Title: Sensory-based quantification of male colour patterns in Trinidadian guppies reveals

nonparallel phenotypic evolution across an ecological transition in multivariate trait space

Running title: Non-parallel evolution of guppy colour patterns

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ABSTRACT

Parallel evolution, in which independent populations evolve along similar phenotypic

trajectories, offers insights into the repeatability of adaptive evolution. Here, we revisit a

classic example of parallelism, that of repeated evolution of brighter males in the Trinidadian

guppy. In guppies, colonisation of low predation habitats is associated with emergence of

'more colourful' phenotypes since predator-induced viability selection for crypsis weakens

while sexual selection by female preference for conspicuity remains strong. Our study differs

from previous investigations in three respects. First, we adopt a multivariate phenotyping

approach to characterise parallelism in multi-trait space. Second, we use ecologically-

relevant colour traits defined by the visual systems of the two selective agents (i.e. guppy,

predatory cichlid). Third, we estimate population genetic structure to test for adaptive

(parallel) evolution against a model of neutral phenotypyc divergence. We find strong

phenotypic differentiation that is inconsistent with a neutral model, but only limited support

for the predicted pattern of greater conspicuity at low predation. Effects of predation regime

on each trait were in the expected direction, but weak, largely non-significant, and explained

little among-population variation. In multi-trait space, phenotypic trajectories of lineages

colonising low from high predation regimes were not parallel. Our results are consistent with

reduced predation risk facilitating adaptive differentiation by female choice, but suggest that

this proceeds in (effectively) independent directions of multi-trait space across lineages.

Pool-sequencing data also revealed SNPs showing greater differentiation than expected

under neutrality and/or associations with known colour genes in other species, presenting

opportunities for future genetic study.

Keywords: Parallel evolution, coloration, guppy, *Poecilia reticulata,* pool-sequencing, colour

pattern analysis

INTRODUCTION

Parallel evolution is defined as the repeated, independent evolution of similar phenotypes

under similar selection regimes in multiple independent populations within closely related

lineages (Schluter 2004; Bolnick *et al.* 2018). Notable examples include the repeated

emergence of stream and lacustrine morphotypes in sticklebacks (McKinnon *et al.* 2004;

Schluter *et al.* 2004), and of thin and thick shelled habitat-specific forms of periwinkle (Butlin

et al., 2014). This phenomenon allows us to ask questions about the predictability of

evolution. Does adaptation frequently recapitulate the same phenotypic 'solutions' to the

same selective challenges? Also, what are the genetic pathways and processes involved?

However, characterizing evolutionary trajectories of complex multivariate phenotypes as

being either parallel or not is too simplistic. In reality, any two trajectories may be more or

less aligned in phenotypic space such that the degree of parallelism lies on quantitative

continuum (Stuart *et al.* 2017; Bolnick *et al.*2018). By characterizing the distribution of

(multivariate) phenotypes within and among-lineages we can quantify parallelism, testing for

similarity in both magnitude and direction of evolutionary trajectories over ecological

transitions. We can also assess the importance of parallel (adaptive) evolution within the

context of other processes contributing to phenotypic divergence (e.g. drift). Here we apply

this approach to a well-known instance of putatively parallel evolution, that of replicated

divergence in male colour patterns across predation regimes in the Trinidadian guppy

(*Poecilia reticulata*) (Endler *et al.* 1987; Reznick *et al.* 1996).

Animal coloration and pattern traits serve many functions including signalling and

crypsis (Cuthill *et al.* 2017) and have long been used to test predictions about the role of

predation in driving parallel evolution (Houde, 1997; Allender *et al.* 2003; Steiner *et al.* 2009).

High predation risk should select for less 'conspicuous' colours and patterns (Endler 1978,

1987; Young *et al.* 2011; Martin *et al.* 2014), but testing this may be sensitive to how colour

phenotypes are quantified. Specific colour traits chosen for analysis often vary across

studies even within species, while quantitative measures based on human perception

(Martin *et al.* 2014) or RGB information (van Belleghem *et al.* 2018; Montenegro *et al.* 2019)

may sometimes lack ecological relevance. The latter concern arises because colour signals

will (co)evolve with the visual systems (and downstream behaviours) of receiver species (i.e.

the sensory drive hypothesis; Endler 1992). For instance, flower traits have coevolved with

hymenopteran vision (Dyer *et al.* 2012), and variation in opsin gene sequence and

expression is linked to colour polymorphism in African cichlids (Seehausen *et al.* 2008).

While this means that human vision could misrepresent colour phenotypes as perceived by

relevant selective agents (e.g. conspecific mates, predators), recent advances have

improved our ability to model colour variation under different visual models (Stevens *et al.*

2007; Endler *et al.* 2018; Troscianko and Stevens 2015; van den Berg *et al.* 2019).

Nevertheless, challenges remain such that colour and pattern phenotypes are necessarily

complex and multivariate. Chromatic (e.g. colour) and achromatic (e.g. luminance) aspects

of a colour signal are commonly considered, but continued reliance on univariate analyses

means that the consequences of trait combinations may be overlooked (Endler and Mappes

2017). Because of these challenges, there has been a call for increased use of spatio-

chromatic phenotyping approaches that integrate variation in colour with pattern (spatial

arrangement) (Endler and Mielke 2005; Endler *et al.* 2018; van den Berg *et al.* 2019).

A more general limitation of many previous studies has been the relatively infrequent

use of population genetic data (but see Steiner *et al.* 2009; Kratochwill *et al.* 2018). Colour

traits are important targets of selective processes, but it does not follow that all divergence

among populations (whether exhibiting parallelism or not) maps adaptively to local selection

regime. Gene flow can sometimes preclude phenotypic divergence between populations

despite differences in selection (Räsänen and Hendry 2008; Nosil 2009), while genetic drift

could cause (non-adaptive) phenotypic divergence that masks parallelism (Stuart *et al.* 2017;

Delisle and Bolnick 2020). Fortunately, patterns of (genome-wide) molecular genetic

differentiation can be used to construct null models against which to test for and isolate the

phenoptypic signal of local adaptation (e.g., Whitlock and Guillaume 2009; Pascoal *et al.*

2018a,b). There are important caveats to this however, for instance when using phenotypic

variation as a proxy for quantitative genetic variation (Pujol *et al.* 2008), or when 'Isolation by

Adaptation' scenarios are plausible (e.g. Nosil *et al.* 2008; Funk *et al.* 2011). Nonethless,

molecular genetic data provide an important opportunity to nuance expectations of

phenotypic structuring among populations hypothesised to have undergone parallel

evolution. In some cases, notably where dense marker data are available, they can also be

used probe the genetic basis of phenotypic differentiation - parallel or otherwise - among

populations (Elmer and Meyer 2011; Gautier *et al.* 2018).

In this study, we revisit the well-documented case of putatively parallel evolution of

colour in the Trinidadian guppy by combining novel colour phenotyping methods, with

estimation of phenotypic divergence among-populations in multivariate trait space and use of

population genomic data. Guppies have provided important insights about many evolutionary

processes (Endler 1980; Houde 1995; Reznick *et al.* 1997; Magurran 2005) with the highly

variable male colour patterns receiving particular attention. Male phenotypes are subject to

antagonistic sexual and viability selection; females prefer more 'conspicuous' male

phenotypes but these also confer greater predation risk. In many rivers in Trinidad, upstream

dispersal of piscivores, notably the pike cichlid *Crenacichla frenata*, is limited by barrier

waterfalls. Downstream habitats are thus characterised by high predation risk (HP), relative

to upstream low predation (LP) sites. Repeated colonization of LP sites has been associated

with phenotypic shifts towards brighter, more colourful males (Endler 1983; Millar *et al.*

2006). This is presumably because the costs of being conspicuous (i.e. predation risk) are

relatively lower in LP populations while the benefits (i.e. attractiveness to females) remain

(Haskins *et al.* 1961; Endler 1980; Houde 1995).

We stress that these patterns of among-population variation in male guppy colour,

and their interpretation with respect to predation risk have proven qualitatively robust (Millar

et al. 2006). However, the diversity of phenotyping methods used across studies has also

limited quantitative comparisons of the extent, and direction of (multivariate) evolution across

lineages. Phenotyping methods have included scoring the presence/absence, number, size

and position of particular colour patches (Endler, 1978; Gotanda *et al.* 2018), spectral

measurements (Kemp *et al.* 2019), and spatial pattern approaches (Endler 2012; Endler *et*

al. 2018). It is also widely acknowledged that the ecological context is more nuanced than

described above (Endler and Houde 1995; Karim *et al.* 2007; Kemp *et al.* 2009). Predation

risk varies within HP and LP contexts not just between them (Endler 1995), but canopy cover

and light environment also differ between upstream and downstream sites, and frequency-

dependent selection also maintains variation within populations (Olendorf *et al.* 2006;

Hughes *et al.* 2013). Finally, not all predictions are fully upheld by experiments. For example,

low predators risk is thought to allow evolution of larger colourful spots on the body (Haskins

et al. 1961). However, in a recent study, parallel increases in melanic (black) spots occurred

across mesocosms lacking predators, but other colour traits changed inconsistently

(Gotanda *et al.* 2018).

