Recovery of hand function after stroke: separable systems for finger strength and

control

Running title: Post-stroke recovery of hand function

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Abstract

Loss of hand function after stroke is a major cause of long-term disability. Hand function

can be partitioned into strength and independent control of fingers (individuation). Here

we developed a novel paradigm, which independently quantifies these two aspects of

hand function, to track hand recovery in 54 patients with hemiparesis over the first year

after their stroke. Most recovery of both strength and individuation occurred in the first

three months after stroke. Improvement in strength and individuation were tightly

correlated up to a strength level of approximately 60% of the unaffected side. Beyond

this threshold, further gains in strength were not accompanied by improvements in

individuation. Any observed improvements in individuation beyond the 60% threshold

were attributable instead to a second independent stable factor. Lesion analysis revealed

that damage to the hand area in motor cortex and the corticospinal tract (CST) correlated

more with individuation than with strength. CST involvement correlated with

individuation even after factoring out the strength-individuation correlation. The most

parsimonious explanation for these behavioral and lesion-based findings is that most

strength recovery, along with some individuation, can be attributed to descending

systems other than the CST, whereas further recovery of individuation is CST dependent.

Keywords:

Finger individuation, strength, stroke, motor recovery, plasticity

Introduction

Human hand function comprises at least two complementary aspects: strength as manifest in a power grip, and control of individual finger movements as in piano playing. The most common observation after stroke is that both are impaired (Kamper and Rymer, 2001; Lang and Schieber, 2003). Weakness presents as difficulties in voluntarily opening of the hand, extending the wrist and fingers against resistance, and producing a strong grip (Colebatch and Gandevia, 1989; Kamper *et al.*, 2003). Loss of finger control manifests as inability to either move a single finger while keeping the others immobile, or to make complex hand gestures, both of which impair the ability to perform tasks such as typing or buttoning a shirt (Kamper and Rymer, 2001; Li *et al.*, 2003; Lang and Schieber, 2004). When strength does recover after stroke, control often remains impaired, causing lasting disability (Heller *et al.*, 1987; Sunderland *et al.*, 1989). However, the relationship between strength and control after stroke remains poorly understood. Separating the effect of stroke on finger strength versus control is a challenge given that most current clinical measurements conflate weakness with deficits in control. In the current study we therefore sought to develop a new paradigm that could measure these two aspects of hand function separately, and to investigate the relationship between strength and control over the time course of hand recovery after stroke. We were specifically interested to test whether these two components recover in a lawful relationship with each other, or whether they recover independently.

Existing behavioral tasks used to assess hand function after stroke, such as the Fugl-Meyer Assessment (FMA) (Fugl-Meyer *et al.*, 1975), the Nine-Hole Peg Task (9NPT) (Sharpless, 1982), and the Action Reach Arm Test (ARAT) (Lyden and Lau,

1991), are not designed to separate deficits in strength and control. To isolate these two

aspects of hand function it is necessary to remove any obligatory relationship between

them (Reinkensmeyer *et al.*, 1992), i.e. derive a control measure that is independent of

strength. Intuitively, a rock climber may have stronger fingers than a pianist, but not

necessarily superior control of individual fingers.

Schieber (1991) devised an individuation task that requires participants to move

individual fingers while keeping the non-moving ones stationary. Movements of the

passive fingers were used as a measure of loss of control. This paradigm however does

not directly track the force relationship between active and passive fingers. In the

paradigm used here, we first measured the maximum voluntary contraction force (MVF)

that a participant could produce with each finger. We then asked participants to produce

isometric forces over four sub-maximal levels with each finger, while keeping the passive

fingers immobile. Even controls show involuntary force production (enslaving) on the

passive fingers, which increases with the required active force level (Li *et al.*, 1998;

Zatsiorsky *et al.*, 2000). The slope of the function of passive finger enslaving on active

force thus provides a measure of individuation that is independent of strength.

Using this paradigm we tracked the recovery of hand strength and finger

individuation in patients over a one-year period after stroke. One possibility is that

strength and control recover independently. For example, a patient may remain quite

weak but have good recovery of individuation, or a patient may recover a significant

amount of grip strength but fail to individuate the digits. Alternatively, recovery may be

such that when strength recovers so does individuation, because either they share a

common neural substrate or repair processes are proceeding in parallel in separate neural

substrates. Finally, lesion analysis allowed us to investigate whether there is any

identifiable anatomical basis for any observed dissociation between strength and control

deficits.

Materials and Methods

Participants

Fifty-four patients with first-time ischemic stroke and hemiparesis (34 male, 20

female; mean age 57.4±14.9 years) were recruited from three centers: The Johns Hopkins

Hospital and Affiliates, Columbia University Medical Center, and The University

Hospital of Zurich and Cereneo Center for Neurology and Rehabilitation. According to

the Edinburgh Handedness Inventory (Oldfield, 1971), Forty-four patients were right-

and 10 were left-handed. All patients met the following inclusion criteria: 1) First-ever

clinical ischemic stroke with a positive DWI lesion within the previous 2 weeks; 2) One-

sided upper extremity weakness (MRC < 5); 3) Ability to give informed consent and

understand the tasks involved. We excluded patients with one or more of the following

criteria: initial UE FMA > 63/66, age under 21 years, hemorrhagic stroke, space-

occupying hemorrhagic transformation, bihemispheric stroke, traumatic brain injury,

encephalopathy due to major non-stroke medical illness, global inattention, large visual

field cut (greater than a quadrantanopia), receptive aphasia (inability to follow 3-step

commands), inability to give informed consent, major neurological or psychiatric illness

that could confound performance/recovery, or a physical or other neurological condition

that would interfere with arm, wrist, or hand function recovery. Due to the exclusion of

aphasic patients, the sample had a bias towards right-sided infarcts (17 left-sided, 37

right-sided; for detailed patient characteristics, see Table 1). The lesion distribution is

shown in Fig. 5A.

We also recruited 14 age-matched healthy control participants (10 male, 4

female; mean age 64±8.2 years; all right-handed) at the three centers. There was no age

difference between the patient and control samples (two-samples t-test $t(65) = 1.60$, $p =$

0.11), nor did the ratio of gender and handedness in the two groups differ (Fisher's exact

test yielded results of p = 0.11 and 0.75, respectively). The healthy controls did not have

any neurological disorder or physical deficit involving the upper limbs. All participants

signed a written consent, and all procedures were approved by Institutional Research

Board at each study center.

