Two stochastic processes shape diverse senescence patterns in a single-cell organism

Short title: Two stochastic processes shape senescence in E. coli

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Abstract

Despite advances in aging research, basic understanding is lacking of the biological processes that shape senescence. We show that for a isogenic *Escherichia coli* bacterial population senescence results from two stochastic processes. A primary random deterioration process within the cell, such as generated by random accumulation of damage, leads to an exponential increase in mortality early in life followed by a late age mortality plateau; a secondary process of stochastic asymmetric transmission of an unknown factor at cell fission influences mortality. This second process is required to explain the difference between the classical mortality plateaus detected for young mothers offspring and the near non-senescence of old mothers' offspring as well as the lack of a mother offspring correlation in age at death. We observed that life span is predominantly determined by underlying stochastic stage dynamics. Our findings support models based on stage-specific actions of alleles for the evolution of senescence.

Introduction

One of the major challenges for biodemographic research on aging is to understand what drives senescence patterns (Vaupel et al. 1998; López-Otín et al. 2013). This challenge is illustrated by the variety of aging models that assume different, as yet unverified generating processes (Kirkwood 2005; Hamilton 1966; Kenneth W Wachter, Steinsaltz, and Evans 2014), and also by an extensive literature on mechanistic approaches (López-Otín et al. 2013). These approaches identify a multitude of underlying biochemical, molecular and organismal mechanisms that relate to the decline of function with age, many of which are rooted in direct or indirect oxidative processes (López-Otín et al. 2013; Kirkwood 2005). Examples include age-related mitochondrial dysfunction, telomere shortening, stem cell exhaustion, genotypic instability, epigenetic alterations, accumulation of damaged proteins and general loss of proteostasis (López-Otín et al. 2013; Tyedmers, Mogk, and Bukau 2010; Lindner and Demarez 2009).

Yet, researchers have not conclusively determined whether such mechanisms are a cause or consequence

of aging. This failure may be due to the complexity of model systems of aging. As a consequence, it is

difficult to relate these mechanisms directly to the observed demographic patterns (Denoth Lippuner,

Julou, and Barral 2014; López-Otín et al. 2013; Tyedmers, Mogk, and Bukau 2010). Only such linkage

— between mechanisms and senescence patterns — can elucidate generating processes that underlie the

various theories and aspects of aging.

Here, we quantify demographic parameters of a simple biological system, isogenic individual *E. coli*

bacteria cells, under highly controlled environmental conditions. We used a high-throughput microfluidic

device (Fig. 1; movie S1) to track individual cells throughout their lives (Gasset-Rosa et al. 2014; Wang

et al. 2010). Two types of cells were tracked: early daughter cells and late daughter cells (Fig.1). A late

daughter cell is the last daughter cell produced by an early daughter and hence an offspring of an old

mother. An early daughter cell is the offspring of a young mother since they were haphazardly extracted

from a population that grew exponentially. According to stable stage population theories such

exponentially growing populations are vastly dominated by young cells (Caswell 2001)(Fig. S1). We

expect early daughters to hold little or no deterioration or damage at birth, whereas late daughters are

likely be born with some damage. In this study, we use damage as a synonym for any unknown aging

factor that leads to deterioration and increased mortality. We use this synonym because many aging

factors are assumed to be accumulated damage caused by oxidative processes (López-Otín et al. 2013;

Kirkwood et al. 2005). A third group of cells, resembling the late daughter type are the last daughter cells

produced by the late daughters, which we call second generation late daughters (Results only in SI1).

Fission of bacteria cells generally results in two equal sized progeny, but the resulting progeny are

phenotypically unequal (Tyedmers, Mogk, and Bukau 2010). This asymmetry among the progeny is

manifested by asymmetry of intracellular content at cell fission and between carrying old and newly

formed cellular poles (Lindner et al. 2008; Stewart et al. 2005). Here, we follow the convention of old

pole and new pole cells, being referred to as mother and daughter cell respectively (Fig. 1), to track the

age of individual cells (Stewart et al. 2005). It has been shown that mother cells have a higher

probability to accumulate misfolded protein and grow slower as compared to their daughter cells, but a

causal relationship to mortality was not established (Lindner et al. 2008). From a theoretical point of

view asymmetry is required to rejuvenate some cells in order to prevent whole population aging that

would occur at perfect symmetric fissions, when damage accumulation accedes damage repair and

dilution due to growth and fission (Ackermann et al. 2007; Evans and Steinsaltz 2007). In our

experiments we built on the concept of cell polarity for both early daughter and late daughter cells to

track the age of cells and distinguish between the mothers (old pole cell) and the daughters (new pole

cell) (Fig.1).

Fig. 1: Left panel: Overview sketch of the microfluidic device with dead-end side channels and the main

laminar flow channel where the media is flushed through. The early daughters (founding initial loaded

mothers) are the bottommost cells of the dead-end channels. Their daughters (new pole progenitor cells)

are located closer to the main laminar flow channel and have more recent poles. When the mother (early

daughter) dies (side channel on the right), its last daughter (1st generation late daughter) is then tracked

throughout its life. Accordingly, 2nd generation late daughter cells are tracked once their mother (1st

generation late daughter) dies. Right top and bottom panel: Phase contrast time sliced (4 min intervals)

side channel images with initial loaded early daughter at the bottom (old pole progenitor). Growth (cell

elongation) and division can be tracked throughout their life as depicted in the segmented cell lineages

(bottom right panel). See also movie S1.

Results and Discussion

Chronological aging in *E. coli*

In our experiments, we followed by automated time-lapse microscopy 516 early daughters throughout

their lives as they divide and age in microfluidic dead-end channels (Fig. 1). We also followed two

generations of late daughters of these 516 early daughters and recorded each cells growth rate and

lifespan. In this, we reveal classical senescence patterns of a decrease in reproduction (Fig. 2A, B) —

indicated by decreased cell elongation and increased size at division with age (Fig. S2, Table S1) — for

both early and late daughters (Figure 2A, B). The observed senescence patterns (Fig. 2C, D) describe

chronological aging in *E. coli* and support previous studies on replicative senescence (time counted as

number of divisions) in this species (Lindner et al. 2008; Wang et al. 2010) (Fig. S2 & Fig. S4). The

main result highlights that early daughters and late daughters differ fundamentally in their senescence

pattern, even though they are isogenic and grown in a highly controlled environment. Only the early

daughters exhibit the classical senescence pattern, marked by an early exponential increase of the probability of death followed by a later age mortality plateau. Late daughters have the same probability of death across most of their lives. i.e. no senescence is observed at the population level. Only late in life does mortality increase, but this increase is largely driven by only one data point, the one for the last age class (>30h). Such a plateau, exhibited by the early daughters — recurrent in many higher organisms including humans (Vaupel et al. 1998) — has not been previously shown for bacteria, potentially indicating deep-rooted features of aging and senescence. In yeast, ambiguous results on senescence patterns have been described (Denoth Lippuner, Julou, and Barral 2014; Minois et al. 2005).

