The Self-Interest of Adolescents Overrules Cooperation in Social Dilemmas

Xiaoyan Wu[1,2,3†]**, Hongyu Fu**[1,2†]**, Gökhan Aydogan**[4]**, Chunliang Feng**[5]**, Shaozheng Qin**[1,2]**, Yi Zeng**[2,6,7,8,9]**, Chao Liu**[1,2]*

*For correspondence:
liuchao@bnu.edu.cn (CL)

[†]These authors contributed equally to this work

Present address: State Key Laboratory of Cognitive Neuroscience and Learning, Beijing Normal University, Beijing, China

[1]State Key Laboratory of Cognitive Neuroscience and Learning, and IDG/McGovernInstitute for Brain Research, Beijing Normal University, Beijing, China; [2]Beijing Key Laboratory of Artificial Intelligence Safety and Superalignment, Beijing, China; [3]Department of Adult Psychiatry and Psychotherapy, University of Zurich, Zurich, Switzerland; [4]Zurich Center for Neuroeconomics, Department of Economics, University of Zurich, Zurich, Switzerland; [5]School of Psychology, South China Normal University, Guangzhou, China; [6]Beijing Institute of AI Safety and Governance, Beijing, China.; [7]Brain-inspired Cognitive AI Lab, Institute of Automation, Chinese Academy of Sciences, Beijing, China.; [8]University of Chinese Academy of Sciences, Beijing, China.; [9]Long-term AI, Beijing, China.

Abstract   Cooperation is essential for success in society. Research consistently showed that adolescents are less cooperative than adults, which is often attributed to underdeveloped mentalizing that limits their expectations of others. However, the internal computations underlying this reduced cooperation remain largely unexplored. This study compared cooperation between adolescents and adults using a repeated Prisoner's Dilemma Game. Adolescents cooperated less than adults, particularly after their partner's cooperation. Computational modeling revealed that adults increased their intrinsic reward for reciprocating when their partner continued cooperating, a pattern absent in adolescents. Both computational modeling and self-reported ratings showed that adolescents did not differ from adults in building expectations of their partner's cooperation. Therefore, the reduced cooperation appears driven by a lower intrinsic reward for reciprocity, reflecting a stronger motive to prioritize self-interest, rather than a deficiency in predicting others' cooperation in social learning. These findings provide insights into the developmental trajectory of cooperation from adolescence to adulthood.

Introduction

Cooperation among individuals facilitates the achievement of shared goals and enhances overall group efficiency (*Fehr and Fischbacher, 2003*; *Nowak, 2006*). For individuals, cooperation skills are key to success in society; this ability is not innate but gradually acquired through socialization (*Warneken, 2018*). Successful cooperation requires individuals to prioritize the common purpose over their personal interests, focusing on collective goals (*Sachs et al., 2004*). Experimental psychology has often used the Prisoner's Dilemma Game (PDG; (*Axelrod and Hamilton, 1981*)) to study human cooperative behaviors. Extensive research has adapted the PDG into a repeated version to explore how people respond to interactive cooperation (*Andreoni and Miller, 1993*; *Embrey et al., 2018*), requiring individuals to adjust their responses dynamically to others and simulating real-life

cooperation more closely (*Axelrod and Hamilton, 1981*). In such social dilemmas, individuals face a trade-off between immediate rewards from defection and long-term benefits from cooperation (*Rilling et al., 2002*). Decision-making in these situations is thought to engage mentalizing abilities, which are functions related to theory of mind that enable individuals to form expectations about others' cooperative intentions (*Rilling et al., 2004*).

Cooperation is not an innate skill but is gradually cultivated and refined through socialization (*House et al., 2020*). Adolescence, in particular, marks a critical developmental phase in the transition to independent social roles (*Steinberg, 2005*). Studies using the PDG consistently show that adolescents cooperate less than adults (*Belli et al., 2012*; *Nava et al., 2023*; *Taheri et al., 2018*). This reduced cooperation is often attributed to an underdeveloped theory of mind, which may lead adolescents to underestimate others' trustworthiness and willingness to cooperate in social learning (*Gutiérrez-Roig et al., 2014*; *Fett et al., 2014*).

However, there are findings that may not support this hypothesis. For example, a previous study found that adolescents' lower cooperation, compared to adults, emerges only when following a partner's cooperation. Conversely, when the partner defected, adolescents' cooperative behaviors resembled those of adult (*Gutiérrez-Roig et al., 2014*). Similarly, a Trust Game study (*Fett et al., 2014*) reported a comparable pattern: adolescents invested less (measured as trust behavior) than adults only when the partner was cooperative. When faced with a non-cooperative partner, both adolescents and adults consistently reduced their trust behaviors. These findings suggest that adolescents, like adults, are able to adjust their behavior in response to others' actions. This selective reduction in adolescent cooperation implies that factors beyond deficits in mentalizing may be at play. Adolescents may prioritize maximizing immediate rewards over long-term reciprocity (*Rilling et al., 2002*). When confident that their partner will cooperate, defection may become the optimal strategy for maximizing self-interest. This hypothesis remains untested, but computational modeling could provide a valuable approach for examining the underlying mental processes behind these behavioral variations (*Farrell and Lewandowsky, 2010*; *Wu et al., 2024a*,b).

This study aimed to investigate variations in cooperative behavior between adolescents and adults and to explore the mental processes underlying these differences using computational modeling. Based on legal criteria for majority and prior empirical work, we adopt 18 years as the boundary between adolescence and adulthood (*Icenogle et al., 2019*; *Tervo-Clemmens et al., 2023*). A total of 127 adolescents and 134 adults participated in the study, playing a repeated Prisoner's Dilemma Game (rPDG) with a presumed human partner, whose behavior was predetermined by a computer program (see *Figure 1*a). The program ensured consistent conditions across age groups. To enhance realism, variability was introduced into the computer-simulated partner's behavior. The rPDG provides a symmetric and simultaneous framework that isolates the motivational conflict between self-interest and joint welfare, avoiding the sequential trust and reputation dynamics characteristic of asymmetric tasks such as the Trust Game (*Rilling et al., 2002*; *King-Casas et al., 2005*). Based on the standard payoff matrix of the rPDG (*Figure 1*b), mutual cooperation maximizes collective interests, while defection maximizes self-interest from an individual perspective. Our focus was on how adolescents respond to their partner's consistent cooperation and defection, aiming to identify potential mental variables contributing to adolescents' lower cooperation.

We developed computational models to investigate the dynamic variables guiding cooperative decisions in the rPDG. The model explicitly incorporates both expectations of the partner's cooperation and the intrinsic reward of reciprocity. A basic reinforcement learning (RL) algorithm was used to model participants' dynamic expectations regarding the partner's cooperation. Drawing on research on asymmetric reward learning in adolescents (*Palminteri et al., 2016*; *Rosenbaum et al., 2022*), we included asymmetric updating for positive (better-than-expected) and negative (worse-than-expected) outcomes. We identified the asymmetric RL learning model as the winning model that best explained the cooperative decisions of both adolescents and adults. Participants' expectations were modeled as a trial-by-trial dynamic variable, represented by parameter p. Following previous studies (*Fareri et al., 2012*, *2015*), a non-monetary reward for cooperation, represented

by parameter ω, reflects individual preferences for mutual cooperation. The term $p{\times}\omega$ quantifies
the intrinsic reward for reciprocity.

We hypothesize that adolescents will exhibit lower overall cooperation compared to adults, con-
sistent with previous studies (*Belli et al., 2012*; *Nava et al., 2023*; *Taheri et al., 2018*). Specifically,
we expect adolescents to demonstrate reduced cooperation after their partner's cooperation but
not following defection (*Gutiérrez-Roig et al., 2014*; *Fett et al., 2014*). Furthermore, we aim to ex-
plore whether this lower conditional cooperation is driven by inappropriate expectations of their
partners (represented by p), a reduced intrinsic reward for reciprocity (represented by $p{\times}\omega$), or a
combination of both.

Results

Adolescents exhibit lower cooperation than adults following partner cooperation, but not defection

In each trial of the rPDG, as shown in *Figure 1*c, participants were presented with two choices:
a triangle representing cooperation and a square representing defection. The choice associated
with each symbol was randomly balanced across participants. They were informed that they were
playing the game simultaneously with another partner. After making their decision, participants
were shown both their own choice and that of their partner. We performed a generalized linear
mixed model (GLMM1) analysis (see Appendix *Table 1*) to examine the effects of each independent
variable and their interactions on the decision to cooperate or defect.

Consistent with most previous studies (*Belli et al., 2012*; *Nava et al., 2023*; *Taheri et al., 2018*),
adolescents cooperated less than adults (b of group = 0.79, 95% CI = [0.311, 1.270], $p = 0.001$;
*Figure 1*d). Following the interaction of group $\times$ previous trial $\times$ partner's choice (b of interaction
= 0.24, 95% CI = [0.126, 0.361], $p < 0.001$), we found that adolescents showed significantly less
cooperation compared to adults only after the partner's cooperation ($t(259)_{\mathrm{group}} = -2.84$, $p = 0.005$,
BF10 = 6.01).

However, such a difference was not significant after the partner's defection ($t(259)_{\mathrm{group}} = -1.86$,
$p = 0.064$, BF10 = 0.69; *Figure 1*d). We also found that adults increased cooperation in response
to their partners' consistent cooperation (the partner cooperated once vs. the partner cooperated
thrice: $t(266)_{\mathrm{adults}} = -2.50$, $p = 0.013$, BF10 = 2.56), but this pattern was not observed in adolescents
($t(252)_{\mathrm{adolescents}} = -1.18$, $p = 0.239$, BF10 = 0.27, see *Figure 1*d).

Nevertheless, both groups significantly decreased cooperation in response to the partner's
continual defection (the partner defected once vs. the partner defected twice: $t(266)_{\mathrm{adults}} = 4.46$,
$p < 0.001$, BF10 $> 10^3$, $t(252)_{\mathrm{adolescents}} = 2.78$, $p = 0.006$, BF10 = 5.21; the partner defected once vs.
the partner defected thrice: $t(266)_{\mathrm{adults}} = 5.56$, $p < 0.001$, BF10 $> 10^3$ for adults, $t(252)_{\mathrm{adolescents}} = 4.32$,
$p < 0.001$, BF10 = 761.12 for adolescents, *Figure 1*e).

