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Application of a novel haplotype-based scan for local adaptation to study high-altitude adaptation in rhesus macaques

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Abstract

When natural populations split and migrate to different environments, they may experience different selection pressures that can lead to local adaptation. To capture the genomic patterns of a local selective sweep, we develop XP-nSL, a genomic scan for local adaptation that compares haplotype patterns between two populations. We show that XP-nSL has power to detect ongoing and recently completed hard and soft sweeps, and we then apply this statistic to search for evidence of adaptation to high altitude in rhesus macaques. We analyze the whole genomes of 23 wild rhesus macaques captured at high altitude (mean altitude > 4000m above sea level) to 22 wild rhesus macaques captured at low altitude (mean altitude < 500m above sea level) and find evidence of local adaptation in the high-altitude population at or near 303 known genes and several unannotated regions. We find the strongest signal for adaptation at EGLN1, a classic target for convergent evolution in several species living in low oxygen environments. Furthermore, many of the 303 genes are involved in processes related to hypoxia, regulation of ROS, DNA damage repair, synaptic signaling, and metabolism. These results suggest that, beyond adapting via a beneficial mutation in one single gene, adaptation to high altitude in rhesus macaques is polygenic and spread across numerous important biological systems.

Impact Summary

When positive selection is ongoing or a beneficial mutation has recently fixed in a population, genetic diversity is reduced in the vicinity of the adaptive allele, and we expect to observe long homozygous haplotypes at high frequency. Here we develop a statistic that summarizes these expected patterns and compares between two

74 populations in order to search for evidence of adaptation that may have occurred in one
75 but not the other. We implement this statistic in a popular and easy-to-use software
76 package, and then apply it to study adaptation to high altitude in rhesus macaques.

77 Extreme environments pose a challenge to life on multiple fronts. Very high-
78 altitude environments are one such example, with low atmospheric oxygen, increased
79 ultraviolet light exposure, harsh temperatures, and reduced nutrition availability. In spite
80 of these challenges, many plants and animals, including humans, have genetically
81 adapted to cope with these hardships. Here we study two populations of rhesus
82 macaques, one living at high altitude and one living close to sea level. We apply our
83 novel statistic to compare their haplotype patterns between them to search for evidence
84 of genetic changes that are indicative of adaptation to their environment.

85 We find evidence for adaptation at a critical gene that helps control physiological
86 response to low-oxygen, one that has been the target of repeated convergent evolution
87 across many species. We also find evidence for positive selection across a range of
88 traits, including metabolic and neurological. This work helps to explain the evolutionary
89 history of the rhesus macaque and furthers our understanding about the ways
90 organisms genetically adapt to high-altitude environments.

91 Introduction

92 Selective sweeps produce regions of reduced genetic diversity in the vicinity of
93 an adaptive mutation. These patterns manifest as long extended regions of
94 homozygous haplotypes segregating at high frequency (Przeworski 2002; Sabeti *et al.*
95 2002; Kim and Nielsen 2004; Garud *et al.* 2015). In the event of a *de novo* mutation that
96 is adaptive in a population, we expect the haplotype it resides on to rapidly rise in

97 frequency in the population (called a ‘hard’ sweep). On the other hand, if an ancestrally
98 segregating neutral or mildly deleterious allele turned out to be adaptive in a new
99 environment, it would likely reside on two or more haplotypes, which would rapidly rise
100 in frequency in the population (called a ‘soft’ sweep) (Hermisson and Pennings 2005;
101 Pennings and Hermisson 2006). As both of these processes happen on a time scale
102 faster than mutation or recombination can act to break up the sweeping haplotypes, we
103 expect to observe long and low diversity haplotypes at high frequency in the vicinity of
104 an adaptive mutation. However, if this mutation either does not exist or is not adaptive in
105 a sister population, we would not expect a sweep to occur and thus we would not
106 expect to observe similar haplotype patterns.

107 To capture these haplotype patterns and contrast them between a pair of
108 populations, we develop XP-nSL, a haplotype-based statistic with good power to detect
109 partial, fixed, and recently completed hard and soft sweeps by comparing a pair of
110 populations. XP-nSL is an extension of nSL (Ferrer-Admetlla *et al.* 2014) and does not
111 require a genetic recombination map for computation. The lack of dependence on a
112 recombination map is important, as other statistics for identifying positive selection are
113 biased towards low-recombination regions (O'reilly *et al.* 2008), but the approach taken
114 by nSL has been shown to be more robust (Ferrer-Admetlla *et al.* 2014). Both nSL and
115 XP-nSL summarize haplotype diversity by computing the mean number of sites in a
116 region that are identical-by-state across all pairs of haplotypes. Whereas nSL contrasts
117 between haplotype sets carrying an ancestral or a derived allele in a single population,
118 XP-nSL contrasts between haplotype sets in two different populations, allowing it to
119 detect local adaptation.

An extreme example of adaptation to a local environment is the transition to high-altitude living. Organisms living at high altitude are confronted with many challenges, including a low-oxygen atmosphere and increased ultraviolet light exposure, and these harsh environments inevitably exert strong selection pressure. Indeed, adaptation to high-altitude living has been studied extensively across many organisms from plants, including monocots (Gonzalo-Turpin and Hazard 2009; Ahmad *et al.* 2016) and dicots (Kim and Donohue 2013; Liu *et al.* 2014; Munne-Bosch *et al.* 2016; Guo *et al.* 2018), to numerous animals including amphibians (Yang *et al.* 2016), canids (Li *et al.* 2014; Wang *et al.* 2014; Wang *et al.* 2020), humans (Bigham *et al.* 2009; Bigham *et al.* 2010; Xu *et al.* 2010; Yi *et al.* 2010; Peng *et al.* 2011; Huerta-Sanchez *et al.* 2013; Huerta-Sanchez *et al.* 2014; Jeong *et al.* 2014), yaks (Qiu *et al.* 2012), birds (Cai *et al.* 2013; Qu *et al.* 2013; Wang *et al.* 2015; Graham and Mccracken 2019), boars (Li *et al.* 2013), mice (Storz *et al.* 2007; Cheviron *et al.* 2012; Schweizer *et al.* 2019; Storz *et al.* 2019; Velotta *et al.* 2020), moles (Campbell *et al.* 2010), antelope (Ge *et al.* 2013), and horses (Hendrickson 2013). Liu *et al.* (2018) recently sequenced and published the whole genomes of 79 wild-born Chinese rhesus macaques collected from multiple sites in China. Among these animals, 23 were sampled from far western Sichuan province in a region with mean altitude > 4000 m above sea level (Liu *et al.* 2018), providing an opportunity to study the genetics of local adaption to high altitude in rhesus macaques.

Rhesus macaques are the second most widely distributed primate, with a range extending from Afghanistan to Vietnam and from a latitude of 15 to 38 degrees north (Fooden 2000). Early ancestors of the macaque migrated out of Africa to the Eurasian continent approximately 7 mya—the earliest catarrhine fossils on the continent are

143 macaque-like (Stewart and Disotell 1998). Modern rhesus macaques trace a recent
144 origin to Southeast Asia, with a major migratory split occurring approximately 162 kya
145 separating the ancestors of modern Indian and Chinese rhesus macaques (Hernandez
146 *et al.* 2007). Macaques have proven to be quite evolutionarily successful, demonstrating
147 ecological flexibility and adaptability via developmental plasticity and behavioral
148 changes (Richard *et al.* 1989; Madrid *et al.* 2018). Other studies have looked at how the
149 rhesus macaque radiation has led to population-level adaptation to climate and food
150 availability (Liu *et al.* 2018).

151 Here we test and evaluate our XP-nSL statistic and apply it to study the genomic
152 consequences of high-altitude living in the rhesus macaque. We use it to compare the
153 haplotype patterns of the 23 animals from the high-altitude population with another 22
154 that Liu *et al.* (2018) sampled in lower-lying regions in eastern China with a mean
155 altitude < 500 m above sea level.

156 Methods

157 Data Preparation

158 Liu *et al.* (2018) generated whole genome sequencing data for 79 Chinese
159 macaques and called all biallelic polymorphic sites according to GATK best practices
160 using rheMac8, identifying 52,534,348 passing polymorphic autosomal sites. We then
161 filter all loci with > 10% missing data leaving 35,639,395 biallelic sites. Next, the
162 program SHAPEIT v4.1.2 (Delaneau *et al.* 2019) was used to phase haplotypes in the
163 full data set with a genetic map that was available for rheMac8 (Bcm-Hgsc 2020).
164 SHAPEIT performs imputation during phasing for any missing genotypes. Our analyses
165 here focus on 45 of the 79 samples from Liu *et al.* (2018), representing 23 from high-

altitude regions of China (*M. m. lasiotis*) and 22 from low-altitude regions of China (*M. m. littoralis*), based on capture location information. See Table S1 for individual IDs used.

Liu *et al.* (2018) also inferred joint demographic histories for their five populations, and we extract the demographic parameters for our two of interest. This demographic history is recapitulated in Fig. 1 with detailed parameters given in Table 1, which are then used for simulations to test XP-nSL.

A Statistic for Detecting Local Adaptation

We developed a cross-population haplotype-based statistic, XP-nSL, to scan for regions of the genome implicated in local adaptation between two populations by extending nSL (Ferrer-Admetlla *et al.* 2014), each of which is defined below.

Consider the sets $A(k)$ and $D(k)$, representing the set of haplotypes at site k carrying the ancestral or derived allele, respectively, and let $n_A(k) = |A(k)|$ and $n_D(k) = |D(k)|$. Next $L_{ij}(k)$ is defined as the number of consecutive sites at which haplotype i and j are identical-by-state (IBS) in the interval containing site k. Then nSL at site k is $nS_L(k) = \log\frac{SL_A(k)}{SL_D(k)}$, where $SL_A(k) = \binom{n_A(k)}{2}^{-1}\sum_{i<j\in A(k)} L_{ij}(k)$ and $SL_D(k) = \binom{n_D(k)}{2}^{-1}\sum_{i<j\in D(k)} L_{ij}(k)$. $SL_A(k)$ and $SL_D(k)$ represent the mean $L_{ij}(k)$ over all pairs of haplotypes carrying either the ancestral or derived allele at locus k, respectively. nSL scores are then normalized genome-wide in site-frequency bins either with respect to the empirical background or neutral simulations with a matching demographic history. The nSL computation is illustrated in Ferrer-Admetlla *et al.* (2014). nSL is implemented in `nsl` (Ferrer-Admetlla *et al.* 2014) and `selscan` v1.1.0+ (Szpiech and Hernandez 2014).

