

Effects of non-invasive brain stimulation on visual perspective taking: A meta-analytic study

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Abstract

Visual perspective taking (VPT) is a critical ability required by complex social interaction. Non-invasive brain stimulation (NIBS) has been increasingly used to examine the causal relationship between brain activity and VPT, yet with heterogeneous results. In the current study, we conducted two meta-analyses to examine the effects of NIBS of the right temporoparietal junction (rTPJ) or dorsomedial prefrontal cortex (dmPFC) on VPT, respectively. We performed a comprehensive literature search to identify qualified studies, and computed the standardized effect size (ES) for each combination of VPT level (Level-1: visibility judgment; Level-2: mental rotation) and perspective (self and other). Thirteen studies (rTPJ: 12 studies, 23 ESs; dmPFC: 4 studies, 14 ESs) were included in the meta-analyses. Random-effects models were used to generate the overall effects. Subgroup analyses for distinct VPT conditions were also performed. We found that rTPJ stimulation significantly improved participants' visibility judgment from the allocentric perspective, whereas its effects on other VPT conditions are negligible. Stimulation of dmPFC appeared to influence Level-1 performance from the egocentric perspective, although it was only based on a small number of studies. Notably, contrary to some theoretical models, we did not find strong evidence that these regions are involved in Level-2 VPT with a higher requirement of mental rotation. These findings not only advanced our understanding of the causal roles of the rTPJ and dmPFC in VPT, but also revealed the efficacy of NIBS on VPT is relatively small. Researchers should also be cautious about the potential publication bias and selective reporting.

Keywords: brain stimulation, visual perspective taking, temporoparietal junction, dorsomedial prefrontal cortex, meta-analysis

1 Introduction

The ability to take another's perspective is crucial for navigating complex social environments. To view the world from the second-person standpoint requires that one distinguishes between the self and the other in relation to the environment (Kessler and Rutherford, 2010; Lieberman, 2007). One social cognitive process that is closely related to this ability is visual perspective-taking (VPT). Dysfunction related to VPT has been observed in multiple clinical disorders, including autism and schizophrenia (Eack et al., 2017). Thus, it is essential to identify cognitive and neural mechanisms underlying the VPT process, as a steppingstone to target interventions for related disorders.

Flavell and colleagues (1977; 1981), identified two levels of VPT. Level-1 VPT refers to the ability to judge an object's visibility from the perspective of both the self and other. Consider, for example, playing hide-and-seek: You need knowledge about what the other person can see to hide from them. Children around the age of 18-24 months (Flavell et al., 1981) as well as chimpanzees (Brauer et al., 2007), dogs (Hare and Tomasello, 2005) and goats (Kaminski et al., 2005) show the ability to make such line-of-sight judgements. Level-2 VPT, on the other hand, enables humans to describe how an object looks from another's perspective and establishes a shared view of the world by creating a common reference frame for spatial localizations (Flavell, 1977; Kessler and Rutherford, 2010; Michelon and Zacks, 2006). For instance, imagine standing in front of a car, while your friend views it from behind: you are aware that although the car is visible to both of you, your friend has a different visual perspective on it (Pearson et al., 2013). Thus, Level-2 VPT has a higher level requirement of embodied rotation compared to its Level-1 counterpart (Martin et al., 2020).

In recent years, researchers have conducted a few neuroimaging studies to assess the neural mechanisms underlying VPT. One candidate region identified for this process is the temporoparietal junction (TPJ), as both the right and left parts of this region appear to play a critical role in multiple processes relevant to VPT, including detecting self-other incongruences, controlling self and other representations, and inhibiting the influence of the non-relevant representation via orienting attention (Bahnmann et al., 2009; Lamm et al., 2016; Quesque and Brass, 2019; Wolf et al., 2010). Indeed, the bilateral TPJ has often been reported across different VPT conditions (Bukowski, 2018; Schurz et al., 2013). Another critical region for integrating self-other processing is the dorsomedial prefrontal cortex (dmPFC). The dmPFC has also been implicated in making judgements about others (Denny et al., 2012), social information processing (Lieberman et al., 2019), and in introspection and assessment of mental states (Dore et al., 2015). In VPT tasks, the dmPFC has been reported when requiring egocentric perspective taking and suppressing the influence of the other's perspective, with a proposed process of imagining movement and suppressing the motor response to physically rotate the body (Bukowski, 2018; Mazzarella et al., 2013; Munzert et al., 2009).

While these neuroimaging studies highlight candidate brain hubs for self-other differentiation and integration in VPT, they are mostly based on correlational methods and thus causal relationships remain to be established (Bell and DeWall, 2018; Lieberman et al., 2019). Fortunately, non-invasive brain stimulation (NIBS) techniques, including transcranial direct current brain stimulation (tDCS) and transcranial magnetic stimulation (TMS), provide an approach to overcome this limitation (Donaldson et al., 2015; Polania et al., 2018). Specifically, tDCS applies weak direct currents to cortical regions. It could either facilitate or inhibit the spontaneous neuronal activity depending on the polarity of the electrode. Typically, anodal and cathodal stimulation has been shown to increase and decrease cortical excitability, respectively (Bell and DeWall, 2018; Brunoni and Vanderhasselt, 2014). TMS, on the other hand, uses a changing magnetic field to induce an ionic current at a brain region based on the principle of electromagnetic induction. The effects of TMS depend on factors including frequency, intensity, and duration of stimulation. For example, single-pulse TMS could depolarize the targeted neurons, whereas high-frequency (e.g., > 10 Hz) repetitive TMS (rTMS) typically disrupts the cortical function during the stimulation (Kobayashi and Pascual-Leone, 2003).

