Aggressiveness as a latent personality trait of domestic dogs: testing local independence and measurement invariance

Conor Goold[1*], Ruth C. Newberry[1]

[1] Department of Animal and Aquacultural Sciences, Norwegian University of Life Sciences, Ås, Akershus, Norway

[*] Corresponding author: conor.goold@nmbu.no (CG)

Abstract

Studies of animal personality attempt to uncover underlying or 'latent' personality traits that explain broad patterns of behaviour, often by applying latent variable statistical models (e.g. factor analysis) to multivariate data sets. Two integral, but infrequently confirmed, assumptions of popular latent variable models in animal personality are: i) behavioural variables are independent (i.e. uncorrelated) conditional on the latent personality traits they reflect (*local independence*), and ii) personality traits are associated with behavioural variables in the same way across individuals or groups of individuals (*measurement invariance*). We tested these assumptions using observations of aggression in four age classes (4 - 10 months, 10 months - 3 years, 3 - 6 years, over 6 years) of male and female shelter dogs (N = 4,743) in 11 different contexts. A structural equation model supported the hypothesis of two correlated (ρ = 0.25; p < 0.001) personality traits underlying aggression across contexts: aggressiveness towards people and aggressiveness towards dogs (comparative fit index: 0.97; Tucker-Lewis index: 0.96; root mean square error of approximation: 0.03). Aggression across contexts was moderately repeatable (towards people: ICC = 0.479, 95% CI: 0.466, 0.491; towards dogs: ICC = 0.303, 95% CI: 0.291, 0.315). However, certain contexts related to aggressiveness towards people (but not dogs) shared significant residual correlations unaccounted for by latent levels of aggressiveness. Furthermore, aggressiveness towards people and dogs in different contexts interacted with sex and age. Thus, sex and age differences in displays of aggression were not simple functions of underlying aggressiveness. Our results illustrate that the robustness of traits in latent variable

models must be critically assessed before making conclusions about the effects of, or factors influencing, animal personality. Our findings are of concern because inaccurate 'aggressive personality' trait attributions can be costly to dogs, recipients of aggression and society in general.

Key words: animal personality assessment; agonistic behaviour; shelter dogs; measurement bias; behavioural phenotyping

Introduction

Studies of non-human animal personality, defined by relatively consistent between-individual differences in behaviour, demonstrate that organisation of the behavioural phenotype is markedly hierarchical and non-independent [1–3]. Observed patterns of behaviour can be explained by broad behavioural dimensions or personality traits, including boldness, activity, exploration, sociability and aggressiveness [4], which are further inter-correlated to form behavioural syndromes [5]. To interpret the complexity inherent in behavioural phenotypes, personality traits and behavioural syndromes are frequently inferred using latent variable statistical models, which reduce two or more measured variables (the *manifest* variables) into one or more lower-dimensional variables (the *latent* variables). Latent variable models have been popular in human psychology for over a century [6], and now comprise a flexible set of methods to derive lower-order variables from multivariate data sets [7].

Because latent variables are unobserved, latent variable models require careful application and interpretation [8–10]. In animal personality, many studies use *formative* models, such as principal components analysis, that construct composite variables comprised of linear combinations of manifest variables. However, formative models impose only weak assumptions about the relationships between latent variables and manifest variables [11]. As such, their utility for inferring the nature of personality traits has been criticised in both animals [12,13] and humans [9,11,14,15]. Alternatively,

researchers have used *reflective* models, such as factor analysis and, increasingly, confirmatory approaches such as structural equation modelling. These models regress measured behaviours on one or more latent variables [11], incorporating measurement error and possibilities to compare *a priori* competing hypotheses (for applications in animal personality, see [1,16–18]).

Whilst reflective models offer a powerful framework to probe the latent variable structure of behaviour, they impose certain constraints on the interpretation and modelling of latent variables that have received scrutiny in human psychology but are rarely discussed in studies of animal personality. Two foundational assumptions of these models are *local independence* and *measurement invariance*. Local independence implies that manifest variables should be independent of each other conditional on the latent variables [19,20]. For example, given a latent variable θ and two binary manifest variables Y_1 and Y_2 that can take the values 0 and 1, the item response theory model asserts that $P(Y_1 = 1, Y_2 = 1|\theta) = P(Y_1 = 1|\theta)P(Y_2 = 1|\theta)$. As such, the latent variables should screen off any dependence between manifest variables. Measurement invariance, on the other hand, implies that the latent variables function the same (i.e. are invariant or equivalent) in different subsets of a population or in the same individuals through time [21–24]. That is, given a fixed level x on a latent variable θ, denoted θ_x, the expected values of manifest variables Y_i, from *i = 1* to *N*, should be the same across a grouping factor π (e.g. sex or different populations), expressed as $E(Y_i \mid \theta_x) = E(Y_i \mid \theta_x, \pi)$. Intuitively, violations of either local independence or

measurement invariance imply that the latent variables do not completely explain variation in manifest variables, which may lead to artefactual conclusions about the differences between individuals as a function of trait scores [25–27]. More generally, investigating violations of local independence and measurement invariance can identify measurement biases that can be rectified to improve the quality of personality assessments.

The use of latent variable modelling is particularly common in studies of personality in domestic dogs (*Canis lupus familiaris*), where the collection of data sets comprised of multiple behavioural variables (e.g. questionnaires) has been prevalent for decades [28–30]. Personality assessments are important both in applied settings to predict the behaviour of dogs at future time points [31] and also to elucidate behavioural traits pertinent to dogs' domestication history [32,33]. Decomposing behaviour into a smaller number of underlying dimensions that explain variation in measured behaviours aids in measuring hypothesised stable behavioural features of individuals from which predictions of future behaviour are estimable and upon which selection may have been focused. Nonetheless, research on personality in dogs has led to different numbers and composition of hypothesised personality traits with little clear consensus on how such traits should be compared within and between studies [34–36]. Most importantly, the predictive value of personality assessments in dogs has been inconsistent [31,37–41], perhaps most prominently in shelter dog personality assessments (e.g. see [31] for a review). The predictive validity of personality assessments is of particular concern in

tests of aggression, where aggression is frequently divided into different aggressiveness traits, including owner-, stranger-, dog- or animal-directed factors [29,35,42,43]. Based on such assessments, if falsely labelled as aggressive in one or more of these categories, the dog may be euthanised unnecessarily whereas, if falsely considered unaggressive, the dog may be more likely to be placed in situations leading to harmful behaviour. Dog bites are a serious public health concern [44], especially for animal shelters rehoming dogs to new owners, and aggressive behaviour is undesirable to many organisations using dogs for various working roles [45].

Why has the prediction of dog personality been difficult? On one hand, the validation of personality tests has been criticised for lacking rigour [34,46], meaning the targeted traits may not be under measurement as intended. In particular, commonly-used 'battery assessments', comprised of different sub-tests conducted sequentially, may not be valid if the responses to one sub-test alter how the animal responds to subsequent sub-tests given an underlying trait (e.g. if test items are 'invasive' to subsequent items; [47]). Moreover, test batteries can be "inherently stress-inducing for test animals" ([31]: 10), so behavioural responses may be functions of the targeted traits and stress responses. On the other hand, personality may simply not be generalisable across all individuals if individual behaviour is differentially dependent on factors such as age (e.g. [35,40]) or fluctuating environmental conditions (e.g. behaviour in a shelter versus out of a shelter; [48]). In reflective latent variable modelling, greater attention to the assumptions of local independence and measurement invariance is, thus, warranted.

