Title: Mental health explains individual deviations from normative range in cognition-associated brain states

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Abstract

Links between cognitive deficits and psychiatric disorders have been studied predominantly at the group level, leaving unique individual characteristics largely unknown. Here, we applied normative modeling to UK Biobank data (N=18,634) and estimated the interplay of large-scale brain networks over time (i.e., dynamic brain state) at the individual level. Abnormality in such brain states was linked to individual variation in mental health. Specifically, brain state measures including fractional occupancy were estimated as a function of general cognitive ability and abnormality scores per participant were quantified to represent the degree of deviations relative to the estimated population norm. We found significant associations between the abnormality scores of several brain states and individual's overall mental health. Our findings suggest potential impact of mental health on dynamic brain states subserving cognitive functions and shed light on the relevant brain mechanisms underlying cognitive deficits in mental illness.

Introduction

Cognitive functions are a set of mental abilities such as problem solving and decision making that are required to perform daily tasks. The ability to execute such functions has been shown to vary across the lifespan [1], and from one condition to the other [2]. In fact, impairment in such ability (i.e., cognitive deficit) is recognized as a common symptom present in the majority of psychiatric disorders, including schizophrenia [3], depression [4] and bipolar disorder [5]. The observed behavioral abnormalities in those disorders are often attributed to various cognitive functions including learning, attention, emotion regulation and cognitive control [6]. Aberrant functional connectivity patterns, in addition to other neurobiological measures, have been linked to cognitive deficits in psychiatric disease groups [7]. A recent meta-analysis study revealed that altered functional connectivity (FC) in large-scale neurocognitive networks were common across psychiatric diagnoses, suggesting a shared mechanism of brain network interactions underlying abnormalities in generalized cognitive ability [8]. These findings point to the FC of large-scale networks as the mechanistic basis of cognitive ability and suggest that the relationship between FC and cognition may vary as a function of mental health.

Interestingly, recent studies on FC suggest that in contrast to traditional static functional connectivity, patterns of large-scale brain networks over time that capture time-varying functional connectivity might be of greater relevance to studying individual differences in cognitive functions [9–11]. Unlike static FC, which averages neural fluctuations over time, dynamic connectivity estimates temporal fluctuations in functional connectivity and provides additional information about temporal reconfiguration of functional resources [12]. Growing evidence suggests that dynamic changes in the resting-state FC are involved in various cognitive processes, including mind-wandering, fluctuations in arousal, vigilance, and perceptual performance [13,14]. Recent studies linking temporal dynamics of large-sale brain networks to cognitive and behavioral traits further revealed that such network temporal properties often had better predictions on cognitive traits than static FC [9], explain unique aspects of cognition [11], and that dynamic connectivity can outperform static FC in terms of capturing trial-wise behavioral variability in cognitive tests [10]. These findings highlight the interactions of brain networks over time as a potentially more sensitive marker than the static FC to capture inter-individual differences in cognition.

Brain network dynamics have also been linked to various psychiatric and neurological disorders. However, almost all reported studies so far have focused on group-level comparisons between patients and healthy controls [15]. Similarly, research into the role of mental illness in

cognitive functions are also predominantly based on case versus control comparisons or clinical subgroup contrasts [16]. These studies therefore only provide group-level inference that essentially overlook unique individual characteristics. To obtain more representative measures of each individual and to eventually identify meaningful biomarkers, we therefore need more personalized approaches [16–18].

In this study, we adopt individualized methods to investigate the interactions of large-scale brain networks over time (i.e., brain network state with temporal information) that are associated with cognitive ability, and their relationships to mental health. Specifically, we leverage normative modeling to estimate individual-specific brain state measures from cognitive ability. Analogous to normative growth charts used in pediatric medicine, normative models locate each individual on the normative range in terms of the estimated variables (e.g., brain state measures) and provide statistical inference at the individual subject level [19]. Deviations from the estimated normative range per individual (i.e., abnormality score) can be then quantified and linked to relevant variables (e.g., individual's mental health). Normative modeling has been increasingly applied to study associations between brain functions and behavior, particularly for clinically relevant conditions where significant heterogeneity may exist that cannot be captured by categorical partitioning of the cohort into traditional patient-control groups [20–23]. Here, we apply normative modeling to a population cohort to investigate whether individual-level abnormality in brain state measures as estimated from cognitive ability can be attributed to overall mental health.