Researchers have increasingly used NIBS to investigate the causal role of different brain regions in VPT in the past few years. For example, anodal tDCS of the right TPJ (rTPJ) has been shown to improve participants'

65 performance when judging an item’s visibility from another’s perspective (Santesteban et al., 2012). Moreover,
66 another study showed that such an improvement could extend to Level-2 allocentric perspective-taking (Martin
67 et al., 2019a). However, there is also opposing evidence in which the rTPJ stimulation increased the impact of
68 perspective discrepancy during Level-1 VPT (Martin et al., 2020). Thus, despite that much effort has been devoted
69 to clarifying the causal relationships between brain regions and VPT, the overall findings to date paint a rather mixed
70 and inconclusive picture.

71 The inconsistency may be partly due to the complexity and heterogeneity of the existing VPT paradigms
72 (Bukowski 2018). As mentioned above, participants may be asked to judge the visibility or location of a target
73 from different perspectives (e.g., self or other), in which distinct underlying cognitive mechanisms may be involved.
74 For example, Level-2 VPT typically requires more embodied processing than Level-1 VPT (Martin et al., 2020).
75 However, there is no consensus yet if stimulating a brain region would selectively influence any VPT conditions.
76 Moreover, with a few exceptions (Martin et al., 2019b,a, 2020), most NIBS studies in this field only focused on one
77 brain region, making it difficult to compare the different regions’ roles in VPT.

78 The current study aims to clarify the causal roles of key brain regions in VPT. Based on the feasibility of the
79 included studies, we focused on studies targeting rTPJ and dmPFC and quantitatively synthesized the effects of
80 stimulation of these two regions on distinct VPT components.

81 2 Methods

82 The meta-analysis was conducted following the Cochrane Handbook for Systematic Reviews of Interventions (Higgins
83 et al., 2019) and PRISMA guidelines for meta-analyses (Liberati et al., 2009). The literature search and review, as
84 well as data extraction, were performed by two co-authors (Y.W.Y and V.C) independently. Discrepancies were
85 resolved by discussion.

86 2.1 Search strategy and eligibility criteria

87 An online literature search was conducted in Pubmed, Web of Science, and ProQuest for full-text articles from
88 January 2000 to June 2020 without language restrictions. The following query syntax was used: (“stimulation” OR
89 “TMS” OR “tDCS” OR “tACS” OR “tPCS” OR “tRNS” OR “TBS”) AND (“perspective taking” OR “perspective-
90 taking” OR “VPT”). To be included in the final meta-analysis, studies had to: (1) perform NIBS, (2) include a VPT
91 task, (3) enroll healthy participants, (4) have a control or sham condition. Studies without full-text available were
92 excluded. Note that, although previous brain stimulation studies mainly focused on rTPJ or dmPFC, we did not
93 explicitly include “rTPJ” OR “dmPFC” during the literature search. However, as a random-effects model requires
94 at least 3 effect sizes (ESs), the sample size limitation did not allow us to perform a meta-analysis on studies that
95 targeted other brain regions.

96 We first identified 27 potentially related studies by checking the title and abstract. Two authors then indepen-
97 dently decided if these studies should be included in the review by reading the full text. The inter-rater reliability
98 for the article selection shows a high agreement (Cohen’s Kappa = 0.87, $p < 0.001$). The authors resolved their dis-
99 agreement about two articles by discussion. The details were listed in Table S1. Finally, a total of 13 studies targeting
100 rTPJ or dmPFC were included in the meta-analyses.

101 2.2 Quality assessment

102 We used the Cochrane risk of bias tool to assess the quality of studies. Ratings (low, high, or unclear risk of bias)
103 were assigned to each study based on the following six criteria: (1) assessments for sequence generation, (2) allocation

104 concealment, (3) blinding of participants and researchers, (4) blinding of outcome assessment, (5) incomplete outcome
105 data, and (6) selective reporting.

106 2.3 Data extraction

107 For each included study, we extracted information regarding the sample size, age, and sex ratio. For intervention
108 characteristics, we extracted the type of NIBS technique, stimulation region, blinding protocol, intensity, duration of
109 active stimulation, valence (excitatory or inhibitory), and study design. For tasks, we extracted the VPT Level (1 or
110 2) and Perspective (Self or Other) information for each effect and focused on these four conditions in the following
111 analyses.

112 As most VPT tasks had an experimental condition, where the object being judged was incongruent from the Self
113 compared to the Other perspective, and a control condition, where the object was congruent from both perspectives,
114 we extracted the means and standard deviation (SDs) of the congruency effect (i.e., incongruent-congruent for RT
115 or congruent-incongruent for accuracy) for each VPT condition whenever possible. If there was no VPT control
116 condition, we extracted the means and SDs of RT or accuracy of the incongruent trials. If the data were only
117 presented in figures, means and SDs were estimated using the WebPlotDigitizer (Rohatgi, 2020). If only the standard
118 error (SE) was available, we calculated the SD with the formula: $SD = SE * \sqrt{n}$.

