Reinstatement of cortical outcome representations during higher-order learning

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Abstract

Naturalistic learning scenarios are characterized by infrequent experience of external feedback to guide behavior. Higher-order learning mechanisms like second-order conditioning (SOC) may allow stimuli that were never experienced together with reinforcement to acquire motivational value in such environments. Despite its explanatory potential for real-world learning phenomena, surprisingly little is known about the neural mechanism underlying such associative transfer of motivational value in SOC. Here, we used multivariate cross-session, cross-modality searchlight classification on functional magnetic resonance imaging data obtained from healthy humans during SOC. We show that visual first-order conditioned stimuli (CS) reinstate cortical patterns representing previously paired gustatory outcomes in the lateral orbitofrontal cortex (OFC). The same OFC region was found to increase functional coupling with amygdala, anterior hippocampus, and medial OFC during SOC. In the amygdala, neural pattern similarity specifically increased between second-order CS and outcomes from early to late stages of SOC. Our data suggest a mechanism by which value is conferred to stimuli that were never paired with reinforcement.

Introduction

Learning in naturalistic settings is characterized by infrequent direct encounters with rewarding or punishing stimuli (Gewirtz and Davis, 2000). Hence, several stimuli or actions often need to be chained together, such that one stimulus serves as a proxy for another stimulus that might eventually predict reward (Gewirtz and Davis, 2000). Exploiting such statistical regularities would allow agents to learn about the value of stimuli or actions that were never directly followed by reinforcement and support instrumental behavior in the absence of reinforcement or in new contexts. Second-order conditioning (SOC) is an example of higher-order learning that allows exactly such associative transfer to stimuli that have never been directly paired with reinforcement (Sharpe et al., 2017). In SOC, a first-order conditioned stimulus (CS_1) is first paired with a motivationally salient event or stimulus (unconditioned stimulus, US). By virtue of this pairing, the CS_1 is capable of evoking conditioned responses (CR, e.g. salivating, as in the presence of food), enabling it to function as conditioned reinforcer (Sharpe et al., 2017). Subsequently, another previously neutral conditioned stimulus (CS_2) is paired with the CS_1. Thereby, CS_2 acquires incentive properties and is, like CS_1, now able to elicit a CR. Despite its relevance for real-life learning phenomena (Gewirtz and Davis, 2000; Parkes and Westbrook, 2011) and decades of behavioral investigations (Barnet et al., 1991; Gewirtz and Davis, 2000; Parkes and Westbrook, 2011; Rizley and Rescorla, 1972), surprisingly little is known about the neural mechanism underlying associative transfer of motivational value in SOC (Parkes and Westbrook, 2011).

At least four potential mechanisms accounting for SOC have been proposed (Barnet et al., 1991; Gewirtz and Davis, 2000; Parkes and Westbrook, 2011; Rizley and Rescorla, 1972). Three of these suggest that CS_2 could become associated with a CR using (i) the associative link between CS_2 and CS_1, (ii) direct pairing of the CS_1-evoked CR with CS_2, or (iii) because CS_1 reactivates a representation of the US which then evokes a CR that becomes paired with CS_2. Behavioral evidence from previous studies suggests that neither of these three hypotheses can account for second-order learning effects (Barnet et al., 1991;

Rizley and Rescorla, 1972). This leaves open a fourth possibility: that CS_2 is directly paired with a neural representation of the US, or more precisely, with the motivational state conveyed by the US during second-order learning ("direct link" hypothesis) (Gewirtz and Davis, 2000; Parkes and Westbrook, 2011). Since the "direct link" hypothesis of second-order conditioning is the only remaining explanatory attempt and has never been tested to date, the present study sought to provide evidence that CS_2 is directly paired with a neural representation of the US. Recent evidence suggests that acquisition of incentive properties by second-order stimuli (CS_2) is related to activity of mesencephalic dopamine neurons that encode a reward prediction error elicited by CS_1 (Maes et al., 2020). Presumably, SOC critically depends on convergence of neural patterns representing CS_2 and US in the amygdala (Gewirtz and Davis, 2000, 1997; Parkes and Westbrook, 2011; Setlow et al., 2002) and hippocampus (Gilboa et al., 2014). However, it remains unclear, how exactly a "direct link" between CS_2 and US representations is formed at the neural level. Additionally, to the best of our knowledge, it is unknown whether human value-based decision-making is affected by motivational value acquired through SOC. This complicates cross-species translation of higher-order learning effects described in animal studies and highlights the need for using similar paradigms in humans.

Building on models of memory reinstatement (Tonegawa et al., 2018), we hypothesized that during SOC, presentation of the CS_1 would trigger reinstatement of the neural pattern representing the US with which it had previously been paired during first-order conditioning (FOC). This would allow linkage of "online" US and CS_2 representations by associative plasticity in amygdala (Gewirtz and Davis, 2000, 1997; Hatfield et al., 1996; Parkes and Westbrook, 2011; Setlow et al., 2002) and hippocampus (Gilboa et al., 2014). Both amygdala and hippocampus, as well as orbitofrontal cortex (OFC) and ventral striatum (Mcdannald et al., 2013) have consistently been shown as the key structures involved in second-order learning and conditioned reinforcement (Gewirtz and Davis, 2000, 1997; Gilboa et al., 2014; Hatfield et al., 1996; Parkes and Westbrook, 2011; Setlow et al., 2002). Our analyses therefore focused on these regions.

To test our predictions, we combine a SOC paradigm and choice preference tests

with cross-session (Stokes et al., 2009), cross-modality searchlight (Kriegeskorte et al.,

2006) classification of functional magnetic resonance imaging (fMRI) data in healthy human

participants. Participants first established Pavlovian associations between visual CS_1 and

appetitively or aversively valued gustatory US. During SOC, participants were exposed to

associations between visual CS_2 and previously learned first-order CS_1. In a subsequent

preference test phase, participants were more likely to select directly (CS_1) and indirectly

(CS_2) appetitively paired stimuli over aversively paired stimuli. These behavioral associative

transfer learning effects were accompanied by first-order CS-related reinstatement of the

neural patterns representing US in the lateral OFC and increased functional coupling

between lateral OFC, amygdala/anterior hippocampus, and medial OFC during SOC.

Furthermore, representations of second-order CS in the amygdala became more similar to

US representations from early to late phases of second-order conditioning, indicating the

acquisition of an association between CS_2 and US patterns.

Results

Behavioral evidence for transfer of motivational value during second-order

conditioning

Participants (N = 49, combined across two experiments) first established Pavlovian visuo-

gustatory associations between a visual CS^+_1 and an appetitive gustatory US^+ (chocolate

milk or orange juice), and between a visual CS^-_1 and an aversive gustatory US^- (quinine-HCl)

during first-order conditioning (FOC, Fig. 1C). During SOC, participants were visually

presented with CS_2^+ followed by CS_1^+, CS_2^- followed by CS_1^-, and a novel pairing of CS_2^n

followed by CS_1^n (Fig. 1D). Since CS_1^n had not been presented during first-order

conditioning, both CS_2^n and CS_1^n were considered neutral CS, serving as control stimuli for

mere exposure effects and for correction of general visual activation effects in fMRI analyses

(see Methods for details). Importantly, participants were not informed about the underlying

associative structure of the tasks but were instructed to perform simple attentional control

Fig. 1. Experimental procedure, task schematic and behavioral results.

A) Experimental procedure: Participants in the fMRI study (N = 29) performed both day 1 and day 2 (B-E), participants in the behavioral study (N = 20) only performed the procedures of day 2 (C-E, outside fMRI). B) Classifier training experiment: After rating of subjective value and intensity of both gustatory US (US$^+$: orange juice and US$^-$: quinine-HCL 0.1 mmol/l solved in purified water), in a total of five runs each US was administered twenty times (40 trials per run, 200 trials in total). Per trial, one US (1 ml bolus per trial) was delivered by a gustometer and infusion lines. The US bolus onset was preceded by a blue square. Trials were separated by an inter-trial-interval (ITI) marked by a grey screen. Participants performed a 0-back-style attentional control task. In 20% pseudo-randomly selected trials, participants were presented with probe trials in which they were asked to indicate which US they had received last ("pleasant" (US$^+$) or "unpleasant" (US$^-$)). C) First-order conditioning: In each trial, a CS_1 was followed by an inter-stimulus interval marked by a fixation cross, and oral infusion of one US (1 ml bolus per trial). Trials were separated by an ITI marked by a grey screen. Each CS_1 was presented 50 times, amounting to 100 trials total. CS_1 were followed by a US with 80% probability. D) Second-order conditioning: In each trial, a CS_2 was followed by an inter-stimulus interval marked by a fixation cross, and a CS_1. CS_2 were followed by a CS_1 fully deterministically. Each CS_2-CS_1 pair was presented 50 times (150 trials total). E) Choice preference test: Following SOC, participants were presented with two separate test phases consisting of repeated binary choices between pairs of CS_1 (right) and pairs of CS_2 (left) to assess behavioral signatures of first- and second-order conditioning. F) Behavioral results (pooled across fMRI and behavioral study, N = 49): Raincloud plots (Allen et al., 2019) showing density of choice probability. CS_1^+ (dark blue, 89%-highest posterior density interval (89%-HPDI): [.58; .65]) and CS_2^+ (light blue, 89%-HPDI: [.60; .67]) were preferred over CS_1^- and CS_2^-, respectively. There was no evidence for choice probabilities of CS_1^n and CS_2^n being different from chance level in novel-to-lure comparison choice trials (dark and light gray, 89%-HPDIs: [.46; .52] and [.42; .48]). Box plot center lines represent sample medians and box bottom/top edges show 25th/75th percentile of the data, respectively.

tasks. This was aimed at leaving participants unaware of the associative learning process.

