Very low depth whole genome sequencing in complex trait association studies

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Abstract

Background

Very low depth sequencing is a cost-effective approach to capture low-frequency and rare variation in complex trait association studies. Here, we perform cohort-wide whole genome sequencing (WGS) at 1x depth coupled to genome-wide association analysis in 2,347 individuals from two isolated populations.

Results

We establish a robust pipeline for calling 1x WGS data, achieving an average minor allele concordance of 97% when compared to genotyping chip data. 9.5% of variants called using 1x WGS are variants with a high predicted quality not captured by genome-wide association study (GWAS) data in the same individuals imputed to a dense haplotype reference panel. Of the 54 association signals arising from genome-wide association analysis of 1x WGS variants with 25 haematological traits (at $p<5\times10^{-7}$), only 57% are recapitulated by the imputed GWAS results in the same samples. Differences in strength of evidence for association are smaller for common than for low-frequency and rare variant signals. We further exemplify power gains by establishing robust evidence for a novel association between rs6489858, an intronic variant in *RPH3A* and increased lymphocyte count (beta=0.13, SE=0.11, $p=8\times10^{-12}$), which replicates in an independent dataset comprising 173,480 samples.

Conclusions

We show that 1x WGS is an efficient alternative to imputed GWAS chip designs for empowering next-generation association studies in complex traits. We demonstrate that population-scale 1x WGS allows the interrogation of a large number of low-frequency and

rare variants missed by classical GWAS array imputation, resulting in potential higher

association power.

Keywords

Whole-genome sequencing, association studies, population isolates

Introduction

Characterisation of the genetic determinants underpinning complex human traits of medical

relevance can help improve our understanding of aetiopathology and point to biological

processes amenable to intervention. Despite great progress in identifying common-

frequency variants with small to modest effect sizes, the allelic architecture of low-

frequency and rare variants for complex traits remains largely unchartered. Power to detect

association is central to genetic studies examining sequence variants across the full allele

frequency spectrum. Whole genome sequencing (WGS)-based association studies hold the

promise of probing a larger proportion of sequence variation compared to genome-wide

genotyping arrays. However, high-depth WGS costs do not yet allow application of the

GWAS paradigm to large-scale sequencing of hundreds of thousands of individuals. As

sample size and haplotype diversity are more important than sequencing depth in

determining power for association studies [1], low-depth WGS has emerged as an

alternative, cost-efficient approach to capture low-frequency variation in large studies.

Improvements in calling algorithms have enabled robust genotyping using WGS at low

depth (4x-8x), leading to the creation of large reference haplotype panels [2, 3], and to the

start of WGS-based association studies [4, 5]. Very low depth (<2x) sequencing has been

proposed as an efficient way to further improve the cost efficiency of sequencing-based

association studies. Simulations have shown that in whole-exome studies, extremely low

sequencing depths (0.1-0.5x) are effective in capturing single-nucleotide variants (SNVs) in

the common (MAF>5%) and low-frequency (MAF 1-5%) categories compared to imputed

GWAS arrays [6]. The CONVERGE consortium demonstrated the feasibility of such

approaches through the first successful case-control study of major depressive disorder in

4,509 cases and 5,337 controls [7].

Studying founder populations can further empower the search for association signals by

allowing the detection of population-specific variants, and of association signals at variants

that have drifted up in frequency compared to cosmopolitan populations, against the

backdrop of a homogeneous environment [8] [9] [10] [11]. Here, we perform very low depth

(1x), cohort-wide WGS in two isolated populations from Greece. We establish a robust 1x

WGS calling pipeline and compare the complement of variants captured to imputed GWAS

in the same samples. As a proof-of-principle, we perform association analysis across

medically-relevant haematological traits and identify a robustly-replicating novel locus

implicated in lymphocyte counts.

Results

As part of the Hellenic Isolated Cohorts (HELIC) study, we whole genome sequenced 990

individuals from the Minoan Isolates (HELIC-MANOLIS) cohort, and 1108 individuals from

the Pomak villages (HELIC-Pomak) at 1x depth, on the Illumina HiSeq2000 platform. In

addition, 249 samples from the MANOLIS cohort were sequenced at 4x depth [12].