Assessment of finger maximum voluntary contraction and of individuation

To achieve good characterization of hand function recovery, the study design

required patient testing at the following five time points post-stroke: within the first 2

weeks (W1, 10±4 days), at 4-6 weeks (W4, 37±8 days), 12-14 weeks (W12, 95±10 days),

24-26 weeks (W24, 187±12 days), and 52-54 weeks (W52, 370±9 days). Healthy controls

were tested at comparable intervals.

At each of the five visits, hand function was tested using an ergonomic device that

measures isometric forces produced by each finger (Fig. 1A). The hand-shaped keyboard

was comprised of ten keys. Force transducers (FSG-15N1A, Honeywell[®]; dynamic range

0-50 N) measured the force exerted by each finger with a sampling rate of 200 Hz. The

data were digitized using National Instrument USB-621x devices interfacing with

MATLAB (The MathWorks, Inc., Natick, MA) Data Acquisition Toolbox. Visual stimuli

of the task were presented on the computer monitor, run by custom-written software

using the Psychophysics Toolbox (Psychotoolbox) in MATLAB environment (Brainard,

1997).

Participants were seated in a comfortable chair, facing the computer monitor.

During the entire procedure, participants rested their two hands on the keyboards with

each finger on top of a key, their wrists were strapped and fixed on a wrist-rest, and their

forearms extended and supported by foam arm rests. Throughout the experiment, ten

vertical gray bars representing the ten digits appeared on top of the screen, and another

ten vertical bars below them instructed the amount of force to be exerted; the required

force level for each finger in each trial was indicated by the height of green filling the

vertical gray bar (Fig. 1B). Participants could monitor the force exerted by all ten digits

in real time by the heights of ten small white horizontal lines moving along the vertical

force bars.

Two separate aspects of finger function were tested: maximal voluntary

contraction force (MVF) and individuation. During each MVF trial, participants were

asked to depress one finger at a time with its maximum strength, and maintain the force

level for two seconds. The participants could press with the other fingers as much as they

wanted as long as maximal force on the instructed finger was achieved. To signal the

start, one force bar corresponding to the instructed finger turned to green. MVF was

measured twice per digit.

In the individuation task, participants had to press each individual finger at a sub-

MVF level of force, while at the same time keeping their other fingers immobile on the

keys. Four target force levels were tested for each digit: 20%, 40%, 60%, and 80% of

MVF, and each level was repeated 4 times. On each trial, a section of a force bar

corresponding to the finger to be depressed turned to green, with the height of the middle

black line representing the target force level and the green region around the middle line

representing the 25% upper and lower bounds around the target force level (Fig. 1B). The

participants were asked to bring the corresponding white line up to the force target line,

and maintain the force level for 0.5 sec. If no response passing the force threshold of 2.5

N was detected within two seconds, the trial was terminated.

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190   Insert Figure 1

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Data analysis

Strength Index. The 95th percentile of the force traces produced across all the

sampled force data points during the finger depressing period in each trial was calculated,

and then averaged across the two MVF trials to obtain a measure of MVF for each digit.

If the force achieved on one of the two trials was below 60% of the force produced on the

other trial, only the larger force was taken as MVF measure (6.5% of the trials were

excluded). The overall strength of the hand was then calculated by averaging across all

five digits. To account for the large inter-subject variability in premorbid strength, all

MVF values were normalized by MVF of the non-paretic hand at W52; estimated using a

mixed-effects model (see below). This normalization provided a Strength Index, with a

value close to 1 implying full recovery. For control participants, one of the hands was

randomly assigned to take the role as the "non-paretic" hand for normalization purposes.

To account for possible laterality effects, the assignment followed the ratio of dominant

to non-dominant hands found in the patients (~10:4).

Individuation Index. If individuation was perfect, a participant should be able to

press the instructed finger without any force being exerted by the passive fingers. For

each time bin t (5ms) in a single trial, we calculated the enslaved deviation of the force of

each passive finger ($F_{t,j}$) from baseline force (BF_j), which was assessed at the beginning

of the trial when a go cue was presented. This deviation was averaged over all bins (T) in

the force trace from the go cue to the end of the trial:

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$$meanDevP = \frac{1}{T}\sum_{t=0}^{T}\sqrt{\sum_{j=passive}(F_{t,j} - BF_j)^2}$$
(1)

where the index j denotes the jth passive finger. A higher *mean deviation* indicates more

enslaving of the passive finger.

For a measure of individuation ability, it is necessary to account for the

relationship between the force deviations of the passive fingers to the force produced by

the active finger. Consistent with previous reports (Li *et al.*, 1998), we observed that

enslaving of passive fingers increases with higher active force (Fig. 1E). The relationship

between the two variables was close to linear. Thus a good measure of individuation is

how much the mean deviation in passive fingers increases for each N of force produced

by the active finger. The ratio of these two variables can be reliably estimated by fitting a

regression line without an intercept. To reduce the influence of outliers, we used robust

regression (Holland and Welsch, 1977). The *slope* of the regression line reflects

individuation ability: The smaller the slope, the better the individuation ability, with the

best case being 0, which means keeping the passive fingers perfectly immobile at any

active force level. Because the regression slope is bounded by zero (as mean deviation is

positive), its distribution is positively skewed. To allow for the use of parametric

statistics the slope was log-transformed. The sign of this value was inverted, so that

higher values would correspond to better function. The negative log slope was calculated

separately for each active finger and then averaged across fingers, giving the

Individuation Index for the hand. As was done for the Strength Index, the Individuation

Index was normalized by each participant's "non-paretic" (randomly assigned for healthy

controls) hand's W52 value as estimated by mixed-effect model to provide the final

Individuation Index for each hand.

Reliability measures for Strength and Individuation. To determine the reliability

of the Strength and Individuation Indices, split-half reliabilities for both measures were

calculated. For the Strength Index, we used one MVF trial per digit in each split. We then

calculated the (normalized) Strength Index on each half of the data independently in the

same way as for the full data set. The correlation between the two halves across all

available sessions and patients was then used as a measure of split-half reliability.

For the Individuation Index, data from each finger was split such that two trials

per force level were assigned to each split. The slope of the regression line and

Individuation Index was then calculated separately for each split, and normalized in the

same way as MVF. We repeated the split multiple times, each time assigning trials at

random and then averaging the split-half correlations from all splits for more reliable

results.

