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1	Title: Empiri	cal dynamic modeling reveals ecological drivers of dengue dynamics
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20 mosquito-borne disease, rainfall, susceptible population size, temperature.

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# 21 Abstract

22	Understanding ecological drivers of mosquito-borne disease is an ongoing public
23	health priority. Previous experiments and mechanistic models suggest mosquito-
24	borne disease transmission involves complex nonlinear interactions between
25	climate and population dynamics. This makes detecting environmental disease
26	drivers at the population level challenging. By analyzing incidence data, estimated
27	susceptible population size, and climate data with methods based on nonlinear time
28	series analysis, collectively referred to as empirical dynamic modeling (EDM), we
29	identified drivers and their interactive effects on dengue dynamics in San Juan,
30	Puerto Rico. Estimated susceptible population size was the strongest causal driver
31	of dengue incidence, and climatic forcing became important above a certain
32	susceptible population size (temperature and rainfall having net positive and
33	negative effects, respectively). Our EDM protocol for measuring and predicting how
34	climate and population dynamics interact to drive epidemics adds to a growing body
35	of empirical studies of complex, nonlinear systems embedded in changing
36	environments.

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# 37 INTRODUCTION

38	Mosquito-borne diseases, and dengue in particular, are (re)emerging globally and
39	spreading to higher latitudes in concert with globalization and climate change
40	(Kilpatrick & Randolph 2012; Ryan <i>et al.</i> 2019). Dengue virus—transmitted
41	primarily by urban <i>Aedes aegypti</i> (Kraemer <i>et al.</i> 2015b)—places half of the global
42	human population in 128 countries at infection risk (Brady et al. 2012; Kraemer et
43	al. 2019) and causes an estimated 390 million annual cases of dengue fever
44	worldwide (Bhatt et al. 2013). In the absence of effective vaccines or treatments
45	(Katzelnick et al. 2017b, a; Sridhar et al. 2018), public health agencies rely on vector
46	control to reduce dengue transmission (Erlanger et al. 2008). Effective vector
47	control measures require understanding the mechanisms linking vector ecology and
48	epidemics to better predict disease outbreaks—a major research challenge.
49	Previous prediction models used phenomenological and mechanistic equation-
50	based approaches (Johansson <i>et al.</i> 2009; Hii <i>et al.</i> 2012; Tran <i>et al.</i> 2013; Liu-
51	Helmersson <i>et al.</i> 2014; Morin <i>et al.</i> 2015; Mordecai <i>et al.</i> 2017; Johnson <i>et al.</i> 2018),
52	which may not fully capture interdependence between predictors. Here, we used a
53	mechanistic, equation-free, data-driven approach that accounts for interdependence
54	to determine ecological drivers and predict dengue outbreaks in an urban
55	environment.

Since mosquitoes have been shown to be sensitive to climate variables, such as
temperature and rainfall (Ibarra *et al.* 2013; Mordecai *et al.* 2019), it is reasonable to
ask whether temperature and rainfall are important drivers of dengue outbreaks.

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59	Although temperature directly affects mosquito and viral traits in laboratory
60	experiments (Watts et al. 1987; Lambrechts et al. 2011), the relationship between
61	temperature and dengue incidence in the field has been ambiguous. Thus,
62	temperature-dependent models have had mixed success predicting the timing and
63	magnitudes of epidemics (Hii et al. 2012; Johansson et al. 2016; Johnson et al. 2018).
64	Similarly, the rainfall-dengue relationship is complex, because the effect of rainfall
65	on mosquitoes depends on local breeding habitat and human behavior. In some
66	settings, rainfall provides more container breeding habitat for mosquitoes, thus
67	increasing mosquito abundance and dengue incidence (Ibarra et al. 2013). By
68	contrast, low rainfall levels could also facilitate dengue transmission by promoting
69	water storage that serves as standing-water habitat for mosquito breeding
70	(Oliveira-lima et al. 2000). Further, heavy rainfall can reduce mosquito abundance
71	by flushing out larvae (Koenraadt & Harrington 2008). The net effect of climate on
72	dengue is a convolution of many different mechanisms of action, and the net
73	outcome depends on specific context details. Moreover, as these details change, the
74	climate-dengue relationship can change.
75	Previous studies using statistical (Johansson <i>et al.</i> 2009; Hij <i>et al.</i> 2012; Liu-
75	r revious studies using statistical (jonalisson et ul. 2005, nil et ul. 2012, hiu
76	Helmersson <i>et al.</i> 2014; Johnson <i>et al.</i> 2018) and mechanistic models (Tran <i>et al.</i>
77	2013; Morin <i>et al.</i> 2015; Mordecai <i>et al.</i> 2017) suggest that temperature and rainfall
78	drive dengue transmission in the field. However, these models rely on parameter
79	estimates from laboratory studies that are engineered to isolate single mechanisms
80	of action, hence producing separable relationships between drivers and dengue,

81 eliminating the complex interdependence at the population level. While laboratory