124 For aggressiveness specifically, if aggression shown in one context or test item increases
125 the likelihood of aggression in others [48], local independence would be violated.
126 Further, interactive effects between aggressiveness and demographic variables, such as
127 age, sex, breed or neuter status (e.g. [49–51]), would violate the assumption of
128 measurement invariance. Whilst van den Berg *et al.* [18] reported measurement
129 invariance of the *stranger-directed aggression* factor from the Canine Behaviour and
130 Research Questionnaire (C-BARQ; [29]) across three breeds of dogs, measurement
131 invariance of aggression assessed with respect to other factors has not been evaluated
132 and, to our knowledge, no authors have confirmed local independence between
133 manifest variables in animal personality assessments.

134

135 In this paper, we assessed local independence and measurement invariance of
136 aggressiveness using a large sample of data on inter-context aggression in shelter dogs.
137 First, we sought to decompose observations of aggression towards people and dogs
138 across contexts into separate aggressiveness traits. Secondly, we assessed whether
139 aggression in different contexts remained associated beyond that explained by latent
140 levels of aggressiveness, testing local independence. Thirdly, we investigated whether
141 the probability of aggression in different contexts believed to be underpinned by the
142 same aggressiveness trait was measurement invariant with respect to sex and age
143 groups. While it may be unrealistic for measurement invariance to hold in all instances,
144 it is important to establish whether it holds across basic biological variables such as age

and sex which are generally applicable to dog populations undergoing personality assessment.

Materials & Methods

Subjects

Observational data on the occurrence of aggression in 4,743 dogs were gathered from Battersea Dogs and Cats Home's (UK) observational and longitudinal dog behaviour assessment records. The data were from a sample of dogs (N=4,990) at the shelter's three rehoming centres during 2014 (including dogs that arrived during 2013 or left in 2015). We selected the records from all dogs that were at least 4 months old, excluding younger dogs because they were more likely to be unvaccinated, more limited in their interactions at the shelter and may have been kennelled in different areas to older dogs. Relevant demographic characteristics of the dogs are provided in Table 1. Although dogs were from a variety of heritages (including purebreeds and mongrels), the analyses here did not explore breed differences because visual assessment to attribute breed to dogs with unknown heritage has been questioned [52].

162

Table 1. Demographic characteristics of the studied dogs.

Variable	**Mean ± SD / *N***
Average age at shelter (years; all ≥ 4 months of age)	3.75 ± 3.03
Total days at the shelter	25.13 ± 41.53
Weight (average weight if multiple measurements; kg)	19.06 ± 10.26
Rehoming centre: London / Old Windsor / Brands Hatch	2897 / 1280 / 566
Males / females	2749 / 1994
Neutered before arrival / neutered at shelter / not neutered	1218 / 1665 / 1502
Relinquished by owners / returned to shelter / strays	2892 / 260 / 1591

163

164 ## Shelter environment

165 The shelter was composed of three different UK rehoming centres: a high-throughput,
166 urban centre based at Battersea, London with capacity for approximately 150-200 dogs;
167 a semi-rural/rural centre based at Old Windsor with capacity for approximately 100-150
168 dogs; and a rural centre based at Brands Hatch with capacity for approximately 50 dogs.
169 All dogs arrived in an intake area of their respective rehoming centre and, when
170 considered suitable for adoption, were moved to a 'rehoming' area that was partially
171 open to the public between 1000 h and 1600 h. All kennels were indoors. Kennels varied
172 in size, but were usually approximately 4m x 2m and included either a shelf and bedding
173 alcove area, or a more secluded bedding area at the back of the kennel (see [53] for
174 more details). At different times throughout the day, dogs had access to indoor runs
175 behind their kennels. In each kennel block area, dogs were cared for (e.g. fed, exercised,

kennel cleaned) by a relatively stable group of staff members, allowing the development of familiarity with staff members and offering some predictability for dogs after arrival at the shelter. Although data on the number of dogs in each kennel were incomplete, in the majority of cases dogs were kennelled singly for safety reasons. The shelter mainly operated between 0800 h and 1700 h each day. All dogs were socialised with staff and/or volunteers each day (often multiple times) unless it was unsafe to do so. They were provided water ad libitum and fed commercial complete dry and/or wet tinned food twice daily (depending on recommendations by veterinary staff). Dogs received daily tactile, olfactory and/or auditory enrichment/variety (e.g. toys, essential oils, classical music, time in a quiet 'chill-out' room).

Data collection

In the observational assessment procedure, trained shelter employees recorded observations of dog behaviour in a variety of contexts as part of normal shelter procedures. The average number of days between successive observations within all contexts and across all dogs was 3.27 (SD = 2.08), and dogs had an average of 9.77 (SD = 13.41) observations within each context. Behavioural observations pertaining to each context were completed using an ethogram specific to that context and recorded in a custom computer system. Multiple observations could be completed each day. The ethogram behaviour that best described a dog's behaviour in a particular context during an observation was recorded by selecting it from a series of drop-down boxes (one for

each context). Although staff could also add additional information in character fields, we did not analyse those comments in this study. The ethogram for each context represented a scale of behaviours ranging from desirable to undesirable considered by the shelter to be relevant to dog welfare and ease of adoption. Contexts had between 10 and 16 possible behaviours to choose from, some of which overlapped between different contexts. Among the least desirable behaviours in each context was aggression towards either people or dogs (depending on context). Aggression was formally defined as "*Growls, snarls, shows teeth and/or snaps when seeing/meeting other people/dogs, potentially pulling or lunging towards them*", distinguished from non-aggressive but reactive responses, defined as "*Barks, whines, howls and/or play growls when seeing/meeting other people/dogs, potentially pulling or lunging towards them*". Observation contexts included both onsite (at the shelter) and offsite (e.g. out in public parks) settings.

For the analyses here, we focused on observations of aggression in nine core onsite contexts: i) *Handling*, ii) *In kennel*, iii) *Out of kennel*, iv) *Interactions with familiar people*, v) *Interactions with unfamiliar people*, vi) *Eating food*, vii) *Interactions with toys*, viii) *Interactions with female dogs*, ix) *Interactions with male dogs*. For the *In kennel* and *Out of kennel* contexts, recording of aggression towards both people and dogs was possible. Although multiple observations could be made, if both occurred at the same time, aggression towards people would be prioritised over aggression towards dogs in the recording process. Therefore, *In kennel* and *Out of kennel* were each divided to reflect

aggression shown towards people and dogs, respectively. This resulted in 11 final aggression contexts, defined in Table 2, which were used as manifest variables in structural equation models to investigate latent aggressiveness traits. Each observation of aggression was recorded in the category that best described the scenario. Nonetheless, certain contexts could occur closely in space and time. For example, the *Handling* context could directly succeed the *Interactions with familiar people* or *Interactions with unfamiliar people* contexts. The sequential occurrence of certain contexts was used to inform tests of local independence, explained below.

We aggregated behavioural observations across time for each dog into a dichotomous variable indicating whether a dog had or had not shown aggression in a particular context at any time while at the shelter. This was performed because the overall prevalence of aggression was low, with only 1.06% of all observations involving aggression towards people and 1.13% towards dogs. Thus, the main difference between individuals was whether they had or had not shown aggression in a particular context during their time at the shelter (see Table S1 for raw counts of aggression by context), and we interpret aggressiveness here as a between-individual difference variable.

Table 2. Behavioural observation contexts.