157   Here, we assess the extent of parallel evolution in male guppy colour patterns across

repeated instances of LP colonisation from HP ancestors. We implement the Quantitative

Colour Pattern Analysis (QCPA) phenotyping approach (van den Berg *et al.* 2019) to model

multivariate phenotypic divergence as perceived by the two hypothesized selective agents:

namely the guppy (sexual selection) and the Trinidadian pike cichlid (viability selection). Our

specific goals are; (i) to determine whether this phenotyping approach recapitulates the

expected finding that LP guppies are more 'conspicuous' than HP guppies across rivers; (ii)

to adopt a geometric perspective (following e.g., Bolnick *et al.* 2018) to quantify the extent of

'parallelism' in multivariate trait space; (iii) to evaluate whether the conclusions differ with

respect to the two visual systems modelled; and (iv) to assess patterns of phenotypic

differentiation (parallel or otherwise) in the context of population genetic structure. We use

genomic information obtained via pool-sequencing (hereafter, Pool-seq, Schlötterer *et al.*

2014) to estimate genome-wide population genetic differentiation. The population genomic

data also allow us to (v) conduct genome wide scans with genetic regions associated with

phenotypic divergence, and thus conduct a preliminary investigation into the genomic basis

of colour patterns in the wild.

MATERIALS and METHODS

I. Sampling of wild fish

Wild guppies were sampled from 16 sites across 7 rivers in the Northern Mountain Range of

Trinidad by seine nets (Table 1) in March 2017. One HP and one or two LP sites were

sampled in each river except the Paria (which contributed a single LP site). Fish were

transported to the University of West Indies (UWI, St Augustine, Trinidad), housed in 200L

aquaria and allowed to acclimate to laboratory conditions for 5-7 days. A subset of males

from each population were photographed (Table 1). Subsequently, different individuals (n ~

100 per population) from twelve of the 16 sites were shipped live to the University of Exeter

(Cornwall, UK). Although primarily for genetic investigations outwith the current study,

paternal sibships produced from these wild-caught fish were used in a pilot to validate the

present phenotyping strategy (see Supplemental Appendix A for details and results of

validation pilot).

II. Phenotyping of male colour traits

Visible and UV-spectrum photography (Figure 1) were carried out using a Samsung NX1000

camera converted to full spectrum, fitted with a Nikkor EL 80mm lens out on live, unsedated

fish housed in a custom designed UV-transparent water filled chamber, as detailed in

Appendix A. After photography, fish were humanely euthanized in a lethal dose of buffered

MS-222 and tissue preserved in RNAlater (Invitrogen). Photographs were analysed using

the QCPA framework, which quantifies colour traits using images and species-specific visual

system parameters (see www.empiricalimaging.com and van den Berg *et al.* 2019). In

addition to chromatic and achromatic signals, QCPA allows incorporation of spatial

information using the geometric arrangement colour patches (with discrimination of adjacent

patches based on the receptor noise estimate ΔS; Vorobyev and Osorio 1998). This has

ecological relevance because, in behavioral trials, female guppies prefer males that are

more conspicuous in colour pattern or have higher discriminant values (Sibeaux *et al.* 2019).

We quantified colour variation over the whole body (excluding head, tail and fins), using four

complementary traits: saturation boundary strength $sat_{\Box S}$, luminance boundary strength

$lum_{\Box S}$, chromaticity boundary strength $chrom_{\Box S}$ and chromaticity boundary strength variation

$CoV_{\Box S}$ (Endler *et al.* 2018). Together, these measures represent the chromatic (saturation

and hue), achromatic (luminance), and spatial (size and position of colour pattern elements)

properties of colour patterns. Simplistically, the first three of these can be thought of as

describing the "grayness" difference, "brightness" difference, and colour contrast between

neighbouring patches respectively. Higher values indicate higher internal contrasts across

colour-patch boundaries and/or more boundaries. The final metric describes the variation in

colour contrast between patches (i.e. it would be higher if there are both high-contrast *and*

low contrast internal boundaries). We used whole body trait measures for two reasons. First,

specific pattern elements (i.e. colour patches at particular body locations) differ greatly within

and among-populations. Thus, designating universally comparable phenotypes from discrete

pattern elements is problematic. Second, female guppies likely select males based on

groups of colour patches not single elements (Cole and Endler 2015).

For each photograph, the four traits were scored under two visual system parameter

sets; a conspecific (guppy) vision model and a predator (pike cichlid, *Crenacichla frenata*)

vision model. Note that *Crenacichla* is used because it is widely viewed as the most

important selective visual predator in guppies, although it is not present at the high predation

sites sampled in the Yarra and Marianne in this study. The parameters sets included user

specified values of (i) receiver spectral sensitivities (the relative abundance of their cones

and rods in the retina), (ii) receiver visual acuity, and (iii) physical distance at which the

receiver might be from a male guppy. The use of two visual systems reflects our view that

perception of male coloration may well differ between receiver species imposing selection

(i.e. a pattern that is more conspicuous to a female guppy may not always be so to a cichlid

predator).

Support for this possibility was provided by a pilot validation study used to check the

suitability of phenotyping methods for detecting genetic variation in male colour and pattern

(see Appendix A for full methods and results). Briefly photographs of males from known

paternal sibships produced using wild caught parents were processed using our QCPA

pipeline. Data analysis confirmed moderate to high sire level repeatabilities (analogous to

heritabilities) for all traits under both visual models as expected (since male colour is partly

Y-linked; Lindholm and Breden, 2002). This confirms the QCPA approach effectively

characterises genetically meaningful phenotypic variation. Importantly, (paternal) family

correlations (r_F) between homologous traits defined under alternative vision models (e.g.,

chrom$_{\square S}$ based on guppy and cichlid vision models) are positive, but also significantly less

than +1 in three of four cases. Using r_F as an estimate of the additive genetic correlation

(Astles et al. 2006), this indicates homologous metrics defined under the two vision models

are genetically distinct (if correlated) traits.

III. Generation of Pool-Seq data

To test characterise population genetic structuring we used a Pool-seq approach to obtain

allele frequency estimates (Gautier *et al.* 2013; Schlötterer *et al.* 2014). While this approach

is cost effective, available funding and tissue samples limited inclusion to 12 of the 16 wild

populations (with guppies from the Turure (TH and TL) and Guanapo (GL1 and GL2) rivers

excluded). With these exceptions, genomic DNA samples were obtained for 40 fish (20

males; 20 females) per population, using a Qiagen DNAeasy kit (Qiagen Co). For each

sample, the concentration and purity of the genomic DNA were measured and using a

Nanodrop spectrophotometer and q-bit (Thermo Fisher Scientific). Purity was further

checked on a 1% agarose gel before sex-specific DNA pools were created for each

population (n = 24; 12 populations by 2 sexes). Pools, each containing equal DNA

concentration from 20 individual fish, were sequenced at the Earlham Institute (Norwich,

UK). Barcoded DNA paired end libraries with insert size of approximately 150bp were

prepared using Illumina Truseq HT library prep and sequenced on 2 lanes using Illumina

NovaSeq (Illumina Inc., California, USA). Raw paired-end reads were checked for quality

using FastQC, and adapters were verified and removed using *cutadapt*. To investigate

among population differences, the raw reads of male and female pools for each population

were merged. Reads were then mapped to the reference guppy genome

(Guppy_female_1.0+MT, GCA_000633615.2, Künster *et al.* 2016) with BWA mem (Li and

Durbin 2010) using default parameters to generate initial BAM files. Aligned reads were then

sorted, marked for duplicates and indexed, to generate final BAM files using Samtools v1.9

(Li *et al.* 2009). All BAM files were merged to create a mpileup file (i.e. samtools mpileup –f),

which was subsequently filtered for indels and then used to generate a sync file

(mpileup2sync.jar; base quality > 20 or --min-qual 20) containing allelic frequency

information for every population using *Popoolation2* (v1.201) (Kofler *et al.* 2011). A clean

sync file was then obtained after indel removal (filter-sync-by-gtf.pl), and subsequently used

for downstream population genomic analyses.

