal

enhancers (Supplementary Table 9). Many recurrent enhancers are distant from genes, so the

nearest genes may not represent their true targets. We used predicted enhancer-gene pairs based

on cross-tissue correlation between enhancer activity and gene expression from Roadmap

Epigenomics (http://khuranalab.med.cornell.edu/roadmap_stringent_enhancers.txt). We found a

recurrent enhancer targeting *ZMIZ1*, a gene missed in previous analysis as it is more than 250 kb

away from the enhancer (Fig. 3d). The region contains two other DNMs in two enhancers, one of

which also has correlated activities with the ZIMZ1 promoter. A target of FMRP, *ZMIZ1* is highly

expressed in the brain and its protein product interacts with neuron-specific chromatin remodeling

complex (nBAF) which is important in regulating synapatic functions[57,58]. Several nBAF members

have been linked to autism, such as ARID1B and BCL11A[59]. The pathogenic potential of *ZMIZ1*

is further implicated by the observation of a *de novo* gene-disrupting translocation in an individual

with intellectual disability[60]. These evidence together strongly support the role of ZMIZ1 in autism,

and also highlight the mechanism that DNMs may increase ASD risk by disrupting distal

regulatory elements.

We next performed recurrent DNM analysis at the level of Topologically associated domains

(TADs)[61]. These are megabase–sized chromatin interaction domains that are stable across cell

types and have been proposed to demarcate transcriptional regulatory units[61]. We reason that if we

identify TADs with excess regulatory SNVs, it may be straightforward to assign target genes of

these regulatory SNVs. Based on estimated mutation rates, we found two TADs with a significant

(at FDR < 0.125) number of regulatory SNVs (Supplementary Fig. 5 and Supplementary Table

10). In both TADs, there are only two or three genes, and we conjecture that SRBD1 and MRSA

are likely the underlying ASD genes in the two TADs, respectively (see Discussion).

Power of mapping ASD risk genes with WGS and WES

We used simulations to address how the power of a DNM-focused WES or WGS study depends

on its sample size and sequencing budget, using our insights of the *de novo* genetic architecture of

ASD (Table 1). We randomly chose 1000 genes as risk genes based on previous estimates of the

total number of ASD risk genes[8,62]; randomly sampled mutations according to mutation rates and

their likely relative risks (causal genes tend to have more deleterious mutations compared to

expectations) and then applied TADA to identify risk genes at q < 0.1 (see Section 6 of

Supplementary Methods for simulation details). We found that the power of our simulated WES

studies is roughly in line with empirical discoveries (Fig. 4a). For example, at N = 2,500 trios, we

would expect to identify about 30-40 ASD genes, similar to the actual number (33) from a previous

study[8].The power of the simulated WGS design is about 20% higher than that of the WES (Fig.

4a). We next investigated whether the additional power gained from WGS is justifiable on the

basis of cost. At the current per-sample cost level (WES: $500 and WGS: $1000), we found that

WES is still more cost-effective (Fig. 4b).

Discussion

Our analysis provides clear evidence that non-coding mutations play an important role in the

etiology of autism. We found elevated rates of *de novo* SNVs in ASD subjects within brain

enhancers that are likely to affect transcription factor binding. Near potential risk genes of ASD

and other neuropsychiatric disorders, there is an excess of ASD-derived DNMs in enhancer regions.

We also found enrichment of splicing-affecting SNVs in ASD and neuropsychiatric genes as well

as in constrained genes. Compared with coding mutations, we found that these non-coding

mutations account for a smaller but substantial proportion of *de novo* ASD risk. Using a statistical

framework, TADA, we were able to leverage both coding and regulatory mutations for gene

mapping and identified several new ASD genes. Our study on non-coding mutations also led to

the characterization of ASD risk units with excess DNMs at the level of enhancers and TAD

regions. Altogether, our findings highlight the importance of regulatory variants in ASD and their

utility in mapping risk genes and functional elements. The analytic framework for *de novo*

mutations we developed in this work is applicable to any trio-based sequencing studies.

