Article Type: Letter

Title: **Egg size variation in a long-lived polyandrous shorebird in the context**

of senescence and breeding phenology

Authors: Luke J. Eberhart-Hertel[*†1], Lourenço Falcão Rodrigues[*†2,3], Johannes Krietsch[4],

Anne G. Eberhart-Hertel[5,6], Medardo Cruz-López[7], Karina Alejandra Vázquez-Rojas[1], Erick

González-Medina[8], Julia Schroeder[3], and Clemens Küpper[*1]

Affiliations:

[1]Research Group Behavioural Genetics and Evolutionary Ecology, Max Planck Institute for

Ornithology, Seewiesen, Germany

[2]Department of Ecology, Autonomous University of Madrid, Madrid, Spain

[3]Department of Life Sciences, Imperial College London, Silwood Park Campus, United

Kingdom

[4]Department of Behavioural Ecology and Evolutionary Genetics, Max Planck Institute for

Ornithology, Seewiesen, Germany

[5]Research Group for Movement Ecology, Senckenberg Biodiversity and Climate Research

Centre, Frankfurt, Germany

[6]Behavioural Ecology, Department of Biology, Ludwig-Maximilians University of Munich,

Planegg-Martinsried, Germany

[7]Institute of Ocean Sciences and Limnology, Universidad Nacional Autónoma de México,

Ciudad de Mexico, Mexico

[8]Conservation Biology Research Group, Department of Anatomy, Cell Biology & Zoology,

University of Extremadura, Badajoz, Spain

[*]Corresponding authors: luke.eberhart@orn.mpg.de, lourenco.falcao@uam.es,

ckuepper@orn.mpg.de,

[†]These authors contributed equally.

Short Running Title: Egg size variation in a polyandrous shorebird

Keywords: ageing, *Charadrius nivosus*, disposable soma, life history, mating system,

reproductive investment, reproductive performance, season, snowy plover, trade-off.

Abstract (279 words):

Anisogamy is a central component of sex role evolution, however, the effect of female-
female mating competition on egg size variation in polyandrous species is unclear. Moreover,
egg size may also be shaped by age-dependent trade-offs between reproductive investments
and somatic maintenance that are responsible for senescence. Here we investigate how
mating behaviour and senescence are associated with egg size variation in female snowy
plovers (*Charadrius nivosus*). Snowy plovers are long-lived shorebirds (longevity record: 20
years) that often produce several nests each year, with females either sequentially changing
partners between breeding attempts or remaining monogamous between attempts. We
examined how age, seasonality, body size, and mating behaviour relate to within- and
between-female variation in egg volume using repeated measures collected over a 15-year
period. We found no evidence of reproductive senescence in egg volume in snowy plover
females. Rather, egg volume, polyandry, and re-nesting were strongly linked to breeding
phenology: early breeding females had a higher likelihood of being polyandrous or replacing
failed clutches, yet these individuals laid smaller eggs likely due to physiological limitations
associated with the early season. Older individuals and local recruits secured the earliest
breeding opportunities in the season suggesting that prior experience could give an edge in
the female-female competition for mates. Larger females laid the largest eggs, as expected,
but there was no relationship between body size and lay date – implying that size may not
provide an advantage in female-female competition. Our findings highlight the existence of
several direct and indirect constraints on female reproductive investment that likely shape
individual variation in lifetime reproductive success. Future research investigating
reproductive senescence of wild populations should consider mating system dynamics when
examining variation in reproductive investment.

Impact Summary (292 words):

The divergent gamete sizes of males and females (anisogamy) is a key principle for sex role

evolution, however, it remains unclear whether and how egg size variation is shaped by

female-female competition in species with reversed sex roles. Furthermore, polygamy is

typically associated with age such that egg size variation may also be shaped by age-

dependent processes including senescence. Here, we use a 15-year longitudinal mark-

recapture dataset of a wild subtropical population of snowy plovers (*Charadrius nivosus*)

breeding in western Mexico to investigate how mating behaviour and senescence are

associated with egg size variation. The snowy plover is a long-lived shorebird characterized

by a flexible polyandrous mating system. This rare breeding behaviour represents a unique

background for investigating senescence in light of individual variation in reproductive

investment. We found no evidence of reproductive senescence in egg volume or polyandry.

Instead, egg volume and polyandry were strongly linked to the seasonal timing of breeding:

early nesting females had a higher likelihood of being polyandrous but laid smaller eggs.

Early nesters also had a higher probability of laying a replacement clutch following breeding

failure. Taken together, this suggests that females are driven to initiate laying as early as

possible despite the cost this has on egg size of their early season clutches. Furthermore,

older individuals and local recruits secured the earliest breeding opportunities in the season

indicating that prior experience gave individuals an advantage in the competition for mates.

Larger females laid the largest eggs, as expected, but there was no relationship between body

size and lay date – implying that size is not an important factor in female-female competition.

We conclude that individual female reproductive performance is regulated by flexible mating

behaviour (i.e., monogamy or polyandry), age- and season-dependent effects, and prior local

experience.

INTRODUCTION

The divergent gamete size of males and females (anisogamy) is fundamental for sex role evolution, however, to what extent female-female competition relates to egg size variation in sex role reversed species remains unclear. At the species level, comparative analysis of birds has shown that a lineage's egg size tends to decrease following the evolution of polyandry (Liker *et al.*, 2001) – supposedly due to the selective advantages that producing smaller eggs has on minimizing the laying period between matings (Goymann *et al.*, 2015) and, hence, maximizing a female's reproductive output given temporal constraints to her breeding schedule (Andersson, 2004). But at the individual level, how does egg size variation relate to the reproductive benefits of female mating behaviour?

A female's breeding schedule depends, in part, on her local resource availability to commence a breeding attempt – resources that include vacant space for reproductive activities (e.g., courting, nesting, etc.), food to launch egg production, and accessible mates (Emlen & Oring, 1977). In sequentially polyandrous species, females compete over these breeding resources to maximize their opportunity for multiple breeding attempts, often resulting in high inter-female variation in reproductive success and breeding schedule (e.g., Colwell & Oring, 1988; Schamel *et al.*, 2004). Females that can reproduce early during the season have more time to replace a failed first attempt (Morrison *et al.*, 2019) and/or to pursue sequential breeding attempts. Moreover, in environments with limited or seasonally dependent resources, breeding early may give offspring favourable conditions to outcompete conspecifics, ultimately reducing the amount of parental care required and maximizing offspring survival prospects (Arnold *et al.*, 2004). However, breeding early may come with the increased physiological stress to egg-laying females, and/or may result in an increased likelihood of brood failure in environments with high stochasticity, such as low food availability, inclement weather, or frequency-dependent predation risk, associated with the early season (Borgmann *et al.*, 2013; Ockendon *et al.*, 2013). Taken together, a seasonal trade-off between quality and quantity exists for females: early nesters may compromise resource investment into their first clutches in the race to maximize time for multiple breeding attempts (Andersson, 2004).

In many organisms, reproductive productivity is age-dependent (Bouwhuis *et al.*, 2009; Hammers *et al.*, 2012; Lemaître *et al.*, 2015; Zhang *et al.*, 2015; Salguero-Gómez *et al.*, 2016; Jankowiak *et al.*, 2018; Dingemanse *et al.*, 2020): an individual's performance

increases over early life to a maximum and is followed by a decline in older ages. The
increase in breeding performance in early life is thought to be an indication of physiological
or competitive inferiority of inexperienced breeders compared to older conspecifics (Curio,
1983), whereas the decline in reproductive performance at old age is considered senescence –
a within-individual age-specific decline in reproductive performance paired with an increase
in mortality risk caused by the loss of physiological and cellular function (Medawar, 1952;
Williams, 1957). A prominent hypothesis explaining the occurrence of senescence is the
disposable soma theory (*sensu* Kirkwood, 1977): individuals are constrained by the amount
of resources they have to invest in survival and reproduction (Kirkwood & Austad, 2000).
Senescence is therefore the consequence of age-dependent trade-offs between energy
investments in reproduction at the expense of somatic repair (Kirkwood & Rose, 1991;
Kirkwood & Austad, 2000; Drenos & Kirkwood, 2005). Thus, an individual's investment in
current reproduction may exacerbate physiological and cellular damage, which might have
knock-on effects on their capability for future reproductive investments at older ages –
ultimately impairing performance.

A key methodological issue for studying senescence in wild populations is that stochastic
extrinsic mortality reduces the frequency of individuals in older age classes, hence making it
challenging to disentangle between- vs. within-individual age-dependent variation – a
phenomenon known as "selective disappearance" (van de Pol & Verhulst, 2006; Nussey *et*
al., 2008). Investigations using longitudinal data to test for senescence are particularly
suitable for this task, as they can control for the confounding effects of selective
disappearance through repeated measures of individuals as they age (van de Pol & Verhulst,
2006; Nussey *et al.*, 2008; Dingemanse *et al.*, 2020).

In oviparous organisms, egg size represents a fundamental measure of female reproductive
investment (Kaplan, 1980; Fox, 1994; Williams, 1994; Starck & Ricklefs, 1998; Moran &
Emlet, 2001; Xu *et al.*, 2019) and is shown to be related to inter- and intra-specific variation
in several life history traits: for example, egg size is associated with developmental mode,
with precocial species typically producing larger eggs than altricial species (Deeming &
Reynolds, 2015). Within individuals, variation in egg size is shown to be highly repeatable
(Christians, 2002), and age-dependent changes in egg or clutch size are often consistent with
senescence although complicated by considerable between-individual variation (Beamonte-
Barrientos *et al.*, 2010; Dingemanse *et al.*, 2020).

Among oviparous animals, shorebirds (part of the order Charadriiformes) produce some of

the largest eggs in relation to body mass due to the needs of their precocial nidifugous young

(Lack, 1968; Rahn *et al.*, 1975). As a clade, shorebirds also exhibit a disproportionately high

prevalence of polyandry (Oring, 1986; Andersson, 2005; Colwell, 2010). Sexual and natural

selection presumably act divergently on egg size: natural selection on offspring viability

favours larger eggs with more nutrients to enhance offspring survival (Blomqvist *et al.*, 1997;

Starck & Ricklefs, 1998; Williams, 2012), whereas sexual selection on polyandry favours

rapid clutch completion to maximize time for multiple nesting attempts – a process that may

indirectly reduce egg size (Liker *et al.*, 2001; Andersson, 2004).

