1	Model Adequacy Tests for Likelihood Models of
2	Chromosome-Number Evolution

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15 Summary

16	•	Chromosome number is a central feature of eukaryote genomes. Deciphering
17		patterns of chromosome-number change along a phylogeny is central to the
18		inference of whole genome duplications and ancestral chromosome numbers.
19		ChromEvol is a probabilistic inference tool that allows the evaluation of several
20		models of chromosome-number evolution and their fit to the data. However,
21		fitting a model does not necessarily mean that the model describes the empirical
22		data adequately. This vulnerability may lead to incorrect conclusions when model
23		assumptions are not met by real data.
24	•	Here, we present a model adequacy test for likelihood models of chromosome-
25		number evolution. The procedure allows to determine whether the model can
26		generate data with similar characteristics as those found in the observed ones.
27	•	We demonstrate that using inadequate models can lead to inflated errors in several
28		inference tasks. Applying the developed method to 200 angiosperm genera, we
29		find that in many of these, the best-fitted model provides poor fit to the data. The
30		inadequacy rate increases in large clades or in those in which hybridizations are
31		present.
32	•	The developed model adequacy test can help researchers to identify phylogenies
33		whose underlying evolutionary patterns deviate substantially from current
34		modelling assumptions and should guide future methods developments.
35	Ke	words: chromEvol, chromosome number, dysploidy, model adequacy, model
36	sel	ection, phylogenetics, polyploidy.

37 Introduction

38 Chromosome number is widely recognized as a key feature of eukaryote genomes. Its 39 popularity in cyto-taxonomical and evolutionary studies has been attributed to its 40 ability to provide a concise description of the karyotype, the ease by which it can be 41 recorded, and its stable phenotype across repeated measurements. Processes that lead 42 to changes in chromosome numbers have direct consequences on central evolutionary 43 processes related to reproductive isolation and speciation, thus providing important 44 information for species determination and phylogenetic relationships (Guerra, 2008; 45 Weiss-Schneeweiss & Schneeweiss, 2013). While chromosome numbers generally 46 exhibit strong phylogenetic signal (e.g. Vershinina & Lukhtanov, 2017; Carta et al., 47 2018), they are also highly dynamic. This variability has been particularly well 48 acknowledged in plants, with counts ranging from n = 2 to n = 720 (Khandelwal, 49 1990; Ruffini Castiglione & Cremonini, 2012), and records showing intraspecific 50 variation in 23% of angiosperm species (Rice *et al.*, 2015). Understanding the 51 underlying processes that gave rise to these changes allows inference of major 52 genomic events that have occurred in the history of a clade of interest and the 53 processes that have shaped its diversification.

Of the various mechanisms underlying chromosome-number change, polyploidy, or 54 55 whole genome duplication (WGD) has received significant attention because of the 56 profound impacts such an event has on the organism. Polyploids often differ markedly from their progenitors in morphological, physiological, or life history characteristics, 57 58 which may contribute to their establishment in novel ecological settings (Stebbins, 59 1971; Levin, 1983; Ramsey & Schemske, 2002; Soltis et al., 2007; Leitch & Leitch, 60 2008; Ramsey & Ramsey, 2014; Spoelhof et al., 2017; Rice et al., 2019). Polyploidy 61 is thus recognized as one of the major processes that has driven and shaped the 62 evolution of higher organisms. A more subtle change in chromosome number is 63 dysploidy, leading to step-wise changes in the number of chromosomes, but typically 64 does not immediately alter the genomic content. Dysploidy occurs via several types of 65 genome rearrangements, leading to ascending or descending dysploidy through 66 chromosome fission or fusion (Weiss-Schneeweiss & Schneeweiss, 2013). 67 Deciphering the pattern of chromosome-number change within a clade allows 68 inferring the number and type of transitions that have occurred along branches of a

69 phylogeny, to estimate ancestral chromosome numbers, and to categorize extant

70 species as diploids or polyploids.

71 In the last decade, several tools that infer changes in chromosome numbers along a 72 phylogeny were developed (Mayrose et al., 2010; Hallinan & Lindberg, 2011; Glick 73 & Mayrose, 2014; Freyman & Höhna, 2017; Zenil-Ferguson et al., 2017, 2018; 74 Blackmon et al., 2019). Among these, the chromEvol probabilistic framework 75 (Mayrose *et al.*, 2010) was the first to incorporate a continuous time Markov process 76 that describes the instantaneous rate of change from a genome with *i* haploid 77 chromosomes to a genome with *j* haploid chromosomes via specific types of 78 dysploidy and polyploidy transitions. Further development of this framework allowed 79 for more intricate types of chromosome-number transitions (Glick & Mayrose, 2014), 80 to differentiate between transitions that coincide with speciation events and those that 81 occur continuously in time along branches of the phylogeny (Freyman & Höhna, 82 2017), and to associate patterns of chromosome-number change with the evolution of 83 a discrete character trait (Zenil-Ferguson *et al.*, 2017; Blackmon *et al.*, 2019). 84 In the chromEvol model, each type of transition is represented by a parameter 85 describing its rate of change. The inclusion (or exclusion) of different parameters 86 entails different hypotheses regarding the pathways by which the evolution of 87 chromosome number proceeded in the clade under study. In a regular application of 88 the chromEvol framework, different models are fitted to the data and the best one is 89 chosen by comparing the relative fit of each model to the data at hand using 90 established model selection criteria, such as the likelihood ratio test or Akaike 91 Information Criterion (AIC; Akaike, 1974). In reality, however, no empirical dataset 92 will meet all the assumptions of any model and thus relaying on the best model (or set 93 of models) may be vulnerable to incorrect conclusions in datasets whose underlying 94 evolutionary process deviate substantially from current modelling assumptions. To 95 prevent such errors, here we develop a model adequacy test that allows determining 96 whether a given model of chromosome-number evolution provides a realistic 97 description of the evolutionary process for reliable inferences. 98

Several assumptions made by existing models of chromosome-number evolution may
be violated when empirical data are analyzed. For example, all models rely on a
memory-less Markovian process, in which the transition rates are only dictated by the

