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24

## 25 **ABSTRACT**

26 The genomic revolution offers renewed hope of resolving rapid radiations in the  
27 Tree of Life. The development of the multispecies coalescent (MSC) model and improved  
28 gene tree estimation methods can better accommodate gene tree heterogeneity caused by  
29 incomplete lineage sorting (ILS) and gene tree estimation error stemming from the short  
30 internal branches. However, the relative influence of these factors in species tree inference  
31 is not well understood. Using anchored hybrid enrichment, we generated a data set  
32 including 423 single-copy loci from 64 taxa representing 39 families to infer the species  
33 tree of the flowering plant order Malpighiales. This order alone includes nine of the top ten  
34 most unstable nodes in angiosperms, and the recalcitrant relationships along the backbone  
35 of the order have been hypothesized to arise from the rapid radiation during the  
36 Cretaceous. Here, we show that coalescent-based methods do not resolve the backbone of  
37 Malpighiales and concatenation methods yield inconsistent estimations, providing  
38 evidence that gene tree heterogeneity is high in this clade. Despite high levels of ILS and  
39 gene tree estimation error, our simulations demonstrate that these two factors alone are  
40 insufficient to explain the lack of resolution in this order. To explore this further, we  
41 examined triplet frequencies among empirical gene trees and discovered some of them  
42 deviated significantly from those attributed to ILS and estimation error, suggesting gene  
43 flow as an additional and previously unappreciated phenomenon promoting gene tree  
44 variation in Malpighiales. Finally, we applied a novel method to quantify the relative

45 contribution of these three primary sources of gene tree heterogeneity and demonstrated  
46 that ILS, gene tree estimation error, and gene flow contributed to 15%, 52%, and 32% of  
47 the variation, respectively. Together, our results suggest that a perfect storm of factors  
48 likely influence this lack of resolution, and further indicate that recalcitrant phylogenetic  
49 relationships like the backbone of Malpighiales may be better represented as phylogenetic  
50 networks. Thus, reducing such groups solely to existing models that adhere strictly to  
51 bifurcating trees greatly oversimplifies reality, and obscures our ability to more clearly  
52 discern the process of evolution.

53

54 Keywords: rapid radiation, triplet frequency, concatenation, coalescent, phylogenomics,  
55 hybrid enrichment, flanking region

56

## 57 **INTRODUCTION**

58 One of the most difficult challenges in systematics is reconstructing evolutionary  
59 history during periods of rapid radiation. During such intervals, few DNA substitutions  
60 accrue, rendering little information for phylogenetic inference. The potentially large  
61 population sizes and close evolutionary relationships create opportunities for widespread  
62 incomplete lineage sorting (ILS) and gene flow, leading to excessive gene tree-species tree  
63 conflict. The tremendous growth of genome-scale data sets, however, has greatly improved  
64 researchers' ability to investigate rapid radiations by providing hundreds to thousands of  
65 unlinked loci. Commonly applied approaches include not only whole genome sequencing  
66 but also RNA-Seq, RAD-Seq, and anchored hybrid enrichment, which in general are cost  
67 effective and efficient across broad taxonomic groups and yield data sets with dense locus

68 and taxon sampling (Lemmon and Lemmon 2013). These approaches are promising and  
69 have been variously applied to successfully resolve a number of recalcitrant clades across  
70 the Tree of Life, including in birds (Prum et al. 2015), mammals (Song et al. 2012), fish  
71 (Wagner et al. 2013), and plants (Wickett et al. 2014).

72 Despite their promise, however, these enormous data sets also introduce new  
73 methodological challenges and complexities. In particular, phylogenomic data sets may  
74 yield strongly supported, yet conflicting or artifactual, results depending on the method of  
75 inference or genomic regions sampled (Song et al. 2012; Jarvis et al. 2014; Xi et al. 2014;  
76 Reddy et al. 2017; Shen et al. 2017). During rapid radiations, ILS can lead to extreme  
77 conditions where the most probable gene tree differs from the topology of the true species  
78 tree, which is referred to as the “anomaly zone” (Degnan and Rosenberg 2006; Rosenberg  
79 and Tao 2008). Such pervasive genealogical discordance, in particular, can result in biased  
80 species tree inference when applying concatenation methods, and produce inconsistent  
81 and conflicting results with strong confidence (Song et al. 2012; Xi et al. 2014). The  
82 multispecies coalescent (MSC) model, which explicitly accommodates gene tree  
83 heterogeneity caused by ILS, in contrast, has been demonstrated to be more reliable under  
84 these circumstances. Most recently, a class of “two-step” summary coalescent methods has  
85 been the focus of substantial development and application (Nakhleh 2013). They are  
86 demonstrated to be statistically consistent under the MSC model and can work efficiently  
87 with genome-scale data (Liu et al. 2009; Liu et al. 2010; Chifman and Kubatko 2014;  
88 Mirarab et al. 2014c). Their application has been successful in resolving mammalian, avian,  
89 and seed plant relationships in cases where concatenation methods have been  
90 demonstrated to be inconsistent (Song et al. 2012; Xi et al. 2013; Reddy et al. 2017).

91           In addition to ILS, gene tree estimation error has also been a major focus of work to  
92 improve the accuracy of phylogenomic inference. This is especially relevant for summary  
93 coalescent methods, which assume the input gene trees to be essentially error-free, e.g.,  
94 Lanier et al. 2014; Mirarab et al. 2014c; Roch and Warnow 2015; Xu and Yang 2016; Blom  
95 et al. 2017. Rapid radiations are particularly challenging in this regard. Here, short internal  
96 branches may yield error-prone gene tree estimation when phylogenetically informative  
97 characters are minimal (Xi et al. 2015). This may be further complicated if such radiations  
98 are ancient and followed by long descendent branches, which may exacerbate long-branch  
99 attraction artifacts (Whitfield and Kjer 2008). Though benchmark studies have  
100 demonstrated the consistency of summary coalescent methods when substantial amounts  
101 of such non-phylogenetic signal are included (Philippe et al. 2011; Roch and Warnow 2015;  
102 Xi et al. 2015; Hahn and Nakhleh 2016), accurate gene tree inference remains of crucial  
103 importance for reliable species tree estimation (Shen et al. 2017). A number of methods  
104 have been developed to mitigate gene tree estimation error, including improving taxon  
105 sampling, applying appropriate models of nucleotide evolution, reducing missing data,  
106 subsampling informative genes, and locus binning (Zwickl and Hillis 2002; Lemmon et al.  
107 2009; Salichos and Rokas 2013; Cox et al. 2014; Mirarab et al. 2014a; Hosner et al. 2015).

108           Beyond ILS and gene tree estimation error, gene flow between non-sister species  
109 can similarly result in gene tree–species tree conflict and lead to incorrect species tree  
110 estimation. Unlike the MSC model, gene flow from a non-sister species leads to an  
111 overrepresentation of the parental allele in the descendants and therefore the frequencies  
112 of the two minor topologies are asymmetrical (Durand et al. 2011). A number of species  
113 network inference methods have been developed to detect and infer gene flow based on

114 such expectation. They either use counts of the shared derived alleles, such as the classic D-  
115 statistic test (Green et al. 2010; Durand et al. 2011), or the gene tree topology as input (e.g.,  
116 Huson et al. 2005; Meng and Kubatko 2009; Yu et al. 2011; Solís-Lemus et al. 2017). The  
117 latter methods are often based on *a priori* evolutionary models and have been increasingly  
118 applied to empirical data sets.

119         During periods of rapid radiation, all of the above phenomena—ILS, introgression,  
120 and gene tree estimation error—may occur simultaneously to obscure phylogenetic signal  
121 (Pease et al. 2016), culminating in a perfect storm confounding phylogenomic inference.  
122 When a limited number of alternative species tree topologies are involved, these  
123 phenomena can be distinguished from each other using methods discussed above (Zwickl  
124 et al. 2014; Arcila et al. 2017; Meyer et al. 2017; Beckman et al. 2018; Glémin et al. 2019).  
125 However, when the rapid radiation generates a cloud of alternative tree topologies, all of  
126 which are weakly supported, such model-based methods become less practical because  
127 priors necessary to test hypothesis of introgression are difficult to determine accurately.  
128 Additional challenges arise from the excessive computational resources required to apply  
129 such network inference methods to data sets involving hundreds of species. Moreover,  
130 following the identification of ILS, introgression, and gene tree estimation error, a more  
131 quantitative assessment characterizing their relative contribution to overall gene tree  
132 variation has not been addressed in any empirical system to our knowledge.