Split-half correlation will underestimate reliability because the variability in each

half will be higher than the variability when using all the data (Guttman, 1945). The

estimate was therefore corrected using the formula

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$$r_{full} = \frac{2r_p}{r_p + 1} \qquad (1)$$

where r_p is the correlation between the two splits.

Stability analysis. To assess whether the relative deficits in strength and

individuation remained stable across different testing time points, or whether there was

meaningful biological change, we calculated the correlation of each measure across

neighboring testing time points. One caution when interpreting these correlations is that

the correlation between two repeated measures will always be smaller than 1 even if the

underlying factor did not change. This is because both measures contain some

measurement noise. To account for this effect, we used the reliability (r_{full}) of the measure

at each time point to compute a noise ceiling, which indicates how much two repeated

noisy measurements should correlate with each other if the underlying variable were

perfectly stable:

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$$r_{noise\ ceiling} = \sqrt{r1_{full} * r2_{full}} \quad . \qquad (2)$$

264   *Statistical analysis and handling of missing data.* Data analysis was performed

265   using custom-written MATLAB and R (R Core Team, 2012) routines. The analysis

266   focused on the Strength and Individuation Indices, but was also performed on standard

267   clinical assessments, FMA, ARAT, and Dynamometry strength measures.

268   The requirement for 5 post-stroke time-points was ambitious, with the

consequence of some missing sessions. A total of 21 patients completed all five time-

points; on average each patient completed 3.6 sessions; thus a total of 75% of the possible

sessions were acquired. To optimally use all the measured data, we employed linear

mixed-effect models. The model specifies joint distributions for observed and missing

observations – then the parameters of those distributions can be estimated by maximizing

the likelihood of the data under the model. There are several advantages to this approach.

First, all the available data can be used and there is no need to exclude any data.

Secondly, it avoids the statistical pitfalls inherent in "filling in" missing observations

with point estimates. Linear mixed-effect models implemented in the *lme4* package in R

(Bates *et al.*, 2014) were used to test the changes in these measures over time. Participant

was taken as a random factor. Time Point (five time points from W1-W52) and Hand

Condition (paretic, non-paretic, and control) were considered fixed factor. The model was

applied to control and patient data separately. Mixed-effect model estimation for group

summary statistics was implemented in MATLAB using the restricted maximum

likelihood method (Laird and Ware, 1982).

Modeling the time-invariant function. To test the hypothesis that there is time-

invariant relationship between strength and individuation, a two-segment piecewise linear

function was fitted. This function had four free parameters: the intercept, the location of

the inflection point, and the slope on each side of the inflection point. Let x be the

predictor with two segments separated by a constant breakpoint c, $x_1 \leq c$ and $x_2 \geq c$. The

linear functions for each segment are

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$$
\begin{aligned}
y_{1i} &= b_{10} + b_{11}x_{1i} + e_{1i} \\
y_{2i} &= b_{20} + b_{21}x_{2i} + e_{2i}
\end{aligned}
\tag{3}
$$

291   The two pieces can be joined at the breakpoint constant c by setting $y_{1i} = y_{2i}$, yielding

292
$$
\begin{aligned}
b_{20} &= b_{10} + (b_{11} - b_{21})c \\
y_{2i} &= b_{10} + (b_{11} - b_{21})c + b_{21}x_{2i} + e_{2i}
\end{aligned}
\tag{4}
$$

Putting the two pieces together, we have the full model

$$y_i = a + b_1 x_i \cdot I(x_i \leq c) + \left[(b_1 - b_2)c + b_2 x_i\right] \cdot I(x_i \geq c) + e_i \qquad (5)$$

where $I(\cdot)$ is an indicator variable, coded as 1's or 0's to indicate the condition satisfied.

The maximum-likelihood (or least-squares) estimates of these parameters were

obtained by using the non-linear optimization routine fminsearch in Matlab. Leave-one-

out cross-validation (Picard and Cook, 1984) was used to evaluate whether this function

changed systematically over time, or whether it was time-invariant. The time-invariant

model with fixed parameters across all time points was compared with a more complex

model that allowed free parameters for each time point. Cross validation provides an

unbiased estimate of a model's ability to predict new data and automatically penalizes

models that are too complex.

Lesion Imaging and Quantification

Imaging acquisition and lesion distribution. Images were acquired using 3T MRI

Phillips scanner and consisted of two DTI datasets (TR/TE=6600/70ms, EPI, 32 gradient

directions, b=700 s/mm2), and an MPRAGE T1-WI (TR/TE=8/3.8ms) sequence. FOV,

matrix, number of slices, and slice thickness were 212×212 mm, 96×96 (zero-filled to

256x256), 60, 2.2mm, respectively, for DTI; and 256×256mm, 256×256, 170, 1.2mm,

respectively, for T1-WI. The DTI were processed using DtiStudio (www.MRIStudio.org)

and the mean diffusion-weighted image (DWI) was calculated.

To define the boundary(s) of the acute stroke lesion(s) for each participant, a

threshold of >30% intensity increase from the unaffected area in the first-time-point

diffusion-weighted image (DWI) extracted from DTI images was applied. A

neuroradiologist (AVF), blind to the patients' clinical information, manually modified the

boundary to avoid false-positive and false-negative areas on RoiEditor

(www.MRIstudio.org). The definitions were double checked by a second rater (MB). The

averaged lesion distribution map across all patients in the current study is shown in Fig.

5A. For the seven patients who had no DTI in the acute phase, lesion definition was

performed on the clinical DWI, which has lower resolution (1x1mm in plane, 4-6mm

thickness). Analysis of white matter ROIs, including the CST, was not performed in

patients.

Region of interest definition and lesion quantification. The focus was on two

ROIs: 1) The cortical gray matter of the hand area in the motor cortex; 2) The entire CST

superior to pyramids, identified by probabilistic maps derived from tract tracing methods

(see below). The percentage volume affected in these regions was correlated with our

main outcome measures, the Strength and Individuation Indices.

