Highlights

Learning Rates Are Not All the Same: The Interpretation of Computational Model Parameters Depends on the Context

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- Efforts in computational cognitive modeling often assume that Reinforcement Learning (RL) modeling parameters will generalize between studies and models, but this is not well established.
- We empirically investigate whether RL parameters generalize between three tasks and models, using a large developmental dataset and a within-participant design.
- We find that RL decision noise/exploration parameters generalize fairly well, but RL learning rates do not.
- Our data support previous conclusions that decision noise/exploration decreases during development (ages 8-17), but suggests that claims about learning rate development cannot be generalized.

Learning Rates Are Not All the Same: The Interpretation of Computational Model Parameters Depends on the Context

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Abstract

Reinforcement Learning (RL) has revolutionized the cognitive and brain sciences, explaining behavior from simple conditioning to problem solving, across the life span, and anchored in brain function. However, discrepancies in results are increasingly apparent between studies, particularly in the developmental literature. To better understand these, we investigated to which extent parameters *generalize* between tasks and models, and capture specific and uniquely *interpretable* (neuro)cognitive processes. 291 participants aged 8-30 years completed three learning tasks in a single session, and were fitted using state-of-the-art RL models. RL decision noise/exploration parameters generalized well between tasks, decreasing between ages 8-17. Learning rates for negative feedback did not generalize, and learning rates for positive feedback showed intermediate generalizability, dependent on task similarity. These findings can explain discrepancies in the existing literature. Future research therefore needs to carefully consider task characteristics when relating findings across studies, and develop strategies to computationally model how context impacts behavior.

Keywords: Computational Modeling, Model Parameter, Reinforcement

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Learning, Development, Generalizability, Interpretability

1 1. Introduction

In recent decades, the cognitive neurosciences have made breakthroughs 2 in computational modeling, showing that reinforcement learning (RL) mod-3 els can explain foundational aspects of human behavior. RL does not only seem to underlie simple cognitive processes such as stimulus-outcome and 5 stimulus-response learning [1, 2, 3], but also complex ones, including goaldirected, temporally-extended behavior [4, 5], meta-learning [6], and abstract problem solving that requires hierarchical thinking [7, 8, 9, 10]. Underlining their centrality in the study of human cognition, RL models have been ap-9 plied across the lifespan [11, 12, 13], and in healthy participants as well those 10 experiencing psychiatric illnesses [14, 15, 16, 17, 18]. RL models are of partic-11 ular interest because they also capture brain function: A specialized network 12 of brain regions, including the basal ganglia and prefrontal cortex, implement 13 computations that mirror specific components of RL algorithms, including 14 action values and reward prediction errors [19, 20, 21, 22, 23, 24, 25]. In sum, 15 explaining behaviors from simple conditioning to complex problem solving. 16 adequate for diverse human populations, based on a compelling theoretical 17 foundation [26], and with strong ties to brain function, RL has experienced 18 a surge in published studies since its inception [27], and emerged as a pow-19 erful and potentially unifying modeling framework for cognitive and neural 20 processing. 21

Despite their increasing popularity, however, not enough attention has 22 been paid to what exactly RL models and model variables (e.g., model pa-23 rameters) measure, and our current assumptions might be imprecise, po-24 tentially slowing further progress. Our recent opinion paper develops this 25 argument in depth [28]. In brief, computational modeling condenses be-26 havioral datasets into a model and a small number of free model parameters 27 [11, 27, 29, 30, 31, 32]. We as researchers often assume that these models and 28 parameters expose mental and/or neural processes, and have the ability to 29 dissect them into specific, unique components (e.g., value updating and deci-30 sion making), thereby measuring participants' inherent characteristics (e.g., 31 individual learning rates). However, we argue in this paper that these as-32 sumptions might be too optimistic and that a careful empirical investigation 33 is required to assess their validity. 34

We focus on two major aspects, which are adopted widely in computa-35 tional modeling [28]: generalizability and interpretability. We define a model 36 variable (e.g., fitted parameter) as *generalizable* if it is consistent across uses, 37 such that a person would be characterized with the same values independent 38 of the specific model or task used to estimate the variable. Generalizability is 39 a consequence of the assumption that parameters are intrinsic to participants 40 (e.g., a person with a high learning rate) rather than task dependent. We 41 further define a model parameter as *interpretable* if it isolates specific and 42 unique elements of cognition, which are often assumed to be implemented in 43 separable neural substrates: Decomposing behavior into model parameters 44 is seen as a way of *carving cognition at its joints*. 45

Assumptions about generalizability and interpretability are rarely stated 46 explicitly, but underlie conclusions across the fields of computational psy-47 chology and neuroscience, and often implicitly guide research efforts. As-48 sumptions of generalizability, for example, inspired many to identify the 49 inherent, task-independent settings of parameters in humans (e.g., empiri-50 cal parameter distributions [33]; relationships between negative and positive 51 learning rates [34]), to characterize the age development of parameters in a 52 task-independent way [11, 12, 13, 35], and to compare parameters between 53 studies in review articles [14, 15, 16, 19, 20, 21, 22, 23, 25], meta-analyses 54 [24, 36, 37], and discussion sections of empirical papers: When model vari-55 ables are compared between different types of studies, there is an implicit 56 assumption of generalization. Relying on interpretability, model variables 57 have been expected to be associated with specific neural substrates (e.g., 58 reward prediction errors and dopamine function [38]), to expose the core of 59 what differentiates participants with psychiatric conditions from healthy ones 60 (e.g., working-memory parameter differences in schizophrenia [39]), and gen-61 erally, to capture processes that are particularly "theoretically meaningful" 62 [14].63

However, inconsistencies in empirical results are emerging across the developmental [13, 40, 41, 42], clinical [15, 16, 17, 18], cognitive, and neuroscientific literature [24, 36, 37, 43], potentially suggesting a lack of generalizability
and/or interpretability, which is also in accordance with different theoretical considerations [27, 28, 44, 45, 46, 47]. Nevertheless, the degree to which
parameters generalize between tasks and are interpretable has not been investigated empirically yet (but see [48] for work on parameter reliability).

This was the goal of the current project. We compared the RL parameters
fit to the same individuals across different learning tasks in a single study.

We used a developmental dataset (291 participants, ages 8-30 years), which 73 allowed us to obtain a large spread of individual differences and address 74 outstanding discrepancies in the developmental psychology literature [13]. 75 The three learning tasks varied on several common dimensions, including 76 feedback stochasticity, task volatility, and memory demands (Fig. 1B), and 77 have previously been used to study RL processes [35, 49, 50]. However, like 78 many tasks in the literature, these tasks likely also engaged other cognitive 79 processes, such as working memory and reasoning. The within-participant 80 design allowed us to test directly whether the same participants showed the 81 same parameters across tasks (generalizability), and the combination of mul-82 tiple tasks shed light on which cognitive processes parameters captured in 83 each task (interpretability). We extensively compared and validated the RL 84 models of each task [27, 30, 51], and previously reported the developmental 85 results separately [35, 49, 50]. 86

We found that the RL parameters that reflect decision noise or exploration 87 (inverse decision temperature β , undirected noise ϵ ; for model details, see 88 section 4.5) were most consistent within individuals across tasks, suggesting 89 that these parameters were most generalizable. Decision noise/exploration 90 parameters also showed a consistent developmental pattern across subjects, 91 declining from age 8-17. RL learning rate parameters (α_+, α_-) , however, 92 were largely inconsistent within individuals across tasks, showing that they 93 did not generalize. Capturing different variance, they likely also reflected 94 different cognitive processes across tasks. Both of these patterns are con-95 sistent with patterns that have started to emerge in the existing literature 96 [13]. Behavioral analyses indicated that task differences, and the associated 97 differences in optimal behavior, might underlie these observed parameter dis-98 crepancies. These results suggest that past computational findings are not 99 as generalizable as often assumed, and that future research needs to address 100 the reasons of the observed discrepancies to move the field forward. 101

102 2. Results

The next section gives a brief overview of the experimental tasks and computational models, before tackling parameter generalizability (section 2.1) and parameter interpretability (section 2.2). Task details are provided in Fig. 1C-E and section 4.4, and computational models and parameter fitting in section 4.5, as well as the original publications [35, 49, 50].

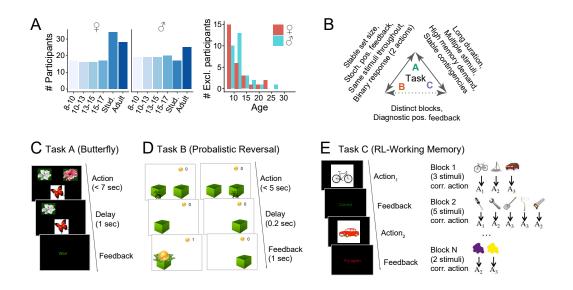


Figure 1: Overview of the experimental paradigm. (A) Participant sample. Left: Number of participants in each age group, broken up by sex (self-reported). Age groups were determined by within-sex age quartiles for participants between 8-17 years. The adult sample is broken up by recruitment type ("Stud.": University undergraduates, receiving course credit for participation. "Adult": Adults recruited from the community using the same methods as the developing participants. Right: Number of participants who participated in the study and whose data were excluded because they failed to reach the performance criterion in at least one task. (B) Pairwise similarities in task design between tasks A, B, and C. Similarities between each pair of tasks are shown above the connecting arrows. Only features are shown that differentiate two tasks from the third. E.g., noting "Stable set size" on the edge between tasks A and B implies that set size was not stable in task C. Task A shared more similarities with tasks B and C than they shared with each other. (C) Procedure of task A ("Butterfly task"). Participants saw one of four butterflies on each trial and selected one of two flowers in response, via button press on a game controller. Each butterfly had a stable preference for one flower throughout the task, but rewards were delivered stochastically (70% for correct responses, 30% for incorrect). For details, see section 4.4 and the original publication [50]. (D) Procedure of task B ("Probababilistic switching"). Participants saw two boxes on each trial and selected one with the goal of finding gold coins. At each point in time, one box was correct and had a high (75%) probability of delivering a coin, whereas the other was incorrect (0%). At unpredictable intervals, the correct box switched sides. For details, see section 4.4 and [49]. (E) Procedure of task C ("Reinforcement learning-working memory"). Participants saw one stimulus on each trial and selected one of three responses. All correct responses and no incorrect responses were rewarded. Stimuli were presented in blocks containing 2-5 different stimuli. The number of stimuli in a block is called set size. The task was designed to disentangle set-size sensitive working memory processes from set-size insensitive RL processes. For details, see section 4.4 and [35].

Depending on the task, RL models contained different parameters, re-108 flecting existing differences in the literature. Task A required participants 109 to learn the correct associations between each of four stimuli (butterflies) 110 and two responses (flowers), through probabilistic feedback (Fig. 1C). The 111 best-fitting model contained three free parameters: learning rate from posi-112 tive outcomes α_+ , inverse decision temperature β , and Forgetting F; and one 113 fixed parameter: learning rate from negative outcomes $\alpha_{-} = 0$ [50]. Task B 114 required participants to adapt to unexpected switches in the action-outcome 115 contingencies of a simple bandit task (only one of two boxes contained a 116 gold coin at any time), based on semi-probabilistic feedback (Fig. 1D). The 117 best-fitting RL model contained four free parameters: α_+ , α_- , β , and choice 118 persistence p [49]. Task C required learning of stimulus-response associations 119 like task A, but over several task blocks with varying numbers of stimuli, 120 and provided deterministic feedback (Fig. 1E). The best model for this task 121 combined RL and working-memory processes, containing RL parameters α_+ 122 and α_{-} ; working-memory parameters capacity K, Forgetting F, and noise ϵ ; 123 and mixture parameter ρ , which determined the relative weights of RL and 124 working memory [35, 52]. 125

To ensure that potential parameter discrepancies in this study were not 126 due to a lack of modeling quality, we employed rigorous model fitting, com-127 parison, and validation [27, 29, 30, 51]: For each task, we compared a large 128 number of competing models, based on different parameterizations and cog-129 nitive mechanisms, and selected the best one based on quantitative model 130 comparison scores, models' ability to reproduce participants' behavior in sim-131 ulation, and other criteria of model fit (e.g., interpretability) [44, 45]. We 132 also used hierarchical Bayesian methods for model fitting and comparison 133 when possible to obtain most accurate parameter estimates [51]. Individual 134 publications provide further details [35, 49, 50]. 135

¹³⁶ 2.1. Part I: Parameter Generalizability

To investigate parameter generalizability, we assessed whether partici-137 pants showed similar parameter values across tasks, and whether different 138 tasks showed the same parameter age trajectories. These within-participant 130 comparisons are crucial to determine whether discrepancies in the previ-140 ous literature were caused by methodological differences (e.g., differences in 141 participant samples, testing procedures, modeling quality, research labs), or 142 could arise from mere differences in task characteristics and computational 143 models, as we hypothesized. 144

145 2.1.1. Differences in Absolute Parameter Values

We first asked whether tasks led to different absolute parameter values (Fig. 2A), using repeated-measures analyses of variance (ANOVAs). When ANOVAs showed significant task effects, we followed up with pairwise, repeated-measures t-tests, using the Bonferroni correction.