Fig. 2: Division rate (divisions per hour) (A, B) and probability of death per hour (C, D) with increasing

age, as well as lifespan distribution (in hours) (E, F) of isogenic *E. coli* cells grown under highly

controlled environmental conditions in a microfluidic device (Fig. 1, movie S1). Age patterns are shown

for early daughters (A, C, E; N=516), and late daughters (B, D, F; N=516). Late daughters are the last

offspring of the founding early daughters. The correlation of early daughter's lifespan (mothers lifespan)

(square root transformed for better visibility; N=516) versus the lifespan of their last daughter cell (late

daughter cells; N=516) is shown in panel (G). For (A, B) hourly means ± standard errors are plotted, for

(C, D) hourly means ± 95% confidence intervals are plotted (confidence intervals are estimated based on

individual level data not the hourly means; hourly means are plotted as open circles, see SI1). Red lines

are models selected based on information criteria (AIC, Table S1, Table S2) for (A, B) or likelihood

optimization for Gamma-Gompertz-Makeham functions, the latter relate closely to the modeling

approach we took below (C, D, see also Fig. 3, simulations below, and SI1; Table S1, Table S2)

(Burnham and Anderson 1998). For lifespan distributions (E, F), mean and median are marked by solid

and dashed lines respectively. For patterns of the second generation late daughters, see Table S2, Table

S3, Table S4, and Fig. S3, Fig. S7, Fig. S8, Fig. S9, Fig. S10, Fig. S11.

The probability of death of newborn late daughters (<=1hour, filled grey data point in Fig. 2D) drops

after the first hour to a level that is lower than their long-lived mothers (early daughters above age 20)

and then remains at that level throughout their lives. This drop in mortality early in life of late daughters

suggests a damage purging effect (see below). Such a drop of mortality after birth might indicate

heterogeneity among newborns amenable to evolutionary selective forces (Yashin, Vaupel, and Iachine

1994), e.g. heterogeneity in maternally-transmitted damage between mothers and daughters. With

increasing age we detected increased variance and increased uncertainty of parameter estimation (Fig. 2

A-E). Such increase is expected for age-structured demographic analyses and well understood for its

declining number of individuals with age (Brillinger 1986; Scherbov and Ediev 2011; Klein and

Moeschberger 2003). The increase in variance is more pronounced for probabilities of death (Fig. 2 C,

D) than for division rates (Fig. 2 A, B), because each individual dies only once while each individual

divides multiple times throughout its life and hence division rates suffer less from smaller sample sizes.

Binning over larger time spans or having an earlier open age bracket would diminish the effect of

increased variance with age.

Damage accumulation and selection

If we assume, as in most physiological theories of aging, that accumulated damage is the determining

factor for the probability of death, then the observed constant probability of death with age for older early

daughters and the late daughters (Fig. 2C, D) indicates an equilibrium distribution of damage among

individuals. Such equilibrium is realized at the population level while the individuals accumulate damage

in a stochastic manner. It equilibrates on one hand the accumulation of new damage within individuals

similar to processes described by random deterioration models (Weitz and Fraser 2001), and on the other

hand, the intracellular repair of damage and purging of damage by two mechanisms. The first mechanism

reduces damage within an individual by asymmetric division of damage at cell fission, which increases

variance in damage among individuals (mother and daughters). The second mechanism selects against

highly damaged cells, i.e. the most damaged cells of a population die, which lowers the average level of

damage in the surviving population, and reduces the variance of damage among individuals (Evans and

Steinsaltz 2007). Intriguingly, this equilibrium (plateau level) is significantly lower for late daughters

(Fig. 2D, see Table S3) than the late-age mortality plateau of the early daughters (Fig. 2C). This indicates

— based on fixed frailty models (Missov and Vaupel 2015) — higher heterogeneity in transmitted

damage of late daughters compared to early daughters. Such increased heterogeneity at birth predicts an

earlier onset of selection against highly damaged cells which would explain why in our simulations

(below) late daughters have on average less damage compared to early daughters at the plateau (see Fig.

3A, B, Fig. S10). Purging of damage through asymmetric division at cell fission has been shown in yeast,

where late daughters showed reduced lifespan, but through further division and presumably damage

dilution, the descendants of these late daughter yeast cells recovered high longevity (Kennedy 1994).

This is also reminiscent of asymmetric partition of protein aggregates in bacterial cells due to their

passive accumulation in bacterial mother cells (Lindner et al. 2008; Coquel et al. 2013).

<u>Stochastic variability in lifespan and lack of mother-daughter correlation of lifespan</u>

We did not find any influence of the lifespan of the mothers on the lifespan of their (last) daughters (Fig.

2G) (Table S6, Fig. S8). Since we are working with an isoclonal population we do not expect any genetic

variability among mother-daughter pairs; therefore, we would not expect a genetically driven mother-

daughter correlation of lifespan caused by additive genetic differences. However, non-genetic factors

such as maternal transmitted damage would predict mother-daughter correlated lifespans except if i)

daughters would be perfectly rejuvenated — all daughters would be born without damage —, or ii) if the

transmission of damage to the daughter is independent of the amount of accumulated damage by the

mother — the mother-daughter transmission of damage is stochastic. Since late daughters display higher

probability of death from birth onwards, which is indicative of some maternally transmitted damage and

obviously shows that late daughters are not perfectly rejuvenated, we can exclude the first explanation

why a mother-daughter correlation of lifespan is lacking. This leaves the second explanation where the

amount of damage transmitted from the mother to the daughter is stochastic and independent of the age

of the mother. Early daughters seem to hold little or no damage at birth, since their mothers have not yet

accumulated much damage, whereas the average late daughter seems to be born with some damage that

raises its probability of death. The lack of a mother-daughter correlation of lifespan is also remarkable

as it contrasts heritability of lifespan in humans and other complex organisms that show similar

senescence patterns (Finch and Kirkwood 2000).

The lack of correlation between mother-daughter lifespan suggests that the assumed mother-daughter transmission of damage is predominantly stochastic. If damage accumulates stochastically within a cell over its life, long-lived mothers might have been lucky and accumulated damage at a lower rate than the average mother, but they still accumulate damage with increasing age. At age of death, long-lived mothers might have similar damage levels compared to short lived mothers. Short-lived mothers, which likely accumulated damage at an exceptional high rate, do not produce daughters with shorter life expectancy. Despite their difference in lifespan, neither long-lived nor short-lived mothers tend to transmit higher or lower amounts of more damage to their daughters, otherwise we would find a negative or positive correlation in mother-daughter lifespan (Fig. 2G). Therefore, our observation that early daughters — which are assumed to hold little or no damage at birth — die at very different ages indicates that the process of damage accumulation within an individual appears to be stochastic and varies highly among individuals. Despite an isogenic population in a highly controlled homogeneous environment, we observe a high variation in lifespan and reproduction among individuals, both for early daughters (Median±stdev lifespan 12 ± 7 hours, Coefficient of Variation, CV 0.57, Fig. 2E) (Fig. S7) and late daughters (Median±stdev 5 ± 7; CV=1.01, Fig. 2F). Such high variation supports substantial influences of stochastic events in shaping lifespan.

Modeling damage accumulation and asymmetric transmission of damage

To support our conclusions based on the empirical results and to gain a better understanding of the underlying stochastic processes that shape our observed senescence patterns, we extended random deterioration process models (e.g. random accumulation of damage, see details in Methods and SI1) (Le Bras 1976; Yashin, Vaupel, and Iachine 1994; Evans and Steinsaltz 2007). Our extended model assumes that all cells accumulate damage unidirectionally at the same rate, i.e. cells move with equal probability

to a higher damage state, mortality increases exponentially with accumulated damage, and each cell

suffers from an age-independent baseline mortality risk. We also assume that early daughters are born

without damage (Fig. 3A). The results of our simulations show that with increasing age population level

damage (and variance) slowly increases at early ages, accelerates after 10 hours and plateaus after 20

hours (Fig. 3A). These patterns in damage accumulation mirror the observed mortality patterns (Fig. 2C).

