Spatial population genetics in heavily managed species: Separating patterns of historical translocation from contemporary gene flow in white-tailed deer

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ABSTRACT

Approximately 100 years ago, unregulated harvest nearly eliminated white-tailed deer (*Odocoileus virginianus*) from eastern North America, which subsequently served to catalyze wildlife management as a national priority. An extensive stock-replenishment effort soon

24   followed, with deer broadly translocated among states as a means of re-establishment. However,

      an unintended consequence was that natural patterns of gene flow became obscured and pre-

26   translocation signatures of population structure were replaced. We applied cutting-edge

      molecular and biogeographic tools to disentangle genetic signatures of historical management

28   from those reflecting spatially heterogeneous dispersal by evaluating 35,099 single nucleotide

      polymorphisms (SNPs) derived via reduced-representation genomic sequencing from 1,143 deer

30   sampled state-wide in Arkansas. We then employed Simpson's diversity index to summarize

      ancestry assignments and visualize spatial genetic transitions. Using sub-sampled transects

32   across these transitions, we tested clinal patterns across loci against theoretical expectations of

      their response under scenarios of recolonization and restricted dispersal. Two salient results

34   emerged: (A) Genetic signatures from historic translocations are demonstrably apparent; and (B)

      Geographic filters (major rivers; urban centers; highways) now act as inflection points for the

36   distribution of this contemporary ancestry. These results yielded a state-wide assessment of

      contemporary population structure in deer as driven by historic translocations as well as ongoing

38   processes. In addition, the analytical framework employed herein to effectively decipher

      extant/historic drivers of deer distribution in Arkansas are also applicable for other biodiversity

40   elements with similarly complex demographic histories.

1 | INTRODUCTION

44   Understanding movement behaviors and patterns of dispersal across the landscape are important

      components of a wildlife management strategy. Quantifying dispersal and connectivity among

46   populations has traditionally been done using direct methods such as radio- and satellite-tracking

(Kays, Crofoot, Jetz, & Wikelski, 2015), yet these methods are labor-intensive, generally only

48   allow for small sample-sizes with results difficult to extrapolate at the population level (Katzner

& Arlettaz, 2020). An ability to do so indirectly via patterns measured from genetic data such as

50   gene flow (i.e., corresponding to migration, demographics) and genetic drift (i.e., effective

population sizes, inbreeding) has made it possible to more accurately assess movement and

52   landscape ecology patterns at a fraction of the per-individual cost (Bossart & Pashley Prowell,

1998; Comte & Olden, 2018; Lowe & Allendorf, 2010). Continual decline of sequencing costs

54   concomitant with computational and analytical advancements now offer an unprecedented

density of sampling and geographic resolution that can link individual movement patterns with

56   landscape genetic associations (Richardson, Brady, Wang, & Spear, 2016), as well as functional

genomic correlations (Schoville et al., 2012). This has spurred a proliferation of approaches,

58   such as deducing natal assignment (Battey, Ralph, & Kern, 2020a), inferring the directionality

and pathways of range expansions (Rengefors et al., 2020), and predicting disease spread from

60   spatial host-pathogen associations (Fountain-Jones et al., 2021; Kozakiewicz et al., 2018).

However, contemporary evaluations examine but a single 'snapshot' from an

62   evolutionary continuum of events that have cumulatively shaped contemporary diversity, a

consequence of which is that genetic patterns may not be interpretable when juxtaposed against

64   modern landscapes (Douglas, Minckley, & DeMarais, 1999). Thus, models that invoke landscape

genetic associations may incorporate implicit assumptions that lack biological reality (Cushman

66   & Landguth, 2010). For example, one common assumption involves dispersal among a series of

discontinuous sub-populations (Lundgren & Ralph, 2019; Petkova, Novembre, & Stephens,

68   2015) that may incorrectly infer discrete causes (such as barrier effects or reproductive

isolation)(Bradburd, Coop, & Ralph, 2018). Likewise, gene flow and genetic drift are often

70   conflated, in that one may generate spurious signals of the other (Battey, Ralph, & Kern, 2020b;

Mazet, Rodríguez, Grusea, Boitard, & Chikhi, 2016). For example, demographies with a history

72   of turbulence may inflate genetic divergences such that population structure appears artificially

discrete and thus suggests restricted gene flow (Austerlitz, Jung-Muller, Godelle, & Gouyon,

74   1997; Excoffier & Ray, 2008; Nei, Maruyama, & Chakraborty, 2014). A classic example of this

phenomenon is the re-expansion of populations following an historical ebb. Upon secondary

76   contact, unique genetic changes that have accumulated during periods of isolation may

seemingly initiate an apparent rapid shift in ancestries when viewed from a landscape genetic

78   context (Zellmer & Knowles, 2009).

More recent events such as anthropogenic habitat fragmentation may also act to obscure

80   genetic patterns across the landscape (Epps, Wasser, Keim, Mutayoba, & Brashares, 2013;

Landguth et al., 2010; Pavlacky, Goldizen, Prentis, Nicholls, & Lowe, 2009). In particular,

82   human-mediated dispersal introduces both diversifying and homogenizing effects that prove

difficult to disentangle (Epps & Keyghobadi, 2015; Seebens et al., 2017). This erodes adaptive

84   and/ or spatial associations in that it disturbs the expected relationships between genetic

differentiation and geographical or environmental proximity (Capinha, Essl, Seebens, Moser, &

86   Pereira, 2015; Einfeldt, Jesson, & Addison, 2020). Well-intended actions, such as translocations

and stockings, may also inadvertently obscure pre-existing landscape genetic patterns (Brown,

88   Hull, Updike, Fain, & Ernest, 2009; Laikre, Schwartz, Waples, & Ryman, 2010; Shephard,

Ogden, Tryjanowski, Olsson, & Galbusera, 2013). This is especially so for heavily-managed

90   species that incur translocations as one aspect of their management (Jahner et al., 2019; Seddon,

Griffiths, Soorae, & Armstrong, 2014)

92   White-tailed deer, one of the most recreationally important species in North America, is a

case in point (Knoche & Lupi, 2012), with populations intensely impacted by both hunter harvest

94   and game management (Waller & Alverson, 1997; Wolverton, Kennedy, & Cornelius, 2007).

Early in the 20th century, deer and other species extensively hunted (e.g., wild turkey) saw

96   widespread extirpations. In a census of the north-central United States, Leopold (1931) noted

that extant deer populations were of low abundance in southern Missouri, an observation that

98   dovetails with estimated numbers of deer in Arkansas during that period (~500 scattered across

isolated refugia at the low ebb; Holder, 1951). Subsequent efforts to bolster resident numbers and

100   re-populate depleted areas in Arkansas involved translocations from within- and out-of-state.

Although successful, these efforts promoted genetic patterns within the state that were based on

102   artificial, rather than natural, movements.

Distributional patterns in white-tailed deer clearly reflect multiple factors. Those extrinsic

104   may include: Rivers, interstate highways, suitable habitat (Robinson, Samuel, Lopez, & Shelton,

2012); agricultural land use and urbanization (Kelly et al., 2014); as well as regional eco-

106   physiographic boundaries (Miller, Miller-Butterworth, Diefenbach, & Walter, 2020). Intrinsic

factors, on the other hand, may point to: Population density (Lutz, Diefenbach, & Rosenberry,

108   2015); age structure/ sex ratio (Long, Diefenbach, Rosenberry, & Wallingford, 2008); and social

hierarchy (Nixon & Mankin, 2016).