Learning rates α_+ and α_- occupied largely distinct ranges across tasks: 150 Values were very low in tasks C (α_+ mean: 0.07, sd: 0.18; α_- mean: 0.03, sd: 151 0.13), intermediate in task A (α_+ mean: 0.22, sd: 0.09; α_- was fixed at 0), 152 and fairly high in task B (α_+ mean: 0.77, sd: 0.11; α_- mean: 0.62, sd: 0.14; 153 for statistical comparisons, see Table 1). Decision noise was high in task B ($\frac{1}{R}$) 154 mean: 0.33, sd: 0.15), but low in tasks A ($\frac{1}{\beta}$ mean: 0.095, sd: 0.0087) and C 155 (ϵ mean: 0.025, sd: 0.032; statistics in Table 1 ignore ϵ because its absolute 156 values were not comparable to $\frac{1}{\beta}$ due to the different parameterization; see 157 section 4.5). Forgetting was significantly higher in task C (mean: 0.19, sd: 158 0.17) than A (mean: 0.056, sd: 0.028; task B was best fit without forgetting). 159 All ANOVAs revealed significant and large task effects, and all follow-up 160

t-tests revealed significant and large pairwise differences (Table 1), showing 161 that absolute parameter values differed substantially between tasks. This 162 shows that the three tasks produced significantly different estimates of learn-163 ing rate, decision noise/exploration, and forgetting for the same participants 164 (Fig. 2B). Interestingly, these parameter differences echoed differences in 165 task demands: Learning rates and noise/exploration were highest in task B, 166 where frequent switches required quick updating and high levels of explo-167 ration. Similarly, forgetting was highest in task C, which posed the largest 168 demands on memory. Using regression models that controlled for age (instead 169 of ANOVA) led to similar results (Table D.9). 170

171 2.1.2. Relative Parameter Differences

However, comparing absolute parameter values between tasks has short-172 comings: It ignores variance between participants, even though between-173 participant variance might be the more meaningful measure because it re-174 flects participants' relationships to each other. The simplest way to inves-175 tigate whether between-participant variance generalized between tasks is to 176 test if individual variance in one task mirrors individual variance in another, 177 using Spearman correlation (suppl. Fig. D.8). Indeed, both α_+ (suppl. Fig. 178 D.8A) and noise/exploration parameters (suppl. Fig. D.8B) were signifi-179 cantly positively correlated between task A and tasks B and C, suggesting 180

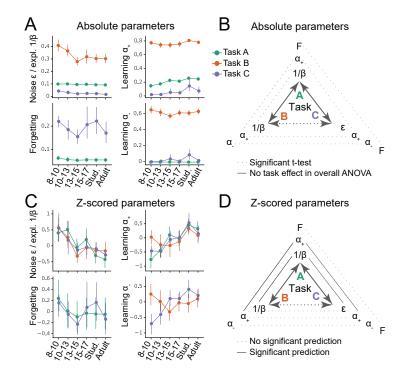


Figure 2: Generalizability of absolute parameter values (A-B) and of parameter age trajectories / z-scored parameters (C-D) between tasks. (A) Fitted parameters over participant age (quartile bins), for all three tasks (A: green; B: orange; C: blue). Parameter values differed significantly between tasks (for statistics, see Table 1). Dots indicate means, error bars specify the confidence level (0-1) for interval estimation of the population mean. (B) Summary of the main results of part (A), visualizing Table 1. Double-sided arrows are replicated from Fig. 1B and show task similarity. Lines show test statistics for absolute parameter values. Dotted lines indicate significant task differences in Bonferroni-corrected pairwise t-tests, which were conducted after observing significant task effects in corresponding ANOVAs. All t-tests were significant, indicating that absolute parameter values differed significantly for each pair of tasks. (C) Parameter age trajectories, i.e., within-task z-scored parameters over participant age bins. Age trajectories can potentially reveal similarities that are obscured by differences in means or variances when assessing absolute parameter values. (D) Summary of the main results of part (C), visualizing Table 4. When parameters in two tasks are connected with a full line, the parameter can be predicted significantly in one task from the other. When parameters are connected with a dotted line, the prediction is not significant. In contrast to absolute parameter values, age trajectories were predictive in several cases, especially for tasks with more similarities (A and B; A and C), compared to tasks with fewer (B and C).

that variance between participants generalized better than absolute values. 181 However, significant correlations were lacking between tasks B and C. This 182 suggests that α_+ and noise/exploration generalized from and to task A, but 183 they did not generalize between tasks B and C, mirroring task similarities 184 (Fig. 1B; also see section 2.2.1; Fig. D.9 shows the correlations between 185 all pairs of features in the dataset.) Note that noise parameters generalized 186 between task A and C despite differences in parameterization (ϵ vs. $1/\beta$), 187 showing robustness in the characterization of choice stochasticity (suppl. Fig. 188 D.8B). 189

190 2.1.3. Parameter Age Trajectories

However, this correlation analysis is limited in its failure to take into 191 account age effects, a known source of variance, such that apparent task 192 similarities could be driven by a shared dependence on age rather than age-193 independent underlying similarities. To address this, we next analyzed pa-194 rameters' age trajectories, which allowed us to abstract away potentially 195 arbitrary differences (e.g., different parameter means and variances across 196 tasks), while conserving potentially meaningful structure in the dataset (i.e., 197 participants' parameter values relative to each other). 198

We obtained age trajectories by z-scoring each parameter within each task (Fig. 2C). To test for differences in age trajectories, we used mixedeffects regression to predict parameters of all tasks from two age predictors (age and squared age) and task (A, B, or C). When this model fit better than the corresponding model without task, task characteristics affected age trajectories, and we added post-hoc models for each pair of tasks.

For α_{-} , the task-based regression model showed a significantly better fit, 205 revealing significant task differences (Table 2). Indeed, α_{-} showed funda-206 mentally different age trajectories in task B compared to C (in task A, α_{-} 207 was fixed): In task B, α_{-} decreased linearly, modulated by a U-shaped cur-208 vature (linear effect of age: $\beta = -0.11$, p < 0.001; quadratic: $\beta = 0.003$. 209 p < 0.001), but in task C, it increased linearly, modulated by an inverse-U 210 curvature (linear: $\beta = 0.32$, p < 0.001; quadratic: $\beta = -0.07$, p < 0.001; 211 Fig. 2C). These differences were reflected in the significant interaction terms 212 of the grand regression model (Table ??). 213

For α_+ , adding task as a predictor did not improve model fit, suggesting that age trajectories did not differ (Table 2). Indeed, age trajectories were qualitatively similar between tasks, showing linear increases that tapered off with age (linear increase: task A: $\beta = 0.33$, p < 0.001; task B: $\beta = 0.052$,

²¹⁸ p < 0.001; task C: $\beta = 0.28$, p < 0.001; quadratic modulation: task A: ²¹⁹ $\beta = -0.007$, p < 0.001; task B: $\beta = -0.001$, p < 0.001; task C: $\beta = -0.006$, ²²⁰ p < 0.001).

For noise/exploration and Forgetting parameters, age trajectories did not differ either (Table 2). For decision noise/exploration, the grand regression model revealed a linear decrease and tapering off in older participants that was consistent across all tasks (Fig. 2C; Table ??), in accordance with previous findings [13]. For Forgetting, the grand model did not reveal consistent age effects (Fig. 2C; Table ??).

In summary, when assessing absolute parameter values (Fig. 2A, 2B), dif-227 ferences in scale obscured existing similarities in age trajectories for noise/exploration 228 parameters and α_{+} (Fig. 2C). For α_{-} , on the other hand, differences existed 229 both in terms of scale (Fig. 2A, 2B) and age trajectories (Fig. 2C). As sug-230 gested by the correlation analysis, patterns of generalization differed between 231 pairs of tasks, such that more generalization was present between tasks that 232 were more similar in terms of task characteristics (A and B; A and C; not B 233 and C). 234

235 2.1.4. Predicting Age Trajectories

So far, we have assessed parameter differences to reveal parameters that 236 do not generalize across tasks. However, the absence of differences only pro-237 vides indirect evidence for generalization. We therefore next assessed how 238 closely parameters were related, using linear regression to predict partici-239 pants' parameters in one task from the values of the same parameter in a 240 different task. We controlled for age by including age and squared age as 241 predictors to ensure that the prediction was driven by parameter similarities 242 beyond age. 243

For both α_{+} and noise/exploration parameters, task A predicted tasks B 244 and C, and tasks B and C predicted task A, but tasks B and C did not pre-245 dict each other (Table 4; Fig. 2D), confirming that α_{+} and noise/exploration 246 generalized from and to task A, but not between tasks B and C, mirroring 247 task similarities (Fig. 1B; also see section 2.2.1). For α_{-} , tasks B and C 248 showed a marginally significant *negative* relationship (Table 4), suggesting 249 that predicting α_{-} in one task from the other would lead to inverse predic-250 tions. Indeed, we previously reported a U-shaped trajectory of α_{-} in task 251 B with minimum in 13-to-15-year-olds [49], but a consistent increase up to 252 early adulthood in task C [50], revealing striking qualitative differences in 253 the estimation of α_{-} when using these two tasks. For Forgetting, tasks A 254

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Parameter	Model	Tasks	F / t	df	p	sig.
$\frac{1}{\beta}$	ANOVA	A, B	830	1	p < 0.001	***
1-	t-test	A vs B	25	246	p < 0.001	***
α_+	ANOVA	A, B, C	2.018	2	p < 0.001	***
	t-test	A vs B	66	246	p < 0.001	***
	t-test	A vs C	12	246	p < 0.001	***
	t-test	B vs C	51	246	p < 0.001	***
α_{-}	ANOVA	B, C	2.357	1	p < 0.001	***
	t-test	B vs C	49	246	p < 0.001	***
Forgetting	ANOVA	A, C	161	1	p < 0.001	***
	t-test	A vs C	49	246	p < 0.001	***

Table 1: Statistics of ANOVAs predicting raw parameter values from task (A, B, C). When an ANOVA showed a significant task effect, post-hoc, Bonferroni-corrected t-tests were added. * p < .05; ** p < .01, *** p < .001.

Table 2: Assessing the existence of age effects on parameter trajectories: Model fits of regression models predicting parameter age trajectories, comparing the added value of including ("AIC with task") versus excluding ("AIC without task") task as a predictor. Differences in AIC scores were tested statistically using F-tests. The best (significantly smaller) AIC scores are highlighted in bold, and their coefficients are shown in Table ??.

Parameter	AIC without task	AIC with task	F(df)	р	sig.
$\frac{1}{\beta}/\epsilon$	2,044	2,054	NA	NA	_
α_+	2,044	2,042	F(4, 245) = 2.34	p = 0.056	_
α_{-}	1,395	$1,\!373$	F(2, 245) = 6.99	p = 0.0011	**
Forgetting	1,406	1,411	NA	NA	_

Table 3: Statistical tests on age trajectories: mixed-effects regression models predicting z-scored parameter values from task (A, B, C), age, and squared age (months). When the task-less model fitted best, the coefficients of this model are shown, showing shared age trajectories (Table 2; $\frac{1}{\beta}/\epsilon$, α_+ , Forgetting). When the age-based model fitted better, pairwise follow-up models are shown (α_-), showing task differences. P-values of follow-up models were corrected for multiple comparison using the Bonferroni correction. * p < .05; ** p < .01, *** p < .001.

Parameter	p < .001. Tasks	Predictor	β	p (Bonf.)	sig.
$\frac{1}{\beta}/\epsilon$	A, B, C	Intercept	1.86	< 0.001	***
1-		Age (linear)	-0.17	0.003	**
		Age (quadratic)	0.004	< 0.001	***
α_+	A, B, C	Intercept	-2.10	< 0.001	***
		Age (linear)	0.20	< 0.001	***
		Age (quadratic)	-0.004	< 0.001	***
α_{-}	B, C	Task (main effect)	4.15	< 0.001	***
		Task * linear age (interaction)	0.43	< 0.001	***
		Task * quadratic age (interaction)	-0.010	< 0.001	***
Forgetting	A, C	Intercept	0.37	0.44	
		Age (linear)	-0.034	0.53	
		Age (quadratic)	0.001	0.63	

²⁵⁵ and C were not predictive of each other (Table 4).

Importantly, these results (Fig. 2D) differ from the previous patterns (Fig. 2C) for Forgetting parameters and α_+ in tasks B and C. This shows that a lack of difference (Fig. 2C) does not imply successful prediction (Fig. 259 2D).

260 2.1.5. Summary Part I

In summary, Part I revealed that (1) different tasks led to different es-261 timates of participants' exploration $(\frac{1}{\beta})$, Forgetting (F), and learning rates 262 (α_+, α_-) , revealing a lack of generalization of absolute parameter values. 263 Intriguingly, absolute parameter values were stable within tasks (reflecting 264 task demands), but varied within participants. (2) In contrast to absolute 265 parameter values, age trajectories of noise/exploration parameters and learn-266 ing rates α_+ were qualitatively similar between tasks, suggesting that pa-267 rameter age trajectories generalized better than absolute values. The age 268 trajectories of learning rates α_{-} , however, differed fundamentally between 260 tasks, highlighting that parameters in the same models can generalize dif-270 ferently. (3) Assessing the parameters with task-consistent age trajectories, 271

²⁷² noise/exploration decreased until early adulthood, in accordance with the ²⁷³ literature [13], while learning rates α_+ increased. (4) Using between-task ²⁷⁴ prediction as the strongest test of generalization, age trajectories of learning ²⁷⁵ rates α_- and Forgetting were not predictive, and noise/exploration param-²⁷⁶ eters and learning rates α_+ could only be predicted between similar tasks, ²⁷⁷ suggesting that generalizability was generally weaker than expected, and ²⁷⁸ might depend on task similarity.

279 2.2. Part II: Parameter Interpretability

Based on these insights, Part II of our investigations focused on parameter interpretability, i.e., the concept that parameters capture specific and unique cognitive processes that are well delineated. We tested parameter interpretability by investigating the relations between different parameters in our dataset, assessing the specificity and distinctiveness of each parameter as well as the relations between parameters and observed patterns of behavior.

286 2.2.1. The Main Axes of Variation

To gain an understanding of what information was captured by each 287 parameter, we employed a data-driven approach, identifying major axes of 288 variance without specifying a priori hypotheses. We used PCA to identify the 289 major axes in our dataset (composed of both behavioral features and model 290 parameters). We then used these axes (principal components; PCs) to in-291 terpret model parameters. To understand the PCs themselves, we analyzed 292 the weights of the behavioral features on each PC (Fig. 3). Detailed infor-293 mation is provided in sections 4.6 (PCA methods), Appendix C (behavioral 294 features), and suppl. Fig. D.10 (additional PCA results). 295

We first examined PC1, the axis of largest variation (25.1%) of explained 296 variance; suppl. Fig. D.10A), to understand the main sources of individual 297 differences in our dataset. Behaviors that indicated good task participation 298 (e.g., high percentage of correct choices) loaded positively on PC1, whereas 299 behaviors that indicated that participants were not on task loaded negatively 300 (e.g., more missed trials, longer response times; Fig. 3A). PC1 comprised 301 measures both in the narrow sense of maximizing task accuracy (e.g., per-302 centage correct choices, measures of task accuracy, win-stay choices), and in 303 the wider sense of reflecting task engagement (e.g., number of missed trials, 304 response times, response time variability). PC1 therefore captured a range 305 of "good performance" indicators, reflecting general task engagement. PC1 306

increased significantly with age, consistent with participants' increasing performance (suppl. Fig. B.6B; age effects of subsequent PCs in suppl. Fig.
D.10; suppl. Table D.8).