IV. Data analysis

The phenotypic data set comprised measures of 8 response variables ($sat_{\square S}$, $lum_{\square S}$,

$chrom_{\square S}$, $CoV_{\square S}$ x 2 visual system models) for 388 male guppies representing 16 distinct

populations (6 HP and 10 LP) from 7 rivers. The genomic data comprised of 24 pools of

DNA-seq data (12 populations x 2 sexes), whose summary statistics are described in the

Supplemental (Table S1). In brief, each pool generated between 362-599M reads with a

mean Phred score of 35.8 and 96x depth of coverage. Between 88-98% of raw reads were

successfully mapped to the guppy female reference genome assembly. Unless otherwise

stated, all statistical analysis was done in R (R Core Team 2019) with ASReml-R (Butler *et*

al. 2017) used to fit mixed models. Full R codes and citations to all packages, bash scripts,

and Javascripts for colour analysis will be provided in the supplemental materials upon

publication.

Testing effects of population and predation regime

We first plotted trait means (by river and overall), to check whether differences between

populations and or predation regimes (HP versus LP) were visually apparent. We then fitted

univariate mixed models to each trait, including a random effect of *population* (i.e. sampling

site) and a fixed effect of predation *regime* (as a two level factor). This allowed us to

estimate variance among *populations* (V_{Pop}) conditional on the fixed effect of *regime* (i.e. as

the among-population variance not explained by *regime*). We scaled traits to standard

deviation units before analysis such that V_{Pop} can be interpreted as an intra-class correlation

unadjusted for fixed effects. We also estimated the adjusted population level repeatability

(R_{Pop}) for each trait (i.e., V_{Pop} as a proportion of total variance conditional on the fixed effect).

Statistical inference on the fixed effect was by Wald F tests and we used likelihood ratio

tests (LRT) to assess the significance of the random effect. For the latter, we assumed that

twice the difference in log-likelihoods between full and reduced (with no random population

effect) is distributed as a 50:50 mix of χ^2_0 and χ^2_1 (subsequently denoted $\chi^2_{0,1}$; Stram and Lee

1994).

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Multivariate models and estimation of pairwise phenotypic distance

We built a multivariate mixed model of all 8 traits. A random effect of population was

included on each trait but with no fixed effects (except intercepts). This yielded an estimate

of the among-population variance-covariance matrix (**D**), containing among-population trait

variances (but no longer conditional on any *regime* effects) on the diagonal, as well as

covariances that were scaled to among-population correlations (r_{Pop}) between all trait pairs.

We assume approximate 95% CI of $r_{Pop} \pm 1.96$ SE.

We then used the fitted multivariate model to predict pairwise phenotypic distances

between multivariate population means by extracting the best linear unbiased predictions

(BLUP) of trait means by population and calculated the 16x16 (since $n_{population}=16$) Euclidean

distance matrix (**E**) among populations in n_{trait}-dimensional trait space. We used all 8 traits

(**E**$_{all}$), but also generated the matrices based on guppy (**E**$_{gvm}$) and cichlid (**E**$_{cvm}$) vision model

traits respectively. Note that uncertainty in BLUP is not reflected in the point estimates of **E**$_{all}$,

E$_{gvm}$ and **E**$_{cvm}$, but we use these matrices to visualise patterns of among-population

divergence not as a basis for robust statistical inference. Specifically, neighbour-joining trees

based on **E**$_{all}$, **E**$_{gvm}$ and **E**$_{cvm}$ were plotted to assess any tendency to cluster by *regime*. For

each population pair, we also plotted the distance based on the guppy vision model against

the corresponding cichlid distance to assess whether (i) a relationship between $\mathbf{E}_{gvm}$ and

$\mathbf{E}_{cvm}$ is apparent, and ii) phenotypic distances tend to be lower for within- versus across-

regime population pairs.

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Testing for association of phenotypic and genetic distance

For the subset of 12 populations with molecular as well as phenotypic data, we used the

sync file to estimate genome-wide differentiation among populations using the fixation index

(F_{ST}). F_{ST} was calculated for single nucleotide polymorphisms (SNPs) in 50k windows using

fst.sliding.pl implemented in *Popoolation2* (parameters: –min-count 2 –min-coverage 4 –

max-coverage 80 –pool-size 80 –window-size 50000 –step-size 50000), which calculates

F_{ST} following Hartl and Clark (1997). The resultant among-population F_{ST} matrix was

compared to phenotypic distance estimates in two ways. First, we used Mantel tests to

assess the correlation with phenotypic distance (represented by $\mathbf{E}_{all}$, $\mathbf{E}_{gvm}$ and $\mathbf{E}_{cvm}$ sub-

setted to the 12 populations with SNP data). Second, following Pascoal *et al.* (2017) we

refitted univariate mixed models to each of the 8 response variables, but this time with no

fixed effects and a modified random effect structure to total V_{Pop} (i.e. not conditional on

regime) into two components: one attributable to divergence under neutral processes (i.e.

drift; $V_{Pop.N}$); and another attributable, under some assumptions, to divergence under

selection (from predation, sexual selection and/or unknown selective agents; $V_{Pop.S}$).

This second approach, follows the same premise as Q_{ST}-F_{ST} comparisons (Leinonen

et al. 2013), and asks whether there is more quantitative phenotypic divergence among-

populations than expected from levels of (putatively) neutral molecular divergence. To

implement it, we ran models with two random effects of population identity, one which

covaries among populations according to their molecular 'relatedness' structure, and a

second which is uncorrelated across populations. For the former we assume phenotypic

covariance between fish sampled from populations *i* and *j* that is attributable to neutral

population effects is equal to $S_{ij} \cdot V_{Pop.N}$ where S_{ij} is genome-wide 'similarity' between the

populations. Since F$_{ST}$ inscreases wih dissimilarity we define $S_{ij} = 1 - {F_{ST_{i,j}}}\big/{F_{ST_{max}}}$ where

$F_{ST_{max}}$ is the highest observed pairwise F$_{ST}$ among the populations sampled. This simply

scales S_{ij} from a maximum of 1 (when i=j) to 0 (for the most differentiated pair of

populations). To test whether trait variation among populations is greater than expected

under neutral divergence alone, we compared the model with both variances to one where

V$_{Pop.S}$ is absent using LRT. Two caveats to this method shoud be noted: first, it assumes that

to a first approximation genome-wide F$_{ST}$ measures neutral differentiation (i.e. from drift and

gene flow); second, since we are modelling phenotypic variation in wild caught fish rather

than in a common-garden experiment, any contributions to population divergence from

phenotypic plasticity are likely to be partitioned into V$_{Pop.S}$ (Pujol *et al.* 2008).

Multivariate models for estimating parallelism among replicate population pairs

Phenotypic analyses described above are 'blind' to the ancestral versus derived

status of HP and LP population pairs with rivers. Therefore, we also adopted recent

methods that use phenotypic change within lineages to quantify parallelism (De Lisle and

Bolnick 2020). We assume that within each river, an HP and LP pair can be viewed as

representing a 'lineage' in which the HP is ancestral. This assumption yielded 9 'lineages'

among the 16 populations with phenotypic data. Fish from the Paria are excluded as there is

no HP site, but the Aripo, Guanapo, and Marianne rivers each contribute two 'lineages' (i.e.

HP-LP comparisons). We describe the approach briefly, using notion from De Lisle and

Bolnick (2020) and referencing their equation numbers. First we calculated, the n$_{trait}$ row by

m$_{lineage}$ column data matrix **X**, where each element of **X** represents $\Delta z_{n,m}$ (the difference in

mean standardized phenotype (z) between HP and LP populations for trait n in lineage m;

Equation 3). Each row of **X** thus represents the vector of phenotypic change from HP to LP

within a lineage in multi-trait space. **X** was used to calculate **C**, the (m x m) among-lineage

correlation matrix of phenotypic change vectors which was then subject to eigen

decomposition (following Equations 5 and 6). To mirror our investigation of the among-

population distance matrices (i.e. $\mathbf{E}_{all}$, $\mathbf{E}_{gvm}$ and $\mathbf{E}_{cvm}$), we calculated and decomposed $\mathbf{C}$

matrices based on $sat_{\square S}$, $lum_{\square S}$, $chrom_{\square S}$ and $CoV_{\square S}$ determined using both vision models

($\mathbf{C}_{all}$), but also using guppy ($\mathbf{C}_{gvm}$) and cichlid vision model ($\mathbf{C}_{cvm}$) traits, separately.

In the case that multivariate trajectories are perfectly parallel across all lineages, the

first eigen vector of $\mathbf{C}$ should explain all variation and all lineages would load on this vector

with the same sign (but differing magnitudes if the extent, rather than direction, of phenotypic

change differed among lineages). In contrast, a uniform distribution of eigen values

(meaning low among-lineage correlations of change vectors), and/or a mixture of positive

and negative lineage-specific loadings on the first eigen vector is expected under non-

parallelism. As suggested by De Lisle and Bolnick (2020), for the case of fewer traits (here 4

or 8 depending on the version of $\mathbf{C}$) than lineages (here 9), we generated a null distribution

of m eigenvalues by simulating 1000 random vectors representing evolutionary change

under independence (among lineages) of multivariate phenotypic trajectories. While

acknowledging that power to reject the null (no parallelism) may be low (De Lisle and Bolnick

2020), observed eigenvalues were then compared to the simulated distribution. An

eigenvalue of $\mathbf{C}$ greater than 95% of values simulated assuming independence is taken as

statistical support for parallelism.