Here, we investigate seasonal- and age-dependent egg size variation using a 15-year

longitudinal mark-recapture dataset of snowy plovers (*Charadrius nivosus*). The snowy

plover is a long-lived shorebird (longevity record: 20 years; Colwell *et al.*, 2017) exhibiting a

rare breeding behaviour characterized by sex-role reversal including facultative sequential

polyandry whereby females regularly desert their broods after hatching to start a new

breeding attempt with another male (Warriner *et al.*, 1986; Eberhart-Phillips *et al.*, 2017;

Kupán *et al.*, 2021). Their breeding season can extend up to six months allowing for multiple

nesting attempts following successful fledging of young or replacing failed nests due to

depredation or flooding (Eberhart-Phillips, 2019). Desertion and re-mating rates decline with

season in polyandrous plovers (Amat *et al.*, 1999; Székely *et al.*, 1999; Cruz-López *et al.*,

2017; Kupán *et al.*, 2021) suggesting strong competition between females over males with

established nesting territories early in the season. First, we examine whether variation in

mating strategy, egg size, and breeding schedule are age-related and follow a pattern typical

for senescence. Following others (i.e., Bouwhuis *et al.*, 2009; Jankowiak *et al.*, 2018), we

hypothesized that a pattern of early-life increase in egg size and advance in lay date, followed

by a peak maximum, and a decline thereafter, would be indicative of senescence. Conversely,

older females may have a higher likelihood of polyandry owing to experience or competitive

advantages over younger conspecifics (Oring *et al.*, 1994). Second, we assess the seasonal

relationship between egg size and a female's potential for multiple clutches in the context of

reproductive trade-offs associated with female-female competition over mates. Here, we

hypothesized that the likelihood of polyandry would be inversely associated with lay date

whereas egg size and lay date should be positively associated: early breeding females would

have the highest rates of seasonal polyandry but would also lay the smallest eggs. Likewise,

we predicted that early breeding females that lost their first clutch would have the highest
rates of re-nesting but lay smaller eggs than late breeding females.

MATERIALS AND METHODS
Data collection
We studied the reproductive effort and breeding schedules of snowy plovers at Bahía de
Ceuta – an important breeding site located on the coast of Sinaloa, western Mexico (23^o54'N,
106^o57'W). Details on the study site and population are provided elsewhere (e.g., Cruz-
López *et al.*, 2017; Eberhart-Phillips *et al.*, 2020a). In brief, we annually monitored breeding
birds from mid-April until early July, and collected mark-recapture data following the
methods described in (Székely *et al.*, 2011). We searched for nests using telescopes and
mobile hides to minimize disturbance. Upon finding a nest, we measured each eggs' length
and width to the nearest tenth of a mm to determine egg size (Figs. S1a, b). Using these egg
dimensions, we calculated egg volume (Fig. S1c) following (Hoyt, 1979) as:

Eq. 1 $$egg\ volume = K \times length \times width^2,$$

where *K* is 0.486, a volume-index constant for snowy plovers determined by (Székely *et al.*,
1994) through the use of an egg volumeter (Hanson, 1954). The modal clutch size of snowy
plovers is three (86.2%) and is the maximum number of eggs we have observed in this
population (Eberhart-Phillips *et al.*, 2020b). We regularly checked incomplete nests until the
clutch was completed and assigned the age of these nests according to the lay date of the last
egg laid (Plaschke *et al.*, 2019). If the clutch was complete upon discovery and had not been
incubated longer than 10 days, we determined its lay date by floating the egg and estimating
the stage of embryonic development (Nosály & Székely, 1993). For successful clutches that
had been incubated for more than 10 days we back-calculated the laying date based on the
hatching date assuming an incubation period of 25 days (Plaschke *et al.*, 2019). In the rare
case that the nest did not hatch and we discovered it after day 10 of incubation, we assumed
that the nest was 17 days old upon discovery (i.e., the midpoint between the minimum age of
11 days and the 25 day incubation period). In summary, the lay dates for 778 (92.5%) nests
were determined through flotation, 45 (5.4%) were back-calculated from hatch date, and 18
(2.1%) were assumed to be 17 days old at discovery.
202
We identified previously marked nesting adults based on their unique colour ring
combination. We captured unmarked adults on their nests during incubation using funnel

traps (Hall & Cavitt, 2012) and assigned a unique colour ring combination for subsequent

recognition. Because snowy plovers have circadian sex roles during incubation (Vincze *et al.*,

2017), we generally targeted females for captures during the day and males during the night.

In the rare circumstance when we were unable to identify parents before hatching, we

attempted to capture parents while they tended chicks. As snowy plovers only show a small

degree of sexual dimorphism (Küpper *et al.*, 2009), we determined the sex of all captured

plovers in the field through a combination of plumage characteristics (Argüelles-Tico *et al.*,

2015), time of capture, and other behavioural cues (e.g., sex-specific brood care; Kupán *et*

al., 2021). For a subset of adults (57.5%), we confirmed sex molecularly from DNA extracted

from blood samples through PCR amplification of Z and W specific DNA regions with two

sex-typing markers: P2/P8 and Calex-31 (Griffiths *et al.*, 1998; Küpper *et al.*, 2007; dos

Remedios *et al.*, 2010).

We visited known active nests every four or five days to determine the status of the nest (e.g.,

active, depredated, etc.) until the 20th day after egg laying and thereafter daily until the eggs

hatched or failed. We weighed chicks shortly after hatching (879 [84.7%] within 24 hours of

hatching, 159 [15.3%] during the second day after hatching) and marked them with an

alphanumeric metal and a single colour ring for subsequent identification in the chance that

these individuals recruited into the breeding population as adults in future years.

For the years 2006 to 2016 all longitudinal data collected has been compiled as part of the

CeutaOPEN project – an open-access database for individual-based field studies in

evolutionary ecology and conservation biology (Eberhart-Phillips *et al.*, 2020a). We accessed

these data directly from the open source repository (Eberhart-Phillips *et al.*, 2020b) and

supplemented them with data from four additional field seasons: 2017–2020. The

CeutaOPEN database is composed of five tables that correspond to our routine data

collection in the field (Székely & Kosztolányi, 2006). Here we used the "Captures",

"Resights", and "Nests" tables. The "Captures" and "Resights" tables contain information

about all the individuals captured and observed, whereas the "Nests" table contains the

morphometric and spatiotemporal information related to each nest monitored. Please refer to

our RMarkdown vignette that connects to *CeutaOPEN* and reproduces all analytical methods

and results presented below (Supplementary File 1).

Statistical Analyses

Age estimation of individuals with unknown origin—Investigating age-dependent processes

in the wild is challenging as the analyses often involve a mix of individuals that are of known

or unknown age (Colchero *et al.*, 2012) – with the former being initially marked at birth (i.e.,

uncensored), and the latter being immigrants of unknown age or those that were born before

the study's first marking occasion (i.e., left-truncated). To estimate the ages of unknown

individuals in our marked population we employed a capture-mark-recapture analysis using

the 'Bayesian Survival Trajectory Analysis' (BaSTA) package in R (v1.9.4, Colchero *et al.*,

2012), which uses a Bayesian hierarchical framework to fit parametric survival functions of

the marked population while accounting for imperfect detection. Furthermore, BaSTA

derives birth year estimates of left-truncated individuals from the population mean of the

projected survival function. As snowy plovers show prominent sex differences in survival

(Eberhart-Phillips *et al.*, 2017; 2018), we used female-specific survival functions for this

study. Due to high natal dispersal, we could not confidently determine the fate of juveniles

marked in our population. To acknowledge this uncertainty, our capture-mark-recapture

sample only included individuals that survived to their first breeding season, i.e., we

constrained first-year survival probability to 1.

In total, our capture-mark-recapture data comprised records of 450 uniquely marked females,

of which 45 hatched locally and subsequently recruited into the adult population as known-

age individuals (Fig. 1a), and the remaining 405 females were adults of unknown age and

origin. We monitored the presence or absence of marked individuals by recapturing or

observing them in the field in all study years except for 2014, amounting to a total of 916

post-birth annual detections of the 450 females in our sample (median annual detections per

adult = 2; mean = 2.04, 1.45 SD). A logistic bathtub-shaped mortality model had the best fit

to our data (Table S1) – revealing that female mortality rate increased until age 5 years, after

which it became constant (Fig. 1b; see Appendix S1 for detailed methods). Using this model,

we extracted the birth year estimate posteriors for each unknown-age individual in the

capture-mark-recapture sample. Note that three individuals (one first encountered as an adult

[CA1579] and two local recruits [CA2036 and CA1526]; Fig. 2) had been already marked

two years prior to the start of our monitoring period (i.e., pre-2006) and were thus added to

our sample after running BaSTA on the 2006–2020 capture-mark-recapture data.

Modelling individual variation in egg volume ("Egg volume model")—Our sample for

studying egg volume dynamics included 2392 eggs from 841 nests belonging to 426 females.

56 (13.1%) females had three or more years of repeated measures (Fig. 2), 83 (19.5%) had

two years of repeated measures, and 286 (67.4%) were measured in a single year.

Furthermore, 43 (10.1%) individuals in our sample were marked as hatchlings but later

recruited as breeding adults in subsequent years (i.e., known age; Fig. 2a), with the remaining

38 (89.9%) individuals being initially marked as adults (i.e., unknown age; Fig. 2b). We

followed common statistical approaches to investigate senescence in birds (e.g., Bouwhuis *et*

al., 2009; 2010; Schroeder *et al.*, 2012; Herborn *et al.*, 2016; Graham *et al.*, 2019;

Dingemanse *et al.*, 2020) by fitting a quadratic function of age to model age-specific trends in

egg volume. We controlled for selective appearance and disappearance of females differing

in average egg volume by fitting 'first observed age' and 'last observed age' as fixed effects

– a method that estimates between-individual age effects introduced by selective

disappearance and appearance (van de Pol & Verhulst, 2006; Dingemanse *et al.*, 2020).

We modelled within-individual age effects on egg volume by fitting a univariate mixed-effect

model, that included linear and quadratic forms of a within-group deviation score for age

(henceforth 'age-deviance'), calculated for individual i at age j as: age_{ij} – [*first observed age*]$_i$

(van de Pol & Verhulst, 2006; Snijders & Bosker, 2011). Tarsus length was also included as a

fixed effect to control for female structural size, and was averaged over an individual's

measurements (i.e., our *a priori* expectation was that tarsus length is static throughout adult

life and that any variation in this trait was due to measurement error) – grand average 24.73

mm (0.73 SD), grand average within-individual standard deviation 0.22 mm (1.13 SD). In

addition to these fixed covariates, we included a quadratic function of lay date to assess

seasonal variation in egg volume as several shorebird studies report seasonal increases

(Skrade & Dinsmore, 2013; Kwon *et al.*, 2018) or decreases in egg volume (Dittmann &

Hötker, 2001; Skrade & Dinsmore, 2013; Kwon et al., 2018; Kubelka et al., 2020; Verhoeven

et al., 2020). To disentangle within- from between-individual effects in lay date, we used the

same logic as with age above: first lay dates of all individuals each year represented the

between-individual seasonal effect, whereas the deviation in lay dates of an individual

relative to its first nest of the season represented the within-individual seasonal effect. We

included random intercepts for nest, individual, and year, and assumed a Gaussian error

distribution of egg volume.