Mortality plateaus, as observed for early daughters, can be explained by either fixed frailty, i.e.

heterogeneity in damage stage at birth, or acquired heterogeneity throughout life (Weitz and Fraser 2001;

Vaupel and Yashin 1985; Missov and Vaupel 2015; Yashin, Vaupel, and Iachine 1994). In our system

heterogeneity in damage stage at birth seems to have little influence, at least for early daughter cells.

Senescence is driven by an age independent baseline mortality and accrued damage as in random

deterioration models (Le Bras 1976; Weitz and Fraser 2001; Yashin, Vaupel, and Iachine 1994). This

conclusion is supported by two results of our model. First, most early daughters die with no or very little

damage (Fig. S10), and second, the observed survival patterns follow closely the survival curve of

simulated cells that did not accumulate any damage throughout life (Fig. 3C).

Contrasting the assumption for the early daughter cells, in our simulation the average simulated late

daughter is born with some damage (~24% are born without damage) (Yashin, Vaupel, and Iachine

1994)(Fig. 3B). For the following results, it is important to note that this heterogeneity in damage at birth

of late daughters is an underestimation since we do not include the exceptionally high first-hour mortality

rate (grey filled data point in Fig. 2D) for the simulations. In our simulations the distribution of this

damage at birth of the late daughters is related to the distribution of the accumulated damage at death of

the early daughter cells (mothers of the late daughters, Fig. S10). For the simulations, we assumed that

the mother to daughter transmission of damage is a fixed fraction of all the damage the mother

accumulated over her lifespan. Note, we only simulated a single binomial division, rather than multiple

asymmetric fissions as others have done (Evans and Steinsaltz 2007). We computed the fixed fraction of

damage transmission between mother and daughter by optimizing the survival pattern predicted by the

simulations to the observed survival curves.

We found that the optimal (simulated) fixed fraction was at a low level of 0.07. This indicates a highly

asymmetric division of damage between mothers and daughters. If damage transmission between

mothers and daughters would be symmetrical (0.5), the predicted survival of the simulations (Fig. 3D

hatched line) does not match the observed survival (Fig. 3D red line). For early daughters, the level of

asymmetry makes little difference because most of them die with no or little accumulated damage (Fig.

3C, Fig. S10). The predicted survival pattern of late daughters at older ages slowly converges to the

simulated survival patterns of undamaged cells (outer right grey line) (Fig. 3D). Therefore, many cells

in our simulation that survive to old ages hold little damage, which supports our interpretation of our

experimental data that late daughter cells are born with diverse damage states, highly damaged cells are

selected against, and at old ages, population-level survival patterns are strongly influenced by cells that

have not accumulated much damage.

Similar selective effects have been described by population level demographic theories as

heterogeneity's ruses, where selection among different damage stages (fixed frailty; i.e. heterogeneity in

damage stage at birth) leads to diverse senescence patterns at the population level (Vaupel and Yashin

1985). Compared to our simple model that only considers a single binary fission (branching) and suggests

a strong asymmetry, more complex and realistic models that include multiple branching (fission) events

show less extreme asymmetry at cell fission to be adaptive (Evans and Steinsaltz 2007). This difference

in optimal asymmetry between our simple and the more complex models can be understood in that the

asymmetric transmission of damage is distributed over multiple fissions in the more complex models,

while for our model it is concentrated at a single event. Even though the optimized fixed transmission

fraction between mother and daughters of our model is low at 0.07, such a fixed fraction would lead to a

correlation in lifespan between mother and daughters of about 0.25 (Fig. 3E fitted line). The observed

data does not support such correlation in lifespan (Fig. 2G) and hence supports our conclusion that the

fraction of transmitted accumulated damage from mother to daughter at cell fission varies stochastically

and is not fixed at a low level of 0.07 as our model assumes.

Fig. 3: Random deterioration model: Simulation results are shown of mean (red solid line, + SD grey

hatched lines) and median (red hatched line) damage state with increasing age (A, B); observed

experimental population level survival curves (red solid line; see also related probability of death Fig. 2

C, D), Gamma-Gompertz-Makeham simulated survival curves with symmetric damage transmission

(hatched black line) as well as simulated asymmetric damage transmission of 7% (solid black lines ±

95% CI in blue) (C, D). Graphs are shown for early daughter cells (A, C) and late daughter cells (B, D).

Thin grey lines in (C, D) depict expected survival curve of simulated cells with different fixed damage

state: outermost (right) curve depicts the highest survival of simulated cells with no damage throughout

their lives, most left survival curve depicts the lowest survival of simulated cells that were born with

maximum damage level of 5000. Lifespan (square root transformed) of simulated 516 early daughter

cells (mothers) versus the lifespan of their simulated last daughter cell (late daughter cells) with 7%

mother to daughter damage transmission (E) (for technical reasons, binning and rounding, due to the use

of Poisson distributions in the simulations [see SI1], only cells that lived at least to age 5 are illustrated).

The blue line illustrates the correlation in lifespan between mothers and daughters (~0.25, with 95% CI

in grey), given a fixed 7% mother to daughter damage transmission at cell fission.

In this study, our experimental setup limits our analysis to two extreme cases: early and late daughter

cells (in fact the last daughter cells). We hypothesize that the results do not only hold for these two

extremes but rather portray the extremes of a continuous process across different aged cells. Two

observations support our claim: first, changes in mortality of early daughters (Fig. 2C) and age patterns

in reproduction (Fig. 2A) are gradual, suggesting a gradual underlying mechanism; second, we do not

observe a pronounced pattern just before death (Fig. S5 & Fig. S6, Table S5). Such a pattern would be

expected if the last daughters are exceptional because their mothers are approaching death, for instance

in the way as predicted under terminal investment theories (Clutton-Brock 1984).

Two stochastic processes shape diverse senescence patterns

We conclude that the diverse senescence patterns, including the classical senescence pattern of a

mortality plateau, are determined by two stochastic processes that relate to underlying (damage) states

and only indirectly to age (Weitz and Fraser 2001; Yashin, Vaupel, and Iachine 1994; K. W. Wachter

1999; Evans and Steinsaltz 2007). The primary process is a random deterioration process, e.g. the

stochastic accumulation of damage throughout life, and the secondary process involves the stochastic

transmission of damage from the mother to their daughters at cell fission. This transmitted damage or

some other unknown aging factor increases the probability of death but is "non-heritable" as we show

by the lack of correlation between mother and daughter lifespan (Weitz and Fraser 2001; Yashin, Vaupel,

and Iachine 1994; K. W. Wachter 1999; Evans and Steinsaltz 2007). Not only additive genetic variation

but also other commonly assumed drivers of senescence, such as epigenetic variability, age itself, or the

(extrinsic) environment can be ruled out as major players in our study. If epigenetic variability had a

significant influence, a positive correlation among mother and daughter lifespan would be seen. If

chronological age determined senescence, early and late daughter cells would show similar mortality

patterns, hence senescence is rather driven by stage dynamics than by chronological age. The highly

controlled microfluidic system creates a uniform environment and hence we can exclude extrinsic

environmental drivers; this does not imply that senescence patterns do not differ under different

environmental conditions, we just investigated only one specific environment. For instance, under

complex media variance in division size and related division rate is increased compared to the minimum

medium we used (Gangan and Athale 2017). Minimum medium also decreases the rate of filamentation,

a stress response that causes continued cell elongate without dividing. Recovery of filamentous cells is

rarely observed under minimum medium but frequent under complex medium (Wang et al. 2010) (see

also SI). The near identical conditions cells experience in the microfluidic system decreases

environmental variability compared to naturally experienced micro-habitats. The laminar flow provides

large amounts of fresh media that constantly diffuses within seconds into the dead-end side channels and

therefore prohibits built-up of micro-environmental-niches. Though, such micro-environments (e.g.

biofilms) are characteristic for the natural conditions in the guts of the mammalian hosts most *E. coli* life

and have evolved under. The microfluidic system we use largely prevents cell-to-cell contact and

chemical communication among cells that can influence population growth (Aoki et al. 2005). The highly

controlled conditions therefore lowers phenotypic variability compared to natural conditions. To what

degree senescence patterns are altered by the environment remains to be explored. Temperature seem to

conserve the shape of senescence patterns and simply scales patterns differently, whereas nutrient

concentrations and nutrient source might alter shapes as has been found in other organisms (Stroustrup

et al. 2016)(Steiner unpublished).

