The effect of recent habitat change on genetic diversity at putatively

adaptive and neutral loci in *Primula veris* in semi-natural grasslands

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Running head: Plant genetic response to habitat change

Abstract

Recent habitat change in semi-natural grasslands due to a lack of management has been

shown to affect the genetic diversity of grassland plants. However, it is unknown how a

change in local environment affects genetic diversity at adaptive loci. We applied RADseq

(restriction-site associated DNA sequencing) to extract > 3,000 SNPs across 568 individuals

from 32 Estonian populations of *Primula veris*, a plant species common to semi-natural

grasslands. We evaluated the effect of recent grassland overgrowth following management

abandonment on the genetic diversity at putatively neutral and adaptive loci, which we

distinguished by applying three methods, i.e., linear and categorical environmental

association analyses, and an F_{ST} outlier test. For validation, we randomised the genotype to

sample assignments. Effects of recent habitat change on genetic diversity differed between

neutral and adaptive SNP sets. Genetic diversity assessed at putatively neutral loci was

similar in open and overgrown habitats but showed a significant difference between these

habitat types at putatively adaptive loci: overgrown (i.e. newly established) habitats

exhibited higher genetic diversity at putatively adaptive loci than open (i.e. old) habitats,

likely due to the exertion of novel selection pressures imposed by new habitat conditions.

This increase in genetic diversity at putatively adaptive loci in the new environment points

to currently ongoing selection processes where genetic adaptation to the old habitat is

potentially lost through altered allele frequencies. Our study suggests that a recent change

in local habitat conditions may not be reflected in neutral loci whereas putatively adaptive

loci can inform about potential selection processes.

Key words: environmental association analysis, genetic differentiation, habitat degradation,

permutation, semi-natural grasslands

Introduction

Habitat change is a consequence of ongoing anthropogenic landscape and climatic change (IPBES, 2018; IPCC, 2019). In particular, intensification of land use and abandonment of traditional management practices during the last century led to a dramatic degradation and isolation of habitats such as European semi-natural grasslands (Cousins, Auffret, Lindgren, & Tränk, 2015; Hooftman & Bullock, 2012). Historically, moderate and continuous human management (e.g. grazing by livestock) has supported development of open semi-natural grassland habitats with increased light availability and elevated niche partitioning, facilitating uniquely high levels of biodiversity (Habel et al., 2013; Wilson, Peet, Dengler, & Pärtel, 2012). Yet, conversion to other land use types or abandonment of traditional management have gradually caused substantial loss in grassland area and increased fragmentation, i.e. due to overgrowth with dense woody vegetation. This has resulted in negative effects on grassland biodiversity (Habel et al., 2013; Picó & Van Groenendael, 2007).

Likewise, the intra-specific genetic diversity of grassland plants suffers from the degradation, fragmentation, and loss of semi-natural grasslands. Due to potentially reduced population sizes and increased landscape barriers, the genetic diversity of many insect-pollinated grassland plant species is compromised by interrupted pollen-mediated gene flow (e.g. DiLeo, Holderegger, & Wagner, 2018; Tewksbury et al., 2002). This likely aggravated gene flow contributes to potentially decreased genetic diversity and increased genetic differentiation in plant populations in affected grasslands (e.g. Honnay & Jacquemyn, 2007; Picó & Van Groenendael, 2007). However, plants might exhibit a delayed response to habitat change (e.g. Aavik et al., 2019; Helm, Hanski, & Pärtel, 2006; Lehtilä et

al., 2016). For instance, life history traits, such as mating system and lifespan, determine the

speed and magnitude of a plant's response to a changed environment (Hamrick & Godt,

1996; Leimu, Mutikainen, Koricheva, & Fischer, 2006). With many grassland plants having a

relatively long lifespan of up to several decades (Ehrlén & Lehtilä, 2002), such longevity

might mask genetic effects induced by habitat change.

Genetic diversity is one of the central parameters in estimating a population's adaptive

potential (Bilska & Szczecińska, 2016). However, because environmental change does not

leave a signature in all parts of the genome (Nei, Suzuki, & Nozawa, 2010), such estimation

requires differentiating between genetic diversity assessed at putatively neutral and

adaptive loci (Bilska & Szczecińska, 2016). Here, we refer to putatively neutral loci when loci

are affected by neutral processes such as gene flow but not by specific environmental

factors. In contrast, putatively adaptive loci explicitly show a response to tested

environmental factors. Well-adapted populations might show a reduced diversity at

adaptive loci due to beneficial mutations going towards fixation, and it is the adaptive

potential at the population level that can be assessed with investigations of genetic diversity

at adaptive loci (Milot et al., 2020). However, the difference in response of genetic diversity

at putatively neutral and adaptive loci to habitat change has been mostly ignored so far, in

particular in the context of recent land use change. The only existing studies that explicitly

account for differences in neutral and adaptive loci in plants concentrated on climatic

factors (Dauphin et al., 2020; Sun et al., 2020). Moreover, in conservation genetics, most

studies focussed on overall or genetic diversity at neutral loci, often using a set of neutral

microsatellite markers, while ignoring adaptive regions of the genome (González et al.,

2019; Wei & Jiang, 2020). However, it is exactly the adaptive regions that are important for

the fate of a population.

With gradual grassland overgrowth, plant populations experience changed

environmental conditions with novel selection pressures such as lower light availability,

changes in soil chemical conditions, and altered pollinator communities (Helm, 2019),

demanding for phenotypic plasticity of individuals or adaptation to the new habitat

conditions. The adaptive potential of a population can be nourished from three sources:

standing genetic variation (Barrett & Schluter, 2008; Radwan & Babik, 2012), gene flow

(Slatkin, 1985) or, in the longer term, spontaneous mutations. Newly introduced barriers to

gene flow should increase the importance of a population's standing genetic variation, and

novel selection pressures induced by habitat change likely trigger a decrease in adaptation

to the former grassland habitat. Thus, habitat change forces plants to either adapt to locally

new habitat conditions predominantly based on their standing genetic variation in an

aggravated gene flow scenario, to disperse to more favourable habitats, or to face local

extinction (e.g. Cheptou, Hargreaves, Bonte, & Jacquemyn, 2017; Frankham, 2005).

In the present study, we were interested in the effect of recent overgrowth of Estonian

semi-natural grasslands with woody vegetation over the past century on the genetic

diversity at putatively neutral and adaptive loci in *Primula veris* populations, a long-lived

grassland specialist plant. Our study is one of the first to test for a land use change effect on

both putatively neutral and adaptive loci in *in-situ* wild plant populations. We applied

double-digest restriction-site associated DNA sequencing (ddRADseq) in 32 populations of *P.*

veris from open and recently overgrown grasslands. We distinguished between putatively

neutral and adaptive loci by performing a combination of environmental association

analyses and F_{ST} outlier tests. We specifically asked whether (1) genetic diversity at neutral

loci of *P. veris* populations is negatively affected by overgrowth of grasslands due to

potentially reduced population size and/or aggravated gene flow; (2) genetic diversity at

adaptive loci exhibits a different response to habitat change than genetic diversity at neutral

loci; and (3) genetic diversity at adaptive loci is actually increased due to ongoing selection

processes where genetic adaptations to the old, open habitat are slowly lost through altered

allele frequencies of the beneficial alleles for open and overgrown habitats.