Context	Definition
Handling	A dog's reaction to informal handling by people (e.g. stroking non-sensitive areas, touching the collar, fitting a harness or lead).
In kennel towards people	A dog's reaction to people approaching or walking past the kennel.
In kennel towards dogs	A dog's reaction to dogs in neighbouring kennels or dogs walking past the kennel.
Interactions with familiar people	A dog's reaction when outside of the kennel to familiar people (interacted with at least once before) approaching, making eye contact, speaking to or attempting to make physical contact with the dog.
Interactions with unfamiliar people	A dog's reaction when outside of the kennel to unfamiliar people (never interacted with before) approaching, making eye contact, speaking to or attempting to make physical contact with the dog.
Out of kennel towards people	A dog's reaction when around people outside of the kennel. Large distances may separate the focal dog and people, and no attempt is made to engage by the people with the focal dog.
Out of kennel towards dogs	A dog's reaction when around dogs outside of the kennel. Large distances may separate the focal dog and other dogs, and the dogs are not encouraged to interact.
Eating food	A dog's reaction when eating food (e.g. from a food bowl, or toy filled with food) to people approaching, in close proximity, or attempting to touch the food container.
Interactions with toys	A dog's reaction when interacting with toys to people approaching within close proximity and/or attempting to touch the toy.
Interactions with female dogs	A dog's reaction during structured interactions with a female dog, including approaching each other, walking in parallel, and interacting off-lead. Both dogs are aware of each other's presence and are in close enough proximity to engage in a physical interaction.
Interactions with male dogs	A dog's reaction during structured interactions with a male dog, including approaching each other, walking in parallel, and interacting off-lead. Both dogs are aware of each other's presence and are in close enough proximity to engage in a physical interaction.

Behavioural observation contexts analysed for the presence or absence of aggression.

Validity of behaviour recordings

Validity of the recording of behaviour was assessed separately from the main data collection as part of a wider project investigating the use of the observational assessment method. Ninety-three shelter employees trained in conducting behavioural observations each watched (in groups of 5 – 10 people) 14 videos, approximately 30 seconds each, presenting exemplars of 2 different behaviours from seven contexts (to keep the sessions concise and maximise the number of participants). For each context, behaviours were chosen pseudo-randomly by numbering each behaviour and selecting two numbers using a random number generator. Experienced behaviourists working at the shelter filmed the videos demonstrating the behaviours. Videos were shown to participants once in a pseudo-random order. After each video, participants recorded on a paper answer sheet the behaviour they thought most accurately described the dog's behaviour based on the ethogram specific to the context depicted. Two of the videos illustrated aggression: one in a combined *Interactions with new* and *familiar people* context (combined because familiarity between specific people and dogs was not universally known) and one in the *In kennel towards dogs* context. The authors were blind to the selection of videos shown and to the video coding sessions with shelter employees.

Data analysis

256 All data analysis was conducted in R version 3.3.2 [54].

257

258 Validity of behaviour recordings

259 The validity of shelter employees' recording of behaviour from videos was assessed by
260 the percentage of participants who identified the 2 videos as showing examples of
261 aggression.

262

263 Missing data

264 Data were missing when dogs did not experience particular contexts while at the
265 shelter. The missing data rate was between 0.06% and 5% for each context, except for
266 the *Interactions with female dogs* and *Interactions with male dogs* categories which had
267 17% and 18% of missing values, respectively (because structured interactions with other
268 dogs did not arise as frequently). Moreover, 16% and 7% of dogs were missing weight
269 measurement and neuter status data, respectively, which were independent variables
270 statistically controlled for in subsequent analyses. We created 5 multiply imputed data
271 sets (using the *Amelia* package; [55]), upon which all of the analyses in the sections
272 below were conducted and results pooled. The multiple imputation took into account
273 the hierarchical structure of the data (observations within dogs), all independent
274 variables reported below, and the data types (ordered binary variables for the context
275 data, positive-continuous for weight measurements, nominal for neuter status). The

data were assumed to be missing at random, that is, dependent only on other variables in the analyses.

Structural equation models

We used structural equation modelling to assess whether aggression towards people (contexts: *Handling*, *In kennel towards people*, *Out of kennel towards people*, *Interactions with familiar people*, *Interactions with unfamiliar people*, *Eating food*, *Interactions with toys*) and towards dogs (contexts: *In kennel towards dogs*, *Out of kennel towards dogs, Interactions with female dogs*, *Interactions with male dogs*) could be explained by two latent aggressiveness traits: aggressiveness towards people and dogs, respectively. We compared a model where the latent variables were orthogonal to a model where the latent variables were allowed to correlate, since positive correlations between different aggressiveness traits in dogs have been reported in dogs [56]. Models were fit using the *lavaan* package [57], with the weighted least squares mean and variance adjusted (WLSMV) estimator and theta/conditional parameterisation, as recommended for categorical dependent variables [7,58,59]. The latent variables were standardised to have mean 0 and variance 1. The results were combined across imputed data sets using the 'runMI' function in the *semTools* package [60]. The fit of each model was ascertained using the comparative fit index (CFI) and Tucker Lewis index (TLI), where values > 0.95 indicate excellent fit, as well as the root mean squared error of

approximation (RMSEA) where values < 0.06 indicate good fit [7]. Parameter estimates were summarised by *p* values (considered significant at $p < 0.05$).

Local independence

We tested the assumption of local independence by re-fitting the best-fitting structural equation model with residual correlations between context variables. To maintain a theoretically driven approach (see [61] regarding the best practice of including residual correlations in structural equation models) and model identifiability, we only tested a predefined set of residual correlations that we believed to be most relevant. First, we allowed correlations between *Handling* with *In kennel towards people*, *Interactions with familiar people*, *Interactions with unfamiliar people* and *Interactions with toys*, respectively. The *Handling* context could directly succeed these other contexts, leading to close temporal-spatial relationships, and whether a dog showed aggression in the *Handling* context may be mediated by a person's decision to handle a dog depending on the dog's behaviour in preceding contexts. The residual correlation between *Handling* and *Eating food* was not estimated because shelter employees would be unlikely to handle a dog while the dog ate its daily meals. The residual correlation between *Handling* and *Out of kennel towards people* was not estimated because any association between *Handling* and *Out of kennel towards people* would be mediated by either the *Interactions with familiar people* or *Interactions with unfamiliar people* context. Therefore, secondly, we estimated the three-way correlations between *Out of kennel*

317 *towards people*, *Interactions with familiar people* and *Interactions with unfamiliar*
318 *people*. Similarly, and lastly, we estimated the three-way correlations between *Out of*
319 *kennel towards dogs, Interactions with female dogs* and *Interactions with male dogs*.

320

321 **Measurement invariance**

322 To test for measurement invariance in each of the latent traits derived from the best
323 fitting structural equation model, we investigated the response patterns across
324 aggression contexts related to the same latent aggressiveness trait using Bayesian
325 hierarchical logistic regression models. Whilst measurement invariance for categorical
326 data can be ascertained in structural equation model frameworks [58,62], item response
327 theory is more commonly applied to dichotomously scored variables. In psychometrics,
328 the 1-parameter item response theory model, or Rasch model, represents the
329 probability that an individual responds correctly to a particular test item as a logistic
330 function of i) each individual's latent ability and ii) the item's difficulty level. The Rasch
331 model can be expressed as a hierarchical logistic regression model [63,64], whereby
332 individual latent abilities are modelled as individual-specific intercepts (i.e. 'random
333 intercepts'), the propensity for a correct answer to an item i is its regression coefficient
334 β_i, and credible interactions between items and relevant independent variables (e.g.
335 group status) indicate a violation of measurement invariance. Here, the dependent
336 variable was the binary score for whether or not dogs had shown aggression in each
337 context and the average probability of aggression across contexts varied by dog,

representing latent levels of aggressiveness. Context type, dog age, dog sex and their interactions were included as categorical independent variables.

Age was treated as a categorical variable, with categories reflecting general developmental periods: i) 4 months to 10 months (juvenile dogs before puberty), ii) 10 months to 3 years (dogs maturing from juveniles to adults), iii) 3 years to 6 years (adults), and iv) 6 years + (older dogs). Broad age categories were chosen due to potentially large differences in developmental timing between individuals. Age was categorised because we predicted that aggression would be dependent on these developmental periods.