Methods

Participants

The UK Biobank is a population cohort dataset with extensive behavioral and demographic assessments, lifestyle, and cognitive measures, as well as high quality imaging data. It recruited adults from the general population across Great Britain between 2006 and 2010 (https://www.ukbiobank.ac.uk). This openly accessible population dataset also aims to obtain multimodal brain imaging for 100,000 individuals before 2023 [43]. We used data from "instance 2" when the initial neuroimaging data were acquired. Participants were included in the analysis if they had resting-state fMRI data and had sufficient data of self-reported mental health questionnaires and of cognitive tasks to derive scores for overall mental health and general cognitive ability (see below). Consequently, data from N=18,634 were analyzed (M_{age}=63.65; N_{female}=9,913; see sample demographics in Supplemental Materials). All participants provided informed consent. UK Biobank has ethical approval from the North West Multi-Centre Research Ethics Committee (MREC). Data access was obtained under UK Biobank application ID 47267.

General Cognitive Ability

Several well-established cognitive tasks were performed when UK Biobank participants visited the assessment center to measure their cognitive functions including IQ, verbal declarative memory, executive function, and non-verbal reasoning. All tests were administered on a touch screen computer on the day of scanning. Following the practice from previous studies, PCA was used to extract the major component across multiple cognitive tasks that are representative of general cognitive ability [44,45]. The summary outcome measures (e.g., mean reaction time, number of correct puzzles) of each task were entered into the PCA with exclusion of items having more than 55% missing values. In case of categorical items, the proportion of each category or level was calculated. This resulted in a final list of 16 items that were entered into the PCA (see full list of items in *SFigure 1*). Prior to PCA, imputation was performed to address missing variables by a regularized iterative PCA algorithm that first imputes the missing values with the initial mean values of variables, followed by a PCA on the completed dataset and then imputes the missing values with the regularized reconstruction formulae order. These steps are iterated until convergence [46]. Just like the g-factor that indicates general intelligence, the first PCA component was derived as a measure for general cognitive ability and individual scores were used to indicate general cognition levels per participant.

Overall Mental Health

On the same day when neuroimaging and cognitive task data were acquired, participants also completed a set of self-report mental health questions on a touch screen. The questionnaire items (40 in total) touched upon lifetime mental health with questions about affective disorders, neuroticism, subjective well-being indicated by happiness, and satisfaction levels with health, family relationships, friends and financial situation (*SFigure 1*). Individual variables from the mental health questionnaire were excluded with more than 55% missing values and categorical variables were transformed into percentages, resulting in a final list of 33 items that were entered into a PCA after the same imputation procedure as described above (see full item list in *SFigure 1*). For the purpose of inferring an overall state of mental health, we used the individual scores for the first PCA component to approximate the *p*-factor that has been proposed to indicate generalized dimensional symptom levels across psychiatric disorders [47,48].

Imaging data acquisition and processing

Resting-state fMRI was obtained using a multiband sequence with an acceleration factor of 8 (TR=0.735; voxel size=2.4x2.4x2.4mm^3). Preprocessing for the resting-state fMRI data included motion correction, grand-mean intensity normalization, high-pass temporal filtering, unwarping and denoising [49]. Full details for imaging acquisition and preprocessing steps can be found in UK Biobank Brain Imaging Documentation (https://biobank.ctsu.ox.ac.uk/crystal/crystal/docs/brain_mri.pdf).

Researchers involved in UK Biobank have centrally developed and shared the imaging-derived phenotypes via the UK Biobank showcase [50], one of which is subject-specific time-series of 100 resting-state nodes. These time-courses were derived from a dual regression analysis using a group-level independent component analyses (ICA) map as the input. Out of 100 dimensions, 55 were classified as non-noise functional network works, resulting in a matrix size of 490 x 55 for each participant's time-series, where 490 is the number of acquired volumes. These data were used in the current study to estimate time-varying functional connectivity.

Dynamic Brain States

We derived brain state measures by fitting a Hidden Markov Model (HMM) to the aforementioned fMRI time-series data. In brief, HMM assumes that brain activity can be characterized as a number of discrete brain states. Under this framework, the brain activity at each time point can be modeled as a mixture of Gaussian distributions with each distribution

corresponding to a different state [10,51]. Importantly, each of the estimated brain states is described by a distinct timeseries and functional connectivity pattern [52]. HMM can be viewed as a probabilistic description of the data in which the hidden information (i.e., brain state) is captured from observable measures (i.e., fMRI time courses). An increasing number of studies have applied HMM to neuroimaging data, demonstrating it as a promising approach for understanding the temporal dynamics of brain functions [52–55].

Similar to previous work applying HMM to fMRI data, we inferred 12 brain states using the resting-state fMRI time-series data from all participants [52,56]. Measures including fractional occupancy (FO) that represents the average state probability across time for each brain state, and the switching rate among brain states that indicates the brain state stability of each participant were quantified at the individual level. These measures were considered as indications of subject-specific brain states (i.e., total of 13 measures from the estimated model with 12 brain states).

