1 Title: Contemporary and historical selection in Tasmanian devils (*Sarcophilus harrisii*) support
2 novel, polygenic response to transmissible cancer
3

Authors: Amanda R. Stahlke[1], Brendan Epstein[2], Soraia Barbosa[1,3], Mark J. Margres[2,4] Austin
Patton[2,5], Sarah A. Hendricks[1], Anne Veillet[1], Alexandra K Fraik[2], Barbara Schönfeld[6], Hamish I.
McCallum[7], Rodrigo Hamede[6], Menna E. Jones[6], Andrew Storfer[2], Paul A. Hohenlohe[1]
Affiliations:
1. Institute for Bioinformatics and Evolutionary Studies (IBEST), University of Idaho,
Moscow, ID 83844, USA
2. School of Biological Sciences, Washington State University, Pullman, WA 99164, USA
3. Department of Fish and Wildlife Sciences, University of Idaho, Moscow ID 83844, USA
4. Department of Integrative Biology, University of South Florida, Tampa, FL 33620, USA
5. Department of Integrative Biology and Museum of Vertebrate Zoology, University of
California, Berkeley, CA 94720, USA
6. School of Natural Sciences, University of Tasmania, Hobart, TAS 7005, Australia
7. Environmental Futures Research Institute, Griffith University, Nathan, QLD 4111,
Australia

18
Abstract

Tasmanian devils (*Sarcophilus harrisii)* are evolving in response to a unique transmissible
cancer, devil facial tumour disease (DFTD), first described in 1996. Persistence of wild
populations and the recent emergence of a second independently evolved transmissible cancer
suggest that transmissible cancers may be a recurrent feature in devils. Here we compared
signatures of selection across temporal scales to determine whether genes or gene pathways
under contemporary selection (6-8 generations) have also been subject to historical selection
(65-85 million years), and test for recurrent selection in devils. First, we used a targeted
sequencing approach, RAD-capture, to identify genomic regions subject to rapid evolution in
approximately 2,500 devils in six populations as DFTD spread across the species range. We
documented genome-wide contemporary evolution, including 186 candidate genes related to
cell cycling and immune response. Then we used a molecular evolution approach to identify
historical positive selection in devils compared to other marsupials and found evidence of
selection in 1,773 genes. However, we found limited overlap across time scales, with historical
selection detected in only 16 contemporary candidate genes, and no overlap in enriched
functional gene sets. Our results are consistent with a novel, multi-locus evolutionary response
of devils to DFTD. Our results can inform management actions to conserve adaptive potential of
devils by identifying high priority targets for genetic monitoring and maintenance of functional
diversity in managed populations.

38
39 Keywords: Rapid evolution, molecular evolution, transmissible cancer, wildlife disease,
40 conservation genomics, RAD capture
41

Introduction

Species are subject to selection by pathogens throughout their evolutionary history, shaping lineage diversification and leading to complex cellular and molecular defensive mechanisms (1). Still, emerging infectious diseases (EIDs) can cause mass mortality and, given sufficient reproduction and genetic variation, initiate rapid adaptive evolution in a naïve host population (2). Although the prevalence and severity of EIDs in wildlife populations is now well-recognized (3-6), we are just beginning to understand the evolutionary impacts of disease in wildlife. We have a relatively short recorded history of infectious disease in wildlife, and therefore a limited ability to predict outcomes or intervene when warranted (7, 8).

High-throughput DNA-sequencing techniques and high-quality annotated reference genomes have revolutionized our ability to monitor and identify mechanisms of evolutionary responses to pathogens (8-10). Inter-specific comparisons of non-synonymous and synonymous variation (dN/dS) within protein-coding regions have long been used to identify positive selection at immune-related loci (11-13). At the population level, rapid evolution in response to disease can be detected by tracking changes in allele frequency before, during, and after the outbreak of disease (14, 15). Intra-specific comparisons across populations can reveal to what extent the evolutionary response to disease is constrained by limited genetic mechanisms or variation for adaptation (16). Reduced representation techniques such as restriction-site associated DNA-sequencing (RADseq) (17) have made the acquisition of genome-wide, time-series genetic data more accessible in non-model systems (18). By integrating these resources and tests of selection at differing temporal scales, we can assess whether species that show rapid evolution in response to contemporary pathogens also show evidence of historical selection to similar pathogens.

A striking example of an EID acting as an extreme selective force in wildlife is devil facial tumour disease (DFTD), a transmissible cancer first described in 1996 afflicting wild Tasmanian devils (*Sarcophilus harrisii*) (19). Tasmanian devils are the largest extant carnivorous marsupial, with contemporary wild populations restricted to the Australian island of Tasmania. As a transmissible cancer, DFTD tumour cells are transmitted between hosts, behaving as a pathogen (20). Transmission typically occurs as devils bite each other during the mating season after devils have reached sexual maturity (21, 22). With few notable exceptions documenting regression (23), DFTD tumors escape recognition, become malignant, and can kill their hosts within six months (24). Starting from a single Schwann cell origin (25), DFTD has now swept across nearly the entire species range (Figure 1A). Devil populations have declined species-wide ~80% (26) with local declines in excess of 90% (27). Nonetheless, population genomic studies have shown that devils are rapidly evolving in response to DFTD (2, 28, 29), and DFTD has been spontaneously cleared (i.e., regressed) in some individuals (23). Long-term field studies and simulation modelling have predicted that cyclical co-existence or DFTD extirpation are more likely scenarios than devil extinction (30). This is particularly alarming because devils have notoriously low genome-wide diversity, attributed to climate- and anthropogenic-induced bottlenecks (31-33). Depleted genetic diversity at immune-related loci has likely further contributed to DFTD vulnerability (34).

Despite transmissible cancers being exceedingly rare across animals, a second independent
transmissible cancer in devils, DFT2, was described in 2014 (35, 36). Comparative and functional
analyses of DFTD and DFT2 showed similar drivers of cancerous mutations and tissue type of
origin (37). Low genetic diversity, chromosomal fragility (38), a reportedly high incidence of
non-transmissible neoplasms (39), and injury-prone biting behaviour (40) may contribute to a
predisposition to transmissible cancers in devils (41). These findings suggest that transmissible
cancers may be a recurring selective force in the Tasmanian devil lineage. If so, this leads to the
hypothesis that the genes and genetic pathways associated with the ongoing evolutionary
response to DFTD may have experienced recurrent historical selection in the devil lineage from
previous transmissible cancers.
96
Because of the threat of DFTD and DFT2 to devil populations, there are ongoing conservation
efforts, including the establishment of a captive devil insurance meta-population. The insurance
population is managed to maintain genome-wide genetic diversity and serve as a source for re-
introductions in an effort to increase genetic diversity and size of wild populations (42). To
inform conservation efforts, it is important to understand what types of genetic variation in
natural populations may allow for evolutionary rescue from disease and maintain adaptive
potential for future threats (43). Given evidence for rapid evolution in response to DFTD,
monitoring of genetic variation at candidate adaptive loci could help evaluate adaptive
potential of wild populations (43, 44). In heavily managed (e.g. captive) populations, loci
associated with an adaptive response to disease could be included in genotyping panels for
maintaining genetic diversity (45).
108
Here we identify targets of selection and signatures of adaptation at both contemporary (6-8
generations) and historical (65-85 million years) scales in Tasmanian devils. First we test for
evidence of contemporary genomic response to selection by genotyping thousands of
individuals sampled at several time points across six populations, using RAD-capture (46) to
target nearly 16,000 loci (47). Next we identify signatures of historical selection in the devil
lineage by comparing across marsupial species with annotated genomic sequence data. Then,
we test for evidence of recurrent selection by examining shared contemporary and historical
signatures of selection, in terms of either specific loci, genes or functional genetic pathways.
117
If transmissible cancer is a novel selective force acting on Tasmanian devils, we expect that
genes under contemporary selection by DFTD will be different from those with a signature of
historical positive selection. Alternatively, if transmissible cancer is a recurrent selective force in
the devil lineage that targets the same set of genes repeatedly, we may expect a conserved
response among populations and an overrepresentation of the same genes or pathways under
both contemporary and historical selection. However, if there are multiple genetic pathways
that could be involved in a response to recurrent transmissible cancers, we may expect a
polygenic response across contemporary populations and little overlap between contemporary
and historical timescales. These alternatives can inform conservation efforts to manage genetic
diversity for resilience in natural devil populations, and any genes or functional pathways that
show both contemporary and historical selection may be relevant to cancer resistance more
broadly.