Materials and Methods

Study species

Primula veris L. (Primulaceae) is an herbaceous perennial rosette-forming hemicryptophyte

most commonly occurring in calcareous grasslands. *Primula veris* prefers open habitats but

can grow under shade with reduced reproduction (Brys & Jacquemyn, 2009). Its average life

span reaches up to 50 years (Ehrlén & Lehtilä, 2002). In Estonia, the study region, *P. veris*

generally flowers in May. The study species is an obligate outbreeder that depends on

insect-pollination (Deschepper, Brys, & Jacquemyn, 2018). Pollen dispersal is spatially

restricted to several meters (Brys & Jacquemyn, 2009). Self-pollination is prevented by

heterostyly with two flower morphs, with low levels of successful intra-morph pollination

(Wedderburn & Richards, 1990). Primary seed dispersal is limited to a few metres from the

maternal plant (Brys & Jacquemyn, 2009).

Study sites and sampling

Study sites were located on dry calcareous grasslands, alvars, on the islands of Muhu and

Saaremaa in Western Estonia (Figure 1). Alvars are semi-natural grasslands on Ordovician

and Silurian bedrock with only a low soil depth (< 20 cm). Management, i.e. grazing

livestock, in the area was abandoned 20 – 90 years ago. Our study sites were part of a large-

scale biodiversity inventory of the European Commission's LIFE+ Nature program restoration

project "LIFE to Alvars" (Helm, 2019), which included monitoring of genetic diversity of

grassland plant species, including *P. veris*, in alvars at different successional stages of

overgrowth (e.g. still open and recently overgrown). The mean temperature in the area is

17°C in summer and -3°C in winter, and the mean annual precipitation is about 680 mm

(EWS, 2020).

We sampled 32 populations (i.e. spatially distinct patches) of *P. veris* distributed across

two regions, Muhu and Saaremaa islands, in the summers of 2015 and 2016 (Figure 1; Table

1). Where possible, we chose pairs of closely located populations (i.e. within pollen- and

seed-mediated gene flow distance) of contrasting habitat types (i.e. open and recently

overgrown grasslands). Nineteen populations were located in open grasslands (i.e. old

habitat; hereafter open habitats) and 13 populations were located in shrubby-overgrown

grasslands (i.e. new habitat; hereafter overgrown habitats), comprising 10 population pairs

with an average distance of 533 m and a minimum distance of 20 m between members of

pairs. Such a paired sampling design has been shown to be efficient in detecting genomic

signatures of local adaptation in environmental association analysis (EAA; Lotterhos &

Whitlock, 2015) and allows the use of categorical EAA approaches (see below). Overgrown

habitats represented mid-successional stages with at least 60% cover of shrubby vegetation,

mostly *Juniperus communis*.

Within each population, we sampled three fresh leaves of 20 random flowering *P. veris*

individuals (where possible) that were at least 50 cm apart. Leaves were stored in silica-gel

until further processing. Approximate population census sizes of *P. veris* were estimated by

assessing the number of both flowering and non-flowering individuals per population.

    Environmental data

To characterize the environment of the study sites, we used data collected within the frame

of the "LIFE to alvars" project (Helm, 2019). We selected 16 *in-situ* measured environmental

variables with the potential to represent contrasting habitat types (open and overgrown),

i.e. "openness", "soil", and "biota". For openness, we considered the total percentual shrub

and tree coverage, respectively, assessed within a radius of 10 m from the center of *P. veris*

populations, and the light availability above and below the herbal layer measured with a Li-

Cor LI-250 Light Meter and a LI-190SA Quantum Sensor (Lincoln, Nebraska, USA), in 1x1 m in

the center of the population. For soil, we considered average soil depth in cm based on ten

random locations taken within a 10 m radius around the central point of *P. veris*

populations. In the same radius, five soil samples were taken from random locations and

pooled for chemical analyses. From each sample, soil pH (KCl solution), available soil

phosphorus (P; extraction with acid ammonium lactate solution), potassium (K), magnesium

(Mg), calcium (Ca), and soil organic content (OC, loss on ignition) were measured. For biota,

we considered butterfly and bumblebee abundance and richness and vascular plant richness

within a 10 m radius around the central points of *P. veris* populations. Butterflies and

bumblebees were monitored using standardised transect counts (Pollard, 1977). Each site

was visited three times for butterflies and two times for bumblebees over two years to

cover phenological aspects of different species (Helm, 2019).

In addition to the 16 *in-situ* measured environmental variables, we extracted climate

data for each population from CHELSA (Karger et al., 2017) with a resolution of 30 arc sec

from the reference period 1979-2013. To increase resolution, we applied a bi-linear

interpolation that accounts for the climate values in surrounding grid cells and the position

of the population within the grid cell. For our analyses, we used temperature (Bio1, annual

mean temperature) and precipitation (Bio12, annual precipitation sum), because they

represent the most comprehensive bioclimatic variables describing the climate in our study

region.

To test which environmental factors significantly differed between the two habitat

types, we performed a (non-paired) two-sample t-test for each of the 18 environmental

variables.

DNA extraction and ddRAD sequencing

Twenty-five mg of leaf material were pulverized with 2.3-mm chrome-steel beads (BioSpec

Products, Bartlesville, USA) in a Mixer Mill 301 (Retsch, Haan, Germany). DNA was extracted

using the LGC sbeadex plant kit (LGC, Berlin, Germany). 400 µl lysis buffer, consisting of 1%

RNAse (100 mg/ml), 0.2% Proteinase K solution (20 mg/ml), and lysis buffer PN were added

to pulverized samples, with an incubation time of 1 h at 65°C on a Thermomixer comfort

(Eppendorf, Hamburg, Germany) at 300 rpm, and followed by a centrifugation at 2500 x g

for 10 min. Lysates were transferred to binding solution, consisting of 420 µl binding buffer

PN and 10 µl sbeadex particle solution. All following steps were conducted on a KingFisher

Flex Purification System (Thermo Fisher Scientific, Waltham, USA), with the specification of

using 400 µl of wash buffer PN1 twice, 400 µl of wash buffer PN2, and eluting purified DNA

in 50 µl elution buffer AMP.

We applied a ddRADseq procedure by customizing an existing ddRADseq protocol

(Westergaard et al., 2019). ddRADseq applies a double restriction enzyme digest followed

by a size-selection of genomic fragments (Peterson, Weber, Kay, Fisher, & Hoekstra, 2012).

RADseq provides a simple and cost-effective method to uncover thousands of polymorphic

markers, both neutral and adaptive, in model and non-model organisms (e.g. Davey et al.,

2011). Because the aim of our study was to identify general patterns of genetic diversity

assessed at neutral and adaptive loci across a high number of samples from many

populations, rather than identifying specific genes involved in adaptation, we chose not to

use whole genome or targeted sequencing. Such methods might be, however, worthwhile

to consider in future analyses following the results of the present study.