110   A first step in designating factors that have shaped deer population structure and dispersal

is to estimate the genetic component within Arkansas deer genomes that still persists as a

112   residual of historic translocations and subsequent genetic drift. Perhaps our most pressing need

with regard to dispersal is the impact these data may have in relation to a widespread and fatal

114   neurodegenerative disease of cervids (chronic wasting disease; CWD), which now represents a

panzootic (Williams and Young 1980; Escobar et al. 2019; Mawdsley 2020). Thus, efforts to

116   contain and mitigate its spread are paramount for wildlife management, not only in North

America, but also globally (Leiss et al., 2017).

1.1 | Separating contemporary dispersal from historical translocation

120   The driving factors for deer dispersal primarily relate to habitat quality and climate, and these in

turn elicit obvious management interest. Region-specific patterns are needed so as to develop

122   broader generalizations regarding species-specific movement ecology (Brinkman, Deperno,

Jenks, Haroldson, & Osborn, 2005), and to evoke management strategies that can optimize

124   hunter-harvest while maintaining population densities.

The goal of our study was to ascertain if signatures of historic translocations are apparent

126   in, and have contributed to, the genetic diversity and structure of deer in Arkansas, and

furthermore, if they can indeed be parsed from a confusing matrix of ongoing gene flow and

128   genetic drift. To do so, we surveyed a broad array of nuclear genomic markers using ddRAD

sequencing (Peterson et al., 2012). Our first objective was to characterize respective patterns of

130   anthropogenic translocation and natural re-colonization via refugial populations, and secondarily

to seek evidence for geographic barriers throughout the state that seemingly prevent or actively

132   filter deer dispersal. However, we hypothesized that the former situation would obscure the

latter, in that long-range anthropogenically-mediated displacement of individuals violates

134   methodological assumptions that stipulate gene flow as occurring in a spatially consistent

manner (Bradburd & Ralph, 2019).

136   Here, we used ancestry assignment probabilities (e.g., for K-number of populations) and

reduced this to a single index (Simpson, 1949) that encapsulated the diversity of ancestries

138   within a given spatial grain. This served to reduce those genetic ancestry components that are

influenced by widespread translocation, to a single interpolated 'surface' representing only

140 changes in ancestry over space (i.e., boundaries between contemporary populations). We also

wished to define the respective roles played by stochastic and deterministic processes in

142 generating these zones, particularly given historic population fluctuations and the internal/

external translocations that conflated the demography of Arkansas deer. To do so, we borrowed

144 components of cline theory to hypothesize how individual loci should transition across these

spaces, relative to a genome-wide average (Barton, 1983; Barton & Hewitt, 1985; Endler, 1973;

146 Hewitt, 2001; Polechová & Barton, 2011; Slatkin, 1973).

Although clinal variation is often taken as evidence for selection (i.e., migration-selection

148 balance; Haldane, 1948), genetic drift in concert with spatially variable gene flow can generate

patterns in individual loci that mimic those seen in adaptive clines (Vasemägi, 2006). With drift

150 operating alone, allele frequencies at each locus may wander 'upward' or 'downward' across

space as a function of the initial allele frequency. Range expansion or dissemination from refugia

152 can also yield a type of 'rolling' founder effect that can mirror a clinal pattern (Excoffier, Foll, &

Petit, 2009; Hallatschek & Nelson, 2008; Hewitt, 2000; Keller et al., 2013). Accordingly, we

154 might predict that multi-locus patterns at population boundaries generated solely due to genetic

drift in expanding populations will be dominated by stochastic directional change, with each

156 fluctuating either above or below the genome-wide average (Santangelo, Johnson, & Ness,

2018).

158 On the other hand, a strong barrier to movement would yield a discontinuity in the rate of

directional change across space (Barton, 2008; Nagylaki, 1976; Slatkin, 1973). The existence of

160 a barrier (either impacting dispersal or as a strong fitness differential) thus implies the presence

of an inflection point that would induce a rapid shift in locus-specific ancestries (the 'width' of

162   this cline is the inverse of the slope at the inflection point; Slatkin 1973; Endler 1977; Fitzpatrick

2013).

164   Given these predictions, we summarized locus-wise patterns using two clinal parameters:

α, which describes the direction of genetic change, and β, which indicates 'width' of a cline. We

166   examined how these two parameters varied, then developed hypotheses regarding the manner by

which it impacted our observed population structure. We then considered the evolutionary

168   implications, and limitations, of these results in the context of management, and in so doing

demonstrated how the methodologies employed herein can be used to make inferences in other

170   heavily managed species regarding demographic factors (i.e., sex-biased and natal dispersal) and

those more extrinsic (i.e., landscape resistance) that drive population structure.

2 | METHODS

174   **2.1 | Sampling and data collection**

During 2016-2019, 1,720 tissues were collected by the Arkansas Game and Fish Commission

176   (AGFC), representing all 75 Arkansas counties. We employed a combination of targeted

sampling, road-kill surveys, and a voluntary state-wide CWD testing program (Chafin et al.,

178   2020). Age and sex were collected where possible, with the former estimated by tooth

development and wear (Severinghaus, 1949). Data for an additional 30 samples were also

180   attained from Wisconsin to test for signals of historically recorded translocation efforts involving

this stock (Holder, 1951). From these, a subset of 1,208 samples were chosen for sequencing.

182   We homogenized tongue or ear tissue (stored at -20°C) and extracted genomic DNA

using QIAamp Fast Tissue kits (Qiagen, Inc), with verification via gel electrophoresis (2%

184   agarose). Samples with sufficient yields of high molecular weight DNA (>200ng) were then

enzymatically fragmented via incubation at 37°C, using high-fidelity *NsiI* and *MspI* restriction

186 enzymes (New England Biolabs, Inc.), following enzyme and size-selection optimization using

in silico digests (Chafin, Martin, Mussmann, Douglas, & Douglas, 2018) of several available

188 reference genomes hosted by NCBI: *Odocoileus virginianus* (GCA_002102435.1), *Capreolus*

capreolus (GCA_000751575.1), and *Capra hircus* (GCF_001704415.1).

190   Digests were purified using Ampure XP beads (Beckman-Coulter, Inc.) and standardized

to 100ng per sample. Unique inline barcodes (Peterson et al., 2012) were then ligated using T4

192 DNA Ligase (following manufacturer protocols; New England Biolabs, Inc.). Samples were then

multiplexed (N=48) prior to automated size selection at 300-450bp using a Pippin Prep (Sage

194 Sciences). Adapter extension was performed over 12 PCR cycles using TruSeq-compatible

indexed primers (Illumina, Inc.) and Phusion high-fidelity *taq* polymerase (New England

196 Biolabs, Inc.) Additional quality controls (e.g., qPCR and fragment analysis) were performed on

final libraries prior to 1x100 single-end sequencing on the Illumina HiSeq 4000 (Genomics and

198 Cell Characterization Facility, University of Oregon/Eugene), with a total of N=96 samples

pooled per lane.