130
Materials and Methods
Contemporary Selection
We used the RAD-capture method (combining RADseq and sequence capture) (46) to conduct
targeted genotyping of single-nucleotide polymorphisms (SNPs) across 2,562 unique individuals
from multiple Tasmanian devil populations, sampled both before and after DFTD appeared in
each population (Figure 1A, Table 1, Supplemental Table S1) (29, 48). We constructed RAD-
capture libraries following Ali et al (2016), using the restriction enzyme *PstI* and a capture array
targeting 15,898 RAD loci selected for membership in one of three functional categories: 1)
those showing signatures of DFTD-related selection from previous work (2), 2) loci close to
genes with known cancer or immune function, and 3) loci widely distributed across the genome
(See 29, 47 for more details on the devopment of this array.). See Supplemental Materials S1
for multiplexing, read processing, and SNP genotyping details.
143
To account for the expected high rates of genetic drift within populations, we used a composite
statistic to compare signatures of selection across populations. We identified candidate SNPs as
the top 1% of a de-correlated composite of multiple signals score (DCMS) (49), which combined
the results of three analyses: change in allele frequency in each population after DFTD (Δaf),
and two methods that estimate strength of selection from allele frequencies at multiple time
points in multiple populations, the method of Mathieson & McVean (14) (*mm*), which allows
the estimated selection coefficient to vary over space; and *spatpg* (15), which allows the
selection coefficient to vary over time and space. Individuals were assigned to generational-
cohorts based on their estimated years of birth (Supplemental Table S1). We estimated Δaf for
five locations at which we had sampling both before and after DFTD was prevalent, according to
DNA collection date and estimated date of birth, combining multiple cohorts when applicable
(Table 1; Supplemental Table S1). Both time-series methods (*mm* and *spatpg*) incorporate
estimates of effective population size, which ranged from 26-37 according to a two-sample
temporal method (50, 51) (Supplemental Table S3). DCMS reduces the signal-to-noise ratio by
combining p-values from different tests at each SNP while accounting for genome-wide
correlation among statistics. We included SNPs with results from at least eleven of the twelve
individual tests (Δaf for five populations, *mm* for all six populations, and *spatpg*) and weighted
based on the statistics with results at that SNP. To characterize the role of standing genetic
variation in rapid evolution (52), we visualized the initial allele frequencies of each population
for each analysis of contemporary evolution (Supplemental Figures S2, S9). We evaluated
repeatability among populations by comparing population-specific p-values of Δaf and *mm* with
the R package dgconstraint for a similarity index called the C-score, where 0 indicates no
similarity between populations (16). See Supplemental Materials S1 for details of each analysis.
167
Historical Selection
We combined existing genomic resources for the South American grey-tailed opossum
(*Monodelphis domestica*) (53) and tammar wallaby (*Notamacropus eugenii*) (54) from the
Ensembl database (55) and the recently published transcriptome assembly of the koala
(*Phascolarctos cinereus*) (56) to identify genome-wide signatures of positive selection in devils,
relative to these other species using the branch-site test of PAML (Phylogenetic Analysis by

Maximum Likelihood) (57, 58). We compiled alignments of orthologous genes and reduced the
marsupial time-calibrated phylogeny of Mitchell et al. (59) to those species for which annotated
full genomes are available (Figure 1B). The branch-site test compares likelihood scores for two
models which estimate dN/dS among site classes of a multi-sequence alignment, allowing
dN/dS to exceed 1 (positive selection) in a proportion of sites along a single branch in the
alternative model. We reduced the potential for false positives by filtering any putative
orthologs with extreme sequence divergence (S > 2), measured as the sum of synonymous
mutations per gene (S), and ensuring alignments of nucleotides were longer than 100 bp (60,
61). We identified historical candidates with the likelihood-ratio test, comparing the likelihoods
of the alternative and neutral models with one degree of freedom and an $\alpha = 0.05$. Historical
candidates were those with estimates of dN/dS > 1 along the devil branch and FDR > 0.05 after
correcting for multiple testing (62). See Supplemental Materials S1 for details regarding
orthology identification and PAML implementation.
187
Recurrent Selection
We refer to genes under both contemporary and historical selection as candidates for recurrent
selection. To test whether genes under contemporary selection differed from genes under
historical selection, we first tested for significant overlap with Fisher's one-tailed test. To test
for differences in the strength of selection, we compared the distributions of dN/dS and the
proportion of sites per gene found under positive selection among candidates for recurrent
selection to all other historical candidates from the genome-wide background using
nonparametric tests of equality, the Kolmogorov-Smirnov test (63), which is more sensitive to
the centre of the distributions, and the Anderson-Darling test (64), which is more sensitive to
extreme values of the distribution and often has more power. To identify and compare key
mechanisms of adaptation among candidate genes from each set, we used gene ontology (GO)
term enrichment analysis using the SNP2GO package (65), the PANTHER web-interface (66),
and in gene sets of the molecular signatures database (MsigDB), using the subset of genes
tested for each test as the respective background set (67). We capitalized on the wealth of
ongoing research in devils and DFTD by comparing our contemporary and historical candidates
to those previously identified using different datasets and analytical approaches (2, 28, 29, 47,
68). See Supplemental Materials S1 for details of these comparisons.
205
Results
Genomic Data
To test for contemporary selection, we sampled a total of 2,562 individuals across six localities
of Tasmania before and after DFTD prevalence (Table 1; Supplemental Table S1; Figure 1A),
with a RAD-capture array (47). After filtering, we mapped a total of 517.7 million reads against
targeted loci. The mean final coverage of targeted loci was 14.8x, with 76.6% of all samples
having coverage of at least 5x (Supplemental Figure S1). After filtering, we retained 14,585 −
22,878 SNPs for downstream analysis, depending on the sampled time point and population.
214
Evidence for contemporary selection
Among each elementary test for selection signatures, 161 − 232 SNPs (depending on
population) were in the top 1% of allele shifts following disease (Δaf), 209 − 217 were in the top

1% of *mm* scores, and 213 were in the top 1% of *spatpg* scores (Supplemental Table S4, Figures
S7-8). Across populations and elementary tests for contemporary selection (Δaf, *mm*, *spatpg*),
p-values were not correlated (Pearson's r < 0.155 for all tests; Supplemental Figure S10). The
computed repeatability indexes for population-specific responses Δaf and *mm* were $C_{\Delta af}$= 4.86
(*p* 1e-04) and C_{mm}= 3.72 (*p* = 1e-04), which implies a low, but significant level of repeatability
(16). In the top 1% of DCMS scores ($\geq$ 1.167), we identified 144 candidate SNPs for
contemporary selection by DFTD; of these, 79 had annotated genes (186 total) within 100 kb
(Figure 2; Supplemental Table S5). The initial frequencies for candidate SNPs were not skewed
toward intermediate frequencies (Supplemental Figure S9).
227
Comparing our contemporary candidates and those previously identified in devils with selection
and genome-wide association analyses (28, 29, 47, 68), we found many overlapping genes
(discussed below). Notably, we found significant enrichment of candidates previously
associated with DFTD-related phenotypes in females (14 genes, *p* = 4.2e-08, Odds ratio=7.3)
(47). Gene ontology enrichment analysis found middle ear morphogenesis (GO:0042474)
significantly enriched among contemporary candidate SNPs (FDR < 0.05). Five candidate SNPs
were within the 100 kb window of two genes associated with this term: EYA1 and PRKRA. Both
EYA1 and PRKRA are involved in cell proliferation and migration and implicated in tumour
suppression and angiogenesis (69-71).
237
Evidence for historical selection
Of the 18,788 genes annotated in the devil reference genome, 6,193 had 1-to-1 orthologs in at
least three of the four marsupial genomes and an appropriate sequence divergence (S<2). Using
the branch-site test for positive selection in PAML, we found a total of 1,773 genes to be
candidates for historical positive selection (Supplementary Table S6). Estimates of dN/dS
spanned the full range of possible values, from 1.05 − 999 and proportion of sites with
substitutions per gene ranged from 0.01 − 0.78 (Fig. 3). The majority of genes were classified as
having a molecular function of binding (GO:0005488) or catalytic activity (GO:0003824); a
plurality involved in cellular processes (GO:0009987) or biological regulation (GO:0065007); and
a plurality as participating in the Wnt signalling pathway (P00057). None of these pathway
classifications were significantly enriched.
249
Recurrent selection
Of the 186 contemporary candidate genes, 68 had 1-to-1 orthologs among the four marsupials
and were tested for historical selection. Sixteen genes showed evidence of historical selection
and are thus candidates for recurrent selection (dN/dS > 1, FDR < 0.05; Supplemental Table S6).
Contemporary candidates were not enriched for historical selection according to Fisher's test
(Odds ratio = 0.0, *p* = 1). Among the 16 recurrent candidates, dN/dS estimates spanned 15.7 −
999 and proportion of sites per gene 0.01 − 0.25 (Fig. 3). According to the Anderson-Darling and
Kolmogorov-Smirnov tests of equality, neither distributions of dN/dS estimates (Fig. 3a; A.D. *p* =
0.86; K.S. *p* = 0.58), nor proportion of sites (Fig. 3b; A.D. *p* = 0.49; K.S. *p* = 0.48) differed
between candidates for recurrent selection (in black) and historical candidates (in red).
260

261 After correcting gene set enrichment for multiple testing (FDR < 0.05) and requiring at least 10
262 genes in the background set, we did not find functional enrichment of any MSigDB gene sets
263 among recurrent candidates or shared between both contemporary and historical sets.
264 Importantly, the permutation test of shared gene sets found *fewer* shared between historical
265 and contemporary selection than expected by chance ($p < 0.001$, Supplemental Figure S11).
266
Discussion
Contemporary Responses to DFTD
Using a targeted set of nearly 16,000 loci, we detected widespread evidence of a response to
selection by DFTD across the Tasmanian devil genome. Our results extend previous work that
has shown genomic evidence of a response to DFTD in wild populations (2, 28, 29, 72). Here we
greatly increased the sample size of individuals and genetically independent populations for
greater power, resulting in strong evidence of a response to selection widely distributed across
the genome. We found greater similarity across populations within analytical approaches than
among methods within populations and relatively low, but significant repeatability across
populations. This result is consistent with rapid, polygenic evolution facilitated by selection for
standing variation within populations that was present prior to disease arrival. This timescale (3
- 8 generations) would likely be too short for new mutations or migration to play a substantial
role in DFTD response, and genetic variation is shared across the species range, despite
geographic population structure (29, 73).