Figure 1. Experiment setup and behavioral results. (**a**) Partner's Cooperation Probability: In the first half of the 120 trials, the partner cooperated 78% of the time; in the second half, cooperation alternated between 20% and 80%. (**b**) Payoff Matrix: Payoffs are 4 for mutual cooperation, 2 for mutual defection, 0 for cooperation when the other defects, and 6 for defecting when the other cooperates. (**c**) Trial Illustration: After a 0.5-second fixation, participants choose a shape (triangle for cooperation, square for defection) within 4 seconds and see both players' choices for 1.5 seconds. (**d-e**) Post hoc Comparisons: Figures 1d and 1e show the participants' cooperation probability on the y-axis. The x-axis represents the consistency of the partner's actions in previous trials (t_{-1}: last trial, $t_{-1,2}$: last two trials, $t_{-1,2,3}$: last three trials). Large red (adolescents) and blue (adults) dots indicate mean probabilities, with black error bars for standard error (SE). Gray dots represent mean probabilities across trials, and green error bars show predicted cooperation rates with SE. Notes: *n.s.* $p > 0.05$; *$p < 0.05$; **$p < 0.01$; ***$p < 0.001$.

Asymmetric RL learning in the social reward model best explains cooperative decisions of adolescents and adults

Computational modeling was used to simulate participants' mental processes during the rPDG. Starting with a baseline model that assumed decisions were made through random selection (Model

1), we compared several alternatives: a win-stay and loss-shift model (Model 2), a reward learning model (Model 3), an inequality aversion model (Model 4), and a social reward model (Model 5). Among these, the social reward model outperformed the others. We then compared a basic RL algorithm (Model 6), an influence learning rule (Model 7), and an asymmetric RL learning rule (Model 8) within the social reward framework. The asymmetric RL learning model best explained the cooperative decisions of both adolescents and adults (see *Figure 2*a for adolescents and *Figure 2*b for adults; methods for details). Model recovery analysis indicated that the asymmetric RL learning within the social reward model was distinguishable from the other models (*Figure 2*c) and accurately captured the behaviors of both adolescents (*Figure 2*d) and adults (*Figure 2*e). The overlap between Models 4 and 5 likely arises because neither model incorporates a learning mechanism, making them less able to account for trial-by-trial adjustments in this dynamic task. For further validation of the best-fitting model, see *Figure 2—figure Supplement 1* for model predictions, *Figure 2—figure Supplement 2* for the distributions of free parameters, *Figure 2—figure Supplement 3* for parameter recovery, *Figure 2—figure Supplement 4* for partial correlation matrices among parameters, and *Figure 2—figure Supplement 5* for group-level posterior distributions from the hierarchical Bayesian estimation for the best-fitting model.

Figure 2. Computational modeling. (**a-b**) Model comparisons for adolescents and adults, respectively. The y-axis represents model fitness based on the Akaike Information Criterion with a correction for sample size (AICc) (*Hurvich and Tsai, 1989*). For each participant, the model with the lowest AICc served as a reference to compute ΔAICc by subtracting it from the AICc of other models (ΔAICc = AICc$_x$ - AICc$_{lowest}$). A lower ΔAICc indicates a better model fit. Protected exceedance probability (PEP) is a group-level measure that assesses the likelihood of each model's superiority over the others (*Rigoux et al., 2014*). (**c**) Model recovery analysis. Each model was used to generate 100 synthetic datasets, and for each dataset, model fitting and comparison were performed. Each column corresponds to one generative model, and each row corresponds to one fitting model. The color in each cell indicates the probability that the synthetic datasets generated by the model in the column were best fit by the model in the row, with a darker color denoting a higher probability. (**d-e**) Model prediction. Sample illustration of the best-fitting model prediction versus data for adolescents and adults, respectively.

Figure 2—figure supplement 1. Model prediction. This figure compares the actual data and presupplement 1 (**a**) and adolescents (**b**). The x-axis represents the trial number, while the y-axis represents the mean cooperation probability of all participants. The shaded area indicates the 95% confidence interval.

Figure 2—figure supplement 2. Distributions of estimated parameters from the best-fitting model. Each panel displays one parameter. The histograms and their kernel fits are represented by color bars and curves, respectively. Red indicates participants in the adolescent sample, and blue denotes those in the adult sample. Parameters have been transformed into a log scale for enhanced visualization.

Figure 2—figure supplement 3. Parameter recovery for the best-fitting model. Each panel represents one parameter. Each dot corresponds to one virtual participant. The value of *r* indicates Pearson's correlation coefficient between the true values (estimated from the participants) and the recovered parameters.

Figure 2—figure supplement 4. Partial correlation matrices among parameters for the best-fitting model. The upper-triangular cells show partial correlations for adults, and the lower-triangular cells show partial correlations for adolescents. Each cell shows the partial Pearson correlation coefficient (controlling for the other parameters). Colors range from green (negative) to violet (positive), with the color bar spanning [-1,1]. Notes: *n.s.* $p > 0.05$; *$p < 0.05$; **$p < 0.01$; ***$p < 0.001$.

Figure 2—figure supplement 5. Group-level posterior distributions from the hierarchical Bayesian estimation for adolescents and adults. Posterior densities are shown separately for adolescents (red) and adults (blue). Δ values indicate the posterior mean difference (Adult – Adolescent) with 95% credible intervals (CrI) and Bayesian *p* values. Compared with adolescents, adults exhibited higher positive learning rates ($\alpha+$) and lower negative learning rates ($\alpha-$), suggesting greater differentiation between learning from positive and negative feedback. Adults also showed lower inverse temperature (β), indicating more exploratory decision behavior, and higher social reward weight (ω), reflecting greater valuation of reciprocity or social outcomes. Notes: *n.s.* $p > 0.05$; *$p < 0.05$; **$p < 0.01$; ***$p < 0.001$.

Distinct learning rates and social preferences between adolescents and adults in repeated cooperation

Although the asymmetric RL learning in the social reward model best explained the behaviors of both adolescents and adults, the two groups exhibited distinct learning dynamics and social preferences for cooperation. Specifically, adolescents applied a higher positive learning rate ($\alpha+$, $t(259) = 2.95$, $p = 0.003$, BF10 = 8.02, *Figure 3*a) to update better-than-expected prediction errors, and a lower negative learning rate ($\alpha-$, $t(259) = -2.62$, $p = 0.009$, BF10 = 3.46, *Figure 3*b) for worse-than-expected prediction errors.

Additionally, a positive correlation was found between participants' age and the negative learning rate ($\alpha-$, $r = 0.21$, $p < 0.001$, *Figure 3*f), while no significant correlation was observed with the positive learning rate ($\alpha+$, $r = -0.09$, $p = 0.16$, *Figure 3*e).

Furthermore, adolescents displayed a weaker preference for cooperation compared to adults (ω, $t(259) = -3.03$, $p = 0.003$, BF10 = 9.92, *Figure 3*c), and their social preferences for cooperation increased with age ($r = 0.20$, $p < 0.001$, *Figure 3*g). Additionally, adolescents exhibited a higher inverse temperature parameter compared to adults, indicating they were more sensitive to utility differences between cooperation and defection (β, $t(259) = 2.14$, $p = 0.034$, BF10 = 1.17, *Figure 3*d). This sensitivity decreased with age, as shown by a negative correlation with age ($r = -0.17$, $p = 0.007$, *Figure 3*h).

Figure 3. Learning Rates and Social Preferences. (**a-d**) Comparison between adolescents and adults for positive learning rate ($\alpha+$), negative learning rate ($\alpha-$), social preference (ω), and inverse temperature (β), respectively. (**e-h**) Correlation between age and positive learning rate, negative learning rate, social preference, and inverse temperature, respectively. Notes: *$p < 0.05$; **$p < 0.01$.

Adolescents compared to adults show no inappropriate expectations but less intrinsic reward for reciprocity

To further explore what underlies the observed decrease in cooperation among adolescents, we focused on two hidden trial-by-trial updating variables: the partner cooperation expectation (p) and the intrinsic reward for reciprocity ($p×\omega$). Additionally, participants' self-reported cooperativeness scores, assessed every 15 trials, provided further insight into their subjective estimation of the partner's willingness to cooperate.

Partner cooperation expectation

We performed a linear mixed model (LMM1, Appendix *Table 2*) on partner cooperation expectation to assess the effects of each independent variable and their interactions. Following the interaction of group × previous trial × partner's choice (b of interaction = 0.03, 95% CI = [0.022, 0.038], $p < 0.001$), we found that the partner cooperation expectation for both adolescents and adults increased with the partner's consistent cooperation (the partner cooperated once vs. the partner cooperated twice: $t(252)_{adolescents} = -2.81$, $p = 0.005$, BF10 = 5.75, $t(266)_{adults} = -4.45$, $p < 0.001$, BF10 > 10^3; the partner cooperated once vs. the partner cooperated thrice: $t(252)_{adolescents} = -3.69$, $p < 0.001$, BF10 =78.00, $t(266)_{adults} = -6.23$, $p < 0.001$, BF10 > 10^3; *Figure 4*a). Additionally, expectations decreased with the partner's consistent defection (the partner cooperated once vs. the partner cooperated twice: $t(252)_{adolescents} = 4.44$, $p < 0.001$, BF10 > 10^3, $t(266)_{adults} = 7.02$, $p < 0.001$, BF10 > 10^3; the partner cooperated once vs. the partner cooperated thrice: $t(252)_{adolescents} = 5.60$, $p < 0.001$, BF10 > 10^3, $t(266)_{adults} = 8.40$, $p < 0.001$, BF10 > 10^3; *Figure 4*b). These results showed that both adolescents and adults held very similar expectations toward their partner's cooperation and did not have significant differences between the groups (b of group = -0.04, 95% CI = [-0.102, 0.021], $p = 0.198$).

Moreover, we performed a LMM2 (Appendix *Table 3*) analysis on participants' self-reported scores regarding the cooperativeness of their partners to examine the effects of each independent variable and their interactions. In line with the expectation of partner cooperation, we observed minimal discrepancy in the self-reported scores on partner cooperativeness between adolescents and adults. Neither the main effect of group nor the interaction achieved statistical significance (b of group = 0.17, 95% CI = [-0.51, 0.85], $p = 0.616$; b of interaction = 0.38, 95% CI = [-0.052, 0.812], $p = 0.085$; *Figure 4*c-d). These results provide evidence that adolescents did not differ from adults in assessing their partner's cooperation.