Our work is one of the first studies that directly assess the role of non-coding regulatory variants

in complex diseases[23]. Previously, our knowledge of non-coding variants came mostly from

GWAS. The challenge with GWAS is that regulatory elements are generally much shorter (~1kb)

than regions of linkage disequilibrium (LD, hundreds of kb on average). It is thus far from

straightforward to assess the quantitative contribution of non-coding variants, or identify specific

regulatory elements underlying diseases from GWAS[63]. In contrast, using DNMs from WGS, we

eliminate the impact of LD and work directly on putative causal mutations. The results of our

findings thus can be interpreted more straightforwardly. For instance, the enhancers with recurrent

DNMs we identified are good candidates for further experimental analysis.

Distinguishing likely pathogenic variants from benign ones is essential when analyzing non-coding

variants. We only observed a modest enrichment of DNMs in H3K27ac enhancers near putative

risk genes, but not across all genes. This suggests that most SNVs within H3K27ac sites are

probably not causative mutations. We tested various methods for annotating variants, including

GERP++[25], PhyloP[41], CADD[24], and Eigen[42], but none was found to enrich pathogenic mutations.

Two possible explanations are: (1) Cross-species evolutionary constraint may be only modestly

correlated with pathogenicity in humans[64]. (2) Two of these methods, CADD and Eigen, were

trained on epigenomic data from mostly non-brain tissues. To better analyze the non-coding

genome, we need a more comprehensive repertoire of regulatory information in brain, within

specific spatial-temporal contexts[65].

A recent study estimated the contributions of common variants, rare inherited variants and DNMs

to the risk of autism, and concluded that most of the autism risk is explained by common variants[44].

Our goal here is not to explain ASD risk *per se*, but use DNMs to explore the relative contributions

of coding and non-coding variants. While risk partition among different types of variants has been

pursued in GWAS[63], linkage disequilibrium limits our ability to do this, as discussed previously.

Our work thus provides independent evidence that non-coding variants make a substantial

contribution to the risk of complex diseases[63]. Comparing with the previous study of autism, our

estimate of ASD risk attributable to coding mutations is somewhat higher (1.9% vs. 1.1%) and in

particular we found a significant contribution from missense DNMs (0.83% vs. 0.04%). We

believe that this higher estimate is largely due to our different modeling assumptions: we treated

all mutations in a category (e.g., missense) as a mixture of causal and non-causal mutations,

whereas the previous study treated all mutations in a category equally (see Methods)[44]. We

estimated that *de* novo coding (1.9%), non-coding (0.9%) and copy number variants (1.46%,

estimated by a previous study[44]) together contribute at least 4.2% of ASD risk. We think that we

significantly under-estimated the contribution of non-coding mutations to ASD risk as we did not

take into account mutations in more distal regulatory regions or in ASD-relevant enhancers not

covered by our datasets, and indels or CNVs disrupting regulatory sequences.

We have identified three new ASD risk genes whose expression patterns and biological functions

are consistent with their putative roles in ASD risk. JUP is a member of the catenin/cadherin

superfamily which has important roles in neuron connections and interactions[66]. It is strongly

expressed in the primate prefrontal cortex and hippocampus[67]. A*PBB1* is an adaptor protein

localized in the nucleus and has a role in Alzheimer's disease. It is down-regulated in ASD

cerebellum than control[68], and its microexons are mis-regulated in the brains of ASD individuals[20].

ARHGAP5 negatively regulates RHO GTPase activity. It is required for normal neural

development, and is implicated in axonal branching and synapse formation[69,70].

The pathogenicity of most of the new ASD genes at FDR < 0.3 is supported by functional and

association studies (summarized in Supplementary Table 8). For example, PLXNB1 is a receptor

in the GTPase signaling pathway that regulates the development of hippocampal neurons[71,72],

remodeling of dendrites[73], and the plasticity of axons[74,75]. *PLXNB1* expression has been found to

be altered in the prefrontal cortex of schizophrenia patients[76]. Another example is DPYSL2, a

multifunctional adaptor protein within the central nervous system[77]. DPYSL2 interacts with

various binding partners to regulate multiple biological processes critical in neuron development,

including axon initiation and elongation[77,78], neurite outgrowth[79], neuronal differentiation [80], and

neurotransmitter release[81]. Functional variants in *DPYSL2* have been found to increase the risk of

schizophrenia[82,83]. Only two of our new ASD genes (*MSL2* and *PPM1D*) have not been

functionally implicated in neuro-development. However, *PPM1D* has recently been identified as

a novel risk gene for intellectual disability[84].