Genotype-phenotype association mapping from pool-seq data

Leveraging the pool-seq data, we tested for relationships between allele frequencies at SNP

loci and phenotypic variation (colour metrics: $sat_{\square S}$, $lum_{\square S}$, $chrom_{\square S}$ and $CoV_{\square S}$) and

predation regime (*high* vs *low*). We did this using genome-wide scans for association in

BayPass (Gautier, 2015). This method identifies loci more with higher than expected

differentiation among populations based on the *XtX* statistic (Gunther and Coop 2013), and

alsotests associations between SNPs and population-specific covariables (while accounting

for background population structure from drift and gene flow). Since direct selection on male

coloration is necessarily sex-limited, we elected to use SNP data from male pools only for

these analyses (rather than combined male and female data used for genome wide F_{ST}

estimates). Major and minor allele counts for each SNP were calculated for each population

from the previously compiled sync file using the *snp-frequency-diff.pl* script with the following

parameters (–min-coverage 75 –max-coverage 200). Using custom awk scripts, allele

counts were extracted and formatted for BayPass's input genotype file (Gautier 2015). The

large number of calculated SNPs from the male pools (ca. 5,977,803 SNPs) mainly on

known linkage groups (LGs) were then subsampled to yield about 100k SNPs along the

whole genome (Mean = 95,911 ± 61.2 SNPs from known LGs only, corresponding to about

4170 per LG) to generate about 100 sub-datasets. Subsets were then analysed individually

using the BayPass core and standard covariate models. The resulting population covariance

matrices of allele frequencies (interpretable as similarity matrices) were generated under the

core model and checked for consistency across all data subsets by calculating the distance

among matrices (hereafter FMD) using the *fmd.dist* function (Figure S1; Table S2). The

mean FMD among all pair wise comparison was 0.2 (with $r > 0.99$ for all pairwise

comparisons).

To identify SNPs that deviated from neutral expectations we used *simulate.baypass*.

This uses the population covariance matrix to simulated 'pseudo-observed datasets' (POD),

assuming SNP neutrality, consistent with the demographic history. This allows a null

distribution of the *XtX* to be generated, the 99% quantile of which was used as a significance

threshold (median *XtX* > 21.2) to determine whether observed SNPs were under selection.

We then used the BayPass IS covariate models, and resultant Bayes Factor (BF) to test for

association between SNPs and population level male colour traits (and/or predation regime)

(Gunther and Coop 2013; Gautier 2015). Evidence for significant association between SNP

and a tested covariate was based on BF > 30db (Gautier 2015).

Finally, we checked whether SNPs identified as deviating from neutrality and/or

associated with a covariate showed enrichment for gene ontology (GO) categories and

functions. We extracted the corresponding identity of genes or nearby genes, which and

blasting against the zebrafish genome (*Danio rerio* NCBI Refseq GCF_000002035.6, Howe

et al. 2013) for annotated orthologs. The annotated orthologs were run in the gene ontology

resource (http://geneontology.org/, Ashburner *et al.* 2000; The Gene Ontology Consortium

2019).

RESULTS

I. Non-significant predation regime effect on and among-population variance in male

colour patterns

Mean values of $chrom_{\Delta S}$, $cov_{\Delta S}$, $sat_{\Delta S}$ and $lum_{\Delta S}$ varied considerable between vision models

and populations (Figure 2a-d). Qualitatively, the guppy vision model detected less variability

among populations and tended to yield higher (more conspicuous) values for $chrom_{\Delta S}$ and

$sat_{\Delta S}$ relative to cichlid vision (with the converse for $cov_{\Delta S}$ and $lum_{\Delta S}$). Differences between

HP and LP regime were not consistent across traits and rivers, although overall (i.e. across

rivers) trait means were higher in LP guppies for all 8 traits.

Univariate mixed models confirmed this pattern, but provided limited statistical

support for differences in trait means between HP and LP (Table 2). Positive *regime*

coeffcients for all traits indicated a tendency to higher conspicuousness at LP sites; but

across the eight models, only $lum_{\Delta S}$ under the guppy vision model was significant (coefficient

(SE): 0.83 (0.37) sdu, P < 0.03; Table 2). A moderately large yet non-significant effect

(+0.57 sdu) was estimated for $chrom_{\Delta S}$ under the cichlid model, while the average estimated

regime effect size of *regime* +0.328 sdu. Support for strong among-population differentiation,

over and above any *regime* effects was unequivocal; LRT comparisons to reduced models

with no population effect yielding P<0.001 for all traits (Table 2). Population level

(conditional) repeatability R_{Pop} ranged from 30% to 71% with an average of 54%. R_{Pop}

estimates were very close to V_{Pop} indicating that *regime* explained little of the among-

population differentiation (if *regime* explained large amounts of trait variance, then in these

models $V_{Pop} + V_R$ would <1 leading to a systematic finding of $R_{Pop}>V_{Pop}$).

II. No clustering of guppy colour patterns by predation regime

The **D** matrix (estimated without conditioning on *regime*) yielded V_{Pop} estimates very similar

to (conditional) estimates obtained from univariate models (Table 3). The correlation

structure within **D** was not universally positive (which is the expectation if populations varied

along a simple axis from greater to less conspicuousness; Table S3). Population-level

correlations between homologous traits defined using guppy and cichlid vision models were

positive but significantly less than +1 (assuming an upper 95% confidence interval of r_{Pop} +

1.96 SE). Specifically r_{pop} (SE) were estimated as 0.65 (0.16), 0.39 (0.23), 0.28 (0.24) and

0.09 (0.28) across the two models for *sat$_{\Delta S}$, lum$_{\Delta S}$, chrom$_{\Delta S}$,* and *CoV$_{\Delta S}$* respectively.

Phenotypic distance matrices derived from the multivariate mixed model fit provided

little support for clustering of populations by *regime* in multi-trait space (Figure 3, Table S4).

Using distance defined from all traits at once and guppy vision models, there was arguably

some patterning (e.g., upper right portion of Fig 3a contains a grouping of 6 LP populations

with 1 HP whereas Fig 3b has a grouping of 6 LP and no HP), but this is less apparent using

just the traits subsets from the predator or cichlid (Figure 3c) vision model.

Pairwise phenotypic distances (using either guppy vision or cichlid vision traits) did

not suggest shorter average pairwise distances between populations within- versus across-

predation regimes (Figure 4). If clustering by *regime* was present, we would expect higher

phenotypic distances for low-high predation (LH) comparisons than for low-low predation

(LL) or high-high predation (HH) (i.e. the ellipse of red LH points in Figure 4 would be shifted

to higher values on x and/or y axes than the blue ones; Figure 4). Two further points emerge

from Figure 4; first, there is no apparent relationship between the pairwise population

distance estimates defined by the two vision models. Population pairs that are more distinct

with respect to guppy vision are not generally more distinct under the cichlid vision model.

This is consistent with the absence of strong positive r_{POP} between homologous traits noted

above. Second, the 95% confidence ellipse of pairwise distances for HH comparisons is

smaller and shifted to lower values (on both axes) relative to HL or LL comparison. This

pattern implies that there may be lower phenotypic variance among high predation

populations that among low predation populations. To investigate further, we fitted post hoc

multivariate mixed models (described Table S5) to regime specific data subsets, and

compared total phenotypic variance among-populations (calculated as the trace of regime

specific **D** matrices). Point estimates corroborated our interpretation of Figure 4, the trace of

D_H being 65% of the trace of D_L. However, unsurprisingly given the number of populations in

each regime, our trace estimates were characterized by high uncertainty and the difference

is not statistically significant.

III. Phenotypic differentiation does not align with putative neutral genetic divergence

Consistent with previous studies (e.g., Suk and Neff 2009; Willing *et al.* 2010; Fraser

et al. 2015), genome-wide pairwise F_{ST} comparisons revealed strong genetic structuring by

river and drainage (Figure 5; Table S6). There was almost no correlation between molecular

differentiation and estimated phenotypic distance (Mantel test of F_{ST} matrix correlation with

E_{all} r=-0.049, P=0.608; with E_{gvm} r=0.026, P=0.426; and with E_{cvm} r=-0.033, P=0.569). A

neutral model of among population divergence was also rejected (at α=0.05) for 6 out of 8

traits using our mixed model strategy to partition V_{Pop} into components attributable to neutral

($V_{Pop.N}$) and selective ($V_{Pop.S}$) components (Table 3). The proportion of trait variance

explained by the latter was considerably higher than by the former for most traits. Under the

cichlid vision model, $sat_{\Delta S}$ and $chrom_{\Delta S}$ proved exceptions to this pattern, with the more

complex model offering no improvement to model fit and $V_{Pop.S}$ being bound to zero. Thus,

with these two exceptions, genome wide-molecular divergence among populations cannot

readily explain the phenotypic differentiation structure.

