A general framework for identifying rare variant combinations in complex disorders

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ABSTRACT

Statistical challenges due to rarity and combinatorial explosion resulting from exhaustive evaluation of rare variant combinations have limited the study of oligogenic etiology for complex disorders. We present *RareComb*, a framework that combines apriori algorithm and statistical inference to identify specific combinations of mutated genes associated with complex phenotypes. Using *RareComb* on 6,189 affected individuals, we identified 718 combinations of mutated genes significantly associated with intellectual disability (ID), and carriers of these combinations showed lower IQ than expected in a replication cohort of 1,878 individuals. These combinations were enriched for nervous system genes, showed complex inheritance patterns, and were depleted in unaffected siblings. We further identified oligogenic combinations associated with multiple comorbid phenotypes, including *COL28A1* and *MFSD2B* mutations for ID and schizophrenia. Our framework enables rare variant analysis in affected individuals lacking diagnosis based on *de novo* mutations, and provides a paradigm for dissecting the genetic basis of complex disorders.

INTRODUCTION

Recent human population growth has led to a rapid increase in the load of rare variants affecting functionally important regions of the genome[1–3]. Thus, rare variants are collectively more abundant in the population compared to common variants, many of which confer significant risk for neurodevelopmental disorders such as autism and intellectual disability[4]. In fact, recent studies have directly implicated rare damaging mutations that are very recent or *de novo* in more than one hundred genes towards neurodevelopmental disorders[5–7]. The ability to establish robust associations between rare variants of high effect size and complex disease has made this class of variants the primary focus of recent studies. However, a much larger class of rare and variably expressive variants that are individually less deleterious but, in combination, exert large effects towards disease is often overlooked. Variants in this category are often transmitted across generations without adverse effects on their carriers until they encounter other similar variants that, when combined, lead to genetic interactions conferring a higher risk for disease than their individual risks[8,9]. While this phenomenon underpins oligogenic models proposed over the years, studies so far have not focused on detecting combinatorial effects of specific sets of rare variants towards disease phenotypes[10–13].

Identifying the effects of specific combinations of rare variants towards disease etiology has been challenging for many reasons. *First*, combinations of rare variants are rarer, and extremely large cohorts are required to observe even a few recurrent instances of specific variant combinations[14]. Prior studies of oligogenic models for rare variants evaded this problem by aggregating variant information at the sample level and comparing the overall burden of rare variants between groups of individuals (such as cases and controls)[6,7,15,16]. *Second*, the combinatorial explosion resulting from even a small set of rare variants makes it difficult to exhaustively evaluate all combinations. While sophisticated frameworks such as network analysis and machine learning provide powerful tools to model the composite effects of thousands of variables on a complex system and predict emergent behaviors and quantitative outcomes, adapting them to exhaustively search and delineate the effects of specific combinations of variables is daunting[17,18]. Furthermore, incorporating an efficient search tool into these frameworks and extending them to detect higher-order combinatorial effects would be nearly impossible. *Third*, even when all combinations of rare variants could be exhaustively evaluated within a large cohort, there is a lack of methods that are sensitive enough to detect

small differences between comparison groups to establish statistical significance. Therefore, an alternate approach that is highly flexible, scalable, and sensitive is necessary to address computational and statistical challenges associated with assessing rare variant combinations.

Here, we present a combinatorial framework called *RareComb* that couples the apriori algorithm[19] with binomial tests to overcome the limitations of data sparsity and high dimensionality, and systematically analyzes patterns of rare variants between groups of interest to identify specific combinations that are significantly associated with phenotypes[20]. We apply our analysis framework to a discovery cohort of 6,189 children with autism to identify genetic interactions involving pairs and triplets of mutated genes that are enriched in individuals with intellectual disability compared to individuals without intellectual disability. We demonstrate that the carriers of mutations in these specific gene pairs and triplets within an independent cohort of 1,878 children have significantly lower-than-expected intelligence quotient (IQ) scores. We also demonstrate the adaptability of our framework by leveraging it to identify mutated gene pairs and triplets significantly associated with two or more comorbid phenotypes among children with autism. Finally, we show how this generalizable and modular framework can be easily extended to identify higher order interactions beyond pairs and triplets of variants. Our stand-alone framework does not depend on *a priori* knowledge and can detect rare patterns from high-dimensional genetic data to generate interpretable results, making it readily applicable for analyzing cohorts of all size ranges to dissect the genetic basis of complex disorders.

RESULTS

We hypothesized that two or more genes disrupted simultaneously by rare deleterious mutations contribute to a highly penetrant phenotype, as in an oligogenic model, or lead to a more severe phenotype than when each of the same genes are disrupted individually. We developed *RareComb* as a framework that combines data mining and statistical analysis to identify specific combinations (such as pairs, triplets, etc.) of rare variants that show significant associations with one or more phenotypes. *RareComb* analyzes an '$n \times p$' sparse Boolean matrix with 'p' genes in 'n' individuals in three discrete steps (**Figure 1**). *First*, it applies the apriori algorithm independently in cases and controls to enumerate the frequency of all simultaneously mutated combinations that meet a pre-set minimum frequency threshold (**Supp. Figure 1**). *Second*, for each qualifying combination of variants, the method derives the expected frequency of simultaneously observing mutations in the constituent genes under the assumption of independence. It then independently quantifies the magnitude of deviation of the observed from the expected frequencies using binomial tests in cases and controls, and uses multiple-testing adjusted p-values to identify combinations that are statistically enriched in cases but not in controls. *Finally*, the method calculates effect sizes using Cohen's d and statistical power at 1% and 5% significance thresholds, to enable prioritization of a high confidence set of combinations that contribute to the phenotype in an oligogenic manner.

RareComb identifies oligogenic combinations associated with ID and autism

We sought to identify pairs and triplets of mutated genes that are significantly associated with intellectual disability (ID) phenotypes by analyzing 6,189 affected individuals from the SPARK[21] cohort for discovery and 1,878 affected individuals from the SSC[22] cohort for validation. To facilitate cross-cohort comparison, we identified 10,217 rare variants (MAF≤1%) that were predicted to be deleterious by multiple methods and observed in both cohorts, and aggregated these variants to genes for the analysis (see **Methods**). We first categorized 1,215 probands from the SPARK cohort diagnosed with ID as cases and 4,974 probands without ID as controls (**Figure 2A**). We then applied *RareComb* to the Boolean matrix and detected 25,602 pairs involving 1,956 mutated genes in cases that were observed at a higher frequency than expected under the assumption of independence. Similarly, analyzing the controls using only the 1,956 genes mutated in cases, *RareComb* identified 148 pairs of mutated genes that were

128 significantly enriched in cases but not in controls (**Supp. Table 1)**, with moderate to high effect
129 sizes (Cohen's d, 0.08-0.15) and adequate statistical power (70%-100% at 5% significance
130 threshold) (**Supp. Figure 2**). These 148 gene pairs belonged to 142 probands, with 74%
131 (105/142) of them carrying more than one significant pair. These observations suggest that an
132 individual can carry multiple combinations, each contributing to the same phenotype at varying
133 effect sizes (**Supp. Figure 3**).