We also found that the two TAD regions with excess regulatory SNVs in ASD are supported by

copy number variant (CNV) studies. In one TAD, recurrent, rare CNVs (Chr2: 45455651-

45984915) spanning the entire *SRBD1* gene (the only protein coding gene disrupted by the CNVs

within this TAD) were reported in ASD subjects[85]. In a later independent study, CNVs in this TAD

region were found to be enriched in ASD cases versus controls[85]. These results suggest that *SRBD1*

is likely the risk gene in this TAD. In the other TAD region, ASD-associated duplication of 8p23.1-

8p23.2 introduces a breakpoint between MSRA and RP1L1[86]. MSRA is a member of the

methionine-sulfoxide reductase system whose function is to alleviate oxidative stress. Increased

exposure to oxidative stress plays an important role in the pathogenesis of ASD[87]. In addition,

GWAS studies have established associations of MSRA with schizophrenia[88] and bipolar disorder[89].

All these evidence together suggests that MSRA is a strong candidate gene for autism.

We investigated the power of WES and WGS studies using simulation. We note that the power of

WGS is significantly underestimated for the same reasons, described previously, that the

contribution of non-coding mutations to ASD risk is under-estimated. Additionally, WGS would

enable much better identification of structural variants, which collectively make a significant

contribution to the autism risk. Finally, we did not account for the cost of recruiting patients, which

would further favor choosing WGS over WES, as WGS requires a smaller cohort to achieve the

same level of power.

In summary, our work established genome-wide DNM analysis as a powerful strategy for mapping

genes and regulatory elements underlying autism. By leveraging multiple types of mutations

perturbing the activity and expression of genes, we gain considerable power over using protein-

coding variants alone. Furthermore, comparing with protein-coding genes, regulatory elements are

typically active in specific cell types and brain regions, thus the knowledge of autism-related

regulatory elements may help us identify important neuronal subtypes and circuits relevant to

autism, a crucial step for understanding its mechanisms[3]. With lowering cost of genome

sequencing and better annotations of regulatory genomes in the brain[65], we anticipate that whole-

genome DNM studies will transform our understanding of neuropsychiatric disorders.

Methods

<u>DNMs from whole genome sequencing data</u>

The detailed information for each WGS dataset is summarized in Table S1. To remove erroneously

called *de novo* SNVs, we excluded 8 individuals with more than 140 (2 times more than the median

of ASD DNMs per individual) DNMs and removed all recurrent DNMs (i.e. exactly the same

mutation in multiple individuals). Our unpublished (Wu et al.) DNM data are from WGS of 32

ASD trios of Han Chinese ancestry (see Section 1 of Supplementary Methods for details, data are

available at http://wwwdev.ebi.ac.uk/eva/?eva-study=PRJEB14713).

Collection of brain cis-regulatory regions

We used H3K27ac sites in fetal and adult brains to define cis-regulatory regions in our main

analyses. Fetal brain sites from human cortex at embryonic stages 5 p.c.w., 7 p.c.w., and 12 p.c.w.

were obtained from a recent study[37]. For each stage, only peak regions consistent between two

biological replicates were selected. Adult brain sites were obtained from Roadmap Epigenomics

Project[38]. They include H3K27ac sites from human angular gyrus, anterior caudate, cingulate

gyrus, middle hippocampus, inferior temporal lobe, mid-frontal lobe and substantia nigra. We used

MACS2 to call peaks from raw data, and only kept peak regions consistent between two biological

replicates for each brain region. We used BEDtools[90] to merge H3K27ac sites from fetal and adult

brain.