XP-nSL is defined similarly, except instead of comparing sets of haplotypes containing ancestral or derived alleles, it compares sets of haplotypes between two different populations. Let $P_1(k)$ and $P_2(k)$, represent the set of haplotypes at site k in population 1 and population 2, respectively, and let $n_{P_1}(k) = |P_1(k)|$ and $n_{P_2}(k) = |P_2(k)|$. Then XP-nSL at site k is $XPnS_L(k) = \log \frac{SL_{P_1}(k)}{SL_{P_2}(k)}$, where $SL_{P_1}(k) = \binom{n_{P_1}(k)}{2}^{-1} \sum_{i<j \in P_1(k)} L_{ij}(k)$ and $SL_{P_2}(k) = \binom{n_{P_2}(k)}{2}^{-1} \sum_{i<j \in P_2(k)} L_{ij}(k)$. $SL_{P_1}(k)$ and $SL_{P_2}(k)$ represent the mean $L_{ij}(k)$ over all pairs of haplotypes in population 1 or population 2 at locus k, respectively. XP-nSL scores are then normalized genome-wide either with respect to the empirical background or neutral simulations with a matching demographic history. The XP-nSL computation is illustrated in Fig. 2 with a toy example. We implement XP-nSL in `selscan` v1.3.0+ (Szpiech and Hernandez 2014) to facilitate wide adoption. It is worth noting that XP-nSL is analogous to XP-EHH (Sabeti *et al.* 2007) as nSL is analogous to iHS (Voight *et al.* 2006).

The goal of these statistics is to capture a signal of extended regions of low diversity on sweeping haplotypes (emblematic of an ongoing or recently completed selective sweep) within a population (nSL) or on sweeping haplotypes in one population versus another (XP-nSL). When XP-nSL scores are positive this suggests evidence for a hard or soft sweep in population 1, and when XP-nSL scores are negative this suggests evidence for a hard or soft sweep in population 2.

Simulations

In order to test the ability of XP-nSL to detect ongoing and recently completed hard and soft sweeps, coalescent simulations were performed conditional on an allele frequency trajectory with the program `discoal` (Kern and Schrider 2016). `discoal`

simulates an allele frequency trajectory for a single non-neutral allele backwards in time and then simulates a neutral coalescent process conditional on this trajectory. This takes advantage of the speed and efficiency of the coalescent while still being able simulate genetic diversity patterns in the vicinity of a non-neutral locus.

All simulations were run with a two-population divergence demographic history (Fig. 1 and Table 1), as inferred by Liu *et al.* (2018), and given by the following `discoal` command line arguments `-p 2 46 44 -en 0 1 0.230410476572876 -en 0.071964666275442 1 0.278345739259351 -en 0.086558359329153 1 1.282885999320505 -ed 0.146214905642895 1 0 -en 0.146214905642895 0 4.089940389782871`. Here 46 haplotypes were sampled from population 0, which has the demographic history of the high-altitude population, and 44 haplotypes were sampled from population 1, which has the demographic history of the low-altitude population. A mutation rate of $\mu = 2.5 \times 10^{-8}$ (Fan *et al.* 2018) was used, along with a recombination rate of $r = 5.126 \times 10^{-9}$, which was computed as the genome-wide mean rate from the rheMac8 recombination map (Bcm-Hgsc 2020). A 500 kb region was simulated, thus giving a scaled mutation rate and scaled recombination parameters for `discoal` as `-t 809.425 -r 165.967`. 5,349 replicates of neutral sequence were simulated under this model, representing approximately the entire macaque genome minus 500 kb. Thus, the total simulated length of all neutral regions plus one selected region is approximately equal to the macaque genome length.

For non-neutral simulations, sweep scenarios were simulated with a positive additive selection coefficient $s \in \{0.01, 0.02, 0.05\}$, which is provided to `discoal` as a scaled selection coefficient $2N_1 s \in \{323.77, 647.54, 1618.85\}$ (`discoal` flag `-a`). Soft

sweeps are simulated as a mutation that arose neutral and turned beneficial at a particular establishment frequency $e \in \{0.01, 0.02, 0.03, 0.04, 0.05, 0.10, 0.20\}$ (`discoal` flag `-f`). Hard sweeps are simulated from a *de novo* mutation that was never neutral, i.e. $e = 0$. Finally, sweeps were conditioned on having either reached a certain frequency in the population at the time of sampling, $f \in \{0.7, 0.8, 0.9, 1.0\}$ (`discoal` flag `-c`), or that the adaptive mutation reached fixation some number of generations prior to sampling, $g \in \{50, 100, 200\}$, which is provided to `discoal` in coalescent units $g/2N_1 \in \{1.5443 \times 10^{-3}, 3.0886 \times 10^{-3}, 6.1774 \times 10^{-3}\}$ (`discoal` flag `-ws`). For the sake of being conservative in our estimation of power (Voight *et al.* 2006; Sabeti *et al.* 2007), it was assumed the actual adaptive mutation remains unsampled (`discoal` flag `-h`) even though whole genome sequencing data are being analyzed. For each combination of parameter values, 500 replicates were simulated.

Neutral Simulations of Mismatched Histories

We also generate neutral simulations for three mismatched demographic histories, in order to study how a mismatched demographic history may influence the power and false positive rates of XP-nSL if used as a normalization baseline. We name these mismatched histories “Rand”, “Under”, and “Over” and generate 5,349 replicates for each one.

For the “Rand” history, each parameter is (uniformly) randomly chosen from within the 95% CI as inferred by (Liu *et al.* 2018) (see Table 1). The parameters are resampled for each replicate of the “Rand” history, and the condition $T_1 < T_2 < T_3$ is enforced. For the “Under” history, the present-day population size parameters N_1 and N_2 are the only ones modified. They are set to $N_1 = 8255$ at the

extreme low end of the 95% CI and $N_2 = 5006$ at the extreme high end of the 95% CI. This represents a scenario where the difference in population sizes is underestimated. For the "Over" history, once again the present-day population size parameters N1 and N2 are the only ones modified. They are set to $N_1 = 26825$ at the extreme high end of the 95% CI and $N_2 = 660$ at the extreme low end of the 95% CI. This represents a scenario where the difference in population sizes is overestimated.

Detecting Local Adaptation in Real Data

From the phased data set, animals captured at high altitude (n = 23) and animals captured at low altitude (n = 22) were subset (see Table S1). Using `selscan` v1.3.0 (Szpiech and Hernandez 2014) to compute raw XP-nSL scores across the genome (`selscan` flags –xpnsl –vcf high-altitude.vcf –vcf-ref low-altitude.vcf), scores were then normalized using the genome-wide empirical background with `selscan`'s `norm` v1.3.0 program (`norm` flag --xpnsl). The low-altitude population was used as the reference population, so positive XP-nSL scores correspond to long homozygous haplotypes and a possible sweep in the high-altitude population compared to the low-altitude population, and vice versa for negative XP-nSL scores. A Manhattan plot of genome-wide normalized XP-nSL scores > 2 is plotted in Fig. S7.

In order to identify regions implicated as potentially adaptive, we search for clusters of extreme scores along a chromosome. Using `selscan`'s companion program `norm v1.3.0`, the genome is divided into non-overlapping 100kb regions and both the maximum XP-nSL score and the fraction of XP-nSL scores > 2 are computed (`norm` flags –xpnsl –bp-win –winsize 100000). `norm` then creates 10 quantile bins (--qbins 10) for windows with more than 10 sites per window (--min-snps 10) and identifies the top

281 1% of windows with the highest fraction of extreme scores (Fig. S6 and Table S4). Each
282 window is then annotated with the ensembl rheMac8 gene list, and a maximum XP-nSL
283 score is assigned to a given gene based on the max-score in the 100 kb window with
284 which it overlaps. If a gene overlaps more than one 100 kb window, it is assigned the
285 top max-score from among the windows.

286 ## Results

287 ### Power Analysis of XP-nSL

288 First, we evaluate the performance of XP-nSL based on simulations. After
289 computing XP-nSL for all sites in all simulations, scores were normalized by subtracting
290 the mean and dividing by standard deviation of the neutral simulations, giving the
291 neutral scores an approximately N(0,1) distribution (Fig. S1).

292 We consider the maximum score in a 100 kb interval as way to identify regions
293 under positive selection similar to Voight *et al.* (2006). To get the null distribution of
294 max-scores, the maximum score is computed in the central 100 kb of each neutral
295 simulation. The distribution of max-scores in neutral simulations had a median of 2.093
296 with 95% of the mass between 0.804 and 3.492, which is represented in Fig. 3 as a
297 solid horizontal black line (median) and two dashed horizontal black lines (95% interval).
298 Next, the maximum score is computed in the central 100 kb of all non-neutral
299 simulations, and the median and 95% intervals are plotted for each parameter
300 combination. Fig. 3 shows good separation between the neutral distribution of max-
301 scores and the distribution of max-scores for a range of non-neutral parameters,
302 suggesting that our statistic can distinguish between neutral and non-neutral scenarios.
303 Note that soft sweeps that start at 0.1 or 0.2 frequency see the least separation from the

304 neutral distribution. To evaluate the power of the max-score statistic, the 99th percentile
305 of the max-score distribution is computed in the neutral distribution ($neutral_{99} = 3.792$)
306 and power is calculated as the mass of each non-neutral max score distribution
307 $> neutral_{99}$. With this approach, even if the entire genome is neutrally evolving, at most
308 1% of the genome will be identified as putatively under selection thus fixing the false
309 positive rate at 1% at most. Results are plotted in Fig. 4A, which shows good power to
310 detect both incomplete and completed sweeps. For soft sweeps that start at a frequency
311 ≤ 0.03, power is $> 75\%$ when the sweep is near or just past fixation; the ability to
312 detect soft sweeps falls off for sweeps that start > 0.03 frequency.
313 Next, we consider that due to linkage disequilibrium consecutive scores will be
314 correlated, and we should therefore expect clusters of extreme scores in true non-
315 neutral regions. We thus consider a window-based approach to identify selected regions
316 similar to Voight *et al.* (2006), where the top 1% windows with a high number of extreme
317 scores are identified. Taking the central 100 kb region of each simulation, the fraction of
318 XP-nSL scores > 2 (representing approximately the highest 2% of all neutral scores) is
319 computed. Since each 100 kb window has a variable number of sites within it, windows
320 with fewer sites are more likely to have a higher fraction of extreme scores by chance.
321 Thus, windows are binned by number of sites into 10 quantile bins, and the top 1% of
322 windows with the highest fraction of extreme scores in each bin establishes the
323 threshold beyond which a window is taken as putatively selected, as in Voight *et al.*
324 (2006). With this approach, even if the entire genome is neutrally evolving, at most 1%
325 of the genome will be identified as putatively under selection thus fixing the false
326 positive rate at 1% at most. Power is computed for each non-neutral parameter set as

327 the proportion of replicates for which the central 100 kb exceeds the 1% threshold as
328 calculated from neutral simulations. The results are plotted in Fig. 4B, which shows
329 improved power over the max-score approach across a wider range of parameters.
330 Indeed, using the window-based method, power to detect soft sweeps improves
331 substantially across the parameter space, with $> 75\%$ power to detect soft sweeps at or
332 near fixation that started at frequency ≤ 0.05.