119 We also contacted the authors for the data and related information that was not reported, such as the correlation
120 between repeated measures.

121 2.4 Data analysis

122 Data analysis was performed using R (version 4.0.1) and the Metafor package (Viechtbauer, 2010). We used the
123 means and SDs to calculate the standardized ESs for each of the four conditions (VPT Level: 1 or 2, Perspective:
124 Self or Other) and each stimulation target (rTPJ, dmPFC), respectively. For between-subject studies, we calculated
125 the standardized mean difference (i.e., Hedge's g). For within-subject studies, we calculated the standardized mean
126 change (Morris and DeShon, 2002). We contacted the authors to ask for the raw data or correlation between repeated
127 measures if the information was not provided in studies.

128 If a study reported both RT and accuracy, we calculated a combined ES and variance using the following equations:

$$129 \quad ES_{comb} = \frac{1}{2}(ES_{RT} + ES_{acc})$$
$$130 \quad var_{comb} = \frac{1}{4}(var_{RT} + var_{acc} + 2corr)\sqrt{var_{RT}}\sqrt{var_{acc}}$$

131 the correlation between RT and accuracy was not provided in studies, we used an assumed $corr = 1$, which was
132 a conservative approach according to Borenstein et al. (2009) and Scammacca et al. (2014).

133 For three studies (Martin et al., 2020; Wang et al., 2016) that reported effects for different rotation angles and
134 body postures during VPT, we focused on the 160-degree condition and calculated a combined ES for both body
135 postures. One study (van Elk et al., 2017) that examined the effects of complex mental body transformation and
136 stimulation sessions (online and offline) on VPT, to ensure its comparability with other studies, we only focused on
137 the z-axis 180-degree condition and combined the ESs of both stimulation sessions. Moreover, the Australian group
138 of (Martin et al., 2019b) was also used in (Martin et al., 2019a), so we only used the data from the South-East Asian
139 group for (Martin et al., 2019b).

140 For consistency, the direction of the ES was defined as positive if the excitatory stimulation increased the incon-
141 gruent RT or decreased the incongruent accuracy, and negative for the inhibitory stimulation. For between-subject
142 tDCS studies that included anodal, cathodal, and sham stimulation groups, only the anodal > sham comparison was
143 used for the main meta-analysis to avoid the data of the sham group being used repeatedly. It should be noted that
144 the anodal stimulation group of (Martin et al., 2019a) is a subset of Martin et al., 2019a, so we only included sham
145 > cathodal comparison for this study.

146 We first performed two separate meta-analyses to examine the overall effect of rTPJ and dmPFC stimulation on
147 VPT, respectively. As some studies included multiple VPT conditions, we used both the two-level (first level: ES,
148 second level: VPT condition) and three-level (third level: study) random-effects model with restricted maximum-
149 likelihood estimator (Cheung, 2014; Konstantopoulos, 2011). A critical difference between the two models is that the
150 former ignored the within-sample variance and treated ESs from different VPT conditions as independent, whereas
151 the latter accounted for potential dependence between ESs from the same study. Model comparison based on Akaike's
152 information criteria (AIC) was conducted to test which model was better given the data. Heterogeneity among the
153 included ESs was assessed using the Q and I² tests. The funnel plot and Egger's test was used to assess publication
154 bias (Egger et al., 1997). If an Egger's test revealed significant publication bias, the trim-and-fill method (Duval and
155 Tweedie, 2000) was used to generate a corrected estimate after accounting for the effects of unpublished studies.

156 To investigate if rTPJ and dmPFC stimulation selectively influence any VPT conditions, we further conducted
157 subgroup analyses for each of the conditions and each stimulation target (except for rTPJ stimulation on VPT Level-2
158 Self condition because of insufficient ESs). Q test was used to statistically compare the aggregate ESs of different
159 subgroups.

160 To assess the reliability of the results, we conducted a few control analyses. First, we used a leave-one-study-out
161 analysis to examine the influence of individual ESs. Second, we conducted a control analysis for the dependent
162 variable measure by replacing each of the combined RT and accuracy ESs with ESs for only RT or accuracy and
163 comparing the effect sizes. In addition, we explored the effects of stimulation timepoint (offline or online), study
164 design (between- or within-subject design), and tDCS electrode size on results for TPJ stimulation. For studies using
165 dmPFC stimulation, we examined the effects of contrast selection (cathodal or anodal) on results, because the studies
166 were otherwise similar in their design.

167 3 Results

168 A total of 15 studies met the inclusion criteria for the qualitative review (Fig 1). Since a random-effects model
169 requires at least 3 effect sizes (ESs), 3 studies were excluded from the meta-analysis because they stimulated regions
170 not tested for in other studies and thus did not provide sufficient ESs required by a meta-analysis. Among the
171 remaining 13 studies, 9 stimulated rTPJ only, 1 stimulated dmPFC only, and 3 stimulated both regions (see Table 1
172 for study overview). After distinguishing 4 unique combinations of VPT Level and Perspective, we obtained 23 ESs
173 for the rTPJ and 14 ESs for the dmPFC targets.