200   Raw reads were demultiplexed using the pyRAD pipeline (Eaton, 2014), and those with

barcode mismatches were discarded. Demultiplexed reads were further filtered by removing

202 those having >4 nucleotides below a quality threshold of 99% accuracy. Reads were then

clustered into putative loci within-individuals, allowing for a maximum distance threshold of

204 15%. This was done using the VSEARCH algorithm (Rognes, Flouri, Nichols, Quince, & Mahé,

2016) as implemented in pyRAD, so as to remove read clusters with 3+ indels, >5 ambiguous

206 consensus nucleotides, or a coverage <20X or >500X. Putative homologs were identified using

among-individual clustering with the same parameters, and additional removal of loci having >2

208   alleles per individual, >70% heterozygosity for any polymorphic site, >10 heterozygous sites, or

<50% individual recovery (see *github.com/tkchafin/scripts* for post-processing and file

210   formatting scripts). To mitigate issues of independence among sites, we additionally sub-

sampled the dataset to one SNP per locus.

2.2 | Derivation of population structure

214   Given known issues with respect to assignment accuracy in datasets dominated by low-frequency

variants (Linck & Battey, 2019), we excluded SNPs exhibiting a minor allele count <2 which

216   corresponds to a global minor-allele frequency of ~1%. This minor-allele count threshold per-

locus (i.e., accounting for total number of sampled alleles), equates to a mean minor-allele

218   frequency threshold of 11% (sd=2%). We then inferred population structure (ADMIXTURE:

Alexander et al., 2009) with parallel processing (ADMIXPIPE: Mussmann et al., 2020). Model

220   selection (i.e. for K, the number of populations) followed a cross-validation approach with results

aggregated from 20 independent replicates (CLUMPAK: Kopelman et al., 2015).

222   Individual-level ADMIXTURE results were summarized as a 'surface' with spatial

discontinuities represented as interpolated assignment probabilities. Here, we constructed state-

224   wide rasters, as representing per-pixel probabilities or 'ancestry proportions,' using Empirical

Bayesian Kriging (ARCMAP 10.7.1, Esri, Inc.). Probability surfaces were then summarized as

226   evenness and diversity of ancestries in each cell using Simpson's index (Simpson, 1949) (where

K=number of statewide sub-populations). Our use of the diversity index was based on a

228   straightforward prediction: Areas representing spatial transitions between populations will have a

correspondingly low certainty of assignment to a given sub-population ancestry (i.e., high

230   diversity). Likewise, those with low inter-population exchange will comprise genetically similar

individuals assigned with high probability to the endemic ancestry (i.e., low diversity). Thus,

232   genetic edges represent a marked transition from one population to another, as identified by site-

wise diversity in assignment probabilities (i.e. 'ancestry diversity'). By lowering the resolution of

234   the genetic diversity surface, we effectively decreased the noise introduced by translocation.

We expect spatial variation in ancestry diversity to be inversely proportional to true rates

236   of gene flow, in that this quantity (as well as the ancestry proportions from which it is computed)

are a product of gene flow averaged over many generations. Of note is the fact that the method

238   only examines local ancestry probabilities and consequently will not be dominated by

translocation-related artefacts (as are methods based on interpolated pairwise genetic distances or

240   F_{ST}). Patterns based on Simpson's diversity were contrasted with those inferred using a form of

2D-stepping stone model (EEMS; Petkova et al. 2015), as run with 2 million MCMC iterations

242   (1 million as burn-in), and sampled every 1,000 iterations [following parameter sweeps that

tuned MCMC acceptance rates to fall between 20% and 50% (Roberts, Gelman, & Gilks, 1997)].

2.3 | Estimating cline parameters

246   We examined the nature of population boundaries by examining how individual loci transitioned

across these regions, relative to the genome-wide average. To do so, we defined eight transects

248   across putative genetic edges, sampling 32-73 individuals per transect ($\bar{x} = 55.38$). Individuals

were chosen to represent each sub-population using a probability threshold applied to

250   ADMIXTURE results. Loci were filtered to remove all SNPs with missing data in >50% of

individuals, and computational time was further restricted by retaining only those loci with a

252   sufficient allele frequency differential among clinal extremes [computed as $\delta > 0.50$; (Gregorius

& Roberds, 1986)]. Locus-wise clinal patterns were then inferred using a Bayesian method

254 developed originally for hybrid zones (BGC: Gompert & Buerkle, 2011, 2012). Open-source

Python code for filtering SNP matrices and generating necessary input are available at:

256 (*github.com/tkchafin/scripts/phylip2introgress.pl* and *phylip2bgc.pl*).

Analyses were performed for each transect across four replicates, each using one million

258 MCMC iterations, discarding the first 500,000 as burn-in, with output thinned to every 500

iterations. Results were summarized by visualizing a 2-D density of cline shape parameters.

260 These are: α (=cline directionality) that describes an increase ($\alpha > 0$) or decrease ($\alpha < 0$) in the

probability of locus-specific ancestry from a parental population; and β (=cline width/ steepness)

262 that defines the rate of transition in probabilities of locus-specific ancestries having either steep

($\beta > 0$) or wide ($\beta < 0$) shapes (Gompert, Parchman, & Buerkle, 2012). In this context, a locus

264 which does not deviate from the genome-wide pattern would have $\alpha=\beta=0$. Deviation in a

directional manner (i.e., an increase or decrease of one ancestry over another) is described by α,

266 whereas deviation in the rate of ancestry change around an inflection point (i.e., sigmoidal) is

described by β. Statistical outliers were designated using the method of Gompert and Buerkle

268 (2011). BGC results were parsed and visualized using the ClinePlotR R package (Martin, Chafin,

Douglas, & Douglas, 2020).

2.4 | Relative dispersal by age and sex

272 Of particular interest in wildlife management is the backwards inference of geographic

positioning from genotypes—that is, the geolocation of 'origination' points for sampled animals.

274 This could be used, for example, to ascertain the geographic origin of poached individuals, or to

estimate post-natal dispersal. To this end, we used the novel, deep-learning method LOCATOR

276 (Battey et al., 2020a) to predict the geographic origin of samples without relying upon explicit

assumptions about population genetic processes underlying spatial genetic differentiation

278   (Bradburd & Ralph, 2019). The analysis was performed iteratively across each individual, using

the remaining samples to train the LOCATOR classifier, with 100 bootstrap pseudo-replicates to

280   assess variance in geolocation. Given computational constraints, we performed the analysis using

a subset of 5000 SNPs having a minor-allele frequency >10%.

282   We estimated relative dispersal distances as the Euclidean distance between sampled

localities and the centroid of predicted coordinates, under the assumption that the distance

284   between predicted and collected locations is the result of lifetime dispersal, at least for samples

for which geolocation variance is low among pseudo-replicates. The results were then partitioned

286   by age, sex, and CWD-status.

Our second approach examined the decay in genetic relatedness as a function of distance

288   from each individual, measured as a Prevosti distance (R-package poppr: Kamvar et al., 2014).

