Double-observer approach with camera traps: Towards an unbiased density

estimation of unmarked animal populations

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Abstract

Camera traps are a powerful tool for wildlife surveys. However, camera traps may not

always detect animals passing in front. This constraint may create a substantial bias in

estimating critical parameters such as the density of unmarked populations. We

proposed the 'double-observer approach' with camera traps to counter the constraint,

which involves setting up a paired camera trap at a station and correcting imperfect

detection with a hierarchal capture-recapture model for stratified populations. We

performed simulations to evaluate this approach's reliability and determine how to

obtain desirable data for this approach. We then applied it to 12 mammals in Japan and

Cameroon. The results showed that the approach could correct imperfect detection as

long as paired camera traps detect animals nearly independently (Correlation coefficient

< 0.2). Camera traps should be installed to monitor a predefined small focal area from

different directions to satisfy this requirement. The field surveys showed that camera

trap might miss animals by 3 %–40%, suggesting that current density estimation models

relying on perfect detection may underestimate animal density by the same order of

magnitude. We hope that our approach will be incorporated into existing density

estimation models to improve their accuracy.

Introduction

Camera trapping is an innovative research technique and has numerous advantages over

classic approaches based on human observations. For example, this technique is

minimally invasive for animals and habitats and can efficiently collect quantitative

records of ground-dwelling mammals, even elusive and nocturnal species, with

relatively low labour costs[1-3]. Moreover, by combining it with well-established

statistical approaches[4,5], wildlife occupancy can be reliably estimated[6,7], abundance or

density determined[8-10], and species richness assessed[11-13]. Furthermore, commercially

available camera traps are becoming cheaper, and their battery duration has also been

improving[14], allowing for large-scale surveys in space and time[15]. Data processing (e.g.

species identification) may require a long time and heavy labour costs, but a novel

technology using machine learning may drastically reduce the cost[16-18]. Given that

camera trapping will be increasingly used to make critical decisions on wildlife

conservation management, researchers must refine this technique further to avoid

arriving at an erroneous conclusion.

One of the primary concerns regarding the application of camera traps may be

that they may not always detect animals passing within the camera detection zone (i.e.

imperfect detection)[19]. Studies have shown that camera sensitivity may be primarily a

function of animal body mass[19,20] and may also be affected by many factors, such as

other animal traits (e.g. typical behaviours), environmental conditions (e.g. temperature

and vegetation), and camera efficiency (e.g. trigger speed)[20-27]. This constraint is one of

the reasons why the number of animal passes captured by camera traps per unit time (i.e.

trapping rate) does not always correlate with animal density and cannot be a reliable

index of animal abundance[28]. A series of statistical models are available to account for

imperfect detections and reliably estimate animal occupancy and abundance[4].

Nonetheless, researchers cannot help assuming perfect or equal detectability in other

situations. In particular, researchers have proposed several analytical approaches to

estimate animal density without recognizing animal individuals[29-37]. However, most of

these models, except for Chandler and Royle[38] and Ramsey *et al.*[36], assume that camera

traps can detect animals passing in a specific area within the camera detection zone with

absolute certainty (i.e. perfect detection). Rowcliffe *et al.*[19] suggested that the effective

area of the camera detection zone (codeterminants of the camera sensitivity) can be

determined using a distance sampling approach by applying detection function models

to data on the position where animals were first detected. Although this approach is an

important step in quantifying camera sensitivity, it also requires perfect detection at a

given point within the camera detection zone, which is still questionable.

A possible yet untested approach to correcting imperfect detection would be to

apply the 'independent double observer approach.' This approach has been used in

point-count surveys based on direct observations by human observers[39-43]. In this

approach, two observers (camera traps in our context) record animals concurrently and

independently, and the detection probability is estimated from the match or mismatch of

observation records. Two analytical models, the capture-recapture and N-mixture

models can be applied to correct imperfect detections[41]. The former requires that the

observers confer with each other regarding each observation, while the latter may be

based solely on the counts by each observer. Nichols $et\ al.$[41] applied these two models to

bird surveys and reported that both models have potential, while the precision of

estimates was higher in the former approach. Given that timestamps within images

captured by camera traps would allow for efficient reconciling of each count, the

capture-recapture model may provide a reliable and efficient means to obtain unbiased

estimates of trapping rates.