For the detailed ddRADseq protocol see supplemental information (Supplemental

Methods). Briefly, standardized DNAs of fully randomized samples were digested and

purified before ligation to a combination of one of 48 EcoRI and 2 TaqI adapters,

respectively, resulting in uniquely tagged barcoded DNA samples. DNA samples with the

same TaqI adapter but different EcoRI adapters (48 samples) were pooled together and size

selected for fragments of 450 bp length. The size-selected sample pools were selected for

fragments containing biotin labelled TaqI adapters. Subsequently, polymerase chain

reaction (PCR) was conducted, PCR products (ddRADseq libraries) were purified and their

DNA concentration was measured to calculate molarity per ddRADseq library. Finally,

samples with distinct TaqI multiplexing indices were combined to produce a final library of

at least 5 nM consisting of 96 samples (2 x 48 uniquely barcoded samples from two

multiplex indices). In addition, for sequencing, we used 15% of a standard Illumina library to

increase index diversity.

Pooled libraries were prepared according to guidelines of the sequencing facility and

sequenced on an Illumina HiSeq2500 at the Functional Genomics Centre Zurich (FGCZ,

Switzerland), using one lane per library with 125 cycles in single-end read (125 bp), high-

output mode. The sample set per library included a negative (no sample DNA) and a positive

(sample replica; different positive controls in different libraries) control to exclude the

possibility of contamination and to calculate the genotyping error of SNPs.

Bioinformatic analysis

Sequence data were demultiplexed, and PCR duplicates were filtered using the functions

"process_radtags" and "clone_filter" of STACKS v1.47, respectively (Catchen et al., 2011;

Catchen, Hohenlohe, Bassham, Amores, & Cresko, 2013). We used trimmed sequencing

reads (generated sequences), applying TRIMMOMATIC v0.36 (Bolger, Lohse, & Usadel, 2014)

with the following conditions: (1) removing Illumina adapter matches allowing a maximum

of two mismatches, (2) removing leading and trailing low quality or N bases below a quality

score of 5, (3) performing a 5 bp sliding window quality check and trimming sequence ends

if quality dropped below 15, and (4) dropping sequences, which were < 50 bp after previous

quality checks. Pre-filtered sequence reads were aligned and mapped against the reference

genome of *P. veris* (Nowak et al., 2015) using BURROWS-WHEELER ALIGNER v0.7.17 (BWA; Li,

2013). SNPs were called using FREEBAYES v1.1.0-54-g49413aa (Garrison & Marth, 2012)

applying default values except a minimum-mapping-quality of 5, a minimum-base-quality of

5, and evaluating the 10 best SNP alleles. We only used SNPs which met quality criteria of

the DDOCENT SNP filtering pipeline (Puritz, Hollenbeck, & Gold, 2014; Puritz, Matz, et al.,

2014) with customizing the following parameters: minimum quality score of 20, minor allele

count of 3, and maximum missing value proportion of 20% across all individuals. Loci

potentially in linkage disequilibrium were filtered using VCFTOOLS 0.1.15 (*geno-r2* function;

Danecek et al., 2011) keeping one random SNP of potentially linked SNP pairs with a

threshold of 0.8. Loci showing an excess of heterozygotes (> 60% of the samples identified

as heterozygotes) were filtered in R v3.4.2 (R Development Core Team, 2017). The

genotyping error of filtered SNPs was calculated by the weighted mean of error rates using

replica samples (i.e. positive controls) with TIGER v1.0 (Wegmann Lab, 2019).

Compilation of adaptive and neutral SNP sets

From the total SNP data set (SNP_overall), we identified putatively adaptive SNPs

(SNP_adapt) that were either (a) linearly associated to environmental factors or (b)

categorically associated to habitat types (see below). An additional putatively adaptive SNP

set (SNP_BayeScan) was compiled by identifying SNPs under potential diversifying or

balancing selection applying an F_{ST} outlier test using BAYESCAN v2.1 (Foll & Gaggiotti, 2008).

The models for BAYESCAN ran with default parameters (10 prior odds, 5,000 iterations with a

thinning interval of 10, a burn in of 50,000 and 20 pilot runs of 5,000 iterations). Potential

F_{ST} outlier loci were extracted for q-values 0.05.

The putatively neutral SNP set (SNP_neutral) was gained by excluding both putatively

adaptive SNP sets (SNP_adapt and SNP_BayeScan) from SNP_overall. Subsequently, we

calculated population genetic diversity parameters for both SNP_neutral and SNP_adapt.

We did not use the SNP_BayeScan set for the estimation of genetic diversity indices at

adaptive loci, because the SNPs identified by BAYESCAN might represent loci putatively

involved in adaptation to other environmental factors than investigated here and were only

used to better define SNP_neutral.

Environmental association analyses

To detect putative signatures of natural selection in open and overgrown habitats, we used

EAA that correlates environmental variation (describing the local habitat) with genetic

variation of a population (Rellstab, Gugerli, Eckert, Hancock & Holderegger, 2015). Here, we

performed EAAs with two types of relationships: linear and categorical.

For the linear analysis, we used latent factor mixed models (*lfmm_ridge* and *lfmm_test*

functions in LFMM v2.0 in R; Caye, Jumentier, Lepeule, & François, 2019), which test for a

linear relationship of the allele frequency (AF) at each SNP with each environmental variable

while accounting for population structure with random latent factors. All 32 populations

were considered, and all 18 environmental variables were used in EAA. We did not remove

correlated variables, because our aim was to identify all SNPs with any sign of

environmental adaptation for the compilation of the different SNP sets. For all subsequent

analyses, however, we concentrated on those SNPs that were associated with

environmental variables that significantly differed among the two habitat types. In LFMM,

the number of latent factors must be set by the user and is recommended to be based on

the number of genetic clusters in the study systems (see below) and the inflation factor λ

(François, Martins, Caye, & Schoville, 2016). To control for false discoveries, we adjusted the

p values per environmental variable using λ and the χ^2 distribution (Caye et al., 2019;

François et al., 2016) and applied the Benjamini-Hochberg algorithm (Benjamini &

Hochberg, 1995) with a false discovery rate (FDR) of 0.05.

For the categorical analysis, we performed three different pairwise analyses on

population AFs of the 20 populations that were sampled in pairs, i.e., geographically close,

but environmentally diverged (open-overgrown; Table 1): a paired t-test, a paired Wilcoxon

signed-rank test, and a sign test. To reduce false positive findings, only SNPs whose

population AFs were significantly different between habitat types in all three categorical

tests were considered for further analyses. For the t- and Wilcoxon tests, we used an α

value of 0.05. In the sign test, we checked whether AF differences between the two

populations of all pairs were consistent (i.e. had the same sign). We considered SNPs

significant if they were consistent in a minimum of eight out of ten comparisons (pairwise

AF differences of 0 were treated as consistent). In these categorical analyses, population

structure is not directly incorporated, but since they test for differences within pairs,

population structure can be ignored, as it is unlikely that population structure would lead to

different signs of AF differences in different pairs. For all further analyses on putatively

adaptive SNPs (SNP_adapt), we concentrated on the SNPs that were (a) associated in LFMM

to those environmental variables that significantly differed among the two habitat types

and/or (b) were significant in all three categorical tests.

