Dorsal anterior cingulate cortex neurons
only partially integrate determinants of value

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ABSTRACT

Evaluation often involves integrating dissimilar determinants of value. A brain region can therefore be placed either before or after a presumed evaluation stage by measuring whether responses of its neurons depend simultaneously on multiple determinants of value. A brain region could also, in principle, show partial integration, which would suggest that it occupies a middle position between (pre-evaluative) non-integration and (post-evaluative) full integration. Existing mathematical techniques cannot distinguish full from partial integration and therefore risk misidentifying regional function. Here we use a new Bayesian regression-based approach to analyze responses of neurons in dorsal anterior cingulate cortex (dACC) to risky offers. We find that dACC neurons only partially integrate across outcome dimensions, indicating that dACC cannot be assigned to a purely post-evaluative position. Neurons in dACC also show putative signatures of value comparison, suggesting that comparison processes do not require evaluation to be complete to proceed.

INTRODUCTION

To evaluate an option, we must consider all of the aspects of the option that influence value and then combine them to generate an integrated value signal (Rangel et al., 2008; Kable & Glimcher, 2009; Busemeyer et al., 2019; Tversky, 1969; Tversky & Kahneman, 1974). For example, a risky prospect (a gamble) may involve two possible outcomes, both of which have some influence on our likelihood of choosing the gamble itself. To evaluate the gamble, we must decide how appealing each of those outcomes is and then somehow combine those component evaluations. This could involve a process of multiplication and addition, as dictated by normative approaches, or could be any of a number of satisficing heuristic approaches (Stewart et al., 2006; Busemeyer and Townsend, 1998; Tversky and Kahneman, 1974; Farashahi et al., 2019). Clearly, this integration process must be reified in the brain—we necessarily process dimensions separately and our actions reflect only the integrated value, so combination has to occur somewhere in the middle (Bowman et al., 2012; Roesch and Olson, 2003; Roesch and Olson, 2004). How it occurs remains an important problem in neuroeconomics.

One important goal in neuroeconomics is identifying where this integration occurs (Rangel et al., 2009). Scientists therefore often seek to categorize brain regions as either preceding or following the integration process. A standard approach is to probe the responses of single neurons and test whether their responses are driven by multiple determinants of value (Padoa-Schioppa and Assad, 2006; Kahnt et al., 2011; Hosokawa et al., 2013; Kennerley et al. 2009; Blanchard et al., 2015; So and Stuphorn, 2010; O'Neill & Schultz, 2018; Raghuraman & Padoa-Schioppa, 2014). Simultaneous selectivity for multiple factors that determine value/utility has therefore become a hallmark of cross-dimensional integration (Padoa-Schioppa, 2011; Hunt et al., 2015; Fellows, 2006; Kim et al., 2008; O'Neill and Schultz, 2018; Strait et al., 2014).

However, this interpretation relies on the assumption that the process of evaluation is

anatomically discrete. Another possibility would be that it unfolds gradually across many

regions, much as complex form is gradually built up over multiple regions of the ventral visual

stream (Yoo and Hayden, 2018; Hunt and Hayden, 2017; Chen and Stuphorn, 2015; Cisek,

2012). That is, reward-sensitive regions may be aligned along a hierarchy, and the dimensions

may become gradually more united the further we travel along it. Critically, the standard

statistical approach of detecting dependence on multiple determinants of value does not

distinguish full (i.e. complete) from partial integration. We recently developed a method that

makes use of Bayesian regression and cross-validation that can distinguish partial from complete

correlations across dimensions in neural ensemble regressions, although we used it for a different

purpose (Azab and Hayden, 2017). Here we use that methods to address this question.

This approach also allows us to address a question about how evaluation relates to

comparison. Conventional approaches choice assume the completion of evaluation before the

start of comparison (Rangel et al., 2008, Padoa-Schioppa, 2011). For example, competitive

inhibition models, the dominant models of economic choice, involve separate pools of integrated

(i.e. post-evaluation) value neurons competing (Chau et al., 2014; Rustichini and Padoa-

Schioppa, 2015; Hunt et al., 2012; Hunt et al., 2015; Louie et al., 2011). However, there is no

special reason why comparison must wait until evaluation is complete (Yoo and Hayden, 2018).

In distributed choice systems, such as beehives, comparison processes can progress even in the

absence of full cross-dimensional integration. For example, individual bees that participate in a

swarm decision may sample only a subset of available determinant dimensions and nonetheless

contribution directly to comparison (Seeley, 2009; Seeley & Buhrman, 1999). It is not clear, in

the human brain, whether evaluation is complete before comparison can begin, or whether

comparison starts soon after evaluation does (Hunt et al., 2014; Cisek, 2012; Chen & Stuphorn,

2015; Strait et al., 2015; Azab & Hayden, 2018; Balasubramani et al., 2018). Cognitive models

of value-based decision-making have, in fact, often proposed that the evaluation and comparison

stages occur in parallel (Busemeyer et al., 2019; Noguchi & Stewart, 2018). Knowing whether

signatures of integration across dimensions is partial or complete can help shed light on this

question.