Intrinsic reward for reciprocity

We performed a LMM3 (Appendix *Table 4*) on the intrinsic reward for reciprocity to assess the effects of each independent variable and their interactions. We found that adolescents appreciated reciprocity less than adults did (b of group = 0.52, 95% CI = [0.224, 0.816], $p < 0.001$).

Following the interaction of group × previous trial × partner's choice (b of interaction = 0.37, 95% CI = [0.318, 0.424], $p < 0.001$), unlike adults, adolescents did not increase their intrinsic reward for reciprocity in response to the partner's consistent cooperation (the partner cooperated once vs. the partner cooperated twice: $t(252)_{adolescents} = -0.96$, $p = 0.336$, BF10 = 0.21, $t(266)_{adults} = -2.13$, $p = 0.034$, BF10 = 1.15; the partner cooperated once vs. the partner cooperated thrice: $t(252)_{adolescents} = -1.38$, $p = 0.170$, BF10 = 0.34, $t(266)_{adults} = -3.08$, $p = 0.002$, BF10 = 11.63; *Figure 4*e).

Moreover, under consistent defection by the partner, evidence for the one-versus-two last trials comparison was inconclusive in adolescents but supported a decrease in adults ($t(252)_{adolescents} = 1.99$, $p = 0.047$, BF10 = 0.90, $t(266)_{adults} = -2.71$, $p = 0.007$, BF10 = 4.27). Importantly, in the one-versus-three last trials comparison, both adolescents and adults consistently showed a decrease in intrinsic reward for reciprocity ($t(252)_{adolescents} = 2.64$, $p = 0.009$, BF10 = 3.66, $t(266)_{adults} = 3.37$, $p < 0.001$, BF10 = 27.47; *Figure 4*f).

In brief, adolescents did not deviate in forming expectations about their partner's willingness to cooperate, but they showed lower social preferences for cooperation and a reduced intrinsic reward for reciprocity. Specifically, compared to adults, adolescents displayed less intrinsic reward for reciprocity and did not increase it in response to consistent cooperation, though their reactions to consistent defection tended to be similar to those of adults.

Figure 4. Analysis of hidden variables from the best-fitting model. (**a-b**) Post-hoc Comparison of LMM1: Interaction of group × previous trial × partner's choice. The y-axis shows participants' expectations of partner cooperation probability (p) from the best-fitting model. (**c-d**) Self-Reported Cooperativeness: Normalized scores on partner cooperativeness for two orders of partner cooperation probability, with adolescents (orange-red line) and adults (blue line). Scores were assessed on a 0-9 scale and normalized to 0-1. The dotted line indicates the presumed partner's cooperation probability, with mean values and standard errors shown. (**e-f**) Post-hoc Comparison of LMM3: Interaction of group × previous trial × partner's choice. The y-axis shows participants' intrinsic reward for reciprocity ($p \times \omega$) from the best-fitting model. The x-axis represents the consistency of the partner's actions in previous trials (t_{-1}: last trial, $t_{-1,2}$: last two trials, $t_{-1,2,3}$: last three trials). Colored dots with error bars indicate mean values with standard errors for adolescents (orange-red) and adults (blue), while small gray dots represent individual participants. Notes: *n.s.* $p > 0.05$; $^*p < 0.05$; $^{**}p < 0.01$; $^{***}p < 0.001$.

Discussion

Cooperation lies at the heart of societal functioning, facilitating the achievement of shared goals and fostering social harmony. In this study, we sought to deepen our understanding of the developmental aspects of cooperation by examining differences in cooperative behavior between adolescents and adults in the context of the rPDG. Our findings shed light on the cognitive and affective processes underlying these behaviors, offering insights into the mechanisms driving cooperative decision-making across different developmental stages.

Consistent with many previous studies (*Fett et al., 2014*; *Gutiérrez-Roig et al., 2014*; *Westhoff et al., 2020*), our results showed that adolescents exhibited lower levels of cooperation compared to adults. However, such lower cooperation was not generally observed during the task, but selectively occurred after their partner cooperated in the previous rounds. Moreover, our results showed that adults increased cooperation in response to their partner's consistent cooperation, such a pattern was not observed in adolescents. However, both age groups decreased cooperation in response to consistent partner defection, indicating shared responses to non-cooperative behavior.

Our results suggest that the lower levels of cooperation observed in adolescents stem from a stronger motive to prioritize self-interest rather than a deficiency in predicting others' cooperation in social learning. In both, the expectation of partner's cooperation estimated from computational modeling and the self-reported measurements, adolescents did not exhibit significant differences from adults. However, adolescents exhibited a weaker preference for (conditional) cooperation compared to adults, resulting in a reduced intrinsic reward for reciprocity. The results are consistent with prior research (*Crone and Dahl, 2012*; *Do et al., 2017*; *Pfeifer and Berkman, 2018*; *Van Den Bos et al., 2010*, *2011*), suggesting that adolescents prioritize immediate gains over long-term benefits, potentially undermining the benefits of cooperation. This tendency aligns with earlier findings that adolescents exhibit heightened sensitivity to reward feedback (*Blakemore and Mills, 2014*; *Crone and Dahl, 2012*; *Davis et al., 2023*; *Do et al., 2017*; *Van Den Bos et al., 2011*; *Van Duijvenvoorde et al., 2015*), which may influence their decision-making in cooperative interactions. Overall, these findings indicate that adolescents' lower cooperation is unlikely to be driven solely by strategic considerations, but may instead reflect differences in the valuation of others' cooperation or reduced motivation to reciprocate. Although defection is the payoff-dominant strategy in the PDG, the selective pattern of adolescents' cooperation and the model comparison results indicate that their reduced cooperation cannot be fully explained by strategic incentives, but rather reflects weaker valuation of social reciprocity.

It has been acknowledged that individuals update positive and negative outcomes by different weights in social cooperation, such asymmetric learning process can be modeled by basic RL algorithm with both positive and negative learning rates (*Garrett and Daw, 2020*; *Rosenbaum et al., 2022*). In this study, we find that an asymmetrical RL algorithm in a social reward model provided best model fits of the behaviors of both, adolescents and adults. Adolescents demonstrated a larger positive learning rate, but a smaller negative learning rate compared to adults, suggesting heightened sensitivity to positive feedback from cooperative behavior and reduced sensitivity to negative feedback from defection. This asymmetrical learning pattern may drive adolescents to focus more on self-beneficial social signals, maximizing immediate gains in response to cooperative behavior. These findings align with (*Van Den Bos et al., 2011*), which highlight adolescents' heightened sensitivity to immediate rewards and less stable trusting behavior compared to adults. Adolescents also showed higher inverse temperature values (β), indicating greater sensitivity to expected value and more value-based choice behavior. Together, these findings suggest that the differentiation between positive and negative learning rates changes with age, reflecting more selective feedback sensitivity in development, while higher (β) values in adolescents indicate greater value sensitivity. This interpretation remains tentative and requires further validation in future research.

Adolescence is characterized by increased self-discovery and egocentrism (*Pfeifer and Berkman, 2018*; *Ting et al., 2019*), leading individuals to prioritize immediate gains over long-term benefits. Consistent with this, the higher value sensitivity (β) observed in adolescents suggests a stronger focus on immediate utility during cooperative exchanges. Consequently, adolescents may be more inclined toward self-serving motives in sustained social interactions (*Pfeifer and Berkman, 2018*). However, these tendencies are not static; as individuals mature into adulthood, their socio-emotional capacities continue to develop (*Worthman and Trang, 2018*), enabling a more balanced integration of short-term rewards and long-term social outcomes (*Crone and Dahl, 2012*; *Wu et al., 2023*).

It is important to note some limitations of this study. First, we used artificial opponents with pre-determined cooperation patterns to better control the stimuli. While this approach allowed us to isolate specific motivations for cooperation (financial vs. social rewards), it's possible that participants might behave differently in more natural settings. Our study serves as an initial step in understanding cooperation motivations in adolescents and adults, and future research could explore these behaviors in more real-world contexts. Second, our study employed the rPDG as the primary task to directly capture cooperation in symmetric multi-round interactions. However, because it is a zero-sum framework that structurally incentivizes defection as the dominant strategy, the rPDG may influence choices beyond participants' intrinsic preferences. For example, one potential interpretation of adolescents' lower cooperation is that they adopt a strategic response to the payoff structure, through leveraging defection as the more rewarding strategy within the game. If this account holds, adolescents should exhibit lower cooperation across all rounds. However, we find that adolescents and adults exhibit similar behavioral patterns when partners defect. By contrast, adolescents cooperate less than adults when partners cooperate, and their cooperation does not increase significantly even when partners cooperate consecutively. Although this pattern is consistent with the interpretation that adolescents' lower cooperation reflects a relatively more self-interested motivation, stronger conclusions about age differences in cooperative preferences require further examination in tasks with varied structures. Third, although both age groups were recruited from Beijing and nearby regions, minimizing major regional and cultural variation, adolescents and adults may still differ in socioeconomic status, financial independence, and social experience. Such contextual differences could interact with developmental processes in shaping cooperative behavior and reward valuation. Future research with demographically matched samples or explicit measures of socioeconomic background will help disentangle biological from sociocultural influences.

In conclusion, our study contributes to an understanding of the developmental aspects of cooperation and the cognitive-affective processes underlying cooperative decision-making. By examining differences in cooperative behavior between adolescents and adults in the rPDG and integrating computational modeling, we offer valuable insights into the mechanisms driving cooperative behavior across different developmental stages. These findings have implications for promoting prosocial behaviors and designing effective socialization interventions during adolescence. By highlighting the importance of reciprocity, our findings offer insights into the developmental trajectory of cooperation from adolescence to adulthood and provide practical implications for enhancing cooperative interactions in real-world contexts.

Methods and Materials

Participants

A total of 261 participants took part in the current study, consisting of 127 adolescents (n = 127, aged 14-17 years, mean ± SD: 16.13 ± 0.63, 44 females) and 134 adults (n = 134, aged 18-30 years, mean ± SD: 21.63 ± 2.88, 79 females). No a priori power analysis was conducted. The sample size was determined based on previous studies investigating cooperation behaviors in adolescents and adults. Adolescents were recruited from the local high school with the consent of their legal

guardians, while adults were recruited through advertisements on the university campus forum. Before the experiment, written informed consent was obtained from all participants and their legal guardians (for adolescents). Participants were included if they had normal or corrected-to-normal vision and no history of psychiatric or neurological illness. Exclusion criteria included any self-reported diagnosed psychiatric or neurological disorder. This study was approved by the ethics committee of Beijing Normal University. No participants dropped out of the experiment, and all data were included in the statistical analysis. Participants received compensation based on their performance in the tasks (see rPDG for details).

