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4   Peter D. Fields[1,2,4], David E. McCauley[3,] and Douglas R. Taylor[1]

[1]Department of Biology, University of Virginia, PO Box 400328, Charlottesville, VA

22904-4328

[2] Current Address: University of Basel, Zoological Institute, Vesalgasse 1, Basel, CH-

4051, Switzerland

[3] Department of Biological Sciences, Vanderbilt University, Nashville, TN 37235

[4] Corresponding author: Email: peter.fields@unibas.ch

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ABSTRACT

Previous studies of North American populations of the invasive plant *Silene latifolia*

showed significant cyto-nuclear linkage disequilibrium (CNLD) between SNP

variants of a mitochondrial gene (*atp*1) and the most common allele at nuclear

microsatellite loci. Fields *et al.* (2014) hypothesized that this CNLD arose partially as

a consequence of admixture that occurred during the colonization of North

American (NA) populations of *S latifolia* via seed dispersal from genetically

differentiated European populations that represent a portion of the native range of

this species. In order to evaluate the plausibility of the admixture hypothesis, as

opposed to metapopulation processes alone, we estimated CNLD for these same loci

using data collected from eastern (EEU) and western (WEU) European populations

of *S. latifolia* known to be genetically differentiated and likely sources of the spread

of the study species to North America. We show that the CNLD found previously in

NA populations of *S. latifolia* can be attributed to admixture of the previously

isolated European demes coupled with decay since that time. Our applied

framework allows the separation of the forces generating and dissolving statistical

associations between alleles in cytoplasmic organelles and the nuclear genome and

may thus be of utility in the study of plant or animal microbiomes.

INTRODUCTION

The magnitude and distribution of genetic variation across the recently established range of an invasive species is a reflection of the demographic history of demes from the native range, colonization events, and other evolutionary processes such as natural selection and genetic drift within the introduced range (Taylor & Keller, 2007, Dlugosch & Parker, 2008, Hufbauer, 2008, Calsbeek et al., 2011, Colautti & Barrett, 2011, Keller et al., 2014). For our purposes genetic variation can be defined as the degree of polymorphism at some number of loci as measured in aggregate across the new range, or as measured both within local populations as well as among those populations (i.e., considering the geographic population genetic structure). Measurements of variation could be extended beyond the standard locus-by-locus variety to include the distribution within and among populations of various multi-locus genotypes. These could include evaluation of non-random associations of alleles across loci (linkage disequilibrium or LD), again considering the genetic composition of all populations in aggregate or on a population-by-population basis.

In order to understand the determinants of genetic variation described above, it would be useful to consider mechanisms of short and long-range gene dispersal, along with the influence of the mating system, on multi-locus genotypic diversity. Note that in many species of seed plants most genes have the potential to disperse in both seeds and pollen. However, because the chloroplast and mitochondrial genomes tend to be extensively (or entirely) maternally inherited (McCauley, 2013) the movement of pollen would appear to be a more important

mechanism for long distance dispersal of nuclear genes as compared to that of

organellar genes. This distinction would appear to be especially relevant to the

emigration of genetic diversity during the early stages of an invasion since a new

population cannot be founded initially by pollen alone and thus requires seed

dispersal (Baker & Stebbins, 1965). On the other hand, long distance seed dispersal

can transport cyto-nuclear genotype combinations that exist in the source

populations (Hu, 2008). That is, when cyto-nuclear linkage disequilibrium (CNLD) is

found in an invasive plant species one must consider whether that CNLD is a

reflection of the genetic composition of the populations from which colonists were

drawn. The possibilities include that the CNLD occurred in some populations in the

native range, or that it was created as colonists from genetically differentiated

populations in the native range underwent admixture at or after the time of the

invasion. However, it is known from theory that CNLD found in neutral variation

should decline by 50% per generation in panmictic populations so CNLD persistence

is problematic whatever the origin (Fields et al., 2014, Hu, 2008). Note also, though,

that numerous species of plants are predominant selfers, and that recurrent self-

fertilization limits the opportunity for the mixing of alleles between genomes. In

addition to helping to reveal the history of populations regarding patterns of gene

flow and mating system, study of CNLD can also be useful when considering the role

of cyto-nuclear co-evolution, including selective phenomenon leading to

coadaptation or conflict between the genomes (Fields et al., 2014, Sloan, 2015, Bock

et al., 2014) and the transfer of genetic information between genomes (Sloan, 2015).