For the SNP_adapt set, we also wanted to know whether we find a non-random pattern

of AF change ("direction") between the two populations of habitat pairs. For each SNP, we

identified the putatively beneficial allele (i.e. major allele; AF > 0.5) for the open habitat and

then calculated the average AF change of this beneficial allele from open to overgrown

habitat (the historical habitat change) within pairs. We then counted how many SNPs

exhibited a beneficial AF decrease or increase from open to overgrown habitats and used an

exact two-sided binomial test in R to check if this pattern deviated from a random 1:1 ratio.

The use of linear and categorical EAAs for extracting putatively adaptive loci in respect

to specific environmental factors to test the effect of habitat differences (which are caused

by the same environmental factors) on genetic diversity indices might result in a circularity

problem. This is because major AF and heterozygosity are not completely independent

(Rosenberg & Jakobsson, 2008) and thus, differences in genetic diversity between habitat

types might reflect the results of EAAs. To address this potential problem, we created 100

SNP_overall data sets with fully randomised genotype to sample assignments in order to

create an empirical null distribution of the calculated parameters and statistics. Each of

those 100 SNP_overall data sets (hereafter randomised data sets) went through the same

process of EAAs and F_{ST} outlier analysis as the original data set to extract putatively neutral

and adaptive loci.

Population genetic diversity

For measurements of genetic diversity at putatively neutral and adaptive loci, we calculated,

for all SNP sets of the original and for SNP_neutral and SNP_adapt of each randomised data

set, observed heterozygosity (H_o) using GENALEX 6.503 (Peakall & Smouse, 2012), and mean

nucleotide diversity (π) using VCFTOOLS within a window of 125 bp over all loci for each

population.

The effect of habitat type (open or overgrown) and SNP set (SNP_neutral or

SNP_adapt), and their interaction, on genetic indices (H_o, π) was tested using linear mixed

effect models with population (1 – 32) and region (Muhu or Saaremaa) as random effects

(lmerTest v3.1-0, *lmer* function; Kuznetsova, Brockhoff, & Christensen, 2017), for the

original and each randomised data set. The latter random effect accounted for potential

differences in landscape history among regions. To quantify the importance of fixed factors

and their interaction we used the Likelihood Ratio Test to obtain *p* values, analysing the

variance between the full and reduced models with Satterthwaite approximation (χ^2 and

associated p values). Here, we were particularly interested in the interaction of habitat type

and SNP set. A significant interaction implies that genetic diversity indices assessed at

putatively neutral and adaptive loci behave differently in open and recently overgrown

habitats. If the interaction term was found significant, we tested for differences of genetic

diversity indices in the different habitats for each SNP set using post-hoc tests with least-

square means (lsmeans v2.30-0, *lsmeans* function; Lenth, 2016). We also tested for an effect

of population size on genetic diversity indices by including it as a fixed effect in the above

mixed effect models. Due to non-significant effects of population size and its interaction, we

only present the results of mixed effect models without population size.

We compared results of the original data set with those of the randomised data sets by

(a) calculating a "permutation p-value" for each factor (habitat type, SNP set and their

interaction), i.e. the number of p-values of the randomised data set model outputs which

were smaller than the p-value of the original data set model output divided by the total

amount of randomisations (100); (b) inspecting whether the proportion of putatively

adaptive SNPs with decreasing and increasing beneficial AF, respectively, of the original data

set is significantly different than the distribution of the proportion of SNP_adapt with

decreasing and increasing beneficial AF of the randomised data sets, using a t-test; and (c)

calculating the mean H_o and π per population over all 100 randomised data sets and

conducting a mixed-effect model analysis as done for the original data set.

Population genetic structure and potential gene flow

Population genetic structure using SNP_neutral and SNP_overall was analysed using

discriminant analysis of principle components (DAPC) in adegenet v2.1.1 in R (Jombart,

2008). DAPC uses uncorrelated principal component analysis (PCA) variables for

discriminant analysis, producing synthetic discriminant functions that maximize between-

group variation while minimizing within-group variation (Jombart, Devillard, & Balloux,

2010). We used cross-validation with 50 replicates to determine the number of principal

components (PCs) to be retained to avoid overfitting. The function *find.clusters* was used to

determine the optimal number of clusters within the data sets. For validation, we also

applied a hierarchical clustering tree analysis on a Nei's genetic distance matrix using mmod

v1.3.3 and the *aboot* function from poppr v2.8.4 in R (Kamvar et al., 2019), with a cut-off of

50 and a bootstrap sample of 1,000.

Pairwise genetic differentiation (F_{ST}) among populations for SNP_neutral and

SNP_overall was calculated using genepop v1.0.5 (Rousset et al., 2017) in R. Potential effects

of geographic distance and habitat type "distance" (open-open, overgrown-overgrown,

open-overgrown/overgrown-open) on genetic differentiation (F_{ST}) were tested for both SNP

sets using multivariate generalized linear mixed models fitted with Markov chain Monte

Carlo techniques (MCMCglmm) with 2,000,000 iterations and 500,000 burnins (MCMCglmm

2.29, *MCMCglmm* function; Hadfield, 2010) to account for non-independence of pairwise

distance data. We present results of the best model according to DIC for each SNP set.

After a first visual inspection of the genetic and geographic distance relationship, we

fitted multiple simple linear functions (package stats 3.4.2, *lm* function; Chambers, 1992) to

further characterize potential isolation by distance (IBD; Van Strien, Holderegger, & Van

Heck, 2015) and to estimate the maximum geographic distance up to which gene flow as

indicated by genetic differentiation might be prevalent compared to other genetic

processes, such as genetic drift. One set of linear models included pairwise F_{ST} values as

response variable and increasing geographic distance as explanatory variable, the other

complementary set of linear models fitted F_{ST} against a constant for the difference in

geographic distance to 100 km (i.e. the maximum distance between study populations). The

threshold for potential gene flow was estimated as the point where the sum of the residual

standard errors of sets of complementary models stayed constant.

Results

Sequencing of ddRADseq fragments yielded on average about 150 M raw sequences per library, with on average about 1.2 M sequences per sample. SNP calling and quality filtering resulted in 4,588 SNPs. From those, 3,084 SNPs remained after excluding loci potentially in linkage equilibrium and with an excess of heterozygotes, in a total of 568 individuals from 32 populations. The genotyping error of quality-filtered SNPs was 0.004. Negative controls did not result in sequences.

Putatively adaptive loci

From the 18 environmental variables describing the habitat of populations, six significantly (t-test, $p \leq 0.05$) differed among the two habitat types (open and overgrown; Figure S 1): shrub coverage, light above and below the herbal layer, butterfly species richness and abundance, and plant species richness.

In the linear EAA, based on the number of clusters in the DAPC ($K = 6$, see below) and based on the fact that the inflation factor λ in un-adjusted p values did not vary considerably from $K = 3\text{-}10$ in all environmental variables, we chose $K = 6$ latent factors for the LFMM analysis. With an FDR of 0.05, we identified eight SNPs being associated to an environmental variable (Table S 1). Only three of them were associated to one of the six variables that significantly differed between the habitat types (butterfly abundance).

