Identifying synergies and antagonisms between plant defense traits in *Asclepias syriaca*

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Abstract

As a general rule, plants defend against herbivores with multiple traits. The defense synergy

hypothesis posits that some traits are more potent when co-expressed with others, compared to

their independent potency. However, this hypothesis has rarely been tested outside of synergies

within a class of particular phytochemicals, and seldom under field conditions. We tested for

synergies between multiple defense traits of common milkweed (*Asclepias syriaca*) by assaying

the performance of two specialist herbivores on plants in natural populations. We employed both

standard regression and exploratory analysis using a novel application of Random Forests that

allowed us to detect synergies between defense traits. In hypothesis testing, we found the first

empirical evidence for a previously hypothesized synergy between one pair of co-expressed

defense traits (latex and cardenolides), but not another (latex and trichomes). When exploring all

potential interactions between pairs of traits we found eight synergies and five antagonisms in

predicting herbivore performance. Half of the identified synergies involved carbon, which is the

basis of several defenses including chemical and physical barriers to feeding, and also essential

nutrients. Our findings suggest that defense synergies could explain co-expression of latex and

cardenolides in milkweeds. This synergy may be common among the diverse plant species that

employ latex as a defense. Future studies should test carbon-based synergies, which our work

suggests are prevalent, as well as the other synergies identified in our exploratory analysis. Our

analytic approach provides a general, flexible framework for more broadly discovering and

predicting the coexpression of traits through their synergistic function.

Keywords: Defense synergy hypothesis; Random Forest; common milkweed (*Asclepias*

syriaca); swamp milkweed beetle (*Labidomera clivicollis*); monarch butterfly (*Danaus*

plexippus); plant-insect interactions; chemical ecology; synergy

Introduction

Plant defenses structure plant and insect communities, shape the way that primary

productivity flows through to the rest of the food web, and can strongly impact the success of

individual plants in the face of herbivory (Kursar et al. 2009, Kant et al. 2015, Coley et al. 2018).

Consequently, ecologists have spent decades developing and evaluating theories to explain when and

how plants should defend themselves (Agrawal 2011). One consistent finding is that individual

plants generally have multiple, diverse defense traits (Duffey and Stout 1996, Agrawal and Fishbein

2006, Salazar et al 2018). Explaining why plants invest in multiple defenses is an open question, and

is particularly important since there is some evidence that selection has favored specific

combinations of defense traits ("defense syndromes" – see Agrawal and Fishbein 2006, da Silva and

Batalha 2011, Sanczuk et al. 2020, cf. Moles et al 2013). Several non-exclusive hypotheses have

been proposed to explain both the prevalence of multi-trait defense strategies and the specific

combinations of defenses that are observed (Agrawal 2011).

Jones and Firn (1991) articulated the problem of multiple defenses clearly: individual

plants often express many different types of secondary metabolites, and it is difficult to explain

why. One of the leading explanations is synergistic interactions between secondary metabolites

(Duffey and Stout 1996, Richards et al. 2016, Wetzel and Whitehead 2020). For example, the

furanocoumarin Xanthotoxin is more effective when mixed with other furanocoumarins than

expected from its performance alone (Berenbaum et al. 1991); similarly, amides in *Piper* plants

are more effective in combination than their performance separately suggests (Scott et al. 2002,

Dyer et al. 2003, 2010, Whitehead and Bower 2014). Synergies have also been detected in

phytochemical defenses against microbes (e.g., Stermitz et al. 2000, Muroi and Kubo 1993) and

fungi (Cipollini and Stiles 1992). On the other hand, there have also been observations of

antagonistic relationships, where compounds reduce the efficacy of each other, including among

furanocoumarins (Diawara et al. 1993), amides (Whitehead and Bowers 2014), and between

pyrrolizidine alkaloids and chlorogenic acid (Liu et al. 2017). Tests for synergies between

secondary metabolites are still relatively rare (Richards et al. 2016), and there have been almost

no tests for synergies between classes of defenses (e.g., physical and chemical defenses), even

though there is every reason to expect them to exist (Rasmann and Agrawal 2009, Agrawal 2011,

Richards et al. 2016, but see Green et al. 2001, Steppuhn and Baldwin 2007).

Studies of synergies between phytochemicals generally use artificial diets with manipulated

concentrations of phytochemicals in an ANOVA design (e.g. Dyer et al. 2003, Berenbaum et al.

1991, Stamp and Yang 1996, Scott et al. 2002). Accordingly, synergy or antagonism is detected

based on the interaction term of the two (or more) treatments (for a discussion on the appropriate null

model for this experimental design, see Pennings 1996, Hay 1996, Nelson and Kursar 1999). For

experimental designs with continuously varying defense traits, the logical extension of this approach

is regression, with the interaction term between two traits capturing potential synergies or

antagonisms (Richards et al. 2016). Nonetheless, implicit in a regression model approach is the

assumption of a *bilinear* interaction: that the synergistic benefit of one additional unit of defense 1

for each unit of defense 2 is the same regardless of the original levels. However, other relationships

may be equally likely. For example, the synergistic benefit of one trait might be a threshold effect;

this could occur if one defense causes a switch in herbivore behavior above a certain level, and the

effect of the second trait is impacted by herbivore behavior. Similarly, we might expect some

synergies with diminishing marginal returns. To capture complex relationships like these, we

introduce tests for synergies and antagonisms using Random Forest regression. Random Forests are a

popular form of machine learning that fits data using an ensemble (effectively, model averaging) of

many decision trees. Because of this structure, Random Forest models can easily capture non-linear

trends in data. Recent innovations have provided a framework for harnessing the flexible nature of

Random Forests for inference about interactions, as we do here (Hooker 2007, Hooker and Mentch

2019). To our knowledge, this approach has not been previously used in ecology, so we include

thoroughly commented code in the supplements.