To define brain-specific enhancers in Fig. 2d, we compared H3K27ac data from human embryonic

cortex with publicly available H3K27ac datasets for the colon, the esophagus and limbs[37,91]. For

each H3K27ac dataset, only replicated peaks were included and any peak with 1-bp overlap to a

Gencode(v19) annotation feature was excluded. A single composite multi-sample enhancer

annotation was generated by merging replicated peaks across all tissue samples. The level of

H3K27ac signal in each region for each sample was quantified by averaging read counts per

kilobase per million mapped reads (RPKM) in each region from each replicate. Each region was

represented by a vector of a length equal to the total number of tissues considered, with each

element representing the RPKM value of marking in that region for a single tissue. Each vector

was standardized to have zero-mean and unit-variance. The matrix of these normalized tissue

quantification values was then subjected to k-means clustering to identify sets of sites exhibiting

the strongest marking in each tissue relative to all other samples in the comparison.

Annotation of regulatory mutations

We annotated DNMs within regulatory regions using various bioinformatics tools. To limit our

analysis to motifs that are most likely to be ASD-relevant, we used JASPAR core vertebrate TF

motifs[92] that meet two criteria: (1) they are enriched in regulatory sequences of gene modules that

are adjacent to regulatory sequences active in developing cortex of human, but not other primates[37];

and (2) their cognate transcription factors are in the top 50% brain-expressed genes (Expression

levels based on mean expression levels across Brainspan samples). To identify motif-changing

SNVs, for each SNV, we generated a 41-bp wild-type sequence centering on the reference allele,

along with a 41-bp mutant sequence centering on the mutated allele. We then used FIMO[93] to

obtain its maximal motif score difference (log(ratio of p-values)) in absolute value between its

wild-type and mutant sequences. A SNV whose maximal absolute motif score difference is in the

top 10% among the difference of all SNVs was defined as a motif SNV. We used ANNOVAR to

annotate coding mutation types and obtain GERP++ scores for all mutations[94]. CADD scores were

downloaded from http://cadd.gs.washington.edu/score, (default parameters, v1.3). Other

annotations that we used include PhyloP scores from the 100-way PhyloP wiggle file from UCSC

(/goldenpath/hg19/phyloP100way/) and Eigen-PC scores[42].

<u>Lists of genes used in burden analyses</u>

Known ASD genes include genes with q-value < 0.3 from a combined analysis of CNVs, indels,

and WES data using TADA[35], SFARI category I (high confidence) and II (strong candidate) genes

(https://sfari.org), AutismKB genes[95] and ASD genes summarized in a previous study[96].

Neuropsychiatric genes are a larger set of genes likely involved in neuropsychiatric disorders,

including genes with TADA q-value < 0.5[35] , SFARI genes (https://sfari.org), AutismKB genes[95],

ASD risk genes summarized in a previous study[96], intellectual disability genes[97], the union of gene

sets enriched with SCZ *de novo* coding mutations[98] and high confidence postsynaptic density

genes[99]. The set of nonASD genes are the 1000 genes with the highest TADA q-values[35]. The size

of this gene set is comparable to that of neuropsychiatric genes. Intolerant genes include genes

with top 5% RVIS[64] and haploinsuffcient genes obtained from two sources, one using copy number

variations (genes with predicted haploinsufficient probability greater than 0.95)[47], and the other

using estimated mutation rate[48]. To define tolerant gene, we started with genes with RVIS scores

in the bottom 10%[64], genes with haploinsufficient probability smaller than 0.1[47], and genes that

were used as control genes for LoF deficient genes[48]. We then removed any genes that were in the

intolerant gene set. To define gene groups based on their expression levels, we used the average

expression level for each gene across all developing brain tissues from BrainSpan data.

<u>Burden analysis of different types of *de novo* mutations</u>

In burden analysis, we accounted for the difference in mutation rate between ASD subjects

(~60/individual) and controls (~39/individual) as follows. Instead of comparing absolute mutation

rates, we normalized the mutation counts in the regions of interest by the mutation counts of some

background sequences that should not be enriched with risk mutations. This relative mutational

frequency is thus comparable between ASD and controls.

Specifically, for the burden analysis of coding sequences, we first calculated the ratio of

nonsynonymous mutations and synonymous mutations for both ASD and control samples. This

ratio can be thought of as the relative frequency of nonsynonymous mutations in cases or controls.