In a recent study of the plant *Silene latifolia*, Fields et al. (2014) identified

high levels of cyto-nuclear linkage disequilibrium (CNLD) between variants of a SNP

that occurs within the mitochondrial gene *atp1* and alleles at a number of nuclear

microsatellite loci. Samples were taken from populations located in southwestern

Virginia, USA, within what is assumed to be the eastern portion of the invasive range

of *S. latifolia* in North America (Keller et al., 2012). A comparison to genotypes found

in the same *S. latifolia* populations sampled 20 years ($\sim$ 7 generations) prior to the

more recent samples revealed only modest change in mean CNLD through time.

Fields *et al.* (2014) speculated that the CNLD that was observed could be traced to

1) metapopulation dynamics, in particular founder effects associated with a

propagule pool (Slatkin, 1977, Wade & McCauley, 1988, Whitlock & McCauley, 1990)

colonization dynamic, and 2) the colonization process by which the European native

S. latifolia became established in this region of the eastern U.S. via seed dispersal.

The relative temporal stability of the observed CNLD could be due to the fact that *S.*

latifolia displays a high degree of local population structure in this area (McCauley,

1994) and that this would limit opportunities for the outbreeding needed for the

erosion of any CNLD that arose when those populations were established. A link

between CNLD and the establishment of populations would be especially likely if the

various European populations contributing colonists (either directly or indirectly

via dispersal from other recently established North American populations)

displayed some degree of cyto-nuclear differentiation that could contribute to CNLD,

especially with admixture.

Extensive recent sampling across much of the putative native range of *S.*

latifolia in Europe has revealed geographic genetic structuring that is especially

noticeable when populations from eastern and western Europe are compared

(Keller et al., 2012, Keller et al., 2014, Taylor & Keller, 2007). Here we investigate

the plausibility that the CNLD in North American populations found by Fields et al.

(2014) could be a consequence of the recent establishment of *S. latifolia* in North

America and admixture between divergent European populations during

hypothetical invasion events. We find that different invasion scenarios, combined

with different forms of cyto-nuclear population differentiation in the native range,

could have enhanced or diminished the CNLD observed in the invasive range.

Application of cyto-nuclear genotyping, in addition to nuclear genotyping alone,

provides additional insights for the origin of statistical associations of genetic

variation among genomic regions, and the relative roles of historical demography

and ongoing metapopulation processes.

Theoretical expectations for shifts in CNLD

Consider a population of individuals that are characterized at two loci, one

cytoplasmic (C) and one nuclear (N) with two alleles at each locus (C_1 & C_2; N_1 & N_2).

This population could be described according to the frequency (p) of one of the

alleles at each locus (i.e. $p_{C,1}$ and $p_{N,1}$) and by the frequency of the cytonuclear

genotypes (i.e. $p_{C,N,1\ etc}$). The disequilibrium between C_1 & N_1, D_{C-N}, could be defined

in this population as $D_{C-N} = [p_{C,N,1} - ((p_{C,1})(p_{N,1}))]$. Thus, D_{C-N} determines whether two

alleles are associated within individuals, more (positive) or less (negative) often

than random and is mathematically equivalent to the covariance between alleles.

Note that the equation for LD is closely related to a covariance (symbolized as D)

with maximum values constrained by the gene frequencies in the sample. Another

standardized measure of CNLD (D') is also commonly reported (e.g. in Fields et al.

(2014)). This standardization constrains values of D' to fall between -1 and 1

(negative and perfect association of alleles, respectively).

Now consider $p_{C,1}$, $p_{N,1}$ and $p_{C,N,1}$ for each of k = 2 or more populations

individually. From this information, the overall CNLD value could be calculated in

two ways. First, D_{C-N} can be calculated for each of the k populations, and assuming

the populations are identical in size, these values could be averaged to obtain $\overline{D}_{C-N}$.

Alternatively, individual genotypes could be pooled across the k populations and

used to calculate values of $p_{C,1}$, $p_{N,1}$, and $p_{C,N,1}$ $p_{C,1}p_{N,1}$ and D_{C-N}^{T} based on the pooled

total set of individual genotypes.

Cytonuclear disequilibrium, and LD in general, may arise from several

sources or by processes operating at different spatial scales. First, at a larger spatial

scale, CNLD will arise naturally as a result of population genetic divergence at both

nuclear and cytoplasmic loci. Individual populations may also display variation in

the population-specific values of D_{C-N} according to how the alleles at the nuclear and

cytoplasmic loci are organized into two-locus genotypes in each population, which

could be a function of a variety of forces operating within populations (e.g. mating

structure, demographic history, selection, etc.).