To defined the CST, each image and respective lesion were mapped to a single

subject adult template, the JHU-MNI atlas (Mori *et al.*, 2008; Oishi *et al.*, 2008, 2009,

2010), using affine transformation followed by dual channel (both b0 and FA maps) large

deformation diffeomorphic metric mapping (LDDMM) (Ceritoglu *et al.*, 2009). This

template has already been segmented into more than 200 regions of interest (ROIs), and

contains probabilistic maps of multiple tracts, including the CST (Zhang *et al.*, 2010). To

ensure accurate mapping, we first used "artificial" images, in which the stroke area was

masked out and substituted by the normal images from the contralateral hemisphere. This

helped to minimize inaccuracies caused by the focal changes in intensity due to the

stroke. The white matter beneath the cortex was identified with a FA-threshold of 0.25.

The segmentation defined in the template, as well as a probabilistic map of the CST, were

then "back-warped" by each subject's deformation field to generate individualized

parcellations.

A different approach was used to define an ROI that would encompass the hand

area of the primary motor cortex. The hand ROI was defined on the average

reconstruction of the cortical surface available in the Freesurfer software (Dale *et al.*,

1999), selecting Brodmann area (BA) 4 based on cytoarchitectonic maps (Fischl *et al.*,

2008). To restrict the ROI to the area of motor cortex involved in the control of the upper

limb, we only included the area 2.5 cm dorsal and ventral of the hand knob (Yousry *et*

al., 1997). The defined ROI was then morphed into MNI space using the surfaces of the

age-matched controls. These ROIs were then brought to the JHU-MNI atlas (in which

each subject and respective stroke area were already mapped, as mentioned above) using

T1-based LDDMM to construct a probabilistic map of the hand area. The probabilistic

map was threshold of 70% to calculate percent-volume affected.

Clinical assessments

At each visit, all participants were also assessed with several clinical outcome

measures. Here we report data for FMA and ARAT. Grip strength was assessed with a

Jamar hydraulic hand dynamometer (Sammons Preston, Rolyan, Bolingbrook, IL, USA).

Strength in the first dorsal interosseous (FDI) and the flexor carpi radialis (FCR) muscles

was assessed using a hand-held dynamometer (Hoggan MiroFET2 Muscle Tester, Model

7477, Pro Med Products, Atlanta, GA, USA).

Results

A total of 54 patients with acute stroke and 14 healthy controls underwent five

testing sessions over a one-year period. Data in the final analysis comprised a total of 251

sessions tested in 53 patients (one patient only completed two blocks of the task, and was

thus removed from further analysis) and 14 controls. Forty-one patients and twelve

controls completed >= 3 sessions. The data were 75% complete, with 25% of the sessions

were missing or unusable. Non-tested sessions were treated as data missing at random

and all available data were used in the statistical analysis (see Materials and Methods).

The Strength and Individuation Indices were reliable.

Finger strength was assessed by measuring the maximum voluntary force (MVF)

for each finger separately and then averaged across all fingers for each hand. MVF for

healthy controls had an average value of 20.35 N (SD = 8.56) for the dominant hand, and

22.76 N (SD = 6.89) for the non-dominant hand. The normalized Strength Index for the

controls' dominant hand was 1.00 (SD = 0.19), and non-dominant hand was 1.17 (SD =

0.25). For patients, the mean for the non-paretic hand was 0.93 (SD = 0.20), and for the

paretic hand it was 0.59 (SD = 0.38). For the paretic hand, Strength Indices did not

correlate with age ($r = 0.04$, $p = 0.75$), nor were they affected by gender ($t(51) = 0.98$, $p =$

0.33) or handedness ($t(51) = 0.10$, $p = 0.92$).

To assess individuation, we measured the amount of involuntary force changes

(enslaving) on the passive fingers for different levels of force production with the active

fingers. The amount of enslaving systematically increased at higher force levels (Fig.

1E). Loss of control at increasing force levels has been shown for the angular position of

the fingers (Li *et al.*, 1998) and the reaching radius of the arm after stroke (Sukal *et al.*,

2007; Ellis *et al.*, 2009). To control for this relationship, we characterized the

Individuation Index as the slope of the function between active force and passive

enslaving. Lower values of Individuation Index indicate more impaired individuation.

Healthy, age-matched controls showed, on average, a normalized Individuation Index of

1.00 (SD = 0.18). This refers to a slope of 0.087 (SD = 0.046), meaning that for a finger

press of 10N the mean deviation of the passive fingers was 0.69N. As was the case for

Strength, Individuation Indices in the paretic hand were not correlated with age ($r = 0.16$,

$p = 0.26$), nor affected by gender ($t(51) = 0.17$, $p = 0.86$) or handedness ($t(51) = 0.34$, $p =$

0.74).

When introducing a new instrument, it is important to first establish its reliability,

i.e., the accuracy with which true differences between subjects and changes within

subject can be determined. We therefore split the data for each session in half, calculated

Strength and Individuation Indices on these two independent data sets, and correlated the

resultant scores across patients and sessions (see methods). The adjusted split-half

reliability across all patients and weeks for the Strength Index was $r_{full} = 0.99$ and 0.94

for the paretic and non-paretic hands respectively, and $r_{full} = 0.89$ for controls, which

indicates good reliability. The adjusted split-half reliability of the Individuation Index of

all patients was $r_{full} = 0.99$ and 0.93 for the paretic and non-paretic hands respectively,

and for controls was $r_{full} = 0.97$.

Consistent with our effort to construct an individuation measure that is

independent of strength, the overall correlation between Individuation and Strength in

controls was very low for both controls (r = -0.19, p = 0.51), and for patients' non-paretic

hand (r = 0.17, p = 0.21).

The Strength and Individuation indices correlated with standard clinical measures.

The Strength and Individuation Indices were compared with existing clinical measures:

the Fugl-Meyer (a measure of impairment) and ARAT (a measure of activity) Table 2

shows the correlations for all four measures obtained from the paretic hand across all

time points. Overall, all correlations were very high (max p = 1.21×10^{-26}), indicating that

all the measures could detect severity of the hand function deficit. The correlation in the

patients between the two clinical measures was 0.91, whereas the correlation between the

Strength and Individuation Indices was 0.73, a significant difference (z = 5.62, p =

2.0×10^{-8}, using z-test with N = 180 (Fisher, 1921)). Given comparable reliabilities for all

measures, this difference unlikely results from measurement noise – rather it suggests

that our Strength and Individuation Indices measure two different aspects of the hand

function, whereas the clinical scales tend to capture a mixture of strength and control.

Recovery of strength and individuation occurred mainly in the first three months after

stroke.

426   We first examined the time courses of recovery for strength and individuation in

the paretic hand. If the two observed variables change in parallel, their recovery may or

may not be mediated by the same underlying process. A difference in the time courses,

however, would provide a strong hint of separate recovery processes for strength and

individuation.