134 We next sought to validate the association of these 148 mutated gene pairs towards
135 intellectual disability. We hypothesized that if the association of the gene pairs with ID in the
136 SPARK cohort were truly significant, carriers of mutations in those gene pairs would tend to
137 have lower than average IQ scores in the independent SSC cohort. We found that 90 of the 148
138 significant pairs identified in the SPARK cohort were observed in at least one proband in the
139 SSC cohort. These 90 mutated gene pairs were carried by 91 unique probands, whose average
140 full-scale IQ scores (average IQ=68.52) were lower than those of all ascertained probands in the
141 SSC cohort (average IQ=86). To assess the significance of this result, we performed 10,000
142 random draws of 91 probands from the SSC cohort to generate a simulated distribution of their
143 average IQ scores. The average IQ of carriers of mutated gene pairs (average IQ=68.52) was
144 significantly lower than the overall distribution of average IQ derived from simulations (average
145 IQ ranged from 73 to 92; empirical p=0) (**Figure 2B**). Furthermore, the average IQ of the 91
146 SSC probands with both mutated genes was significantly lower than the average IQ of 1,252
147 carriers of mutations in only one of the two genes (68.5 versus 82.8; Kolmogorov-Smirnov p =
148 1.302×10^{-16}) (**Figure 2C**). When each of the 90 combinations was evaluated individually,
149 carriers of mutations in both genes for 73% (66/90) of the combinations showed lower IQ than
150 individuals with mutations in individual genes of the same combination, with 39/90 remaining
151 significant after multiple testing correction (**Supp. Table 2; Supp. Figure 4**). These results
152 provide evidence for synergistic effects of deleterious mutations within specific pairs of genes
153 towards ID phenotypes.

154 We also applied *RareComb* to identify gene triplets associated with intellectual disability
155 using the two cohorts and repeated the simulations to identify 1,593 significant combinations in
156 the SPARK cohort. We selected 570 high-confidence triplets (with ≥90% statistical power at 5%
157 significance threshold; **Supp. Table 3**) and found that 79 probands in the SSC cohort carried at
158 least one of these deleterious triplets. The average IQ score of individuals carrying significant

159 gene triplets (average IQ score=73) was significantly lower than a distribution of average IQ
160 scores from 10,000 draws of 79 SSC probands (average IQ score=82.5; min=72, max=94;
161 empirical p=0.0011; see **Supp. Figure 5**). This result reiterated that carriers of mutations in the
162 significant gene combinations have lower IQ than a random group of probands. Our results also
163 demonstrate the ability of the framework to identify higher order combinations of mutations that
164 are significantly associated with specific phenotypes in individuals with complex disorders.
165

166 **Oligogenic combinations are enriched for specific inheritance patterns**

167 As individual variants can arise *de novo* or be inherited maternally or paternally, variants in pairs
168 of genes can have six possible patterns of transmission (**Supp. Figure 6A**). We identified a total
169 of 926 occurrences of the 148 pairs of mutated genes enriched among SPARK probands with ID
170 (n=142 probands), of which inheritance could be determined without ambiguity for 887
171 instances. We found that one variant occurred *de novo* and the other variant was inherited from
172 the mother in 244/887 instances (27.5%). Similarly, both mutated genes were inherited from the
173 mother in 226/887 instances (25.4%) or occurred *de novo* in 221/887 instances (24.9%), while
174 the remaining fraction (~22%) of variant pairs were either inherited from both parents, inherited
175 from the father, or transmitted *de novo* and paternally. To assess the significance of our
176 observations, we performed simulations to establish a baseline expectation of proportions for
177 each category of parental inheritance pattern. We selected 926 pairs of genes in 1000 random
178 draws of all possible mutated gene pairs among SPARK probands and calculated the fraction of
179 instances that fell into each of the six transmission categories. The observed proportion was
180 higher than the simulated proportions for instances when both variants occurred *de novo* (24.9%
181 versus 17%, empirical p=0) and when one variant was *de novo* and the other was inherited
182 maternally (27.5% versus 25%, p=0.028) (**Figure 3A**). We repeated this analysis for 7,596
183 children affected with autism in the SPARK cohort compared to 11,740 unaffected parents and
184 identified 110 gene pairs significantly associated with autism (**Supp. Table 4**). Similar to the
185 results obtained for the ID phenotype, we found that both variants of a gene pair were more
186 likely to occur *de novo* (24% versus 18%, empirical p=0) or one variant occurring *de novo* and
187 the other inherited maternally (33% versus 26%, p=0) than expected based on simulation studies
188 (**Supp. Figure 7**). The enrichment of *de novo* or maternally inherited variants for significant
189 gene pairs aligns with published reports that severely affected children tend to carry multiple *de*

190 *novo* mutations or inherit pathogenic rare variants from mildly affected or unaffected carrier
191 mothers[16,23,24].
192 We then assessed whether the mutated gene pairs associated with ID were also found in
193 siblings of carrier probands. Restricting our analysis to families with unaffected siblings whose
194 probands had mutations in ID-enriched gene pairs, we found that both variants were present in
195 the corresponding sibling for only 53/219 (24.2%) instances of gene pairs, while 102/219
196 (46.6%) had variants in only one of the two genes and 64/219 (29.2%) instances had no variants
197 in either of the genes in the siblings (**Supp. Figure 6B)**. Using simulations, we found a
198 significantly higher proportion of instances with only one of the two variants present in siblings
199 compared to the expected values (46.6% versus 38.5%, p=0.007). Furthermore, the proportion of
200 observed instances with neither of the variants present in siblings (29.2% versus 33.1%,
201 empirical p=0.098) or both variants present in siblings (24.2% versus 28.4%, p=0.079) was
202 lower than expected (**Figure 3B)**. The observation that only a small fraction of unaffected
203 siblings carried both mutated gene pairs suggests a strong association of these gene pairs with ID
204 phenotypes. These results suggest that mutations in pairs of genes significantly associated with a
205 severe phenotype in probands are more likely to occur individually than simultaneously in
206 unaffected siblings of the same family.
207
208 **Genes forming oligogenic combinations are distinct from canonical autism genes**
209 We expanded our analysis to include all 16,556 mutated genes in the SPARK cohort, as opposed
210 to genes with mutations present in both the SPARK and SSC cohorts, and identified 52
211 significant gene pairs (**Supp. Table 5)** and 230 triplets associated with the ID phenotype (with
212 ≥90% statistical power at 1% significance threshold; **Supp. Table 6)**. Due to the expanded
213 search space, the mutated gene pairs showed more significant p-values from the binomial tests
214 when compared to those obtained from the more restricted set of variants overlapping both
215 SPARK and SSC cohorts (**Supp. Figure 8**). Mutated genes within these combinations included
216 several genes related to nervous system development, such as *NIN*, *HDC*, *NGF*, and *BRD8.*
217 Furthermore, 5/52 pairs and 59/230 triplets contained at least one gene associated with autism in
218 the SFARI database, including *FGFR1*, associated with multiple disorders including Kallmann
219 syndrome[25] and Pfeiffer syndrome[26]; *RELN*, associated with temporal lobe epilepsy[27]; *SYNE1*,
220 associated with spinocerebellar ataxia[28,29]; and *PNPLA7*, associated with autism and ID[30]. Thus,