101 current number of chromosomes of the lineage. Thus, for example, the transition rate 102 from n = 10 to n = 9 is not affected by the duration of time the lineage possessed 10 103 chromosomes, nor by the sequence of events that had led to it. However, because 104 rates of descending dysploidy may increase following WGD (Wood et al., 2009; 105 Wendel, 2015; Soltis *et al.*, 2016), the transition from n = 10 to n = 9 is more probable 106 if n = 5 was the ancestral state compared to n = 11. Additionally, most models assume 107 that the transition rates are similar across the phylogeny, although in practice the 108 transition patterns may be rather different in some sub-clades compared to others, as 109 has been demonstrated, for example, in Cyperaceae (Márquez-Corro et al., 2019). 110 Finally, all current models are based on a phylogenetic structure and thus ignore the 111 possibility of hybridizations. Notably, allopolyploidy, one of the main types of 112 polyploidy, is defined by such reticulate evolutionary events and the biases caused by 113 their presence is rather unexplored.

114 One aspect of understanding the reliability of a model and interpreting its results is to 115 quantify its adequacy for the data and the question at hand. The aim of model 116 adequacy tests is to determine the absolute fit of a model to the data, rather than to 117 compare its relative fit among a set of models. With some variations, the general 118 procedure of such tests is composed of several steps: first, given an empirical dataset, 119 obtain the best-fitting model and its parameter values. Next, use that model to 120 generate multiple simulated datasets. Then, compute several test statistics that 121 describe various characteristics of the data on each simulated dataset and on the 122 empirical dataset. If the empirical values of the test statistics fall outside the range of 123 variation encompassed by the simulated data, then it may be concluded that the model 124 cannot provide an adequate description of the data at hand. To date, model adequacy 125 approaches are established for several types of data and inference tasks, including 126 those related to sequence evolution (Bollback, 2002; Brown, 2014; Duchêne et al., 127 2015; Chen et al., 2019) and continuous valued organismal traits (Slater & Pennell, 128 2013; Pennell *et al.*, 2015). However, both are inappropriate for data and analyses 129 concerning the evolution of chromosome numbers as the former rely on statistics 130 derived from many sites, while the latter rely on Brownian motion statistics.

In the following, we first provide the details of the developed model adequacyframework for likelihood models of chromosome-number evolution. We then use

- simulations to assess the type I error rate and to explore the consequences of using
- 134 inadequate models in several common inference tasks, such as ancestral
- reconstructions of chromosome numbers and ploidy-level inference. Finally, we apply
- the developed procedure to a large cohort of angiosperm genera, as well as to clades
- 137 that are expected to violate model assumptions.

138 Methodological Description

139 Model adequacy framework for chromosome-number evolution

140 Given chromosome count data and a compatible phylogeny (together denoted as D), 141 chromEvol can be used to assess the fit of various models $(M_1, M_2, ..., M_N; N$ denotes 142 for the number of models) to D. Each model differs with respect to the included rate 143 parameters or the constraints placed on them $[\theta(M_1), \theta(M_2), \dots, \theta(M_N)]$. The most 144 general model considered here includes six free parameters (Glick & Mayrose, 2014) 145 and assumes that five types of events are possible: a single chromosome-number 146 increase (ascending dysploidy with rate λ) or decrease (descending dysploidy with 147 rate δ), WGD (i.e. exact duplication of the number of chromosomes with rate ρ), 148 demi-polyploidy (multiplications of the number of chromosomes by 1.5 with rate μ), 149 and base-number transitions (the addition to the genome by any multiplication of an 150 inferred base number, where β , is the inferred base number and v is its respective 151 transition rate). A combination of these parameters allows a range of models to be 152 evaluated (Table 1 shows the various models considered here). We note that the 153 chromEvol software also allows the ascending and descending dysploidy rates to 154 depend on the current number of chromosomes, but this option was not evaluated 155 here.

156 In a common application of chromEvol, several models are fitted to D, the optimal 157 model is selected based on its relative fit using established model selection criteria 158 (e.g. AIC), and subsequent inference tasks are performed based on this model. The 159 model adequacy test can be carried out to any model of interest, whether or not it is 160 the most fitted one. The general aim of this test is to examine whether a specified 161 model, M_x , is able to generate data that are similar to D. Our model adequacy 162 procedure is based on parametric bootstrapping (Goldman, 1993; Efron & Tibshirani, 163 1994), where the observed data are compared to a background distribution generated

164 from simulations. These simulations are generated under the specified model, whose parameters, $\hat{\theta}(M_{\chi})$, were optimized with respect to D and the respective probabilities 165 166 of chromosome-numbers inferred at the root of the phylogeny (exact details of the 167 simulation procedure are given in the Supporting Information). Comparing between 168 true and simulated data is performed using a set of test statistics, which reflects 169 various characteristics of the data. First, the test statistics $(T_1, T_2, ..., T_m; m)$ denotes for 170 the number of statistics) are computed for the true data D. Second, multiple datasets 171 are simulated under the specified model and its inferred parameters. For each 172 simulated dataset, the same set of test statistics is computed, resulting in a distribution 173 for each test statistic $(T_{s1}, T_{s2}, ..., T_{sm})$. If the empirical value of the test statistic falls 174 within the range of variation encompassed by the simulated data (herein defined as the 2.5th and 97.5th percentiles), the model is considered as capable of generating data 175 similar to the original ones and is thus inferred as adequate. Otherwise, it is inferred 176 177 as inadequate. A schematic illustration of the developed model adequacy test is 178 presented in Fig. 1.