Imputation-based genotype refinement was performed on the two cohort-wide datasets

using a combined reference panel of 10,422 haplotypes from MANOLIS 4x WGS, the 1000

Genomes [2] and UK10K [4] projects.

Variant calling pipeline

We established a variant calling, quality control (QC) and genotype refinement pipeline for

very low depth WGS by benchmarking nine pipelines that make use of state of the art

bioinformatics tools (Methods). Our optimised approach allowed the capture of 80% of true

low-frequency (MAF 1-5%) variants and 100% of true common-frequency (MAF>5%) SNVs

prior to imputation-based refinement, when compared to variants present on the Illumina

OmniExpress and HumanExome chips genotyped in the same samples. In order to assess

sensitivity and specificity of SNV calls pre-imputation, we estimate the false positive and

false negative rate by comparing 1x WGS variant calls with high-depth WGS data (see

Methods). We estimate that 12% of 1x sites with at least one heterozygote call are false

positives, whereas 24.6% of the sites called with high-depth WGS are not recapitulated

using 1x data.

In order to improve the false negative rate and genotype accuracy, we performed genotype

refinement and imputation using a large reference panel containing haplotypes from 4,873

cosmopolitan samples as well as the phased haplotypes from the 249 MANOLIS samples

sequenced at 4x depth. After imputation and QC, we captured 95% of rare, 99.7% of low-

frequency and 99.9% of common variants compared to the Illumina OmniExpress and

HumanExome GWAS chips, with an average minor allele concordance of 97% across the

allele frequency spectrum (Methods and Figure 1). By comparing 1x calls with those

produced by whole-exome sequencing in 10 individuals across both cohorts, we estimate a

false-positive rate of 2.4% post-imputation in the coding parts of the genome (Methods).

Comparison of variant call sets with an imputed GWAS

The genotype refinement and imputation step yielded 30,483,136 and 29,740,259 non-

monomorphic SNVs in 1,239 MANOLIS and 1,108 Pomak individuals, respectively. The

number of variants discovered using 1x WGS is nearly twice as high as that from array-based

approaches. In a subset of 982 MANOLIS individuals with 1x WGS, we called 25,673,116

non-monomorphic SNVs using 1x WGS data, compared to 13,078,518 non-monomorphic

SNVs in the same samples with OmniExpress and ExomeChip data imputed up to the same

panel [9]. The main differences are among rare variants (MAF<1%) (Figure 2): 13,671,225

(53.2%) variants called in the refined 1x WGS are absent from the imputed GWAS, 98% of

which are rare. 82% of these rare unique SNVs are singletons or doubletons, and therefore

9.5% of all variants called in the 1x WGS dataset were unique variants with MAC>2.

Experimental validation of genotypes

We performed experimental genotyping of 65 variants (23 common, 18 low-frequency and

24 rare) in a subset of 1087 and 859 samples in the MANOLIS and Pomak cohorts,

respectively, using the Agena Biosciences MassARRAY technology. On average, minor allele

concordance was 76% and positive predictive value was 82%. As expected, these values

differ between MAF categories (Additional File 1: Table S1). Minor allele and genotype

concordance between 1x calls and this set of directly assayed genotypes were in line with

those computed genome-wide between 1x calls and GWAS data (Figure 1).

Association analysis

As a proof of principle, we performed single-point association analysis across 25 haematological traits with 14,948,665 and 15,564,905 variants with MAC>2 in MANOLIS and Pomak, respectively (Methods). We used an empirical genetic relatedness matrix calculated on high-confidence genotypes to account for relatedness within the two isolated cohorts.