133         Using anchored hybrid enrichment (Lemmon et al. 2012), we generated a large  
134 phylogenomic data set including 423 single-copy nuclear loci with 64 taxa to infer  
135 relationships of the flowering plant clade Malpighiales. The order Malpighiales comprise ca  
136 7.8% of eudicot diversity (Magallon et al. 1999) and include more than 16,000 species in

137 ~36 families (Stevens and Davis 2001). Species in Malpighiales encompass astonishing  
138 morphological and ecological diversity ranging from epiphytes (Clusiaceae), submerged  
139 aquatics (Podostemaceae), to emergent rainforest canopy species (Callophyllaceae). The  
140 order also includes numerous economically important crops with sequenced genomes, e.g.,  
141 rubber (*Hevea*), cassava (*Manihot*), flax (*Linum*), and aspen (*Populus*). Despite their  
142 ecological and economic importance, the evolutionary history of Malpighiales remains  
143 poorly understood. While analyzing chloroplast genome sequences has greatly improved  
144 the resolution of this clade, relationships among its major subclades remain uncertain (Xi  
145 et al. 2012), and analyses using nuclear genes lack resolution along the spine of the clade  
146 (Davis et al. 2005; Wurdack and Davis 2009). According to Smith et al. (2013), this region  
147 of the Malpighiales phylogeny has been implicated in nine of the top ten most unstable  
148 nodes across all angiosperms, including Pandaceae, Euphorbiaceae, Linaceae, the most  
149 recent common ancestor (MRCA) of Salicaceae and Lacistemataceae, the MRCA of  
150 Malpighiaceae and Elatinaceae, as well as the MRCA of putranjivoids, phyllanthoids,  
151 chrysobalanoids, and rhizophoroids *sensu* Xi et al. (2012). In short, Malpighiales have been  
152 coined one of the “thorniest nodes” in the angiosperm tree of life (Soltis et al. 2005). A long-  
153 standing hypothesis for this lack of resolution has been attributed to the clade’s rapid  
154 radiation during the Albian and Cenomanian (112–94 million years ago [Ma]; Davis et al.  
155 2005; Wurdack and Davis 2009; Xi et al. 2012). This radiation has produced a phylogeny  
156 characterized by extremely short internal branches along the backbone of the phylogeny,  
157 followed by long branches subtending most crown group families. This is particularly  
158 problematic because, as we summarize above, short internal branches represent species  
159 tree anomaly zones where ILS may be pervasive and gene tree estimation error is high (Liu

160 et al. 2015; Roch and Warnow 2015; Edwards et al. 2016). Incongruent phylogenetic  
161 signals between organelle and nuclear genes also support introgression associated with the  
162 origin of this order (Sun et al. 2015).

163         The development of next-generation sequencing, the MSC model that accommodate  
164 ILS, and best practices to reduce gene tree estimation error offers a unique opportunity to  
165 re-examine Malpighiales in the context of resolving rapid radiations. Here, we apply both  
166 concatenation and coalescent-based methods for phylogenomic analyses and evaluate the  
167 relationships and consistency of nodal resolution under a variety of conditions. We also  
168 apply simulations to explore the impact of ILS and gene tree estimation error based on the  
169 empirical parameters of our inferred species tree. We further apply a triplet analysis to  
170 detect gene flow and identify hotspots of reticulate evolution in the species tree. And  
171 finally, we develop a novel method to quantitatively assess the contribution of three  
172 primary sources of gene tree variation in Malpighiales—ILS, gene tree estimation error,  
173 and gene flow.

174

## 175 **MATERIALS AND METHODS**

### 176 *Taxon Sampling*

177         We sampled a total of 56 species in the order Malpighiales, representing 39 families  
178 and all major clades *sensu* Wurdack and Davis (2009) and Xi et al. (2012) (Table S1).  
179 Species were sampled to represent the breadth of Malpighiales diversity. Four species from  
180 the order Celastrales and two species from the order Oxalidales were sampled as closely  
181 related outgroups (Chase et al. 2016). Two species from the order Vitales were also  
182 included as more distantly related outgroups (Chase et al. 2016, Table S1).



183

184 *Library Preparation, Enrichment, and Locus Assembly*

185 Data were collected at the Center for Anchored Phylogenomics at Florida State  
186 University (<http://www.anchoredphylogeny.com>) using the anchored hybrid enrichment  
187 method (Lemmon et al. 2012; Buddenhagen et al. 2016). This method targets universally  
188 conserved single-copy regions of the genome that typically span 250 to 800 base pairs (bp),  
189 thus mitigating the confounding effect of paralogy in gene tree estimates. Briefly, total  
190 genomic DNA was sonicated to a fragment size of 300–800 bp using a Covaris E220  
191 Focused-ultrasonicator. Library preparation and indexing was performed following the  
192 protocol in Hamilton et al. (2016). A size-selection step was also applied after blunt-end  
193 repair using SPRI select beads (Beckman-Coulter Inc). Indexed samples were then pooled  
194 and enriched using the Angiosperm v1 kit (Agilent Technologies Custom SureSelect XT kit  
195 ELID 623181; Buddenhagen et al. 2016). The resulting libraries were sequenced on an  
196 Illumina HiSeq 2500 System using the PE150 protocol.

197 Quality-filtered sequencing reads were processed following the methods described  
198 in Hamilton et al. (2016) to generate locus assemblies. Briefly, paired reads were merged  
199 prior to assembly following Rokyta et al. (2012). Reads were then mapped to the probe  
200 region sequences of the following reference genomes: *Arabidopsis thaliana* (Malvales,  
201 *Arabidopsis* Genome Initiative 2000), *Populus trichocarpa* (Malpighiales, Tuskan et al.  
202 2006), and *Billbergia nutans* (Poales, Buddenhagen et al. 2016). Finally, the assemblies  
203 were extended into the flanking regions. Consensus sequences were generated from  
204 assembly clusters with the most common base being called when polymorphisms could be  
205 explained as sequencing error.

206

## 207 *Orthology Assignment*

208           Orthologous sequences were determined following Prum et al. (2015) and Hamilton  
209 et al. (2016). The assembled sequences were grouped by locus and a pairwise distance was  
210 calculated as the percent of shared 20-mers. Sequences were subsequently clustered based  
211 on this distance matrix using the neighbor-joining algorithm (Saitou and Nei 1987). When  
212 more than one cluster was detected for a target region, each cluster was treated as a  
213 different locus in subsequent analyses. Clusters including less than 50% of the species were  
214 discarded.

215

## 216 *Sequence Alignment, Masking, and Site-subsampling*

217           Each locus was first aligned using MAFFT v7.023b (Kato and Standley 2013) with  
218 “--genafpair --maxiterate 1000” flags imposed. Alignments were end trimmed and  
219 internally masked to remove misassembled or misaligned regions (Buddenhagen et al.  
220 2016). Firstly, conserved sites were identified in each alignment where >40% of the  
221 nucleotides at that site were identical across species. For end trimming, sequences for each  
222 gene accession were scanned from both ends towards the center until more than fourteen  
223 nucleotides in a sliding window of 20 bp matched the conserved sites. Once the start and  
224 end of each sequence was established, the internal masking then required that >50% of the  
225 nucleotides in a sliding window of 30 bp matched the conserved sites. Regions that did not  
226 meet this criterion were masked. Finally, we removed any gene sequence in the alignment  
227 with >50% ambiguous nucleotide composition. We also required all locus alignments to  
228 contain *Leea guineense* (Vitales) for rooting purposes.

229 To further explore the phylogenetic utility of the flanking regions of hybrid  
230 enrichment data, we applied three increasingly stringent site-subsampling strategies using  
231 trimAl v1.2 (Capella-Gutiérrez et al. 2009) following our masking steps described above. To  
232 construct our “low-stringency data set”, we set the gap threshold to be 0.8 (-gt 0.8) in  
233 trimAl to remove sites containing >20% indels or missing data for each alignment. This  
234 data set includes the highest percentage of flanking regions and resulted in the longest  
235 alignments. We then applied a site composition heterogeneity filter to this “low-stringency  
236 data set” to create our “medium-” and “high-stringency data set” by setting the minimum  
237 site similarity score to be 0.0002 and 0.001 (e.g., -st 0.001), respectively. This has the effect  
238 of removing especially rapidly evolving sites within flanking regions for which we expect  
239 higher composition heterogeneity. The resulting “medium-” and “high-stringency data set”  
240 thus include lower percentage of flanking regions.

241

#### 242 *Gene Tree Estimation*

243 To infer individual gene trees for coalescent-based analyses, we applied maximum  
244 likelihood (ML) as well as Bayesian Inference (BI). To estimate ML trees, we used RAxML  
245 v8.1.5 (Stamatakis 2014) under the GTR+ $\Gamma$  model with 20 random starting points. We  
246 chose the GTR+ $\Gamma$  model because it accommodates rate heterogeneity among sites, while the  
247 other available GTR model in RAxML, the GTRCAT model, is less appropriate due to our  
248 small taxon sampling size (Stamatakis 2014). Statistical confidence of each gene tree was  
249 assessed by performing 100 bootstrap (BP) replicates. We additionally inferred the  
250 Bayesian posterior distribution of gene trees using MrBayes v3.2.1 (Ronquist and  
251 Huelsenbeck 2003). We only applied BI to the low-stringency data set due to computational

252 cost and this data set yielded the best resolved gene trees (see Results below). We applied  
253 the GTR+ $\Gamma$  model with two independent runs for each gene. Each run included four chains,  
254 with the heated chain at temperature 0.20 and swapping attempts every 10 generations.  
255 Initially, four million generations were used with 25% burn-in period, sampled every 1,000  
256 generations. Runs that failed to reach the targeted standard deviation of split frequencies  
257  $\leq 0.02$  were rerun with the same settings but with 10 million generations, sampled every  
258 5,000 generations until attaining a standard deviation of split frequencies  $\leq 0.02$ . We  
259 randomly sampled 100 trees in the posterior distribution of inferred gene trees to conduct  
260 bootstrap replication in the coalescent analyses (Table S2). Trees sampled from the  
261 posterior distribution are more similar to the optimum Bayesian tree than those sampled  
262 from the non-parametric bootstrapping. Therefore, we also expect higher support values in  
263 the species tree.