Here, the assumption is that recently dispersed individuals will be, on average, more genetically

290   dissimilar from resident individuals, whereas resident individuals having an appreciable

reproductive output will be less so. These calculations were limited to individuals that had

292   neighboring samples within a 5km radius, thus removing individuals from sparsely sampled

regions where the relationship between relatedness and distance would be unreliable. We also

294   note that the traditional aging method employed herein (Severinghaus 1949) has an accuracy

seemingly reduced in older deer, potentially suggesting caution in the interpretation of results

296   (Cook & Hart, 1979; Gee, Holman, Causey, Rossi, & Armstrong, 2002; Mitchell & Smith,

1991).

3 | RESULTS

300 **3.1 | Data processing**

Our raw data represented N=1,143 samples, including N=29 from Wisconsin (Table S1). We

302 removed N=83 that had missing metadata, discrepancies with coordinates, or <50% of loci

present. Assembly in pyRAD yielded an average of 25,584 loci per sample (σ=8,639). After

304 removing loci present in <50% of samples and excluding those containing potential paralogs

(e.g., excessive heterozygosity or >2 alleles per locus), our final dataset contained 35,420 loci,

306 from which 2,655,584 SNPs were catalogued. Of these, 54,102 were excluded as singletons. To

limit signal redundancy, we then condensed the data to one SNP per locus, yielded a final matrix

308 of 35,099 SNPs for analyses of population structure.

310 **3.2 | Population structure and 'ancestry surfaces'**

Cross-validation, performed on N=20 replicates each for subpopulation model (K=1-20),

312 revealed the optimal number of clusters as K=8. Spatial orientation of these samples (Fig. 1)

provided a geographic definition, with some subpopulations qualitatively defined by apparent

314 landscape features, such as the Arkansas River Valley as the southern extent of subpopulations

k=3 and k=6.

316   Two ancestries (k1 and k8) largely dominated south of the Arkansas River, bounded by

Interstate 30 to the north and the Ouachita River to the south (Figs. 1-2), each of which supports

318 the argument that genetic structure is defined by large geographic barriers. The southwestern

portion of the state has two ancestral assignments (k1 and k3), with the latter having mixed

320 representation in the north-central section (potentially an artefact of weak differentiation rather

than true shared ancestry). The southeastern section is dominated by a single gene pool (k8),

322 which coincidentally subsumed all Wisconsin samples, a strong signature of genetic variability

as a residual of historic translocations.

324   A greater amount of locally endemic structure occurred north of the Arkansas River in

the Ozark Mountains, where six sub-populations were evident. The most broadly distributed ($k5$)

326 was to the east in the Mississippi alluvial plains, extending westward across the mainstem of the

White River then northward towards the confluence of the Black and White rivers, where it

328 grades into several distinct yet loosely defined sub-populations (Figs. 1-2). The northwestern

corner of the state was the most heterogeneous, with four primarily endemic gene pools (k=2,3,4

330 and 7; Fig. 2). The northern-most of these was approximately bounded by the White River (Fig.

2) and graded westward into an area with high levels of mixed assignment (Fig. 1). The

332 remaining northwestern region was defined by several gene pools showing spatially weak

transitions, suggesting reduced gene flow but with geographic and/ or environmental boundaries

334 reasonably porous.

    Effective migration surfaces (EEMS) failed to capture any discernable pattern relating to

336 spatially defined population structure (Fig. S1). Geographic breaks separating sub-populations

(=genetic edges) were captured instead by reducing interpolated assignment probabilities ($n = k$)

338 as a continuous Simpson's diversity index (n=1) (Fig. 3). This, in turn, reflects a dependence on

homogeneity of local assignments, rather than global patterns compounded by long-distance

340 transplants.

342 **3.3 | Intraspecific genomic clines**

Genomic clines varied substantially among transects (Figs. 4, S2). Most inter-population

344 comparisons within northwest Arkansas, to include $k2$ x $k3$ (both eastern and southern transition

zones), $k2$ x $k6$, and $k4$ x $k7$, indicated variation primarily restricted to cline directionality (α),

346   with cline steepness (β) at a minimum. The variation in locus-wise pattern for these cases

(hereafter termed 'α-dominant'), indicated a directional change in the representation of reference

348   populations across the transect, but without a noticeable 'inflection' point, as implied by non-

zero cline width (β).

350   Remaining comparisons, including one additional transect from northwest Arkansas ($k3$ x

$k4$), primarily reflected variation in cline width (β), in that loci varied most prominently

352   regarding steepness of transition around an inflection point respective to genome-wide ancestry.

Although two transects showed minor exception (i.e., $k3$ x $k6$ varied along both parameters; $k4$ x

354    $k7$ each showed minimal variation), the contrasting variation between α-dominant versus β-

dominant transects suggests that different processes underly ancestry transitions.

3.4 | Estimating dispersal using geolocation analysis

358   Individuals from densely sampled regions could be assigned to geographic origin using the

'deep-learning' approach, with CWD-positive individuals assigned with a mean bootstrap

360   distance from centroid prediction generally <15km (Fig. S3). However, we did find assignment

error (e.g., among bootstraps) was elevated in low-density sampling regions (Figs. 5, S4), which

362   resulted in higher estimated individual dispersal distances (Fig. S4A, B). Given that variance in

dispersal estimates dropped considerably below ~25km (Fig. S4C), a conservative threshold of

364   10km was chosen and all individuals having a mean bootstrap-centroid distance above that were

removed for the purposes of dispersal estimates. After filtering, N=110 samples remained (Fig.

366   5A). A higher error threshold (i.e., 20km) allowed a greater number of total samples to be

incorporated (N=264; Fig. 5B), with several 'roadkills' appearing as long-distance transfers.

368 Although these results are low-precision assignments, they do underscore the capacity of the

method regarding the identification of transported individuals (e.g., poached or illegally dumped

370 deer, or carcasses transported across state lines or regional management zones).

Geo-located results for individuals passing a strict error filter demonstrated a dispersal

372 distance for males approximately double that of females across all age classes (statistically

significant only for the Y2-2.5 class due to low sample sizes; Fig. 6). This pattern was

374 established as early as the Y1-1.5 group, indicating apparent male dispersal by that age. Smaller

dispersal distances were found for fawns across both sexes (Fig. 6), again corroborating *a priori*

376 biological expectations.

Patterns of genetic dissimilarity also showed an age x sex effect, with greater genetic

378 dissimilarity for neighboring Y0-1 males than females, and with a shift towards reduced

dissimilarity in males >5 (Fig. S5). This again supports the argument that male deer in Arkansas

380 have dispersed by the Y1-1.5 age class, and potentially reflects age-biased reproduction. Males

contributed disproportionately to their local gene pools by age 5 (i.e., producing offspring with

382 resident females), thereby creating a pattern of lower genetic dissimilarity among neighboring

individuals, regardless of distance.

4 | DISCUSSION

386 **4.1 | Genetic footprints of historic management**

The prolonged history of hunter-harvest, and subsequent long-range translocations into Arkansas

388 and the surrounding region (Ellsworth et al., 1994; Holder, 1951), provide a sideboard to our

study in that artificial long-distance movements such as these violate the assumptions inherent

390 with many spatially-explicit methods. However, we found estimates of ancestry diversity and

probability surfaces to be qualitatively robust in that they firmly recapitulated the record of

392   historic translocations (Holder, 1951). This, in turn, necessitates that our results be placed within

an historic context.