A constraint in applying the capture-recapture model to the double-observer

approach may be the assumption that two observers (i.e. camera traps) detect animals

independently. If this assumption is violated, detection probability is overestimated,

underestimating the number of animal passes. Regrettably, in practice, the detection

probability by paired camera taps may depend on numerous unknown factors (e.g.

animal body mass, ambient temperature, etc.), resulting in correlated detections. One

viable method to deal with this issue would assume a beta distribution instead of a

Bernoulli distribution for the detection processes[44,45]. However, this method may not be

a panacea, and thus it is necessary to determine to what extent of the heterogeneity this

modelling approach allows for[46].

This study tested the potential of a hierarchical capture-recapture model using

the Bayesian framework to correct imperfect detections and estimate the number of

animal passes. Firstly, we introduced the hierarchical capture-recapture model for

stratified populations. Secondly, Monte Carlo simulations were performed to evaluate

the reliability of this approach, focusing on to what extent the model can accommodate

the correlated detections. Thirdly, we performed additional simulations to determine

desirable camera selections and placements for this approach. Fourthly, we applied the

models to the datasets obtained from two different habitats, Cameroon and Japan, and

quantified the detection probability of 12 mammals with varying body sizes.

Results

Testing the effectiveness of the hierarchical capture-recapture model

The results of the Monte Carlo simulations showed that the hierarchal capture-recapture

models could provide good estimates with a reasonable confidence interval coverage, as

long as a pair of camera traps detects a passing animal nearly independently (Table 1).

On the other hand, when the correlation coefficient was > 0.2, detection probability was

overestimated, underestimating the number of animal passes. This pattern did not differ

largely between the lower (0.8) and higher (0.4) detection probabilities (Table 1).

Determining suitable camera traps and their installations

The simulations mimicking the detection processes of moving animals showed that

camera models with a lower trigger speed produced a more correlated detection history

(Table 2, Fig. 1). The degree of the correlations also depended on the camera

placements. Even when using a camera model with high trigger speed, the detection

histories of paired camera traps were highly correlated when monitoring the entire field

of view ($r = 0.53$) or the small focal area from the same direction ($r = 0.27$). However,

the correlation significantly decreased when monitoring the small focal area from the

different directions ($r = 0.18$).

Field surveys

120 In total, 12 species were recorded more than 10 times (7 in Japan and 5 in Cameroon),

121 which were targets for our analyses. Although none of the species was perfectly

122 detected, the detection probability was estimated relatively high (> 0.8) except for

123 Japanese field mice (95% CI: 0.56-0.64) and tree pangolins (0.62-0.91). The analysis

124 results are summarised in Fig. 2 and Supplementary Table S1.

Discussion

The Monte-Carlo simulation showed that the hierarchical capture-recapture model

assuming correlated detections (i.e. using a beta-binomial distribution) provided

unbiased estimates of detection probability and the number of animal passes, as long as

the correlation coefficients were ≤ 0.2. However, as the correlation increased, the model

produced more biased estimates. Regrettably, data obtained from two camera traps may

not be sufficient to estimate the parameters of a beta distribution reliably. Indeed, when

we assumed installing ten camera traps at each station, the model could provide

unbiased estimates even with data having a correlation coefficient of 0.8

(results not shown). Nonetheless, this installation is not practical, and it is more realistic

to carefully design a survey to obtain as much independent data as possible.

The simulation mimicking the detection processes of moving animals

suggested that camera traps should monitor a small focal area from different directions

to obtain nearly independent data. A pair of camera traps facing the same direction have

the same hazard landscape, so they often fail to detect animals passing through the

periphery of the focal area. On the other hand, camera traps installed in different

directions can compensate for weak areas with a lower detection probability by one

another. In addition, it may be critical to use a camera model with high trigger speed to

avoid missing fast-moving animals. There may be other viable approaches to address

heterogeneity in detection probability. For example, one could measure the distance and

angle of animals from the camera trap and incorporate them into the model as covariates.

