# A single mechanism for global and selective response inhibition under the influence of motor preparation

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Abbreviated title: Response inhibition and motor preparation

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Conflict of interest: The authors declare no competing financial interests.

**Acknowledgments:** This work was supported by the University of Oslo and the Peder Sather Center, an International Research and Educational Collaboration between UC Berkeley and Norway. RBI was supported by a grant from the National Institutes of Health, NS092079. IG was supported by a grant from the National Institutes of Health, TR002370. We thank Claudia Tischler and Vincent Ngo for their assistance during data collection, and Assaf Breska and Nicole Swann for their helpful comments.

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# Abstract

In our everyday behavior, we frequently cancel one movement while continuing others. Two competing models have been suggested for the cancellation of specific actions: 1) An abrupt global inhibition followed by the restart of the other previously initiated actions, or 2) the parallel operation of distinct global and selective inhibitory mechanisms. To evaluate these models, we examined behavioral and physiological markers of proactive control, motor preparation, and response inhibition using a combination of task performance measures, electromyography, electroencephalography, and motor evoked potentials elicited by transcranial magnetic stimulation. Healthy participants performed two versions of a stop signal task with cues incorporating proactive control: A unimanual task involving the initiation and inhibition of a single movement, and a bimanual task involving the selective stopping of one of two prepared responses. Stopping latencies, motor evoked potentials, and frontal beta power (13-20 Hz) did not differ between the uni- and bimanual tasks. However, evidence for selective proactive control before stopping was manifest in the bimanual condition by changes of mu oscillations (9-14 Hz) over the motor cortex. Moreover, mu oscillations before stop signal presentation influenced subsequent stopping success. Altogether, our results favor a single inhibitory mechanism with the net behavioral output depending on the levels of action-specific motor preparation.

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# Significance statement

Response inhibition is a core function of cognitive flexibility and movement control. Previous research has suggested separate mechanisms for selective and global inhibition, yet the evidence is controversial. Additional research has examined the influence of preparation for action stopping, or what is called proactive control, on stopping performance, yet the neural mechanisms underlying this interaction are unknown. We combined transcranial magnetic stimulation, electroencephalography, electromyography and behavioral measures to compare selective and global inhibition models and to investigate markers of proactive control. The results favor a single inhibitory mechanism over separate selective and global mechanisms, but indicate a vital role for preceding motor activity in determining whether and which actions will be stopped.

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# Introduction

Flexible movement control requires the ability to inhibit a specific component of multiple actions. One form of inhibitory control can be studied with the stop signal task, where the primary instruction is to respond as quickly as possible to a go signal. On a fraction of trials, a stop signal is presented shortly after the go signal, and participants are instructed to inhibit their response. In a bimanual-selective version of this task, the go signal requires simultaneous responses with both hands and cancellation of only one of the two responses.

Several markers of response inhibition have been identified. A behavioral stopping latency index is captured by the stop signal reaction time (SSRT), estimated based on the independent horse race model (Logan and Cowan, 1984; Band et al., 2003). A peripheral inhibition measure is the latency of partial response electromyography (prEMG; Raud and Huster, 2017), detectable when an initiated response is interrupted. Motor evoked potentials (MEPs), elicited with single-pulse transcranial magnetic stimulation (TMS) over the motor cortex, provide a physiological index of corticospinal excitability during response preparation, execution, and inhibition (Bestmann and Krakauer, 2015). Electroencephalography (EEG) measures provide a global measure of cortical activity, with the stop signal evoking power increase in the beta band (13-20 Hz) over frontal electrodes (Picazio et al., 2014; Wagner et al., 2018).

Two models have been offered to explain selective stopping. One model proposes that inhibition globally influences the motor system via the cortical-subthalamic nucleus, hyperdirect pathway (Wessel and Aron, 2017). Reduced MEP amplitudes in task-irrelevant muscles during a unimanual stop task support such global inhibition (Badry et al., 2009; Cai et al., 2012; Greenhouse et al., 2012; Majid et al., 2012; Wessel et al., 2013). In the bimanual task, selective stopping can be achieved by a global inhibitory signal to interrupt all movements, followed by the re-activation of the required response hand (Coxon et al., 2007; Macdonald et al., 2014; Cowie et al., 2016). This inflates responding hand

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reaction times (RT) and causes a transient change in its EMG profile (Aron and Verbruggen, 2008; Raud and Huster, 2017; Figure 1A).

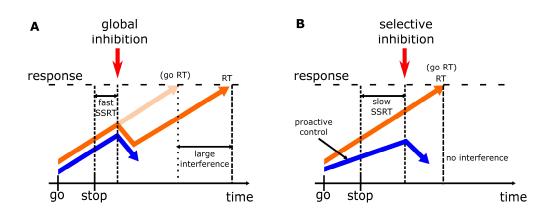
The second model posits distinct global and selective mechanisms (Figure 1A-B; Aron, 2011). Here, slow selective inhibition, implemented via the indirect cortico-striatal pathway, is activated by response-specific cues, prolonging the SSRTs, but reducing behavioral interference (Aron and Verbruggen, 2008; Claffey et al., 2010). This dual-mechanism model has been challenged by studies that found null or opposite effects regarding the SSRT differences between cued and uncued trials (Smittenaar et al., 2013; Lavallee et al., 2014; Smittenaar et al., 2015; Raud and Huster, 2017).

Proactive control can be engaged with response-specific cues, indicating that a subsequent stop signal, if presented, would require stopping only one component of a multi-effector response. These cues result in anticipatory activation of the stopping network (Chikazoe et al., 2009; Swann et al., 2012) or modulation of attentional, sensory, and motor mechanisms (Elchlepp et al., 2016; Langford et al., 2016). In the motor system, down-stream effects of proactive control are evident in the hand-specific modulation of corticomotor excitability (Claffey et al., 2010; Jahfari et al., 2010; Cai et al., 2011; Greenhouse et al., 2012), and reduction in motor mu (9-14 Hz) and beta (15-25 Hz) desynchronization (i.e. decreased power) that precede motor actions (Liebrand et al., 2017, 2018). Reduced mu and beta desynchronization have also been associated with more successful stopping (Mazaheri et al., 2009; Krämer et al., 2011).

We present a detailed multi-modal investigation of inhibitory control using the stop signal task, varying the demands on stopping in terms of response specificity and proactive control. Motor preparation and inhibition indices were extracted from behavioral, EMG, EEG, and MEP data, acquired simultaneously during task performance. We addressed three questions: 1) Do inhibition indices in the uni- and bimanual stop signal task dissociate between the two models of inhibition? 2) Does motor preparation influence subsequent inhibition success? 3) Is motor preparation affected by proactive control?

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**Figure 1.** Models for selective stopping in the bimanual-selective stop signal task. The single inhibition model posits that stopping is achieved by the global inhibition mechanism (A), while the dual-mechanism model posits separate global and selective inhibition mechanisms (both A and B). The primary task is to produce go responses with both hands. In case of a stop trial, only one hand's response needs to be stopped. Orange lines represent runners for the response hand and the blue lines represent runners for the stop hand. **A** represents global inhibition affects both runners, resulting in fast stop signal reaction times (SSRT) but delayed responding hand reaction time (RT) compared to a standard go response (represented by the light orange arrow), which leads to a large interference effect. **B** represents selective inhibition that is facilitated by proactive control in preparation for the stop signal. This leads to slow SSRTs, but no interference effect, as the responding hand is not affected by inhibition. Here, the effect of proactive control is presented as slower accumulation rate of the stop runner (blue arrow); however, it could also be depicted as a lower starting point for the stop runner or as hand-specific response threshold modulation.

# Materials and methods

## Sample

All participants provided informed consent under a protocol approved by the internal review board of the University of California, Berkeley. Participants were right-handed, had no psychiatric or neurological disorders, and no contraindications to TMS or EEG. The TMS coil was positioned over the EEG cap and electrodes which added an extra 2 cm to the distance between the coil and the skull and necessitated higher TMS intensities to produce MEPs. As such, we restricted testing to individuals with TMS resting motor thresholds (RMTs) below 50% of the maximum stimulator output when not wearing the EEG cap, since the intensity required with the cap in place would have exceeded the maximum stimulator output. Altogether, 33 participants were screened for their stimulation threshold. Ten were excluded due to high thresholds and two withdrew during the thresholding

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procedure. Of the twenty-one who continued with the experiment, three participants did not complete all sessions, two were excluded from the analysis due to stop accuracies below 30% in one of the sessions, and one participant's EEG was not recorded due to technical problems. Thus, the final sample included 15 participants for behavioral/EMG/MEP analyses (age 18-29, mean age 21, 7 females, 8 males) and 14 for EEG analysis.

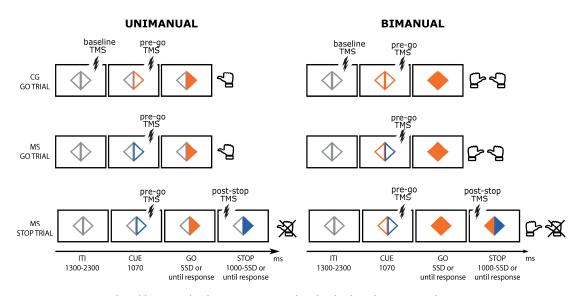


Figure 2. Unimanual and bimanual-selective stop signal task. The baseline TMS pulses were given 200 ms prior to the cue, pre-go TMS pulses 100 ms prior to the go signal and post-stop TMS pulses 150 ms after the stop signal. TMS was applied on approximately 27% of trials. All trial types were divided into go/stop and left/right trials (only go/stop right trials shown on figure). Responses were given as thumb presses. CG – certain go; MS – maybe stop.