333 We next consider how XP-nSL power compares to nSL, XP-EHH, and F_{ST}. We
334 compare to nSL since XP-nSL is an extension of it, to XP-EHH as it is a similar
335 haplotype-based two-population selection statistic, and to F_{ST} as it a popular two-
336 population method used to infer local adaptation. For all simulations, nSL and XP-EHH
337 are computed using `selscan` v1.3.0 (Szpiech and Hernandez 2014). Normalization,
338 identification of top windows, and power calculation was done as described above for
339 XP-nSL. F_{ST} is computed using `VCFtools` v0.1.16 (Danecek *et al.* 2011), which
340 implements Weir and Cockerham's formulation (Weir and Cockerham 1984), in 100kb
341 windows for all simulations. The 99th percentile F_{ST} value was determined from the
342 central 100kb window among all neutral simulations. Power was then computed for
343 each non-neutral parameter set as the proportion of replicates for which the F_{ST} value of
344 the central 100kb window is greater than the neutral threshold as calculated from
345 neutral simulations.

346 Fig. 5 shows the difference in power between XP-nSL and each statistic (raw
347 power for nSL, XP-EHH, and F_{ST} shown in Fig. S4) over the simulated parameter
348 space, where positive values indicate that XP-nSL has more power than the comparison
349 statistic. XP-nSL is compared to nSL in Fig. 5A, which shows XP-nSL improves on nSL

350 across nearly the entire parameter space and especially for sweeps that have fixed in
351 the recent past. XP-nSL is compared to XP-EHH in Fig. 5B, which shows XP-nSL
352 improving on XP-EHH for soft sweeps ($e > 0$). Finally, XP-nSL is compared to F_{ST} in
353 Fig. 5C, which shows XP-nSL improving on F_{ST} mostly for incomplete sweeps, although
354 F_{ST} performed better post-fixation for certain soft sweep scenarios ($e \geq 0.05$).

355 **Caveats for Using Neutral Simulations to Normalize Real Data**

356 In principle, one could use matched neutral simulations as a normalization
357 baseline when analyzing real data. However, we can only recommend this approach
358 when the populations being studied have very well characterized (1) joint demographic
359 histories, (2) mutation rates, (3) and recombination rates, as a mismatch can skew the
360 power and false positive rates of the statistic. To illustrate this point, we generated three
361 mismatched sets of neutral simulations "Rand", "Under", and "Over" (see Methods) to
362 use as a normalization baseline for our original simulations.

363 When neutral simulations are normalized with the correct demographic history,
364 they approximately follow a standard normal distribution (Fig. S1 and Fig. S2), however
365 the distribution of neutral scores gets badly distorted when one of the mismatched
366 histories is used (Fig. S2). These distortions have practical consequences for making
367 inferences. Power was calculated for XP-nSL using the window-based method
368 described above but using each mismatched history as a normalization baseline (Fig.
369 S3). The false positive rate for each scenario was also estimated by calculating the
370 proportion of neutral simulations that are identified as under selection for each scenario
371 (Table S2). Fig. S3A-B shows that for the "Rand" and "Under" normalization scenarios,
372 power was greatly reduced across the whole parameter space. Only hard sweeps with

373 the strongest selection coefficients post-fixation were likely to be identified. False
374 positive rates for these scenarios were at 0 (Table S3), whereas for a matched history it
375 was estimated at 9.908×10^{-3}. Fig.S3C shows that, for the "Over" normalization
376 scenario, power was uniformly excellent, however when analyzing neutral simulations,
377 the false positive rate was estimated at 0.9538. This suggests that using such a badly
378 matched demographic history for normalization creates a near complete inability to
379 distinguish between neutral and selected regions.

380 **Identifying Evidence for Adaptation to High Altitude in Rhesus Macaques**

381 Next, we analyzed the pair of rhesus macaque populations using XP-nSL,
382 searching for evidence of local adaptation in the high-altitude population. Using the low-
383 altitude population as the reference population, normalized $XPnS_L > 0$ corresponds to
384 longer and higher frequency haplotypes in the high-altitude population, with very large
385 positive scores and clusters of large scores suggesting evidence for positive selection.
386 Dividing the genome into 100 kb windows, the maximum XP-nSL score of that region is
387 assigned to each gene in it (see Methods), thus multiple genes may have the same
388 max-score by virtue of being in the same 100kb window. Genes overlapping multiple
389 100kb windows were assigned the top max-score among the windows.

390 Using the per-gene max-scores, PANTHER (Mi *et al.* 2019) gene ontology
391 categories were tested for enrichment of high scores, where significance suggests an
392 enrichment of signals of positive selection among genes involved. Significant terms
393 related to regulation of ion transport and synaptic signaling (Table 2), each of which are
394 affected by hypoxic conditions (Karle *et al.* 2004; Corcoran and O'connor 2013).

395 From the genomic regions that were identified to contain a high proportion of
396 extreme positive scores (see Methods), 303 annotated genes were found across 270
397 regions. A permutation test (10,000 replicates) that randomly shuffles 270 100kb
398 regions around the genome indicates that this is substantially fewer than one would
399 expect by chance ($p = 1.4 \times 10^{-3}$; Fig. S5), indicating that the method is not simply
400 randomly picking gene regions. These regions, their characteristics, and the genes
401 contained therein are given in Table S2. A PANTHER (Mi *et al.* 2019) gene ontology
402 overrepresentation test indicates a 9.04-fold enrichment of genes associated with
403 monooxygenase activity ($FDR = 4.47 \times 10^{-2}$).
404 The monooxygenases in the selected regions include FMO2, FMO5, CYP2C8,
405 CYP2C9, CYP2C93, and ENSMMUT00000011129. These genes are important for the
406 metabolism of oxygen and the generation of reactive oxygen species (ROS) (Krueger
407 and Williams 2005). Under the physiological stress of a low-oxygen environment, ROS
408 levels increase and cause oxidative damage, and, in humans, long-term adaptation to
409 high altitudes includes adaptation to oxidative damage (Janocha *et al.* 2017). Indeed,
410 AOX1 is also identified in our top regions, mutations in which have been shown to affect
411 ROS levels in humans (Foti *et al.* 2017).
412 The genome-wide top ten 100 kb windows based on the percentage of extreme
413 XP-nSL scores are summarized in Table 3, and these windows overlap several genes,
414 including EGLN1. The EGLN1 locus is directly adjacent to the single strongest selection
415 signal identified in the entire genome (Fig. 6). This region has the third highest cluster of
416 extreme scores (Table 3), contains the highest XP-nSL score in the entire genome
417 (chr1:207,698,003, $XPnS_L = 6.54809$), and contains six of the top ten genome wide XP-

418 nSL scores (colored dark red in Fig. 6). EGLN1 is a regulator of oxygen homeostasis
419 (To and Huang 2005) and is a classic target for adaptation to low-oxygen levels, having
420 repeatedly been the target of adaptation in numerous organisms living at high altitude
421 around the world (Bigham *et al.* 2009; Bigham *et al.* 2010; Jeong *et al.* 2014; Graham
422 and Mccracken 2019). In addition to EGLN1, other genes related to lung function,
423 oxygen use, and angiogenesis had evidence for local adaptation between low- and
424 high-altitude populations: TRPM7, RBPJ, and ENSMMUT00000040566 (Table S2).
425 TRPM7 downregulation in a hypoxia-induced rat model was associated with pulmonary
426 hypertension (PAH) (Xing *et al.* 2019). ENSMMUT00000040566 is a MAPK6 ortholog,
427 which interacts with EGLN3 (Rodriguez *et al.* 2016), and both it and RBPJ are involved
428 in angiogenesis (Ramasamy *et al.* 2014).
429 Due to the reduced oxygen levels at high altitudes, we expect genes involved in
430 metabolism and respiration may be under positive selection. Indeed, MDH1 encodes a
431 critical enzyme in the citrate cycle (Tanaka *et al.* 1996) and is found in the top ten
432 genome-wide regions (Table 3). A paralog of MDH1, MDH1B, has been previously
433 identified as a target of selection in humans living at high altitude (Yi *et al.* 2010).
434 ACADM and COX15 are also found in putatively adaptive regions (Table S2). Mutations
435 in and differing expression levels of ACADM are related to oxidative stress and
436 mitochondrial dysfunction in human disease (Xu *et al.* 2018). COX15 is involved in
437 oxidative phosphorylation (Alston *et al.* 2017), and cytochrome c oxidase (COX) genes
438 have previously been identified as under selection in primates relative to other
439 mammals (Osada and Akashi 2012).

440 High-altitude environments present a particular metabolic challenge to organisms
441 that must maintain a stable internal body temperature (Rosenmann and Morrison 1974;
442 Hayes and Chappell 1986; Chappell and Hammond 2004), a result of increased oxygen
443 demand from aerobic thermogenesis conflicting with lower oxygen availability. It has
444 been shown that highland deer mice have adapted by increased capacity to metabolize
445 lipids compared to lowland deer mice (Cheviron *et al.* 2012), and previous studies in
446 rhesus macaques have shown there may be drastic differences in diets between high-
447 and low-altitude populations (Zhao 2018). Indeed, in the high-altitude macaque
448 population studied here, genes related to lipid and fat metabolism (DOCK7,
449 ST6GALNAC5, ANGPTL3, and ACACA) were found in putative adaptive regions.
450 Across human populations, these genes are all responsible for varying blood levels of
451 fatty acids (Guo *et al.* 2016; Dewey *et al.* 2017; Hebbar *et al.* 2018). Furthermore,
452 ACACA has been shown to vary fatty acid blood concentrations and be differentially
453 expressed in highland versus lowland swine populations (Shang *et al.* 2019).
454 STXBP5L also appears in the top ten regions (Table 3) and is involved in
455 vesicular trafficking and neurotransmitter release (Kumar *et al.* 2015). As primate brains
456 use large amounts of oxygen and energy to function (Osada and Akashi 2012)
457 signatures of selection on neurological genes may be expected across populations
458 living at altitudes with differing oxygen levels. In addition to STXBP5L, several genes
459 related to neural development and synaptic formation (JAG2, TRPM7, DOCK7, NSG2,
460 AUTS2) were identified (Table S2). JAG2 is involved in the Notch signaling pathway.
461 While Notch signaling is involved in many developmental and homeostatic processes,
462 its role in neuronal differentiation in the mammalian brain is notable in this context

463 (Cardenas *et al.* 2018). TRPM7, in addition to its association with PAH, plays a role in
464 hypoxic neuronal cell death (Aarts *et al.* 2003).

