

# 1 **Scaling between stomatal size and density in forest plants**

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3 Congcong Liu<sup>1,2#</sup>, Christopher D. Muir<sup>3#</sup>, Ying Li<sup>1</sup>, Li Xu<sup>1</sup>, Mingxu Li<sup>1</sup>, Jiahui Zhang<sup>1,2</sup>, Hugo

4 Jan de Boer<sup>4</sup>, Lawren Sack<sup>5</sup>, Xingguo Han<sup>6</sup>, Guirui Yu<sup>1,2</sup>, Nianpeng He<sup>1,2,7\*</sup>

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6 <sup>1</sup> Key Laboratory of Ecosystem Network Observation and Modeling, Institute of Geographic  
7 Sciences and Natural Resources Research, Chinese Academy of Sciences, Beijing 100101,  
8 China

9 <sup>2</sup> University of Chinese Academy of Sciences, Beijing 100049, China

10 <sup>3</sup> School of Life Sciences, University of Hawai'i at Mānoa, Honolulu, 96822, USA

11 <sup>4</sup> Copernicus Institute of Sustainable Development' Department of Environmental Sciences,  
12 Utrecht University, the Netherlands, 80125, Utrecht.

13 <sup>5</sup> Department of Ecology and Evolutionary Biology, University of California, Los Angeles,  
14 90025, USA

15 <sup>6</sup> State Key Laboratory of Vegetation and Environmental Change, Institute of Botany, Chinese  
16 Academy of Sciences, Beijing 100093, China

17 <sup>7</sup> Institute of Grassland Science, Northeast Normal University, and Key Laboratory of Vegetation  
18 Ecology, Ministry of Education, Changchun 130024, China

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20 # These authors contributed equally to this work.

21 \*Corresponding author Nianpeng He (henp@igsnr.ac.cn).

22 Tel.: +86-10-64889263

23 Fax: +86-10-64889399

24 **Abstract**

25 **The size and density of stomatal pores limit the maximum rate of leaf carbon gain and**  
26 **water loss ( $g_{\max}$ ) in land plants. The limits of  $g_{\max}$  due to anatomy, and its constraint by the**  
27 **negative correlation of stomatal size and density at broad phylogenetic scales, has been**  
28 **unclear and controversial. The prevailing hypothesis posits that adaptation to higher  $g_{\max}$  is**  
29 **typically constrained by geometry and/or an economic need to reduce the allocation of**  
30 **epidermal area to stomata (stomatal-area minimization), and this would require the**  
31 **evolution of greater numbers of smaller stomata. Another view, supported by the data, is**  
32 **that across plant diversity, epidermal area allocated to guard cells versus other cells can be**  
33 **optimized without major trade-offs, and higher  $g_{\max}$  would typically be achieved with a**  
34 **higher allocation of epidermal area to stomata (stomatal-area increase). We tested these**  
35 **hypotheses by comparing their predictions for the structure of the covariance of stomatal**  
36 **size and density across species, applying macroevolutionary models and phylogenetic**  
37 **regression to data for 2408 species of angiosperms, gymnosperms, and ferns from forests**  
38 **worldwide. The observed stomatal size-density scaling and covariance supported the**  
39 **stomatal-area increase hypothesis for high  $g_{\max}$ . A higher  $g_{\max}$  involves construction costs**  
40 **and maintenance costs that should be considered in models assessing optimal stomatal**  
41 **conductance for predictions of water use, photosynthesis, and water-use efficiency as**  
42 **influences on crop productivity or in Earth System models.**

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44 Stomatal pores are critical determinants of the function of plants and the composition of  
45 the atmosphere (1). The stomatal conductance to diffusion of water vapor and CO<sub>2</sub> ( $g_s$ )  
46 influences a broad spectrum of ecological processes at leaf, community, and ecosystem scales,  
47 including photosynthesis, net primary production, and water use efficiency (2, 3). Theoretically,  
48 stomata can regulate  $g_s$  through evolutionary or plastic shifts in stomatal size or numbers (4) or  
49 through short-term stomatal aperture changes (5). The  $g_s$ , and its typical operational value ( $g_{op}$ ),  
50 can thus vary from near zero with stomata fully closed and  $g_{max}$  with stomata fully open. The  
51  $g_{max}$  is a fundamental anatomical constraint, and across species measured under controlled  
52 conditions,  $g_{op}$  and  $g_{max}$  are correlated (6, 7). Because of their importance in controlling leaf  
53 water and CO<sub>2</sub> fluxes, stomatal anatomy can provide critical information in global vegetation and  
54 crop models (8-11) toward the current grand challenge of understanding how crops and forest  
55 trees are optimized for carbon gain versus water use. Yet, there has been substantial debate about  
56 the anatomical underpinnings of the evolution of higher  $g_{max}$ , and its associated costs.

57 The  $g_{max}$  is a mathematic function of underlying anatomical traits stomatal density ( $D_s$ ,  
58 number of pores per unit epidermal area) and size ( $A_s$ , area of guard cells surrounding each pore).  
59 Indeed, these traits are widely used to study the adaptation and competition of plants because  
60 they are reliable indicators of  $g_{max}$ (12-18). Further an inverse relationship between  $A_s$  and  $D_s$   
61 across diverse plant species has been recognized since 1865 (19). A prevalent view in the  
62 literature established by Franks and Beerling (20) is that the negative  $A_s$  and  $D_s$  relationship and  
63 the cost of stomatal area place a strong constraint on the evolution of  $g_{max}$ . According to an early  
64 version of the “stomatal area minimization hypothesis” a packing limit geometry constrains  $g_{max}$ ,  
65 because the total fraction of epidermal area allocated to stomata ( $f_s$ ) cannot exceed unity:

$$f_s = D_s A_s. \quad (1)$$

66 This, in turn generates the negative  $A_s$  and  $D_s$  relationship such that the evolution of larger  
67 numbers of stomata would necessitate reduction in their size. Thus, higher  $g_{\max}$  can only be  
68 achieved by the evolution of larger numbers of smaller stomata. This early packing geometry  
69 argument was rendered moot given observations that for functional leaves, maximum  $f_s$  is  
70 usually far lower than unity (33.6% in our data; solid line in Fig. 1a). First, because stomata need  
71 to be spaced out by epidermal cells to open and close properly (21), and second, because the  
72 development of higher  $D_s$  can occur through the increased differentiation of epidermal cells into  
73 stomata (i.e., achieving higher stomatal differentiation rate, or stomatal index (22, 23), such that  
74 stomatal numbers can be independent of sizes. Yet the stomatal area minimization hypotheses for  
75 the evolution of higher  $g_{\max}$  and its association with the negative  $A_s$  and  $D_s$  relationship was  
76 reached by a different argument: that to minimize stomatal construction and maintenance costs  
77 (24), plants evolving higher  $g_{\max}$  must do so with a reduced  $f_s$ , and this maximization of  $g_{\max}$   
78 relative to  $f_s$  would in turn generate the negative  $A_s$  and  $D_s$  relationship (20).

79 To see why, note that a leaf's  $g_{\max}$  is determined by stomatal anatomy:

$$g_{\max} = bmD_sA_s^{0.5}, \quad (2)$$

80 where  $b$  and  $m$  are biophysical and morphological constants, respectively (22) (see Methods for  
81 equations to calculate these constants). By Eq. 2, a higher  $g_{\max}$  can be achieved with a smaller  
82 total stomatal area by increasing stomatal number and reducing stomatal size, because smaller  
83 stomata also have a shorter channel for diffusion. For example, consider two leaves with  
84 stomatal densities 250 and 200 pores  $\text{mm}^{-2}$  and stomatal areas  $150 \mu\text{m}^2$  and  $187.5 \mu\text{m}^2$ . They  
85 have identical  $f_s$ , but  $g_{\max}$  at 25 °C is 11% greater for the leaf with smaller stomata (1.32 versus  
86  $1.47 \text{ mol m}^{-2} \text{ s}^{-1}$ ). Thus, selection for higher  $g_{\max}$  would result in more numerous, smaller

87 stomata, to minimize epidermal allocation to stomata and the evolution of higher  $g_{\max}$  is strongly  
88 associated with the negative  $A_s$  and  $D_s$  relationship.

89 The ‘stomatal-area minimization’ hypothesis is controversial, however, as it is at odds  
90 with data in the literature that instead support an opposite, ‘stomatal-area increase’ hypothesis,  
91 i.e., that  $g_{\max}$  should increase with  $f_s$  during evolution. Conversely, selection decreased  $g_{\max}$   
92 would be associated with decreased  $f_s$ . The positive covariance of  $g_{\max}$  with  $f_s$  has been shown in  
93 many studies and has been utilized in many papers in the literature that have indeed used  $f_s$  as a  
94 proxy for  $g_{\max}$  (25, 26). According to the “stomatal-area increase” hypothesis, selection for  
95 higher  $g_{\max}$  is much stronger than that to minimize cost, leading to greater surface allocation,  
96 even if this incurs a cost. Under this scenario, the negative  $A_s$  and  $D_s$  relationship would not act  
97 directly as a constraint on  $g_{\max}$ . Yet, selection for higher  $g_{\max}$  would generate a distinctive  
98 covariation between these constituent anatomical traits. First, no relationship of  $A_s$  and  $D_s$  is  
99 absolutely required, which is consistent with data in the literature for species sets for which no  
100 relationship is found (27); as less than 50% of the leaf surface is typically taken up by stomata  
101 and many qualitatively different relationships between stomatal size and density across species  
102 are geometrically possible, including negative, zero, and positive covariances (ellipses in Fig.  
103 1a). Yet, on average, a specific covariation would be expected if many combinations of  $A_s$  and  $D_s$   
104 have similar fitness through their effect on either  $g_{\max}$  or  $f_s$ , as we derive below.

105 It is critical to distinguish between these hypotheses for the evolution of  $g_{\max}$  and the  
106 potential for  $f_s$  to constrain the observed stomatal size-density relationship. Implications of  
107 stomatal-area minimization are that  $g_{\max}$  is ultimately constrained by the costs of high  $f_s$ , that  
108 such costs are minimized, and that evolving higher  $f_s$  would be slowed by costs associated with  
109 allocating too much epidermal area to stomata<sup>27–32</sup>. By contrast, the stomatal-area increase

110 hypothesis implies that selection on stomatal size and density primarily optimizes  $g_{\max}$ , which  
111 varies across environments, and greater  $g_{\max}$  incurs stomatal construction costs and opportunity  
112 costs of epidermal space. Testing these hypotheses will further reveal how the evolution of high  
113  $g_{\max}$  relates to the general inverse stomatal size-density relationship.

114         Indeed, these hypotheses can be tested against data for diverse species by considering in  
115 detail the covariation among  $D_S$  and  $A_S$ , for which they make different predictions. Under both  
116 hypotheses,  $D_S$  and  $A_S$  are constituents of composite traits,  $f_S$  or  $g_{\max}$  (Eq. 1-2; Fig. 1b). We  
117 investigated how stomatal size-density scaling would differ between the hypotheses using  
118 models of macroevolutionary landscapes (28-31). We used the Ornstein-Uhlenbeck (OU) model  
119 originally derived from quantitative genetics for intraspecific (population) trait microevolution  
120 by Lande (32), and developed by Hansen (29) and others (28) for macroevolutionary interspecific  
121 trait variation. In the macroevolutionary OU model, interspecific trait variation expands through  
122 time until it reaches a stationary distribution around a long-term average (29). Within each  
123 species, microevolutionary forces (selection, genetic drift, mutation, and migration) and the  
124 environment drive genetic and plastic trait variation, respectively, and species' trait means  
125 should be near their current adaptive optimum. The across-species distribution that becomes  
126 stationary in the OU model is thus dependent on these independent shifts in species' optimum  
127 trait values. At stationarity, an OU process leads to stable trait mean and variance, setting the  
128 overall phenotypic constraint. Fitness tradeoffs likely limit the breadth of values for adaptive trait  
129 optima, given that extreme trait values will rarely optimize competing functions (33). Notably,  
130 the specific mechanisms for constraints on trait values are not specified but are implicit in the  
131 application of Ornstein-Uhlenbeck (OU) process to model evolution phenomenologically.