These questions are especially interesting to ask in the dorsal anterior cingulate cortex

(dACC). Among value-sensitive or economic regions, there is a broad consensus that this is an

anatomically or hierarchically late region (Heilbronner and Hayden, 2016; Rushworth et al.,

2011; Schall et al., 2002; Wunderlich et al., 2009). That is, it may serve as a repository for feed-

forward signals generated by other value-sensitive regions, such as orbitofrontal cortex and

ventromedial prefrontal cortex, and have a direct role in influencing motor and premotor

processing that other economic regions lack. As a consequence, it stands to reason that, among

value-sensitive regions it should be most likely to carry fully-integrated signals. Indeed, current

debates about the economic function of the dACC are generally split between (late) post-

comparison selection role and (middle) post-evaluative comparison role (Hare et al. 2011;

Wunderlich, Rangel & O'Doherty, 2009; Rangel & Hare, 2010; Azab & Hayden, 2017).

Here, we analyzed responses of neurons in dACC during a two-option token gambling

task (Azab & Hayden, 2017; Azab & Hayden, 2018). We find that while neurons encode both

large and small prospective outcomes of each gamble, their tuning functions for these outcomes

are similar but are demonstrably not identical. In other words, dACC responses reflect partial

integration across dimensions, not complete integration. We also observed putative signatures of

comparison in these neurons, suggest that comparison does not require evaluation to be complete

120 to proceed. These results suggest that dACC, even though it is normally placed late in the choice

121 hierarchy, does not follow the completion of integration. Moreover, we believe these results are

122 more consistent with distributed than with highly modular accounts of economic choice.

Figure 1: ***A:*** *Example trial from token-gambling task.* **B.** *Behavior for each subject, fit to a sigmoid function. Subjects choose the left option more often as its value increases, as would be expected given understanding of the task. EV: expected value of gamble.* **C:** *Regions of interest (for exact coordinates, see Methods). Figure adapted with permission from Azab & Hayden, 2018.*

RESULTS

Behavior is informed by integrated value

Details of the *token risky choice task* are provided in the **Methods** and are illustrated in **Figure 1A** (see also Azab and Hayden, 2017 and 2018). On each trial, a macaque subject chose between two gambles. Gambles offered the possible gain or loss of virtual tokens. Tokens accumulated across trials; once subjects accrued six, they received a large water reward. Gambles consisted of two possible outcomes (which we denote "large" and "small"—the small one was most often negative or zero). Gambles appeared asynchronously; this staggered presentation allowed us to examine neuronal responses to the first gamble independent of the value of the second.

Gamble values influence subjects' choices

Regressor	Subject B (n = 66 sessions)		Subject J (n = 74 sessions)	
	Mean regression coefficient	Wilcoxon sign-rank Z-statistic (p-value)	Mean regression coefficient	Wilcoxon sign-rank Z-statistic (p-value)
Left offer – win	1.40	7.06 (p = 1.64 x 10^{-12})	1.15	7.47 (p = 7.773 x 10^{-14})
Left offer – loss	0.486	6.97 (p = 3.26 x 10^{-12})	0.293	7.22 (p = 6.35 x 10^{-13})
Left offer – probability of win	5.12	7.06 (p = 1.64 x 10^{-12})	3.70	7.47 (p = 7.73 x 10^{-14})
Right offer – win	-1.33	-7.06 (p = 1.65 x 10^{-12})	-1.33	-7.47 (p = 1.67 x 10^{-12})
Right offer – loss	-0.412	-6.83 (p = 8.37 x 10^{-12})	-0.237	-7.06 (p = 1.67 x 10^{-12})

Right offer – probability of win	-4.85	-7.06 (p = 1.64 x 10^{-12})	-4.09	-7.47 (p = 7.73 x 10^{-14})

Table 1: *Subjects' choices were influenced by all values characterizing both gambles. Average regression coefficients from a multiple logistic regression model of choice (left=1, right=0) against the variables in the regressor column, for each behavioral session. The distribution of regression coefficients across sessions for every behavioral variable listed deviated significantly from zero. A 'loss' within a gamble was always less than or equal to the win outcome of that gamble, and may or may not be a negative value. The magnitude of the regression coefficient indicated the extent to which this variable influenced choice, while its sign indicated whether higher values of this variable (positive vs. negative) favored choice of the left vs. right gamble, respectively. Adapted with permission from Azab & Hayden, 2018.*

Details of the subjects' behavior are provided elsewhere (Azab and Hayden, 2017; Azab & Hayden, 2018; Farashahi et al., 2018). Briefly, both subjects preferred the option with the higher expected value (subject B: 80.3%; subject J: 75.1%; two-sided binomial test: both P < 0.0001). Behavior was significantly influenced by all three components (the two outcomes and the probability) of both gambles (see **Table 1**).