In line with previous population genomic studies (2, 28, 29, 47, 68), our analysis of
contemporary evolution detected a putatively adaptive response related to the immune
system, cell adhesion, and cell-cycle regulation (Supplemental Table S5). Our GO enrichment
result for middle ear development (GO:0042474) among contemporary candidates may
highlight selection for interactions with the peripheral nervous system and cell proliferation.
Genes annotated with nervous system associations may indicate selection for behavioural
changes (28), or highlight importance and vulnerability of peripheral nerve repair by Schwann
cells in devils, given the prevalence of biting and Schwann cell origin of DFT (41). Significant
overlap for genes associated with devil infection status (case-control), age, and survival (47)
among our contemporary candidates is a strong indicator that these contemporary candidates
likely confer relevant phenotypic change. We also confirmed five (CRBN,
ENSSHAG00000007088, THY1, USP2, C1QTNF5) of seven candidates identified previously (2) in
a genome scan for loci under selection from DFTD in three of the same populations (Freycinet,
Narawntapu, and West Pencil Pine). In contrast, we identified those five and only two more
(TRNT1 and FSHB) of 148 candidates from a re-analysis of that same dataset which studied
population-specific responses (28).
297
Among genes that have been associated with host variation responsible for tumour regression
on devils (68, 74), we found only JAKMIP3, a Janus kinase and microtubule binding protein (74),
in our list of contemporary candidates. However, we found devil regression candidates TL11,
NGFR, and PAX3, which encodes a transcription factor associated with angiogenesis (75); as
well as GAD2, MYO3A, and unannotated ENSSHAG00000009195 (74), among population-
specific candidates for allele frequency change (Δaf), possibly reflecting differences in test

sensitivities. Overall, the paucity of candidates shared between our contemporary analysis and
regression studies suggests that regression may not be the dominant form of phenotypic
response to DFTD; to date tumour regression has only been detected in a few populations (74)
not represented in our study.
308
Historical selection in the devil lineage
With our genome-wide molecular evolution approach (57), we found widespread historical
positive selection across the devil genome in about 28% of all 6,249 orthologs tested
(Supplemental Table S6). The branch-site test is known to be less conservative than related
models, particularly when divergence among species is large (76), but the rates of historical
selection we found in devils are similar to those described in other taxa; e.g. 23% of genome-
wide orthologs among 39 avian species using a similar approach (13).
316
We did not find preferential positive selection for immunity-related genes, as has been shown
in primates (1), eutherian mammals more generally (77), and birds (13). Instead, we found the
highest proportion of pathways under historical selection to be functionally classified within the
Wnt pathway, a signalling cascade regulating cell adhesion and implicated in carcinogenesis
(78). As genomic resources grow and improve in marsupials (10), interspecific analyses for
positive selection at finer scales may reveal more recent and specific selection targets in
Tasmanian devils. Our ability to detect historical selection due to transmissible cancer in devils
could be improved by genome assembly efforts among more closely-related Dasyuridae, as well
as complementary annotation.
326
Comparing Contemporary and Historical Timescales
Remarkably few transmissible cancers have been discovered in nature (79, 80), and yet two of
those independent clonally-transmitted cancers have been discovered in Tasmanian devils in
less than 20 years. This and the observed rapid evolutionary response to disease suggest that
transmissible cancers may be a recurrent event in devils. We found no significant overlap of
historical and contemporary selection at either individual genes or functional gene sets. This
does not rule out the possibility of prior transmissible cancers in devils; but it suggests that if
transmissible cancers have been a recurrent feature of devil evolution prior to DFTD, they did
not generally impose selection on the same set of genes or genetic pathways that show a
contemporary response to DFTD. Nonetheless, the 16 candidate genes showing both historical
and contemporary evidence for selection (Supplemental Table S7) raise interesting targets for
understanding adaptively important variation in devils.
339
The 16 candidate genes for recurrent selection (Supplemental Table S7) are generally related to
three main themes: transcription regulation, the nervous system, and the centrosome. Four of
these candidates for recurrent selection were previously associated with disease-related
phenotypes (47). We additionally found 82 historical candidates previously identified in the top
1% of SNPs associated with disease-related phenotypes with three represented in the top 0.1%
associated with large-effect sizes for female case-control and survival (47). This overlap lends
support to the hypothesis of recurrent selection by transmissible cancers, but was not
significant (p = 1, odds ratio= 0). Both our contemporary selection analysis and the genome-

wide association study (GWAS) approach used by Margres and colleagues (47) are statistically
limited by small populations, sample size, and the time scale over which DFTD-related selection
has occurred. By considering the complement of these results together, the overlapping
historical, GWAS, and contemporary candidates may still be targets of recurrent selection along
similar functional axes, potentially including transmissible cancer.
353
The low prevalence of candidates for recurrent selection and lack of shared functional gene set
enrichment between both contemporary and historical signatures of selection suggest a novel
response to DFTD compared to historic selection in the devil lineage. However, there are
alternative hypotheses. For example, there could be redundancy in genetic mechanisms
underlying resistance to transmissible cancers, potentially as a result of repeated selection for
resistance, allowing selection to act across many loci (81). That is, the low genetic diversity
observed in devils could be the result of widespread historical purifying selection resulting from
transmissible cancers or other diseases (82), or historical bottlenecks due to climate change and
habitat loss (31-33), that prevent a response to selection under DFTD at loci that are still
associated with disease phenotypes.
364
The widespread contemporary evolution we found in devils reflects the recent prediction (83)
that response to an emergent disease is most likely controlled by many genes conferring
quantitative resistance (84), for example by reducing the within-host growth rate of tumors.
DFTD is predicted to become less virulent in the short-term (30, 85). If DFTD persists long-term
in the devil population with ongoing coevolution, it may lead to diversifying selection for
specific, qualitative host resistance mechanisms (83). Indeed, phylodynamic analysis of DFTD as
it spread across Tasmania supports the hypothesis that devils may be mounting a response;
transmission rates have decayed such that DFTD appears to be shifting from emergence to
endemism (85). Although host-genomic variation was not jointly considered in that study, the
combined evidence of multiple studies demonstrating rapid evolutionary response of devils to
DFTD, including this one, support these interpretations.
376
Conservation Implications
Calls have been made to consider the historical context of adaptation when proposing
conservation management solutions based on genomic results (86). Our analysis of historical
selection largely supports the hypothesis that DFTD is a newly emerging and novel selective
force, distinctly shaping today's remaining wild devils. The targets of novel selection that we
identified (Fig. 2, Supp Table S4) and their functional roles should be considered for
prioritization of monitoring and conservation in light of DFTD. At the same time, the wide
distribution of contemporary candidates across the genome also highlights the importance of
standing genetic variation to continue to respond to unique selective forces, including local
environmental factors (29). Genomic monitoring could be useful for maintaining both
functional diversity at candidate loci and genome-wide variation in captive populations (45, 87,
88) and in the wild. Multiple genomic tools are available for targeted monitoring of large sets of
loci (e.g. 89, 90) and could be used to track adaptive evolution and potential in the form of
genetic diversity (43). However, before management decisions are made for specific genes,

further work would need to identify favoured alleles and fitness effects for the genes we
identified (Supplemental Table S5) .

DFTD has yet to reach devils in the far west (Fig. 1a) and continues to circulate throughout the
island. To maintain long-term adaptive capacity in the face of similar recurrent selective forces
including DFT2 and potential future transmissible cancers, our results warrant (1) the
monitoring of genetic variation in broad functional groups and (2) management strategies to
maintain genetic diversity across those broad groups. Although these populations were not
subject to DFT2 at the time of writing, an important and interesting future direction should
examine the evolutionary response to DFT2 and could compare loci under selection by the two
independent transmissible cancers. This study could provide a list of candidate loci for
development of a genotyping panel for either purpose, with flexibility to target many or fewer
loci. At the same time, given urgent and unpredictable present-day threats including not just
emerging diseases but environmental change and population fragmentation, it is important
that monitoring and population management also focus on maintaining genetic variation across
the genome.