One might predict how the cyto-nuclear genetic structure of a set of k

populations determines $\mathbf{D^T_{C\text{-}N}}$ in a population originating from mixing individuals

from those populations. It can be shown that

$$D^T_{C\text{-}N} = [\bar{D}_{C\text{-}N} + Cov(p_{C,1}, p_{N,1})], \quad (1)$$

where $Cov(p_{C,1}, p_{N,1})$ is the covariance of the k population-specific cytoplasmic and

nuclear allele frequencies across populations contributing equally to the mixture (p.

550, Hedrick, 2011). Note, this equation does not apply directly to D'.

Figure 2 illustrates the application of this equation to several hypothetical

sets of $k = 2$ source populations. From this equation it can be seen that mixing

individuals from more than one source can result in a $D^T_{C\text{-}N}$ in the new population

higher than the average $D^T_{C\text{-}N}$ in the source populations if $Cov(p_{C,1}, p_{N,1}) > 0$ is

(Figure 2, scenario 1). $Cov(p_{C,1}, p_{N,1})$ would be positive if certain combinations of

cytoplasmic and nuclear alleles coexist in the same populations more often than a

random association across populations. Mixing can diminish $D_{C\text{-}N}$ if $Cov(p_{C,1}, p_{N,1})$ is

negative (Figure 2, scenario 2; scenario 3), as would be the case if certain cyto-

nuclear alleles co-occur less often than random in the source populations. It can

have no effect on $D^T_{C\text{-}N}$ as compared to average population-specific CNLD (ignoring

stochastic sampling) if $Cov(p_{C,1}, p_{N,1}) = 0$, as (Figure 2 scenario 4; scenario 5).

$Cov(p_{C,1}, p_{N,1}) = 0$ would be the case if the nature of cyto-nuclear associations is

consistent across populations.

CNLD in the European range of S. latifolia

Samples of *S. latifolia* were collected from 85 individuals distributed across

12 populations in the species' native European range (Figure 1), a subset of the total

sample used by Keller et al. (2012). Genomic DNA was extracted as described in

Keller et al. (2012) using methods adapted from Edwards et al. (1991) and Slotta et

al. (2008). Individuals were genotyped at 14 microsatellite markers derived from *S.*

latifolia (slat18, slat32, slat33, slat48, slat72, slat85) described by Molecular Ecology

Resources Primer Development et al. (2010), and (SL_eSSR04, SL_eSSR06,

SL_eSSR09, SL_eSSR12, SL_eSS16, SL_eSSR20, SL_eSSR29, SL_eSSR30) described by

Moccia et al. (2009). Fragment analysis was conducted using an ABI 313xl

automated sequencer, with genotypes determined using GENEMAPPER v.3.0

software (Applied Biosystems) and binned using the program TANDEM (Matschiner

& Salzburger, 2009).

We followed the methods described by Keller et al. (2012) in order to

distinguish the two distinct, post-glacial demes in *S. latifolia*'s native range. Briefly,

individuals are assigned to Eastern and Western Europe (EEU and WEU,

respectively) based upon a suture zone running north from the Alps to the North

Sea, and has been described previously as different routes of expansion for *S.*

latifolia out of distinct Mediterranean refugia following Holocene warming (Hewitt,

1996, Petit et al., 2003, Taylor & Keller, 2007, Keller et al., 2009, Keller et al., 2012).

Using these *a priori* expectations, Keller et al. (2012) applied Bayesian clustering

using the program STRUCTURE v.2.2 (Pritchard et al., 2000, Falush et al., 2003),

assuming an admixture model with allele frequencies correlated and remaining

parameters set as a default over 10 independent runs with 200,000 MCMC

iterations and a burn-in period of 50,000 iterations. Here we utilize cluster

assignments assuming K=2 representing EEU and WEU, respectively (Keller et al.,

2012, Taylor & Keller, 2007).

In order to estimate CNLD, we followed the approach of Fields et al. (2014).

The same genomic samples used for the microsatellite analysis were assayed for a

SNP in the mitochondrial gene *atp1* using a PCR/RFLP method (McCauley & Ellis,

2008, Fields et al., 2014). PCR products were digested with *AluI* (New England

Biolabs), electrophoresed on 4% Metaphor agarose gel (Lonza Inc, Rockland, ME,

USA), and stained with ethidium bromide in order to resolve the two possible

alleles.