221 most genes forming combinations are not involved in canonical autism or ID disorders,
222 suggesting synergistic effects of these genes without prior association to disease.
223 We also performed gene ontology enrichment analysis for genes within the combinations
224 and identified seven out of nine significantly enriched GO terms to be exclusively associated
225 with nervous system-related functions, including synthesis and metabolism of catecholamines,
226 axon/neuron regeneration, and neuron generation and differentiation (**Supp. Figure 9**)[31].
227 Furthermore, the differences in the type and specificity of GO terms enriched for significant
228 pairs versus triplets were apparent, with genes forming pairs involved in nervous system function
229 and genes forming triplets associated with both nervous system as well as other biological
230 processes. We next assessed the enrichment and depletion of Human Phenotype Ontology (HPO)
231 terms for genes forming significant pairs towards ID phenotypes[32]. *First*, we calculated the
232 fraction of all 4,484 genes within the HPO database associated with each HPO term. *For*
233 *example*, 30% (1,366/4,484) of all genes in HPO were associated with ID. We compared these
234 expected values calculated for each HPO term with the corresponding fractions observed within
235 the 95 genes forming 52 ID-associated pairs using binomial tests. Interestingly, genes associated
236 with HPO terms related to neurodevelopmental phenotypes, such as ID, global developmental
237 delay, seizure, and microcephaly, were significantly depleted within the set of 95 genes forming
238 gene pairs (**Supp. Table 7**). *Next*, we evaluated whether genes within each of the 52 significant
239 pairs shared one or more common HPO phenotype or disease. Of the 52 pairs, only one pair
240 (*DNASE1* & *MTR*) shared an HPO phenotype ("epilepsy"). This was significantly lower than the
241 expected value obtained from the distribution of the number of shared HPO phenotypes between
242 all possible pairs of genes in the HPO database (1/52, 1.9% ID gene pairs compared to 31.5% of
243 all HPO gene pairs shared one HPO phenotype, $p= 2.2\times10^{-16}$; one-sided binomial test) (**Supp.**
244 **Figure 10; Supp. Table 8**). We note that the 4,484 genes within HPO are potentially biased
245 towards well-studied disorders, making pairs of genes drawn from HPO more likely to share
246 phenotypes than random pairs of genes from the genome. Overall, GO and HPO analyses show
247 that genes forming oligogenic combinations are involved in neuronal processes but have not
248 been previously connected to neurodevelopmental phenotypes, indicating the novelty of the
249 associations between these genes and ID phenotypes.
250
251

252 **Identifying variant combinations towards specific patterns of comorbid phenotypes**
253 We adapted our framework to identify significant associations of two or more genotypes with
254 multiple comorbid phenotypes. To identify novel comorbid associations, we eliminated
255 phenotypes that were highly correlated with each other, such as ADHD and reading disorder[33].
256 We analyzed variant profiles of 6,189 autism probands from the SPARK cohort with records of
257 comorbid features, including 1,215 individuals with ID, 1,825 with anxiety and depression, and
258 332 with schizophrenia features. We assessed for significant co-occurrences of two or more
259 mutated genes with two or more of the above phenotypes (**Figure 4**). Using one-tailed binomial
260 tests to compare the observed frequency of combinations of genotypes and phenotypes to the
261 expected frequency, we first identified 169 significant associations between pairs of mutated
262 genes and two comorbid phenotypes as well as 82 combinations of three mutated genes and two
263 comorbid phenotypes (**Supp. Tables 9 & 10**). As some of these significant genotype-phenotype
264 combinations can be confounded by high degree of co-occurrence of mutated genes, we next
265 calculated genotype-only p-values using binomial tests for all significant genotype-phenotype
266 associations. For 32/169 combinations of two mutated genes and two comorbid phenotypes and
267 5/82 combinations of three mutated genes and two comorbid phenotypes, the composite
268 genotype-phenotype p-values were significant while genotype-only p-values were not
269 significant, suggesting stronger associations between these variant combinations and phenotypes.
270 *For example*, even when variants in genes *COL28A1* and *MFSD2B* did not co-occur more
271 frequently than expected under the assumption of independence, these mutated genes co-
272 occurred more frequently than expected among probands with ID and schizophrenia phenotypes.
273 Loss-of-function and rare missense mutations in *COL28A1* have been reported in individuals
274 with autism[34,35], and *MFSD2A, a* paralog of *MFSD2B,* has been directly implicated in an
275 autosomal recessive disorder associated with progressive microcephaly, spasticity and brain
276 imaging abnormalities[36]. Similarly, we found *ARVCF* and *FAT1* to be significantly associated
277 with ID and schizophrenia, with *ARVCF* mapping within the 22q11.2 DiGeorge syndrome
278 region[37], while rare *de novo* mutations in *FAT1* being associated with autism and
279 schizophrenia[6,38]. Finally, we found that the mutations in genes *ABCA4*, *DNAH10* and *MC1R*
280 significantly co-occurred in individuals with ID and anxiety/depression phenotypes. These
281 results demonstrate the utility of identifying higher-order associations between genotypes and
282 phenotypes in complex disorders such as autism.

DISCUSSION

Current rare variant analysis strategies are geared towards either searching for individual variants of high effect size whose influence on the phenotype is evident, such as *de-novo* gene-disruptive mutations, or comparing rare variant burden to explain collective effects on phenotypes[7,39,40]. The wider space between these two extremes of the analysis spectrum that involves combinations of rare variants has largely remained understudied. Although digenic diseases and multi-hit models of complex diseases have been used to provide post-hoc explanations for an observed phenomenon, they are not equipped to serve as a framework to actively search and identify rare variant combinations that fit oligogenic models for specific phenotypes[9,12,13]. While machine learning has become the de-facto approach for disease outcome predictions, the lack of holy-grail predictors and reduced interpretability due to data sparsity makes it less fit to detect combinatorial effects[17]. In addition, the common practice of evaluating feature importance metrics of machine learning classifiers falls short of the objective to identify combinations of features that exert higher effect on the phenotype than evident from their independent effects[17,18]. Furthermore, prior studies to assess combinatorial effects have been inherently biased due to their need to minimize the search space by restricting the analysis to only a subset of genes chosen based on *a priori* knowledge[41–43]. Here, we provide a proof-of-concept analytical framework that remains agnostic to prior evidence and performs exhaustive searches to identify combinatorial effects among rare variants while retaining high granularity of data and interpretability of results.