In all three tasks, noise/exploration loaded negatively on PC1 (Fig. 3A), 310 showing that elevated decision stochasticity was associated with poorer per-311 formance in all tasks. Forgetting parameters also loaded negatively, support-312 ing a negative role for performance. α_{+} showed positive loadings in all three 313 tasks, suggesting that faster integration of positive feedback was associated 314 with better performance. Intriguingly, α_{-} loaded positively in task C, but 315 negatively in task B, suggesting that performance increased when partici-316 pants integrated negative feedback faster in task C, but decreased when they 317 did the same in task B. This distinction can be interpreted in terms of task 318 demands: Negative feedback was diagnostic in task C, but non-diagnostic in 319 task B (Fig. 1B), such that repeating choices after negative feedback ("Lose-320 stay" behavior) was hurtful in the former (negative loading on PC1 for task 321 C), but can be beneficial in the latter (positive loading on PC1 for task B; 322 Fig. 3A). 323

Having gained insight into parameters' roles for task engagement by an-324 alyzing PC1, we next turned to PC2 and PC3. To facilitate their interpreta-325 tion, we flipped the loadings of all PC2 and PC3 features that were negative 326 on PC1, to make them intrepretable with respect to task engagement (for 327 methodological details, see section 4.6). This pre-processing revealed that 328 PC2 and PC3 encoded task contrasts: PC2 contrasted task B to task C 329 (loadings on corresponding features were positive / negative / near-zero for 330 tasks B / C / A; Fig. 3B). PC3 contrasted task A to both B and C (load-331 ings on corresponding features were positive / negative for task A / tasks 332 B and C; Fig. 3C; missed trials and response times did not show task con-333 trasts, suggesting that these features did not differentiate between tasks). 334 The ordering of PC2 and PC3 shows that participants' behavior differed 335 more between tasks B and C (PC2: 8.9% explained variance) than between 336 B or C and A (PC3: 6.2%; suppl. Fig. D.10), in accordance with descriptive 337 task characteristics (Fig. 1B). This shows that after task engagement, the 338 main variation in our dataset arose from task differences. 330

Intriguingly, noise/exploration parameters, α_+ , and α_- reproduced the task contrasts of PC2 and PC3, showing positive or negative loadings based on the task in which they were measured (Fig. 3B, 3C). This means that these parameters differed sufficiently between tasks to be discriminable (as opposed to, e.g., response times and numbers of missed trials, which did not show

task contrasts, suggesting that they were not discriminable between tasks).
Each parameter therefore captured enough task-specific variance to make it
possible to be identified with the correct task. This degree of differentiability
would not be expected if parameters captured the same processes in each
task, in which case they would capture the same variance and not show
task differences. Taken together, PC2 and PC3 confirmed that each of these
parameters captured task-unique processes.

Taken together, the PCA revealed that (1) the main axes of variation in 352 the dataset were task engagement (PC1) and task differences (PC2-PC3). 353 (2) Noise/exploration, Forgetting, α_+ , and α_- all were related to task en-354 gagement (PC1). Whereas the relation was consistent between tasks for the 355 former three, it was task-dependent for α_{-} and mirrored specific task de-356 mands. (3) Noise/exploration, α_+ , and α_- all captured enough task-specific 357 variance to be correctly identified with the corresponding task, showing that 358 they captured different processes depending on the task (PC2-PC3). 359

360 2.2.2. Parameters and Cognitive Processes

Whereas the previous analysis revealed that all parameters contained 361 task-specific information, it did not specify how much information was task-362 specific and how much was shared. For example, noise/exploration param-363 eters contained enough task-specific information to make it possible to de-364 termine in which task they were measured (PC2-PC3; Fig. 3B, 3C), but 365 they also showed similar associations with engagement across tasks (PC1; 366 Fig. 3A), similar age trajectories (Fig. 2C), and were mutually predictive 367 (Fig. 2D). To quantify these patterns, we need to understand how much of 368 each parameter's variance was unique and how much was shared between 360 parameters and between tasks. 370

To achieve this, we probed how much of each parameter's variance was 371 explained by other parameters, using regression. We assumed that param-372 eters reflected one or more cognitive processes, such that shared variance 373 between parameters would imply overlapping cognitive processes. If param-374 eters reflected similar cognitive processes across tasks, then the same param-375 eter should dominate this analysis (e.g., when using parameters in task A 376 to predict $\frac{1}{\beta}$ in task B, task A's $\frac{1}{\beta}$ should show the largest regression co-377 efficient). However, if parameters captured different processes across tasks. 378 this would not be the case (e.g., all parameters of task A might predict 379 task B's $\frac{1}{\beta}$ equally). We used repeated, k-fold cross-validated Ridge regres-380 sion to avoid overfitting, obtaining unbiased out-of-sample estimates of the 381

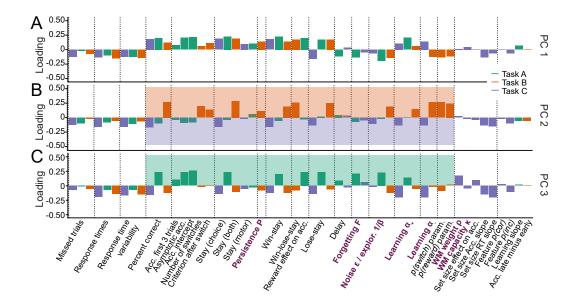


Figure 3: Identifying the major axes of variation in the dataset. A PCA was conducted on the entire dataset (39 behavioral features and 15 model parameters). The figure shows the factor loadings of the first three PCs. RL model parameters are highlighted in purple on the x-axis. Behavioral features are explained in detail in Appendix A and Appendix B. (A) PC1 captured broadly-defined task engagement, with negative loadings on features that were negatively associated with performance (e.g., number of missed trials) and positive loadings on features that were positively associated with performance (e.g., percent correct trials). (B-C) PC2 (B) and PC3 (C) captured task contrasts. PC2 loaded positively on features of task B (orange box) and negatively on features of task C (blue box). PC3 loaded positively on features of task A (green box) and negatively on features of tasks B and C. We flipped the loadings of features that were negative on PC1 when showing PC2 and PC3 to better visualize the task contrasts (section 4.6).

means and variances of explained variance R^2 and regression coefficients w(for methods, see section 4.7).

We first assessed the overall patterns of prediction, and found that all 384 significant coefficients highlighted shared variance between tasks A and B or 385 tasks A and C, but never between tasks B and C, mirroring our previous 386 results (Fig. 2D; section 2.1.2) and patterns of task similarity (Fig. 1B). 387 This means that no parameters in tasks B or C played a significant role in 388 predicting parameters in the other, while both tasks' parameters were predic-389 tive (and being predicted by) parameters in task A. This further highlights 390 the potential role of task similarity in parameter generalizability. 391

We next focused on noise/exploration parameters. Noise/exploration 392 parameters in tasks B and C showed significant coefficients when predict-393 ing noise/exploration in task A, but the inverse was not true, such that 394 noise/exploration in task A did not show significant coefficients when pre-395 dicting noise/exploration in tasks B or C (Fig. 4A; Table 5). The first 396 result shows that noise/exploration parameters captured variance (cognitive 397 processes) in task A that they also captured in tasks B and C. The second re-398 sult shows that noise/exploration parameters captured additional cognitive 399 processes in tasks B and C that they did not capture in task A. Further-400 more, prediction accuracy increased when combining tasks B and C's param-401 eters to predict noise/exploration in task A, showing that noise/exploration 402 parameters in tasks B and C captured partly non-overlapping aspects of 403 noise/exploration in task A (Fig. 4B, left-most set of bars, compare pur-404 ple to orange and blue). This highlights both specificity in terms of which 405 cognitive processes were captured by noise/exploration parameters across 406 tasks (prediction between similar tasks), and some lack thereof (prediction 407 was just one-way; no prediction between dissimilar tasks). Furthermore, 408 noise/exploration in task A was predicted by Persistence and α_{-} in task B, 409 and by α_{-} and working-memory weight ρ in task C (Fig. 4A; Table 5). This 410 shows that some processes that noise/exploration parameters captured in 411 task A were captured by different parameters in the other tasks, revealing a 412 lack of distinctiveness in noise/exploration parameters. 413

We next assessed learning rates. Specificity was evident in that learning rate α_+ in task A showed a significant regression coefficient when predicting learning rates α_+ and α_- in task C, and learning rate α_- in task C showed a significant coefficient when predicting learning rate α_+ in task A (Fig. 4A; Table 5). However, a lack of specificity was evident in task B: When predicting α_+ in task B, no parameter of any task showed a significant coefficient

(including α_+ in other tasks; Table 5), and it was impossible to predict vari-420 ance in task B's α_+ even when combining all parameters of the other tasks 421 (Fig. 4B, "Task B" panel). This reveals that α_+ captured fundamentally 422 different cognitive processes in task B compared to the other tasks. The case 423 was similar for parameter α_{-} , which strikingly was inversely related between 424 tasks A and B (Table 5), and impossible to predict in task B from all other 425 parameters (Fig. 4B). This shows a lack of specificity, implying that learning 426 rates did not reflect a consistent core of cognitive processes across tasks. 427

We then turned to the distinctiveness of learning rate parameters. Learn-428 ing rate α_+ in task A was predicted indistinctly by all parameters of task B 429 (with the notable exception of α_+ itself; Fig. 4A; Table 5), suggesting that 430 the cognitive processes that α_+ captured in task A were captured by an inter-431 play of several parameters in task B. Furthermore, task A's α_+ was predicted 432 by task C's working-memory parameters ρ and K (Fig. 4A; Table 5), suggest-433 ing that α_+ captured a conglomerate of RL and working-memory processes 434 in task A that was isolated by different sets of parameters in task C [52]. 435 In support of this interpretation, no variance in task C's working-memory 436 parameters could be explained by any other parameters (Fig. 4B), revealing 437 that they captured unique cognitive processes, likely working memory. Task 438 C's RL parameters, on the other hand, could be explained by parameters in 439 other tasks (Fig. 4B), suggesting they captured overlapping RL processes. 440

441 2.2.3. Parameters and Behavior

Faced with mounting evidence for parameter inconsistencies, we lastly 442 aimed to uncover whether parameters shared any consistent similarities across 443 tasks. The previous sections showed that parameters likely captured differ-444 ent (neuro)cognitive processes across tasks (e.g., different internal character-445 istics of learning and choice). However, computational models are funda-446 mentally models of behavior, so we argued that parameters might capture 447 similar behavioral features (e.g., similar tendencies to stay after positive feed-448 back). Even though related, (neuro)cognitive processes and behavioral pat-449 terns should not be equated (Fig. 5). For example, different (neuro)cognitive 450 mechanisms (e.g., prefrontal cortical reasoning, basal ganglia value learning, 451 hippocampal episodic memory) might underlie the same behavioral pattern 452 (e.g., lose-stay behavior) in different tasks, depending on the characteris-453 tics (e.g., stable versus volatile contingencies; deterministic versus stochastic 454 feedback). 455

⁴⁵⁶ To investigate this possibility, we assessed the relationships between model

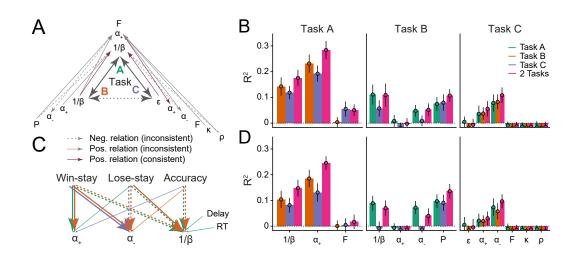


Figure 4: Assessing parameter interpretability by analyzing shared variance. (A) Parameter variance (cognitive processing) that is shared between tasks. Each arrow shows a significant regression coefficient when predicting a parameter in one task (e.g., α_+ in task A) from all parameters of a different tasks (e.g., P, α_- , α_+ , and $\frac{1}{\beta}$ in task B). The predicted parameter is shown at the arrow head, predictors at its end. Full lines indicate positive regression coefficients, and are highlighted in purple when connecting two identical parameters; dotted lines indicate negative coefficients; non-significant coefficients are not shown. Table 5 provides the full statistics of the models summarized in this figure. (B) Variance of each parameter that was also captured by parameters of other models. Each bar shows the percentage of explained variance (R^2) when predicting one parameter from all parameters of a different task/model, using Ridge regression. Part (A) of this figure shows the coefficients of these models. The x-axis shows the predicted parameter, and colors differentiate between predicting tasks. Three models were conducted to predict each parameter: One combined the parameters of both other tasks (pink), and two kept them separate (green, orange, blue). Larger amounts of explained variance (e.g., Task A $\frac{1}{\beta}$ and α_{-}) suggest more shared processes between predicted and predicting parameters; the inability to predict variance (e.g., Task B α_+ ; Task C working-memory parameters) suggests that distinct processes were captured. Bars show mean R^2 , averaged over k data folds (k was chosen for each model based on model fit, using repeated cross-validated Ridge regression; for details, see section 4.7); error bars show standard errors of the mean across folds. (C) Relations between parameters and behavior. The arrows visualize Ridge regression models that predict parameters (bottom row) from behavioral features (top row) within tasks (full statistics in Table 6). Arrows indicate significant regression coefficients, colors denote tasks, and line types denote the sign of the coefficients, like before. All significant within-task coefficients are shown. Task-based consistency (similar relations between behaviors and parameters across tasks) occurs when arrows point from the same behavioral features to the same parameters in different tasks (i.e., multiple arrows). (D) Variance of each parameter that was explained by behavioral features; corresponds to the behavioral Ridge models shown in part (C).

parameters and behavioral features across tasks. Using regularized Ridge re-457 gression like above, we predicted each model parameter from five selected 458 behavioral features (Appendix A, Appendix C) of each of the three tasks 459 (15 predictors; for regression methods, see section 4.7). One possible outcome 460 of this analysis is "absolute consistency": parameters might capture the same 461 behavioral pattern within and across tasks (e.g., noise/exploration of each 462 task might capture task A accuracy). This outcome would be expected if pa-463 rameters captured the same cognitive processes across tasks, and behavioral 464 features were a direct reflection of cognitive processes. Another possible out-465 come is "absolute inconsistency" (e.g., in every task, noise/exploration might 466 capture different behavioral features). This outcome would suggest that pa-467 rameters captured unrelated cognitive and behavioral features in each task. 468 Crucially, a third possible outcome is "task-based consistency": Parameters 469 might capture the same behavioral features, but only within tasks (e.g., in 470 each task, learning rates might capture the win-stay behavior of that task, 471 but not of other tasks). This outcome would suggest that parameters gen-472 eralized in terms of which behavioral features they reflected, but behavioral 473 features—like (neuro)cognitive processes—differed between tasks. 474

Focusing on noise/exploration parameters, $\frac{1}{\beta}$ in tasks A and B was pre-475 dicted by task A win-stay behavior, revealing absolute consistency (Table 476 6). $\frac{1}{\beta}$ was also predicted by accuracy, win-stay, and lose-stay behavior within 477 both tasks A and B, but not across tasks, revealing task-based consistency 478 (Fig. 4C; Table 6). For learning rates, α_+ in tasks A and B was predicted 470 by the corresponding win-stay behavior, and α_{-} in tasks B and C was neg-480 atively predicted by the corresponding lose-stay behavior, and positively by 481 the corresponding win-stay behavior (Fig. 4C; Table 6), revealing task-based 482 consistency. The consistency of α_{-} is especially noteworthy given the abun-483 dance of discrepancies in previous sections. 484

Taken together, noise/exploration parameters, α_+ , and α_- captured similar behavioral features across tasks (Fig. 4C), despite differences in cognitive processing (Fig. 4A, 4B), captured information (Fig. 3B, 3C), age trajectories (Fig. 2C, 2D), and absolute values (Fig. 2A, 2B). Notably, the observed discrepancies reflected task characteristics (Fig. 1B, 3A) for both parameters (Fig. 1B, 3A) and behavior (suppl. Fig. B.6B), suggesting that task characteristics shaped behavioral responses and model parameters.