Genome-wide significance was set at $p<1.0x10^{-9}$ based on the effective number of traits and tested variants (see Methods). In the discovery sample, one association met this threshold in the Pomak dataset. rs35004220, located in an intron of the haemoglobin B (*HBB*) gene, was associated with six red blood cell traits (haemoglobin, mean corpuscular haemoglobin, mean corpuscular volume, red cell distribution width in volume and percent, red blood cells) (Additional File 1 and 2: Figure S1 and Table S2). A total of 5,090 variants with association $p<1.0x10^{-5}$ in the discovery stage corresponding to 556 and 465 independent signals in the MANOLIS and Pomak cohorts, respectively (Additional File 1: Table S2 and S3), were carried forward to *in silico* replication using data from a large meta-analysis of 173,480 samples from the UK Biobank and INTERVAL studies [5]. Out of the 3,336 variants for which replication data were available, 52.4% had a concordant direction of effect compared to the discovery stage ($p=2.9x10^{-3}$, one-sided binomial test). Upon meta-analysis of the discovery and replication data, we identify a previously unreported, genome-wide significantly associated signal (Figure 3). The G-allele of rs6489858 (EAF=0.40) at 12q24.13 is associated with increased lymphocyte count in MANOLIS (beta=0.022, SE=0.004, $p=2.29x10^{-9}$ in the discovery and replication meta-analysis). We found evidence of heterogeneity (I^2=95.4%, Q-statistic $p=2x10^{-6}$) and therefore applied a random effects meta-analysis model [13]

(beta=0.13, SE=0.11, p=8x10^{-12}). rs6489858 is located in an intronic region of the rabphilin

3A (*RPH3A*) gene, which encodes a peripheral membrane protein involved in protein

transport and synaptic vesicle traffic.

Comparison of association summary statistics with imputed GWAS

1x WGS calls a larger number of variants than imputed GWAS of the same samples. To

evaluate how this difference affects association study power, we compared the association

results of all independent suggestive signals at p<5x10^{-7} from the 1x WGS with the imputed

GWAS results for the same variants (Figure 4). Among the 54 variants significantly

associated at this threshold in the 1x WGS, 17 (31%) were not observed in the imputed

GWAS study. Rare (MAF<1%) variant signals are more poorly captured by the imputed

GWAS, with 10 out of 16 signals (62.5%) being missed, however, for 12 (70%) of the missed

variants, a tagging SNV at r^2>0.8 was available in the imputed GWAS. For the signals where

imputed GWAS results are indeed present, the majority (62.2%) do not meet our

significance threshold, an effect which is more marked in the rare and low-frequency

(16/23, 69%) than in the common (7/14, 50%) signals. This observation persists when

considering tagging variants at r^2>0.8: twenty-seven (55%) out of the 49 taggable 1x WGS

signals have a p-value above the significance threshold. Generally, differences in p-values

between the two studies were smaller for common than for low-frequency and rare variant

signals (5.5 and 2.0 on the log-scale, p=6x10^{-3}, two-sample t-test).

Discussion

In this work, we empirically demonstrate the relative merits of very low depth WGS both in

terms of variant discovery and association study power for complex quantitative traits

compared to GWAS approaches. However, the advantages of 1x WGS have to be weighed

against compute and financial cost considerations. As of January 2017, 1x WGS on the HiSeq

4000 platform was approximately half of the cost of a dense GWAS array (e.g. Illumina

Infinium Omni 2.5Exome-8 array), 1.5 times the cost of a sparser chip such as the Illumina

HumanCoreExome array, and a third of the cost of WES at 50x depth. By comparison, 30x

WGS was 21 or 16 times more costly depending on the sequencing platform (Illumina HiSeq

4000 or HiSeqX, respectively). The number of variants called by 1x WGS is lower than high-

depth WGS, but is in the same order of magnitude, suggesting comparable disk storage

requirements for variant calls. However, storage of the reads required an average 650Mb

per sample for CRAMs, and 1.3Gb per sample for BAMs.

Genome-wide refinement and imputation of very low depth WGS generates close to 50

times more variants than a GWAS chip. The complexity of the imputation and phasing

algorithms used in this study is linear in the number of markers, linear in the number of

target samples and quadratic in the number of reference samples [14], which results in a 50-

fold increase in total processing time compared to an imputed GWAS study of equal sample

size. Therefore, parallelisation plays a crucial role in managing computational load. For

example, in MANOLIS the genome was divided in 13,276 chunks containing equal number of

SNVs, which took an average of 31 hours each to refine and impute. The total processing

time was 47 core-years (Methods and Additional File 2: Figure S2). Parallelisation allowed

processing the 1,239 MANOLIS samples in under a month.