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82	studies provide robust validation of a mechanism (Lambrechts et al. 2011), the fixed
83	relationships (correlations) taken from them do not necessarily translate into
84	proper causal inference for nonlinear field systems (Sugihara et al. 2012). Even if
85	causality exists between two variables in such a system, their correlation can switch
86	signs during different time periods, resulting in a net correlation of zero (Deyle <i>et al.</i>
87	2016b). This temporal variation in the direction of correlation results from the
88	nonlinear, state-dependent relationship between the variables (i.e., the importance
89	or direction of one effect depends on the state of another variable). Conversely, even
90	if two variables are consistently correlated, the association could be spurious due to
91	a confounder—a third variable that drives two otherwise unrelated variables. Thus,
92	covariation among variables poses a problem for identifying causal drivers. Both
93	temperature and rainfall follow seasonal patterns in most regions of the world and
94	often covary, making it difficult to separate their effects. Since ecological systems
95	are often nonlinear with covarying drivers, it is difficult to isolate causality from
96	field data.

97 Disease incidence also depends nonlinearly on (potentially climate-driven) 98 transmission rates, because epidemic growth slows as the population of susceptible 99 individuals is exhausted (Rypdal & Sugihara 2019). Thus, susceptible population 100 size is an important driver of infectious disease dynamics (Anderson & May 1979; 101 Kraemer et al. 2015a; Rypdal & Sugihara 2019). Variation in susceptible host 102 availability may influence the effects of climate on dengue dynamics, but such 103 interactive effects are difficult to detect in observational data. A lack of rigorous 104 methods has hindered research on potential interdependence of climate and

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susceptible population size as drivers of dengue incidence. Because correlations are
unreliable and variables are interdependent in nature, disentangling climate from
other drivers in disease systems remains challenging (Thai & Anders 2011; Morin *et al.* 2013).

109 To overcome these challenges, we used empirical dynamic modeling (EDM)

110 (Sugihara *et al.* 2012) to identify and model causal mechanisms driving dengue

111 epidemics (see <u>http://tinyurl.com/EDM-intro</u> for an introduction). EDM differs from

112 statistical and equation-based mechanistic models in two key ways. First, unlike

113 statistical approaches where inferences about cause-and-effect relationships are

114 based on fixed independent pairwise associations between system variables, EDM is

115 based on reconstructing the system *dynamics*. This allows relationships among

116 system variables to change through time to reflect that interactions among variables

117 are changing. Second, unlike equation-based mechanistic models, EDM does not

118 require assumptions about the functional form of the model, but instead

119 reconstructs the dynamic attractor empirically from time series observations. An

120 attractor is a geometric object (i.e., curve or manifold) that embodies the rules for

121 how relationships among variables change with respect to each other through time

122 depending on system state (specific location on the attractor). Like a set of

123 equations, the geometric attractor encompasses the dynamics of a system, and thus

124 can provide an agnostic (without an assumed set of equations) yet mechanistic

125 understanding of the system that is empirically based. Although traditional

126 equation-based mechanistic models can be constructed to account for nonlinearity,

127 these approaches require *a priori* assumptions about the identity and the form of

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the causal relationships between variables that may not hold; however, EDM doesnot rely on such assumptions.

130	Here, we use EDM and a recent approach for inferring the susceptible population
131	size from incidence data (Rypdal & Sugihara 2019) to answer three questions: (1)
132	Do temperature, rainfall and/or inferred susceptible population size drive dengue
133	incidence? (2) Can we predict dengue dynamics using climate data and inferred
134	susceptible population size? (3) What is the approximate pattern of each causal
135	relationship, and how is this relationship influenced by the other drivers? We
136	examine the hypothesis that temperature and rainfall drive dengue incidence in
137	complex ways that depend on susceptible population size. The EDM protocol
138	presented here—identifying drivers, their predictive power, and state-dependent
139	functional responses—can be applied to other complex, ecological systems.

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## 140 **METHODS**

## 141 **Data and proxy for susceptible population size**

- 142 We obtained time series of weekly observations of dengue incidence (total number
- 143 of new dengue cases of all serotypes), temperature (°C), and rainfall (mm) in San
- 144 Juan, Puerto Rico, 1990–2009 (Figure 1a–c) from the National Oceanic and
- 145 Atmospheric Administration on November 10, 2016
- 146 (<u>http://dengueforecasting.noaa.gov/</u>). Direct measurement of susceptible
- 147 populations is not feasible, so we used a recently developed method for inferring a
- 148 proxy for the susceptible population size (Rypdal & Sugihara 2019). This method
- 149 uses incidence data during the inter-outbreak period to construct a time series for
- the susceptible population. Although few dengue cases occur during the inter-
- 151 outbreak period, incidence during this time contains information about the
- 152 susceptible population size in the next outbreak. Because during an inter-outbreak
- 153 period the disease system fluctuates around a disease-free equilibrium, a linear
- approximation of the incidence rate can be made where the coefficient, i.e., the
- 155 leading eigenvalue,  $\lambda$ , scales linearly with the susceptible population size. If we
- assume a simple Susceptible-Infected-Recovered (SIR) model (Kermack &
- 157 McKendrick 1927) for the disease system, then

$$\lambda = \beta \tilde{S} - \gamma \tag{1}$$

158 where  $\tilde{S}$  is a parameter representing the susceptible population size,  $\beta$  is the force of 159 infection, and  $\gamma$  is the recovery rate. Thus,  $\lambda$  is linearly related to the proportion of