174 3.1 Quality assessment

175 The quality assessment showed that all the 16 studies used a random assignment to allocate participants to different
176 stimulation conditions. A total of 6 tDCS and 5 TMS studies used the within-subject design, whereas the remaining
177 4 tDCS and 1 TMS studies used the between-subject design. However, none of the studies included an explicit
178 statement about the allocation concealment, yielding potential biases related to this criterion.

179 Regarding blinding of participants and researchers, 6 studies were double-blinded. The remaining 10 studies did
180 not report if blinding was used, resulting in the unclear risk of bias regarding this criterion. All studies used sham
181 procedures. Specifically, 9 tDCS studies used a procedure by turning off the electric current shortly after stimulation
182 onset, with a length ranging from 15 to 60 seconds. The remaining 1 tDCS study used anodal stimulation of the
183 occipital cortex with the same duration and intensity as an active control condition (Santesteban et al., 2015). For
184 TMS studies, 2 used a sham coil and played loud sounds mimicking TMS discharges via earphones during both active
185 and sham stimulation (Gooding-Williams et al., 2017; Wang et al., 2016). Another 3 studies stimulated the vertex
186 (Guise et al., 2007; Qureshi et al., 2020; Soutschek et al., 2016) and 1 study stimulated the occipital cortex for control

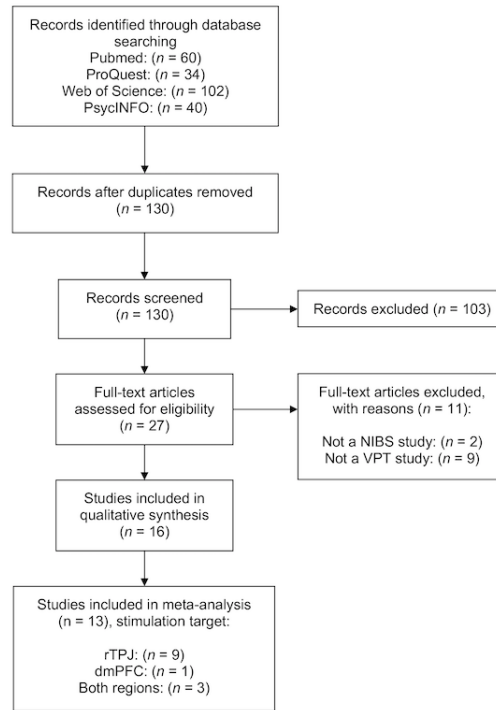


Figure 1: PRISMA flow diagram of literature search strategy

187 (Santiesteban et al., 2017).

188 The risk of incomplete outcome data (e.g., attrition bias) was low for all of the included studies except one,
189 in which 11 participants were excluded due to technical issues related to tDCS (van Elk et al., 2017), although it
190 included a relatively large sample of participants (n = 58).

191 Finally, regarding selective reporting, 7 studies reported both RT and accuracy measures, 7 and 2 studies only
192 reported RT or accuracy respectively. Thus, the 9 studies that only reported one dependent variable may be associated
193 with a high risk of selective reporting. Moreover, 5 studies reported congruency effects, whereas the remaining 11
194 studies reported data from specific conditions (e.g., incongruent trials). Therefore, the research degrees of freedom
195 in data analysis and results reporting appears to be high.

196 3.2 Effects of rTPJ stimulation

197 The two-level random-effects model showed that the overall effect of rTPJ stimulation on VPT was not significant
198 (ES = -0.10, 95% CI: [-0.22, 0.03], Z = -1.49, p = 0.14), with high dispersion and residual heterogeneity (I2Level-2=
199 62.34%, Q(22) = 60.08, p < 0.001). Egger's test (Z = -1.75, p = 0.08) indicated that the publication bias was not
200 significant (Fig S1). The three-level random-effects model considering the dependence between ESs from the same
201 study showed a slightly larger but not significant effect (ES = -0.18, 95% CI: [-0.44, 0.08], Z = -1.40, p = 0.17).
202 In this model, 76.12% (I2Level-3) of the total variation can be attributed to between-study, 5.10% to within-study
203 heterogeneity (I2Level-2) and 18.77% to sampling variance (I2Level-1). Model comparison slightly preferred the
204 three-level random-effects model (AIC = 25.23) over the two-level model (AIC = 27.04), reflecting the dependency
205 between ESs from the same study.