394   In the early 20th century, following many decades of over-hunting, the Arkansas deer

population declined to <500 individuals (Holder, 1951). In response, Arkansas Game and Fish

396   Commission (AGFC) implemented an extensive restocking program (1941-1951) involving as its

basis three primary sources. The first was Howard County deer farm (southwestern Arkansas;

398   Fig. 7), established from locally transplanted central Arkansas individuals (Wynn, 1943), and

now located within the epicenter of population $k1$ (Figs. 2, 7). Its ancestry is shared elsewhere in

400   the state, thus establishing it as an epicenter for local translocations.

A second major source was the Sylamore District in the Ozark Mountains (Wood, 1944;

402   Holder, 1951), where individuals were naturally abundant (Fig. S7). Individuals from this region

have a consistently higher probability of assignment to $k3$ (Figs. 2, 7), as records indicate ~81%

404   of repopulated individuals in the Gulf Coastal Plain originated from there (Holder, 1951; Karlin,

Heidt, & Sugg, 1989; Wood, 1944). Our analyses agree in that many individuals from

406   southwestern Arkansas (Figs. 1, 6) reflect a mixed assignment to the 'Sylamore' cluster. They

also comprised ~36% of stocking efforts in the Mississippi Delta region (Karlin et al., 1989).

408   Our results also indicate some mixed assignment of individuals to $k3$, although a more

widespread representation is found in $k5$. AGFC surveys (~1942-1945) indicated deer as being

410   relatively abundant in the southeastern Delta region (Figs. 7, S7), seemingly linked with re-

population efforts following the disastrous 1927 Mississippi flood. This event, coupled with

412   over-hunting, nearly extirpated deer in the region, save small numbers sustained by local

sportsmen (Holder, 1951; Fig. S7). We hypothesize deer in the Delta region largely result from

414   those efforts, given the contemporary homogeneity of ancestry assignments in this region.

A third (extraneous) source was the Sandhill Game Farm (Babcock, WI; Wood 1944).

416   Records indicate that ~64% of deer released into the Mississippi Delta region originated out-of-

state, the majority from Wisconsin (Holder, 1951; Karlin et al., 1989). Our Wisconsin samples

418   were unanimously assigned to a gene pool prominently represented in the southern Delta region

(k8; Fig. 2), firmly establishing its genetic legacy as extending from imported deer.

420   Mitochondrial haplotypes putatively originating from Wisconsin were also uncovered in

Missouri, Kentucky, and Mississippi (Budd, Berkman, Anderson, Koppelman, & Eggert, 2018;

422   DeYoung et al., 2003; Doerner et al., 2005).

All genetic 'clusters' that lacked spatial cohesion in our analyses can be connected to the

424   three major stocking sources involved in earlier restoration efforts (per historic records). The

remaining sub-populations, primarily in the Ozarks (k2, 4, 6, and 7) may represent natural re-

426   colonization from refugial populations (Figs. 2, 7), an hypothesis supported by early census data

(Figs. S7, S8). If this was not recognized, then any analysis of deer genetic structure across

428   Arkansas would instead infer unusually high rates of gene flow among these geographically

distant regions. This, in turn, would preclude any analysis aimed at linking variation in genetic

430   structure to environmental factors.

432   **4.2 | Contemporary genetic structure**

A primary focus in our study was the contemporary genetic structure of Arkansas deer,

434   especially regarding how these data promote an understanding of landscape resistance to

dispersal (Hemming-Schroeder, Lo, Salazar, Puente, & Yan, 2018; Kelly et al., 2014). Yet

436   several prerequisites are apparent in this regard. For example, one apparent question previously

constrained by technological limitations is the degree to which potential patterns of genetic

438   variability in deer have been conflated by anthropogenically-mediated translocations.

We addressed this issue by utilizing next-generation molecular techniques to derive

440   highly variable markers across the genome. We also implemented advanced geospatial

procedures that visualized the spatial transitions inherent within deer ancestry in such a way that

442   translocation histories were not a limiting factor. We did so by interpolating our assignment

probabilities from ADMIXTURE, then applying Simpson's diversity index as a means of

444   reducing the probabilities to a single vector. We also polarized stochastic *versus* deterministic

processes by examining locus-specific patterns within our transitions.

446   Previous studies concluded that inferences from landscape genetic analyses were limited

at best, due to complex interactions between historic translocations and subsequent population

448   growth (Leberg et al. 1994; Leberg and Ellsworth 1999; Budd et al. 2018). However, many of

these studies relied upon legacy molecular markers (e.g., mtDNA or reduced panels of

450   microsatellite DNA markers) that capture substantially less polymorphism than do next-

generation methods (Hodel et al., 2017; Jeffries et al., 2016; Lemopoulos et al., 2019). Recent

452   studies at more refined spatial scales have supported large-scale geographic barriers (e.g. rivers,

highways) as being semipermeable to gene flow (Kelly et al., 2014; Locher et al., 2015; Miller et

454   al., 2020; Robinson et al., 2012). This, despite the potential complication caused by re-stocking

efforts, and in accordance with radio-telemetry data (Peterson et al., 2017).

456   Several scale-dependent issues may modulate the degree to which translocations

obfuscate landscape genetic patterns. First, as spatial scale increases, so does the probability of

458   sampling individuals whose genetic dissimilarity reflects translocation rather than natural gene

flow via dispersal. Second, the most obvious environmental effects that drive deer dispersal have

460   the greatest probability of being recovered in an analysis, even if they represent spurious signals

driven by translocation. However, landscape-level analyses of dispersal can still be informative,

462   despite those rapid transitions between historic and contemporary conditions (Epps &

Keyghobadi, 2015) that may obfuscate relationships. For example, large-scale environmental

464   features also drive spatial genetic patterns in invasive species, despite reduced levels of genetic

diversity (Lopez, Hurwood, Dryden, & Fuller, 2014; Sacks, Brazeal, & Lewis, 2016; Zalewski,

466   Piertney, Zalewska, & Lambin, 2009). Such patterns are also rapidly manifested within

populations occupying habitats recently modified by anthropogenic developments, such as large

468   urban centers (Beninde et al., 2016; Combs, Puckett, Richardson, Mims, & Munshi-South, 2018;

Kimmig et al., 2020).

470   The analytical artefacts derived from translocation are expected to be not only scale-

dependent but also a function of specific analytical assumptions. For example, classic tests of

472   'isolation-by-distance'(IBD) implicitly assume a negative relationship between geographic

proximity and pairwise patterns of genetic distance (Meirmans, 2012; Rousset, 1997). Yet,

474   translocations are demonstrably inconsistent with this expectation. Introduced individuals reflect

levels of genetic similarity that are demonstrably inconsistent with spatial and/ or environmental

476   distance (i.e., 'resistance') that now separates them. This, in turn, incorporates spurious signals

within the landscape inferences subsequently derived.

478   The assignment method used in our study yielded patterns of genetic similarity that

reflected translocations as depicted in historic records (Fig. 7). We sought to more appropriately

480   expose landscape features that could potentially modulate deer dispersal by effectively

'removing' this artificial signal. Our approach (i.e., reducing the dimensionality of assignment

482   probabilities) produced relatively straightforward predictions about the ensuing metric: Areas

with uninterrupted movement of individuals should have increased homogeneity regarding

484   interpolated ancestry assignments. By the same logic, regions representing barriers to movement

among regions should likewise demarcate transitions in interpolated ancestries. We acknowledge

486   that spatial assignments can also be vulnerable to artefacts, such as an over-fitting of discrete

clusters within otherwise continuous populations (Bradburd et al., 2018). Yet despite this, they

488   are demonstrably effective at identifying linear barriers to gene flow (Blair et al., 2012).