We calculated the observed (H_O) and expected (H_E) heterozygosity of our

nuclear genetic markers using the software GenoDive version 2.0b21 (Meirmans &

Van Tienderen, 2004). Estimates of genetic substructure using hierarchical F-

statistics were calculated using the software FSTAT 2.9.3.2 (Goudet, 2002), with

significant deviations from panmixia assessed by testing for Hardy–Weinberg

Equilibrium with 10,000 permutations and ∂=0.05.

We estimated nuclear-nuclear LD (hereafter, nuclear LD) following the

methods described in Fields et al. (2014). Statistical significance of nuclear LD

between pairs of loci for post-glacial demes, as well as within the full European

sample, was determined using a Monte-Carlo approximation of Fisher's exact test

implemented in the software Arlequin (Excoffier & Lischer, 2010), which assumes a

null hypothesis of random allelic assortment. Arlequin uses a Markov chain

extension of Fisher's exact test for *RxC* contingency tables (Slatkin, 1994, Li &

Merila, 2010). 100,000 alternative tables were explored by the Markov chain (Slate

& Pemberton, 2007, Li & Merila, 2010).

We follow the methods described in Fields et al. (2014) to estimate cyto-

nuclear D and D'. Briefly, we followed the approach of Basten and Asmussen (1997),

using the program CNDm to estimate a standardized estimate an allelic D and D'

between each nuclear locus and atp1 mtDNA locus. CNDm uses a Monte-Carlo

approach to approximate Fisher's exact test for RxC contingency tables and tests for

significant deviations from the null hypothesis of no allelic association (Basten &

Asmussen, 1997). For this analysis, all nuclear loci were treated as bi-allelic by

pooling all alleles other than the most common allele in the full European sample

into a single composite allele (as in Latta et al., 2001).

RESULTS AND DISCUSSION

Cyto-nuclear genotypes were obtained for a total of 85 European (EU)

individuals (52 individuals from the Western European Deme (WEU) and 33

individuals from the Eastern European Deme (EEU)). For the *atp1* SNP the most

common variant over the full European range occurred at a frequency of 0.535

(0.929 in EEU, and 0.254 in WEU). Table 1 presents a summary of global estimates

of nuclear genetic diversity, population substructure (F_{IT}) and a summary of nuclear

LD for the EEU, WEU, and full EU sample. All microsatellite loci were polymorphic (>

1 allele) in both regions, and ranged between 2 and 21 alleles. As in Fields et al.

(2014), individual loci showed a large range of H_O (0.061-0.761, mean = 0.38) and

H_E (0.134-0.902, mean = 0.63). Average F_{IT} for the whole EU sample was 0.498

(0.17-0.94), with all loci exhibiting significant ($P < 0.05$) deviations from panmixia,

corroborating deviations between H_O and H_E. The observed F_{IT} was larger than that

reported in Fields et al. (2014), which is to be expected given differences in scale

and the opportunity for admixture that took place during the invasion of the North

American continent (Keller et al., 2012). The average number of loci pairs showing

significant nuclear-nuclear LD was approximately equal in WEU and EEU (3.86 and

5.71, respectively), though the full sample showed approximately twice as much

(9.86). Table 2 presents the observed distribution of D and D' values for EEU, WEU,

and the full EU sample, respectively, for each of 14 nuclear microsatellites and the

mitochondrial *atp1* SNP. Similarly to Fields et al. (2014), overall mean levels of LD

were consistent, while specific locus values varied widely.

The main goal of this study was to ask whether the levels of CNLD seen by

Fields et al (2014) in a North American (NA) sample of *S. latifolia* could plausibly

have resulted from admixture of colonists from the European range of this species,

Given this, it would seem useful to compare values of D and/or D' obtained from

North America (NA) to those from Europe. To that end we have included the pooled

values of *D'* from Table 2 of Fields *et al.* (2014) in their Table 2, along with

associated values of D calculated from the same data but not presented in Fields et

al. (2014).

From our Table 2 it can be seen that the range of *D'* for NA was -.372 to.487

(mean = .086). For EU the range was -.536 .to .557 (mean = .043). In a similar

comparison of D values the range for NA was -.078 to .1042 (mean = .016). For EU

the range was -.075 to 0.52 (mean = -.009)

Given a fair degree of locus-to-locus variation in the nuclear locus-specific

values of D' and D in EU and NA, an additional method a comparing the similarity of

CNLD in EU and NA would be to examine the consistency of these locus-specific

values by calculating their product-moment correlation (r). That correlation is r = -

.208 for D' and r = -0.397 for D. Neither r value is significantly different from zero

under a null hypothesis of $r = 0$ ($p > 0.05$, df = 12 in each case), suggesting drift-like

dynamics are responsible for shifting CNLD rather than selective agents.