In the categorical EAA comparing the two habitat types, we identified 99 SNPs with the paired t-test ($p \leq 0.05$), 95 with the Wilcoxon test ($p \leq 0.05$), and 557 SNPs with the sign test. Seventy-four SNPs were identified in all three pairwise tests, but none of them overlapped with the three LFMM SNPs (Figure S 2). The 74 SNPs from categorical EAAs and the three

SNPs from linear EAAs were used for further analyses (SNP_adapt = 77 SNPs). Using

randomised data sets, the mean number of SNP_adapt was 82 (± 20 SD) SNPs.

In the 77 SNPs that were putatively involved in adaptation to habitat type, 53 SNPs

showed a decrease of the average beneficial AF (for the old, open habitat) in the new,

overgrown compared to the open habitat (Figure 2). The binomial test revealed that this

pattern was significantly different from a random expectation ($p < 0.01$). However, AF

differences between habitat types were small; average AF change was 0.09 (range 0.04-

0.16) in the 53 SNPs with decreasing, and 0.08 (range 0.03-0.15) in the 24 SNPs with

increasing beneficial AF. Yet, the maximum AF difference found between two populations of

a pair in any SNP was 0.56 (Figure S 3). Using randomised data sets, the mean proportion of

SNP_adapt loci with decreasing beneficial AF was 52 (± 9 SD) %, and the mean proportion of

SNP_adapt loci with increasing beneficial AF was 48 (± 9 SD) % (Figure S 4).

The BayeScan analysis resulted in 391 potential F_{ST} outlier loci (SNP_BayeScan). These

SNPs, together with those from SNP_adapt, were removed from the SNP_overall to create

SNP_neutral (2619 loci). There was almost no overlap between SNP_adapt and

SNP_BayeScan (Figure S 2).

Population genetic diversity

There was a significant interaction effect of habitat type and SNP set on observed

heterozygosity (H_o; Table 2). For SNP_neutral, H_o ranged from 0.21 to 0.30 across all study

populations (Table 1). There was no significant difference of H_o between populations in

open and overgrown habitats (Figure 3a; post-hoc test: $p = 0.91$). For SNP_adapt, H_o ranged

from 0.18 to 0.31 across all study populations (Table 1). Importantly, there was a significant

difference of H_o between populations in open and overgrown habitats (post-hoc test:

$p < 0.05$), with populations in overgrown habitats exhibiting increased H_o compared to

populations in open habitats (Figure 3a). The random factors region and population

accounted for 45.1% and 32.8% of variation in the data for H_o. For the mean of randomised

data sets, there was no significant interaction nor habitat effect on H_o ($p = 0.58$ and $p = 0.24$,

respectively) but a significant SNP set effect ($p < 0.001$; Figure S 5a and Figure S 6a). The

"permutation p-value" for H_o for the factor habitat type was $p = 0.05$, for SNP set it was

$p = 0.7$, and for the interaction it was $p = 0.04$.

For nucleotide diversity (π), there was a marginally (non-)significant interaction effect of

habitat type and SNP set ($p = 0.056$; Table 2). For SNP_neutral, π ranged from 0.0024 and

0.0033 across all study populations (Table 1). There was no significant difference of π

between populations in open and overgrown habitats (Figure 3b; post-hoc test: $p = 0.97$).

For SNP_adapt, π ranged from 0.0022 and 0.0032 across all study populations (Table 1). π of

populations in overgrown habitats was higher than π of populations in open habitats, but

this difference was not significant (Figure 3b; post-hoc test: $p = 0.21$). The random factors

region and population accounted for 0% and 68.4% of variation in the data for π. For the

mean of randomised data sets, there was no significant interaction nor habitat effect on π

($p = 0.36$ and $p = 0.73$, respectively) but a significant SNP set effect ($p < 0.001$; Figure S 5b

and Figure S 6b). The "permutation p-value" for π for the factor habitat type was $p = 0.14$,

for SNP set it was $p = 0$, and for the interaction it was $p = 0.32$.

Population genetic structure and potential gene flow

DAPC of SNP_neutral identified six genetic clusters across the 32 populations with 200 PCs

retained (Figure 4a). The first two discriminant functions from DAPC analysis and the

hierarchical clustering tree analysis highlighted a differentiation by geographic regions,

mainly by the islands Muhu and Saaremaa (Figure 4b,d). Importantly, the separation of

genetic clusters was not based on habitat types (Figure 4c,d).

Pairwise F_{ST} values assessed using SNP_neutral were mostly moderate with an average

of 0.10 (± 0.05 SD) ranging from 0.01 to 0.25. We found a positive significant relationship

between pairwise F_{ST} and geographic distances among all populations ($p_{MCMCglmm}$ < 0.001;

Table S 2), indicating isolation by distance (IBD; Figure 5). Model fitting showed that values

of residual standard error (RSE) of the models roughly reached a plateau between 15 to

30 km in geographic distance between populations (Figure S 7), indicating a potential

threshold up to where genetic differentiation is driven by geographic distance and gene flow

rather than random (genetic) processes. The habitat type did not explain patterns in genetic

differentiation (Table S 2).

Results of all analyses using SNP_overall were highly similar to when using SNP_neutral.

For completeness, these results are presented as supplemental information (Table S 2 – S 4,

Figure S 7 – S 10).

Discussion

Habitat degradation due to abandoned management and related loss and isolation of

European semi-natural grasslands (Auffret et al., 2018; Habel et al., 2013) during the last

century has been shown to negatively impact the biodiversity of these grasslands, both at

the species and the genetic level (e.g. Helm et al., 2006; Picó & Van Groenendael, 2007). Yet,

the effect of abandoned management, which can result in gradual overgrowth of grasslands

with woody vegetation, on the genetic diversity at adaptive loci of grassland plants has

remained unknown. Here, we examined the effect of recent overgrowth of abandoned

semi-natural calcareous grasslands (Estonian alvars) on the genetic diversity at putatively

neutral and adaptive loci of the perennial herb *Primula veris*. Our study revealed that *P.*

veris populations in the new, overgrown habitats likely had a similar level of genetic

diversity at putatively neutral loci as in the old, open habitats, despite substantial change in

environmental conditions. Genetic diversity at putatively adaptive loci, however, was higher

in the new, overgrown compared to the old, open habitats. Neutral genetic structure and

gene flow as indicated by neutral genetic differentiation was not (yet) affected by grassland

overgrowth. We are among the first to demonstrate how recently changed non-climatic

selection pressures are potentially in the process of changing adaptive genetic patterns of

wild populations. Most other studies concentrated on the effects of climate on plants (e.g.

Dauphin et al., 2020; Sun et al., 2020) or used manipulative experiments in otherwise

natural habitats (Laurentino et al., 2020). Hence, our study is an example of potential *in-situ*

"adaptation in action", where genetic diversity at adaptive loci is increasing due to a slow

loss of previous genetic adaptations.