Coincidentally, our approach using summarized ancestries revealed numerous linear subdivisions

490   that aligned with major landscape barriers, such as rivers in Arkansas (Fig. 3). Other recognized

transitions approximately corresponded with large urban centers. Our argument is that

492   approaches involving an assumption of stepwise gene flow (implicit with pairwise genetic

distances) may generate conclusions that either lack the necessary nuances or are altogether

494   incorrect. This is apparent when the 'surfaces' inferred using Simpson's diversity of ancestry

assignments (Fig. 3) are contrasted with EEMS results (Fig. S1).

4.3 | Separating stochastic *versus* deterministic processes in clinal patterns

498   When deriving clinal patterns, it becomes difficult to disentangle contemporary demographic

processes from analytical idiosyncrasies. As a means of clarification, we generated multi-locus

500   genomic patterns under the expectation that the relative importance of stochastic *versus*

deterministic processes could be derived from the manner by which genome-wide averages

502   deviated from locus-wise clines (Barton, 1983; Barton & Hewitt, 1985; Slatkin, 1973; Vasemägi,

2006). We found eight transects that varied either in the steepness of genetic transitions (β), or

504   the directionality of allelic frequency change (α). Interestingly, 'α-dominant' transects were

found within the Ozark region, previously hypothesized as being naturally re-colonized from

506   local refugia (i.e., $k2$ x $k3$E/S, $k2$ x $k6$, and $k4$ x $k7$; Figs. 4, S2).

Under this hypothesis, genetic drift along the leading edge of population expansions

508   would generate two genetic reverberations: The random, directional fixation of alleles, and an

apparent inflation of population structure. Both were observed in our data (Fig. 1, 4). The

510   hypothesis is also supported by state-wide surveys (1940's) that found deer in habitat associated

with each Ozark sub-population during, and immediately following, the low ebb of deer in

512   Arkansas (Figs. 7, S8). The transects likewise displayed near-zero β values, suggesting the

absence of a 'true' barrier to dispersal. Here, the expectation would be that a hard barrier

514   between discrete populations would reflect changes in allele frequencies around an inflection

point.

516   One interpretation is that genetic drift at the edge of an expanding refugium creates a

'rolling' founder effect, with alleles either over- ($\alpha>0$) or under-represented ($\alpha<0$) across the re-

518   colonized territory (Excoffier & Ray, 2008; Hallatschek & Nelson, 2008). This phenomenon—

termed 'gene surfing'—can generate wholly neutral patterns that appear to be adaptive (Peischl,

520   Dupanloup, Bosshard, & Excoffier, 2016; Travis et al., 2007), and may explain the presence of

numerous α-outlier loci in those transects (Fig. 4).

522   Interestingly, all our β-dominant transects crossed major rivers: Transect $k3$ x $k6$ spanned

the Arkansas River; $k3$ x $k5$ the White River; and $k1$ x $k8$ the Ouachita River. Transect $k3$ x $k4$

524   crosses the smaller and shallower Buffalo River, which may explain the depressed variability in

cline steepness (Fig. 4). The coincidence of these patterns with major rivers suggests barriers to

526   individual movement, in that non-zero β values implicate variability in the 'steepness' of clinal

transitions around an inflection point [i.e., with rates of change in allele frequencies either steep

528   ($\beta>0$) or wide ($\beta<0$)].

Classically, clinal patterns are established either by selection or hard boundaries to gene

530   flow (Endler, 1973; Slatkin, 1973), but our analyses (Fig. 4) revealed little evidence of selection-

driven outliers. However, spatially variable gene flow, when coupled with drift, can yield clinal

532   patterns across a considerable proportion of the genome without necessitating selection

(Vasemägi, 2006). Selection could feasibly be involved, at least for regions where different deer

534   subspecies were translocated, in that subspecific crosses have elicited fitness impacts such as

dystocia (abnomal maternal labor due to shape, size, or position of the fetus; Galindo-Leal and

536   Weber 1994). The genetic costs of inter-subspecific stocking are recognized (Hopken, Lum,

Meyers, & Piaggio, 2015), and manifested as an anomalous variation in breeding time among

538   other southern-recovered deer populations (Sumners et al., 2015). Some limited evidence also for

reproductive isolation among mule deer subspecies (*Odocoileus hemionus*) as driven by

540   pheromonal differences (Müller-Schwarze & Müller-Schwarze, 1975).

542   **4.4 | Management implications**

Our results identified a diagnosable genetic signature of historic translocations within the

544   genomes of extant Arkansas deer populations. This simultaneously underscores the success of

early restocking efforts, while also reiterating long-standing concerns about the genetic and/ or

546   phenotypic impacts of anthropogenic translocations (Meffe & Vrijenhoek, 1988). It also

reinforces the need to formalize as an explicit management consideration the intraspecific

548   taxonomy of deer, both to better understand the (yet to be seen) evolutionary consequences of

past translocations, and to provide a more proactive baseline going forward (Cronin, 2003;

550   Gippoliti, Cotterill, Groves, & Zinner, 2018).

Resource agencies, in particular, should prioritize those populations that retain endemic

552   genetic diversity, and by so doing, acknowledge the importance of preserving evolutionary

legacy (Crandall, Bininda-Emonds, Mace, & Wayne, 2000). Such recognition would capitalize

554   on the many years of natural selection that have operated on those populations prior to

anthropogenic interference, which may have generated adaptive genetic variation necessary to

556   withstand changing conditions, or mediated an evolutionary response to disease (e.g., in the

PRNP gene; Chafin et al., 2020). This further emphasizes the importance of integrating genomic

558   methods for heavily managed species (Flanagan et al., 2017), given their recognized lack of

intraspecific adaptive variation [i.e., a 'Darwinian shortfall' in biodiversity conservation (Diniz-

560   Filho, Loyola, Raia, Mooers, & Bini, 2013)].

It is noteworthy that the current boundaries for the AGFC 'Deer Management Units'

562   (DMUs) (Meeker et al., 2019) are remarkably consistent with many of the population boundaries

identified in our study. This validates the continued use of these ecosystem-based management

564   units for the application of locally-appropriate management and harvest regimes.

Another valuable application of our results would be disease containment, particularly

566   given that CWD is a significant challenge globally for the management of cervids (Rivera,

Brandt, Novakofski, & Mateus-Pinilla, 2019; Uehlinger, Johnston, Bollinger, & Waldner, 2016).