Post-experimental tests for explicit knowledge indeed revealed that participants were

unaware of the (indirect) associations between CS_2 and US (Supplementary Results and

Supplementary Table 1). In a subsequent choice preference test (Fig. 1E), participants

performed binary decisions between pairs of CS_1 and CS_2. Choice trials were interspersed

with lure decisions between CS_1^n or CS_2^n and novel fractals or kanjis, respectively, that had

only been seen during pre-task rating. We reasoned that choice probabilities (CP) should

exceed the indifference criterion (CP > 0.5) if associative direct (CS_1) and associative

transfer (CS_2) learning effects occurred. Additionally, we assumed that choice probability

should not exceed the indifference criterion in trials involving novel-to-lure stimuli, if choice

behavior was not exclusively driven by mere exposure or novelty-dependent response

biases. Combined across the two experiments, participants showed a preference both for the

appetitive first- and second order stimuli, CS_1^+ and CS_2^+, respectively, over the aversive CS_1^-

and CS_2^- (Fig. 1F). Highest posterior density intervals (HPDI) of choice probabilities (binomial

model) for CS_1^+ versus CS_1^- [.58; .65] and for CS_2^+ versus CS_2^- [.60; .67] did not overlap (0%

overlap) with the defined region of practical equivalence (ROPE, CP = [.45; .55]), an interval

of parameter values representing the null hypothesis of no difference from chance level.

These findings show that our conditioning procedure reliably induced both direct associative

(CS_1) and associative transfer learning effects (CS_2). Importantly, there was no evidence for

both choice probabilities of CS_1^n and CS_2^n being different from chance level in novel-to-lure

comparison choice trials (Fig. 1F, HPDIs: [.46; .52] and [.42; .48], 100% and 48% ROPE

overlap, respectively). This pattern of results rules out an explanation of CS_1^+ and CS_2^+

preference arising from mere-exposure or novelty. Importantly, very similar results were

obtained when repeating the analyses separately for the behavioral and fMRI sample

(Supplementary Figure 1).

First-order CS reinstate neural US patterns during second-order conditioning

After establishing behavioral evidence for associative transfer learning, we reasoned that a

prerequisite for the observed learning effect would be reinstatement of the neural US pattern

by the paired CS_1 to establish a direct associative link between CS_2 and US. Importantly, it

should be noted that US reinstatement during SOC could only be faithfully attributed to the

respective CS_1, but not to CS_2, since only CS_1 had been directly paired with the US, and CS_2

had not previously been experienced. To test for cortical reinstatement of neural patterns

representing US during second-order conditioning, we used functional magnetic resonance

imaging (fMRI) and a cross-session (Stokes et al., 2009), cross-modality searchlight

(Kriegeskorte et al., 2006) classification approach (multivariate pattern analysis, MVPA). To

obtain unbiased estimates of the neural patterns representing our gustatory US, without the

confounding influence of associations to a learned first-order CS, we first performed an fMRI

classifier training experiment on day 1 (Fig. 1B) during which the US^+ and US^- were

presented. A multivariate pattern classifier (linear support vector machine, C-SVM) was

trained on the fMRI data from this session. On day 2, during SOC, we used the weights of

the classifier trained on gustatory neural patterns to predict the class label of the visual CS

(Fig. 1D) that had been paired with the US during FOC (Fig. 1C, see Methods and Fig. 2A for

schematic of the classification approach). In other words, we used data from one sensory

modality, assessed during a first day to train a classification model and tested the model's

generalizability to unseen data from another sensory modality on a second day. We found

evidence for reinstatement of US patterns in a region in the left lateral orbitofrontal cortex

(lOFC). In this region, it was possible to predict the class labels of CS neural patterns based

on US pattern information obtained during classifier training on the previous day.

Classification with a 3-mm searchlight revealed a small-volume corrected (Neurosynth

(Yarkoni et al., 2011) meta-analytic functional mask for term "taste") cluster of above-chance

level (0.5) classification accuracy (extracted cluster mean = 0.56, SD = 0.07) in the left

lateral OFC (lOFC, Fig. 2B, peak voxel at MNI [x = –21, y = 30, z = –17], z = 2.26, P = 0.012,

random-effect cluster corrected (Stelzer et al., 2013), 50,000 iterations, one-tailed). The

location of this lOFC cluster is consistent with this region's well-documented role in gustatory

processing, particularly in representing motivational (Rolls, 2006, 2000; Small et al., 1999)

aspects of gustatory sensation and taste memory (Kobayashi et al., 2004). To ensure that

our results are not dependent on this particular choice of ROI, we repeated the same

analysis using two different, independent ROIs. The first was an anatomical mask of lateral orbitofrontal cortex (Harvard-Oxford atlas), the second was obtained from an independent gustatory mapping study by Benz and colleagues (K. Benz, personal communication, 12/2019). Again, we found significant CS decodability in both ROIs (Supplementary Results).

When we split participants into a high- and low-bias group, depending on their preference for the CS_2^+, we observed that significant classification of CS labels in lOFC was possible in the high bias group ($Z = 3.43$, $P = .006$, $U3_1 = .88$, one-sample Wilcoxon signed-rank test, two-tailed), but not in the low bias group ($Z = 1.44$, $P = .151$, $U3_1 = .58$, one-sample Wilcoxon signed-rank test, two-tailed, Fig. 2C).

Interaction between lOFC, amygdala, and medial OFC during second-order conditioning

To form an associative link between CS_2 and US, the reinstated US patterns need to be projected from their cortical storage site to regions like amygdala and hippocampus, allowing for convergence of US and CS_2 information. We investigated whole-brain BOLD signal covariation of the cluster in the left lOFC with the whole brain during SOC using two separate psychophysiological interaction (PPI) analyses. In the first PPI, we investigated general covariation differences in response to CS^- and CS^+ relative to CS^n, reasoning that in the former, but not in the latter, reinstated US representations need to be linked with the CS. To this end, we contrasted CS^- and CS^+ trials with CS^n trials., where the reinstated US is presumably paired with CS_2. We found that covariation of BOLD signal in the left lOFC with a cluster in the left hippocampus, extending to amygdala and medial temporal lobe (Fig. 3A, peak voxel at MNI [x = –38, y = --11, z = –24], $Z = 3.75$, $P = 0.003$, whole-brain corrected) and a cluster in the right inferior temporal gyrus, extending to the right hippocampus (peak

Figure 2. Cross-session, cross-modality classification analysis.
A) Analysis approach: For cross-session, cross-modality classification, a multivariate classifier was trained on the spatial activation patterns (*t* maps) of US^+ versus US^- and tested on CS_2^+/CS_1^+ versus CS_2^-/CS_1^- (CS_2^n/CS_1^n related activation were subtracted to control for general visual effects) during SOC within 3-mm searchlight spheres across the whole brain, resulting in one whole-brain map of classification accuracies per participant. Additionally, we repeated this procedure 100 times per participant with randomly permuted class labels in the training data set (US^+ versus US^-) to create 50,000 chance level group maps for random-effect cluster-statistics(Stelzer et al., 2013). B) Classification accuracy exceeded chance level (0.5) in a small-volume corrected cluster in the left lateral orbitofrontal cortex (lOFC, peak voxel at [x = −21, y = 30, z = −17], *z* = 2.26, *P* = 0.012, one-tailed, random-effect cluster-statistics, 50,000 iterations). C) Average classification accuracy in the lOFC cluster (extracted cluster mean = 0.56, *SD* = 0.07) was significantly above chance level in the high bias group of participants showing higher than chance preference for CS_2^+ (dark purple raincloud, *Z* = 3.43, *P* = 0.006, $U3_1$ = 0.88, one-sample Wilcoxon signed-rank test, two-tailed), but not in the low bias group (light purple raincloud, *Z* = 1.44, *P* = 0.151, $U3_1$ = 0.58, one-sample Wilcoxon signed-rank test, two-tailed). Color bar represents *Z*-values. Black dots indicate sample means and error bars represent standard errors of the means of the data, respectively.

voxel at MNI [x = 40, y = --32, z = −20], Z = 3.55, *P* = 0.007, whole-brain corrected) was

higher in CS^- and CS^+ compared to CS^n trials. Despite the fact that we employed an

approach with high statistical sensitivity for group-level cluster-based inference (FSL's

FLAME1+2 (Beckmann et al., 2003)), a robust estimation technique that yields low family-

wise errors rates even at liberal activation thresholds (*Z* > 2.3, corresponding to *p* < 0.05)

(Eklund et al., 2016), we would like to note that the aforementioned result did not survive

whole-brain correction at more conservative thresholds (e.g. *Z* > 2.6). Furthermore, in the

same PPI analysis, we found a positive correlation between second-order choice preference

and functional covariation of the left lOFC with a region in the right and left lateral prefrontal

cortex (Supplementary Fig. 3, right peak voxel at MNI [x = 49, y = 41, z = 7], Z = 3.80, *P* <

0.001, whole-brain corrected; left peak voxel at MNI [x = −45, y = 46, z = −4], Z = 3.93, *P* =

0.010, whole-brain corrected). The more participants preferred CS_2^+ (versus CS_2^-), the

stronger these regions' BOLD signal covaried during SOC.

Some studies hint at complex interactions between OFC and amygdala that differ

between appetitive and aversive stimuli (Morrison et al., 2011). Therefore, we next asked

whether there were any specific differences between CS^- and CS^+ trials in lOFC BOLD signal

covariation with other regions. We found that covariation of the BOLD signal in the left lOFC

was higher in CS^+ compared to CS^- trials in a cluster located in the medial OFC, extending to

subgenual anterior cingulate cortex (Fig. 3B, peak voxel at MNI [x = 10, y = 50, z = –1], Z =

3.90, P < 0.001, whole-brain corrected) and a cluster in the left anterior insula, extending to

left caudo-lateral OFC and temporal pole (peak voxel at MNI [x = –51, y = 18, z = –11], Z =

3.33, P = 0.004, whole-brain corrected). Given both the difficulties in inferring directionality

from fMRI data and the reciprocal nature of connections between lOFC and both amygdala

and medial OFC, a directionality of our PPI results cannot be determined.