465 DNA damage, including double strand breaks and pyrimidine dimerization, can
466 manifest as a result of oxidative stress (Ye *et al.* 2016) or increased exposure to UV
467 radiation (Zhang *et al.* 2000; Greinert *et al.* 2012), both of which increase at high
468 altitudes. Three DNA damage repair genes are in our set of genes identified in positively
469 selected regions. The ring finger protein RNF138 is in our list of top ten genomic
470 windows (Table 3) and has been shown to promote DNA double-strand break repair
471 (Ismail *et al.* 2015). Furthermore, the DNA polymerase POLH also appears in a
472 putatively adaptive region (Table S2) and is known to be able to efficiently bypass
473 pyrimidine dimer lesions (Zhang *et al.* 2000). Finally, PAXIP1 appears in Table S2 as
474 well and has been shown to promote repair of double strand breaks through
475 homologous recombination (Wang *et al.* 2010).

476 Interestingly, within the putatively selected region that includes PAXIP1 is an
477 uncharacterized long non-coding RNA (lncRNA), ENSMMUT00000081951, that was
478 recently annotated in the rheMac10 genome build. This lncRNA has high sequence
479 similarity to a lncRNA on the same synteny block in humans called PAXIP1-AS1. In
480 human pulmonary artery smooth muscle cells, knockdown of PAXIP1-AS1 leads to an
481 abnormal response to PAH where migration and proliferation of cells is reduced, and
482 overexpression of PAXIP-AS1 leads to apoptosis resistance (Jandl *et al.* 2019).
483 Although, we note that any link between PAH and ENSMMUT00000081951 in rhesus
484 macaques is highly speculative at this point.

485 **Discussion**

486 When populations split and migrate, they may adapt in different ways in response
487 to their local environments. Genetic adaptations that arise and sweep through the
488 population leave a characteristic genomic pattern of long haplotypes of low diversity and
489 high frequency. We develop a two-population haplotype-based statistic, XP-nSL, to
490 capture these patterns. With good power to detect hard and soft sweeps that occur in
491 one population but not another, XP-nSL can identify positively selected regions of the
492 genome likely the result of local adaptation. We apply this statistic to genomes sampled
493 from a pair of wild-born populations of Chinese rhesus macaques, inferred to have
494 diverged approximately 9500 generations ago, one of which lives at high altitude in far
495 western Sichuan province, and the other that lives close to sea level.
496 Life at high altitude presents extreme biological challenges, including low
497 atmospheric oxygen, increased UV exposure, harsh winters, and reduced nutrition
498 availability, which create strong selection pressure. Organisms that survive and persist
499 are likely to be carrying genetic mutations that confer an advantage for living in such
500 harsh environments. Common targets for adaptation to such an environment include
501 genes related to hypoxia, regulation of ROS, DNA damage repair, and metabolism
502 (Cheviron and Brumfield 2012; Witt and Huerta-Sanchez 2019; Storz and Cheviron
503 2021). Indeed, in the high-altitude macaque population, we identify a strong signal of
504 positive selection at the EGLN1 locus (Fig. 6), a classic target for adaptation to low-
505 oxygen environments, in addition to 302 other genes, many of which are related to the
506 myriad environmental selection pressures expected in high-altitude environments. As
507 has been suggested previously for other organisms (Cheviron and Brumfield 2012;
508 Bigham and Lee 2014; Simonson 2015; Witt and Huerta-Sanchez 2019; Storz and

Cheviron 2021), these results suggest that, rather than a single adaptive mutation at a single locus, adaptation to this extreme environment by rhesus macaques is polygenic and acts through multiple biological systems.

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Author Contributions

ZAS and LSS conceived of the study. ZAS performed all simulations and genomic analyses and implemented novel statistics. ZAS, TEN, and NPB characterized gene functions and ontologies with contributions from LSS. ZAS wrote the manuscript with contributions from NPB, TEN, and LSS. All authors read and approved of the manuscript.

Data Accessibility

Macaque whole genome VCFs are available at http://dx.doi.org/10.5524/100484. Selection scan data available at https://doi.org/10.5061/dryad.kkwh70s40.

Tables

Table 1. Demographic parameters used for simulations with 95% confidence intervals from (Liu *et al.* 2018). T values are given in number of generations before present. N values represent diploid effective population size.

Parameter	Value	95% CI
T_1	4660	$(651, 5034)$
T_2	5605	$(3962, 10468)$
T_3	9468	$(4563, 14047)$
N_1	16188	$(8255, 26825)$
N_2	3730	$(660, 5006)$
N_3	4506	$(572, 124135)$
N_4	20768	$(1710, 88242)$
N_5	66210	$(20581, 301430)$

Table 2. Gene ontology enrichment analysis results based on maximum XP-nSL scores per gene. Significant GO terms are enriched for high XP-nSL scores.

PANTHER GO-Slim Biological Process	Gene Ontology ID	p-value	FDR
anterograde trans-synaptic signaling	GO:0098916	4.43×10^{-5}	1.52×10^{-2}
chemical synaptic transmission	GO:0007268	4.43×10^{-5}	1.83×10^{-2}
regulation of transport	GO:0051049	3.60×10^{-5}	1.85×10^{-2}
trans-synaptic signaling	GO:0099537	6.32×10^{-5}	1.86×10^{-2}
synaptic signaling	GO:0099536	1.03×10^{-4}	2.66×10^{-2}
ion transport	GO:0006811	1.33×10^{-4}	2.75×10^{-2}
transmembrane transport	GO:0055085	1.57×10^{-4}	2.94×10^{-2}
regulation of ion transport	GO:0043269	1.30×10^{-4}	2.97×10^{-2}
metal ion transport	GO:0030001	2.21×10^{-4}	3.04×10^{-2}
regulation of localization	GO:0032879	2.13×10^{-4}	3.14×10^{-2}
regulation of ion transmembrane transport	GO:0034765	2.66×10^{-4}	3.23×10^{-2}
inorganic ion transmembrane transport	GO:0098660	3.01×10^{-4}	3.45×10^{-2}
sodium ion transport	GO:0006814	3.36×10^{-4}	3.65×10^{-2}
ion transmembrane transport	GO:0034220	4.50×10^{-4}	4.42×10^{-2}
regulation of transmembrane transport	GO:0034762	4.40×10^{-4}	4.54×10^{-2}
cation transport	GO:0006812	4.87×10^{-4}	4.57×10^{-2}

Table 3. The top ten 100 kb genomic regions as ranked by percentage of scores greater than 2. Concatenated region represents the genomic region merged with adjacent top 1% regions. Max score represents the max XP-nSL score in the concatenated region. Genes gives all genes overlapping the concatenated region.

Genomic Region	% > 2	Flanking Regions	Concatenated Region	Max Score	Genes
*chr2:42300001-42400000	81.62%	3	chr2:42200001-42600000	5.42043	STXBP5L
chr3:99100001-99200000	78.18%	3	chr3:99100001-99500000	5.54282	-
chr1:207600001-207700000	***77.39%***	***1***	***chr1:207600001-207800000***	***6.54809†***	***EGLN1, TSNAX***
chr10:69700001-69800000	76.60%	4	chr10:69700001-70200000	4.89388	PITPNB, ENSMMUT00000079195.1, TTC28
chr13:64300001-64400000	74.01%	1	chr13:64300001-64500000	4.17728	WDPCP, MDH1, ENSMMUT00000070474.1
chr7:96400001-96500000	73.86%	1	chr7:96400001-96600000	4.58886	FAM177A1, PPP2R3C, ENSMMUT00000039911.2, ENSMMUT00000027928.3
chr10:70400001-70500000	70.24%	0	chr10:70400001-70500000	3.69192	TTC28
*chr2:42400001-42500000	65.86%	3	chr2:42200001-42600000	5.42043	STXBP5L
chr7:120400001-120500000	62.54%	0	chr7:120400001-120500000	5.86114	KIAA0586, ENSMMUT00000059161.1
chr18:20600001-20700000	62.42%	0	chr18:20600001-20700000	5.37386	RNF138

***These regions are contained in the same concatenated region.**
†Top genome-wide score. This region contains 6 of the top 10 genome-wide scores.

Table S1. Population classification by individual ID.
See TableS1.xlsx

Table S2. List of genomic windows with the top 1% highest fraction of extreme XP-nSL scores.
See TableS2.xlsx

Table S3. Estimates of false positive rates for various demographic history normalization scenarios.

Demographic History	Estimated False Positive Rate
Matched	9.908×10^{-3}
"Rand"	0
"Under"	0
"Over"	0.9538

Table S4. Bin boundaries (# of sites), number of windows per bin, and top 1% thresholds for the XP-nSL analysis of rhesus macaques.