#### Task and procedure

The participants performed two versions of the stop signal task – unimanual and bimanual – with response-specific cues (Figure 2). The tasks were performed in separate experimental sessions and the order of the tasks was counterbalanced across participants. Participants sat at a viewing distance of about 50 cm in front of the computer screen (refresh rate 80 Hz), hands laid palms down on pillows placed on a table in front of them so that the thumbs (responding effectors) rested on a keyboard used to collect responses. The experimental protocol was implemented with Psychtoolbox 3.0.

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Two hollow triangles were present on the screen at all times and served as the fixation point during the inter-trial intervals. In the unimanual task, the change of the outline color (from gray) of a single triangle signaled the participant to prepare a response with either the left or right thumb, depending on which triangle changed color. The specific color of the outline of the triangle cued the participant whether the given trial was a certain go trial (CG) or whether it may be a stop trial (MS). The triangle then filled with color (the go signal) to signal participants to execute their prepared response. On a fraction of trials, the triangle changed color again (the stop signal) signaling participants to stop their prepared or initiated response.

The bimanual task was similar, except that both triangle outlines changed color after the inter-trial interval, cueing participants to prepare thumb responses with both hands. In the CG condition, both triangle outlines were of the same color, whereas in the MS condition the triangle outlines were of different colors with one color indicating which hand's response may have to be stopped. On go trials, both triangles filled with the same color, instructing the participants to produce a simultaneous response with the left and right hands. On stop trials, only the previously cued triangle changed color again, instructing the participants to stop the cued hand while continuing the response with the other hand. The colors for go and stop stimuli (and corresponding CG and MS cues) were blue and orange shown on a gray background, with the color mapping counterbalanced between participants.

Each trial started with the presentation of the cue for 1070 ms, indicating the response hand(s) and possibility of stopping. The cue was followed by the go signal. On go trials, the go signal remained visible until 100 ms after a response was registered or until the end of a response window of 1000 ms. On stop trials, the relevant triangle changed to the stop signal color after the stop signal delay (SSD). The SSD was adjusted according to a staircase procedure separately for stop left and stop right hand trials. The SSD increased in steps of 47 ms if stopping was successful on the previous trial, and otherwise decreased by 47 ms if stopping failed. The stop signal was presented until 100 ms after a

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response (unsuccessful stop) or until the end of the response window (successful stop). The inter-trial interval varied between 1300-2300 ms.

In the unimanual task, the trials were divided equally into left and right hand trials. In the bimanual task, all go trials were the same (bimanual go), but stop trials were equally divided into stop left and stop right hand trials. All trials were further divided into those with and without TMS. Each task consisted of 200 CG-go trials, 680 MS-go trials, and 340 stop trials (see Table 1 for the number of trials included in each analysis modality). The stop signal was presented on 33% of the MS trials and 26% of all trials. To prevent premature responding during the cue period, 88 catch trials (20 CG and 68 MS, corresponding to 10% of go trials) were added, in which the go signal was not presented after the cue. The total number of trials amounted to 1308 per task, divided into 10 blocks (each lasting about 8 minutes) with pauses of self-determined durations between the blocks. The order of trials in each block was pseudorandomized.

Participants received instantaneous feedback if they did not respond within the response window of 1000 ms in go trials or if their response was longer than 800 ms to prevent strategic slowing due to stop signal expectancy. A red exclamation mark appeared on the screen in the bimanual task if the difference between left and right hand responses exceeded 70 ms. Finally, block-wise feedback was given, prompting the participants to be faster if the reaction times of the previous block on average were slower than 600 ms, or to be more accurate if stopping accuracy fell under 45%.

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				GO TRIALS				
	UNIMANUAL				BIMANUAL			
	cert	certain go maybe stop		certain go	maybe stop			
Total	200 68		580	200	680			
Witho	out TMS							
BEH	97 (9)		611 (21)		91 (6)	592 (21)		
EMG	86	86 (13) 580		0 (35)	85 (9)	563 (35)		
EEG	75 (12)		472 (58)		70 (10)	455 (50)		
With 1	TMS (pre-go	o)*						
	selected	unselected	selected	unselected	selected-go	maybe stop	selected-go	
MEP	29 (5)	29 (5)	31 (3)	31 (3)	60 (8)	31 (3)	31 (4)	
				STOP TRIAL	S			
	UNIMANUAL				BIMANUAL			
	suce	cessful	unsuccessful success		successful	cessful unsuccessful		

		UNIMANUAL			BIMANUAL			
	suc	cessful	unsi	uccessful	successful	ι	unsucce	ssful
Total		340			340			
Witho	ut TMS							
BEH	ç	91 (9)		89 (10)		96 (7)		
EMG	8	87 (8)		n.a.	86 (9)	n.a.		
EEG	6	68 (11)		65 (14)		66 (9)		
With TMS (post-stop)								
	stop	unselected	stop	unselected	stop	response	stop	response
MEP	19 (6)	32 (7)	n.a.	31 (8)	23 (7)	19 (6)	n.a.	n.a.

Table 1. Number of trials (standard deviations in brackets) per condition. Total refers to the total number of trials presented across all modalities and the rest refers to the mean number of trials per condition, participant, and analysis modality (standard deviations in the brackets) that were included in the analysis after data cleaning. Baseline MEP trials and catch trials are omitted from the table. The MEPs are further divided according to whether the right hand was selected or unselected for the response (unimanual task) or selected for stopping or responding (bimanual task). \*Maybe stop pre-go TMS pulses were given both in go and stop trials with a ratio of 30/10 to prevent the prediction of the upcoming stop signal from application of the pre-go pulse. N.a. – not analyzed.

# **TMS procedure**

Single TMS pulses were delivered with a 70 mm figure-of-eight coil driven by a Magstim 200-2 magnetic stimulator. The coil was positioned tangentially over the left motor cortex in the posterior-anterior direction and the position was optimized to elicit MEPs in the right abductor pollicis brevis (APB) muscle. The stimulation intensity was set to 115% of the RMT, defined as the intensity where 5 out of 10 consecutive pulses elicited a MEP with an amplitude of at least 50  $\mu$ V. This threshold was determined after EEG preparation to account for the increased distance between the coil and the

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scalp due to the electrodes. The coil rested on a plastic spacer placed between the coil and the EEGcap to reduce mechanical EEG artifacts due to the pressure of the coil (Ruddy et al., 2018).

TMS pulses during the task were administered pseudorandomly between all trials with an interstimulus-interval between 4 and 54 seconds. The total number of pulses per session was 360. However, some pulses were omitted (~12% per session) when a pulse was scheduled by pseudorandomization with an inter-stimulation interval <4 sec or when the participant terminated the trial before the scheduled pulse time (e.g. premature responses or when an unsuccessful stop response was executed before the scheduled post-stop pulse time). These trials were re-coded offline as no-TMS trials. To establish a baseline measure of corticospinal excitability, on 40 CG-go trials TMS pulses were scheduled 200 ms prior to the cue (Figure 2). To measure excitability changes during response preparation, the pre-go pulses were scheduled 100 ms before go onset (corresponding to 970 ms after the cue) on 80 CG and 80 MS trials (further divided into 60 go and 20 stop trials). To measure changes in excitability in response to the stop signal, a post-stop TMS pulse was scheduled at 150 ms following the stop signal on 160 trials, as we expected inhibition to be present around this time in stop trials, based on previous TMS and EMG results (van den Wildenberg et al., 2010; Macdonald et al., 2014; Raud and Huster, 2017; Atsma et al., 2018). All TMS trials were divided equally into left and right hand trials (except for the bimanual CG-go trials).

# EEG and EMG data acquisition

EMG was recorded from the left and right hand abductor pollicis brevis with bipolar Ag/Cl surface electrodes (Delsys, Inc.) placed in parallel to the belly of the muscle. The ground electrode was placed over the ulna of the right elbow. The data were recorded with a sampling rate of 4000 Hz and bandpass filtered between 20-450 Hz on data acquisition. EEG was recorded with the Active2 system from 64 channels (BioSemi) with a sampling rate of 2048 Hz (low-pass filter 417 Hz). Additional electrodes were placed on the nose-tip, mastoids, next to the outer canthis and below the right eye.

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#### Data preprocessing and feature extraction

**Behavior**. The following variables were determined for trials without TMS pulses: go accuracy, go reaction time (RT), stop accuracy, stop signal delay (SSD), stop signal reaction time (SSRT), unsuccessful stop trial RT. Additional variables determined for bimanual task stop trials included the responding hand RT and behavioral interference, calculated as the difference between the responding hand RT and the MS go RT. To fully characterize the task performance, we also calculated the percentage of go trials with premature responses, erroneous responses (incorrect hand responses in the unimanual task and responses with only one hand in the bimanual task), omissions, and asynchronous responses (RT difference between left and right hand larger than 70 ms in the bimanual task). In correct go trials, very fast (<100 ms) go RTs were discarded from the RT calculation (0.1 %). In the bimanual task, the RT was obtained by averaging over both hands. The SSRT was calculated by the integration method with replacements for go omissions and including go errors (Verbruggen et al., 2019). The data from left and right hand trials were averaged for statistical analysis.

**MEPs.** MEP amplitudes were measured as the peak-to-peak amplitude in the EMG response in the 10-100 ms window following the TMS pulse. Trials were rejected when the average rectified EMG activity 200 ms prior to the pulse was over 10 μV. Additionally, individual trial data were visualized, and trials with EMG activity just before or at the time of the TMS pulse were discarded (26% on average across all conditions). Trials with a MEP amplitude greater than +/-2 standard deviations from the mean were discarded (5% on average). The MEPs were divided into conditions depending on the time-point of stimulation (baseline, pre-go, post-stop) and the hand selected for response in a given trial. As only the left hemisphere was stimulated, this resulted in the 'selected hand' (respond/stop right) and 'unselected hand' condition (respond/stop left) in the unimanual task. In the bimanual task, this resulted in CG go (all CG trials), MS go (respond-right/maybe-stop-left MS trials), and MS maybe-stop conditions (respond-left/maybe-stop-right MS trials). The pre-go and post-go MEPs were normalized by calculating the percentage change from the average baseline MEP. Differences in raw MEP

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amplitudes for each time point and response type were compared against baseline MEPs using separate paired t-tests with a Bonferroni-Holm correction across all conditions.