This approach is called MRDS (Mark-Recapture Distance Sampling), a well-developed

theoretical framework. However, this approach does not always function well[46] and also

requires an intensive field survey. The installations proposed here may be much simpler

and could reduce the labour costs.

The results of our field survey showed that it is critical to account for possible

imperfect detection in actual analyses. Although the estimated detection probability was

relatively high (>0.8) for most species, camera traps could not detect any species

completely. In particular, the detection probability was lower for field mice and tree

pangolins (Fig. 1), possibly reflecting the smaller body mass (field mice) or the scaly

hairs preventing heat radiation from the body interior (tree pangolins). Given that the

size and position of the focal area were specified to maximize the detection probability,

the sensitivity of the camera model might be less than one throughout the detection zone

for these species. This result warns against applying the current density estimation

models without accounting for imperfect detection. In the present study, trapping rates

were underestimated by 4 %–36%, which led directly to underestimating animal density

by the same order of magnitude. Therefore, the double-observer approaches proposed

here should be incorporated into the existing density estimation models relying on

perfect detections.

We admit that a shortcoming of this approach remains. In particular, it is

necessary to have more cameras available, which may be a constraint on

implementation. Nonetheless, researchers may be able to challenge these constraints in

various ways. For example, installing a pair of cameras at every camera station may not

be necessary, as long as the variance in detection probability among camera stations is

not too large. Instead, one may choose to install a pair of camera traps in locations with

high trapping rates. It may be effective that the study period may be divided into two

parts: one in which two cameras are installed to estimate the detection probability and

one in which one camera is installed to estimate the trapping rate. Furthermore, if our

approach were used at different locations, it would be possible to extrapolate the results

to new sites by incorporating environmental conditions and animal characteristics as

covariates. Given that many surveys estimating the density of marked populations use

paired camera traps at each camera station (to recognize individual animals reliably), it

may also be possible to roughly assess the detection probability from currently available

data using the approach proposed here.

180   This study showed that the double-observer approach, combined with

hierarchical capture-recapture models using Bayesian frameworks, might be an effective

option for estimating detection probabilities and the number of animal passes. We also

suggest that commercially available camera traps have higher detection probability

within a small focal area but still do not perfectly detect animals. The hierarchical

capture-recapture model used here can estimate the distribution of detection probability

and the number of animals passing concurrently, and thus, it is readily incorporated into

the current density estimation models. We hope that our approach will be incorporated

into them to improve their accuracy.

Methods

Model framework

The capture-recapture model applied here is the hierarchical model for stratified

populations proposed by Royle *et al.*[47]. The model aims to estimate local population

size using capture-recapture data from multiple independent locations. In the following,

we briefly describe the model in our context, including addressing heterogeneity in

detection probability.

197   Let us consider that we establish S independent camera stations in a survey area.

Then, we install K camera traps at each station to monitor exactly the same focal area

(totally $S \times K$ camera traps will be used). We assume that these camera traps detect

animals within the focal areas N_T times in total. For animal pass i ($i = 1, 2, 3, …, N_T$),

we will obtain (1) at which station the animal is detected (hereafter station identity; g_i),

and (2) how many of the K cameras at the station were successful in detecting the

animal pass (hereafter detection history; y_i). The hierarchal capture-recapture model

uses these two vector data, g_i and y_i.

205   Let the number of the animal passes at station s be N_s ($s = 1, 2, 3, …, S$). Then,

we assume that N_s follows a Poisson distribution with a parameter λ. In this case, the

probability of passage i occurring at station s is expected to be $\frac{\lambda}{\lambda \times S}$. Thus, station

identity, g_i, can be modelled as follows:

$$g_i \sim \text{Categorical} \left(\frac{\lambda}{\lambda \times S}\right)$$

When the number of the animal passes at station s, N_s, may have larger variation than

expected from the Poisson case, we may assume a negative binomial distribution model

or may give a random effect to the parameter of the Poisson distribution at the camera

station level.