Models included additional demographic variables (see Table 1) that may mediate the probability of aggression: body weight (average weight if multiple measurements were taken), total number of days spent at the shelter, the rehoming centre at which dogs were based (London, Old Windsor, Brands Hatch), neuter status (neutered before arrival, neutered at the shelter, not neutered) and source type (relinquished by owner, returned to the shelter after adoption, stray). Categorical variables were represented as sum-to-zero deflections from the group-level intercept to ensure the intercept represented the average probability of aggression across categorical levels. Weight and total days at the shelter were mean-centered and standardised by 2 standard deviations. Due to the potentially complex relationships between these variables and

aggression (e.g. interactive effects between neuter status and sex; [49]), which could also include violations of measurement invariance, we decided not to interpret their effects inferentially. Instead, they were included to make the assessment of measurement invariance between sexes and age groups conditional on variance explained by potentially important factors.

For comparability to other studies in animal personality, behavioural repeatability was calculated across contexts in each model using the intraclass correlation coefficient (ICC), calculated as $\frac{\sigma_\beta^2}{\sigma_\beta^2+\sigma_\epsilon^2}$, where σ_β^2 represented the between-individual variance of the probability of aggression (i.e. the variance of the random intercepts), and σ_ϵ^2 was $\pi^2/3$, the residual variance of the standard logistic distribution [65].

Computation

Models were computed using the probabilistic programming language Stan version 2.12 [66], using Hamiltonian Monte Carlo, a type of Markov Chain Monte Carlo (MCMC) algorithm, to sample from the posterior distribution (model code supplied in Supporting Information). Prior distributions for all independent variables were normal distributions with mean 0 and standard deviation 1, attenuating regression coefficients towards zero for conservative inference. The prior on the overall intercept parameter was normally distributed with mean 0 and standard deviation 5. The standard deviation of dog-

379 specific intercept parameters was given a half-Cauchy prior distribution with mean 0
380 and shape 2. Each model was run with 4 chains of 2,000 iterations with a 1,000 step
381 warm-up period. The Gelman-Rubin statistic (ideally < 1.05) and visual assessment of
382 traceplots were used to assess MCMC convergence and we checked the accuracy of the
383 posterior predicted probabilities of aggression against the raw data. Regression
384 coefficients were expressed as odds ratios and were summarised by their mean and 95%
385 Bayesian credibility interval (CI, i.e. the 95% most probable parameter values). To
386 compare levels of categorical variables and their interactions, we computed the 95% CI
387 of the differences between the respective posterior distributions.

388

389 ***Model selection & parameter inference***

390 Models were run on each imputed data set and their respective posterior distributions
391 were averaged to attain a single posterior distribution for inference. Adopting a
392 Bayesian approach allowed the estimation of interaction parameters (i.e. testing
393 measurement invariance) without requiring corrections for multiple comparisons as in
394 frequentist null hypothesis testing [67]. Nonetheless, models included a large number of
395 estimated parameters. Two strategies were employed to guard against over-fitting of
396 models to data. First, we selected the model with the best out-of-sample predictive
397 accuracy given the number of parameters based on the Widely Applicable Information
398 Criterion (WAIC; using the R package *loo* [68]). Four variants of each model were
399 computed: two-way interactions between contexts and age and contexts and sex,

respectively (model 1), a single interaction with sex but not with age (model 2), a single interaction with age but not with sex (model 3), and no interactions (model 4). All models included the mediating independent variables above. Second, to avoid testing point-estimate null hypotheses, the effect of a parameter was only considered credibly different from zero if the odds ratio exceeded the region of practical equivalence (ROPE; see [69]) around an odds ratio of 1 from 0.80 to 1.25. An odds ratio of 0.80 or 1.25 indicates a 20% decrease or increase (i.e. 4/5 or 5/4 odds), respectively, in the odds of an outcome, frequently used in areas of bioequivalence testing (e.g. [70]), which we here considered to be small enough to demonstrate a negligible effect in the absence of additional information. If a 95% CI fell completely within the ROPE, the null hypothesis of no credible influence of that parameter was accepted; if a 95% CI spanned the ROPE (i.e. included part of the ROPE), then the parameter's influence was left undecided [69].

Ethics statement

Permission to access and utilise the data was given by the shelter after signing a non-disclosure agreement. Approval from an ethical review board was not required for this study.

Data accessibility

The data used in this study are protected by a non-disclosure agreement. Researchers can access the data by contacting Battersea Dogs and Cats Home.

Results

Validity of behaviour recordings

For the video showing aggression towards people, 51.61% of participants identified the behaviour as aggression and 41.94% identified the behaviour as non-aggressive but reactive behaviour (see definitions above). For the video showing aggression towards dogs, 52.69% identified the behaviour correctly and 44.09% identified the behaviour as non-aggressive but reactive behaviour. For the 12 other videos not showing aggression, only 1 person coded a video as aggression towards people and 3 people coded videos as aggression towards dogs.

Structural equation models

The raw tetrachoric correlation matrix for aggression between contexts is presented in Table S2. Both structural equation models demonstrated excellent fit, with the model

435 with correlated latent variables fitting marginally better (CFI: 0.97; TLI: 0.96; RMSEA:
436 0.03) than the model with uncorrelated variables (CFI: 0.96; TLI: 0.95; RMSEA 0.04). All
437 regression coefficients of the model with correlated latent variables were positive and
438 significant (Table 3), and latent variables shared a significant positive correlation (ρ =
439 0.25; $p < 0.001$; Table 3).

440

441 Local independence

442 Allowing the pre-defined residuals to co-vary in the structural equation model resulted
443 in marginally better fit (CFI = 0.98; TLI = 0.97; RMSEA: 0.03). A significant negative
444 correlation was observed between *Handling* and *In kennel towards people* contexts (ρ =
445 -0.67; p = 0.003; Table 4), and *Handling* and *Interactions with unfamiliar people* contexts
446 (ρ = -0.55; p = 0.01; Table 4). Significant positive correlations were observed between
447 *Handling* and *Interactions with toys* contexts (ρ = 0.15; p = 0.04; Table 4), and *Out of*
448 *kennel towards people* with the *Interactions with unfamiliar people* context (ρ = 0.27; p
449 < 0.001; Table 4). No significant residual correlations between contexts reflecting
450 aggressiveness towards dogs were observed (Table 4).

451

452

Table 3. Structural equation model parameter estimates.

Parameter	Estimate	SE	*z* value	*p* value
Handling[a]	0.83	0.06	15.07	< 0.001
In kennel towards people[a]	1.31	0.09	15.00	< 0.001
Out of kennel towards people[a]	0.33	0.07	12.60	< 0.001
Interactions with familiar people[a]	0.99	0.07	14.81	< 0.001
Interactions with unfamiliar people[a]	1.58	0.12	12.97	< 0.001
Eating food[a]	0.73	0.06	13.02	< 0.001
Interactions with toys[a]	0.54	0.07	8.04	< 0.001
In kennel towards dogs[b]	0.75	0.06	11.74	< 0.001
Out of kennel towards dogs[b]	0.50	0.04	11.87	< 0.001
Interactions with female dogs[b]	0.94	0.07	11.19	< 0.001
Interactions with male dogs[b]	0.89	0.07	12.66	< 0.001
Correlation: People ~ Dogs	0.25	0.03	7.50	< 0.001

Parameter estimates from the best-fitting structural equation model, explaining aggression by two correlated latent variables.

[a] Contexts reflecting aggressiveness towards people

[b] Contexts reflecting aggressiveness towards dogs

453

454

Table 4. Tests of local independence.