What is the impact of the covariance of cyto-nuclear allele frequencies across

EEU and WEU on EU D values? Recall from eq. 1 and Fig 1 that represent various

"scenarios" that such covariance can either increase or decrease the magnitude of

EU D relative to the simple mean of site-specific D values. Inspection of Fig. 3A

suggests that all of the scenarios are approximated, depending on the nuclear locus

in question. Locus-specific Shifts in CNLD D values range from a shift of .058 (locus

SLeSSR12) to a decrease of -.1062 (locus Slat 33) (mean shift in D = -.002).

Examination of Table 2 and Figure 3A suggests a large role of differing

frequency of the most common cytoplasmic allele, wherein the greater frequency of

the allele in EEU leads to low values of D for most loci, and a larger range of D in

WEU. Figure 3B describes a similar pattern for D', wherein the near fixation of one

cytoplasmic allele in the EEU deme generates a very strong signal of CNLD across 12

of 14 nuclear loci. However, both Figure 3A and 3B imply just how drastically

admixture of two highly divergent demes can shift the overall mean and variance of

CNLD. Thus, our results are in least partial support of our previous hypothesis that

the CNLD found in our North American *S. latifolia* samples could result, in part, from

admixture of colonists from two or more sources in the native range of *S. latifolia*

that contributed to the expansion of the range of this species into eastern North

America.

CONCLUSION

Linkage disequilibrium, either nuclear-nuclear or cyto-nuclear, provides a useful

summary of the demographic processes taking place within and among species. While

estimation of nuclear-nuclear LD is common, estimation of CNLD is much less so. This

is at least in part due to a much less extensive exploration, both theoretical and

empirically, of CNLD. This state seems particularly problematic given the general

realization of the role of cyto-nuclear interactions in important evolutionary processes,

including local adaptation (e.g. Bock et al., 2014) and speciation (e.g. Trier et al., 2014),

as well as co-evolutionary dynamics more generally (Sloan, 2015). Fields et al. (2014)

provided a temporal study of the how CNLD will shift over time as a result of

metapopulation processes. The present study shows that the incomplete admixture of *S.*

latifolia in North America also contributed to CNLD, as well as outlining how allele

frequencies in the species native range resolved into present patterns of CNLD in the

invaded range. Application of the presented framework, and modifications therein,

should allow for a fuller understanding the role of selective and non-selective forces in

generating patterns of LD both within and among genomes.

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REFERENCES

Baker, H. G. & Stebbins, G. L. 1965. *The genetics of colonizing species*. Academic Press,
New York.

Basten, C. J. & Asmussen, M. A. 1997. The exact test for cytonuclear disequilibria.
Genetics **146**: 1165-1171.

Bock, D. G., Andrew, R. L. & Rieseberg, L. H. 2014. On the adaptive value of
cytoplasmic genomes in plants. *Molecular Ecology* **23**: 4899-4911.

Calsbeek, B., Lavergne, S., Patel, M. & Molofsky, J. 2011. Comparing the genetic
architecture and potential response to selection of invasive and native
populations of reed canary grass. *Evolutionary Applications* **4**: 726-735.

Colautti, R. I. & Barrett, S. C. H. 2011. POPULATION DIVERGENCE ALONG LINES OF
GENETIC VARIANCE AND COVARIANCE IN THE INVASIVE PLANT LYTHRUM
SALICARIA IN EASTERN NORTH AMERICA. *Evolution* **65**: 2514-2529.

Dlugosch, K. M. & Parker, I. M. 2008. Founding events in species invasions: genetic
variation, adaptive evolution, and the role of multiple introductions.
Molecular Ecology **17**: 431-449.

Edwards, K., Johnstone, C. & Thompson, C. 1991. A simple and rapid method for the
preparation of plant genomic DNA for PCR analysis. *Nucleic Acids Research*
19: 1349-1349.

Excoffier, L. & Lischer, H. E. L. 2010. Arlequin suite ver 3.5: a new series of programs
to perform population genetics analyses under Linux and Windows.
Molecular Ecology Resources **10**: 564-567.

Falush, D., Stephens, M. & Pritchard, J. K. 2003. Inference of population structure

using multilocus genotype data: Linked loci and correlated allele frequencies.

Genetics **164**: 1567-1587.

Fields, P. D., McCauley, D. E., McAssey, E. V. & Taylor, D. R. 2014. Patterns of cyto-

nuclear linkage disequilibrium in Silene latifolia: genomic heterogeneity and

temporal stability. *Heredity* **112**: 99-104.