406
Conclusion
Our results suggest that the contemporary evolutionary response to DFTD is mostly novel
compared to the genome-wide signature of historical selection. Comparing the degree of
overlap and distributions among contemporary and historical candidates did not support
recurrent selection on a common set of genes in response to transmissible cancer. Our work
contributes to mounting evidence of possible mechanisms by which devil populations are
persisting and rapidly evolving in the face of DFTD despite overall low genetic diversity and
population bottlenecks (2, 23, 47, 72, 91). Broadly, this type of approach can be applied to
analyses of novel threats in wildlife populations in the current era of anthropogenic global
change to guide monitoring and management actions focused on genetic adaptive potential.

417
418

Data and Script Accessibility
Demultiplexed sequence data has been deposited at NCBI under Bio-Project PRJNA306495
(http://www.ncbi.nlm.nih.gov/bioproject/?term=PRJNA306495) and BioProject PRJNA634071
(http://www.ncbi.nlm.nih.gov/bioproject/?term=PRJNA634071). Code and tabular results are
available at https://github.com/Astahlke/contemporary_historical_sel_devils and on Dryad.

Acknowledgements
The authors would like to acknowledge Michael R. Miller, Sean M. O'Rourke, Cody G. Wiench,
for their contributions to the RAD capture data; and Rebecca Johnson and Denis O'Meally for
providing the koala transcriptome data. This work was funded by NSF grant DEB-1316549 and
NIH grant R01-GM126563 to A.S., P.A.H., M.E.J., and H.I.M. as part of the joint NSF-NIH-USDA
Ecology and Evolution of Infectious Diseases program. Genomics and bioinformatics were
supported by an Institutional Development Award (IDeA) from the National Institute of General
Medical Sciences of the NIH under grant number P30 GM103324. This work was supported
by the Bioinformatics and Computational Biology Program at the University of Idaho in

partnership with IBEST (the Institute for Bioinformatics and Evolutionary Studies). Sample
collection was funded under Australian Research Council grants DE170101116, DP110102656,
LP0989613 and LP0561120 to R.H., M.E.J., H.I.M., and A.S., an ARC Future Fellowship
FT100100031 to M.E.J., and multiple awards from the University of Tasmania Foundation Eric
Guiler Research Grant.

Authors' contributions

ARS, HIM, MEJ, AS, and PAH conceived and designed the study. RH and MEJ conducted
fieldwork and sampling. ARS, BE, SB, SAH, AV, BS conducted genomics labwork. ARS, BE, SB, AP,
SAH, AKF, and PAH conducted bioinformatic analyses. ARS wrote the manuscript with
contributions from all authors.

Tables and Figures

Table 1. Number of adults sampled before and after the year of first detection of DFTD at each

site. See Supplemental Table S1 for sample size for each year at each locality.

Location	Year of First Detection	Samples Before	Samples After	Total
wukalina/Mt. William	1996	0	155	155
Freycinet	2001	300	382	682
Forestier	2004	131	552	683
Fentonbury	2005	99	169	268
West Pencil Pine	2006	52	348	400
Narawntapu	2007	224	150	374
Total		806	1756	2562

Figure 1. A) Map of the six contemporary sampling locations relative to disease prevalance over time (red lines) with the year of first detection labled at each site. B) Reduced, unrooted time-calibrated phylogeny (59) of marsupials used to estimate genome-wide historical selection on the devil lineage with estimated divergence times (Ma) indicted along edges. Devil cartoon by David Hamilton. Wallaby, koala, and opossum digital images retrieved from http://www.shutterstock.com/amplicon. From top to bottom: The tammar wallaby (*Notamacropus eugenii*), koala (*Phascolarctos cinereus*), Tasmanian devil (*Sarcophilus harrisii*), and South American grey-tailed opossum (*Monodelphis domestica*).

Figure 2. Results of each elementary test of contemporary selection across populations and the
composite scores for final candidate (filled points) and noncandidate (opaque grey points)
SNPs, ordered by chromosome and colored by population when applicable. From top to
bottom: Change in allele frequency (*Δaf*), Mathieson and McVean (*mm*) (14), *spatpg* (15), de-
correlated composite of multiple signals (*DCMS*) (49).

Figure 3. A) Estimates of dN/dS and B) the proportion of sites under positive selection (57) for
historical candidates across the genome-wide background (red; N=1,982) and candidates for
recurrent selection (genes with significant results for both historical and contemporary
selection; black). Each point represents the respective result for a single gene. C) Distribution of
-log(FDR) for historical selection across all 6,193 genes tested (gray squares, non-significant at
both scales; black squares, significant contemporary and non-significant historical; red circles,
non-significant contemporary and significant historical; black circles, significant at both scales).
472
473

475
476
477   References
478
1.   Barreiro LB, Quintana-Murci L. From evolutionary genetics to human immunology: how
selection shapes host defence genes. Nat Rev Genet. 2010;11(1):17-30.
2.   Epstein B, Jones M, Hamede R, Hendricks S, McCallum H, Murchison EP, et al. Rapid
evolutionary response to a transmissible cancer in Tasmanian devils. Nature communications.
2016;7:12684.
3.   De Castro F, Bolker B. Mechanisms of disease□induced extinction. Ecology Letters.
2005;8(1):117-26.
4.   Smith KF, Sax DF, Lafferty KD. Evidence for the role of infectious disease in species
extinction and endangerment. Conservation biology. 2006;20(5):1349-57.
5.   Jones KE, Patel NG, Levy MA, Storeygard A, Balk D, Gittleman JL, et al. Global trends
in emerging infectious diseases. Nature. 2008;451(7181):990-3.
6.   Lafferty KD. The ecology of climate change and infectious diseases. Ecology.
2009;90(4):888-900.
7.   McCallum H, Dobson A. Detecting disease and parasite threats to endangered species
and ecosystems. Trends in ecology & evolution. 1995;10(5):190-4.
8.   Storfer A, Kozakiewicz CP, Beer MA, Savage AE. Applications of Population Genomics
for Understanding and Mitigating Wildlife Disease. 2020.
9.   McKnight DT, Schwarzkopf L, Alford RA, Bower DS, Zenger KR. Effects of emerging
infectious diseases on host population genetics: a review. Conservation genetics.
2017;18(6):1235-45.
10.   Brandies P, Peel E, Hogg CJ, Belov K. The Value of Reference Genomes in the
Conservation of Threatened Species. Genes. 2019;10(11):846.
11.   Li W-H. Unbiased estimation of the rates of synonymous and nonsynonymous
substitution. Journal of molecular evolution. 1993;36(1):96-9.
12.   Yang Z. Likelihood ratio tests for detecting positive selection and application to primate
lysozyme evolution. Molecular biology and evolution. 1998;15(5):568-73.
13.   Shultz AJ, Sackton TB. Immune genes are hotspots of shared positive selection across
birds and mammals. Elife. 2019;8:e41815.
14.   Mathieson I, McVean G. Estimating selection coefficients in spatially structured
populations from time series data of allele frequencies. Genetics. 2013;193(3):973-84.
15.   Gompert Z. Bayesian inference of selection in a heterogeneous environment from genetic
time-series data. Mol Ecol. 2016;25(1):121-34.
16.   Yeaman S, Gerstein AC, Hodgins KA, Whitlock MC. Quantifying how constraints limit
the diversity of viable routes to adaptation. PLoS genetics. 2018;14(10):e1007717.
17.   Baird NA, Etter PD, Atwood TS, Currey MC, Shiver AL, Lewis ZA, et al. Rapid SNP
discovery and genetic mapping using sequenced RAD markers. PloS one. 2008;3(10):e3376.
18.   Andrews KR, Good JM, Miller MR, Luikart G, Hohenlohe PA. Harnessing the power of
RADseq for ecological and evolutionary genomics. Nature Reviews Genetics. 2016;17(2):81.
19.   Hawkins CE, Baars C, Hesterman H, Hocking GJ, Jones ME, Lazenby B, et al. Emerging
disease and population decline of an island endemic, the Tasmanian devil Sarcophilus harrisii.
Biological Conservation. 2006;131(2):307-24.