132           Given that the stomatal-area minimization and increase hypotheses differ in their  
133 prediction of how the species variation in composite traits ( $f_S$  and  $g_{\max}$ ) are constrained by their  
134 constituent traits ( $D_S$  and  $A_S$ ), examination of the trait evolution can indicate which hypothesis  
135 was supported. The OU model can indicate *which* composite trait,  $f_S$  or  $g_{\max}$ , is primarily  
136 constrained. In both cases, analogous quantitative theory shows that constraint on composite  
137 traits imposed by stabilizing selection limits variation in constituent traits(34), and constraint on  
138  $f_S$  results in a different covariance structure of  $D_S$  versus  $A_S$  than a primary constraint on  $g_{\max}$ .  
139 Note that both  $f_S$  and  $g_{\max}$  show similar mathematical dependence on  $D_S$  and  $A_S$ :

$$Z_S = \lambda D_S A_S^\beta, \quad (3)$$

140 where composite stomatal trait  $Z_S$  (i.e.,  $f_S$  or  $g_{\max}$ ) is proportional to the product of constituent  
141 stomatal traits, with scaling exponent  $\beta$  multiplied by a scalar  $\lambda$ , which reflects stomatal  
142 dimension proportionalities and physical diffusion factors (22). For  $g_{\max}$ ,  $\lambda = bm$  and  $\beta = 0.5$   
143 (Eq. 1); for  $f_S$ ,  $\lambda = 1$  and  $\beta = 1$  (Eq. 2). Since all traits are log-normally distributed<sup>31</sup>, and the  
144 OU model assumes Gaussian traits, we log-transformed Eq. 3:

$$z_S = \log(\lambda) + d_S + \beta a_S, \quad (4)$$

145 where lowercase variables indicate log-transformation of uppercase counterparts. Log-  
146 transformation also has the advantage of simplifying variance decomposition by linearizing the  
147 equation and enables traits measured on different scales to be directly compared in their  
148 proportional changes. Using random variable algebra, the variance in  $z_S$  is defined as:

$$\text{Var}(z_S) = \text{Var}(d_S) + \beta^2 \text{Var}(a_S) + 2\beta \text{Cov}(d_S, a_S). \quad (5)$$

149 Using the variance-covariance of  $d_S$  and  $a_S$ , we can find the scaling exponent  $\beta$  that minimizes  
150  $\text{Var}(z_S)$ :

$$\beta = -\frac{\text{Cov}(d_S, a_S)}{\text{Var}(d_S)} \quad (6)$$

151 Notably, the right-hand side of Eq. 6 is the negative of the ordinary linear regression slope of  
152 log-stomatal size against log-density. Thus, for any dataset,  $\beta$  can be estimated using ordinary  
153 regression methods, but a negative slope estimate will result in a positive value of  $\hat{\beta}$ . The  
154 stomatal-area minimization hypothesis predicts that  $\hat{\beta} = 1$  because  $f_S$  constrains  $d_S$  and  $a_S$  (Eq.  
155 1), whereas the stomatal-area increase hypothesis predicts that  $\hat{\beta} = 0.5$  because  $g_{\max}$  constrains  
156  $d_S$  and  $a_S$  (Eq. 2). Note that the above prediction assumes that the primary constrained composite  
157 trait will also be the least variable composite trait, which allowed to identify the relationship  
158 between  $\beta$  and trait (co)variance in Eq. 6. We evaluated this assumption using forward-time,  
159 individual based, macroevolutionary quantitative genetic simulations (Supplementary  
160 Information). In each simulation, 1000 independent lineages evolve toward a moving optimal  
161 composite trait until stationarity following an OU process. The simulations confirm that the  
162 constrained composite trait is the least variable and that ordinary regression on interspecific trait  
163 means can accurately identify the simulated  $\beta$ . Estimates of  $\beta$  are not substantially affected by  
164 microevolutionary details about mutational and genetic covariances or geometric constraints on  
165  $f_S$  (Fig. S2-S5).

166 We estimated stomatal size-density scaling in 2408 forest plant species from new field-  
167 collected samples over 28 sites in China and global synthesis of data from the literature (Fig. 2)  
168 and estimated the scaling exponent  $\beta$  using OU phylogenetic multiple regression with group  
169 (Angiosperm, Pteridophyte, Gymnosperm) and growth form (tree, shrub, herb) as covariates (see  
170 Methods).

171 Stomatal size-density scaling among forest plant species was consistent with a primary  
172 constraint on  $g_{\max}$  (stomatal-area increase hypothesis,  $\beta = 0.5$ ). Given the variance in stomatal



173 density, the covariance between size and density among forest species minimizes the variance in  
174  $g_{\max}$ . This implies that selection for higher  $g_{\max}$  results in increased stomatal area allocation, and  
175 not minimizing area allocation (Fig. 3). There is no evidence that scaling differs between major  
176 groups, Angiosperms, Gymnosperms, and Pteridophytes (Fig. 3a; Table S1), but  $g_{\max}$  is 49% (17-  
177 88% 95% CI;  $P = 0.001$ ) and 14% (1-30% 95% CI;  $P = 0.04$ ) higher in Angiosperms than  
178 Gymnosperms and Pteridophytes, respectively (Table S2). Trees also have 18% (8-28% 95% CI;  
179  $P < 0.0001$ ) and 48% (39-59% 95% CI;  $P < 0.0001$ ) greater  $g_{\max}$  than shrubs and herbs,  
180 respectively (Table S2). The across-species mean and variance in  $\log(g_{\max})$  are nearly invariant  
181 across latitude, temperature, and precipitation gradients, indicating that most of the variation in  
182  $g_{\max}$  occurs for species of contrasting ecology within rather than between forest sites, a finding  
183 similar to that for other key functional traits such as leaf mass per area and wood density (35)  
184 (Fig. 4).