Bin Boundaries (min # sites, max # sites)	# Windows in Bin	Top 1% Threshold (Fraction of Scores > 2)
(11, 731)	2672	0.346755
(732, 987)	2680	0.262945
(988, 1155)	2676	0.207089
(1156, 1285)	2675	0.196678
(1286, 1390)	2663	0.226605
(1391, 1484)	2687	0.19274
(1485, 1576)	2665	0.163597
(1577, 1681)	2648	0.169275
(1682, 1824)	2666	0.164025
(1825, 5766)	2666	0.153176

Figures

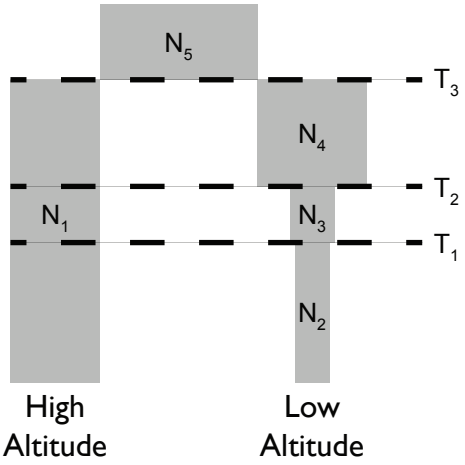
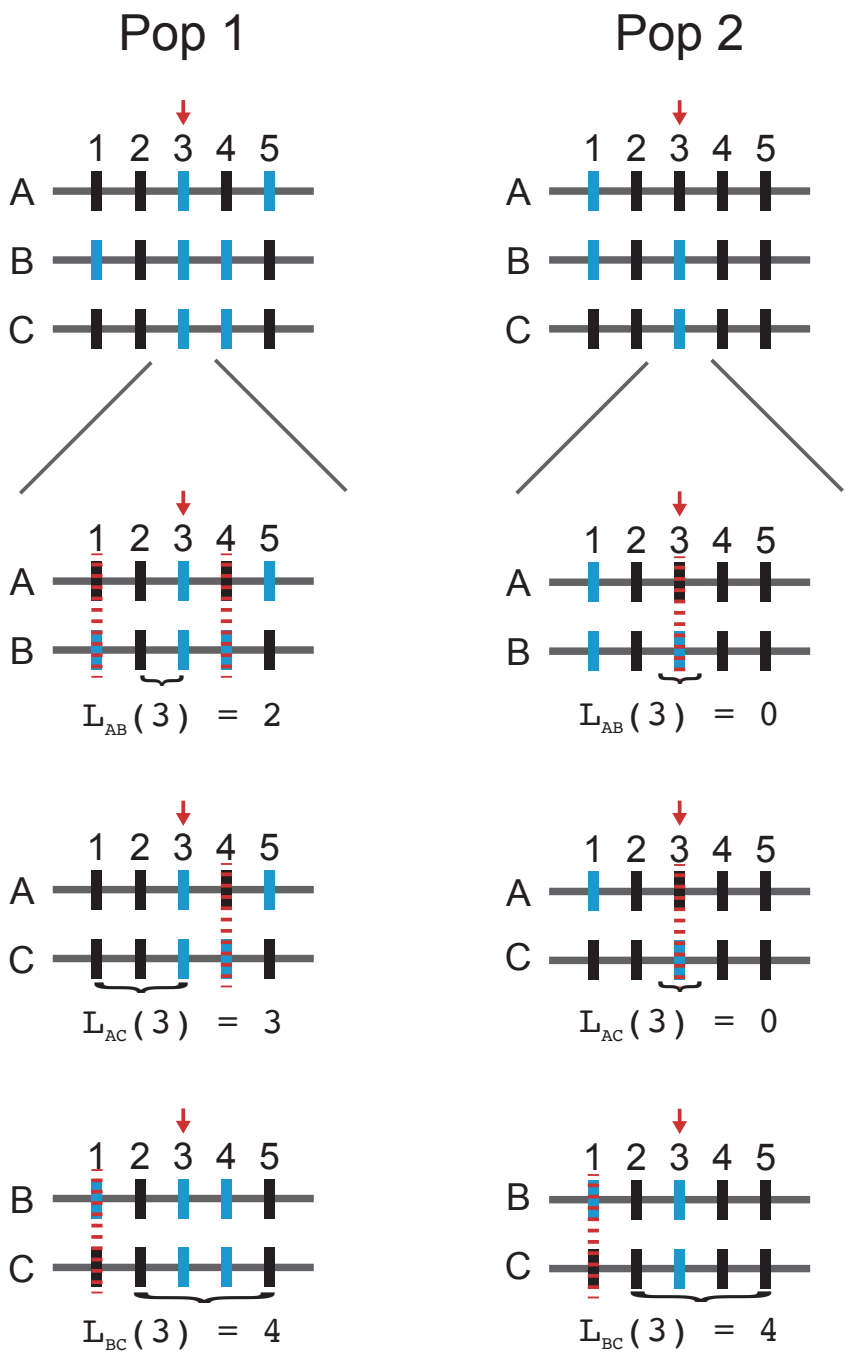


Figure 1. A representation of the demographic history for our high- and low-altitude populations as inferred by Liu *et al.* (2018).



$$SL_{P1}(3) = (2+3+4)/3 = 3 \qquad SL_{P2}(3) = (0+0+4)/3 = 4/3$$

$$XPnSL(3) = \log(SL_{P1}(3)/SL_{P2}(3)) = \log(9/4) = 0.3522$$

Figure 2. A toy example illustrating the computation of XP-nSL at a single site in two populations with three haplotypes (grey horizontal bars labeled A-C) and five sites (vertical bars labeled 1-5) where different alleles are colored blue or black. XP-nSL is calculated at site 3 (marked by red arrow). In each population, for each pair of haplotypes, the number of identical-by-state (IBS) sites are counted extending out from and including the test site (red arrow) until reaching a non-IBS site (marked by red dotted line). Within each population, the mean number of IBS sites is calculated across all pairs of haplotypes, and then the log-ratio of the mean from each population is computed to get XP-nSL at site 3.

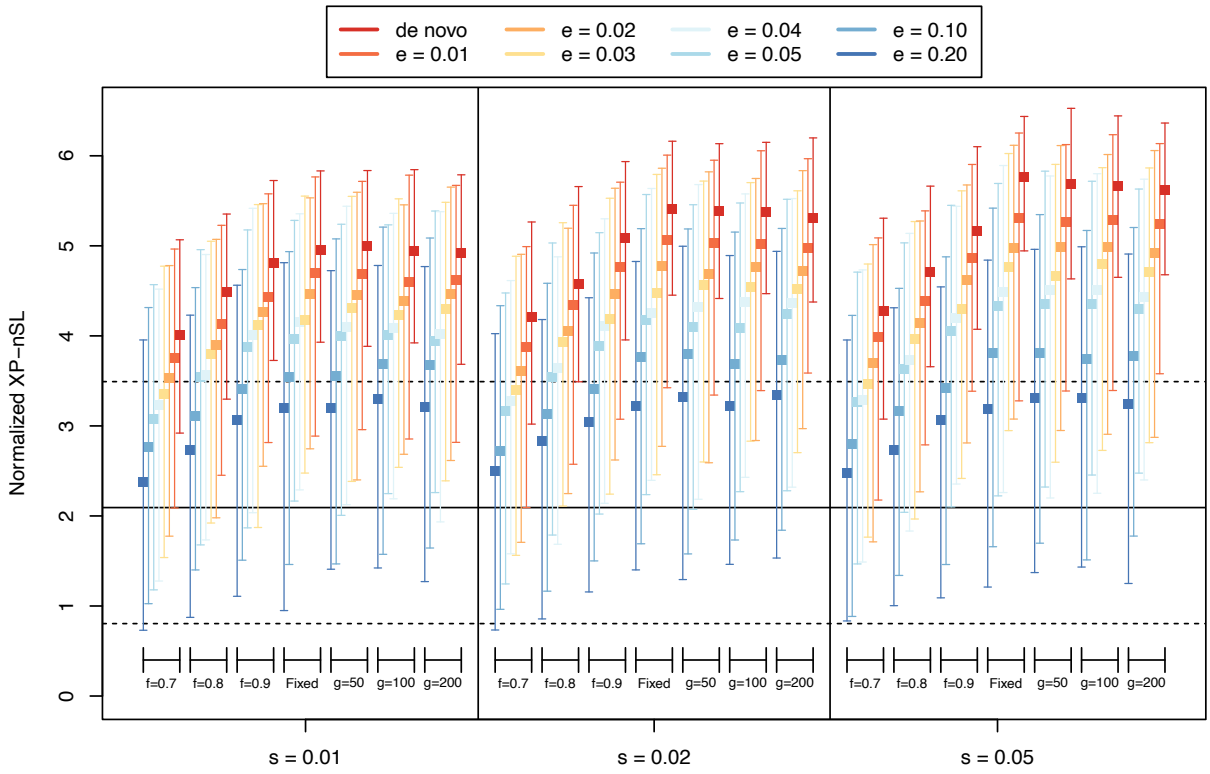


Figure 3. The distribution of maximum XP-nSL scores from simulations across various parameters, represented by medians and intervals containing 95% of the mass of the distribution. Neutral simulations are represented by the black solid horizonal line (median) and the black dashed horizonal line (95% interval). Non-neutral simulations represented by a colored box (median) and error bars (95% interval). The parameters are e (frequency at which selection begins, e > 0 indicates soft sweep), f (frequency of selected mutation at sampling), g (number of generations since fixation), and s (selection coefficient).

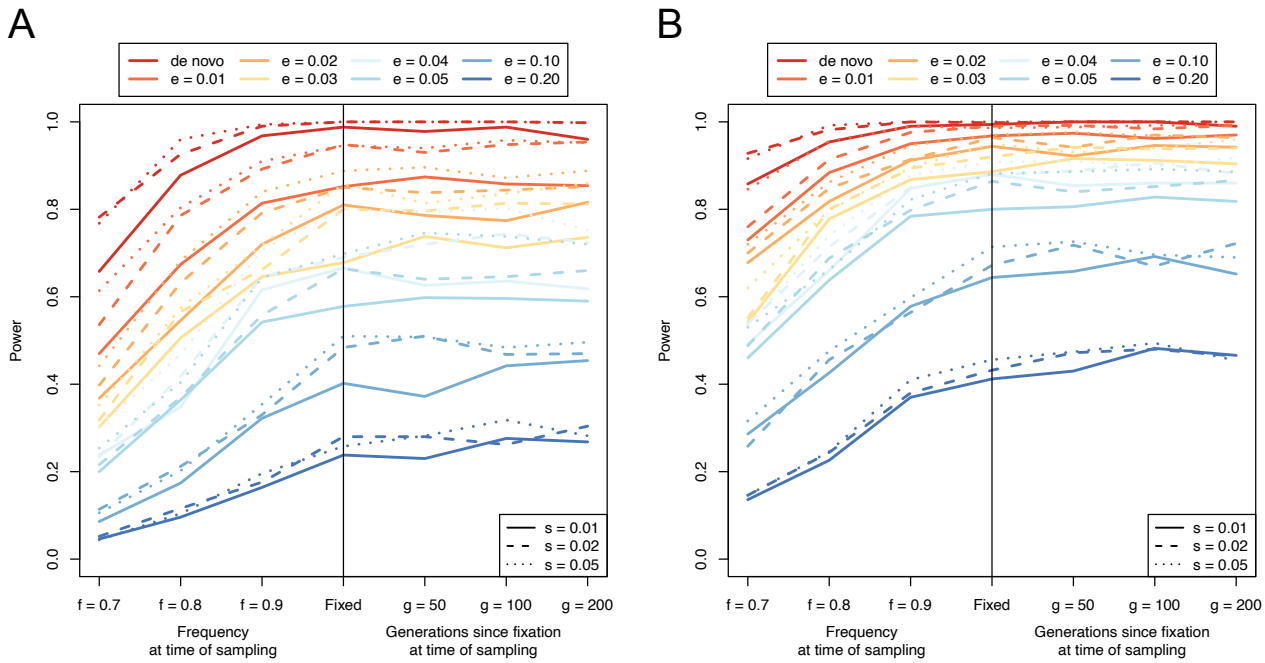


Figure 4. Power curves for (A) the max-score approach and (B) the window-based approach to identifying sweeps. The parameters are e (frequency at which selection begins, e > 0 indicates soft sweep), f (frequency of selected mutation at sampling), g (number of generations since fixation), and s (selection coefficient).