264

#### 265 *Species Tree Inference Using Concatenation and Coalescent-based Methods*

266 Our trimmed gene matrices were concatenated and analyzed using both RAxML and  
267 ExaML v3.0.18 (Kozlov et al. 2015). In our RAxML analyses, the species trees were inferred  
268 under the GTR+ $\Gamma$  model with 100 rapid bootstrapping followed by a thorough search for  
269 the ML tree. In ExaML analyses, species trees were inferred under the GTR+ $\Gamma$  model with  
270 20 random starting points. We then conducted 100 bootstrap replicates to evaluate nodal  
271 support. Partitions for both analyses were selected by PartitionFinder v2.1.1 based on AICc  
272 (Akaike Information Criterion) criteria using the heuristic search algorithm “rcluster”  
273 (Lanfear et al. 2012). We also conducted BI for species tree estimation as implemented in  
274 PhyloBayes (Lartillot et al. 2013). For BI analysis, we applied the CAT-GTR model, which

275 accounts for across-site compositional heterogeneity using an infinite mixture model  
276 (Lartillot and Philippe 2004). Two independent Markov chain Monte Carlo (MCMC)  
277 analyses were conducted for each concatenated nucleotide matrix. Convergence and  
278 stationarity from both MCMC analyses were determined using bpcomp and tracecomp  
279 from PhyloBayes. We ran each MCMC analysis until the largest discrepancy observed  
280 across all bipartitions was smaller than 0.1 and the minimum effective sampling size  
281 exceeded 200 for all parameters in each chain.

282 To infer our species tree using coalescent-based models, we obtained ML gene trees  
283 and BI consensus trees for each locus. MP-EST (Liu et al. 2010) and ASTRAL-II (Mirarab  
284 and Warnow 2015) were subsequently used to perform species tree inference using  
285 optimally estimated gene trees. Statistical confidence at each node was evaluated by  
286 performing the same species tree inference analysis on 100 ML bootstrap gene trees or  
287 trees sampled from our Bayesian posterior distributions. The resulting 100 species trees  
288 estimated from bootstrapped samples were summarized onto the species tree inferred  
289 from ML gene trees using the option “-f z” in RAxML.

290

291 *Simulation of gene alignments with realistic parameters of ILS and gene tree estimation error*

292 To investigate the impact of ILS and gene tree estimation error on the accuracy of  
293 species tree inference we simulated sequences assuming a known species tree. Here, the  
294 tree topology estimated by MP-EST with the low-stringency data set (analysis No. 15 in  
295 Table S2) was invoked as the known species tree. We chose this best-supported MP-EST  
296 topology because the branch lengths are estimated in coalescent unit, which is an essential  
297 parameter for ILS simulation. We thus applied this species tree to all of the downstream

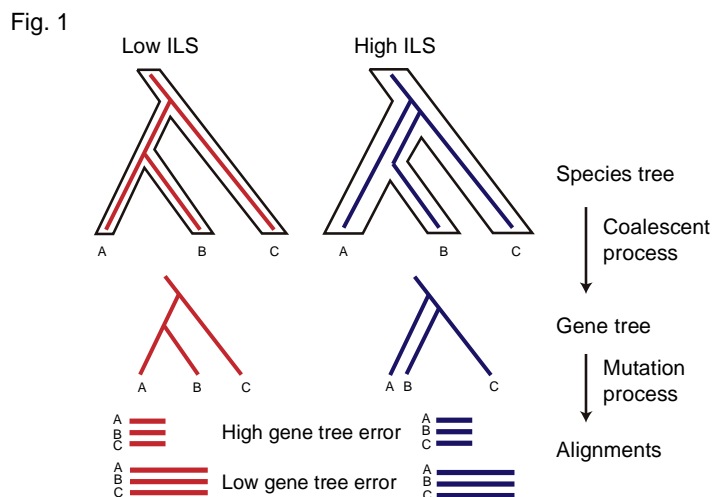
298 simulation-based analyses, including the triplet test for MSC model fitness and relative  
299 importance analysis.

300 To simulate conditions of high and low levels of ILS, we modified the key population  
301 mutation parameter “theta” when generating gene trees under the coalescent model using  
302 the function “sim.coal.tree.sp.mu” in the R package Phybase (Liu and Yu 2010). Theta was  
303 set to be 0.01 and 0.1 to reflect low and high ILS, respectively. The range of theta was  
304 determined based on our empirical data sets by following two steps. First, we inferred the  
305 branch lengths of the species tree in mutation units in RAxML using the fixed topology of  
306 the MP-EST species tree and the concatenated low-stringency data set. Second, theta for  
307 each branch was calculated by dividing the branch lengths estimated from RAxML  
308 (mutation units) by that estimated from MP-EST (coalescent units). The other input for  
309 Phybase, the ultrametric species tree, was generated from this RAxML phylogeny using the  
310 function “chronos” in the R package ape (Paradis et al. 2004). In addition, we set the  
311 relative mutation rates to follow a Dirichlet distribution with alpha equal to 5.0. This alpha  
312 reflected the large variance in gene mutation rates. Finally, 1,500 non-ultrametric gene  
313 trees were simulated separately for each theta.

314 From these simulated gene trees, DNA alignments of different lengths were  
315 subsequently generated to reflect various levels of gene tree estimation error since  
316 alignment length is easy to manipulate and shorter alignments correspond to higher error  
317 rates (Mirarab et al. 2014b). We used bppsuite (Guéguen et al. 2013) to simulate  
318 alignments under the GTR+ $\Gamma$  model. Parameters of the model, including the substitution  
319 matrix, base frequency, and the gamma rate distribution were extracted from the RAxML

320 phylogeny above inferred from the low-stringency data set. For each gene tree we  
321 generated alignments of 300, 400, 500, 1,000, and 1,500 bp.

322 As a result, fifty data sets were generated by including 100, 200, 500, 1,000, and  
323 1,500 simulated loci of five length categories and two theta categories (Table S3, Fig. 1).  
324 Species trees were inferred using the concatenation and coalescent methods as described  
325 above under these varying levels of ILS and gene tree estimation error. Finally, we  
326 quantified gene tree–species tree discordance and species tree error by measuring the RF  
327 distance between an estimated gene tree or species tree to the true species tree.



328

329 **Figure 1** Simulation of ILS and gene tree estimation error. ILS was simulated through the  
330 coalescent process by setting low (0.01) and high (0.1) theta values. DNA alignments were  
331 subsequently generated through the mutation process based on simulated gene trees. Five  
332 alignments were generated for each gene tree with lengths of 300, 400, 500, 1000, and  
333 1500 bp (only two are shown in the graph). Shorter alignment lengths increase in gene tree  
334 estimation error.

335

336 In order to assess the sensitivity of our simulation results to the choice of input  
337 species tree and theta values, we additionally examined gene tree–species tree discordance  
338 among bootstrapped samples. We simulated 1,500 gene trees for each of the 100 MP-EST  
339 bootstrapped species trees. Gene trees were simulated directly from each species tree  
340 using the “sim.coal.mpest” function in the R package Phybase (Liu and Yu 2010). This  
341 method does not require *a priori* theta parameters as was imposed in our simulation above  
342 and so alleviates concerns of applying erroneous theta values. We subsequently quantified  
343 the gene tree–species tree discordance for each bootstrap replicate as described above. We  
344 did not use these gene trees to simulate alignments because these gene trees are  
345 ultrametric (Liu and Yu 2010) and thus not suitable for such purpose.

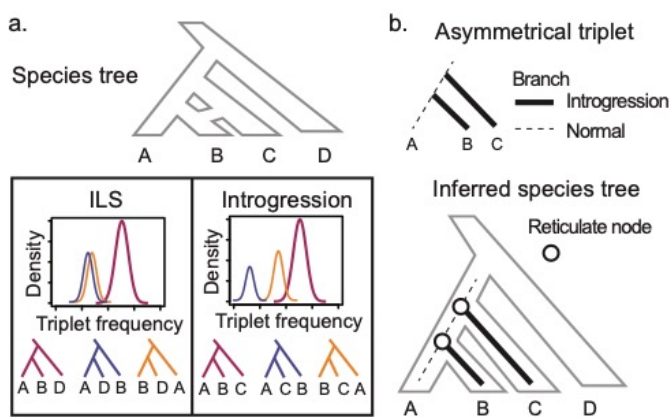
346

#### 347 *A Test of the MSC Model Using Triplet Frequencies*

348 To determine the fit of the MSC model to our empirical data we additionally  
349 examined the triplet frequency for all 423 ML genes trees inferred from our low-stringency  
350 data set using a custom R script available on Github  
351 ([http://github.com/lmcai/Coalescent\\_simulation\\_and\\_gene\\_flow\\_detection](http://github.com/lmcai/Coalescent_simulation_and_gene_flow_detection)). We used the  
352 asymmetrical triplet frequency as evidence for introgression (Fig. 2a). This metric has been  
353 widely applied in parsimony, likelihood, and Bayesian based species network inference  
354 methods to detect sources of gene flow (Nakhleh 2013) . Our method differs from these  
355 methods in two aspects: first, the statistical significance of asymmetry in triplet frequency  
356 is determined by a null distribution simulated from the empirical data. We took into  
357 account variations from ILS and missing data, thus reducing the false positive rate. Second,  
358 unlike other model-based species network inference methods, after identifying



359 significantly asymmetrical triplets, we used a novel method to summarize and visualize the  
360 distribution of lineages involved in gene flow on a species tree without optimizing the  
361 global network (Fig. 2b). As a result, our methods can be easily scaled to genomic data  
362 involving hundreds of taxa.