20. Pearse AM, Swift K. Allograft theory: transmission of devil facial-tumour disease.
Nature. 2006;439(7076):549.
21. Hamilton DG, Jones ME, Cameron EZ, McCallum H, Storfer A, Hohenlohe PA, et al.
Rate of intersexual interactions affects injury likelihood in Tasmanian devil contact networks.
Behavioral Ecology. 2019;30(4):1087-95.
22. Hamede RK, McCallum H, Jones M. Seasonal, demographic and density□related
patterns of contact between Tasmanian devils (Sarcophilus harrisii): Implications for
transmission of devil facial tumour disease. Austral Ecology. 2008;33(5):614-22.
23. Pye R, Hamede R, Siddle HV, Caldwell A, Knowles GW, Swift K, et al. Demonstration
of immune responses against devil facial tumour disease in wild Tasmanian devils. Biol Lett.
2016;12(10).
24. Siddle HV, Kreiss A, Tovar C, Yuen CK, Cheng Y, Belov K, et al. Reversible epigenetic
down-regulation of MHC molecules by devil facial tumour disease illustrates immune escape by
a contagious cancer. Proceedings of the National Academy of Sciences. 2013;110(13):5103-8.
25. Murchison EP, Tovar C, Hsu A, Bender HS, Kheradpour P, Rebbeck CA, et al. The
Tasmanian devil transcriptome reveals Schwann cell origins of a clonally transmissible cancer.
Science. 2010;327(5961):84-7.
26. Lazenby BT, Tobler MW, Brown WE, Hawkins CE, Hocking GJ, Hume F, et al. Density
trends and demographic signals uncover the long□term impact of transmissible cancer in
Tasmanian devils. Journal of Applied Ecology. 2018;55(3):1368-79.
27. McCallum H, Jones M, Hawkins C, Hamede R, Lachish S, Sinn DL, et al. Transmission
dynamics of Tasmanian devil facial tumor disease may lead to disease□induced extinction.
Ecology. 2009;90(12):3379-92.
28. Hubert J-N, Zerjal T, Hospital F. Cancer-and behavior-related genes are targeted by
selection in the Tasmanian devil (Sarcophilus harrisii). PloS one. 2018;13(8):e0201838.
29. Fraik AK, Margres MJ, Epstein B, Barbosa S, Jones M, Hendricks S, et al. Disease
swamps molecular signatures of genetic-environmental associations to abiotic factors in
Tasmanian devil (Sarcophilus harrisii) populations. Evolution. 2020;n/a(n/a).
30. Wells K, Hamede RK, Jones ME, Hohenlohe PA, Storfer A, McCallum HI. Individual
and temporal variation in pathogen load predicts long□term impacts of an emerging infectious
disease. Ecology. 2019;100(3):e02613.
31. Brüniche–Olsen A, Jones ME, Burridge CP, Murchison EP, Holland BR, Austin JJ.
Ancient DNA tracks the mainland extinction and island survival of the Tasmanian devil. Journal
of biogeography. 2018;45(5):963-76.
32. Brüniche-Olsen A, Jones ME, Austin JJ, Burridge CP, Holland BR. Extensive population
decline in the Tasmanian devil predates European settlement and devil facial tumour disease.
Biology letters. 2014;10(11):20140619.
33. Patton AH, Margres MJ, Stahlke AR, Hendricks S, Lewallen K, Hamede RK, et al.
Contemporary demographic reconstruction methods are robust to genome assembly quality: A
case study in Tasmanian Devils. Molecular biology and evolution. 2019;36(12):2906-21.
34. Siddle HV, Kreiss A, Eldridge MDB, Noonan E, Clarke CJ, Pyecroft S, et al.
Transmission of a fatal clonal tumor by biting occurs due to depleted MHC diversity in a
threatened carnivorous marsupial. Proceedings of the National Academy of Sciences.
2007;104(41):16221-6.

35.   Pye RJ, Pemberton D, Tovar C, Tubio JMC, Dun KA, Fox S, et al. A second
transmissible cancer in Tasmanian devils. Proceedings of the National Academy of Sciences.
2016;113(2):374-9.
36.   James S, Jennings G, Kwon YM, Stammnitz M, Fraik A, Storfer A, et al. Tracing the rise
of malignant cell lines: Distribution, epidemiology and evolutionary interactions of two
transmissible cancers in Tasmanian devils. Evolutionary applications. 2019;12(9):1772-80.
37.   Stammnitz MR, Coorens THH, Gori KC, Hayes D, Fu B, Wang J, et al. The origins and
vulnerabilities of two transmissible cancers in Tasmanian devils. Cancer Cell. 2018;33(4):607-
19.
38.   Deakin JE, Bender HS, Pearse A-M, Rens W, O'Brien PCM, Ferguson-Smith MA, et al.
Genomic restructuring in the Tasmanian devil facial tumour: chromosome painting and gene
mapping provide clues to evolution of a transmissible tumour. PLoS genetics.
2012;8(2):e1002483.
39.   Peck SJ, Michael SA, Knowles G, Davis A, Pemberton D. Causes of mortality and severe
morbidity requiring euthanasia in captive Tasmanian devils (Sarcophilus harrisii) in Tasmania.
Australian veterinary journal. 2019;97(4):89-92.
40.   Hamede RK, McCallum H, Jones M. Biting injuries and transmission of T asmanian
devil facial tumour disease. Journal of Animal Ecology. 2013;82(1):182-90.
41.   Patchett AL, Flies AS, Lyons AB, Woods GM. Curse of the devil: molecular insights into
the emergence of transmissible cancers in the Tasmanian devil (Sarcophilus harrisii). Cellular
and Molecular Life Sciences. 2020.
42.   Hogg CJ, Lee AV, Srb C, Hibbard C. Metapopulation management of an Endangered
species with limited genetic diversity in the presence of disease: the Tasmanian devil Sarcophilus
harrisii. International Zoo Yearbook. 2017;51(1):137-53.
43.   Hohenlohe PA, McCallum HI, Jones ME, Lawrance MF, Hamede RK, Storfer A.
Conserving adaptive potential: lessons from Tasmanian devils and their transmissible cancer.
Conservation Genetics. 2019;20(1):81-7.
44.   Wright BR, Farquharson KA, McLennan EA, Belov K, Hogg CJ, Grueber CE. A
demonstration of conservation genomics for threatened species management. Molecular Ecology
Resources. 2020.
45.   Grueber CE, Peel E, Wright B, Hogg CJ, Belov K. A Tasmanian devil breeding program
to support wild recovery. Reproduction, Fertility and Development. 2019;31(7):1296-304.
46.   Ali OA, O'Rourke SM, Amish SJ, Meek MH, Luikart G, Jeffres C, et al. RAD Capture
(Rapture): Flexible and efficient sequence-based genotyping. Genetics. 2016;202(2):389-400.
47.   Margres MJ, Jones ME, Epstein B, Kerlin DH, Comte S, Fox S, et al. Large☐effect loci
affect survival in Tasmanian devils (Sarcophilus harrisii) infected with a transmissible cancer.
Molecular ecology. 2018;27(21):4189-99.
48.   Hamede RK, Pearse AM, Swift K, Barmuta LA, Murchison EP, Jones ME. Transmissible
cancer in Tasmanian devils: localized lineage replacement and host population response. Proc
Biol Sci. 2015;282(1814).
49.   Ma Y, Ding X, Qanbari S, Weigend S, Zhang Q, Simianer H. Properties of different
selection signature statistics and a new strategy for combining them. Heredity. 2015;115(5):426.
50.   Jorde PE, Ryman N. Unbiased estimator for genetic drift and effective population size.
Genetics. 2007;177(2):927-35.

51.   Do C, Waples RS, Peel D, Macbeth GM, Tillett BJ, Ovenden JR. NeEstimator v2:
re□implementation of software for the estimation of contemporary effective population size (Ne)
from genetic data. Molecular ecology resources. 2014;14(1):209-14.
52.   Kelly JK, Hughes KA. Pervasive linked selection and intermediate-frequency alleles are
implicated in an evolve-and-resequencing experiment of Drosophila simulans. Genetics.
2019;211(3):943-61.
53.   Mikkelsen TS, Wakefield MJ, Aken B, Amemiya CT, Chang JL, Duke S, et al. Genome
of the marsupial Monodelphis domestica reveals innovation in non-coding sequences. Nature.
2007;447(7141):167-77.
54.   Renfree MB, Papenfuss AT, Deakin JE, Lindsay J, Heider T, Belov K, et al. Genome
sequence of an Australian kangaroo, Macropus eugenii, provides insight into the evolution of
mammalian reproduction and development. Genome biology. 2011;12(8):1-26.
55.   Aken BL, Achuthan P, Akanni W, Amode MR, Bernsdorff F, Bhai J, et al. Ensembl
2017. Nucleic acids research. 2016;45(D1):D635-D42.
56.   Johnson RN, O'Meally D, Chen Z, Etherington GJ, Ho SYW, Nash WJ, et al. Adaptation
and conservation insights from the koala genome. Nature genetics. 2018;50(8):1102.
57.   Yang Z. PAML 4: phylogenetic analysis by maximum likelihood. Molecular biology and
evolution. 2007;24(8):1586-91.
58.   Yang Z. PAML: a program package for phylogenetic analysis by maximum likelihood.
Bioinformatics. 1997;13(5):555-6.
59.   Mitchell KJ, Pratt RC, Watson LN, Gibb GC, Llamas B, Kasper M, et al. Molecular
phylogeny, biogeography, and habitat preference evolution of marsupials. Molecular biology and
evolution. 2014;31(9):2322-30.
60.   Anisimova M, Bielawski JP, Yang Z. Accuracy and power of the likelihood ratio test in
detecting adaptive molecular evolution. Molecular biology and evolution. 2001;18(8):1585-92.
61.   Bielawski JP, Yang Z. Maximum likelihood methods for detecting adaptive protein
evolution. Statistical methods in molecular evolution: Springer; 2005. p. 103-24.
62.   Shaffer JP. Multiple hypothesis testing. Annual review of psychology. 1995;46(1):561-
84.
63.   Massey Jr FJ. The Kolmogorov-Smirnov test for goodness of fit. Journal of the American
statistical Association. 1951;46(253):68-78.
64.   Scholz FW, Stephens MA. K-sample Anderson–Darling tests. Journal of the American
Statistical Association. 1987;82(399):918-24.
65.   Szkiba D, Kapun M, von Haeseler A, Gallach M. SNP2GO: functional analysis of
genome-wide association studies. Genetics. 2014;197(1):285-9.
66.   Mi H, Muruganujan A, Ebert D, Huang X, Thomas PD. PANTHER version 14: more
genomes, a new PANTHER GO-slim and improvements in enrichment analysis tools. Nucleic
acids research. 2018;47(D1):D419-D26.
67.   Subramanian A, Tamayo P, Mootha VK, Mukherjee S, Ebert BL, Gillette MA, et al.
Gene set enrichment analysis: a knowledge-based approach for interpreting genome-wide
expression profiles. Proceedings of the National Academy of Sciences. 2005;102(43):15545-50.
68.   Wright B, Willet CE, Hamede R, Jones M, Belov K, Wade CM. Variants in the host
genome may inhibit tumour growth in devil facial tumours: evidence from genome-wide
association. Sci Rep. 2017;7(1):423.