Our findings indicate that two simple stochastic processes can create complex senescence patterns. Easily

understood, simplistic, and plausible arguments behind evolutionary theories of aging might lure readers

into the misconception that chronological age in itself plays an important role in senescence. Such

simplistic assumptions of age-specific drivers must be approached with caution and seem not applicable

to this study. Selective forces likely decline with time in our system, but the decline might be dominated

by underlying stage dynamics. Selection on heterogeneity among individuals is rather generated by such

stage dynamics than by increasing chronological age. In this sense, selective forces and their decline vary

considerably among individuals, even though genetic load that relates to the accumulation of deleterious

mutations should be negligible in an isoclonal population as ours. The substantial effect of stochastic

events on life histories is also indicated by the large variability in life histories despite excluding any

genetic or environmental variability. Such large stochastic variation supports arguments behind neutral

theories of life history (Steiner and Tuljapurkar 2012) that suggest that stochastic variability can account

for substantial fractions of the total variability in fitness components. Our interpretation of stochastic

events determining individual life courses is consistent with findings of significant stochastic influences

at the molecular (Elowitz et al. 2002) and protein level (Balázsi, Van Oudenaarden, and Collins 2011;

Tyedmers, Mogk, and Bukau 2010) and adds to the growing interest in such phenotypic heterogeneity

and individuality (Ackermann 2015; Davidson and Surette 2008; Wolf et al. 2007).

Future challenges will include determining whether these stochastic processes are neutral or adaptive,

and what drives their evolution (Steinsaltz and Evans 2004; Kærn et al. 2005; Vera et al. 2016; Norman

et al. 2015). For the basic aging processes that drive senescence, the causal relationships that drive age

patterns are currently unknown (López-Otín et al. 2013; Lindner and Demarez 2009). The complex

senescence patterns found in this study of a simple model organism under highly controlled conditions

emphasizes the challenges to quantify contributions of well-defined determinants of aging in the complex

systems on which most aging research is focused (López-Otín et al. 2013). In light of the growing

evidence that stochastic processes can have cascading effects across all levels of higher organisms (Finch

and Kirkwood 2000; Elowitz et al. 2002; Balázsi, Van Oudenaarden, and Collins 2011) new avenues in

aging research may require a shift towards the underlying stochastic processes that drive such stage

dynamics in simple systems like bacteria and perhaps beyond. Identifying the underlying currently

unknown (damage) states remains another challenge that we believe requires a combined quantitative

demographic and mechanistic approach, because of the high level of stochastic influences. Promising

steps in such directions have been initiated by theoretically exploring stage-specific alleles shaping

senescence patterns (Kenneth W Wachter, Steinsaltz, and Evans 2014) and by exploring transcription

factor signal dynamics at the single cell level across increasing parts of the lifespans of many individuals

(Norman et al. 2013).

Methods

We collected data in two independent sets of experiments. We loaded *E. coli* K12-derived MG1655 strain

cells into the designed microfluidic (PDMS) chip (Wang et al. 2010; Gasset-Rosa et al. 2014) from an

exponentially growing culture in supplemented minimum media M9. During each experiment, we

acquired 77 hours of time lapse phase-contrast imaging (15 frames/hour for each of the 2x44 fields

followed) using a temperature-controlled (at 43°C) inverted microscope. We used 43°C to accelerate the

aging process and thereby shorten the time the system needed to run under stable conditions. Such

stability was particularly important to accurately track late daughter and second generation late daughter

cells. Increasing temperature up to 43 °C scales senescence patterns, but does not alter the shape of

patterns. Above 43 °C patterns got unstable and cells were not viable (results not presented). The rod-

shaped bacteria cells grew in dead end side channels of a microfluidic chip, with the focal cells trapped

at the end of the side channels, and we tracked them over their lifespan (Fig. 1; movie S1). We used

customized image analysis to generate the demographic data (lifespan, cell elongation rate, cell size, and

time of each division).

We assured by starting the experiments with exponentially growing cells that the initial loaded cells are

descendants of young mothers (i.e. they are early daughter cells) (Fig. S1). At the end of their lives, these

early daughter cells produced a last daughter (late daughters) that then became the next bottom-most cell

trapped at the end of the respective growth channel. Therefore, we could directly compare mother and

daughter cells. Note that the late daughters are not born at the same time (main text Fig. 2 E, F). The late

daughter cells produced another generation of late daughter cells at the end of their lives (second

generation late daughter cells) for which the results are shown in Fig. S2-8, Table S1-3. Statistical

analysis on the empirical data were done in program R (R Core Team 2016) using general linear,

generalized linear, and non-linear models. Models were selected based on information criteria (AIC)

(Burnham and Anderson 1998) or based on differential evolution algorithm for global optimization (R

package DEoptim).

For the extended random deterioration model, we first estimated parameters by fitting a fixed frailty

model — a Gamma-Gompertz-Makeham (GGM) model — to the observed survival data of the early

daughter cells (Missov and Vaupel 2015; Yashin, Vaupel, and Iachine 1994). We then translated these

fitted GGM model parameters to an extended random deterioration process model, an extended LeBras

type deterioration model (Le Bras 1976). In doing so, we took advantage of mathematical similarities

between the two types of models, even though they are biologically distinct (Yashin, Vaupel, and Iachine

1994; Missov and Vaupel 2015). With these random deterioration model parameters we could estimate

the probability matrix of an individual being at stage i at age x. Microsimulations based on these models

provided stage at death distribution. For the late daughters, we assumed that they are born at a scaled

version of the same damage stage in which their mothers died. We also assumed that late daughters

accumulate damage at the same rates as early daughters do, i.e. same probabilities to transition to a higher

damage state of early and late daughters. We further assumed that the amount of accumulated damage in

late daughters had the same effect on mortality than on early daughters, that is, early and late daughter

cells are not fundamentally different except that late daughters are born in different damage stages, while

early daughters start their lives without damage. Our model simplifies the biological system substantially

in as much as no repair of damage or purging of damage through asymmetric division at cell fission is

considered. Damage accumulates unidirectionally, mortality increases exponentially with accumulated

damage, and each cell suffers from an age-independent baseline mortality risk (Makeham term).

For more detailed methodological information on strains and growth conditions, time-lapse imaging,

image analysis, determining death and estimating demographic parameters, statistical analyses, the

simulations and more please see the supplemental information.

Acknowledgements

We thank D.A. Roach, D. Steinsaltz, T. Coulson, Y. Yang and S. Tuljapurkar for discussions and

comments.