**EMG**. EMG was downsampled to 512 Hz using a polyphase anti-aliasing filter. Continuous data were epoched relative to the cue onset with an epoch length of -300 to 2600 ms. A moving average procedure with a window size of +/-5 datapoints was applied to the root mean square of the signal, and this was normalized by dividing each datapoint with the average baseline period of -300 to 0 ms relative to the cue onset. The epochs of all trials were then concatenated for each participant and hand, and the data was z-scored across all conditions. The data was then re-epoched into the conditions of interest, with an epoch length of -300 to 1800 ms for cue-locked activity, and -300 to 600 ms for stop-locked activity.

An automatic artifact rejection and EMG-burst identification algorithm was applied to the epochs, where an EMG burst in a given trial was determined whenever the z-scored RMS data exceeded a threshold of 1.2. This threshold was chosen by visual inspection of several datasets and was held constant for all datasets. Trials were rejected automatically if the average rectified baseline activity exceeded 10  $\mu$ V or if an EMG burst was detected prior to the go stimulus onset. All trials were further visualized and trials with excessive tonic muscle activity were discarded. Altogether, 3.28% trials per condition were discarded on average at this step.

The following dependent variables were extracted: 1) prEMG burst frequency: the percentage of successful stop trials with EMG activity above threshold relative to the total number of successful stop trials; 2) EMG onset latency: the latency of the first data-point exceeding the EMG threshold; 3) EMG peak latency: the time point at which the EMG activity reached its maximum. Onset latencies were calculated relative to the go-signal onset, while the peak latency of the prEMG was calculated relative to the stop-signal onset. Trials in which the amplitude or latency measure exceeded +/-2.5 standard deviations of each participant's mean were discarded from the analysis, as well as trials where prEMG peaked before the stop signal or more than 500 ms after the stop signal, as these were likely not

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related to stop signal processing (altogether 2.48% per condition on average). To capture potential interference in the responding hand EMG in the bimanual task at the single trial level, the Δpeak was calculated by subtracting the prEMG peak latency of the stopped hand from the EMG peak latency of the responding hand (see Figure 3A for schematics). All variables were averaged over left and right hand trials prior to statistical analysis.

**EEG**. EEG was analyzed with *eeglab* v14.1.2 (Delorme and Makeig, 2004). Trials with TMS pulses were discarded from the analysis, and the data -5 to 500 ms relative to each TMS pulse were replaced using a cubic interpolation to avoid possible filtering artefacts that could contaminate the trials directly before or after the pulse. The data were low-pass filtered using a hamming window filter with a 6db cut-off at 50 Hz (transition bandwidth 20 Hz), resampled to 512 Hz, and high-pass filtered with a 6db cut-off at 0.1 Hz (transition bandwidth 0.8 Hz). The data were epoched relative to the cue onset, and noisy trials (i.e. high muscle noise, trials with noise from the TMS coil) were rejected manually during visual inspection (21% on average). At this step, noisy channels were identified and temporarily discarded (1.25 electrodes or 1.95% on average). An independent component analysis was run on the remaining data and artifactual components reflecting eye-movements and muscle activity were identified and rejected semi-automatically (5.39 components or 8.57% per dataset on average) using the *eeglab* plugin SASICA (Chaumon et al., 2015). An interpolation procedure was used to replace data from previously identified noisy electrodes. Lastly, to reduce the volume conduction in the scalp EEG, a current source density interpolation was estimated by a spatial Laplacian transformation using the current density toolbox for Matlab (Kayser and Tenke, 2006).

For time-frequency analyses, the data were segmented into condition-specific epochs of -1200 to 3000 ms. Time-frequency decompositions for 1-40 Hz were calculated using Morlet wavelets with the number of cycles linearly increasing from 1 to 20. The cue-locked activity was normalized by dividing the power by the average of the frequency-specific baseline between -400 to -100 ms relative to cue-

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onset. The stop-locked activity was similarly normalized to the average of -400 to -100 ms before the stop signal.

Frontal beta (13-20 Hz) activity was extracted from 100 to 150 ms after the stop signal, averaged over the frequencies and electrodes of interest: left prefrontal (F5, F3, FC5, FC3), frontocentral (F1, Fz, F2, FC1, FCz, FC2) and right prefrontal (F4, F6, FC4, FC6). Electrodes and frequencies were determined based on previous studies showing frontal beta effects in the stop signal task (Liebrand et al., 2018; Wagner et al., 2018). The time-window was selected on the rationale that brain activity prior to the peripheral inhibition latency (i.e. prEMG latency at ~150 ms) would be the best indicator of inhibitory control.

Sensorimotor mu (9-14) and beta power (15-25 Hz) were extracted from the cue-locked activity in prego (-200 to 0 ms) and post-go (100 to 200 ms) time-windows on go trials, and post-go time-window (100 to 200 ms) on stop trials. The pre-go time-window was selected so that it would match the timing of the pre-go TMS pulses (-100 ms relative to go) to allow for comparisons between the EEG and MEP results. The post-go time window was selected so that it would correspond to the time after the initial sensory processing and before response execution (i.e. EMG onset). Frequencies and electrodes of interest were determined based on previous studies (Krämer et al., 2011; Liebrand et al., 2018), corresponding to left (C5, C3, CP5, CP3) and right (C4, C6, CP4 and CP6) sensorimotor regions. Data for EEG features were relabeled as contra- or ipsilateral, depending on the response hand, and averaged over left and right hand trials in the analyses.

#### Experimental design and statistical analysis

Group-level statistical analyses were performed using repeated measures ANOVAs (rmANOVA) and paired samples t-tests. The alpha level for all tests was set to 0.05. All significant main and interaction effects were followed up by paired t-tests with Bonferroni-Holm corrections for multiple comparisons.

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Given that there were normality violations for several dependent variables in many of the statistical tests, all p-values were determined by permutation tests with 10 000 permutations and marked as p<sub>prm</sub> in the results section. The permutation rmANOVAs were performed with the *permuco* package in the R Project for Statistical Computing (version 3.5.2), using the Kherad-Pajouh-Renaud method for the permutation of nuisance variables (Kherad-Pajouh and Renaud, 2015). The permutation t-tests were performed using the *RVAideMemoire* package in R. The bivariate correlations represent Pearson's coefficients and the corresponding permutation p-values were calculated using the package *wPerm*.

The statistical design is described in more detail in the following sections, which reference the three main research questions outlined in the introduction.

**1)** Do indices of inhibition in the uni- and bimanual stop signal task dissociate between the two models of inhibition? We evaluated the two models of inhibition by comparing the indices of inhibition (SSRT, prEMG latency, post-stop MEP, frontal beta power) between the two tasks. Further, the extent of inhibition was evaluated by the MEPs of the unselected hand in the unimanual task, and behavioral (RT costs on the responding hand) and EMG interference (Δpeak) on the responding hand in the bimanual task.

The global inhibition account predicts reduced unselected hand MEPs and a large interference effect in the bimanual task, with no differences in the inhibition indices (SSRT, prEMG latency, post-stop MEP, frontal beta power) between the tasks. The dual-mechanism account predicts slower inhibition latencies (SSRT and prEMG) in the bimanual task due to the activation of the slower indirect pathway (as opposed to the fast hyperdirect pathway in the unimanual task), accompanied by weaker suppression of MEP amplitudes of the stopped hand shortly after the stop signal – a pattern consistent with a mechanism for slow selective stopping.

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Differences between the tasks in SSRTs and prEMG peak latencies (relative to the stop signal) were assessed through paired samples t-tests. We additionally compared the prEMG onset latencies (relative to the go-signal) with a paired t-test to account for potential task differences in the timing of motor initiation. Given that significant differences were found in both prEMG peak and onset latencies, the prEMG peak latency analysis was repeated after residualizing for the prEMG onsets. That is, a linear regression model with prEMG onsets as a predictor was fit for each participant's prEMG peak latencies across all trials in both tasks, and the residuals for both tasks were compared statistically by paired t-test.

Given the potential relationship between the prEMG onsets and peak latencies, we tested this formally at the single-trial level. We further tested whether the stopping latency (prEMG latency) is associated with stopping extent ( $\Delta$ peaks) at the single trial level. These associations were tested using linear mixed effects regression models to control for the influence of individual participants. The regressions were fitted with the R package *nlme*. Given the individual differences in the general response speed, we considered the model with random intercepts as the baseline model and tested the improvement in model fit after including random slopes with chi-square tests. Both unstandardized and standardized coefficients are reported, marked as b and  $b_{std}$ , respectively.

The global inhibition model predicts reduced corticospinal excitability after the stop signal in an effector that is not selected for a stop-response. This was tested by comparing all MEPs (including the post-stop unselected hand) to the baseline MEPs. In contrast, the dual-mechanism model predicts differences in the post-stop MEPs between the two tasks: the corticomotor excitability in the stopped hand should be reduced sooner after the stop signal in the unimanual task compared to the bimanual task, as inhibition is expected to occur faster. The post-stop MEPs were tested with a one-way rmANOVA. Given that the TMS pulse was always applied over the left motor cortex, there were four levels in the rmANOVA: unimanual stop (i.e. MEPs measured from the right hand on right hand stop

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trials), unimanual unselected (i.e. MEPs measured from the right hand on left hand stop trials), bimanual stop (i.e. stop right, respond left) and bimanual response (i.e. stop left, respond right).