179 In our implementation, four test statistics were calculated given the chromosome-180 number data of extant taxa and the corresponding phylogeny: (1) Variance; higher 181 values in the simulated data relative to the observed ones may point to some 182 constraints that were not accounted for by the model (e.g. hard bounds on the number 183 of chromosomes in the genome), or to errors in the parameter estimation process. (2) 184 Shannon's entropy (Shannon, 1948); Lower entropy of the observed data than 185 predicted by the model is indicative of higher-than-expected concentration of 186 genomes with certain haploid numbers. This could be due to selective constraints, or 187 to a very low variability exhibited in certain subclades of the phylogeny, such that 188 specific states are clumped into large blocks of the tree more than expected. (3) 189 Parsimony score; the most parsimonious number of character transitions across the 190 phylogeny is calculated based on Fitch (1971). If the parsimony score of the observed 191 data are lower than expected it means that the model assumes more transitions than 192 actually occurred. This could occur due to rate heterogeneity across the tree. For 193 example, if chromosome-number transitions occur more frequently in one subclade 194 relative to the rest of the phylogeny, this could be accommodated by inferring higher values of the transition rates. (4) Parsimony versus time (Pars^{Time}); the parsimonious 195 196 number of transitions are computed per branch using the accelerated transformation

197 criterion (ACCTRAN; Farris, 1970). The regression line between the divergence 198 times (computed from the root to the end of the branch) and their parsimony scores is 199 calculated, and the slope of this line is taken as the test statistic. This statistic is 200 similar in spirit to that employed by Pennell et al. (2015) for testing the adequacy of 201 models for continuous trait evolution. Under a time-homogenous model, as 202 implemented in chromEvol, we expect no relationship between the divergence times 203 and the number of transitions. Violations of this assumption suggests that transitions 204 are either concentrated around the root or occur more frequently towards the tips. We 205 note that aside from these four statistics, two additional ones were computed (the range and the number of unique counts). These two statistics were found to be highly 206 correlated with the other test statistics ($r^2 = 0.85$ between range and variance and $r^2 =$ 207 208 0.74 between unique counts and entropy, when computed over the 200 empirical 209 datasets; detailed below), and thus we chose to discard them from further analyses. 210 The coefficient of determinations between all pairs of the four remaining test statistics 211 was below 0.40 (Supporting Information Table S1). Because the four test statistics are 212 not independent and researchers might be interested in revealing the specific aspects 213 of the data that differ from expectations, we followed Pennell et al. (2015) and did not 214 apply a multiple testing correction. Thus, in all analyses presented here a model is 215 considered as adequate only if all four statistics fall within the boundaries of the 216 simulated distribution.

217 **Performance assessment using simulations**

Simulations were conducted to examine the performance of the model adequacyprocedure. Given an input phylogeny and a set of model parameters, simulated

220 chromosome numbers were generated as previously described in Mayrose *et al.*

221 (2010). As the number of simulation conditions is infinite, we concentrated on eight

scenarios that vary in terms of data size (the number of tips in the phylogeny and the

- 223 observed chromosome-number distribution) and the inferred pattern of chromosome-
- number change (Table 2). The phylogenies, chromosome counts, and model
- 225 parameters were taken from empirical datasets previously analyzed using chromEvol
- 226 (Glick *et al.*, 2016; Rice *et al.*, 2019), thus representing realistic data characteristics.
- 227 For each simulation scenario, a total of 100 replicates were generated. Each simulated
- dataset was then fitted to a set of four models: D_{ys} , $D_{ys}D_{up}$, $D_{ys}B_{num}$, $D_{ys}D_{up}B_{num}$. For

each simulation scenario, one of these models was the generating model (i.e. the
model that was used to simulate the data) and three were non-generating models. We
note that these models share both common and distinct aspects of the parameter space,
such that some, but not all, models are nested within each other (Table 1). Finally, the

adequacy of each model to the simulated data was assessed.

234 Inference errors of adequate and inadequate models

235 The consequences of using an adequate versus inadequate model were evaluated by comparing the errors of four common inference tasks: (1) the chromosome number at 236 237 the root of the phylogeny; calculated as the deviation from 1.0 of the posterior 238 probability assigned to the true (i.e. simulated) chromosome number at the root. (2) 239 The total number of dysploidy events across the phylogeny; calculated as the relative error between the inferred and simulated number of events: $2 * \frac{|x_1 - x_2|}{x_1 + x_2}$, where x_1 and 240 241 x_2 are the simulated and inferred number of dysploidy events, respectively. In case 242 both x_1 and x_2 equal zero the error was assigned as zero. (3) The total number of 243 polyploidization events across the phylogeny; the relative error was calculated similar 244 to the total number of dysploidy events. Duplication events, demi-duplications, and 245 base-number transitions were regarded as polyploidization events. (4) Ploidy level 246 assignments. The ploidy-level inference of tip taxa, as either diploids or polyploids, 247 was based on the procedure described in Glick & Mayrose (2014). The assignments of 248 all tips were compared between the inferred and true values. The number of falsely 249 inferred taxa, divided by the total number of taxa, was used as the error measure.

250 In this analysis, six of the eight simulation scenarios in Table 2 were examined. The 251 two scenarios excluded were those generated under the simple D_{vs} model, for which 252 not all inference tasks are relevant. To eliminate possible confounding effects between 253 the specific model used for inference and the magnitude of the error, in this evaluation 254 a single non-generating model (D_{vs}D_{up} or D_{vs}B_{num}) was fitted to the data per 255 simulation scenario (Supporting Information Table S2). For each simulation scenario, 256 300 replicates were generated. For each replicate, the phylogeny and the simulated 257 chromosome counts were given as input to the model adequacy test and the dataset 258 was determined as either adequate or inadequate. A one-sided *t*-test was conducted to

259 determine whether the error of a certain inference task is significantly larger in the

260 inadequate set compared to the adequate set.