69. Wu K, Li Z, Cai S, Tian L, Chen K, Wang J, et al. EYA1 phosphatase function is
essential to drive breast cancer cell proliferation through cyclin D1. Cancer research.
2013;73(14):4488-99.
70. Cai S, Cheng X, Liu Y, Lin Z, Zeng W, Yang C, et al. EYA1 promotes tumor
angiogenesis by activating the PI3K pathway in colorectal cancer. Experimental cell research.
2018;367(1):37-46.
71. Meurs EF, Galabru J, Barber GN, Katze MG, Hovanessian AG. Tumor suppressor
function of the interferon-induced double-stranded RNA-activated protein kinase. Proceedings of
the National Academy of Sciences. 1993;90(1):232-6.
72. Brüniche-Olsen A, Austin JJ, Jones ME, Holland BR, Burridge CP. Detecting selection
on temporal and spatial scales: a genomic time-series assessment of selective responses to devil
facial tumor disease. PloS one. 2016;11(3).
73. Hendricks S, Epstein B, Schönfeld B, Wiench C, Hamede R, Jones M, et al. Conservation
implications of limited genetic diversity and population structure in Tasmanian devils
(Sarcophilus harrisii). Conservation Genetics. 2017;18(4):977-82.
74. Margres MJ, Ruiz-Aravena M, Hamede R, Jones ME, Lawrance MF, Hendricks SA, et
al. The genomic basis of tumor regression in Tasmanian devils (Sarcophilus harrisii). Genome
biology and evolution. 2018;10(11):3012-25.
75. Fang WH, Wang Q, Li HM, Ahmed M, Kumar P, Kumar S. PAX 3 in neuroblastoma:
oncogenic potential, chemosensitivity and signalling pathways. Journal of cellular and molecular
medicine. 2014;18(1):38-48.
76. Zhang J. Frequent false detection of positive selection by the likelihood method with
branch-site models. Molecular Biology and Evolution. 2004;21(7):1332-9.
77. Kosiol C, Vinař T, da Fonseca RR, Hubisz MJ, Bustamante CD, Nielsen R, et al. Patterns
of positive selection in six mammalian genomes. PLoS genetics. 2008;4(8).
78. Nelson WJ, Nusse R. Convergence of Wnt, ß-catenin, and cadherin pathways. Science.
2004;303(5663):1483-7.
79. Metzger MJ, Reinisch C, Sherry J, Goff SP. Horizontal transmission of clonal cancer
cells causes leukemia in soft-shell clams. Cell. 2015;161(2):255-63.
80. Strakova A, Murchison EP. The cancer which survived: insights from the genome of an
11000 year-old cancer. Curr Opin Genet Dev. 2015;30:49-55.
81. Casanova J-L, Abel L, editors. Human genetics of infectious diseases: Unique insights
into immunological redundancy2018 2018: Elsevier.
82. Guiler E. The Tasmanian devil: St. David's Park Pub.; 1992.
83. Miller IF, Metcalf CJE. Evolving resistance to pathogens. Science.
2019;363(6433):1277-8.
84. Gandon S, Michalakis Y. Evolution of parasite virulence against qualitative or
quantitative host resistance. Proceedings of the Royal Society of London Series B: Biological
Sciences. 2000;267(1447):985-90.
85. Patton AH, Lawrance MF, Margres MJ, Kozakiewicz CP, Hamede R, Ruiz-Aravena M,
et al. A transmissible cancer shifts from emergence to endemism in Tasmanian devils. Science.
2020;370(6522).
86. Kardos M, Shafer ABA. The Peril of Gene-Targeted Conservation. Trends Ecol Evol.
2018;33(11):827-39.

87.   McLennan EA, Grueber CE, Wise P, Belov K, Hogg CJ. Mixing genetically
differentiated populations successfully boosts diversity of an endangered carnivore. Animal
Conservation. 2020.
88.   Wright B, Morris K, Grueber CE, Willet CE, Gooley R, Hogg CJ, et al. Development of
a SNP-based assay for measuring genetic diversity in the Tasmanian devil insurance population.
BMC Genomics. 2015;16:791.
89.   Campbell NR, Harmon SA, Narum SR. Genotyping-in-Thousands by sequencing (GT-
seq): A cost effective SNP genotyping method based on custom amplicon sequencing. Mol Ecol
Resour. 2015;15(4):855-67.
90.   von Thaden A, Nowak C, Tiesmeyer A, Reiners TE, Alves PC, Lyons LA, et al.
Applying genomic data in wildlife monitoring: Development guidelines for genotyping degraded
samples with reduced single nucleotide polymorphism panels. Molecular Ecology Resources.
2020;20(3).
91.   Margres MJ, Ruiz-Aravena M, Hamede R, Chawla K, Patton AH, Lawrance MF, et al.
Spontaneous tumor regression in Tasmanian devils associated with RASL11A activation.
Genetics. 2020.
712
713

Supplementary materials S1
Methods
Rapture Sequencing for Contemporary Selection
We multiplexed 672-868 individuals per lane of Illumina NextSeq, obtaining 2.4 billion 150 base
pair (bp) paired-end reads. We later re-sequenced 2,379 individuals from those Rapture
libraries on an Illumina HiSeq 4000 to increase coverage and confidence in genotype inference.
Finally, we also incorporated data from standard RAD sequencing libraries on four additional
individuals from two of the six populations, West Pencil Pine and Fentonbury (see 1, 2 for
details.).
723
For each individual, we merged all available reads from across sequencing efforts. Then reads
were de-multiplexed and low-quality reads were removed using process_radtags in Stacks using
the '--bestrad' option which checks for the single restriction enzyme cut site on either read; this
step also removed any reads without recognizable barcodes or cut sites (3). The Stacks
clonefilter program was used to remove potential PCR duplicates (4). Using bowtie2 (5), reads
were aligned to the *S. harrisii* reference genome Devil_ref v7.0 (6), downloaded from Ensembl
in June 2014. Population allele frequencies can often be estimated with greater accuracy and
reduced bias than individual genotypes (7). To do this, we used ANGSD v0.910 (8). For each set
of individuals, we calculated genotype likelihoods and estimated allele frequencies within
regions using the settings in Supplemental Table S2. Regions on the X chromosome were
excluded.
735
Spatial and temporal analysis of contemporary selection
First, we calculated allele frequency change after DFTD infection. Within the five locations for
which we had sampling both before and after DFTD appearance, we estimated the magnitude
and direction of allele frequency change at SNPs with data from at least 10 individuals at both
time points and a minor allele frequency (MAF) ≥ 0.05 and a likelihood-ratio test p-value (for
presence of a SNP from ANGSD) $\leq 10^{-6}$ in at least one of the time points. For this analysis,
individuals were assigned to "before" or "after" time points based on the date of DNA
collection. Table 1 presents the first year of DFTD detection in each population and samples
collected from individuals after those years were considered "after." DFTD could have been
present at very low frequency prior to detection in some of these populations, but we believe
using these dates of detection still provides a good estimate of pre-DFTD allele frequencies. We
performed an arcsine (Fisher's angular) transformation on the estimated allele frequencies to
reduce bias induced by the allele frequency spectrum (9). The SNPs were ranked by the
magnitude of change, and the fractional rank was used as a pseudo- p-value for the composite
statistic (described in the Main Text).
751
Then, we identified SNPs with estimates of strong selection using two time-series approaches
which account for a population structure: the method of Mathieson & McVean (10) (hereafter
mm), which allows the estimated selection coefficient to vary over space; and *spatpg*, which
allows the selection coefficient to vary over time as well as space (11). Individuals were divided
into two-year cohorts based on their year of birth, starting with 1997 and ending with 2012
(Supplemental Table S1). Only SNPs with MAF ≥ 0.05, minor allele count ≥ 3, and p-value from