We first tested whether *choices* reflect the integrated values of the gambles or the components of each gamble independently. That is, whether the weighted product of outcome magnitudes and probabilities explained behavior over and above these gamble components individually. We fit choices to two multiple linear regression models, one including the three individual components of each gamble and another that included their interaction. For both subjects, inclusion of the integrated value predictor improved model fit (Wilcoxon sign-rank test between Akaike Information Criterion, AIC, subject B: z=7.06; subject J: z=7.47; p<0.0001 for both subjects). These results provide important evidence that our subjects behaviorally integrate the values. In other words, they indicate that the final common pathway (the oculomotor control neurons) encode integrated value, and therefore indicate that the integration must occur before or at that point (Roesch and Olson, 2003; Roesch and Olson, 2004).

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Figure 2: *Neural responses to gamble components. **A:** Responses of example neurons to*
the large outcome (left), probability of large outcome (middle), and small outcome (right) of the
*first gamble. Gray region: period of significant modulation. **B:** Proportion of neurons modulated*
by the components of the first gamble throughout the course of the trial. This fraction was
computed by fitting a stepwise regression model to the normalized firing rate over a 500 ms
sliding window against the components of the attended gamble (left) or the chosen gamble
(right), along with other task-relevant variables (see Methods). Apparent effects before time zero
reflect the width of the analysis window.
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Responses of dACC neurons carry value-relevant information

We recorded 129 neurons from dACC in two subjects (subject B: 55 neurons, subject J:

74 neurons; see **Figure 1B** for region of interest). Basics of modulation in response to task-

relevant variables are presented in detail in previous manuscripts and are not repeated here (Azab

& Hayden, 2017 and 2018). We use the same epoch we used in our previous papers, a 500 ms

window starting 100 ms after events. Example cells are shown in **Figure 2**. For the first example

cell, lower low values result in higher firing rates (t-test on firing rates, p<0.001). For the second

example cell, higher high values result in higher firing rates (p<0.005). For the third example

cell, firing depended on the probability of the large outcome (p<0.001).

Responses of 17.05% of neurons were affected by the size of the large outcome

(n=22/129, one-sided binomial test: p<0.0001). Responses of 20.2% were affected by the

probability of that outcome (n=26/129, p<0.0001), and 8.53% were affected by the small

outcome (n=11/129, p=0.059) of the same gamble in the first epoch. Note that the lack of a

significant effect for the small outcome raises the possibility that dACC simply does not encode

this parameter. However, more sensitive tests do show an effect. Specifically, we performed a

stepwise linear regression model; a significant fraction of neurons showed dependence on the

small outcome in both epochs (epoch 1: 34.1%, n = 44/129, P < 0.0001; epoch 2: 55.0%, n =

71/129, two-sided binomial test: P < 0.0001).

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Figure 3: *Large and small outcomes are encoded in partially-separable formats in the*
recorded population. ***A:*** *Correlation between formats of large and small outcomes for the first*
gamble in the first epoch. (left) Scatter-plot showing the correlation between regression
coefficients for large and small outcomes for each neuron. Red line indicates the positive
correlation between regression coefficients. Shaded red region indicates the 95% credible
interval on the data correlation. Black dashed line indicates the estimated ceiling correlation
between these variables in this epoch. (right) Results of permutation test. Black histogoram
indicates the distribution of the estimated ceiling correlation from permuted data. Red dashed
line indicates degree of correlation observed in the data. ***B:*** *Correlation between variable*
formats for the second gamble in the second epoch. ***C.*** *Correlation between variable formats for*
the first gamble in the second epoch. ***D:*** *Correlation between variable formats for the chosen*
*gamble in the pre-choice epoch. *: p < 0.05, ** p < 0.01.*
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dACC neurons use partially integrated formats to encode gamble outcomes

We have used the term *ensemble coding format* (or just *format*) to refer to the vector of

tuning coefficients for all recorded neurons (Strait et al., 2014; Wang and Hayden, 2017). We

then correlate formats across conditions to determine how they relate. Bayesian regression

allows us to draw samples from the likely format (rather than the single estimate conventional

regression provides, Gelman et al., 2013). All correlations below use Spearman (i.e. rank-order)

correlation, which is more robust to outliers than Pearson correlation, and is thus a more

conservative method. This analysis approach has a good deal of conceptual resemblance to

representational similarity analysis (RSA; Kriegeskorte et al., 2008); the main practical

difference is that it uses parameter tuning (specifically, regression weights) rather than raw

response (e.g. firing rates), making it more appropriate for the questions we want to ask here (see

Discussion).

We found that ensemble coding formats for large and small outcomes of each offer are

positively correlated (r=0.28, p<0.001, 1000 samples). That is, our population uses overlapping

(i.e. more similar than expected by chance) neural codes to signal the values of the large and

small outcomes; this is a signature of cross-dimensional integration (Strait et al., 2014;

Blanchard et al., 2015).