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160 the population that is susceptible over time, and can be used as a proxy (see

- 161 Supporting Information for details).
- 162 Since the weekly incidence data are discrete we need to obtain the discrete-time
- 163 eigenvalue  $\tilde{\lambda} = e^{\lambda \Delta t}$ . To infer  $\tilde{\lambda}$  from the weekly incidence data I(t), we performed
- 164 linear regression by fitting the statistical model  $I(t + \Delta t) = \tilde{\lambda} I(t)$  for 12 time
- 165 points in a 12-week running window ( $\Delta t = 1$  week). The model is robust to the

166 window size (Rypdal & Sugihara 2019). In the discrete case, when  $\tilde{\lambda} < 1$  the system

- 167 is stable (inter-outbreak period) and when  $\tilde{\lambda} \ge 1$  then the system is unstable
- 168 (outbreak period). Here, we treated the resulting time series of  $\tilde{\lambda}$  as a proxy for the
- susceptible population size, or "susceptibles index" for short (Figure 1d).

### 170 Statistical analyses

171 All analyses were conducted in R version 3.5.1 (R Development Core Team 2018).

172 We performed pairwise cross-correlations on the time series to investigate time-

173 lagged relationships between potential drivers (i.e., temperature, rainfall, and

174 susceptibles index) and dengue incidence using the tseries package version 0.10-

175 45 (Trapletti & Hornik 2018). We calculated the interannual mean to obtain the

176 seasonal variability for each variable. Determining whether the variables follow

177 seasonal trends is important for EDM analyses. Before performing EDM, we

- 178 normalized each time series to zero mean and unit variance to remove
- 179 measurement unit bias, ensuring the variables would be comparable and the

180 attractor would not be distorted.

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# 181 Empirical dynamic modeling (EDM)

182	All EDM analyses were performed using package rEDM version 0.7.1 (Ye <i>et al.</i> 2018).
183	EDM includes approaches to infer a system's mechanistic underpinnings and predict
184	its dynamics. EDM uses time series data of one or more variables to construct an
185	attractor in state space (Figure S1). This is called univariable or multivariable state
186	space reconstruction (SSR) for an attractor built using a single or multiple variables,
187	respectively. Properties of the attractor are assessed to examine characteristics of
188	the system (Deyle & Sugihara 2011).
189	EDM is sensitive to stochasticity, and should be applied in systems where there is
190	evidence of deterministic dynamics (Cummins et al. 2015). Some stochasticity is
191	allowed (e.g., there can be stochastic drivers in the system), but the system cannot
192	be entirely stochastic (i.e., there must be low-dimensional deterministic structure to
193	most variables). To test for low-dimensional deterministic dynamics we performed
194	univariable SSR for each variable and used simplex projection (Sugihara & May
195	1990)—a type of nearest neighbor regression—to check that the prediction skill
196	decreases with time to prediction, an indicator of deterministic dynamics (Figures
197	S2a and S4; see Supporting Information for details). To test for nonlinear state
198	dependence of a variable—the motivation behind EDM—we used the S-map test for
199	nonlinearity (Sugihara 1994) (Figures S2b,c and S5; see Supporting Information for
200	details).

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# 201 EDM: Convergent cross-mapping

202	We used an EDM approach called convergent cross-mapping (CCM) (Sugihara et al.
203	2012) to identify drivers of dengue incidence. If two variables are causally related,
204	then a multivariable attractor—where each variable in the system represents a
205	dimension that traces the dynamics of the system—can be semi-reconstructed using
206	lagged versions of just one of the variables (Figure S1). Based on Takens' Theorem,
207	this univariable "shadow attractor" preserves the structural and dynamic properties
208	of the original multivariable attractor (Takens 1981) [see video
209	https://youtu.be/QQwtrWBwxQg (Sugihara et al. 2012)]. The concept behind CCM
210	is that if temperature causes dengue incidence, then information about temperature
211	will be embedded in the dynamics of dengue, such that the shadow attractor
212	produced using only dengue dynamics allows us to accurately reconstruct
213	temperature in the past. However, the converse scenario would not be true: since
214	dengue does not cause temperature, the shadow attractor constructed using
215	temperature data should not contain information to accurately reconstruct dengue
216	incidence (see Supporting Information for details).
217	The critical criterion for testing the existence of causality using CCM is checking that
218	the cross-mapping skill monotonically increases and plateaus (i.e., converges) with
219	the length of the response variable data series used in cross-mapping. The cross-
220	mapping skill, $ ho$ , is the Pearson's correlation coefficient between predicted driver
221	values using the univariable SSR of the response variable, and the observed driver
222	values. We used the Kendall's $ au$ test as a significance test for convergence using the

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223	Kendall package version 2.2	(McLeod 2011). This test checks whether cross-
	1 ()	

224 mapping skill has a significant monotonic increasing trend. If  $\tau > 0$  then there is 225 convergence (Grziwotz *et al.* 2018).