Plasticity of association between second-order CS and US in the amygdala

We reasoned that if the amygdala uses the reinstated cortical outcome pattern in lOFC to

acquire an association between CS_2 and the respective US, one would expect similarity

between the neural patterns evoked by CS_2 and US, respectively. This similarity should

increase over the course of second-order conditioning, as the association is being acquired.

Importantly, during the classifier training experiment, there was evidence for US

representations in the amygdala. We found differential neural patterns in response to US^-

and US^+ in the amygdala (among other regions) using an exploratory leave-one-run-out

cross-validated classification analysis on the neural pattern recorded during the classifier

training experiment (Supplementary Fig. 4). Using a least-squares separate (LS-S) approach

(Mumford et al., 2012) to deconvolve single-trial estimates, we first computed overall neural

pattern similarity (Kriegeskorte et al., 2008) between the pattern evoked by CS_2 (during SOC)

and their respective US (during classifier training) in the bilateral amygdala. There was

significant pattern similarity between CS_2^- and US^- (t_{28} = 3.38, P = 0.002, $U3_1$ = 0.79; but not

Figure 3. Psychophysiological interaction analyses and neural pattern similarity results.
A) BOLD covariation was higher in CS^- and CS^+ trials than in CS^n trials between the left lOFC and a cluster in the anterior hippocampus, extending to (basolateral) amygdala and medial temporal lobe (peak voxel at MNI [x = −38, y = −11, z = −24], Z = 3.75, P = 0.003, whole-brain corrected) and a cluster in the right inferior temporal gyrus, extending to the right hippocampus (peak voxel at MNI [x = 40, y = −32, z = −20], Z = 3.55, P = 0.007, whole-brain corrected). B) Covariation of BOLD signal in the left lOFC was higher in CS^+ trials compared to CS^- trials in a cluster located in the medial OFC, extending to subgenual anterior cingulate cortex (peak voxel at MNI [x = 10, y = 50, z = −1], Z = 3.90, P < 0.001, whole-brain corrected) and a cluster in the left anterior insula, extending to left caudo-lateral OFC and temporal pole (peak voxel at MNI [x = −51, y = 18, z = −11], Z = 3.33, P = 0.004, whole-brain corrected). C & D) Template-based neural pattern similarity between classifier training experiment patterns of both US and least-squares separate estimated, pooled (early 25 and late 25 SOC trials) single-trial estimates of neural patterns representing CS_2 (C) and CS_1 (D), corrected for the respective pooled CSn neural patterns in a bilateral amygdala ROI. C) CS_2^- and US^- patterns became more similar from early to late trials (t_{28} = 1.88, P = 0.035, Cohen's d = 0.35, paired-samples t-test, one-tailed). However, there was also no evidence for a difference between early and late trial similarity for CS_2^+ and US^+ (t_{28} = 0.50, P = 0.310, Cohen's d = 0.09, paired-samples t-test, one-tailed). D) There was no evidence for change in similarity for CS_1^--US^- (t_{28} = 0.50, P = 0.310, Cohen's d = 0.09, paired-samples t-test, one-tailed) or CS_1^+ and US^+ neural pattern similarity (t_{28} = 0.50, P = 0.310, Cohen's d = 0.09, paired-samples t-test, one-tailed). Black dots indicate sample means and error bars represent standard errors of the means of the data, respectively. Color bars represent Z-values.

between CS_2^+ and US^+, t_{28} = 0.74, P = 0.464, $U3_1$= 0.55). We reason that this could be due

to differential motivational aspects of the gustatory outcomes we used. The US^- consisted of

an aversively tasting quinine solution, which has been shown to produce innate avoidance

responses that are not thought to be generated by learning (Yiannakas and Rosenblum,

2017). In contrast, US^+ was orange juice, a compound stimulus that consists of many taste

features (e.g. sour and fruity components), next to sweet sensation, which also produces

innate responses (Yiannakas and Rosenblum, 2017). Presumably, this heterogeneity of taste

facets in US^+ had induced more variability in taste responses in our participants. Additionally,

it appears likely that in a society offering sugary and highly palatable foods and drinks in

abundance, US^+ (an orange juice, most likely a very familiar taste to most of our participants)

might not have been perceived as rewarding, or novel enough to represent a motivationally

relevant and salient event for our participants. It might well be that participants did not have

experience with the aversive taste of quinine, rendering it a novel and salient gustatory

sensation, presumably resulting in only the US^- being considered behaviorally and

motivationally relevant. Thus, it is possible that the observed preference for CS_1^+ and CS_2^+

was indeed due to an aversion against CS_1^- and CS_2^-.

If the observed neural pattern similarity between CS_2 and US indeed reflects the

formation of an associative link during second-order learning, one would expect the observed

similarity to increase from early to late stages of learning, as the association is being

acquired and the CS comes to predict the US. We therefore compared our measure of neural

pattern similarity between early and late trials of SOC. As expected, neural pattern similarity

between CS_2^- and US^- in the bilateral amygdala ROI increased from early to late trials of

SOC ($t_{28} = 1.88$, $P = 0.035$, Cohen's $d = 0.35$, paired-samples t-test, one-tailed, Fig. 3C).

This change in similarity was not observed for the first-order CS_1^+-US^+ or CS_1^--US^- pattern

similarity (Fig. 3D), nor for second-order CS_2^+-US^+ similarity (all $P > .310$, paired-samples t-

tests, one-tailed, Fig. 3C).

Discussion

Using a second-order conditioning paradigm, we show associative transfer effects during

higher-order conditioning, even when participants are unaware of the underlying associative

structure of the experiment. Participants were more likely to select directly and indirectly

appetitively paired stimuli over aversively paired stimuli, closely resembling rodent studies

describing choice biases consistent with second-order conditioning (Maes et al., 2020;

Sharpe et al., 2017). This suggests that humans, similar to rodents, implicitly acquire

preferences through higher-order transfer learning mechanisms – a finding that goes beyond

previous studies promoting acquisition of explicit associative relationships between stimuli

(Jara et al., 2006; Pauli et al., 2019; Wang et al., 2020; Wimmer and Shohamy, 2012).

Notably, choice biases for second-order conditioned stimuli emerged in the absence of

explicit knowledge of the underlying higher-order associative structure. To the best of our

knowledge, this is the first report demonstrating that human value-based decision making is

affected motivational value implicitly conferred via second-order conditioning. The

associative transfer learning observed in the present study may be beneficial, as it allows an

agent to exploit the underlying relational structure and assign value to states for which direct

experience with reinforcement is not available. It may however also be maladaptive when it

leads to pursuit of suboptimal options, for example if states only share contiguous

occurrence but lack correlated reward-predictive properties.

Using a multivariate cross-session, cross-modality searchlight classification approach,

we found reinstatement of neural outcome representations by visual first-order CS that had

previously been directly paired with gustatory outcomes during first-order conditioning. This

reinstatement of outcome patterns occurred in a region in the lateral OFC. Notably, the exact

same rostrolateral portion of OFC has previously been implicated in representing stimulus-

outcome associations (Jocham et al., 2016; Klein-Flügge et al., 2013; Luettgau et al., 2020)

and in correctly assigning credit for a reward to the causal stimulus choice (Jocham et al.,

2016). It is also part of the network most consistently involved in taste processing as

obtained from the Neurosynth data base (Yarkoni et al., 2011) and plays a well-documented

role in representing motivational (Rolls, 2006, 2000; Small et al., 1999) aspects of gustatory

sensation and taste memory (Kobayashi et al., 2004). Importantly, we had trained the

classifier on the gustatory US prior to pairing them with visual first-order CS (on a separate

day). Thus, decoding of the visual CS during second-order learning cannot be spuriously driven by visual responses elicited by the US or the visually presented swallowing cue during classifier training. Importantly, even though the classifier was tested on the entire activation pattern elicited by the first-/second-order CS pairs, US reinstatement during SOC can only be faithfully attributed to the respective CS_1, since only CS_1, but not CS_2, had been directly paired with the outcome and should thus be capable of eliciting a US representation (Wagner, 1981). In "direct link" models of second-order conditioning, reinstatement of the neural pattern representing the outcome constitutes a necessary prerequisite to establish an associative link between second-order CS and the motivational state conveyed by gustatory outcomes. Our findings are consistent with a body of lesion studies in animals that indicate a key role of OFC in second-order conditioning and in responding for conditioned reinforcement (Cousens and Otto, 2003; Gallagher et al., 1999). They also align with the account of OFC representing abstract relational structure, or a state space (Niv, 2019; Wikenheiser and Schoenbaum, 2016) and recent evidence suggesting that OFC neurons encode relevant second-order combinations of past state variables to guide decision-making (Nogueira et al., 2017).