But does it reflect a *full* cross-dimensional integration? Mathematically speaking, why is

the measured correlation coefficient less than 1.0? It could reflect intermediate levels of

integration or it could reflect complete integration that appears intermediate due only to

measurement noise. (Recall that neurons are inherently noisy, so we can be certain that such

noise is a major factor). To differentiate these possibilities, we repeated the same analysis on a

shuffled dataset (where the values of the two variables of interest are shuffled *within* trial). This

tells us the maximum correlation observable between the encoding of two variables given the

noise in our data, which we call the *ceiling correlation*.

The ceiling correlation for large and small outcomes in the first epoch was r=0.45

(standard deviation = 0.047; **Figure 3A**). This result indicates that noise is high in our dataset (as

it is in most neural recordings) and that good deal of the explanation for the low correlation we

observed is simply noise. However, our observed correlation (r = 0.28, see above) is nonetheless

significantly lower than this value (p=0.001, bootstrap significance test, 1,000 permutations).

Thus, although large and small outcomes are encoded in similar formats, these formats are not

identical, as one would expect if dACC outputs reflected a completed value computation.

Mathematically speaking, the difference in formats necessarily reflects a qualitative (non-affine)

difference in tuning—a consistent gain shift or baseline shift would not be detected by our

method and would not generate a correlation significantly different from ceiling. In other words,

dACC neurons show evidence simultaneously of vectors corresponding integration and non-

integration, and these vectors are combined in our sample of neurons.

We found a similar pattern when considering the second gamble in the *second offer*

epoch (t=100 to 600 ms after the second offer appears). Specifically, although we did not find

evidence of integration across the two dimensions (r=0.059, p=0.060, compared to zero), we

found that whatever integration existed was less than the ceiling correlation value (r=0.17;

p=0.028; Figure 3B). Thus, it appears that dACC does not fully integrate either offer. (Note that,

because it is a null result, the lack of integration observed for the second offer does not prove no

integration occurred, although it is consistent with that possibility.)

We next asked whether full integration occurs when the information about the gamble is

transferred to working memory. It appears not. The coding of the first offer during the second

offer epoch (when it was no longer on screen) resulted in the same pattern: partial but not

complete integration (r=0.18, p<0.001 for comparison with zero. Ceiling correlation was r=0.27,

which was larger than the observed value, p=0.031; **Figure 3C**).

dACC continues to show partial integration around the time of choice

We next examined the 500 ms immediately preceding choice fixation (*pre-choice epoch*).

Whereas above we focused on offered gambles, here we focused on the *chosen gamble*, with the

reasoning that identifying an option for choice would potentially result in full integration. The

formats for integration of the two stakes is, indeed, positive (r=0.16, p<0.001). However, as like

in previous epochs, this correlation is less than ceiling (r=0.25, p=0.047; Figure 3D). These

results thus show that separation in large- and small-outcome formats persists at least until right

before choice is indicated.

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Figure 4: *Large outcomes and their probabilities are encoded in integrated formats in*
*the recorded population. **A:** Correlation between formats of large outcomes and probabilities for*
the first gamble in the first epoch. (left) Scatter-plot showing the correlation between regression
coefficients for large outcomes and probabilities for each neuron. Red line indicates the positive
correlation between regression coefficients. Shaded red region indicates the 95% credible
interval on the data correlation. Black dashed line indicates the estimated ceiling correlation
between these variables in this epoch. (right) Results of permutation test. Black histogoram

Formats used to encode stakes and probabilities are consistent with full integration

for those dimensions

For the first gamble in the first epoch, we find that the large outcome and its probability

were encoded in similar formats (r=0.54, p<0.001). This degree of correlation was not

significantly less than our estimated ceiling (ceiling r=0.58, p=0.135; Figure 4A). While we

cannot definitively conclude that there is full integration of these variables, the difference

between these results and the previous ones is striking. Indeed, for the first offer in the second

epoch (i.e. when it is in working memory), large outcomes and probabilities are also encoded in

similar formats (r=0.44, p<0.001) that are no different from ceiling (r=0.46, p=0.39; Figure 4B).

For the second offer, large outcomes and their probabilities are also encoded in similar formats

(r=0.37, p<0.001), and again no different from ceiling (r=0.41, p=0.22; Figure 4C). The chosen

gambles' large outcomes and probabilities are encoded in similar formats in the 500 ms

preceding choice (r=0.49, p<0.001), and again no different from ceiling (r=0.47, p=0.63; Figure

4D). Although the lack of a difference is not sufficient to prove identity, the matching patterns

for all four conditions, and their striking differences with the stakes analyses above, suggest that

large stakes and probability may be fully integrated in dACC.

These results are important because they indicate that the lack of integration observed for

large and small stakes is not a consequence of our task design, of the psychology of our subjects,

or of a limitation in our analysis approach. It appears that dACC is capable of fully integrating

some variables and maintaining some separation in others, and our methods are capable of

detecting this partial integration when it occurs.