261 Application to empirical datasets

262 To demonstrate the usability of the model adequacy framework, we applied it to a 263 dataset of 200 angiosperm genera, which were randomly selected from a large 264 database consisting of thousands of plant genera, excluding genera with no variations 265 in chromosome numbers as well as those with less than 5 species with both 266 phylogenetic and chromosome-numbers information. The initial database was used, in 267 part or as a whole, in several previous analyses (e.g. Glick et al., 2016; Zhan et al., 268 2016; Salman-Minkov et al., 2016; Rice et al., 2019). From this database we also 269 selected 40 angiosperm genera that each contains at least one allopolyploid species, 270 based on data from Barker et al. (2016). Due to overlaps between these two sets, a 271 total of 233 unique datasets were analyzed. Full details of the reconstruction of the 272 original database are described in Rice et al. (2019). Briefly, for each genus, the 273 OneTwoTree pipeline (Drori *et al.*, 2018) was used to automatically reconstruct the 274 phylogeny using publicly available sequence data as appear in GenBank (Benson et 275 al., 2013). Chromosome numbers for all species were retrieved from the Chromosome 276 Counts Database (CCDB; Rice *et al.*, 2015). These data were given as input to 277 chromEvol, which was executed on the six models detailed in Table 1. Additionally, 278 we applied similar procedures to seven clades of higher taxonomical ranks, including 279 five families, one subfamily, and one tribe. The evolution of chromosome numbers in 280 these clades using chromEvol was previously examined in several studies (Supporting 281 Information Table S3).

282 Implementation and availability

- 283 The model adequacy procedure was implemented in Python and R (R Core Team,
- 284 2013). The source codes and running instructions are available at
- 285 <u>https://github.com/MayroseLab/chromEvol_model_adequacy</u>. The obligatory inputs
- are three files obtained through a chromEvol run of the examined model: the
- summary results file, the tree with the inferred ancestral reconstruction in a NEWICK
- format, and the original counts file in FASTA format. The program outputs, for each
- test statistic examined, its value computed from the empirical data, the percentile in

which it falls within the simulated distribution, and the 2.5th and 97.5th percentiles of the simulated distribution as indicative for the upper and lower bounds expected under the modelling assumptions. The model adequacy test is also available for on-line use through the chromEvol web-server (<u>http://chromevol.tau.ac.il/</u>), which is currently in a Beta version.

295 **Results**

In this work we developed a statistical framework for testing the adequacy of 296 297 likelihood models of chromosome-number evolution. In essence, the method tests 298 whether a specified model is capable of generating data that are similar to the data at 299 hand. If not, the model is considered as providing inadequate description of the data, 300 suggesting that other processes than those modeled have driven the evolution of 301 chromosome numbers along the examined phylogeny. We first evaluated the 302 performance of the model adequacy framework using simulations. We then applied it 303 to a large number of real datasets derived from dozens of angiosperm genera, as well 304 as to seven clades of higher taxonomic ranks, that together vary greatly in their extent 305 of divergence time and patterns of chromosome number variation.

306 Framework validation

307 Simulations were used to validate the developed model adequacy approach. Several 308 simulation scenarios were examined, whose phylogenies and simulated parameters 309 were derived from real data analyses and cover various data characteristics (Table 2). 310 In each scenario, a single model was used to generate the data. Given the simulated 311 data, the generating model and three additional models were fitted to the data, and 312 their adequacies were examined. The four examined models are indicated by the type 313 of transitions they allow for: D_{vs}, D_{vs}D_{up}, D_{vs}B_{num}, and D_{vs}D_{up}B_{num} (Table 1). In total, 314 eight different simulation scenarios were examined; two for each type of generating 315 model.

316 We first examined the type I error rate, i.e. inferring the generating model as

317 inadequate. Our results indicated that when considering a single test statistic

independently, the error rate is around the expected value of 0.05 (average = 0.02,

319 across the eight simulation scenarios and four test statistics; Supporting Information

320 Table S4). Combining multiple test statistics together, we consider a model as 321 inadequate if one or more of the statistics fell outside the margins of the simulated distributions (see Methodological Description). Under this definition, the percentage 322 323 of generating models that were inferred as inadequate varied between 0.04 and 0.13 324 across the eight simulation scenarios (Table 3). When Bonferroni correction for 325 multiple testing was applied, the type I error rate dropped to an average of 0.008. We 326 note however, that the four test statistics are not independent, violating the assumption 327 of this correction.

328 We next examined the capability of the adequacy test to detect models that deviate 329 from that of the generating models. Three types of model misspecification were 330 examined: over-parameterization, under-parameterization, and miss-parameterization. 331 In the case of over-parameterization, the tested model allows for additional types of 332 chromosome-number change (as represented by extra free parameters) than those used 333 to generate the data. This corresponds to cases where the generating model is nested 334 within the tested model (e.g. the generating model is D_{vs}D_{up} while the tested model is 335 $D_{vs}D_{uv}B_{num}$). Our results indicated that the performances of over-parameterized 336 models are very similar to that of the generating models (Table 3). The few 337 discrepancies were the result of either (1) inaccurate parameter estimates of the more 338 general model due to the extra degrees of freedom; (2) the optimization procedure 339 reaching suboptimal regions of the parameter space (we note that while chromEvol 340 allow for more thorough likelihood optimization search, which should reduce such 341 instances, this was not attempted here due to the large number of simulations 342 employed); (3) very similar parameter estimates obtained using the two models, but 343 slight deviations of the test statistics led one model to be inferred as inadequate while 344 the other one as adequate.

345 In the case of under-parameterized models, the tested model allows for fewer types of 346 transitions than the generating model (e.g. the generating model is $D_{vs}D_{up}$ while the 347 tested model is D_{vs}). As may be expected, in all simulation scenarios the under-348 parameterized models were more frequently inferred as inadequate compared to the 349 generating models. The adequacy rate was very low when the tested model allowed 350 only for dysploid transitions while in reality polyploid transitions (either WGD and/or 351 base-number transitions) have occurred (Table 3; all cases where the tested model is 352 D_{vs}). The adequacy rates were higher when the generating model allowed for multiple

353 types of polyploid transitions (i.e. D_{vs}D_{up}B_{num} allowing for both exact duplications 354 and base-number transitions), while the tested model allowed for a subset of these 355 $(D_{ys}D_{up} \text{ and } D_{ys}B_{num} \text{ that allow only for duplications or base-number transitions,}$ 356 respectively). Comparing the adequacy of the two under-parametrized models ($D_{vs}D_{up}$ 357 and $D_{ys}B_{num}$), the $D_{ys}B_{num}$ model that incorporated base-number transitions had higher 358 adequacy rates compared to the D_{vs}D_{up} model that allowed for exact duplications, as 359 the former allows for several transitions that frequently include also exact 360 duplications (e.g. in case the base number is 8, both $8 \rightarrow 16$ and $8 \rightarrow 24$ transitions are 361 allowed).

