The fate of standing variation and new mutation under climate change

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Abstract:

Many species face existence threat under anthropogenic climate change, and standing genetic

variation was proposed as a way for sessile species to adapt to novel environments. However,

it is still unknown whether standing genetic variants, being adaptive to current environmental

variability, are sufficient to guarantee future survival. Here we investigate the relative

importance of standing variation versus new mutations and their relative effect sizes from the

past to infer the future. While theoretical and experimental evolution studies provided valuable

insights, examples in nature are scarce. In the wild banana *Musa itinerans* in Taiwan, we found

enrichment of standing variation with environmental associations. New mutations exert larger

effect size per variant in precipitation-related climatic variables, where Taiwan contains

extreme environments beyond the ancestral climatic range, and novel alleles have stronger

association with novel environments. This pattern is reversed for temperature-related variables,

where Taiwan is within the ancestral environmental range. Despite such differences,

anthropogenic climate change impacts both types of variants strongly. The results constitute

one of the few examples in nature demonstrating that the effect sizes of adaptive variants differ

under distinct environmental pressures, and the patterns support theoretical predictions that

natural selection favors new mutations with larger effect sizes in a novel environment where

the population is far from the adaptive peak. Despite the importance of new mutations, however,

the pace of anthropogenic climate change may not allow the accumulation of such mutations.

Introduction:

Anthropogenic climatic change posts an imminent threat to most organisms. For large and

sessile plant species with long generation time, the speed of migration may not keep up with

environmental change, and therefore phenotypic plasticity and genetic variation in the

population may allow their survival under novel environments[1-4]. Adaptive genetic variation

originates from standing variants before or new mutations after environmental change. Since

anthropogenic climate change greatly outpaces natural mutation, the amount of standing

genetic variation is therefore critical for the rapid response of a population to changing

environments[5-7]. It remains unclear, however, whether standing variants are sufficient to

guarantee future survival, given that they are mostly adaptive to the present range of climatic

variability. While it may be difficult to perform manipulative experiments in the field to

compare the effects of new mutations (NM) and standing variation (SV), one could investigate

NM and SV during climatic change in the past.

Adaptation could happen through genetic variants that differ in their origins (NM or SV)

or effect sizes (Mendelian genes with major effects or polygenic variants with minor effects).

However, how these factors interact and respond to environmental pressures remains relatively

uninvestigated. For example, does SV and NM differ in their relative number or effect sizes

towards environmental adaptation, and how does this relationship change with different types

of environmental factor? For adaptive new mutations that were fixed when facing environment

change, Fisher first predicted primarily small allelic effects[8] while Kimura emphasized

intermediate effects[9]. Orr, later considering the entire adaptive walk, concluded the evolution

towards a novel adaptive peak should first happen through fixation of large-effect mutations

and later by small-effect polymorphisms[10]. While this was supported by some studies, the

majority of these are microbial experimental evolution in well-controlled environments[11,12],

and few have specifically compared the effects of NM and SV. To test whether this idea holds

in nature, empirical investigations on natural populations are needed.

Taiwan is well-suited for such studies: Unlike oceanic islands such as Hawaii, Taiwan is

a continental island where most species originated from the East Asian continent with recurrent

gene flow[13]. The land bridge between Taiwan and China during the glacial maximum allowed

exchange of SV, and the isolation during interglacial periods enabled the development of NM.

Here we investigate the genomic basis of environmental adaptation of a wild banana species,

Musa itinerans, whose habitats in Taiwan are considered peripheral from ancestral area

reconstructions[14], providing an opportunity to distinguish SV from NM, as well as their

response to ancestral versus novel adaptive landscapes. We investigated how past events (SV

versus NM) influence present adaptation and whether local adaptation may persist under future

anthropogenic climate change.

Results:

Environmental adaptation in *Musa itinerans*

We first sampled *Musa itinerans* at 24 populations across Taiwan (Fig. 1a; Supplementary
Fig. 1a; Supplementary Table 1) and investigated the population structure using 14
microsatellites (Supplementary Table 2). Environmental Niche Modeling (with 483 occurrence
points from field survey and Google Street View) reported species distribution (Fig. 1b) in line
with the previous statement that *Musa itinerans* inhabits sunny valleys, watersheds, and
hillsides with gentle slopes[15]. Populations differentiated mostly between east and west
(Supplementary Fig. 1b). The most unsuitable environments lay within the Central Mountain
Range and the southwestern plains (Fig. 1b), respectively corresponding to low annual mean
temperature (BIO1) and low precipitation of driest quarter (BIO17), the two most important
bioclimatic variables determining species distribution (MaxEnt permutation importance[16]: 36.7
for BIO1 and 27.3 for BIO17).

To test for local adaptation, we examined the pattern of "isolation by adaptation[17]", a
process where differential local adaptation restricted effective gene flow and promoted genetic
differentiation among populations, by dissecting geographic and environmental effects on
genetic differentiation. The strait-line "fly-over" geographical distance, calculated as the
straight-line distance between locations, does not explain patterns of genetic differentiation
(Mantel's $r = 0.146$ and $P = 0.062$). However, if we considered that Central Mountain Range
lacks corridors for *M. itinerans* to disperse (Fig. 1b), this fly-over geographical distance could
be too unrealistic. We therefore used resistance distance, calculated from the route with least
resistance among populations on the niche suitability map (Fig. 1b), to represent the "realized"
geographical distance (Fig. 1c) and found that genetic differentiation was significantly
associated with resistance (Mantel's $r = 0.226$ and $P = 0.006$). The environmental Mahalanobis
distance of bioclimatic variables also showed strong association with genetic differentiation
(Mantel's $r = 0.298$ and $P = 0.005$). Given that the environmental distance could be strongly
dependent on geography, we performed Partial Mantel test to control the geographical effect.
After controlling for realized geographic distance (resistance distance), genetic differentiation
still correlates with the Mahalanobis environmental distance (Mantel's $r = 0.250$ and $P = 0.012$),
suggesting differential local adaptation is associated with genetic variation.