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Supporting information SI1

Strains and growth conditions

We fabricated the microfluidics chips as previously described (Gasset-Rosa et al. 2014). We grew *E. coli* K12 MG1655 (Blattner 1997) strain derivative (intC::pSulA-yfp) cells at 43°C in filtered minimal medium supplemented with 0.4% Glucose and 0.2% Casamino Acids (hereafter M9). We diluted 100µl of the overnight culture into 50ml of M9 and then grew it at 43°C to exponential growth phase (OD600 ~0.2). We centrifuged the cells and resuspended the cell pellet in 100µl of M9 to load them into the chip by injection into the main (feeding) channel followed by centrifugation (15 minutes at ~168g). We applied a continuous laminar flow (2.7ml/h) of M9 supplemented with 1.5% Polyethylene Glycol (PEG P3015 Sigma Aldrich) through the main channel throughout the experiment. The temperature was kept constant at 43°C (see below).

Time-lapse imaging

We followed bacterial growth in the microfluidics chip by phase-contrast time-lapse imaging at a rate of 4min/frame using a MetaMorph (Molecular Devices)-controlled inverted Nikon microscope (Eclipse Ti, 100xobjective, CoolSNAP HQ2 camera) with a temperature-control chamber (Live Imaging Services). We continuously scanned 44 positions, each comprising 18 channels, for 77 hours, in two independent experimental sets.

Image analysis

We used a custom image analysis to segment all cells within the side channels per frame, the software measured the cells' location and size within the time series and generated cell lineages. To this end, phase

contrast time-lapse images were lowpass-filtered and background flattened (MetaMorph Molecular

Devices software) to increase contrast. We further processed the images by a customized ImageJ plugin

for cell segmentation to crop each of the growth channels, stabilize the time sequenced images, adjust

the brightness and to create a binary image by thresholding and filtering (median filter). Finally, the

ImageJ plugin water-sheds the time sliced images. We then applied a final round of segmentation in

Matlab by correcting errors in segmentation by checking for minimum (0.8) and maximum (1.4) cell

elongation rates for each cell (size time-4 min, size time+4min). This way, we recorded for any time

points (at 4 minute intervals) the exact location of a cell in the side channel (the pixel coordinate of the

center of mass of the cell area), the side channel of the cell, and if a cell divided. By the position within

the side channel we determined which cell was the mother (old pole progenitor) cell (bottommost cell),

their most recent daughter (new pole progenitor) cell and so forth. We measured the length of a cell as

the largest distance of the rod shaped cell in the orientation of the side channel. We applied a minimum

"cell length" of ~1µm to exclude artefacts originating mainly from small deformations in the growth

channels, or some shadows/reflections generated at the dead end of the growth channel.

Age at death

We defined age at death as the time when a cell stopped growing and dividing for at least one hour and

20 min. None of the cells divided after such a long division arrest or resumed growth.

Early and late daughter cells

We defined the early daughter population as the cells that were initially loaded into the microfluidic chip

and settled at the dead-end of the side channels. They originated largely from young mothers because the

population was in exponential growth phase when we loaded cells into the chip (Jagers 1978). The age

distribution of these early daughter cells is, because of the exponential growth phase, negatively

exponentially distributed. We show their expected age distribution in Fig. S1, based on a Leslie matrix

generated from the survival and division rates of the early daughters (Fig. 2A & C). Due to the

exponential growth phase, we assume a stable age distribution (the right eigenvector corresponding to

the dominant eigenvalue of the Leslie matrix) (Caswell 2001). We tracked these early daughter cells to

the end of their lives (Fig. 1, movie S1). The last daughters that were produced by these early daughter

cells are the late daughter population. These late daughters are all of age zero but originate from mother

cells (the early daughters) that had accumulated damage over their lives, and that were only one division

from dying. We again tracked these late daughters throughout their lives. We focus mainly on these first

two types of cells (early and late daughters), though as expected, the next generation of late daughter

cells (second generation late daughters) follow the late daughter patterns (Fig. S4-6).

Fig. S1: Expected (stable) age distribution (in hours) of early daughters (initial loaded cells), estimated

from a Leslie matrix parameterized with the demographic rates of the early daughter cells. Due to the

theory of stable age populations, the distribution of the ages of the mothers of these early daughters

should be exactly as the (stable) age distribution shown here of the initial loaded (early daughter cells)

at the start of the experiments.

Estimating demographic parameters

Image analysis (see above) provided us the basis for estimating the age-specific demographic parameters:

division rate, cell elongation rate, size at division, and survival rate. We excluded all cells that were not

tracked throughout their lives. This concerned mostly cells that filamented (cell division arrested without

growth arrest) and were flushed out of the side channels when they reached a size longer than 25 µm

(length of the side channel). We also excluded all (early and late daughter) cells that died at chronological

age 0 (never started to grow or divide). This was practiced because loading into the chip might be

damaging and we wanted to exclude such death. This approach was conservative, because fewer early

daughters never started growing or divided compared to late daughters (who could not have been

damaged by the loading). We used data from 516 early daughters that produced 516 late daughters which

in turns produced 298 (second generation) late daughters that were tracked throughout their lives.

Statistical analysis

We did the statistical analysis using the R software (R Core Team 2016). We estimated and plotted hourly

rates, rather than rates on the 4 min intervals because data would be sparse for 4 min intervals at older

ages; already the hourly rates at old ages suffer from sparse estimates and increased uncertainty. We also

clustered all data of ages beyond 30 h for the same reason. To guide our model selection (Burnham and

Anderson 1998), we used AIC (Akaike Information Criteria) and considered better support between

models when the ΔAIC was more than 2. For the parameter estimation of the GGM and extended random

deterioration models, we used the DEoptim R package and the optim function within the stats R package.

We calculated the CI (Fig. 2 C & D) by bootstrapping 1000 times. This bootstrapping is based on

individual level data not the hourly means that are plotted as open circles in Fig. 2 C, D. We choose such

an likelihood based global optimization approach to closely relate to the GGM modeling approach we

used for the simulations. Using model comparison (AIC) based on the hourly rates (results not shown)

selects for models with sigmoidal shapes like the GGM likelihood optimization, i.e. an early exponential

increase in mortality followed by a late age mortality plateau (Fig. 2 C). For early daughters, other

exponential or linear functions receive less support. For late daughters model comparisons selected non-

senescence (Fig. 2 D, flat mortality) over linear or exponential models.

Lifespan distribution second generation late daughters

In the main article, we focused on early and late daughter cells. Qualitatively, patterns of the second gen

eration late daughters (the last daughters produced by the late daughters) correspond to those of their m

others (the late daughters), though data becomes sparse uncertainty increases and some deviations occur

. Fig. S7 illustrates lifespan distributions of the 298 second generation late daughters, their mothers (late

daughters), and grandmothers (early daughters). Kolmogorov-Smirnov tests (two-sided) verifies that bo

th lifespan distributions between early and late daughters ($D = 0.3691$, $p < 0.00001$), and early to secon

d generation late daughters are significantly different ($D = 0.4094$, $p < 0.00001$), while those between la

te daughters and second generation late daughters do not differ ($D = 0.0973$, $p = 0.1189$).