Modelling seasonal variation in polyandry potential ("Polyandry model")—Our sample for

studying seasonal polyandry dynamics included 426 females for which the identity of their

mates had been verified through observation. We defined observed polyandry as a binomial variable that scored an individual as being monogamous or polyandrous each year based on our observations of them having one or multiple breeding partners, respectively (see Fig. 2 for an example of the sampling distribution). By definition, all polyandrous cases bred at least twice within a season, but also 12.4% of monogamous females were observed breeding more than once. Overall, we observed 92 cases of polyandry from 76 females over the 15-year period (annual average incidence of observed polyandry: 11.8%, range: 0–25.6%). Monogamous females remained with the same partner for another breeding attempt only after their initial attempt had failed. To assess the relationship between the likelihood of polyandry and lay date and age, we fitted a binomial linear mixed effects model that tested the likelihood of polyandry predicted by the fixed effects of lay date (i.e., of an individual's first nest of the season), age-deviance (see above), and first observed age. We included individual and year as random effects.

Modelling seasonal variation in re-nesting potential ("Re-nesting model")—Our sample for studying seasonal re-nesting dynamics included 177 females for which the fate of their first nest of the season had been verified as a failure. We defined re-nesting as a binomial variable that scored an individual as being a re-nester or a single-nester each year based on our observations of them attempting to re-nest after the loss of their first clutch or not, respectively. Overall, we observed 64 cases of re-nesting from 55 females over the 15-year period following a failed attempt (annual average incidence of observed re-nesting: 27%, range: 0–58.3%). Most cases of re-nesting are monogamous in this population (92.4%; see Fig. 2 for an example of the sampling distribution). To assess the relationship between the likelihood of re-nesting and lay date, we fitted a binomial linear mixed effects model that tested the likelihood of re-nesting predicted by the fixed effect of lay date (i.e., of an individual's first nest of the season). We included individual and year as random effects.

*Modelling individual variation in lay dat*e *("Lay date model")*— Modelling the age effects of first nest lay date followed the same logic as the above egg volume model, with a univariate mixed-effect structure that included age-deviance, age-deviance-squared, first observed age, last observed age, and average tarsus length as fixed covariates, and individual and year as random intercepts. Furthermore, recruitment status was also fitted as a two-level fixed effect describing if a breeding female hatched locally ("local recruit") or was first encountered as an adult of unknown origin ("immigrant"). Our sample for studying lay date dynamics used the

same nest-level sample as the polyandry model above, however, as we were interested in how
the recruitment status of an individual influenced breeding phenology, we excluded data from
2006 as this was the first year of our study when all birds were first individually marked. This
resulted in 568 nests from 376 females. We visualized the distribution of lay dates to confirm
normality and to assess the population-level variance in breeding schedule – an indication of
inter-female breeding asynchrony and the intensity of competition for mates (Andersson,
2004).

Evaluating effect sizes and uncertainty—We used the "lme4" (Bates et al., 2015), "rptR"
(Stoffel *et al.*, 2017) and "partR2" (Stoffel *et al.*, 2020) packages in R version "Bunny-
Wunnies Freak Out" (R Core Team, 2020) to conduct our statistical modelling and assessed
homoscedasticity by visually examining the residuals (see Fig. S6). For each of the four
mixed-effect models described above, we evaluated uncertainty in our parameter estimates by
simulating 1000 parametric bootstraps via the "partR2::partR2" function (Stoffel *et al.*,
2020). Likewise we derived nest-, individual-, and year-level repeatabilities (i.e., intra-class
correlations) by simulating 1000 parametric bootstraps of the four mixed-effect models using
"rptR::rpt". We report fixed effects as standardized regression coefficients (i.e., beta weights)
and repeatability as the 'adjusted repeatability' – interpreted as the repeatability of a given
hierarchical group after controlling for fixed effects (Nakagawa & Schielzeth, 2010).

To ensure that intercepts of our age-dependent models represented the reproductive
performance for the earliest age at reproduction (i.e., age 1 in snowy plovers, Page et al.,
2009), we fitted age as '*age* – 1' – otherwise it would represent reproduction as age 0, which
is an empirically meaningless estimate. For the "Egg volume model" and "Lay date model"
we ran an additional simulation that acknowledged uncertainty in the BaSTA age estimate of
a given individual: we bootstrapped each model 1000 times, with every iteration randomly
drawing a birth year estimate for unknown aged individuals from their posterior distributions
provided by BaSTA. For both simulations, we evaluated the influence of birth year
uncertainty by examining the effect size distribution of the 1000 bootstraps in relation to the
95% confidence interval for effect sizes of the original model that used the median birth year
estimate from BaSTA.

RESULTS

The modal clutch size of the 841 clutches was 3 eggs (724 nests, 86.1%; 2-eggs: 103 nests,

12.3%, 1-egg: 14 nests, 1.6%). Average egg length was 3.09 cm (0.10 cm SD, Fig. S1a) and

width was 2.24 cm (0.05 cm SD, Fig. S1b), which translated into an average egg volume of

7.59 cm^3 (0.46 cm^3 SD). The average egg volume of a clutch strongly predicted the average

hatch weight of the subsequent brood (β [95% CIs]: 0.63 [0.55–0.70]; R$^2_{marginal}$ = 0.370

[0.310–0.436]; Figs. S1c and S3, Table S2, see Appendix S2 for methods). Based on

BaSTA's estimated birth year, 185 of the 383 unknown-age females in our sample were first

observed nesting at age 1 (48.3%), 120 at age two (31.3%), 72 at age three (18.8%), five at

age 4 (1.3%), and one at age 5 (0.3%). Of the 43 locally hatched females in our sample, 29

first nested at age one (67.4%), six were first observed nesting at age two (14.0%), two at age

3 (4.7%), three at age 4 (7.0%), three at ages 5, 7, and 8, respectively (7.0%). The average

local tenure of all females in the sample was 1.57 years (2.15 SD) with an average age span

of 3.12 years (2.03 SD, median: 3, range: 1–14 years) and an average of 1.56 years of

observed ages per female (1.04 SD, median: 1, range: 1–8 age-specific observations).

Females in our sample were typically observed nesting every consecutive year since their

first observation, however, some individuals skipped years (Fig. 2, average yearly interval

between nesting attempts = 1.07, 0.27 SD). On average, females made 1.43 (0.56 SD) nesting

attempts per season (median = 1, range 1 to 3).

Individual variation in egg volume

Overall, mixed effects accounted for 71.2% ([68.2, 74.3] 95%CIs) of variation in egg

volume, with fixed effects explaining 7.9% ([5.1, 12.3] 95%CIs) of this variation (Table S3).

Females were highly repeatable in their egg volumes between clutches: r = 0.47 ([0.41, 0.53]

95%CIs; Fig. 3, Table S3). Furthermore, eggs within the same clutch were moderately

repeatable in volume (r = 0.18 [0.14, 0.22]; Fig. 3, Table S3). Senescence in egg volume was

not supported (β_{age} [95% CIs]: 0.00 [-0.06, 0.6], β_{age}^2: -0.05 [-0.09, -0.01]; semi-partial R^2 of

senescence function = 0.003 [0, 0.05]; Fig. 3, Table S3). Furthermore, we found no support

for selective (dis)appearance of individuals according to egg volume, as the 95% CIs for first

and last observed ages of reproduction overlapped zero (Fig. 3, Table S3). The bootstrap

analysis incorporating the individual birth-year posteriors estimated from BaSTA (Fig. S8a–

c) confirmed these results – all iterations fell within the 95% CI of the main model. The

strongest fixed effect explaining egg volume variation was the structural size of the mother

(β_{tarsus} [95% CI]: 0.23 [0.15, 0.29]; semi-partial R^2 of female tarsus = 0.05 [0.02, 0.09]; Figs.

3 and S7b): larger females laid larger eggs than smaller females (model predicted difference:

0.58 cm^3 [0.34, 0.81] 95%CI). The second strongest effect was the between-individual

quadratic season function (Fig. 4c): eggs were smallest at the start of the season (model

prediction: 7.18 cm^3 [6.98, 7.37] 95%CI) and largest shortly after the middle of the season

(model prediction: 7.68 cm^3 [7.60, 7.76] 95%CI). Average egg volume also increased

between sequential clutches within individuals but with smaller magnitude than the

population-level trend (β_{within} [95%CI]: 0.10 [0.06, 0.14], predicted increase of 0.21 cm^3

[0.04, 0.38]; Fig. 3).

Seasonal variation in polyandry and re-nesting potential

A female's probability of being polyandrous was strongly dependent on the lay date of their

first nest with early breeders being more likely to be polyandrous than late breeders (β [95%

CIs]: -2.25 [-3.12, -1.84]; $R^2_{marginal}$ = 0.37 [0.27, 0.50]; Figs. 4a and S4, Table S4). Likewise,

a female's likelihood of re-nesting following a failed attempt decreased with advancing lay

date of the first nest (β [95% CIs]: -1.77 [-3.15, -1.30]; $R^2_{marginal}$ = 0.39 [0.25, 0.58]; Figs. 4e

and S5, Table S5). The lay date distribution of polyandrous females was bimodal, with peaks

in the first and second nests occurring 11.1 days before and 29.1 days after the unimodal

seasonal peak for monogamous females (Fig. 4b). Likewise, the lay date distribution of

renesting females was also bimodal, with peaks in the first and replacement clutches

occurring 9.7 days before and 28.4 days after the unimodal seasonal peak for single nesters

(Fig. 4d). Females had low repeatability in polyandry among years (adjusted individual cross-

year repeatability (r [95% CIs] = 0.01 [0, 0.15]; Table S4, Fig. S4) and we found no evidence

of age-dependent polyandry (Fig. S4).

Individual variation in lay date

Females had moderate repeatability in the lay date of their first nest among years (r = 0.19

[0.06, 0.32] 95%CI; Fig. 3, Table S6). We found strong support for the effect of origin on

first nest lay date: females that locally hatched and later recruited into the breeding

population initiated nests 7.80 days earlier (95% CI: [5.09, 10.50]) on average compared to

conspecifics whose origin was unknown (Figs. 3 and 5b). The next strongest fixed effect was

the within-individual age function predicting the lay date of a female's first nest in the

season: young individuals laid later nests compared to their older conspecifics with lay date

advancing by ~2.17 days per year until age six (95% CI: [1.41, 2.93]; Fig. 5a), after which

the uncertainty in the trend became unwieldly in the oldest age classes of our sample (Fig.