Standing variation versus new mutations

To identify genomic regions associated with environmental adaptation, we performed
pooled sequencing for each population. SNPs were separated into standing variation (SV: both
alleles exist in Taiwan and China) or new mutations (NM, polymorphic only in Taiwan). SV
outnumbered NM in both adaptive (identified with Bayenv[18]) and non-adaptive SNPs, and after
controlling for the overall number of SNPs in SV and NM, SV were further enriched among
adaptive polymorphisms (Supplementary Table 3). However, since adaptive SNPs also have

higher minor allele frequency (MAF) (Supplementary Fig. 2a), this pattern could be
confounded: SV are more likely to have higher MAF than NM, and SNPs with higher MAF
may be more likely detected as adaptive due to higher statistical power. We therefore
performed the same test with a subset where the adaptive and non-adaptive SNPs have similar
allele frequencies (ranging from the first quantile of adaptive MAF to the third quantile of non-
adaptive MAF separately for each bioclimatic variable; Supplementary Fig. 2a). In this case as
well, SV are still disproportionately abundant (Supplementary Table 4), suggesting SV are
more likely than NM to become environment-associated SNPs. Another potential confounding
factor is the geographic extent of variants: if most NM resulted from mutations restricted to a
few local populations, the limited distribution prevents environment association for NM. We
therefore compared the number of Taiwanese populations containing the minor alleles for SV
and NM, respectively. Contrary to the direction predicted by the aforementioned confounding
factor, the geographic extent of minor alleles for SV is slightly smaller than NM (15.8
populations for SV and 16.3 for NM, $P < 0.001$).

In addition to SNP number, do NM and SV differ in their directions of effect? Under the
null hypothesis that (1) the effects of NM are equally likely to facilitate adaptation to the
ancestral or novel environments and (2) natural selection is equally likely to fix NM facilitating
adaptation towards either direction, we expected no enrichment of new alleles in either
environment. When we separated the Taiwanese populations into those within the Chinese
ancestral environmental range and those with novel environments, frequencies of putatively
adaptive new alleles in NM SNPs were higher in the latter set of populations, with precipitation
of driest quarter (BIO17) and precipitation of coldest quarter (BIO19) showing the strongest
effect (Fig. 2a; Supplementary Fig. 2b). Given that the directions of mutation effects should be
random, these results suggest that new variants facilitating adaptation to novel environments
are more likely to be retained by selection.

The results above suggest SV might be more important than NM in terms of enriched
number of variants with environment association, and adaptive NM, while lower in number,
are more associated with novel environments. On the other hand, the number of candidate SNPs
does not necessarily reflect the overall importance of SV over NM, since the effect size also
needs to be considered. To investigate the effect size of SV and NM in environmental
adaptation, we compared their Bayes factors from Bayenv and focused on two bioclimatic
variables that were important in determining species distribution, annual mean temperature
(BIO1) and precipitation of driest quarter (BIO17). The two variables exhibit opposite patterns
(Fig. 2c): While in BIO17 and other related precipitation variables (Supplementary Fig. 3a)
NM consistently had higher Bayes factor and therefore stronger effect size than SV, in BIO1
and other related temperature variables (Supplementary Fig. 3a) we observed the reverse. The
same trend was observed when we estimated the effect size with gradient forest[19,20] (Fig. 2b, c;
Supplementary Fig. 3b, c): BIO17 was the most important factor for differential local

adaptation, and NM had stronger effects than SV. On the other hand, BIO1 was the least
important factor where SV had stronger effects. Finally, the "importance" estimated by
gradient forest is analogous to r^2, representing the amount of allele frequency variation
explained by environmental gradients. Assuming a simple linear relationship between allele
frequency and environment, the value of r^2 only represents how well each data point (a
population) fits along the regression line. We were, however, also interested in the regression
slope: the amount of allele frequency changes along environmental gradients (Supplementary
Fig. 3d). Again, BIO17 had the largest overall slope among all bioclimatic variables
(Supplementary Fig. 3e), with NM being significantly higher than SV (Fig. 2c). BIO1 had the
lowest overall slope, again with the reversed pattern. Given that the MAF of adaptive NM and
SV SNPs are similar, there is no need to control for allele frequency in these tests
(Supplementary Fig. 4a). Therefore, NM with larger effect size per SNP (as estimated by
Bayenv Bayes factor, gradient forest r^2 weighted importance, and gradient forest slope) were
associated with the adaptation to novel environments outside of the ancestral niche space,
consistent with previous population genetics modeling results[8-10].

The observed patterns could be integrated with the unique climate of Taiwan. In
comparison to the rest of the species range, northern Taiwan experiences northeastern
monsoons during winter and has higher precipitation during the typical winter dry season (Fig.
3b). This pattern has been maintained since at least the last glacial maximum (Fig. 3d). The
novel environments might impose novel adaptive optima to the immigrant population from
China. The response to selection imposed by these environmental gradients is strong (with
highest Bayes factor, r^2, and slope among all bioclimatic variables; Supplementary Fig. 3a, c,
e) especially for NM, where new alleles are strongly associated with novel environments
(Supplementary Fig. 2b). More importantly, for this major driver of adaptation (BIO17), the
greatest increment of gradient forest importance lies between 200 mm and 300 mm (Fig. 2b),
a range also distinguishing the novel Taiwanese versus ancestral Chinese environments (Fig.
3b). This suggests that the majority of differential local adaptation is associated with such
novel-versus-ancestral environmental differences. Annual mean temperature (BIO1) is the
other extreme: the environmental gradient within Taiwan is well within the ancestral Chinese
environmental range (Fig. 3a), which can also be traced back to the last glacial maximum (Fig.
3c). It is likely that SV already contained genetic variants adaptive to such environmental
gradients and are therefore more important than NM (Fig. 2c). In summary, we observed
adaptation happening through the assortment of SV for a new territory with similar adaptive
landscape and optimum (BIO1; Figs. 2c, 3a, c). For adaptation to novel environments and a
new adaptive landscape (BIO17; Figs. 2c, 3b, d), NM with larger effect sizes were more likely
favored by natural selection. Our results are therefore consistent with Orr's model[3], providing
one of the few examples in nature.