568   The spread of CWD will be more rapid within rather than among populations, and the

cumulative risk of disease spread to new populations will depend on both dispersal and disease

570   prevalence. Thus, potential mitigation efforts should work to constrain disease transmission

within CWD-affected areas as well as limiting the outward dispersal of individuals from those

572   areas. One approach to reduce dispersal is to reduce the density of yearling males, which often

comprise >50% of emigrating individuals (McCoy, Hewitt, & Bryant, 2005; Nelson, 1993;

574   Nixon, Hansen, Brewer, & Chelsvig, 1991). Our patterns of age- and sex-biased dispersal

distances (Fig. 5, 6) largely recapitulate these expectations at a local level. Removal of young

576   males to slow dispersal also may be compatible with the application of male-focused harvest

strategies predicted to reduce disease prevalence (Potapov, Merrill, Pybus, & Lewis 2016). A

578   similar action would be to reduce local population densities that serve to drive dispersal, in that

robust population numbers elevate social pressure as well as resource competition (Long et al.,

580   2008; Lutz et al., 2015; Shaw, Lancia, Conner, & Rosenberry, 2006). Other harvest strategies

have been implemented that effectively decrease emigration while focusing on meeting specific

582   demography goals including: actively decreasing overall population density, elevating the

proportion of mature males in the population, and increasing the female component of the sex

584   ratio (Brothers and Ray, 1975; Hamilton et al., 1995; Shaw et al., 2006; Long et al., 2008;

McCoy et al., 2005; Lutz et al., 2015). However, mature male deer tend to demonstrate higher

586   CWD prevalence than other demographic groups, and management strategies that maximize this

component of the population may exacerbate disease dynamics (Miller, Runge, Holland, &

588   Eckert 2020). Additionally, our findings indicate that disease surveillance strategies should

include efforts that focus on detecting 'breaches' between genetically distinct regions (those

590   bordered by major rivers and urban centers, as herein; Fig. 3) to further direct management and

response efforts.

4.5 | Utility of genomic approaches for white-tailed deer management

594   Beyond providing a nuanced view of historical management processes, the genomic methods

used herein offer several improvements beyond 'legacy' approaches. For example, they provide

596   vastly increased statistical power for individual-level genetic analyses, such as estimating

relatedness (Lemopoulos et al., 2019), or parentage assignment (Flanagan & Jones, 2018). This

598   is also accomplished at a fraction of the cost-per-marker (Puckett, 2017). Evidence for this is

found in our geolocation results (Fig. 5, S3-S4), where a lower dispersion of predicted natal

600   locations were achieved when compared with previous attempts using microsatellites (Green et

al., 2014). The genome is also more densely sampled and these data can be used in conjunction

602   with the analytical methods used herein to infer both environmental associations and adaptive

genetic variation (Martin et al., 2020), or detect anthropogenically-mediated change through time

604   (Chafin et al., 2019).

SNP-based methods have a particular advantage with regards to scalability (i.e., such as

606   in a multi-state consortium of wildlife management agencies). Large datasets (as herein) may

allow a subset of highly consistent, maximally informative markers to be extracted that are

608   broadly applicable across states and regions. In ddRAD (and other RADseq methods), much

sequencing effort may be wasted on low-occupancy or systematically-missing loci, a result of

610   variability introduced during: Library preparation (Escudero et al., 2014); Bioinformatic

processing (Eaton, 2014); Mutation-disruption (Rubin et al., 2012), or; Insufficient sequencing

612   coverage (Huang & Knowles, 2016). Additionally, many loci may be monomorphic, contain

low-frequency mutations, or be uninformative at the scale examined. Methods to optimize cost-

614   efficiency can be applied to consistently assay target loci (e.g., Chafin et al., 2018) by using

targeted-enrichment approaches [e.g., Rapture (Hoffberg et al., 2016)] or those amplicon-based

616   [GT-Seq (Campbell, Harmon, & Narum, 2014)], thus reducing the total costs to <$10 per

sample.

618   Selecting an appropriate approach depends upon desired throughput capacity, yet all

share the need for an initial database of genetic variation from which to assays can be generated.

620   For white-tailed deer, our approach would first be applied to generate a broad geographic

reference database (i.e., at a national or multi-regional scale), followed by assay design for a *post*

622 *hoc* method (such as that of Campbell, Harmon, & Narum, 2014). This would allow our

methodology to be deployed so as to generate a sampling sufficiently dense for management-

624 oriented population genetic analyses such as natal geolocation. This, in turn, would facilitate the

detection of illegal carcasses, deer trafficking, migrant detection, or characterization of historic

626 translocation (as herein). Moreover, such an approach would be fully replicable across

laboratories and agencies, thereby facilitating inter-agency and multi-state collaboration.

628   Although dense spatial and genomic sampling may offer a composite resolution (e.g.,

Chafin et al., 2020; Linck et al., 2019), it does not necessarily supplant other approaches. For

630 example, sampling density was shown herein to be a particular limiting factor in geolocation

analyses, with high prediction error in regions of the state sparsely sampled (Fig. 5). Although

632 we did not directly compare our methodology with that based on a microsatellite dataset, we

suspect greater error in the latter given decreased information content associated with the much

634 reduced dataset. Genetic input from unsampled populations (e.g., long-distance migrants from

out-of-state) found at edges of our sampling areas could also inflate assignment uncertainty (see

636 Fig. 3), as could extremely uneven sampling (Lawson, van Dorp, & Falush, 2018). We are also

unsure of the degree to which long-distance migrants could influence our clinal analyses (e.g.,

638 via a 'smoothing' effect; Smith and Weissman 2020), nor fluctuation that may occur over brief

time scales (as observed in other study systems; Chafin et al., 2019). Thus, we caution that the

640 deployment of these methods at multi-state or regional scales should focus on dense, spatially

expansive sampling that employs both spatial and temporal replicates, where possible (Short Bull

642 et al., 2011). Both priorities are facilitated via the aforementioned enrichment methods that

drastically reduce per-sample cost-efficiency.

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AUTHOR CONTRIBUTIONS

All authors contributed to study design and conceptualization. MCG, CRM, and JRB coordinated
sample collection and curation of samples; TKC, BTM, ZDZ, and MRD contributed to molecular
work; TKC, ZDZ, MRD, and MED planned the analysis approach; TKC and BTM wrote code

for analysis and contributed exploratory analyses; TKC performed bioinformatic work,

668   assembled data, and performed analyses; TKC and ZDZ prepared figures; TKC and ZDZ drafted

the manuscript and all authors revised the manuscript and contributed to the final version.

DATA AVAILABILITY STATEMENT

672   Raw sequences are accessioned in the NCBI Sequence Read Archive (SRA) under BioProject

PRJNA690954; Assembled SNP data and sample metadata is available via the Open Science

674   Framework (doi: 10.17605/OSF.IO/T82RV); codes and custom scripts developed in support of

this work are also available as open-source via GitHub under the GNU Public License:

676   github.com/tkchafin (and as cited in-text).

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Figures and Tables

Figure 1: **Ancestry proportions for N=1,143 Arkansas white-tailed deer**, as inferred using the program ADMIXTURE applied to 35,099 single-nucleotide polymorphisms (SNPs) generated via ddRAD-seq. Samples are represented as pie charts plotted at absolute collection coordinates, with colors of assignment probabilities proportional to a particular subpopulation.

Figure 2: **Assignment probabilities for eight populations of Arkansas white-tailed deer (k=1 through k=8)**, as inferred from ADMIXTURE applied to 35,099 SNPs, interpolated using Empirical Bayesian Kriging of assignments for N=1,143 individuals. P(k)=1.0 corresponds to 100% probability of ancestry per raster cell, and P(k)=0.0 corresponds to 0%. Individual samples are represented as black dots.