Residual correlations	Estimate	SE	*z* value	*p* value
Handling ~ In kennel towards people[a]	-0.67	0.22	-3.01	0.003
Handling ~ Interactions with familiar people[a]	0.15	0.09	1.72	0.09
Handling ~ Interactions with unfamiliar people[a]	-0.55	0.21	-2.66	0.01
Handling ~ Interactions with toys[a]	0.15	0.07	2.06	0.04
Out of kennel towards people ~ Interactions with familiar people[a]	0.04	0.08	0.51	0.61
Out of kennel towards people ~ Interactions with unfamiliar people[a]	0.27	0.09	3.20	0.001
Interactions with familiar people ~ Interactions with unfamiliar people[a]	0.003	0.11	0.02	0.98
Out of kennel towards dogs ~ Interactions with female dogs[b]	-0.65	0.63	-1.04	0.30
Out of kennel towards dogs ~ Interactions with male dogs[b]	-0.41	0.67	-1.08	0.27
Interactions with female dogs ~ Interactions with male dogs[b]	-0.28	0.57	-0.50	0.62

Estimated residual correlations between *a priori* defined structural equation model parameters.

[a] Contexts reflecting aggressiveness towards people

[b] Contexts reflecting aggressiveness towards dogs

455

456 ## Measurement invariance

457 Separate models were run for contexts reflecting aggressiveness towards people and

458 aggressiveness towards dogs. For all models, all Gelman-Rubin statistics were all < 1.01,

459 effective sample sizes for all parameters were > 1000, and traceplots showed good
460 mixing. Posterior predictive checks of model estimates reflected the raw data (Figs 1
461 and 2). The full measurement invariance model (model 1) including interactions
462 between contexts and sex and contexts and age groups had the best out-of-sample
463 predictive accuracy for both the aggressiveness towards people and aggressiveness
464 towards dog models, respectively, illustrated by the lowest WAIC values (Table 5). Since
465 some models included numerous interactions, we provide an overall summary of the
466 main results below (Figs 1 and 2) with full parameter estimates provided in Tables S3
467 and S4.

Table 5. Bayesian hierarchical model selection using WAIC.

Model	Aggressiveness towards people	Aggressiveness towards dogs
Model 1	13405.6 ± 179.0	15257.2 ± 133.1
Model 2	13506.3 ± 179.6	15381.4 ± 133.4
Model 3	13426.3 ± 179.1	15285.3 ± 133.0
Model 4	13521.7 ± 179.5	15407.6 ± 133.4

Mean ± standard error of the Widely Applicable Information Criteria (WAIC) values per model and latent variable (aggressiveness towards people and dogs, respectively), used to assess measurement invariance: model 1 (age x context and sex x context interactions), model 2 (sex x context interaction only), model 3 (age x context interaction only), model 4 (no interactions). Models with the lowest WAIC values are estimated to have the best out-of-sample predictive accuracy.

468

469 **Aggressiveness towards people**

470 The odds of aggression towards people, across categorical predictors and for an average
471 dog of mean weight and length of stay at the shelter, were 0.022 (CI: 0.021 to 0.024), a

probability of approximately 2%. On average, aggression was most likely in the *In kennel towards people* context (OR = 0.054; CI: 0.049 to 0.058) and least probable in the *Interactions with toys* context (OR = 0.008; CI: 0.007 to 0.009).

Aggression was less likely across contexts for females than males (OR = 0.719; CI: 0.668 to 0.770), although there were also credible interactions between sex and contexts (Fig 1a; Table S3). Whereas males and females had similar odds of aggression in the *Out of kennel towards people* context, smaller differences were observed between *Out of kennel towards people* and *Handling* (OR = 0.578; CI: 0.481 to 0.682), *Eating food* (OR = 1.812; CI: 1.495 to 2.152) and *Interactions with familiar people* (OR = 1.798; CI: 1.488 to 2.126) contexts in females compared to males. Aggression in the *Interactions with unfamiliar people* context was also similar between males and females, while larger differences were observed between *Interactions with unfamiliar people* and *Handling* (OR = 0.616; CI: 0.530 to 0.702), *Eating food* (OR = 0.594; CI: 0.506 to 0.686) and *Interactions with familiar people* (OR = 0.598; CI: 0.513 to 0.687) contexts in females compared to males.

Apart from lower odds of aggression in 4 to 10 month olds compared to 10 month to 3 year old dogs (OR = 0.638; CI: 0.565 to 0.705), there was no simple influence of age group on aggressiveness. Between the 4 to 10 months old and 3 to 6 years old groups, differences between the odds of aggression across contexts varied due to an increase of

493 aggression in certain contexts but not others (Fig 1b; Table S4). Aggression in *In kennel*
494 *towards people* and *Interactions with unfamiliar people* contexts particularly increased,
495 leading to larger differences between, for example, *In kennel towards people* and *Eating*
496 *food* (OR = 0.524; CI: 0.400 to 0.642) and *Eating food* and *Interactions with unfamiliar*
497 *people* (OR = 1.721; CI: 1.403 to 2.059) contexts for 10 month to 3 year olds compared
498 to 4 to 10 month olds, and between *In kennel towards people* and *Out of kennel towards*
499 *people* (OR = 0.470; CI: 0.355 to 0.606) and *Out of kennel towards people* and
500 *Interactions with unfamiliar people* (OR = 2.051; CI: 1.608 to 2.543) contexts in 3 to 6
501 year olds compared to 4 to 10 month olds. In 3 to 6 year old compared to 10 month to 3
502 year old dogs, aggression increased in the *Handling* and *Eating food* contexts but
503 decreased in the *Out of kennel towards people* context, resulting in larger differences
504 between, for instance, *Handling* and *Out of kennel towards people* (OR = 0.526; CI: 0.409
505 to 0.631) and *Out of kennel towards people* and *Interactions with unfamiliar people* (OR
506 = 2.349; CI: 1.891 to 2.925), and smaller differences between *Eating food* and
507 *Interactions with familiar people* (OR = 0.576; CI: 0.468 to 0.687).

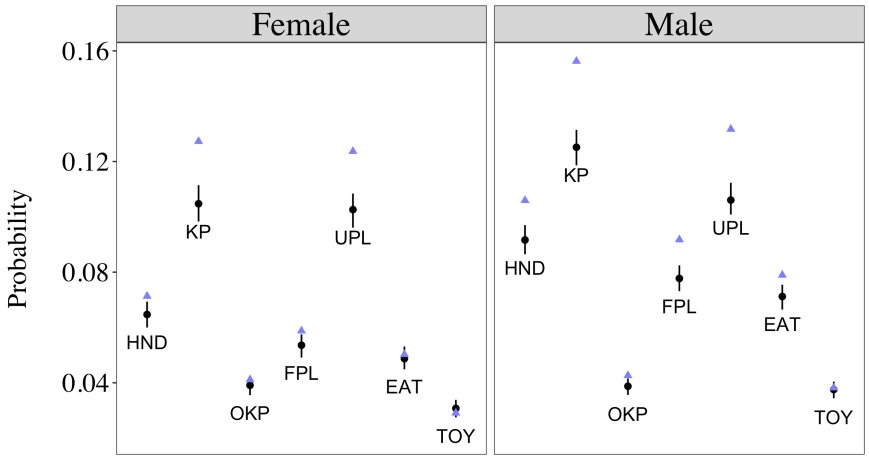
508

509 Dogs over 6 years old demonstrated qualitatively different response patterns across
510 certain contexts than all other age groups. While aggression was most probable in *In*
511 *kennel towards people* and *Interactions with unfamiliar people* contexts for dogs aged 4
512 months through 6 years, dogs over 6 years old were most likely to show aggression in
513 the *Handling* context, leading to interactions between, for example, *Handling* and *In*
514 *kennel towards people*, and between *Handling* and *Interactions with unfamiliar people*

515 contexts compared to the other age groups (Fig 1b; Table S3). Aggression when *Eating*
516 *food* and in *Interactions with toys* contexts also increased compared to that expressed
517 by younger dogs, resulting in credible differences between, for instance, *Eating food* and
518 *Interactions with familiar people* contexts between dogs aged 10 months to 3 years and
519 over 6 years (OR = 0.379; CI: 0.300 to 0.465) and between *Out of kennel towards people*
520 and *Interactions with toys* contexts between over 6 year olds and all other age groups
521 (Table S3).