5a). Notably, female size did not affect lay date (Fig. 3 and S7a), nor did between-individual

effects of first or last age at breeding (Fig. 3). The bootstrap analysis incorporating the
individual birth-year posteriors estimated from BaSTA (Fig. S8d–f) confirmed these results –
all iterations fell within the 95% CI of the main model.

DISCUSSION

Identifying trade-offs between reproductive effort and survival in wild organisms is central
for understanding the evolutionary mechanisms of senescence (Lemaître *et al.*, 2015).
Consistent with previous work (Christians, 2002), we found that egg size was highly
repeatable for individual females, even after controlling for their structural size. However,
our data show that egg size variation in snowy plovers is not a senescent trait – but is rather a
seasonally dynamic trait likely linked to female-female competition to breed early. The
distribution of lay dates in this snowy plover population extended over a 112-day period,
indicating high phenological asynchrony within the breeding population, a pre-cursor for
intra-sexual competition (Andersson, 2004). Early nesting females had a much higher
likelihood of being sequentially polyandrous or re-nesting than late breeders, likely because
early nesting allows enough time for multiple chances to re-mate or lay replacement clutches
(Morrison *et al.*, 2019). Early season nesters had on average smaller eggs than females
starting to breed at a later date suggesting a trade-off between mating strategy and gamete
investment (Andersson et al. 2004). This was supported by the observed within-individual
effect: females generally increased egg volume between consecutive nesting attempts, albeit
the effect size was small. A seasonal increase in egg volume at both between- and within-
individual levels may indicate that maternal investment during early breeding attempts is
impaired by a combination of several mutually non-exclusive physiological constraints
including, but not limited to, low food availability (Steigerwald *et al.*, 2015), pressure to
initiate and complete egg-laying quickly (Birkhead & Nettleship, 1982), and carry-over costs
associated with intense competition (Duckworth, 2006) or migratory status (Crossin *et al.*,
2010). In contrast, late breeders can take advantage of a slower pace to gather more resources
and maximize investment during egg production – while having only enough time for a
single breeding attempt.

Despite being long-lived and investing substantially in reproduction year-after-year, we
found no evidence of age-dependent trade-offs in egg size or polyandry potential in this
snowy plover population. Older females tended to initiate nests earlier in the season
compared to their younger conspecifics – indicating age-dependent competitive ability that

could reflect local experience. However, this age-dependent variation in lay date followed a
non-linear pattern: lay date advanced with each year of age until a peak at age six, after
which the age-lay date trajectory was unclear because of limited sampling in older age
classes. Moreover, locally recruited females (i.e., hatched locally) bred earlier than immigrant
females, further suggesting a competitive advantage for individuals with prior experience at
the breeding site (which would also apply to older females that bred at the study site before).
Importantly, polyandry was not repeatable within individuals – likely due to stochastic socio-
ecological dynamics, such as local mate availability and breeding success, which have been
suggested to influence mating tactics in plovers (Carmona-Isunza *et al.*, 2017; Eberhart-
Phillips *et al.*, 2017; Halimubieke *et al.*, 2020).

Several studies of oviparous organisms have observed age-dependent variation in egg size,
with some studies finding a positive relationship (Cooch *et al.*, 1992; Flint & Sedinger, 1992;
Robertson *et al.*, 1994; Warner *et al.*, 2016; Verhoeven *et al.*, 2020) and others observing a
negative relationship (Reid, 1988; Potti, 1993; Ito, 1997). However, earlier studies often
failed to account for selective disappearance (e.g., Cooch *et al.*, 1992; Flint & Sedinger,
1992; Potti, 1993) and thus complicate the interpretation of individual- vs. population-level
effects. Recent longitudinal studies document an increase in egg size in early life, followed
by peak and then a late-life decline (Bouwhuis *et al.*, 2009; Jankowiak *et al.*, 2018). The
early-life increase in egg size may indicate individual improvement through more efficient
foraging abilities, better mate choice, or improved predator avoidance (Forslund & Pärt,
1995). Although we did not find clear statistical support for a relationship between age and
egg volume, we did find strong evidence of females nesting earlier as they aged. We suspect
that this early-life delay in lay date is related to the inferior competitive abilities or lack of
experience that young females have when faced with the seasonal competition for early
breeding opportunities (Oring *et al.*, 1994) – especially when they are migrants and arrive
later then local winter residents (Grist *et al.*, 2017).

In snowy plovers, a completed clutch equals approximately 60% of a female's body mass
(Page *et al.*, 2009) – a substantial investment that young females may not be physiologically
ready for at the start of their first breeding season. Such a constraint would force yearlings to
commence egg laying after the earliest attempts by more mature females – especially when
they hatched late or experienced poor conditions during the juvenile stage. We have recorded
females first breeding as early as 10 months after hatching (Eberhart-Phillips *et al.*, 2020b) –

an age at which females may be unable to invest heavily into reproduction. However, we

have observed some of these locally recruited females already being polyandrous in their first

year meaning that they can be competitive in the female-female competition over mates.

Given that we have documented females still breeding at an age of 13 years (Fig. 2), it is

remarkable that we find no evidence of senescence. Such within-individual consistency over

life becomes relevant when considering the developmental mode of snowy plovers. For

plovers with their nidifugous chicks, small egg volume differences may have significant

ramifications for chick survival (Starck & Ricklefs, 1998), as chicks are not fed by the

parents but rather must forage for themselves immediately after hatching. In shorebirds,

larger chicks that hatch from larger eggs typically survive better than smaller chicks

(Blomqvist et al. 1997), likely owing to the extra nutrients provided by a large egg that make

up for the body mass reduction during the first few days of life when foraging efficiency is

reduced due to learning (Ricklefs, 1968). Moreover, chicks of polyandrous females are

typically cared for solely by their father, forcing them to rely more on their intrinsic reserves

than the added benefits of biparental care. Consequently, comparatively small differences in

egg size could have large knock-on effects for chick survival (Williams, 1994; Starck &

Ricklefs, 1998) – meaning that a females' potential to save resources by reducing egg volume

is limited.

Most studies conducted on temperate or high latitude breeding shorebirds have found a

negative association between time of the season and egg size (Byrkjedal & Kalas, 1985;

Sandercock *et al.*, 1999; Kubelka *et al.*, 2020) although in polyandrous red-necked

phalaropes (*Phalaropus lobatus*) egg size increased across the breeding season (Kwon *et al.*,

2018, n.b., the presented effect size is small). However, many investigations of seasonal egg

size dynamics did not disentangle whether the observed changes were due to within- or

between-individual effects. For example, laying schedules associated with female quality

and/or age could be responsible for much of the observed seasonal variation in egg size. A

study including 15 arctic shorebirds suggested that indeed between-individual variation may

account for more of the seasonal variation in egg size than within-individual variation

(Weiser *et al.*, 2018). We observed a between-individual quadratic effect of time of season on

egg size: early and late season clutches had on average smaller eggs than those nesting at the

middle of the season. The within-individual effect complemented the population-level trend,

with eggs of sequential nests being larger than those of first clutches.

Past studies have linked polyandry and sex-role reversal to reduced female gamete size
(Slotow, 1996; Andersson, 2004), as smaller eggs would permit females to produce several
clutches in rapid succession (Liker *et al.*, 2001). Although it is intuitive to interpret our
results as support for polyandrous females limiting their reproductive investment into the first
clutch by producing small eggs to enable them to quickly produce a second clutch, this is an
unlikely explanation for the observed seasonal variation in egg size. Rather, we argue that the
observed seasonal variation in egg size is a result of mating strategy and environmental
conditions that both favour early breeders. Clutches produced early and late during the season
generally contained smaller eggs (Fig. 4c). Although early breeding by females may be
constrained by obtaining adequate energy reserves for egg production, it provides a head-start
in the competition for mates and allows these females to exploit sequential mating
opportunities later in the season or lay replacement clutches following stochastic failure of
their first attempts (Morrison *et al.*, 2019). Nest failures may put early females on a
monogamous trajectory as re-nesting after nest loss almost always happens with the same
mate (Halimubieke *et al.*, 2019). However, if the clutch is successful, early breeding females
almost always desert broods within a few days of hatching (Kupán et al. 2021) and will
become polyandrous. The reduced egg size at the end of the season can be explained by late
nesting females being under a tight schedule to complete nesting before impending high tides
and precipitation flood the breeding grounds (Plaschke *et al.*, 2019), however, it should be
noted that the late season effect size for egg volume is small and within our margin of
measurement error.

Egg size may also be tied to the survival prospects of the young. Notably in this population,
chick survival is especially high for nests laid at the beginning of the season despite the
smaller egg volumes of early clutches – suggesting a higher resource availability and lower
predation rate for chicks hatching from early clutches (Cruz-López *et al.*, 2017; Kupán *et al.*,
2021). As the season progresses, chick survival declines (Cruz-López *et al.*, 2017) and many
chicks originating from late nesting attempts die from starvation as the local water bodies that
sustain invertebrate prey communities dry out. At the time of laying, late nesting females
undergo oogenesis during the local peak in resource availability, likely producing the larger
eggs found in the latter half of the season – a maternal effect that may help to counter
reduced chick survival when conditions deteriorate 25 days later after hatch. All things
considered, we think that the seasonal constraints of resource availability for breeding plovers

are a better explanation for the differences in reproductive investment by females than the mating strategy *per se*.

One limitation of our study is that some snowy plover females show high breeding dispersal and can produce sequential nests hundreds of kilometres apart (Stenzel *et al.*, 1994; D'Urban Jackson *et al.*, 2020). As our population is open to immigration and emigration, we have likely missed documenting polyandrous breeding attempts at unmonitored neighbouring nesting sites – meaning that we underestimate the true extent of polyandry. Yet, our observed laying distribution (Fig. 4b) suggests that this limitation is not a major concern because we would have otherwise expected a larger share of seemingly monogamous females breeding at the beginning and end of the breeding season – instead we see that monogamous breeders tend to nest in the middle of the season. This is also true for re-nesting activity following failed attempts: birds that did not lay replacement clutches tended to nest in the middle of the season (Fig. 4d). Furthermore, we acknowledge that our relatively small sample of known-age individuals presents a limitation to our study reflecting the challenge of studying an open wild population with high natal dispersal.

In conclusion, we show that egg size variation in snowy plovers is highly repeatable within females and remains stable over life despite a substantial cumulative maternal investment. Egg size variation may be driven by local seasonal fluctuations in resource availability and a reproductive advantage for early nesting females. Our results suggest that senescence is not a major driver of age-dependent dynamics of egg size – a surprising result that is inconsistent with the disposable soma theory. Yet, our results show that prior experience gives older and local females a competitive advantage over younger and naïve conspecifics scrambling for early breeding opportunities that are at the heart of the polyandrous mating strategy (Andersson 2004). We suggest that future research should explore whether and how polygamy and senescence interact to affect reproductive output and what consequences these factors have on offspring survival. Furthermore, future reproductive senescence studies should be wary about cryptic collinearity between age and seasonal trait variation because age-dependent patterns may be misinterpreted when seasonal dynamics are unaccounted for. Studying patterns of senescence in wild populations with flexible mating systems may help shed light on how variation in mating strategies shapes individual life history trajectories and lifetime reproductive success.