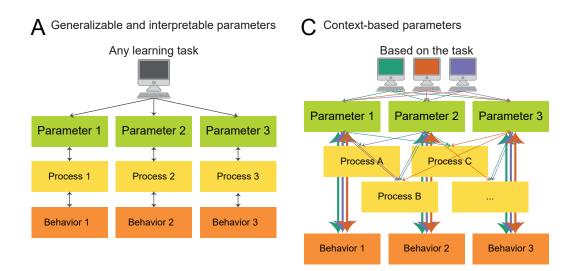


Figure 5: What do model parameters measure? (A) View based on generalizability and interpretability. In this view, which is implicitly taken by much current computational modeling research, models are fitted in order to reveal individuals' intrinsic model parameters, which reflect clearly delineated, separable, and meaningful (neuro)cognitive processes, a concept we call *interpretability*. Interpretability is evident in that every model parameter captures a specific cognitive process (bidirectional arrows between each parameter and process), and that cognitive processes are separable from each other (no connections between processes). Task characteristics are treated as irrelevant, a concept we call *generalizability*, such that parameters of any learning task (within reason) are expected to capture similar cognitive processes. (C) Updated view, based on our results, that acknowledges the role of context (e.g., task characteristics, model parameterization, participant sample) in computational modeling. Which cognitive processes are captured by each model parameter is influenced by context (green, orange, blue), as shown by distinct connections between parameters and cognitive processes. Different parameters within the same task can capture overlapping cognitive processes (not interpretable), and the same parameters can capture different processes depending on the task (not generalizable). However, parameters likely capture consistent behavioral features across tasks (thick vertical arrows).

492 3. Discussion

The current study subjected the generalizability and interpretability of RL models to a scrutinous empirical investigation, using a developmental sample. We found weaker levels of generalizability and interpretability than would be expected based on current research practices, such as comparing parameters between studies that use different tasks and models [28].

Interestingly, patterns of generalizability and interpretability varied be-498 tween parameters: Exploration/noise parameters showed considerable gener-490 alizability in the form of correlated variance and age trajectories. The decline 500 in exploration/noise we observed between ages 8-17 was also consistent with 501 previous studies reviewed in [13]. Interpretability of exploration/noise pa-502 rameters was mixed: Despite evidence for specificity in some cases (overlap 503 in parameter variance between some tasks), it was missing in others (lack of 504 overlap between other tasks), and crucially, parameters lacked distinctiveness 505 (substantial overlap in variance with other parameters). 506

Learning rate from negative feedback, on the other hand, showed a sub-507 stantial lack of generalizability: parameters were less consistent within par-508 ticipants than within tasks, and age trajectories differed both quantitatively 509 and qualitatively. This result is consistent with discrepancies in learning rate 510 parameters across developmental studies [13]. Learning rates from positive 511 and negative feedback combined were interpretable to a limited degree (over-512 lap in variance between some tasks). However, interpretability was overshad-513 owed by a lack of specificity (lack of shared core variance) and distinctive-514 ness (fundamental entangling with several other parameters, most notably 515 working-memory parameters). 516

These within-participant findings are consistent with patterns that are 517 emerging from comparisons of studies published by different labs [13]. Our 518 within-participant design allowed us to go beyond these between-study find-519 ings by confirming that the same participants can show different parame-520 ters when tested using different tasks. The within-participant consistency of 521 noise/exploration parameters strengthens our confidence that these indeed 522 decrease with age [13, 53, 54]. The inconsistency of learning rate parameters 523 leads to the unexpected, but important conclusion that we cannot measure 524 an individual's "intrinsic learning rate" using RL modeling, and that we 525 cannot draw general conclusions about "the development of learning rates" 526 that apply to all RL contexts, using current methods. 527

⁵²⁸ Our findings also help us clarify the source of parameter inconsistencies

in the previous literature, which could indicate replication problems and 529 technical issues: For example, model misspecification [13], lack of model 530 comparison and validation [27, 30], inappropriate fitting methods [29, 51], 531 and lack of parameter reliability due to suboptimal methods [48] have all 532 been suggested as potential sources of inconsistencies. However, our results 533 show that discrepancies are expected even with a consistent methodological 534 pipeline, and using up-to-date modeling techniques (detailed model compari-535 son, validation, and hierarchical Bayesian model fitting where possible). This 536 should encourage the field of computational modeling to study the external 537 factors that drive such inconsistencies, and are currently undescribed by RL 538 methods, with more rigor. 539

540 3.1. Limitations

One limitation of our results is that regression analyses might be con-541 taminated by parameter cross-correlations (in sections 2.1.2, 2.1.3, 2.1.4), 542 which would reflect modeling limitations (fewer true degrees of freedom than 543 model parameters), and not necessarily shared cognitive processes. For ex-544 ample, parameters α and β are mathematically related in the regular RL 545 modeling framework [26, 29], and we observed significant correlations be-546 tween parameters within tasks for two of our three tasks (suppl. Fig. D.7). 547 This indicates that caution is required when interpreting correlation results. 548 However, correlations were also present between tasks (suppl. Fig. D.8), 540 suggesting that within-model trade-offs were not the only explanation for 550 shared variance, and that shared cognitive processes likely also played a role. 551 Furthermore, correlations between parameters within models are frequent 552 in the existing literature, and do not prevent researchers from interpreting 553 parameters—in this sense, the existence of similar correlations in our study 554 allows us to address the question of generalizability and interpretability in 555 similar circumstances as in the existing literature. 556

557 3.2. Moving Forward

With this research, we do not intend to undermine RL modeling, but to improve its quality. Computational model parameters potentially provide highly valuable insights into (neuro)cognitive processing—we just need to refrain from assuming that the identified processes are necessarily and inherently specific, distinct, and "theoretically meaningful" [14] (interpretable). Parameters with the same names also do not automatically transfer between

tasks or models, and are less interchangeable than we often implicitly assume [28]. At the same time, the behavioral features that are captured by parameters seem to generalize well between tasks.

In the long term, we need to understand why RL parameters differ between tasks. We suggest three potential, not mutually exclusive answers:

1. **Optimality**. Variance in RL parameters may reflect how participants 569 adapt their behavior to task demands, an explanation proposed by 570 [13]. For example, participants might tune learning rates to task char-571 acteristics (e.g., adopting lower learning rates in stable than volatile 572 contexts [55]), rather than learning rates reflecting intrinsic "settings" 573 (e.g., 10-year-olds having a learning rate of 20%; 16-year-olds of 40%). 574 An optimality-based view would also explain why learning rates dif-575 fer between deterministic and stochastic tasks, which require different 576 amounts of behavioral change in response to feedback, to reach opti-577 mal performance. Age differences can potentially be explained because 578 optimal settings likely differ between ages because they interact with 579 different environments, or because different ages might have different 580 capacities to shift internal settings when shifting from task to task. 581 More research is needed, however, to determine whether parameter op-582 timality can explain all inconsistencies in the literature. For example, 583 our finding that participants showed the most optimal parameter val-584 ues in the intermediate age range in task B [49], whereas optimality 585 increased monotonously with age in tasks A and C [35, 50], is difficult 586 to reconcile with this view. 587

2. Modulatory processes. RL Parameters may vary as a function of 588 modulatory processes that are not well captured in current RL models. 589 Modulatory processes have been described in cognition and neurobiol-590 ogy and likely serve to shift functional outputs (e.g., hunger increasing 591 motivation) [56, 57, 58]. Some modulatory processes reflect the inte-592 gration of external contextual information: for example, uncertainty 593 affects dopamine neuron firing [59, 60, 61]. In addition, environments 594 with different degrees of uncertainty have been shown to elicit different 595 learning rates [55]. It is thus possible that neuromodulation by task 596 uncertainty could modulate RL processes, reflected in RL parameters. 597 In our data, feedback stochasticity and task volatility likely contribute 598 to such uncertainty-related modulation. However, other factors like 599 task similarity (low versus high), task characteristics (e.g., volatility 600

[55, 49], feedback stochasticity, memory load [35, 52], feedback valence
and conditioning type [24]), and choice of model parameters (e.g., forgetting [35, 50], counter-factual learning [49], negative and positive
learning rates [34, 62, 63]), also seem to affect RL parameters, but are
independent of uncertainty. More research is needed to systematically
investigate the factors the contribute to modulatory processes, and how
they impact cognition and computation.

3. RL processes are multifaceted. RL Parameters capture a multi-608 tude of separate processes, whose composition differs across tasks (Fig. 609 5B). RL algorithms are framed in the most general way to allow ap-610 plication to a wide range of contexts, including AI, neuroscience, and 611 psychology [26, 28]. As behavioral models, their use has spanned be-612 haviors from simple conditioning [1, 38] to complex decision making 613 [4, 6, 7, 8, 9, 22, 64], meaning that the same parameters capture cog-614 nitive processes that vary considerably in type and complexity: Pro-615 cesses can include the slow acquisition of implicit preferences [1], long-616 term memory for such preferences [65], quick recognition of contingency 617 switches [49, 66], selection of abstract high-level strategies [7, 9, 67], 618 meta-learning [6], habitual and goal-directed decision making [5], work-619 ing memory or episodic memory-guided choice [52, 68, 69], and many 620 others. This list alone outnumbers the list of typical RL model param-621 eters, suggesting that RL parameters capture different (combinations 622 of) cognitive processes depending on the paradigm. Similar arguments 623 have also been made for behavioral analyses [70]. 624

625 3.3. Conclusion

Our research has important implications for fields that focus on individ-626 ual differences, including developmental and clinical computational research. 627 The current study should be seen as a proof of concept that many contextual 628 factors impact computational modeling, and larger studies will be necessary 629 to quantify these effects and determine their structure. Other areas of model-630 ing besides the RL framework should be subjected to a similar investigation. 631 It is possible, for example, that generalizability differs for sequential sam-632 pling [71, 72], Bayesian inference [49, 73, 74], model-based versus model-free 633 RL [48, 75, 76], or other models. 634

In sum, our results suggest that relating model parameters to cognitive constructs and real-world behavior might require us to carefully account for task variables, and environmental variability in general. This ties into the

bigger picture of understanding how neurocognitive processes are shared be-638 tween tasks [77], and reflects a larger pattern of realization in psychology 639 that we cannot objectively assess an individual's cognitive processing while 640 ignoring subjective context. We have shown that in lab studies, different task 641 contexts recruit different system settings within an individual; similarly, real-642 life environment, its changes during development, and past environment [78] 643 may also modulate which cognitive processes we recruit. Heightened aware-644 ness and systematic study of contextual variables will therefore be a valuable 645 future investment as we work to measure and accommodate diversity in cog-646 nitive processes. 647

648 4. Methods

649 4.1. Study Design

Our sample of 291 participants was balanced between females and males. 650 and all ages (8-30 years) were represented equally (Fig. 1A, left). Partici-651 pants completed four computerized tasks, questionnaires, and a saliva sample 652 during the 1-2 hour lab visit (see section 4.3). To reduce noise, we excluded 653 participants based on task-specific performance criteria (see section 4.2). Due 654 to worse performance, more younger than older participants were excluded, 655 which is a caveat for the interpretation of age effects (note however that 656 these exclusions cannot account for the observed age effects but act against 657 them; Fig. 1A). Our tasks—A ("Butterfly task" [50, 79]), B ("Probabilis-658 tic Switching" [66, 49]), and C ("Reinforcement learning-Working memory" 659 [35, 52])—were all classic reinforcement learning tasks: on each trial, partic-660 ipants chose between several actions in an effort to earn rewards, which were 661 presented as binary feedback (win/point or lose/no point) after each choice. 662

The tasks varied on several common dimensions (Fig. 1B), which have 663 been related to discrepancies in behavioral and neurocognitive results in the 664 literature [24, 36, 37]. For example, in one task (task C), positive feedback 665 was deterministic, such that every correct action led to a positive outcome, 666 whereas in the two other tasks (tasks A and B), positive feedback was stochas-667 tic, such that some correct actions led to positive and others to negative 668 outcomes. A different set of two tasks (B and C) provided diagnostic pos-669 itive feedback, such that every positive outcome indicated a correct action, 670 whereas in the third (A), positive feedback was non-diagnostic, such that 671 positive outcomes could indicate both correct and incorrect actions. Two 672 tasks (A and C) presented several different stimuli/states for which correct 673

actions had to be learned, whereas the third (B) only presented a single 674 one. Overall, task A shared more important similarities with both tasks B 675 and C than either of these shared with each other, allowing us to explore 676 whether task similarity played a role in parameter generalizability and in-677 terpretability. A comprehensive list of task differences is shown in Fig. 1B, 678 and each task is described in more detail in section 4.4. Section Appendix 679 B explains the most prominent findings of each task individually, and shows 680 several behavioral measures over age. 681