185 Our results overturn the prevailing view that the evolution of high  $g_{\max}$  across diverse  
186 species is constrained by size-density scaling and minimized stomatal area allocation. Instead,  
187 the covariance between stomatal size and density supports stomatal area allocation increasing  
188 with the evolution of high  $g_{\max}$ . Thus, limits on the fraction of epidermis allocated to stomatal  
189 ( $f_s$ ) are a secondary consequence of limits on  $g_{\max}$ . Our novel analysis developed from  
190 quantitative genetic and macroevolutionary theory could distinguish the  $g_{\max}$  evolution  
191 hypotheses. Notably, our  $\beta$  exponent for the scaling of  $d_s$  and  $a_s$  depends on using  
192 (phylogenetic) least squares regression, and thus, the results of studies reporting stomatal scaling  
193 slopes using standardized major axis (SMA) regression (which minimizes residual variance in  
194 both  $d_s$  and  $a_s$ ) would need to be recalculated to test against our findings (see Supplementary  
195 Information). Although estimated scaling using standard phylogenetic regression approaches (see

196 Methods), it is more appropriate to interpret our results not as minimizing residual variance, but  
197 rather estimating the  $\beta$  consistent with the covariance structure of stomatal size and density (Fig.  
198 1).

199 Our results have at least two important implications for understanding the evolutionary  
200 anatomical mechanisms of high  $g_{\max}$  and its consequences for the stomatal size-density scaling  
201 relationship. First, the finding that size-density scaling does not constrain the evolution of higher  
202  $g_{\max}$  implies that stomatal cost is not a constraint on high  $g_{\max}$  and thus a different constraint on  
203 the evolution of extreme values of  $g_{\max}$  across environments. Very high  $g_{\max}$  may be rare because  
204 the  $g_{\text{op}}:g_{\max}$  ratio is constrained in a region of maximal control to respond rapidly to changing  
205 environments (36). Additionally or alternatively, a high  $g_{\max}$  may also be linked with a high  
206 wilting point thereby setting a physical upper limit to leaf gas exchange and a high risk of  
207 hydraulic failure (37) if open stomata face transiently high atmospheric drought. Other possible  
208 costs include detrimental consequences of high  $g_{\max}$  for stomatal movements and diffusion, as  
209 well as energetic costs of opening closing more and/or larger stomata (38, 39). Future work  
210 should prioritize identifying the fitness costs and functional trade-offs that constrain the  
211 evolution of high  $g_{\max}$ . Second, if  $g_{\max}$  is the primary constraint, this implies that space allocation  
212 to stomata is relatively unimportant, such that plants could allocate a greater fraction of their  
213 epidermal area to stomata than they currently do without countervailing selection. Thus, if  
214 stomatal size and density can be manipulated independently, anatomies with the same  $g_{\max}$ , but  
215 different  $f_s$ , would have similar fitness in the same environment. This finding also clarifies the  
216 evolution of stomata across major plant lineages, and refutes the hypothesis that smaller stomata  
217 were required to increase  $g_{\max}$  in angiosperms (20). All three major land plant lineages have  
218 similar variance in  $g_{\max}$  (Fig. 3b); angiosperms have higher  $g_{\max}$  than gymnosperms and

219 pteridophytes on average due to their higher  $d_s$  for a given  $a_s$ , not because of differences in the  
220 scaling relationship. The higher stomatal density of angiosperms would be linked to increases in  
221 leaf water transport capacity, for example, by decreasing the distance between vein and stomata,  
222 allowing stomata to stay open<sup>40</sup>. The primary constraint on maximum stomatal conductance  
223 appears to be that selection rarely favors extreme values, implying that vegetation and crop  
224 models should incorporate nonepidermal costs of extreme trait values to predict optimal  $g_{\max}$  for  
225 the prediction of photosynthetic carbon gain and transpiratory water loss across scales.

## 226 **Methods**

### 227 *Stomatal trait data from global forests*

228 The stomatal dataset of global forests represents a total of 2408 plant species from natural  
229 forests, including novel field data collected from Chinese forest communities and a compilation  
230 of published literature values.

231 Our field data were collected from 28 typical forest communities occurring between 18.7 °N and  
232 53.3 °N latitude in China. The field sites were selected to cover most of the forest types in the  
233 northern hemisphere, including cold-temperate coniferous forest, temperate deciduous forest,  
234 subtropical evergreen forest, and tropical rain forest (Fig. 2). In total, we sampled 28 forest sites.  
235 We used the Worldclim database (40) to extract additional data on mean annual temperature  
236 (MAT) and precipitation (MAP) over the period 1960-1990 using latitude and longitude. Among  
237 these forests, mean annual temperature (MAT) ranged from -5.5-23.2 °C, and mean annual  
238 precipitation (MAP) varied from 320 to 2266 mm. The field investigation was conducted in July-  
239 August, during the peak period of growth for forests. Sampling plots were located within well-  
240 protected national nature reserves or long-term monitoring plots of field ecological stations, with  
241 relatively continuous vegetation. Four experimental plots (30 × 40 m) were established in each  
242 forest.