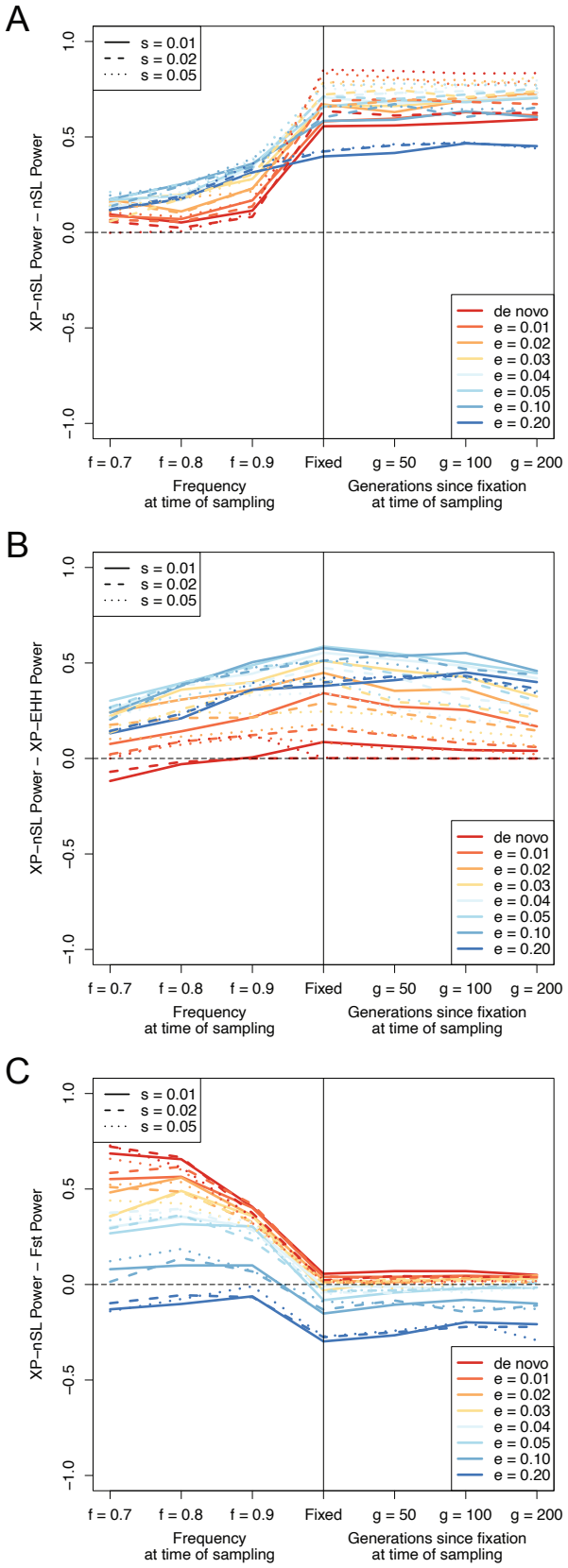

A
s = 0.01
s = 0.02
s = 0.05
de novo
e = 0.01
e = 0.02
e = 0.03
e = 0.04
e = 0.05
e = 0.10
e = 0.20
XP−nSL Power − nSL Power
1.0
0.5
0.0
−0.5
−1.0
f = 0.7
f = 0.8
f = 0.9
Fixed
g = 50
g = 100
g = 200
Frequency
at time of sampling
Generations since fixation
at time of sampling
B
XP−nSL Power − XP−EHH Power
C
XP−nSL Power − Fst Power

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Figure 5. Difference in power between XP-nSL and (A) nSL, (B) XP-EHH, and (C) F_{ST}. Values above 0 indicate XP-nSL has more power, and values below 0 indicate XP-nSL has less power. The horizontal black dotted line marks 0. The parameters are e (frequency at which selection begins, e > 0 indicates soft sweep), f (frequency of selected mutation at sampling), g (number of generations since fixation), and s (selection coefficient).

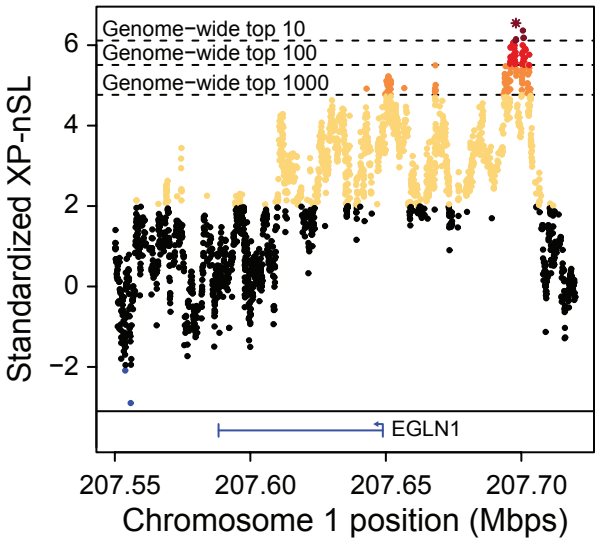


Figure 6. XP-nSL scores in the vicinity of the EGLN1 locus. This locus contains the genome wide top score (star) and six of the top ten genome wide scores (dark red).

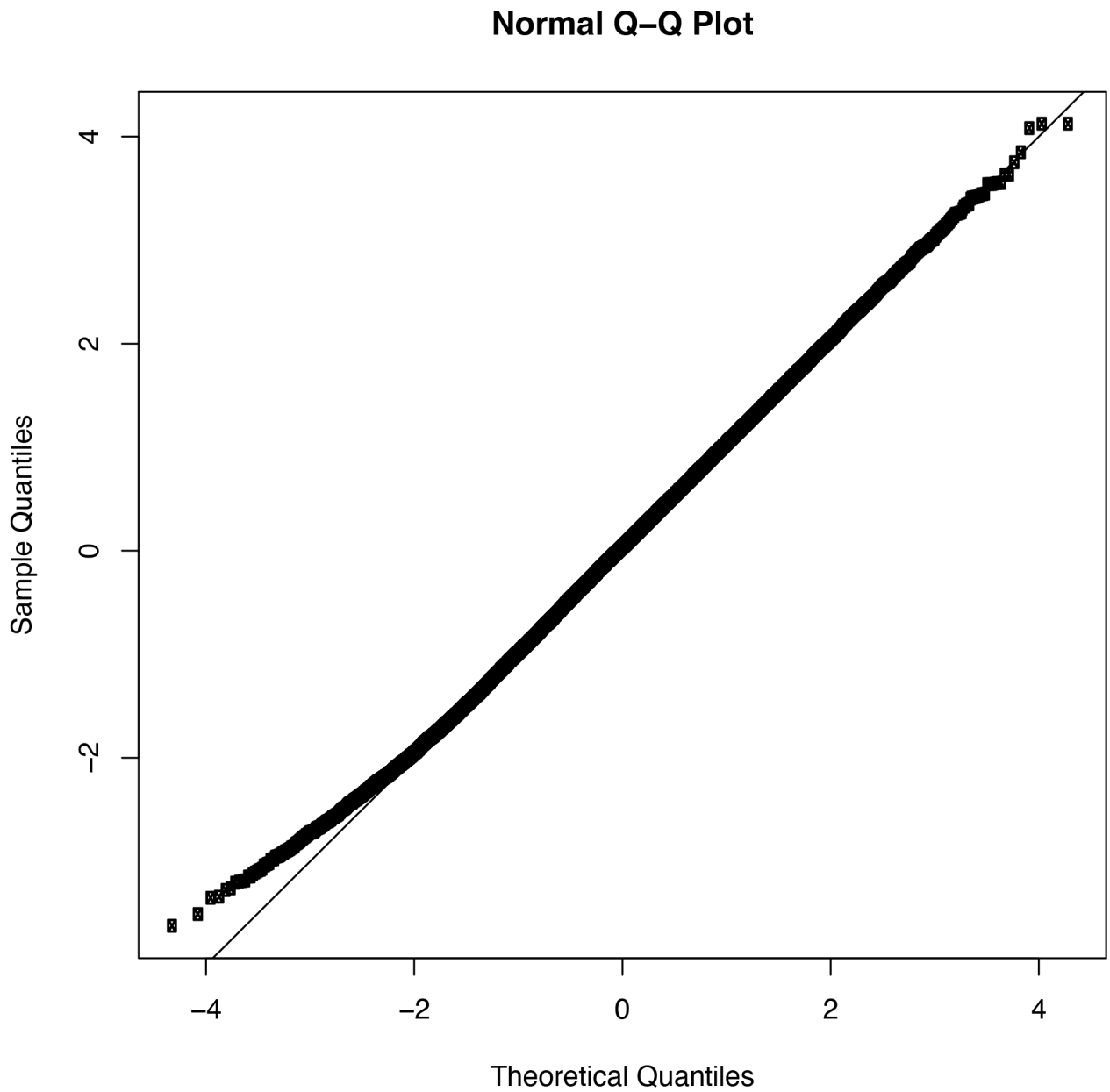


Figure S1. A normal quantile-quantile plot of neutral XP-nSL scores showing generally good adherence to a standard normal distribution. Due to autocorrelation along the genome, only every 1000th score is plotted.