The detection history y_i can be modelled using a data augmentation procedure[47].

Specifically, the original detection y_i is artificially augmented by many $M - n$ passes

with all-zero histories (i.e. not detected by any camera). The augmented data will

consist of the passage that occurred but was not detected by any camera (false zero),

which occurs with probability ψ, and the passage that did not occur (structural zeros)

with the probability $1 - \psi$. A set of latent augmentation binary variables, $z_1, z_2, \ldots z_M$, is

introduced, which denotes the false zero ($z = 1$) and the structural zero ($z = 0$). That is

$$z_i \sim \text{Bernoulli} (\psi).$$

The augmented data Y_i ($y_1, y_2 \ldots y_n$, 0, 0, … 0) with zero inflation can be modelled

conditional on the latent variables z_i.

$$Y_i \mid z_i = 1 \sim \text{Binomial} (J, p)$$

When accounting for the heterogeneity of detection among animal passes, it can be

accommodated using a beta distribution as follows;

$$Y_i \mid z_i = 1 \sim \text{Binomial } (J, p_i)$$

$$p_i \sim \text{Beta}(\textit{alpha}, \textit{beta})$$

Note that the expected total number of animal passes can be expressed as $\lambda \times S$. Thus,

ψ can be fixed as follows:

$$\psi = \frac{\lambda \times S}{M}$$

For more details of the models, see Royle et al. [47].

Testing the effectiveness of the hierarchical capture-recapture model

We performed Monte Carlo simulations to evaluate the effectiveness of the hierarchical

capture-recapture model. Because the model reliability has been confirmed well[48], we

here focused on the effects of heterogeneity in detection probability on the accuracy and

precision of the estimates.

We assumed that the number of detections by camera traps followed a negative

binomial distribution with a mean of 5.0 and dispersion parameter 1.27, which deprived

the actual data on an ungulate in African rainforests[34]. We also assumed two camera

traps at 30 stations (i.e. 60 camera traps in total). We generated detection histories (i.e.

the number of camera traps successfully detecting animals in each animal passage)

using a beta-binomial distribution with the expected detection probability at 0.8 or 0.4.

We varied the correlation coefficients (= 1/(alpha + beta + 1),) from 0.1 to 0.5 in 0.1

increments. The scale parameters of the beta distributions for each scenario are shown

in Table 1.

We estimated the model parameters using the Markov chain Monte Carlo

(MCMC) implemented in JAGS (version 3.4.0) in all the simulations. We assumed that

the number of animal passes followed a negative binomial distribution, and the

detection probability did a beta distribution. Since the estimation of scale parameters (α

and β) of the beta distribution, we transformed as $\alpha = p*phi$ and $\beta = p*(1 - phi)$ (p is an

expected detection probability). Then we used weakly informative prior (gamma

distribution with shape = 10 and rate = 2) for *phi* while a non-informative uniform

distribution from 0 to 1 for the detection probability[49]. We generated three chains of

3000 iterations after a burn-in of 1000 and thinned by 5. The convergence of models

was determined using the Gelman–Rubin statistic, where values <1.1 indicated

convergence. These procedures were repeated 300 times. The JAGS code is available as

supplementary material (Supplementary R1).

Determining a suitable camera installation

The above simulations suggested that a key to safely applying the hierarchical

capture-recapture model may avoid correlated detections. To determine a preferred

survey design to secure independent detections, we performed additional simulations.

Specifically, we tested how the position and the trigger speed of camera traps may affect

the independence of detections by considering the process by which cameras detect

moving animals.