ANGSD $\leq 10^{-6}$ in at least five population / cohort combinations were tested. For *mm*, we
assumed the same effective population sizes as Epstein et al. (2016), and otherwise assumed
similar effective population sizes where previous estimates did not exist (Supplemental Table
S3). For *spatpg*, variance effective population sizes were estimated within the program with a
bounded prior between 25 and 40. We created input allele count files for both *spatpg* and *mm*
by multiplying allele frequencies estimated in ANGSD and the number of individuals and
rounded to the nearest whole number. Due to computational limits, the dataset was randomly
divided into 18 separate *spatpg* runs. Following *spatpg* recommendations in the manual (11),
we calculated a support value for each SNP by taking the proportion of the posterior
distribution (i.e. proportion of MCMC steps) for which the regression coefficient β, was non-
zero ($0 < \beta$ or $0 > \beta$, whichever was smaller), where β describes the association between
allele change and presence/absence of DFTD within a population. We multiplied these support
values by two and treated them as pseudo- p-values when calculating a genomic inflation factor
and the composite statistic.
772
Following the recommendations in Francois and colleagues (12), we adjusted the p-values of
each test to reduce false positives. If the distribution of p-values was not uniform, we divided
the Z-scores by the inflation factor. The inflation factor was calculated as the ratio between the
median Z-scores and the expected median Z-scores for a χ^2 distribution with one degree of
freedom. The genomic inflation factor varied from $0.25 - 0.58$ for *mm* analyses of each
population. For *spatpg*, we found an inflation factor of 0.44.
779
Using the adjusted p-values and pseudo p-values from the individual analyses, we calculated
the DCMS statistic (13). This statistic combines the p-values from different tests at each SNP
while accounting for genome-wide correlation among statistics. For each SNP, we used a weight
based on only the statistics with results at that SNP, and we only included SNPs with results
from at least eleven of the twelve individual tests (Δaf for five populations, *mm* for all six
populations, and *spatpg*).
786
We then divided DCMS by the number of defined tests to get a mean composite score and
ranked SNPs by this mean score. Because DCMS is not defined when one of the p-values is one
or zero, we replaced p-values of one with 0.99999, and values of zero (only occurred for
spatpg) with 0.00005 before performing the calculation. In accordance with previous linkage
disequilibrium estimates (2), we identified candidate genes within 100 kb of top SNPs, using
bedtools (version 2.26.0) (14); and supplemented annotations of novel genes with the Ensembl
Compara gene family pipeline (15).
794
Historical Selection
To test for historical selection we used the branch-site test of PAML (Phylogenetic Analysis by
Maximum Likelihood; 16, 17), implemented in the Bio.Phylo toolkit (18) of BioPython (19). The
branch-site test estimates the ratio of nonsynonymous-synonymous mutation rates (dN/dS)
among aligned codons using a phylogenetic tree to allow for the appropriate evolutionary
model to be employed. In the neutral model, all site classes and branches are constrained to

dN/dS ≤ 1. In the alternative model, dN/dS of Site Class 2 is allowed to exceed 1 for only the
foreground branch, while constraining the background branches to dN/dS ≤ 1.
803
The devil, wallaby, and opossum orthologous genes and respective sequences were mined from
the Ensembl database with BioMart (20, 21). For the koala, we used the orthologs identified
with blastx version 2.2.27+ (22) supplied by Johnson and colleagues (23). If splicing variants
were available, only the first (most common) variant was retained for downstream data
preparation. Only 1-to-1 orthologs were retained (i.e., paralogs were excluded). Orthologous
gene tables were then reduced to genes with at least three of four possible sequences present
and the respective species were pruned from the greater phylogeny. For the koala, we used an
open reading frame finder, getorf from EMBOSS, (24) to generate amino acid sequences. The
peptide sequence alignments were generated with MUSCLE version 3.8.31 (25, 26), then used
to guide alignments of nucleotides with tranalign.
814
Functional Enrichment of Genes Under Selection
For contemporary candidate SNPs, we used the SNP2GO package (27) in the R environment
(version 3.4.3). We filtered the most recent Gene Transfer File
(Sarcophilus_harrisii.DEVIL7.0.100) and Ensembl gene ID GO term associations downloaded
from Ensembl May 7, 2020 to include only genes which were within 100 kb of targeted loci and
account for the biased subset of targeted genes. We limited enrichment analysis to GO terms
with a minimum of one association in the reference set and allowed an extension window of
100 kb from a candidate SNP. For GO term enrichment analysis of historical selection, we used
the PANTHER web-interface with HUGO gene names (28) and the reference set defined as all
genes that were tested in PAML (29).
825
826 We only compared the distributions among genes for which dN/dS > 1 and were statistically
827 significant according to the likelihood ratio test. To account for the bias induced by targeted
828 sequencing among contemporary candidates, we defined the reference set as all genes that
829 could have been detected in both tests for a given overlapping or enrichment analysis
830
We capitalized on the wealth of ongoing research in devils and DFTD by comparing our
contemporary and historical candidates to those previously identified using different datasets
and analytical approaches (2, 30-33) using Fisher's Exact Test implemented in the R package
GeneOverlap (34). We then tested for overrepresentation of contemporary and historical
candidates in gene sets of the molecular signatures database (MsigDB) (35). MSigDB contains
several libraries of gene sets which allowed us to gain further insight to pathways that may be
under selection in devils. We compared our contemporary and historical candidate gene lists to
gene sets from the MsigDB Hallmark, Curated, Computational, Oncogenic Signatures, and
Immunologic Signatures (36, 37). We built 2x2 contingency tables for each set of genes under
positive selection and in each of the tested gene sets of MsigDB. Despite limitations, this
overrepresentation method is straight-forward and flexible for non-model organisms and
targeted sequencing. To identify intersecting gene sets, we first converted all Ensembl gene IDs
of interest to HUGO annotations with Biomart. Second, we created appropriate background
sets (of length N) by intersecting each MsigDB gene set with the respective list of all genes that

were tested for selection. Lastly, from these contingency tables we computed overlaps, using
the hypergeometric distribution (dhyper(c(0:x), m, n, x, log = FALSE), where c(0:x) is a vector of
quantiles representing the number of genes both under selection and found in an MsigDB gene
set; m is the number of genes in the candidate gene list, n is the number of genes in the
candidate gene list but not in the MSigDB gene set, and x is the number of candidate genes in
the MsigDB gene set list. After accounting for multiple testing with the Benjamini-Hochberg
correction (38), we considered the overrepresentation result statistically significant with
adjusted p-values < 0.05 and background gene sets greater than or equal to ten genes. Finally,
we performed a permutation test to establish a null expectation for the rate of shared gene
sets between contemporary and historical selection and compared the resultant empirical null
to our observed proportion of shared gene sets. To do this, we randomly selected the same
number of candidate genes with HUGO annotations (112) from the list of all HUGO annotated
contemporary candidates (3,920) 1,000 times, with replacement, and performed gene set
overlap analysis as above with only the known gene sets significantly overlapping with the
historical candidates.

Figures

Supplemental Figure S1. Mean coverage of individuals across populations at targeted loci

Figure S2. The folded allele frequency spectra for each population before DFTD became prevalent. wuaklina/Mt. William is not presented because it was first sampled in 2004, eight years after DFTD was first described at that locality.

Figure S3. Un-adjusted p-values for all SNPs of each population analysed with *mm* (10).

Figure S4. Adjusted p-values (12) for all SNPs of each population analysed with *mm* (10).

Figure S5. Un-adjusted p-values for all SNPs of each populations analysed with *spatpg* (11).

Figure S6. Adjusted p-values (12) for all SNPs of each population analysed with *spatpg* (10, 11).

Figure S7. Allele frequency change (Δaf) for each population separately. SNPs in the top 1% are indicated by more opaque points. The threshold line for the top 1% within each population is indicated by a dashed line.

Figure S8. Signatures of selection as detected by *mm* for each population separately. SNPs in the top 1% are indicated by more opaque points. The threshold line for the top 1% within each population is indicated by a dashed line.

Figure S9. Binned initial allele frequency distributions for DCMS candidates and non-candidates across all populations with samples before DFTD became prevalent.

Supplemental Figure S10. Correlations among elementary statistics: afchange = allele frequency change; mm = Mathieson and McVean (10), spatpg (11). Correlations are clustered by similarity along the x-axis.

Figure S11. Histogram of shared MSigDB (35) gene set overlaps between contemporary and historical candidates, i.e. the signature of ongoing selection, in a permutation test of 1000 draws with replacement. We observed no shared gene set overlaps between contemporary and historical candidates. Of the 1000 permutations, 100% had fewer overlapping sets than observed.

Tables

Supplemental Table S1. Number and year of sampling across seven localities.

Supplemental Table S2. ANGSD (8) genotype calling settings.

Supplemental Table S3. Estimates of effective population size.