We use our framework to identify gene pairs and triplets significantly associated with intellectual disability and show that several constituent genes are associated with nervous system processes. These mutated gene combinations are more likely to be inherited maternally or occur *de novo*, are depleted in unaffected siblings from the same family, and are less likely to involve canonical autism or ID genes, suggesting that genes forming significant combinations are less deleterious on their own but manifest effects only when combined with other similar genes carrying rare mutations. While previous studies have linked aggregate rare variant burden towards intellectual disability[44,45], our results fine map the association to specific combinations of constituent genes contributing to the burden. We propose a novel paradigm for dissecting the complexity of genetic disorders, where an affected individual carries multiple combinations of rare variants, and each combination contributes to either the same phenotype or distinct

314 phenotypes at varying effect sizes (**Figure 5**). A limitation of our method is that it tends to be
315 biased towards genes that are mutated frequently enough to be observed in a combination. This
316 limitation can be addressed by fixing specific primary variants of interest irrespective of their
317 frequency and screening for "second-hit" modifiers that significantly co-occur with the primary
318 variant, such as the co-occurrence of *RBM8A* variants in distal 1q21.1 deletion carriers
319 manifesting thrombocytopenia-absent-radius syndrome and *TBX6* variants in 16p11.2 deletion
320 carriers with scoliosis[46,47].
321 Our method is fast and scalable, allows for fine-tuning combinatorial searches based on
322 frequency, statistical power, and multiple testing criteria, and can be adapted to enable
323 computational approximations to further improve run time and assess higher-order combinations
324 beyond triplets. While larger sample sizes are generally required for detecting smaller frequency
325 differences, we note that our framework achieves reliable statistical power even with modest
326 sample sizes, implying that our framework could be applied to exome sequencing studies of
327 other neurodevelopmental disorders that have not been explored for combinatorial effects. This
328 approach can also be used to address a variety of research questions involving rare event
329 combinations, including searching for protective effects of rare variants where simultaneous
330 mutations are enriched in controls but not in cases, and finding combinations that exhibit specific
331 enrichment or depletion patterns in more than two phenotypic groups. In summary, we provide a
332 conceptual framework and the necessary tools to identify the oligogenic basis for complex
333 disorders such as autism and intellectual disability, which hitherto was restricted to the analysis
334 of canonical disorders such as Hirschsprung disease[48] and Bardet-Biedl syndrome[12].
335
336

MATERIALS AND METHODS

We developed *RareComb* to address computational and statistical challenges associated with combinatorial analysis of rare variants. *RareComb* first uses the apriori algorithm to efficiently count the frequencies of co-occurring variant combinations. It then uses one-tailed binomial tests to compare the observed frequency of each variant combination to the expected frequency derived under the assumption of independence among the constituent variants within each combination (**Figure 1**). This method can be applied to identify variant combinations that are significantly enriched in cases but not in controls. In studies involving multiple comorbid phenotypes, this method can also be used to detect associations between specific combinations of variants and one or more (comorbid) phenotypes (see **Supplementary Note**). The general principles of our method, built using the basic axioms of probability theory, can be easily extended to a variety of problems involving rare higher-order combinations (**Supp. Figure 11**).

Identifying frequencies of rare variant combinations

RareComb utilizes the apriori algorithm to efficiently calculate frequencies of variant combinations from sparse Boolean matrices (of 0s and 1s) (**Supp. Figure 12A**). The apriori algorithm has been successfully applied to analyze consumer behavior, where identifying products frequently purchased together could benefit a company[49,50]. While an algorithm that is used to derive insights from patterns within highly frequent events (i.e. frequent itemset mining) might not seem like a good fit to analyze rare variant combinations, its ability to perform disciplined search based on both built-in and user-specified constraints makes it an ideal counting tool. *For example*, the apriori algorithm avoids enumerating each of the 50 million pairs or 167 billion triplets from just 10,000 variants, and instead prunes the search-space based on user-defined criteria such as minimum frequency threshold and size of combinations (pairs, triplets, etc.) (**Supp. Figure 12B**). *RareComb* applies an additional constraint to the algorithm to limit its search to co-occurring events, which further reduces the search space (see **Supplementary Note)**. *For example*, when considering variants A and B, only the frequency of the presence of both variants (A=1 & B=1) is counted, and not absence of either or both variants (A=1 & B=0; A=0 & B=1; or A=0 & B=0).

368 **Statistical Inference**

369 *RareComb* utilizes the p-values of one-tailed binomial tests to establish the magnitude of
370 enrichment for each rare variant combination (**Figure 1**). For each combination, *RareComb*
371 formulates null and alternate hypotheses for the binomial test by considering the event of
372 observing all constituent variants together within a group of individuals as success and all other
373 possibilities as failure in a binomial trial:

374 $$H_0 : \pi = \pi_0$$

375 $$H_a : \pi > \pi_0$$

376 where,

377 π = Probability of *observing* all constituent rare variants of a combination together within
378 a cohort, i.e., P(A=1 & B=1)

379 π_0 = *Expected* probability derived from the frequency of individual variants of a
380 combination, under the assumption of independence, i.e., P(A=1) * P(B=1).

381 *RareComb* then compares the null binomial distribution derived using the sample size of the
382 group (n) and the expected probability (π_0) (i.e., X ~ Binom(n, p = π_0)) with the observed
383 probability (π), and calculates the probability of observing rare variants occurring together at
384 least as frequently as they were observed within the cohort (i.e. p-value).

385 In case-control analyses, this method is applied independently to each group, and the p-
386 values between them are compared. The combinations exhibiting enrichment in both cases and
387 controls, likely due to proximity of variants in linkage disequilibrium, are eliminated, following
388 which the p-values in cases are adjusted for multiple-testing to identify statistically significant
389 combinations that exhibit enrichment in cases but not in controls. Finally, the effect sizes are
390 calculated using Cohen's d and the statistical power is measured using 2-sample 2-proportion
391 tests, as additional metrics to prioritize the final set of significant rare variant combinations. In
392 genotype-comorbid phenotype association analyses, the method is applied just once to the entire
393 cohort, with multiple-testing adjusted p-values serving as a sufficient metric to identify high
394 quality associations between genotypes and two or more co-occurring phenotypes.

395

396 **Statistical power and computational performance of the method**

397 We measured the relationship between sample size and statistical power for both binomial and 2-
398 sample 2-proportion tests used in the framework. It took 1,356 samples for the binomial test to

