I Inbreeding is associated with shorter early-life

2 telomere length in a wild passerine

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21 ABSTRACT

22 Inbreeding can have negative effects on survival and reproduction, which may be of 23 conservation concern in small and isolated populations. However, the physiological 24 mechanisms underlying inbreeding depression are not well-known. The length of telomeres, 25 the DNA sequences protecting chromosome ends, has been associated with health or fitness 26 in several species. We investigated effects of inbreeding on early-life telomere length in two 27 small island populations of wild house sparrows (Passer domesticus) known to be affected by 28 inbreeding depression. Using genomic and pedigree-based measures of inbreeding we found 29 that inbred nestling house sparrows have shorter telomeres. This negative effect of inbreeding 30 on telomere length may have been complemented by a heterosis effect resulting in longer 31 telomeres in individuals that were less inbred than the population average. Furthermore, we 32 found some evidence of stronger effects of inbreeding on telomere length in males than 33 females. Thus, telomere length may reveal subtle costs of inbreeding in the wild and 34 demonstrate a route by which inbreeding negatively impacts the physiological state of an 35 organism already at early life-history stages.

36

37 INTRODUCTION

38 Inbreeding may have significant detrimental effects on survival, reproduction, and 39 resistance to disease and other stressors in wild populations (Keller & Waller, 2002). Such 40 decline in fitness resulting from an increase in genome-wide homozygosity is known as 41 inbreeding depression (Charlesworth & Willis, 2009) and is of major concern in small and 42 isolated populations, in particular of endangered species (Bozzuto, Biebach, Muff, Ives, & 43 Keller, 2019; Harrisson et al., 2019; Hedrick & Kalinowski, 2000). Increased homozygosity 44 can lead to reduced fitness due to expression of deleterious recessive alleles ("dominance 45 hypothesis") or increased homozygosity with heterozygote at loci advantage 46 ("overdominance hypothesis", Charlesworth & Willis, 2009). Regardless of the genetic basis
47 for inbreeding depression, it is difficult to identify and quantify the physiological
48 mechanisms underlying the fitness costs of inbreeding (Fox & Reed, 2011; Kristensen,
49 Pedersen, Vermeulen, & Loeschcke, 2010; Losdat, Arcese, Sampson, Villar, & Reid, 2016).

50 Telomeres are short DNA tandem repeats that are found at the tips of most eukaryotic 51 chromosomes (Blackburn & Gall, 1978; Červenák, Sepšiová, Nosek, & Tomáška, 2021). Telomeres shorten during cell division (Harley, Futcher, & Greider, 1990), but may also 52 53 shorten due to several other reasons including physiological processes generating oxidative stress (Barnes, Fouquerel, & Opresko, 2019; Monaghan & Ozanne, 2018; Reichert & Stier, 54 55 2017; von Zglinicki, 2002). The high guanine content of telomeres (50%) makes them 56 particularly vulnerable to oxidative stress (Kawanishi & Oikawa, 2004). Short telomeres can 57 trigger apoptosis and telomere attrition is considered a hallmark of aging (López-Otín, 58 Blasco, Partridge, Serrano, & Kroemer, 2013), although the causal involvement of telomere 59 shortening in organismal senescence is not well understood (Simons, 2015). However, 60 telomere length (TL) may reflect the cumulative stress experienced by an individual 61 (Bateson, 2016; Monaghan, 2014), and TL or TL shortening are associated with health or 62 fitness in several species (Barrett, Burke, Hammers, Komdeur, & Richardson, 2013; 63 Chatelain, Drobniak, & Szulkin, 2020; Froy et al., 2021; Heidinger, Kucera, Kittilson, & 64 Westneat, 2021; Wilbourn et al., 2018). Thus, TL is increasingly used as a biomarker of 65 somatic integrity in studies of physiological or evolutionary ecology (Bateson & Poirier, 66 2019; Haussmann, 2010; Pepper, Bateson, & Nettle, 2018; Young, 2018).