OFC clearly does not act in isolation. A plethora of studies suggests that interactions between OFC, amygdala and ventral striatum (Baxter et al., 2000; Cousens and Otto, 2003; Fiuzat et al., 2017; Morrison et al., 2011) might also support second-order learning processes. We found evidence for the region of OFC that showed reinstatement of outcome representations to increase BOLD signal covariation with amygdala, anterior hippocampus, and medial OFC during second-order learning. Neurons in OFC and amygdala show complex, bi-directional interactions during acquisition and reversal of outcome-predictive associations (Morrison et al., 2011) and functional disconnection of OFC and amygdala using asymmetric lesions produce deficits in flexibly adjusting behavior to changes in stimulus value (Baxter et al., 2000; Fiuzat et al., 2017). We also observed that representations of visual second-order CS in amygdala became more similar to gustatory US representations from early to late phases of second-order conditioning. This indicates the gradual

development of an associative link between CS_2 and US in this region. Indeed, both

amygdala and hippocampal lesions impair second-order learning (Gewirtz and Davis, 1997;

Gilboa et al., 2014). The hippocampus may play a specific role, since animals with lesions

show impaired acquisition of second-order contingencies, while acquisition of first-order

stimulus-outcome associations is spared (Gilboa et al., 2014).

Our study has at least two implications for current theories of learning and decision-

making. First, our study is – to the best of our knowledge – the only report so far of

behavioral evidence for associative transfer of motivational value during human second-order

conditioning. Although there is a rich literature on second-order conditioning, direct evidence

for this phenomenon in humans using similar procedures as in previous animal work was

lacking. The present study thus enables a closer comparison between humans and other

species during higher-order learning. Similar to the present results, a recent fMRI study in

humans employed a sequential Pavlovian conditioning paradigm to provide initial evidence

that Pavlovian stimulus-stimulus associations might be learned and represented in a model-

based fashion by the lateral OFC (Pauli et al., 2019). This result is consistent with evidence

that rat midbrain dopamine neurons encode both reward and identity prediction errors

(Takahashi et al., 2017). However, in the paradigm used by Pauli and colleagues (Pauli et

al., 2019) two CS and a US are presented in close temporal proximity (or even overlapping)

within the same trial. This differs from our study in two important ways. Firstly, unlike SOC,

this procedure does not require any reinstatement of US representations, nor any transfer of

value, but rather imposes different temporal offsets between CS and US. This contrasts with

SOC, where the CS_2 is presented in a separate block that is entirely devoid of US

presentations. Secondly, for the same reasons, it is difficult to dissociate the activation

patterns related to the proximal CS (CS_1 in second-order conditioning) and the US. In

contrast, in the present study, the classifier was trained on the (gustatory) US on a separate

day and then tested on the (visual) CS presented during the second-order learning phase of

the SOC paradigm.

A second implication of our study pertains to the neural mechanisms of associative

transfer learning in SOC, which thus far have remained elusive. Our results suggest that,

during second-order conditioning, outcome representations, and presumably features

pertaining to the motivational value of outcomes, are reinstated in the lateral OFC. Our data

suggests that reinstated outcome representations could be communicated between lOFC

and amygdala/anterior hippocampus for associative linking of previously neutral stimuli with

reinstated representations of the outcomes. It should be noted that the present design does

not allow us to dissociate whether the reinstated US representations detected by the

multivariate classification algorithm pertain to the reinforcement value, the intensity or

sensory features of the US – or a combination of these features. However, the question

which of these features was reinstated by CS_1 and contributed to the associative transfer of

motivational value to CS_2 is beyond the scope of the present study.

It has been proposed that relational knowledge of events and behavior characterized

by considering the state-transition probabilities of the environment might emerge from

associative learning mechanisms (Wang et al., 2018). Second-order conditioning can be

thought of as an instantiation of such a learning process in which a currently absent, but

previously experienced and motivationally relevant task state (receipt of an outcome) is

reinstated in lOFC to enable propagation of motivational relevance to task states that are

never directly experienced in contiguity with reward or punishment. Higher-order learning

paradigms impose a relational structure of events on agents and thus could be cast as an

inference problem, very much like in transitive inference (A→B, B→C, hence A→C) (Dusek

and Eichenbaum, 1997). Credit assignment processes of this kind have been shown to

involve the lOFC (Jocham et al., 2016; Walton et al., 2010). The lOFC reinstatement of

outcome representations observed in the present study could thus pose a necessary

prerequisite for value attribution to task states or stimuli that have never been directly

followed by an outcome, but nevertheless possess outcome-predictive properties.

How the reinstatement of cortical US patterns found in our study relates to the

observation in rats that midbrain dopamine neurons acquire temporal difference error signals

in response to CS$_2$ (Maes et al., 2020) presents an important question for future studies.

Furthermore, it would be of great interest to elucidate the directionality and exact content of

information flow between lOFC and amygdala/anterior hippocampus, or between lOFC and

medial OFC, and whether this transfer of information is supported by phase coherence in

theta oscillations (Benchenane et al., 2010; Knudsen and Wallis, 2020; Young and Shapiro,

2011).

Taken together, our data lend support to the "direct link" hypothesis of second-order

conditioning (Barnet et al., 1991; Gewirtz and Davis, 2000; Rizley and Rescorla, 1972),

stating that second-order CS and representations of outcomes – events that had never been

explicitly paired during the individual's learning history – can be linked by exploiting the

relational structure of events. Importantly, the present study enables a closer comparison

between humans and other species during higher-order learning. In conclusion, our results

suggest a neural mechanism by which outcome representations can be propagated to stimuli

that are never experienced in contiguity with reinforcement, allowing for credit assignment in

real-world learning scenarios with infrequent direct encounters with rewards or punishments.

Methods

Participants

Participants were recruited from the local student community of the Otto-von-Guericke University Magdeburg, Germany by public advertisements and via online announcements. Only participants indicating no history of psychiatric or neurological disorder, no regular intake of medication known to interact with the central nervous system and participants not reporting quinine intolerance were included. Participants in both samples had normal or corrected-to-normal vision and did not report experience with Japanese kanjis or Chinese characters. All participants provided informed written consent before participation and received monetary compensation for taking part in the study. The study was approved by the local ethics committee at the medical faculty of the Otto-von-Guericke University Magdeburg, Germany (reference number: 101/15) and conducted in accordance with the Declaration of Helsinki.

32 healthy adult volunteers (age: M = 24.16, SD = 3.61 years, 15 males) participated in the fMRI study, 20 healthy adult volunteers (age: M = 23.40, SD = 3.07 years, 9 males), participated in the behavioral study. In the fMRI study, two participants were excluded from statistical analyses due to self-reports of having fallen asleep during the second-order conditioning scanning run. One additional participant had to be excluded due to a scanner malfunction and corruption of three of the five classifier training experiment scanning runs, thus leaving a total of N = 29 participants for final analyses.

Learning experiment – ratings

The learning experiment was performed on day 2 in the fMRI sample, in the behavioral sample, only the learning experiment was performed in a single session. Participants received written instructions for the experiment and were instructed once again on the computer screen. All experiments were programmed in MATLAB 2012b (v8.0.0.783, MATLAB and Statistics Toolbox Release 2012b, The MathWorks, Inc., Natick, MA, USA), using Psychophysics Toolbox (Brainard, 1997) (version 3). Before and after the learning

experiment, participants rated ten different round, greyscale fractal images (300 x 300 pixels) serving as first-order conditioned stimuli (CS_1), ten white Japanese kanjis (Tamaoka et al., 2017) (250 x 250 pixels) serving as second-order conditioned stimuli (CS_2) and gustatory stimuli (aversive: quinine-HCl 0.2 mmol/l solved in purified water, appetitive: either chocolate milk (Nesquik, Nestlé, Switzerland) or orange juice (Milde Orange, EDEKA, Germany) in the behavioral study, only orange juice in the fMRI study, serving as unconditioned stimuli (US). Each participant received the same kind and amount of US per trial during the experiment. Ratings of subjective value/liking were assessed with the number buttons on a German (QWERTZ) computer keyboard from 1 (not liked) to 9 (very much liked). In the fMRI sample, participants additionally rated gustatory stimuli regarding their subjective intensity levels from 1 (low intense) to 9 (very intense). Three fractals and kanjis rated closest to 5 (equivalent to "neutral") were selected for first- and second order conditioning and their order was randomized before being associated with the US in first-order conditioning or with the first-order CS_1 in second-order conditioning. To ensure motivational salience of the gustatory US, participants were instructed to abstain from food for 12 hours and on average reported having fasted for 13.28 (SD = 2.71) hours. Participants reported intermediate levels of hunger before the task on a paper-pencil visual analog scale (VAS), ranging from 0 ("not hungry") to 100 ("very hungry"), M = 58.38 (SD = 29.61).

Learning experiment – first-order conditioning

In first-order conditioning (FOC), participants were presented with CS^+_1 followed by the appetitive US^+, and CS^-_1 followed by the aversive US^-. In a typical trial, a CS_1 (2000 ms) was followed by an inter-stimulus interval (1000 ms) marked by a fixation cross, and oral infusion of one US (1 ml bolus per trial). Each CS_1 was presented 50 times, amounting to 100 trials total. CS_1 were followed by a US with 80% probability (40 trials of each CS_1-US pair, 10 trials of CS_1-no US per CS). US were delivered by a MATLAB code-controlled custom-made gustometer consisting of two high pressure single syringe pumps (AL-1000HP, World Precision Instruments, Saratoga, FL) operating 50 ml Luer lock syringes. Syringes were

attached to Luer lock infusion lines (1,40 m length, 2 mm inner diameter) that participants held centrally in their mouths like drinking straws. In the fMRI study, infusion line position order (Q-O (N = 17) and O-Q (N = 15) for quinine (Q) and orange juice (O)) was counterbalanced across participants. In the fMRI study, the US bolus onset was preceded (500 ms) by a blue square that was presented for 2500 ms on the screen (Figure 1B). Participants were instructed to only swallow the US bolus upon offset of the blue square. Each trial was separated by an inter-trial-interval (ITI) marked by a grey screen. The ITI per trial was drawn from a discretized γ-distribution (shape = 6, scale = 1) truncated for an effective range of values between 3500 ms and 10,000 ms. Participants took self-paced breaks after each 10th trial during which they could drink water. Importantly, the instructions did not contain information about the underlying associative structure of the experiment, aiming at leaving participants unaware of the associative learning process. Instead, participants were instructed to perform a simple attentional control task, during which they should react as quickly and as correctly as possible by pressing the "y" button upon seeing the CS_1 colored in red. Each CS_1 was colored red in 10 % of the trials (90 % of trials grayscale image) and color did not predict US contingency. Performance during the task was rewarded with a bonus of 1 € (if > 70 % correct answers). Participants performed very well in the attentional control task (overall probability of correct answers: M = .99, SD = .05). In the fMRI sample, both ratings and first-order conditioning were performed outside the MRI scanner.