One key point of this study is the correct designation of SV or NM. It is possible that some
SV SNPs were mis-assigned as NM if we missed an allele in China, most likely for SNPs with
low MAF. We addressed this issue with the following: (1) The Taiwanese populations were
nested within China in the phylogenetic tree (Supplementary Fig. 4b) and contained much less
genetic variation (Supplementary Fig. 4c). Due to the stronger genetic drift in Taiwan than
China, it is less likely that an originally SV SNP would retain both alleles in Taiwan but lose
one in China. (2) In the extreme case, assuming 50% of NM SNPs were mis-assigned from SV,
we performed 100 new analyses, each randomly assigning 50% of NM SNPs back to SV. These
new analyses yielded similar results, with NM having higher effect sizes than SV in
precipitation-related variables (Supplementary Fig. 5a). (3) Since MAF are correlated between
Taiwan and China (Spearman's rank correlation $\rho = 0.35$, $P < 0.001$), we performed analyses
with top 50% MAF SNPs in Taiwan, thereby reducing the chance of missing minor alleles in
China. The results are qualitatively the same (Supplementary Fig. 5b-d).
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Fate of adaptive variants under future climate change
In addition to understanding how past events (SV versus NM) affected present adaptation,
we are also concerned with how these factors affect the future of this species under
anthropogenic climate change. We predicted 16 future outcomes and used Bayenv to
investigate the future fate of currently adaptive SNPs. Currently adaptive SNPs retaining high
association with future environments are classified as "retention", while those no longer
associated with future environments are "disruption". Different from the present pattern of SV
being enriched in adaptive SNPs, we saw no clear tendency for any set of adaptive SNPs
enriched towards retention or disruption (Supplementary Table 5), suggesting both SV and NM
will be affected by climate change.
Under the 16 future scenarios we identified two extreme conditions: CCSM4-2070-
RCP8.5 as the niche-expansion extreme (Fig. 4a-d) and MIROC-2070-RCP4.5 as the niche-
contraction extreme (Fig. 4e-h). Intending to investigate whether SV and NM have different
fates under the two contrasting future predictions of this species, we used the genetic offset
value from the gradient forest results to estimate genetic mismatch for SV and NM separately,
which is associated with the magnitude of allele frequency turnover perturbed by future
climatic conditions[20] (Fig. 4b, f). The "genetic offset difference" was calculated by subtracting
genetic offset caused by SV from that of NM (Fig. 4c, g). We found no general consensus of
whether or where in Taiwan NM or SV SNPs are more affected by climate change (Fig. 4c, g),
suggesting NM and SV might both be affected by future anthropogenic climate change despite
their distinct genetic architecture in adaptation to current environments.
Traditional species distribution models build the species-wide niche model from all
occurrence records of a species, thereby assuming all populations in a species reacting equally
to the environment and overlooking local adaptation due to within-species polymorphism[6].

Here we propose a concept, "extinction risk", integrating species-level and population-level
responses to future climate change simply by dividing genetic offset (estimated from gradient
forest) by suitability (estimated from MaxEnt). While the two future conditions exhibit slight
differences in suitability and genetic offset, the extinction risks are surprisingly similar, with
very high risk in western Taiwan (Fig. 4d, h). In Western Taiwan, the expansion-extreme model
suggested high suitability for the species as a whole but also high local population mismatch,
thereby having similar trends of extinction risk as the contraction-extreme model. Taken
together, our results emphasize the importance of considering both whole-species suitability as
well as adaptation of local populations when considering the effect of anthropogenic climate
change on the fate of sessile species.
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Discussion:
Assessing the within-species variation in climate association is the crucial first step to
understand species susceptibility to fluctuating environment, and the relative importance of
standing variation (SV) and new mutations (NM) in adaptation has long been debated[7,21-25]. In
this study, we investigate how past genetic variation (SV and NM) contribute to the adaptation
of present environments as well as how these factors together affect the future fate of a species
under anthropogenic climate change.
For a population facing environmental change and therefore a novel adaptive landscape,
previous population genetics models have documented the effect size distribution of new
mutations fixed by natural selection. Considering the entire adaptive walk of a population
facing novel environments, Orr's model predicted the distribution of effect sizes where early
substitutions have larger effect than later ones[10]. Consistent with Orr's prediction, we show
that NM have stronger effect size than SV in precipitation-related variables, where Taiwan
exceeds the ancestral climatic range in China (and therefore the migrating population was far
away from the optimum in the new adaptive landscape). This pattern is reversed for
temperature-related variables, where Taiwan has similar environmental range as China. Here
we provide another perspective to recent research showing that SV contributes to adaptation[7,25].
We show that SV indeed dominate over NM in number[7]. However, the effect size of SV and
NM hinges on environmental conditions: natural selection may prefer new mutations
contributing to the adaptation to the new rather than the old environment, and the effect sizes
of NM tend to be higher than SV under such conditions. Our results imply that standing genetic
variation in a population may not be sufficient for the adaptation to anthropogenic climate
change.
Ecological niche models are widely used to predict future species distribution, but such
models often do not account for the differential adaptation of populations within species. Our
observed difference between the major determinant of species-wide range distribution (BIO1)
and the driver of differential adaptation of populations within species (BIO17) demonstrates