As a proof of principle, we used 1x WGS in samples from isolated populations and identified

a novel association with lymphocyte count, previously missed by large-scale GWAS in

cosmopolitan populations [15-17]. The signal had a much larger effect size in the isolated

population discovery cohort (beta=0.25 standard deviation increase in the discovery

samples, beta=0.02 in the replication cohorts for rs6489858). In the subsample of 1,225

individuals with both 1x and GWAS data, minor allele concordance was 99.5% for

rs6489858, and the association p-value with LYM was in the same order of magnitude

($p=7.5 \times 10^{-7}$ in the imputed GWAS, $p=3.2 \times 10^{-7}$ in the 1x WGS data). rs6489858 is located

260kb from *PTPN11*. In juvenile myelomonocytic leukemia, the RAS/MAPK pathway is

frequently deregulated due to somatic mutations in *PTPN11* [18]. *PTPN11* is also involved in

LEOPARD syndrome, metachondromatosis and Noonan syndrome. Animal models of this

gene show diverse and severe phenotypes including hematopoietic abnormalities such as

abnormal leukopoiesis [19]. *DTX1* is located 280kb away from the index variant. Deltex-1,

the cytoplasmic protein product of this gene, is a regulator of the Notch pathway and plays

an important role in the development of B and T lymphocytes [20]. Animal models of this

gene show various hematopoietic and immune related abnormalities [21] due to interfering

with T cell development.

Conclusions

We show that very low depth whole-genome sequencing allows the accurate assessment of

most common and low-frequency variants captured by imputed GWAS designs and achieves

denser coverage of the low-frequency and rare end of the allelic spectrum, albeit at an

increased computational cost. This allows very low depth sequencing studies to identify

signals also discoverable by imputed chip-based efforts, and to discover significantly

associated variants missed by GWAS imputation [22]. As sequencing technologies continue

to evolve, higher sequencing depths will provide accurate genotyping across the full range

of the allelic spectrum, enabling comprehensive exploration of human phenotype

associations through rare variant aggregation tests.

Materials and methods

Cohort details

The HELIC (Hellenic Isolated Cohorts; www.helic.org) MANOLIS (Minoan Isolates) collection

focuses on Anogia and surrounding Mylopotamos villages on the Greek island of Crete. All

individuals were required to have at least one parent from the Mylopotamos area to enter

the study. The HELIC Pomak collection focuses on the Pomak villages, a set of isolated

mountainous villages in the North of Greece. Recruitment of both population-based

samples was primarily carried out at the village medical centres. The study includes

biological sample collection for DNA extraction and lab-based blood measurements, and

interview-based questionnaire filling. The phenotypes collected include anthropometric and

biometric measurements, clinical evaluation data, biochemical and haematological profiles,

self-reported medical history, demographic, socioeconomic and lifestyle information.

Sequencing

Sequencing and mapping for the 995 MANOLIS samples at 1x depth has been described

before [22], as well as for 250 MANOLIS samples at 4x [9]. 1166 samples from the Pomak

were sequenced at 1x depth using the same protocol using Illumina HiSeq 2000 and Illumina

HiSeq 2500 sequencers. For comparison, 5 samples from each cohort were whole-exome

sequenced at an average depth of 75x.

Read mapping and variant calling

Following generation of raw reads on the Illumina HiSeq 2000 and HiSeq 2500 sequencing

machines, reads were converted from BCL format to BAM format using the Illumina2BAM

(https://github.com/wtsi-npg/illumina2bam) software. Illumina2BAM was again used to de-

multiplex lanes that had been sequenced so that the tags were isolated from the body of

the read, decoded, and could be used to separate out each lane into lanelets containing

individual samples from the multiplex library and the PhiX control. The quality scores were

then recalibrated using the purity recalibration algorithm [23] using the PhiX data for

reference. Read mapping was then carried out using the BWA backtrack algorithm version

0.5.10 using the GRCh37 1000 Genomes phase III reference (also known as hs37d5). PCR

and   optically   duplicated   reads   were   marked   using   Picard   MarkDuplicates

(http://broadinstitute.github.io/picard).