ACKNOWLEDGEMENTS

We dedicate this article to our colleague and friend René Beamonte-Barrientos who tragically

passed away in March 2016. René worked on senescence in birds including snowy plovers

and inspired us to conduct this study. We thank Salvador Ángel del Gómez, Miguel Guevara-

Medina and Ivan Guardado for assistance with fieldwork, and Daniel Galindo-Espinosa for

logistic support. We also thank Martin Bulla, Niels Dingemanse, Wolfgang Forstmeier, Bart

Kempenaers, James Tolliver, and Martin Stoffel for their constructive discussions and

statistical advice. Fieldwork was conducted under permit from SEMARNAT

(SGPA/DVGS/2631/19). A full list of funding for fieldwork is provided on

www.chorlito.org. LJE-H was funded by the German Science Foundation (DFG Eigene

Stelle grant: 415037502). AGE-H was funded by the European Union's Horizon 2020

research and innovation programme under the Marie Skłodowska-Curie grant agreement No

793077 and the German Science Foundation (DFG Eigene Stelle grant: 448621093). MCL

was supported by a doctoral grant from CONACyT (248125/378124) and the Cornell Lab of

Ornithology 'Coastal Solutions Fellowship'. CK was supported by the Max Planck Society.

AUTHOR CONTRIBUTIONS

LJE-H, CK, and JS: conceived and supervised the study

LJE-H, LF, CK, MC-L, KAV-R, and EG-M: collected the data

LJE-H, LF, and AGE-H: conducted the analysis and wrote the accompanying RMarkdown

document (Supplementary File 1)

LJE-H, LF, JK, AGE-H, and CK: interpreted the results

LJE-H, LF, and CK: wrote the manuscript

All authors helped to revise the manuscript and approved the final version.

DATA ACCESSIBILITY

We provide all computer code and documentation as a html vignette written in Rmarkdown

(Supplementary File 1) together with all the raw datasets needed to reproduce our modeling

and analyses – these files can be found in this project's Open Science Framework repository

(https://doi.org/10.17605/OSF.IO/UCW6J) and GitHub page

(https://leberhartphillips.github.io/snowy_plover_eggs/).

LITERATURE CITED

Amat, J.A., Fraga, R.M. & Arroyo, G.M. (1999) Brood desertion and polygamous breeding
617   in the Kentish plover *Charadrius alexandrinus*. *Ibis*, **141**, 596–607.
618   Andersson, M. (2004) Social polyandry, parental investment, sexual selection, and evolution
619   of reduced female gamete size. *Evolution*, **58**, 24–34.
620   Andersson, M. (2005) Evolution of classical polyandry: three steps to female emancipation.
621   *Ethology*, **111**, 1–23.
622   Argüelles-Tico, A., Küpper, C., Kelsh, R.N., Kosztolányi, A., Székely, T. & van Dijk, R.E.
623   (2015) Geographic variation in breeding system and environment predicts melanin-
624   based plumage ornamentation of male and female Kentish plovers. *Behavioral Ecology*
625   *and Sociobiology*, **70**, 49–60.
626   Arnold, J.M., Hatch, J.J. & Nisbet, I.C. (2004) Seasonal declines in reproductive success of
627   the common tern *Sterna hirundo*: timing or parental quality? *Journal of Avian Biology*,
628   **35**, 33–45.
629   Bates, D., Mächler, M., Bolker, B. & Walker, S. (2015) Fitting linear mixed-effects models
630   using lme4. *Journal of Statistical Software* **67**,1–48.
631   Beamonte-Barrientos, R., Velando, A., Drummond, H. & Torres, R. (2010) Senescence of
632   maternal effects: aging influences egg quality and rearing capacities of a long‑lived
633   bird. *The American Naturalist*, **175**, 469–480.
634   Birkhead, T.R. & Nettleship, D.N. (1982) The adaptive significance of egg size and laying
635   date in thick-billed Murres *Uria lomvia*. *Ecology*, **63**, 300–306.
636   Blomqvist, D., Johansson, O.C. & Götmark, F. (1997) Parental quality and egg size affect
637   chick survival in a precocial bird, the lapwing *Vanellus vanellus*. *Oecologia*, **110**, 18–
638   24.
639   Borgmann, K.L., Conway, C.J. & Morrison, M.L. (2013) Breeding phenology of birds:
640   mechanisms underlying seasonal declines in the risk of nest predation. *PLoS ONE*, **8**,
641   e65909.
642   Bouwhuis, S., Charmantier, A., Verhulst, S. & Sheldon, B.C. (2010) Trans-generational
643   effects on ageing in a wild bird population. *Journal of Evolutionary Biology*, **23**, 636–
644   642.
645   Bouwhuis, S., Sheldon, B.C., Verhulst, S. & Charmantier, A. (2009) Great tits growing old:
646   selective disappearance and the partitioning of senescence to stages within the breeding
647   cycle. *Proceedings of the Royal Society B: Biological Sciences*, **276**, 2769–2777.
648   Byrkjedal, I. & Kalas, J.A. (1985) Seasonal variation in egg size in golden plover *Pluvialis*
649   *apricaria* and dotterel *Charadrius morinellus* populations. *Ornis Scandinavica*, **16**,
650   108–112.
651   Carmona-Isunza, M.C., Ancona, S., Székely, T., Ramallo-González, A.P., Cruz-López, M.,
652   Serrano-Meneses, M.A., *et al.* (2017) Adult sex ratio and operational sex ratio exhibit
653   different temporal dynamics in the wild. *Behavioral Ecology*, **28**, 523–532.
654   Christians, J.K. (2002) Avian egg size: variation within species and inflexibility within
655   individuals. *Biological Reviews of the Cambridge Philosophical Society*, **77**, 1–26.

Colchero, F., Jones, O.R. & Rebke, M. (2012) BaSTA: An R package for Bayesian
estimation of age-specific survival from incomplete mark-recapture/recovery data with
covariates. *Methods in Ecology and Evolution*, **3**, 466–470.

Colwell, D.M.A. (2010) *Shorebird Ecology, Conservation, and Management*. University of
California Press.

Colwell, M.A. & Oring, L.W. (1988) Breeding biology of Wilson's Phalarope in southcentral
Saskatchewan. *The Wilson Bulletin*, 567–582.

Colwell, M.A., Feucht, E.J., McAllister, S.E. & Transou, A.N. (2017) Lessons learned from
the oldest Snowy Plover. *Wader Study*, **124**, 157–159.

Cooch, E.G., Lank, D.B., Rockwell, R.F. & Cooke, F. (1992) Is There a Positive Relationship
between Body Size and Fecundity in Lesser Snow Geese? *The Auk*, **109**, 667–673.

Cox, D.R. & Oakes, D. (1984) *Analysis of Survival Data*. Chapman and Hall, London.

Crossin, G.T., Trathan, P.N., Phillips, R.A., Dawson, A., Le Bouard, F. & Williams, T.D.
(2010) A carryover effect of migration underlies individual variation in reproductive
readiness and extreme egg size dimorphism in macaroni penguins. *The American*
Naturalist, **176**, 357–366.

Cruz-López, M., Eberhart-Phillips, L.J., Fernández, G., Beamonte-Barrientos, R., Székely,
T., Serrano-Meneses, M.A., *et al.* (2017) The plight of a plover: viability of an
important snowy plover population with flexible brood care in Mexico. *Biological*
Conservation, **209**, 440–448.

Curio, E. (1983) Why de young birds reproduce less well? *Ibis*, **125**, 400–404.

D'Urban Jackson, J., Bruford, M.W., Székely, T., DaCosta, J.M., Sorenson, M.D., Russo, I.-
R.M., *et al.* (2020) Population differentiation and historical demography of the
threatened snowy plover *Charadrius nivosus* (Cassin, 1858). *Conservation Genetics*,
21, 387–404.

Deeming, D.C. & Reynolds, S.J. (2015) Nests, eggs, and incubation new ideas about avian
reproduction. *Avian Biology Research*, **8**, 193.

Dingemanse, N.J., Moiron, M., Araya-Ajoy, Y.G., Mouchet, A. & Abbey-Lee, R.N. (2020)
Individual variation in age-dependent reproduction: Fast explorers live fast but senesce
young? *Journal of Animal Ecology*, **89**, 601–613.

dos Remedios, N., Lee, P.L., Székely, T., Dawson, D.A. & Küpper, C. (2010) Molecular sex-
typing in shorebirds: a review of an essential method for research in evolution, ecology
and conservation. *Wader Study Group Bulletin*, **117**, 109–118.

Drenos, F. & Kirkwood, T.B.L. (2005) Modelling the disposable soma theory of ageing.
Mechanisms of Ageing and Development, **126**, 99–103.

Duckworth, R.A. (2006) Behavioral correlations across breeding contexts provide a
mechanism for a cost of aggression. *Behavioral Ecology*, **17**, 1011–1019.

Eberhart-Phillips, L.J. (2019) Plover breeding systems: diversity and evolutionary origins. In
The population ecology and conservation of Charadrius plovers (ed. by Colwell, M.A.
& Haig, S.M.). pp. 63–88.

Eberhart-Phillips, L.J., Cruz-López, M., Lozano-Angulo, L., Gómez del Ángel, S., Rojas-
Abreu, W., Bucio-Pacheco, M., *et al.* (2020a) CeutaOPEN, individual-based field
observations of breeding snowy plovers *Charadrius nivosus*. *Scientific Data*, **7**, 149.

Eberhart-Phillips, L.J., Cruz-López, M., Lozano-Angulo, L., Gómez del Ángel, S., Rojas-
Abreu, W., Bucio-Pacheco, M., *et al.* (2020b) CeutaOPEN v1.5. Open Science
Framework.

Eberhart-Phillips, L.J., Küpper, C., Carmona-Isunza, M.C., Vincze, O., Zefania, S., Cruz-
López, M., *et al.* (2018) Demographic causes of adult sex ratio variation and their
consequences for parental cooperation. *Nature Communications*, **9**, 1–8.

Eberhart-Phillips, L.J., Küpper, C., Miller, T.E.X., Cruz-López, M., Maher, K.H., Remedios,
dos, N., *et al.* (2017) Sex-specific early survival drives adult sex ratio bias in snowy
plovers and impacts mating system and population growth. *Proceedings of the National
Academy of Sciences of the United States of America*, **114**, e5474–e5481.