206 To test the effects of rTPJ stimulation on specific task conditions, we further ran 4 subgroup meta-analyses

for each unique combination of VPT Level and Perspective (Fig 3). Our results showed that rTPJ stimulation significantly improved participants' performance on Level-1 VPT Other condition (ES = -0.39, 95% CI: [-0.76, -0.02], Z = -2.09, p = 0.04). Egger's test (Z = -4.24, p < 0.001) indicated a high risk of publication bias in these studies. The trim-and-fill method that considered this bias yielded a negligible effect (ES = -0.07). The effects on Level-1 VPT Self (ES = -0.08, 95% CI: [-0.27, 0.11], Z = -0.84, p = 0.40), Level-2 VPT Self (ES = 0.14, 95% CI: [-0.15, 0.44], Z = 1.01, p = 0.34) and Level-2 VPT Other condition (ES = -0.06, 95% CI: [-0.27, 0.15], Z = -0.56, p = 0.58) are small and insignificant. Egger's tests for these three subgroups did not show significant publication bias either (ps > 0.25). To further test if the effects on Level-1 VPT were stronger than the other three conditions, we conducted a three-level meta-regression with the task condition as a moderator. The Wald tests only showed a significant difference between Level-1 Other and Level-2 Self conditions (Z = 2.36, p = 0.02).

3.3 Effects of dmPFC stimulation

The two-level random-effects model showed that the overall effect of dmPFC on VPT was significant, showing a slight increase in reaction time or error rate after stimulating the dmPFC (ES = 0.09, 95% CI: [0.02, 0.17], Z = 2.41, p = 0.02, Fig S2), with very low heterogeneity (I² Level-2 < 0.01%, Q(13) = 9.29, p = 0.75). Egger's test (Z = 1.20, p = 0.22) suggests that the risk of the potential publication bias was low (Fig S1). The three-level model yielded similar results (ES = 0.09, 95% CI: [0.01, 0.18], Z = 2.42, p = 0.03). As expected, the model comparison preferred two-level random-effects model (AIC = -11.20) than the three-level counterpart (AIC = -9.19).

Again, we ran subgroup meta-analyses for each VPT condition respectively (Fig 4). We found a significant effect of dmPFC stimulation in the Level-1 VPT Self condition (ES = 0.18, 95% CI: [0.004, 0.36], SE = 0.09, Z = 2.01, p = 0.04). None of the other three conditions showed a significant effect: Level-1 VPT Other (ES = 0.05, 95% CI: [-0.09, 0.21], Z = 0.69, p = 0.49), Level-2 VPT Self (ES = 0.16, 95% CI: [-0.02, 0.34], Z = 1.75, p = 0.08) and Level-2 VPT Other condition (ES = 0.04, 95% CI: [-0.09, 0.18], Z = 0.67, p = 0.50). Egger's tests for these subgroups did not show significant publication bias (ps > 0.50). We also conducted the three-level meta-regression and Wald tests as above but found no significant difference between the subgroups (ps > 0.23).

3.4 Supplementary analysis

We first performed sensitivity analyses regarding the selection of the dependent variable for each stimulation target. In the analyses mentioned above, we used the combined RT and accuracy ESs for studies that reported both measures. The sensitivity analyses showed that the results of the meta-analyses remained similar if we only used RT or accuracy ESs for those studies (Table S3). Moreover, we conducted stricter sensitivity analyses by doing the above analyses only based on RT or accuracy data, which did not show significant differences between models based on different dependent variables (Table S3). However, it should be noted that the effects of rTPJ stimulation on Level-2 VPT Other condition became negligible in the RT-only model (ES = -0.10). Since the sample sizes were smaller when focusing on a single dependent variable, more evidence is needed to confirm these findings. We also compared the overall effects of studies using different stimulation timepoints (online and offline), tDCS electrode sizes, and study designs (within- and between-subject designs) and did not find significant differences either (Table S4). Finally, the leave-one-study-out analyses showed that the overall effects of the main and subgroup analyses were relatively stable. The key findings were not driven by any individual studies. The detailed results were listed in Table S5.

4 Discussion

This meta-analytic study examined how rTPJ and dmPFC stimulation influenced VPT across 13 studies. The results showed that the rTPJ was mainly involved in allocentric visibility judgment. The dmPFC appeared to play a role in