The burden of nonsynonymous mutations is then the ratio between the relative frequency of

nonsynonymous mutation of ASD subjects and that of control subjects:

$$\text{Coding burden} = \frac{ASD\ nonsynonymous\ SNVs/ASD\ synonymous\ SNVs}{control\ nonsynonymous\ SNVs/control\ synonymous\ SNVs}$$

The p-value of the burden was calculated from Fisher's exact test.

We defined splicing SNVs as SNVs whose delta-splicing scores predicted by SPIDEX[26] are in the

lower 10% of all the SNVs in our studies. At this cutoff, enriched splicing SNVs near known ASD

genes have been observed previously[26]. We defined non-spicing mutations as SNVs whose delta-

splicing scores are in the middle 10%. The burden of splicing SNVs was calculated in a manner

similar to coding mutation, using synonymous SNVs as background mutations.

For the burden of regulatory SNVs, we followed the same strategy but used all regulatory SNVs
as background instead of synonymous mutations. The burden of regulatory SNVs with a certain
annotation (or in a certain group of genes) is expressed as:

$\text{Regulatory SNV burden} = \dfrac{\textit{ASD regulatory SNVs with a certain annotation/all ASD regulatory SNVs}}{\textit{control regulaotrySNVs with a certain annotation/all control regulatory SNVs}}$

<u>Burden analysis by simulation</u>

To estimate the burden of ASD regulatory SNVs in a gene set, we sampled the number of
regulatory ASD SNVs of a gene using a multinomial distribution, with the total number given by
the observed number of ASD regulatory SNVs. The multinomial probability of a gene is
proportional to the total mutation rate of H3k27ac regulatory regions within 10 kb of the TSS of
that gene. We performed simulations 10,000 times and generated the null distribution of the total
number of mutations that land in the gene set of interest. We then used this distribution to derive
the empirical enrichment p-value and fold of enrichment (observed mutation number in the gene
set divided by the mean of simulation distribution). To estimate the burden of motif-changing
SNVs, we first shuffled the positions of ASD regulatory SNVs within all H3K27ac regulatory
regions (under the same mutation-rate-based multinomial distribution used above), and then, for
each new position, generated a mutant allele based on the mutational spectrum of all ASD SNVs.
We counted the number of motif-changing SNVs from each of 10000 simulations, and generated
a null distribution of this number, which is used to obtain enrichment p-values of the observed
number of motif-changing SNVs.

<u>Mutation rates of regulatory sequences</u>

We applied the tri-nucleotide mutation model developed by a previous study[45] to the whole-
genome to derive the base-level mutation rate. The mutation rate for each functional unit (e.g., one
enhancer) is defined as the sum of the per-base mutation rates[27] in the unit. To approximate the
mutation rate of splicing SNVs for each gene, we counted the number of splicing SNVs across
genes in controls, and used the ratio between this number and the total number of synonymous

mutations in controls as a scaling factor. We then multiplied this scaling factor to the synonymous

mutation rate of each gene to estimate its splicing SNV mutation rate.

<u>Estimating the relative risks of de novo mutations</u>

We used a strategy developed previously to estimate the average relative risk of a given type of

mutation (section 6 in the Supporting Text of TADA). It is also described in the user guide of the

software implementation of TADA[27]. Note that the relative risk of a type of mutation (e.g. missense)

is interpreted as the average relative risk of a missense mutation in a causal gene. Let λ be the

burden estimated for the mutational types of interest across all genes. Because causal genes

represent only a small proportion of all genes, the relative risks in causal genes should be higher

than λ. Let π be the fraction of genes that are risk genes, the relative risk of causal genes can be

calculated as $\frac{\lambda-1}{\pi} + 1$. We used $\pi = 0.06$ in our analysis of coding mutations, which corresponds

to there being 1000 ASD risk genes[9,27]. For regulatory SNVs and splicing SNVs, the genome-

wide burdens of these two types of mutations are not significant, so we instead limited our analysis

to previously indicated ASD genes with TADA q-value <0.5[35]. We then converted the burden in

these genes to relative risks with $\pi = 0.5$ using the same equation. When estimating the relative

risk of splicing SNVs, we removed SNVs that were predicted to be nonsynonymous mutations or

that fell into the brain H3K27ac regions within 10kb of TSSs.