363  
364 **Figure 2** Identification of reticulate evolution using triplet frequency. (a) Theoretical  
365 expectations of triplet frequency distribution under the multi-species coalescent (MSC)  
366 model with and without introgression. In case of incomplete lineage sorting (ILS),  
367 symmetrical distributions of the frequency of two minor topologies are expected owing to  
368 deep coalescence (left). In case of introgression, one of the minor topologies will occur with  
369 higher frequency due to gene flow (right). (b) Mapped asymmetrical triplets to species tree  
370 to identify reticulate nodes.

371  
372 In order to identify a triplet with significantly asymmetrical frequencies, we  
373 generated a null distribution of triplet frequencies for each triplet using simulated gene  
374 trees under the MSC model. For each of the 100 MP-EST BP species trees, we simulated 423  
375 gene trees using the “sim.coal.mpest” function in Phybase. For each set of simulated gene

376 trees, we then generated missing data for each species by pruning that species randomly  
377 among all gene trees so that the number of sampled genes of that species was the same as  
378 the empirical data. We subsequently counted triplet frequency for these gene trees in each  
379 bootstrap replicate. This simulated distribution reflects the variation of triplet frequency  
380 arising from ILS, estimation error, sampling error, and missing data. A triplet in the  
381 empirical data was identified to be significantly asymmetrical if the difference between the  
382 two less frequent triplets exceeded the maximum difference under simulated conditions.  
383 Such triplets potentially violate the assumptions of the MSC model, and point towards gene  
384 flow especially as an additional factor influencing gene tree heterogeneity, though ancestral  
385 population structure (Slatkin and Pollack 2008) and biases in substitution or gene loss can  
386 produce asymmetrical triplet as well (see Discussion below).

387

### 388 *Identifying hotspots of reticulate evolution using the Reticulation Index*

389 We developed a relative measurement statistic, the 'Reticulation Index', to quantify  
390 the intensity of introgression at each node. First, for each asymmetrical triplet, we mapped  
391 the two inferred introgression branches to the species tree (Fig. 2b). Second, for each node  
392 on the species tree, we counted the number of introgression branches that were mapped to  
393 it. These raw counts were then normalized by the total number of triplets associated with  
394 that node. The resulting percentage is the Reticulation Index for each node. The R script for  
395 calculating the Reticulation Index and visualizing the result on a species tree is available in  
396 the above Github repository.

397

398 *A Novel Method to Quantify Gene Tree Variation Due to ILS, Gene Tree Estimation Error, and*  
399 *Gene Flow*

400           Untangling the effects of ILS, gene tree estimation error, and gene flow is  
401 challenging since they all lead to gene tree–species tree discordance. Here, based on a  
402 multiple regression model (Grömping 2006), we assign shares of relative importance to ILS,  
403 gene tree estimation error, and gene flow in generating gene tree variation by variance  
404 decomposition.

405           For all 63 internal nodes in our species tree, we separately estimated the level of ILS,  
406 gene tree estimation error, and gene flow for each node. ILS is represented by our  
407 estimates of theta. Gene flow is represented by the Reticulation Index for each node. To  
408 infer the level of gene tree estimation error at each node, we additionally simulated 423  
409 gene alignments of 446 bp (median alignment length in low-stringency data set) from the  
410 MP-EST species tree, but each with unique substitution model parameters estimated from  
411 the empirical alignments. This simulation and phylogeny inference followed the same  
412 strategy of alignment simulation described above (Fig. 1). We subsequently inferred  
413 phylogenies for these alignments and summarized them on the species tree to obtain the  
414 BP value at each node. Here, the BP values represent the gene tree variation generated by  
415 estimation error.

416           The gene tree variation in the empirical data is obtained by summarizing bootstrap  
417 trees from each of the 423 loci in our low-stringency data set onto the species tree. The  
418 resulting BP values represented observed gene tree variation at each node. We then  
419 inferred the relative contribution of ILS, estimation error, and gene flow in explaining gene  
420 tree variation using linear regression methods implemented in the R package relaimpo

421 (Grömping 2006). We used four different methods, “lmg”, “last”, “first”, and “Pratt”, to  
422 decompose the relative importance of the three regressors (Lindeman 1980; Pratt 1987).  
423 All of these methods are capable of dealing with correlated regressors and “lmg” is the  
424 most robust method among them (Grömping 2006). We applied the functions “boot.relimp”  
425 and “booteval.relimp” to estimate the relative importance and their confidence interval by  
426 bootstrapping 100 times.

427  
428 *Testing the utility of the triplet-frequency-based method using a genomic data set from yeast*

429 To further validate our introgression detection method using the triplet frequency  
430 distribution, we applied it to the benchmark multi-locus yeast data set from Salichos and  
431 Rokas (2013). We obtained the 1,070 gene trees and inferred a species tree using MP-EST.  
432 We also conducted 100 bootstrap replicates of species tree inference using the bootstrap  
433 gene trees. We then applied our triplet method to identify asymmetrical triplets as  
434 described above (*A Test of the MSC Model Using Triplet Frequencies*). Finally, all  
435 asymmetrical triplets were mapped to the inferred species tree and the Reticulation Index  
436 for each node was calculated and visualized as described above (*Identifying hotspots of*  
437 *reticulate evolution using the Reticulation Index*).

438

## 439 **RESULTS**

### 440 *Hybrid Enrichment*

441 We successfully captured and sequenced 423 of our 491 targeted loci. The resulting  
442 data matrix was densely sampled and included only 12% missing data. One hundred and  
443 one loci included at least 61 taxa (>95% occupancy) and only four loci had more than 19

444 missing species (>30%). The locus sampling per taxon varied from 423 (*Leea guineense*) to  
445 278 (*Ouratea sp.* and *Lophopyxis maingayi*, Table S4). After applying site subsampling, the  
446 alignment lengths ranged from 190 to 885 bp (median 446 bp) for the low-stringency data  
447 set, 157 to 791 bp (median 376 bp) for the medium-stringency data set, and 112 to 751 bp  
448 (median 271 bp) for the high-stringency data set (Table S4). In all data sets, the number of  
449 parsimony informative sites and the average nodal support was significantly positively  
450 correlated with alignment length ( $p$ -value <1e-5, Fig. S1).

451

#### 452 *Flanking Regions Increase Gene Tree and Species Tree Resolution*

453 We observed increasing mean BP support among gene trees as increasingly larger  
454 percentages of the flanking region were included. The average gene tree nodal support  
455 from our low-stringency data set (42 ML BP) was significantly higher than nodal support  
456 estimates for the medium (39 ML BP,  $p$ -value = 1.2e-89 in paired t-test) and high-  
457 stringency data sets (35 ML BP,  $p$ -value = 1.3e-77, Fig. S2a).

458 These increases in gene tree resolution also contributed to increased species tree  
459 resolution as well as species tree inference congruency. For both concatenation and  
460 coalescent analyses, species trees estimated from the low stringent data set with highest  
461 amount of flanking regions, always resulted in the highest average BP support (Fig. S2b,c,  
462 Table S2) and the lowest pairwise RF distances (Fig. S2d) indicating increased statistical  
463 consistency when adding flanking regions.

464

#### 465 *Malpighiales Species Tree Resolution*

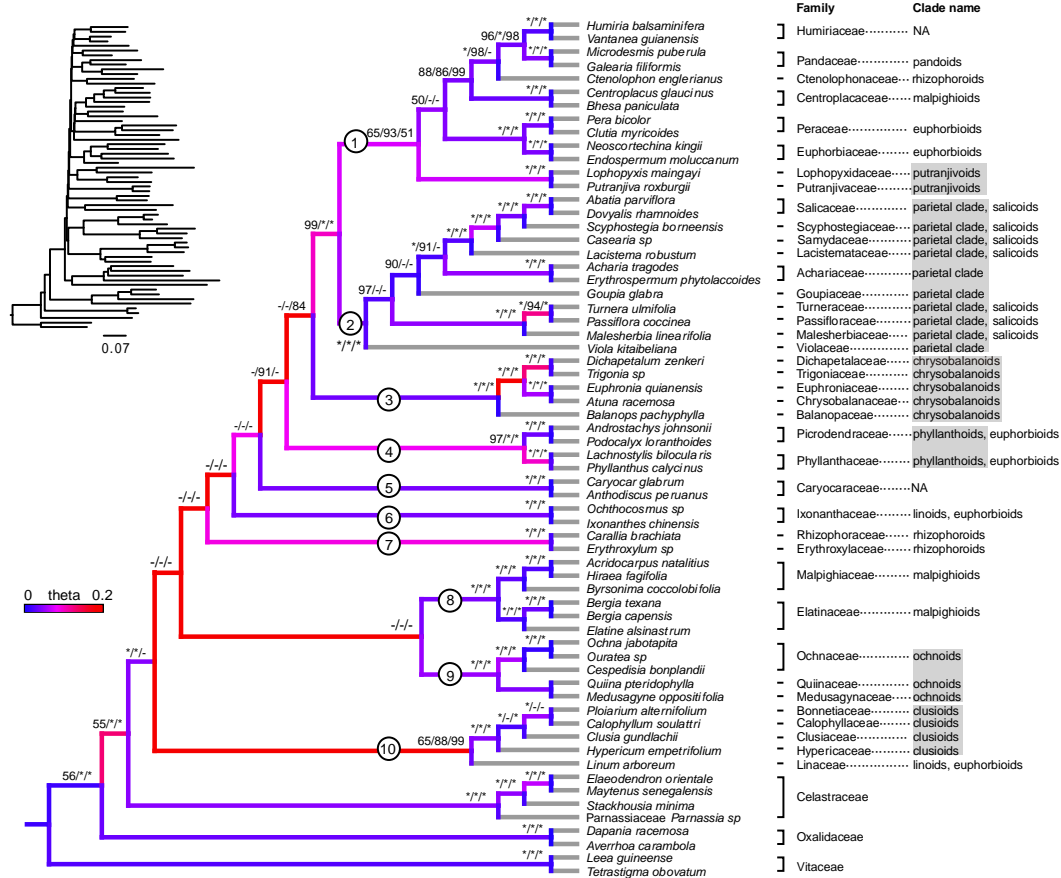
466 We observed significantly higher average species tree nodal support in  
467 concatenation compared to coalescent reconstructions (Table S2,  $p$ -value =  $2.62e-28$  in  
468 paired  $t$ -test). However, our results also suggest statistical inconsistency across data sets  
469 when applying concatenation (Fig. S3). The higher pairwise weighted Robinson–Foulds  
470 distance (WRF) in concatenation indicate more well-supported conflicts among species  
471 trees, which further supports mounting evidence that coalescent methods are more  
472 consistent when reconstructing species tree relationships involving extensive ILS (i.e., the  
473 anomaly zone, Degnan and Rosenberg, 2006, Rosenberg and Tao, 2008). In addition, we did  
474 not find locus subsampling based on locus length, number of PI sites, or gene tree quality  
475 help increase species tree resolution (see Supplementary Note 1, Table S2).