399 achieve a statistical power of 80% to establish statistical enrichment between expected and
400 observed co-occurrence frequencies of 0.1% and 0.5% (**Supp. Figure 13**). This number
401 increased to 6,469 when the test needed to be more sensitive to compare frequencies of 0.3% and
402 0.5%. Similarly, it took 7,840 samples for the 2-sample 2-proportion test to achieve 80% power
403 to establish statistical difference between co-occurrence frequencies of 2% and 0.5% observed in
404 two groups (**Supp. Figure 14**). The sample size requirement increased to 14,633 to differentiate
405 frequencies of 1.5% and 0.5% at 80% statistical power. These results align with the known
406 relationship between sample size and statistical power, and indicate that our method can be
407 reliably applied to analyze reasonably modest-size cohorts.
408 We also measured the run times for the case-control analysis to identify significant pairs
409 and triplets of mutated genes using simulated data of three discrete sizes of samples (5,000,
410 10,000, and 50,000 individuals) and genes (5,000, 10,000, and 15,000 genes). The apriori
411 algorithm was run on single-core CPUs with 256 GB memory and was constrained to analyze
412 combinations observed in at least 0.15% of the samples. Given the memory-intensive nature of
413 the apriori algorithm implemented in the 'arules' package, 256 GB was chosen to maintain
414 uniformity[51]. However, smaller input files could be processed successfully using much less
415 memory. As expected, the runtimes were proportional to the size of the combination (pairs
416 versus triplets) and the number of input variables (**Supp. Figure 15**). While the increase in run
417 time with the increase in sample size is apparent for pairs, lower runtimes observed with running
418 50,000 samples compared to 5,000 samples for triplets can be attributed to stochasticity of the
419 input data. Overall, the analysis of gene pairs took between one minute and 12 minutes while
420 triplets took between two minutes and 150 minutes. Since several factors influence the runtime
421 of the method, a trial-and-error approach to determine an optimal minimum frequency threshold
422 for co-occurring events can help identify relevant combinations without resulting in insufficient
423 memory due to combinatorial explosion.
424
425 **Samples**
426 We used whole exome sequencing data from 6,189 affected males from the Simons Foundation
427 Powering Autism Research (SPARK)[21] and 1,878 affected males from 2,247 simplex families
428 from the Simons Simplex Collection (SSC)[52] cohort from the Simons Foundation Autism
429 Research Initiative (SFARI)[53]. We selected only male probands for our analysis to avoid any

430 confounding effect due to gender or ascertainment bias[54,55]. While diagnosis information for
431 intellectual disability (ID), anxiety, attention deficit hyperactivity disorders (ADHD),
432 schizophrenia, language and sleep disorders were encoded as binary variables for the SPARK
433 samples, full-scale intelligence quotient (IQ) scores were available for the SSC cohort.
434
435 **Data preparation and quality control**
436 Variant Call Format (VCF) files obtained from exome sequencing data were annotated using
437 ANNOVAR[56] for rsID information and variant frequency using ExAC[57] and gnomAD[58]. To
438 overcome the limitations of using a single method to predict pathogenicity, the effects of non-
439 synonymous mutations were annotated using 11 prediction methods: SIFT[59], Polyphen2[60]
440 (HDIV), Polyphen2 (HVAR), LRT[61], MutationTaster[62], MutationAssessor[63], FATHMM[64],
441 MetaSVM[65], PROVEAN[66], REVEL[67], and CADD[68]. Briefly, all missense, stop-loss/gain, and
442 start-loss/gain variants within exonic, 3', and 5' UTR regions with minor allele frequencies ≤1%
443 identified based on both ExAC and gnomAD databases were selected. Then, variants with allele
444 depth of ≥15 and allele balance between 25% and 75% for heterozygous variants and > 90% for
445 homozygous variants were selected as high-quality variants. Deleteriousness of the variants were
446 measured and reported differently by each prediction method. REVEL provided a score between
447 0 and 1, with higher scores indicating higher level of deleteriousness, while Polyphen2 and
448 MutationAssessor classified variants into one of three categories. *For example*, Polyphen2
449 classified variants as 'Deleterious', 'Possibly damaging', or 'Tolerated', while MutationAssessor
450 classified variants as 'High', 'Medium', or 'Low'. The other nine methods classified variants as
451 either 'Deleterious' or 'Tolerated'. Pathogenicity reported by each tool was encoded as a binary
452 variable, with the categories 'Possibly damaging' and 'Medium' encoded as 0.5. Thus, the
453 composite pathogenicity score derived from the 10 tools could range between 0 and 10. Missense
454 variants with a cumulative score of ≥4 and stop-loss/gain predicted as 'deleterious' either based
455 on CADD score (CADD phred >30) or MutationTaster were considered deleterious for all
456 analyses. Indels and other smaller structural variants were not considered, as their functional
457 impact could not be easily assessed.
458
459
460

461 **Gene Ontology (GO) and Human Phenotype Ontology (HPO) enrichment analyses**
462 Gene Ontology term enrichment analyses were performed using the 'Gene Ontology API'
463 accessed using the 'post' command of the python package 'requests' (python version 3.7)[31]. All
464 analyses were performed using parameters for *homo sapiens* (organism = '9606') to identify
465 biological processes enrichment (annotDataSet = 'GO:0008150') using binomial tests. HPO
466 enrichment analyses were performed using data from the 'genes_to_phenotype' file obtained
467 from the HPO website[32]. Since enrichment of phenotypes is not automatically evaluated by HPO,
468 we used customized R scripts to derive baseline expectations that could be compared against the
469 actual observations to determine significance using the p-values from binomial tests.
470
471 **Statistical analysis**
472 All statistical analyses were performed using R v3.6.1 (R Foundation for Statistical Computing,
473 Vienna, Austria)[69] and Python (v3.7)[70]. All data-related plots were generated using the R
474 package ggplot2[71].
475
476 **Software Availability**
477 *RareComb* is available as an open-source (https://github.com/girirajanlab/RareComb) R package
478 that can be downloaded from the Comprehensive R Archive Network (CRAN) repository[72]. It
479 can also be installed into development environments via interfaces such as Rstudio[73] using the
480 command *install.packages('RareComb')*. The tool provides several functionalities that allow
481 users to run the types of analyses described in this manuscript. The functionalities are as follows:
482 (1) Identify rare event combinations statistically enriched within a single group; (2) Identify rare
483 event combinations statistically enriched in cases but not in controls; (3) Identify rare event
484 combinations enriched in cases but depleted in controls; (4) Identify statistically enriched rare
485 event combinations that include at least one element from an user-supplied list; and (5) Identify
486 genotypes statistically enriched within individuals manifesting two or more comorbid
487 phenotypes. Each functionality takes a Boolean matrix as input and provides a set of user-
488 adjustable parameters to customize the analysis, and delivers the results in a tabular format as csv
489 files. Detailed instructions on the available functionalities and parameters built into *RareComb*
490 and their usage can be found on the GitHub page or CRAN website. A shiny app illustrating the
491 ideas behind *RareComb* is available online at https://girirajanlab.shinyapps.io/RareComb/ [74].

DECLARATIONS

Ethics approval and consent to participate

As these data were de-identified, all our samples were exempt from IRB review and conformed to the Helsinki Declaration. No other approvals were needed for the study.

Consent for publication

All authors agree and consent for publication of the manuscript.

Competing interests

The authors declare that no competing interests exist in relation to this work.

Authors' contributions

VK and SG conceived the project. VK performed the analyses, generated the plots/images, and wrote and revised the manuscript; SG supervised the research and wrote and revised the manuscript. All authors read and approved the final draft of the manuscript.