<u>Partitions of *de novo* ASD risk into coding and non-coding mutations</u>

We treated ASD liability (risk) as a continuous trait, and estimated the percentage of variance in

ASD liability explained by four types of mutations (Fig. 2f). The variation of ASD liability

explained by the *j-th* type of mutations could be expressed as the equation below (see

Supplementary Methods):

$$V_j = \beta_j^2 * p_j * \left(1 - p_j\right)$$

where β_j is the effect size of the *j-th* type of mutations at the liability scale and p_j is the probability

that an individual carries a mutation of type *j*. Note that only causal mutations contribute to ASD

liability, so both β_j and p_j are defined for mutations affecting causal genes. We calculated β_j from

the relative risk of j-th type of mutation using standard quantitative genetic calculations. To obtain

p_j, we calculate the total mutation rate of type j mutations (note that an individual has two copies

of genomes), and then multiply this by 0.06 (fraction of ASD risk genes) to obtain the rate of causal

mutations of type j (see Section 4 of Supplementary Methods for details).

<u>TADA analysis to map ASD risk genes</u>

A detailed description of TADA can be found in Section 5 of Supplementary Methods and in

previous publications[8,27]. To combine evidence of the coding and non-coding mutations for a

particular gene, we multiplied the pre-computed coding-mutation Bayes factors from published

work[35] by the non-coding Bayes factors derived from our study. If a gene does not have brain

H3K27ac regions within 10 kb of its TSS, we assigned its non-coding Bayes factor to be 1. We

translated the combined Bayes factor of each gene into a q-value using Bayesian FDR control.

To define new ASD genes at a particular q-value cutoff, we chose the genes with a combined q-

value below the cutoff and a coding q-value above the cutoff. We then filtered out new ASD genes

that do not have H3K27ac regions within 10kb of TSSs. Note that even if a gene has no evidence

from non-coding mutations, its q-value could change, because the q-value depends on the

distribution of Bayes factors of all genes. We also lost a few genes at each q-value cutoff: These

genes have no regulatory SNVs and thus receive a small penalty to their Bayes factors.

<u>DAWN analysis</u>

DAWN (Detecting Association With Networks) algorithm[49,50] is a guilt-by-association-based gene

prediction algorithm. The fundamental assumption underlying the algorithm is that risk genes are

part of a functional cluster and thus, tightly connected. A gene has a high posterior risk probability

if it is interacts in a network with other risk genes or has a high prior risk probability or both. We

used TADA to assign each gene a p-value for the prior evidence of being an ASD risk gene. For

the underlying network, we constructed partial co-expression networks for two spatio-temporal

windows, namely the mid-fetal prefrontal cortex (PFC) and the infancy mediodorsal cerebellar

cortex (MD-CBC), which are indicated as high risk windows for ASD[51]. BrainSpan microarray

dataset is used as the source for spatio-temporal gene expression data

(http://developinghumanbrain.org).

DAWN was run separately for each above-mentioned network. We used regularization parameter

(lambda) = 0.12, p-value cutoff = 0.1 and correlation thresholds 0.7 for PFC and 0.85 for MD-

CBC, respectively. In Table 2, posterior risk scores (q-values) are shown for the candidate genes.

A Dash means that the corresponding gene is not co-expressed with other risk genes in any of the

spatio-temporal windows.

<u>GeneMania analysis</u>

GeneMania[52] is a tool for studying interactions among genes in a network using various types of

information, such as gene co-expression and protein-protein interactions (PPIs). We first studied

the connection between our candidate genes with high-confidence ASD genes (genes with coding

TADA FDR < 0.1 and genes in SFARI categories I and II, https://sfari.org) using co-expression

data. We then chose 30 top neighboring genes for each of our new candidate genes using multiple

networks, including co-localization, genetic interactions, shared pathway, physical interactions,

predicted interactions and shared protein domains.