Elevated frontal beta power in scalp EEG may occur early enough to serve as a marker of response inhibition. We tested the beta power shortly after the stop signal (100-150 ms, averaged over both tasks and successful and unsuccessful stops) with a one-sample t-test against zero. In addition, if frontal beta is an EEG indicator of stopping, we would expect stronger beta power in successful, compared with unsuccessful stop trials. Regarding task differences, the dual-mechanism account may predict reduced beta power in the bimanual task following the stop signal due to the slower selective inhibition mechanism, compared to the fast global inhibition in the unimanual task. However, as there is no strict correspondence between EEG amplitude and latency measures and we are the first to compare frontal beta power effects were tested with a three-way rmANOVA, factors SUCCESS (unsuccessful, successful), TASK (unimanual, bimanual) and LOCATION (frontal contralateral, central, frontal ipsilateral). 'Contralateral' refers to the electrodes positioned contralateral to the stopped hand (and thus the electrodes placed ipsilateral to the hand still responding in stop trials of the bimanual task).

**2) Does motor preparation influence subsequent inhibition success?** Given the putative roles of sensorimotor mu and beta activity in action preparation (Neuper et al., 2006; Brinkman et al., 2016) and proactive control (Liebrand et al., 2017, 2018), we hypothesized that mu and beta activity before the stop signal would predict stopping success. We expected weaker desynchronization in these signals preceding successful stopping compared to unsuccessful stopping. In the bimanual task, we further expected a lateralization effect, with weaker desynchronization on successful stop trials in the hemisphere contralateral to the stopped hand.

The sensorimotor mu and beta power in stop trials were extracted from the time-window after the go signal, but prior to the stop signal. The presence of desynchronization was tested with one-sample t-

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tests against zero (separately for the tasks and corrected for multiple comparisons with Bonferroni-Holm method), where the power in the electrodes contralateral to the responding hand was averaged over successful and unsuccessful stops. The condition differences were then tested using rmANOVAs incorporating the success of stopping and location. This was done separately for the tasks, as the differences in the go instructions did not allow for a balanced statistical comparison regarding the hemispheric effects. Thus, in the unimanual task, we tested this with a two-factor rmANOVA with the factors SUCCESS (successful, unsuccessful) and LOCATION (contralateral, ipsilateral). In the bimanual task, each hand/hemisphere has a different simultaneous assignment on each MS trial (respond with one hand, stop the other hand). Therefore, the rmANOVA still used the factors SUCCESS (successful, unsuccessful) and LOCATION, but the levels for the LOCATION referred to the hemisphere contralateral to the response hand (contra-go) and contralateral to the stop hand (contra-stop).

**3)** Is motor preparation affected by proactive control? In the current study, proactive control is relevant in conditions in which the participants are informed that a forthcoming trial may require stopping. We tested the effect of proactive control on response preparation by comparing trials with maybe stop (MS) and certain go (CG) cues. Behaviorally, we expected the MS cue would lead to response slowing, with participants adopting a more conservative strategy in anticipation of a potential stop signal. The effects of the cue on go RT and EMG onsets in go trials were tested with rmANOVAs including the factors CUE (certain go (CG), maybe stop (MS)) and TASK (unimanual, bimanual).

We expected the RT slowing to be paralleled by changes in the corticomotor excitability during the cue period. Pre-go MEPs are typically reduced in both hands in the delayed response task, where a response certainly needs to be executed (here captured by the CG cue; Hasbroucq et al., 1999; Duque and Ivry, 2009; Labruna et al., 2014; Greenhouse et al., 2015; Duque et al., 2017). Larger reductions in pre-go MEPs were shown to correlate with faster go RTs (Hannah et al., 2018). The direction of this modulation in the MS condition, however, is less clear. The added uncertainty about the upcoming

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response in the MS condition might weaken this preparatory suppression, resulting in increased MEP amplitudes compared to the CG condition (Duque and Ivry, 2009). Alternatively, applying inhibition proactively may result in reduced excitability (and thus decreased MEP amplitudes) in the MS condition compared to the CG condition. Such reduction has been previously observed in the bimanual stop task, where the MEP amplitudes where smaller in the hand that was cued to stop, compared to the hand that was cued to respond (Cai et al., 2011). We tested the effects of the cue and response hand on pre-go MEPs with rmANOVAs computed separately for each task. In the unimanual task, this resulted in a two-way ANOVA with factors CUE (CG, MS) and HAND (selected, unselected). In the bimanual task this resulted in a single combined CUE-HAND factor with levels CG, MS go (i.e. respond right, maybe stop left), and MS maybe stop (i.e. maybe stop right, respond left). MEPs in the bimanual task were always from the selected hand since the right hand was involved in responding on every trial.

Lastly, we expected that the MS cue would produce less mu and beta desynchronization compared to the CG cue, given the added uncertainty (Liebrand et al., 2017, 2018). The presence of desynchronization was tested by one-sample t-tests against zero (separately for the tasks, as well as pre-go and post-go periods, and corrected for multiple comparisons with Bonferroni-Holm method), where the power in the electrodes contralateral to the responding hand was averaged over CG and MS cue conditions. The effects of the cue on sensorimotor mu and beta power were tested with rmANOVAs, separately for each task. Here, we additionally tested the difference between pre-go and post-go activity, to determine whether the effects were evident in the cue-delay period. This analysis was limited to go trials. In the unimanual task, the factors were CUE (CG, MS), LOCATION (contralateral, ipsilateral), and TIME (pre-go, post-go). In the bimanual task, the levels for a combined CUE-LOCATION factor were CG bilateral, MS contra-go, MS contra-stop. In the CG-bilateral, activity was averaged over left and right hemisphere electrodes given the bilateral response, while the MS contra-go refers to the electrodes contralateral to the responding hand, and MS contra-stop refers to the electrodes contralateral to the hand that may need to be stopped.

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# Results

The results are presented with respect to the three main research questions, starting with the differences between uni- and bimanual stopping, followed by the effect of motor activity on stopping success in both tasks, and ending with the effects of proactive control on motor preparation in both tasks.

# 1) Do inhibition indices in the uni- and bimanual stop signal task dissociate between the two models of inhibition?

We compared the inhibition indices (SSRTs, prEMG peak latencies, frontal beta power) between the uni- and bimanual stop signal task. To recapitulate, the global inhibition model predicts suppressed post-stop MEPs of the unselected hand in the unimanual task and prevalent interference (measured by the RT costs on the responding hand and the EMG Δpeak in stop trials) in the bimanual task, but does not predict any differences in the latencies or strength of inhibition. In contrast, the dual-mechanism model predicts minimal interference (RT costs and Δpeak close to zero in the bimanual task) and slower inhibition (SSRT, prEMG peak latency) in the bimanual task, accompanied by reduced post-stop MEP suppression. Frontal beta power may be reduced as well, due to the slower selective inhibition mechanism in the bimanual task.

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	Unimanual	Bimanual
Go trials		
CG accuracy (%)	95.38 (6.18)	95.58 (4.07)
omissions	1.31 (2.51)	0.86 (1.83)
errors*	0.25 (0.74)	0.35 (0.67)
premature	2.98 (4.37)	2.33 (2.37)
asynchronous (>70 ms)	-	0.90 (1.08)
MS accuracy (%)	97.29 (3.27)	94.67 (3.35)
omissions	2.11 (2.97)	0.95 (1.20)
errors*	0.25 (0.28)	1.25 (1.26)
premature	0.31 (0.31)	0.33 (0.33)
asynchronous (>70 ms)	-	2.79 (1.76)
CG RT (ms)	373 (84)	394 (92)
MS RT (ms)	527 (87)	536 (76)
CG asynchrony (ms)	-	14 (3)
MS asynchrony (ms)	-	15 (4)
Stop trials		
Stop accuracy (%)	50.61 (4.30)	50.77 (3.19)
Unsuccessful stop RT (ms)	447 (93)	478 (74)
SSD (ms)	298 (95)	299 (91)
SSRT (ms)	227 (24)	232 (24)
Response hand RT (ms)	-	652 (76)
Behavioral interference (ms)	-	116 (38)

Table 2. Means and standard deviations of the behavioral parameters in both tasks for trials without TMS. CG – certain go; MS – maybe stop; RT – reaction time; SSD – stop signal delay; SSRT – stop signal reaction time. \*In the unimanual task, errors refer to the trials where the response was executed using the wrong hand; in the bimanual task, errors refer to trials where a response was executed with only one hand.

No differences between the SSRTs of the uni- and bimanual tasks. Task performance was in accordance with expectations based on the horse race model, with stopping accuracies at 51% in both tasks (see Table 2 for all behavioral results) and faster unsuccessful stop RTs than go RTs in all individuals and at the group level (unimanual: t(14) = -16.87,  $p_{prm} < 0.001$ , d = -4.36; bimanual: t(14) = -15.96,  $p_{prm} < 0.001$ , d = -4.12). The SSRTs were 227 and 232 ms for the uni- and bimanual tasks, respectively (Figure 3B), and the difference was not significant (t(14) = -0.93,  $p_{prm} = 0.377$ , d = -0.24).

# Delayed prEMG peaks in the bimanual task, possibly resulting from dependencies between go and

stop processes. Successful stopping was defined as trials in which a key press was not detected.

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However, on approximately 30% of these trials (an average of 24 trials per task for each participant), EMG activity was detected in the stopped hand, or what we refer to as prEMG (Table 3). The onset of prEMG relative to the go signal provides an index of the speed of the go process on these trials, while the latency of the decline of this signal (i.e. the prEMG peak latency) relative to the stop signal provides an index of the speed of the inhibitory process (Figure 3A). The bivariate correlations between the SSRT and prEMG were r(13) = 0.31,  $p_{orm} = 0.252$  (unimanual) and r(13) = 0.35,  $p_{orm} = 0.215$  (bimanual).

The prEMG peak latency was later in the bimanual task compared to the unimanual task (170 vs 144 ms, t(14) = -3.99,  $p_{prm} = 0.003$ , d = -1.03, Figure 3A, 3B). However, prEMG onset was also significantly later in the bimanual (429 ms) than in the unimanual task (395 ms; t(14) = -2.20,  $p_{prm} = 0.048$ , d = -0.57). Thus, the results suggest both go and stop processes were delayed in the stop trials of the bimanual task.