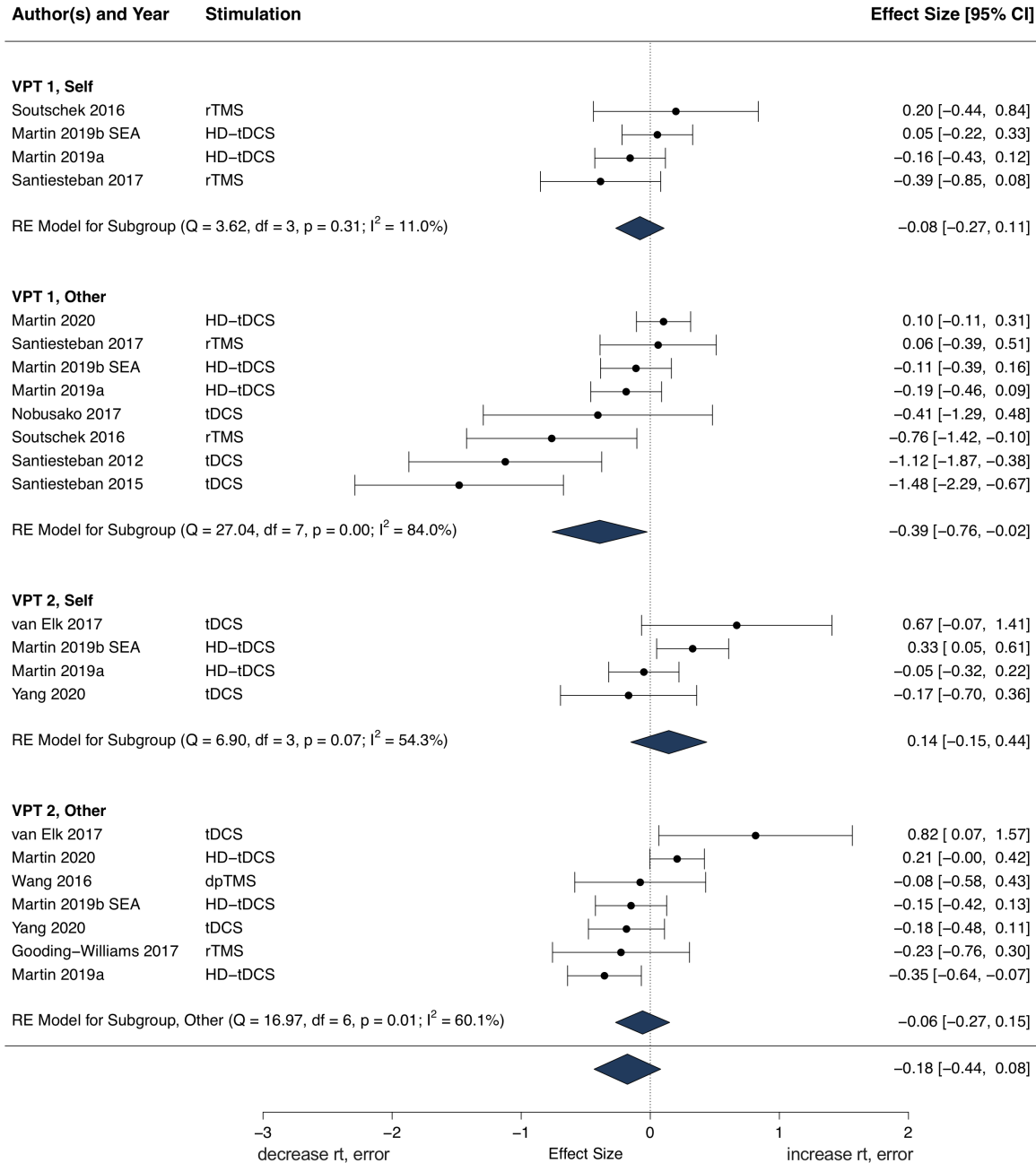


Figure 2: The effect of rTPJ stimulation on different VPT conditions. The excitatory stimulation of the rTPJ (vice versa for the inhibitory stimulation) significantly increased participants' performance (i.e., shorter RT or lower error rate) in Level-1 VPT Other condition (ES = -0.39). The effects of rTPJ on other VPT conditions are negligible. Congruency Effect = incongruent-congruent for RT, congruent-incongruent for accuracy. SEA = South-East Asian participants.

247 processes related to the egocentric perspective. Importantly, the overall effects of rTPJ and dmPFC stimulation on
248 most VPT conditions were negligible. These findings not only advanced our understanding of the neural mechanisms
249 underlying VPT, but also systematically evaluated the efficacy of NIBS on VPT and the implications for its practical
250 use.

251 One main finding of our meta-analysis is that excitatory stimulation of the rTPJ increased performance in Level-1
252 VPT Other condition: Participants' error rate and reaction time decreased during line-of-sight or visibility judgements
253 when their own perspective was incongruent with the other's perspective. Level-1 VPT requires participants to trace
254 the line of sight between the self and target object and does not rely on deliberate movement simulation (Kessler and
255 Rutherford, 2010). Our findings thus suggest that rTPJ plays a critical role in suppressing the egocentric perspective
256 when taking the other's perspective (Santiesteban et al., 2012). Notably, the ability to overcome one's self-centered
257 perspective implemented in the posterior TPJ was also recruited in choosing delayed and prosocial rewards (Soutschek
258 et al., 2016). Moreover, two sub processes have been proposed in Level-1 VPT: (1) perspective calculation, which is
259 the fast, automatic, and cognitively efficient calculation of someone else's perspective, and (2) perspective selection,
260 which is the effortful selection of either one representation, depending on task demands (Apperly and Butterfill, 2009;
261 Qureshi et al., 2020; Todd et al., 2019). Therefore, a promising direction for future studies is to elucidate the effects
262 of rTPJ stimulation on these two subprocesses of Level-1 VPT.

263 Notably, the subgroup analysis showed that the aggregate effect of rTPJ stimulation on Level-2 VPT Other
264 condition was negligible. Because of the proposed role of the rTPJ in Theory of Mind (Krall et al., 2015); Saxe and
265 Wexler (2005) and its implications in multisensory integration between proprioceptive and visual inputs (Blanke and
266 Mohr, 2005; Ionta et al., 2011), it was suggested that the rTPJ might be critical for Level-2 allocentric perspective
267 taking, which has a high-level requirement of embodied rotation (Martin et al., 2020). The current study, however,
268 did not provide strong evidence for this hypothesis and thus cast doubt on the rTPJ's involvement in embodied
269 processes during VPT. A plausible explanation is that the relative contribution of the rTPJ-centered network to
270 Level-2 VPT is smaller than its Level-1 counterpart, as Level-2 VPT is more complex and may rely on the coordinated
271 effort of more distinct networks, although this finding remains to be confirmed due to relatively small sample size
272 and high heterogeneity in NIBS methods. Moreover, the Bayesian statistics (Schmalz et al., 2021; van de Schoot et al.,
273 2021) is able to provide a more thorough investigation into null results when more VPT studies are accumulated in
274 the future.