Experimental procedure

All participants completed the experiments in a laboratory setting with multiple participants present. They were informed that they were participating in a multiple-round interaction game with an anonymous partner. In the instructions section, we referred to the interaction game as the rPDG and refrained from using the terms 'cooperate' and 'defect' to minimize the influence of social expectations, biases and promote comparability across studies. Participants were instructed to believe that their partner was also playing the game at the same time. Compensation for their participation was based on the tokens earned during the game, with 10% of the rounds randomly selected for payment calculation at an exchange rate of 1 token to 1 yuan. Participants were explicitly informed in advance about this incentive mechanism. Prior to the formal experiment, participants underwent a quiz and several practice rounds to ensure a full understanding of the task. Following the experiment, participants completed a Social Value Orientation (SVO) task to assess their prosocial personality traits. The entire procedure lasted approximately 60 minutes. Blinding was not applicable in this study, as all participants interacted with a computer-controlled partner. To minimize potential bias, the partner's behavior patterns and stimulus meanings were randomized across participants. A detailed protocol is available upon request.

The repeated prisoner's dilemma game

Similarly to the classic version of PDG, rPDG involves two players. Consistent with the standard payoff matrix of the PDG (*Figure 1*b), when both players cooperated (defected), they each received 4 tokens (2 tokens). If the players made different decisions, the one who cooperated received 0 tokens, while the one who defected received 6 tokens. Participants were told that their partner was another human participant in the laboratory and that they would interact with the same partner across all rounds. However, in reality, the actions of the partner were predetermined by a computer program. This setup allowed for a clear comparison of the behavioral responses between adolescents and adults. Participants were not informed of the total number of rounds in the rPDG. In order to enhance the realism of the partner's response, we manipulated the variability in the partner's decision making. The partner's cooperation probability remained stable at 78% for half of the trials. In the other half of the trials, the partner's cooperation probability varied, switching between 20%, 80%, and 20% for each set of 20 trials. The order of these two sessions was counterbalanced between participants. During the rPDG, participants were asked every 15 rounds to evaluate their partner's cooperativeness using a 10-point scale, where 0 represents "no cooperation" and 9 represents "very high cooperation." The question posed to the participants was "How cooperative do you think your partner is at the moment?"

Behavioral data analysis

All statistical analyses were conducted in MATLAB R2023a (RRID:SCR 001622). GLMM was implemented using the 'fitglme' function in MATLAB. Interaction contrasts were performed for significant interactions and, when higher-order interactions were not significant, pairwise or sequential contrasts were performed for significant main effects. Post hoc comparisons were conducted using Bayes factor analyses with MATLAB's bayesFactor Toolbox version v3.0, with a Cauchy prior scale $\sigma = 0.707$ (*Krekelberg, 2024*).

₃₆₃   GLMM1: Participant's choices (cooperate or defect) of all trials are the dependent variable; fixed
₃₆₄ effects include an intercept, the main effects of group (adolescents or adults), previous trial (last
₃₆₅ one trial, last two trials and last three trials), partner's choice (cooperation or defection), and all
₃₆₆ possible interaction effects of the independent variables. Gender (male and female) and timing
₃₆₇ (trial number from 1 to 120) were also included as the control variables. Random effects include
₃₆₈ correlated random slopes of group, previous trial, partner's choice, gender, trial number and ran-
₃₆₉ dom intercept for participants. The group, previous trial, partner's choice, gender are the category
₃₇₀ variables. The trial number is a continuous variable. See Appendix *Table 1* for the statistical results
₃₇₁ of GLMM1.

₃₇₂   LMM2: Participants' self-reported score on partner's cooperativeness is the dependent vari-
₃₇₃ able; fixed effects include an intercept, the main effects of group (adolescents or adults), the order
₃₇₄ of the sessions (regarding the partner's cooperation involved fixed 78% cooperation probability,
₃₇₅ followed by shifting into 20%, 80%, and 20% for each 20, or vice versa), the interaction of group ×
₃₇₆ order . Gender (male and female) and timing (trial number from 1 to 120) were also included as the
₃₇₇ control variables. Random effects include correlated random slopes of group, gender, timing and
₃₇₈ random intercept for participants. The group, previous trial, partner's choice, gender are the cate-
₃₇₉ gory variables. The trial number is a continuous variable. See Appendix *Table 3* for the statistical
₃₈₀ results of LMM2.

₃₈₁ **Behavioral modelling**
₃₈₂ We systematically developed models based on various assumptions regarding participants' decision-
₃₈₃ making processes in the rPDG.

₃₈₄ ## Model 1: The baseline model
₃₈₅ We modeled each participant's choices in each trial (i.e., whether to cooperate) as outcomes from
₃₈₆ a Bernoulli distribution, where the cooperation probability is controlled by a parameter, $b \in [0, 1]$.
₃₈₇ For each participant, the probabilities of cooperation ($q(\text{cooperation})$) and defection ($q(\text{defection})$)
₃₈₈ are denoted as follows:

$$q(\text{cooperation}) = b \tag{1}$$

$$q(\text{defection}) = 1 - b \tag{2}$$

₃₈₉ ## Model 2: Win-stay & loss-shift model
₃₉₀ The model assumes that individuals adopt a tit-for-tat strategy in decision-making. Participants
₃₉₁ are likely to repeat their previous choice with a probability of $1 - \frac{\epsilon}{2}$ if they won, and $\frac{\epsilon}{2}$ if they lost
₃₉₂ in the last trial, where ϵ represents the choice variability. Winning and losing are defined based on
₃₉₃ the payoff outcomes of 4 or 6 (win) and 0 or 2 (loss), respectively.

$$q_{t+1} = q_t \left(1 - \frac{\epsilon}{2}\right) \delta + q_t \left(\frac{\epsilon}{2}\right)(1 - \delta) \tag{3}$$

₃₉₄   where q_t denotes the probability of repeating the previous choice and $1 - q_t$ denotes the prob-
₃₉₅ ability of shifting to another option at trial t.

₃₉₆ ## Model 3: Reward learning model
₃₉₇ This model assumes that participants make decisions by comparing the values of choosing coop-
₃₉₈ eration and defection. The values of the two options are updated using a RL algorithm:

$$V_{c_{t+1}} = V_{c_t} + \alpha(R_t - V_{c_t}) \tag{4}$$

$$V_{d_{t+1}} = V_{d_t} + \alpha(R_t - V_{d_t}) \tag{5}$$

where V_c (V_d) denotes the value of cooperation (defection) option. R represents the reward feedback, which can be 0, 2, 4, or 6, depending on the payoff matrix. $R_t - V_{c_t}$ ($R_t - V_{d_t}$) represents the reward prediction error for the cooperation (defection) option, and α is the learning rate. Participants' choices are modeled by a softmax function:

$$q(\text{cooperate})_t = \frac{1}{1 + e^{\beta(V_{d_t} - V_{c_t})}} \tag{6}$$

where q_t denotes the participants' probability of cooperation and β denotes the inverse temperature. The lower the value of the inverse temperature, the greater the sensitivity to the different values between options.

Model 4: Inequality aversion model

The model assumes that participants' decisions aim to reduce both disadvantageous and advantageous inequality between themselves and their partners:

$$U_{c_t} = c_{\text{self}} - \varphi \max(c_{\text{other}} - c_{\text{self}}, 0) - \nu \max(c_{\text{self}} - c_{\text{other}}, 0) \tag{7}$$

$$U_{d_t} = d_{\text{self}} - \varphi \max(d_{\text{other}} - d_{\text{self}}, 0) - \nu \max(d_{\text{self}} - d_{\text{other}}, 0) \tag{8}$$

where U_c (U_d) denotes the utility of cooperation (defection). φ represents aversion to disadvantageous inequality and ν represents aversion to advantageous inequality. c_{self} and c_{other} denote the expected payoffs for cooperation to oneself and the partner, respectively, while d_{self} and d_{other} denote the expected payoffs for defection to oneself and the partner. p denotes participants' partner cooperation expectation. The model assumes that participants did not update the inferred cooperation probability based on feedback; p is fixed at 0.5.

Based on the payoff matrix, the payoffs for participants and their partners are calculated using the following functions:

$$c_{\text{self}} = 4p \tag{9}$$

$$c_{\text{other}} = 6 - 2p \tag{10}$$

$$d_{\text{self}} = 4p + 2 \tag{11}$$

$$d_{\text{other}} = 2 - 2p \tag{12}$$

Participants' choices are modeled by a softmax function:

$$q(\text{cooperate})_t = \frac{1}{1 + e^{\beta(U_{d_t} - U_{c_t})}} \tag{13}$$

where q_t denotes the participants' probability of cooperation and β denotes the inverse temperature. The lower the value of the inverse temperature, the greater the randomness in decisions.

Model 5: Social reward model

The model assumes that participants make decisions by comparing the expected payoff of cooperation and defection based on the payoff matrix and an additional subjective bonus from cooperation:

$$U_{c_t} = p(4 + \omega) \tag{14}$$

$$U_{c_t} = 4p + 2 \tag{15}$$

where ω represents an additional social reward associated with cooperation.

Model 6: Social reward model with RL algorithm

The model, building on Model 5, assumed that participants update their expectations of partner cooperation trial-by-trial, based on the partner's previous decisions, using a basic RL algorithm:

$$U_{c_t} = p_t(4 + \omega) \tag{16}$$

$$U_{d_t} = 4p_t + 2 \tag{17}$$

where p_t denotes participants' expectation of partner cooperation probability at trial t and is updated by the following function:

$$p_{t+1} = p_t + \alpha(P_t - p_t) \tag{18}$$

where α is the learning rate applied to the prediction error, $(P_t - p_t)$. P_t represents the partner's decision at trial t, equating to 1 if the partner cooperates and 0 if the partner defects.