The delayed go and stop processes in the bimanual task raise the possibility that there are dependencies between the two processes, and that these dependencies may differ between the tasks. This is important for the evaluation of the latency differences between the tasks, as dependencies would violate assumptions of the horse race model that are necessary for a reliable estimation of SSRTs. We explored this possibility at the single trial level on a post-hoc basis using mixed linear regression models, where prEMG peak latency was predicted by the prEMG onset latency and task type, with participants' intercepts and slopes as random effects. We fit models predicting prEMG peaks from onsets separately in both tasks (Figure 3C). In the unimanual task, the fixed effect indicated a positive relationship between the peaks and onsets (b = 0.24, 95% CI = 0.12, 0.36, t(339) = 3.93, p < 0.001, b<sub>std</sub> = 0.56) and considerable variability in the slopes across participants (sd = 0.21, 95% CI = 0.14, 0.32). There was a significant improvement in model fit when including the variability of slopes, compared to the random intercept model ( $\chi^2(1) = 87.76$ , p < 0.001), indicating that the relationship between peaks and onsets was relatively heterogeneous among the participants (Figure 3C, left panel). In the bimanual task, the fixed effects also indicated a positive relationship between peaks and onsets was relatively heterogeneous among the participants (Figure 3C, left panel). In the bimanual task, the fixed effects also indicated a positive relationship between peaks and

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onsets (b = 0.36, 95% CI = 0.28, 0.44, t(348) = 8.76, p < 0.001, b<sub>std</sub> = 0.69), but relatively smaller variability in the slopes (sd = 0.10, 95% CI = 0.03, 0.28). The improvement in model fit was not significant after including variable slopes ( $\chi^2(1) = 2.73$ , p = 0.255), suggesting that the association between the onsets and slopes were relatively similar across participants (Figure 3C, right panel).

Given the positive association between prEMG onsets and peaks, we next asked if the task differences in prEMG peak latencies could be explained by the differences in the onsets. We tested the differences between the peaks again, but this time after residualizing for the differences in onset times. The histograms for prEMG peaks before and after residualization (Figure 3D) reveal that the initial differences between the tasks were greatly diminished after residualization. The average residuals still indicated somewhat delayed inhibition in the bimanual task, but this difference was not significant at the chosen alpha level (t(14) = -1.99, p<sub>prm</sub> = 0.062, d = -0.51).

	Unimanual	Bimanual	
Overt EMG			
CG go onset (ms)	253 (90)	284 (94)	
MS go onset (ms)	402 (91)	424 (80)	
MS stop onset (response hand; ms)	-	511 (76)	
MS unsuccessful stop onset (ms)	337 (101)	368 (75)	
prEMG			
prEMG frequency (%)	29.65 (15.24)	31.50 (15.37)	
prEMG onset (relative to go; ms)	395 (91)	429 (93)	
prEMG peak (relative to stop; ms)	144 (27)	170 (22)	
Δpeak (ms)	-	150 (50)	

Table 3. Mean and standard deviation of overt EMG (where button presses were registered) and prEMG in successful stop trials (where no button press was registered). CG – certain go; MS – maybe stop.

**Prevalent and variable interference in the bimanual task.** The RTs of the responding hand on bimanual successful stop trials were delayed compared to go RTs (MS cue), giving rise to a mean behavioral interference effect of 116 ms (one-sample t-test against zero: t(14) = 11.93, p < 0.001). This interference was also evident in the EMG of the continuing response hand (grey lines on Figure 3A,

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right panel). There was a halt in the progression of the EMG activity at around 170 ms, followed by

Α Go-locked EMG Stop-locked EMG ∆peak 3 CG go stop (stop hand) unimanual stop hand MS go stop (response hand) bimanual stop hand EMG amplitude (z) unimanual bimanual bimanual response hand 2 EMG amplitude (z) onset onset peak 2 0 400 800 0 800 400 -400 -200 0 200 400 600 800 time (ms) time (ms) time (ms) unimanual В С 600 bimanual unimanual bimanual 500 500 prEMG peak (ms) 400 400 400 ms 300 300 200 200 200 100 100 0 0 prEMG prEMG SSRT 400 600 800 200 200 400 600 800 onset peak prEMG onset (ms) prEMG onset (ms) D Ε bimanual unimanual bimanual 80 200 400 Δpeak (ms) <sup>60</sup> 40 20 150 200 100 0 20 50 01 0 200 400 -200 -100 0 100 200 0 100 200 300 400 prEMG peak (ms) residualized prEMG peak prEMG peak (ms)

the recovery of the EMG activity leading to the key press.

Figure 3. **A**. EMG time courses time-locked to the go signal (left panels) and the stop signal (right panel). The red arrows reflect the parametrization of prEMG onset, peak and bimanual task  $\Delta$ peak. The solid vertical lines represent the time-locking event and the dashed vertical lines represent the average SSD. **B**. Average prEMG onsets, peaks and SSRTs across participants. The dots represent individual participants and the gray lines connect the same participants in the uni- and bimanual task. Asterisks indicate significant differences between tasks. **C**. Single trial correlations between prEMG onset and prEMG peak. The black lines indicate the fixed effects of prEMG onset on prEMG peak, and gray lines indicate the random effects (intercept and slope) for each participant. **D**. Histograms of prEMG peak latencies across all trials and participants before (left panel) and after (right panel) residualizing for the prEMG onset. **E**. Single trial correlation between prEMG peak and  $\Delta$ peak latency in the bimanual task. The horizontal dashed line marks zero  $\Delta$ peak; data points in the vicinity of the line correspond to trials with small or no interference. CG – certain go; MS – maybe stop.

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EMG allows for the investigation of the interference effect and its associations at a single trial level. To achieve this, we extracted the  $\Delta peak$  (the difference between responding hand EMG peak and stopped hand prEMG peak) from all trials where prEMG was detected. The Δpeak was on average 150 ms (one-sample t-test against zero: t(14) = 11.69, p < 0.001), but varied within trials between -141 and 460 ms, with a negative  $\Delta$ peak indicating that the EMG in the responding hand peaked before the prEMG in the stopped hand. At the group level, the  $\Delta peak$  correlated positively with the behavioral interference (r(13) = 0.85,  $p_{prm} < 0.001$ ). Further, the single trial  $\Delta peak$  correlated negatively with prEMG peak latency, as tested by the regression model with random intercepts (Figure 3E; b = -0.71, 95% CI = -0.86, -0.56, t(348) = -9.31, p < 0.001,  $b_{std} = -0.43$ ), indicating that later prEMG peaks occurred on trials with with reduced interference. This relationship remained significant after residualizing the peaks by prEMG onsets (b = -0.71, 95% CI = -0.91, -0.51, t(348) = -7.00, p < 0.001,  $b_{std} = -0.32$ ). There was no significant improvement in the model fits after introducing random slopes, indicating that this association was relatively stable across participants ( $\chi^2(1) = 0.63$ , p = 0.730). The large single-trial variability in the interference effect together with its associations with the stopping latency suggests that successful stopping in the bimanual task may include a mix of trials with different behavioral strategies, which also play a role in regulating inhibition speed.

**Post-stop MEPs are reduced in the unimanual unselected hand, but do not differ between the uniand bimanual tasks for the stopped hand.** Resting motor thresholds were 71% of maximum stimulator output in both tasks (standard deviation 5.91 in unimanual and 7.17 in bimanual task), and the raw baseline MEP amplitudes did not differ between the two tasks (unimanual: 0.93 (0.71) mV; bimanual: 0.84 (0.72) mV; t(14) = 0.74, p<sub>prm</sub> = 0.473, d = 0.19). The post-stop MEPs, probed by a TMS pulse applied over left motor cortex 150 ms after stop signal presentation, were measured on trials where either the left or right hand was stopped, resulting in four different post-stop MEP conditions: unimanual stop (i.e. MEPs measured from the right hand on right hand stop trials), unimanual unselected (i.e. MEPs measured from the right hand on left hand stop trials), bimanual stop (i.e. stop right, respond left), and bimanual response (i.e. stop left, respond right; Figure 4A). As predicted by

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the global inhibition model, MEPs in the unimanual unselected condition were suppressed compared to baseline (t(14) = -4.33, p<sub>prm</sub> = 0.002, d = -1.12); however, all other conditions were on par with the baseline (all p<sub>prm's</sub> > 0.796, absolute values of d's < 0.50). There were considerable differences between the conditions (one-way ANOVA: F(3,42) = 10.99, p<sub>prm</sub> < 0.001,  $\eta_p^2$  = 0.44). The post-hoc tests revealed that MEPs in the unimanual unselected condition were suppressed compared to all other conditions (all p<sub>prm</sub>'s <= 0.019, d's > 0.81), and MEPs in the bimanual response condition were elevated compared to all other conditions (all p<sub>prm</sub>'s <= 0.040, d's > 0.66). Critically, there were no differences in MEP amplitudes between the unimanual stop and bimanual stop conditions (t(14) = 0.41, p<sub>prm</sub> = 0.698, d = 0.11).

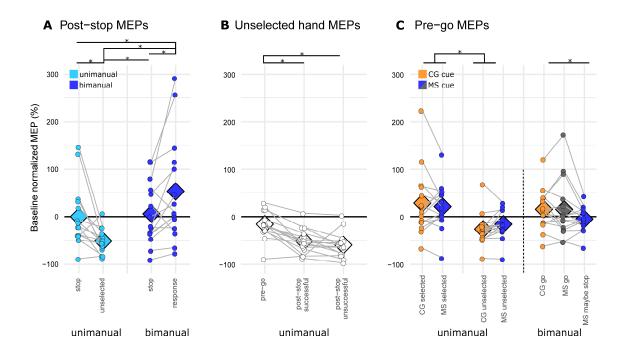


Figure 4. Baseline normalized MEPs measured after the stop signal (A), before go signal and after the stop signal in only the unselected hand (B), and before the go signal (C). The diamonds represent average values across participants, the dots represent single participants, and the gray lines connect the same participants in selected conditions. The dashed vertical line on (C) represents the separation between unimanual and bimanual task conditions. CG – certain go; MS – maybe stop. \*represent significant differences between conditions.