Inbreeding depression can be caused by reduced immune response (Charpentier,
Williams, & Drea, 2008; Reid, Arcese, & Keller, 2003) and higher maintenance metabolism
(Ketola & Kotiaho, 2009), which increases oxidative stress (de Boer et al., 2018a; Okada,
Blount, Sharma, Snook, & Hosken, 2011). Thus, inbred individuals may experience higher

71 levels of oxidative stress (Kristensen, Sørensen, Kruhøffer, Pedersen, & Loeschcke, 2005; 72 Pedersen et al., 2008) and thus have shorter telomeres (von Zglinicki, 2002). We therefore 73 hypothesize that TL could provide an integrative measure of the somatic costs associated 74 with inbreeding depression in wild populations, with inbred individuals having shorter 75 telomeres than outbred individuals. However, the few studies investigating associations 76 between inbreeding and TL have found equivocal results. In line with our expectations, 77 Bebbington et al. (2016) found that homozygosity was negatively associated with TL in wild 78 Seychelles warblers (Acrocephalus sechellensis) and Seluanov et al. (2008) reported that 79 telomeres were shorter in inbred laboratory strains of Norway rats (*Rattus norvegicus*) in 80 captivity compared to a single wild-caught rat. Many domesticated species are generally 81 assumed to be more inbred than their wild counterparts (Bosse, Megens, Derks, de Cara, & 82 Groenen, 2018; Moyers, Morrell, & McKay, 2018; Wiener & Wilkinson, 2011). However, 83 several studies have found that telomeres were longer in inbred domesticated strains of laboratory mice (Mus spp. and Peromyscus spp., Hemann & Greider, 2000; Manning, 84 85 Crossland, Dewey, & Van Zant, 2002; Seluanov et al., 2008), in domesticated strains of pearl millet (Pennisetum glaucum, Sridevi, Uma, Sivaramakrishnan, & Isola, 2002), in 86 87 domesticated inbred chicken (Gallus gallus, O'Hare & Delany, 2009), and across several 88 species of domesticated mammals (Pepke & Eisenberg, 2021) compared to non-domesticated 89 species. However, there were no clear differences in TL between inbred and wild leporid 90 strains (Forsyth, Elder, Shay, & Wright, 2005). Other studies found no association between 91 pedigree-based inbreeding coefficients and TL or telomere attrition in humans (Homo 92 sapiens, Mansour et al., 2011), wild sand lizards (Lacerta agilis, Olsson, Wapstra, & Friesen, 93 2018), or wild natterjack toads (*Epidalea calamita*, Sánchez-Montes et al., 2020). Becker et 94 al. (2015) reported a weak non-significant but positive association between inbreeding and 95 TL in wild white-throated dippers (Cinclus cinclus).

96 These contrasting results suggest that the telomere dynamics of captive, domesticated 97 species living in a controlled environment may not be representative of wild, free-living populations (Chatelain et al., 2020; Pepke & Eisenberg, 2021; Weinstein & Ciszek, 2002). 98 99 For instance, captive populations may be less vulnerable to inbreeding because inbreeding 100 depression is greater under stressful environmental conditions (Fox & Reed, 2011; Reed, 101 Briscoe, & Frankham, 2002). Furthermore, captivity may in itself provide conditions that 102 change the telomere dynamics of the populations (Eisenberg, 2011), e.g. Hemann and Greider 103 (2000) attributed the longer telomeres of inbred mice to effects of captive breeding and not 104 inbreeding *per se*. For instance, TL shortening rates may increase during metabolically costly 105 processes such as reproduction (Sudyka, Arct, Drobniak, Gustafsson, & Cichoń, 2019; Wood 106 et al., 2021) and inbreeding may reduce fecundity (Keller & Waller, 2002). Such effects have 107 been suggested to explain the observation of longer adult TL in some inbred domesticated 108 species (Eisenberg, 2011), which could be resolved by measuring TL in early-life. 109 Furthermore, most of the studies of domesticated animals compared TLs of different 110 populations or species and their results may not be extrapolated to natural variation in TL and inbreeding levels within wild populations. Indeed, TL can vary considerably within species 111 112 (Tricola et al., 2018) and across closely related species (Pepke, Ringsby, & Eisenberg, 2021) 113 in the wild. Finally, it is not known if outbreeding could be accompanied by a heterosis effect 114 (hybrid vigor, e.g. Charlesworth & Willis, 2009) acting on TL. For instance, the observed 115 fitness benefits of outcrossing inbred populations (Frankham, 2015) could be reflected in TL 116 restoration (Nuzhdin & Reiwitch, 2002; Ozawa et al., 2019).