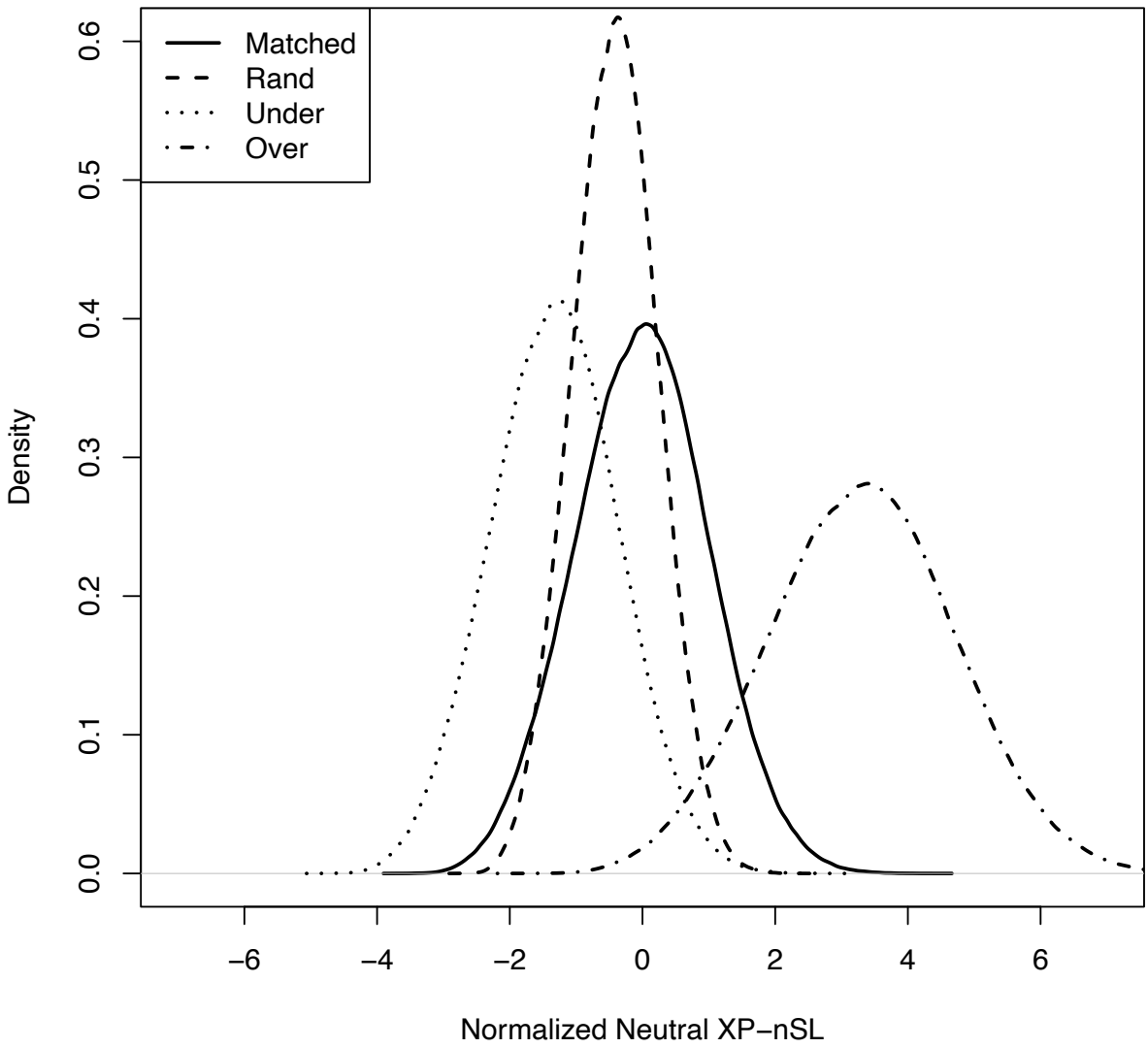
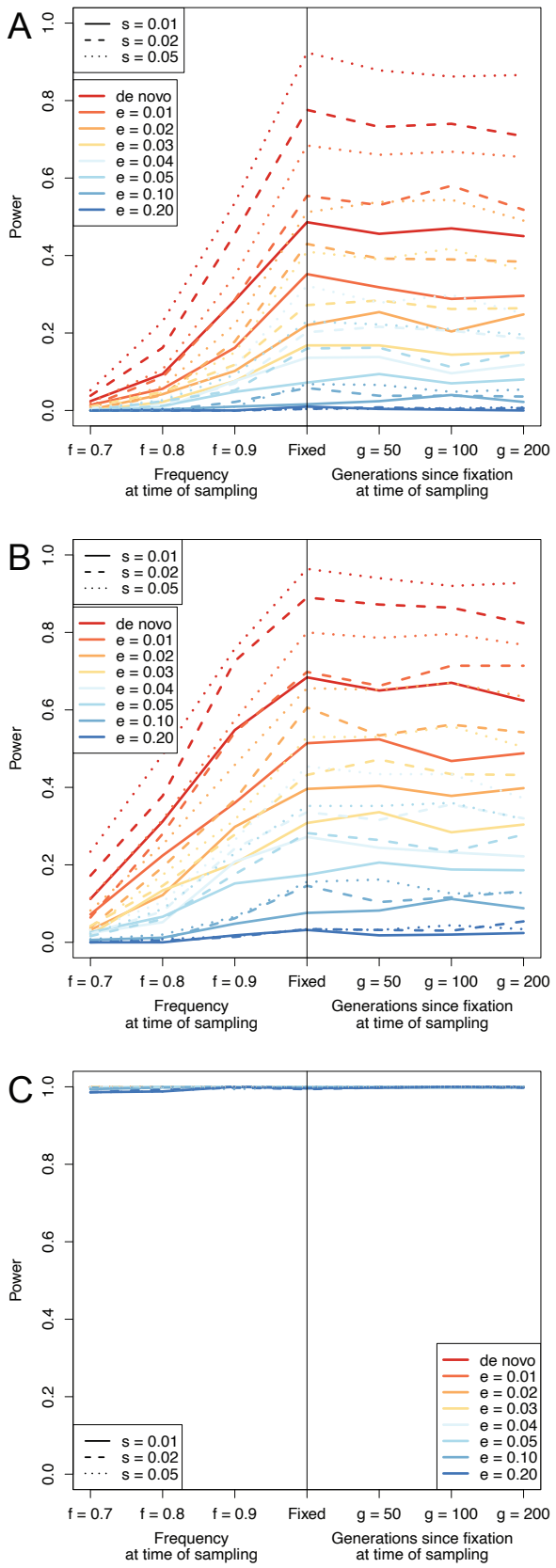


Figure S2. The distribution of neutral XP-nSL scores normalized with a matched demographic history (solid line), normalized with the “Rand” demographic history (dashed line), normalized with the “Under” demographic history (dotted line), and normalized with the “Over” demographic history (dash-dot line). Normalizing with the wrong demographic history can dramatically shift the distribution of neutral XP-nSL scores.


A
B
C
Power
0.0
0.2
0.4
0.6
0.8
1.0
s = 0.01
s = 0.02
s = 0.05
de novo
e = 0.01
e = 0.02
e = 0.03
e = 0.04
e = 0.05
e = 0.10
e = 0.20
f = 0.7
f = 0.8
f = 0.9
Fixed
g = 50
g = 100
g = 200
Frequency
at time of sampling
Generations since fixation
at time of sampling

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707 **Figure S3.** XP-nSL power using mismatched demographic histories for normalization.
708 (A) Using the “Rand” history. (B) Using the “Under” history. (C) Using the “Over” history.
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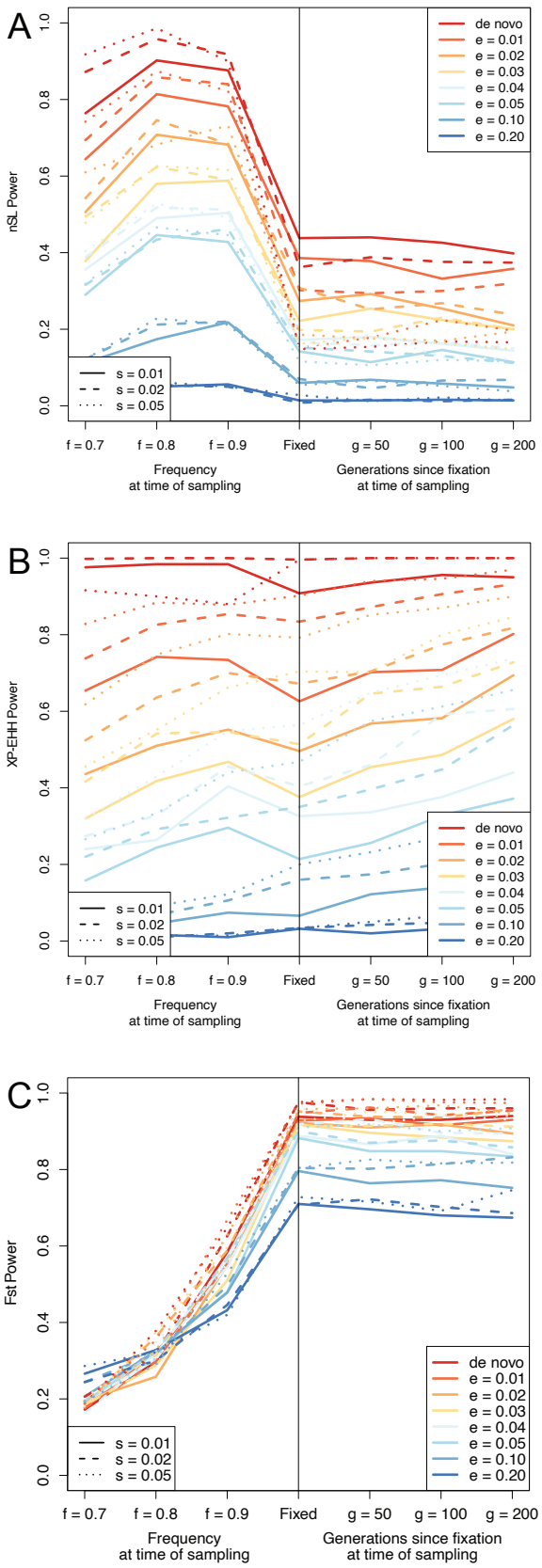
A
nSL Power
0.0
0.2
0.4
0.6
0.8
1.0
de novo
e = 0.01
e = 0.02
e = 0.03
e = 0.04
e = 0.05
e = 0.10
e = 0.20
s = 0.01
s = 0.02
s = 0.05
f = 0.7
f = 0.8
f = 0.9
Fixed
g = 50
g = 100
g = 200
Frequency
at time of sampling
Generations since fixation
at time of sampling
B
XP-EHH Power
0.0
0.2
0.4
0.6
0.8
1.0
de novo
e = 0.01
e = 0.02
e = 0.03
e = 0.04
e = 0.05
e = 0.10
e = 0.20
s = 0.01
s = 0.02
s = 0.05
f = 0.7
f = 0.8
f = 0.9
Fixed
g = 50
g = 100
g = 200
Frequency
at time of sampling
Generations since fixation
at time of sampling
C
Fst Power
0.0
0.2
0.4
0.6
0.8
1.0
de novo
e = 0.01
e = 0.02
e = 0.03
e = 0.04
e = 0.05
e = 0.10
e = 0.20
s = 0.01
s = 0.02
s = 0.05
f = 0.7
f = 0.8
f = 0.9
Fixed
g = 50
g = 100
g = 200
Frequency
at time of sampling
Generations since fixation
at time of sampling