The MEPs showed considerable within-condition variability, particularly in the stopped hand during both tasks and the response hand in the bimanual task. MEPs measured in the stopped hand in both

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tasks were not below baseline. This was likely due to the proximity of the pulse timing to the motor execution processes. Indeed, a large number of trials had to be excluded due to contamination of the EMG time-courses with voluntary muscle activity (64% of successful stop trials). When compared to the prEMG onset in trials without TMS pulses, the timing of the pulse (150 ms after stop) was after the timing of the average prEMG onset (124 ms after stop), suggesting the available stopped hand MEPs are sampled from trials with late response initiation (producing large MEPs) and/or early inhibition (producing reduced MEPs). Thus, the post-stop MEPs likely reflect the net effect of simultaneous excitatory and inhibitory processes within the motor control system that are difficult to dissociate from each other and that contribute to the relatively large variability, both across trials and participants.

The reduced post-stop MEP amplitudes in the unselected hand are in line with the account of a global inhibitory mechanism. However, MEP suppression relative to baseline has been observed during a cued preparatory interval in a delayed response task, similar to that used here (Duque et al., 2017). While here, the pre-go unselected hand MEPs were not significantly different from baseline (see Figure 4B-C and the section about pre-go MEP results for details), they were still negative, and the suppressed post-stop MEPs could reflect a continuation of earlier inhibitory processes engaged during the cue-delay period. We therefore contrasted directly the unselected pre-go MEPs with the unselected post-stop MEPs from successful and unsuccessful stop trials (Figure 4B). There was a significant main-effect for the omnibus rmANOVA (F(2, 28) = 26.73, p<sub>prm</sub> < 0.001,  $\eta_p^2$  = 0.66). The post-hoc tests indicated that both post-stop conditions were further suppressed when compared to the pre-go MEPs (Figure 4B; both p<sub>prm</sub>'s < 0.001, d's > 1.58). Interestingly, the degree of unselected hand MEP suppression was not different between successful and unsuccessful stop trials (t(14) = 1.26, p<sub>prm</sub> = 0.229, d = 0.33).

No differences in frontal beta between uni- and bimanual tasks, nor between successful and unsuccessful stopping. Frontal beta may be a cortical index of response inhibition, and may therefore

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differentiate between the tasks. Specifically, we looked at the early differences prior to the prEMG peak latency, expecting stronger beta in successful than unsuccessful trials. In addition, lower beta power in the bimanual task may be indicative of a slower and more selective inhibitory mechanism. An increase in beta power was observed in all stop trials, though it appeared to be strongest at later time-points (Figure 5) and was not significantly different from zero in the time-window of 100-150 ms after the stop signal (avg. beta over all conditions = 0.16 db (sd = 0.40), t(13) = 1.48, p = 0.164). In contrast to our predictions, there were no significant differences between the tasks (F(1,13) = 0.60,  $p_{prm} = 0.455$ ,  $\eta_p^2 = 0.04$ ), hemispheres (F(2,26) = 0.41,  $p_{prm} = 0.673$ ,  $\eta_p^2 = 0.03$ ), nor between successful and unsuccessful stop trials (F(1,13) = 1.26,  $p_{prm} = 0.285 \eta_p^2 = 0.09$ ).

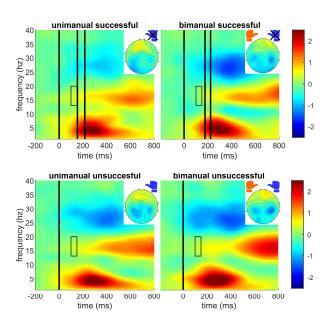


Figure 5. Stop trial time-frequency results time-locked to the stop signal. The images are of the average signal from frontocentral (F1, Fz, F2, FC1, FCz, FC2) electrodes. Vertical lines indicate, from left to right, stop signal onset, average prEMG latency, and average SSRT (only stop onset marked on unsuccessful stop trials). Rectangles indicate the time- and frequencies-of-interest included in the frontal beta analysis (13-20 Hz at 100-150 ms). The topographies show the distribution of beta values across all electrodes at the extracted time-window. The data for left and right hand trials are averaged in all panels, with the electrodes flipped in relevant conditions so that the depicted left hemisphere is contralateral to the stop hand. This is also reflected by the hand symbols, where blue refers to the stop-hand and orange refers to the response-hand, and the crosses above the blue hands indicate that these responses were successfully inhibited.

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**Summary I.** We found no evidence for distinct global and selective stopping mechanisms. There were no differences between unimanual and bimanual stopping in behavior, corticomotor excitability for the stopped hand, or frontal beta power. The delayed prEMG peak in the bimanual task relative to the unimanual task, when considered on its own, is consistent with the dual-mechanism model. However, this delay was greatly diminished, and the between-task comparison no longer reached statistical significance, after accounting for differences in the prEMG onsets. The prevalent interference effect in the bimanual task and the suppression of the unselected hand during stopping, evident in reduced MEP amplitudes, support a single global inhibition mechanism model.

#### 2) Does motor preparation influence subsequent inhibition success?

We hypothesized that the motor cortical activity prior to the stop signal is associated with subsequent stopping success. Specifically, we expected weaker sensorimotor mu and beta desynchronization on successful than on unsuccessful stop trials prior to stop signal presentation. Such a pattern could arise from a less mature go process and/or the influence of proactive control. These effects were expected to be localized to the contralateral hemisphere of the selected hand in the unimanual task. In the bimanual task, we expected a differential lateralization effect, with weaker desynchronization in the hemisphere contralateral to the stopped hand, as opposed to the hemisphere contralateral to the response hand.

Weaker mu desynchronization on successful stop trials. Mu desynchronization was prominent in all trials (Figure 6A; one-sample t-tests against zero: unimanual t(13) = -3.60, p = 0.006; bimanual t(13) = -3.05, p = 0.009). In the unimanual task, post-go mu desynchronization was weaker prior to successful than unsuccessful stop trials (Figure 6B; SUCCESS: F(1,13) = 12.73, p<sub>prm</sub> = 0.004,  $\eta_p^2$  = 0.49). There was also a main effect of LOCATION, indicating that mu desynchronization was stronger in the contralateral hemisphere (LOCATION: F(1,13) = 20.49, p<sub>prm</sub> < 0.001,  $\eta_p^2$  = 0.61). In the bimanual task, there was a significant interaction effect of SUCCESS and LOCATION (F(1, 13) = 14.28, p<sub>prm</sub> = 0.004,  $\eta_p^2$  = 0.52), but no significant main effects. The interaction reflected that mu desynchronization was weaker prior to

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successful stop trials in the hemisphere contralateral to the stopped hand, compared to the hemisphere contralateral to the response hand (Figure 6B, right panel; t(13) = 3,41,  $p_{prm} = 0.020$ , d = 0.91). In contrast, mu activity was similar in both hemispheres on unsuccessful stop trials (t(13) = 0.48,  $p_{prm} \sim 1.000$ , d = 0.12). The contralateral mu activity was also numerically weaker on successful than on unsuccessful stop trials, but this difference did not reach significance at the selected alpha level (t(13) = 3.08,  $p_{prm} = 0.052$ , d = 0.82). This pattern indicates the mu activity was weaker in the hemisphere contralateral to the hand that was cued to stop, suggesting that hand-specific motor activity influences subsequent stopping.

Weaker contralateral beta desynchronization in successful stop trials in the unimanual task. The post-go beta power was also reduced in both tasks (Figure 6A; one sample t-tests against zero: unimanual t(13) = -6.33, p < 0.001; bimanual t(13) = -5.11, p < 0.001), yet the effects observed here differed from those observed in the mu band. Significant effects were limited to the unimanual task where there was a main effect of LOCATION (F(1,13) = 6.184, p<sub>prm</sub> = 0.025,  $\eta_p^2$  = 0.32), and an interaction of SUCCESS and LOCATION F(1,13) = 10.30, p<sub>prm</sub> = 0.005,  $\eta_p^2$  = 0.44; Figure 6C). The latter reflected the finding that beta desynchronization in the contralateral hemisphere was weaker prior to successful stop trials compared to unsuccessful stop trials (t(13) = 3.04, p<sub>prm</sub> = 0.019, d = 1.05). Further, beta desynchronization was stronger in the contra- than ipsilateral hemisphere prior to unsuccessful (t(13) = 5.88, p<sub>prm</sub> = 0.004, d = 1.57), but not prior to successful stop trials (d(13) = 0.11, p<sub>prm</sub> ~ 1.000, d = 0.03).

**Summary II.** The state of motor activity prior to stop signal onset influenced subsequent stopping success. In the unimanual task, both mu and beta desynchronization were weaker prior to successful compared to unsuccessful stopping. Selective stopping in the bimanual task exhibited a distinct pattern of mu lateralization: mu desynchronization contralateral to the stopped hand was weaker than in the hemisphere contralateral to the responding hand prior to successful stop trials, while no lateralization effect was observed prior to the unsuccessful stop trials.