For both measures, most of the recovery appeared to occur within the first 12

weeks after stroke (Fig. 2A-B). A model with a fixed effect of Week and a random effect

of Subject was built to evaluate this statistically. An effect of Week was tested with a

likelihood ratio test against the null model with the random effect only. Results indicate

that both the Strength and Individuation Indices significantly improved over time

(Strength: $\chi^2 = 47.65, p = 5.10 \times 10^{-12}$; Individuation: $\chi^2 = 18.58, p = 1.63 \times 10^{-5}$). Paired t-

tests between adjacent time points showed significant improvement (after Bonferroni

correction) of the Strength Index up to week 12; whereas the Individuation Index only

showed a significant improvement between weeks 4-12 (see detailed statistics in Fig. 2A-

B). A similar recovery curve was found for the standard clinical measures of motor

function (detailed statistics in Fig. 3).

To directly compare the time courses of between the two indices at the early stage

of recovery, we z-normalized scores of the two variables and then investigated the change

in the scores for the time intervals W1-4 vs. W4-12 (Fig. 2C). This analysis suggests that

strength may recover mostly in the first four weeks, while individuation recovery may

occur equally in both time periods. Repeated-measures ANOVA over z-scores for

Strength and Individuation Indices during the two time intervals yielded a significant

interaction ($F(1,25) = 6.82, p = 0.015$, Fig. 2C). Thus, despite overall similarity, there

was a significant difference in the time courses of recovery of strength and individuation,

with strength showing faster early recovery.

That most improvement in both strength and individuation occurred over the first

12 weeks is also apparent in the correlations between adjacent testing time points for each

variable across individuals (Fig. 2D). The correlation between weeks 1 and 4 for the

Individuation Index was significantly lower than it was for subsequent time points (W1-4

vs. W24-52: $z = -4.23$, $p = 0.000023$), and this difference for Strength Index was

marginally significant ($z = -1.83$, $p = 0.067$), using z-test with N = 28 and 33 (Fisher,

1921). Thus, the position of the patients on the mean recovery curve changed more

during the first 4 weeks than in the last 6 months. This correlation difference cannot be

attributed to measurement noise, as both measures had stable reliabilities at all time

points (dashed line). Instead, the lack of stability of these measures during early recovery

is indicative of meaningful biological change.

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463   Insert Figure 2

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Consistent with previous findings (Noskin *et al.*, 2008), the non-paretic hand also

showed mild impairment in the first month after stroke. A likelihood ratio test of the

mixed-effect model showed a significant effect of Week for Strength ($\chi^2 = 7.86$, $p =$

0.0051), and a more subtle effect for Individuation ($\chi^2 = 4.12$, $p = 0.042$) (Fig. 2A-B).

This increase in performance is unlikely to be related to a general practice effect, because

the Strength Index in healthy controls decreased slightly over time ($\chi^2 = 4.54$, $p = 0.033$),

perhaps due to reduced effort, whereas the Individuation Index for healthy controls was

maintained at a similar level over the whole year ($\chi^2 = 0.33$, $p = 0.56$).

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474   Insert Figure 3

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In summary, most recovery of both strength and individuation occurred in the first

three months after stroke, with stabilization of recovery around 3-6 months. The data also

suggest a slight difference in the time course, with strength recovering faster than

individuation in the first month.

Evidence for a time-invariant relationship between strength and control.

The time course analysis only provides weak evidence for partial independence of

the recovery processes for strength and individuation. Therefore we undertook a closer

examination of the relationship between the two variables at each testing time-point (Fig.

4A). Although patients tended to move from the lower left corner to the upper right

corner of this space over the time course of recovery, the overall shape of the strength-

individuation impairment relationship seemed to be remarkably preserved across weeks.

At lower strength levels, there was a clear correlation between strength and individuation;

whereas once above ~60% of normal strength level, the two variables were unrelated,

producing a distinct curvilinear shape for the overall function (Fig. 4B).

To formally test the time invariance suggested by visual inspection, we first found

a function to describe the strength-individuation relationship. We used data from all time

points and evaluated the goodness of fit of a piecewise function with two linear segments

connected at an inflection point, using leave-one-out cross-validation (see Materials and

Methods). Cross-validation automatically penalizes models that are too complex. This

functional form gave us a good fit to the data (cross-validated $R^2 = 0.53$, Fig. 4B). We

also explored first- to fourth- order polynomial functions. All four models resulted in a

worse fit (cross-validated $R^2 < 0.49$) than the piece-wise linear function.

We then tested for "time-invariance" of this strength-control relationship, that is,

whether the function shape changed across weeks. Again, using leave-one-out cross-

validation, the time-invariant model with fixed parameters across all weeks was

compared with a model that allowed the parameters to change for each week (time-

varying model). The cross-validated R^2 for the time-varying two-segment piecewise

linear function was 0.45, a worse fit than the time-invariant model.

These results suggest that there is a time-invariant recovery relationship between

strength and individuation after stroke, which consists of two parts: up to a certain level

of strength (60.7% of non-paretic hand), the Strength and Individuation Indices are

strongly correlated ($r = 0.74, p = 6.61 \times 10^{-18}$); after strength exceeds this threshold, the

two variables are no longer correlated ($r = -0.17, p = 0.11$; Fig 4B). This lack of

correlation cannot be attributed to a ceiling effect for the Individuation Index, because for

both patients and controls there was still a considerable variability, and the reliability of

Individuation Index was very high. This indicates that our measure has enough dynamic

range and sensitivity to detect inter-individual differences even in the healthy population.

514   -----------------------------

515   Insert Figure 4

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Overall, our results suggest that recovery can be captured as traversal along a

time-invariant function relating strength and individuation. Differences in recovery arise

because patients vary substantially in the distance they move along this function: some

patients with initial severe impairment made a good recovery, moving past the inflection

point of 60.7% strength (exemplified by the yellow dot in Fig. 4A). Other severely

impaired patients failed to reach the inflection point (red dot in Fig. 4A). Finally, some

mildly impaired patients started off beyond the inflection point and showed a good range

of individuation capacity.

A second process contributed additional recovery of finger individuation.