Fig. S7: Lifespan distribution (in hours) of isogenic *E. coli* cells grown under identical environmental

conditions in a microfluidic device (Fig. S1). 1A – Founding early daughter cells (Median±stdev lifespan

12 ± 6.8 hours, Coefficient of Variation, CV 0.57) (N=298); 1B – Late daughters (last daughter) of the

founding early daughters of 1A (Median±stdev lifespan 4.4 ± 7.0 hours, Coefficient of Variation, CV

1.03) (N=298); 1C – Second generation late daughters (last daughter of the late daughters of 1B;

(Median±stdev lifespan 3.4 ± 7.1 hours, Coefficient of Variation, CV 1.14) (N=298). (Fig. S1 & S3).

Hence, early daughters are the mothers of late daughters and grandmothers of second generation late

daughters.

Heritability or cross generation correlation in lifespan

In order to estimate correlation of lifespan between mother cells (early daughters) and their last (late)

daughter cells, we used linear models on square root transformed data for age at death (lifespan) for the

mothers' lifespan as response variable and their last daughters' lifespan as explanatory variable, or

intercept only models (null model) as comparative models.

Model selection (Table S6) does not allow distinguishing between the two alternative models and we ca

nnot completely rule out that the longer-lived mothers produce slightly shorter-lived late daughters (Fig

. 2G). In any case, the effect size is weak and such a negative correlation would suggest that longer-live

d mothers accumulate more damage that is then partly transmitted to their late daughter. Similar results

hold for late daughters to second generation late daughters (Fig. S8). We cannot completely rule out a w

eak tendency that long-lived (late daughter) mothers produce last daughters (second generation late dau

ghters) that live slightly shorter (Table S6). In any case, our simulations with a low fixed transmission r

ate of 7% of damage transmitted between mother and daughters would lead to much higher correlation i

n lifespan as observed (Fig. 3E main text).

Fig. S8: Correlation between lifespan (Square root-transformed) of the 298 late daughters [mothers]

versus the lifespan of their last daughter cell, the second generation late daughters (Table S6).

Table S6: Model selection for correlation between lifespan (Square root transformed, Gaussian error)

Model	Slope	Std. Error	T value	p-value	AIC	ΔAIC
Early daughters [516 mothers] and daughter cell (last daughter of early daughter)						

sqrt(Lifespan Daughters) ~ sqrt(Lifespan mothers)	-0.10	0.05	-1.8	0.07	1774.1	
sqrt(Lifespan Daughters) ~ intercept only					1775.5	1.3

Late daughters [mothers 298] and second generation late daughters
(last daughter cell produced by late daughters)

sqrt(Lifespan Daughters) ~ sqrt(Lifespan mothers)	- 0.09	0.06	-1.55	0.12	1061.1	
sqrt(Lifespan Daughters) ~ intercept only					1061.5	0.4

Lifespan and lifetime reproductive success: reproductive and chronological aging

Previous aging studies on *E. coli* (Wang et al. 2010; Stewart et al. 2005) and other bacteria (Ackermann,

Stearns, and Jenal 2003) or many yeast studies (Denoth Lippuner, Julou, and Barral 2014) have focused

on replicative aging and patterns have been described for time measured in number of divisions. Such

age measures can be difficult to compare directly to chronological age if the time between divisions

varies. Nonetheless, we detected a strong correlation between lifespan and the lifetime reproductive

success (number of divisions an individual undergoes throughout its life), but there remains some

variation, particularly for longer lived individuals (Fig. S4).

Fig. S4: Correlation between lifespan and the lifetime reproductive success (cumulative number of

divisions an individual undergoes throughout its life) for left panel) the 516 early daughters, mid panel)

the 516 late daughters, and right panel) the 298 second generation late daughters.

Division rate

724 We fitted linear models on hourly average division rates (at the cellular level) as response variable and

725 age as explanatory variable. A model with a linear and quadratic term was best supported for the early

726 (Fig. 2A), late (Fig. 2B) and second generation late (Fig. S3A) daughters (Table S2).

728 Table S2: Model selection division rate

Model	Slope±Std. Error	Quadratic term±StdE	AIC	ΔAIC
Early daughters				
Divison rate ~intercept only			19663	24
Divison rate ~age	-0.011±0.002		19642	21
Divison rate ~age+age^2	**0.005±0.0078**	**-0.0007±0.0003**	**19638**	
Late daughters				
Divison rate ~intercept only			12432	24
Divison rate ~age	-0.009±0.004		12428	20
Divison rate ~age+age^2	**0.039±0.011**	**-0.002±0.0005**	**12408**	
Second generation late daughters				
Divison rate ~intercept only			5494	63.3
Divison rate ~age	0.009±0.004		5490.3	59.6
Divison rate ~age+age^2	**0.090±0.011**	**-0.004±0.0004**	**5430.7**	

Mortality rate

We fitted Gamma-Gompertz-Makeham related models to the survival data. GGM models were tested for

assumptions among parameters, for instance, whether shape and scale parameters are independent or

defined by a rate parameter. For late (Fig. 2D) and second generation late daughters (Fig. S3B), we

optimized the scaling of the mother to daughter state relationship (transmission factor) based on the age

at death state distribution of early daughters (or late daughters for second generation late daughters) (Fig.

S10). We assumed that accumulation of damage, which is the stage transition probabilities did not differ

among early, late and second generation late daughters. We also assumed that the effect of the stage was

not different for early, late or second-generation late daughters. These assumptions allowed us to use the

translated GGM estimates of the early daughters also for late and second generation late daughters.

Fig. S3: Age-specific division rate (number of divisions per hour) (A) and age-specific hourly mortality

rate (B) of the 298 second generation late daughters. For (A) means ± standard errors are plotted, for (B)

means ± 95% Confidence Intervals are plotted. Red lines are based on model selection using information

criteria (AIC) or the GGM model estimates (Table S2).

Difference in motrality rates at old ages for early and late daughters

The late age mortality plateau detected in early daughters (Fig. 2C) and the flat mortality hazard pattern

of the late daughters might indicate similar mortality rates at older ages (>19 hours). To test for this we

compared models that distinguished between early and late daughters but only included cells that

survived to an age of at least 20 hours. Late daughter cells have lower mortality rates ($q_x=0.1248\pm0.0191$

mean$\pm$Std.Err.) than early daughters at old ages ($q_x=0.2005\pm0.0168$ mean$\pm$Std.Err.) (Table S3) . This

result comprises 73 early daughters and 41 late daughters that lived longer than 19 hours.

Table S3: Model selection for age-specific mortality rates (q_x) of old (>19h) early daughters and old
(>19h) late daughters, based on binomial models.

Model	Estimate±Std. Error	Quadratic term	AIC	ΔAIC
Mortality ~intercept only			95.79	6.09
Mortality ~cell type	**-0.591±0.211**		**89.7**	

Cell elongation rates

We fitted linear models for cell elongation rates for early, late and second generation late daughters. For

all three cell types a model with a linear and quadratic term was best supported, though the curvature

was minimal for early daughters (Table S1). Early and late daughters showing clear decline in cell

elongation rate with age, while second generation late daughters first increase in cell elongation rate

before levelling off and maybe declining at old ages (Fig. S2).

Fig. S2 Cell elongation and cell size:

A-C) Age-specific cell elongation rate (fractional elongation) and D-F) cell size at division (in pixels,

1pixel=0.064 µm) of the A&D) 516 early daughter cells, B&E) 516 late daughter cells, and C&F) 298

second generation late daughter cells. Means ± standard errors are plotted, red lines are based on model

selection using information criteria (AIC) (Table S1).

Table S1: Model selection for age-specific cell elongation rates and cell size at division.