Based on the cross-correlation analyses (Figure S6), we applied a 9-week time lag

between temperature and incidence, an averaged lag of 3–9 weeks for rainfall, and a

228 5-week lag for the susceptibles index. The rainfall variable is a proxy for standing

water as mosquito breeding habitat, and thus we were interested in the

accumulation of water over time. We refer to these as "ecological lags" and treat

them as proxies for the time delay of cause-and-effect, since they account for the

ecological processes in the causal chain of events (ecological lags are separate from

233 lags used in SSR). These ecological time lags are consistent with results from other

field studies, which showed that temperature and rainfall predict dengue cases 6–12

weeks ahead (Chen *et al.* 2010; Ibarra *et al.* 2013).

### 236 EDM: Null models

237 For CCM, we assessed the strength of the evidence for causal effects of potential

drivers on dengue using two null models that control for the seasonal trend

observed in all variables (Figure 2). These null models address the sensitivity of

240 CCM to periodic fluctuations (i.e., seasonality), which can make two variables appear

to be causally linked when instead they are simply synchronized by an unobserved

seasonal variable (Cobey & Baskerville 2016; Deyle *et al.* 2016a). In the first null

- 243 model ("seasonal null model"), we preserved the seasonal signal in the null, but
- randomized the interannual anomalies, and compared model performance with the

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245	true time series (Deyle <i>et al.</i> 2016a). In the second, more conservative null model
246	("Ebisuzaki null model"), we conserved any serial correlation (beyond seasonal) and
247	randomized the phases of Fourier transformed time series, and compared model
248	performance with the true time series (Ebisuzaki 1997).
249	We also repeated the CCM method in the nonsensical, reverse-causal direction (e.g.,
250	to test whether incidence drives climate) as a control for potential spurious
251	relationships generated by non-causal covariation (e.g., due to seasonality). This
252	addresses the issue of synchrony, in which CCM can indicate bidirectional causality
253	when one direction is false or nonsensical (Baskerville & Cobey 2017; Sugihara et al.
254	2017).

### 255 **EDM: Forecast improvement**

256 We examined the predictive power of each driver—or combination of drivers—on

257 dengue incidence by assessing how well we can predict dengue dynamics using

temperature, rainfall, and susceptibles index. We used a combination of univariable

259 SSR (i.e., with incidence data) and multivariable SSR to determine the improvement

260 of forecasting (using simplex projection) when including each driver or combination

261 of drivers (Dixon *et al.* 1999; Deyle *et al.* 2013, 2016a).

We investigated the potential forecast improvement of dengue incidence using
temperature, rainfall, susceptibles index, and their combined effect. In addition, we
investigated the predictive power of dengue incidence using multivariable SSR with

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just the potential driver variables: temperature, rainfall, and susceptibles index (seeSupporting Information).

## 267 EDM: Scenario exploration

268 In nonlinear systems, drivers generally do not have a constant effect. The effect is 269 state-dependent—the strength and possibly the direction of the effect depends on 270 the current state of the system. Scenario exploration with multivariable EDM 271 allowed us to assess the effect of a small change in temperature or rainfall on 272 dengue incidence, across different states of the system. The outcome of these small 273 changes allowed us to deduce the relationship between each climate driver and 274 dengue incidence and how they depend on the system state. For each time step t we 275 used S-maps (Sugihara 1994; Deyle *et al.* 2016a) to predict dengue incidence using a 276 small increase  $(+\Delta X/2)$  and a small decrease  $(-\Delta X/2)$  of the observed value of driver 277 X(t) (temperature or rainfall). For each putative climate driver, the difference in dengue predictions between these small changes is  $\Delta Y = Y(t+1) \left[ X(t) + \frac{\Delta X(t)}{2} \right] -$ 278  $Y(t+1)\left[X(t)-\frac{\Delta X(t)}{2}\right]$ , where Y(t+1) is a function of X and all other state variables, 279 280 and we used  $\Delta Y / \Delta X$  to approximate the effect of driver X at time t. We repeated this 281 over all time steps in our time series data for both temperature and rainfall to 282 recover their approximate relationships with dengue incidence at different states of 283 the system. Scenario exploration analyses were repeated across several model 284 parameterizations to address potential sensitivity to parameter settings (see 285 Supporting Information).