682 4.2. Participant Sample

683 4.2.1. Sample Overview

All procedures were approved by the Committee for the Protection of Hu-684 man Subjects at the University of California, Berkeley. We tested 312 partic-685 ipants: 191 children and adolescents (ages 8-17) and 55 adults (ages 25-30) 686 were recruited from the community and completed a battery of computer-687 ized tasks, questionnaires, and saliva samples; 66 university undergraduate 688 students (aged 18-50) completed the four tasks as well, but not the question-680 naires or saliva sample. Community participants of all ages were prescreened 690 for the absence of present or past psychological and neurological disorders; 691 the undergraduate sample indicated the absence of these. Compensation for 692 community participants consisted in \$25 for the 1-2 hour in-lab portion of 693 the experiment and \$25 for completing optional take-home saliva samples; 694 undergraduate students received course credit for participation in the 1-hour 695 study. 696

697 4.2.2. Participant Exclusion

Two participants from the undergraduate sample were excluded because 698 they were older than 30, and 7 were excluded because they failed to indicate 699 their age. This led to a sample of 191 community participants under 18, 57 700 undergraduate participants between the ages of 18-28, and 55 community 701 participants between the ages of 25-30. Of the 191 participants under 18. 702 184 completed task B, and 187 completed tasks A and C. Reasons for not 703 completing a task included getting tired, running out of time, and technical 704 All 57 undergraduate participants completed tasks B and C and issues. 705 55 completed task A. All 55 community adults completed tasks B and A, 706 and 45 completed task C. Appropriate exclusion criteria were implemented 707 separately for each task to exclude participants who failed to pay attention 708 and who performed critically worse than the remaining sample (for task A, 709

see [50]; task B [49]; task C [35]). Based on these criteria, 5 participants
under the age of 18 were excluded from task B, 10 from task A, and none
from task C. One community adult participant was excluded from task A,
but no adult undergraduates or community participants were excluded from
tasks B or C.

Because this study related the results across all three tasks, we only 715 included participants who were not excluded in any task, leading to a final 716 sample of 143 participants under the age of 18 (male: 77; female: 66), 51 717 undergraduate participants (male: 17; female: 34), and 53 adults from the 718 community (male: 25; female: 28), for a total of 247 participants (male: 119; 719 female: 128). We entirely excluded the fourth task of our study from the 720 current analysis, which was modeled after a rodent task and used in humans 721 for the first time [80], because the applied performance criterion led to the 722 exclusion of the majority of our developmental sample. We split participants 723 into quantiles based on age, which were calculated separately within each 724 sex. 725

726 4.3. Testing Procedure

After entering the testing room, participants under 18 years and their 727 guardians provided informed assent and permission: participants over 18 728 provided informed consent. Guardians and participants over 18 filled out 729 a demographic form. Participants were led into a quiet testing room in view 730 of their guardians, where they used a video game controller to complete 731 four computerized tasks. The first task was called "4-choice" and assessed 732 reversal learning in an environment with 4 different choice options, with a 733 duration of approximately 5 minutes (designed after [80]). This task was 734 excluded from the current analysis (see section 4.2.2). The second task was 735 C ("Reinforcement learning-Working memory") and took about 25 minutes 736 to complete [52, 35]. After the second task, participants between the ages of 737 8-17 provided a saliva sample (for details, see [35]) and took a snack break (5-738 10 minutes). After that, participants completed task A ("Butterfly task"). 739 which took about 15 minutes [79, 50], and task B ("Probabilistic Switch-740 ing"), which took about 10 minutes to complete [49]. At the conclusion of 741 the tasks, participants between 11 and 18 completed the Pubertal Develop-742 ment Scale (PDS [81]) and were measured in height and weight. Participants 743 were then compensated with \$25 Amazon gift cards. The PDS questionnaire 744 and saliva samples were administered to investigate the role of pubertal mat-745 uration on learning and decision making. Pubertal analyses are not the focus 746

of the current study and will be or have reported elsewhere [35, 49, 50]. For
methodological details, refer to [35]. The entire lab visit took 60-120 minutes,
depending on the participant.

750 4.4. Task Design

751 4.4.1. Task A ("Butterfly task")

The goal of task A was to collect as many points as possible, by guessing correctly which of two flowers was associated with each of four butterflies. Correct guesses were rewarded with 70% probability, and incorrect guesses with 30%. The task contained 120 trials (30 for each butterfly) that were split into 4 equal-sized blocks, and took between 10-20 minutes to complete. More detailed information about methods and results can be found in [50].

758 4.4.2. Task B ("Probabilistic Switching")

The goal of task B was to collect golden coins, which were hidden in two 759 green boxes. The task could be in one of two states: "Left box is correct" 760 or "Right box is correct". In the former, selecting the left box led to reward 761 in 75% of trials, while selecting the right box never led to a reward (0%). 762 Several times throughout the task, task contingencies changed unpredictably 763 and without notice (after participants had reached a performance criterion 764 indicating they had learned the current state), and the task switched states. 765 Participants completed 120 trials of this task (2-9 reversals), which took ap-766 proximately 5-15 minutes. For more information and additional task details, 767 refer to [49]. 768

769 4.4.3. Task C ("Reinforcement Learning-Working Memory")

The goal of task C was to collect as many points as possible by pressing 770 the correct key for each stimulus. Pressing the correct key deterministically 771 led to reward, and the correct key for a stimulus never changed. Stimuli 772 appeared in blocks that varied in the number of different stimuli, with set 773 sizes ranging from 2-5. In each block, each stimulus was presented 12-14 774 times, for a total of 13 * set size trials per block. Three blocks were presented 775 for set sizes 2-3, and 2 blocks were presented for set sizes 4-5, for a total of 10 776 blocks. The task took between 15-25 minutes to complete. For more details, 777 as well as a full analysis of this dataset, refer to [35]. 778

779 4.5. Computational Models

For all tasks, we used RL theory to model how participants adapted their behavior in order to maximize reward. RL models assume that agents learn a policy $\pi(a|s)$ that determines (probabilistically) which action a to take in each state s of the world [26]. Here and in most cognitive RL models, this policy is based on action values Q(a|s), i.e., the values of each action a in each state s. Agents learn action values by observing the reward outcomes, r_t , of their actions at each time step t. Learning consists in updating existing action values $Q_t(a|s)$ using the "reward prediction error", the difference between the expected reward $Q_t(a|s)$ and the actual reward r_t :

$$Q_{t+1}(a|s) = Q_t(a|s) + \alpha(r_t - Q_t(a|s))$$

How much a learner weighs past action value estimates compared to new out-780 comes is determined by parameter α , the learning rate. Small learning rates 781 favor past experience and lead to stable learning over long time horizons, 782 while large learning rates favor new outcomes and allow for faster and more 783 flexible changes, focusing on shorter time horizons. With enough time and 784 in a stable environment, the RL updating scheme guarantees that value es-785 timates will reflect the environment's true reward probabilities, and thereby 786 allow for optimal long-term choices [26]. 787

In order to choose actions, most cognitive RL models use a (noisy) "softmax" function to translate action values Q(a|s) into policies p(a|s):

$$p(a_i|s) = \frac{exp(\beta \ Q(a_i|s))}{\sum_{a_i \in A} exp(\beta \ Q(a_j|s))}$$

A refers to the set of all available actions (tasks A and B have 2 actions, 788 task C has 3), and a_i and a_j to individual actions within the set. How 789 deterministically versus noisily this translation is executed is determined by 790 exploration parameters β , also called inverse decision temperature, and/or 791 ϵ , the decision noise (see below). Small decision temperatures $\frac{1}{\beta}$ favor the 792 selection of the highest-valued actions, enabling exploitation, whereas large 793 decision temperatures select actions of low and high values more evenly, 794 enabling exploration. Parameter ϵ adds undirected noise to action selection, 795 selecting random action with a small probability ϵ on each trial. 796

⁷⁹⁷ Besides α , β , and noise, cognitive RL models often include additional ⁷⁹⁸ parameters to better fit empirical behavior in humans or animals. Com-⁷⁹⁹ mon choices include Forgetting—a consistent decay of action values back to

baseline—, and Persistence—the tendency to repeat the same action independent of outcomes, a parameter also known as sticky choice or perseverance [63]. In addition, cognitive models often differentiate learning from positive versus negative rewards, splitting learning rate α into two separate parameters α_+ and α_- , which are applied to only positive and only negative outcomes, respectively [34, 40, 82, 83, 84, 85, 86, 87, 88]. The next paragraphs introduce these parameters in detail.

In task A, the best fitting model included a forgetting mechanism, which was implemented as a decay in Q-values applied to all action values of the three stimuli (butterflies) that were not shown on the current trial:

$$Q_{t+1}(a|s) = (1-f) * Q_{t+1}(a|s) + f * 0.5.$$

⁸⁰⁷ The free parameter 0 < f < 1 reflects individuals' tendencies to forget.

In task B, free parameter P captured choice persistence, which biased choices on the subsequent trial toward staying (P > 0) or switching (P < 0). P modifies action values Q(a|s) into Q'(a|s), as follows:

$$Q'_t(a|s) = Q_t(a|s) + P \iff a_t = a_{t-1}$$
$$Q'_t(a|s) = Q_t(a|s) \iff a_t \neq a_{t-1}$$

In addition, the model of task B included counter-factual learning parameters α_{C+} and α_{C-} , which added counter-factual updates based on the inverse outcome and affected the non-chosen action. For example, after receiving a positive outcome (r = 1) for choosing left (a), counter-factual updating would lead to an "imaginary" negative outcome $(\bar{r} = 0)$ for choosing right (\bar{a}) .

$$Q_{t+1}(\bar{a}|s) = Q_t(\bar{a}|s) + \alpha_{C+}(\bar{r} - Q_t(\bar{a}|s)) \iff r = 1$$
$$Q_{t+1}(\bar{a}|s) = Q_t(\bar{a}|s) + \alpha_{C-}(\bar{r} - Q_t(\bar{a}|s)) \iff r = 0$$

⁸⁰⁸ \bar{a} indicates the non-chosen action, and \bar{r} indicates the inverse of the received ⁸⁰⁹ outcome, $\bar{r} = 1 - r$. The best model fits were achieved with $\alpha_{C+} = \alpha_+$ and ⁸¹⁰ $\alpha_{C-} = \alpha_-$, so counter-factual learning rates are not reported in this paper.

In tasks A and B, positive and negative learning rates are differentiated in the following way:

$$Q_{t+1}(a|s) = Q_t(a|s) + \alpha_+(r_t - Q_t(a|s)) \iff r_t = 1$$
$$Q_{t+1}(a|s) = Q_t(a|s) + \alpha_-(r_t - Q_t(a|s)) \iff r_t = 0$$

In the best model for task A, only α_+ was a free parameter, while α_- was fixed to 0. In task C, α_- was a function of α_+ , such that $\alpha_- = b * \alpha_+$, where b is the neglect bias parameter that determines how much negative feedback is neglected compared to positive feedback. Throughout the paper, we report $\alpha_- = b * \alpha_+$ for task C.

In addition to an RL module, the model of task C included a workingmemory module with perfect recall of recent outcomes, but subject to forgetting and capacity limitations. Perfect recall was modeled as an RL process with learning rate $\alpha_{WM+} = 1$ that operated on working-memory weights W(a|s) rather than action values. On trials with positive outcomes (r = 1), the model reduces to:

$$W_{t+1}(a|s) = r_t$$

On trials with negative outcomes (r = 0), multiplying $\alpha_{WM+} = 1$ with the neglect bias b leads to potentially less-than perfect memory:

$$W_{t+1}(a|s) = W_t(a|s) + b * (r_t - W_t(a|s))$$

Working-memory weights W(a|s) were transformed into action policies $p_{WM}(a|s)$ in a similar way as RL weights Q(a|s) were transformed into action probabilities $p_{RL}(a|s)$, using a softmax transform combined with undirected noise:

$$p(a_i|s) = (1-\epsilon) * \frac{\exp(\beta \ Q(a_i|s))}{\sum_{a_i \in a} \exp(\beta \ Q(a_j|s))} + \epsilon * \frac{1}{|a|}$$

|a| = 3 is the number of available actions and $\frac{1}{|a|}$ is the uniform policy over these actions; ϵ is the undirected noise parameter.

Forgetting was implemented as a decay in working-memory weights W(a|s) (but not RL Q-values):

$$W_{t+1}(a|s)_{t+1} = (1-f) * W_t(a|s)_t + f * \frac{1}{3}$$

Capacity limitations of working memory were modeled as an adjustment in the weight w of $p_{WM}(a|s)$ compared to $p_{RL}(a|s)$ in the final calculation of action probabilities p(a|s):

$$w = \rho * (min(1, \frac{K}{ns}))$$
$$p(a|s) = w * p_{WM}(a|s) + (1 - w) * p_{RL}(a|s)$$

The free parameter ρ is the individual weight of working memory compared to RL, *ns* indicates a block's stimulus set size, and *K* captures individual differences in working-memory capacity.

We fitted a separate RL model to each task, using state-of-the-art meth-821 ods for model construction, fitting, and validation [30, 27]. Models for tasks 822 A and B were fitted using hierarchical Bayesian methods with Markov-Chain 823 Monte-Carlos sampling, which is an improved method compared to maximum 824 likelihood that leads to better parameter recovery, amongst other advantages 825 [89, 90, 91]. The model for task C was fitted using classic non-hierarchical 826 maximum-likelihood because model parameter K is discrete, which renders 827 hierarchical sampling less tractable. In all cases, we verified that the model 828 parameters were recoverable by the selected model-fitting procedure, and 829 that the models were identifiable. Details of model-fitting procedures are 830 provided in the original publications [35, 49, 50]. 831

For additional details on any of these models, as well as detailed model comparison and validation, the reader is referred to the original publications (task A: [50]; task B: [49]; task C: [35]).