Cell elongation rates				
Model	Slope±Std. Error	Quadratic term	AIC	ΔAIC
Early daughters				
Elongation rate ~interc. only			-21415	407
Elongation rate ~age	-0.002±0.0001		-21820	2

Model	Slope±Std. Error	Quadratic term	AIC	ΔAIC
Elongation rate ~age+age^2	**-0.002±0.0002**	**-0.00002±0.00001**	**-21822**	
Late daughters				
Elongation rate ~interc. only			-10264	45
Elongation rate ~age	-0.00102±0.00016		-10304	5
Elongation rate ~age+age^2	**0.00013±0.00047**	**-0.000052±0.00002**	**-10309**	
Second generation late daughters				
Elongation rate ~interc. only			-4042.4	32.1
Elongation rate ~age	0.0014735±0.0002959		-4065.1	10.4
Elongation rate ~age+age^2	**0.00429±0.00088**	**-0.00012±0.000035**	**-4074.5**	
Cell size at division				
Model	**Slope±Std. Error**	**Quadratic term**	**AIC**	**ΔAIC**
Early daughters				
Size at division ~interc. only			53741	1378
Size at division ~age	1.72±0.04		52431	68
Size at division ~age+age^2	**2.797±0.135**	**-0.049±0.006**	**52362**	
Late daughters				
Size at division ~interc. only			28592	101
Size at division ~age	0.75±0.09		28525	34
Size at division ~age+age^2	2.204±0.259	-0.069±0.012	28491	
Second generation late daughters				
Size at division ~interc. only			8381.7	276.2
Size at division ~age	3.036±0.223		8215.9	110.4
Size at division ~age+age^2	**8.919±0.576**	**-0.293±0.027**	**8105.5**	

Size at division

We fitted linear models for cell size at division for early, late, and second generation late daughters. For

all three cell types, a model with a linear and quadratic term was best supported, though the curvature

was minimal for early and late daughters (Table S1). Early and late daughters showed clear senescence

in cell elongation rate while the second-generation late daughters selected models might be over-fitted

(Fig. S2).

Reverse time analyses: Hours before death

Chronological aging follows individuals from birth to death and time counted starts at birth. Various

hypotheses about the evolution of life histories, aging and senescence, in particular, terminal investment

strategies (Charlesworth 1994; Clutton-Brock 1984; Williams 1966), assume that individuals can sense

that they are approaching death and that an optimal strategy for an individual might be to invest remaining

resources into reproduction rather than maintenance when approaching immediate death. As Fig. 2 &

Fig. 3 in the main text illustrates, age at death is highly variable, for that chronological senescence

patterns and terminal investment strategies might provide very different insights. We investigated

whether patterns in reproduction (division rate), cell elongation rate, or size at division differ significantly

when time is counted backward (remaining lifespan), starting with death. We would expect that the last

hour before death would show significant changes compared to the more spread out chronological

senescence patterns (Fig. 2, Fig. S3 & S4). In general, there seems to be little evidence that terminal

investment is happening to a larger degree, at least such strategies if existent, do not show larger effects

compared to chronological aging and classical senescence (Fig. S5 & S6, Table S4 & S5).

Division rate for hours before death

When individual cells approach death, their division rates decrease (Fig. S5), and these senescence

patterns are similar compared to the (chronological) senescence patterns (Fig. 2A, B) (Table S4 & Table

S2).

Fig. S5: Hourly division rate for hours before death [remaining life time] (death = time 0) for left panel

the 516 early daughters cells, mid panel the 516 late daughters, and right panel the 298 second generation

daughters. Means ± standard errors are plotted, red lines are based on model selection using information

criteria (AIC) (Table S4).

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804 Table S4: Model selection for division rate in reverse time (starting with age at death=0).

Model	Slope±Std. Error	Quadratic term	AIC	ΔAIC
Early daughters				
Division rate ~interc. only			19567	69
Division rate ~reverse time	0.006±0.002		19561	64
Division rate ~ reverse time + reverse time ^2	**0.062±0.0070**	**-0.002±0.0003**	**19497**	
Late daughters				
Division rate ~interc. only			12171	8
Division rate ~reverse time	0.006±0.003		12170	7
Division rate ~ reverse time + reverse time ^2	**0.036±0.017**	**-0.001±0.0004**	**12163**	
Second generation daughters				
Division rate ~interc. only			5474.5	47.8
Division rate ~reverse time	0.015±0.003		5456	29.3
Division rate ~ reverse time + reverse time ^2	**0.069±0.017**	**-0.002±0.0004**	**5426.7**	

Cell elongation rate for hours before death

The cell elongation rate does not drastically fall off when approaching death (no initial steep increase in

cell elongation rate close to the time of death) (Fig. S6, Table S5). The senescence pattern is, at least for

early and late daughters, less pronounced compared to that observed for chronological senescence (Fig.

S2). Second generation daughters seem to reduce cell growth rates before death more substantially, but

this reduction already starts a few hours before death and is not in line with terminal investment theories.

Fig. S6: (Fractional) Cell elongation rate and cell size at division (in pixels, 1pixel=0.064 μm) for hours

before death (death = time 0) for A & D) the 516 early daughters, B & E) the 516 late daughters, and C

& F) the 298 second generation late daughters. Means ± standard errors are plotted, red lines are based

on model selection using information criteria (AIC) (Table S5).

Table S5: Model selection for cell elongation and cell size at division rate in reverse time (starting with
age at death=0).

Model	Slope±Std. Error	Quadratic term	AIC	ΔAIC
Cell elongation (reverse time)				
Early daughters				
Elongation rate ~interc. only			-22138	250
Elongation rate ~reverse time	0.001±0.00009		-22288	100
Elongation rate ~ reverse time + reverse time ^2	**0.004±0.0003**	**-0.0001±0.00001**	**-22388**	
Late daughters				
Elongation rate ~interc. only			-10529	65
Elongation rate ~reverse time	0.001±0.0001		-10561	33
Elongation rate ~ reverse time + reverse time ^2	**0.003±0.0004**	**-0.0001±0.00002**	**-10594**	
Second generation late daughters				
Elongation rate ~interc. only			-4082.8	122.1
Elongation rate ~reverse time	0.003±0.0003		-4167.5	37.4
Elongation rate ~ reverse time + reverse time ^2	**0.008±0.0009**	**-0.0002±0.00003**	**-4204.9**	
Cell size at division (reverse time)				
Early daughters				
Size at division ~interc. only			53679	386
Size at division ~reverse time	-0.952±0.0486		53299	6
Size at division ~ reverse time + reverse time ^2	-1.337±0.145	0.018±0.006	53293	
Late daughters				
Size at division ~interc. only			28629	26
Size at division ~reverse time	-0.438±0.088		28606	3
Size at division ~ reverse time + reverse time ^2	**-0.990±0.258**	**0.026±0.011**	**28603**	
Second generation late daughters				
Size at division ~interc. only			8182.3	54.9

Size at division ~reverse time	1.254±0.236		8156.6	29.2
Size at division ~ reverse time + reverse time ^2	**4.531±0.627**	**-0.158±0.028**	**8127.4**	

Cell size at division for hours before death

Cell size at division increased with chronological age, but mainly for early daughters (Fig. S2). One

mechanism that might lead to such an increase in cell size at division is the formation of filamentous

cells: cells that do continue to elongate but stop dividing. Such filamentation is usually seen as a stress

response and increases mortality of the cell (Mattick, Rowbury, and Humphrey 2003; Wang et al. 2010).