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# 286 **RESULTS**

# 287 **Detecting drivers**

288	EDM showed that temperature, rainfall, and susceptibles index drive dengue
289	incidence since the convergence criterion was met (Kendall's $ au > 0$ , p < 0.01) in all
290	three CCM cases (Figure 3a–c). Rainfall and susceptibles index were significant
291	drivers of dengue incidence beyond seasonality, as their effects were
292	distinguishable from seasonal and Ebisuzaki null models (Figures 3b, c and S8b, c; p
293	< 0.05). This implies significant causal effects of both rainfall and susceptible
294	population size on dengue, which are not obscured by a confounder with a periodic
295	signal. However, temperature was not a significant driver beyond seasonality
296	(Figures 3a and S8a; $p > 0.05$ ). This implies we cannot rule out the possibility that
297	the apparent forcing of temperature on dengue is due to a confounder with a
298	seasonal signal. However, if there is no such confounder, then the seasonal trend in
299	temperature, which accounts for most temperature variation in San Juan, drives the
300	seasonal trend observed in dengue incidence (i.e., seasonality of temperature drives
301	seasonality of dengue). Compared to the other drivers, the predictive skill of the
302	temperature null model was relatively high (Figure 3a), suggesting that temperature
303	seasonality in the null model was predictive. This further supports the notion that
304	seasonal temperature may be driving dengue dynamics.

305 As expected, EDM tests for causality in the nonsensical directions—incidence
306 driving temperature or rainfall—were not significant (i.e., no convergence; Figure

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307	S7, black lines). This further supports that temperature and rainfall drive dengue
308	incidence, because their causal relationships were not confounded by spurious
309	bidirectionality. Further, the null models for the nonsensical directions of causality
310	(Figure S7, grey lines) also displayed no convergence (completely flat), as expected
311	(i.e., seasonality of dengue incidence does not drive seasonality of temperature or
312	rainfall). However, seasonality (or any periodicity) of temperature, rainfall and
313	susceptibles index drive dengue dynamics, shown by convergence of the seasonal
314	and Ebisuzaki null models (grey lines in Figures 3 and S8).

#### 315 **Predictive power of drivers**

316 Dengue incidence was highly predictable using univariable SSR of incidence data

alone (Adjusted R<sup>2</sup> = 0.8922,  $\rho$  = 0.9446; Figure 4a). The predictive power of dengue

318 incidence improved only slightly when temperature, rainfall, and susceptibles index

319 were added in a multivariable SSR model (Adjusted R<sup>2</sup> = 0.8927,  $\rho$  = 0.9448; Figure

4d). These results suggest that dengue incidence time series contain information

321 about these drivers, resulting in the high predictability of the attractor in

322 univariable SSR (Figure 4a).

323 Dengue dynamics were also highly predictable using only the driver time series (i.e.,

324 temperature, rainfall, and susceptibles index) in a multivariable SSR model

325 (Adjusted  $R^2 = 0.5044$ ,  $\rho = 0.7102$ ; Figure 4c), where timing and magnitude of

326 epidemics were captured reasonably well. However, the model using only

327 temperature and rainfall data did not predict dengue incidence as well (Adjusted R<sup>2</sup>

328 = 0.0533,  $\rho$  = 0.2309; Figure 4b). Thus, adding the susceptibles index increases

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329	predictive power compared to climate variables alone, particularly for predicting
330	the magnitude of outbreaks (Figure 4c). However, temperature and rainfall did
331	capture the timing (seasonality) of the epidemics (Figure 4b). This supports the
332	notion that seasonality of temperature and rainfall is important for explaining the
333	seasonality of dengue, and that susceptible population size is important for
334	determining epidemic size. All SSR models (Figure 4a–d) had significant ( $p < 0.001$ )
335	F statistics >> 1 from ANOVA, rejecting the null hypothesis (i.e., no relationship

between predicted and observed dengue incidence).

### 337 State-dependent functional responses

338 By investigating the rate of change of dengue incidence as a function of climate

across system states, we found that temperature had a small *positive* median effect

340 (0.0035, Wilcox p < 0.001) on dengue incidence. A positive effect is expected

341 (Mordecai *et al.* 2017) for the temperature range in Puerto Rico (Figure S9, black

dashed lines), although the effect was occasionally much stronger, both positive and

343 negative (Figure 5a, b). The large negative effects occurred only at the highest

344 temperature values (as predicted by mechanistic models of temperature-dependent

transmission), reinforced by a lower quantile regression with a strongly negative

- 346 slope (Figure 5b, bottom dashed red line). However, positive effects occurred across
- 347 the whole temperature range.

Rainfall had a small *negative* median effect (-0.0022, Wilcox p < 0.001), but

349 occasionally had very large negative effects (Figure 5a, c). These large, negative

350 effects of rainfall on dengue occurred when there was less than 100 mm of rain per

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351	week (Figure 5c), in line with expectations that <i>low</i> amounts of rainfall could lead to
352	high number of dengue cases, since people tend to store water in containers during
353	periods of drought, enabling mosquitoes to breed (Oliveira-lima et al. 2000). There
354	are also small positive effects of rainfall on dengue (Figure 5c), suggesting that
355	overall the results showed competing effects of low-moderate rain providing
356	standing water for mosquito breeding and humans storing standing water where
357	mosquitoes can breed when there is drought or low rain.