275 In addition to the rTPJ, the dmPFC is also closely related to complex social cognition (Lieberman et al., 2019),
276 particularly in merging the self- and other-related information to guide social decision-making (Schurz and Perner,
277 2015; Wittmann et al., 2016). In the context of VPT, we found that the dmPFC stimulation significantly decreased
278 participants' performance during Level-1 egocentric perspective taking, possibly by increasing the salience of irrelevant
279 information from the allocentric perspective (Martin et al., 2019a). The effects of dmPFC stimulation on allocentric
280 perspective taking are rather negligible. Taken together, the dmPFC might be recruited to integrate the external
281 information into one's own perspective, especially when embodied rotation is less required (i.e., Level-1). This
282 interpretation is consistent with findings that the excitatory dmPFC stimulation decreased the self-reference effect in
283 episodic memory (Martin et al., 2019a). However, findings should be regarded as preliminary and interpreted with
284 caution, because they are based on a relatively small number of studies from the same research group.

285 The current study also revealed some general problems related to NIBS studies in this field. First, there is
286 no consensus on the selection of dependent variables. Both RT and accuracy were widely used to reflect VPT
287 performance. Moreover, some studies calculated the differences between incongruent and congruent trials, whereas
288 others just analyzed data from specific conditions (e.g., incongruent trials). This flexibility may increase the risks
289 of selective reporting. Therefore, we recommend researchers report both RT and accuracy measures for all task
290 conditions and perform a multiverse analysis (Steegen et al., 2016) to comprehensively evaluate the effects of NIBS
291 on VPT. Second, previous studies showed that trait factors, such as baseline perspective-taking ability (Fini et al.,

292 [2017] or empathetic understanding (Bukowski et al., 2020), may modulate the effect of NIBS (of other brain regions)
293 on spatial or emotional perspective taking. However, most studies only focused on the stimulation effects at the group
294 level without taking individual differences into account. It is particularly important for between-subject studies since
295 the effects may be attributed to differences on a dispositional factor rather than stimulation itself. Finally, although
296 the potential of NIBS as an intervention for VPT-related disorders has been proposed by some researchers (Martin
297 et al., 2020; Santiesteban et al., 2012), our findings cast doubt on its efficacy. For example, even the largest effect
298 we found (i.e., rTPJ stimulation on Level-1 VPT Other condition) is relatively small and may be associated with
299 publication bias. Therefore, it appears still immature to apply this approach to practical use at the current stage.

300 As one of the first meta-analytic studies that examined NIBS on VPT tasks, the present study has some limitations.
301 First, the sample sizes for some subgroup analyses are limited. The current findings thus should be interpreted with
302 caution. Results are expected to be more robust and reliable with future NIBS studies on VPT.

303 Second, the meta-analyses are mainly based on the evidence from tDCS studies. Due to the sample size limitation,
304 we are unable to directly compare the effects of different NIBS techniques (e.g., tDCS vs. TMS) on a certain VPT
305 condition in the current study. This issue is likely to be addressed when more TMS studies are available.

306 Third, both the rTPJ and dmPFC are heterogeneous regions. At least three subregions have been identified in
307 the TPJ (Mars et al., 2012). The posterior region might be recruited during control of self and other representations
308 and the anterior region during attentional reorientation (Corbetta et al., 2008; Krall et al., 2015). Similarly, the
309 dorsal and ventral parts of the dmPFC appear to mainly involve in other- and self-related processes, respectively
310 (Lieberman et al., 2019). Due to the spatial precision limitation of stimulation (e.g., two-electrode tDCS), most
311 studies did not specify which subregions of the rTPJ or dmPFC were stimulated. Future studies may address this
312 issue with the assistance of neuronavigation and more focal NIBS techniques (Donaldson et al., 2015).

313 Finally, most of the included studies focused on the role of an individual brain region in VPT. However, neu-
314 roimaging evidence suggests that complex social cognitive processes, such as VPT, depend on the interactions of
315 multiple brain areas (Schurz et al., 2013). Particularly, the effects of tDCS may be not limited to the targeted area
316 either, because the current travels along the path between anodal and cathodal electrodes (Stagg and Nitsche, 2011).
317 Therefore, another future avenue for research is to elucidate how NIBS influences interactions between brain networks
318 during VPT.

319 5 Conclusions

320 The current meta-analytic study found that the rTPJ and dmPFC appeared to be causally involved in allocentric and
321 egocentric Level-1 VPT, respectively. The effects of the stimulation of both regions on Level-2 VPT were negligible,
322 suggesting that neither was necessary embodied processing. These findings contribute to a better understanding of
323 the neural mechanisms of VPT and show the limitations and future directions of the NIBS technique as a potential
324 intervention for patients with related deficits.