the need to consider local adaptation. Incorporating species distribution modeling (predicting
the future range of the whole species) and genetic offset (predicting the genetic mismatch of
each local population to future environments), we identified the potential risk of western
lowland populations despite the seemingly distinct species distribution modeling results in
different future climate scenarios. In other words, despite in some future scenarios the range of
environments suitable for the whole species might increase, giving the impression that the
species benefits from anthropogenic climate change, such change might be detrimental to each
local population uniquely adapted to a much narrower environmental range than the whole
species. Finally, we separately calculated the genetic offset originated from the mismatch of
SV or NM to future environments and showed that both will be affected by anthropogenic
climate change, regardless of their distinct genetic architecture towards environmental
adaptation. Since we distinguished SNPs into SV or NM based on past environmental change,
all of these genetic variants will be standing variation when facing future climate change. Novel
mutations facilitating the adaptation to new environments may happen in the future, although
they will be strongly outpaced by anthropogenic climate change.
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Methods:
Sample collection and DNA extraction.
Field work was conducted during 2017 (August - December) and 2018 (January - May).
We sampled *Musa itinerans* at 24 sites across Taiwan (Supplementary Fig. 1a; Supplementary
Table 1). Fresh leaves were harvested from nine to fifteen individuals at each site. Total
genomic DNA was extracted using the standard CTAB extraction method[26]. Since other
commercial *Musa* species were also grown in Taiwan, we developed an indel marker for
species delimitation. From previous studies[27-30], we identified a 6-bp insertion specific for the
Taiwanese *M. itinerans* in the *atpB-rbcL* region of chloroplast. We designed a primer pair (5'-
GAAGGGGTAGGATTGATTCTCA-3'; 5'-CGACTTGGCATGGCACTATT-3') and used
amplicon size to confirm all collected samples are Taiwanese *M. itinerans*.
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Simple sequence repeat genotyping and analysis.
SSR primer sequences used in this study were originally developed for the genus *Musa*[31,32],
which were then applied on *Musa itinerans*. Previously documented primer sequences were
first searched against the *Musa acuminata* DH-Pahang genome version 2[33] on Banana Genome
Hub (https://banana-genome-hub.southgreen.fr/) to check specificity as having only one
amplicon, resulting in 26 primer pairs. These primers were then experimented to check
specificity on *M. itinerans*, resulting in 14 pairs (Supplementary Table 2). We modified each
pair of primers by capping the 5' end of forward primers with M13 sequences
(CACGACGTTGTAAAACGAC) and inflorescent molecules[34]. SSR amplicons were run
through capillary electrophoresis and the length of each allele was recorded.

Population structure of 20 populations (Supplementary Table 1) was analyzed with 14
SSR markers (Supplementary Table 2). Lowland populations (C35H, WFL, THNL, PTWT,
P199H, MLLYT, HDPG, TTL, NAJY, HLCN, NXIR, and DFR), east transect populations
(TPS300, TPS500, TPS700, and TPS900), and west transect populations (XT400, XT700,
XT1200, and XT1500) were used in the analysis. We inferred the ancestry of 244 individuals
with STRUCTURE 2.3.4[35,36], parameterizing a run to have (1) run length of burnin and after-
burnin period of 100,000, (2) admixture ancestry model, and (3) independent allele frequency
model, further setting 20 runs for each K value.
To investigate the association among genetic, geographical, and environmental distance,
we generated these distance matrices. Genetic distance was calculated by GenAlEx 6.503[37,38]
from 14 SSR markers; straight geographical distance (the fly-over distance) was generated by
ArcGIS 10.5 (http://desktop.arcgis.com/en/); environmental distance was measured as
Mahalanobis distance to address the correlation among nine bioclimatic variables (below). In
addition to the fly-over geographical distance which assumes organism dispersal ignores
landscapes, we further calculated as resistance distance the cumulative cost along the least cost
path (below). Matrix association was examined under Mantel and Partial Mantel tests.
Statistical significance was examined with 1,000 permutations. We performed Mantel tests on
(1) genetic distance *vs*. fly-over distance, (2) genetic distance *vs*. resistance distance, (3) genetic
distance *vs*. Mahalanobis environmental distance, and Partial Mantel tests on (4) genetic
distance *vs*. Mahalanobis environmental distance while controlling for resistance distance.
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Species distribution modeling.
Current and future species distribution models were built for *Musa itinerans* using
presence-only data (483 occurrence points) obtained from field survey and Google Street view.
Occurrence points were then reduced to 204 cells by the removal of co-occurring presence data
within the same 1×1 km grid. MaxEnt 3.4.1[16], implemented with the maximum entropy
modeling approach, reports an overall niche suitability and the importance of predictors by
analyzing the presence-only data as well as background (psuedo-absence) data distribution[39-
41]. We downloaded from WorldClim database version 1.4 (http://worldclim.org/) spatial layers
of 19 present-day bioclimatic variables based on high-resolution monthly temperature and
rainfall data[42]. Layers were selected at spatial resolution of 30 arc-second and with a mask that
ranges 119.25-122.47°E and 21.76-25.49°N covering Taiwan. Variables showing high
dependence (Pearson's correlation coefficient > 0.9 calculated from ENMTools[43]) from each
other were removed, resulting in nine final variables: BIO1－mean annual temperature, BIO2
－mean diurnal range, BIO3－isothermality, BIO7－temperature annual range, BIO12－
annual precipitation, BIO15－precipitation seasonality, BIO16－precipitation of wettest
quarter, BIO17－precipitation of driest quarter, and BIO19－precipitation of coldest quarter
(Supplementary Table 1).

Present species distribution model was constructed using the default optimization settings
in MaxEnt, except the regularization set to three. We tested the predictive model by ten-fold
cross-validation which was carried out by randomly partitioning the data into ten equally sized
subsets and then replicating models while omitting one subset in turn. In each turn, the
predictive model was built using nine subsets as training data and evaluated using the other
subset as test data. The output of the predictive model is the probability of presence, or called
suitability, and we averaged the ten runs to have an averaged suitability.