IV. Non-parallelism in guppy colour pattern among lineages

Decompositions of the three **C** matrices revealed a highly skewed distribution of

eigenvalues (Table 4). However, since the rank of **C** is limited to the smaller of n (traits, here

either 4 or 8) and m (lineages, here 9) all three 9 x 9 **C** matrices will necessarily be rank

deficient. In this context, eigen values of zero are inevitable and not biologically informative.

Moreover, the first eigenvectors were not particularly dominant, capturing 55%, 51% and

40% of $\mathbf{C}_{gvm}$, $\mathbf{C}_{cvm}$ and $\mathbf{C}_{all}$ respectively, while lineages loaded on this first eigenvector with a

mixture of positive and negative signs. The structure of the $\mathbf{C}$ matrices was thus not

consistent with parallel phenotypic evolution among these lineages in 8-trait space, or in

either of the 4-trait spaces. Random vectors provided no evidence that the first eigen values

were larger than expected under a null model of independent evolutionary trajectories across

lineages (Figure S2). The second eigen values of $\mathbf{C}_{gvm}$ and $\mathbf{C}_{all}$ were greater than 95% of

simulated values, which suggested some overdispersion and thus deviation from complete

independence. However, lineages also loaded on this second eigenvector with a mix of

signs too (in all three $\mathbf{C}$ matrices) so we cannot interpret this as parallel evolution in a

direction defined by the second axis of $\mathbf{C}$. We suspect that this signal of non-independence

arises because data from single high predation populations in Aripo, Guanapo, and

Marianne rivers each contributed to the vectors of phenotypic change for two lineages (i.e.

HP vs LP comparisons within-river).

V. Genome-phenotype association analysis reveals that differentiated SNPs are found

in genes related to cell morphogenesis and neural development.

527   The core BayPass model revealed 85 SNPs more differentiated among populations

than expected under neutrality at the 0.01% POD significance threshold (Figure 6). Of these,

61 and 24 were located within genic and non-genic regions, respectively (summarized in

Table 5). Some of the genes were orthologous to zebrafish genes documented to be

involved in cell morphogenesis and cell projections, colour patterning and pigmentation (e.g.

xpc, rpl, rilp, netrins). Some were also within genes involved in neural development, such as

sensory axon guidance (e.g. *ptprfa*). Using orthologous zebrafish genes determined by

blasting, our GO analysis found no enrichment for any gene ontology category.

535   Using the BayPass IS model, we found a total of 195 SNPs associated with a tested

covariate using a BF threshold > 30db (colour traits: guppy vision model = 77; cichlid vision

model = 101; predation = 17, summarized in Table 6). Interestingly, there was no overlap

with the 85 differentiated SNPs identified under the core model. Taken at face value, this

suggests SNPs identified by the core model as being more differentiated than expected are

not related to colour pattern differences among populations. SNPs significantly associated

with covariates in the IS model were scattered throughout the genome (not clustered on

specific chromosomes), and found within both non-genic and genic regions (~42% in non-

genic regions). With a single exception (LG22: 9903588; associated with both measures of

sat$_{\Delta S}$) SNP-colour trait associations differed between between the guppy and cichlid visual

perceptual models, consistent with the finding that traits defined under the two models have

different genetic underpinnings. As with the *XtX*-based results from the core model, many

SNPs associated with covariates under the IS model were found in genes previously

implicated with cell morphogenesis and neuronal functions (e.g., *hapln2*; Table S7). There

was no enrichment of genes related to any functional processes using putative genes

identified from SNPs associated with traits under the guppy visual model. However, we

found enrichment of genes for nervous system development, anatomical structure

morphogenesis, and cellular process using SNPs associated with the traits defined usng the

cichlid visual model ($P < 0.0001$, Table S8). Interestingly, some of the non-genic SNPs with

elevated BF values were near genes that have known roles in teleost patterning (e.g. *bnc2,*

cdh11, for full list see Table S9 and S10).

DISCUSSION

Using colour traits defined by the visual systems of two major agents of selection on colour

in male guppies (conspecifics, cichlid predator), we found an overall tendency towards

higher, more conspicuous, phenotypic means in LP guppies as predicted. However, (i) trait

means are not systematically higher at LP for within-river comparisons, (ii) statistical support

for predation regime effects are weak, and (iii) regime effects explain little of the total among-

population differentiation. Moreover, when modelling colour variation in multivariate

phenotypic space we find (iv) little support for clustering of populations by predation regime,

and (v) no evidence for parallel evolution of lineages. Nevertheless, putatively neutral

patterns of genome-wide molecular differentiation did not readily explain the among-

population phenotypic structure. This suggests that adaptive evolutionary processes have

caused divergence of male colour among populations. Here, we first discuss these

phenotypic patterns and their interpretation in relation to the evidence for parallel evolution of

brighter male colouration across the high to low predation ecological transition in guppies.

We then comment on the results of our genomic analyses that revealed some SNPs

differentiated among populations in (or close) genes implicated in pigmentation, patterning

and neuronal development in other fish species.

Patterns of among-population phenotypic differentiation

Traits derived from the novel QCPA phenotyping pipeline offer only qualified support for the

longstanding view that guppies from LP populations are particularly colourful and

conspicuous (Haskins 1961; Endler 1980; Millar *et al.* 2006). Thus, trait means are higher at

LP overall, but this is not a statistically robust pattern. Nor is it found consistently across the

HP to LP transition within rivers. The lack of stronger patterns is perhaps somewhat

surprising, although we note that previous guppy studies have also reported similarly

inconsistent differences across regimes within rivers using transplant (Dick *et al.* 2018;

Kemp *et al.* 2009, 2018) and predator-manipulation experiments (Gotanda *et al.* 2018).

Colour differentiation between populations is strong using the QCPA phenotyping approach

(approximately half of all phenotypic variance is among-populations), but predation regime

does not explain much among-population variance (in single traits or multivariate

phenotype). This conclusion is further supported by the finding that populations do not

obviously cluster by predation regime in multi-trait space, regardless of whether this space is

defined in 8 dimensions (using all traits defined from guppy and cichlid visual models) or 4

(i.e. using just guppy or just cichlid perception). Nor does correlation structure in the

population variance-covariance matrix (**D)** supports the presence of a single major axis of

variation running from low (i.e. low trait values) to high conspicuity.

593   In principle these results could occur even with parallel evolution of replicate lineages

594  across the HP-LP transition, if recent coancestry, drift and/or ongoing gene flow between

595  populations are masking the phenotypic signature of parallelism. We think multiple lines of

596  evidence argue against this possibility. First, among the 12 populations with SNP data we

597  found no correlation of pairwise genome wide F_{ST} values and phenotypic distance. Second,

598  our use of mixed models to partition among-population variance suggest not only that neutral

599  processes are insufficient to explain among-population differentiation, but also that adaptive

600  divergence explains most variation (discussed further below). Third, and perhaps most

601  importantly, when isolating the phenotypic change from HP to LP within rivers, the inferred

602  directions of phenotypic evolution are not parallel among lineages. This analysis assumes

603  that river can be used as a proxy for lineage (*sensu* de Lisle and Bolnick 2020), but that

604  appears reasonable based on present and previous population genetic analyses (Willing *et*

605  *al.* 2010; Blondel *et al.* 2020).

606   Thus, viewing male colour as a complex multivariate phenotype seen through the

607  vision systems of biological agents of selection, we find no quantitative support for parallel

608  evolution across the HP to LP regime transition. Rather, our results suggest that – in this

609  phenotypic space – colour patterns evolve approximately independently in each river.

610  However, the fact that genome wide molecular genetic structure does not predict phenotypic

611  structuring argues against the idea that divergence of colour patterns could be primarily

612  neutral. How can these results be reconciled with our existing understanding of colour

613  evolution in guppies? In fact, we suggest the lack of parallelism, coupled to apparently

614  greater variation among LP than among HP populations, is consistent with the widely-held

615  view that reduced selection pressure from visual predators facilitates evolutionary

616  divergence of male traits by female choice. In guppies, the evidence that females prefer

617  more conspicuousness male patterns that also increase predation risk is abundant (reviewed

618  in Houde 1997). However, this (univariate) conceptualisation of female choice masks the fact

619  that, in multi-trait space, there may be many different directions that increase conspicuity to

620  females.