Normative modeling of brain state measures as a function of cognition
To obtain statistical inferences at the individual level with respect to the normative range of brain state measures, we used a Gaussian Process Regression (GPR) to predict HMM-derived measures from the general cognitive ability scores. GPR is a non-parametric probabilistic Bayesian method that can provide reliable estimates of uncertainty for the prediction and has been widely applied in machine learning field [57]. Using GPR, we estimated each individual measure of brain states separately and quantified subject-level deviations from the estimated normative range of per measure (i.e., abnormality score) for each individual participant. Abnormality scores were calculated as the difference between estimated brain state measure value from the normative model and the actual value derived from HMM, which was further divided by the estimated standard deviation from the normative model (i.e., Z scores) [20].

To infer the normative range with the current sample size (N>18K), we leveraged local approximate GP modeling. In contrast to full GPR that requires to obtain a full set of predictive equations, approximate GPR models involves approximating the predictive equations at a particular location (i.e., X) via a subset of the data. As the remote elements from the subset to the location X have vanishingly small influence on prediction, ignoring these elements in order to work with much smaller matrices can be computationally efficient with little impact on prediction accuracy [58]. A 10-fold cross validation was employed to ensure generalizability of the estimates, where data were partitioned into training and testing sets across participants per fold (i.e., one out of the total ten subsamples was held for testing per fold). One-sample t-tests

were used to examine which of the brain state measures had abnormality scores that on average were significantly different from zero, using the adjusted P-value (i.e., p<0.05/13=0.00385). Only the brain state measures for which the group mean abnormality was significantly different from zero were included in subsequent analyses, linking brain state abnormality to mental health.

Linking brain state abnormality to mental health

In order to examine whether inter-individual differences in the degree of abnormality in brain state measures can be attributed to overall mental health, we performed linear regression models. Abnormality scores were included as the independent variable and the overall mental health scores the dependent variable. We further considered age, sex, education and income levels, as well as head motion during resting-state imaging acquisition (i.e., the mean framewise displacement) as nuisance covariates and included those in the analyses that showed significant correlations with overall mental health scores. Multiple comparisons were accounted for the Beta values of abnormality scores (i.e., predictor) in each model using Bonferroni correction (i.e., adjusted p<0.05/7=0.007). To examine whether any abnormality scores of brain state measures exhibited unique association with mental health over and above each other, we further conducted a multiple regression model including all brain state abnormality scores that passed the one-sample t-test.

Implementation of statistical analyses

All statistical analyses were conducted using R version 3.6.1 (R Core Team, 2019). Imputation prior to PCA was conducted for missing values using the R function imputePCA from missMDA package [46], and function prcomp from basic stats pckage [59] was used for the PCA analyses. Local approximate GPR was performed using the laGP package [60].

Results

General cognitive ability and mental health state

To summarize general cognitive ability and overall mental health, we performed principal component analyses across a set of cognitive tests and mental health self-report questions. For cognitive tasks, the first principal component derived from all items explained 33.75% of variance, and the first principal component derived from mental health questionnaire items explained 24.08% of variance. Individual scores for these first components per participant were used to indicate general cognitive ability and overall mental health, respectively. This approach has been widely applied to approximate general intelligence (i.e., g-factor) and generalized dimensional symptom levels across psychiatric disorders (i.e., p-factor). The loadings (i.e., eigenvalues) of each individual item on the principal components for cognitive ability and mental health are shown in *SFigure 1*.

Brain state measures

Hidden Markov Modeling (HMM) was performed on individual timeseries data from 55 components as defined with independent component analysis to estimate the interactive patterns of brain networks over time (i.e., brain states). Twelve brain states were inferred from the model, each of which is represented by a multivariate Gaussian distribution, as described by the mean and covariance. Different brain networks, including the default mode and sensorimotor networks were engaged in these inferred brain states (*Figure 1*; also see full brain activation maps in *SFigure 2*). Per participant, fractional occupancy (FO) of each inferred brain state that represents the state probability over time and the switching rate (SR) across these brain states that indicates state stability were quantified (*Figure 2*).

Figure 1. Dynamic Brain States Inferred from Hidden Markov Model. A number of brain states common to all participants were estimated with participant-specific time-courses per state that indicate when each brain state is active. These states are characterized by their mean activation and functional matrix. Here, mean activation maps of four example brain states are thresholded to illustrate the top 10% voxels engaged in each brain state, along with their functional connectivity matrices (i.e., lower triangular portion). Unthreholded mean activation maps for all 12 brain states can be found in Supplemental Material.

Figure 2. HMM derived brain state measures. Fractional occupancy (FO) per inferred brain state (i.e., indicated by numbers after FO as in FO1) and switching rate (SR) across all brain states were derived for each individual participant to indicate their brain network interactions over time. Data from six randomly selected individual participants were overlaid on the sample mean to illustrate individual variation in these measures.