To predict the species distribution under different scenarios of future climate change, we
projected the present-day model onto eight future climatic conditions combining two periods
(2050 and 2070) and four Representative Concentration Pathways (RCP 2.6, RCP 4.5, RCP 6.0,
and RCP 8.5). Future climatic layers were obtained from the WorldClim database at spatial
resolution of 30 arc-second and were developed based on two general circulation models: the
Community Climate System Model[44], CCSM, and the Model for Interdisciplinary Research on
Climate[45], MIROC. Species distribution models for the future were carried out using the same
settings described above.

To estimate the least cost path between populations, we first generated the resistance
surface by taking the reciprocal of suitability. Resistance and suitability is simply a monotonic
transformation in which locations with higher suitability exhibit lower resistance. Pairwise
least cost path was then measured among 20 populations from the resistance surface, performed
by SDM Toolbox v2.3[46]. While least cost path is the single line with least overall cost, we also
constructed the least cost corridor between populations, allowing 1%, 2%, or 5% higher cost
than the least cost value. In essence, the least cost corridors represent the realized dispersal
routes of organisms along suitable habitats.
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Sequencing library construction and SNP identification.
We conducted whole genome pooled-sequencing[47] for each population (Supplementary
Table 1), resulting in 24 pooled-sequencing libraries. Equal amount of DNA from ten
individuals at each population were pooled, except for the PTWT population where only nine
individuals were available. A library with 300-400 bp insert size for each pool was prepared
using NEBNext Ultra II DNA Library Prep Kit (New England Biolabs). Libraries were then
sequenced with 150 bp paired-end on the HiSeq X Ten platform.

Illumina reads were then trimmed with SolexaQA[48], followed by the removal of adaptor
sequences with cutadapt[49], subsequently mapped to the *Musa itinerans* reference genome
assembly   ASM164941v1[50]   with   BWA   0.7.15[51].   Picard   Tools
(http://broadinstitute.github.io/picard) was used to mark duplicated read pairs, and the
genotypes were called following GATK 3.7 best practice[52].

For the 24 pooled samples, we filtered out sites with (1) more than two alleles, (2) indels,
(3) quality (QUAL) < 30, (4) quality by depth (QD) < 2, (5) call rate < 0.74, and (6) depth (DP)

> genome-wide average depth plus three standard deviations, resulting in 4,200,177 SNPs.

SNPs with (1) minor allele frequency (MAF) < 0.05, (2) missing data in any of the pooled-seq

sample, and (3) DP per sample < 20 were further filtered out, resulting in 1,256,894 SNPs.

To investigate the relationship between Taiwanese and Chinese *M. itinerans*, we

downloaded public data from 24 Chinese accessions (SRR6382516 - SRR6382539)[53]. SNPs

were called using all 24 Chinese and the 24 Taiwanese samples together following the pipeline

described above. We did not perform any site filtering for this joint data set since the main

objective is to investigate whether specific SNPs in Taiwan also existed in China as SV. This

dataset has 18,442,853 SNPs. SRR6382532 was excluded due to high missing rate. Only when

evaluating the averaged expected heterozygosity between Taiwanese and Chinese populations

did we filter out sites with (1) indels and (2) QUAL < 30, resulting in 15,591,923 SNPs.

To assess the phylogeny of our Taiwanese populations and Chinese accessions, we

downloaded *Musa acuminata* sequence (SRR7013754) as an outgroup. SNPs were called using

one *M. acuminata* species, 24 Chinese and 24 Taiwanese samples together following the

pipeline described above. We filtered out sites with (1) more than two alleles, (2) indels, (3)

QUAL < 30, and (4) call rate < 0.9, resulting in 12,693,687 SNPs. This dataset also excluded

SRR6382532.

Environmentally-associated SNP identification.

We used Bayenv 2.0[18] to search for SNPs highly associated with environmental variables.

Bayenv estimates the relationship between SNPs and environments while controlling the

whole-genome population structure from a subset of loose linkage-disequilibrium SNPs. Loose

linkage-disequilibrium SNPs were formed by sampling (1) one SNP from scaffolds more than

10 kb and less than 100 kb, (2) two SNPs from scaffolds more than 100 kb and less than 500

kb, (3) three SNPs from scaffolds more than 500 kb and less than 1000 kb, and (4) four SNPs

from scaffolds more than 1000 kb. We then, for each bioclimatic variable, defined as the

adaptive SNPs ones exhibiting top 1% Bayes factor and top 5% rho value (a nonparametric

correlation coefficient capable to reduce outlier effects).

We further investigated the fate of currently adaptive SNPs under anthropogenic climate

change, performing the same Bayenv analyses of currently adaptive SNPs using future climatic

conditions. We included two time periods (2050 and 2070) and four Representative

Concentration Pathways (RCP 2.6, RCP 4.5, RCP 6.0, and RCP 8.5) from two general

circulation models, CCSM[44] and MIROC[45], resulting in 16 future climatic conditions. If a

currently adaptive SNP remains strongly associated with environments, it should exhibit Bayes

factor above the current threshold. We then defined as "retention" a currently adaptive SNP

constantly exhibiting Bayes factor above the current adaptive threshold in all future scenarios,

and defined as "disruption" a currently adaptive SNP exhibiting Bayes factor above the current

adaptive threshold in none of the future scenarios.