362 In the case of miss-parametrization, the tested and generated models are not nested

363 within each other and thus their parameters only partially overlap. For the set of

models examined here, this fits the case where the generating model is $D_{ys}D_{up}$ while

the tested model is $D_{ys}B_{num}$, or vice versa. When the tested model was $D_{ys}B_{num}$, it

366 obtained similar adequacy rates to those of the generating $D_{ys}D_{up}$ model. In contrast,

367 and similar to the results detailed in the case of under-parameterized models, the

368 $D_{ys}D_{up}$ model was inferred as inadequate a large number of times when the generating 369 model was $D_{ys}B_{num}$.

370 Inference errors of adequate and inadequate models

371 A central usage of probabilistic models of chromosome number evolution is their 372 inference capabilities, such as ancestral reconstructions of chromosome numbers, or 373 predicting the branches in which dysploidy and polyploidy events have most likely 374 occurred. Still, it is unclear whether the use of inadequate models would deteriorate 375 the performance of such inference tasks. To this end, simulations were used to 376 compare the errors of the following four common inference tasks when adequate and 377 inadequate models are employed: (1) the chromosome number at the root of the 378 phylogeny; (2) the total number of inferred dysploidy events; (3) the total number of 379 inferred polyploidization events, and (4) inferring the ploidy level of tip taxa as either 380 diploid or polyploidy (see Methodological Description for details regarding the error 381 computed for each inference task).

Our results demonstrated that the use of inadequate models frequently leads to largerinference errors, although under some simulation scenarios the inference errors of

384 inadequate models were similar to that obtained using adequate models. For example, 385 the error in the inference of the root chromosome number was significantly larger in 386 the case of inadequate models under two simulation scenarios, but was non-387 significantly different in the other four (Fig. 2). Similarly, in two out of the six 388 simulation scenarios, the error of inferring the ploidy level of extant taxa was 389 significantly larger when computed using inadequate versus adequate models. In this 390 case, the magnitude of the error was relatively low whether adequate or inadequate 391 models were applied: when inadequate models were applied, the mean error was 4.6% 392 across all simulation scenarios, reaching up to 12% under the Brassica simulation 393 scenario. In comparison, the mean error was 2% when adequate models were applied, 394 reaching up to 6% of erroneous inferences under the Hordeum simulation scenario. 395 Larger differences in the errors between adequate and inadequate models were 396 observed in inferring the total number of polyploidizations, and even more so in 397 inferring the total number of dysploidy events. For both these inference tasks, 398 significant differences between adequate and inadequate models were obtained for 399 three out of the six simulation scenarios. Generally, the relative error in inferring the 400 total number of dysploidy events was larger compared to that of inferring the total 401 number of polyploidizations (the mean relative error was roughly twice for dysploidy 402 compared to polyploidy transitions, both in the adequate set and the inadequate set; 403 Fig. 2).

404 Application to empirical datasets

We applied the model adequacy framework to 200 datasets, each corresponding to a 405 406 single randomly-selected angiosperm genus. First, we performed a standard model 407 selection procedure based on the AIC (Akaike, 1974) to evaluate the relative fit of 408 each of the six chromEvol models to the data. In 24% of the datasets, the simple D_{vs} 409 model, which allows for dysploid transitions only, was selected. The model that was 410 most frequently selected was D_{ys}D_{up} (28%), while models that allow for demi-411 polyploidy transitions and those that allow for base-number transitions were selected 412 in 27% and 21% of the datasets, respectively (Fig. 3a). Next, we applied the model adequacy test to the best model identified for each dataset. We found that in 74% of 413 414 the genera, the model that was chosen as best by the AIC was inferred to provide an 415 adequate description of the data. Applying the model adequacy test to all six models

416 per dataset (whether or not selected as best), we found that models that allow for 417 fewer types of transitions were more frequently predicted as inadequate (Fig. 3b). For 418 example, the D_{vs} model that allows only for dysploidy transitions was adequate in 419 only 28% of the 200 datasets, models that additionally allow for one type of 420 polyploidy, either duplication or base-number transition, were adequate 60% and 64% 421 of the cases, respectively, while the three models that incorporate two types of 422 polyploidy transitions $(D_{vs}D_{up}D_{em}, D_{vs}D_{up}D_{em}^{*}, and D_{vs}D_{up}B_{num})$ were inferred as 423 adequate most frequently. The adequacy rates of all models were generally related to 424 the complexity of the model that was selected as optimal. Thus, when the most 425 complex models were selected $(D_{vs}D_{up}D_{em} \text{ and } D_{vs}D_{up}B_{num})$, the adequacy rates of all 426 models - including that of the chosen model - were low (33% and 47%, 427 respectively), while when the least complex model was selected, the adequacy rates of 428 all models was high (70%; Supporting Information Table S5). 429 Next, we examined the model adequacy procedure in groups that have evolved via 430 reticulate evolution at some point in their histories. In these clades, the underlying 431 assumption of the chromEvol framework, in which evolution proceeds along a 432 phylogenetic structure, is violated, at least to some extent. This analysis was 433 performed on 40 genera that were identified in the literature to include allopolyploid 434 species, and thus hybridizations were reported to occur (data taken from Barker et al., 435 2016). In the majority of these genera (24 out of 40), the model that was selected as 436 optimal according to the AIC was found by our model adequacy procedure as inadequate. This adequacy rate is significantly lower ($p \ll 0.05$; χ^2 test) compared to 437

438 a random set of 193 genera in which allopolyploidy was not reported (the 200 genera

analyzed above, omitting seven that include a reported allopolyploid species).