The fact that recovery of both strength and individuation could be captured by a

single time-invariant function that relates them is compatible with the hypothesis of a

single underlying process that drives recovery of both aspects of hand function. It is

possible, however, that an additional process injects further recovery, which determines a

patient's position relative to the mean recovery function in strength-individuation space.

If such a process exists, a given patient should occupy a consistent position above or

below the mean recovery function across time points.

To test this hypothesis we investigated the residuals of the Individuation Index for

each patient at each time point after subtracting out the mean two-segment piecewise-

linear recovery function. If the variability around this mean function were purely due to

noise, we should observe no consistent week-by-week correlation between residuals for

each patient. Alternatively, if the residuals were to be correlated across weeks, it would

indicate that some patients were consistently better at individuation than that predicted

from the function, and others were consistently worse, suggesting an additional factor

mediating individuation recovery (black arrows in Fig. 6).

Correlations of residuals from adjacent time points across patients were initially quite

low. However, from week 4 onwards, most patients' distances from the mean function

remained stable (Fig. 4C-D). This consistent structure in residuals provides evidence for

an extra factor contributing to recovery of individuation. The consistent pattern of

residuals at later time points could not be attributed to pre-morbid inter-individual

differences, because both the Strength and Individuation Indices were normalized to the

non-paretic hand. The low week-by-week correlations between early time points argues

that the later correlations do not simply reflect sparing of a particular neural system after

the stroke. If this had been true, the correlation between the Individuation residuals

should have remained constant across all time points. Furthermore, the lower early

correlation cannot be attributed to measurement noise, as reliabilities for the early

measurement points were high (Fig. 4D). Rather, the initially low but then increasing

correlation indicates an additional recovery process operating above the lower bound of

the strength-individuation function (Fig. 6). This process is mostly active in the first three

months after stroke and determines how well individuation recovers above that expected

from the time-invariant recovery function.

Lesions involving motor cortex and the corticospinal tract correlated more with

individuation than strength.

To investigate the underlying neural substrates of recovery processes, we

correlated the location and size of the lesion with the Strength and Individuation Indices.

We were especially interested in the particular role of the corticospinal tract (CST).

While both corticospinal and corticoreticular projections originate in part from the

precentral gyrus and are intermingled to some degree, cortical projections to the reticular

formation have a more widespread bilateral origin from other pre-motor areas (Keizer

and Kuypers, 1989), whereas direct corticospinal projections to ventral horn neurons

primarily arise from the anterior bank of the precentral gyrus/central sulcus, i.e. "new

M1" (Rathelot and Strick, 2009; Witham *et al.*, 2016). We therefore predicted that extent

of the damage to the hand area of the primary motor cortex, and to the white matter ROI

that characterizes the most likely course of the CST (see Materials and Methods) would

correlate more with Individuation, and less with Strength. Furthermore, lesions in these

areas should correlate with individuation recovery over and above the level expected

from the mean recovery function.

As hypothesized, the extent of involvement by the lesion of the cortical hand area

correlated significantly with the Individuation Index at all time points. For the CST, all

correlations were significant after week 1 (Fig. 5B-D). While both lesion measures also

correlated with the Strength Index, these correlations were weaker (repeated-measures

ANOVA showed a significant main effect for behavioral measure ($F(1,3) = 146$, $p =$

0.001). This difference was not due to measurement noise, as the Strength and

Individuation Indices had comparable reliabilities. Furthermore, percent lesion

involvement also significantly correlated with the Individuation Index, after accounting

for the average Strength-Individuation relationship ($p < 0.05$ for correlations after week

24 for cortical hand area, and after week 12 for CST). Indeed, at W52, the correlations

with the residuals were as high as with the Individuation Indices themselves ($r = 0.61$ vs.

$r = 0.57$ for the cortical hand area, $r = 0.51$ vs. $r = 0.54$ for the CST). Together these

results suggest that Individuation recovery is most heavily determined by the sparing in

the hand area of the primary motor cortex and of direct CST projections, while strength

recovery may also depend on other spared descending pathways.

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591   Insert Figure 5

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Discussion

In a large-scale longitudinal study, we tracked recovery of two independent

behavioral components of hand function: strength and finger control. Patients were tested

at five time points over a one-year period after stroke, using a novel paradigm that

separately measures maximum voluntary contraction force (a measure of strength) and

finger individuation ability (a measure of control), and crucially controls for any

obligatory dependency between these two measures. This approach allowed us to

determine how recovery of strength and control interrelate. Our main question was to ask

whether there is a causal relationship between strength and control at the level of

recovery mechanisms, after the two variables had been experimentally uncoupled. If they

are truly dissociable, then hypothetically patients could show perfect control of individual

fingers, even with significant weakness (except for complete hemiplegia, in which case

no individuation measure would be obtainable).

We showed that involuntary movements in passive fingers (enslaving) increased

with the level of force production of the active finger. This phenomenon is analogous to

what Dewald and colleagues (Sukal *et al.*, 2007; Ellis *et al.*, 2009) have described for the

paretic arm: a decrease in arm reaching workspace as the force requirement to resist

gravity increases. We showed that the ratio of enslaving and active force can account for

this dependency and thereby provides a sensitive measure of finger control independent

of the level of force deficit.

We first examined the time courses of recovery for strength and individuation.

Consistent with what has been described with traditional clinical scales (Duncan *et al.*,

1992; Jørgensen *et al.*, 1995; Krakauer *et al.*, 2012), both measures showed most

recovery over the first three months after stroke. This similarity between the time courses

does not, however, necessarily imply that recovery of strength and individuation is

dependent on a single underlying neural substrate or mechanism. It remains possible that

recovery of these two components occurs in parallel because of commonalities in basic

tissue repair mechanisms post-ischemia but they are nevertheless independent modules.

Indeed, we found a small but robust difference in the time course of recovery of strength

compared to control: finger strength showed a faster rate of change compared to

individuation over the first month. This finding raises the interesting possibility that

different neurological substrates underlie recovery of strength and individuation.