243 Leaves from trees, shrubs, and herbs were collected within and around each plot. For trees,  
244 mature leaves were collected from the top of the canopy in four healthy trees and mixed as a  
245 composite sample. Eight to 10 leaves from the pooled samples were cut into roughly 1.0 × 0.5  
246 cm pieces along the main vein, and were fixed in formalin-aceto-alcohol (FAA) solution (5 ml  
247 38 % formalin, 90 ml 75 % ethanol, 5 ml 100 % glacial acetic acid, and 5 ml 37 % methanol)

248 (41). In the laboratory, three small pieces were randomly sampled, and each replicate was  
249 photographed twice using a scanning electron microscopy (Hitachi SN-3400, Hitachi, Tokyo,  
250 Japan) on the lower surface at different positions. We focused on the lower epidermis (42),  
251 because a previous study has demonstrated that most of leaf upper epidermis has no stomata for  
252 forest plants (43).

253 In each photograph, the number of stomata was recorded, and  $D_S$  was calculated as the  
254 number of stomata per unit leaf area. Simultaneously, five typical stomata were selected to  
255 measure stomatal size using an electronic image analysis equipment (MIPS software, Optical  
256 Instrument Co. Ltd., Chongqing, China).

257 Peer-reviewed papers on leaf stomata were collected using an all-databases search of Web  
258 of Science ([www.webofknowledge.com](http://www.webofknowledge.com)) from 1900 to 2018 using “forest” and “stomata” as a  
259 topic, in line with the principle of “natural forest, non-intervention, species name” (i.e. we did  
260 not use data from controlled experiments or where taxonomic data was missing). A total of 90  
261 papers (see Supporting Table S3) which met our requirements, yielding  $D_S$  and  $L$  measurements  
262 from 413 plant species (Fig. 2) from which we calculated  $g_{\max}$  and  $f_S$ .  $f_S$  is proportional to the  
263 stomatal pore area index (SPI), which defined as the product of  $D_S$  and stomatal length ( $L$ )  
264 squared (25), because  $A_S = mL^2$  (22).

265 We calculated  $g_{\max}$  (Equation 1) to water vapor at a reference leaf temperature ( $T_{\text{leaf}} = 25^\circ$   
266 C) following Sack and Buckley (22). They defined a biophysical and morphological constant as:

$$b = \frac{D_{\text{wv}}}{v}$$
$$m = \frac{\pi c^2}{j^{0.5}(4hj + \pi c)}$$

267  $b$  is the diffusion coefficient of water vapor in air ( $D_{wv}$ ) divided by the kinematic viscosity of dry  
268 air ( $\nu$ ).  $D_{wv} = 2.49 \times 10^{-5} \text{ m}^2 \text{ s}^{-1}$  and  $\nu = 2.24 \times 10^{-2} \text{ m}^3 \text{ mol}^{-1}$  at  $25^\circ$  (44). For kidney-  
269 shaped guard cells,  $c = h = j = 0.5$ ; for dumbbell-shaped guard cells in the Poaceae,  $c = h =$   
270  $0.5$  and  $j = 0.125$ . We used the species average  $g_{\max}$  and  $f_s$  for all analyses.

### 271 *Phylogenetic regression*

272 By positing that the least variable composite of stomatal size and density indicates the trait  
273 with the most constraint (Fig. 1), we identify a new way to estimate the scaling exponent  $\beta$  (Eq.  
274 6) using linear regression estimates, and also accounted for phylogenetic nonindependence. We  
275 used the Plant List (<http://www.theplantlist.org>) to confirm species names, then we assembled a  
276 synthetic phylogeny using S.PhyloMaker (45). We fitted phylogenetic regression models using  
277 the **phylolm** version 2.6 package in R (46). As we derived in the main text, the scaling exponent  
278  $\beta$  can be estimated from the slope of the regression of  $a_s$  on  $d_s$ , where  $\hat{\beta} = -\text{slope}$ . We estimated  
279 separate scaling exponents for major groups, Angiosperms, Pteridophytes, and Gymnosperms.  
280 We also estimated different intercepts, corresponding with different average  $g_{\max}$  values, for  
281 functional types (herbs, shrubs, and trees) and grasses, because of their unique stomatal anatomy.  
282 We used the “OUrandomRoot” model of trait evolution. 95% confidence intervals for all  
283 parameters were estimated from 1000 parametric bootstrap samples generated by simulating  
284 from the best-fit model and re-fitting.  $P$ -values for coefficients are based on  $t$ -tests. We used the  
285 same methods to test whether  $g_{\max}$  (log-transformed for homoskedasticity) was affected by  
286  $|\text{latitude}|$ , MAP, MAT, group (Angiosperms, Pteridophytes, Gymnosperms), and/or functional  
287 type (herb, shrub, tree). One gymnosperm species, *Torreya fargesii*, had substantially lower  
288 stomatal size than would be predicted from its density (Fig. 3a). These results of the paper did  
289 not change if this outlier was excluded because the confidence intervals for stomatal-density

290 scaling are very wide for Gymnosperms regardless. Therefore, we excluded this species from

291 statistical analyses but show it in the figure for completeness. All data were analyzed in R

292 (47)version 4.0.5

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435 **Author contributions**