268   The simulation was performed using similar procedures taken by Rowcliffe *et*

269   *al.* [19]. We assumed that the sensor of camera traps has a two-dimensional detection

270   surface defining the instantaneous 'risk' (as analogous to the risk of mortality in

271   survivorship analysis) of an animal being detected at any given location within the

272   camera's field of view (FOV). The instantaneous risk landscape was defined with

273   respect to distance r and angle relative to the camera as follows:

$$h = a\left(1 - \exp\left(-\left(r/s\right)^{-b}\right)\right)\exp\left(-\theta/c\right)$$
274

where a defines the maximum risk close to the sensor, s and b define the position and

shape of decline in risk with distance, respectively, and c defines the rate of decline in

risk with angle. We set a, s, b, and c at 0.2, 3.0, 5.0, and 0.5, respectively. Although

there is no empirical evidence for this particular function form, the simulated results

were comparable to those observed. Note that our interest is not in estimating the actual

values of the correlation but instead in obtaining information that will help us decide

what cameras to install and how to install them.

We made an animal pass through the risk landscape in a straight line in a

random direction. The animal movement speed (ms^{-1}) followed a log-normal

distribution with a mean ($\pm SD$) of 0.18 ± 0.15 m, which roughly accord with the speed

of red duikers in our study sites (Y Nakashima, unpublished data). We then generated a

random animal's position to be detected, considering the cumulative risk of detection

and the camera trap's trigger speed (for the details, see Rowcliffe et al.[19]).

We considered three designs of camera placements (Fig. 1). The first one is to

install two camera traps at the same position (i.e. mounted on the same tree) and in the

same direction to monitor the entire field of view within 10 m from the cameras (ins. 1).

The second and third one is assumed to monitor a specific small area within the FOV.

The focal area was a small equilateral triangle with a side length of 1.9m and was

centred within the FOV. The nearest vertex was set to be 1.9m away from the cameras.

This area corresponds to the highest detection probability in the camera model used in

our field study. This focal area is monitored from the same direction (ins. 2) or different

angles of 60 degrees (ins. 3). We then considered using camera models with a fast

trigger speed (0.1 s) and a slow (1.5 s) for each installation. Finally, we generated the

detection history of 500 times animal passes and calculated the correlation coefficients

between the two camera traps for each scenario.

Field surveys

We conducted a field survey in and around the Boumba-Bek and Nki National Parks in

southeast Cameroon (October 2018 to January 2019) and the Boso Peninsula in Japan

(August-November 2018). The study area of Cameroon consists primarily of evergreen

and semi-deciduous forests. The annual rainfall is approximately 1,500 mm, and the

mean annual temperature is *ca*. 24°C. Typically, the dry season occurs from December

to February and the rainy season is from March to November. The Boso Peninsula was

in the southern part of Chiba Prefecture in central Japan (35N °N, 140E). The vegetation

consists of either broad-leaved evergreen forests (*Castanopsis sieboldii* and *Quercus*

spp.) or coniferous plantations (*Cryptomeria japonica* and *Chamaecyparis obtusa*). The

monthly mean temperature was 20.6 ± 5.9 °C during the study period, with the highest

in August (27.4 °C) and the lowest (13.6 °C) in November.

The simulations suggest that camera traps monitor the predefined focal area

from different positions (see below). According to the results, we used the camera traps

with a high trigger speed (0.15 s) (Browning Strike Force Pro, BTC-5HDP, Browning,

Missouri, US) at both study sites. We regarded two camera traps to monitor the same

equilateral triangle area as in the simulation from different directions by 60 degrees (the

right panel of Fig. 1). We surrounded the focal area with a white rope and manually

filmed it with camera traps as a reference. The rope was removed after filming to avoid

disturbing the animal behaviour. We set camera traps at approximately 0.7 m above

ground without baits or lures. We used the 'video mode' and designated the video length

as 20 s and the delay period between videos at 1 s (minimum delay period in this

product). In Japan, we established seven camera stations at least 2 km apart and

installed two camera traps at each station. In Cameroon, we set 26 camera stations at

least 2-km apart from each other. Since the number of camera stations is not enough to

estimate the expected number of animal passes in Japan, we focused on the detection

probability

We determined whether animals passed within or outside the focal area by

superimposing videos and the reference image. We used only images of animals

crossing the focal area for subsequent analyses. We then matched each detection from

the two cameras to determine whether two camera traps successfully recorded an animal

pass. We applied the model to the detection probability of the species detected more

than 10 times in both Japan and Cameroon. The analysis was limited to images in which

the animal species were reliably identified. We plotted the estimated detection

probability against the median value of body mass (kg), drawn from Ohdachi et al.[50] for

animals in Japan and Kingdon[51] for those in Cameroon.