Overgrowth of semi-natural grasslands has often a negative impact on genetic diversity

of grassland-specialist plants due to a decrease in habitat quality and potential development

of barriers for gene flow (Aavik & Helm, 2018; Picó & Van Groenendael, 2007). However, in

our study, neutral genetic diversity of *P. veris* was similar in open and overgrown habitats,

indicating that habitat and landscape change do not necessarily restrict gene flow among

populations of *P. veris*. Still, an environmental impact on the neutral part of a genome might

only show after several generations (e.g. Landguth et al., 2010). *Primula veris* is a perennial

plant with an average lifespan of up to 50 years and can persist with reduced reproduction

even when the environment has changed due to overgrowth (Ehrlén & Lehtilä, 2002).

Hence, genetic diversity of such populations potentially reflects the state before their

habitats started to change (Reinula, 2018). In addition, heterozygosity indices, as used in our

study, have been suggested to respond more slowly to environmental change as compared

to, for example, measures of inbreeding (Lloyd, Campbell, & Neel, 2013; Lowe, Boshier,

Ward, Bacles, & Navarro, 2005). Deschepper et al. (2017), who examined patterns of neutral

genetic diversity of *P. veris* in grassland and forest populations in Belgium, also reported no

significant difference in H_o between habitat types. Yet, their study system, i.e. forests,

represents a late-successional stage, whereas ours represents a mid-successional stage

(shrubby overgrowth). Consequently, neutral genetic patterns of *P. veris* might need a very

long time (up to centuries) to reach an equilibrium with the new overgrown environment.

In contrast to genetic diversity at putatively neutral loci, genetic diversity assessed at

putatively adaptive loci differed between habitat types in our study. Populations in recently

overgrown grasslands showed higher genetic diversity at putatively adaptive loci compared

to populations in open grasslands (Figure 3a,b). This is most likely caused by changed

selection pressures in the new, overgrown habitats compared to the old, open habitats. The

open grasslands used in our study have been in an open state for several hundreds of years,

whereas ongoing overgrowth started only about 90 years ago (Helm et al., 2006). Hence,

populations of *P. veris* in open grasslands have experienced homogenous selection

pressures for a very long time, which increased the frequencies of alleles beneficial for an

open habitat and led to reduced heterozygosity and genetic diversity at adaptive loci. The

majority of putatively adaptive SNPs (53 out of 77) exhibited a smaller average beneficial

(i.e. major) AF for open habitat conditions in populations from the new, overgrown

compared to those from the old habitat. This led to an increase in heterozygosity and

genetic diversity in populations of the new, overgrown habitats (Figure 2). Considering *P.*

veris' potential longevity, this indicates that populations in overgrown habitats are still

adapting to their new selection pressures (e.g. lower light availability, reduced or altered

pollinator community), i.e. many alleles potentially beneficial for the new, overgrown

habitat are still far from fixation (i.e. homozygosity). Consequently, the genetic diversity at

putatively adaptive loci of *P. veris* populations in overgrown grasslands has not yet been

reduced. A similar pattern was found in the conifer tree *Pinus cembra*, where populations in

the core of the current niche exhibited a decreased genetic diversity at adaptive loci, but no

difference in the one at neutral loci, compared to populations at the niche margin, which

most likely present unstable or novel habitats (Dauphin et al., 2020).

An increase of non-beneficial alleles for the open habitat in overgrown populations, and

thus an increase in heterozygosity at putatively adaptive loci in overgrown populations, can

be achieved in two ways: (1) the alternative alleles at SNP loci were either already present in

lower frequencies in populations of open habitats (i.e. standing genetic variation), or (2)

arrived to overgrowing grasslands by gene flow, before they were subject to positive

selection in the overgrown habitat. As shown in Figure S 3, beneficial alleles in populations

of the open habitat were rarely fixed, strongly pointing towards the importance of standing

genetic variation. This implies that even populations with reduced genetic diversity at

putatively adaptive loci (e.g. in open grasslands) may possess the ability to react to habitat

changes due to the low, but crucial amount of standing genetic variation (e.g. Morris,

Bowles, Allen, Jamniczky, & Rogers, 2018). Still, we found gene flow as indicated by genetic

differentiation potentially spreading alleles between habitats which, in theory, can

contribute to the increased genetic diversity at adaptive loci in overgrown populations. The

potential gene flow distances found in our study (up to 30 km; Figure 4, Figure 5) stand in

marked contrast to the potential dispersal ranges of pollen and seeds of *P. veris* for which

very short distances up to 12 m and 0.5 m, respectively, have been reported (Antrobus &

Lack, 1993; Richards & Ibrahim, 1978). Yet, pollinating insects of *P. veris* have been shown to

occasionally travel up to 2 km (Kreyer, Oed, Walther-Hellwig, & Frankl, 2004; Zurbuchen,

Bachofen, Müller, Hein, & Dorn, 2010) and dispersal distances are often underestimated in

ecological studies (e.g. Bullock, Shea, & Skarpaas, 2006). Historical rotational grazing of

domestic animals and movement of wild animals (e.g. deer, moose, wild boar) might also

have facilitated seed dispersal for longer distances in *P. veris* which is not adapted to

zoochory (Plue, Aavik, & Cousins, 2019). Overall, standing genetic variation but also gene

flow are capable of supplying *P. veris* populations undergoing habitat changes with new or

alternative alleles fostering adaptation to new habitat conditions in our study region.

Importantly, the fact that we detected a significant effect of habitat type on genetic

diversity (H_o) when using the putatively adaptive SNP set but not with the neutral SNP set

suggests the need to examine genetic diversity at neutral and adaptive loci separately when

studying the genetic response of plant species to environmental changes. Besides, in the

overall SNP set (Figure S 10), there was no significant effect of habitat type on genetic

diversity, which indicates that genetic diversity of an overall SNP set represents rather

neutral genetic patterns (Dauphin et al., 2020). In conservation genetics and restoration,

most assessments have been based on overall or neutral genetic diversity so far (González

et al., 2019; Wei & Jiang, 2020), even though its stand-alone importance is questionable

(Teixeira & Huber, 2021) and it is likely the putatively adaptive regions of a genome that are

important for the fate of a population in a changed environment.

The different effects of habitat type on genetic diversity at putatively neutral and

adaptive loci could partly be due to the unevenness in the number of loci at neutral and

putatively adaptive regions (2,619 versus 77, respectively). However, the ddRADseq

procedure used in our study should result in a "random" and representative subset of both,

neutral and adaptive SNPs. Here, we were not particularly interested in the actual molecular

mechanisms underlying adaptation, but in general patterns at putatively adaptive loci,

which should also become visible with 77 SNPs if these patterns are of substantial nature.

Additionally, the reliability of the detected SNPs can be assumed high, because the

genotyping error for our SNPs was low (0.004) and the SNP sets were identified using a draft

genome (Hoban et al., 2016), which covers about 63% of the whole 479.22 Mb genome of

the study species *P. veris* (Nowak et al., 2015).