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408	**The gradient forest method and genetic offset.**
409	We used a novel method, gradient forest[19,20], to estimate the effect of environmental
410	gradients on allele frequency differences among populations. Gradient forest is a regression-
411	tree based machine-learning algorithm using environmental variables to partition SNP allele
412	frequencies. The analysis was done separately for each SNP. The resulting "importance"
413	measures how much of the variation in allele frequency was explained by partitioning the
414	populations based on a specific value in an environmental variable. By making multiple
415	regression trees (thus generating a random forest) for a SNP, the goodness-of-fit r^2 of a random
416	forest is measured as the proportion of variance explained by this random forest, which is then
417	partitioned among environmental variables in proportion to their conditional importance. Such
418	SNP-wise importance of each environmental variable is then averaged across SNPs belonging
419	to the standing variation (SV) or new mutations (NM) group, resulting in the overall importance
420	(of each environmental variable). In the end, one could obtain the relation curve between
421	environmental gradient and cumulative importance (analogous to the cumulative r^2, proportion
422	of allele frequency differences among populations explained by environments). This curve has
423	two properties. First, the highest point of the cumulative importance curve denotes the overall
424	association between a climatic variable and allele frequency, and we used this to represent the
425	effect size of these SNPs. Second, when traversing along the environmental gradient, a sudden
426	increase of cumulative importance at some environmental value (for example, 20°C) denotes
427	populations on either side of this environmental cutoff have very different allele frequency
428	compositions. In other words, this represents a threshold factor for local adaptation.
429	One can use this cumulative importance curve to estimate the effect of future
430	environmental change on local populations. In the example above, a population's local
431	temperature increased from 19°C to 21°C due to climate change would require larger allele
432	frequency shift than another population whose local temperature changed from 17°C to 19°C.
433	The "genetic offset"[10] could then be calculated as the Euclidean distance between cumulative
434	importance corresponding to the contemporary environmental value and that corresponding to
435	the future environmental value, considering all bioclimatic variables together. Genetic offset
436	can then be considered to be the magnitude of genetic change needed for a population to be
437	still adaptive in the face of climate change.
438	
439	**Regression slope.**
440	The regression slope is not given by gradient forest, since it only reports the r^2 importance
441	estimate. Thus, we introduced the simple linear regression $y=\alpha+\beta x$ to measure the regression
442	slope. We took y as the allele frequency, x as the standardized bioclimatic variable, and β
443	(slope) as the measurement of the amount of allele frequency changes along environmental
444	gradients. By fitting simple linear regression with the general least-square approach, β can

445   then be expanded to $r_{xy} \frac{s_y}{s_x}$, where r_{xy} is the correlation coefficient (the square root of gradient

446   forest measured "importance") between x (environment value) and y (allele frequency), and

447   s_x and s_y are the standard deviation of x and y.

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 Figures:

Fig. 1 | Sample distribution and niche modeling of *Musa itinerans*. a, Sampling sites are
distributed along the latitudinal and altitudinal gradient. Lowland populations are represented
as white circles; transect populations are represented as white triangles and squares. **b**,
Suitability is derived from MaxEnt niche modeling. **c**, Least-cost-corridor landscape is
constructed from pairwise least-cost paths among 20 populations from the resistance surface
(the reciprocal of niche suitability).

Fig. 2 | The environment-dependent enrichment of new alleles and the distinct effect sizes of standing variation (SV) and new mutations (NM) in contrasting climatic factors. BIO1 reflects annual mean temperature, and BIO17 indicates precipitation of driest quarter. **a**, Mean frequency of new alleles among NM SNPs compared between populations that have local environments within or outside of the ancestral climatic range. New alleles are enriched in novel environments (**$P < 0.01$, ***$P < 0.001$, t-test). Values on the horizontal axis denote the number of Taiwanese populations within the ancestral or novel environmental range. **b**, Gradient forest cumulative r^2 importance is shown along environmental gradients. **c**, Effect sizes as estimated from Bayes factor, gradient forest r^2 importance, and gradient forest slope all show that SV exhibit higher effect sizes in BIO1 but the reverse in BIO17 (**$P < 0.01$, ***$P < 0.001$, Wilcoxon rank-sum test for Bayes factor and t-test for r^2 importance and slope).

618 **Fig. 3 | Annual mean temperature (BIO1) and precipitation of driest quarter (BIO17) for**
619 **the present and last glacial maximum. a**, **b**, Present environments. **c**, **d**, Last-glacial-
620 maximum environments, showing the environments on the extent of land. Maps on all panels
621 have the same range in latitude and longitude.
622

Fig. 4 | The fate of *Musa itinerans* under two extreme future conditions. a-d, Niche-expansion extreme (CCSM4-2070-RCP8.5) reports the species-average niche suitability from MaxEnt modeling (**a**), the genetic mismatch of locally adaptive populations to future environments based on the genetic offset from gradient forest (**b**), genetic offset difference calculated as the genetic offset experienced by new mutations minus that from standing variation (**c**), and extinction risk estimated as genetic offset divided by niche suitability (shown as 32 bins with equal grid counts; **d**). **e-h,** Niche-contraction extreme (MIROC-2070-RCP4.5) reports the suitability (**e**), genetic offset (**f**), genetic offset difference (**g**), and extinction risk (**h**). Grids with present niche suitability < 0.2 are excluded and colored in gray in **b-d** and **f-h**.

Data Availability:

Population pooled sequencing reads are available under NCBI BioProject PRJNA575344.

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Author Contributions:

C-RL designed the research. Both authors performed experiments and analyses and wrote the
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Competing Interest Declaration:

None

Supplementary Fig. 1 | Sampling site and population structure of Taiwanese *Musa itinerans* under STRUCTURE *K* = 2. a, Solid circles represent collection locations corresponding to their coordinates and elevation: Gray circles indicate populations of low elevation (LE); blue circles indicate populations of Xitou transect (XT); red circles indicate populations of Taipingshan transect (TPS). **b**, Individual ancestry is plotted on the right side, while population ancestry is plotted on map with a pie chart. Map template is provided by *Cheng-Tao Lin.

*Cheng-Tao Lin (2018) QGIS template for displaying species distribution by horizontal and vertical view in Taiwan. URL: https://github.com/mutolisp/distrmap_tw.qgis. DOI: 10.5281/zenodo.1493690

Supplementary Fig. 2 | Allele frequency distribution. a, Minor allele frequency distribution
of adaptive and non-adaptive SNPs. **b**, Frequency of new alleles in Taiwanese populations
within the ancestral or novel environmental range. Statistical significance from Student's *t*-test
between the novel and ancestral environmental range for each bioclimatic variable is shown.
Values on the horizontal axis denote the number of Taiwanese populations within the ancestral
or novel environmental range.