In this study, we utilized a long-term metapopulation study to examine how inbreeding affects early-life TL in wild house sparrows (*Passer domesticus*). Inbreeding has been shown to reduce fitness components such as recruitment probability, adult lifespan, and both annual and lifetime reproductive success in this metapopulation (Billing et al., 2012;

121 Jensen, Bremset, Ringsby, & Sæther, 2007; Niskanen et al., 2020), but the physiological 122 effects underlying these phenomena remain unknown. We expect that inbred individuals will 123 have shorter telomeres if TL is a general biomarker of somatic integrity and health (e.g. 124 Bebbington et al., 2016; Boonekamp, Simons, Hemerik, & Verhulst, 2013; Wilbourn et al., 125 2018). The effects of inbreeding on TL might be sex-specific (Benton et al., 2018; Billing et 126 al., 2012; de Boer et al., 2018a; de Boer, Eens, & Müller, 2018b) or depend on environmental 127 conditions (Armbruster & Reed, 2005; Szulkin & Sheldon, 2007). However, TL is negatively 128 associated with body size or growth rate within many species (Monaghan & Ozanne, 2018; 129 Ringsby et al., 2015) and may change with age (Hall et al., 2004; Remot et al., 2021) or vary 130 between sexes (Barrett & Richardson, 2011; Remot et al., 2020) and habitat quality 131 (Angelier, Vleck, Holberton, & Marra, 2013; McLennan et al., 2021; Wilbourn et al., 2017). 132 We therefore account for body size (measured as tarsus length), age, sex, and habitat type, 133 and test for an interaction between inbreeding levels and sex or habitat type, when 134 investigating the association between TL and inbreeding. We use three different measures of 135 inbreeding; marker-based estimates (n=371) which are better at capturing homozygosity and 136 inbreeding caused by distant ancestors not included in a pedigree, and pedigree-based 137 estimates (Kardos, Taylor, Ellegren, Luikart, & Allendorf, 2016) for which larger samples 138 size may be obtained from long-term field studies (n=1195). Finally, to investigate a potential 139 heterosis effect on TL, we test if the association between TL and inbreeding is different 140 among outbred and inbred individuals.

141

142 MATERIAL AND METHODS

143 *Study system*

144 This study was conducted in two natural populations of house sparrows in northern 145 Norway. On the island of Hestmannøy (66°33'N, 12°50'E), the sparrows live around dairy

146 farms, where they nest inside barns in cavities or nest boxes. The island is characterized by 147 cultivated grassland, mountains, forest, and heathland. On the island of Træna (66°30'N, 148 12°05'E), 34 km further from the mainland, the sparrows live in gardens of a small human 149 settlement and nest in nest boxes. This island is dominated by heathland, sparse forest, and 150 gardens. The natural breeding environment for house sparrows is human habitation (Hanson, 151 Mathews, Hauber, & Martin, 2020) and they have evolved their commensal relationship with 152 humans for millennia (Ravinet et al., 2018). While human presence or farming provide the 153 natural basis of existence for house sparrows (Ringsby, Sæther, Jensen, & Engen, 2006), 154 demographic characteristics, breeding densities, and inbreeding rates are comparable to other 155 small isolated wild animal populations (Araya-Ajoy et al., 2021; Jensen et al., 2007; 156 Niskanen et al., 2020). In the years 1994-2013 (on Hestmannøy) and 2004-2013 (on Træna), 157 nestlings at the age of 5-14 days were ringed with a unique combination of color rings for 158 identification. Nestlings were also blood sampled by brachial venipuncture, and 159 tarsometatarsus (tarsus) was measured with slide calipers to the nearest 0.01 mm. Tarsus 160 length is here used as an index of body size (Rising & Somers, 1989; Senar & Pascual, 1997). Blood samples (25 µL) were stored in 96% ethanol at room temperature in the field and at -161 162 20°C in the laboratory until DNA extraction (described in Pepke et al., submitted 2021b). 163 Birds that were resignted or recaptured in the year following hatching (i.e. from 1995-2014 164 on Hestmannøy and from 2005-2014 on Træna) were categorized as first-year survivors.