The extent to which female choice drives Fisherian runaway evolution of traits that

are 'arbitrary' (not under viability selection) versus 'costly' has been extensively debated

(Kokko *et al.* 2002). However, even if females do prefer costly male phenotypes in all

populations, this need not translate into closely aligned vectors of sexual selection on

multivariate phenotype. Indeed, Endler and Houde (1995) demonstrated substantial

geographic variation in female preferences for colour patches, e.g. black, orange, and colour

contrast. Among-individual (female) differences in preference, a prerequisite for heritable

variation that would allow trait-preference coevolution, have also been demonstrated (Brooks

and Endler 2001; Brooks 2002). Lastly, we note that frequency dependent selection, with

females preferring novel male phenotypes is well documented in guppies (Hughes *et al.*

2013), and will generally impose vectors of directional selection that differ in space (i.e.

across populations) as well as time (e.g. as rare phenotypes become more common).

Our conclusion that predation regime has no consistent effect on the direction of

evolution in multivariate trait holds irrespective of the vision model used to define

phenotypes. However, the finding of population level correlations significantly less than +1

between homologous traits illustrates the wider potential for QPCA to offer new insights. The

same objective colour phenotype may be perceived differently by prospective mates and

potential predators. Here, guppy colour patterns appear generally more conspicuous in the

chromatic channel for conspecifics, whereas the pike cichlid is more sensitive to achromatic

(luminance) elements (Weadick *et al.* 2018). Such differences highlight the value of adopting

colour measures appropriate to the visual systems of hypothesised selective agents (Endler

1978; Endler *et al.* 2005; Endler and Mielke, 2005), and have implications for the way we the

evolution of colour signals with multiple receiving species. We also acknowledge however,

that that our study is naive to any abiotic factors (e.g., substrate, water colour, light

transmission, and canopy cover) that may affect *in situ* perception of phenotype by guppies

and/or their predators. Local abiotic conditions are an explicit part of the wider sensory drive

hypothesis (Endler 1980, 1992), and Kemp *et al.* (2018) found that canopy cover influenced

the relative abundance of iridescent versus melanic colour patches in guppies. While we

have no evidence to suggest this, it is at least possible that conditioning among-population

(or lineage) variance on abiotic covariates would reveal greater support for parallel evolution

in the colour phenotype as measured here.

Insights from genome-scans and SNP associations

Our pool-sequencing data provide several genetic insights that complement the phenotypic

analyses. First, based on detected associations with the phenotypic traits defined here, male

colour patterns are probably highly polygenic with genes distributed across the genome

rather than being restricted to the sex chromosomes or otherwise clustered. Associated

SNPs are found both outside and within genic regions, and involve cis-acting elements. For

instance, a SNP (LG:31463600) associated with differences in guppy-specific saturation was

found in an intergenic region near basonuclin 2 (*bnc2*), a gene implicated in the maintenance

of extracellular environments within which pigment cells driving patterning reside (Lang *et al.*

2009). Third, the minimal overlap between SNPs associated with homologous traits defined

under the two vision models is consistent with the population level phenotypic correlations

being less that one, suggesting that homolgous traits defined under different models are

genetically distinct. However, since colour genetics have often been studied using human

visual perception (Tripathi *et al.* 2009), it also highlights the possibility that ecologically-

important genes may have been previously overlooked. Conversely, we did not detect any

trait-SNP association consistency with colour genes previously identified in guppies, i.e.

csf1ra and *kita* (Kottler *et al.* 2013). This may again be explained by our choice of defined

phenotype; focusing here on measures of overall body patterns and/or conspicuity (as

perceived by selective agents), rather than features of individual pigmented patches (as

perceived by humans). Fourth, no SNPs significantly associated with tested covariables

were actually among the set identified as being significantly more differentiated than

expected under neutrality in the core BayPass model. Taken at face value, this could

indicate that greater (adaptive) genetic divergence among populations has occurred for

specific traits or aspects of phenotype (e.g. life history) that were not quantified in this study.

To our knowledge, this is the first study to investigate among-population natural

variation in male guppy patterns at the genomic level. Consequently some brief, and

necessarily tentative, comments on specific SNPs identified are perhaps warranted. Among

those more differentiated than expected under neutrality, several SNPs were detected in the

protein tyrosine phosphatase receptor of types *Fa* (*ptprfa*) and N2 (*ptprn2*). Basic functions

of these genes include cell proliferation and epithelial cell-cell adhesion. They are

considered to counteract tyrosinase kinase receptors (Xu and Fisher, 2012), which play

known roles in melanocyte and melanophores development (*Kita* and *cKit*, respectively;

Alexeev and Yoon 2006; Larsson and Parichy 2019). It is not known if tyrosine phosphatase

receptors carry a direct pigmentation function, although this is plausible given their

documented involvement in cell differentiation, oncogenic- events, and sensory guidance to

the skin (Wang *et al.*2012). Several SNPs deviating from neutral expectaions were also

identified in less well-known genes including *polg* and *trio,* which are part of regulatory

networks influencing melanocyte differentiation (Seberg *et al.* 2017; Park *et al.* 2018).

Among SNPs associated with tested covariates (trait means and predation regime), we also

find several in (or near) genes with known roles in fish colour. For instance, one SNP

association with guppy vision sat$_{\Delta S}$ was near *ablim3*, a gene thought to be involved in

pigment cell movement in fishes (Ahi *et al.* 2020). Another was found within *cx30.3*, part of a

family of genes implicated in zebrafish skin development and pattern formation (Tao *et al.*

2010; Irion *et al.* 2014).

Conclusion

Trinidadian guppies ostensibly represent a classic example of parallel (or

convergent) evolution, with repeated divergence in male colour patterns between upstream

and downstream populations due partly to predation conditions. We find that the well-

described tendency for greater conspicuousness under low predation holds qualitatively true

when we analyse single traits defined under guppy and predator vision models. However,

statistical support for this is weak, and repeated colonisation of low predation habitats has

not lead to the parallel evolution of conspicuous phenotypes when these are characterised in

quantitative, multivariate, phenotypic space. Instead, we suggest that colonisation of LP

habitat may reduce selective constraints on phenotypic space imposed by predation,

facilitating population-specific divergence under sexual selection. Together, our results are

nonetheless consistent with the standard view that reduced selection by predators allows

male colour traits to evolve under sexual selection imposed by female choice. However, the

common assertion that female choice in guppies selects for 'brighter' or more 'conspicuous'

males, masks the fact that – in multivariate trait space - colour is evolving in rather different

directions across populations and lineages.

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DATA ACCESSIBILITY

Raw phenotypic and genomic data as well as R and Java scripts will be available publicly upon the acceptance of the manuscript in Dryad, GeneBank, and appropriate data repositories.

AUTHOR CONTRIBUTIONS

LY and AW designed the study with input from DPC and JT. LY, DPC and AW performed the research. JT and IR contributed analytical tools. LY and AW analysed the data. LY and AW wrote the paper with input from DCP, IR, and JT.

TABLES AND FIGURES

Table 1. Description of collection site (drainage, river, and regime type) and their GPS coordinates. Sample size denotes fish that were photographed.

Drainage	River	Regime	Site code	Lat (N)	Long (W)	n_{males}
Caroni	Aripo	high	AH	10° 39' 02"	61° 13' 26"	26
Caroni	Aripo	low	AL	10° 41' 08"	61° 13' 56"	29
Caroni	Aripo	low	AL2	10° 41' 10"	61° 13' 30"	23
Caroni	Guanapo	high	GH	10° 39' 28"	61° 15' 13"	26
Caroni	Guanapo	low	GL	10° 42' 42"	61° 16' 01"	5
Caroni	Guanapo	low	GL2	10° 42' 20"	61° 15' 48"	26
Northern	Marianne	high	MH	10° 45' 54"	61° 18' 15"	25
Northern	Marianne	low	ML	10° 44' 58"	61° 17' 13"	23
Northern	Marianne	low	PML	10° 46' 40"	61° 18' 04"	28
Northern	Paria	low	PL	10° 44' 57"	61° 15' 54"	24
Northern	Yarra	high	YH	10° 47' 23"	61° 21' 12"	24
Northern	Yarra	low	YL	10° 44' 25"	61° 19' 17"	26
Oropouche	Quare	high	QH	10° 39' 53"	61° 11' 35"	27
Oropouche	Quare	low	QL	10° 40' 33"	61° 11' 49"	26
Oropouche	Turure	high	TH	10° 40' 32"	61° 09' 54"	26
Oropouche	Turure	low	TL	10° 41' 31"	61° 10' 18"	24

Table 2. Estimated regime (fixed) and population (random) effects from univariate mixed models of each trait. Regime effects denote the consequence of low relative to high predation regime with inference by F-tests. Among-population variance (V_{Pop}) is conditional on fixed effects and was tested by likelihood ratio test. Also shown is the population level repeatability (R_{Pop}).