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Contributions

YN conceived the ideas and designed the methodology; SH and YG collected the data;

KM and YG performed video analyses; YN wrote the manuscript. All authors

contributed critically to the drafts and approved the final manuscript for publication.

Ethics declarations

Competing interests

The authors declare no competing interests.

Data availability

All the data used here is available in Supplementary Information.

 Figure legends

Fig. 1. A schematic diagram showing the installations of three focal areas assumed in

the simulation mimicking detection processes of moving animals. The circular sector

shows the hazard landscape of the detection zone. The open and filled circle shows the

positions of camera traps. In ins. 1 (left panel), the focal area was defined as the entire

field of view within 10 m from the cameras. In ins. 2 (centre panel), the focal was

restricted to an equilateral triangle with a side of 1.9 m. The distance from the camera to

the nearest vertex was assumed to be 1.9 m (shown in white lines). Finally, in ins. 3

(right panel), the same equilateral triangle was monitored from different angles of 60

degrees. The camera traps were assumed to have a sensor detection range of 42°. In the

right panel, the landscape of the second camera trap was shown in a grey polygon.

Fig. 2. The estimated detection probability of 12 species within a small focal area (1.56

m^2) in Cameroon (5 species) and Japan (7 species) using the capture-recapture models

for stratified populations. Error bar shows the 95% credible interval.

Table

Table 1. Summary of the Monte Carlo simulations to test the reliability of the

hierarchical capture-recapture model. Results are the mean of estimated median detection

probability *P,* the expected number of animal passes (lambda), and their 95% credible

interval (CI) coverage of the densities. Parameter estimation of the capture-recapture

model was performed using the Markov chain Monte Carlo (MCMC) method, and their

variances and credible limits were calculated as the posterior summary.

Detection probability	Correlation Coefficient	*P*		*lambda*	
		Mean	CI Coverage (%)	Mean	CI Coverage (%)
0.8	0.1	0.79	94.0	5.13	90.7
	0.2	0.82	91.7	5.01	92.3
	0.3	0.85	71.0	4.85	90.0
	0.4	0.89	42.7	4.67	87.3
	0.5	0.92	13.7	4.55	89.7
0.4	0.1	0.37	96.3	5.48	86.3
	0.2	0.42	98.3	4.75	87.7

0.3	0.48	85.0	4.32	87.3
0.4	0.56	38.0	3.65	63.3
0.5	0.63	11.3	3.10	31.7

Table 2. Summary of the Monte Carlo simulations to test the reliability of the hierarchical capture-recapture model.

Results are the mean of estimated median detection probability *P* and the expected number of animal passes (*lambda*) and

their 95% credible interval (CI) coverage of the densities. Parameter estimation of the capture-recapture model was

performed using the Markov chain Monte Carlo (MCMC) method, and their variances and credible limits were calculated

as the posterior summary.

	Setting values				Estimated values			
			Scale parameters			*P*		*lambda*
	Correlation						CI Coverage		CI Coverage
Detection probability	Coefficient	alpha	beta	Mean			Mean	
					(%)			(%)
0.8	0.1	7.20	1.80	0.79	94.0		5.13	90.7
	0.2	3.20	1.80	0.82	91.7		5.01	92.3

	0.3	1.87	0.47	0.85	71.0	4.85	90.0
	0.4	1.20	0.30	0.89	42.7	4.67	87.3
	0.5	0.80	0.20	0.92	13.7	4.55	89.7
0.4	0.1	0.40	0.60	0.37	96.3	5.48	86.3
	0.2	3.60	5.40	0.42	98.3	4.75	87.7
	0.3	1.60	2.40	0.48	85.0	4.32	87.3
	0.4	0.93	1.40	0.56	38.0	3.65	63.3
	0.5	0.40	0.60	0.63	11.3	3.10	31.7

532   **Figure**

534   Figure-1

536   Figure-2