Goudet, J. 2002. FSTAT, a program to estimate and test gene diversities and fixation

indices (version 2.9.3.2). *Available from*

http://www2.unil.ch/popgen/softwares/fstat.htm.

Hedrick, P. W. 2011. *Genetics of Populations,* 4th ed. Jones and Bartlett, Boston.

Hewitt, G. M. 1996. Some genetic consequences of ice ages, and their role in

divergence and speciation. *Biological Journal of the Linnean Society* **58**: 247-

276.

Hu, X.-S. 2008. Barriers to the Spread of Neutral Alleles in the Cytonuclear System.

Evolution **62**: 2260-2278.

Hufbauer, R. A. 2008. Biological invasions: Paradox lost and paradise gained. *Current*

Biology **18**: R246-R247.

Keller, S. R., Fields, P. D., Berardi, A. E. & Taylor, D. R. 2014. Recent admixture

generates heterozygosity–fitness correlations during the range expansion of

an invading species. *Journal of Evolutionary Biology* **27**: 616-627.

Keller, S. R., Gilbert, K. J., Fields, P. D. & Taylor, D. R. 2012. Bayesian inference of a

complex invasion history revealed by nuclear and chloroplast genetic

diversity in the colonizing plant, *Silene latifolia*. *Molecular Ecology* **21**: 4721-

4734.

Keller, S. R., Sowell, D. R., Neiman, M., Wolfe, L. M. & Taylor, D. R. 2009. Adaptation

and colonization history affect the evolution of clines in two introduced

species. *New Phytologist* **183**: 678-690.

Latta, R. G., Linhart, Y. B. & Mitton, J. B. 2001. Cytonuclear disequilibrium and genetic

drift in a natural population of ponderosa pine. *Genetics* **158**: 843-850.

Li, M. H. & Merila, J. 2010. Extensive linkage disequilibrium in a wild bird population.

Heredity **104**: 600-610.

Matschiner, M. & Salzburger, W. 2009. TANDEM: integrating automated allele

binning into genetics and genomics workflows. *Bioinformatics* **25**: 1982-

1983.

McCauley, D. E. 1994. Contrasting the distribution of chloroplast DNA and allozyme

polymorphism among local populations of Silene alba: implications for

studies of gene flow in plants. *Proceedings of the National Academy of*

Sciences of the United States of America **91**: 8127-8131.

McCauley, D. E. 2013. Paternal leakage, heteroplasmy, and the evolution of plant

mitochondrial genomes. *New Phytologist* **200**: 966-977.

McCauley, D. E. & Ellis, J. R. 2008. Recombination and linkage disequilibrium among

mitochondrial genes in structured populations of the gynodioecious plant

Silene vulgaris. *Evolution* **62**: 823-832.

Meirmans, P. G. & Van Tienderen, P. H. 2004. GENOTYPE and GENODIVE: two

programs for the analysis of genetic diversity of asexual organisms.

Molecular Ecology Notes **4**: 792-794.

Moccia, M. D., Oger-Desfeux, C., Marais, G. A. & Widmer, A. 2009. A white campion

(*Silene latifolia*) floral expressed sequence tag (EST) library: annotation, EST-

SSR characterization, transferability, and utility for comparative mapping.

BMC Genomics **10**: 243.

Molecular Ecology Resources Primer Development, C., Abdoullaye, D., Acevedo, I.,

Adebayo, A. A., Behrmann-Godel, J., Benjamin, R. C., Bock, D. G., Born, C.,

Brouat, C., Caccone, A., Cao, L.-Z., Casado-AmezÚA, P., CatanÉO, J., Correa-

Ramirez, M. M., Cristescu, M. E., Dobigny, G., Egbosimba, E. E., Etchberger, L.

K., Fan, B. I. N., Fields, P. D., Forcioli, D., Furla, P., Garcia de Leon, F. J., GarcÍA-

JimÉNez, R., Gauthier, P., Gergs, R., GonzÁLez, C., Granjon, L., GutiÉRrez-

RodrÍGuez, C., Havill, N. P., Helsen, P., Hether, T. D., Hoffman, E. A., Hu, X.,

Ingvarsson, P. K., Ishizaki, S., Ji, H., Ji, X. S., Jimenez, M. L., Kapil, R., Karban, R.,

Keller, S. R., Kubota, S., Li, S., Li, W., Lim, D. D., Lin, H., Liu, X., Luo, Y.,