When using linear and categorical EAA to test effects of habitat type on genetic

diversity, we must consider a potential circularity issue. Nevertheless, the results of the

randomised data sets indicate that, for H_o, the habitat type related results (including the

interaction with SNP set) of the original data are likely not concerned by circularity. Yet, this

is the case for π. Potentially, since our study sites are just in the beginning of a habitat

change process (i.e. in terms of perennial plants like *P. veris*), more time might be needed

until a stronger habitat change effect can be detected by π. Importantly, even though the

number of putatively adaptive loci using randomised data sets is in the range of putatively

adaptive loci found using the original data set, the putatively adaptive loci from the original

data set do not show a random pattern with respect to their change in beneficial AF (Figure

S 4). In addition, applying the same mixed-effect model statistic to the mean randomised

data set, for both H_o and π, there is no habitat type or interaction effect anymore (Figure S

5), as can be expected when breaking the apparent link between genotype and habitat type

sample assignment of the original data set. Therefore, caution is needed when interpreting

the results, especially those using π, but the overall conclusion of potential early stages of

"adaption in action" processes remains. To our knowledge, there is no previous study

applying EAA to examine putatively neutral and adaptive genetic diversity that used

permutations for validation. In contrast, such validations are common practice in genome

wide association studies (e.g. Kreiner, Tranel, Weigel, Stinchcombe, & Wright, 2021; Tibbs

Cortes, Zhang, & Yu, 2021) which, together with our results, might call for their wider

application in landscape genomic analyses.

In addition to genetic diversity, population size is another factor that is affected by

habitat change, and which also directly influences genetic diversity (Leimu et al., 2006). The

census sizes of our *P. veris* populations were similar in open and overgrown habitats and

were not associated with genetic diversity at neutral and adaptive loci. In contrast to our

results, a meta-analysis showed a clear relationship between population size and (neutral)

genetic diversity, including long-lived and self-incompatible plant species, such as *P. veris*

(Leimu et al., 2006). Populations in our study exhibited a minimum of about 100 individuals,

which is larger than the overall census for small populations in the study by Leimu et al.

(2006). Consequently, our *P. veris* populations might be still sufficiently large to counteract

population size driven effects on genetic diversity.

Conclusions

Our landscape genomic investigation of *Primula veris* in Estonian semi-natural grasslands is

one of the first to demonstrate the potential effects of land use change on the genetic

diversity at putatively adaptive loci of *in-situ* wild plant species. We show that the effect of

recent overgrowth of grasslands might not be genetically manifested when considering

neutral SNPs, even after almost a century of ongoing environmental changes. Yet, genetic

diversity assessed at putatively adaptive loci was higher in populations in overgrown

compared to open habitats, most probably due to allele frequency changes of standing

genetic variation. Thus, even populations in degraded and fragmented habitats may possess

the ability to adapt to habitat changes due to their important standing genetic variation in

addition to potential allele immigration due to gene flow. However, permutation analyses

indicated that caution is needed when interpreting results based on linear and categorical

EAA, especially when working with nucleotide diversity.

For perennial long-lived plant species, such as *P. veris*, long time spans might pass

before habitat change can be detected at neutral regions of the genome whereas habitat

effects at adaptive loci could be noticeable faster. Consequently, a repeated monitoring of

genetic diversity at both neutral and adaptive loci and further investigations of

contemporary gene flow at different spatial scales would be highly valuable to identify

potential genetic consequences of recent and ongoing environmental change in natural and

semi-natural habitats. In addition, extending our results to whole-genome and targeted

sequencing approaches would be vital to reliably identify genes and gene networks

putatively involved in the adaptation of *P. veris* to habitat change, and to assess the relative

importance of loss of previous and gain of new genetic adaptations to the altered

environment.

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Data Accessibility

Sequence data used in this study will be made available at the European Nucleotide Archive

(ENA) upon acceptance (ERS5253979 – ERS5254546). R-scripts, genotypic and

environmental data will be provided at the Dryad Digital Repository upon acceptance (xxx).

Author Contributions

T.A., S.T. and A.H designed the conceptual approach and carried out field work. S.T. and I.R.

conducted laboratory work. S.T. and N.Z. performed bioinformatic analyses. S.T. and C.R.

analysed the data. R.H. contributed with discussing the results in a broader ecological

context. S.T. wrote the manuscript with major contributions from C.R. All authors read,

commented and approved the final version of the manuscript.

Competing interests

The authors declare no competing interests.

Supplemental Information

Supplemental Methods

Figure S 1 Environmental variables and their response to habitat type (open – overgrown).

Figure S 2 Venn diagram of shared putatively adaptive loci by different methods.

Figure S 3 SNP allele frequencies and their behaviour in open and overgrown habitats.

Figure S 4 Proportion of putatively adaptive SNPs with de- and increasing beneficial AF of
the original and randomised data sets.

Figure S 5 Mean of genetic and nucleotide diversity (H_o and π) at putatively neutral and
adaptive SNPs of 100 randomised data sets.

Figure S 6 Genetic and nucleotide diversity (H_o and π) at putatively neutral and adaptive
SNPs per population of 100 randomised data sets.

Figure S 7 Residual standard error results for IBD analyses, measured using the overall
(3,084 loci) and neutral (2,619 loci) SNP sets.

Figure S 8 Genetic structure of *Primula veris* populations in the study area measured using
the overall set of loci (3,084 loci).

Figure S 9 Isolation by distance pattern measured at the overall (3,084 SNPs) set of loci for
Primula veris populations.

Figure S 10 Genetic and nucleotide diversity (H_o and π) at the overall set of SNPs (3,084 loci).

Table S 1 Number of associations between SNPs and environmental variables.

Table S 2 Results of generalized mixed effect models for the effect of geographic distance
and habitat type distance on pairwise genetic differentiation (F_{ST}) for the neutral (2,619 loci)
and overall (3,084 loci) SNP sets.

Table S 3 Genetic diversity measures using the overall set of SNPs (3,084 loci) for the studied
populations of *Primula veris*.

Table S 4 Results of linear mixed effect models for H_o, π for the overall SNP set (3,084 loci) of
the studied *Primula veris* populations.

Tables

Table 1 Location information and genetic diversity measurements for the studied populations of *Primula veris* in Muhu and Saaremaa, Estonia.
Habitat – affiliation to open or overgrown grassland; pair – affiliation to a pair of closely situated populations (0 indicates affiliation to no pair);
N – sample size after SNP filtering; population size – estimated number of individuals per population; H_o – observed heterozygosity; π –
nucleotide diversity.