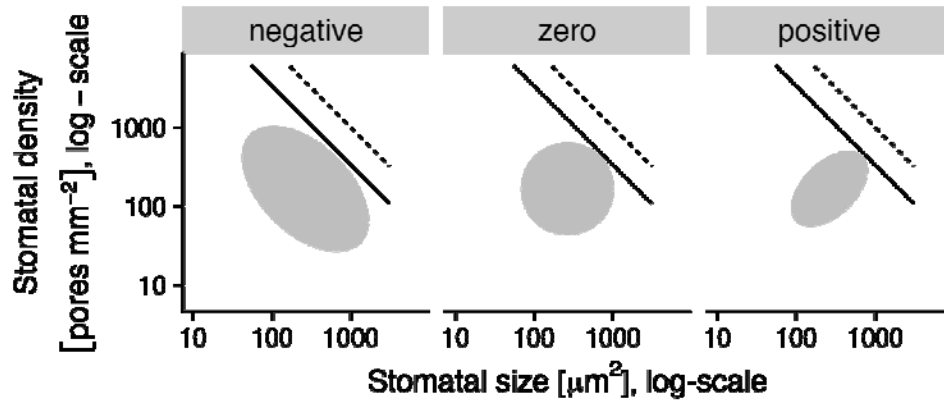
436 N.H. and G.Y. designed field sampling; N.H., C.D.M, and C.L. conceived the initial ideas; C.L.,  
437 N.H., Y.L. J.Z., Z.Z., M.L. and L.X collected the data; C.L. wrote the first draft, and C.D.M.  
438 contributed the final mathematical derivations, data analysis, and wrote the final manuscript; L.S.,  
439 H.J.B., C.L., N.H., G.Y., and X.H. revised the manuscript. All authors gave final approval for  
440 publication.

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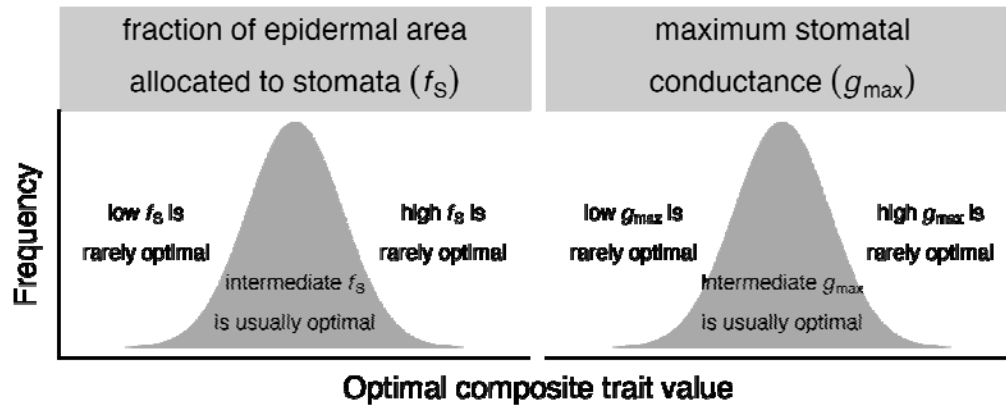
442 **Supplementary information**

443 **Figures**

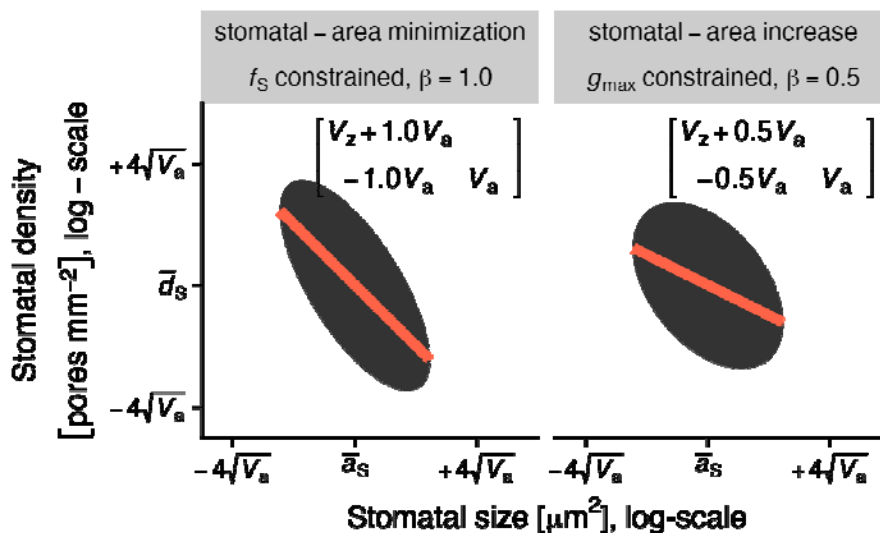
**a Many relationships between stomatal size and density are geometrically possible**



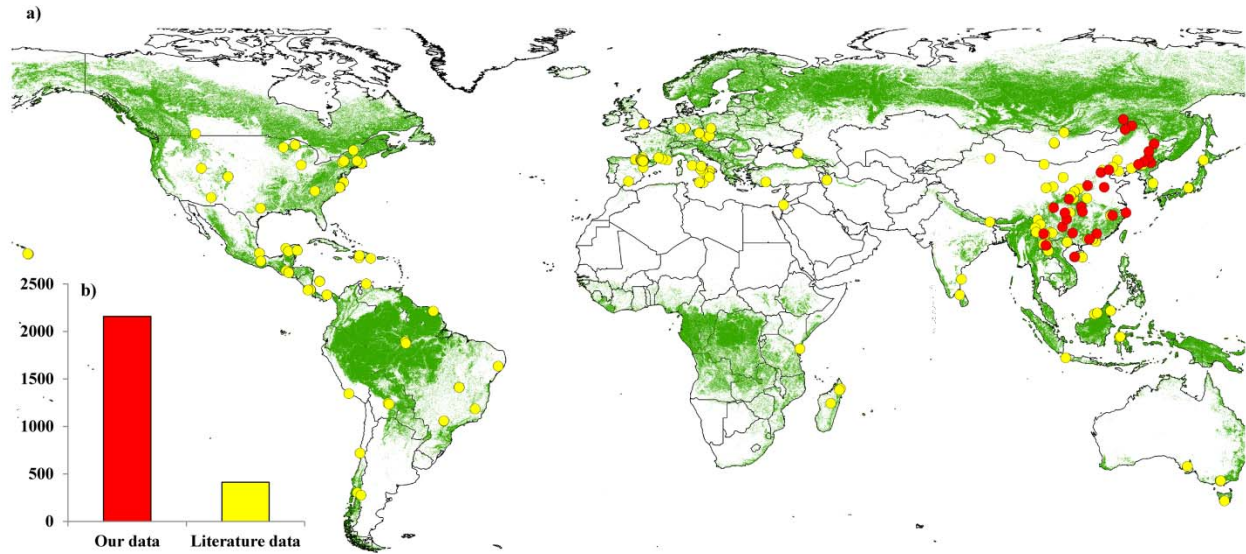
**b Hypothesis: covariation in stomatal size and density is shaped by constraint on a composite trait**