The common milkweed, *Asclepias syriaca*, is a well-characterized system for studying plant

defenses. The species produces toxic cardenolides that are potent defenses and also severely limit the

species of herbivores that can feed on *A. syriaca*, leading to a simplified herbivore community

(Agrawal 2017). In addition to cardenolides, *A. syriaca* also employ other defenses: trichomes on the

leaf surfaces may reduce traction and insect access to feeding (Agrawal et al. 2009), latex exuded

from channels (laticifers) in damaged tissues can cement insect mouth parts (Agrawal and Konno

2009), and leaf toughness and low nutrient content may reduce insect preference or performance

(Agrawal and Fishbein 2006). Furthermore, expression of these defensive traits varies substantially

in natural conditions within and between species, allowing experiments to capture broad ranges of

defense traits in a field setting.

Two putative defense synergies have been hypothesized for milkweed plants (Agrawal

2011). First, freshly hatched monarch butterfly caterpillars were observed to first carefully cut

trichomes in a small feeding region, then bite into the leaf tissue (Zalucki et al. 2002). At this point

latex exudes, and the caterpillar must struggle free before it can consume leaf tissue. Many

caterpillars fail to survive this first exposure to latex (Zalucki et al. 2001). Thus, trichomes and latex

may act synergistically as one defense follows the other before feeding can commence; further, these

two traits also show correlated evolution, suggesting that selection has favored their joint expression

(Agrawal and Fishbein 2006). Second, studies of *A. syriaca* latex have found that it contains

cardenolides that mirror the concentrations in leaves (Agrawal et al. 2014), so it may be a vehicle to

deliver toxins directly into the mouth of hungry caterpillars, while also physically gumming up

mouthparts. This could be another defensive benefit of co-expression, in which the chemical and

physical barriers associated with cardenolides and latex synergize. There are indications that this

mechanism may be present in other plants with latex and secondary metabolites, as reduced levels of

secondary metabolites in latex led to increased herbivory in the common dandelion (*Taraxacum*

officinale) (Huber et al. 2016). While latex-by-trichome and latex-by-cardenolide have been

proposed in papers across the last two decades, to our knowledge there has been no direct empirical

evidence for either.

In the current study we evaluate plant defense synergies using data from herbivore

performance on common milkweed *A. syriaca* under field conditions, using both well-

established regression approaches and a novel application of Random Forests. Unlike the vast

majority of existing studies of synergies (which use artificial diets), we measure herbivore

performance under field conditions, relying on natural variation in defense traits. Using

observational field studies sacrifices controlled conditions, but provides a realistic context for

plant defenses, and allows us to study defenses that are not easily manipulated in the lab. Using

three data sets for two specialist chewing herbivores (one of which was previously published),

we test the defense synergy hypothesis for two predicted synergies: latex-by-cardenolide and

latex-by-trichome. We also explore other potential synergies and antagonisms among >5 traits.

Our methods provide a road map for future tests of the synergy hypothesis between classes of

plant defenses, and for other evaluations of trait synergies in ecology.

Material and Methods

Study Species

A. syriaca is a long-lived herbaceous plant native throughout eastern North America,

often in recently disturbed habitats including roadsides and fields. It employs a number of

defenses against herbivores, including a sticky latex that exudes upon damage, an abundance of

trichomes, particularly on the underside of leaves, and toxic cardenolides throughout plant

tissues (Malcolm and Zalucki, 1996; Zalucki et al., 2001). The monarch butterfly, *Danaus*

plexippus, and the swamp milkweed beetle, *Labidomera clivicollis*, are both *Asclepias* specialists

that feed on the leaf tissues of *A. syriaca* in their larval stages. Both of these herbivores have

modifications to their sodium pumps that increase their resistance to cardenolides (Dobler et al.

2012), but do not provide complete immunity (e.g. Zalucki et al. 2001, Agrawal 2005, Tao et al.

2016, Jones et al. 2019).

Field experiments

In September 2015, we conducted an experiment to assay the growth and survival of

neonate monarch caterpillars, *Danaus plexippus,* on 117 *A. syriaca* plants near Ithaca, NY (42.39

N, 76.39 W). Caterpillars were from a laboratory colony (see Jones et al 2019). Plants in this

natural population were selected haphazardly, using distance and visual phenotype to avoid

selecting ramets from the same plant; we also avoided ramets that showed high levels of damage

or early senescence. We placed two freshly hatched monarch caterpillars on one of the top pair of

fully expanded leaves and enclosed them in a fine mesh bag. We used the other leaf in that pair to

measure plant traits, including taking latex and toughness measurements, collecting a leaf disk for

measuring trichome density and specific leaf area (SLA). We then removed that paired leaf to be

dried for measurement of cardenolides, carbon, and nitrogen.

Latex, leaf toughness, specific leaf area, and cardenolides were quantified following the

methods of Hahn et al. (2019). In brief, latex was measured by removing 2-3 mm of the leaf tip

and measuring the weight of the resulting latex exudate; leaf toughness was measured using a

penetrometer (Chatillon type 5/6) once on either side of the leaf and averaging the results; SLA

(area per dry mass) was measured by drying and weighing a 31.67 mm^2 leaf punch; cardenolides

were separated and measured from dried and ground leaf tissue using High Performance Liquid

Chromatography (HPLC) following the methods of Züst et al. (2019). Trichomes density was

measured by first imaging the underside of leaf discs with the Ziess SteREO Discovery.V20, then

using Fiji, a variant of imageJ, to draw vertical and horizontal transects through the middle of

the image and count all hairs that crossed the transects (Schindelin et al., 2012). These transect

counts were highly correlated with transect densities obtained by counting all hairs in a quadrant

of the leaf disc, but were faster and more robust to overlapping hairs (XXX, unpublished).

Ground and dried leaf samples were analyzed at the Cornell Isotope Laboratory (COIL) for

concentrations of carbon and nitrogen.

Monarch caterpillars generally remained on the plants for 6 days, and then were removed

and weighed. In a few cases, field conditions prevented collection on day six, and caterpillars

were instead collected on day 7; where appropriate, including day of collection as a covariate in

statistical models accounted for this discrepancy. Missing caterpillars were presumed dead. In a

several cases, plants died during the experiment, necessitating removing them from the study, or

technical issues prevented us from obtaining one or more measurements of plant traits. In total, we

obtained 165 weighed larvae and 173 measures of survivorship for 92 plants.