Vision model	Trait	Fixed effects				Random effects			
		Regime (SE)	F	DF	P	V_{Pop} (SE)	R_{Pop} (SE)	$X^2_{0,1}$	P
Guppy	sat_{DS}	0.148 (0.353)	0.177	1,13.9	0.681	0.439 (0.177)	0.418 (0.1)	150	<0.001
	lum_{DS}	0.833 (0.366)	5.171	1,13.7	0.04	0.484 (0.192)	0.539 (0.1)	204	<0.001
	$chrom_{DS}$	0.271 (0.434)	0.39	1,14	0.543	0.689 (0.267)	0.638 (0.091)	307	<0.001
	CoV_{DS}	0.230 (0.302)	0.58	1,13.8	0.459	0.307 (0.130)	0.301 (0.091)	80.5	<0.001
Cichlid	sat_{DS}	0.270 (0.449)	0.361	1,14	0.557	0.744 (0.286)	0.713 (0.08)	403	<0.001
	lum_{DS}	0.054 (0.401)	0.018	1,13.7	0.896	0.578 (0.230)	0.514 (0.101)	186	<0.001
	$chrom_{DS}$	0.570 (0.433)	1.728	1,14	0.21	0.690 (0.266)	0.682 (0.085)	357	<0.001
	CoV_{DS}	0.245 (0.400)	0.375	1,13.2	0.551	0.556 (0.231)	0.512 (0.106)	136	<0.001

Table 3. Estimates of among-population repeatabilities (R_{Pop}) derived from univariate mixed models that partition total population effects into components attributable to putative neutral ($R_{Pop.N}$) versus selective ($R_{Pop.S}$) processes. Also shown are LRT comparisons to a reduced model in which all among-population variance is assumed to have a neutral basis. Note where a variance component, and so the corresponding repeatability, was bound to zero to keep it in allowable parameter space no estimate of the SE is possible.

Vision model	Trait	R_{Pop} (SE)	$R_{Pop.N}$ (SE)	$R_{Pop.S}$ (SE)	$X^2_{0,1}$	P
	sat_{DS}	0.385 (0.109)	0.000 (-)	0.385 (0.109)	3.787	0.026
Guppy	lum_{DS}	0.414 (0.144)	0.073 (0.083)	0.342 (0.196)	4.537	0.017
	$chrom_{DS}$	0.654 (0.136)	0.031 (0.092)	0.623 (0.208)	4.95	0.013
	cov_{DS}	0.240 (0.089)	0.000 (-)	0.240 (0.089)	9.651	0.001
	sat_{DS}	0.418 (0.109)	0.418 (0.109)	0.000 (-)	0	0.5
Cichlid	lum_{DS}	0.364 (0.108)	0.000 (-)	0.364 (0.108)	5.286	0.011
	$chrom_{DS}$	0.400 (0.108)	0.400 (0.108)	0.000 (-)	0	0.5
	cov_{DS}	0.495 (0.114)	0.000 (-)	0.495 (0.114)	8.581	0.002

Table 4. Spectral decomposition for among independent lineages for phenotypic parallelism. Note: Highlighted cells indicates that that dimension is significantly above 95% threshold.

(a) guppy-vision colour

		Eigenvectors								
		q1	q2	q3	q4	q5	q6	q7	q8	q9
	eigenvalues (upper 95%)	5.06	3.25	2.13	1.26	0.00	0.00	0.00	0.00	0.00
	eigenvalues (obs)	4.97	3.38	0.40	0.26	0.00	0.00	0.00	0.00	0.00
	% Variance explained	0.55	0.38	0.04	0.03	0.00	0.00	0.00	0.00	0.00
	aripo	-0.43	0.08	-0.13	0.50	0.72	0.00	0.00	0.00	0.17
	aripo1	-0.41	0.17	0.29	-0.30	0.05	-0.01	0.71	-0.23	-0.25
	guanapo	0.35	0.29	-0.46	-0.18	0.21	-0.56	0.32	0.28	0.04
	guanapo1	0.18	-0.48	-0.20	0.45	-0.07	0.18	0.40	0.24	-0.50
lineage	marianne	-0.07	0.52	-0.32	-0.16	0.08	0.44	-0.26	0.10	-0.56
	marianne1	0.01	0.50	0.47	0.51	-0.29	-0.26	-0.02	0.32	-0.12
	quare	-0.42	-0.16	0.01	-0.28	-0.08	0.13	0.01	0.80	0.21
	turure	-0.38	-0.28	-0.04	-0.10	-0.02	-0.61	-0.38	-0.07	-0.49
	yarra	-0.40	0.14	-0.56	0.24	-0.58	-0.03	0.12	-0.21	0.23

(b) cichlid-vision colour

		Eigenvectors								
		q1	q2	q3	q4	q5	q6	q7	q8	q9
	eigenvalues (upper 95%)	5.06	3.19	2.12	1.26	0.00	0.00	0.00	0.00	0.00
	eigenvalues	4.59	2.42	1.85	0.14	0.00	0.00	0.00	0.00	0.00
	% Variance explained	0.51	0.27	0.21	0.02	0.00	0.00	0.00	0.00	0.00
	aripo	-0.16	0.56	-0.25	0.19	0.57	0.00	0.00	0.00	0.48
	aripo1	-0.43	0.20	0.10	-0.42	0.32	-0.37	0.21	-0.12	-0.54
	guanapo	-0.38	-0.34	-0.17	-0.32	-0.03	-0.29	-0.17	0.61	0.35
lineage	guanapo1	0.09	-0.50	-0.44	0.03	0.22	-0.37	-0.08	-0.59	0.12
	marianne	-0.28	0.02	-0.59	-0.15	-0.29	0.43	0.53	-0.06	-0.01
	marianne1	-0.41	0.26	-0.18	0.02	-0.36	0.09	-0.72	-0.26	-0.11

	q1	q2	q3	q4	q5	q6	q7	q8	q9
quare	0.44	0.08	-0.10	-0.76	0.20	0.32	-0.26	-0.05	0.06
turure	0.33	0.44	-0.14	-0.16	-0.50	-0.59	0.16	-0.03	0.17
yarra	-0.31	-0.06	0.54	-0.25	-0.16	0.09	0.18	-0.44	0.54

(c) Combined

		Eigenvectors								
		q1	*q2*	*q3*	*q4*	*q5*	*q6*	*q7*	*q8*	*q9*
	eigenvalues (upper 95%)	3.66	2.57	1.87	1.36	0.96	0.62	0.33	0.11	0.00
	eigenvalues	3.63	2.97	1.06	0.71	0.42	0.11	0.09	0.01	0.00
	% Variance explained	0.40	0.33	0.12	0.08	0.05	0.01	0.01	0.00	0.00
	aripo	-0.15	-0.50	-0.25	-0.17	0.37	-0.41	-0.27	0.36	0.36
	aripo1	-0.41	-0.24	0.27	-0.04	0.53	0.53	0.00	0.09	-0.36
	guanapo	-0.37	0.38	0.01	-0.25	0.03	0.28	-0.39	-0.32	0.57
	guanapo1	0.23	0.37	-0.08	-0.66	0.44	-0.27	0.11	-0.14	-0.28
lineage	marianne	-0.38	-0.04	-0.49	-0.38	-0.50	0.11	-0.13	0.20	-0.39
	marianne1	-0.47	-0.09	-0.31	0.14	0.11	-0.22	0.62	-0.45	0.08
	quare	0.29	-0.40	0.08	-0.50	-0.19	0.39	0.44	0.00	0.35
	turure	0.32	-0.40	-0.32	0.06	0.07	0.14	-0.39	-0.64	-0.19
	yarra	-0.25	-0.28	0.65	-0.24	-0.29	-0.41	-0.14	-0.29	-0.14

Table 5. Location (LG and bp position) of significantly differentiated SNPs among populations and putative underlying genes. Shaded cells indicate non-genic regions within which SNPs are found