522

a



b

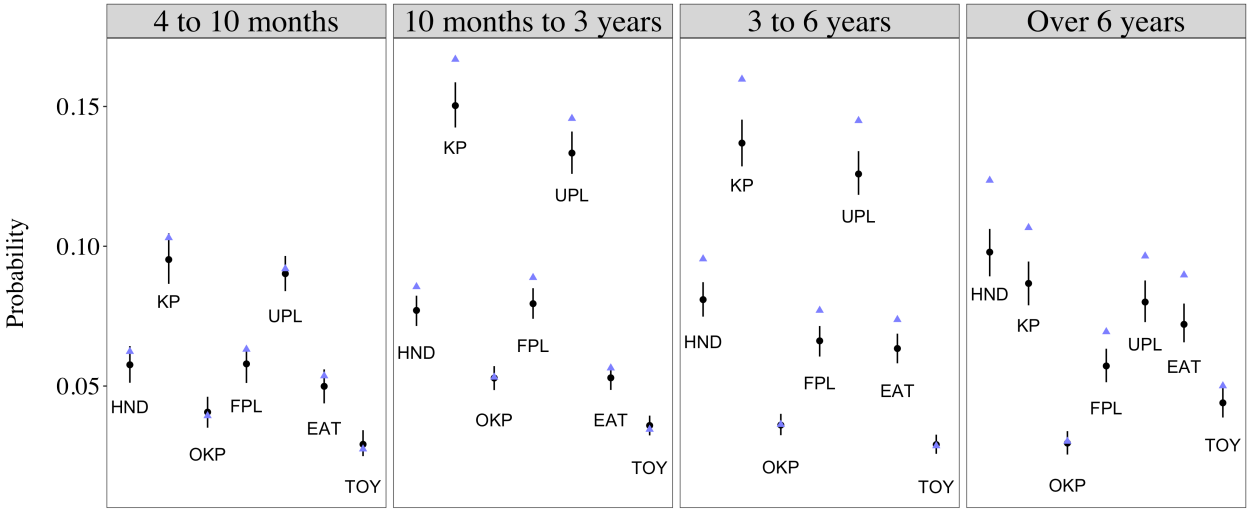


Fig 1. Predicted probabilities of aggression towards people in different contexts by sex (panel a) and age groups (panel b). Black points and vertical lines show mean and 95% credibility intervals of model parameter estimates; blue triangles show raw sample data. Abbreviations used in the figure: HND (*Handling*); KP (*In kennel towards people*); OKP (*Out of kennel towards people*); FPL (*Interactions with familiar people*); UPL (*Interactions with unfamiliar people*); EAT (*Eating food*); TOY (*Interactions with toys*).

523

524

525 **Aggressiveness towards dogs**

526 The odds of aggression towards dogs, across categorical predictors and for an average
527 dog of mean weight and length of stay at the shelter, was 0.176 (CI: 0.168, 0.184),
528 corresponding to a probability of approximately 15%. Dogs were most likely to show
529 aggression in the *Interactions with male dogs* context (OR = 0.297; CI: 0.198 to 0.217)
530 and least likely in the *In kennel towards dogs* context (OR = 0.099; CI: 0.094 to 0.104; Fig
531 2; Table S4).

532

533 No credible mean-level differences existed between females and males (OR = 1.187; CI:
534 1.128 to 1.250). However, the difference in aggression between the *Interactions with*
535 *female dogs* and *Interactions with male dogs* contexts was smaller for females (OR =
536 1.542; CI: 1.400 to 1.704; Fig 2a; Table S4), as were the differences between *Interactions*
537 *with male dogs* and *In kennel towards dogs* (OR = 0.661; CI: 0.590 to 0.732) and *In*
538 *kennel towards dogs* and *Out of kennel towards dogs* (OR = 1.420; CI: 1.269 to 1.587).
539 Females were also more likely to show aggression in *Interactions with female dogs* than
540 *Out of kennel towards dogs* compared to males (OR = 1.444; CI: 1.301 to 1.603).

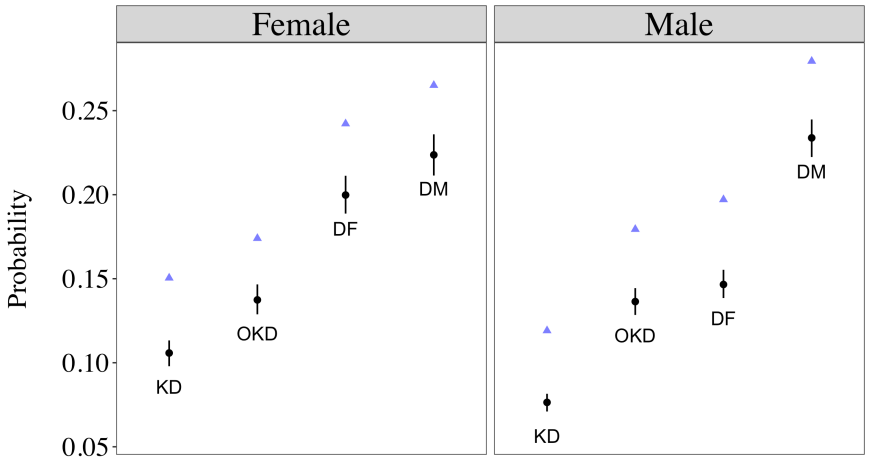
541

542 Dogs aged 4 to 10 months old had credibly lower odds of aggression towards dogs than
543 older dogs across contexts (Fig 2b; Table S4). However, contexts and age also showed
544 interactive effects. In particular, aggression in *Interactions with female dogs* and

545 *Interactions with male dogs* contexts tended to increase relative to other contexts. For
546 instance, the relationship between *Interactions with female dogs* and *Out of kennel*
547 *towards dogs* contexts reversed in direction between 4 to 10 month and 10 month to 3
548 year olds (OR = 0.595; CI: 0.495 to 0.688) as did the relationship between *Interactions*
549 *with male dogs* and *Out of kennel towards dogs* contexts (OR = 0.499; CI: 0.422 to
550 0.575). The relationship between *In kennel towards dogs* and *Out of kennel towards*
551 *dogs* contexts also changed across age groups (Fig 2b; Table S4). Four to 10 months old
552 were more likely to show aggression in *Out of kennel towards dogs* than *In kennel*
553 *towards dogs* contexts, but the difference was smaller in 10 months to 3 year olds (OR =
554 0.608; CI: 0.505 to 0.728) and in over 6 year olds (OR = 0.396; CI: 0.316 to 0.481). The
555 latter relationship was reversed in 3 to 6 year olds compared to 4 to 10 month old dogs
556 (OR = 0.277; CI: 0.227 to 0.331) and 10 month to 3 year old dogs (OR = 0.456; CI: 0.396
557 to 0.516).