Learning experiment – second-order conditioning

For second-order conditioning (SOC), participants were presented with CS_2^+ followed by CS_1^+, CS_2^- followed by CS_1^- and CS_2^n followed by CS_1^n. CS_1^n had not been presented during first-order conditioning and thus was not paired with any US. The novel CS_2^n-CS_1^n pair served to control for attentional and novelty effects. In each trial (Figure 1D), a CS_2 (2000 ms) was followed by an inter-stimulus interval (500 ms) marked by a fixation cross, and a CS_1 (2000 ms). CS_2 were followed by a CS_1 deterministically. Each trial was separated by an

ITI marked by a grey screen. The ITI per trial was drawn from a discretized γ-distribution (shape = 7, scale = 1) truncated for an effective range of values between 4000 ms and 10,000 ms. Each CS_2-CS_1 pair was presented 50 times, amounting to 150 trials total. Again, instructions did not explicitly mention relational structures to be learned, but participants were instructed to perform a simple attentional control task. Participants were instructed to respond as quickly and as correctly as possible by pressing the "y" button (behavioral study) or with the right index finger on an MRI-compatible 4-button response box (fMRI study) upon seeing the CS_2 tilted by a 45° angle. Each CS_2 was tilted in 10 % of its presentations. Performance during the task was rewarded with a bonus of 1 € (if > 70 % correct answers). Participants performed very well in the attentional control task (overall probability of correct answers: M = .97, SD = .07). In the fMRI study, SOC was performed inside the MRI scanner.

Learning experiment – choice preference test

Following SOC, participants were presented with two separate test phases consisting of repeated binary choices (30 trials) between pairs of CS_1 and pairs of CS_2 to assess behavioral signatures of first- and second-order conditioning. Each choice between CS_1^+ and CS_1^-, and between CS_2^+ and CS_2^- was presented ten times (twelve times in the fMRI sample). We hypothesized preference for both CS_1^+ and CS_2^+ over CS_1^- and CS_2^-, respectively, as expressed by choice probabilities above indifference criterion (CP > 0.5). Decision trials were interleaved with choices between the novel CS_1^n or CS_2^n and lure stimuli (fractal or kanjis, respectively) that had only been seen during pre-task rating in pseudo-random order. These novel-to-lure comparison trials were intended to rule out response biases related to mere exposure to the stimuli experienced during conditioning. We reasoned that if choices were reflecting acquired and transferred value and could not only be attributed to mere exposure-dependent response biases, choice probability in trials involving novel-to-lure stimuli should not exceed the indifference criterion. Choice options were presented for 1500 ms on the right- and left-hand side of the screen. Order (left/right) of choice options was counterbalanced. If participants did not respond within this time-window, a time-out message

was displayed, and the respective trial was repeated at the end of the choice preference test.

Participants selected choice options by pressing the "y" (left option) or "m" (right option)

button (behavioral sample, left or right index finger on MR-compatible response box in the

fMRI sample). Participants were instructed to select the CS they preferred/liked more.

Importantly, participants were never presented with the US related to their chosen or

unchosen CS.

Behavioral analyses

Data were analyzed using MATLAB 2019a (v9.6.0.1072779, The MathWorks, Inc., Natick,

MA, USA) and RStudio (RStudioTeam, 2019) (version 3.6.3, RStudio Team, Boston, MA)

using custom analysis scripts. We calculated an average choice probability (CP) per first-

and second-order CS per subject including all binary decisions in which the respective CS

was present (1 = selection of the respective target CS, 0 = selection of the alternative CS) for

CS_1^+ and CS_2^+ (versus CS_1^- and CS_2^-, respectively). We hypothesized above-chance level

CPs for both CS_1^+ and CS_2^+. For both first-order and second-order novel-to-lure comparison

trials, we calculated an average CP per stimulus and expected no difference from chance

level for these choices. CPs for each CS in the fMRI and behavioral study were pooled and

jointly analyzed. For hypothesis testing, we calculated 89%-highest posterior density

intervals (89%-HPDI) for each choice probability using the R package rethinking (McElreath,

2020). The HPDI represents the parameter values most consistent with the data. We

specified a weakly informative prior probability distribution (10000 steps), an aggregated

binomial distribution including the number of choices for the respective stimulus and the total

number of choice trials involving the respective CS as likelihood function (Equation 1) and

randomly drew 100000 samples from a grid-approximated (10000 steps) normalized

posterior probability distribution. For hypothesis testing, we specified a region of practical

equivalence (ROPE), i.e. an interval of parameter values (CP = [.45; .55]) representing the

null hypothesis. Specifically, we expected that the 89%-HPDI would not overlap with the

ROPE for choices of CS_1^+ and CS_2^+ (versus CS_1^- and CS_2^-, respectively), allowing to reject

the null hypothesis. However, the ROPE would overlap with the 89%-HPDI for both first-order and second-order novel-to-lure comparison trials. Overlap between HPDIs and ROPEs was quantified using the R package bayestestR (Makowski et al., 2019).

$$CP_i \sim \text{Binomial}(N_i, p_i)$$

$$p_i \sim \text{Normal}(0.5, 0.1)$$

(1)

where CP_i is the binomially distributed choice probability for CS_i (CS_1^+, CS_2^+, CS_1^n or CS_2^n). N_i is the total number of trials in which CS_i was present and p_i is the proportion of choices of CS_i.

fMRI – classifier training experiment

The classifier training experiment was always performed on day 1 of two consecutive testing days. This was done to acquire unbiased estimates of the neural patterns representing the two US for training of a multivariate classification algorithm (multivariate pattern analysis, MVPA), before any association of the US with first- or second order CS had been acquired. Participants received oral instructions for the classifier training experiment and were instructed once again on the projection screen in the MRI. Before and after the experiment, participants rated gustatory US (aversive: quinine-HCL 0.1 mmol/l solved in purified water and appetitive: orange juice) according to subjective value/liking and intensity via button presses with their right and left index and middle fingers on a MRI-compatible 4-button response box. Ratings ranged from 1 (not liked/not intense) to 4 (very much liked/very intense) and were indicated by e.g. pressing the left middle finger corresponding to rating 1 or pressing the right index finger corresponding to rating 3. As on day 2, participants were instructed to abstain from food for 12 hours and on average reported to have fasted for 13.59 (*SD* = 2.18) hours. Participants reported intermediate levels of hunger before the experiment on a paper-pencil VAS, *M* = 48.50 (*SD* = 31.28). In each of the five total runs of the classifier training experiment, each US was administered twenty times (40 trials per run, 200 trials in total). Per trial, one US (1 ml bolus per trial) was delivered by a MATLAB code-controlled custom-made gustometer and Luer lock infusion lines (8.30 m length, 2 mm inner diameter).

Participants held infusion lines centrally in their mouths like drinking straws. Additionally, infusion lines were fixed at absorbent materials attached to throat, chin and cheeks. The US bolus onset was preceded (500 ms) by a blue square that was presented in total for 4000 ms on the screen. Participants were instructed to only swallow the US bolus upon offset of the blue square. Each trial was separated by an ITI marked by a grey screen. The ITI was drawn from a uniform distribution with 5 discrete steps (range: 3000 – 7000 ms). Both US were presented in a pseudo-random order, thus reducing the influence of potentially confounding low-level features of the schedule (e.g. number of same/different US repetitions, different ITI lengths following each US). During each run of the classifier training experiment, participants performed a 0-back style attentional control task. After a pseudo-random 20% of trials, participants were presented with probe trials in which they were asked to indicate which US they had received last ("pleasant" (US$^+$) or "unpleasant" (US$^-$)) via button presses with their right and left index fingers on an MRI-compatible response box. Correct responses were rewarded with 0.05 € and incorrect responses or time-out trials (without a response by the participant within 2500 ms after onset of the probe trial) resulted in a 0.05 € penalty which would be summed up as a bonus upon completion of the experiment. On average, participants earned a bonus of 1.86 € (*SD* = .11) during the classifier training experiment. Performance during the 0-back attentional control task was generally high (overall probability of correct answers, excluding time-out trials: *M* = .93, *SD* = .05).

fMRI – acquisition

During SOC, one run and during the classifier training experiment, five runs of fMRI were acquired on a 3 Tesla Siemens PRISMA MR-system (Siemens, Erlangen, Germany), using a 64-channel head coil. Blood oxygenation level dependent (BOLD) signals were acquired using a multi-band accelerated T2*-weighted echo-planar imaging (EPI) sequence (multi-band acceleration factor 2, repetition time (TR) = 2000 ms, echo time (TE) = 30 ms, flip angle = 80°, field of view (FoV) = 220 mm, voxel size = 2.2 × 2.2 × 2.2 mm, no gap). Per volume, 66 slices covering the whole brain, tilted by approximately 15° in z-direction relative to the

anterior–posterior commissure plane were acquired in interleaved order. The first 5 volumes

of the functional imaging time series were automatically discarded to allow for T1 saturation.