710

711 **Figure S4.** Power curves for (A) nSL, (B) XP-EHH, and (C) F_{ST}. The parameters are e
712 (frequency at which selection begins, e > 0 indicates soft sweep), f (frequency of
713 selected mutation at sampling), g (number of generations since fixation), and s
714 (selection coefficient).
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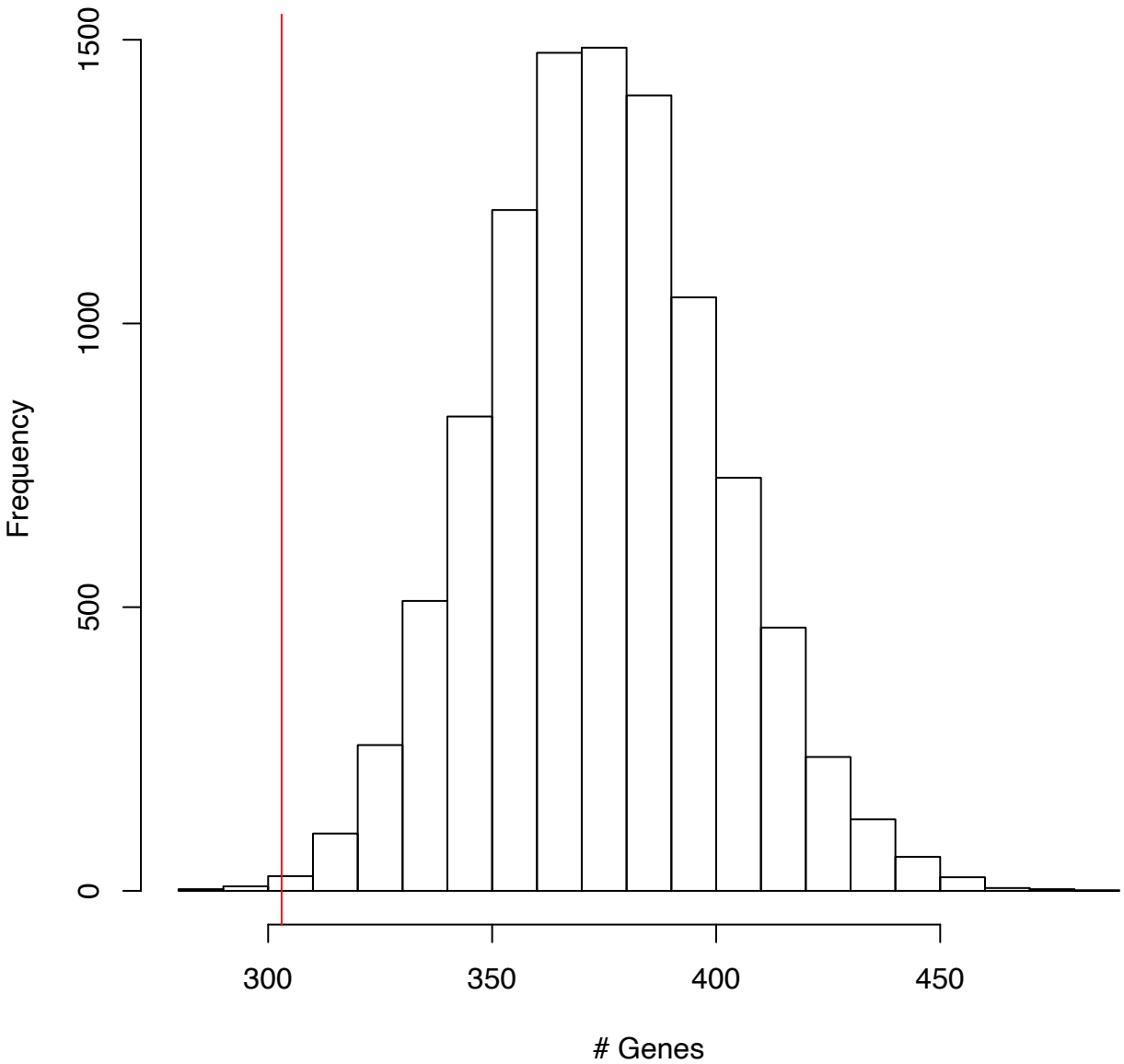

Figure S5. A permutation test (10,000 replicates) that shuffles 270 100kb regions around the macaque genome and counts the number of unique genes overlapping. The red vertical line marks the 303 genes found in the real data analysis. The probability of observing 303 or fewer genes is 1.4×10^{-3}, indicating the analysis is not randomly choosing gene regions.

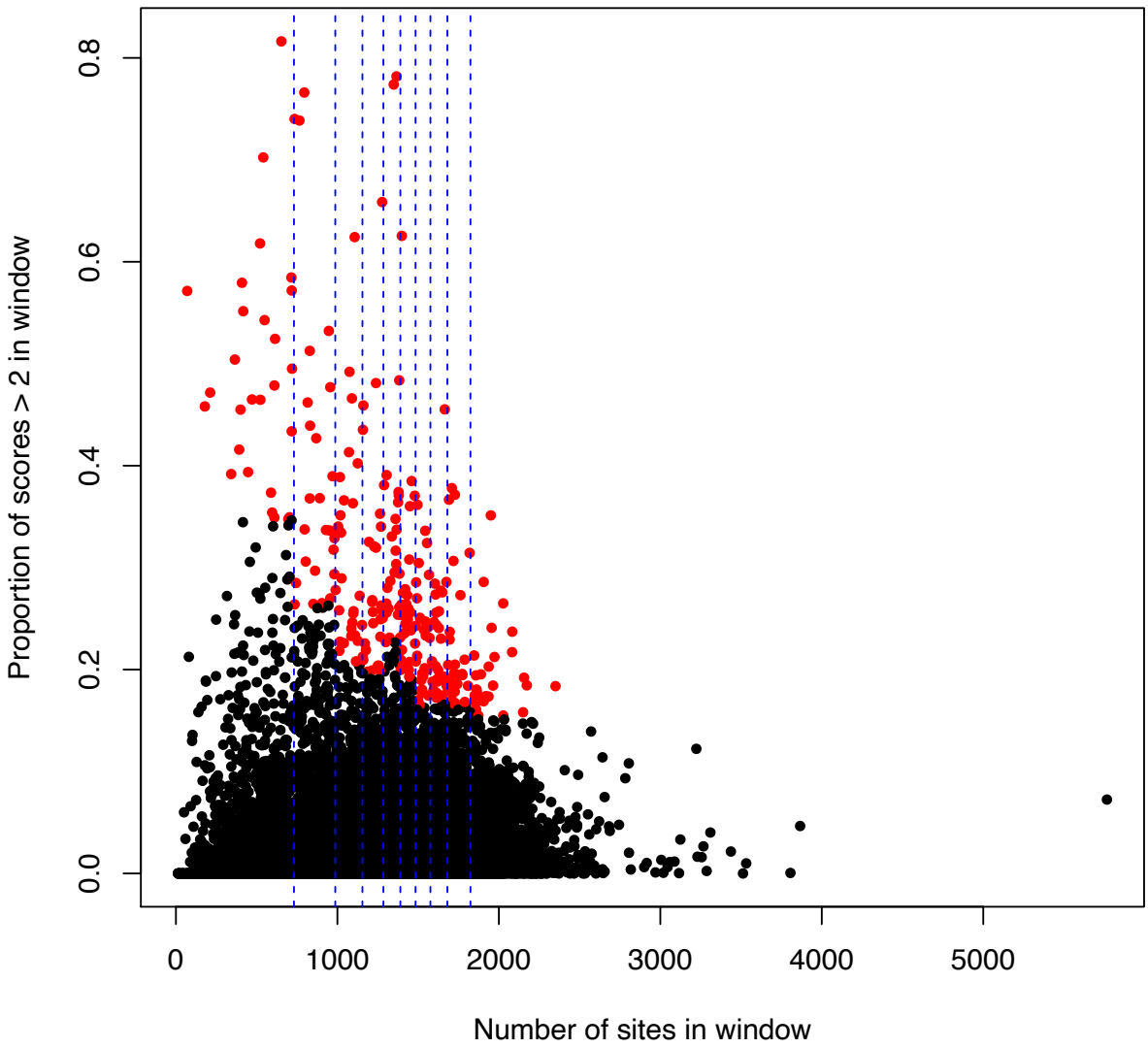

Figure S6. Proportion of scores > 2 versus number of sites in window. Blue vertical dashed lines indicate bin boundaries. Each circle is a window, red dots indicate a proportion of scores > 2 beyond the 1% threshold for that bin.

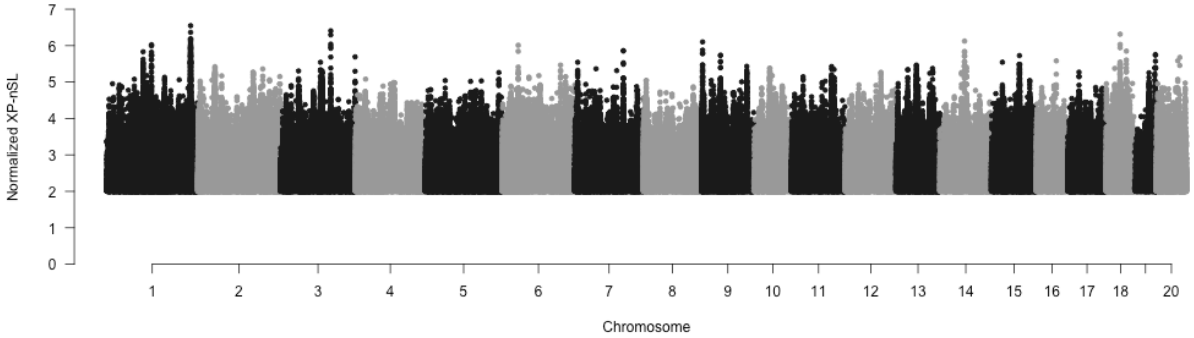

Figure S7. A Manhattan plot of normalized XP-nSL scores across the genome. Due to a very large number of points, only scores > 2 were plotted.

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791

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