Figure 3: Simpson's diversity of interpolated ancestry proportions for N=1,143 Arkansas white-tailed deer. Each raster cell was assigned eight values equal to the expected proportion of ancestry corresponding to the eight genetic clusters (Fig. 1), then summarized using Simpson's diversity index.

Figure 4: Relationships between genomic cline parameters contrasted among eight sub-sampled transects of individuals spanning population boundaries in Arkansas white-tailed deer (Fig. 3). Contour plots show relative densities of SNPs varying in cline steepness (β), representing the gradient of the clines, and cline directionality (α), representing bias in SNP ancestry. Outlier loci are highlighted in blue.

Figure 5: Summary of geo-location predictions in LOCATOR as a random subset of N=490

Arkansas white-tailed deer at two error thresholds: 10km (=A) and 20km (=B), thus constraining

results to N=110 (=A) and N=264 (=B). Error calculated as mean distance of bootstrap

predictions from the centroid predicted for each individual (see Fig. S4). A black dot denotes the

predicted location of an individual and a colored dot indicates the 'true' location (colors

proportional to measurement error in km).

**Figure 6: Inferred dispersal distances, partitioned by sex and age class for N=1,143

Arkansas white-tailed deer**. Dispersal distances were calculated as the difference (in km)

between the predicted and true locations, excluding all individuals with a mean prediction error

of 10km (see Figs. 14-15). Sample sizes for each group are given in black, with the mean in red

below each box plot. Within-age, p-values are reported for two-sample t-tests comparing means

from males and females.

**Figure 7: Proposed translocation pathways and refugial populations for Arkansas white-

tailed deer**, showing (left) translocation events supported by genetics along (dashed line), as

well as those with anecdotal support given limited historical records (solid), from three primary

stocking sources: Sandhill Game Farm in Wisconsin (W); 'Sylamore District' state refuge sites

(S); and the Howard County Game Farm (H). Also shown are major inhabited regions of white-

tailed deer (right) estimated from 1942-47 game surveys, excluding small regions (see full

version in Online Supplementary Material). Putative inhabited regions are also annotated

according to their hypothesized associations with contemporary genetic clusters (Figs. 1-2).

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Assignment probability

0.00 - 0.20	0.21 - 0.40	0.41 - 0.60	0.61 - 0.80	0.81 - 1.00

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 Supplementary Figures and Tables

1072 **Table S1:** Numbers (=N) of Arkansas white-tailed deer genotyped across ~35,000 SNP loci.
1073 Samples were collected from 75 counties (=County) from 2016-2019. Code indicates standard 2-
1074 letter county abbreviation. Samples removed due to missing data are not included.
1075

County	Code	N	County	Code	N
Arkansas	AR	11	Lee	LE	9
Ashley	AS	1	Lincoln	LI	8
Baxter	BA	19	Little River	LR	10
Benton	BE	19	Logan	LO	32
Boone	BO	34	Lonoke	LN	11
Bradley	BR	2	Madison	MA	39
Calhoun	CA	14	Marion	MR	39
Carroll	CR	50	Miller	MI	8
Chicot	CH	3	Mississippi	MS	2
Clark	CL	7	Monroe	MO	10
Clay	CY	15	Montgomery	MN	2
Cleburne	CE	8	Nevada	NE	13
Cleveland	CV	5	Newton	NW	140
Columbia	CO	4	Ouachita	OU	12
Conway	CN	11	Perry	PE	8
Craighead	CG	7	Phillips	PH	9
Crawford	CW	6	Pike	PI	8
Crittenden	CT	8	Poinsett	PO	10
Cross	CS	10	Polk	PL	3
Dallas	DA	7	Pope	PP	62
Desha	DE	11	Prairie	PR	10
Drew	DR	14	Pulaski	PU	11
Faulkner	FA	14	Randolph	RA	8
Franklin	FR	20	Saline	SA	11
Fulton	FU	9	Scott	SC	2
Garland	GA	14	Searcy	SE	32
Grant	GR	5	Sebastian	SB	17
Greene	GE	16	Sevier	SV	12
Hempstead	HE	9	Sharp	SH	16
Hot Spring	HS	3	St. Francis	SF	7
Howard	HO	11	Stone	ST	15
Independence	IN	7	Union	UN	6
Izard	IZ	6	Van Buren	VB	23
Jackson	JA	2	Washington	WA	17
Jefferson	JE	13	White	WH	10
Johnson	JO	45	Woodruff	WO	9
Lafayette	LA	14	Yell	YE	31
Lawrence	LW	17	**Total**		**1,143**

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Figure S1: **Effective migration rates and intra-population diversity** (log$_{10}$ scale) for Arkansas white-tailed deer, calculated from effective migration surfaces (EEMS). Rates are plotted according to colored bin, with divisions calculated as natural breaks using the Jenks algorithm in ArcMAP.

1084   **Figure S2: Sampled transects employed for genomic cline analysis of Arkansas white-tailed deer,** with selected groups within
each coalesced within colored hulls: Reference population 1 (p1; red); Reference population 2 (p2; green); and putative admixed
1086   individuals (popA; blue).

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Figure S3: LOCATOR Geo-located results for CWD-positive Arkansas white-tailed deer derived via LOCATOR. Individuals are represented by dots, pairs of which are connected by a single line. A black dot denotes predicted location, while a colored dot indicates 'true' (observed = sampled) location (color proportional to distances separating each). Inserted figures offer full prediction results for three selected samples, with bootstrap estimates demarked by a 95% contour.

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Figure S4: Effects of error thresholds on inferred dispersal distances for Arkansas white-tailed deer, as predicted by LOCATOR analysis. Inferred dispersal kernels were computed at several error thresholds: (A) Error computed as mean distance of predicted coordinates for each individual at 100 bootstrap estimates from predicted centroid (=interpretation of localized vs. dispersed predictions). Dispersal distances calculated as difference (in km) between predicted centroid versus 'true' (observed) location. Inferred dispersal distances generally increase for individuals with larger prediction error; (B) Larger stochastic variation in centroid location for individuals with lower predicted precision among bootstraps; (C) Standard deviation of dispersal distances for the latter computed in a sliding window of 10km along the x-axis (prediction error).

Figure S5: Spatial patterns of genetic dissimilarity among Arkansas white-tailed deer partitioned by age and sex. Genetic distances between individuals and neighbors derived from 5,000 randomly sampled SNPs depicted across physical distance (x-axis). Results depict different cohorts of males (A) and females (B) by age.

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1112   **Figure S6**: Census estimates for Arkansas white-tailed deer from surveys in 1942-1946.

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1128   **Figure S7: Subpopulation probabilities overlain with 'deer occupied territories' in Arkansa circa 1942-1947** (in green). Polygons representing deer occupation were manually compiled
1130   from the Wildlife and Cover Map of Arkansas (1942-47), prepared by Arkansas Game and Fish Commission in cooperation with U. S. Fish and Wildlife Service as one aspect of a Federal Aid
1132   Project (drawn by Flaun M. Tolar). Map accessed from the University of Arkansas Library Arkansas Collection (Special Collections). A higher resolution geo-referenced version available
1134   at a later date (reproduction/ copy permissions granted by University of Arkansas Libraries).

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