476 Our most well resolved species trees estimation inferred with ASTRAL and MP-EST  
477 uncovered ten major subclades of Malpighiales (Clade 1 to 10 in Fig. 3). These relationships  
478 corresponded to families or closely related clades of families, five of which have previously  
479 been identified using plastid genome (Fig. 3, Xi et al. 2012). Five new clades were  
480 supported with  $\geq 50$  BP,  $>0.90$  PP. Three of these newly identified clades are in conflict ( $>70$   
481 BP) with the plastid phylogeny from Xi et al. (2012) and are discussed more extensively  
482 below. Interrelationships among these ten major subclades, however, were not well  
483 resolved ( $<50$  BP).

484

#### 485 *Simulated Levels of ILS and Gene Tree Estimation Error Reflects Empirical Data*

486 The 5% and 95% quantiles of theta were inferred to be 0.0254 and 0.176,  
487 respectively, with a median of 0.0688. High theta was mostly found along the backbone of  
488 the species tree, indicating the likelihood of extensive ILS within this region of the tree (Fig.



489

490 **Figure 3** Species phylogeny of Malpighiales derived from MP-EST with complete low-

491 stringency data set (analysis No. 15 in Table S2). Gene trees are estimated using MrBayes.

492 Branches are colored by the inferred population mutation parameter theta. Warmer colors

493 indicate higher theta and thus higher level of ILS. Terminal branches are colored grey due

494 to lack of data to infer theta. BP values from best-resolved MP-EST/ASTRAL/RaxML

495 analyses (analysis No. 15, 17, and 11 in Table S2) are indicated above each branch; an

496 asterisk indicates 100 BP support; a hyphen indicates less than 50 BP. Branch lengths

497 estimated from RaxML by fixing the species tree topology are presented at the upper left

498 corner. The eleven major clades highlighted in the discussion are identified with circled

499 numbers along each relevant branch. The clade affiliation for each family based on the

500 plastid phylogeny (Xi et al. 2012) is indicated on the right. Clades identified by Xi et al.  
501 (2012) that are also monophyletic in this study are highlighted using gray shades.

502  
503 3). This is likely an overestimation of theta since all topological variations are attributed to  
504 coalescent process including the ones originate from mutational variance (Huang and  
505 Knowles 2009). We therefore set the theta parameter to be 0.01 and 0.1 in our coalescent  
506 gene tree simulation, which reflected the left and right tails of low and high ILS estimated  
507 from empirical data.

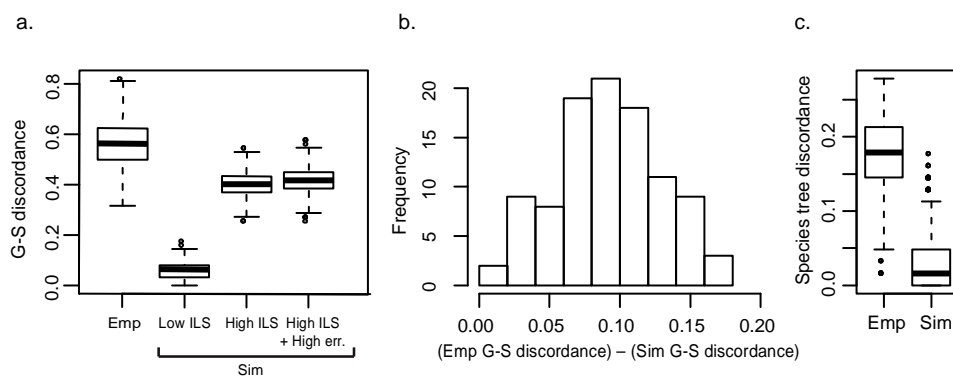
508 In our simulation, the average gene tree estimation error was 0.319 for alignments  
509 of 300bp, 0.261 for 400bp, 0.221 for 500bp, 0.133 for 1000bp and 0.098 for 1500bp under  
510 low ILS and 0.340 for alignments of 300bp, 0.286 for 400bp, 0.241 for 500bp, 0.161 for  
511 1000bp and 0.120 for 1500bp under high ILS. Here, an RF distance of 0 signifies error-free  
512 reconstruction versus 1 indicating that none of the true nodes are recovered. Gene tree  
513 estimation error was therefore lower in low ILS ( $p$ -value=6.08e-16 in Student's  $t$ -test), but  
514 was still significantly higher than that estimated from empirical data ( $p$ -value= 4.24e-65 in  
515 Student's  $t$ -test; see Supplementary Note 2; Fig. S4).

516

#### 517 *Simulation Yields Consistent and Accurate Species Tree Estimation*

518 In our empirical analyses, the low-stringency data set yielded the lowest average  
519 gene tree–species tree conflict of 0.563 among the other data sets. In our simulations, the  
520 highest average gene tree–species tree conflict observed was 0.507, by setting theta = 0.1  
521 and alignment length = 300 bp. Therefore the lowest empirical gene tree–species tree  
522 discordance was still significantly higher than the simulated conditions with extremely





523

524 **Figure 4** Extensive gene tree discordance in empirical versus simulated data. (a) Gene  
525 tree–species tree (G-S) discordance in the empirical (Emp) and simulated (Sim) data  
526 assuming fixed theta in simulation. Discordance is measured by RF distance between  
527 inferred gene trees and the species tree. Under various simulated conditions of ILS (e.g.,  
528 ‘Low ILS’, theta = 0.01 and ‘High ILS’, theta = 0.1) and gene tree estimation error (‘High ILS  
529 + High err.’, theta = 0.1, alignment length=300bp), the simulated gene tree–species tree  
530 discordance is significantly lower than that from empirical data. (b) Gene tree–species tree  
531 discordance is higher in empirical versus simulated conditions without setting theta a  
532 priori. For each BP data set, gene tree–species tree discordance is measured and compared  
533 in both empirical and simulated data sets. Positive values indicate higher gene tree–species  
534 tree discordance in our empirical data. (c) Species tree estimation discordance in empirical  
535 data (left) and simulated data (right).

536

537 high level of ILS and gene tree estimation error ( $p$ -value =  $2.2e-16$ , Student’s  $t$ -test, Fig. 4a).

538 The same conclusion also applies when simulating gene trees directly from species tree  
539 without setting theta *a priori* (Fig. 4b).