Emlen, S.T. & Oring, L.W. (1977) Ecology, sexual selection, and the evolution of mating
systems. *Science*, **197**, 215–223.

Flint, P.L. & Sedinger, J.S. (1992) Reproductive implications of egg-size variation in the
black brant. *The Auk*, **109**, 896–903.

Forslund, P. & Pärt, T. (1995) Age and reproduction in birds - hypotheses and tests. *Trends in
Ecology & Evolution*, **10**, 374–378.

Fox, C.W. (1994) the influence of egg size on offspring performance in the seed beetle,
Callosobruchus maculatus. *Oikos*, **71**, 321–325.

Gompertz, B. (1812) On the nature of the function expressive of the law of human mortality,
and on a new mode of determining the value of life contingencies. *Philosophical
Transactions of the Royal Society of London B: Biological Sciences*, **115**, 513–583.

Goymann, W., Makomba, M., Urasa, F. & Schwabl, I. (2015) Social monogamy vs.
polyandry: ecological factors associated with sex roles in two closely related birds
within the same habitat. *Journal of Evolutionary Biology*, **28**, 1335–1353.

Graham, J.L., Bauer, C.M., Heidinger, B.J., Ketterson, E.D. & Greives, T.J. (2019) Early-
breeding females experience greater telomere loss. *Molecular Ecology*, **28**, 114–126.

Griffiths, R., Double, M.C., Orr, K. & Dawson, R.J.G. (1998) A DNA test to sex most birds.
Molecular Ecology, **7**, 1071–1075.

Grist, H., Daunt, F., Wanless, S., Burthe, S.J., Newell, M.A., Harris, M.P., *et al.* (2017)
Reproductive performance of resident and migrant males, females and pairs in a
partially migratory bird. *Journal of Animal Ecology*, **86**, 1010–1021.

Halimubieke, N., Kupán, K., Valdebenito, J.O., Kubelka, V., Carmona-Isunza, M.C., Burgas,
D., *et al.* (2020) Successful breeding predicts divorce in plovers. *Scientific Reports*, **10**,
15576–13.

Halimubieke, N., Valdebenito, J.O., Harding, P., López, M.C., Meneses, M.A.S., James, R.,
et al. (2019) Mate fidelity in a polygamous shorebird, the snowy plover (*Charadrius
nivosus*). *Ecology and Evolution*, **28**, 523.

Hall, L.K. & Cavitt, J.F. (2012) Comparative study of trapping methods for ground-nesting
shorebirds. *Waterbirds*, **35**, 342–346.

Hammers, M., Richardson, D.S., Burke, T. & Komdeur, J. (2012) Age-dependent terminal
declines in reproductive output in a wild bird. *PLoS ONE*, **7**, e40413.

Hanson, H.C. (1954) Apparatus for the study of incubated bird eggs. *The Journal of Wildlife
Management*, **18**, 191–198.

Herborn, K.A., Daunt, F., Heidinger, B.J., Granroth-Wilding, V.H.M., Burthe, S.J., Newell,
M.A., *et al.* (2016) Age, oxidative stress exposure and fitness in a long-lived seabird.
Functional Ecology, **30**, 913–921.

Hoyt, D.F. (1979) Practical methods of estimating volume and fresh weight of bird eggs. *The*
Auk, **96**, 73–77.

Ito, K. (1997) Egg-size and -number variations related to maternal size and age, and the
relationship between egg size and larval characteristics in an annual marine gastropod,
Haloa japonica (Opisthobranchia; Cephalaspidea). *Marine Ecology Progress Series*,
152, 187–195.

Jankowiak, Ł., Zyskowski, D. & Wysocki, D. (2018) Age-specific reproduction and
disposable soma in an urban population of Common Blackbirds *Turdus merula*. *Ibis*,
160, 130–144.

Kaplan, R.H. (1980) The Implications of Ovum Size Variability for Offspring Fitness and
Clutch Size within Several Populations of Salamanders (Ambystoma). *Evolution*, **34**,
51–64.

Kirkwood, T.B.L. (1977) Evolution of ageing. *Nature*, **270**, 301–304.

Kirkwood, T.B.L. & Austad, S.N. (2000) Why do we age? *Nature*, **408**, 233–238.

Kirkwood, T.B.L. & Rose, M.R. (1991) Evolution of senescence: late survival sacrificed for
reproduction. *Philosophical Transactions of the Royal Society B-Biological Sciences*,
332, 15–24.

Kubelka, V., Sládeček, M., Zámečník, V., Vozabulová, E. & Šálek, M. (2020) Seasonality
predicts egg size better than nesting habitat in a precocial shorebird. *Ardea*, **107**, 239.

Kupán, K., Székely, T., Cruz-López, M., Seymour, K. & Küpper, C. (2021) Offspring
desertion with care? Chick mortality and plastic female desertion in Snowy Plovers.
Behavioral Ecology.

Küpper, C., Augustin, J., Kosztolányi, A., Burke, T., Flguerola, J. & Székely, T. (2009)
Kentish versus snowy plover: phenotypic and genetic analyses of *Charadrius*
alexandrinus reveal divergence of Eurasian and American subspecies. *The Auk*, **126**,
839–852.

Küpper, C., Horsburgh, G.J., Dawson, D.A., ffrench-Constant, R., Székely, T. & Burke, T.
(2007) Characterization of 36 polymorphic microsatellite loci in the Kentish plover
(*Charadrius alexandrinus*) including two sex-linked loci and their amplification in four
other Charadrius species. *Molecular Ecology Notes*, **7**, 35–39.

Kwon, E., English, W.B., Weiser, E.L., Franks, S.E., Hodkinson, D.J., Lank, D.B., *et al.*
(2018) Delayed egg-laying and shortened incubation duration of Arctic-breeding
shorebirds coincide with climate cooling. *Ecology and Evolution*, **8**, 1339–1351.

Lack, D.L. (1968) *Ecological adaptations for breeding in birds*.

Lemaître, J.-F., Berger, V., Bonenfant, C., Douhard, M., Gamelon, M., Plard, F., *et al.* (2015)
Early-late life trade-offs and the evolution of ageing in the wild. *Proceedings of the*
Royal Society B: Biological Sciences, **282**, 20150209.

Liker, A., Reynolds, J.D. & Székely, T. (2001) The evolution of egg size in socially
polyandrous shorebirds. *Oikos*, **95**, 3–14.

Medawar, P.B.P. (1952) *An Unsolved Problem of Biology*. University College London,
London.

Moran, A.L. & Emlet, R.B. (2001) Offspring size and performance in variable environments: field studies on a marine snail. *Ecology*, **82**, 1597–1612.

Morrison, C.A., Alves, J.A., Gunnarsson, T.G., Þórisson, B. & Gill, J.A. (2019) Why do earlier-arriving migratory birds have better breeding success? *Ecology and Evolution*, **9**, 8856–8864.

Nakagawa, S. & Schielzeth, H. (2010) Repeatability for Gaussian and non‐Gaussian data: a practical guide for biologists. *Biological Reviews*, **85**, 935–956.

Nosály, G. & Székely, T. (1993) Clutch and egg-size variation in the Kentish Plover (*Charadrius alexandrinus*) during the breeding season. *Aquila*, **100**, 161–179.

Nussey, D.H., Coulson, T., Festa-Bianchet, M. & Gaillard, J.M. (2008) Measuring senescence in wild animal populations: Towards a longitudinal approach. *Functional Ecology*, **22**, 393–406.

Ockendon, N., Leech, D. & Pearce-Higgins, J.W. (2013) Climatic effects on breeding grounds are more important drivers of breeding phenology in migrant birds than carry-over effects from wintering grounds. *Biology Letters*, **9**, 20130669.

Oring, L.W. (1986) Avian Polyandry. In *Current Ornithology, vol 3*, Current Ornithology, vol 3 (ed. by Johnston, R.F.). pp. 309–351.

Oring, L.W., Reed, M.J. & Alberico, A.J. (1994) Mate acquisition tactics in polyandrous spotted sandpipers (*Actitis macularia*): the role of age and experience. *Behavioral Ecology*, **5**, 9–16.

Page, G.W., Stenzel, L.E., Warriner, J.S., Warriner, J.C. & Paton, P.W. (2009) Snowy Plover (*Charadrius nivosus*). *The Birds of North America Online*, 71935–72002.

Pinder, J.E., Wiener, J.G. & Smith, M.H. (1978) The Weibull distribution: a new method of summarizing survivorship data. *Ecology*, **59**, 175–179.

Plaschke, S., Bulla, M., Cruz-López, M., Gómez del Ángel, S. & Küpper, C. (2019) Nest initiation and flooding in response to season and semi-lunar spring tides in a ground-nesting shorebird. *Frontiers in Zoology*, **16**, 15–11.

Pletcher, S. (1999) Model fitting and hypothesis testing for age-specific mortality data. *Journal of Evolutionary Biology*, **12**, 430–439.

Potti, J. (1993) Environmental, ontogenetic, and genetic variation in egg size of Pied Flycatchers. *Canadian Journal of Zoology-revue Canadienne De Zoologie - CAN J ZOOL*, **71**, 1534–1542.

R Core Team. (2020) *R: A Language and Environment for Statistical Computing*. R Foundation for Statistical Computing, Vienna, Austria. https://www.R-project.org/.

Rahn, H., Paganelli, V.C. & Ar, A. (1975) Relation of avian egg weight to body weight. *The Auk*, **92**, 750–765.

Reid, V.W. (1988) Age-specific patterns of reproduction in the glaucous-winged gull: increased effort with age. *Ecology*, **69**, 1454–1465.

Ricklefs, R.E. (1968) Patterns of growth in birds. *Ibis*, **110**, 419–451.

Robertson, G.J., Cooch, E.G., Lank, D.B., Rockwell, R.F. & Cooke, F. (1994) Female age and egg size in the lesser snow goose. *Journal of Avian Biology*, **25**, 149.

Salguero-Gómez, R., Jones, O.R., Jongejans, E., Blomberg, S.P., Hodgson, D.J., Mbeau-Ache, C., *et al.* (2016) Fast-slow continuum and reproductive strategies structure plant

life-history variation worldwide. *Proceedings of the National Academy of Sciences of*
the United States of America, **113**, 230–235.

Sandercock, B.K., Lank, D.B. & Cooke, F. (1999) Seasonal declines in the fecundity of
arctic-breeding sandpipers: different tactics in two species with an invariant clutch size.
Journal of Avian Biology, **30**, 460.

Schamel, D., Tracy, D.M. & Lank, D.B. (2004) Male mate choice, male availability and egg
production as limitations on polyandry in the red-necked phalarope. *Animal Behaviour*,
67, 847–853.

Schroeder, J., Burke, T., Mannarelli, M.E., Dawson, D.A. & Nakagawa, S. (2012) Maternal
effects and heritability of annual productivity. *Journal of Evolutionary Biology*, **25**,
149–156.