Model 7: Social reward model with influence model

The model is based on Model 6 and includes an additional assumption that participants update their expectation of the partner's cooperation by considering not only the partner's previous decisions but also the influence of their own previous decisions on the partner's subsequent decisions. This aspect is referred to as second-order belief and is updated by the following function:

$$p_{t+1} = p_t + \alpha(P_t - p_t) + \kappa(Q_t - q_t') \tag{19}$$

$$q_t' = \frac{2}{\omega} \ln\left(\frac{1}{p_t} - 1\right) \frac{1}{\beta\omega} \tag{20}$$

where Q_t represents the participants' decision at trial t, equating to 1 if the participants cooperate and 0 if participants defect. q_t' represents the participants' inferred cooperation probability of themselves from the partner's perspective in trial t, which was inferred from function 13. Therefore, $(Q_t - q_t')$ denotes the second-order prediction error, and κ is the second-order learning rate that governs the updating of second-order belief.

Model 8: Social reward model with asymmetric RL rule

The model, based on Model 6, assumes that participants asymmetrically update positive expectation errors (better than expected) and negative prediction errors (worse than expected) using two distinct learning rates:

$$p_{t+1} = p_t + \alpha_+ \delta(PE) + \alpha_-(1 - \delta)PE \tag{21}$$

$$PE = P_t - p_t \tag{22}$$

$$\delta = \begin{cases} 1, & \text{if } PE > 0 \\ 0, & \text{if } PE < 0 \end{cases} \tag{23}$$

Model fitting and model comparison

We used maximum likelihood estimation to fit models to each participant's choices across all trials. The likelihood function, based on the binomial distribution, captured the association between each participant's choices and each model's predictions. To minimize the negative log-likelihood, we employed MATLAB's (MathWorks) *fmincon* function. To enhance the likelihood of finding the global minimum, we repeated the parameter search process 500 times, using different starting points. In

addition, we tested Model 9 (social reward model with Pearce–Hall learning, i.e., dynamic learning rate; see Appendix Analysis for details; and also see *Figure Supplement 6*).

For model evaluation, we first used the AICc, which accounts for the model's complexity and the number of observed data points(*Hurvich and Tsai, 1989*). The second metric was the protected exceedance probability from group-level Bayesian model selection (Rigoux et al., 2014), providing a measure of the likelihood that a specific model is superior to other models under consideration. We chose to use the AIC as the metric of goodness-of-fit for model comparison for the following statistical reasons. First, BIC is derived based on the assumption that the "true model" must be one of the models in the limited model set one compares (*Burnham and Anderson, 2002*; *Gelman and Shalizi, 2013*), which is unrealistic in our case. In contrast, AIC does not rely on this unrealistic "true model" assumption and instead selects out the model that has the highest predictive power in the model set (*Gelman et al., 2014*). Second, AIC is also more robust than BIC for finite sample size (*Vrieze, 2012*).

The log-likelihood is calculated as the following function:

$$L = \sum_{t=1}^{T} \log(q_t | \alpha+, \alpha-, \omega, \beta) \tag{24}$$

where q_t represents the probability of participants' decision at trial t, equating to $q(\text{cooperation})$ if participants cooperate and $q(\text{defection})$ if participants defect.

Model identifiability and parameter recovery analyses

We performed a model identifiability analysis to ensure that model comparisons were not compromised by model misidentification. For each model, we generated synthetic datasets using parameters estimated from the data of all participants. We then fitted each alternative model to its corresponding synthetic dataset and identified the best-fitting model through model comparison. To test robustness, we repeated this procedure 100 times, calculating the percentage of instances where each model was recognized as the best model across all synthetic datasets generated by that specific model. Consistently high percentages indicated model identifiability. Additionally, we assessed parameter recovery for the best-fitting model (model 8: social reward model with an asymmetric RL rule). This assessment involved calculating the Pearson correlation between the parameters estimated from the 100 synthetic datasets (recovered parameters) and the parameters used to generate these datasets. A higher correlation coefficient between the recovered and the estimated parameters suggested non-redundancy in the parameter space (*Figure 2—figure Supplement 3*).

Hidden mental variables analysis

LMM1: Participants' expectation of partner's cooperation probability that estimated from the winning model, the variable *p*, is the dependent variable. The fixed and random effects remain the same as GLMM1. See Appendix *Table 2* for the statistical results of LMM1.

LMM3: Participants' intrinsic reward for reciprocity that estimated from the winning model, *p×ω*, are the dependent variable. The fixed and random effects remain the same as GLMM1. See Appendix *Table 4* for the statistical results of LMM3.

Data analysis

All analyses were performed with custom code in MATLAB, version R2023a (RRID:SCR 001622). All scripts used for analysis are available at https://doi.org/10.5281/zenodo.15046430.

Acknowledgments

This project has received funding from the Brain Science and Brain-like Intelligence Technology - National Science and Technology Major Project (2021ZD0200500), the National Natural Science

Foundation of China (32441109, 32271092, 32130045), the Open Research Fund of the State Key Laboratory of Cognitive Neuroscience and Learning (CNLYB2404), the Beijing Major Science and Technology Project under Contract (Z241100001324005), and the Opening Project of the State Key Laboratory of General Artificial Intelligence (SKLAGI20240P06). We thank Christian C. Ruff and Xiangjuan Ren for insightful discussions.

References

Andreoni J, Miller JH. Rational cooperation in the finitely repeated prisoner's dilemma: Experimental evidence. The economic journal. 1993; 103(418):570–585. doi: https://doi.org/10.2307/2234532.

Axelrod R, Hamilton WD. The evolution of cooperation. science. 1981; 211(4489):1390–1396. doi: https://doi.org/10.1126/science.7466396.

Belli SR, Rogers RD, Lau JY. Adult and adolescent social reciprocity: Experimental data from the Trust Game. Journal of adolescence. 2012; 35(5):1341–1349. doi: https://doi.org/10.1016/j.adolescence.2012.05.004.

Blakemore SJ, Mills KL. Is adolescence a sensitive period for sociocultural processing? Annual review of psychology. 2014; 65(1):187–207. doi: https://doi.org/10.1146/annurev-psych-010213-115202.

Burnham KP, Anderson DR. Model selection and multimodel inference: A practical information-theoretic approach. 2nd ed. New York, NY: Springer; 2002. doi: https://doi.org/10.1007/b97636.

Crone EA, Dahl RE. Understanding adolescence as a period of social–affective engagement and goal flexibility. Nature reviews neuroscience. 2012; 13(9):636–650. doi: https://doi.org/10.1038/nrn3313.

Davis MM, Modi HH, Skymba HV, Finnegan MK, Haigler K, Telzer EH, Rudolph KD. Thumbs up or thumbs down: neural processing of social feedback and links to social motivation in adolescent girls. Social Cognitive and Affective Neuroscience. 2023; 18(1):nsac055. doi: https://doi.org/10.1093/scan/nsac055.

Do KT, Moreira JFG, Telzer EH. But is helping you worth the risk? Defining prosocial risk taking in adolescence. Developmental cognitive neuroscience. 2017; 25:260–271. doi: https://doi.org/10.1016/j.dcn.2016.11.008.

Embrey M, Fréchette GR, Yuksel S. Cooperation in the finitely repeated prisoner's dilemma. The Quarterly Journal of Economics. 2018; 133(1):509–551. doi: https://doi.org/10.1093/qje/qjx033.

Fareri DS, Chang LJ, Delgado MR. Effects of direct social experience on trust decisions and neural reward circuitry. Frontiers in neuroscience. 2012; 6:148. doi: https://doi.org/10.3389/fnins.2012.00148.

Fareri DS, Chang LJ, Delgado MR. Computational substrates of social value in interpersonal collaboration. Journal of Neuroscience. 2015; 35(21):8170–8180. doi: https://doi.org/10.1523/JNEUROSCI.4775-14.2015.

Farrell S, Lewandowsky S. Computational models as aids to better reasoning in psychology. Current Directions in Psychological Science. 2010; 19(5):329–335. doi: https://doi.org/10.1177/0963721410386677.

Fehr E, Fischbacher U. The nature of human altruism. Nature. 2003; 425(6960):785–791. doi: https://doi.org/10.1038/nature02043.

Fett AKJ, Gromann PM, Giampietro V, Shergill SS, Krabbendam L. Default distrust? An fMRI investigation of the neural development of trust and cooperation. Social cognitive and affective neuroscience. 2014; 9(4):395–402. doi: https://doi.org/10.1093/scan/nss144.

Garrett N, Daw ND. Biased belief updating and suboptimal choice in foraging decisions. Nature communications. 2020; 11(1):3417. doi: https://doi.org/10.1038/s41467-020-16964-5.

Gelman A, Carlin JB, Stern HS, Dunson DB, Vehtari A, Rubin DB. Bayesian data analysis. 3rd ed. Boca Raton, FL: Chapman and Hall/CRC; 2014. doi: https://doi.org/10.1201/b16018.

Gelman A, Shalizi CR. Philosophy and the practice of Bayesian statistics. British Journal of Mathematical and Statistical Psychology. 2013; 66(1):8–38. doi: https://doi.org/10.1111/j.2044-8317.2011.02037.x.

Gutiérrez-Roig M, Gracia-Lázaro C, Perelló J, Moreno Y, Sánchez A. Transition from reciprocal cooperation to persistent behaviour in social dilemmas at the end of adolescence. Nature communications. 2014; 5(1):4362. doi: http://dx.doi.org/10.1038/ncomms5362.

House BR, Kanngiesser P, Barrett HC, Broesch T, Cebioglu S, Crittenden AN, Erut A, Lew-Levy S, Sebastian-Enesco C, Smith AM, et al. Universal norm psychology leads to societal diversity in prosocial behaviour and development. Nature Human Behaviour. 2020; 4(1):36–44. doi: https://doi.org/10.1038/s41562-019-0734-z.

Hurvich CM, Tsai CL. Regression and time series model selection in small samples. Biometrika. 1989; 76(2):297–307. doi: https://doi.org/10.1093/biomet/76.2.297.

Icenogle G, Steinberg L, Duell N, Chein J, Chang L, Chaudhary N, Di Giunta L, Dodge KA, Fanti KA, Lansford JE, Oburu P, Pastorelli C, Skinner AT, Sorbring E, Tapanya S, Uribe Tirado LM, Alampay LP, Al-Hassan SM, Takash HMS, Bacchini D. Adolescents' cognitive capacity reaches adult levels prior to their psychosocial maturity: Evidence for a "maturity gap" in a multinational, cross-sectional sample. Law and Human Behavior. 2019; 43(1):69–85. doi: https://doi.org/10.1037/lhb0000315.