Supplementary Fig. 3 | Profiling of effect sizes. a, Bayenv Bayes factor distribution for adaptive new mutations and standing variation (*$P < 0.05$, **$P < 0.01$, ***$P < 0.001$, Wilcoxon rank-sum test). **b**, Cumulative r^2 importance from gradient forest along environmental gradients. **c**, Gradient forest r^2 importance distribution (***$P < 0.001$, Welch's t-test). **d**, Example figure showing the relationship between r^2 importance and slope. r^2 indicates the extent that allele frequency fits a linear model, while slope indicates the amount of allele frequency changes along the linear relationship. The graphs indicate one should also investigate regression slopes in addition to the gradient forest r^2. Values on the horizontal axis show the range of standardized environmental variables. **e**, The distribution of regression slopes when one regresses adaptive SNP allele frequency onto environmental gradients (**$P < 0.01$, ***$P < 0.001$, Welch's t-test).

Supplementary Fig. 4 | Evolutionary analyses on Taiwanese and Chinese *Musa itinerans.*

a, Minor allele frequency distribution of adaptive SNPs in Taiwan. **b**, Phylogeny of Taiwanese

and Chinese *Musa itinerans*. The gray-colored indicates Taiwanese lowland populations (LE);

the blue-colored indicates Taiwanese populations of Xitou transect (XT); the red-colored

indicates Taiwanese populations of Taipingshan transect (TPS); the black-colored indicates

Chinese accessions and the outgroup *Musa acuminata*. **c**, Table showing the expected

heterozygosity among Taiwanese and Chinese populations.

Supplementary Fig. 5 | Multi-analyses addressing ascertainment of standing variation and new mutation. a, The distribution of gradient forest r^2 importance across 100 re-sampling trials. In each re-sampling trials, a random set of 50% new mutations were assigned as standing variation, and the mean r^2 importance was reported for each trial. All comparisons show strong statistical significance from Student's t-test between new mutations and introduced standing variation (***$P < 0.001$). **b-d**, Profiling of effect sizes of adaptive SNPs with top 50% minor allele frequency. Distribution of Bayes factor (***$P < 0.001$, Wilcoxon rank-sum test; **b**), gradient forest r^2 importance (*$P < 0.05$, ***$P < 0.001$, Welch's t-test; **c**), and regression slopes of adaptive SNP allele frequency onto environmental gradients (***$P < 0.001$, Welch's t-test; **d**) is analyzed between new mutations and standing variation for each bioclimatic variable.

Supplementary Table 1 | Population coordinates and bioclimatic information

Site	Latitude	Longitude	BIO1	BIO2	BIO3	BIO7	BIO12	BIO15	BIO16	BIO17	BIO19
TPS300	24.60488	121.5359	208	63	35	180	2,970	46	1,223	364	364
TPS400	24.58135	121.5124	194	65	36	177	2,769	50	1,145	301	301
TPS500	24.56765	121.4992	199	65	36	179	2,763	50	1,142	299	299
TPS600	24.55172	121.5090	193	65	36	176	2,709	51	1,124	289	289
TPS700	24.54638	121.5125	185	65	37	174	2,699	49	1,107	302	302
TPS900	24.54021	121.5126	178	66	38	173	2,757	46	1,110	336	336
XT400	23.74292	120.7386	217	77	42	180	2,468	86	1,395	87	127
XT500	23.72780	120.7641	206	77	44	173	2,600	84	1,420	97	141
XT700	23.71137	120.7782	199	77	45	169	2,412	83	1,311	94	138
XT900	23.69294	120.7857	189	77	47	162	2,357	83	1,287	97	145
XT1200	23.67333	120.7854	176	77	48	159	2,478	82	1,346	97	148
XT1500	23.66936	120.7710	160	78	48	160	2,637	81	1,402	98	149
C35H	24.69907	121.1528	202	65	34	188	2,500	55	1,073	192	289
WFL	25.06752	121.7998	203	58	31	185	3,440	29	1,229	606	749
THNL	24.08273	120.8021	212	74	40	184	2,188	83	1,220	78	138
PTWT	22.58684	120.6489	227	78	49	158	2,912	98	1,784	79	87
P199H	22.24302	120.8465	223	73	50	144	3,114	92	1,856	178	178
MLLYT	24.33526	120.7864	211	71	37	190	2,056	76	1,084	79	148
HDPG	24.88347	121.5019	197	62	33	185	3,201	37	1,201	475	475
TTL	23.45230	120.6134	215	79	45	174	3,515	98	2,154	81	123
NAJY	24.43352	121.7602	216	62	35	174	2,860	61	1,328	271	296
HLCN	23.90623	121.4928	213	68	40	168	2,127	46	849	251	255
NXIR	23.40887	121.4503	227	71	43	163	2,147	57	938	211	211
DFR	23.10356	121.2724	206	75	48	155	2,017	68	1,026	166	166