**c Prediction: stomatal size-density covariation depends on which composite trait is constrained**



445 **Fig. 1 | Competing hypotheses for stomatal size-density scaling make different predictions**  
446 **about the trait covariance structure.** Maximum stomatal conductance ( $g_{\max}$ ) and the fraction of  
447 epidermal area allocated to stomata ( $f_s$ ) are composite traits determined by stomatal density and  
448 size. On a log-scale, they are the sum of log-stomatal density ( $d_s$ ) and log-stomatal size ( $a_s$ )  
449 times a scaling exponent ( $\beta$ ), 0.5 for  $g_{\max}$  and 1.0 for  $f_s$  (see Methods). **a.** Many scaling  
450 relationships between stomatal size and density are possible as long as  $f_s$  does not exceed 1  
451 (dashed line) or more realistically a value less than 1 to allow space between stomata (solid line,  
452  $f_s = 0.34$ , the maximum value in our data set). The grey ellipses represent different possible  
453 scaling relationships with the same mean trait values in our data set ( $\bar{A}_s = 263 \mu\text{m}^2$ ,  $\bar{D}_s =$   
454  $168 \text{ pores mm}^{-2}$ ). These are 95% quantile of covariance ellipses for a bivariate normal with  
455 trait correlations of -0.5, 0, and 0.5 and trait variances of 0.75, 0.55, and 0.45 for ‘negative’,  
456 ‘zero’, and ‘positive’ relationships, respectively. **b.** We hypothesized that size-density scaling is  
457 determined by constraint on either  $g_{\max}$  (stomatal-area increase; left panel) or  $f_s$  (stomatal-area  
458 minimization; right panel). Under either hypothesis, the optimal composite trait varies but  
459 extreme values of the composite trait are rarely optimal. **c.** Both hypotheses predict negative  
460 size-density scaling but with different covariance relationships. If the interspecific means ( $\bar{d}_s, \bar{a}_s$ )  
461 and variances ( $V_d, V_a$ ) of stomatal density and size, respectively, are measured, the covariance  
462 between them ( $V_{d,a}$ ) is equal to  $-\beta V_a$ . Under the stomatal-area increase (left panel) and stomatal-  
463 area minimization (right panel) hypotheses,  $\beta$  should be 0.5 and 1, respectively. The ellipse is the  
464 0.95 quantile of covariance ellipse associated with the covariance matrix (upper right corner of  
465 the plot); the orange line is the scaling exponent fit through the constituent trait means.



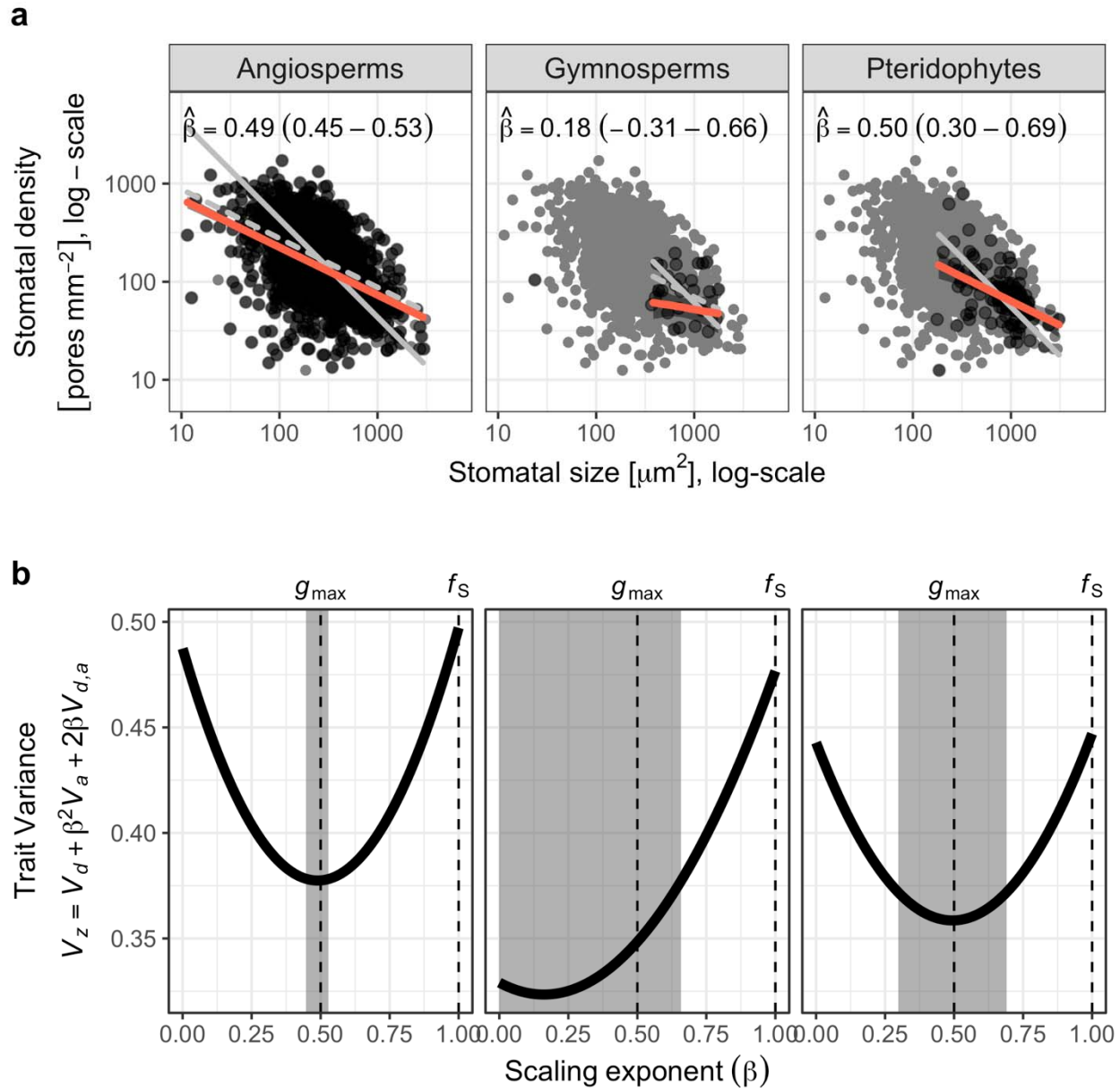
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467