Previous studies (conducted under different media and temperature than ours) have reported higher

filamentation rates with increasing age (Wang et al. 2010). Under our growth conditions, cells that

entered a filamentation stage rarely recovered from that stage and did not return to a regular dividing

smaller cell, though they not always stalled dividing. To test if such filamentous cells drive the

chronological senescence pattern in size at division, we investigated the cell size for the hours before

death. For early and late daughters, we detected an increase in cell size at division before death, but this

increase is already initiated hours before the actual death and not a final increase in size due to

filamentation (Fig. S6, Table S5).

Random deterioration model

Our random deterioration model builds on a LeBras type cascading failure model(Le Bras 1976). In such

discrete state models, individuals increase from a current state i to a state $i+1$ at a rate proportional to i,

and mortality increases proportionally to state i. We can think of such states as accumulating damage, or

model the reverse process, where an individual starts with some vitality and this vitality then decreases

with increasing state i (random deterioration) (Weitz and Fraser 2001). In this study, we assume

accumulating damage that stands symbolically for an unknown aging factor. These random deterioration

models are mathematically closely linked to demographic frailty models that do not assume a stochastic

stage progression among individuals, but rather heterogeneity of individual is defined by fixed frailty at

birth (Yashin, Vaupel, and Iachine 1994). Such frailty, an age independent mortality that follows a

gamma distribution (Missov and Vaupel 2015), can be added to the classical models in demography, a

Gompertz- model, which describes an age-dependent exponential increase in mortality. These models

can also be extended by a Makeham term, an age-independent baseline mortality, which further improves

model fit, at least for humans and various other complex species.

We first estimated parameters of such a Gamma-Gompertz-Makeham (GGM) model, $Zae^{bx}+c$ fitted to

the survival data of the early daughters. Z is the random frailty variable that follows a gamma distribution

across individuals. The distribution has a mean 1 and variance of σ^2. This definition of a gamma

distribution with a mean of 1 implies that we have to set the scale parameter equal to the shape parameter

of the gamma distribution. The rate of aging, b, is the same for all individuals, x determines the age in

hours, a describes the baseline mortality of the Gompertz part of the model, and c describes the Makeham

term, the age independent mortality.

We then transformed these estimates of the GGM model to a random deterioration model according to

Yashin et al. (Yashin, Vaupel, and Iachine 1994), making use of mathematical similarities between the

GGM and random deterioration models.

Let $t(i) = \lambda_0 + i\,\lambda$, $m(i) = \mu_0 + i\mu$ (for i greater than or equal to 0) denote the state transition rate and the

death rate, respectively.

$\mu_0 = a+c$, the death rate at birth i.e. state $i=0$,

$\lambda_0 = \dfrac{b}{\sigma^2} - a$, the transition rate from state $i=0$ to state $i=1$,

$\lambda = b - \sigma^2 a$, the change in the state transition rate between state i and $i+1$, for $i>0$

$\mu = \sigma^2 a$, the change in the death rate at state i, for $i>0$

Resulting estimates for early daughters were $\mu_0=0.056$, $\lambda_0 = 0.178$, $\lambda = 0.598$, and $\mu = 0.00023$.

Using these estimates for early daughters, we estimated the probability of observing an individual at state

i at age x, $P(x,i)$,

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$$P_i(x) = \begin{cases} P_0(0)e^{-(\lambda_0+\mu_0)x} , & i = 0 \\ \dfrac{P_0(x)}{i!}\left[\dfrac{\lambda - \lambda e^{-(\lambda+\mu)x}}{\lambda + \mu}\right]\prod_{k=1}^{i}\left(\dfrac{\lambda_0}{\lambda} + (k-1)\right), & i > 0 \end{cases}$$

We then built a matrix of $P(x,i)$ for which we limited the age, x, between 0 and 31 and the states were

limited between 0 and 5000. The column sum of this matrix equals the survival probability at any given

age x to age $x+1$.

Microsimulations of the random deterioration model

In order to simulate the (damage) state at death distribution we used a microsimulation. As a step-by-

step calculation would take too long, we modelled the state at death distribution by drawing of a Poisson

distribution with a rate of $\dfrac{1}{\mu_0+\hat{\imath}\mu}$, where $\hat{\imath} = i + \dfrac{1}{\mu_0+i\mu}$. In doing so, we estimated the expected state at

death, rather than tracking individuals transitioning through damage states and their associated mortality

risk as a step-by-step calculation would do. A step-by-step simulation would take a long time because

individuals can easily end up in very high damage states (very long tailed distributions).

Estimating parameters for the late daughters

For the late daughters, we used the same state transition and survival parameters (μ_0=0.056, $\lambda_0 = 0.178$,

$\lambda = 0.598$, and $\mu = 0.00023$) that we estimated for the early daughters. We did so because early and

late daughter cells should biologically not be fundamentally different – aside of starting with different

damage levels at birth. They should accumulate damage at the same rate and should experience mortality

risk just based on their damage state. We assumed that the late daughters are born at the same stage, or a

scaled version of the same stage, in which the mothers died, therefore let

$$\lambda_{0,late} = nint\left(s\left(\lambda_{0,early} + i\lambda\right)\right)$$

$$\mu_{0,late} = nint\left(s\left(\mu_{0,early} + i\mu\right)\right)$$

Where subscript *late* and *early* denotes late and early daughter estimates and $s \in (0,1)$ and $nint(.)$ is

the nearest integer function. So based on the estimated damage state distribution of the mothers at death

(early daughters), we also knew the damage state distribution at birth of the daughters (late daughters)

[see below for optimally scaling this distribution]. Using the state transition probabilities of the early

daughters and the probabilities of death, we estimated the survival of the late daughters. We also

estimated the survival functions of cells that did not change their damage state with age. Each grey line

in main text Fig. 3, describes such a fixed damage class survival function.

In order to determine the transmission fraction of accumulated damage between mothers (early

daughters) and daughters (late daughters), we optimized the scaling of the mother state at death

distribution (fraction of mother-daughter transmission) to the daughter age at birth distribution. The

optimization was based on observed survival of the late daughters, by minimizing the squared deviations

from the observed survival function. This optimal scaling factor was estimated at 7%.

Similar to the late daughters, we simulated the damage state distribution at birth, the damage state

distribution at death, and calculated the matrix $P(x,i)$ for the second generation late daughters.

Fig. S9: Observed population level survival curve (red line), GGM simulated survival curve with

symmetric damage transmission (hatched black line) as well as GGM simulated survival curve with

asymmetric damage transmission of 7% (solid black line ± 95% CI in blue) for early (A), late (B) and

second generation late daughters (C). Graphs A & B are identical to Fig. 1C & D. Thin grey lines in

depict expected survival curve of cells with different fixed damage state. That is, for instance, the

outermost thin grey line in B depicts the survivorship curve of a hypothetical cohort that starts without

damage and never accumulates any damage. The most left survival curve illustrates the low survival of

cells that were born with maximum damage level of 5000.

Fig. S10: Density distribution of damage state at death of early daughters, late daughters, and second

generation late daughters. Most cells die with little accumulated damage. Note, the not so smooth tails

of the distributions are a result of drawing from a Poisson distribution in our simulations.

Fig. S11: GGM model results: mean (red solid line, + SD grey hatched lines) and median (red hatched

line) damage state with increasing age for early daughter cells (A), and late daughter cells (B) and second

generation late daughter cells (C). Panel A & B are identical to Panel Fig. 1A, B.

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