In September 2016, we conducted an experiment to assay growth and survival of neonate

swamp milkweed beetles, *Labidomera clivicollis,* on 129 *A. syriaca* plants with the same

methods as above, with following exceptions. First, we used a separate field site approximately

300 meters away, at Dunlop Meadow (a Cornell University Natural Area). We did not measure carbon and

nitrogen concentrations, and were unable to measure SLA. Due to unusual weather patterns, many of

the plants experienced accelerated senescence during the experiment; we did not include these

samples. Because *L. clivicollis* aggregate in their early life stages, we placed 5 larvae in each

enclosure. In a few cases, loss of larvae during placement left us with only 4 larvae in an

enclosure. In total, we obtained 134 weighed larvae and 464 records of survivorship for 95 plants.

For comparison, and to demonstrate the applicability of our approach with existing data,

we also analyzed a previously published data set of *D. plexippus* bioassays (Agrawal 2005).

These *A. syriaca* were planted outdoors in a common garden outside of Toronto in 2001, and an

assay of *D. plexipus* performance was carried out in 2002. The plant traits were measured in

largely the same way as the 2015 and 2016 experiments (see Agrawal 2005); notable differences

were that only one herbivore was placed on each plant, cardenolide concentrations are reported

as a single aggregated concentration, and trichomes were counted on quadrats of both sides of a

leaf disk rather than transects of the underside of the leaf.

Analysis

Following Tredennick et al. (2021), we separated our analysis into two sections: hypothesis

testing, and exploration. First, we carried out strict hypothesis testing on previously proposed

synergies between *A. syriaca* defenses (latex-by-trichome, latex-by-cardenolide) using regression

models, including main effects of the two traits of interest as well as their interaction. Then, we

carried out exploratory analysis of other trait interactions using regression models and Random

Forests. Our *a priori* hypothesis testing was conducted using conventional significance tests.

Exploratory analyses used regression models involving many tests for synergies (all pairwise

combinations of measured traits), and are thus susceptible to false positives. As such, we used P

values as a general benchmark for the relative support for different synergies; the P value for any

one test should be interpreted with caution. Using Random Forests did not provide P values, but

allowed us to identify more complicated (ie non-bilinear) synergistic or antagonistic interactions

between defense traits. In all cases, we used log-transformed larval mass divided by days (i.e.,

daily growth rate assuming exponential growth) as our metric of herbivore growth. For all

analyses, we rescaled plant traits so that their standard deviations were 1, but did not further

adjust the mean. All analyses and simulations were carried out in the programming language R

(R Core Team 2021).

In total we had three data sets, each with two measures of herbivore response (survival

and growth). For simplicity, in the rest of this paper we will refer to these data sets as "Monarchs

2005", "Monarchs 2015", and "Clivicollis 2016" (Monarchs 2005 are labeled by publication date

– the experiment was carried out in 2002). These three data sets have some overlap in the

defense traits measured, but often these traits are not directly comparable (e.g., trichomes

counting methods for Monarchs 2005 differed from the more recent experiments). Additionally,

the experiment in Agrawal (2005) was set up on a single day with a single bioassay herbivore,

rather than a blocked design with two (Monarch 2015) or five (Clivicollis 2016) larvae per plant,

implemented across multiple days. For this reason, only Monarch 2015 and Clivicollis 2016

analyses required mixed effects models to capture the hierarchical structure of the data. Three

data sets and two response variables result in six repetitions of each analysis. Although this leads

to unavoidable complications in the description of each analysis and the associated analyses

scripts, these are merely technical details of doing appropriate regression analyses for each

experiment. Fundamentally, for each experiment we (a) use regression modeling to test

previously hypothesized interactions between defense traits, (b) use regression modeling to

examine the data for evidence of other interactions between defense traits, and (c) use Random

Forests as an alternative approach to (a) and (b) that requires no assumptions about the linearity

of interactions.

Trait Correlations

We computed pairwise across-plant trait (phenotypic) correlations for each of our

experiments, using Pearson correlation coefficients. Where we identified substantial correlations

between particular cardenolide compounds (Monarchs 2015 and Clivicollis 2016) we used

principal component analysis (PCA) to identify a single measure (the first PCA) that represented

the majority of the variation in the covarying cardenolides. All further analysis used these first

PCAs in place of the individual cardenolides in question.

Main effects of defenses

We evaluated the efficacy of plant traits by fitting regression models for each

(experiment x response) combination, using an upward testing approach: fit the null model with

only intercept and blocking factors as appropriate, test whether model fit is improved by the

inclusion of individual plant traits. Growth was fit using linear regression (Monarchs 2005) or

linear mixed models (Monarchs 2015, Clivicollis 2016) including random effects for day and

plant identity. For survival, we carried out logistic regression using generalized linear models

(Monarch 2005) or generalized linear mixed models (Monarch 2015, Clivicollis 2016) with the

same model structure.

Hypothesis testing

For each hypothesized synergy (latex-by-trichomes, latex-by-cardenolide for each

cardenolide compound), we fit a regression model that included main effects for the two focal

traits, and an interaction between those two traits (and random effects of plant and date for the

Monarch 2015 and Clivicollis 2016 data). We then used Anova() in the **car** package to test the

significance of the interaction term (Fox and Weisberg 2019). A significant interaction term

represented a trait synergy or antagonism between the traits. Because the response was insect

performance, a negative coefficient for the interaction term represented a defense trait synergy,

while a positive term represented an antagonism. The magnitude of the slope coefficient can be

interpreted as follows. Because all plant traits were scaled for analysis to have standard deviation

of 1, an interaction coefficient of $\beta>0$ means that a one-standard-deviation increase in either trait

increases by β the effect of a one-standard-deviation increase in the other (similarly, a negative

coefficient implies a decrease in the effect of the second trait, if the first is increased). When

survival is the response, this interpretation applies to logit survival, rather than survival itself

(i.e., it applies to the linear predictor in a logistic regression model); when growth is the response

it applies to growth.