708 **Supplementary Table 2 | SSR primer information**

Locus	Primer (5' to 3')	Chromosome*	Position*
CIR436	ATAAGCTCATATGGGTACAGTCACA CTGCAGCAACCAAATTTATTTCT	1	4,917,979 - 4,918,003 4,918,057 - 4,918,080
CIR276	CTCCTCCATAGCCTGACTGC TGACCCACGAGAAAAGAAGC	1	13,588,759 - 13,588,778 13,588,847 - 13,588,866
CIR1113	ACTCTCGCCCATCTTCATCC ACTTATTCCCCCGCACTCAA	1	28,009,948 - 28,009,967 28,010,173 - 28,010,192
Ma-3-132	TCCCTCTTCAACCAAAGCAC AACGCGAATGTGTGTTTTCA	2	20,365,901 - 20,365,920 20,366,041 - 20,366,060
CIR646	AACACCGTACAGGGAGTCAC GATACATAAGGCAGTCACATTG	2	23,940,966 - 23,940,985 23,941,276 - 23,941,297
Ma-1-17	AGGCGGGGAATCGGTAGA GGCGGGAGACAGATGGAGT	2	24,383,391 - 24,383,408 24,383,488 - 24,383,506
CIR332a	ATGACCTGTCGAACATCCTTT TCCCAACCCCTGCAACCACT	3	8,978,575 - 8,978,595 8,978,831 - 8,978,848
Ma-3-48	CCCGTCCCATTTCTCA TTCGTTGTTCATGGAATCA	5	32,672,883 - 32,672,898 32,673,018 - 32,673,036
CIR631a	ATTAGATCACCGAAGAACTC ATCTTTTCTTATCCTTCTAACG	6	33,942,751 - 33,942,770 33,943,017 - 33,943,038
CIR16a	TCATCTCACAATGCTTTCATAGTT TGGTTGAGTAGATCTTCTTGTGT	8	1,234,609 - 1,234,632 1,234,700 - 1,234,722
Ma-3-103	TCGCCTCTCTTTAGCTCTG TGTTGGAGGATCTGAGATTG	8	40,127,782 - 40,127,800 40,127,910 - 40,127,929
CIR348b	ACAGAATCGCTAACCCTAATCCTCA CCCTTTGCGTGCCCCTAA	10	27,228,162 - 27,228,186 27,228,325 - 27,228,342
Ma-3-139	ACTGCTGCTCTCCACCTCAAC GTCCCCCAAGAACCATATGATT	10	30,237,591 - 30,237,611 30,237,717 - 30,237,735
CIR550a	ACCGCACCTCCACCTCCTG TGCTGCCTTCATCGCTACTA	10	31,914,219 - 31,914,237 31,914,459 - 31,914,478

710 Primer sequences were searched against the *Musa acuminata* DH-Pahang version 2[33] on

711 Banana Genome Hub (https://banana-genome-hub.southgreen.fr/).

712 *The chromosome and position indicate locations on *Musa acuminata* where primer sequences

713 were found by blast.

715   Supplementary Table 3 | Number of new-mutation (NM) or standing-variation (SV) SNPs

716   with (adaptive) or without (non-adaptive) significant associations with bioclimatic

717   variables

Variable	Non-adaptive NM	Non-adaptive SV	Adaptive NM	Adaptive SV	P	Odds ratio
BIO1	293,348	954,367	816	7,113	8.92e-169	2.68
BIO2	293,274	953,198	890	8,282	8.43e-213	2.86
BIO3	293,551	953,901	613	7,579	6.47e-256	3.80
BIO7	293,644	953,981	520	7,499	1.32e-282	4.44
BIO12	293,452	954,272	712	7,208	3.12e-203	3.11
BIO15	293,168	951,447	996	10,033	2.33e-281	3.10
BIO16	293,826	957,605	338	3,875	1.95e-123	3.52
BIO17	293,461	954,969	703	6,511	1.61e-166	2.85
BIO19	293,516	955,326	648	5,142	1.06e-107	2.44

719   Statistical significance from χ^2 test is shown for each bioclimatic variable. Odds ratio is

720   calculated as (adaptive SV / adaptive NM) / (non-adaptive SV / non-adaptive NM).

722 **Supplementary Table 4 | Number of new-mutation (NM) or standing-variation (SV) SNPs**
723 **with (adaptive) or without (non-adaptive) significant associations with bioclimatic**
724 **variables (controlled for minor allele frequency)**

Variable	Non-adaptive NM	Non-adaptive SV	Adaptive NM	Adaptive SV	P	Odds ratio
BIO1	44,920	260,113	238	2,573	1.27e-20	1.87
BIO2	23,475	148,796	124	1,073	1.18e-03	1.37
BIO3	39,432	234,613	182	2,689	2.21e-34	2.48
BIO7	41,020	242,313	182	2,794	4.35e-38	2.60
BIO12	37,266	224,455	220	2,514	3.36e-20	1.90
BIO15	32,906	201,729	290	2,647	1.36e-10	1.49
BIO16	33,685	206,816	113	1,282	2.88e-10	1.85
BIO17	12,258	72,175	64	614	2.26e-04	1.60
BIO19	11,237	66,547	41	362	1.81e-02	1.50

725
726 The number of four sets of SNPs whose minor allele frequency (MAF) ranges from the first
727 quantile of adaptive MAF to the third quantile of non-adaptive MAF separately for each
728 bioclimatic variable is shown. Statistical significance from χ^2 test is shown for each
729 bioclimatic variable. Odds ratio is calculated as (adaptive SV / adaptive NM) / (non-adaptive
730 SV / non-adaptive NM).
731

 Supplementary Table 5 | Number of currently adaptive new-mutation (NM) or standing-
 variation (SV) SNPs that remain (retention) or lose (disruption) significant associations
 with environments under all future climate-change scenarios

Variable	Disruption NM	Disruption SV	Retention NM	Retention SV	P	Odds ratio
BIO1	92	635	127	1,176	5.11e-02	1.30
BIO2	34	374	348	2,010	6.85e-04	0.53
BIO3	21	269	119	2,277	1.32e-01	1.50
BIO7	29	251	128	2,393	4.53e-04	2.20
BIO12	65	481	136	1,509	1.37e-02	1.50
BIO15	37	365	437	2,982	4.79e-02	0.69
BIO16	26	289	25	237	6.93e-01	0.85
BIO17	37	354	159	1,182	2.21e-01	0.78
BIO19	28	235	45	480	4.14e-01	1.30

 Statistical significance from χ^2 test is shown for each bioclimatic variable. Odds ratio is
 calculated as (retention SV / retention NM) / (disruption SV / disruption NM).