358 The results on climate effects suggest that the strength and direction of the effects of

climate on dengue dynamics depend on the state of the system. A potential cause of

360 state-dependent climate effects on dengue dynamics is the variation in the

361 susceptible population size over time (Figure 6a, b). Outbreaks do not occur when

there are too few susceptible people to get infected. As expected, when the

363 susceptible population size was small ( $\tilde{\lambda} < 0.85$ ) incidence was insensitive to

364 climate (Figure 6c, e). By contrast, when the susceptible population size was large

365 ( $\tilde{\lambda} > 0.85$ ), temperature and rainfall effects on dengue incidence appeared (Figure

366 6d, f). The gradual increase and decrease of the rate of change of dengue as a

367 function of temperature (Figure 6d, red solid lines) aligned well with the sigmoidal

368 part (Figure S9, black dashed lines representing the temperature range in our

369 study) of the unimodal temperature response curve (assuming transmission

370 primarily by *Ae. aegypti*) developed previously (Mordecai *et al.* 2017). This is an

371 important finding, since evidence of climate functional responses for disease

372 dynamics is rare due to the difficulty of obtaining appropriately informative field

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- data. It is possible that if we had temperature data ranging across a larger
- 374 spectrum—possibly by assembling data across multiple climates—that the
- 375 empirical functional response derived from EDM would also look unimodal.

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# 376 **DISCUSSION**

377	EDM can identify causal drivers, quantify their predictive power, and approximate
378	functional responses in nonlinear, state-dependent ecological systems, which few
379	other methods can do. Infectious disease dynamics are often seasonal, adding
380	another layer of complexity. Here, we used seasonal null models and a well-studied
381	ecological system to disentangle nonlinear and interactive mechanisms driving
382	disease dynamics. First, EDM detected rainfall, susceptible population size, and
383	plausibly temperature (mostly via its seasonality) as drivers of dengue incidence.
384	The seasonal variation in incidence was more attributed to climate, while the
385	interannual variation in incidence was more explained by the susceptible
386	population (Figure 3). Second, EDM provided a predictive model based on these
387	three drivers that had a reasonably good fit to dengue incidence data (R² = 0.50, $\rho$ =
388	0.71; Figure 4c). Dengue dynamics were also highly predictable from incidence data
389	alone ( $R^2 = 0.89$ , $\rho = 0.94$ ; Figure 4a); thus, robust to missing state variables. This
390	implies that EDM methods could be powerful for forecasting epidemics, provided
391	that surveillance efforts continue to report weekly case data. Third, EDM revealed
392	that climate effects on dengue appeared once the susceptible population size
393	exceeded a threshold ( $ ilde{\lambda}$ > 0.85; Figure 6).

The fact that climate effects are observed before the onset of an outbreak, when  $\tilde{\lambda} =$ 1 (Rypdal & Sugihara 2019), suggests that rainfall, and possibly temperature, have an effect on the timing of an impending epidemic. Thus, although rainfall and temperature might not influence the dynamics of dengue during an inter-outbreak

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398	period when the susceptible population size is small, climate could act as a catalyst
399	to spark an epidemic once the susceptible population size is large enough. This
400	resonates with the notion that climate could drive the force of infection, $eta$ , thus
401	influencing the susceptibles index, $\lambda$ (Rypdal & Sugihara 2019). The timing of an
402	outbreak, when $\lambda \geq 0$ (or in the discrete case when $  ilde{\lambda} \geq 1$ ), could be attributed to
403	the changes in $eta$ caused by seasonal climatic drivers (Rypdal & Sugihara 2019).
404	Further, seasonality of temperature and rainfall had higher predictive skill than
405	seasonality of susceptibles index (Figures 3 and S8, grey solid lines); however,
406	adding susceptibles index dramatically improved dengue forecasts due to more
407	accurate epidemic magnitudes (Figure 4b, c). Thus, climate may be mostly
408	responsible for the <i>timing</i> of seasonal epidemics, while susceptible population size
409	may mostly determine the epidemic <i>magnitude</i> . Using the same dataset, Johnson <i>et</i>
410	al. (2018) found that mechanistic models could predict the timing of seasonal
411	epidemics, but that a phenomenological machine learning component was needed to
412	capture interannual variation in epidemic magnitude. Our work suggests that the
413	unobserved size of the susceptible population was a key missing link for predicting
414	magnitude variation across years.
415	We showed that susceptible population size is an important modifier of climate

We showed that susceptible population size is an important modifier of climate
effects on dengue (Figure 6). This climate-susceptible population interdependence
might be generalizable across other infectious diseases. For example, for influenza
dynamics, population density in cities—potentially a proxy for susceptible
population density—modulated climate effects on disease transmission (Dalziel *et al.* 2018). Given complex immune and serotype dynamics of dengue (Katzelnick *et*