In order to ensure the quality of the large quantity of BAMs produced for the project, an

automatic quality control system was used to reduce the number of data files that required

manual intervention. This system was derived from the one originally designed for the

UK10K project (http://www.uk10k.org) and used a series of empirically derived thresholds

to assess summary metrics calculated from the input BAMs. These thresholds included:

percentage of reads mapped; percentage of duplicate reads marked; various statistics

measuring INDEL distribution against read cycle and an insert size overlap percentage. Any

lane that fell below the "fail" threshold for any of the metrics were excluded; and any lane

that did not fall below these thresholds for any of the metrics was given a status of "pass"

and allowed to proceed into the later stages of the pipeline.

Passed lanelets were then merged into BAMs corresponding to the libraries for each sample

and duplicates were marked again with Picard MarkDuplicates after which they were then

merged into BAMs on a per sample basis. Finally sample level bam improvement was

carried out using GATK 1.6[24, 25] and samtools[26] from git commit

72d6457f7f361c323f62bd2d3170980132ba2113. This consisted of re-alignment of reads

around known and discovered INDELs followed by base quality score recalibration both

using the GATK, lastly samtools calmd was applied and indexes were created. Known

INDELs for realignment were taken from Mill Devine and 1000G Gold set and the 1000G

phase low coverage set both part of the Broad's GATK resource bundle version 2.2. Known

variants for BQSR were taken from dbSNP 137 also part of the Broad's resource bundle.

The input BAM files are fed into samtools mpileup to create all-sites BCF files, which are

piped into bcftools view to create variant-only VCF files containing genotype calls. We split

the genome into chunks of 100,000 base pairs, and separate these chunks into SNV and

INDEL files. We run GATK UnifiedGenotyper to calculate site-level annotations.

Variant filtering

Variant quality score recalibration was performed using GATK VQSR v.3.1.1. However, using

the default parameters for the VQSR mixture model yields poor filtering, with a Ti/Tv ratio

dropoff at 83% percent sensitivity and a Ti/Tv ratio of 1.8 for high-quality tranches

(Additional File 2: Figure S3.a). We therefore ran exploratory runs of VQSR across a range of

values for the model parameters, using the dropoff point of the transition/transversion

(Ti/Tv) ratio below 2.0 as an indicator of good fit (Additional File 2: Figure S4). A small

number of configurations outperformed all others, which allowed us to select an optimal set

of parameters. For the chosen set of parameters, false positive rate is estimated at 10%±5%

(Additional File 2: Figure S3.b). Indels were excluded from the dataset out of concerns for

genotype quality. We found that the version of VQSR, as well as the annotations used to

train the model, had a strong influence on the quality of the recalibration (Additional File 2:

Figure S4 and Supplementary Text).

Comparison with Platinum genomes

For quality control purposes, reads from 17 of the well-characterised Platinum Genomes

sequenced by Illumina at 50x depth [3], and downsampled to 1x depth using samtools [26]

were included in the merged BAM file. VQSR-filtered calls were then compared to the high-

confidence call sets made available by Illumina for those samples. 524,331 of the 4,348,092

non-monomorphic variant sites were not present in the high-confidence calls, whereas

1,246,403 of the 5,070,164 non-monomorphic high-confidence were not recapitulated in

the 1x data. This corresponds to an estimated false positive rate of 12% and false negative

rate of 24.6%. Both unique sets had a much higher proportion of singletons (corresponding

to MAF < 2.9%) than the entire sets (57.9% vs 19.9% of singletons among 1x calls and 51% vs

18.1% among high-confidence calls), which suggests that a large fraction of the erroneous

sites lies in the low-frequency and rare part of the allelic spectrum. However, genotype

accuracy is poor, to the point where it obscures peculiarities in the distribution of allele

counts (Additional File 2: Figure S5). Due to them being present in the 1000 genomes

reference panel, we remove the 17 Platinum Genomes prior to imputation.

Genotype refinement and imputation

Reference Panel

Phased haplotypes from 1092 samples from the 1000 Genomes Project Phase 1 study were

merged with 3781 7x WGS samples from the UK10K [4] TwinsUK [27] and ALSPAC [28]

studies, and with 249 MANOLIS samples sequenced at 4x depth [9] using SHAPEIT v2 [29]

and converted to VCF format. Alleles in the reference panel were flipped so as to

correspond to the reference allele in the called dataset. Positions where the alleles differed

between the called and reference datasets were removed from both sources. Indels were

filtered out due to poor calling quality.