Siler, W. (1979) A Competing Risk Model for Animal Mortality. *Ecology*, **60**, 750–757.

Skrade, P. & Dinsmore, S.J. (2013) Egg-size investment in a bird with uniparental incubation
by both sexes. *The Condor*, **115**, 508–514.

Slotow, R. (1996) Black coucal *Centropus grillii* egg volume predicts their polyandrous
mating system. *Journal of Avian Biology*, 171–173.

Snijders, T. & Bosker, R.J. (2011) *Multilevel analysis: An introduction to basic and*
advanced multilevel modeling. 2nd edn. Sage.

Spiegelhalter, D.J., Best, N.G., Carlin, B.P. & Van Der Linde, A. (2002) Bayesian measures
of model complexity and fit. *Journal of the Royal Statistical Society: Series B*
(Statistical Methodology), **64**, 583–639.

Starck, J.M. & Ricklefs, R.E. (1998) *Avian growth and development. Evolution within the*
altricial-precocial spectrum. New York.

Steigerwald, E.C., Igual, J.-M., Payo-Payo, A. & Tavecchia, G. (2015) Effects of decreased
anthropogenic food availability on an opportunistic gull: evidence for a size-mediated
response in breeding females. *Ibis*, **157**, 439–448.

Stenzel, L.E., Warriner, J.C., Warriner, J.S., Wilson, K.S., Bidstrup, F.C. & Page, G.W.
(1994) Long-distance breeding dispersal of snowy plovers in western North America.
Journal of Animal Ecology, **63**, 887–902.

Stoffel, M.A., Nakagawa, S. & Schielzeth, H. (2017) rptR: repeatability estimation and
variance decomposition by generalized linear mixed effects models. *Methods in*
Ecology and Evolution, **8**, 1639–1644.

Stoffel, M.A., Nakagawa, S. & Schielzeth, H. (2020) partR2: Partitioning R2 in generalized
linear mixed models. *bioRxiv*, 2020.07.26.221168.

Székely, T. & Kosztolányi, A. (2006) Practical guide for investigating breeding ecology of
Kentish plover *Charadrius alexandrinus*. *University of Bath*.

Székely, T., Argüelles-Tico, A., Kosztolányi, A. & Küpper, C. (2011) Practical guide for
investigating breeding ecology of Kentish plover, 1–16.

Székely, T., Cuthill, I.C. & Kis, J. (1999) Brood desertion in Kentish plover sex differences
in remating opportunities. *Behavioral Ecology*, **10**, 185–190.

Székely, T., Kozma, J. & Piti, A. (1994) The volume of snowy plover eggs (el volumen de
los huevos de *Charadrius alexandrinus*). *Journal of Field Ornithology*, **65**, 60–64.

van de Pol, M. & Verhulst, S. (2006) Age☐dependent traits: a new statistical model to
separate within☐ and between☐individual effects. *The American Naturalist*, **167**, 766–
773.
Verhoeven, M.A., Loonstra, A.H.J., McBride, A.D., Tinbergen, J.M., Kentie, R.,
Hooijmeijer, J.C.E.W., *et al.* (2020) Variation in egg size of black-tailed godwits.
Ardea, **107**, 291.
Vincze, O., Kosztolányi, A., Barta, Z., Küpper, C., Alrashidi, M., Amat, J.A., *et al.* (2017)
Parental cooperation in a changing climate: fluctuating environments predict shifts in
care division. *Global Ecology and Biogeography*, **26**, 347–358.
Warner, D.A., Miller, D.A.W., Bronikowski, A.M. & Janzen, F.J. (2016) Decades of field
data reveal that turtles senesce in the wild. *Proceedings of the National Academy of
Sciences of the United States of America*, **113**, 6502–6507.
Warriner, J.S., Warriner, J.C., Page, G.W. & Stenzel, L.E. (1986) Mating system and
reproductive success of a small population of polygamous snowy plovers. *The Wilson
Bulletin*, **98**, 15–37.
Weiser, E.L., Brown, S.C., Lanctot, R.B., Gates, H.R., Abraham, K.F., Bentzen, R.L., *et al.*
(2018) Life-history tradeoffs revealed by seasonal declines in reproductive traits of
Arctic-breeding shorebirds. *Journal of Avian Biology*, **49**, jav–01531.
Williams, G.C. (1957) Pleiotropy, natural selection, and the evolution of senescence.
Evolution, **11**, 398.
Williams, T.D. (1994) Intraespecific variation in egg size and egg composition in birds:
effects on offspring fitness. *Biological Reviews of the Cambridge Philosophical
Society*, **69**, 35–59.
Williams, T.D. (2012) *Physiological adaptations for breeding in birds*. Princenton.
Xu, F., Yang, W. & Li, Y. (2019) Enlarged egg size increases offspring fitness of a frog
species on the zhoushan archipelago of China. *Scientific Reports*, **9**, 11653.
Zhang, H., Vedder, O., Becker, P.H. & Bouwhuis, S. (2015) Age-dependent trait variation:
the relative contribution of within-individual change, selective appearance and
disappearance in a long-lived seabird. *Journal of Animal Ecology*, **84**, 797–807.

Figure 1. Logistic-bathtub mortality function for snowy plover females: a) frequency
distribution of age-specific observations of 45 known-aged females and b) age-dependent
mortality hazard.

Figure 2. Mating strategy and clutch number of female snowy plovers according to age. Each
row shows an individual female in the population for which we have at least three years of

observations (note that our analysis also includes females with one or two years of
observation, but given space constraints only individuals with a minimum of three years are
plotted in this graph). Panel A) shows known-aged females which were born locally, whereas
B) shows females that were initially captured as adults and are therefore of unknown age.
Points illustrate the age at which we collected observations of egg volume, with the size of
the point corresponding to the number of clutches measured at a given age, and the colour
indicating if we observed the female mating with one or two distinct males (i.e., in case of
multiple clutches at a given age). The light grey buffer around unknown-age females
indicates the 95% CrI of the ages for an individuals' observed period (i.e., lower limit
indicates the minimum age the individual could have entered the population and the upper
limit indicates the maximum age of an individual's last observation based on BaSTA's birth
year posterior). The dark grey lines indicate the period for which an individual was observed
alive (i.e., in some cases we encountered an individual in the field and confirmed its survival,
but we did not observe its nest to be able to measure the eggs. Note also that the age at first
encounter of all known-aged individuals is 0).

Figure 3. Sources of egg volume (left) and lay date (right) variation. Top row: standardized effect sizes (± 95% CI) of fixed effects. Middle row: variance explained by fixed effects (note: the term 'Senescence' describes the collective variation explained by the linear and quadratic within-individual age effects in the top row; the term 'Selective appearance' and 'Selective disappearance' describe the variation explained by the between individual first- and last-breeding age fixed effects of the top row, respectively; the term 'Within ind. seasonality' describes the variation explained by the within individual lay date effect in the top row, and the term 'Between ind. seasonality' describes the collective variation explained by the linear and quadratic lay date effects in the top row). Bottom row: adjusted repeatabilities (± 95% CI) of random effects.

Figure 4. Phenology of mating behaviours, re-nesting, egg laying, and egg size in 426 female snowy plovers breeding at Bahía de Ceuta. A) Relationship between polyandry potential and lay date of a female's first nest of the season. Each datum is the lay date of an individual's first nest and their observed local mating behaviour of each year. B) Lay date distributions of all nests for females that were polyandrous (yellow; 92 nests from 76 females) or monogamous (green; 573 nests from 403 females). C) Seasonal variation in egg volume – trend shows the between-individual polynomial function of the model prediction. Each datum

is an egg's volume (cm^3) and lay date. D) Lay date distributions of all nests for females that

laid replacement clutches (orange; 64 nests from 127 females) or did not re-nest (purple; 146

nests from 134 females) following failure of their first nest. E) Relationship between re-

nesting potential and lay date of a female's first nest of the season. As with panel (A), each

datum is the lay date of an individual's first nest and their observed local re-nesting activity

of each year. Late date is standardized for each year across all panels.

Figure 5. Age- and origin-dependent breeding phenology of female snowy plovers. A)
Within-individual variation in age-specific nest initiation date – as females gained more
experience in the local population, they started nesting earlier, however this trend reversed at
older ages (albeit with high uncertainty). Each datum represents an individual's 'age-
deviance' (i.e., a within-group centred measure of the number of years since the individual's
first observed local breeding attempt, see *Methods* for more details) and the lay date of its
first nest each year. B) Origin-specific variation in nest initiation date – females that hatched
locally and recruited into the breeding population (orange) tended to nest earlier than birds
originating from elsewhere (purple). Inner-most distributions show the model estimates and
95% CI, outer-most box plots show the inter-quartile ranges of the raw data (point-cloud).

SUPPLEMENTARY MATERIAL

Table S1. Hierarchical capture-mark-recapture models used to describe mortality patterns of
snowy plovers at Bahía de Ceuta, Mexico, between 2006 and 2020. *k:* number of modelled
parameters; ΔDIC: difference in deviance information criterion (*DIC*) between a given model
and the top model. *DIC* value of top model was 5279.1.

Mortality function	Shape	k	ΔDIC
Logistic	Bathtub	7	0.00
Logistic	Simple	4	90.91
Logistic	Makeham	5	97.01
Weibull	Bathtub	6	139.36
Weibull	Simple	3	211.98
Weibull	Makeham	4	226.01
Gompertz	Bathtub	6	328.12
Gompertz	Simple	3	401.06
Gompertz	Make	4	445.12
Exponential	Simple	2	679.94

959   **Table S2.** Relationship between the average egg volume of a clutch and the average chick

960   weight at hatching. See Appendix S2 for methods.

Chick hatch weight	Mean estimate	95% confidence interval
Fixed effects β (standardized)		
Average egg volume of clutch	0.60	[0.55, 0.66]
Partitioned R^2		
Total Marginal R^2	0.37	[0.3, 0.43]
Total Conditional R^2	0.47	[0.38, 0.56]
Random effects σ^2		
Mother identity	0.01	[0, 0.02]
Year	0.01	[0, 0.02]
Residual	0.09	[0.08, 0.11]
Adjusted repeatability r		
Mother identity	0.05	[0, 0.12]
Year	0.05	[0.01, 0.11]
Residual	0.90	[0.81, 0.98]
Sample sizes n		
Year	14	
Individuals	425	
Observations (i.e., Nests)	840	

Table S3. (*displayed on next page*) Sources of egg size variation. Note: within the

Partitioned R^2 sub-table the term 'Senescence' describes the collective variation explained by

the linear and quadratic within-individual age effects in the top sub-table, the term 'Selective

appearance' and 'Selective disappearance' describe the variation explained by the between

individual first- and last-breeding age fixed effects of the top sub-table, respectively; the term

'Within ind. seasonality' describes the variation explained by the within individual lay date

effect in the top sub-table, and the term 'Between ind. seasonality' describes the collective

variation explained by the linear and quadratic lay date effects in the top sub-table.