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421	al. 2017b), total population density might not work as a proxy for the susceptible
422	population density in dengue dynamics. It has been difficult for previous
423	mechanistic models to capture susceptible dynamics for dengue. By inferring the
424	susceptibles index from incidence data (Rypdal & Sugihara 2019), we were able to
425	capture the strong influence of the susceptible population size on dengue dynamics,
426	which in turn moderated the effect of climate drivers on dengue dynamics. We
427	showed that climate-dengue relationships were only detectable once we accounted
428	for seasonal and interannual variation in susceptible availability (Figure 6d, f). This
429	is expected from theory (Kermack & McKendrick 1927; Xu et al. 2017), but
430	demonstrating it empirically is a unique contribution of this study.
121	Even when accounting for suscentible availability the offects of temperature and
431	Even when accounting for susceptible availability, the effects of temperature and
432	rainfall on dengue were still strongly state-dependent (Figure 6d, f). The remaining
433	variation in temperature and rainfall effects on dengue—given that the susceptible
434	population is large enough for an outbreak—may be partially explained by variation
435	in temperature and rainfall over time and space that is not captured by weekly
436	climate averages, and by interactions between temperature and rainfall. Further,
437	any subtle lagged effects of temperature or rainfall on dengue are not captured by a
438	single ecological lag (a 9-week or an averaged 3–9-week lag, respectively). We
439	expect multiple time lags to operate based on the different ecological processes that
440	generate reported cases: mosquito population growth, mosquitoes getting infected
441	and biting susceptible humans, development of symptoms, and care-seeking. The
442	ecological lags are simple proxies for this complex set of ecological processes.

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443	EDM is like any other quantitative method in that it is appropriately applied when
444	its assumptions are met. EDM is sensitive to stochasticity and synchrony (e.g., via
445	seasonality), so it is important to first determine whether a system is governed by
446	deterministic dynamics, and then to use null models that account for synchrony
447	(Ebisuzaki 1997; Deyle <i>et al.</i> 2016a; Chang <i>et al.</i> 2017). Nevertheless, EDM is a
448	useful approach for testing causality between hypothesized drivers and response
449	variables in a dynamic system, and recent work allows EDM to perform well with
450	stochasticity (Cenci & Saavedra 2018; Cenci et al. 2019). To infer strongly supported
451	causal relationships, it is also important to consider the mechanistic hypotheses
452	underlying the system. For dengue, there are complementary lines of evidence
453	about the drivers of incidence from experimental studies (Watts et al. 1987;
454	Lambrechts et al. 2011), mechanistic models parameterized with data (Otero et al.
455	2006; Mordecai <i>et al.</i> 2017), phenomenological studies across settings (Johansson <i>et</i>
456	al. 2009; Hii et al. 2012; Ibarra et al. 2013), and EDM from this study. In particular,
457	the 19-year weekly time series dataset used here has been investigated using
458	multiple methods (Johansson et al. 2019); however, most of these methods are
459	either not mechanistic or include simple mechanisms that do not account for the
460	joint influence of climate and susceptible dynamics. EDM gives us the opportunity to
461	infer mechanisms, and assess their support in other studies using different
462	approaches. In sum, EDM has limitations, but the limitations of other approaches are
463	at least as great, and may not capture state-dependent effects. When combined with
464	supporting evidence from other approaches, EDM can be a powerful test for

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465 causality, prediction, and functional relationships between drivers and response466 variables.

467	Connecting climate and dengue at the population level is challenging, because causal
468	relationships are likely to be nonlinear and state-dependent. Thus, a toolbox of
469	rigorous methods for testing hypotheses, understanding mechanisms, and making
470	predictions is essential for understanding disease dynamics in complex, natural
471	populations. Ultimately, understanding how climate-driven vector-borne diseases
472	are influenced by other variables, such as susceptible population size, is important
473	for optimizing vector control under critical conditions where climate might spark
474	epidemics. EDM is most appropriate for deterministic systems, which are common
475	in disease ecology. The mechanisms inferred from EDM could be applied to
476	understand and predict future ecological responses to changing environments,
477	including dengue epidemics in a world undergoing global change.

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## 493 AUTHORSHIP

- 494 NN, EAM, GS, and MR conceived the idea for the study and designed the analyses.
- 495 MR inferred the proxy for susceptible population size time series. NN, ERD, MSS,
- 496 AJM, and MLC performed the analyses. NN wrote the first draft of the manuscript. All
- 497 authors contributed to manuscript revisions and gave approval for publication.

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# 646 **FIGURES**





649 incidence (i.e., total number of cases per week), (b) weekly average temperature, (c)

- total weekly rainfall, and (d) a proxy for susceptible population size (see Supporting
- 651 Information for details) in San Juan, Puerto Rico.