Pipeline

As described previously [22], we used Beagle v.4 [30] to perform a first round of imputation-

based genotype refinement on 1,239 HELIC MANOLIS variant callsets, using a previously

described [9] reference panel composed of 10,244 haplotypes from the 1000 Genomes,

UK10K and MANOLIS 4x reference sequences. This was followed by a second round of

reference-free imputation, using the same software. The same pipeline was applied to 1166

individuals from the Pomak cohort.

Evaluation of pipelines

The authors of SHAPEIT [29] advise to phase whole chromosome when performing pre-

phasing in order to preserve downstream imputation quality.   This approach is

computationally intractable for the 1x datasets, where the smallest chromosomes contain

almost 7 times more variants than the largest chromosomes in a GWAS dataset.

For benchmarking purposes, we tested 13 genotype refinement pipelines involving Beagle

v4.0 [30] and SHAPEIT2 [29] using a 1000 Genomes phase 1 reference panel, which we

evaluated against minor allele concordance. All pipelines were run using the vr-runner

scripts (https://github.com/VertebrateResequencing/vr-runner). Pipelines involving Beagle

with the use of a reference panel ranked consistently better (Additional File 2: Figure S6),

with a single run of reference-based refinement using Beagle outperforming all other runs.

IMPUTE2 performed worst on its own, whether with or without reference panel; in fact the

addition of a reference panel did not improve genotype quality massively. Phasing with

Beagle without an imputation panel improved genotype quality, before or after IMPUTE2.

Halving the number of SNVs per refinement chunk to 2,000 (including 500 flanking

positions) resulted in only a modest loss of genotype quality in the rare part of the allelic

spectrum (Additional File 2: Figure S7), while allowing for a twofold increase in refinement

speed. Genotype quality dropped noticeably for rare variants when imputation was turned

on (Additional File 2: Figure S7), but remained high for low-frequency and common ones. A

reference-free run of Beagle allowed to phase all positions and remove genotype

missingness with no major impact on quality and a low computational cost. We also tested

thunderVCF [31] for phasing sites, however, the program took more than 2 days to run on

5,000 SNV chunks and was abandoned.

Variant-level QC

Beagle provides two position level imputation metrics, allelic R-squared and dosage R-

squared. Both measures are highly correlated (Additional File 2: Figure S8.a). Values

between 0.3 and 0.8 are typically used for filtering [32]. In both 1x datasets 59% and 91% of

imputed variants lie below those two thresholds, respectively. The distribution of scores

does not provide an obvious filtering threshold (Additional File 2: Figure S8.b) due to its

concavity. Since most imputed variants are rare and R-squared measures are highly

correlated with MAF, filtering by AR2 and DR2 would be similar to imposing a MAF

threshold (Additional File 2: Figure S8.c and d.). Moreover, due to a technical limitation of

the vr-runner pipelines, imputation quality measures were not available for refined

positions at the time, only imputed ones. Therefore, we did not apply any prior filter in

downstream analyses, but used imputation metrics as well as variant quality scores to

prioritise variants post-association.

Sample QC

Due to the sparseness of the 1x datasets, sample-level QC was performed after imputation.

58 individuals were removed from the Pomak cohort due to contamination and sample

swap issues. 5 samples were excluded from the MANOLIS 1x cohort and 1 sample from the

4x cohort following PCA-based ethnicity checks.

Comparison with WES

A set of high confidence genotypes was generated for the 5 exomes in MANOLIS using filters

for variant quality (>200), call rate (AN=10, 100%) and depth (250). These filters were

derived from the respective distributions of quality metrics (Additional File 2: Figure S9).

When compared to 5 whole-exome sequences from each cohort, imputed 1x calls

recapitulated 77.2% of non-monomorphic, high-quality exome sequencing calls.

Concordance was high, with only 3.5% of the overlapping positions exhibiting some form of

allelic mismatch. When restricting the analysis to singletons, 9105 (58%) of the 15,626 high-

quality singletons in the 10 exomes were captured, with 21% of the captured positions

exhibiting false positive genotypes (AC>1). To assess false positive call rate, we extracted 1x

variants falling within the 71,627 regions targeted by the Agilent design file for WES in

overlapping samples, and compared them to those present in the unfiltered WES dataset.