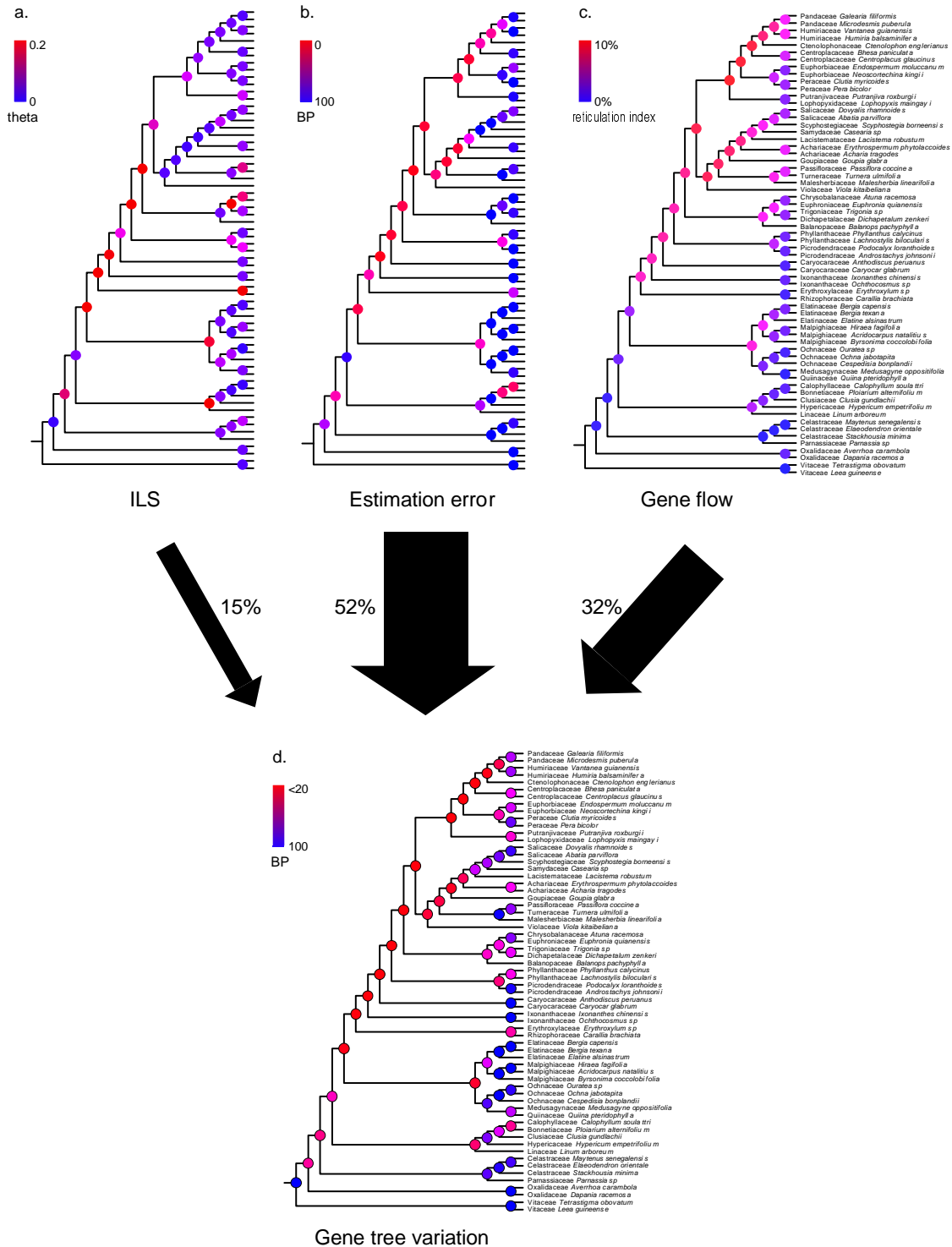
540 Moreover, even under such simulated conditions of extremely high ILS and gene  
541 tree estimation error, both concatenation and coalescent-based methods yielded consistent

542 and accurate species tree estimation with no more than 12 nodes ( $< 0.10$  RF distance, Fig.  
543 4c, Fig. S5) failing to be recovered. The performance of coalescent-based methods is mainly  
544 affected by gene tree estimation error (Fig. S5). Under the highest gene tree estimation  
545 error (300bp), both ASTRAL and MP-EST require 1000 loci to recover the true species tree.  
546 For concatenation methods, ML estimations are robust under low ILS levels, which is  
547 consistent with previous findings (Mirarab et al. 2014b; Tonini et al. 2015). We were able  
548 to recover the correct species tree with the smallest data set (100 loci with 300bp in length)  
549 under low ILS ( $\theta = 0.01$ ). However, major challenges and inaccurate species trees are  
550 generated under high ILS. Under such conditions, it requires the largest data set (1500 loci  
551 with  $\geq 400$  bp length) to recover the true species tree (Fig. S5).

552

553 *MSC Model Fitness and the Relative Contribution of ILS, Gene Tree Estimation Error, and Gene*  
554 *Flow to Gene Tree Variation*

555 Among all 41,664 triplets we examined, 553 (1.3%) have significant asymmetrical  
556 minor frequencies. The node with the highest Reticulation Index is the MRCA of Clade 1 and  
557 Clade 2 (the MRCA of Salicaceae and Euphorbiaceae; Fig. 5c). 10.3% of the triplets  
558 associated with this node are significantly asymmetrical. According to our relative  
559 importance decomposition analysis, ILS, gene tree estimation error, and gene flow explain  
560 57.5% of the gene trees variation using the lmg algorithm ( $R^2 = 0.575$ ). When scaling these  
561 three factors to sum 100%, gene tree estimation error is the most dominant factor, which  
562 explains 52% of the gene tree variation (Fig. S6). The second most significant factor is gene  
563 flow, which explains 32% of the gene tree variation. And ILS explains the least variation



564

565 **Figure 5** Relative contributions of ILS, estimation error, and gene flow across Malpighiales.

566 (a) ILS. Nodes are colored by inferred population mutation parameter  $\theta$ . (b) Gene

567 estimation error. Nodes are colored by BP values, which represent percentage of recovered  
568 nodes from simulation (see Materials and Methods). (c) Gene flow. Nodes are colored by  
569 Reticulation Index. (d) Gene tree variation. Nodal BPs reflect nodal recovery in gene trees.  
570 Percentages of gene tree variation ascribed to ILS, estimation error, and gene flow are  
571 indicated by black arrows.

572  
573 (15%). The relative ranks of these three factors are consistent among regression methods  
574 and bootstrap replicates (Fig. S6).

575 Further investigation revealed significant negative correlation ( $p$ -value  $2.2e-16$ )  
576 between the overall gene tree variation and species tree resolution (Fig. S7a). All of the  
577 contributors to gene tree variation—ILS, tree estimation error, and introgression—are  
578 strongly negatively correlated with species tree resolution ( $p$ -value  $<6.6e-4$ ). We observed  
579 the highest level of ILS, introgression and gene tree estimation error for the most  
580 recalcitrant nodes along the backbone of the phylogeny using our methods (Fig. S7b–d).  
581 This further corroborates our conclusion that a combination of all three factors contribute  
582 to this low resolution. We did not find significant correlation between the estimated level of  
583 introgression and ILS, suggesting that our triplet method can effectively distinguish these  
584 two phenomena. However, both ILS and introgression are positively correlated to gene tree  
585 estimation error ( $p$ -value  $< 6.8e-3$ ).

586  
587 *The triplet-frequency-based method identified three hotspots of introgression in yeasts*

588 Our species tree of yeast inferred using MP-EST is identical to the topology reported  
589 in the original study by Salichos and Rokas (2013). We identified 116 asymmetrical triplets

590 among the 1,771 triplets in the yeast species tree. These triplets revealed three hotspots of  
591 introgression that correspond to those identified by Yu and Nakhleh (2015): in the MRCA  
592 of *Saccharomyces kluyveri* and *Kluyveromyces waltii*, the MRCA of *Zygosaccharomyces rouxii*  
593 and *Saccharomyces castellii*, and the MRCA of *Candida guilliermondii* and *Debaryomyces*  
594 *hansenii* (Fig. S8). The first two hotspots of reticulation (the MRCA of *S. kluyveri* and *K.*  
595 *waltii*, the MRCA of *Z. rouxii* and *S. castellii*) reflect the donor and recipient lineage of one of  
596 the two reticulation branches identified by Yu and Nakhleh (2015). The third introgression  
597 hotspot involving the MRCA of *C. guilliermondii* and *D. hansenii* reflects the second  
598 reticulation branch inferred in Yu and Nakhleh (2015).

599

## 600 **DISCUSSION**

601 Our results indicate that despite extensive phylogenomic data, the early branching  
602 order of Malpighiales remains uncertain. We attribute this to a combination of factors—a  
603 perfect storm—involving ILS, gene tree estimation error, and gene flow. Below we highlight  
604 our findings in four subsections: the phylogenetic utility of flanking regions in sequence  
605 capture data, novel phylogenetic relationships gleaned for Malpighiales, an efficient  
606 method to investigate gene flow in large data sets, and a novel simulation-based method to  
607 decompose gene tree variation into various contributing factors.

608

### 609 *Flanking Regions Greatly Enhance Phylogenetic Resolution*

610 Hybrid enrichment probes are designed to capture highly conserved anchor regions  
611 as well as the more variable flanking regions adjacent to these anchors. Despite the  
612 perceived utility of these flanking regions in mammals (McCormack et al. 2012) and more

613 recently in plants (Fragoso-Martínez et al. 2017), assumptions of the enhanced  
614 phylogenetic utility of these flanking regions have not been tested explicitly to our  
615 knowledge. Here, we observed significantly higher average ML BP across gene trees,  
616 increased species tree resolution, and most importantly, increased species tree estimation  
617 congruency as flanking regions were increasingly added (Fig. S2). This suggests that longer  
618 loci, favoring more phylogenetically informative flanking regions, should be prioritized in  
619 future anchored hybrid enrichment kit designs. These flanking regions represent genomic  
620 regions under nearly neutral selection where mutation rates are high, and thus appear to  
621 be a rich source of phylogenetic utility. It has been demonstrated that the inclusion of genes  
622 with higher mutation rates can greatly enhance phylogenetic resolution, even deep within  
623 organismal phylogenies (Hilu et al. 2003; Lanier et al. 2014). Our site-subsampling strategy,  
624 which includes increasingly larger proportions of these more rapidly evolving flanking  
625 regions provides the first empirical evidence that these regions are particularly informative  
626 for resolving phylogenetic relationships at shallow and deeper phylogenetic depths.