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## Figure 3. Susceptible population size and climate drive dengue incidence. 662 663 Cross-mapping between dengue incidence and potential drivers—temperature with a 9-week lag (a; red), rainfall with an averaged 3–9-week lag (b; blue), and proxy for 664 susceptible population size $(\tilde{\lambda})$ with a 5 week lag (c; purple)—display significant 665 666 (Kendall's test $\tau > 0$ : p < 0.01) convergence in cross-mapping skill (i.e., $\rho$ increases and reaches a flat asymptote) as the number of time series data points increases 667 668 (sign of causality). Cross-mapping skill is the Pearson's correlation coefficient, $\rho_{i}$ 669 between predicted driver values using the univariable state space reconstruction of 670 the response variable, and the observed driver values. Rainfall and susceptibles 671 index showed significant forcing above and beyond seasonal signal (p < 0.05), 672 because cross-mapping of the true time series (blue and purple) are distinguishable 673 from their respective null models (grey). The red, blue and purple shaded regions 674 represent the 0.025 and 0.975 quantiles of bootstrapped time series segments. The 675 grey shaded regions represent the 0.025 and 0.975 quantiles of the seasonal null 676 distributions obtained from 500 runs of randomized time series with conserved 677 seasonal trends (Deyle et al. 2016a). The grey line represents the median of the null 678 distribution.

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680 **Figure 4. Predictive power of climate and susceptibles index** ( $\tilde{\lambda}$ ). Forecasting results showing observed (black) and predicted (green) values of dengue incidence 681 682 (cases per week) using state space reconstruction (SSR). Univariable SSR with just 683 incidence data (a) illustrates that dengue incidence is strongly predictable ( $R^2$  = 684 0.8922,  $\rho$  = 0.9446). Multivariable SSR using only temperature and rainfall (b) 685 predicts dengue incidence less well ( $R^2 = 0.0533$ ,  $\rho = 0.2309$ ), especially the 686 magnitudes of the outbreaks, but the seasonal trend is captured. Multivariable SSR 687 using temperature, rainfall, and the proxy for susceptible population size  $\tilde{\lambda}$  (c) 688 predicts incidence well ( $R^2 = 0.5044$ ,  $\rho = 0.7102$ ). Forecast predictions improve 689 slightly ( $R^2 = 0.8927$ ,  $\rho = 0.9448$ ) when the drivers temperature, rainfall, and 690 susceptibles index are added to incidence SSR (d). All R<sup>2</sup> values represent adjusted 691  $R^2$ . All SSR models (Figure 4a–d) had significant (p < 0.001) F statistics >> 1 based 692 on ANOVA and the null hypothesis was rejected (H<sub>0</sub>: there is no relationship 693 between predicted and observed dengue incidence).

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Figure 5. Temperature and rainfall show mixed functional responses of

696 dengue incidence. Scenario exploration quantifies the variable effect of changes in 697 drivers on dengue. Boxplots show that the median effects of rainfall (Rain) and 698 temperature (Temp) are small (close to zero), but drivers occasionally have strong 699 impacts (a). To investigate climate driver functional responses, we plotted the rate 700 of change of dengue incidence as a function of temperature (b) and the rate of 701 change of dengue incidence as a function of rainfall (c). Red and blue lines represent 702 regression on the median for temperature and rainfall, respectively. The dashed red 703 and blue lines represent regression on the 0.05 and 0.95 quantiles of temperature 704 and rainfall, respectively. Temperature has an overall positive effect on dengue 705 incidence (median regression line of the rate of change is positive), but can also 706 have large negative and positive effects (a, b). Rainfall has an overall negative effect 707 (median regression line of the rate of change is negative), but can also have small 708 positive effects (a, c). Some of this effect variation can be explained by the 709 susceptible population size (Figure 6).

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711Figure 6. Temperature and rainfall effects on dengue vary depending on the712susceptible population size. Effects of climate drivers (i.e., temperature and713rainfall) are investigated in relation to the proxy for susceptible population size,  $\tilde{\lambda}$ .714Plotting the effect of changes in temperature (a) and rainfall (b) against the715susceptible population size shows that driver effects are split around the threshold716of  $\tilde{\lambda} = 0.85$  (purple dashed line). The red and blue lines represent the median717regression of temperature and rainfall effects, respectively (a-f). The dashed red

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718	and blue lines represent the 0.5 and 0.95 quantile regressions of temperature and
719	rainfall effects, respectively (a–f). Neither driver has an effect when there is not a
720	sufficiently large susceptible population size ( $ ilde{\lambda} <$ 0.85) for an outbreak (c, e).
721	However, above a certain susceptible population size the climate effects are
722	observed: temperature with mostly a (possibly sigmoidal, see Figure S9) positive
723	effect (d) and rainfall with a negative effect (f). However, even when driver effects
724	are split at the evident threshold of $\tilde{\lambda}=0.85$ (c–f), there are still many occurrences
725	when the susceptible population size is sufficient large ( $ ilde{\lambda} > 0.85$ ) but temperature
726	and rainfall have no effect. In certain cases, temperature has even a negative effect
727	on dengue (d).