<u>Enhancers with recurrent *de novo* mutations</u>

For this analysis, we used brain H3K27ac regions not overlapping with exons. We observed 25

enhancers with at least two *de novo* SNVs in ASD samples. We performed simulations to assess

its signfiicance. In each simulation, we randomly re-distributed *de novo* SNVs of all brain

enhancers, following a multinomial distribution. The multinomial probability of an enhancer is the

ratio between the mutation rate of that enhancer and the sum of mutation rates across all enhancers.

We performed simulations 10,000 times and obtained the distribution of the number of enhancers

with recurrent SNVs.

TADs with recurrent *de novo* SNVs

For each TAD region, we calculated its regulatory mutation rate as the sum of per-base mutation

rates of brain H3K27ac sites within the TAD. Under the null hypothesis, the count of regulatory

SNVs follows a Poisson distribution, whose rate is the regulatory mutation rate times twice of the

sample size[27,45]. We then calculated the Poisson *p*-value of each TAD region and used the

Benjamini-Hochberg procedure to control FDR.

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Contributions

Y.L.[Yuwen] and X.H. led the study design and the writing of the manuscript; Y.L.[Yuwen]

performed all the major analyses; E.C. performed network analysis with DAWN and GeneMania.

Y.L.[Yanyu], N.K. and I.I.-L contributed to burden analyses; J.L., Y.M., Y.W., Y.J., Z.L., J.W.,

K.X., and Z.S.S. generated new WGS data from ASD trios, called DNMs, and calculated mutation

rates for non-coding regions; R.M., M.K., E.G., and J.N. provided developing brain enhancer data,

tissue-specific enhancers and contributed ideas to the burden analysis and TAD analysis; All

authors contributed to the writing of the manuscript.

Competing financial interests

The authors declare no competing financial interests.

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Figure 1

Figure 1. Overview of study design. We consider four mutational types, including LoF: Loss-of-Function coding mutations; Mis3: probably damaging nonsynonymous coding mutations predicted by Polyphen-2; SplSNV: SNVs predicted to affect pre-mRNA splicing by SPIDEX; RegSNV: non-coding SNVs that fall into brain H3K27ac regions within 10 kb of TSSs. RR denotes average relative risks of different types of mutations in risk genes. In the Mutation Count/ Expected Count Box, each cell contains the observed mutation count (colored in red) and the expected count (colored in black) (the numbers are for illustration only). TADA combines information from all mutation types related to a gene to test how likely it is to be a risk gene. The average relative risk of a mutation type is used in TADA to weigh the contribution of that type.

Figure 2

Figure 2. Burdens analyses of different types of de novo mutations. The error bars in (a), (b), and (c) represents the 95% confidence intervals of burdens, based on the lower bounds and upper bounds of the odds ratios from Fisher's exact tests. We did not show error bars in (d) and (e) as they are too large to show. (a) De novo nonsynonymous mutations from WGS studies in multiple gene sets. See Methods for the definitions of these gene sets. (b) De novo SNVs in H3K27ac regions within 10 kb of TSSs. In the barplot, we show burdens of promoter/5'UTR SNVs and regulatory SNVs (including all SNVs in brain H3K27ac regions within 10 kb of TSSs). (c) Regulatory SNVs and distal regulatory SNVs (10 kb ~ 50 kb to TSSs) in different gene sets. (d) Comparison between burdens of SNVs in brain-specific H3K27ac regions and of those in other-tissue-specific H3K27ac regions. (e) Burdens of SNVs that are predicted to affect splicing vs. SNVs not predicted to have such effects (non-splicing SNVs). (f) Partition of de novo ASD risk into coding and non-coding mutations.

Figure 3

Figure 3. Predicted risk genes and enhancers of ASD. (a) GeneMania network analysis of the three new ASD genes. Red circles represent new ASD genes while grey circles represent known ones. The figure shows only the co-expression networks: two genes are connected if their co-expression across multiple tissues reaches a threshold. Only connections between the two gene sets are shown. (b) A de novo SNV in the TSS-proximal regulatory region of PLXNB1 disrupts a RFX2 binding motif. (c) Distribution of the number of enhancers with recurrent (more than two) de novo SNVs from 10,000 simulations. The vertical red arrow marks the observed number of enhancers with recurrent de novo SNVs from ASD data. (d) A plausible distal enhancer (marked by a star) with recurrent ASD SNVs has correlated activities with ZMIZ1 promoter. Grey curves represent correlated activities between enhancers and ZMIZ1 promoter.