This final point is further emphasized by our next analysis. In the first epoch, the small

outcomes and probabilities were encoded in similar formats ($r=0.11$, $p=0.002$), which are less

correlated than ceiling ($r=0.30$, $p=0.001$). Results are not conclusive for the same gamble in the

second epoch: there is no correlation between the formats for small outcomes and large-outcome

probabilities ($r=0.050$, $p=0.11$), but this correlation is no different from our estimated ceiling,

which is very low ($r=0.13$, $p=0.057$). The small outcome and large-outcome probability of the

second gamble are encoded in weakly-correlated formats in the second epoch ($r=0.063$,

$p=0.049$), and this correlation is no less than our estimated ceiling (also low, ceiling $r=0.13$,

$p=0.12$). The low stakes and probability components of the chosen gamble were encoded in

positively-correlated formats in the pre-choice epoch ($r=0.17$, $p<0.001$); no less than ceiling

($r=0.15$, $p=0.65$). The lack of full integration between small outcomes and probabilities in the

first epoch (which is highly significant and survives Bonferroni correction) highlights the

qualitatively different way in which large and small offer outcomes are treated.

DISCUSSION

We examined the neural encoding of risky offers in dACC. We made use of a recently-

developed statistical method that can disambiguate variable encodings that are partially

correlated from ones that are fully correlated but different due solely to noise (Azab and Hayden,

2017). We find, first, that while dACC neurons encode both large and small prospective

outcomes of individual gambles on the screen and in working memory, the population tuning

formats for the two outcomes are only partially correlated—neither orthogonal nor collinear.

The "mid-evaluation" signal we observe is a type that is predicted by hierarchical

(sometimes called serial) models of value, which involve multiple small steps leading from a

space of options to one of actions rather than a single value estimation locus (Hunt & Hayden,

2017; Yoo & Hayden, 2018). In such models, reward and value processing occur along an

anatomical hierarchy and computation within that hierarchy implements a gradual transformation

from stimulus information to motor commands, rather than a series of categorical steps (Hunt et

al., 2012; Hunt & Hayden, 2017; Cisek, 2012; Chen & Stuphorn, 2015; Cisek & Kalaska, 2010).

More specifically, we have proposed that the relationship between different hierarchical levels

may be one of untangling—of rotating representations into a format more usable for effector

systems, while retaining information (Yoo and Hayden, 2018; DiCarlo et al.. 2012). In the

domain of economic choice the tangled format involves the two non-integrated dimensions and

the untangled format is an integrated value representation. The untangling theory, which is based

on how the visual stream works, proposes that each anatomical stage untangles information from

the earlier stage and allows for intermediate amounts of untangling.

The first major effect we observe suggests that dACC follows the initiation of value

integration but precedes its completion. Together these results suggest that dACC cannot be

assigned either a purely pre- or post-evaluation stage in choice. Moreover, in conjunction with

other results showing evidence for value comparison in dACC (e.g. Azab & Hayden, 2017; Azab

& Hayden, 2018; Klein-Flugge & Bestmann, 2012; Hare et al. 2011; Wunderlich et al., 2009;

Rangel & Hare, 2010), these findings indicate that value integration does not need to be

complete before value comparison begins (Hunt et al., 2014).

354   These results complement our recent findings using this dataset (Azab & Hayden, 2017).

In that study, we concluded that dACC occupies a middle role in the *comparison* of values. Our

complementary results here suggest that dACC occupies a middle role in the *computation* of

values. Juxtaposed, the two results have a larger implication: that the completion of integration is

not essential to the initiation of comparison. That is, it may be impossible, even in theory, to

draw a line, either in anatomical space or in time, between evaluation and comparison stages of

choice (Yoo & Hayden, 2018).

361   The identification of abstract amodal value signals is a desideratum of neuroeconomics

(Shizgal, 1997; Padoa-Schioppa, 2011; O'Doherty, 2014; Montague & Berns, 2002; Levy &

Glimcher, 2012). Our results suggest that, if these exist, they are most likely to occur only

downstream of dACC. We envision two major possibilities about this site of value

representation. First, value representations may only be complete in the premotor or motor

system (see Cisek, & Kalaska 2010, Cisek, 2007, and Thura & Cisek, 2016 for related

arguments). Second, that may be no single location at which value is represented abstractly and

completely (Chen & Stuphorn, 2015; Cisek, 2012; Yoo & Hayden, 2018; Hunt et al., 2015;

Balasubramani et al., 2018).

370   Ours is not the first study to consider representations of multi-attribute offers (O'Neill &

Schultz, 2018; Raghuraman & Padoa-Schioppa, 2014; Kennerley et al., 2009; Strait et al., 2014;

Blanchard et al., 2015; Padoa-Schioppa & Assad, 2006; Hunt et al., 2015). A broad finding from

these studies is that neurons in several regions show some overlap in the way they encode the

multiple dimensional determinants of value. The major difference is that we show that

integration of attributes into value representations is not an all-or-none phenomenon. The

distinction between attributes that are not integrated vs. somewhat integrated may not be

meaningful for readout, since this information is extractable in either case—future work is

needed to know what information is read out of dACC responses.