440 Finally, we evaluated the model adequacy procedure on a set of seven groups whose 441 taxonomic rank is higher than the genus level, thus representing clades whose 442 divergence time is generally older than those inspected above. The evolution of 443 chromosome numbers in these clades likely violates the time homogeneity assumption 444 of chromEvol, in which the transition pattern is similar across the phylogeny. For four 445 of these seven clades, the model that was chosen as optimal according to AIC did not 446 provide adequate description of the data according to the model adequacy test 447 (Supporting Information Table S3) and in one additional case the empirical values of

two test statistics were placed close to the lower boundaries of the simulated
distributions (falling in the 0.027 and 0.043 percentiles). Taken together, the last two
analyses indicate that the model adequacy procedure can identify cases in which the
evolution of chromosome numbers is driven by processes that deviate from the basic
modelling assumptions of the chromEvol framework.

453 **Discussion**

454 For over a century, the determination of chromosome numbers has played a vital role 455 in studying evolutionary and genomic processes in plants. Probabilistic models of 456 chromosome-number change are a relatively recent addition to the research toolbox 457 available to study the evolution of major genomic processes. As the usage of such 458 models increases, so does the need to assess their validity when applied to real data. 459 Here, we developed a model adequacy test for likelihood models of chromosome-460 number evolution. We focused our analysis on those models implemented in the 461 chromEvol software (Glick & Mayrose, 2014), but the procedures are general and can 462 be implemented in other platforms that use variations to the chromEvol model 463 (Freyman & Höhna, 2017; Zenil-Ferguson et al., 2017; Blackmon et al., 2019). The 464 developed test is based on the parametric bootstrapping approach (Goldman, 1993; 465 Efron & Tibshirani, 1994) in which observed data are compared to a simulated 466 distribution generated by the examined model. Using multiple test statistics that 467 describe various characteristics of the data, the test allows to determine whether the 468 model can generate data that are similar to those found in the observed ones.

469 Our simulation results indicate that the model adequacy framework has an acceptable 470 type I error rate (i.e. inferring as inadequate a model that was used to generate the 471 data). However, higher type I errors were found in models that allow for base-number 472 transitions ($D_{vs}B_{num}$ and $D_{vs}B_{num}D_{up}$). This suggests that these models might not be 473 appropriate in all cases. The current implementation of such models assumes the same 474 rate for all possible base-number transitions (e.g. given a base number of $\beta = 7$, the 475 additions of 7, 14, or 21 chromosomes are equally likely). Alternatively, it may be 476 more appropriate to place a probability distribution over the possible base-number 477 transitions. This will allow, for the example of $\beta = 7$, higher rates for additions by 7 478 chromosomes compared to those by 21.

479 Our simulation results also demonstrate that the adequacy rate of over-parameterized 480 models, which allows for more types of transitions than those that truly occurred, is 481 similar to that of the generating models. While it is expected that the accuracy of 482 inferring the model parameters will decrease as overly-complexed models are 483 evaluated, in many cases the auxiliary parameters were optimized to very low values, 484 resulting in a process that is nearly identical to the generating model. Thus, it seems 485 that the flexibility offered by complex models does not necessarily lead to their 486 disadvantage, at least for some inference tasks, as has been recently demonstrated for 487 models of nucleotide sequence evolution (Abadi et al., 2019). In other cases of model 488 violations, either for under-parameterized or miss-parameterized models, when the 489 rate parameters deviated substantially from the original ones (e.g. dysploidy rates an 490 order of magnitude larger than the simulated rates), the model adequacy framework 491 detected such cases as inadequate. This suggests that the adequacy test is capable of 492 detecting models that are completely wrong. In other cases, the nature of model 493 misspecification affected the outcome. In the simulations examined here, $D_{vs}B_{num}$ was 494 more frequently adequate than $D_{vs}D_{up}$, both in the case of under-parameterization (i.e. 495 when the generating model was $D_{vs}D_{up}B_{num}$ such that both models miss one type of 496 transition) and miss-parameterization. Nevertheless, we note that the $D_{vs}B_{num}$ model 497 may not fit well in large phylogenies with high dysploidy rates. In its current 498 implementation, the model assumes that a single base number typifies a clade. 499 However, if there is a high dysploidy rate, each subclade of the phylogeny may be 500 characterized by its own base number or by multiple base numbers, which will 501 necessitate more complex modelling options.

502 We further tested the consequences of using an inadequate model by examining the 503 errors of several inference tasks. First, we found that the difference in inference error 504 between adequate and inadequate models depended on the simulation scenarios: in 505 some simulation scenarios the use of inadequate models resulted in significantly 506 inflated inference errors compared to the use of adequate models, in some scenarios it 507 affected only certain inference tasks and not others, while in others the difference was 508 negligible for all tasks. Second, we found that some inference tasks are much more 509 sensitive to model misspecification than others. The assignment of extant taxa as 510 diploids or polyploids was the inference task that was least affected from using an 511 inadequate model, and in general, the error of this inference task was very low (in all

scenarios, the ploidy level of 88% or more of the taxa were correctly identified). This
indicates that determining the ploidy levels of extant taxa is generally robust to model
misspecification. On the other hand, the error of determining the number of events
that had occurred – either dysploid or polyploid transitions – can be substantial when
inadequate models are employed.

517 Applying the model adequacy test to hundreds of angiosperm genera, we found that in 518 the majority of the cases the best-fitted model provided sufficient approximation to 519 the evolutionary processes underlying the data and was determined as adequate. 520 However, in roughly one fourth of the examined genera, this selection turned out to be 521 inadequate, suggesting that there is ample room for future modelling improvements. 522 Indeed, we found high rates of model inadequacy when applying the developed 523 procedures to two types of clades that are expected to violate basic modelling 524 assumptions: first, clades in which allopolyploidy events are known to occur, thus 525 violating the assumption that evolution proceeds via a phylogenetic structure; second, 526 in the case of large and diverse clades in which a single transition process is fitted to 527 the entire phylogeny, following the time homogeneity assumption, is insufficient. 528 These results thus indicate that promising future developments would be to focus on 529 analytical procedures based on phylogenetic networks (Nakhleh, 2010), rather than on 530 bifurcating phylogenies, and to further incorporate time-heterogeneous processes. 531 Phylogenetic model adequacy tests have been previously developed for other data 532 types and inference tasks, although their use has not been widely adopted. This could 533 be due to the apparent limited benefit offered to a researcher when all examined 534 models are deemed inadequate when applied to a clade of interest. We argue, 535 however, that model adequacy tests are of practical use to methods developers and 536 end users alike, and should thus be regularly practiced as part of a broader model

537 assessment routine. For researchers interested in data analysis, inadequate models can

538 hint on errors in the input data, which should thus be more carefully inspected. In the

case studied here, possible sources of errors include those in the assumed

540 phylogenetic hypothesis, in the collection of chromosome counts, or in taxa sampling.