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Study	Target	Type	Electrode Size	Study design	N	NIBS control	Intensity	Off-/ On-line	Stimulation	Duration, pulse	DV	VPT level	Perspective
Conson et al. (2015) *	right dlPFC (F4), left dlPFC (F3)	tDCS	5*7cm	Cross-over	16	Sham	1.0mA	Off-line	Anodal, cathodal	15min	RT	2	Self, Other
Gooding-Williams et al. (2017) *	rTPJ	rTMS	7cm	Within	14	Sham	80% RMT	On-line	Repetitive (Inhibitory for 10 Hz)	15 pulses at 6 (not used in meta-analyses) or 10 Hz	RT	2	Other
Guisse et al. (2007) *	right FC (FP2), left FC (FP1)	TMS	7cm	Within	7	Sham (to CZ)	90% RMT	On-line	Single pulse (inhibitory)	96 single pulses (24/condition)	RT	1	Self, Other
Martin et al. (2019a)	rTPJ (CP6), dmPFC (15% from Fz to FPz)	HD-tDCS	Center: 2.5 cm, Return: 7.5/9.8cm	Cross-over	52	Sham	1.0mA	Off-line	Anodal	20min	RT CE, accuracy	1, 2	Self, Other
Martin et al. (2019b)	rTPJ (CP6), dmPFC (65% from Cz to Fpz)	HD-tDCS	Center: 2.5 cm, Return: 7.5/9.8cm	Cross-over	52	Sham	1.0mA	Off-line	Anodal	20min	RT CE, accuracy	1, 2	Self, Other
Martin et al. (2020)	rTPJ (CP6), dmPFC (65% from Cz to Fpz)	HD-tDCS	Center: 2.5 cm, Return: 7.5/9cm	Cross-over	88	Sham	1.0 mA	Off-line	Anodal	20min	RT	1, 2	Other
Martin et al. (2017)	rTPJ (CP6), dmPFC (15% from Fz to FPz)	HD-tDCS	Center: 2.5 cm, Return: 9.2/11.5cm	Cross-over	40	Sham	1.0mA	Off-line	Anodal, cathodal	20min	RT	1, 2	Self, Other
Nobusako et al. (2017)	rTPJ (CP6), IFC (FC6)	tDCS	5*7cm	Between	10/group	Sham	1.0mA	Off-line	Anodal	20min	RT, Accuracy	1	Other
Qureshi et al. (2020) *	right dlPFC (F4)	TMS	7cm	Cross-over	31	Vertex stimulation	80% RMT	Off-line	cTBS (inhibitory)	Three-pulse bursts 50 Hz,600 pulses	RT, IES, Accuracy	1	Self, Other
Santiesteban et al. (2012)	rTPJ (CP6)	tDCS	5*7cm	Between	17: anodal, 17: cathodal, 15: sham	Sham	1.0mA	Off-line	Anodal, cathodal (ref.: vertex)	20min	Accuracy	1	Other
Santiesteban et al. (2015)	rTPJ (CP6), ITPJ (CP5)	tDCS	5*7cm	Between	15/group	OZ stimulation	1.0mA	Off-line	Anodal (ref.: vertex)	20min	Accuracy, RT	1	Other
Santiesteban et al. (2017)	rTPJ (MNI: 54, -47, 26)	rTMS	7cm	Within	19	Mid-occipital stimulation	110% RMT	On-line	inhibitory	6 pulses/trial 10Hz	RT	1	Self, Other
Soutschek et al. (2016)	rTPJ (MNI: 50, -60, 32)	TMS	7cm	Between	20: rTPJ, 18: vertex, 21: S1	Vertex, S1 stimulation	80% RMT	Off-line	cTBS (inhibitory)	Three-pulse bursts 50Hz, 600 pulses	Accuracy	1 2	Other Self
van Elk et al. (2017)	rTPJ (CP6)	tDCS	5*7cm	Between	16: anodal, 15: cathodal, 14: sham	Sham	1.0mA	1/2 On-line, 1/2 Off-line	Anodal, cathodal (ref.: C3)	20min	RT	2	Self, Other
Wang et al. (2016)	rTPJ (MNI: 50, -60, 32)	TMS	7cm	Within	15	Sham	110% RMT	On-line	Dual pulse (inhibitory)	80 dual pulses	RT	2	Other
Yang et al. (2020)	rTPJ (CP6), ITPJ (CP5)	tDCS	5*7cm	Cross-over	45	Sham	2.0mA	On-line	Anodal (ref.: vertex)	29min (10min rest)	RT,accuracy, IES, switching cost	2	Self, Other

333 Table 1: * These studies were not included in the meta-analyses because they targeted different regions. CE: congruency ef-
 334 fect; dlPFC: dorsolateral prefrontal cortex; dmPFC: dorsomedial prefrontal cortex; ITPJ: left temporoparietal junction; min:
 335 minutes; NIBS: Non-invasive brain stimulation; ref: reference; RMT: resting motor threshold; rTPJ: right temporoparietal
 junction.

331 7 Conflicts of interest

336 None

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