4.6. Principal Component Analysis (PCA)

The PCA in section 2.2.1 included 15 model parameters (α_{+} and noise/exploration 836 in each task; Forgetting and α_{-} in two tasks; Persistence in task B; four 837 working-memory parameters in task C; see section 4.5) and 39 model-free 838 features, including simple behavioral features (e.g., overall performance, re-839 action times, tendency to switch), results of behavioral regression models 840 (e.g., effect of stimulus delay on accuracy), and the model parameters of an 841 alternative Bayesian inference model in task B. All behavioral features, in-842 cluding their development over age, are described in detail in Appendix C 843 and suppl. Fig. B.6B. For simplicity, section 2.2.1 focused on the first three 844 PCs only; the weights, explained variance, and age trajectories of remaining 845 PCs are shown in suppl. Fig. D.10. 846

PCA is a statistical tool that decomposes the variance of a dataset into 847 so-called "principal components" (PCs). PCs are linear combinations of a 848 dataset's original features (e.g., response times, accuracy, learning rates), 840 and explain the same variance in the dataset as the original features. The 850 advantage of PCs is that they are orthogonal to each other and therefore 851 capture independent aspects of the data. In addition, subsequent PCs ex-852 plain subsequently less variance, such that selecting just the top PCs of a 853 dataset retains the bulk of the variance and the ability to reconstruct the 854

dataset up to random noise. When using this approach, it is important to understand which concept each PC captures. So-called factor loadings, the original features' weights on each PC, can provide this information.

PCA performs a *change of basis*: Instead of describing the dataset using 858 the original features (in our case, 54 behaviors and model parameters), it cre-859 ates new features, PCs, that are linear combinations of the original features 860 and capture the same variance, but are orthogonal to each other. PCs are 861 created by eigendecomposition of the covariance matrix of the dataset: the 862 eigenvector with the largest eigenvalue shows the direction in the dataset in 863 which most variance occurs, and represents the first PC. Eigenvectors with 864 subsequently smaller eigenvalues form subsequent PCs. PCA is related to 865 Factor analysis, and often used for dimensionality reduction. In this case, 866 only a small number of PCs is retained whereas the majority is discarded, in 867 an effort to retain most variance with a reduced number of features. 868

We highlight the most central behavioral features here; more detail is pro-869 vided in Appendix A and Appendix C. Response to feedback was assessed 870 using features "Win-stay" (percentage of trials in which a rewarded choice 871 was repeated), and "Lose-stay" (percentage of trials in which a non-rewarded 872 choice was repeated). For task B, we additionally included "Win-lose-stay" 873 tendencies, which is the proportion of trials in which participants stay after a 874 winning trial that is followed by a losing trial. This is an important measure 875 for this task because the optimal strategy required staying after single losses. 876

We also included behavioral persistence measures in all tasks. In tasks A and C, these included a measure of action repetition (percentage of trials in which the previous key was pressed again, irrespective of the stimulus and feedback) and choice repetition (percentage of trials in which the action was repeated that was previously selected for the same stimulus, irrespective of feedback). In task B, both measures were identical because every trial presents the same stimulus.

We further included task-specific measures of performance. In task A, 884 these were: the average accuracy for the first three presentations of each 885 stimulus, reflecting early learning speed; and the asymptote, intercept, and 886 slope of the learning progress in a regression model predicting performance 887 (for details about these measures, see [50]. In task B, task-specific mea-888 sures of performance included the number of reversals (because reversals 889 were performance-based); and the average number of trials to reach criterion 890 after a switch. In tasks A and C, we also included a model-independent 891 measure of forgetting. In task A, this was the effect of delay on performance 892

in the regression model mentioned above. In task C, this was the effect of 893 delay in a similar regression model, which also included set size, the number 894 of previous correct choices, and the number of previous incorrect choices, 895 whose effects were also included. Lastly for task C, we included the slope 896 of accuracy and response times over set sizes, as measures of the effect of 897 set size on performance. For task B, we also included the difference between 898 early (first third of trials) and late (last third) performance as a measure of 899 learning. To avoid biases in the PCA toward any specific task, we included 900 equal numbers of behavioral features for each task. 901

To facilitate the interpretation of PC2 and PC3, we normalized the loadings (PCA weights) of each feature (behavioral and model parameter) with respect to PC1, flipping the loadings of all features in PC2 and PC3 that loaded negatively on PC1. This step ensured that the directions of factor loadings on PC2 and PC3 were interpretable in the same way for all features, irrespective of their role for task performance, and revealed the encoding of task contrasts.

909 4.7. Ridge Regression

In sections 2.2.2 and 2.2.3, we use regularized, cross-validated Ridge re-910 gression to determine whether parameters captured overlapping variance, 911 which would point to an overlap in cognitive processes. We used Ridge re-912 gression to avoid problems that would be caused by overfitting when using 913 regular regression models. Ridge regression regularizes regression weight pa-914 rameters w based on their L2-norm. Regular regression identifies a vector 915 of regression weights w that minimize the linear least squares $||y - wX||_2^2$. 916 Here, $||a||_2^2 = \sqrt{\sum_{a_i \in x} a_i^2}$ is the L2-norm of a vector a, vector y represents 917 the outcome variable (in our case, a vector of parameters, one fitted to each 918 participant), matrix X represents the predictor variables (in our case, either 919 several behavioral features for each participant [2.2.2], or several parame-920 ters fitted to each participant 2.2.3), and vector w represents the weights 921 assigned to each feature in X (in our case, the weight assigned to each pre-922 dicting behavioral pattern or each predicting parameter). 923

When datasets are small compared to the number of predictors in a regression model, *exploding* regression weights w can lead to overfitting. Ridge regression avoids this issue by not only minimizing the linear least squares like regular regression, but also the L2 norm of weights w, i.e., by minimizing $||y-wX||_2^2 + \alpha * ||w||_2^2$. Parameter α is a hyper-parameter of Ridge regression,

which needs to be chosen by the experimenter. To avoid bias in the selection 920 of α , we employed repeated cross-validated grid search. At each iteration of 930 this procedure, we split the dataset into a predetermined number $s \in [2, 3, 3]$ 931 ..., 8] of equal-sized folds, and then fitted a Ridge regression to each fold, us-932 ing values of $\alpha \in [0, 10, 30, 50, 100, 300, \dots, 10,000, 100,000, 1,000,000]$. For 933 each s, we determined the best value of α based on cross-validation between 934 folds, using the amount of explained variance, R^2 , as the selection criterion. 935 To avoid biases based on the random assignment of participants into folds, 936 we repeated this procedure n = 100 times for each value of α . To avoid biases 937 due to the number of folds, the entire process was repeated for each s, and 938 the final value of s was selected based on \mathbb{R}^2 . We used the python package 939 "scikit learn" [92] to implement the procedure. 940

We conducted three models per parameter to determine the relations be-941 tween parameters: predicting each parameter from all the parameters of each 942 of the other two tasks (2 models); and predicting each parameter from all 943 parameters of both other tasks combined (1 model; Fig. 4A). We conducted 944 the same three models per parameter to determine the relations between pa-945 rameters and behaviors, predicting each parameter from behavioral features 946 of the other tasks (Fig. 4A). In addition, we conducted a fourth model for 947 behaviors, predicting each parameter from the behaviors of all three tasks 948 combined, to assess the contributions of all behaviors to each parameter (Fig. 949 4C). Meta-parameters s and α were allowed to differ (and differed) between 950 models. The final values of R^2 (Fig. 4B and 4D) and the final regression 951 weights w (Fig. 4A and 4C; Table 6) were determined by refitting the winning 952 model. 953

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965 References

- [1] W. Schultz, P. Dayan, P. R. Montague, A Neural Substrate
 of Prediction and Reward, Science 275 (5306) (1997) 1593–1599.
 doi:10.1126/science.275.5306.1593.
- J. P. O'Doherty, P. Dayan, J. Schultz, R. Deichmann, K. Friston, R. J.
 Dolan, Dissociable Roles of Ventral and Dorsal Striatum in Instrumental Conditioning, Science 304 (5669) (2004) 452–454, publisher:
 American Association for the Advancement of Science Section: Report. doi:10.1126/science.1094285.
- 974 URL https://science.sciencemag.org/content/304/5669/452
- J. Gläscher, A. N. Hampton, J. P. O'Doherty, Determining a role for ventromedial prefrontal cortex in encoding action-based value signals during reward-related decision making, Cerebral Cortex (New York, N.Y.: 1991) 19 (2) (2009) 483-495. doi:10.1093/cercor/bhn098.
- [4] J. Ribas Fernandes, A. Solway, C. Diuk, J. T. McGuire, A. G.
 Barto, Y. Niv, M. Botvinick, A Neural Signature of Hierarchical Reinforcement Learning, Neuron 71 (2) (2011) 370–379. doi:10.1016/j.neuron.2011.05.042.
- [5] N. Daw, S. Gershman, B. Seymour, P. Dayan, R. Dolan, Model-Based
 Influences on Humans' Choices and Striatal Prediction Errors, Neuron
 69 (6) (2011) 1204–1215. doi:10.1016/j.neuron.2011.02.027.
- [6] J. X. Wang, Z. Kurth-Nelson, D. Kumaran, D. Tirumala, H. Soyer,
 J. Z. Leibo, D. Hassabis, M. Botvinick, Prefrontal cortex as a metareinforcement learning system, Nature Neuroscience 21 (6) (2018) 860–
 868. doi:10.1038/s41593-018-0147-8.
- [7] M. K. Eckstein, A. G. E. Collins, Computational evidence for hierarchically structured reinforcement learning in humans, Proceedings of the National Academy of Sciences 117 (47) (2020) 29381–29389. doi:10.1073/pnas.1912330117.
- URL https://www.pnas.org/content/117/47/29381
- ⁹⁹⁵ [8] M. Botvinick, Hierarchical reinforcement learning and decision ⁹⁹⁶ making, Current Opinion in Neurobiology 22 (6) (2012) 956–962.

997 998		doi:10.1016/j.conb.2012.05.008. URL http://linkinghub.elsevier.com/retrieve/pii/S0959438812000876
999 1000 1001	[9]	A. G. E. Collins, E. Koechlin, Reasoning, Learning, and Creativity: Frontal Lobe Function and Human Decision-Making, PLOS Biology 10 (3) (2012) e1001293. doi:10.1371/journal.pbio.1001293.
1002 1003 1004 1005 1006	[10]	D. M. Werchan, A. G. E. Collins, M. J. Frank, D. Amso, Role of Prefrontal Cortex in Learning and Generalizing Hierarchical Rules in 8-Month-Old Infants, The Journal of Neuroscience 36 (40) (2016) 10314-10322. doi:10.1523/JNEUROSCI.1351-16.2016. URL http://www.jneurosci.org/lookup/doi/10.1523/JNEUROSCI.1351-16.2016
1007 1008 1009 1010 1011	[11]	W. van den Bos, R. Bruckner, M. R. Nassar, R. Mata, B. Eppinger, Computational neuroscience across the lifespan: Promises and pitfalls, Developmental Cognitive Neuroscience (Oct. 2017). doi:10.1016/j.dcn.2017.09.008. URL http://linkinghub.elsevier.com/retrieve/pii/S1878929317301068
1012 1013 1014 1015 1016	[12]	F. Bolenz, A. M. F. Reiter, B. Eppinger, Developmental Changes in Learning: Computational Mechanisms and Social Influ- ences, Frontiers in Psychology 8, publisher: Frontiers (2017). doi:10.3389/fpsyg.2017.02048. URL https://www.frontiersin.org/articles/10.3389/fpsyg.2017.02048/full
1017 1018 1019 1020 1021	[13]	K. Nussenbaum, C. A. Hartley, Reinforcement learning across development: What insights can we draw from a decade of re- search?, Developmental Cognitive Neuroscience 40 (2019) 100733. doi:10.1016/j.dcn.2019.100733. URL http://www.sciencedirect.com/science/article/pii/S1878929319303202
1022 1023 1024 1025	[14]	Q. J. M. Huys, T. V. Maia, M. J. Frank, Computational psychiatry as a bridge from neuroscience to clinical applications, Nature neuroscience 19 (3) (2016) 404-413. doi:10.1038/nn.4238. URL https://www.ncbi.nlm.nih.gov/pmc/articles/PMC5443409/
1026 1027 1028 1029	[15]	R. A. Adams, Q. J. M. Huys, J. P. Roiser, Computational Psychiatry: towards a mathematically informed understanding of mental illness, Journal of Neurology, Neurosurgery & Psychiatry 87 (1) (2016) 53– 63, publisher: BMJ Publishing Group Ltd Section: Neuropsychiatry.

1030 1031		doi:10.1136/jnnp-2015-310737. URL https://jnnp.bmj.com/content/87/1/53
1032 1033 1034 1035	[16]	T. U. Hauser, GJ. Will, M. Dubois, R. J. Dolan, Annual Re- search Review: Developmental computational psychiatry, Jour- nal of Child Psychology and Psychiatry 60 (4) (2019) 412–426. doi:https://doi.org/10.1111/jcpp.12964.
1036 1037 1038 1039	[17]	WY. Ahn, J. R. Busemeyer, Challenges and promises for trans- lating computational tools into clinical practice, Current Opinion in Behavioral Sciences 11 (2016) 1-7. doi:10.1016/j.cobeha.2016.02.001. URL https://www.sciencedirect.com/science/article/pii/S2352154616300237
1040 1041 1042 1043 1044	[18]	L. Deserno, R. Boehme, A. Heinz, F. Schlagenhauf, Reinforcement Learning and Dopamine in Schizophrenia: Dimensions of Symptoms or Specific Features of a Disease Group?, Frontiers in Psychiatry 4, publisher: Frontiers (2013). doi:10.3389/fpsyt.2013.00172. URL https://www.frontiersin.org/articles/10.3389/fpsyt.2013.00172/full
1045 1046 1047 1048 1049	[19]	M. J. Frank, E. D. Claus, Anatomy of a decision: Striato-orbitofrontal interactions in reinforcement learning, decision making, and reversal., Psychological Review 113 (2) (2006) 300-326. doi:10.1037/0033- 295X.113.2.300. URL http://doi.apa.org/getdoi.cfm?doi=10.1037/0033-295X.113.2.300
1050 1051	[20]	Y. Niv, Reinforcement learning in the brain, Journal of Mathematical Psychology 53 (3) (2009) 139–154.
1052 1053 1054	[21]	D. Lee, H. Seo, M. W. Jung, Neural Basis of Reinforcement Learning and Decision Making, Annual review of neuroscience 35 (2012) 287– 308. doi:10.1146/annurev-neuro-062111-150512.
1055 1056 1057 1058	[22]	J. P. O'Doherty, S. W. Lee, D. McNamee, The structure of reinforcement-learning mechanisms in the human brain, Current Opinion in Behavioral Sciences 1 (2015) 94–100. doi:10.1016/j.cobeha.2014.10.004.
1059 1060 1061	[23]	P. W. Glimcher, Understanding dopamine and reinforcement learning: The dopamine reward prediction error hypothesis, Proceedings of the National Academy of Sciences 108 (3) (2011) 15647–15654.