558

a



b

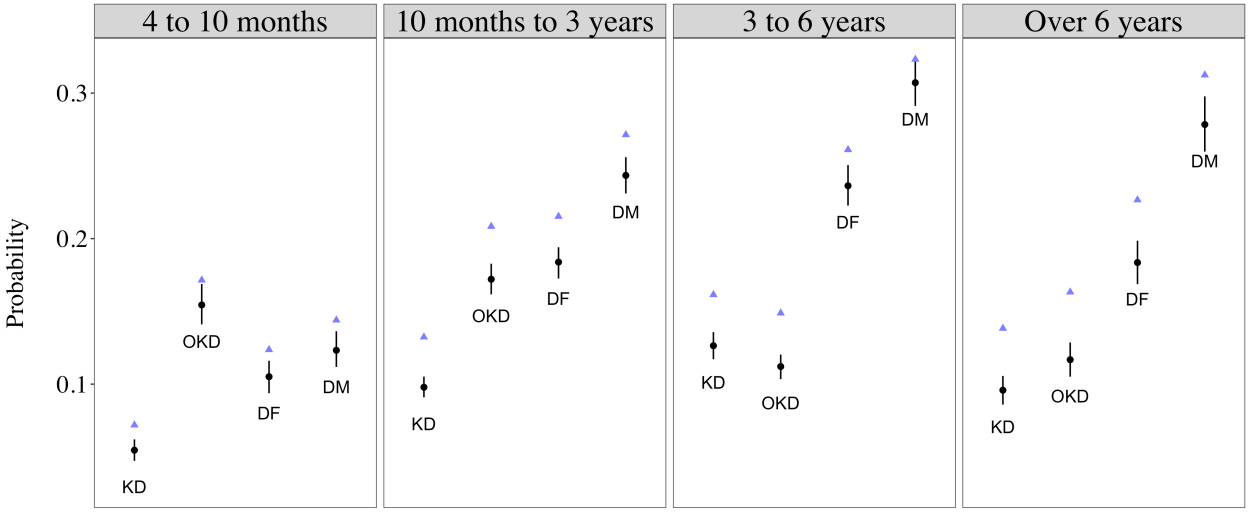


Fig 2. Predicted probabilities of aggression towards dogs in different contexts by sex (panel a) and age groups (panel b). Black points and vertical lines show mean and 95% credibility intervals of model parameter estimates; blue triangles show raw sample data. Abbreviations used in the figure: KD (*In kennel towards dogs*); OKD (*Out of kennel towards dogs*); DF (*Interactions with female dogs*); DM (*Interactions with male dogs*).

559

560

Repeatability

Both aggressiveness towards people and dog showed moderate repeatability across contexts (ICC_{people} = 0.479; CI: 0.466, 0.491; ICC_{dogs} = 0.303; CI: 0.291, 0.315), although aggressiveness towards people was more repeatable than aggressiveness towards dogs ($ICC_{difference}$ = 0.176; CI: 0.158, 0.192).

Discussion

In this study, we have examined whether local independence and measurement invariance hold for hypothesised latent aggressiveness traits in shelter dogs. Observational recordings of aggression directed towards people and dogs across different shelter contexts were explained by two positively correlated latent variables, and behaviour across contexts was moderately repeatable. These results are consistent with the definition of animal personality as behaviour that shows moderately consistent between-individual differences across time or contexts [4], and characterised by multiple observed behaviours being decomposed into lower-dimensional behavioural traits. Yet, subsequent investigations indicated violations of local independence and measurement invariance, questioning the validity of the latent variables as homogeneous personality traits. While a number of factors may contribute to the low predictive validity of certain dog personality assessments [48], ensuring the robustness of inferences made about personality traits is critical, especially considering the large

number of traits that have been proposed [34,35]. Given the popularity of latent variable models, such as exploratory and confirmatory factor analysis (e.g. [17,29,71]), to understand the organisation of personality in dogs, ascertaining local independence and measurement invariance of personality traits should be routine practice, as it has become in human (personality) psychology (e.g. [26,72–74]).

Local independence implies that the association between manifest variables is greater than that explained by the latent variable. Here, local independence was investigated between manifest variables that were believed to have close temporal-spatial relationships. While local independence was confirmed for contexts reflecting aggressiveness towards dogs, contexts reflecting aggressiveness towards people showed a violation of local independence. Aggression in the *Handling* context shared a negative residual correlation with both the *In kennel towards people* and *Interactions with unfamiliar people* contexts, while positive residual correlations were present between *Handling* and *Interactions with toys*, and *Out of kennel towards people* and *Interactions with unfamiliar people* contexts (Table S4). Violations of local independence may arise through shared method variance [75–78] or unmodelled latent variables influencing manifest variables [79,80]. For example, if a dog showed aggression when an unfamiliar person approached, it may be less likely to be handled by that person, inducing a negative residual correlation conditional on latent levels of aggression as was observed here between the *Handling* and *Interactions with unfamiliar people* contexts. Likewise, the *Handling* context and the *Interactions with toys* contexts are similar in that both

603 require dog and person to be close together, and interacting with toys and handling may
604 co-occur at the same time, inducing a positive residual correlation between contexts.
605
606 While authors have argued that greater standardisation and validation of personality
607 assessments is key to ensuring the accurate measurement of underlying traits [34,46], it
608 may be untenable to avoid dependencies between testing contexts. Displays of
609 aggression in one sub-test will likely change how people conduct future sub-tests with
610 the same dog, regardless of test standardisation. Moreover, the hierarchical structure of
611 animal personality, including the presence of behavioural syndromes, makes the
612 isolated measurement of one trait unrealistic [75]. For instance, the positive residual
613 correlation between *Out of kennel towards people* and *Interactions with unfamiliar*
614 *people* may be mediated by additional traits of interest to personality researchers, such
615 as fearfulness or anxiety [29,81], if dogs who are fearful of interacting with unfamiliar
616 people are more likely to show aggression beyond that described by a latent
617 aggressiveness trait. Some human psychologists have argued that violations of local
618 independence are a natural consequence of the organisation of behaviour as a complex
619 dynamic system [82,83], which unfolds with respect to time- and context-dependent
620 constraints [84]. Thus, awareness of local independence and its violation could facilitate
621 closer understanding of the dynamics driving personality test responses beyond
622 explanations purely based on personality traits.
623

624 While different subsets of a population may differ in mean levels of trait expression,
625 interactions between behavioural responses and those subsets indicate that the same
626 phenomenon is not under measurement across groups [23,24]. Using a Bayesian
627 hierarchical model analogous to the Rasch model, we found that the probability of
628 aggression across contexts was still dependent on sex and age conditional on latent
629 levels of aggressiveness towards people and dog (Figs 1 and 2; Tables S3 and S4),
630 indicating the violation of measurement invariance. Female dogs, for example, were
631 more likely than males to show aggression in *Out of kennel towards people* and
632 *Interactions with unfamiliar people* contexts relative to other contexts (Fig 1a). Females
633 also demonstrated similar odds of aggression during *Interactions with female dogs* and
634 *Interactions with male dogs*, whereas males were more likely to show aggression
635 towards male than female dogs (Fig 2a). Thus, latent levels of aggressiveness did not
636 easily explain differences in aggression across contexts. As with local independence,
637 different behavioural variables unaccounted for in this study may also result in
638 violations of measurement invariance. For instance, while dogs up to 6 years old were
639 most likely to show aggression in *In kennel towards people* and *Interactions with*
640 *unfamiliar people contexts*, dogs over 6 years old demonstrated aggression most
641 commonly in the *Handling* context, which may reflect an increase in pain-motivated
642 aggression. Dogs over 6 years old also showed an increase in aggression in the *Eating*
643 *food* and *Interactions with toys* contexts relative to other age groups, suggesting that
644 older dogs in shelter populations may be less tolerant during close interactions with

645 people (i.e. handling, people in the vicinity of their food and toys) compared to other
646 contexts.