To ensure close spatial alignment of the acquired slices during both sessions, the AutoAlign

technique provided by the vendor was applied. At the end of both testing days, a B_0

magnitude and phase map was acquired to estimate field maps and B_0 field distortion during

preprocessing (TR = 660 ms, TE 1 = 4.92 ms, TE 2 = 7.38 ms, flip angle = 60°, FoV = 220

mm). Additionally, before task-based fMRI on both days, a high-resolution three-dimensional

T1-weighted anatomical map (TR = 2500 ms, TE = 2.82 ms, FoV = 256 mm, flip angle = 7°,

voxel size = 1 × 1 × 1 mm, 192 slices) covering the whole brain was obtained using a

magnetization-prepared rapid acquisition gradient echo (MPRAGE) sequence. This scan was

used as anatomical reference to the EPI data during the registration procedure. For all cross-

session classification analyses, SOC EPI data was referenced to orientation of the classifier

training experiment.

fMRI – data preprocessing

All fMRI preprocessing steps were performed using tools from the Functional Magnetic

Resonance Imaging of the Brain (FMRIB) Software Library (FSL, v5.0 and v6.0) (Jenkinson

et al., 2012). Preprocessing for each run of both classifier training task and SOC task

included motion correction using rigid-body registration to the central volume of the functional

time series (Jenkinson et al., 2002), correction for geometric distortions using the field maps

and an n-dimensional phase-unwrapping algorithm (B_0 unwarping) (Jenkinson, 2003), slice

timing correction using Hanning windowed sinc interpolation and high-pass filtering using a

Gaussian-weighted lines filter of 1/100 Hz. EPI images were registered to the high-resolution

structural image using affine linear registration (boundary-based registration) and then to

standard (MNI) space using linear (12 degrees of freedom) and nonlinear registration

(Andersson et al., 2007a, 2007b). Functional data was not spatially smoothed. We applied a

conservative independent components analysis (ICA) to identify and remove obvious

artefacts. Independent components were manually classified as signal or noise based on

published guidelines (Griffanti et al., 2017), and noise components were removed from the

functional time series.

fMRI – searchlight classification analyses

As conditioning enhances CS's ability to pre-activate neural ensembles representing an

associated US, we expected that a multivariate classification algorithm trained on neural

patterns representing the US (during the classifier training task) would be able to correctly

predict the class of a paired CS_1 during SOC. Hence, we used a cross-session (Stokes et al.,

2009), cross-modality searchlight (Kriegeskorte et al., 2006) classification approach to

identify brain regions in which the CS class during SOC could be predicted based on training

the classifier during the classifier training experiment on day 1. General linear models

(GLMs) were fitted into pre-whitened data space to account for local autocorrelations

(Woolrich et al., 2001). For the searchlight classification analyses, the individual level (first

level) GLM design matrix per run and participant of the classifier training experiment included

four box-car regressors in total. Two regressors coded for onsets and durations of both US

(each modelled as single events of 4000 ms duration) and two regressors coded onsets and

durations of left and right button presses (delta stick functions on the recorded time of

response button presses) and the six volume-to-volume motion parameters from motion

correction during preprocessing were entered to account for residual head motion. The SOC

run per participant included five box-car regressors in total. Three regressors coding for

onsets and durations of all combinations of CS_2-CS_1 (CS_2^+/CS_1^+, CS_2^-/CS_1^-, CS_2^n/CS_1^n) trials

(each modelled as single events of 4500 ms duration, due to the 100 % contingency of CS-

CS pairs), one regressor coding onsets and durations of right button presses (delta stick

functions on the recorded time of response button presses), one regressor coding onset and

duration of the within-run pause (45 sec), and the six volume-to-volume motion parameters

from motion correction during preprocessing were entered.

Additionally, we performed an exploratory multivariate searchlight classification to

distinguish between the spatial activation patterns (t maps) of US^+ versus US^- across the

whole brain during the classifier training experiment. This analysis resulted in one whole-brain map of classification accuracies per participant. The classification analysis was performed using a 5-fold cross-validation scheme, iteratively training the classification algorithm on the difference between US^+ versus US^- on 4 (out of 5 runs) labeled training data sets and then assessing the out-of-sample predictive accuracy of the classifier on the remaining, left-out data set (1 out of 5 runs). For this analysis, the individual level (first level) GLM design matrix per run and participant of the classifier training experiment included four box-car regressors in total. Two regressors coded for onsets and durations of both US (each modelled as single events of 4000 ms duration) and two regressors coded onsets and durations of left and right button presses (delta stick functions on the recorded time of response button presses) and the six volume-to-volume motion parameters from motion correction during preprocessing were entered to account for residual head motion.

Regressors were convolved with a hemodynamic response function (γ-function, mean lag = 6 s, SD = 3 s). Each first level GLM included two or three (for classifier training experiment or SOC, respectively) contrasts to estimate individual per run t-statistic maps for each US or each CS_2-CS_1 pair (for classifier training experiment or SOC, respectively). For SOC, the activation pattern of CS_2^n/CS_1^n was subtracted from CS_2^+/CS_1^+ and CS_2^-/CS_1^- activation patterns. This was intended to account for general activation expected to result from visually presented CS and the visually presented swallowing cue coinciding with US presentation, which could have confounded classification accuracies.

All classification-based analyses were conducted in subject native space. Per participant t-statistic maps were subjected to linear support vector machine (C-SVM, cost parameter C = 1) classification using the MATLAB-based multivariate pattern analysis toolbox CoSMoMVPA (Oosterhof et al., 2016). For cross-session, cross-modality classification, a classifier was trained on the spatial activation patterns of US^+ versus US^- and tested on CS_2^+/CS_1^+ versus CS_2^-/CS_1^- during SOC within 3-mm searchlight spheres across the whole brain. For the exploratory multivariate searchlight classification between US^+ versus US^- patterns, we similarly employed a 3-mm searchlight sphere across the whole

brain. Each searchlight sphere classification accuracy was mapped to the center voxel of the sphere, resulting in one whole-brain map of classification accuracies per participant. Additionally, in cross-session, cross-modality classification we repeated this procedure 100 times per participant with randomly permuted class labels in the training data set (US$^+$ versus US$^-$) to create chance level maps. Before group level statistics, the normalization parameters obtained from preprocessing were applied to both whole-brain classification accuracy maps and chance level maps for normalization to MNI space. The resulting normalized whole-brain maps were spatially smoothed with a Gaussian kernel (5 mm full-width half-maximum).

In cross-session, cross-modality classification, we employed small-volume correction (P_{SVC}) to assess significance of clusters with above-chance level (0.5) classification accuracy in the lateral orbitofrontal cortex (lOFC), our a priori region-of-interest (ROI). This region has been implicated in higher-order gustatory processing, such as motivational aspects and discrimination of taste (Kobayashi et al., 2004; Miranda, 2012; Small et al., 1999), but also representation and adaptive changes of stimulus-outcome associations (Boorman et al., 2016; Jocham et al., 2016; Klein-Flügge et al., 2013; Luettgau et al., 2020). We thus reasoned that this brain region would be a well-suited candidate for associative coupling between visual conditioned stimuli and gustatory outcomes and the proposed associative transfer learning. We employed two different approaches for ROI definition: 1) an independent anatomical mask of the lateral OFC (Harvard-Oxford Atlas) and 2) a meta-analysis-based explicit mask for functional activations related to the term "taste" in the Neurosynth data base (Yarkoni et al., 2011) (www.neurosynth.org), thresholded at $z > 7$. This value was chosen so that only meta-analytic clusters of activation, but not single voxels, survived thresholding. The Neurosynth-based mask encompassed, aside from the lateral OFC, functional activation clusters in bilateral anterior insula. This approach aimed at reducing inferential limitations related to arbitrary ROI definition.

Within these ROIs, we computed group-level random-effect cluster-statistics corrected for multiple comparisons in small volumes as implemented in CoSMoMVPA (Oosterhof et al., 2016). In brief, for 50,000 iterations, we randomly selected one chance

level map per participant, calculated a group level z-statistic map for the respective

permutation and finally compared the resulting cluster sizes drawn from this empirical null

distribution (50,000 samples) to the clusters in the "real" classification accuracy map (Stelzer

et al., 2013). Since classifications accuracies below chance level are generally limited in

interpretability, we considered clusters significant if $P_{SVC} < .05$ (one-tailed). Please note that

due to the exploratory nature and illustrative purpose of the exploratory multivariate

searchlight classification between US^+ versus US^- patterns, no multiple comparisons

correction using random-effects cluster-statistics was performed in this analysis.

Additionally, we split participants into a high- or low-bias group, depending on their

preference for the CS_2^+ (vs. CS_2^-) and compared classification accuracies within the small-

volume corrected cluster. The cluster ROI was built in MNI space and the ROI was then

back-projected into subject native space using inverse normalization parameters obtained

during preprocessing to extract individual averaged classification accuracies from the ROI.