Because the interaction term coefficient is interpreted on the scale of the linear predictor

rather than the response, its magnitude cannot be used to compare the strength of synergies

between survival and growth responses, or to make comparisons with equation-free models like

Random Forests. To provide a more interpretable and robust measure of the strength of synergies

and antagonisms, for each trait pair and response (growth or survival) we calculated the

"predictable response range" and the percent of response variation explained by the interaction.

As these metrics are central to our Random Forest analyses, we fully describe their calculation in

Exploration: Random Forest, below. For a given pair of plant traits, the response range shows

how much changing the focal traits (within the range of values observed in the field) impacts

herbivores; the percent of variation explained by the interaction term identifies how important

the synergy or antagonism is relative to the direct effects of the traits. Notably, these measures

are based on the predictable variation in the response (i.e., the fitted response surface described

the regression equations) and are in intuitive units (growth rate or survival probability, percent

variation explained). Because Random Forests are more flexible models, we don't expect the

interaction term to explain as much of the variation, but this is unavoidable.

Exploration: regression

Using the same framework as in the *a priori* regression tests above, we looked for

synergies or antagonisms for all pairwise combinations of measured traits. We use the P value as

an indicator for potential synergies/antagonisms, but we do not interpret P-values as we would

for an individual hypothesis test, because this process involves multiple comparisons that would

produce multiple false positives even if there are no interactions. Instead, the P values provide a

benchmark of relative support for the interactions tested.

Exploration: Random Forest

Like most machine learning algorithms, Random Forests can be very effective at

prediction (e.g. "How fast do caterpillars grow on a plant with these defense traits?"), but do not

directly provide information (analogous to regression coefficients) about the basis for those

predictions (e.g. "Which plant traits most strongly affect caterpillar growth rate?"). Hooker

(2007) provides methodology for interrogating a fitted Random Forest model to determine the

presence or absence of interactions among variables, and to quantify the importance of

interactions if they are present. We have adopted this approach for both Random Forest models

and our regression analyses (above).

For a given experiment and response (growth or survival), we first fit a Random Forest

model that included all available plant traits (Random Forests capture interactions without

needing explicit terms to represent them). Following Hooker (2007), we then treated this fitted

model as if it were an experimental system, in which we could measure herbivore performance

on hypothetical plants having any specified combination of defense trait values. We conducted

simulated experiments, using the fitted Random Forest to generate the "response" (predicted

herbivore performance) for a set of hypothetical plants with differing defense traits. The

predicted herbivore performance was thus the result of any linear and non-linear patterns in the

real data that the Random Forest model captured. We carried out simulated experiments for each

possible defense trait pair, and we analyzed the resulting simulated data to determine if the

response was additive (no interactions) or non-additive (synergy or antagonism). In the case of

our regression analyses (above), we carried out the same procedure using our fitted regression

models. The procedure for simulated experiments and their analysis was as follows:

1.   For each real plant measured, we generated 100 (Random Forest) or 2500

(regression model) hypothetical plants, in which all trait values were the same as the real plant,

except that values of the two traits of interest were varied across a 10x10 (Random Forest) or

50x50 (regression model) grid ranging from the smallest to largest values observed in any real

plant. For an empirical data set of size N, this gave us 100N (Random Forest) or 2500N

(regression model) hypothetical plants. Because the outputs of Random Forests are piecewise

constant at fine scales, it is important to avoid partitioning the trait space too finely (Hooker,

personal communication). With regression models that don't have such limitations, we found

that using the predictions of a 50x50 grid of plant traits produced results similar to higher

resolution grids, and could be computed in a reasonable amount of time (CBE, unpublished

simulations).

2.   Using the fitted model (Random Forest or regression), we predicted average

herbivore performance on those hypothetical plants. The range of herbivore performance

(survival or growth) across these predictions gives us our "predictable response range" (PRR),

which characterizes the amount of variation in expected herbivore performance (i.e, average

growth or average survival of multiple herbivores exposed to that trait combination) that can be

attributed to varying the focal traits. Correlations among the traits we studied are sufficient weak

that the rectangle defined by the range of variation in each trait is a good approximation to the

possible trait combinations for real plants (see Fig. S1).

3.   We fit a two-way ANOVA model to the predicted response of the form

response ~ trait1 + trait2 + trait1:trait2

where the trait values of trait 1 and trait 2 were treated as factors. This meant that the

interaction term captured any non-additivity for a given combination of trait values. In the case

of Random Forests this captured non-additivity even if the independent response to each trait

was nonlinear, or if synergies or antagonisms were not consistent across different levels of the

traits. There were several complications in the actual implementation, particularly with Random

Forests, which we describe in the Supplements. One key feature for Random Forests was that we

used a weighted ANOVA, where each observation from the Random Forest trial was weighted

by how close it was (in terms of defense traits) to an actual empirical observation. This focused

our results on meaningful patterns driven by the data, rather than spurious patterns driven by

extrapolation of the Random Forest model outside the range of the data (Hooker and Mentch

2019).

4.   From the fitted ANOVA model, we determined what percent of the total Sum of

Squares explained by the ANOVA model we could attribute to the interaction term. This

provides a measure of the importance of the synergy or antagonism in intuitive units. In the case

of Random Forests, for which there is not a P value equivalent, we used this percent as our

measure for identifying important synergies, and chose 15% as a threshold for meaningful

synergies or antagonisms to report.

5.   Step 4 identified trait pairs with notable non-additivities, but did not characterize

the interaction as a synergy or antagonism. For regression models, the sign of the interaction

coefficient identified whether the interaction was a synergy or an antagonism, but Random

Forests did not provide such a metric. For that purpose, we generated diagnostic plots. These (a)

visualized the Random Forest predictions for a grid of points, and (b) compared how those

predictions differed from the expectations of an additive ANOVA model, which was like the

model in step 3 but contained only main effects. We then assessed whether the resulting pattern

was a synergy, antagonism, or unclear, by comparing those diagnostic plots to diagnostic plots

from simulated data with known trait interactions (see Random Forest simulations below, and

Fig. 2).