468 **Fig. 2 | Geographic distribution of sampling sites (a) and the number of plant species (b) in**

469 **this study.**





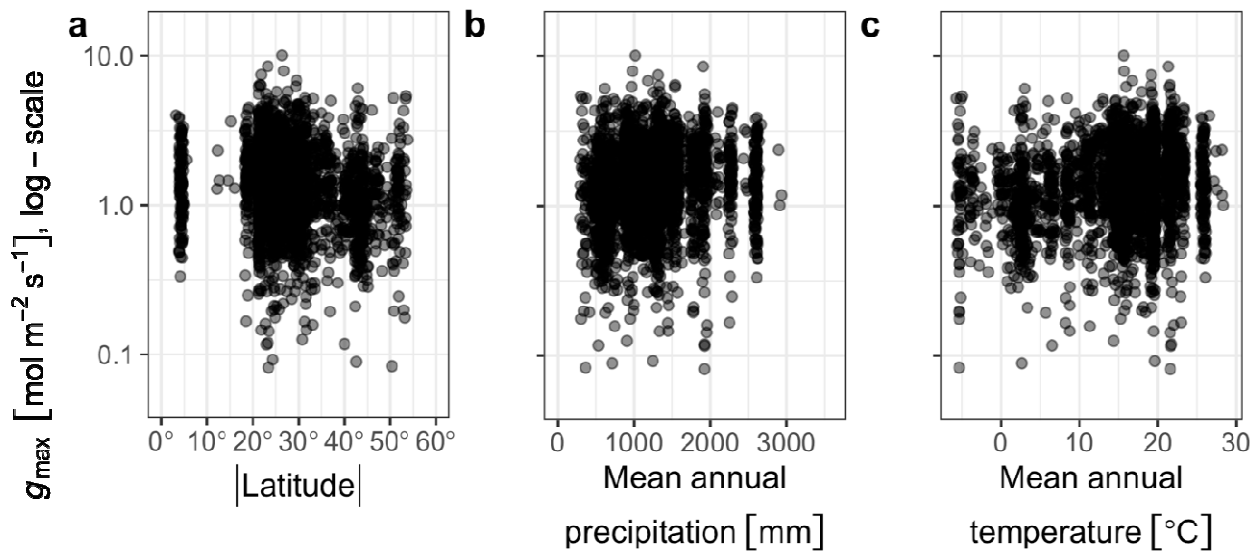
470

471 **Fig. 3 | Stomatal size-density scaling is consistent with stomata-area increase but not area-**  
 472 **minimization. a.** In both angiosperms (left panel) and pteridophytes (right panel), the scaling  
 473 exponent ( $\hat{\beta}$ ) estimated as the phylogenetic linear regression slope of stomatal size against  
 474 density (Methods) is close to 0.5 as predicted by the stomatal-area increase hypothesis, but much  
 475 less than 1.0, as predicted by the stomatal-area minimization hypothesis. For comparison, thin  
 476 gray lines in the background show predicted slopes for each group when  $\beta = 1.0$  (solid line) and



477  $\beta = 0.5$  (dashed line). The bootstrap 95% confidence intervals are in parentheses and shown  
478 graphically by the width of the grey rectangle in **b**. Dark points represent species mean trait  
479 values from the focal group; grey background points are from all groups for comparison. Orange  
480 line and ribbon are the estimated phylogenetic regression line and the 95% bootstrap confidence  
481 intervals. Scaling in gymnosperms (middle panel) is not significantly different from 0 or 0.5, but  
482 the confidence intervals do not include 1.0. **b**. The variance of the composite trait ( $V_z$ ) is  
483 minimized near  $\beta = 0.5$ , as predicted under the stomatal-area increase hypothesis (dashed-line  
484 under  $g_{\max}$ ) but not where  $\beta = 1.0$  as predicted by the stomatal-area minimization hypothesis  
485 (dashed-line under  $f_s$ ).

486



487

488 **Fig. 4 | Anatomical maximum stomatal conductance varies little with latitude, mean annual**

489 **precipitation, or mean annual temperature.** Each point is the species' mean  $|\text{latitude}|$  (a.),

490 mean annual precipitation (b.), or mean annual temperature (c.) on the  $x$ -axis and the maximum

491 stomatal conductance ( $g_{\max}$ ) on the  $y$ -axis (log-scale). Based on phylogenetic multiple regression,

492 the relationship between  $\log(g_{\max})$  and mean  $|\text{latitude}|$  ( $P = 0.69$ ) and mean annual temperature ( $P =$

493  $0.10$ ) are not significant; the relationship with mean annual precipitation is significant ( $P =$

494  $0.009$ ) but weak since the total model  $R^2$  including all climate, lineage, and growth explanatory  
495 variables is only 0.11.

496

497