1062 1063 1064 1065 1066	[24]	J. Garrison, B. Erdeniz, J. Done, Prediction error in reinforce- ment learning: A meta-analysis of neuroimaging studies, Neu- roscience & Biobehavioral Reviews 37 (7) (2013) 1297-1310. doi:10.1016/j.neubiorev.2013.03.023. URL http://www.sciencedirect.com/science/article/pii/S0149763413000833
1067 1068 1069 1070	[25]	P. Dayan, Y. Niv, Reinforcement learning: The Good, The Bad and The Ugly, Current Opinion in Neurobiology 18 (2) (2008) 185-196. doi:10.1016/j.conb.2008.08.003. URL https://linkinghub.elsevier.com/retrieve/pii/S0959438808000767
1071 1072	[26]	R. S. Sutton, A. G. Barto, Reinforcement Learning: An Introduction, 2nd Edition, MIT Press, Cambridge, MA; London, England, 2017.
1073 1074 1075 1076	[27]	 S. Palminteri, V. Wyart, E. Koechlin, The Importance of Falsification in Computational Cognitive Modeling, Trends in Cognitive Sciences (6) (2017) 425-433. doi:10.1016/j.tics.2017.03.011. URL https://linkinghub.elsevier.com/retrieve/pii/S1364661317300542
1077 1078 1079 1080 1081	[28]	M. K. Eckstein, L. Wilbrecht, A. G. E. Collins, What do Rein- forcement Learning Models Measure? Interpreting Model Parameters in Cognition and Neuroscience, psyArxivType: article (May 2021). doi:10.31234/osf.io/e7kwx. URL https://psyarxiv.com/e7kwx/
1082 1083 1084	[29]	N. D. Daw, Trial-by-trial data analysis using computational models, Decision Making, Affect, and Learning: Attention and Performance XXIII (2011). doi:10.1093/acprof:oso/9780199600434.003.0001.
1085 1086 1087 1088	[30]	R. C. Wilson, A. G. Collins, Ten simple rules for the computational modeling of behavioral data, eLife 8 (2019) e49547, publisher: eLife Sciences Publications, Ltd. doi:10.7554/eLife.49547. URL https://doi.org/10.7554/eLife.49547
1089 1090 1091 1092 1093	[31]	O. Guest, A. E. Martin, How Computational Modeling Can Force The- ory Building in Psychological Science, Perspectives on Psychological Science (2021) 1745691620970585Publisher: SAGE Publications Inc. doi:10.1177/1745691620970585. URL https://doi.org/10.1177/1745691620970585

1094 1095 1096 1097 1098	[32]	G. Blohm, K. P. Kording, P. R. Schrater, A How-to-Model Guide for Neuroscience, eNeuro 7 (1), publisher: Society for Neuro- science Section: Research Article: Methods/New Tools (Jan. 2020). doi:10.1523/ENEURO.0352-19.2019. URL https://www.eneuro.org/content/7/1/ENEURO.0352-19.2019
1099 1100 1101 1102	[33]	S. J. Gershman, Empirical priors for reinforcement learning models, Journal of Mathematical Psychology 71 (2016) 1-6. doi:10.1016/j.jmp.2016.01.006. URL http://www.sciencedirect.com/science/article/pii/S0022249616000080
1103 1104 1105 1106	[34]	T. Harada, Learning From Success or Failure? - Positiv- ity Biases Revisited, Frontiers in Psychology 11 (Jul. 2020). doi:10.3389/fpsyg.2020.01627. URL https://www.ncbi.nlm.nih.gov/pmc/articles/PMC7396482/
1107 1108 1109 1110 1111	[35]	 S. L. Master, M. K. Eckstein, N. Gotlieb, R. Dahl, L. Wilbrecht, A. G. E. Collins, Disentangling the systems contributing to changes in learning during adolescence, Developmental Cognitive Neuroscience 41 (2020) 100732. doi:10.1016/j.dcn.2019.100732. URL http://www.sciencedirect.com/science/article/pii/S1878929319303196
1112 1113 1114 1115	[36]	Z. A. Yaple, R. Yu, Fractionating adaptive learning: A meta-analysis of the reversal learning paradigm, Neuroscience & Biobehavioral Reviews 102 (2019) 85-94. doi:10.1016/j.neubiorev.2019.04.006. URL http://www.sciencedirect.com/science/article/pii/S0149763418308996
1116 1117 1118 1119 1120	[37]	X. Liu, J. Hairston, M. Schrier, J. Fan, Common and distinct networks underlying reward valence and processing stages: A meta-analysis of functional neuroimaging studies, Neuroscience and Biobehavioral Re- views 35 (5) (2011) 1219–1236. doi:10.1016/j.neubiorev.2010.12.012. URL https://www.ncbi.nlm.nih.gov/pmc/articles/PMC3395003/
1121 1122 1123 1124	[38]	W. Schultz, A. Dickinson, Neuronal Coding of Prediction Errors, Annual Review of Neuroscience 23 (1) (2000) 473-500. doi:10.1146/annurev.neuro.23.1.473. URL http://www.annualreviews.org/doi/10.1146/annurev.neuro.23.1.473
1125 1126 1127	[39]	A. G. E. Collins, J. K. Brown, J. M. Gold, J. A. Waltz, M. J. Frank, Working Memory Contributions to Reinforcement Learning Impairments in Schizophrenia, Journal of Neuroscience 34 (41) (2014)

1128 13747–13756, publisher: Society for Neuroscience Section: Articles. 1129 doi:10.1523/JNEUROSCI.0989-14.2014.

1130

URL https://www.jneurosci.org/content/34/41/13747

[40] A. H. Javadi, D. H. K. Schmidt, M. N. Smolka, Adolescents adapt more slowly than adults to varying reward contingencies, Journal of Cognitive Neuroscience 26 (12) (2014) 2670–2681. doi:10.1162/jocn_{a0}0677.

[44] S.-J. Blakemore, T. W. Robbins, Decision-making in the adolescent brain, Nature Neuroscience 15 (9) (2012) 1184–1191, number: 9 Publisher: Nature Publishing Group. doi:10.1038/nn.3177.

1134 URL http://www.nature.com/articles/nn.3177

[42] S. DePasque, A. Galván, Frontostriatal development and probabilistic reinforcement learning during adolescence, Neurobiology of Learning and Memory 143 (2017) 1–7. doi:10.1016/j.nlm.2017.04.009.

¹¹³⁸ URL http://www.sciencedirect.com/science/article/pii/S107474271730062X

[43] A. Mohebi, J. R. Pettibone, A. A. Hamid, J.-M. T. Wong, L. T. Vinson, T. Patriarchi, L. Tian, R. T. Kennedy, J. D. Berke, Dissociable dopamine dynamics for learning and motivation, Nature 570 (7759) (2019) 65–70, number: 7759 Publisher: Nature Publishing Group. doi:10.1038/s41586-019-1235-y.
URL https://www.nature.com/articles/s41586-019-1235-y

[44] W. R. Uttal, On some two-way barriers between models and mechanisms,
 Perception & Psychophysics 48 (2) (1990) 188–203. doi:10.3758/BF03207086.
 URL https://doi.org/10.3758/BF03207086

[45] B. Webb, Can robots make good models of biological behaviour?, Behav-¹¹⁴⁸ ioral and Brain Sciences 24 (6) (2001) 1033–1050, publisher: Cambridge ¹¹⁴⁹ University Press. doi:10.1017/S0140525X01000127.

¹¹⁵⁰ URL http://www.cambridge.org/core/journals/behavioral-and-brain-sciences/article/

[46] D. J. Navarro, Between the Devil and the Deep Blue Sea: Tensions Between ¹¹⁵² Scientific Judgement and Statistical Model Selection, Computational Brain

¹¹⁵³ & Behavior 2 (1) (2019) 28–34. doi:10.1007/s42113-018-0019-z.

¹¹⁵⁴ URL https://doi.org/10.1007/s42113-018-0019-z

[457] T. Yarkoni, The generalizability crisis, The Behavioral and brain science-¹¹⁵⁶ sPublisher: Behav Brain Sci (Dec. 2020). doi:10.1017/S0140525X20001685.

1157 URL https://pubmed.ncbi.nlm.nih.gov/33342451/

- [48] V. M. Brown, J. Chen, C. M. Gillan, R. B. Price, Improving the Reliability
- ¹¹⁵⁹ of Computational Analyses: Model-Based Planning and Its Relationship
- ¹¹⁶⁰ With Compulsivity, Biological Psychiatry: Cognitive Neuroscience and
- ¹¹⁶¹ Neuroimaging 5 (6) (2020) 601–609. doi:10.1016/j.bpsc.2019.12.019.
- $\label{eq:urb} {\rm URL\ https://www.sciencedirect.com/science/article/pii/S2451902220300161}$
- M. K. Eckstein, S. L. Master, R. E. Dahl, L. Wilbrecht, A. G. E. Collins,
 ¹¹⁶⁴ Understanding the Unique Advantage of Adolescents in Stochastic, Volatile
 ¹¹⁶⁵ Environments: Combining Reinforcement Learning and Bayesian Inference,
 ¹¹⁶⁶ bioRxiv (2020) 2020.07.04.187971Publisher: Cold Spring Harbor Laboratory
- ¹¹⁶⁷ Section: New Results. doi:10.1101/2020.07.04.187971.
- ¹¹⁶⁸ URL https://www.biorxiv.org/content/10.1101/2020.07.04.187971v1
- [50] L. Xia, S. Master, M. Eckstein, L. Wilbrecht, A. G. E. Collins, Learning under uncertainty changes during adolescence, in: Proceedings of the Cognitive Science Society, 2020.
- [574] M. D. Lee, How cognitive modeling can benefit from hierarchical
 ¹¹⁷³ Bayesian models, Journal of Mathematical Psychology 55 (1) (2011)
 ¹¹⁷⁴ 1–7. doi:10.1016/j.jmp.2010.08.013.
- 1175 URL https://linkinghub.elsevier.com/retrieve/pii/S0022249610001148
- [52] A. G. E. Collins, M. J. Frank, How much of reinforcement learning is working memory, not reinforcement learning? A behavioral, computational, and neurogenetic analysis: Working memory in reinforcement learning, European Journal of Neuroscience 35 (7) (2012) 1024–1035. doi:10.1111/j.1460-9568.2011.07980.x.
- [53] L. H. Somerville, S. F. Sasse, M. C. Garrad, A. T. Drysdale, N. Abi Akar,
 C. Insel, R. C. Wilson, Charting the expansion of strategic exploratory behavior during adolescence, Journal of Experimental Psychology: General 146 (2)
 (2017) 155–164, place: US Publisher: American Psychological Association.
 doi:10.1037/xge0000250.
- [54] A. Gopnik, Childhood as a solution to explore–exploit tensions, Philosophical Transactions of the Royal Society B: Biological Sciences 375 (1803)
 (2020) 20190502, publisher: Royal Society. doi:10.1098/rstb.2019.0502.
- URL https://royalsocietypublishing.org/doi/10.1098/rstb.2019.0502

- 555 T. E. J. Behrens, M. W. Woolrich, M. E. Walton, M. F. S. Rushworth, Learning the value of information in an uncertain world, Nature Neuroscience 1191 10 (9) (2007) 1214–1221. doi:10.1038/nn1954. 1192
- URL https://www.nature.com/articles/nn1954 1193
- 156 K. C. Berridge, The debate over dopamine's role in reward: the case for incentive salience, Psychopharmacology 191 (3) (2007) 391–431. 1195 doi:10.1007/s00213-006-0578-x. 1196
- URL https://doi.org/10.1007/s00213-006-0578-x 1197
- A. J. Yu, P. Dayan, Uncertainty, Neuromodulation, and Attention, Neuron 1598 46 (4) (2005) 681–692. doi:10.1016/j.neuron.2005.04.026.
- 1199
- URL http://www.sciencedirect.com/science/article/pii/S0896627305003624 1200
- S. Bouret, S. J. Sara, Network reset: a simplified overarching theory of locus 1508 coeruleus noradrenaline function, Trends in Neurosciences 28 (11) (2005) 1202 574–582. doi:10.1016/j.tins.2005.09.002. 1203
- [59] S. J. Gershman, Dopamine, Inference, and Uncertainty, Neural Computation 29 (12) (2017) 3311–3326. doi:10.1162/neco_{a0}1023. URLhttp://www.mitpressjournals.org/doi/abs/10.1162/neco_{a0}1023C. K.Starkweather, S. 1204
- [60] S. J. Gershman, N. Uchida, Believing in dopamine, Nature Reviews Neuroscience 20 (11) (2019) 703–714, number: 11 Publisher: Nature Publishing 1206 Group. doi:10.1038/s41583-019-0220-7.
- 1207
- URL https://www.nature.com/articles/s41583-019-0220-7 1208
- K. Katahira, The statistical structures of reinforcement learning with 1602
- asymmetric value updates, Journal of Mathematical Psychology 87 (2018) 1210
- 31-45. doi:10.1016/j.jmp.2018.09.002. 1211

URL http://www.sciencedirect.com/science/article/pii/S0022249617302407 1212

- M. Sugawara, K. Katahira, Dissociation between asymmetric value up-1613 dating and perseverance in human reinforcement learning, Scientific Re-1214 ports 11 (1) (2021) 3574, number: 1 Publisher: Nature Publishing Group. 1215 doi:10.1038/s41598-020-80593-7. 1216
- URL https://www.nature.com/articles/s41598-020-80593-7 1217

- [64] C. Diuk, A. Schapiro, N. Córdova, J. Ribas-Fernandes, Y. Niv, M. Botvinick, Divide and Conquer: Hierarchical Reinforcement Learning and Task Decomposition in Humans, in: Computational and Robotic Models of the Hierarchical Organization of Behavior, Springer, Berlin, Heidelberg, 2013, pp. 271–291. doi:10.1007/978-3-642-39875-9₁2.
- [65] A. G. E. Collins, The Tortoise and the Hare: Interactions between Reinforcement Learning and Working Memory, Journal of Cognitive Neuroscience 30 (10) (2018) 1422–1432. doi:10.1162/jocn_{a0}1238.
- [66] L.-H. Tai, A. M. Lee, N. Benavidez, A. Bonci, L. Wilbrecht, Transient stimulation of distinct subpopulations of striatal neurons mimics changes in action value, Nature Neuroscience 15 (9) (2012) 1281–1289. doi:10.1038/nn.3188.
- [67] M. Donoso, A. G. E. Collins, E. Koechlin, Foundations of human reasoning in the prefrontal cortex, Science 344 (6191) (2014) 1481–1486. doi:10.1126/science.1252254.
- URL http://www.sciencemag.org/cgi/doi/10.1126/science.1252254
- [68] A. M. Bornstein, K. A. Norman, Reinstated episodic context guides
 sampling-based decisions for reward, Nature Neuroscience 20 (7) (2017) 997–
 1003. doi:10.1038/nn.4573.
- 1228 URL https://www.nature.com/articles/nn.4573
- [69] O. M. Vikbladh, M. R. Meager, J. King, K. Blackmon, O. Devinsky,
 D. Shohamy, N. Burgess, N. D. Daw, Hippocampal Contributions to ModelBased Planning and Spatial Memory, Neuron 102 (3) (2019) 683–693.e4.
 doi:10.1016/j.neuron.2019.02.014.