165

Telomere length measurements

Relative erythrocyte telomere length (TL) was measured in DNA derived from whole blood samples (n=2746 nestlings) using the qPCR method (Cawthon, 2002) as described in Pepke et al. (*submitted* 2021a). For this study, we included only individuals with two known parents and at least two known grandparents, or for which genomic inbreeding coefficients could be estimated (described below), resulting in a sample size of n=1370 individuals 171 (n=1161 from Hestmannøy and n=209 from Træna). TL was determined relative to the 172 amount of a non-variable gene (GAPDH) and a reference sample (Criscuolo et al., 2009). All 173 samples were randomized and run in triplicates on 96-well plates. All samples were 174 processed within a few months by the same researcher (MLP) to reduce technical effects. 175 Relative TL was computed using qBASE (Hellemans, Mortier, De Paepe, Speleman, & 176 Vandesompele, 2007) while controlling for inter-run variation. All individual plate 177 efficiencies were within $100\pm10\%$ (see Pepke et al., submitted 2021a). Sex was determined 178 by amplification of the CHD-gene as described in Jensen et al. (2007).

179 Microsatellite pedigree construction

180 Microsatellite (MS) pedigrees (n=1857 individuals from Hestmannøy and n=342 from 181 Træna including non-phenotyped ancestors) were constructed based on 13 polymorphic 182 microsatellite markers using CERVUS 3.0 (Kalinowski, Taper, & Marshall, 2007) as 183 described in Billing et al. (2012). Maximum pedigree depth was 13 generations. We 184 calculated inbreeding coefficients (F_{PED}), which estimate the expected proportion of an 185 individual's genome that is identical by descent (IBD), based on the MS pedigree for 186 individuals with two known parents and at least two known grandparents (n=1057 from 187 Hestmannøy and *n*=138 from Træna, Table 1) using the R package *pedigree* (Coster, 2012). 188 We also selected a subset of individuals with at least two full ancestral generations (i.e. four 189 known grandparents) to only include the most robust estimates of F_{PED} (n=313 from 190 Hestmannøy and *n*=7 from Træna).

191 Genomic inbreeding estimation

Starting from year 1997 (Hestmannøy) or 2004 (Træna), birds that survived until recruitment (n=275 from Hestmannøy and n=96 from Træna) were genotyped for 200,000 Single Nucleotide Polymorphisms (SNPs) as described in Lundregan et al. (2018). Two genomic inbreeding coefficients were then estimated using 118,810 autosomal SNPs not in

196 strong linkage disequilibrium, as described in Niskanen et al. (2020). The weighted average 197 homozygosity over all loci from the genomic relationship matrix (F_{GRM}) was estimated for 198 the whole metapopulation simultaneously using the GCTA software (Yang, Lee, Goddard, & 199 Visscher, 2011). F_{GRM} gives more weight to homozygotes of the minor allele than of the 200 major allele, and it is an estimate of the correlation between homologous genes of the two 201 gametes of an individual relative to the current population (Yang et al., 2011). F_{GRM} can be 202 negative if the probability that the two homologous genes of an individual are IBD is smaller 203 than that of two homologous genes being drawn at random from the reference population 204 (Wang, 2014; Yang et al., 2011). Thus, the individuals with the smallest estimates of F_{GRM} 205 are expected to be outbred (hybrids) because of e.g. mating involving immigrants (Wang, 206 2014). The proportion of the genome within runs-of-homozygosity (F_{ROH} ranging from 0 to 207 1, McQuillan et al., 2008) was estimated using the PLINK software (Purcell et al., 2007). 208 ROH arise through mating of individuals that are IBD, and may therefore be used to estimate 209 inbreeding (Curik, Ferenčaković, & Sölkner, 2014).