Figure 4

Figure 4. Comparison of power between WES and WGS from simulations. Power is measured as the expected number of discovered ASD risk genes at q < 0.1, and is obtained at each level of sample size (a) or sequencing cost (b).

Table 1. Mutational exposure (number of mutations from ASD risk genes per subject) and relative risk of different types of *de novo* mutations.

Mutation Class	Mutational Exposure	Relative Mutational Exposure(%)	Relative risk	Variance of ASD liability explained (100%)
Mis3	0.0175	11.6	4.7	0.83
Loss-of-function	0.00405	2.68	20	1.08
Regulatory SNV	0.115	76.1	1.9	0.72
Splicing SNV	0.0145	9.60	2.1	0.14

Table 2. Mutational counts and evidence of three new ASD genes. In the evidence rows, Y means overlap with a gene set and N otherwise.

Gene name	JUP	APBB1	ARHGAP5	Enrichment p-value
LoF	1	1	1	
Mis3	0	0	1	
Regulatory SNV	4	3	1	
HI	Y	Y	Y	**7.90e-4**
RVIS (%)	2.74	19.93	20.02	**1.11e-2**
ExAC zscore (%)	30.62	14.25	9.69	1.38e-1
FMRP targets	N	Y	N	1.25e-1
BrainSpan expression (%)	19.29	3.41	22.03	**1.49e-2**
DAWN	5.00e-4	-	2.00e-4	**4.06e-4**

Lower RVIS and ExAC zscores percentiles correspond to higher constraint.
Lower BrainSpan percentiles correspond to higher brain expression.
Enrichment p-values were calculated by hypergeometric tests. In RVIS, ExAC zscore, Brainspan, we tested the enrichment of new ASD genes in genes in the lower quartiles. In DAWN analysis, we tested the enrichment of new ASD genes in genes with DAWN q-value < 0.05. The DAWN q-value for each gene in the table is the minimum of the q-values of that gene in two brain regions, mid-fetal prefrontal cortex (PFC) and infancy mediodorsal cerebellar cortex (MD-CBC).

Table 3. Biological processes enriched among neighborhood genes of the three new ASD genes.

Term	P-value	Adjusted P-value	Overlap	Neighbor Genes
Adherens junction organization	1.17e-05	1.06e-02	4/69	DSP, CTNND1, CTNNA1(APBB1); THY1(ARHGAP5)
Negative regulation of neurogenesis	3.59e-05	1.27e-02	5/192	APP, NOTCH1, LRP4(JUP); CTNNA1(APBB1); THY1(ARHGAP5)
Neuromuscular junction development	5.00e-05	1.27e-02	3/34	APP, ERBB2, LRP4(JUP)
Negative regulation of nervous system development	5.57e-05	1.27e-02	5/211	APP, NOTCH1, CTNNA1(APBB1); LRP4(JUP); THY1(ARHGAP5)
Negative regulation of cell development	9.92e-05	1.51e-02	5/239	APP, NOTCH1, LRP4 (JUP); CTNNA1(APBB1); THY1(ARHGAP5)
Kidney development	1.00e-04	1.52e-02	4/122	MME, TSHZ3, LRP4 (JUP); ZBTB16(APBB1)
Neuron remodeling	1.64e-04	1.98e-02	2/8	APP(JUP); RND1(ARHGAP5)
Forebrain development	1.74e-04	1.98e-02	3/53	APP, NOTCH1, APLP2(JUP)
Negative regulation of neuron differentiation	2.01e-04	2.04e-02	4/147	APP, NOTCH, LRP4(JUP); THY1(ARHGAP5)
Mating behavior	2.41e-04	2.20e-02	2/10	APP, APLP2(JUP)

Neighborhood genes are the top 30 connected genes from GeneMania analysis for the three new ASD genes. In the last column, after each set of neighbor genes (separated by semicolon) in parenthesis is their common connected new ASD gene. In the "Overlap" column, the number of neighborhood genes in each GO category is followed by the total number of genes in this GO category.