647

648 Investigating factors that predict a dog's personality, or whether a dog's personality
649 predicts other outcomes of interest (e.g. future behavioural scores; [39,40]), is also of
650 substantive interest to researchers. Persson *et al.* [32] found interactive effects between
651 sex and age on human-directed social behaviour traits. Moreover, Asp *et al.* [56] found
652 that sex and age interacted with breed in explaining differences in C-BARQ trait scores.
653 To ensure the robustness of their conclusions, however, researchers should ensure that
654 the traits being investigated satisfy model assumptions, not only for statistical accuracy,
655 but so that interpreting the difference between individuals as a function of trait scores is
656 both feasible and meaningful.

657

658 Local independence and measurement invariance are assumptions of confirmatory,
659 reflective latent variable models, but are not required assumptions for formative
660 models, which posit that the latent variable is simply a linear composite of manifest
661 variables, rather than a causal, underlying variable [20]. While formative models such as
662 principal components analysis (e.g. [32,33,85]) may, as a result, appear attractive and
663 continue to be used in dog personality studies, their use has been discouraged. Principal
664 components analysis will always result in lower-dimensional variables comprised of
665 linear combinations of manifest variables, even when those manifest variables are

666 uncorrelated random variables (e.g. see [12]). Consequently, finding principal
667 components that underlie behavioural data is neither surprising nor evidence for the
668 discovery of domain-general personality traits. Crucially, for personality traits to be of
669 use in understanding the organisation of dog behaviour or be considered as predictors
670 of future dog behaviour, they should, arguably, hold causal status. In human psychology,
671 Schimmack [86] and Borsboom [6] note the importance of interpreting latent variables
672 as causal variables for understanding unobserved constructs, such as personality traits.
673 To this end, reflective models, especially confirmatory approaches such as structural
674 equation modelling or item response theory, present a more powerful framework to
675 distinguish signal from noise in multivariate behavioural data [20,86] and are concurrent
676 with a theoretical interpretation of personality traits as causal variables underpinning
677 animal behaviour [16]. Increasing the popularity of such approaches could be
678 particularly helpful in evaluating the reproducibility of dog personality traits across
679 existing studies.

680

681 Although we have identified violations of both local independence and measurement
682 invariance, we remain cautious about hypothesising *a posteriori* about their causes. A
683 problem for the wider perspective of animal personality research is that personality
684 traits are typically defined operationally, based on the statistical repeatability of
685 quantifiable behaviour [77,87,88]. As has been discussed in human personality
686 psychology, operational definitions are ontologically ambiguous [89,90]. That is, while
687 operational definitions facilitate experimentation in animal personality [4], they are

688 ambiguous with respect to the biological mechanisms underlying trait expression. For
689 example, Budaev and Brown remark that boldness, defined as a propensity to take risks,
690 could encompass a range of distinct personality traits, each with a different biological
691 basis [75]. Whilst reflective latent variable models allow researchers to test hypotheses
692 about the relatedness of measured behaviours via one or more underlying traits, they
693 have also been criticised as ambiguous [81]. For example, it is uncertain what reflective
694 latent variables may represent in biological organisation [89] or even whether they are
695 features individuals possess or simply emergent features of between-individual
696 differences [91,92]. Such considerations highlight the importance of research on the
697 proximate mechanisms of personality [87] and longitudinal data analyses to separate
698 between- from within-individual behavioural variation [93,94].

699

700 A number of authors have emphasised the poor predictive value of aggression tests in
701 shelter dogs compared to tests of other traits [37,39,48]. The low occurrence of
702 aggression can make its accurate measurement difficult [39], and some studies actively
703 exclude dogs that have shown aggression in the shelter (e.g. [39]). The probability of
704 observing aggression recorded in this study was low, especially in contexts reflecting
705 aggressiveness towards people (Fig 1). Nonetheless, evaluation of the validity of the
706 behavioural recordings indicated that shelter employees might mistake observations of
707 aggression for non- aggressive responses (e.g. over-excitement and frustration when
708 seeing people/dogs), meaning that the true probability of aggression was potentially
709 under-estimated (although incorrectly coding other behaviours as aggression also

occurred, albeit rarely). Infrequent occurrence and/or recording of aggression may also limit accurate predictions of future behaviour. Patronek and Bradley [48] demonstrate using simulation that the low prevalence of aggression inflates the chance that aggression shown in a shelter assessment represents a false positive. In general, our results support this conclusion in the sense that aggression may be shown differentially across contexts not explained by latent levels of aggressiveness. Violations of local independence and measurement invariance as found here indicate, further, that it is not only the difference between false and true positives and negatives, but the validity of inferring homogeneous personality traits by which to compare individual dogs, that needs careful consideration. Consequently, we agree with recommendations to establish the efficacy of longitudinal, observational assessments rather than relying on a single assessment made using a traditional test battery [31,39,48]. This approach will prioritise the cumulative understanding of a dog's context-dependent behaviour and help to guide decisions about the potential risk a dog poses to humans and other animals.

Conclusion

This study has tested the assumptions of local independence and measurement invariance of personality traits in shelter dogs. Using structural equation modelling, aggression across behavioural contexts was explained by two, correlated latent variables and demonstrated repeatability. Nevertheless, significant residual correlations

731 remained between certain behavioural contexts related to aggressiveness towards
732 people, violating the assumption of local independence. In addition, aggression in
733 different contexts showed differential patterns of response across sex and age,
734 indicating a lack of measurement invariance. Violations of local independence and
735 measurement invariance imply that aggressiveness towards people and dogs may not
736 be enough to explain patterns of aggression in different contexts, or that inferences
737 based on these hypothesised personality traits may in fact be misleading. We encourage
738 researchers to more closely assess the measurement assumptions underlying reflective
739 latent variable models before making conclusions about the effects of, or factors
740 influencing, personality.

741

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744 data for this study.

745

746 References

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Supporting Information

Table S1. Counts of aggression per context. The number of dogs who had 0, 1, and > 1 counts of aggression.

Table S2. Tetrachoric correlations between aggression contexts. Tetrachoric correlations between aggression contexts on the raw binary data, before the multiple imputation. Abbreviations used: HND (*Handling*); FPL (*Interactions with familiar people*); UPL (*Interactions with unfamiliar people*); KD (*In kennel towards dogs*); KP (*In kennel towards people*); OKD (*Out of kennel towards dogs*); OKP (*Out of kennel towards people*); EAT (*Eating food*); TOY (*Interactions with toys*); DM (*Interactions with male dogs*); DF (*Interactions with female dogs*).

Table S3. Bayesian hierarchical model parameter estimates for aggression towards people in different contexts. Mean and 95% credibility interval (CI) estimates for all parameters from the Bayesian hierarchical logistic model assessing measurement invariance for contexts reflecting aggressiveness towards people. Differences between levels of categorical variables are indicated by '.v.' in the parameter name; interactions are denoted with '*' in the parameter name. The decision rule for each parameter is given except for those variables not interpreted inferentially: YES = 95% CI falls

completely outside the region of practical equivalence (ROPE); NULL = 95% CI falls completely inside the ROPE; ROPE = 95% CI partly covers the ROPE.

Table S4. Bayesian hierarchical model parameter estimates for aggression towards dogs in different contexts. Mean and 95% credibility interval (CI) estimates for all parameters from the Bayesian hierarchical logistic model assessing measurement invariance for contexts reflecting aggressiveness towards dogs. Differences between levels of categorical variables are indicated by '.v.' in the parameter name; interactions are denoted with '*' in the parameter name. The decision rule for each parameter is given except for those variables not interpreted inferentially: YES = 95% CI falls completely outside the region of practical equivalence (ROPE); NULL = 95% CI falls completely inside the ROPE; ROPE = 95% CI partly covers the ROPE.