Machordom, A., Martin, A. P., Matthysen, E., Mazzella, M. N., McGeoch, M. A.,

Meng, Z., Nishizawa, M., O'Brien, P., Ohara, M., Ornelas, J. F., Ortu, M. F.,

Pedersen, A. B., Preston, L., Ren, Q. I. N., Rothhaupt, K.-O., Sackett, L. C., Sang,

Q., Sawyer, G. M., Shiojiri, K., Taylor, D. R., Van Dongen, S., Van Vuuren, B. J.,

Vandewoestijne, S., Wang, H., Wang, J. T., Wang, L. E., Xu, X.-L., Yang, G., Yang,

Y., Zeng, Y. Q., Zhang, Q.-W., Zhang, Y., Zhao, Y. & Zhou, Y. A. N. 2010.

Permanent Genetic Resources added to Molecular Ecology Resources

Database 1 August 2009–30 September 2009. *Molecular Ecology Resources*

10: 232-236.

Petit, R. J., Aguinagalde, I., de Beaulieu, J.-L., Bittkau, C., Brewer, S., Cheddadi, R.,

Ennos, R., Fineschi, S., Grivet, D., Lascoux, M., Mohanty, A., Müller-Starck, G.,

Demesure-Musch, B., Palmé, A., Martín, J. P., Rendell, S. & Vendramin, G. G.

2003. Glacial Refugia: Hotspots But Not Melting Pots of Genetic Diversity.

Science **300**: 1563-1565.

Pritchard, J., Stephens, M. & Donnelly, P. 2000. Inference of population structure

using multilocus genotype data. *Genetics* **155**: 945-959.

Slate, J. & Pemberton, J. M. 2007. Admixture and patterns of linkage disequilibrium

in a free-living vertebrate population. *Journal of Evolutionary Biology* **20**:

1415-1427.

Slatkin, M. 1977. Gene flow and genetic drift in a species subject to frequent local

extinctions. *Theoretical Population Biology* **12**: 253-262.

Slatkin, M. 1994. Linkage disequilibrium in growing and stable populations. *Genetics*

137: 331-336.

Sloan, D. B. 2015. Using plants to elucidate the mechanisms of cytonuclear co-

evolution. *New Phytologist* **205**: 1040-1046.

Slotta, T. A. B., Brady, L. & Chao, S. 2008. High throughput tissue preparation for

large-scale genotyping experiments. *Molecular Ecology Resources* **8**: 83-87.

Taylor, D. R. & Keller, S. R. 2007. Historical range expansion determines the

phylogenetic diversity introduced during contemporary species invasion.

Evolution **61**: 334-345.

Trier, C. N., Hermansen, J. S., Sætre, G.-P. & Bailey, R. I. 2014. Evidence for Mito-

Nuclear and Sex-Linked Reproductive Barriers between the Hybrid Italian

Sparrow and Its Parent Species. *PLoS Genet* **10**: e1004075.

Wade, M. J. & McCauley, D. E. 1988. Extinction and recolonization: their effects on

the genetic differentiation of local populations. *Evolution* **42**: 995-1005.

Whitlock, M. C. & McCauley, D. E. 1990. Some population genetic consequences of

colony formation and extinction: genetic correlations within founding

groups. *Evolution* **44**: 1717-1724.

Figure 1. European sampling localities. Grey dots indicate samples from the

Western European (WEU) deme and black dots indicate samples from Eastern

European (EEU) deme.

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Figure 2. Hypothetical consequences for CNLD *D* values when populations that

differ in their respective values of *D*, due to variable genetic compositions,

contribute individuals to form an admixed deme. Five types of mixtures/scenarios

are described in the text.

Figure 3. Locus by locus CNLD Values of D (3A) or D' (3B) obtained from single

European sources (EEU and WEU) and from pooling samples from those sources

(EU). The shift predicted from eq. 1 can be visualized by noting that the D value seen

in EU are not always the average of the locus-specific D found in EEU and WEU.

A

B

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Table 1. Global estimates of genetic diversity, population substructure, and linkage disequilibrium across Europe. M=repeat size of the

marker; N=number of observed alleles; P=allele frequency of the most frequent allele. H_O and H_E are observed and expected

heterozygosity, respectively, of each sampled microsatellite locus; F_{IT} is the correlation of alleles within individuals relative to the

continental sample as a whole, and can be interpreted as an approximate measure of deviation from panmixia. Statistical deviations

from panmixia are indicated by the symbols *P<0.05, **P<0.01 and ***P<0.001) and LD_{EEU} is an indication of the number of other

nuclear loci for which each focal nuclear locus exhibits a non-random association in the EEU deme; LD_{WEU} is an indication of the number

of other nuclear loci for which each focal nuclear locus exhibits a non-random association in the WEU deme; and LD_{EU} is an indication of

the number of other nuclear loci for which each focal nuclear locus exhibits a non-random association across the whole European

sample, based upon an approximation of Fisher's exact test of random association and a P-value £ 0.05.