Both groups' average extracted classification accuracies were separately tested against

chance level (0.5) using one-sample Wilcoxon signed-rank tests. We report measures of

effect size Cohen's $U3_1$ for one-sample Wilcoxon signed-rank tests (range: 0 – 1, .5 indicating

no effect), calculated in the MATLAB-based Measures-of-Effect-Size-toolbox (Hentschke,

2020).

fMRI – functional covariation analyses

We investigated functional covariation between the MVPA-identified lOFC cluster surviving

multiple comparisons (thresholded at $z > 1.65$, or $P < .05$, one-tailed) during SOC as a seed

region and the whole-brain using a psychophysiological interaction (PPI) analysis in FSL. For

each participant, we set up two separate first-level PPI GLMs. The first PPI GLM was aimed

at investigating general differences in BOLD signal covariation for both CS_2^+/CS_1^+ and CS_2^-

$/CS_1^-$ compared with CS_2^n/CS_1^n. The first PPI GLM contained the following regressors: 1) the

BOLD timeseries (preprocessed functional time series) of the lOFC seed as physiological

regressor, 2) onsets and durations of CS_2^+/CS_1^+ and CS_2^-/CS_1^- pairs (magnitude coded as 1)

and CS_2^n/CS_1^n (magnitude coded as –1) as psychological regressor and 3) the interaction between 1) and 2) as psychophysiological interaction regressor. First-level regressor 2) was convolved with a hemodynamic response function (γ-function, mean lag = 6 s, SD = 3 s). The second PPI GLM aimed at investigating specific differences in covariation between CS_2^+/CS_1^+ and CS_2^-/CS_1^-. The second PPI GLM contained the following regressors: 1) the BOLD timeseries (preprocessed functional time series) of the lOFC seed as physiological regressor, 2) onsets and durations of CS_2^+/CS_1^+ (magnitude coded as 1) and CS_2^-/CS_1^- pairs (magnitude coded as –1) as psychological regressor and 3) the interaction between 1) and 2) as psychophysiological interaction regressor. First-level regressor 2) was convolved with a hemodynamic response function (γ-function, mean lag = 6 s, SD = 3 s).

In addition to the main effects of the PPI regressors, we also investigated whether task-related functional covariation of lOFC with the rest of the brain was related to our behavioral index of second-order conditioning (choice probability of CS_2^+ vs. CS_2^-). For this, we entered the averaged second-order choice preference test data per subject as an additional regressor to the group level GLMs. In all analyses, contrast images from the first level were then taken to the group level and analyzed using mixed-effects analyses in FLAME1+2 (Beckmann et al., 2003). We used cluster-based correction with an activation threshold of $Z > 2.3$ and a cluster-extent threshold of $P < .05$ at whole-brain level to control for multiple comparisons.

fMRI – neural pattern similarity analyses

In addition to our classification-based approach, we investigated whether there was evidence for changes of neural pattern similarity across second-order conditioning. Specifically, we used a least-squares separate (LS-S) approach (Mumford et al., 2012) to deconvolve single-trial estimates of neural patterns representing CS_2 and CS_1 and performed a template-based neural pattern similarity analysis (Wimber et al., 2015) (a variant of representational similarity analysis, RSA (Kriegeskorte et al., 2008)) between patterns of the two US during the classifier training experiment and all trial-specific CS_2/CS_1 patterns during the SOC run. We

estimated two subject-specific GLMs per trial, one for both CS_1 and CS_2, containing the

following regressors: 1) onset and duration of the respective trial to be estimated (each

modelled as a single event of 2000 ms duration), 2) onsets and durations of the remaining

CS_1 or CS_2 trials, respectively, 3) onsets and durations of all other CS_1 or CS_2 trials,

respectively. Additionally, one regressor coding onsets and durations of right button presses

(delta stick functions on the recorded time of response button presses), one regressor coding

onset and duration of the within-run pause (45 sec), and the six volume-to-volume motion

parameters from motion correction during preprocessing were entered. For example, the

GLM of the 25th trial of CS_1^+, contained one regressor (onset and duration) modelling this

particular trial and one separate regressor (onsets and durations) for all remaining CS_1 trials

but the 25th trial of CS_1^+ (149 other trials in total) and one separate regressor (onsets and

durations) modelling all CS_2 trials. This procedure is assumed to allow for better

deconvolution of events happening within close temporal proximity than standard least-

squares-based approaches, which generally do not distinguish between overall and trial-

specific error terms (Abdulrahman and Henson, 2016; Mumford et al., 2012).

Each first level GLM included one contrast to model activation related to the

respective trial of interest versus baseline. The a priori ROIs of the bilateral amygdala were

built in MNI space and back-projected into subject native space using inverse normalization

parameters obtained during preprocessing. We used these individual ROIs for spatially

constrained multivoxel pattern extraction from the respective contrast t-value maps.

Similarity-based analyses were carried out using CoSMoMVPA (Oosterhof et al., 2016). We

employed $1-$Pearson's product-moment correlation coefficient $(1-r)$ as a measure of

pairwise dissimilarity between trial-specific neural patterns and the US template. Within-

subject pairwise neural dissimilarity was subtracted from 1 (to create a measure of neural

pattern similarity) and Fisher-Z transformed to closer approximate normally distributed data.

CS_1^n-US^-/CS_2^n-US^- and CS_1^n-US^+/CS_2^n-US^+ early and late neural pattern similarity were

subtracted from CS_1^--US^-/CS_2^--US^- and CS_1^+-US^+/CS_2^+-US^+ early and late neural pattern

similarity, respectively. We hypothesized that neural pattern similarity would show

dissociable changes from early (first 25) to late (last 25) trials, as both CS_2^+ and CS_2^- should

become more similar to the respective indirectly associated US, indicating associative

learning transfer from CS_1 to CS_2. Contrarily, CS_1^+ and CS_1^- should, if anything, become less

similar to the respectively paired US (e.g. due to extinction). Average change of early to late

CS-US neural pattern similarity was analyzed at the group level with paired-samples t-tests.

Due to the expected negative change for CS_1^+ and CS_1^- and the expected positive change

for CS_2^+ and CS_2^- across trials, we used one-tailed tests accordingly.

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Competing interests

Authors declare no competing interests.

Data availability

The behavioral data, extracted classification accuracies and neural pattern similarity matrices that support the findings of this study will be made publicly available at GitHub upon publication. The neuroimaging raw data that support the findings of this study are available upon reasonable request from the corresponding author (LL). The neuroimaging raw data are not publicly available due to them containing information that could compromise research participant privacy/consent.

Code availability

Custom analysis code for the reported behavioral data analyses and multivariate fMRI analyses will be made publicly available at GitHub upon publication.

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1092 # Supplementary Information

1093 **Reinstatement of cortical outcome representations during higher-order learning**

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Supplementary Results

Explicit knowledge of associative structure

Post-experimental paper-pencil tests for explicit knowledge of the associative structure

revealed that the number of participants that had indeed realized the indirect associations

between CS_2 and US was significantly below the value that would be expected under a

random guessing assumption (29%, P = 0.004, binomial test vs. 0.5). Participants could also

not reliably indicate which CS_2 and which US had been indirectly linked in the experiment

(mean correct responses: M = 0.31, SD = 0.68, maximum of 2 correct answers possible,

Supplementary Table 1).

CS and US ratings

During the ratings prior to the learning experiment, the subjective values/liking of the three

kanjis selected as CS_2 did not differ significantly from each other (main effect of stimulus:

$F_{2,96}$ = 0.07, P = .993, η^2_p < .001, rmANOVA, Supplementary Fig. 2A). However, there was a

significant main effect of stimulus for the three fractals selected as CS_1 (main effect of

stimulus: $F_{2,96}$ = 3.25, P = .043, η^2_p = .06, rmANOVA, Supplementary Fig. 2B) that was driven

by higher rating of CS_1^n than CS_1^- (t_{48} = 2.26, P = .028, paired-samples t-test). This difference

most likely resulted from the selection process and the inherent rank-ordering of the CS

according to subjective value ratings. However, there were no significant differences

between CS_1^n and CS_1^+ (t_{48} = 1.68, P = .101, paired-samples t-test). Most importantly, no

rating differences were observed between CS_1^- and CS_1^+ (t_{48} = 0.44, P = .659, paired-

samples t-test). Overall, US valence ratings did not differ from pre to post rating across both

US (main effect of time (pre/post): $F_{1,48}$ = 0.34, P = .564, η^2_p = .01, rmANOVA), but the

appetitive US^+ was – as expected – rated as significantly more pleasant than the aversive

US^- (main effect of stimulus: $F_{1,48}$ = 466.64, P < .001, η^2_p = .91, rmANOVA). There was also

a significant interaction effect between stimulus and time, indicating that US^+ was rated as

more pleasant and US^- was rated as less pleasant from pre to post (interaction effect of

stimulus x time: $F_{1,48}$ = 7.32 P = .009, η^2_p = .13, rmANOVA, Supplementary Fig. 2C).

Importantly, no subject rated US^- as more pleasant than US^+ (pre or post). In the fMRI study,

intensity ratings did not differ between US overall (main effect of stimulus: $F_{1,28}$ = 1.05, P =

.315, η^2_p = .04, rmANOVA), but US were rated as significantly more intense at the post rating

(main effect of time: $F_{1,28}$ = 4.89, P = .035, η^2_p = .15, rmANOVA, Supplementary Fig. 2D).

There was also a significant interaction effect between stimulus and time, indicating that US^-

intensity rating increased more strongly than US^+ from pre to post (interaction effect of

stimulus x time: $F_{1,28}$ = 5.62, P = .025, η^2_p = .17, rmANOVA). The average temperature

(°Celsius) of US^+ (M = 22.49, SD = 1.46) was significantly higher than temperature of US^- (M

= 22.37, SD = 1.52) before being loaded into the syringes (t_{27} = 2.22, P = .035, paired-samples t-test). However, we would argue that this minor difference in average temperatures (M = .12) is unlikely to be perceived by participants. Additionally, this difference does not meaningfully influence the observed results, since no neural activation was recorded during first-order conditioning and the multivariate classification analysis was trained on the data resulting from the fMRI classifier training experiment.