627

### 628 *Sequence Capture Data Confirms Malpighiales Relationships and Identifies Novel Clades*

629 We assessed the performance of hybrid enrichment markers by evaluating support  
630 for major clades previously identified from plastome sequences (Xi et al., 2012; Fig. 3). The  
631 majority of the well-supported (>90 BP) clades identified by Xi et al. (2012) are  
632 corroborated in our analyses with high confidence (>97 BP). These include the parietal,  
633 clusioid, phyllanthoid, ochroid, chrysobalanoid, and putranjivoid subclades. With rare  
634 exception, relationships within these clades were also identical to those by Xi et al. (2012).  
635 In the case of the parietal and clusioid clades, internal resolutions were less well supported

636 owing to conflicting topologies recovered among coalescent and concatenation methods  
637 (low nodal support indicated by ‘-’ in Fig. 3). Within the parietal clade, for example, the  
638 monophyly of the salicoids *sensu* Xi et al. (2012, Fig. 3) is supported by the RAxML  
639 phylogeny with moderate support (69 BP) but is not supported in any of the coalescent  
640 methods.

641         Additionally, we discovered several noteworthy clades that conflict with those  
642 reported by Xi et al. (2012). The euphorbioids, malpighioids, and rhizophoroids were  
643 paraphyletic in all of the best resolved MP-EST, ASTRAL, and RAxML analyses (Fig. 3). The  
644 euphorbioids—including Euphorbiaceae, Peraceae, Lophopyxidaceae, Linaceae, and  
645 Ixonanthaceae—were split into four polyphyletic groups. In particular, Linaceae was  
646 placed as sister to the clusioid clade in all of the best resolved coalescent and concatenation  
647 analyses (Fig. 3). The affiliation of Linaceae to the clusioids instead of to other members of  
648 the euphorbioids is also supported in a recent transcriptomic study of this group with less  
649 dense taxon sampling (Cai et al. 2019). Within malpighioids, Centroplacaceae is confidently  
650 placed (>86 BP) with Humiriaceae, Pandaceae, and Ctenolophonaceae (Fig. 3) instead of  
651 with Malpighiaceae and Elatinaceae. This relationship is partially supported by Wurdack et  
652 al. (2004) in which Centroplacaceae was placed with Pandaceae, although with low support.  
653 Within the rhizophoroids, Ctenolophonaceae was well nested (>98 BP for coalescent  
654 methods) within a clade including Euphorbiaceae and Pandaceae (Clade 1 in Fig. 3) rather  
655 than with Rhizophoraceae and Erythroxlaceae.

656

657 *ILS and Gene Tree Estimation Error Alone Are Insufficient to Explain the Lack of Species Tree*

658 *Resolution in Malpighiales*

659 Our simulations to explore gene tree heterogeneity encompass the full  
660 distributional range of ILS and gene tree estimation error inferred from the empirical data,  
661 and clearly demonstrate that the data we have assembled should be sufficient to resolve  
662 Malpighiales species tree relationships. Specifically, despite our inability to estimate a well-  
663 resolved species tree from our empirical data, we were able to recover a species tree with  
664 very high confidence in simulation (mean nodal support >91 BP). This is true even when  
665 ILS ( $\theta = 0.1$ ) and gene tree estimation error (alignment length = 300bp) were set to the  
666 highest levels inferred from our empirical data. Such extreme levels of  $\theta$ , in particular,  
667 are ten times higher than empirical estimations from *Arabidopsis* and *Drosophila* (0.01–  
668 0.001 in both cases; Drost and Lee 1995; Fischer et al. 2017). Even when down sampling  
669 our data set under these extreme conditions to include a mere 100 loci, both concatenation  
670 and coalescent analyses recover the true species with no more than 10% error (Fig. S5). In  
671 addition, we observed far fewer conflicts among species trees reconstructed from different  
672 methods and data partitions in simulation versus from those estimated from the empirical  
673 data (Fig. 4c). These results suggest that ILS and gene tree estimation error alone are  
674 insufficient to explain the lack of resolution along the spine of Malpighiales, and suggest  
675 that additional factors likely contribute to gene tree heterogeneity.

676  
677 *Gene Flow Compromises Malpighiales Species Tree Resolution: A Novel Method for Assessing*  
678 *Gene Tree Heterogeneity*

679 Beyond ILS and gene tree estimation error, gene tree heterogeneity is also  
680 attributable to two other common biological factors: gene duplication and gene flow (Yang  
681 2006). As we demonstrate above, the first two factors alone are insufficient to explain this



682 lack of resolution. Orthology assignment problems owing to gene duplications are also  
683 highly unlikely for two reasons. First, our sequence capture data set was specifically  
684 designed for single copy nuclear loci across land plants (Buddenhagen et al. 2016). Second,  
685 large-scale genome duplication identified in Malpighiales all occurred *subsequent* to the  
686 explosive radiation where discordance is localized (Cai et al. 2019). Thus, biased gene loss  
687 arising from genome duplications are unlikely to hinder our ability to resolve backbone  
688 relationships in the order. Additional analytical artifacts not reflected in our assessment  
689 include homolog calls, alignment error, and most importantly, misspecification of DNA  
690 substitution models, all of which can compromise species tree estimation. Though these  
691 analytical errors may explain some discordance, it is quite possible that conflicts are  
692 attributed to additional biological phenomena.

693         Gene flow has yet to receive attention in phylogenomic studies, especially at deep-  
694 time phylogenetic scales. It is estimated that at least 25% of plant species and 10% of  
695 animal species hybridize (Mallet 2007) and various network inference methods have been  
696 developed to assess gene flow in phylogenies (Nakhleh 2013). These methods have  
697 provided valuable insights into reticulate evolution, including those associated with the  
698 rapid radiations in wild tomatoes and heliconius butterflies (Pease et al. 2016; Edelman et  
699 al. 2019). However, the performance of these methods often relies on accurate species tree  
700 estimation and the generation of a handful of alternative species tree topologies to conduct  
701 hypothesis testing. However, when alternative topologies are too numerous to evaluate,  
702 such as along the backbone of Malpighiales, existing tools become quite limited. In  
703 particular, these methods are computationally expensive and amenable only to small data  
704 sets. For example, maximum likelihood can only be applied to networks involving fewer

705 than 10 taxa and three reticulations (Yu and Nakhleh 2015). We leveraged the theoretical  
706 predictions of triplet frequencies to make inferences about gene flow by summarizing the  
707 distribution of lineages involved in horizontal processes using our novel measurement  
708 statistic, the Reticulation Index. Our method can effectively identify hotspots of reticulate  
709 evolution, including both the donor and recipient lineage, in large clades and in deep time,  
710 and provide valuable guidance to empirical studies. We further validated the application of  
711 our Reticulation Index using the yeast data set from Salichos and Rokas (2013). The three  
712 hotspots we identified in the yeast phylogeny (Fig. S8) correspond precisely to the two  
713 reticulation branches previously inferred by Yu and Nakhleh (2015), thus demonstrating  
714 the promise of our method for applications in larger phylogenies like Malpighiales.

715 In Malpighiales, the Reticulation Indices are especially high in deeper parts of the  
716 phylogeny, suggesting that certain clades may contribute substantially to this phenomenon  
717 (Fig. 5c). In particular, we hypothesize that the overabundance of asymmetrical triplets  
718 observed within Clades 1 (MRCA of Euphorbiaceae and Putranjivaceae) and Clade 2 (MRCA  
719 of Salicaceae and Violaceae) result from ancient and persistent gene flow between early  
720 diverging members of these lineages. Specifically, Clade 1 contains six paralogous lineages  
721 from the plastid phylogeny (Xi et al. 2012) and is a major hotspot for plastid-nuclear  
722 conflict. Such conflict is widely recognized as an indicator of introgression (Soltis and  
723 Kuzoff 1995; Baum et al. 1998). Moreover, members of two clades, the putranjivoids and  
724 Pandaceae, have previously been implicated in the top three most unstable nodes of all  
725 angiosperms (Smith et al. 2013). We hypothesize that this may be attributed to the  
726 chimeric nature of their ancestral genealogy resulting from gene flow. The Reticulation  
727 Index is also significantly negatively correlated with species tree resolution (Fig. S7d),

728 suggesting that introgression is an important barrier for robust species tree estimation in  
729 Malpighiales. For the most recalcitrant nodes where almost no bootstrap replicates recover  
730 the same topology, we also observed the highest values of inferred introgression. In the  
731 meantime, no correlation is identified between the estimated level of ILS and introgression,  
732 suggesting that our methods can effectively distinguish ILS and introgression. However,  
733 nodes with strong introgression signals also have higher gene tree estimation error (Fig.  
734 S7e). One possible explanation for such correlation is that the short branch lengths created  
735 by introgression may lead to elevated estimation error at these nodes.

736 To better characterize gene tree variation attributable to ILS, gene tree estimation  
737 error and gene flow, we devised a novel regression method to parse variation attributable  
738 to these analytical and biological factors. Our method of decomposing gene tree variation  
739 revealed that the majority of variation is due to estimation error (52%), while gene flow  
740 and ILS account for 32% and 15%, respectively. This decomposition analysis is based on  
741 estimations of ILS, gene tree error, and gene flow through simulation and is subject to  
742 common limitations of regression analyses. As a result, errors from the simulation and  
743 regression analysis can render the absolute values of these percentages less reliable.  
744 Regardless, the relative influence of biological and analytical aspects of gene tree variation  
745 as interpreted from these metrics can shed important light on empirical investigations and  
746 the development of enhanced species tree inference methods. For example, though gene  
747 tree error is to blame for the majority of gene tree variation in our test case, gene flow still  
748 plays a significant role in gene tree variation. Therefore, a species network inference  
749 method that accommodates gene flow is essential to better understand the early  
750 evolutionary history of Malpighiales. Application of this method to other taxonomic groups

751 will also reveal the key factors contributing to recalcitrant relationships and provide  
752 guidance for phylogenomic marker design targeting at specific questions.