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MAIN FIGURES

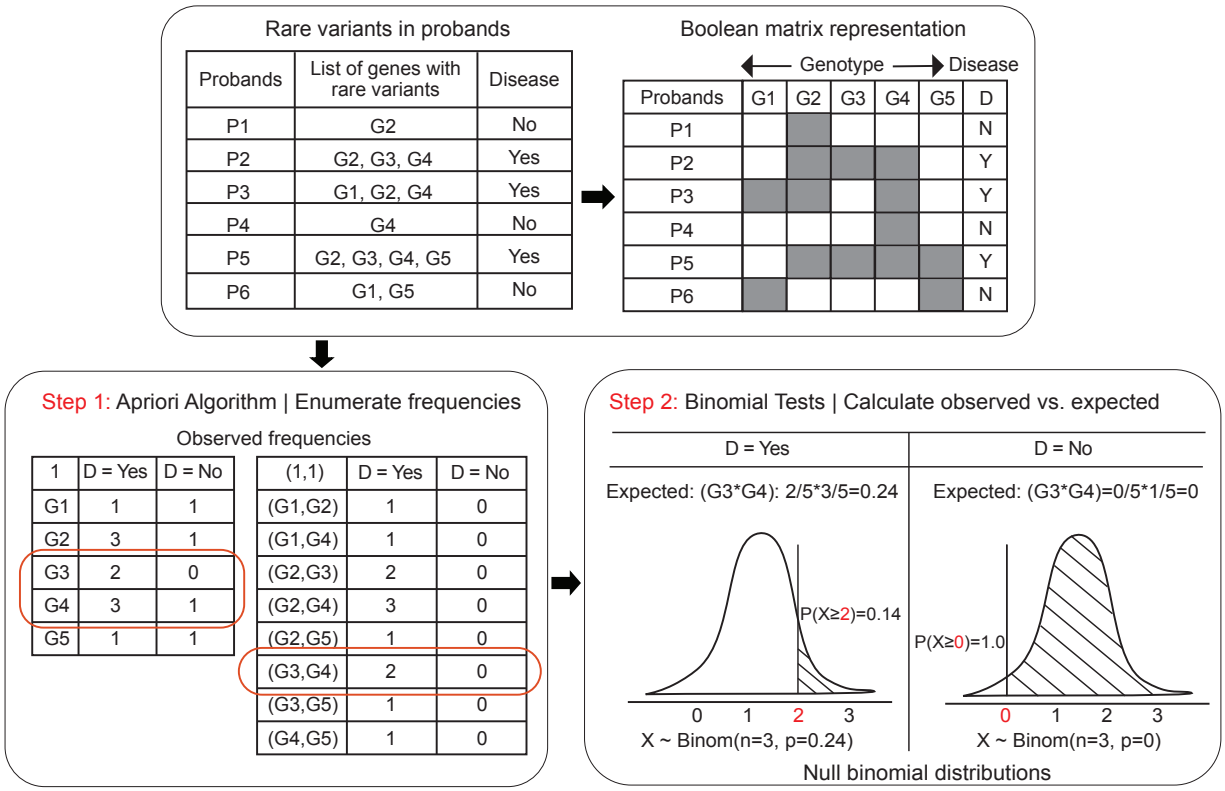


Figure 1: Conceptual overview of combinatorial analyses using *RareComb*. A Boolean representation of genotype (mutated genes, G1, G2, etc) and disease status for probands (P1, P2, etc) is shown. In *step 1*, the apriori algorithm is applied to the Boolean input matrix to calculate the frequencies of individual (for example, G1) and simultaneous occurrences of events (G1 and G2) that meet the user-specified criteria, including the size of combinations (pairs, triplets, etc.) and minimum frequency threshold of simultaneous occurrences. In *step 2*, independently in case and control groups, for each combination, the binomial test is applied to compare the observed frequency of simultaneous occurrence of events with its corresponding null binomial distribution of the expected frequencies calculated under the assumption of independence. Binomial test for gene pair G3 and G4 is shown as an example.