903   Supplemental Table S4. Number of shared genes within 10000 bp of the top 1% of SNPs for the
904   DCMS list of contemporary candidates and each intermediate test. The total number of genes
905   in each list is found in the diagonal. Populations are abbreviated as follows: FEN = Fentonbury,
906   FOR = Forestier, FREY = Freycinet, wuk = wukalina/Mt. William, NAR= Narawntapu, and WPP =
907   West Pencil Pine.
908
909   Supplemental Table S5. Annotated Tasmanian devil gene IDs of within 1000 bp of SNPs in the
910   top 1% of *DCMS* scores; i.e. candidates for contemporary selection. Also provided at:
911   https://github.com/Astahlke/contemporary_historical_sel_devils/blob/master/contemporary/angsd_2019-01-
912   18/next/composite_stat/2019-2-22/results/annotation_top1.0/composite.snps.everything.top.genes.100000bp.txt.
913
Supplemental Table S6. Also provided at:
https://github.com/Astahlke/contemporary_historical_sel_devils/blob/historical/ sig_paml_results_21-1-13.csv . All
PAML (16, 17) branch site test results for the 1,773 genes with inferred historical positive
selection. Columns indicate: ensemble_gene_id = Ensembl gene ID for Tasmanian devils;
likelihood = likelihood of the three-ratio model; p0 = proportion of sites belonging to site class
0, f0= dN/dS estimates for devil branch in site class 0 ; p1= proportion of sites in site class 1; f1=
dN/dS estimates for devil branch in site class 0; p2 = proportion of sites belonging to site class
2a; f2 = dN/dS estimates for devil branch in site class 2a; p3 = proportion of sites belonging to
site class 2b; dN/dS estimates for devil branch in site class 2b; f3 = dN/dS estimates for devil
branch in site class 2b FDR = likelihood-ratio test p-values adjusted for multiple testing (38);
p2a_2b = proportion of sites belonging to either cite class 2a or 2b
Site Class 0 = Background: 0< dN/dS<1; Foreground: 0< dN/dS<1
Site Class 1 = Background: dN/dS=1; Foreground: dN/dS=1
Site Class 2a = Background: dN/dS<1; Foreground: dN/dS≥1
Site Class 2b = Background: dN/dS=1; Foreground: dN/dS≥1
929
Supplemental Table S7. Sixteen candidate genes for both historical positive selection (p < 0.05
in PAML branch-site test) and a response to contemporary selection from DFTD (top 1% DCMS
scores). Novel genes are annotated here by the Ensembl Compara gene family pipeline (15). *
denotes genes previously associated with disease-related phenotypes in many of the same
devils (32).

Figure S1. Mean coverage of individuals across populations at targeted loci.

936
937
938

939 Figure S2. The folded allele frequency spectra for each population before DFTD became
940 prevalent. wuaklina/Mt. William is not presented because it was first sampled in 2004, eight
941 years after DFTD was first described at that locality.

943 Figure S3. Un-adjusted p-values for all SNPs of each population analysed with *mm (10)*.

945　Figure S4. Adjusted p-values (12) for all SNPs of each population analysed with *mm (10)*.

946
947
948

949   Figure S5. Un-adjusted p-values for all SNPs across all populations analysed with *spatpg (11).*

950
951

952 Figure S6. Adjusted p-values (12) for all SNPs across all populations analysed with *spatpg (10,*
953 *11).*

955 Figure S7. Allele frequency change (Δaf) for each population separately. SNPs in the top 1% are
956 indicated by more opaque points. The threshold line for the top 1% within each population is
957 indicated by a dashed line.

958
959

Figure S8. Signatures of selection as detected by *mm* for each population separately. SNPs in
the top 1% are indicated by more opaque points. The threshold line for the top 1% within each
population is indicated by a dashed line.

Figure S9. Binned initial allele frequency distributions for DCMS candidates and non-candidates
across all populations with samples before DFTD became prevalent.

Figure S10. Correlations among elementary statistics: afchange = Allele frequency change (Δaf);
mm = Mathieson & McVean. Correlations are clustered by similarity along the x-axis.

Figure S11. Histogram of shared MSigDB (35) gene set overlaps between contemporary and
historical candidates, i.e. the signature of ongoing selection, in a permutation test of 1000
draws with replacement. We observed no shared gene set overlaps between contemporary and
historical candidates. Of the 1000 permutations, 100% had more overlapping sets than
observed.

 Supplemental Table S1. Number and year of sampling across six localities.

Population	Year	Number of individuals
wukalina/Mt. William	2004	11
wukalina/Mt. William	2005	28
wukalina/Mt. William	2006	20
wukalina/Mt. William	2007	21
wukalina/Mt. William	2008	17
wukalina/Mt. William	2009	14
wukalina/Mt. William	2010	8
wukalina/Mt. William	2011	15
wukalina/Mt. William	2012	4
wukalina/Mt. William	2013	10
wukalina/Mt. William	2014	8
Freycinet	1999	107
Freycinet	2000	122
Freycinet	2001	71
Freycinet	2002	65
Freycinet	2003	38
Freycinet	2004	60
Freycinet	2005	54
Freycinet	2006	27
Freycinet	2007	32
Freycinet	2008	21
Freycinet	2009	7
Freycinet	2010	10
Freycinet	2011	10
Freycinet	2012	18
Freycinet	2013	12
Freycinet	2014	26
Forestier	2004	131
Forestier	2005	46
Forestier	2006	168
Forestier	2007	98
Forestier	2008	93
Forestier	2009	107
Forestier	2010	13
Forestier	2012	26
Forestier	2013	1
Fentonbury	2004	47

Fentonbury	2005	52
Fentonbury	2006	58
Fentonbury	2007	63
Fentonbury	2008	37
Fentonbury	2009	11
West Pencil Pine	2006	52
West Pencil Pine	2007	71
West Pencil Pine	2008	33
West Pencil Pine	2009	57
West Pencil Pine	2010	33
West Pencil Pine	2011	84
West Pencil Pine	2012	62
West Pencil Pine	2013	43
West Pencil Pine	2014	13
Narawntapu	1999	33
Narawntapu	2003	9
Narawntapu	2004	46
Narawntapu	2005	27
Narawntapu	2006	64
Narawntapu	2007	45
Narawntapu	2008	63
Narawntapu	2009	30
Narawntapu	2010	27
Narawntapu	2011	15
Narawntapu	2012	22

978   Supplemental Table S2. ANGSD genotype calling settings
979

Option	Setting	Description
-minMapQ	40	Minimum read mapping quality
-minQ	25	Minimum base Phred score
-baq	2	Perform extended base quality adjustment around indels (Li 2010)
-GL	2	Use the (old) GATK genotype likelihood model
-doMaf	2	Estimate allele frequencies assuming one known allele
-doMajorMinor	4	Use the reference allele as the known allele

981   Supplemental Table S3. Estimates of effective population size.
982

Population	Ne estimate
Fentonbury	35
Forestier	35
Freycinet	34
Narawntapu	37
West Pencil Pine	26

991 Supplemental Table S4. Number of shared genes within 10000 bp of the top 1% of SNPs for the DCMS list of contemporary
992 candidates and each intermediate test. The total number of genes in each list is found in the diagonal. Populations are abbreviated
993 as follows: FEN = Fentonbury, FOR = Forestier, FREY = Freycinet, wuk = wukalina/Mt. William, NAR= Narwantapu, and WPP = West
994 Pencil Pine.

	DCMS	FEN Δaf	FEN mm	FOR Δaf	FOR mm	FREY Δaf	FREY mm	wuk mm	NAR Δaf	NAR mm	spatpg	WPP Δaf	WPP mm
DCMS	247	14	60	34	67	28	32	41	17	55	39	14	26
FEN Δaf	-	165	20	24	31	15	8	19	26	21	9	4	16
FEN mm	-	-	279	33	22	23	22	8	42	38	19	18	20
FOR Δaf	-	-	-	301	30	20	28	28	25	48	17	34	22
FOR mm	-	-	-	-	283	14	14	38	28	26	21	28	28
FREY Δaf	-	-	-	-	-	299	19	8	29	12	31	24	42
FREY mm	-	-	-	-	-	-	289	16	16	53	13	16	26
wuk mm	-	-	-	-	-	-	-	295	21	48	39	8	22
NAR Δaf	-	-	-	-	-	-	-	-	282	24	21	46	42
NAR mm	-	-	-	-	-	-	-	-	-	314	36	21	30
spatpg	-	-	-	-	-	-	-	-	-	-	284	15	21
WPP Δaf	-	-	-	-	-	-	-	-	-	-	-	329	29
WPP mm	-	-	-	-	-	-	-	-	-	-	-	-	232

995　Supplemental Table S5. Annotated Tasmanian devil gene IDs of within 1000 bp of the top 1% of
996　composite SNPs; i.e., candidates for contemporary selection. Also provided at:
997　https://github.com/Astahlke/contemporary_historical_sel_devils/blob/master/contemporary/angsd_2019-01-
998　18/next/composite_stat/2019-2-22/results/annotation_top1.0/composite.snps.everything.top.genes.100000bp.txt