We have proposed that what we call a value may be better thought of as evidence in favor

of the proposition that the option will be chosen (Azab & Hayden, 2017; see also Hayden &

Moreno-Bote, 2018). From this perspective, there is no reason to expect that valuation must be

complete before comparison can begin. Each dimension can provide its own evidence in favor or

against the choice of an option. That evidence can be summed in one specific place or it can be

fed into a slowly evolving action plan without ever being integrated in one place. We and others

have previously argued that the analogy to bee swarm decision-making is helpful for

understanding choice (Hunt & Hayden, 2017; Eisenreich et al., 2017; Couzin, 2009; Passino et

al., 2008; Seeley et al., 2012; Pirrone et al., 2018). Individual bees evaluate hive sites in an

extremely noisy fashion (Seeley, 2009; Seeley & Buhrman, 1999). Indeed, an individual scout

may only assess one or two of the several dimensions along which hives vary. Even if no

individual bee performs a full cross-dimensional integration, the hive as a whole can still make a

good decision—without ever computing anything like value—through a quorum sensing

procedure. And the hive as a whole takes several days to select a hive site. In the middle of this

time, the hive may be said to have made a partially-complete evaluation process.

Similarly, some *cognitive* (rather than neural) models of value-based decision-making do

not make this assumption of discrete, serial stages; ordinal within-attribute comparisons

oftentimes contribute to the construction of each option's value. In this framework, comparison

would, in some sense, *precede* evaluation, and all value signals are entirely relative (Noguchi &

Stewart, 2018; Busemeyer et al., 2019). These models—also ones of value-based decision-

making, albeit at a different level of granularity—clearly demonstrate the importance of

integrating different levels of analysis when exploring the inner workings of a cognitive process,

and the many ways that a system (biological or otherwise) could potentially instantiate it.

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MATERIALS AND METHODS

The data used here were analyzed and summarized in other manuscripts (Strait et al., 2015; Azab & Hayden, 2017; Azab & Hayden, 2018; Farashahi et al., 2018).

Surgical Procedures

All procedures were approved by the University Committee on Animal Resources at the University of Rochester and were designed and conducted in compliance with the Public Health Service's Guide for the Care and Use of Animals. Two male rhesus macaques (*Macaca mulatta*: subject B age 5y. 7mo.; subject J age 6y. 7mo. at the start of recording) served as subjects in both studies. We used standard procedures, as described previously (Strait et al., 2014). A small prosthesis for holding the head was used. Animals were habituated to laboratory conditions and then trained to perform oculomotor tasks for liquid reward. A Cilux recording chamber (Crist Instruments, Hagerstown, Maryland, USA) was placed over the dACC and attached to the calvarium with ceramic screws. Appropriate anaesthesia was used at all times; induction was performed with ketamine and isoflurane was used for maintenance. For surgical induction, we used 10–15 mg/kg of ketamine, 0.25 mg/kg of midazolam, and 2– 4 mg/kg of propofol. For maintenance, we used isoflurane, ad lib level, set depending on active monitoring procedure. For systemic antibiotics, we used cefazolin and for topical application, we used standard veterinary triple antibiotic. For analgesics, we used meloxicam, and, when judged necessary by veterinary staff, buprenorphine. Post-operative care included close monitoring and restoration of fluid intake. Animals received appropriate analgesics and antibiotics after all procedures. Position was verified by magnetic resonance imaging with the aid of a Brainsight system (Rogue Research Inc., Montreal, Quebec, Canada). Throughout both behavioral and physiological recording sessions, the chamber was kept sterile with regular antibiotic washes and sealed with sterile caps.

Recording Site

We approached dACC through a standard recording grid (Crist Instruments). We defined dACC according to the Paxinos atlas (Paxinos et al, 2000). Roughly, we recorded from a ROI lying within the coronal planes situated between 29.50 and 34.50 mm rostral to interaural plane, the horizontal planes situated between 4.12 to 7.52 mm from the brain's dorsal surface, and the sagittal planes between 0 and 5.24 mm from medial wall (Figure 1B). Our recordings were made from a central region within this zone. We confirmed recording location before each recording session using our Brainsight system with structural magnetic resonance images taken before the experiment. Neuroimaging was performed at the Rochester Center for Brain Imaging, on a Siemens 3T MAGNETOM Trio Tim using 0.5 mm voxels. We confirmed recording locations by listening for characteristic sounds of white and gray matter during recording, which in all cases matched the loci indicated by the Brainsight system.

Electrophysiological Techniques

Single electrodes (Frederick Haer & Co., Bowdoin, Maine, USA; impedance range 0.8 to 4 MU) were lowered using a microdrive (NAN Instruments, Nazaret Illit, Israel) until waveforms of between one and three neuron(s) were isolated. Individual action potentials were isolated on a Plexon system (Plexon, Inc., Dallas, Texas, USA). Neurons were selected for study solely on the basis of the quality of isolation; we never pre-selected based on task-related response properties. All collected neurons for which we managed to obtain at least 250 trials were analyzed.