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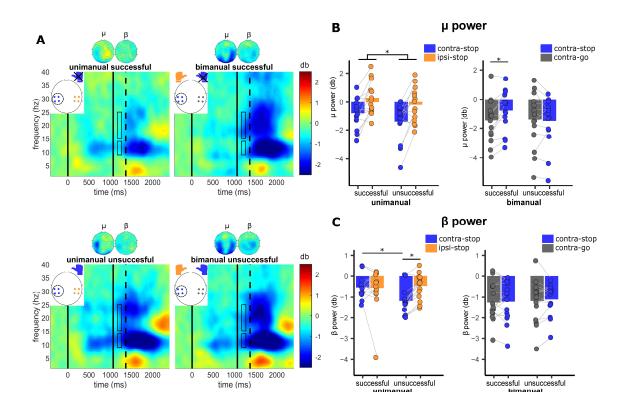


Figure 6. Time-frequency results on stop trials, time-locked to the cue onset. The spectrograms represent the average signal from the circled electrodes on the white topographical plots. The electrode colors correspond to the hemispheres shown in B and C; the blue hands symbolize the stop response (top: successful, and bottom: unsuccessful), and the orange hands symbolize the (correctly) executed response. The three black vertical lines represent, from left to right, cue onset, go onset, and average stop signal onset. Rectangles indicate the time-and frequencies-of-interest included in the analysis (9-14 Hz (mu) and 15-25 Hz (beta) at 100-200 ms after go onset). The topographies show the distribution of mu and beta values across all electrodes at the extracted time-window. **B**. Average mu power. **C**. Average beta power.

# 3) Is motor preparation affected by proactive control?

The effect of proactive control on motor activity was tested by contrasting the trials with cues indicating that the prepared response may need to be stopped (MS) to the trials where the cue indicated the certain execution of the prepared response (CG). We expected the MS cue to lead to slower go RTs, altered MEP amplitudes during the cue-delay period, and reduced motor mu and beta desynchronization.

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**Delayed go RTs in the MS cue condition.** Relative to the CG condition, RTs were delayed by ~150 ms in the MS condition (Table 2; CUE: F(1,14) = 33.78,  $p_{prm} < 0.001$ ,  $\eta_p^2 = 0.71$ ). This effect was also evident in the EMG traces (Table 3, F(1,14) = 33.43,  $p_{prm} < 0.001$ ,  $\eta_p^2 = 0.70$ ). The increase in RT in the MS condition was not different between the tasks (TASK: F(1, 14) = 1.01,  $p_{prm} = 0.326$ ,  $\eta_p^2 = 0.07$ ; EMG onsets: F(1,14) = 3.01,  $p_{prm} = 0.104$ ,  $\eta_p^2 = 0.18$ ). Importantly, there were no systematic RT differences in the bimanual go trials of the MS condition between the hand that was cued to respond compared to the hand that was cued to potentially stop (536 and 535 ms, respectively; t(14) = 1.66,  $p_{prm} = 0.122$ , d = 0.43), and on average, the two responses were initiated in close synchrony (mean absolute difference 15 ms in both cue types; t(14) = -1.40,  $p_{prm} = 0.174$ , d = -0.36). In sum, proactive control led to substantial response slowing of comparable magnitude in both tasks.

Hand-specific cue modulation of the MEPs in the bimanual task, but no cue effects in the unimanual task. Turning to the TMS data, we did not observe suppression of corticospinal excitability during the delay period for either type of cue (Figure 4C). Relative to baseline, MEP amplitudes tended to increase in the selected hand (both tasks) during the delay period, and to decrease in the unselected hand (unimanual tasks); however, these changes were not statistically different from baseline (all  $p_{prm}$ 's > 0.940, absolute values of d < 0.46).

In the unimanual task, MEPs elicited in the unselected hand were smaller than those elicited in the selected hand (HAND: F(1,14) = 8.90,  $p_{prm} = 0.009$ ,  $\eta_p^2 = 0.39$ ), with no difference between the CG and MS conditions (CUE: F(1,14) = 0.07,  $p_{prm} = 0.811$ ,  $\eta_p^2 < 0.01$ ). In the bimanual task, a one-way rmANOVA comparing the three conditions, CG, MS go hand, MS maybe stop hand, showed a significant main effect (F(2, 28) = 3.29,  $p_{prm} = 0.045$ ,  $\eta_p^2 = 0.19$ ). Post hoc tests indicated that CG MEPs were larger than MEPs in the hand that potentially would have to be stopped in the MS condition (t(14) = 3.12,  $p_{prm} = 0.014$ , d = 0.81), but indicated no other differences (p's > 0.186, d's < 0.47).

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In sum, we did not observe the expected reduction in corticospinal excitability during the delay period, and most relevant to the current study, there were no consistent differences between the CG and MS conditions. The one exception was in the bimanual task where MEPs were smaller in the maybe stop hand in the MS condition relative to the CG condition.

**Reduced mu desynchronization in the MS cue condition.** Sensorimotor mu and beta desynchronization were tested between the cue types and hemispheres. To evaluate the time course of the desynchronization effects, we included time-window (pre- vs post-go signal) as an additional factor in the analyses (Figure 7A). One-sample t-tests against zero confirmed the presence of mu desynchronization both in the pre-go (unimanual t(13) = -4.95, p = 0.001; bimanual t(13) = -3.09, p = 0.011) and post-go period (unimanual t(13) = -4.34, p = 0.002; bimanual t(13) = -3.33, p = 0.011).

In the unimanual task, mu desynchronization was stronger following CG cues compared to MS cues (CUE: F(1, 13) = 19.90,  $p_{prm} < 0.001$ ,  $\eta_p^2 = 0.60$ ), and stronger over the contralateral hemisphere compared to the ipsilateral hemisphere (LOCATION: F(1, 13) = 26.12,  $p_{prm} < 0.001$ ,  $\eta_p^2 = 0.67$ ; Figure 7B, left panel). This desynchronization was evident during the delay period and persisted following the imperative stimulus with no difference between the pre- and post-go periods (Figure 7D; TIME: F(1,13) = 2.81, p = 0.119,  $\eta_p^2 = 0.18$ ).

In the bimanual task, since both hands are responding simultaneously, the terms ipsilateral and contralateral are not unambiguously defined. Therefore, both cue and hemispheric effects were tested with a rmANOVA with a combined CUE-LOCATION factor (levels CG bilateral, MS contra-go, MS contra-stop). This analysis revealed a significant main effect (F(2, 16) = 5.81, p<sub>prm</sub> = 0.009,  $\eta_p^2$  = 0.31; Figure 7B, right panel). The post-hoc tests indicated that the mu desynchronization was stronger in the CG than in the MS contra-stop condition (t(13) = 2.61, p<sub>prm</sub> = 0.045, d = 0.70), paralleling the changes in the pre-go MEP amplitudes. Here, lateralization was evident also within the MS cue condition. Mu desynchronization in the hemisphere contralateral to the go response was stronger

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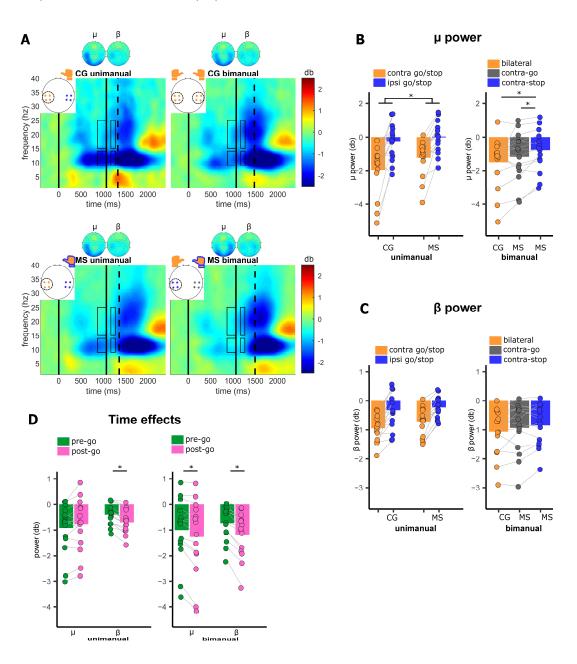
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compared to the hemisphere contralateral to the maybe-stop response (t(13) = 3.34,  $p_{prm} = 0.017$ , d = 0.89). This is interesting, given that the bimanual response was executed simultaneously, with no RT differences between the hand that was cued to respond vs. the hand that was cued to potentially stop. Unlike the unimanual task, there was a significant main effect of TIME (Figure 7D; F(1, 13) = 9.29,  $p_{prm} = 0.008$ ,  $\eta_p^2 = 0.42$ ) with stronger mu desynchronization in the post-go than in the pre-go period.

No cue effects on motor beta desynchronization. Beta desynchronization was also evident in both tasks (Figure 7A and 7C) during pre-go (one-sample t-tests against zero: unimanual t(13) = -5.93, p < 0.001; bimanual t(13) = -3.97, p = 0.002) and post-go periods (unimanual t(13) = -7.98, p < 0.001 ; bimanual t(13) = -5.00, p < 0.001). Beta desynchronization was stronger later in the trial, after the go signal (Figure 7D; TIME unimanual: F(1, 13) = 81.23, p<sub>prm</sub> < 0.001,  $\eta_p^2$  = 0.86; bimanual: F(1, 13) = 39.42, p < 0.001,  $\eta_p^2$  = 0.75). Beta desynchronization was stronger in the contra- than in the ipsilateral hemisphere, relative to the selected hand in the unimanual task (LOCATION: F(1, 13) = 45.02, p<sub>prm</sub> < 0.001,  $\eta_p^2$  = 0.78). No cue effects were found in either task in the tested time-windows (CUE unimanual: F(1,13) = 3.03, p = 0.105,  $\eta_p^2$  = 0.19; CUE-LOCATION bimanual: F(2, 26) = 3.14, p<sub>prm</sub>= 0.059,  $\eta_p^2$  = 0.19).

**Summary III.** Proactive control influenced motor preparation, evident in the pronounced RT increase on MS trials in both tasks. MEPs during the delay period were attenuated in the MS condition in the bimanual task, but only in the hand that was cued to maybe-stop. In the unimanual condition, MEP amplitudes in the pre-go period did not differ between the two types of cues. Thus, the MEP changes were restricted to the condition in which participants prepared for selective stopping, consistent with previous studies (Claffey et al., 2010; Cai et al., 2011). As expected, mu desynchronization was weaker in MS than in CG trials, paralleling the RT differences. Moreover, mu desynchronization was weaker in the hemisphere corresponding to the hand cued for possible stopping compared to the freely responding hand in the bimanual task. This laterality effect is consistent with the observed pattern of task-specific MEP modulation. In contrast, beta power did not differ between the cue types in either task and developed later in the trial.