Normative range of and abnormality in brain states

Normative modeling with Gaussian process regression was performed to predict each of thirteen brain state measures separately (i.e., FO for each of the 12 brain states and SR across all brain states) from general cognitive ability scores. This approach enabled quantification of individual-specific abnormality scores per brain state measure (i.e., deviations from the estimated normative range). Tests on the group mean of such abnormality scores revealed that the average level of abnormality in FO for seven brain states was significantly different from zero (all p's<0.0006). Importantly, large individual variation was observed in these abnormality scores (e.g., some participants deviated much more than the others or in different directions), demonstrating unique individual patterns that were not represented by the group mean (*Figure 3*).

Figure 3. Abnormality in brain states. As shown in the boxplot, averaged deviations (i.e., red dots) in fractional occupancy (FO) for seven brain states were significantly different from zero at the group level with substantial individual variation in both magnitude and direction. Boxes here indicate the inter-quantile range (i.e., middle 50% of the data) and black horizontal lines inside each box represent the median values. Outliers for each FO measure are shown as gray dots beyond the either end of the boxes.

Individual abnormality in brain states and mental health

To capture the individual level relationship between the cognition-derived brain state measures and mental health, we linked the abnormality scores of each of the seven FOs that were significantly different from zero to the overall mental health scores. Four of these FO abnormality scores showed significant associations with overall mental health after Bonferroni correction for multiple comparisons ($t_{FO1}=-5.66$; $t_{FO4}=-3.26$; $t_{FO11}=-3.63$; $t_{FO12}=2.93$; all p's<0.0035; *Figure 4*). We further included all four FO abnormality scores in one regression model and the results indicated unique predictive effects of three FO abnormality scores on overall mental health ($t_{FO1}=-5.01$; $t_{FO4}=-4.31$; $t_{FO11}=-2.02$; all p's<0.009).

Figure 4. Overall mental health and abnormality in brain states. Individual deviations from the normative range (i.e., abnormality) of fractional occupancy for brain state 1 (FO1), 4 (FO4) 11 (FO11) and 12 (FO12) were significantly associated with individual's overall mental health. Brain images are mean activation maps of each brain state, thresholded to illustrate the top 10% voxels.

As a comparison, we also conducted a more traditional moderator effect analysis to assess whether mental health moderates the relationship between cognitive ability and HMM-derived brain state measures. Importantly, these tests focused on the commonalities across all participants, assuming that interactions between mental health and cognitive ability will impact brain states in the same way for each individual. However, no significant moderator effects were observed (all p's>0.15).

Discussion

In this study, we adopted a personalized approach to quantify individual brain state measures that indicate brain network interactions over time, as a function of cognitive ability. Our aim was to test whether individual's deviation in brain states from the normative range can be explained by their mental health. Applying normative modeling to a population cohort, we identified several brain states with significant deviations across all participants from the normative range. More importantly, we demonstrated that the degree of abnormality in the fractional occupancy of four brain states could be attributed to the inter-individual difference in overall mental health.

While abnormality in cognitive functions and alterations in the relevant neural correlates have been implicated in a variety of neuropsychiatric conditions, the vast majority of investigations so far has predominantly adopted a case-control approach or divided samples into phenotypical subgroups (e.g., based on symptom scores), which essentially overlooks unique individual characteristics. In the current study, our approach revealed associations between mental health and brain network interaction patterns that were specific to each individual participant's cognitive ability, whereas more traditional moderator effect tests that focus on commonalities across participants failed to capture such relationships. These findings echo the goal to adopt more personalized approaches in precision psychiatry [24] and are well in line with the National Institute of Mental Health's Research Domain Criteria initiative (RDoC) to take into account individual's variability rather than diagnostic groups [25].

In our findings, significant deviations across all participants were observed in several brain states that involve the default mode and sensorimotor networks (*SFigure 1*). Interestingly, whereas the DMN has been consistently associated to various cognitive processes [26–28] and altered connectivity of this network has also been suggested to underly multiple mental illness conditions with substantial cognitive deficits [29], sensorimotor networks are predominantly involved in motor related processes. However, increasing evidence shows that sensorimotor networks can also contribute to disorder symptoms related to cognitive functions via its connections with other large-scale brain networks such as the DMN [30–32].

Our findings further revealed associations between the degree of subject-level abnormality in brain state measures (i.e., FO of several brain states) and inter-individual differences in overall mental health. Interestingly, these associations were mostly negative. Since the overall mental health score was derived from the first principal component over mental health questions and it loaded negatively on most of those questions (*SFigure 1B*), the observed negative associations between abnormality in FOs and overall mental health imply that the

further away individual's FO deviate from the population norm, the higher the mental health scores they have in general (i.e., more symptoms). These results imply that one's mental health may impact brain network interactions over time that provide support to cognitive functions and thus point to potential brain mechanisms likely underlying cognitive deficits across a spectrum of major neuropsychiatric disorders. As mental health questions included in UKB dataset primarily assess mood disorders and neuroticism [33], our findings here can be particularly relevant for understanding the neural mechanisms for these disorders.