541 Inadequacy could also point to additional attributes that should be considered in the

- analysis. For example, if all models that assume a time-homogenous transition
- 543 process fail, it could suggest that patterns of chromosome-number change are
- dependent on an organismal trait (e.g. the plant growth form), that if accounted for,

545 using more complex models (e.g. Zenil-Ferguson et al., 2017; Blackmon et al., 2019) 546 would enhance the analysis. For researchers interested in large scale analyses that 547 include multiple datasets, where the in-depth examination of each inadequate dataset 548 is not feasible, the filtration of such clades is one obvious possible direction. For some 549 inference tasks, such as the identification of ploidy levels of extant taxa, the effect of 550 using an inadequate model is rather negligible, indicating that the treatment of the 551 flagged clades should be tuned to the analysis in question. For the developers, the 552 frequent application of model adequacy tests should provide interesting test cases on 553 which new models are trained. Moreover, when a model is deemed inadequate, the 554 test statistics that fail to align may point to processes absent from existing models, 555 which could be included in the future. Model adequacy should thus take a vital part in 556 this recurrent chain of scientific progress in which new methods are developed, 557 regularly used, and then replaced by more advanced alternatives.

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562 Author Contribution

IM and AR conceived the study; AR built the tool and analyzed the data; AR and IMwrote the manuscript; IM supervised the study.

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679 **Tables**

- **Table 1.** The set of chromEvol models examined in this study, together with their rate
- 681 parameters.

Model	Model parameters ¹	Description	Nested in ²
D _{ys}	λ, δ	Dysploidy (descending or ascending)	$D_{ys}D_{up}, D_{ys}D_{up}D_{em}^{*},$ $D_{ys}D_{up}D_{em}, D_{ys}B_{num},$ $D_{ys}D_{up}B_{num}$
$D_{ys}D_{up}$	λ, δ, ρ	Dysploidy and duplication	$D_{ys}D_{up}D_{em}^{*},$ $D_{ys}D_{up}D_{em}, D_{ys}D_{up}B_{num}$
$D_{ys}D_{up}D_{em}^{*}$	λ , δ , $ ho = \mu$	Dysploidy, constraining equal rates of duplication and demi-polyploidy	
$D_{ys}D_{up}D_{em}$	λ, δ, ρ, μ	Dysploidy, duplication, and demi- polyploidy	$D_{ys}D_{up}D_{em}^{*}$
$D_{ys}B_{num}$	λ, δ, β, ν	Dysploidy and base number transition	$D_{ys}D_{up}B_{num} \\$
$D_{ys}D_{up}B_{num}$	λ, δ, ρ, β, ν	Dysploidy, base number transition, and duplication	

¹ The model parameters are the base number (β), and rates of ascending dysploidy (λ),

descending dysploidy (δ), duplication (ρ), demi-duplication (μ), and base number

² In case all parameters of the model are a subset of other models, the more complex
models are indicated.

⁶⁸⁴ transition (v).

Genus	Number	Generating model	Model parameters ¹				
Genus	of taxa		λ	δ	ρ	β	v
Aloe	120	$D_{ys}D_{up}$	0 (0)	0.34 (1)	2.61 (8)		
Phacelia	53	$D_{ys}D_{up}$	0.20 (2)	2.33 (21)	0.67 (6)		
Lupinus	77	\mathbf{D}_{ys}	0.85 (7)	9.53 (76)			
Hypochaeris	38	\mathbf{D}_{ys}	1.14 (5)	0.43 (2)			
Brassica	36	$D_{ys}B_{num}$	1.24 (11)	0.70 (6)		8	0.55 (5)
Pectis	49	$D_{ys}B_{num}$	0 (0)	0.40 (2)		12	0.55 (3)
Crepis	81	$D_{ys}D_{up}B_{num} \\$	2.41 (19)	0.99 (8)	0.26 (2)	8	0.18 (1)
Hordeum	36	$D_{ys}D_{up}B_{num} \\$	0 (0)	0 (0)	1.80 (5)	7	1.36 (4)

Table 2. The eight simulation scenarios examined in this study.

¹ In parentheses: average number of simulated events across the tree.

Table 3. The inadequacy rates of the four tested models in the various simulation

Simulation	Generating		Tested Mod	lels ¹	
scenario	Model	D _{ys} D _{up}	D _{ys}	$D_{ys}B_{num}$	$D_{ys}D_{up}B_{num}$
Aloe	$D_{ys}D_{up}$	0.07	1.00	0.06	0.08
Phacelia	$D_{ys}D_{up} \\$	0.04	0.99	0.04	0.03
Lupinus	D_{ys}	0.06	0.08	0.06	0.07
Hypochaeris	\mathbf{D}_{ys}	0.03	0.06	0.04	0.05
Brassica	$D_{ys}B_{num}$	0.14	0.98	0.05	0.06
Pectis	$D_{ys}B_{num}$	0.86	1.00	0.12	0.07
Crepis	$D_{ys}D_{up}B_{num} \\$	0.27	0.95	0.11	0.08
Hordeum	$D_{ys}D_{up}B_{num} \\$	0.77	1.00	0.30	0.13

690 scenarios examined (100 simulations per tested model per scenario).

¹The diagonal (white cells) are cases where the generating model is also the tested

model. Dark grey represents over-parametrized models, light grey under-parametrized

693 models, and patterned cells miss-parametrized models.