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**Figure 7**. **A**. Go trial time-frequency results, time-locked to the cue onset. The spectrograms depict average signal from the encircled electrodes on the white topographical plots. The electrode colors correspond to the hemispheres shown in B and C; all hand symbols are orange depicting that the response was always given, but the blue outline indicates which hands were additionally prepared for the maybe stop response. Vertical lines indicate, from left to right, cue onset, go onset, and average go-EMG onset. Rectangles indicate the time- and frequencies-of-interest included in the analysis (9-14 Hz (mu) and 15-25 Hz (beta) at -200-0 and 100-200 ms relative to go onset. The topographies show the distribution of mu and beta values across all electrodes averaged over both pre-go and post-go time-windows. **B**. Average mu power (averaged over pre-go and post-go period). **C**. Average beta power (averaged over pre-go and post-go period). **D**. Pre-go and post-go mu and beta power, averaged over hemispheres and cue conditions. CG – certain go, MS – maybe stop.

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# Discussion

# Single inhibition mechanism with predominantly global extent

We evaluated the explanatory power of two alternate models of response inhibition. As summarized in Table 4, the results were more consistent with the single global inhibition model as indicated by stopping interference observed in the behavior and EMG time-courses, as well as the suppressed MEPs of the unselected hand after the stop signal. There were no differences in the SSRTs between uni- and bimanual stopping, stopped hand MEP amplitudes, or frontal beta power, at odds with the the predictions of the dual-mechanism model in which inhibition can be selectively directed. An initial difference in prEMG between the tasks diminished after controlling for dependencies between go and stop processes, suggesting that the hypothesized delays previously associated with selective inhibition may be driven by differences in the go process or the interactions between going and stopping.

Global inhibition	
Interference in the responding hand RT in bimanual successful stop trials	$\checkmark$
Interference in the responding hand EMG in the bimanual successful stop trials	$\checkmark$
Suppressed unselected hand post-stop MEP amplitude in the unimanual task	$\checkmark$
Dual-mechanism	
Delayed inhibition latency in the bimanual task (SSRT and prEMG latency)	×
Weaker post-stop MEP suppression in the bimanual task	×
Reduced post-stop beta power in the bimanual task	×

Table 4. Predictions regarding putative inhibition indices (SSRT, prEMG, MEP amplitudes, frontal beta power) in the context of the two models. The marks on the right indicate whether these predictions were confirmed (checkmark) or not confirmed (X) by the data.

While the preponderance of evidence supports a single global inhibition mechanism, a few effects were indicative of selective stopping. We devised a novel single-trial marker, the  $\Delta$ peak, for quantifying interference in the bimanual task for trials that exhibit prEMG. Despite the prevalent interference in the EMG at the group level, the single-trial  $\Delta$ peak indicated large variability, including trials with no interference. This is in line with previous findings showing that interference can be eliminated under specific conditions (Xu et al., 2015) or when a small delay is required between the

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bimanual left and right hand (in contrast to simultaneous) responses that necessitates decoupling of the bimanual response (Wadsley et al., 2019). These results highlight that the stop-restart behavior in the bimanual stop signal task is not mandatory, but a preferred option when selectivity is not incentivized by the task requirements. Our single-trial EMG results further indicate that stopping consists of a mix of processes, including trials where the entire action plan was aborted (large  $\Delta$ peak), trials where there was selective response preparation and/or inhibition (small or no  $\Delta$ peak), and trials where the motor plan was aborted before motor initiation (no prEMG; De Jong et al., 1990; McGarry and Franks, 1997).

These results motivate an alternative to the global inhibition model, one that highlights flexible control, the extent and speed of which vary on a trial-by-trial basis depending on the motivational and task context. Inhibition may be one component of a continually revised complex motor plan, incorporating facilitatory and inhibitory drives. These dynamics could depend on the wider sensorimotor control system spanning the dorsal attentional, motor, and premotor areas (Cisek and Kalaska, 2010; Mirabella, 2014). One overarching function of cortical inhibition in this model could be to set the gain for motor representations or their thresholds, with the extent of inhibition changing according to uncertainty about the response alternatives or requirements for proactive control (Greenhouse et al., 2015).

#### Proactive control, motor preparation, and response inhibition

Proactive control biases sensory and attentional systems in expectation of a stop signal. This results in slower go-trial performance, but heightened sensitivity for stop stimulus (Elchlepp et al., 2016; Langford et al., 2016). Our results speak to the role of anticipatory regulation on the motor system. First, we observed that motor cortical activity differed between successful and unsuccessful stop trials prior to stop signals. Second, mu activity was affected by proactive cues in a hand-specific way, similarly to the pre-go MEP modulation in the bimanual task. These results suggest functional interactions between proactive control, motor preparation, and response inhibition. Successful

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inhibition may arise with respect to acquired stimulus-stop associations in a quick and reflexive manner with sensory, attentional, and motor parameters pre-set by proactive control (Verbruggen et al., 2014a, 2014b). In the motor system, this may be achieved by regulating response thresholds or motor gain through premotor-motor cortical projections or via fronto-basal ganglia networks (Jahfari et al., 2012; Stuphorn and Emeric, 2012).

## EEG signatures of response inhibition and motor preparation

A number of EEG markers have been proposed to index response inhibition. However, most of these occur relatively late in the trial. The prEMG latencies in this study are consistent with recent findings showing that inhibition occurs around 140-180 ms after the stop signal (Raud and Huster, 2017; Atsma et al., 2018; Hannah et al., 2019; Huster et al., 2019; Jana et al., 2019; Thunberg et al., 2019; Raud et al., 2020). Beta power increases in the right inferior frontal gyrus and pre-supplementary motor area within 100 ms following the stop signal (Swann et al., 2009, 2012); thus, frontal beta in the scalp EEG may be a temporal marker of inhibition. However, we did not find increased beta in the time-window of 100-150 ms after stopping. Moreover, it did not differentiate between selective and global stopping, nor between successful and unsuccessful stop trials. The latter is in line with intracranial recordings of beta power in the pre-SMA (Swann et al., 2012) and a study showing that the number of frontal beta bursts (a potentially more sensitive marker than power fluctuations) did not differ between successful and unsuccessful stopping before 200 ms (Wessel, 2020). As such, a fast, reactive inhibition mechanism may invariably be engaged whenever a stop signal is detected, while the success of stopping may depend on the perceptual speed and the contemporaneous state of the motor cortex.

Regarding the cortical signatures of activity within the motor system, our results suggest a dissociation between the mu and beta rhythms. We observed weaker mu desynchronization accompanied by delayed go RTs in the cued condition. In the bimanual task, we observed weaker mu desynchronization in the hemisphere contralateral to the maybe-stop hand, paralleled by reduced MEP amplitudes in the hand cued for stopping. Weaker mu desynchronization may therefore reflect the engagement of

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proactive control or its downstream effects on the motor system, in line with the hypothesis that cortical oscillations in this frequency band reflect more general inhibitory functions not those limited to the motor cortex (Jensen and Mazaheri, 2010). Accordingly, Muralidharan et al. (2019) reported a power increase relative to the baseline in the hemisphere contralateral to the stop hand in a frequency range incorporating both mu and beta bands, which they suggest was reflective of proactive selective inhibition of the motor cortex. Here, mu activity was negative in both hemispheres. It is thus important to acknowledge the possibility of relative differences between the two hemispheres arising as a marker of lateralized disinhibition associated with the portion of the upcoming movement that can be freely executed.

Motor beta activity did not differ between the cue types nor between the two hemispheres in the bimanual task. This is in contrast with previous studies that report differences between the cues in unimanual inhibition tasks, particularly in the ipsilateral hemisphere (Liebrand et al., 2017, 2018). We did observe stronger beta desynchronization after the go signal relative to the cue-delay period. Thus, beta may be directly related to the motor output upon the execution of the movement, being less affected by proactive dynamics during action planning.

#### Limitations

While our results agree with the previous reports on uni- and bimanual stopping, there are important differences to note. The validity of the SSRT relies on the assumption of independent go and stop processes, yet we demonstrated and quantified the violation of this assumption at the single trial level. The validity of existing methods for calculating SSRT has been called into question based on observations of similar violations (Ozyurt et al., 2003; Bissett and Logan, 2014; Gulberti et al., 2014; Verbruggen and Logan, 2015), biases produced by subtle changes in the go RT distributions (Verbruggen et al., 2013), and the inability to account for attentional lapses (Matzke et al., 2017; Heathcote et al., 2019). Altogether, these critical findings call for a re-conceptualization of SSRT differences observed between different conditions.

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We did not replicate the robust preparatory inhibition of corticomotor excitability in a hand selected for a forthcoming response reported in a number of studies (e.g., Duque et al., 2017). Several design features of our study differed from previous studies (e.g., the inclusion of stop trials, the target muscle, stimulation hemisphere, pulse timing, and inter-stimulation interval among others). As such, preparatory inhibition may not generalize across different task contexts (Quoilin et al., 2019).

#### Conclusions

Our results favored a single inhibition mechanism over a model involving independent selective and global mechanisms. We further observed that the success of stopping differed according to the state of sensorimotor activity in the mu frequency band preceding the stop signal. Moreover, the scalp distribution of sensorimotor mu activity corresponded to TMS measurements of hand-specific reduction of corticomotor excitability in the hand cued for possible stopping. The results, taken as a whole, are consistent with a model emphasizing the global extent of inhibition. However, the novel use of EMG activity to quantify stopping interference at the single trial level demonstrated that selective stopping can be achieved in certain trials, pointing to greater flexibility of inhibitory mechanisms in terms of their spatial extent.

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