 Table S3. (*caption on previous page*)

Egg volume	Mean estimate	95% confidence interval
Fixed effects β (standardized)		
Mother tarsus length	0.23	[0.15, 0.29]
Within ind. linear age	−0.00	[-0.06, 0.06]
Within ind. quadratic age	−0.05	[-0.09, -0.01]
Between ind. first breeding age	0.04	[-0.04, 0.14]
Between ind. last breeding age	0.02	[-0.1, 0.15]
Within ind. lay date	0.10	[0.06, 0.14]
Between ind. linear lay date	0.16	[0.11, 0.21]
Between ind. quadratic lay date	−0.07	[-0.12, -0.01]
Partitioned R^2		
Total Marginal R^2	0.08	[0.05, 0.12]
Total Conditional R^2	0.71	[0.68, 0.74]
Mother tarsus length	0.05	[0.02, 0.09]
Senescence	0.00	[0, 0.05]
Selective appearance	0.00	[0, 0.05]
Selective disappearance	0.00	[0, 0.05]
Between ind. seasonality	0.01	[0, 0.06]
Within ind. seasonality	0.01	[0, 0.05]
Random effects σ^2		
Nest / Individual	0.04	[0.03, 0.04]
Individual	0.10	[0.08, 0.12]
Year	0.01	[0, 0.02]
Residual	0.06	[0.06, 0.07]
Adjusted repeatability r		
Nest / Individual	0.18	[0.14, 0.22]
Individual	0.47	[0.41, 0.53]
Year	0.04	[0.01, 0.08]
Residual	0.31	[0.28, 0.35]
Sample sizes n		
Years	14	
Individuals	426	
Nests	841	
Observations (i.e., Eggs)	2392	

970 **Table S4.** Relationship between the lay date of an individual's first nest and their likelihood
971 of polyandry each year. Fixed effect size of 'first nest lay date' is the standardized estimate
972 on the logit scale.

Polyandry probability	Mean estimate	95% confidence interval
Fixed effects β (logit standardized)		
First nest lay date	−2.25	[-3.12, -1.84]
Age since first breeding	−0.00	[-0.32, 0.28]
First breeding age	0.22	[-0.04, 0.53]
Partitioned R^2		
Total Marginal R^2	0.37	[0.27, 0.5]
First nest lay date	0.37	[0.26, 0.5]
Age since first breeding	0.00	[0, 0.23]
First breeding age	0.00	[0, 0.23]
Total Conditional R^2	0.39	[0.28, 0.55]
Random effects σ^2		
Individual	0.07	[0, 1.41]
Year	0.18	[0, 0.63]
Adjusted repeatability r		
Individual	0.01	[0, 0.15]
Year	0.02	[0, 0.07]
Sample sizes n		
Years	14	
Individuals	426	
Observations (i.e., Nests)	665	

973 **Table S5.** Relationship between the lay date of an individual's first nest and their likelihood
974 of re-nesting following breeding failure in a given year. Fixed effect size of 'first nest lay
975 date' is the standardized estimate on the logit scale.

Re-nesting probability	Mean estimate	95% confidence interval
Fixed effects β (logit standardized)		
First nest lay date	−1.77	[-3.15, -1.3]
Partitioned R^2		
Total Marginal R^2	0.39	[0.25, 0.58]
First nest lay date	0.39	[0.25, 0.58]
Total Conditional R^2	0.41	[0.27, 0.63]
Random effects σ^2		
Individual	0.06	[0, 1.74]
Year	0.09	[0, 0.71]
Adjusted repeatability r		
Individual	0.01	[0, 0.21]
Year	0.02	[0, 0.1]
Sample sizes n		
Years	13	
Individuals	177	
Observations (i.e., Nests)	209	

976 **Table S6.** Sources of lay date variation. Note: within the *Partitioned R^2* sub-table the term
977 'Senescence' describes the collective variation explained by the linear and quadratic within-
978 individual age effects in the top sub-table, the term 'Selective appearance' and 'Selective
979 disappearance' describe the variation explained by the between individual first- and last-
980 breeding age fixed effects of the top sub-table, respectively.

Lay date of first nest	Mean estimate	95% confidence interval
Fixed effects β (standardized)		
Mother tarsus length	0.02	[-0.07, 0.1]
Local recruit	−0.41	[-0.68, -0.15]
Within ind. linear age	−0.26	[-0.41, -0.11]
Within ind. quadratic age	0.04	[0.01, 0.08]
Between ind. first breeding age	0.06	[-0.04, 0.17]
Between ind. last breeding age	−0.04	[-0.19, 0.11]
Partitioned R^2		
Total Marginal R^2	0.06	[0.04, 0.12]
Total Conditional R^2	0.23	[0.13, 0.38]
Mother tarsus length	0.00	[0, 0.06]
Local recruit	0.02	[0, 0.08]
Senescence	0.02	[0, 0.07]
Selective appearance	0.01	[0, 0.06]
Selective disappearance	0.00	[0, 0.06]
Random effects σ^2		
Individual	60.69	[18.94, 117.67]
Year	2.00	[0, 12.89]
Residual	283.38	[229.95, 326.82]
Adjusted repeatability r		
Individual	0.19	[0.06, 0.32]
Year	0.01	[0, 0.03]
Residual	0.80	[0.67, 0.93]
Sample sizes n		
Years	13	
Individuals	376	
Observations (i.e., Nests)	568	

Figure S1. Egg size variation for snowy plovers breeding in Bahía de Ceuta, Mexico
between 2006 and 2020. Egg volume was estimated using an allometric equation combining
length (A) and width (B) measurements collected in the field (see *Methods*). C) Average egg
volume of a clutch was highly correlated with the clutch's average chick weight at hatch (β
[95% CIs]: 0.628 [0.552–0.704]; $R^2_{marginal}$ = 0.370 [0.310–0.436]). Ribbon shows the 95% CI
of the model prediction, vertical and horizontal error bars visualize within clutch variation in
egg volume and chick hatch weight (1 SD); see Appendix S2 for modelling methods.

Figure S2. Bayesian model diagnostics of the Logistic bathtub shaped survival model. a)
Serial autocorrelation of iterations within chains, b) chain convergence (*R-hat*) of model, and
c) trace plots of chain dynamics.

chick model

991   **Figure S3.** Sources of variation in the average hatch weight of chicks.

Polyandry model

A

Fixed effects

First nest lay date

First breeding age

Age since first breeding

Standardized effect size (logit β) ± 95% CI

B

Semi-partial R^2

Marginal R^2

First nest lay date

First breeding age

Age since first breeding

Variance explained (R^2) ± 95% CI

C

Intra-class correlation

Individual

Year

Adjusted repeatability (r) ± 95% CI

Figure S4. Sources of variation in a female's probability of being polyandrous each year.

Figure S5. Sources of variation in a female's probability of re-nesting after a failed attempt
each year.

Figure S6. Homoscedastic spread of residuals from the A) 'egg volume' and B) 'lay date'
models.

Figure S7. Relationship between a female's structural size (i.e., her tarsus length) and A) the

lay date of her first nest of the season, and B) the volume of her eggs.

Figure S8. Visualization of the effect of uncertainty in the age estimate provided by BaSTA
on age-dependent trends in egg volume (top row) and lay date (bottom row) dynamics. Black
trends show the model predictions for the 1000 bootstraps of the BaSTA age estimate
posteriors (see *Methods*). Panels A and D show the within-individual trends of the 'age-
deviation' score – as expected, these measures are not impacted by uncertainty in the BaSTA
age estimate because they are centred for each individual (i.e., the absolute age is irrelevant).
Panels B and E show the between-individual trend of the 'age at first breeding' (i.e., selective
appearance), and panels C and F show the between-individual trend of the 'age at last
breeding' (i.e., selective disappearance). Yellow trends and grey ribbons visualize the 95%
CI of the model predictions when using the median age estimate provided by BaSTA (i.e., the
effect sizes of the 'egg volume' and 'lay date' models shown in Fig. 3).

Supplementary File 1. Commented RMarkdown vignette detailing the analytical steps

needed to reproduce all our models and results. Click here to view on your internet browser:

https://raw.githack.com/leberhartphillips/snowy_plover_eggs/main/Rmd/Supplementary_File

_1/Supplementary_File_1.html

Appendix S1.

To explain how mortality risk is associated with age, we compared four commonly used

survival functions: exponential, Gompertz, logistic, and Weibull models. The exponential

function keeps survival constant across age (Cox & Oakes, 1984), whereas the other

functions allow for age-dependent variation in survival: the Gompertz model is an

exponential function in which age-specific mortality is scaled by baseline mortality

(Gompertz, 1812; Pletcher, 1999) and the Weibull model is a power function which assumes

that baseline and age-specific mortality rates are independent (Pinder *et al.*, 1978). In

addition to the stand-alone versions of these functions, we considered two alternative forms

of the Gompertz, logistic, and Weibull functions – "Makeham" and "bathtub" shapes. The

Makeham shape constrains the survival function to converge to a constant, rather than zero,

as age increases (Pletcher, 1999), whereas the bathtub shape enables concavity in the

function, such that mortality could decrease at early ages, but increase later in life (Siler,

1979).

We used four parallel simulations to run the Markov chain Monte Carlo (MCMC)

optimization procedure in BaSTA with 800,000 iterations, a 100,000 burn-in period, and a

thinning interval of 2000 to minimize serial autocorrelation in the chain (see Fig. S2 for

simulation diagnostics). We ranked the survival models according to their deviance

information criterion (*DIC*) and determined that the logistic model with a bathtub shape fitted

our data best (i.e., lowest *DIC*, (Spiegelhalter *et al.*, 2002); Table S1). This logistic model

revealed that female mortality rate increased until age 5 years, after which it became constant

(Fig. 1b). Using the top model, we extracted the point estimate and 95% credible interval

birth year for each individual in the capture-mark-recapture sample.

Appendix S2.

In theory, the selective benefits of larger eggs is that the subsequent hatchlings will be larger

and have higher survival owing to more intrinsic reserves provided by the mother (Blomqvist

et al., 1997). To link egg size variation to potential fitness consequences of subsequent

offspring we evaluated the predicted positive relationship between egg volume and chick

weight using the egg dataset described above but reduced observations to the nest level and

filtered to only include nests that had chicks measured within one day of hatching, resulting

in 456 nests from 276 females. As it was unclear which chick came from which egg, each

datum represented the nest-level average of chick weights and egg volumes. We included

random intercepts for mother identity and year, and assumed a Gaussian error distribution of

egg volume.