King-Casas B, Tomlin D, Anen C, Camerer CF, Quartz SR, Montague PR. Getting to know you: reputation and trust in a two-person economic exchange. Science. 2005; 308(5718):78–83. doi: https://doi.org/10.1126/science.1108062.

Krekelberg B, Matlab Toolbox for Bayes Factor Analysis (v3.0). Zenodo; 2024. doi: https://doi.org/10.5281/zenodo.13744717.

Li J, Schiller D, Schoenbaum G, Phelps EA, Daw ND. Differential roles of human striatum and amygdala in associative learning. Nature Neuroscience. 2011; 14(10):1250–1252. doi: https://doi.org/10.1038/nn.2904.

Nava F, Margoni F, Herath N, Nava E. Age-dependent changes in intuitive and deliberative cooperation. Scientific Reports. 2023; 13(1):4457. doi: https://doi.org/10.1038/s41598-023-31691-9.

Nowak MA. Five rules for the evolution of cooperation. Science. 2006; 314(5805):1560–1563. doi: https://doi.org/10.1126/science.1133755.

Palminteri S, Kilford EJ, Coricelli G, Blakemore SJ. The computational development of reinforcement learning during adolescence. PLoS Computational Biology. 2016; 12(6):e1004953. doi: https://doi.org/10.1371/journal.pcbi.1004953.

Pearce JM, Hall G. A model for Pavlovian learning: Variations in the effectiveness of conditioned but not of unconditioned stimuli. Psychological Review. 1980; 87(6):532–552. doi: https://doi.org/10.1037/0033-295X.87.6.532.

Pfeifer JH, Berkman ET. The development of self and identity in adolescence: Neural evidence and implications for a value-based choice perspective on motivated behavior. Child Development Perspectives. 2018; 12(3):158–164. doi: https://doi.org/10.1111/cdep.12279.

Rigoux L, Stephan KE, Friston KJ, Daunizeau J. Bayesian model selection for group studies—revisited. Neuroimage. 2014; 84:971–985. doi: https://doi.org/10.1016/j.neuroimage.2013.08.065.

Rilling JK, Sanfey AG, Aronson JA, Nystrom LE, Cohen JD. The neural correlates of theory of mind within interpersonal interactions. Neuroimage. 2004; 22(4):1694–1703. doi: https://doi.org/10.1016/j.neuroimage.2004.04.015.

Rilling JK, Gutman DA, Zeh TR, Pagnoni G, Berns GS, Kilts CD. A neural basis for social cooperation. Neuron. 2002; 35(2):395–405. doi: https://doi.org/10.1016/s0896-6273(02)00755-9.

Rosenbaum GM, Grassie HL, Hartley CA. Valence biases in reinforcement learning shift across adolescence and modulate subsequent memory. ELife. 2022; 11:e64620. doi: https://doi.org/10.7554/eLife.64620.

Sachs JL, Mueller UG, Wilcox TP, Bull JJ. The evolution of cooperation. The Quarterly Review of Biology. 2004; 79(2):135–160. doi: https://doi.org/10.1086/383541.

Stan Development Team. Stan Reference Manual, Version 2.33; 2023, https://mc-stan.org.

Steinberg L. Cognitive and affective development in adolescence. Trends in cognitive sciences. 2005; 9(2):69–74. doi: https://doi.org/10.1016/j.tics.2004.12.005.

Taheri M, Rotshtein P, Beierholm U. The effect of attachment and environmental manipulations on cooperative behavior in the prisoner's dilemma game. PLOS ONE. 2018; 13(11):e0205730. doi: https://doi.org/10.1371/journal.pone.0205730.

Tervo-Clemmens B, Calabro FJ, Parr AC, et al. A canonical trajectory of executive function maturation from ado-
lescence to adulthood. Nature Communications. 2023; 14:6922. https://doi.org/10.1038/s41467-023-42540-8,
doi: 10.1038/s41467-023-42540-8.

Ting F, He Z, Baillargeon R. Toddlers and infants expect individuals to refrain from helping an ingroup
victim's aggressor. Proceedings of the National Academy of Sciences. 2019; 116(13):6025–6034. doi:
https://doi.org/10.1073/pnas.1817849116.

Van Den Bos W, Van Dijk E, Westenberg M, Rombouts SARB, Crone EA. Changing Brains, Changing Per-
spectives: The Neurocognitive Development of Reciprocity. Psychological Science. 2011; 22(1):60–70. doi:
https://doi.org/10.1177/0956797610391102.

Van Den Bos W, Westenberg M, Van Dijk E, Crone EA. Development of trust and reciprocity in adolescence.
Cognitive Development. 2010; 25(1):90–102. doi: https://doi.org/10.1016/j.cogdev.2009.07.004.

Van Duijvenvoorde ACK, Huizenga HM, Somerville LH, Delgado MR, Powers A, Weeda WD, Casey BJ, Weber EU,
Figner B. Neural Correlates of Expected Risks and Returns in Risky Choice across Development. The Journal
of Neuroscience. 2015; 35(4):1549–1560. doi: https://doi.org/10.1523/JNEUROSCI.1924-14.2015.

Vrieze SI. Model selection and psychological theory: A discussion of the differences between the Akaike Infor-
mation Criterion (AIC) and the Bayesian Information Criterion (BIC). Psychological Methods. 2012; 17(2):228–
243. doi: https://doi.org/10.1037/a0027127.

Warneken F. How children solve the two challenges of cooperation. Annual Review of Psychology. 2018;
69:205–229. doi: https://doi.org/10.1146/annurev-psych-122216-011813.

Westhoff B, Molleman L, Viding E, Van Den Bos W, Van Duijvenvoorde ACK. Developmental asymmetries in
learning to adjust to cooperative and uncooperative environments. Scientific Reports. 2020; 10(1):21761.
doi: https://doi.org/10.34894/Z1OYYA.

Worthman CM, Trang K. Dynamics of body time, social time and life history at adolescence. Nature. 2018;
554(7693):451–457. doi: https://doi.org/10.1038/nature25750.

Wu X, Ren X, Liu C, Zhang H. The motive cocktail in altruistic behaviors. Nature Computational Science. 2024;
4:659–676. doi: https://doi.org/10.1038/s43588-024-00685-6.

Wu X, Zhu R, Gong X, Luo Y, Liu C. Social incentives foster cooperation through guilt aversion: An
effect that diminishes with primary psychopathic traits. PsyCh Journal. 2023; 12(3):389–398. doi:
https://doi.org/10.1002/pchj.641.

Wu X, Fu H, Zhang T, Bao D, Hu J, Zhu R, Feng C, Gu R, Liu C. A cognitive computational mechanism for mutual
cooperation: The roles of positive expectation and social reward. Acta Psychologica Sinica. 2024; 56(9):1299–
1312. doi: https://doi.org/10.3724/SP.J.1041.2024.01299.

Additional files

Data and code availability

621 All datasets and code are available at https://doi.org/10.5281/zenodo.15046430.

Appendix Analysis

Model 9: Social reward model with Pearce–Hall (PH) learning algorithm

To examine whether trial-by-trial learning rate adaptation improves model performance, we ex-
tended the social reward model by incorporating a Pearce–Hall (PH) dynamic learning rule(***Pearce***
and Hall, 1980; ***Li et al., 2011***). In this model, participants update their belief about their partner's
cooperation probability ($\hat{r}_t$) using a learning rate α_t that changes adaptively as a function of recent
prediction errors:

$$\hat{r}_{t+1} = \hat{r}_t + \alpha_t (r_t - \hat{r}_t), \quad \alpha_{t+1} = \alpha_t + \lambda(|PE_t| - \alpha_t),$$

where r_t denotes the observed partner cooperation (or reward), $PE_t = r_t - \hat{r}_t$, and λ controls the rate
at which α_t adapts to the magnitude of recent prediction errors. This formulation allows the learn-
ing rate to increase following surprising outcomes and decrease when outcomes are predictable,

independent of the valence of feedback. As in previous models, decision values were computed
based on the expected utilities of cooperation ($U^{coop} = \hat{r}_t(4+\omega)$) and defection ($U^{def} = \hat{r}_t \times 4 + 2$), and
choice probabilities were obtained via a softmax function governed by the inverse temperature pa-
rameter β. Model comparison indicated that the PH dynamic learning model did not outperform
the best-fitting asymmetric RL model (Model 8) in either age group (see *Figure Supplement 6*).

Hierarchical Bayesian Estimation

To complement the maximum likelihood estimation analyses, we additionally implemented a hier-
archical Bayesian estimation for the best-fitting model. The hierarchical model incorporated both
group-level (adolescent and adult) and individual-level structures to improve the stability and iden-
tifiability of parameter estimation.

At the individual level, each participant s was characterized by four parameters: the positive
learning rate ($\alpha+_s$), negative learning rate ($\alpha-_s$), social reward weight (ω_s), and inverse temperature
(β_s). Each parameter was assumed to follow a normal distribution centered on the corresponding
group-level mean with a group-specific standard deviation:

$$\alpha_s^+ \sim \mathcal{N}(\mu_{\alpha_{g[s]}^+}, \sigma_{\alpha_{g[s]}^+}),$$
$$\alpha_s^- \sim \mathcal{N}(\mu_{\alpha_{g[s]}^-}, \sigma_{\alpha_{g[s]}^-}),$$
$$\omega_s \sim \mathcal{N}(\mu_{\omega_{g[s]}}, \sigma_{\omega_{g[s]}}),$$
$$\beta_s \sim \mathcal{N}(\mu_{\beta_{g[s]}}, \sigma_{\beta_{g[s]}}),$$

646 where $g[s] \in \{\text{adolescent}, \text{adult}\}$ denotes the group of participant s.