Because Random Forest regression struggles with missing data, we removed observations

with one or more missing plant trait measurement. Random Forests also do not have appropriate

ways to account for random effects, so we do not include the date or plant identity terms that

were present in the regression analyses (Monarch 2015, Clivicollis 2016).

Diagnostic plots with known interaction patterns

To determine whether patterns in the Random Forests diagnostic plots corresponded to

synergies or antagonisms, we simulated data with known synergies and antagonisms, and

generated the resulting diagnostic plots. In brief, we defined defense efficacy as a specified

function of two hypothetical traits X and Y, where larger values indicate better defense (e.g. the

first column in Fig 2a-f). We then used that function to simulate herbivore performance at a 10 X

10 grid of trait values, analogous to our fitted Random Forest predictions above (e.g. the second

column in Fig 2a-f). For this set of simulated herbivore performance data, we fit an additive

regression model of the form (response ~ trait X + trait Y), with traits X and Y treated as factors.

This is exactly the same as the model fitted in step 5 above. We then looked at how the simulated

herbivore performance compared to the predictions of this additive model (e.g. the third column

in Fig 2a-f). This process produced plots that were comparable to those produced in step 5 above,

but with the difference that the underlying synergistic or antagonistic relationship between the

two defense traits was known.

We assumed that traits X and Y were scaled to have values between 0 and 1, and defined

herbivore performance as 10 minus the defense efficacy; as we were looking for qualitative

patterns, these specific values do not qualitatively change our results (CBE, unpublished

simulations). We considered many potential trait synergies and antagonisms, but here present

results for three synergies (linear, threshold, diminishing marginal returns) and three

antagonisms (linear, threshold, decaying). For more details, and code to simulate other synergies

and antagonisms, see the Supplements.

Results

Trait Correlations

We identified three cardenolide compounds that were highly correlated in the Monarch

2015 experiment (peaks that eluted at 17.4, 17.7, and 18.4 minutes; pairwise Pearson correlation

coefficients ranged from r = 0.77 to 0.86) and a pair of correlated cardenolide compounds in the

Clivicollis 2016 experiment (peaks 10.7 and 18.1; pairwise correlation coefficient of r = 0.73).

To avoid misleading results from highly correlated predictors, we used Principal Component

Analyses (PCA) on these two sets of cardenolides, and used the first component as a single plant

trait to represent the three peaks in 2015 (labeled "card15.pca") and the two peaks in 2016

(labeled "card16.pca"). We also found numerous plant traits that were mildly-to-moderately

correlated (Figs. S1) which may reflect underlying genetic structure, or correlated expression in

response to biotic or environmental drivers. We did not further collapse our data with PCA, but

where appropriate we note the consequences of correlated traits on our results.

While it was not the focus of our study, we highlight several significant patterns of trait

co-expression in our field studies (Fig S1). In 2015, several of the physical traits showed positive

phenotypic correlations – notably, leaf toughness was correlated with both latex (r = 0.17) and

trichomes (r = 0.49). Percent leaf carbon was positively correlated with the two traits that clearly

incorporate carbon, trichomes (r = 0.49) and leaf toughness (r = 0.21). In 2016, physical defense

traits (trichomes, toughness, and latex) all had significant positive correlation with each other

(latex and toughness, r = 0.38; latex and trichomes, r = 0.13; toughness and trichomes, r = 0.30),

and were generally negatively correlated or essentially uncorrelated with cardenolides.

Main effects of defenses

Trichomes had a negative main effect on Monarch 2005 growth (slope = -0.024, p =

0.025), and both percent carbon (p = 0.006) and carbon:nitrogen ratio (p = 0.015) similarly had

negative main effects Monarch 2005 survival. These predictors were only somewhat correlated

(Pearson correlation coefficient = 0.26), and a model fit using carbon and carbon:nitrogen was

much better than fitting a model with carbon and nitrogen alone (likelihood ratio test, p < 0.001).

The model with both carbon (slope = -0.393) and carbon:nitrogen ratio (slope = -0.151) had

negative coefficients for both, suggesting that even when plants had high carbon (bad for

monarch survival), having a higher carbon:nitrogen ratio provided further defense against

monarchs.

In contrast, percent carbon (slope = 0.028, p = 0.028) and trichomes (slope = 0.013, p =

0.086) had positive main effects on monarch growth in 2015, and no other plant traits had

significant effects. Using AIC model selection, a model with only carbon was better than either a

model with only trichomes (ΔAIC = -2.60) or a model with trichomes and carbon (ΔAIC = -0.37,

in this case the models are statistically indistinguishable using AIC). There were no significant

predictors of Monarch 2015 survival. Leaf toughness had a marginally significant positive main

effect on *L. clivicollis* growth (slope = 0.017, p=0.074), and cardenolide peaks 10.2 and 8.4 had

negative main effects on *L. clivicollis* survival (slope = -0.969, p = 0.041; slope = -0.8116, p =

0.059 respectively).

Hypothesis testing

We found no evidence for the hypothesized synergies between defense traits for growth

or survival in Monarch 2005. We did find a marginally significant synergy between latex and

cardenolides (peak 18.6) on Monarch 2015 survival (slope = -0.724, p = 0.077; Predictable

Response Range (PRR) = [0.0114, 0.9999], 69% of PRR explained by synergy), such that plants

were more defended when latex and cardenolides levels were both high than either defense trait

would suggest in isolation (Fig. 1a; Table 1). We also found a significant synergy between latex

and cardenolides (peak 10.2) on *L. clivicollis* growth (slope = -0.038, p = 0.033; PRR = [-1.2719,

-0.7359], 30% of PRR explained by synergy) (Marginal for survival, slope = -2.085, p = 0.100;

PRR = [0, 0.9824], 19% of PRR explained by synergy) (Fig 1b-c, Table 1).

Exploration: regression

Due to the many models fit in our exploratory analyses, p values in exploration should

not be interpreted as the exact probability of the data under the null hypothesis. However, p

values do represent relative statistical support for the different synergies, and we use p<0.1 as a

threshold for reporting here; for the complete results of the analyses, see the supplementary code.