210 *Statistical analyses*

211 To test whether TL was affected by inbreeding, we fitted linear mixed models 212 (LMMs) using the package *lme4* (Bates, Mächler, Bolker, & Walker, 2015) in R v. 3.6.3 (R 213 Core Team, 2020). TL (response variable) was log_{10} -transformed to conform to the 214 assumption of normally distributed residuals and the models were fitted with a (continuous) 215 fixed effect of one of the inbreeding coefficients (F_{PED} [n=1195], F_{PED} with at least two full 216 generations known [n=320], F_{GRM} [n=371], or F_{ROH} [n=371], see Table 1 for sample size 217 details). Since genomic estimators of inbreeding (F_{GRM} and F_{ROH}) were only available for 218 recruits (first-year survivors), we tested whether the relationship between TL and F_{PED} varied 219 between survivors ("1", n=206) and non-survivors ("0", n=989) by including an interaction 220 effect between F_{PED} and first-year survival. Tarsus length increases with nestling age, so 221 tarsus length was age-corrected by taking the residuals from a regression of tarsus length on 222 age and age squared. This allowed us to include both tarsus length and age in the models 223 describing variation in TL. Thus, age-standardized tarsus length, fledgling age at sampling (in 224 number of days), hatch day (ordinal date mean centered across years), population identity 225 (categorical: Hestmannøy or Træna), and sex (categorical: male or female) were included as 226 fixed effects in all models. We tested whether the effect of inbreeding on TL varied between 227 sexes and populations by including two-way interaction terms between the inbreeding 228 coefficient and sex or population identity. Random intercepts were fitted for year and brood 229 identity to account for the non-independence of nestlings from the same year and brood. This 230 also controls for within-brood effects of inbreeding levels (Olsson et al., 2018). We then 231 tested whether the inclusion of the inbreeding coefficient and interaction terms improved the 232 baseline model (without the inbreeding coefficient) by comparing the resulting 5 candidate 233 models using Akaike's information criterion corrected for small sample sizes (AICc, Akaike, 234 1973; Hurvich & Tsai, 1989). Akaike weights (w) and evidence ratios (ER) were calculated to 235 determine the relative fit of models to the data (Burnham & Anderson, 2002). To investigate 236 heterosis effects on TL, we tested if the slopes of the regression between F_{GRM} and TL 237 differed between individuals that were more inbred than on average (F_{GRM} > mean F_{GRM}) and 238 individuals that were less inbred than average ($F_{GRM} < \text{mean } F_{GRM}$). We did this by testing if 239 the inclusion of a regression break point at the mean F_{GRM} improved the models by 240 comparing the resulting 9 candidate models using AICc. Models were validated visually 241 using diagnostic plots of residuals, and model parameters are from models refitted with 242 restricted maximum likelihood (REML). Estimates are reported with standard errors (SE) and 243 95% confidence intervals (CI). Regression lines were visualized using ggplot2 (Wickham, 244 2016).

246 **RESULTS**

247 Individual MS pedigree-based inbreeding coefficients (F_{PED}) varied from 0.000-0.250 248 (mean 0.007, 16.9% non-zero values). The highest ranked model explaining variation in TL 249 included a negative effect of F_{PED} , but only slightly improved the fit of the baseline model $(\Delta_{2:1}AICc=0.8$ [subscripts denote which ranked models are compared], $w_1=0.36$, 250 251 $ER_1 = w_1/w_2 = 1.49$, Table S1 in the supporting information). Thus, there was a tendency for TL 252 to be shorter in more inbred sparrows ($\beta_{F_{PED}}$ =-0.169±0.101, CI=[-0.366, 0.028], n=1195, 253 Fig. 1a and Table 2). The model ranked third ($\Delta_{3:I}AICc=1.3$) indicated that TL was less 254 associated with F_{PED} in males than in females ($\beta_{F_PED*sex[female]}$ =-0.167±0.196, CI=[-0.549, 255 (0.216]), while the model ranked fourth ($\Delta_4 AICc=1.9$) indicated that TL was less associated 256 Hestmannøy population than in the with F_{PED} in the Træna population 257 $(\beta_{F_PED*island[Hestmannøv]}=0.115\pm0.314, CI=[-0.498, 0.728])$. However, due to high uncertainty 258 in these parameter estimates, these effects are not deemed reliable.

When only including individuals with at least 2 full ancestral generations known (33.8% non-zero values), the model with F_{PED} was ranked second ($\Delta_{2:1}AICc=1.1$, $\beta_{F_PED}=-$ 0.205±0.198, CI=[-0.588, 0.189], *n*=320, Fig. 1b, Table S2-3) and the baseline model was highest ranked.