694 **Figure Legends**

Fig 1. A schematic illustration of the model adequacy framework for likelihoo	A schematic illustration of the model a	adequacy framework for likelihood
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696 models of chromosome-number evolution. In the case illustrated here, the model is

- adequate because none of the test statistics lies in the tails of the simulated
- 698 distribution.
- 699 Fig. 2. The mean inference errors obtained under adequate and inadequate models for
- roo each simulated scenario. Each row presents the error of a different inference task.

From top to bottom: inferring the total number of polyploid events across the tree,

inferring the total number of dysploid events across the tree, ploidy level assignments

of extant taxa, the probability of the chromosome number at the root of the

704 phylogeny. Each column denotes a different simulation scenario. For each scenario,

- 300 simulations were conducted and runs were partitioned to adequate and inadequate
- models. The violin plots represent the distribution of the errors obtained for the
- adequate (light grey, right) and inadequate (dark grey, left) sets. The black dot within
- each distribution denotes its mean. Asterisk indicates significant difference between

709 the two groups (*, p < 0.05 and ***, p < 0.01).

710 Fig. 3. Application of the model adequacy test to 200 angiosperm genera. (a) A bar

711 plot representing the frequency of selection according to the AIC of each of the six

tested models in the 200 examined angiosperm genera. The height of each bar is

partitioned according to the percentage of genera that were determined as adequate

(light blue) or inadequate (red). (b) The adequacy rate of each model when applied to

all genera, regardless of whether the model was selected (n = 200).

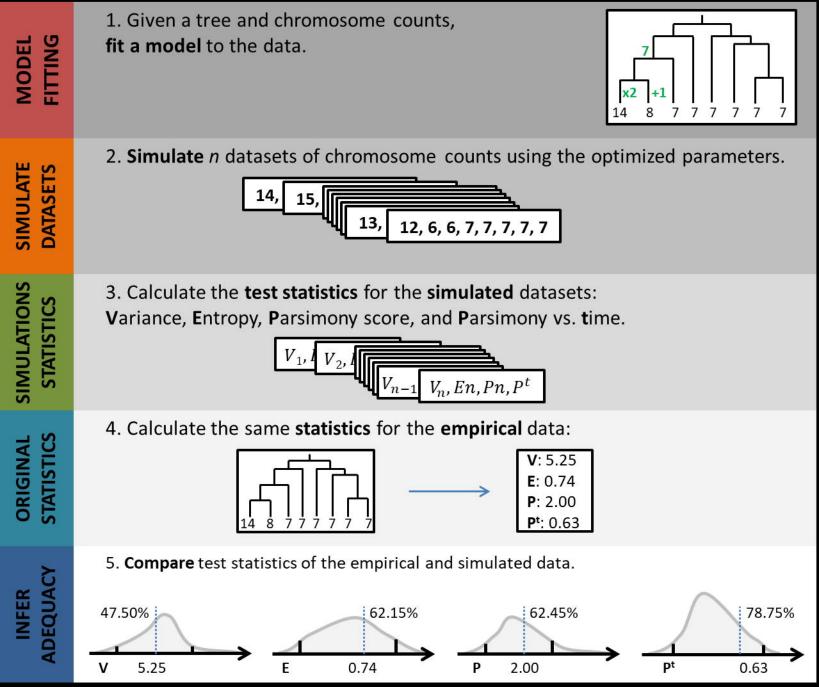
716 Supplementary information

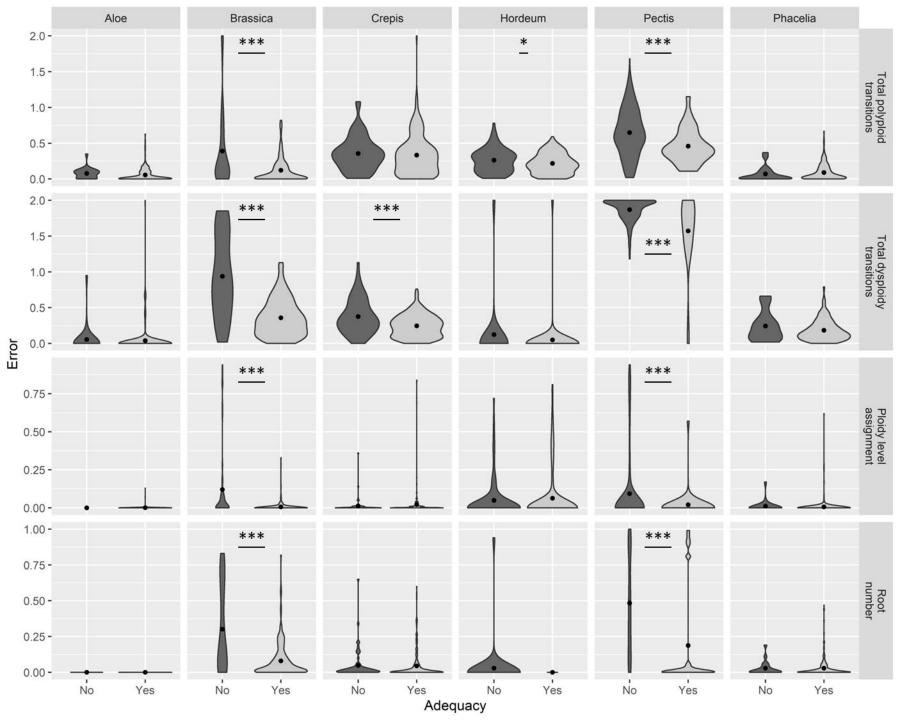
717 Supplementary Information Methods:

718 Item 1: Description of the simulation procedures.

719 Supplementary Information Tables:

- 720 Table S1: Pearson's *r* coefficient between each pair of statistics.
- Table S2: The generating and fitted model for each simulation scenario used in the
- comparison of inference error between adequate and inadequate models.
- Table S3: Details of the seven plant clades, whose taxonomic rank is above the genus
- 724 level, examined in this study.
- 725 Table S4: Type I error rates for each test statistic per simulation scenario.
- Table S5: Adequacy rates of all models, including those of the chosen models.





(a)