647 At the group level, the hyperparameters μ and σ were assigned weakly informative priors to
648 regularize estimation while allowing flexibility across age groups. The likelihood for each trial t fol-
649 lowed a Bernoulli distribution based on the model-predicted cooperation probability $p_{s,t}$, computed
650 via the softmax (logistic) transformation of the utility difference:

$$y_{s,t} \sim \text{Bernoulli}(p_{s,t}), \quad p_{s,t} = \frac{1}{1 + \exp[-\beta_s(U_{s,t}^{coop} - U_{s,t}^{def})]},$$

651 where $U_{s,t}^{coop}$ and $U_{s,t}^{def}$ denote the expected utilities of cooperation and defection, respectively, up-
652 dated trial-by-trial using asymmetric learning rates:

$$PE_{s,t} = r_{s,t} - \hat{r}_{s,t-1}, \quad \hat{r}_{s,t} = \begin{cases} \hat{r}_{s,t-1} + \alpha_s^+ \cdot PE_{s,t}, & \text{if } PE_{s,t} > 0, \\ \hat{r}_{s,t-1} + \alpha_s^- \cdot PE_{s,t}, & \text{if } PE_{s,t} < 0. \end{cases}$$

The hierarchical Bayesian model was implemented in `Stan` (*Stan Development Team, 2023*) and
fitted using the `sampling()` function in `RStan`. Four independent Markov chains were run with 4,000
iterations each, including 1,000 warm-up samples for adaptation. Convergence was evaluated us-
ing the potential scale reduction statistic $\hat{R}$, with all parameters showing $\hat{R} < 1.01$, indicating good
convergence and stable sampling (see *Figure Supplement 7a*). Trace plots for the group-level pa-
rameters ($\alpha+$, $\alpha-$, ω, and β) further confirmed well-mixed chains (see *Figure Supplement 7b*).

Posterior group-level parameter estimates (μ parameters) were then directly compared be-
tween adolescents and adults. The hierarchical Bayesian results closely replicated those obtained
from the MLE approach, demonstrating consistent age-group differences in parameters (see *Fig-
ure Supplement 5*).

Robustness analyses and extensions of main findings

664 To further examine the robustness of the findings, we first reconduct GLMM1 and LMM1–3 by
665 replacing the categorical Group factor (adolescents vs. adults) with a continuous age predictor,
666 yielding GLMM$_{sup}$1 and LMM$_{sup}$1-3. Overall, the patterns replicated those observed for the Group
667 effect: the Previous trial × Partner's choice × Age interaction was significant in in GLMM$_{sup}$1 (*b* =

0.02, 95% CI = [0.001, 0.036], p = 0.038, see Appendix *Table 5*) and $LMM_{sup}3$ (b = 0.04, 95% CI = [0.031, 0.046], p < 0.001, see Appendix *Table 8*), and marginally significant in $LMM_{sup}1$ (b = 0.0012, 95% CI = [-1.63×10^{-7}, 0.002], p = 0.050, see Appendix *Table 6*). Consistent with the nonsignificant effect of the group in LMM2 (see Appendix *Table 3*), the corresponding age effect in $LMM_{sup}2$ (see Appendix *Table 7*) was also not significant (b = 0.03, 95% CI = [-0.168, 0.222], p = 0.784). Furthermore, based on GLMM1, LMM1, and LMM3, we further included the phase factor (stable versus changing phase) as a fixed effect with random slopes, yielding $GLMM_{sup}2$ and $LMM_{sup}4$–5. This specification accounts for potential effects of the experimental manipulation of the partner's cooperation, which differed between the first and second halves of the trials. We also replicated the main results: the Previous trial × Partner's choice × Age interaction was significant in $GLMM_{sup}2$ (b = 0.25, 95% CI = [0.128, 0.366], p < 0.001, see Appendix *Table 9*), $LMM_{sup}4$ (b = 0.03, 95% CI = [0.022, 0.037], p < 0.001, see Appendix *Table 10*), and $LMM_{sup}5$ (b = 0.37, 95% CI = [0.314, 0.417], p < 0.001, see Appendix *Table 11*). Besides, to account for the potential influence of individual differences in prosociality on participants' cooperative behavior and the reward for reciprocity, we also extended GLMM1 and LMM3 by adding the measured SVO as a fixed effect with random slopes, yielding $GLMM_{sup}3$ and $LMM_{sup}6$. The results showed that higher SVO positively predicted greater cooperation (b = 0.02, 95% CI = [0.007, 0.027], p < 0.001, see Appendix *Table 12*), whereas its effect on the reward for reciprocity was not significant (b = 0.006, 95% CI = [-0.002, 0.014], p = 0.137, see Appendix *Table 13*). Importantly, the primary findings remained unchanged after controlling for SVO.

Finally, we conducted an exploratory analysis to examine the relationship between participants' expectations of partner's cooperation probability and their own cooperative behavior. Specifically, we estimated $GLMM_{sup}4$, in which participants' choices served as the dependent variable. The fixed effects model included trial number, gender, group, cooperation expectation, and the interaction between group and cooperation expectation. Random effects included correlated random slopes of trial number, gender, group, cooperation expectation, and a random intercept for participants. We showed that participants' cooperation expectations positively predicted cooperative behavior (b = 7.90, 95% CI = [7.258, 8.459], p < 0.001, see Appendix *Table 14*), indicating that their cooperation depends on their estimates of partner' belief. In addition, we found the interaction between group and cooperation expectation was not significant (b = 0.015, 95% CI = [-0.358, 0.387], p = 0.938), suggesting that this social learning process likely operates similarly in adolescents and adults and is not driven by age-group differences.

 Appendix Tables

Table 1. Statistical results for cooperation decision (GLMM1).

Fixed effects	Estimated beta value	SE	t value	p value
(Intercept)	0.47	0.22	2.10	$p = .036$
Timing	-0.001	0.001	-1.01	$p = .311$
Gender	0.14	0.15	0.93	$p = .355$
Group	0.79	0.24	3.23	$p = .001$
Previous trial	-1.81	0.08	-23.43	$p < .001$
Partner's choice	-0.73	0.10	-7.21	$p < .001$
Group × Previous trial	-0.41	0.10	-3.89	$p < .001$
Group × Partner's choice	-0.19	0.11	-1.77	$p = .076$
Previous trial × Partner's choice	1.12	0.04	25.43	$p < .001$
Group × Previous trial × Partner's choice	0.24	0.06	4.05	$p < .001$

Coding of variables. Trial number: integer sequence from 2 to 120; gender: female = 0, male = 1; group: adolescents = 0, adults = 1; previous trials: last one trial = 1, last two trials = 2, last three trials = 3; partner's choice: cooperation = 1, defection = 0.

Source Data: Download Data

Table 2. Statistical results for partner cooperation expectation (LMM1).

Fixed effects	Estimated beta value	SE	t value	p value
(Intercept)	0.59	0.03	17.52	$p < .001$
Trial number	<0.001	<0.001	0.02	$p = .982$
Gender	-0.01	0.03	-0.36	$p = .718$
Group	-0.04	0.03	-1.29	$p = .198$
Previous trial	-0.25	0.01	-48.97	$p < .001$
Partner's choice	-0.06	0.01	-4.90	$p < .001$
Group × Previous trial	-0.04	0.01	-6.10	$p < .001$
Group × Partner's choice	-0.02	0.01	-2.73	$p = .006$
Previous trial × Partner's choice	0.16	0.002	53.89	$p < .001$
Group × Previous trial × Partner's choice	0.03	0.004	7.31	$p < .001$

Coding of variables is consistent with Table 1.

Source Data: Download Data

Table 3. Statistical results for self-reported perceived partner cooperativeness (LMM2).

Fixed effects	Estimated beta value	SE	t value	p value
(Intercept)	5.82	0.27	21.93	$p < .001$
Group	0.17	0.35	0.50	$p = .616$
Gender	-0.44	0.11	-3.84	$p < .001$
Order of sessions	0.09	0.15	0.60	$p = .549$
Rating number	-0.13	0.02	-5.55	$p < .001$
Group × Order	0.38	0.22	1.72	$p = .085$

Coding of variables. Group: adolescents = 0, adults = 1; gender: female = 0, male = 1; order of sessions: stable to volatile = 0, volatile to stable =1; rating number: integer sequence from 1 to 8.

Source Data: Download Data

Table 4. Statistical results for intrinsic reward for reciprocity (LMM3).

Fixed effects	Estimated beta value	SE	t value	p value
(Intercept)	2.43	0.14	17.04	$p < .001$
Trial number	-0.001	0.001	-0.60	$p = .551$
Gender	0.08	0.13	0.59	$p = .554$
Group	0.52	0.15	3.44	$p < .001$
Previous trial	-1.23	0.03	-36.18	$p < .001$
Partner's choice	-0.27	0.08	-3.24	$p = .001$
Group × Previous trial	-0.57	0.05	-11.95	$p < .001$
Group × Partner's choice	-0.24	0.05	-4.43	$p = .006$
Previous trial × Partner's choice	0.79	0.02	40.64	$p < .001$
Group × Previous trial × Partner's choice	0.37	0.02	13.70	$p < .001$

Coding of variables is consistent with Table 1.

Source Data: Download Data

Table 5. Statistical results for cooperation decision with age (GLMM$_{sup}$1).

Fixed effects	Estimated beta value	SE	t value	p value
(Intercept)	-1.55	0.73	-2.12	$p = .034$
Trial number	-0.001	0.001	-0.94	$p = .346$
Gender	-0.06	0.18	-0.34	$p = .732$
Previous trial	-0.49	0.15	-3.30	$p = .001$
Partner's choice	-0.60	0.33	-1.83	$p = .068$
Age	0.10	0.04	2.49	$p = .013$
Previous trial × Partner's choice	0.89	0.17	5.10	$p < .001$
Previous trial × Age	-0.02	0.01	-2.00	$p = .045$
Partner's choice × Age	-0.01	0.02	-0.68	$p = .499$
Previous trial × Partner's choice × Age	0.02	0.01	2.07	$p = .038$

Age was treated as a continuous variable, and all other variables were coded as in Table 1.

Source Data: Download Data

Table 6. Statistical results for partner cooperation expectation with age (LMM$_{sup}$1).

Fixed effects	Estimated beta value	SE	t value	p value
(Intercept)	0.76	0.12	6.52	$p < .001$
Trial number	-2.53×10^{-7}	0.0002	-0.002	$p = .999$
Gender	-0.01	0.03	-0.32	$p = .746$
Previous trial	-0.10	0.01	-10.51	$p < .001$
Partner's choice	-0.03	0.03	-1.05	$p = .294$
Age	-0.014	0.006	-2.27	$p = .023$
Previous trial × Partner's choice	0.15	0.01	13.24	$p < .001$
Previous trial × Age	-4.76×10^{-7}	0.0004	0.001	$p = .999$
Partner's choice × Age	-0.002	0.001	-1.74	$p = .082$
Previous trial × Partner's choice × Age	0.0012	0.0006	1.96	$p = .050$

Coding of variables is consistent with Table 5.

Source Data: Download Data

Table 7. Statistical results for self-reported perceived partner cooperativeness with age (LMM$_{sup}$2).