Closer examination of the two variables revealed a time-invariant non-linear

relationship between strength and individuation in the paretic hand. This function has two

distinct parts: individuation and strength were highly correlated below a strength

threshold of ~60% of the non-paretic side; beyond this point, they were uncorrelated. The

shape of this function remained the same across all time points. Recovery of hand

function could be characterized as movement along this invariant function: patients with

good recovery traveled further along the function, whereas patients with poor recovery

remained in the first segment. The strong correlation between strength and individuation

for severely impaired patients is consistent with a single system mediating recovery of

both. Indeed, in our cohort there was no patient with relatively good strength but severe

impairment of individuation, which also suggests that recovery of finger control

correlates with recovery of strength in patients with severe hemiparesis. However, two

pieces of behavioral evidence suggest that strength and finger control might rely on

partially separate mechanisms of recovery. First, a correlation between strength and

individuation was absent for the subset of well-recovered patients – i.e. patients with a

Strength Index above 60%. This breakdown in correlation cannot be attributed to a

ceiling effect for Individuation. Secondly, analysis of the residuals around the mean

recovery function revealed that patients differed consistently in the amount of their

individuation recovery relative to the level predicted by their strength recovery. Notably,

their positioning relative to the mean recovery curve seemed to be set early in the

recovery process and then remained relatively stable at later time points.

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648   Insert Figure 6

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Thus we propose that recovery of strength and individuation relies on at least

partially separate systems. One system primarily contributes strength, but also has some

limited control capacity. The isolated contribution of this system would determine the

lower bound of the data points in the strength-individuation plot (dashed line in Fig. 6):

when a patient regains some strength, he or she automatically regains a limited amount of

control with it. However, the amount of individuation is limited and does not increase

above a certain level. This would explain both the strong correlation between strength

and individuation for the severely impaired patients, and the fact that no patient occupied

the lower right corner of strength-individuation space, i.e. no patients had good strength

but minimal control.

The second system would then add additional control capacities to the first system

(vertical arrows in Fig. 6). Patients with a strong contribution from this second system

may gain full recovery of individuation; patients with no or only partial contribution from

the second system may recover completely in strength, but not in individuation.

Importantly the recovery of this second system also occurs early after stroke,

subsequently a patient's relative position above or below the mean recovery function

remains relatively fixed (Fig. 4D).

The lesion analysis adds support to the two-systems model for recovery suggested

by the behavioral data. A wealth of evidence in humans and non-human primates

implicates the role of CST in finger control, especially the monosynaptic cortico-

motoneuronal connections originating from "new" M1 (Lawrence and Kuypers, 1968;

Porter and Lemon, 1993; Rathelot and Strick, 2009). Notably, these connections do not

generate high levels of force but rather finely graded forces riding on top of larger forces

(Maier *et al.*, 1993). Consistent with this idea, lesions in the gray matter of the hand areas

in M1- the main origin of corticospinal projections- as well as the CST, correlated more

with impaired individuation than with strength.

In contrast, finger strength may rely on other neural pathways, including the

reticulospinal tract (RST), which can support strength and gross movements (Buford and

Davidson, 2004; Davidson and Buford, 2004). Although the RST has been found to

participate in some degree of finger control, its functional range is limited and biased

towards flexor muscles (Riddle *et al.*, 2009; Baker, 2011).

Recovery after stroke is likely to result from the dynamic interplay between the

CST and other descending pathways, particularly the RST. In this scenario, the

correlation between strength and control at low levels of strength may represent the state

of both the residual CST and of cortical projections to reticular nuclei in the brainstem. In

this framework, recovery along the lower bound of the invariant function would represent

the contribution of the RST and other non-CST descending pathways. Those patients

with less damage to the CST would consistently ride above this function.

The dichotomy proposed here may be too simplistic. While the origin of the

corticoreticular inputs is more diffuse (Keizer and Kuypers, 1989) and bi-laterally

organized (Buford and Davidson, 2004; Sakai *et al.*, 2009; Soteropoulos *et al.*, 2012),

many of the projections to the reticular formation arise from the primary motor cortex

(Catsman-Berrevoets and Kuypers, 1976; Jones and Wise, 1977). Thus, our lesion ROIs

will have included the corticoreticular tract to some degree, possibly explaining the

lower, but nevertheless significant correlation with strength. Furthermore, it is very likely

that direct corticospinal projections contribute to hand strength to some degree.

Interestingly, there was a small degree of impairment, especially in strength, in

the hand ipsilesional to the stroke. This finding confirms previous reports of deficits in

the non-paretic hand using clinical scales, e.g. muscle weakness measured by

dynamometry (Colebatch and Gandevia, 1989), and dexterity measured with the Nine

Hole Peg Test (9HPT) in (Noskin *et al.*, 2008). This ipsilesional impairment is consistent

with positing a strength role for the RST because it projects bilaterally.

A limitation of the current study is that the paradigm is designed to assess

weakness and enslaving in finger flexors, but not extensors. Because finger extensors

play an important role in finger individuation, and have been particularly associated with

the CST, it is possible that individuation in the extensors would also be more CST-

dependent and the dual systems we are implying for the flexor might not apply to the

same degree.

Conclusions

Here we found that hand function after stroke can be partitioned into strength and

strength-independent control. Most recovery of both these components occurred in the

first three months after stroke, although strength continued to improve for up to six

months. At any time point after stroke, strength and strength-independent control were

related by an invariant curvilinear function: strength and some degree of control are

correlated up to a certain strength level and then control saturates; some subjects showed

additional improvement in individuation riding on top of the main recovery function. The

results suggest that hand recovery is supported by two separable systems: one that mainly

contributes to the generation of large forces, as in the power grip, and another that is

responsible for more precise control of the digits at all levels of force. This behavioral

and imaging evidence for two systems contributing to recovered hand function after

stroke is consistent with the known characteristics of the CST and RST.

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Table 1. *Patient characteristics: age* (*years*)*, sex, paretic side, initial FMA* (*Fugl-Meyer*

arm score, maximum 66)*, initial MoCA* (*Montreal Cognitive Assessment, maximum 30*)*.*

Table 2. *Correlation between Strength Index, Individuation Index, FMA* (*Fugl-Meyer*

arm score, maximum 66)*, and ARAT* (*Action Reach Arm Test, maximum 57*)*. All four*

measures are highly correlated; however Strength and Individuation Indices show the

weaker correlation compared to that between FMA and ARAT.