Whereas recent neuroimaging studies have mainly applied normative modeling to clinical cohorts and studied extreme deviations in estimated brain traits in relation to symptom levels [20,34,35], our study extended the application of normative modeling to a large population cohort and demonstrated a continuum in the relationship between individual abnormality in brain traits and their overall mental health state. Our findings align well with prior evidence that cognition can be compromised as a function of mental health and highlights potential mechanisms in the brain.

Our findings should be interpreted in light of several limitations. First, the sample of the current study has a specific lifespan (i.e., between 45-81) and is heavily biased towards a specific ethnicity (i.e., White), which may limit the generalization of our results to other populations with different characteristics. Nevertheless, with careful controls for the potential confounding factors including age, sex, education and income levels, as well as head motion during the fMRI data acquisition, our results provide evidence that dynamic interplays of large-scale brain networks may contribute to cognitive deficits due to mental health. Second, HMM was utilized in the current study to infer dynamic brain states that reoccur over time. It is important to note that HMM-derived measures appear to contain information of both time varying and static functional connectivity, and that our implementation of HMM (i.e., model estimation based on the mean and covariance) may capture additional network characteristics (i.e., node amplitude) [11]. Although it is beyond the scope of the current study to tease apart the unique contributions of these network properties, previous studies did suggest that temporal dynamics of large brain networks are potentially more sensitive to explain inter-individual differences in cognitive traits [9–11]. Lastly, in our study head motion showed correlations with almost all observed abnormalities in brain state measures, as well as with overall mental health. This may appear concerning as head motion is typically considered and modeled as artefacts particularly in the resting-state fMRI studies. However, increasing studies have observed its association with typical functional connectivity networks in the brain [36,37], demonstrated different patterns of head motion for patients versus healthy control groups or in different age

groups [38–40] and suggested that head motion might have genetic roots [41,42]. These findings imply that head motion may contain extra information beyond mere noise that is relevant for our interpretations of imaging studies and that estimating the tendency of head motion in relation to the variables of interest may offer more insight into individual level brain signatures. Nevertheless, the reported findings here are carefully controlled for head motion effects in two-folds: ICA-FIX was performed at the individual subject level during preprocessing, and the averaged framewise displacement was further modeled as a between-subject covariate in all statistical analyses. Furthermore, we assessed the robustness of the reported associations between abnormality in brain state measures and individual's overall mental health by estimating the potential impact of multicollinearity among predictors (i.e., mental health and motion). Our results indicate very little influence of head motion (see Supplemental Results). Hence, the relationship between brain state abnormality and overall mental health was significant while accounting for head motion effects.

In conclusion, we show in the current study that individualized brain state measures for cognitive ability can be quantified and that the degree of abnormality in such brain states can be attributed to individual differences in overall mental health. These results shed light on the potential brain mechanisms underlying cognitive deficits in mental illness conditions, which may be investigated in patient cohorts in the future.