Fixed effects	Estimated beta value	SE	t value	p value
(Intercept)	5.18	1.89	2.74	$p = .006$
Gender	-0.37	0.19	-1.96	$p = .050$
Order of sessions	2.31	1.21	1.90	$p = .057$
Age	0.03	0.10	0.27	$p = .784$
Rating number	-0.13	0.02	-6.65	$p < .001$
Order × Age	-0.10	0.07	-1.57	$p = .116$

Age was treated as a continuous variable, and all other variables were coded as in Table 3

Source Data: Download Data

Table 8. Statistical results for partner cooperation expectation with age (LMM$_{sup}$3).

Fixed effects	Estimated beta value	SE	t value	p value
(Intercept)	2.02	0.58	3.49	$p < .001$
Trial number	-0.001	0.001	-0.60	$p = .546$
Gender	-0.07	0.13	-0.50	$p = .619$
Previous trial	-0.22	0.06	-3.59	$p < .001$
Partner's choice	-0.11	0.18	-0.62	$p = .533$
Age	0.02	0.03	0.61	$p = .540$
Previous trial × Partner's choice	0.25	0.08	3.28	$p = .001$
Previous trial × Age	-0.02	0.003	-5.24	$p < .001$
Partner's choice × Age	-0.02	0.01	-1.85	$p = .065$
Previous trial × Partner's choice × Age	0.04	0.004	9.82	$p < .001$

Coding of variables is consistent with Table 5.

Source Data: Download Data

Table 9. Statistical results for cooperation decision with phase (GLMM$_{sup}$2).

Fixed effects	Estimated beta value	SE	t value	p value
(Intercept)	1.06	0.24	4.34	$p < .001$
Trial number	-0.01	0.001	-4.46	$p < .001$
Gender	0.17	0.15	1.10	$p = .271$
Group	0.60	0.19	3.15	$p = .002$
Previous trial	-0.64	0.04	-16.49	$p < .001$
Partner's choice	-0.75	0.10	-7.73	$p < .001$
Phase	-0.66	0.10	-6.45	$p < .001$
Group × Previous trial	-0.16	0.05	-3.15	$p = .002$
Group × Partner's choice	-0.19	0.11	-1.77	$p = .077$
Previous trial × Partner's choice	1.04	0.04	23.44	$p < .001$
Group × Previous trial × Partner's choice	0.25	0.06	4.07	$p < .001$

Phase was dummy-coded (0 = stable, 1 = changing), and all other variables were coded as in Table 1.

Source Data: Download Data

Table 10. Statistical results for partner cooperation expectation with phase (LMM$_{sup}$4).

Fixed effects	Estimated beta value	SE	t value	p value
(Intercept)	0.77	0.03	28.09	$p < .001$
Trial number	-5.86×10^{-6}	9.25×10^{-5}	-0.06	$p = .949$
Gender	-0.02	0.03	-0.96	$p = .337$
Group	-0.06	0.03	-2.36	$p = .018$
Previous trial	-0.08	0.002	-33.62	$p < .001$
Partner's choice	-0.07	0.01	-6.65	$p < .001$
Phase	-0.14	0.01	-19.16	$p < .001$
Group × Previous trial	-0.013	0.003	-4.02	$p < .001$
Group × Partner's choice	-0.025	0.007	-3.11	$p = .002$
Previous trial × Partner's choice	0.15	0.003	50.67	$p < .001$
Group × Previous trial × Partner's choice	0.03	0.004	7.40	$p < .001$

Coding of variables is consistent with Table 9.

Source Data: Download Data

Table 11. Statistical results for intrinsic reward for reciprocity with phase (LMM$_{sup}$5).

Fixed effects	Estimated beta value	SE	t value	p value
(Intercept)	3.57	0.14	25.57	$p < .001$
Trial number	-0.0011	0.0006	-2.01	$p = .045$
Gender	0.02	0.13	0.18	$p = .860$
Group	0.25	0.13	1.96	$p = .050$
Previous trial	-0.38	0.02	-22.73	$p < .001$
Partner's choice	-0.33	0.07	-4.59	$p < .001$
Phase	-0.83	0.06	-14.76	$p < .001$
Group × Previous trial	-0.19	0.02	-8.79	$p < .001$
Group × Partner's choice	-0.27	0.05	-5.18	$p < .001$
Previous trial × Partner's choice	0.70	0.02	37.41	$p < .001$
Group × Previous trial × Partner's choice	0.37	0.03	13.99	$p < .001$

Coding of variables is consistent with Table 9.

Source Data: Download Data

Table 12. Statistical results for cooperation decision with SVO (GLMM$_{sup}$3).

Fixed effects	Estimated beta value	SE	t value	p value
(Intercept)	-0.63	0.22	-2.88	$p = .004$
Trial number	-0.002	0.001	-1.16	$p = .247$
Gender	0.18	0.14	1.27	$p = .202$
Group	0.56	0.18	3.11	$p = .002$
Previous trial	-0.71	0.04	-18.17	$p < .001$
Partner's choice	-0.77	0.10	-7.41	$p < .001$
SVO	0.02	0.005	3.35	$p < .001$
Group × Previous trial	-0.14	0.05	-2.84	$p = .005$
Group × Partner's choice× Age	-0.16	0.11	-1.44	$p = .150$
Previous trial × Partner's choice	1.13	0.04	25.18	$p < .001$
Group × Previous trial × Partner's choice	0.25	0.06	4.06	$p < .001$

SVO was treated as a continuous variable, and all other variables were coded as in Table 1.

Source Data: Download Data

Table 13. Statistical results for intrinsic reward for reciprocity with SVO (LMM$_{sup}$6).

Fixed effects	Estimated beta value	SE	t value	p value
(Intercept)	2.03	0.15	13.73	$p < .001$
Trial number	-0.001	0.00 1	-0.71	$p = .479$
Gender	0.02	0.12	0.18	$p = .855$
Group	0.26	0.13	2.07	$p = .039$
Previous trial	-0.45	0.02	-25.75	$p < .001$
Partner's choice	-0.31	0.08	-3.65	$p < .001$
Phase	0.006	0.004	1.49	$p = .137$
Group × Previous trial	-0.18	0.02	-8.02	$p < .001$
Group × Partner's choice	-0.17	0.06	-3.03	$p = .002$
Previous trial × Partner's choice	0.80	0.02	40.26	$p < .001$
Group × Previous trial × Partner's choice	0.37	0.03	13.38	$p < .001$

Coding of variables is consistent with Table 12.

Source Data: Download Data

Table 14. Statistical results for cooperation decision predicted by cooperation expectation (GLMM$_{sup}$4).

Fixed effects	Estimated beta value	SE	t value	p value
(Intercept)	-4.57	0.35	-12.92	$p < .001$
Trial number	-0.002	0.001	-2.25	$p = .024$
Gender	0.20	0.28	0.71	$p = .475$
Group	1.37	0.34	4.06	$p < .001$
Cooperation expectation	7.90	0.33	24.05	$p < .001$
Group × Cooperation expectation	0.01	0.19	0.08	$p = .938$

Coding of variables. Trial number: integer sequence from 2 to 120; gender: female = 0, male = 1; group: adolescents = 0, adults = 1. Cooperation expectation was treated as a continuous variable.

Source Data: Download Data

700 **Appendix Figures**

Figure supplement 6. Model comparison results for (a) adults and (b) adolescents, including the newly added M9 (Social Reward & Pearce–Hall learning). Lower ΔAICc values indicate better model fits. The dynamic learning rate model (Model 9: Social Reward model with dynamic RL algorithm) did not outperform the best-fitting model (Model 8) in either group.

Figure supplement 7. Convergence diagnostics for the hierarchical Bayesian model. (**a**) Distribution of $\hat{R}$ (Rhat) values across all model parameters. The majority of $\hat{R}$ values are below the conservative convergence threshold of 1.01 (red dashed line), indicating stable and well-mixed MCMC chains. The gray shaded area highlights the region where $\hat{R} \leq 1.01$. (**b**) Trace plots for the group-level parameters (four chains) in adolescents (left, red box) and adults (right, blue box). Each line represents the sampled posterior values of one chain across iterations, with overlapping traces and stable fluctuations confirming adequate convergence and mixing for all key parameters ($\alpha+$, $\alpha-$, ω, β).

Figure 2—figure supplement 1. Model prediction. This figure compares the actual data and presupplement 1 (**a**) and adolescents (**b**). The x-axis represents the trial number, while the y-axis represents the mean cooperation probability of all participants. The shaded area indicates the 95% confidence interval.

Figure 2—figure supplement 2. Distributions of estimated parameters from the best-fitting model. Each panel displays one parameter. The histograms and their kernel fits are represented by color bars and curves, respectively. Red indicates participants in the adolescent sample, and blue denotes those in the adult sample. Parameters have been transformed into a log scale for enhanced visualization.

Figure 2—figure supplement 3. Parameter recovery for the best-fitting model. Each panel represents one parameter. Each dot corresponds to one virtual participant. The value of r indicates Pearson's correlation coefficient between the true values (estimated from the participants) and the recovered parameters.

Figure 2—figure supplement 4. Partial correlation matrices among parameters for the best-fitting model. The upper-triangular cells show partial correlations for adults, and the lower-triangular cells show partial correlations for adolescents. Each cell shows the partial Pearson correlation coefficient (controlling for the other parameters). Colors range from green (negative) to violet (positive), with the color bar spanning [-1,1]. Notes: *n.s.* $p > 0.05$; *$p < 0.05$; **$p < 0.01$; ***$p < 0.001$.

Group-level posterior distributions (μ parameters)

Adolescent vs Adult; annotation shows Adult − Adol posterior difference

Figure 2—figure supplement 5. Group-level posterior distributions from the hierarchical Bayesian estimation for adolescents and adults. Posterior densities are shown separately for adolescents (red) and adults (blue). Δ values indicate the posterior mean difference (Adult – Adolescent) with 95% credible intervals (CrI) and Bayesian p values. Compared with adolescents, adults exhibited higher positive learning rates ($\alpha+$) and lower negative learning rates ($\alpha-$), suggesting greater differentiation between learning from positive and negative feedback. Adults also showed lower inverse temperature (β), indicating more exploratory decision behavior, and higher social reward weight (ω), reflecting greater valuation of reciprocity or social outcomes. Notes: *n.s.* $p > 0.05$; $*p < 0.05$; $**p < 0.01$; $***p < 0.001$.