There was a tendency for the negative effect of F_{PED} on TL to be weaker in first-year survivors (*n*=206, mean TL=0.95±0.02, mean F_{PED} =0.010±0.003) than in non-survivors (*n*=989, mean TL=0.97±0.01, mean F_{PED} =0.007±0.001, $\beta_{F_PED*first-year survival}$ =0.304±0.201, CI=[-0.089, 0.697], *n*=1195, Fig. 1c, Table S4). This effect was uncertain with a CI overlapping zero. This suggests that the following analyses using genomic estimators of inbreeding in recruits were not biased towards stronger inbreeding effects in recruits.

Genomic inbreeding coefficient (F_{GRM}) estimates varied from -0.200 to 0.300 (mean 0.016). The highest ranked model ($\Delta_{2:1}AICc=2.1$, Table S5) showed that TL was shorter in

271 more inbred sparrows (β_{F_GRM} =-1.517±0.293, CI=[-2.150, -0.920], *n*=371, Fig. 1d, and Table 272 3). In addition, the effect of F_{GRM} on TL was stronger in the Træna population 273 ($\beta_{F_GRM*island[Hestmannøy]}$ =0.824±0.339, CI=[0.142, 1.529], Table 3) and in males 274 ($\beta_{F_GRM*sex[female]}$ =0.644±0.314, CI=[0.034, 1.262], Table 3).

275 Including a break point at the mean F_{GRM} improved the model compared to a model 276 with no break point (comparing models without interaction terms which were ranked 8 and 5: 277 $\Delta_{8:5}AICc=4.5$, see Table S6). The highest ranked model ($\Delta_{2:1}AICc=3.1$, Table S6) revealed a 278 strong negative association between TL and F_{GRM} among individuals with F_{GRM} <0.016 but no 279 significant association among inbred individuals with F_{GRM} >0.016 (Fig. 1e and Table 4). This 280 indicates that a heterosis effect resulting in longer telomeres in outbred individuals may 281 explain the negative association found between inbreeding and TL. This model also included 282 an interaction term suggesting that this heterosis effect was stronger in the Træna population 283 (Table 4).

The runs-of-homozygosity inbreeding coefficient (F_{ROH}) estimates varied from 0.000-0.240 (mean 0.010, 73% non-zero values). The best model provided evidence for a negative effect of F_{ROH} on TL (β_{F_ROH} =-1.148±0.512, CI=[-2.144, -0.153], n=371, Fig. 1f, Table S7 and 5). This model also indicated that the negative effect of F_{ROH} tended to be stronger in males (β_{F_ROH} *sex [female]=0.915±0.610, CI=[-0.270, 2.102]).

Overall, F_{PED} was not a good predictor of genomic estimators of inbreeding (Fig. S1a,c; Pearson's r_P =0.05, n=371), but its relationships with F_{GRM} and F_{ROH} were improved when including only individuals with at least two generations known (Fig. S1b,d; r_P >0.30, n=59). F_{GRM} and F_{ROH} were strongly correlated (Fig. S1e,f; r_P =0.7, n=371).

293

294 **DISCUSSION**

295 We found evidence across multiple complementary measures of inbreeding that more 296 inbred house sparrow nestlings had shorter telomeres (Fig. 1). Individual differences in TL 297 are established early in life (Entringer, de Punder, Buss, & Wadhwa, 2018), are heritable 298 (Dugdale & Richardson, 2018; Pepke et al., *submitted* 2021a), and are positively associated 299 with fitness in some species (Heidinger et al., 2012; Wilbourn et al., 2018). Thus, short 300 telomeres in more inbred individuals may therefore underpin a physiological basis of 301 inbreeding depression in fitness components that has been found in this species (Billing et al., 302 2012; Jensen et al., 2007; Niskanen et al., 2020) and in other wild animal populations (Keller 303 & Waller, 2002).