Reference

1. Wadsworth SJ, Corley RP, Munoz E, et al. CATSLife: A Study of Lifespan Behavioral Development and Cognitive Functioning. *Twin Res Hum Genet*. 2019. doi:10.1017/thg.2019.49
2. Nakagome K. Cognitive impairment in psychiatric disorders. *Psychiatry Clin Neurosci*. 2017. doi:10.1111/pcn.12517
3. Mosiołek A, Gierus J, Koweszko T, Szulc A. Cognitive impairment in schizophrenia across age groups: A case-control study. *BMC Psychiatry*. 2016. doi:10.1186/s12888-016-0749-1
4. McIntyre RS, Cha DS, Soczynska JK, et al. Cognitive deficits and functional outcomes in major depressive disorder: Determinants, substrates, and treatment interventions. *Depress Anxiety*. 2013. doi:10.1002/da.22063
5. Bora E, Özerdem A. Meta-Analysis of longitudinal studies of cognition in bipolar disorder: Comparison with healthy controls and schizophrenia. *Psychol Med*. 2017. doi:10.1017/S0033291717001490
6. Millan MJ, Agid Y, Brüne M, et al. Cognitive dysfunction in psychiatric disorders: Characteristics, causes and the quest for improved therapy. *Nat Rev Drug Discov*. 2012. doi:10.1038/nrd3628
7. Etkin A, Gyurak A, O'Hara R. A neurobiological approach to the cognitive deficits of psychiatric disorders. *Dialogues Clin Neurosci*. 2013. doi:10.31887/dcns.2013.15.4/aetkin
8. Sha Z, Wager TD, Mechelli A, He Y. Common Dysfunction of Large-Scale Neurocognitive Networks Across Psychiatric Disorders. *Biol Psychiatry*. 2019. doi:10.1016/j.biopsych.2018.11.011
9. Liégeois R, Li J, Kong R, et al. Resting brain dynamics at different timescales capture distinct aspects of human behavior. *Nat Commun*. 2019;10(1). doi:10.1038/s41467-019-10317-7
10. Eichenbaum A, Pappas I, Lurie D, Cohen J, D'Esposito M. Differential Contributions of Static and Time-Varying Functional Connectivity to Human Behavior. *Netw Neurosci*. 2020. doi:10.1101/2020.07.14.203273
11. Vidaurre D, Llera A, Smith SM, Woolrich MW. Behavioural relevance of spontaneous, transient brain network interactions in fMRI. *Neuroimage*. 2021. doi:10.1016/j.neuroimage.2020.117713
12. Lurie DJ, Kessler D, Bassett DS, et al. Questions and controversies in the study of time-varying functional connectivity in resting fMRI. *Netw Neurosci*. 2020;4(1):30-69. doi:10.1162/netn_a_00116
13. Cohen JR. The behavioral and cognitive relevance of time-varying, dynamic changes in functional connectivity. *Neuroimage*. 2018;180(April 2017):515-525. doi:10.1016/j.neuroimage.2017.09.036
14. Gonzalez-Castillo J, Hoy CW, Handwerker DA, et al. Tracking ongoing cognition in individuals using brief, whole-brain functional connectivity patterns. *Proc Natl Acad Sci*. 2015;112(28):8762-8767. doi:10.1073/pnas.1501242112
15. Valsasina P, De La Cruz MH, Filippi M, Rocca MA. Characterizing rapid fluctuations of resting state functional connectivity in demyelinating, neurodegenerative, and psychiatric conditions: From static to time-varying analysis. *Front Neurosci*. 2019. doi:10.3389/fnins.2019.00618
16. Marquand AF, Wolfers T, Mennes M, Buitelaar J, Beckmann CF. Beyond Lumping and Splitting: A Review of Computational Approaches for Stratifying Psychiatric Disorders. *Biol Psychiatry Cogn Neurosci Neuroimaging*. 2016. doi:10.1016/j.bpsc.2016.04.002
17. Alda M. Personalized psychiatry: Many questions, fewer answers. *J Psychiatry Neurosci*. 2013. doi:10.1503/jpn.130221
18. Fernandes BS, Williams LM, Steiner J, Leboyer M, Carvalho AF, Berk M. The new field of "precision psychiatry." *BMC Med*. 2017. doi:10.1186/s12916-017-0849-x
19. Marquand AF, Kia SM, Zabihi M, Wolfers T, Buitelaar JK, Beckmann CF. Conceptualizing mental disorders as deviations from normative functioning. *Mol Psychiatry*. 2019;24(10):1415-1424. doi:10.1038/s41380-019-0441-1
20. Marquand AF, Rezek I, Buitelaar J, Beckmann CF. Understanding Heterogeneity in Clinical Cohorts Using Normative Models: Beyond Case-Control Studies. *Biol Psychiatry*. 2016;80(7):552-561. doi:10.1016/j.biopsych.2015.12.023
21. Ziegler G, Ridgway GR, Dahnke R, Gaser C. Individualized Gaussian process-based prediction and detection of local and global gray matter abnormalities in elderly subjects. *Neuroimage*. 2014. doi:10.1016/j.neuroimage.2014.04.018
22. Zabihi M, Floris DL, Kia SM, et al. Fractionating autism based on neuroanatomical normative modeling. *Transl Psychiatry*. 2020. doi:10.1038/s41398-020-01057-0
23. Lv Q, Yan M, Shen X, et al. Normative Analysis of Individual Brain Differences Based on a Population MRI-Based Atlas of Cynomolgus Macaques. *Cereb Cortex*. 2021. doi:10.1093/cercor/bhaa229
24. Manchia M, Pisanu C, Squassina A, Carpiniello B. Challenges and future prospects of precision medicine in psychiatry. *Pharmgenomics Pers Med*. 2020. doi:10.2147/PGPM.S198225
25. Insel T, Cuthbert B, Garvey M, et al. Research Domain Criteria (RDoC): Toward a new classification framework for research on mental disorders. *Am J Psychiatry*. 2010. doi:10.1176/appi.ajp.2010.09091379
26. Sormaz M, Murphy C, Wang HT, et al. Default mode network can support the level of detail in experience

during active task states. *Proc Natl Acad Sci U S A*. 2018. doi:10.1073/pnas.1721259115

27. Vatansever D, Menon DK, Stamatakis EA. Default mode contributions to automated information processing. *Proc Natl Acad Sci U S A*. 2017. doi:10.1073/pnas.1710521114