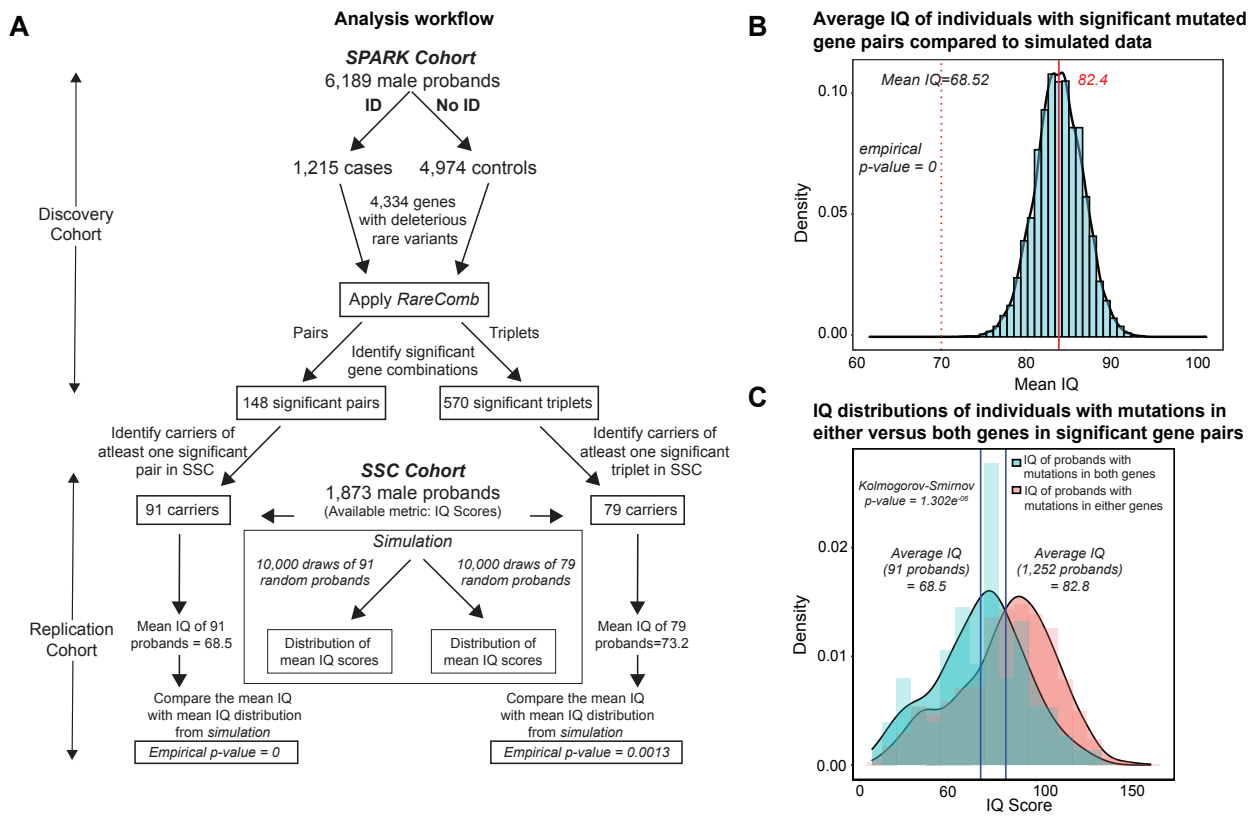
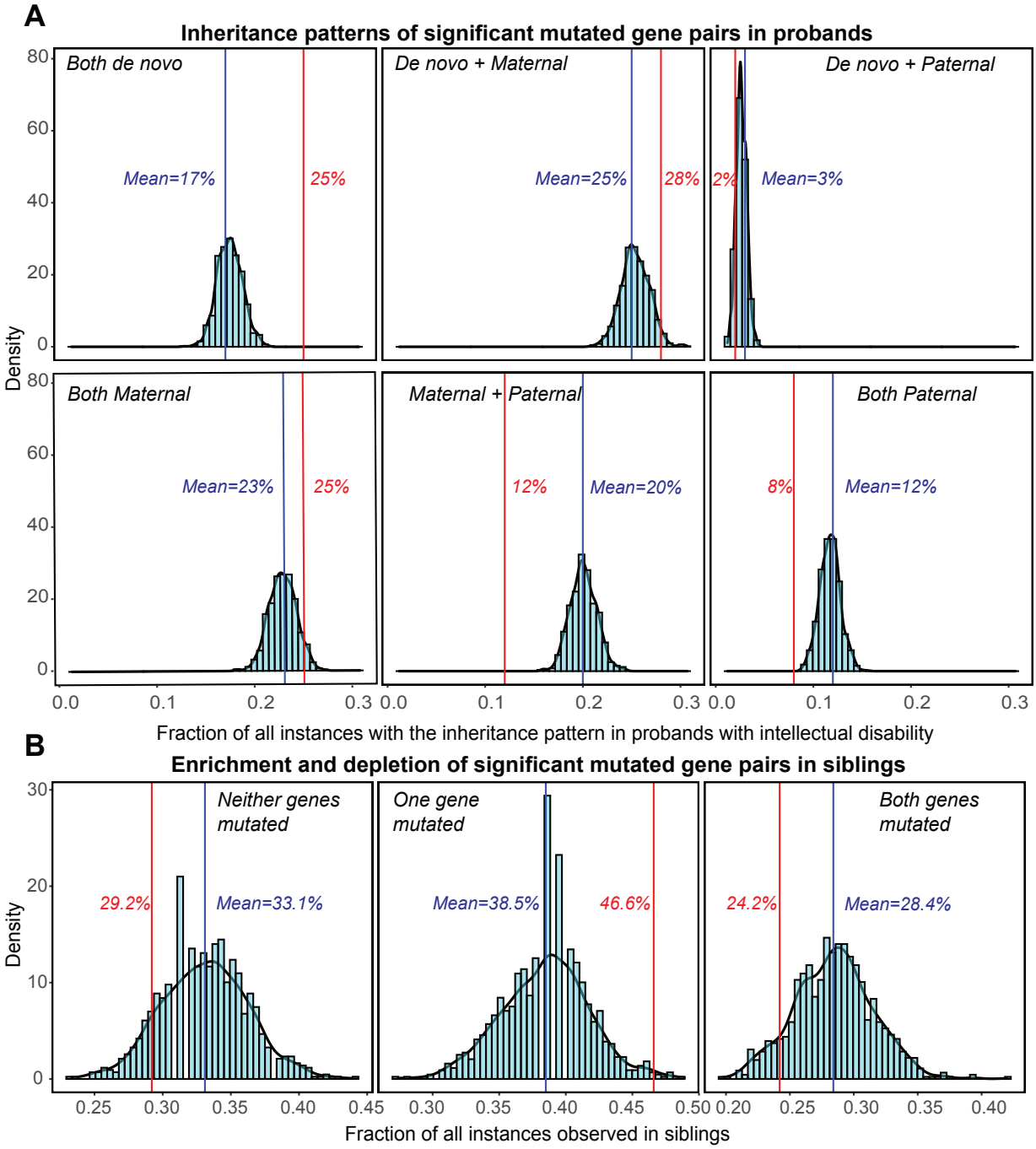


Figure 2: Combinations of rare variants contributing to intellectual disability (ID) phenotype. (A) An outline of the approach used to identify and validate mutated gene pairs and triplets enriched in probands with ID is shown. We tested whether mutated gene pairs identified as significant in one cohort (SPARK) are also associated with severe phenotypes in an independent cohort (SSC). To test this, we obtained the mean IQ score of individuals from the SSC cohort carrying significant combinations identified from the SPARK cohort. Empirical p-values were then calculated based on the deviation of the mean IQ from the distribution of mean IQ scores obtained from 10,000 random draws in the simulation. **(B)** The mean IQ of individuals with mutated gene pairs in the SSC cohort was significantly lower (empirical p-value=0) when compared to the distribution of mean IQ scores obtained from the simulation. **(C)** Histogram shows the distributions of IQ scores of SSC probands who carried mutations in either genes versus both constituent genes of the significant gene pairs. The distributions were significantly different from each other (p-value = 1.302×10^{-6}, Kolmogorov-Smirnov test).



741

742 **Figure 3: Analysis of parental and sibling inheritance patterns of significant gene pairs**
743 **associated with ID. (A)** Fraction of all instances of significant gene pairs observed within each
744 of the six possible parental inheritance patterns (red) compared against 1,000 simulations is
745 shown (blue). During each simulation, random mutated gene pairs from the SSC cohort were
746 selected, the inheritance status of the mutations was identified, and the fraction of those instances

747 belonging to one of the six pre-defined categories was calculated. Comparing the observed
748 fractions with the simulated fractions indicate statistical enrichment for two specific inheritance
749 patterns: both variants being *de novo*, and one variant being *de novo* and the other transmitted
750 from the mother. **(B)** Histograms show the carrier status of significant gene pairs in siblings of
751 carrier probands (red) compared against 1,000 simulations (blue). Among significant pairs, both
752 genes were mutated in only 24.2% of all siblings (compared to 28.4% in simulations), whereas
753 one of the two genes was mutated in 46.6% of all siblings (compared to 38.5% in simulations).
754 These results show that mutations are more likely to be observed in just one of the two genes
755 within the gene pairs and are less likely to be observed simultaneously in siblings of carrier
756 probands.
757
758