103,717 variants were called in these regions from WES sequences, compared to 58,666

non-monomorphic positions in the 1x calls. 1,419 (2.4%) of these positions were unique to

the 1x dataset, indicating a low false-positive rate in exonic regions post-imputation.

Genetic relatedness matrix

In order to correct for genetics relatedness within the two isolated cohorts, we calculated a

genetic relatedness matrix using GEMMA [33]. Given the isolated nature of the population

and the specificities of the sequencing dataset, we used different variant sets to calculate

kinship coefficients. Using the unfiltered 1x variant dataset produced the lowest coefficients

(Figure 10.a), whereas well-behaved set of common SNVs [34] produced the highest, with

an average difference of 3.67×10^{-3}. Filtering for MAF lowered the inferred kinship

coefficients. Generally, the more a variant set was sparse and enriched in common variants,

the higher the coefficients were. However, these differences only had a marginal impact on

association statistics, as evidenced by a lambda median statistic difference of 0.02 between

the two most extreme estimates of relatedness when used for a genome-wide association

of triglycerides in Pomak (Additional File 2: Figure S9.b). For our association study, we used

LD-pruned 1x variants filtered for MAF<1% and Hardy Weinberg equilibrium $p<1x10^{-5}$ to

calculate the relatedness matrix.

Phenotype preparation

Twenty-five haematological phenotypes were prepared, with certain traits measured only in

one cohort or the other (Additional File 2: Table S4), and high levels of correlation for some

traits(Additional File 2: Figure S11). Full details of the trait transformation, filters and

exclusions are described in Additional File 2: Table S4. The 'transformPhenotype'

(https://github.com/gmelloni/transformPhenotype) R script was used to apply a

standardised preparation for all phenotypes. If gender differences were significant

(Wilcoxon rank sum P < 0.05), the phenotype was stratified accordingly. Following trait-

specific exclusions and adjustments, outliers were filtered out based on 3 standard

deviations (SD) away from the mean where necessary. Traits not normally distributed were

transformed to normality using an inverse normal transformation, after testing for a

number of power transformation including logarithmic. For all traits age and age^2 were

added as covariates as necessary and standardised residuals were used. If male and female

phenotypes were prepared separately these were standardised before combining the

residuals. Summary statistics for all of the traits are provided in Additional File 2: Table S5.

Single-point association

Pipeline

Association analysis was performed on each cohort separately using the linear mixed model

implemented in GEMMA [33] on all variants with minor allele count (MAC) greater than 2

(14,948,665 out of 30,483,158 variants in MANOLIS and 15,564,905 out of 29,740,281 for

Pomak). We used the aforementioned centered kinship matrix. GC-corrected p-values from

the likelihood ratio test (p_lrt) are reported. Singletons and doubletons are removed due to

overall low minor allele concordance.

Estimating the significance threshold

We determine the significance threshold by calculating $\alpha_{adj} = \frac{0.05}{N_{eff} \times k_{eff}}$, where N_{eff} is the

effective number of SNVs after correcting for LD and k_{eff} is the effective number of traits

tested after correcting for correlation. We estimated k_{eff} using two different methods. The

first method selects the number of principal components (PCs) in a principal component

analysis (PCA) of standardised, normalised traits that explain 95% of total trait variance. This

yielded $k_{eff} = 8$ for both MANOLIS and Pomak ($M = 20$ and $M = 18$, respectively). The

second method uses the Kaiser method on the eigenvalues of the trait correlation matrix to

calculate k_{eff} [35], and gives $k_{eff} = 9$ for MANOLIS and $k_{eff} = 8$ for Pomak.

For N_{eff}, we extrapolate the number of SNVs based on calibration curves[36] that provide

the number of independent SNVs given the total number of tested SNVs (assuming

MAF>0.5%). This gives $N_{eff} = 5{,}145{,}236$ for MANOLIS ($\alpha_{adj} = 1.08 \times 10^{-9}$) and $N_{eff} =$

5,361,759 for Pomak ($\alpha_{adj} = 1.16 \times 10^{-9}$). Performing LD-pruning using PLINK [37] yields

8,123,367 variants with MAC>2 for Pomak ($\alpha_{adj} = 7.7 \times 10^{-10}$) and 6,833,823 variants for

MANOLIS ($\alpha_{adj} = 8.12 \times 10^{-10}$). We define genome-wide significance at $\alpha_{adj} = 1.0 \times 10^{-9}$,

which is reasonably close to these estimates.