28. Andrews-Hanna JR, Smallwood J, Spreng RN. The default network and self-generated thought: component processes, dynamic control, and clinical relevance. *Ann N Y Acad Sci*. 2014;1316(1):29-52. doi:10.1111/nyas.12360

29. Doucet GE, Janiri D, Howard R, O'Brien M, Andrews-Hanna JR, Frangou S. Transdiagnostic and disease-specific abnormalities in the default-mode network hubs in psychiatric disorders: A meta-analysis of resting-state functional imaging studies. *Eur Psychiatry*. 2020. doi:10.1192/j.eurpsy.2020.57

30. Shumskaya E, van Gerven MAJ, Norris DG, Vos PE, Kessels RPC. Abnormal connectivity in the sensorimotor network predicts attention deficits in traumatic brain injury. *Exp Brain Res*. 2017. doi:10.1007/s00221-016-4841-z

31. Martino M, Magioncalda P, Huang Z, et al. Contrasting variability patterns in the default mode and sensorimotor networks balance in bipolar depression and mania. *Proc Natl Acad Sci U S A*. 2016. doi:10.1073/pnas.1517558113

32. Doucet GE, Bassett DS, Yao N, Glahn DC, Frangou S. The role of intrinsic brain functional connectivity in vulnerability and resilience to bipolar disorder. *Am J Psychiatry*. 2017. doi:10.1176/appi.ajp.2017.17010095

33. Davis KAS, Coleman JRI, Adams M, et al. Mental health in UK Biobank – development, implementation and results from an online questionnaire completed by 157 366 participants: a reanalysis. *BJPsych Open*. 2020;6(2):e18. doi:10.1192/bjo.2019.100

34. Zabihi M, Oldehinkel M, Wolfers T, et al. Dissecting the Heterogeneous Cortical Anatomy of Autism Spectrum Disorder Using Normative Models. *Biol Psychiatry Cogn Neurosci Neuroimaging*. 2019. doi:10.1016/j.bpsc.2018.11.013

35. Lv J, Di Biase M, Cash RFH, et al. Individual deviations from normative models of brain structure in a large cross-sectional schizophrenia cohort. *Mol Psychiatry*. 2020. doi:10.1038/s41380-020-00882-5

36. Zeng LL, Wang D, Fox MD, et al. Neurobiological basis of head motion in brain imaging. *Proc Natl Acad Sci U S A*. 2014. doi:10.1073/pnas.1317424111

37. Bright MG, Murphy K. Is fMRI "noise" really noise? Resting state nuisance regressors remove variance with network structure. *Neuroimage*. 2015. doi:10.1016/j.neuroimage.2015.03.070

38. Haller S, Monsch AU, Richiardi J, Barkhof F, Kressig RW, Radue EW. Head Motion Parameters in fMRI Differ Between Patients with Mild Cognitive Impairment and Alzheimer Disease Versus Elderly Control Subjects. *Brain Topogr*. 2014. doi:10.1007/s10548-014-0358-6

39. Kong XZ, Zhen Z, Li X, et al. Individual differences in impulsivity predict head motion during magnetic resonance imaging. *PLoS One*. 2014. doi:10.1371/journal.pone.0104989

40. Pardoe HR, Kucharsky Hiess R, Kuzniecky R. Motion and morphometry in clinical and nonclinical populations. *Neuroimage*. 2016. doi:10.1016/j.neuroimage.2016.05.005

41. Hodgson K, Poldrack RA, Curran JE, et al. Shared Genetic Factors Influence Head Motion During MRI and Body Mass Index. *Cereb Cortex*. 2017. doi:10.1093/cercor/bhw321

42. Engelhardt LE, Roe MA, Juranek J, et al. Children's head motion during fMRI tasks is heritable and stable over time. *Dev Cogn Neurosci*. 2017. doi:10.1016/j.dcn.2017.01.011

43. Littlejohns TJ, Holliday J, Gibson LM, et al. The UK Biobank imaging enhancement of 100,000 participants: rationale, data collection, management and future directions. *Nat Commun*. 2020;11(1):2624. doi:10.1038/s41467-020-15948-9

44. Davies G, Lam M, Harris SE, et al. Study of 300,486 individuals identifies 148 independent genetic loci influencing general cognitive function. *Nat Commun*. 2018. doi:10.1038/s41467-018-04362-x

45. Lyall DM, Cullen B, Allerhand M, et al. Cognitive test scores in UK biobank: Data reduction in 480,416 participants and longitudinal stability in 20,346 participants. *PLoS One*. 2016. doi:10.1371/journal.pone.0154222

46. Josse J, Husson F. missMDA: A package for handling missing values in multivariate data analysis. *J Stat Softw*. 2016. doi:10.18637/jss.v070.i01