Locus	M	N	H_O	H_E	F_{IT}	LD_{EEU}	LD_{WEU}	LD_{EU}
SL_eSSR04	3	12	0.678	0.806	0.172	1	7	8
SL_eSSR06	3	13	0.516	0.818	0.37	5	8	9
SL_eSSR09	3	11	0.324	0.502	0.408	6	9	11
SL_eSSR12	2	15	0.656	0.876	0.294	6	6	11
SL_eSSR16	3	8	0.415	0.528	0.485	2	0	10
SL_eSSR20	3	5	0.051	0.258	0.819	7	9	11
SL_eSSR29	3	21	0.761	0.902	0.205	3	8	11
SL_eSSR30	3	18	0.362	0.797	0.544	3	6	10
slat_18	3	11	0.307	0.809	0.651	4	6	11
slat_32	4	6	0.42	0.789	0.487	4	3	7
slat_33	3	2	0.061	0.134	0.94	2	0	12
slat_48	3	2	0.11	0.25	0.562	3	4	5
slat_72	3	13	0.436	0.818	0.504	5	4	9
slat_85	3	14	0.2	0.536	0.77	3	10	13
Overall		*10.79*	*0.38*	*0.63*	*0.498*	*3.86*	*5.71*	*9.86*

467 **Table 2.** Values of Cyto-nuclear linkage disequilibrium (D or D') between the mitochondrial ATP1 SNP and the most common allele (at frequency p) of each of 14 nuclear microsatellite markers calculated from sample taken at two localities in Europe (EEU & WEU) and a mixture of

468 those European samples (EU). Also presented are D and D″ calculated from North American (NA) samples reported by Fields *et al.* (2014).

Locus	EEU			WEU			EU			NA		
	P	D	D′	P	D	D′	P	D	D′	P	D	D′
SL_eSSR04	0.329	0	0	0.237	-0.0288	-0.25	0.276	-0.0099	-0.0682	0.498	0.0255	0.1737*
SL_eSSR06	0.268	-0.0069	-0.061	0.395	0.0481	0.25	0.342	0.0137	0.1146***	0.262	-0.0257	-0.1751**
SL_eSSR09	0.744	0.0234	0.2361	0.623	-0.1202	-0.641	0.672	-0.0579	-0.4504**	0.743	0.0312	0.2161**
SL_eSSR12	0.232	-0.0386	-0.3889	0.061	0.0337	1	0.133	0.0223	0.387***	0.315	0.0448	0.3243***
SL_eSSR16	0.207	-0.022	-0.2254	0.816	0.0048	0.0526	0.561	-0.0513	-0.3266***	0.544	0.0539	0.3258***
SL_eSSR20	0.902	-0.011	-1	0.798	-0.0481	-0.5	0.842	-0.0237	-0.4466***	0.953	-0.0078	-0.3721
SL_eSSR29	0.25	-0.0413	-0.4762	0.026	-0.0144	-1	0.119	-0.0064	-0.0836***	0.203	0.0254	0.2841**
SL_eSSR30	0.427	0.0317	0.4107	0.298	-0.024	-0.1515	0.352	0.0057	0.042	0.389	0.0117	0.0691
slat_18	0.038	-0.022	-0.5926	0.429	-0.0481	-0.25	0.268	-0.0636	-0.4032***	0.402	0.0125	0.071
slat_32	0.351	0.0069	0.1316	0.278	0.0288	0.2143	0.308	0.0217	0.2088	0.316	-0.0561	-0.3161***
slat_33	0.11	-0.0262	-0.3016	0.965	0	0	0.607	-0.0748	-0.5362***	0.607	0.1042	0.4868***
slat_48	0.808	0.0096	0.0833	0.9	0.0337	0.6364	0.862	0.0174	0.1981	0.582	0.0475	0.2415***
slat_72	0.325	0.0193	0.3889	0.219	0.0673	0.5833	0.263	0.0518	0.5573*	0.308	0.0026	0.0193
slat_85	0.878	0.0358	0.3714	0.145	-0.0577	-0.75	0.458	0.0322	0.1995***	0.516	-0.0256	-0.1492*