LG	position	Median XtX	annotated gene	gene description	zebrafish ortholog
LG01	20679153	21.22	ctnna2	catenin (cadherin-associated protein), alpha 2	ctnna2
LG01	21630340	21.84	na		-
LG01	24267902	21.50	na		
LG02	735314	22.89	na		
LG02	3366912	21.21	mid1	midline 1	mid1
LG02	11968486	21.46	na		
LG02	18552459	21.67	LOC103480034	teleost multiple tissue opsin 2b	tmtops2b
LG02	18738097	21.30	LOC103480073	DCN1, defective in cullin neddylation 1, domain containing 2a	dcun1d2a
LG02	23502075	21.57	LOC103481844	potassium voltage-gated channel, subfamily H, member 3	kcnh3
LG02	26822612	21.43	LOC103457726	kalirin-like	kalrnb
LG02	27317080	21.23	na		
LG02	27737287	23.24	tnfsf13b	TNF superfamily member 13b	tnfsf13b
LG03	4026992	21.52	na		
LG03	10136707	22.11	LOC103462047	UPF0469 protein KIAA0907 homolog	uncharacterised
LG04	1388450	21.33	LOC108166264	collagen alpha-1 chain-like	col24a1
LG04	19135875	21.27	LOC103464057	sec1 family domain-containing protein 2	scfd1
LG04	22158698	21.54	LOC103463697	uncharacterised	uncharacterised
LG04	29560803	21.61	elavl4	ELAV like neuron-specific RNA binding protein 4	elavl4
LG04	30040468	21.89	LOC103464223	protein tyrosine phosphatase receptor type Fa	ptprfa
LG04	30061932	21.98	LOC103464223	protein tyrosine phosphatase receptor type Fa	ptprfa
LG04	30108556	21.22	LOC103464223	protein tyrosine phosphatase receptor type Fa	ptprfa
LG04	30110320	21.47	LOC103464223	protein tyrosine phosphatase receptor type Fa	ptprfa
LG04	30492932	21.26	LOC103464192	epidermal growth factor receptor substrate 15-like 1	eps15l1a
LG04	31086053	21.28	pde6d	phosphodiesterase 6D, cGMP-specific, rod, delta	pde6d
LG06	1451327	21.65	dusp6	dual specificity phosphatase 6	dusp6
LG06	7012709	22.19	tmem117	transmembrane protein 117	tmem117
LG06	15480429	21.76	polg	polymerase (DNA directed), gamma	polg
LG06	21928581	21.39	LOC103466341	netrin-4-like	ntn4

LG06	21928809	21.65	LOC103466341	netrin-4-like	ntn4
LG07	1303358	21.33	na		
LG07	9267800	21.91	xpc	xeroderma pigmentosum, complementation group C	xpc
LG07	22765503	21.85	dcaf1	DDB1 and CUL4 associated factor 1	DDB1- and CUL4-associated factor 1-like
LG08	2558818	22.29	LOC103468161	carbonic anhydrase-related protein 10-like	ca10a
LG08	9964866	21.68	clg8h16orf72	linkage group 8 C16orf72 homolog	zgc:66160
LG08	16159265	21.51	lfng	LFNG O-fucosylpeptide 3-beta-N-acetylglucosaminyltransferase	lfng
LG09	5667793	21.49	na		
LG09	7471619	21.61	LOC103469702	protein kinase, AMP-activated, beta 1 non-catalytic subunit, b	prkab1b
LG09	12715576	21.56	LOC103469911	uncharacterised	zgc:152968
LG09	26617339	21.21	LOC103470491	transient receptor potential cation channel, subfamily M, member 3	trpm3
LG09	27956069	21.37	na		
LG09	29207832	21.68	dtx1	deltex 1, E3 ubiquitin ligase	dtx1
LG09	33808366	21.49	prdm6	PR domain containing 6	prdm6
LG10	20059227	21.77	na		
LG10	22172121	21.26	LOC103471396	TSC22 domain family, member 3	tsc22d3
LG10	25748269	22.48	na		
LG11	5785615	21.36	LOC103472064	uncharacterised	uncharacterised
LG11	6043719	21.31	trio	trio Rho guanine nucleotide exchange factor	Trio
LG11	6221092	24.40	rpl14	ribosomal protein L14	rpl14
LG11	11324398	21.83	LOC103472245	calcitonin receptor	calcr
LG11	17282316	21.44	na		
LG11	19054623	21.34	LOC103472636	uncharacterised	uncharacterised
LG11	27311868	21.51	LOC103472993	cyclic GMP-AMP synthase-like	uncharacterised
LG11	28122311	21.85	LOC108166719	uncharacterised	uncharacterised
LG12	3570482	21.77	LOC104373253	macrophage mannose receptor 1-like	rilpl1
LG12	13100673	21.67	LOC103473585	matrix metallopeptidase 17a	mmp17a
LG12	16344388	21.39	na		

LG12	19606315	22.08	na		
LG12	23073848	21.24	LOC103474004	immunoglobulin superfamily, member 9a	igsf9a
LG12	23074083	21.44	LOC103474004	immunoglobulin superfamily, member 9a	igsf9a
LG13	14610669	21.79	tusc5	trafficking regulator of GLUT4 (SLC2A4) 1a	trarg1a
LG14	2734379	22.45	na		
LG14	13371933	21.39	na		
LG14	17753953	21.38	med13	med13	med13a
LG15	6448226	21.22	na		
LG15	7538997	22.53	na		
LG16	12769623	21.42	pianp	PILR alpha associated neural protein	pianp
LG17	963224	21.33	LOC103478874	TOX high mobility group box family member 4 a	tox4a
LG17	4792866	21.46	lrp8	low density lipoprotein receptor-related protein 8, apolipoprotein e receptor	lrp8
LG17	15468027	21.77	na		
LG17	30749000	22.27	slc39a6	solute carrier family 39 member 6	slc39a6
LG18	1955592	21.43	LOC103480181	neurexin-2-beta-like	uncharacterised
LG18	2195221	22.31	na		
LG19	1718516	21.38	rab11fip3	RAB11 family interacting protein 3 (class II)	rab11fip3
LG19	7450869	21.63	LOC103481309	uncharacterised	uncharacterised
LG19	14860722	21.29	ttyh2	tweety family member 2	ttyh2
LG19	21916746	21.34	LOC108167282	uncharacterised	uncharacterised
LG20	12634650	22.26	LOC103482504	uncharacterised	uncharacterised
LG20	12638410	21.91	LOC103482504	uncharacterised	uncharacterised
LG20	23266788	21.93	ptprn2	protein tyrosine phosphatase receptor type N2	ptprn2
LG21	1839279	21.90	LOC103457165	AT rich interactive domain 1B (SWI1-like)	arid1b
LG22	10792343	21.27	na		
LG22	13499891	21.28	kiz	kizuna centrosomal protein	kiz
LG23	1038549	21.87	LOC103459116	protein FAM180A-like	FAM180A
LG23	1038658	21.34	LOC103459116	protein FAM180A-like	FAM180A
LG23	10260495	22.01	na		

Figure 1. Example of photography setup using calibrated photography using visible and UV spectrum (grayscale). Black and white blocks are 5% and 95% reflectance standards, respectively.

Figure 2. Distribution of population mean trait values (prior to conversion to SDU) by predation regime and population under guppy and pike cichlid vision models for (a) colour saturation ($sat_{\Delta S}$), (b) luminance ($lum_{\Delta S}$), (c) chromaticity ($chrom_{\Delta S}$), (d) coefficient of variation in chromaticity ($chrom_{\Delta S}$). Lines link populations of differing regime within rivers. Also shown are the overall trait means by predation regime (red circle) with standard error. Asterisk (*) denotes significant difference in overall mean between regimes. Note that Paria (light blue square) does not have a corresponding high predation population.

Figure 3. Neighbour joining trees of populations based on estimated phenotypic distance matrices in n-dimensional trait space. Tree shown are determined using (a) all eight colour measures, (b) the four traits derived from the guppy-and (c) the four traits derived from the cichlid-vision model. Red and blue labels denote high and low predation *regime* populations, respectively.

Figure 4. Between-population distance estimates based on guppy (x-axis) and cichlid (y-axis) vision models. Each point denotes a single pairwise population comparison. Red points denote cross-regime distances between a high and a low predation site and blue points denote pairwise distances within predation regimes (dark blue for high, light blue for low), Ellipses illustrate approximate 95% confidence limits of the distributions and the 1:1 line is also shown.

Figure 5. Neighbor joining tree based on whole-genome population F_{ST} distance, clustering by rivers, then drainages, consistent with the independent evolution of each replicate. Branch length represents Fst. Populations: AH = Aripo High; AL = Aripo Low; AL2 = Aripo Low 2; GH = Guanapo High; MH = Marianne High; ML = Marianne Low; PML = Petite Marianne; PL = Paria; QH = Quare High; QL = Quare Low; YH = Yarra High; YL = Yarra Low. Blue denotes Low predation regime, whereas Red High.

Figure 6. Manhattan plot of genome-wide differentiation among 12 male populations of guppies. The red dotted line and points indicate the 0.01% (21.2 *XtX*) threshold and significantly differentiated SNPs, respectively. Red dots denote significantly differentiated SNPs. Labels above dots denote putative candidate genes implicated cell morphogenesis linked to color patterns.

Visible spectrum

UV spectrum

Fig. 5

Fig. 6

SUPPLEMENTARY FIGURES

Correlation map based on $\hat{\Omega}$

Hier. clust. tree based on $\hat{\Omega}$ ($d_{ij}=1-\rho_{ij}$)

Spectral decomposition for color patterns (guppy)

Spectral decomposition for color patterns (cichlid)

Spectral decomposition for color patterns (all)