47. Caspi A, Houts RM, Belsky DW, et al. The p factor: One general psychopathology factor in the structure of psychiatric disorders? *Clin Psychol Sci*. 2014. doi:10.1177/2167702613497473

48. Caspi A, Moffitt TE. All for one and one for all: Mental disorders in one dimension. *Am J Psychiatry*. 2018. doi:10.1176/appi.ajp.2018.17121383

49. Alfaro-Almagro F, Jenkinson M, Bangerter NK, et al. Image processing and Quality Control for the first 10,000 brain imaging datasets from UK Biobank. *Neuroimage*. 2018. doi:10.1016/j.neuroimage.2017.10.034

50. Miller KL, Alfaro-Almagro F, Bangerter NK, et al. Multimodal population brain imaging in the UK Biobank prospective epidemiological study. *Nat Neurosci*. 2016;(September). doi:10.1038/nn.4393

51.   Baker AP, Brookes MJ, Rezek IA, et al. Fast transient networks in spontaneous human brain activity. *Elife*. 2014. doi:10.7554/eLife.01867

52.   Vidaurre D, Smith SM, Woolrich MW. Brain network dynamics are hierarchically organized in time. *Proc Natl Acad Sci U S A*. 2017;114(48):12827-12832. doi:10.1073/pnas.1705120114

53.   Vidaurre D, Abeysuriya R, Becker R, et al. Discovering dynamic brain networks from big data in rest and task. *Neuroimage*. 2018;180(June 2017):646-656. doi:10.1016/j.neuroimage.2017.06.077

54.   Zhang G, Cai B, Zhang A, et al. Estimating Dynamic Functional Brain Connectivity with a Sparse Hidden Markov Model. *IEEE Trans Med Imaging*. 2020. doi:10.1109/TMI.2019.2929959

55.   Li X, Li Y, Cui J. Estimating interactions of functional brain connectivity by Hidden Markov models. In: *Lecture Notes in Computer Science (Including Subseries Lecture Notes in Artificial Intelligence and Lecture Notes in Bioinformatics)*. ; 2018. doi:10.1007/978-3-030-05090-0_34

56.   Kottaram A, Johnston LA, Cocchi L, et al. Brain network dynamics in schizophrenia: Reduced dynamism of the default mode network. *Hum Brain Mapp*. 2019. doi:10.1002/hbm.24519

57.   Rasmussen CE, Williams CKI. *Gaussian Processes for Machine Learning*.; 2018. doi:10.7551/mitpress/3206.001.0001

58.   Gramacy RB, Apley DW. Local Gaussian Process Approximation for Large Computer Experiments. *J Comput Graph Stat*. 2015. doi:10.1080/10618600.2014.914442

59.   R Core Team. R: A language and environment for statistical computing. *R Found Stat Comput*. 2019.

60.   Gramacy RB. LaGP: Large-scale spatial modeling via local approximate Gaussian processes in R. *J Stat Softw*. 2016;72. doi:10.18637/jss.v072.i01

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Author Contributions

WZ designed the research plan and analyzed the data with critical input from JB. WZ and JB wrote the manuscript. DV consulted on the methods and interpretation and reviewed the manuscript.

Competing Interest Statement

The authors declare no competing interests.

Code Availability

R scripts for all statistical analyses can be found on GitHub (https://github.com/PersonomicsLab/laGP_nm). Matlab scripts for running HMM are also available on GitHub (https://github.com/OHBA-analysis/HMM-MAR).

Data Availability

The UK Biobank data that support the findings of this study can be accessed by researchers on application (https://www.ukbiobank.ac.uk/register-apply/). Variables derived specifically for this study will be returned to the UK Biobank for future applicants to request.

Diversity Statement

Recent work in several fields of science has identified a bias in citation practices such that papers from women and other minority scholars are under-cited relative to the number of such papers in the field. With the help of cleanBib tool (Dworkin et al., 2020), here we sought to proactively consider choosing references that reflect the diversity of the field in thought, form of contribution, gender, race, ethnicity, and other factors. As estimated by this tool, our initial reference contained 7.02% woman(first author)/woman(last author), 14.7% man/woman, 19.3% woman/man, and 58.99% man/man, which now has been improved to 10.63% woman(first)/woman(last), 15.46% man/woman, 21.67% woman/man, and 52.24% man/man. Similarly, our initial reference contained 10.57% author of color (first)/author of color(last), 14.66% white author/author of color, 26.73% author of color/white author, and 48.04% white author/white author, which has been improved to 15.51% author of color (first)/author of color(last), 15.2% white author/author of color, 24.31% author of color/white author, and 44.98% white author/white author.

These estimations are limited due to reference database limitations and may not be indicative of every possible case. Codes for predictions and more detailed information can be found on GitHub (https://github.com/dalejn/cleanBib).