711
Eye Tracking and Reward Delivery
Eye position was sampled at 1,000 Hz by an infrared eye-monitoring camera system (SR
Research, Ottawa, Ontario, Canada). Stimuli were controlled by a computer running Matlab
(Mathworks) with Psychtoolbox (Brainard, 1997) and Eyelink Toolbox (Cornelissen, Peters &
Palmer, 2002). Visual stimuli were colored rectangles on a computer monitor placed 57 cm from
the animal and centered on its eyes. A standard solenoid valve controlled the duration of juice
delivery. The relationship between solenoid open time and juice volume was established and
confirmed before, during, and after recording.
720
Behavioral Task
Subjects performed a two-option gambling task. The task was conceptually similar to
risky choice tasks we have used previously (Strait et al., 2014), with two major differences: (1)
monkeys gambled for virtual tokens—rather than liquid—rewards, and thus (2) outcomes could
be losses as well as wins. Previous training history for these subjects included a set shifting task
(Sleezer et al., 2016), a simple choice task (Heilbronner and Hayden, 2016), an attentional task
(Hayden and Gallant, 2013), and several foraging and foraging-inspired tasks (see Calhoun and
Hayden, 2014; Hayden, 2018)
Two offers were presented on each trial. Each offer was represented by a rectangle 300
pixels tall and 80 pixels wide (11.35° of visual angle tall and 4.08° of visual angle wide). 20% of
options were safe (100% probability of either 0 or 1 token), while the remaining 80% were
gambles. Safe offers were entirely red (0 tokens) or blue (1 token). The size of each portion
indicated the probability of the respective reward. Each gamble rectangle was divided
horizontally into a top and bottom portion, each colored according to the token reward offered.
Gamble offers were thus defined by three parameters: two possible token outcomes, and
probability of the top outcome (the probability of the bottom was strictly determined by the
probability of the top). The top outcome was 10%, 30%, 50%, 70% or 90% likely on gamble
offers.
Six initially unfilled circles arranged horizontally at the bottom of the screen indicated
the number of tokens to be collected before the subject obtained a liquid reward. These circles
were filled appropriately at the end of each trial, according to the outcome of that trial. When 6
or more tokens were collected, the tokens were covered with a solid rectangle while a liquid
reward was delivered. Tokens beyond 6 did not carry over, nor could number of tokens fall
below zero.
On each trial, one offer appeared on the left side of the screen and the other appeared on
the right. Offers were separated from the fixation point by 550 pixels (27.53° of visual angle).
The side of the first offer (left and right) was randomized by trial. Each offer appeared for 600
ms and was followed by a 150 ms blank period. Monkeys were free to fixate upon the offers
when they appeared (and in our observations almost always did so). After the offers were
presented separately, a central fixation spot appeared and the monkey fixated on it for 100 ms.
Following this, both offers appeared simultaneously and the animal indicated its choice by
shifting gaze to its preferred offer and maintaining fixation on it for 200 ms. Failure to maintain
gaze for 200 ms did not lead to the end of the trial, but instead returned the monkey to a choice
state; thus, monkeys were free to change their mind if they did so within 200 ms (although in our
observations, they seldom did so). A successful 200 ms fixation was followed by a 750 ms delay,
after which the gamble was resolved and a small reward (100 μL) was delivered—regardless of