Pop ID	Region	Site	Habitat	Pair	Latitude [°N]	Longitude [°E]	N	Population size	Neutral SNP set		Adaptive SNP set	
									H_o	π	H_o	π
1	Saaremaa	Asva1	open	1	58.445	23.061	6	100	0.22	0.0033	0.23	0.0032
2	Saaremaa	Asva2	overgrown	1	58.419	23.064	20	200	0.23	0.0027	0.25	0.0026
3	Saaremaa	Atla	overgrown	0	58.302	21.936	10	100	0.30	0.0031	0.30	0.0026
4	Saaremaa	Kahtla1	open	0	58.466	23.240	20	200	0.25	0.0028	0.23	0.0023
5	Muhu	Koguva	overgrown	2	58.611	23.091	20	2000	0.26	0.0027	0.27	0.0024
6	Muhu	Koguva	open	2	58.611	23.091	19	500	0.27	0.0028	0.26	0.0024
7	Saaremaa	Kõruse	overgrown	3	58.447	21.939	20	1000	0.27	0.0029	0.25	0.0025
8	Saaremaa	Kõruse	open	3	58.447	21.939	19	1000	0.26	0.0028	0.23	0.0023
9	Muhu	Lõetsa1	overgrown	4	58.650	23.313	20	1000	0.28	0.0029	0.31	0.0027
10	Muhu	Lõetsa1	open	4	58.650	23.314	19	1000	0.27	0.0029	0.25	0.0025
11	Muhu	Lõetsa2	overgrown	0	58.644	23.345	20	100	0.28	0.0029	0.31	0.0027
12	Saaremaa	Lõu	open	0	58.122	22.201	20	5000	0.26	0.0031	0.22	0.0026
13	Muhu	Mäla	overgrown	5	58.579	23.270	20	800	0.24	0.0029	0.27	0.0028
14	Muhu	Mäla	open	5	58.579	23.271	15	500	0.26	0.0030	0.24	0.0026
15	Saaremaa	Neeme	open	0	58.484	21.947	20	1500	0.27	0.0029	0.22	0.0024
16	Saaremaa	Neeme	open	0	58.499	21.927	18	3000	0.28	0.0029	0.28	0.0028
17	Muhu	Nõmmküla	open	0	58.669	23.209	19	1000	0.27	0.0029	0.30	0.0026
18	Muhu	Nõmmküla	open	0	58.667	23.204	20	500	0.26	0.0028	0.28	0.0025
19	Saaremaa	Orinõmme	open	6	58.585	23.024	20	3500	0.27	0.0028	0.26	0.0024
20	Saaremaa	Orinõmme	overgrown	6	58.584	23.024	20	5000	0.26	0.0028	0.27	0.0024
21	Muhu	Paenase	overgrown	7	58.641	23.159	20	1000	0.29	0.0029	0.30	0.0025

22	Muhu	Paenase	open	7	58.641	23.154	19	1000	0.28	0.0028	0.30	0.0024
23	Saaremaa	Türju	overgrown	0	57.980	21.993	17	100	0.26	0.0028	0.25	0.0025
24	Saaremaa	Vanamõisa	open	0	58.243	22.674	19	1000	0.21	0.0024	0.20	0.0023
25	Saaremaa	Vanamõisa	open	8	58.225	22.685	15	500	0.21	0.0025	0.18	0.0022
26	Saaremaa	Vanamõisa	overgrown	8	58.222	22.697	7	100	0.22	0.0031	0.25	0.0032
27	Muhu	Võiküla1	open	0	58.545	23.385	20	500	0.28	0.0028	0.28	0.0024
28	Muhu	Võiküla2	overgrown	9	58.551	23.309	20	10000	0.27	0.0030	0.29	0.0027
29	Muhu	Võiküla2	open	9	58.551	23.309	19	10000	0.27	0.0029	0.27	0.0026
30	Saaremaa	Võrsna	open	0	58.389	22.747	11	100	0.22	0.0027	0.23	0.0027
31	Muhu	Üügu	overgrown	10	58.675	23.222	17	500	0.28	0.0029	0.29	0.0028
32	Muhu	Üügu	open	10	58.671	23.238	19	4000	0.27	0.0029	0.27	0.0025
Open mean								1800	0.26	0.0029	0.25	0.0025
Overgrown mean								1700	0.27	0.0029	0.28	0.0027

Table 2 Results of linear mixed effect models for observed heterozygosity (H_o), and
nucleotide diversity (π) of *Primula veris* populations of the original data set. Fitted
parameters (± SE), t-values and significance (p_{Model}) is given for the full model of each
genetic index, H_o and π. Factors: habitat (H): open and overgrown; SNP_set (S): SNP_neutral
and SNP_adaptive. RF_{region} and RF_{pop} denote the use of region (Muhu and Saaremaa) and
population (1–32) as random factors (RF). The importance of the fixed factor interaction is
given as χ^2 and *p*-values for each genetic measurement. For H_o and π, the importance of the
fixed factor interaction is given for the comparison of the full model (with H * S) with the
next simplest model (H + S).

Effect	Value	SE	t	p_{Model}	χ^2	p
H_o ~ H * S + RF_{region} + RF_{pop}					Importance of H * S	
Habitat	0.027	0.008	3.304	0.002		
SNP_set	0.008	0.005	1.793	0.083		
habitat:SNP_set	-0.021	0.007	-2.942	0.006	8.110	0.004
π ~ H * S + RF_{region} + RF_{pop}					Importance of H * S	
Habitat	$1.4e^{-4}$	$7.0e^{-5}$	2.009	0.051		
SNP_set	$3.4e^{-4}$	$3.5e^{-5}$	9.573	<0.001		
habitat:SNP_set	$-1.1e^{-4}$	$5.6e^{-5}$	-1.985	0.056	3.949	0.047

Figures

Figure 1 Map of the study area on Muhu and Saaremaa islands in Western Estonia. *Primula*
veris populations are depicted with blue filled circles when occurring in open semi-natural
grassland and with orange filled triangles when occurring in recently overgrown historical
semi-natural grasslands.

Figure 2 Patterns of allele frequency (AF) change of the 77 putatively adaptive SNPs
(SNP_adapt) averaged across pairs of populations for each SNP. For each SNP, the putatively
beneficial allele frequency for the open habitat (AF > 0.5) is shown. SNPs were derived from
linear and categorical environmental association analyses. SNPs with decreasing/increasing
AF in overgrown habitats (orange triangles) compared to open habitats (blue circles) are
indicated with dark grey/light grey lines, respectively. For detailed results, see Figure S 3.

Figure 3 Boxplots of observed heterozygosity (H_o; A) and nucleotide diversity (π; B) at putatively neutral and adaptive loci in open (blue, circles) and overgrown (orange, triangles) grasslands. Factors: habitat (H): open and overgrown; SNP_set (S): SNP_neutral and SNP_adaptive. Significance values: [ns] $p > 0.05$; [+] $p = 0.056$; * $p \leq 0.05$; ** $p < 0.01$; *** $p < 0.001$.

Figure 4 Genetic structure of *Primula veris* populations in the study region. (A) Result of the
discriminant analysis of principle components on the putatively neutral SNP set resulting in
six geographical clusters (C; named by cardinal directions). The distribution of the six
clusters per discriminant function is shown in panel B. (D) Result of the hierarchical
clustering tree analysis with population codes (see Table 1) and colour coding by cluster.

Figure 5 Relationship between genetic differentiation (F_{ST}) and geographic distance for all
possible pairs of 32 *Primula veris* populations, measured at putatively neutral loci.
Population pairs with a geographic distance (D) less or equal to 30 km are visualized as
green squares. Population pairs with a geographic distance greater than 30 km are given as
yellow triangles. The effect of geographic distance on F_{ST} is highly significant ($p < 0.001$) up
to a threshold of 30 km.