We found a notable nitrogen-by-trichome antagonism (slope = 0.0151, p =0.0323; PRR = [-

0.9313, -0.6643], 62% of PRR explained by antagonism), cardenolide-by-toughness synergy

(slope = -0.012, p = 0.0497; PRR = [-0.8795, -0.7016], 92% of PRR explained by synergy), and

carbon-by-trichome synergy (slope = -0.0124, p = 0.0412; PRR = [-0.9715, -0.6036], 60% of

PRR explained by synergy) for Monarch 20005 growth (Table S1). We also found a carbon-by-

nitrogen synergy for monarch survival (slope = -0.515, p = 0.009; PRR = [0.0035, 0.9997], 70%

of PRR explained by synergy). In 2015, we found toughness-by-carbon (slope = -0.0143, p

=0.326; PRR = [-0.9416, -0.6251], 39% of PRR explained by synergy) and nitrogen-by-

cardenolide 10.5 (slope = -0.140, p = 0.0579; PRR = [-0.8371, -0.6368], 68% of PRR explained

by synergy) synergies, and a toughness-by-cardenolide 10.5 antagonism (slope = 0.014, p

=0.0611; PRR = [-0.8297, -0.6475], 70% of PRR explained by antagonism) on Monarch growth.

We did not find any new notable interactions on Monarch 2015 survival (latex-by-cardenolide

18.6 synergy was present in the *a priori* testing). In 2016, we found a latex-by-toughness synergy

(slope = -0.0238, p = 0.0375; PRR = [-0.9955, -0.7689], 99% of PRR explained by synergy) and

a toughness-by-cardenolide PCA antagonism (slope = 0.0181, p = 0.060; PRR = [-1.0627, -

0.6207], 31% of PRR explained by antagonism) on *L. clivicollis* growth (in addition to the

previously identified latex-by-cardenolide 10.2 synergy). We did not identify any new notable

interactions affecting *L. clivicollis* survival.

Diagnostic plots with known interaction patterns

When applied to simulated data with known underlying interactions between defense

traits, synergies always produced diagnostic plots in which the bottom left and top right regions

showed herbivores underperforming predictions from an additive model (an ANOVA model

with no interaction terms fit to the simulated data), and the top left and bottom right showed

herbivores over-performing predictions from an additive model (e.g. Fig. 2a-c). The reverse was

true for antagonisms: herbivores always over-performed predictions from an additive model in

the bottom left and top right, and underperformed predictions from an additive the top left and

bottom right regions (e.g. Fig. 2d-f). These patterns held true for all of the synergies and

antagonisms we evaluated (XXX, unpublished simulations).

Exploration: Random Forests

We found strong evidence of non-additivity for nine trait pairs, for which at least 15% of

the sum of squares of the linear model was explained by the interaction term. Of these, three

showed strong patterns of synergies: carbon-by-nitrogen for Monarchs 2005 survival (also

identified with regression models above) (PRR = [0.36123, 0.8194], 37% explained by

interaction), latex-by-carbon for Monarchs 2015 growth (PRR = [-0.8484, -0.7229], 19%

explained by interaction), and toughness-by-carbon in Monarch 2015 (also identified with

regression models above) (PRR = [-0.8683, -0.7259], 15% explained by interaction), (Fig. 3a-c).

Three interactions also appeared to be antagonisms, all from Monarchs 2005: toughness-by-

trichome on survival (PRR = [-0.8766, -0.7679], 23% explained by interaction), and toughness-

by-carbon (PRR = [0.394, 0.7830], 20% explained by interaction) and latex-by-cardenolide on

survival (PRR = [0.4877, 0.7191], 21% explained by interaction) (Fig. 3d-f). For the three

synergies we identified in the *a priori* hypothesis testing above, the Random Forest plots show

the same diagnostic patterns of synergy, even though the Sum of Squares explained in these

cases did not exceed our 15% threshold for the Random Forest approach (Fig S2). Of the six

strong synergies and antagonisms identified using Random Forests, only two of them had been

detected in the regression analyses, suggesting the remaining four have non-linear functional

forms.

Discussion

Using surveys of herbivore performance in the field, and both well-established and novel

statistical tools, we found evidence for the plant defense synergy hypothesis, including some

potential new synergies and antagonisms. We found the first direct empirical support for the

previously hypothesized synergy between latex and cardenolides, suggesting that secondary

metabolites may serve to amplify the physical dangers of latex, or non-exclusively, that latex

serves as a focused delivery system for chemical defense. In contrast, we found no support for

the hypothesized synergy between latex and trichomes. Here, however, we caution that the

hypothesized mechanism for latex-by-trichome synergies focuses on mortality, and our Monarch

2015 experiment had such low mortality that there was very little statistical power to detect

associations between caterpillar deaths and specific plant traits.

In our exploration of all possible interactions, we identified eight potential synergies

between classes of defenses, and five potential antagonisms. Because these potential interactions

were discovered by searching for all possible interactions, they should not be accorded the same

level of confidence as those supported through strict hypothesis testing; nonetheless, they do

represent potentially important trait interactions that can be tested in the future. Half of the

identified potential synergies were between carbon and other traits (carbon-by-nitrogen, carbon-

by-trichome, carbon-by-latex, carbon-by-toughness). This is especially noteworthy given that we

were unable to measure carbon in the 2016 Clivicollis data, so those synergies came from only

two of the three field surveys. Percent carbon in leaves is a crude bulk measure, as much of the

leaf is itself carbon, and many defense traits have high carbon content. Thus, carbon may be

capturing investment in carbon-rich defenses, but can also measure overall plant allocation; as

such, carbon has played an important role in defense theory. Two long-standing plant defense

hypotheses address the allocation of carbon to plant defenses: the carbon/nutrient balance

hypothesis (Bryant et al. 1983), and the growth-differentiation balance hypothesis (GDBH)

(Hermes and Mattson 1992). While these hypotheses have mixed empirical support and have lost

favor in the literature, they provide a key insight: plants should invest in carbon-rich defenses

when they are nitrogen-limited and have excess carbon (Massad et al. 2012). In this context,

defense synergies involving carbon are particularly important, as nitrogen-limited plants may

have excess carbon that could increase the efficacy of the other synergistic defenses. If there are

consistent synergies between carbon and other classes of defense across plant species, we expect

to see interspecific patterns of defense co-expression driven by carbon availability, as a natural

extension of existing plant defense hypotheses. We also note that two of those carbon synergies

(carbon-by-trichome, carbon-by-leaf-toughness) are synergies between leaf carbon content, and

carbon-rich defenses. This might reflect a general benefit for investing in suites of carbon-rich

defenses, but could also represent uncaptured aspects of these traits (e.g., variation in trichome

length in addition to density, variation in leaf toughness to shearing in addition to punctures).