the outcome of the gamble—to sustain motivation. This small reward was delivered within a 300
ms window. If 6 tokens were collected, a further delay of 500 ms was followed by a large liquid
reward (300 μL) within a 300 ms window, followed by a random inter-trial interval (ITI)
between 0.5 and 1.5 s. If 6 tokens were not collected, subjects proceeded immediately to the ITI.
761   Each gamble included at least one positive or zero-outcome, ensuring that every gamble
762   carried the possibility of a win (or at least no change in tokens). This decreased the number of
763   trivial choices presented to subjects, and maintained motivation.
764
765   **Statistical Methods**
766   **Peri-stimulus time histograms (PSTHs)** were constructed by aligning spike rasters to
767   the presentation of the first offer and averaging firing rates across multiple trials. Firing rates
768   were calculated in 20 ms bins but were generally analyzed in longer epochs. For latency
769   analyses, we used PSTHs with firing rates calculated in 5 ms bins to grant us finer resolution to
770   detect shorter latencies.
771   Firing rates were **normalized** by subtracting the mean and dividing by the standard
772   deviation of the neuron's entire PSTH (i.e. z-scoring).
773   We tested for correlations between variables of interest, since these could produce
774   spurious results in our format and population correlation analyses. Although gamble values were
775   chosen independently, safe gambles (where the large and small outcome were of the same value)
776   introduced a spurious correlation between large and small outcomes. We thus exclude trials
777   where the gamble of interest was a safe gamble, as we have done in previous manuscripts (refs).
778   We tested for single-unit modulation using a **multiple generalized linear regression**
model, including the following task-relevant variables: large outcomes, small outcomes, and
probability of the large outcomes for gambles 1 (and two, if that gamble had been presented), the
number of tokens collected as of the beginning of the trial, and the side the first offer appeared
on. We also included the side of the chosen offer for analyses in the pre-choice and post-choice
epochs. We use the same variables to fit stepwise linear regression models to assess small-
outcome modulation.
785   **Analysis epochs** were chosen *a priori*, before data analysis began, to align with epochs
we have chosen for previous studies (Strait et al., 2014; Azab & Hayden, 2017; Azab & Hayden,
2018). The first and second offer epochs were defined as the 500 ms epoch beginning 100 ms
after the offer was presented. The pre-choice epoch was the 500 ms epoch before choice was
indicated using an express saccade. The post-choice epoch was the 500 ms epoch immediately
after choice fixation was complete, and before reward feedback was revealed. These epochs were
chosen before data collection began; indeed, they originally were selected for a previous study
on a different brain region, and to account for hypothesized delays in neural responding (Strait et
al., 2014). We maintained the same epochs here as a mechanism to reduce the likelihood of p-
hacking.
795   All **fractions of neurons** were tested for significance using a one-sided binomial test.
796   When testing for significant encoding in the population, we use an alpha level of 0.05 to indicate
797   chance (i.e. the number of neurons that would exhibit significant modulation at random).
798   **Format and population correlation analyses** were performed in the following manner.
We use **beta correlation analyses** to assess whether neurons represented two variables (or the
same variable at different time periods) using similar / orthogonal / opposing formats, in
overlapping / orthogonal / distinct populations. To do this, we first found the regression
coefficient associated with the variable of interest per neuron. We estimate this using a multiple

linear regression model including the same variables detailed above (for spatial location
analyses, we split trials by the side of offer appearance, and thus do not include that variable in
our regression model). We then combined the regression coefficients associated with the variable
of interest into a vector of the same length as the number of neurons in our sample. This vector
indicates the relative strength (after normalizing) and direction of modulation for each individual
neuron in the population, in response to a particular variable in a particular epoch. We call this
the population "format". We compared different formats by finding the Spearman correlation
coefficient between them.
Note that this approach provides a more sensitive method of examining population
properties than conventional approaches, which involve determining which cells cross a
significance threshold and then using those as focal cells for further analyses. By taking on all
cells regardless of their response, our method accomplishes two things. First, it uses all available
information, even that information that is not sufficient by itself to achieve significance. Second,
it avoids introducing a hard categorical boundary, which can introduce false positives in data
(Blanchard et al., 2017; Maxwell & Delaney, 1993).
We extend this method to account for the noise inherent in estimating each neuron's
encoding of each variable of interest, which existing methods do not account for (Azab &
Hayden, 2017). We used a Bayesian regression to obtain a probabilistic distribution over each
regression coefficient for each neuron (rather than an individual coefficient estimate per neuron;
this is akin to taking into account the confidence interval on each regression coefficient
estimate). We sampled 1,000 regression coefficients from this distribution for each neuron, to
obtain a probability distribution of (1,000) potential formats for the population, for each of two
task conditions. We then performed the correlation analyses on each of these samples across task
conditions, thus generating a probability distribution of (1,000) correlation coefficients. This is a
more robust estimate of the correlation between formats, as it takes into account the uncertainty
inherent in estimating any individual regression coefficient, and allows us to view the spread of
the distribution of this correlation when this significant source of noise is taken into account.
The Spearman correlation coefficient between signed regression coefficients indicated
whether variables were represented in a similar *format* i.e. directionality of tuning across the
population. (Note that we used Spearman instead of Pearson because it is more robust to outliers
and is therefore more statistically conservative). A positive correlation indicated a preservation
of directionality, while a negative correlation suggested variables were represented in opposing
directionality of firing rate modulation. No correlation suggests orthogonal formats, but we draw
no strong conclusions from these.
Similarly, the correlation coefficient between unsigned regression coefficients indicated
whether similar neuronal populations tended to be involved in encoding the two variables in
question, regardless of their direction of modulation. A positive correlation indicated overlapping
populations, while a negative correlation indicated separate ones. A lack of correlation suggests
orthogonal populations (i.e. encoding one variable does not affect the neuron's likelihood of
encoding the other variable), although given the lack of interpretability of null results, it does not
constitute evidence for lack of true correlation.
It is important to note that the analysis method we use here is rooted in previous analysis
techniques used by our lab and others. The closest relative is Representational Similarity
Analysis (RSA; Kriegeskorte, Mur & Bandettini, 2008; Hunt et al., 2018). In RSA, neural
representations (vectors of neural responses) associated with different behavioral conditions are
compared and their similarity characterized to facilitate comparison with different data sets,

sometimes collected using different means. Rather than comparing raw neural responses, though,

our method relies on comparing regression coefficients; this allows us to hone in on the signal

associated with a particular variable, allowing us to compare the neural responses to different

variables at the same point in time.