Patient	Age at stroke	Gender	Paretic Side	Initial impairment (FMA)	Initial MoCA
1	57	M	R	48	27
2	24	M	L	35	23
3	67	F	R	16	23
4	74	F	R	39	17
5	61	F	L	48	26
6	59	F	R	60	28
7	57	M	R	54	27
8	66	M	L	65	25
9	42	F	R	5	18
10	65	M	L	30	25
11	66	F	L	60	19
12	51	M	L	34	25
13	63	F	L	57	26
14	55	M	L	0	26
15	56	M	L	38	25
16	56	M	L	64	24
17	64	F	R	20	16
18	60	F	R	55	21
19	64	M	L	63	25
20	25	F	L	42	29
21	39	F	L	47	20
22	46	M	L	9	27
23	53	F	L	4	29
24	66	M	L	59	24
25	71	M	L	4	26
26	52	M	L	53	24
27	46	M	R	4	21
28	46	M	L	49	30
29	71	M	L	6	24
30	47	M	R	57	10
31	45	M	L	8	27
32	55	F	L	19	25
33	68	F	L	61	NaN
34	65	M	L	32	28
35	51	F	L	63	26
36	42	M	R	54	25
37	58	M	L	4	24
38	41	F	L	4	23
39	35	M	L	4	29
40	68	M	L	52	27
41	76	M	L	53	18

42	86	M	L	54	20
43	48	M	L	16	25
44	74	M	R	5	25
45	80	F	R	9	24
46	64	F	L	58	19
47	22	M	R	63	27
48	88	F	R	55	28
49	22	M	R	63	27
50	87	F	R	50	28
51	84	M	R	30	26
52	53	M	R	30	29
53	54	M	L	59	21
54	58	M	R	61	23

	Strength Index	Individuation Index	FMA	ARAT
Strength Index		0.73	0.76	0.74
Individuation Index			0.68	0.72
FMA				0.91

Figure 1. *Strength and Individuation task.* ***(A)*** *Ergonomic hand device. The participant's*

fingers are securely placed on the keys using Velcro straps. ***(B)*** *Computer screen*

showing the instructional stimulus, which indicates both which finger to press and how

much force to produce (height of the green bar). In the MVF task, maximal force was

required; in the Individuation task a specific force level had to be reached. ***(C***, ***D)***

Example trials from two healthy control participants during the Individuation task. Four

trials are shown, one at 20% and one at 80% of MVF for the two participants. In this

case the fourth finger (red) was the active finger. Note the higher level of enslaving of the

passive fingers for higher active force level. ***(E)*** *Mean deviation from baseline in the*

passive fingers plotted against the force generated by the active finger for (C) and (D).

Increased enslaving with higher active force levels is clearly visible. The Individuation

Index is the -log(slope) of the regression line between active force and passive mean

deviation, measured as root mean square (RMS) force from baseline force produced by

passive fingers.

Figure 2. *Temporal profiles of recovery for strength and individuation.* ***(A-B)*** *Group*

recovery curves for the Strength and Individuation Indices for patients and controls.

Asterisks indicate significant week-to-week change for the paretic hand (Bonferroni

corrected p-values for each segments of Strength Index: p(W1-4) = 0.0045, p(W4-12) =

0.0082, p(W12-24) = 0.068, p(W24-52) = 0.87; Individuation Index: : p(W1-4) = 0.81,

p(W4-12) = 0.0024, p(W12-24) = 1.92, p(W24-52) = 2.91). ***(C)*** *Rate of change (i.e.,*

change per week) in Z-normalized Strength and Individuation Indices during the first two

intervals (Week 1 to 4 and Week 4 to 12). The two intervals show a significant interaction

between strength and individuation, indicating faster initial improvement of strength; (***D***)

Week-to-week correlations between adjacent time points for the Strength and

Individuation Indices. Dashed lines are the noise ceilings based on the within-session

split-half reliabilities.

Figure 3. *Temporal recovery profiles measures with clinical assessments. (A) Fugl-*

Meyer for the arm (FMA) and (B) hand (FMH); (C) ARAT; (D-F) strength for hand grip,

FDI, and FCR muscles, as measured by Dynamometry. All measures showed significant

change over time for the paretic hand. FMA: $\chi^2 = 37.73$*, p =* 8.13×10^{-10}*; FMH:* $\chi^2 =$

29.03, p = 7.14×10^{-8}*; ARAT:* $\chi^2 = 36.33$*, p =* 1.67×10^{-9}*; grip:* $\chi^2 = 33.02$*, p =* 9.21×10^{-9}*;*

FDI: $\chi^2 = 19.21$*, p =* 1.67×10^{-5}*; FCR:* $\chi^2 = 28.47$*, p =* 9.50×10^{-8}*.*

Figure 4. *Time-invariant impairment function relating strength and control.* (***A***) *Scatter*

plots for Individuation against Strength Indices at each time point. Each black dot is one

patient's data; blue dots and ellipse indicates the mean and standard error for controls at

the time point. Two patients' data are highlighted: one with good recovery (yellow dot)

and one with poor recovery (red dot). (***B***) *Scatter plot with data from all time points*

superimposed with the best fitting two-segment piecewise-linear function with one

inflection point at Strength Index = 0.607. (***C***) *Residuals from each week subtracting out*

the mean impairment function (B, red line). The tendency of the residuals to stay above

or below the typical Strength-Individuation relationship indicates that there are stable

factors that determine Individuation recovery over and above strength recovery. (***D***)

Correlations of residuals from (C) across adjacent time points increased over time

(Bonferroni corrected p(W1-4) = 2.12, p(W4-12) = 0.00064, p(W12-24) = 0.0024,

$p(W24-52) = p = 3.39 \times 10^{-6}$ *). Dashed line is the noise ceilings based on the within-*

session split-half reliabilities.

Figure 5. *Lesion distribution and correlation with behavior. **(A)** Averaged lesion*

distribution mapped to JHU-MNI space (see Materials and Methods), with lesion flipped

*to one hemisphere. Color bar indicates patient count. **(B)** Correlation of behavior*

measures (Strength and Individuation Indices) at each time point with the percentage of

damaged cortical gray matter within the M1 hand area ROI, corticospinal tract (CST).

(D) Mean of week-by-week correlations between the two behavior measures and percent

lesion volume measures for the cortical gray matter hand area and CST ROI. Black

asterisks indicate significant correlations (tested against zero), and red asterisks indicate

a significant difference between the correlation for Strength and Individuation for each

week (p<0.005).

Figure 6. *A schematic diagram of the hypothesis of two recovery systems. The first*

system (basic strength recovery) underlies strength recovery and a restricted amount of

individuation recovery. This system therefore defines the lower bound (dashed line) of the

space occupied by recovering patients (gray clound). A second system (additional

inidividuation recovery) adds further individuation abilities on top of the basic strength

recovery.