Signal prioritisation

Signals were extracted using the peakplotter software (https://github.com/wtsi-
team144/peakplotter) using a window size of 1Mb.

Replication

The discovery and validation studies were conducted in different populations. This can

affect the strength of associations of genetic variants and lead to heterogeneity in effect

sizes [38, 39]. Therefore, we assessed heterogeneity and carried out a random effects meta-

analysis when there was evidence of heterogeneity at $p<0.05$. We estimated heterogeneity

using I2 and Q statistics. We used the method described in [13] for the random effects

meta-analysis, as it was shown to have higher power to detect associations than the

conventional random effects method.

Declarations

Ethics approval and consent to participate

The study was approved by the Harokopio University Bioethics Committee and informed

consent was obtained from every participant.

Consent for publication

Not applicable.

Availability of data and materials

The following HELIC genotype and WGS datasets have been deposited to the European

Genome-phenome Archive (https://www.ebi.ac.uk/ega/home): EGAD00010000518;

EGAD00010000522; EGAD00010000610; EGAD00001001636, EGAD00001001637. The

peakplotter software is available at https://github.com/wtsi-team144/peakplotter, the

transformPhenotype app can be downloaded at https://github.com/wtsi-

team144/transformPhenotype.

Competing interests

The authors declare that they have no competing interests.

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award WT091310.

Author's contributions

AG performed variant call set quality control, evaluated and ran imputation pipelines,

performed single-point association with the help of KH, DS and XG and was a major

contributor in writing the manuscript. KK performed follow-up on the single-point signals

and contributed to writing the manuscript. LS analysed the genotype data and ran signal

discovery scripts. RM performed sample quality control on the Pomak dataset. GEMM

contributed to the transformPhenotypes tool. AEF under the supervision of GD and with the

help of ET and MK, coordinated the sample collection. GR performed bioinformatics

analyses on selected variants. JS performed quality control on the 4x sequences. PD

designed the runner pipelines used for imputation. BK maintained the phenotype database.

MP performed variant calling. HE, WJA, TJ, AB and NZ provided replication results in the UK

Biobank meta-analysis. EZ supervised the project. All authors read and approved the final

manuscript.

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and UK10K_COHORT_ALSPAC (the Avon Longitudinal Study of Parents and Children). A full

list of the investigators who contributed to the generation of the data is available from

www.UK10K.org, last accessed April 29, 2016. The GATK3 program was made available

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Inc. This research has been conducted using the UK Biobank Resource using Application

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    Figures

Figure 1: Concordance and call rate for 1x WGS genotypes. Genotype (blue circles) and

minor allele (yellow circles) concordance is computed for 1239 samples in MANOLIS against

merged OmniExpress and ExomeChip data. Call rate is assessed for the refined (purple) and

refined plus imputed (green) datasets.

Figure 2: Unique variants called by sequencing and imputed GWAS. Variants unique to

either dataset, arranged by MAF bin. Both datasets are unfiltered apart from

monomorphics, which are excluded. MAF categories: rare (MAF<1%), low-frequency (MAF

1-5%), common (MAF>5%).

Figure 3: Association between rs6489858 and lymphocyte count in MANOLIS.

Figure 4: Imputed GWAS results for association signals found in the 1x WGS at p<5x10^{-7}.

Purple dots represent significant results in the 1x analysis. Orange dots, if present, denote

the p-value of the same SNP in the imputed GWAS study. Absence of a dot indicates the

variant was not found in the imputed GWAS dataset. Red dashes indicate the minimum p-

value among all tagging SNPs in the imputed GWAS (r^2>0.8).

Additional Files

File name	File Format	Title	Description
Additional File 1.xlsx	Excel (xlsx)	Supplementary Tables	Tables S1 to S5
Additional File 2.pdf	PDF	Supplementary Figures and Text	Figures S1 to S11, Supplementary Text

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