Three of the identified potential synergies (latex-by-carbon, latex-by-cardenolide, latex-

by-toughness) involved latex. Latex is present in nearly 10% of angiosperm species, has no

known function in plants other than defense and is itself a cornucopia of multiple physical and

chemical traits (Agrawal and Konno, 2009). Latex can be quite harmful to herbivores directly

(e.g. Zalucki & Brower 1992), but its commonness across plant families is all the more

understandable if it synergizes with other common defenses. Here, our measure of latex was the

wet mass exuded upon leaf damage, which typically correlates negatively with insect

performance, but also represents multiple finer-grained traits. Particularly notable were the latex-

by-cardenolide synergies identified in this study; these synergies had been previously

hypothesized (Agrawal 2011) and we found support for them in the natural populations tested.

Note that typical measures of leaf cardenolides in *A. syriaca* are independent of latex; in other

words, residual latex in the leaves, once cut, contributes relatively little to the cardenolides

measured (Agrawal et al. 2014). The hypothesized mechanisms for latex-by-cardenolide

synergies – that latex could serve as a delivery system for toxic phytochemicals, and toxic

phytochemicals make latex even more dangerous to herbivores – is not a special feature of *L.*

clivolis and *D. plexippus* feeding on *A. syriaca.* The latex of many plants contains secondary

metabolites, often at elevated levels compared to other plant tissues (Agrawal and Konno 2009).

A recent study on the common dandelion found that herbivore damage increased when

sesquiterpene levels in latex were reduced; while this study was not explicitly addressing

synergies, their findings imply latex-by-sesquiterpene synergy (Huber et al. 2016). To our

knowledge, our study is the first to find explicit evidence for latex-by-defense-metabolite

synergies, but we speculate that such synergies may be common across plant taxa.

Several pairs of defense traits were identified as both synergies and antagonisms in

different contexts (e.g. cardenolide-by-toughness decreased herbivore growth synergistically in

Monarchs 2005, but interacted antagonistically to impact growth in Monarchs 2015 and

Clivicollis 2016). It would perhaps be more satisfying to observe consistent effects of defense

trait pairings, as this would imply a universal mechanism underlying the synergy or antagonism;

in this context, trait pairs showing both synergies and antagonisms could be concerning. While

contradictions in our findings might reflect spurious relationships identified with our exploratory

approaches, there are several reasons not to discount these results. First, it is not unusual for field

experiments of plant defenses to find that defenses are inconsistent in their efficacy between

years or other contexts (e.g. inconsistent efficacy of latex and cardenolides as defense of

milkweeds, Agrawal and Fishbein 2006, and inconsistent effects of protease inhibitors and

nicotine as a defense of wild tobacco, Mitra et al. 2008); indeed, this has been proposed as a

mechanism that favors multi-trait defense strategies (Rasmann and Agrawal 2009). Studies of

synergies between phytochemical have also found that some pairs of traits can be synergistic in

some contexts and antagonistic in others (e.g. Whitehead and Bower 2014). Second, two of the

trait pairs with apparent contradictions include "cardenolide" as a single class of defense

(measured as totally concentrations in Monarchs 2005). However, the synergies and antagonisms

never involved the same specific cardenolide compound, and so while they suggest that

cardenolides as a class may produce either synergies or antagonisms in combination with other

traits, this may not be the case with individual compounds.

Multivariate plant defense strategies have been explored empirically through

observations of trait co-expression. This can be done at the level of between-species comparisons

(e.g. Agrawal and Fishbein, Moles et al. 2013, Muniz da Silva and Batalha 2011, Sanczuk et al.

2020), between-population comparisons (e.g. Barton 2014), or between individuals within a

population (Agrawal et al. 2014). Trait co-expression can provide indirect support for synergies, but

these same signals might represent other forms of adaptive advantage or disadvantage for specific

combinations. The statistical framework we provide here helps identify, explicitly, when two traits

are associated with synergistic (or antagonistic) impacts on herbivores. Among *Asclepias* spp.,

comparative work has found interspecific patterns of co-expression of latex and trichomes, but

comparisons between species did not show a correlation between latex and cardenolide expression

(Agrawal and Fishbein 2006).

Tests of synergies between classes of defenses (rather than between secondary metabolites

within a class) are rare, but ours is not the first. There have been a few studies in aquatic systems

finding synergies between structural components (calcium carbonate, spicules) and secondary

metabolites in artificial diets simulating seaweeds (Hay et al. 1994) and sponges (Hill et al. 2005,

Jones et al 2005). Simpson and Raubenheimer (2001) found that tannic acids had an effect on

locust only for some levels of protein-to-carbohydrate ratios in the artificial diets, suggesting that

plant digestion inhibitors might synergize with strategies of low nutrient availability. Perhaps the

strongest test for cross-class synergies in terrestrial plants used gene knockouts in *Nicotiana*

attenuata to show that protease inhibitors (TPI, a form of digestion inhibitor) synergized with

toxic nicotine to defend against generalist *Spodoptera exigua* (Stepphun and Baldwin 2007).