¹²³³ URL https://www.sciencedirect.com/science/article/pii/S0896627319301230

M. E. van der Schaaf, E. Warmerdam, E. A. Crone, R. Cools, Distinct
linear and non-linear trajectories of reward and punishment reversal learning during development: relevance for dopamine's role in adolescent decision making, Developmental Cognitive Neuroscience 1 (4) (2011) 578–590.
doi:10.1016/j.dcn.2011.06.007.

[73] N. Sendhilnathan, M. Semework, M. E. Goldberg, A. E. Ipata, Neural Correlates of Reinforcement Learning in Mid-lateral Cerebellum, Neuron 106 (1)
(2020) 188–198.e5. doi:10.1016/j.neuron.2019.12.032.

- [72] S. D. McDougle, A. G. E. Collins, Modeling the influence of working memory,
 reinforcement, and action uncertainty on reaction time and choice during
 instrumental learning, Psychonomic Bulletin & Review 28 (1) (2021) 20–39.
 doi:10.3758/s13423-020-01774-z.
- ¹²⁴⁶ URL https://doi.org/10.3758/s13423-020-01774-z

[73] A. Radulescu, Y. Niv, I. Ballard, Holistic Reinforcement Learning: The
Role of Structure and Attention, Trends in Cognitive Sciences 23 (4) (2019)
278–292. doi:10.1016/j.tics.2019.01.010.

¹²⁵⁰ URL https://www.sciencedirect.com/science/article/pii/S1364661319300361

Konovalov, I. Krajbich, 1754 А. Neurocomputational **Dynamics** of Sequence Learning, Neuron 98(6)(2018)1282-1293.e4. 1252 doi:10.1016/j.neuron.2018.05.013. 1253

- ¹²⁵⁴ URL http://www.sciencedirect.com/science/article/pii/S0896627318303854
- [75] W. Kool, F. A. Cushman, S. J. Gershman, When Does Model-Based Control
- ¹²⁵⁶ Pay Off?, PLOS Computational Biology 12 (8) (2016) e1005090, publisher:
- ¹²⁵⁷ Public Library of Science. doi:10.1371/journal.pcbi.1005090.
- ¹²⁵⁸ URL https://journals.plos.org/ploscompbiol/article?id=10.1371/journal.pcbi.100509
- [746] C. F. d. Silva, T. A. Hare, Humans are primarily model-based learners in
 the two-stage task, bioRxiv (2020) 682922Publisher: Cold Spring Harbor
 Laboratory Section: New Results. doi:10.1101/682922.
- 1262 URL https://www.biorxiv.org/content/10.1101/682922v4

[767] I. W. Eisenberg, P. G. Bissett, A. Zeynep Enkavi, J. Li, D. P. MacKinnon, L. A. Marsch, R. A. Poldrack, Uncovering the structure of selfregulation through data-driven ontology discovery, Nature Communications 106 (1) (2019) 1–13. doi:10.1038/s41467-019-10301-1.

- ¹²⁶⁷ URL https://www.nature.com/articles/s41467-019-10301-1
- W. C. Lin, K. Delevich, L. Wilbrecht, A role for adaptive developmental
 plasticity in learning and decision making, Current Opinion in Behavioral
 Sciences 36 (2020) 48–54. doi:10.1016/j.cobeha.2020.07.010.
- 1271 URL http://www.sciencedirect.com/science/article/pii/S2352154620301121
- [79] J. Davidow, K. Foerde, A. Galvan, D. Shohamy, An Upside to Reward Sen-¹²⁷³ sitivity: The Hippocampus Supports Enhanced Reinforcement Learning in

- Adolescence, Neuron 92 (1) (2016) 93-99. doi:10.1016/j.neuron.2016.08.031.
 URL http://linkinghub.elsevier.com/retrieve/pii/S0896627316305244
- [80] C. Johnson, L. Wilbrecht, Juvenile mice show greater flexibility in multiple
- 1277 choice reversal learning than adults, Developmental Cognitive Neuroscience
- 1278 1 (4) (2011) 540–551. doi:10.1016/j.dcn.2011.05.008.
- URL http://linkinghub.elsevier.com/retrieve/pii/S1878929311000533
- [234] A. C. Petersen, L. Crockett, M. Richards, A. Boxer, A self-report measure 1281 of pubertal status: Reliability, validity, and initial norms, Journal of Youth
- and Adolescence 17 (2) (1988) 117–133. doi:10.1007/BF01537962.
- ¹²⁸³ URL https://doi.org/10.1007/BF01537962
- [82] A. Christakou, S. J. Gershman, Y. Niv, A. Simmons, M. Brammer, K. Rubia, Neural and psychological maturation of decision-making in adolescence and young adulthood, Journal of Cognitive Neuroscience 25 (11) (2013) 1807– 1823. doi:10.1162/jocn_{a0}0447.
- [263] W. van den Bos, M. X. Cohen, T. Kahnt, E. A. Crone, Striatum-Medial
 Prefrontal Cortex Connectivity Predicts Developmental Changes in
 Reinforcement Learning, Cerebral Cortex 22 (6) (2012) 1247–1255.
 doi:10.1093/cercor/bhr198.
- URL https://academic.oup.com/cercor/article/22/6/1247/307075
- [264] M. J. Frank, L. C. Seeberger, R. C. O'Reilly, By Carrot or by Stick: Cognitive
 Reinforcement Learning in Parkinsonism, Science 306 (5703) (2004) 1940–
 1943. doi:10.1126/science.1102941.
- [35] R. D. Cazé, M. A. A. van der Meer, Adaptive properties of differential learning rates for positive and negative outcomes, Biological Cybernetics 107 (6)
 (2013) 711-719. doi:10.1007/s00422-013-0571-5.
- S. Palminteri, E. J. Kilford, G. Coricelli, S.-J. Blakemore, The Computational Development of Reinforcement Learning during Adolescence, PLoS Computational Biology 12 (6) (Jun. 2016). doi:10.1371/journal.pcbi.1004953.
 URL https://www.ncbi.nlm.nih.gov/pmc/articles/PMC4920542/
- [87] G. Lefebvre, M. Lebreton, F. Meyniel, S. Bourgeois-Gironde, S. Palminteri,
- Behavioural and neural characterization of optimistic reinforcement learning,
 Nature Human Behaviour 1 (4) (2017) 0067. doi:10.1038/s41562-017-0067.
- ¹³⁰² URL http://www.nature.com/articles/s41562-017-0067

- [868] W. Dabney, Z. Kurth-Nelson, N. Uchida, C. K. Starkweather, D. Has¹³⁰⁴ sabis, R. Munos, M. Botvinick, A distributional code for value in
 ¹³⁰⁵ dopamine-based reinforcement learning, Nature 577 (7792) (2020) 671–675.
 ¹³⁰⁶ doi:10.1038/s41586-019-1924-6.
- 1307 URL http://www.nature.com/articles/s41586-019-1924-6
- [&9] A. Gelman, J. B. Carlin, H. S. Stern, D. B. Dunson, A. Vehtari, D. B. Rubin,
 Bayesian Data Analysis, 3rd Edition, Chapman and Hall/CRC, Boca Raton,
 2013.
- [90] K. Katahira, How hierarchical models improve point estimates of model parameters at the individual level, Journal of Mathematical Psychology 73
 (2016) 37–58. doi:10.1016/j.jmp.2016.03.007.
- [91] S. Watanabe, A Widely Applicable Bayesian Information Criterion, Journal
- ¹³¹⁵ of Machine Learning Research 14 (Mar) (2013) 867–897.
- 1316 URL http://www.jmlr.org/papers/v14/watanabe13a.html
- [92] F. Pedregosa, G. Varoquaux, A. Gramfort, V. Michel, B. Thirion, O. Grisel,
 M. Blondel, P. Prettenhofer, R. Weiss, V. Dubourg, J. Vanderplas, A. Passos,
 D. Cournapeau, M. Brucher, M. Perrot, Duchesnay, Scikit-learn: Machine
 Learning in Python, Journal of Machine Learning Research 12 (85) (2011)
 2825–2830.
- 1322 URL http://jmlr.org/papers/v12/pedregosa11a.html
- D. A. Peterson, C. Elliott, D. D. Song, S. Makeig, T. J. Sejnowski, H. Poizner,
 Probabilistic reversal learning is impaired in Parkinson's disease, Neuro science 163 (4) (2009) 1092-1101. doi:10.1016/j.neuroscience.2009.07.033.
 URL http://www.sciencedirect.com/science/article/pii/S0306452209012068
- R. Swainson, R. D. Rogers, B. J. Sahakian, B. A. Summers, C. E. Polkey,
 T. W. Robbins, Probabilistic learning and reversal deficits in patients with
 Parkinson's disease or frontal or temporal lobe lesions: possible adverse
 effects of dopaminergic medication, Neuropsychologia 38 (5) (2000) 596–612.
 doi:10.1016/S0028-3932(99)00103-7.
- 1332 URL http://www.sciencedirect.com/science/article/pii/S0028393299001037
- [95] J. A. Waltz, J. M. Gold, Probabilistic reversal learning impairments in
 ¹³³⁴ schizophrenia: Further evidence of orbitofrontal dysfunction, Schizophrenia

- 1335 Research 93 (1) (2007) 296–303. doi:10.1016/j.schres.2007.03.010.
- 1336 URL http://www.sciencedirect.com/science/article/pii/S092099640700120X
- [96] D. P. Dickstein, E. C. Finger, M. A. Brotman, B. A. Rich, D. S. Pine,
- ¹³³⁸ J. R. Blair, E. Leibenluft, Impaired probabilistic reversal learning in youths
- ¹³³⁹ with mood and anxiety disorders, Psychological Medicine 40 (7) (2010)
- 1340 1089–1100. doi:10.1017/S0033291709991462.
- ¹³⁴¹ URL http://www.cambridge.org/core/journals/psychological-medicine/article/impaire
- [97] R. Cools, L. Clark, A. M. Owen, T. W. Robbins, Defining the Neural Mech-
- ¹³⁴³ anisms of Probabilistic Reversal Learning Using Event-Related Functional
- ¹³⁴⁴ Magnetic Resonance Imaging, Journal of Neuroscience 22 (11) (2002) 4563–
- ¹³⁴⁵ 4567. doi:10.1523/JNEUROSCI.22-11-04563.2002.
- 1346 URL https://www.jneurosci.org/content/22/11/4563
- [98] R. Cools, M. J. Frank, S. E. Gibbs, A. Miyakawa, W. Jagust, M. D'Esposito,
- ¹³⁴⁸ Striatal Dopamine Predicts Outcome-Specific Reversal Learning and Its Sen-
- ¹³⁴⁹ sitivity to Dopaminergic Drug Administration, Journal of Neuroscience 29 (5)
- 1350 (2009) 1538–1543. doi:10.1523/JNEUROSCI.4467-08.2009.
- ¹³⁵¹ URL https://www.jneurosci.org/content/29/5/1538
- [99] F. Lourenco, B. Casey, Adjusting behavior to changing environmental
- ¹³⁵³ demands with development, Neuroscience & Biobehavioral Reviews 37 (9)
- 1354 (2013) 2233–2242. doi:10.1016/j.neubiorev.2013.03.003.
- $_{\tt 1355}$ URL https://linkinghub.elsevier.com/retrieve/pii/S0149763413000638
- [100] A. Izquierdo, J. L. Brigman, A. K. Radke, P. H. Rudebeck, A. Holmes, The ¹³⁵⁷ neural basis of reversal learning: An updated perspective, Neuroscience 345
- 1358 (2017) 12–26. doi:10.1016/j.neuroscience.2016.03.021.
- ¹³⁵⁹ URL http://www.sciencedirect.com/science/article/pii/S030645221600244X
- [104] A. G. E. Collins, B. Ciullo, M. J. Frank, D. Badre, Working Memory Load ¹³⁶¹ Strengthens Reward Prediction Errors, The Journal of Neuroscience 37 (16)
- 1362 (2017) 4332–4342. doi:10.1523/JNEUROSCI.2700-16.2017.
- ¹³⁶³ URL http://www.jneurosci.org/lookup/doi/10.1523/JNEUROSCI.2700-16.2017
- [102] A. G. E. Collins, M. A. Albrecht, J. A. Waltz, J. M. Gold, M. J.
 ¹³⁶⁵ Frank, Interactions Among Working Memory, Reinforcement Learning,
 ¹³⁶⁶ and Effort in Value-Based Choice: A New Paradigm and Selective
 ¹³⁶⁷ Deficits in Schizophrenia, Biological Psychiatry 82 (6) (2017) 431–439.

¹³⁶⁸ doi:10.1016/j.biopsych.2017.05.017.

¹³⁶⁹ URL http://www.sciencedirect.com/science/article/pii/S0006322317316190