In the fMRI classifier training experiment, overall US valence ratings were higher in pre than in post rating (main effect of time (pre/post): $F_{1,28}$ = 11.67, P = .002, η^2_p = .29, rmANOVA), but the appetitive US$^+$ was, as intended, rated as significantly more pleasant than the aversive US$^-$ across ratings (main effect of stimulus: $F_{1,28}$ = 185.28, P < .001, η^2_p = .87, rmANOVA). There was no interaction effect between stimulus and time ($F_{1,28}$ = 0.28, P = .602, η^2_p = .01, rmANOVA). Importantly, no subject rated US$^-$ as more pleasant than US$^+$ (pre or post). Intensity ratings for US$^+$ were higher than for US$^-$ across ratings (main effect of stimulus: $F_{1,28}$ = 15.14, P < .001, η^2_p = .35, rmANOVA). There was no difference between US ratings from pre to post (main effect of time: $F_{1,28}$ = 0.02, P = .899, η^2_p < .001, rmANOVA). However, there was also a marginal interaction effect between stimulus and time, indicating that US$^-$ intensity rating increased while US$^+$ intensity rating decreased from pre to post (interaction effect of stimulus x time: $F_{1,28}$ = 3.93, P = .057, η^2_p = .12, rmANOVA). Average temperature of US$^+$ (M = 21.71, SD = 1.93 °C) and US$^-$ (M = 21.68, SD = 1.95) did not differ before being loaded into the syringes (t_{27} = 0.71, P = .486, paired-samples t-test).

fMRI searchlight classification

In the main article, we present evidence for reinstatement of US patterns in a region in the left lateral orbitofrontal cortex (lOFC). In this region, significant decoding of CS identity was possible on the basis of the US pattern obtained during classifier training on the previous day using classification with a 3-mm searchlight and small-volume correction in a ROI defined by a meta-analysis based on the term "taste" in the Neurosynth data base. To ensure that our results are not dependent on this particular choice of ROI, we repeated the same analysis using two different, independent ROI. The first was an anatomical mask of the left lateral orbitofrontal cortex (Harvard-Oxford atlas), the second was obtained from an independent gustatory mapping study by Benz and colleagues (K. Benz, personal communication, 12/2019). Using random-effect cluster-statistics with 50,000 iterations, we found similar results in the left lOFC anatomical ROI (peak voxel at MNI [x = –21, y = 30, z = –17], z = 1.96, P = .024, corrected, one-tailed) and in the mask from Benz and colleagues (peak voxel at MNI [x = –21, y = 30, z = –17], z = 1.84, P = .033, corrected, one-tailed).

Association between second-order CS and US in the lOFC

When repeating the analyses performed in the amygdala within the lOFC cluster in which it was possible to predict CS labels based on US patterns, we did not observe significant pattern similarity between CS_2^- and US^- ($t_{28} = 0.06$, $P = 0.955$, $U3_1 = 0.52$), nor between CS_2^+ and US^+ ($t_{28} = 0.71$, $P = 0.489$, $U3_1 = 0.52$). There was no indication of increases of pattern similarity in this region (CS_2^- and US^-: $t_{28} = 0.02$, $P = 0.491$, Cohen's $d = 0.005$; CS_2^+ and US^+: $t_{28} = 0.33$, $P = 0.629$, Cohen's $d = 0.06$).

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Supplementary Fig. 1. Extended behavioral results.
A) Choice probabilities in the behavioral study (N = 20): Raincloud plots showing density of
choice probability. CS_1^+ (dark blue, 89%-highest posterior density interval (89%-HPDI): [.62;
.72]) and CS_2^+ (light blue, 89%-HPDI: [.65; .77]) were preferred over CS_1^- and CS_2^-,
respectively. HPDIs did not overlap (0% overlap) with the defined region of practical
equivalence (ROPE, CP = [.45; .55]). There was no evidence for choice probabilities of CS_1^n
and CS_2^n being different from chance level in novel-to-lure comparison choice trials (dark and
light gray, 89%-HPDIs: [.42; .52] and [.46; .56], 68.3% and 94.8% ROPE overlap,
respectively). B) Choice probabilities in the fMRI study (N = 29). CS_1^+ (dark blue, 89%-HPDI:
[.54; .63]) and CS_2^+ (light blue, 89%-HPDI: [.65; .77]) were preferred over CS_1^- and CS_2^-,
respectively. HPDIs did only minimally overlap (4.6% and 0.8% overlap, respectively) with
the defined ROPE. There was no evidence for choice probabilities of CS_1^n being different
from chance level in novel-to-lure comparison choice trials (dark gray, 89%-HPDIs: [.47; .55],
99.2% ROPE overlap). However, participants neglected CS_2^n compared to control stimuli
(light gray, 89%-HPDI [.36; .44], 0% ROPE overlap). Box plot center lines represent sample
medians and box bottom/top edges show 25th/75th percentile of the data, respectively.
1198

1199
1200
Supplementary Fig. 2. CS and US ratings.
A) During the ratings prior to the learning experiment, pooled across both studies ($N = 49$)
the subjective values/liking of the three kanjis selected as CS_2 did not differ significantly from
each other ($F_{2,96} = 0.07$, $P = .993$, $\eta^2_p < .001$, rmANOVA). B) There was a significant main
effect of stimulus for the three fractals selected as CS_1 ($F_{2,96} = 3.25$, $P = .043$, $\eta^2_p = .06$,
rmANOVA) that was driven by higher rating of CS_1^n than CS_1^- ($t_{48} = 2.26$, $P = .028$, paired-
samples t-test). However, there were no significant differences between CS_1^n and CS_1^+ ($t_{48} =$
1.68, $P = .101$, paired-samples t-test). Most importantly, no rating differences were observed
between CS_1^- and CS_1^+ ($t_{48} = 0.44$, $P = .659$, paired-samples t-test). C) US valence ratings
did not differ from pre to post rating across both US (main effect of time (pre/post): $F_{1,48} =$
0.34, $P = .564$, $\eta^2_p = .01$, rmANOVA), but the appetitive US^+ was – as expected – rated as
significantly more pleasant than the aversive US^- (main effect of stimulus: $F_{1,48} = 466.64$, $P <$
.001, $\eta^2_p = .91$, rmANOVA). There was also a significant interaction effect between stimulus
and time, indicating that US^+ was rated as more pleasant and US^- was rated as less pleasant
from pre to post (interaction effect of stimulus x time: $F_{1,48} = 7.32$ $P = .009$, $\eta^2_p = .13$,
rmANOVA). D) In the fMRI study ($N = 29$), intensity ratings did not differ between US overall
(main effect of stimulus: $F_{1,28} = 1.05$, $P = .315$, $\eta^2_p = .04$, rmANOVA), but US were rated as
significantly more intense at the post rating (main effect of time: $F_{1,28} = 4.89$, $P = .035$, $\eta^2_p =$
.15, rmANOVA). There was also a significant interaction effect between stimulus and time,
indicating that US^- intensity rating increased more strongly than US^+ from pre to post
(interaction effect of stimulus x time: $F_{1,28} = 5.62$, $P = .025$, $\eta^2_p = .17$, rmANOVA). Bar plots
represent sample means and error bars indicate standard errors of the means, respectively.
1223

x = 49   z = 6   y = 41

2.3   3.9

Supplementary Fig. 3. Psychophysiological interaction analyses: Association with CS_2^+ versus CS_2^- choice probability.

A positive correlation was observed between second-order choice preferences (CS_2^+ versus CS_2^-) and functional covariation between the left lOFC and a region in the right and left lateral prefrontal cortex (right peak voxel at MNI [x = 49, y = 41, z = 7], Z = 3.80, $P < 0.001$, whole-brain corrected; left peak voxel at MNI [x = –45, y = 46, z = –4], Z = 3.93, $P = 0.010$, whole-brain corrected). The more participants preferred CS_2^+ (versus CS_2^-), the higher these regions' activation covaried during second-order conditioning. Color bar represents Z-values.

x = –21   z = –16   y = 30

3.1 ▬▬▬▬▬▬▬▬▬▬▬▬▬▬▬▬▬▬ 4.5

Supplementary Fig. 4. US⁺ versus US⁻ classification.

Group average (N = 29) cross-validated predictive accuracy for US⁺ versus US⁻ activation patterns during the classifier training experiment. As an exploratory analysis, we employed multivariate searchlight classification (3-mm searchlight spheres) to distinguish between the spatial activation patterns (t maps) of US⁺ versus US⁻ across the whole brain, resulting in one whole-brain map of classification accuracies per participant. This classification was performed using a 5-fold cross-validation scheme, iteratively training the classification algorithm on the difference between US⁺ versus US⁻ on labeled training data sets (4 out of 5 runs) and then assessing the out-of-sample predictive accuracy of the classifier on the remaining, left-out data set (1 out of 5 runs). Data was averaged across participants and smoothed with a Gaussian kernel with 5 mm FWHM. Please note that due to the exploratory nature and illustrative purpose of this analysis, no multiple comparisons correction using random-effects cluster-statistics was performed. Color bar represents uncorrected Z-values.

 Supplementary Table 1. Explicit knowledge of CS-US associations

CS$_1$–US (N = yes)	CS$_1$–US (Mean + SD)	CS$_2$–US (N = yes)	CS$_2$–US (Mean + SD)
38 (78 %)	.98 (.95)	14 (29 %)	.31 (.68)

Note: Number (N) and percentage (in parentheses) of participants indicating explicit knowledge of CS–US associations between directly paired CS$_1$ and US (providing an answer to the question which CS$_1$ was associated with which US, regardless of being correct or incorrect) and indirectly paired CS$_2$ and US (responding "yes" to the question whether there was a relationship between CS$_2$ and US). These numbers of participants were significantly above (CS$_1$–US) and below (CS$_2$–US) the numbers expected under the assumption of random guessing, $P < .001$ and $P = .004$, respectively (Binomial test vs. 0.5). Additionally, we provide the mean (and standard deviation, SD) number of correct answers for CS–US associations between CS$_1$ and US, and CS$_2$ and US. Since participants were presented with two CS–US associations during conditioning, the maximum number of correct answers is 2 in both cases.