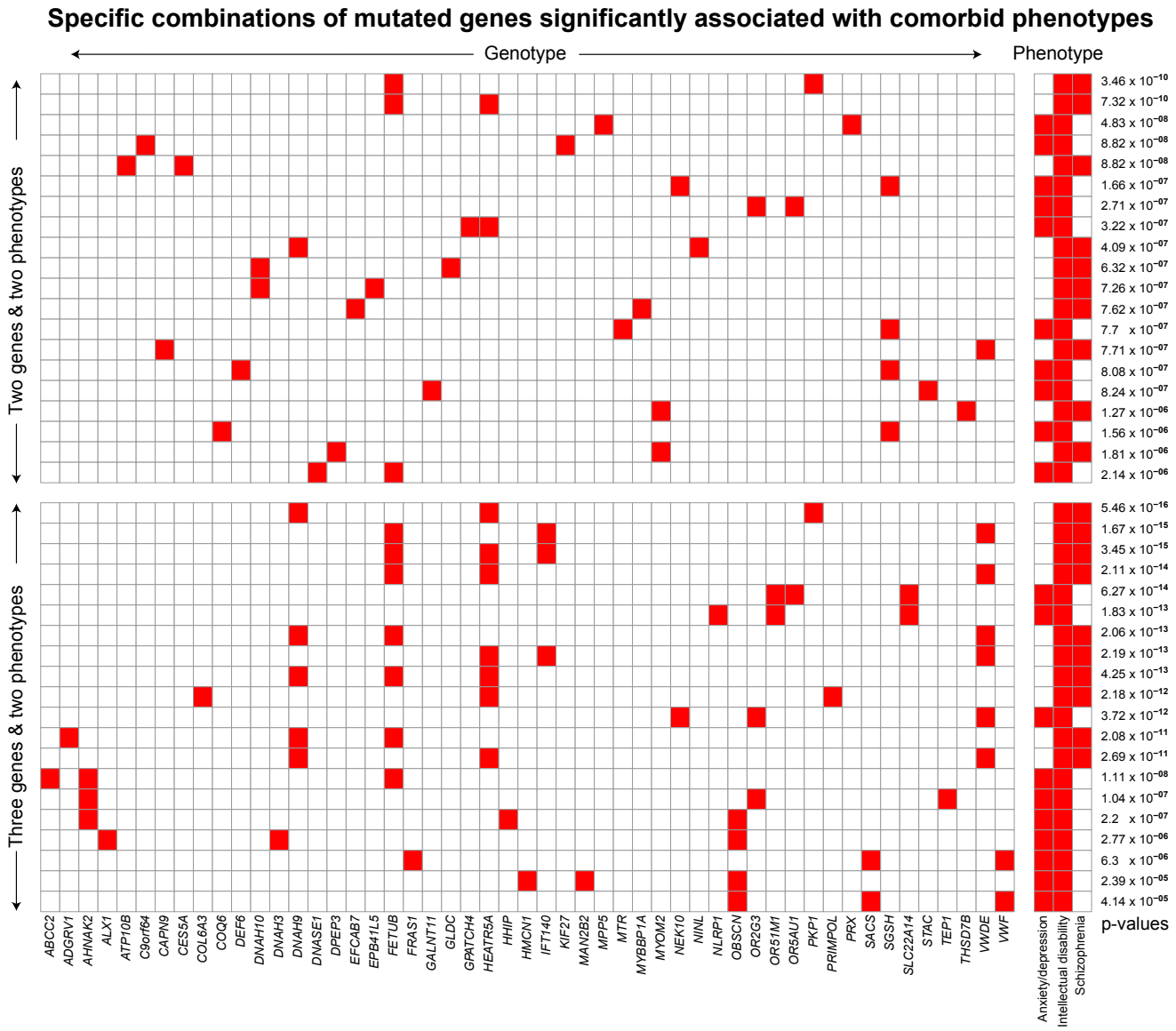


Figure 4: Analysis of comorbid phenotypes using *RareComb*. We analyzed the genotypes of probands with anxiety/depression, ID, or schizophrenia. The heatmap shows combinations of two or three mutated genes that were significantly enriched in individuals with specific patterns of comorbid phenotypes compared to the expected frequency under the assumption of independence.

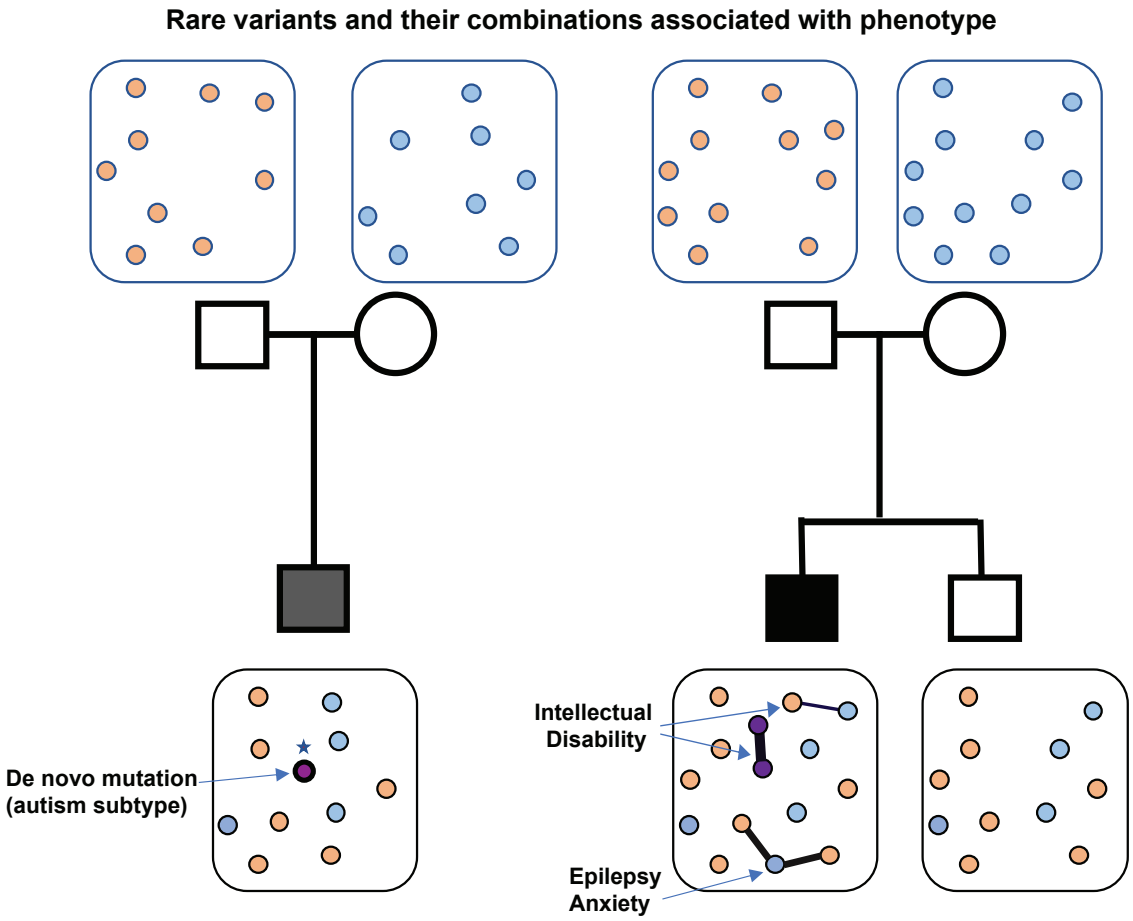


Figure 5: Rare variant models for complex disorders. The schematic shows two models for the genetic etiology of complex disorders. Circles represent rare variants present that are either *de novo* or inherited from a parent. On the left, individual high-effect *de novo* variants are strongly associated with a phenotype of interest. On the right, rare variants within an individual combine in multiple ways and contribute towards distinct phenotypes. The thickness of the connecting lines denotes effect sizes, and an affected individual can carry multiple oligogenic combinations of rare variants, each of which contributes to the same or distinct phenotypes. This extension of the oligogenic model enables further dissection of the genetic architecture of complex disorders.