Follow-up work in the same system found that TPI reduced caterpillar response to simulated

predation, suggesting a potential defense synergy between parasitoid-attracting volatiles and TPI;

however, field experiments found plants with higher TPI actually had *reduced* parasitism

(Schuman et al. 2012). Similar to Stepphun and Baldwin (2007), Green et al. (2001) found phytic

acid (an anti-nutritive defense) reduced insect metabolism of defensive xanthotoxins, implying a

potential synergy with phytic acid interfering with specialist herbivore counter-defenses;

however, this was not extended to measure the ultimate effect of this potential synergy on

herbivore survival or growth.

The rarity of synergy tests between classes of defenses is likely due in part to the

logistical challenges involved in such experiments. The overwhelming majority of synergy tests

to date have used artificial diets (e.g. Whitehead et al. 2021), and non-metabolite defenses like

latex, leaf toughness, or trichomes do not lend themselves to such controlled experiments.

Similarly, the most elegant approaches to date for testing phytochemical synergies identify an

interaction index based on the dose necessary to cause 50% mortality (ED_{50}) (Nelson and Kursar

1999, Tallarida 2000, Richards et al. 2016). These tools are well-suited to artificial diet

experiments of secondary metabolites in which an ED_{50} is certain to exist, but are difficult to

implement when defenses cannot be manipulated and even high levels of putatively defensive

traits have no guarantee of causing mortality, as with many specialist insects.

Conclusion

The regression approach we use for hypothesis testing is a natural extension of ANOVA

tests for interactions, with the advantage that it can easily be applied to studies with naturally

occurring trait levels which can vary continuously (Richards et al. 2016). However, regression

models with an interaction term inherently assume that the interaction term is bilinear (i.e.,

proportional to the product of the two defense trait levels). Our application of Random Forests,

while more complicated than traditional ANOVA analyses, captures the same essential feature: it

identifies when two traits have a different effect in combination than we would expect from their

separate effects. However, the Random Forests analysis has the key advantage that it makes no

assumptions about the functional forms of the independent effects and the interaction. This

approach to leverage Random Forests to identify interactions has been outlined in the statistical

literature (Hooker 2007, Hooker and Mensch 2019), but to our knowledge has not yet been used

in an ecological context. Our simulations of known synergies and antagonisms showed

consistent diagnostic patterns (Fig. 2), suggesting that this approach could be used broadly to

identify non-additive synergies and antagonisms in many contexts. Here we have demonstrated

the value of combining two approaches: finding strong evidence with regression models for a

previously hypothesized synergy between latex and cardenolides, and finding four synergies and

antagonisms using Random Forests which were not identified using regression models.

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Table 1: Summary of testing a priori hypotheses; metrics are presented only for interactions with

P values higher than 0.01. PRR is the "Predictable Response Range", as described in the

methods.

Experiment	Interaction	P value	slope	PRR	PRR units	% PRR explained by interaction
Monarch 2005 Growth	latex x cardenolide	0.319				
	latex x trichomes	0.1272				
Monarch 2005 Survival	latex x cardenolide	0.7178				
	latex x trichomes	0.7496				
Monarch 2015 Growth	latex x trichomes	0.36				
	latex x cardenolide 10.5	0.782				
	latex x cardenolide 18.6	0.93				
	latex x card15 pca	0.427				
Monarch 2015 Survival	latex x trichomes	0.837				
	latex x cardenolide 10.5	0.473				
	latex x cardenolide 18.6	**0.0767**	-0.7236	0.0114 -- 0.9999	survival probability	69%
	latex x card15 pca	0.365				
Clivicolis 2016 Growth	latex x trichomes	0.28				
	latex x cardenolide 8.4	0.842				
	latex x cardenolide10.2	**0.0333**	-0.03736	-1.2719 -- -0.7359	growth rate	30%
	latex x card16 pca	0.665				
Clivicolis 2016 Survival	latex x trichomes	0.538				
	latex x cardenolide 8.4	0.974				
	latex x cardenolide10.2	**0.1**	-2.085	0 -- 0.9824	survival probability	19%
	latex x card16 pca	0.501				

Figures

Figure 1: Visualizations of synergies for the three significantly supported *a priori* synergies

between latex and cardenolides. Heatmaps show the projections for survival or growth from

fitted models across the empirically observed ranges of the synergistic traits, assuming all other

traits are held constant and blocking factors are at their average. (a) Survival of Monarchs in

2015 is dramatically lower when cardenolides and latex are expressed together. (b) Cardenolides

and latex suppress growth of *L. clivicollis* far more when combined than either does separately.

(c) *L. clivicollis* survival is largely unaffected when latex or cardenolide 10.2 are expressed

separately, but any amount of combination leads to a precipitous decline in survival.

Figure 2: Simulations of diagnostic plots with known synergies or antagonisms. For each

subplot, we first show the per-unit benefit of investment in trait Y (when Y is at 0.5) as trait X is

varied; then the predictions of herbivore performance (this could represent growth or survival

depending on units and rescaling); then the difference in herbivore performance between

additivity (e.g. defenses are independent) and the actual relationship, where positive values show

herbivores over-performing the additive model, and negative values show herbivores

underperforming the additive model. (a-c) for synergistic relationships, herbivores underperform

the additive model when both defense traits are high or both are low, and over-perform when one

trait is high and the other is low. (d-f) for anti-synergistic relationships, herbivores over-perform

the additive model when both traits are high or both are low, and over-perform when one trait is

high and the other is low.

Figure 3: visualization of non-additivities identified using Random Forest analyses (see Fig. 2

for reference plots). For each of the six trait combinations, the left panel shows growth or

survival predictions for a 10x10 grid of trait values based on the fitted Random Forest model,

with points identifying trait values of observed plants (predictions for grid cells with no observed

plants are not plotted). The paired non-additivity panel shows the difference between Random

Forest predictions, and what would be predicted with no trait interactions. Because the no-

interaction predictions are based on an ANOVA model with traits as factors, rows and columns

with a single cell are always fit exactly, and so are removed from the non-additivity panels. The

non-additivity panel can be used to identify synergies or antagonisms (see Fig. 2). (a-c) notable

defense trait synergies; (d-f) notable defense trait antagonisms.

Figure 1

Figure 3