753 Our results suggest that a confluence of factors—ILS, gene tree estimation error, and  
754 gene flow—influence this lack of resolution and contribute to a perfect storm inhibiting our  
755 ability to reconstruct branching order along the back of the Malpighiales phylogeny. Gene  
756 flow, in particular, is a potentially potent, and overlooked factor accounting for this  
757 phenomenon. Despite a relatively small percentage of asymmetrical triplets attributed to  
758 gene flow (1.3% of all triplets), they appear to contribute substantially to gene tree  
759 heterogeneity based on our relative importance decomposition (32%). Our approach of  
760 interrogating this phenomenon using triplet frequencies and the relative importance  
761 analyses can elucidate factors that give rise to gene tree variation. These approaches are  
762 likely to be especially useful for investigating the causes of recalcitrant relationships,  
763 especially at deeper phylogenetic nodes, and to highlight instances where relationships are  
764 better modeled as a network rather than a bifurcating tree.

765

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773

774 **FIGURE CAPTIONS**

775

776 **Figure 1** Simulation of ILS and gene tree estimation error. ILS was simulated through the  
777 coalescent process by setting low (0.01) and high (0.1) theta values. DNA alignments were  
778 subsequently generated through the mutation process based on simulated gene trees. Five  
779 alignments were generated for each gene tree with lengths of 300, 400, 500, 1000, and  
780 1500 bp (only two are shown in the graph). Shorter alignment lengths increase in gene tree  
781 estimation error.

782

783 **Figure 2** Identification of reticulate evolution using triplet frequency. (a) Theoretical  
784 expectations of triplet frequency distribution under the multi-species coalescent (MSC)  
785 model with and without introgression. In case of incomplete lineage sorting (ILS),  
786 symmetrical distributions of the frequency of two minor topologies are expected owing to  
787 deep coalescence (left). In case of introgression, one of the minor topologies will occur with  
788 higher frequency due to gene flow (right). (b) Mapped asymmetrical triplets to species tree  
789 to identify reticulate nodes.

790

791 **Figure 3** Species phylogeny of Malpighiales derived from MP-EST with complete low-  
792 stringency data set (analysis No. 15 in Table S2). Gene trees are estimated using MrBayes.  
793 Branches are colored by the inferred population mutation parameter theta. Warmer colors  
794 indicate higher theta and thus higher level of ILS. Terminal branches are colored grey due  
795 to lack of data to infer theta. BP values from best-resolved MP-EST/ASTRAL/RAxML  
796 analyses (analysis No. 15, 17, and 11 in Table S2) are indicated above each branch; an

797 asterisk indicates 100 BP support; a hyphen indicates less than 50 BP. Branch lengths  
798 estimated from RAxML by fixing the species tree topology are presented at the upper left  
799 corner. The eleven major clades highlighted in the discussion are identified with circled  
800 numbers along each relevant branch. The clade affiliation for each family based on the  
801 plastid phylogeny (Xi et al. 2012) is indicated on the right. Clades identified by Xi et al.  
802 (2012) that are also monophyletic in this study are highlighted using gray shades.

803  
804 **Figure 4** Extensive gene tree discordance in empirical versus simulated data. (a) Gene  
805 tree–species tree (G-S) discordance in the empirical (Emp) and simulated (Sim) data  
806 assuming fixed theta in simulation. Discordance is measured by RF distance between  
807 inferred gene trees and the species tree. Under various simulated conditions of ILS (e.g.,  
808 ‘Low ILS’, theta = 0.01 and ‘High ILS’, theta = 0.1) and gene tree estimation error (‘High ILS  
809 + High err.’, theta = 0.1, alignment length=300bp), the simulated gene tree–species tree  
810 discordance is significantly lower than that from empirical data. (b) Gene tree–species tree  
811 discordance is higher in empirical versus simulated conditions without setting theta *a*  
812 *priori*. For each BP data set, gene tree–species tree discordance is measured and compared  
813 in both empirical and simulated data sets. Positive values indicate higher gene tree–species  
814 tree discordance in our empirical data. (c) Species tree estimation discordance in empirical  
815 data (left) and simulated data (right).

816  
817 **Figure 5** Relative contributions of ILS, estimation error, and gene flow across Malpighiales.  
818 (a) ILS. Nodes are colored by inferred population mutation parameter theta. (b) Gene tree  
819 estimation error. Nodes are colored by BP values, which represent percentage of recovered

820 nodes from simulation (see Materials and Methods). (c) Gene flow. Nodes are colored by  
821 Reticulation Index. (d) Gene tree variation. Nodal BPs reflect nodal recovery in gene trees.  
822 Percentages of gene tree variation ascribed to ILS, estimation error, and gene flow are  
823 indicated by black arrows.

824  
825 **Figure S1** Number of PI sites and mean gene tree BP is positively correlated with  
826 alignment length in high/medium/low-stringency data sets. (a,b) Correlation between  
827 number of PI sites (a) or mean gene tree BP (b) with alignment lengths inferred from the  
828 high-stringency data set. (c,d) Correlation between number of PI sites (c) or mean gene tree  
829 BP (d) with alignment lengths inferred from the medium-stringency data set. (e,f)  
830 Correlation between number of PI sites (e) or mean gene tree BP (f) with alignment lengths  
831 inferred from low-stringency data set. Pearson's  $R^2$  is presented at lower right corner of  
832 each plot.

833  
834 **Figure S2** Increased gene tree and species tree resolution as more flanking sites are  
835 included in the analysis. (a) Distribution of mean gene tree BP in high/medium/low-  
836 stringency data sets. (b,c) Increased species tree BP in concatenation (b) and coalescent  
837 analysis (c). Analyses with same locus subsampling are connected by lines. (d) Increased  
838 species tree inference consistency reflected by pairwise RF distance.

839  
840 **Figure S3** Species tree discordance is more sensitive to site and locus subsampling in  
841 coalescent (black) versus concatenation analyses (grey). Left, distribution of pairwise  
842 species tree distances derived from all coalescent (black) and concatenation analyses (grey)

843 measured by RF distance. Right, distribution of pairwise species tree distances from  
844 coalescent (black) and concatenation (grey) analyses measured by weighted RF (WRF)  
845 distance (weighted by nodal support).

846  
847 **Figure S4** Gene tree estimation error in empirical and simulated data. Gene tree estimation  
848 error is measured by RF distance to the 'true gene tree' for both empirical and simulated  
849 data sets. In both cases, gene tree estimation error is negatively correlated with alignment  
850 length.

851  
852 **Figure S5** Species tree estimation error in simulated data sets. Species tree estimation  
853 error is measured by RF distance from inferred species in each analysis to the known  
854 species tree. Results derived from alignments of varying lengths (300, 400, 500, 1000, 1500  
855 bp) are marked by different color and shape. (a,b) Species tree estimation error of ExaML  
856 under low (a) and high (b) ILS. (c,d) Species tree estimation error of MP-EST under low (c)  
857 and high (d) ILS. (e,f) Species tree estimation error of ASTRAL-II under low (e) and high (f)  
858 ILS.

859  
860 **Figure S6** Relative importance of ILS, gene tree estimation error, and gene flow in  
861 generating gene tree variation based on four regression methods. Percentages are  
862 normalized to sum 100%. 95% confidence intervals are represented by bars.

863  
864 **Figure S7** ILS, gene tree estimation error, and introgression contribute to low species tree  
865 resolution in Malpighiales. Species tree resolution is represented by nodal support from the



866 MP-EST phylogeny in Figure 3 from the main text. The other statistics reflect the variables  
867 presented in Figure 5. The  $p$ -value of the Pearson's correlation test is indicated in the upper  
868 right corner in each panel. (a) Significant negative correlation between gene tree variation  
869 and species tree resolution. (b) Significant negative correlation between ILS and species  
870 tree resolution. (c) Significant negative correlation between gene tree estimation error and  
871 species tree resolution. (d) Significant negative correlation between introgression and  
872 species tree resolution. (e) Significant positive correlation between gene tree estimation  
873 error and introgression. (f) No significant correlation between gene tree estimation error  
874 and ILS.

875

876 **Figure S8** Hotspots of reticulate evolution in baker's yeast. Species phylogeny is inferred  
877 from MP-EST with the 1,070 genes trees in Salichos and Rokas (2013). Nodes are colored  
878 by Reticulation Index. Black thick arrows indicate inferred reticulation by Yu and Nakhleh  
879 (2015) for comparative purpose.

880

881 **Table S1** Voucher and GenBank information for 64 species in Malpighiales, Celastrales,  
882 Oxalidales, and Vitales used for anchored hybrid enrichment.

883

884 **Table S2** Species tree estimation strategies using various phylogenetic estimation methods  
885 and phylogenetic subsampling methods (see Supplementary Note 1) with high-, medium-,  
886 and low-stringency data sets.

887

888 **Table S3** Coalescent and mutational parameters of simulated data sets.

889

890 **Table S4** Summary statistics of 423 loci in high/medium/low-stringency data sets,  
891 including number of captured taxa, alignment length, number of PI sites, and mean gene  
892 tree BP.

893

894

895

## 896 **LITERATURE CITED**

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