304 The effect of inbreeding on TL in house sparrows was negative across all measures of 305 inbreeding, but strongest when using genomic levels of inbreeding (Fig. 1d-f), probably 306 because they are better at capturing homozygosity causing inbreeding depression compared 307 to using a pedigree-based estimator (Fig. 1a-c, Alemu et al., 2021; Huisman, Kruuk, Ellis, 308 Clutton-Brock, & Pemberton, 2016; Kardos et al., 2016). Mating between full siblings or 309 between parent and offspring (F=0.25) resulted in a severe reduction in (relative) TL of 58% 310 (F_{GRM}) , 48% (F_{ROH}) or 11% (F_{PED}) compared to breeding between unrelated individuals 311 (Tables 2, 3, and 5). TL may be under strong selection in natural populations (Voillemot et 312 al., 2012). Consequently, strong inbreeding depression is expected for fitness components or 313 traits that are under strong selection (Bérénos, Ellis, Pilkington, & Pemberton, 2016; DeRose 314 & Roff, 1999), The analyses using genomic estimators of inbreeding were limited to recruited 315 individuals, but the negative effect of inbreeding on TL may be even stronger if very inbred 316 individuals, presumably with short telomeres, do not survive their first year and were thus 317 excluded from our analyses (Jensen et al., 2007; Wilbourn et al., 2018). There was a tendency 318 for such an effect when using pedigree-based levels of inbreeding (Fig. 1c and Table S4). We 319 also found some evidence that inbreeding had stronger negative effects on TL in males than

females (Tables 3 and 5). Such sex-specific effects of inbreeding are known from other species (de Boer et al., 2018a; de Boer et al., 2018b; Janicke, Vellnow, Sarda, & David, 2013), but have rarely been observed early in life. There was a weak tendency for longer TL in males than females (Tables 2-5), which has been observed in similar house sparrow populations (Pepke et al., *submitted* 2021b). Thus, males may be better buffered against the effects of inbreeding on TL. However, no sex-specific differences in inbreeding depression were observed in adult sparrows across this study metapopulation (Niskanen et al., 2020).

327 Increased inbreeding may be accompanied by population decline in small populations 328 (Bozzuto et al., 2019; Chen, Cosgrove, Bowman, Fitzpatrick, & Clark, 2016; Feng et al., 329 2019), which can drive populations to extinction (O'Grady et al., 2006; Saccheri et al., 1998; 330 Wright, Tregenza, & Hosken, 2007). Niskanen et al. (2020) showed that inbreeding 331 depression in adult sparrows in our study system varied little across years or across the 332 different island environments inhabited by these house sparrows. Hence, the strength of 333 inbreeding depression is similar between populations, but due to harboring more inbred 334 individuals, the relative effect is stronger in smaller populations (Niskanen et al., 2020). Small declining populations may be characterized by gradual population-wide and trans-335 336 generational telomere erosion. For instance, Dupoué et al. (2017) observed shorter TL along 337 an extinction risk gradient in populations of common lizards (Zootoca vivipara) that are 338 disappearing from low altitudes at their southern range limit, presumably due to climate 339 warming (Sinervo et al., 2010). Combined, these results suggest that TL may represent a 340 potential physiological biomarker or molecular tool in conservation genetics addressing the 341 viability of some small animal populations (Bebbington et al., 2016; Bergman et al., 2019; 342 Dupoué et al., 2017; Madliger, Franklin, Love, & Cooke, 2020).

The negative effect of F_{GRM} on TL (Fig. 1d) was stronger among individuals that were less related than the average population (Fig. 1e). This suggests that longer telomeres in 345 outbred individuals may partly be attributed to a general heterosis effect (Charlesworth & 346 Willis, 2009) involving mating between immigrants and native individuals (Dickel et al., 347 2021; Ebert et al., 2002). In our study metapopulation, the proportion of dispersers among 348 recruits can be high among the island populations (0.2 on average ranging from 0.0-1.0 349 across years and islands, Ranke et al., 2021; Saatoglu et al., 2021), and hence most islands are 350 not strongly differentiated (Niskanen et al., 2020). We found that the negative effect of F_{GRM} 351 on TL was stronger in the Træna population (Table 3-4). Træna is known to have a higher 352 proportion of immigrants than Hestmannøy (Ranke et al., 2021), which may contribute to a 353 stronger effect of heterosis in this population (Table 4). Furthermore, the gardens of Træna 354 expose the sparrows to a different environment than the farms on Hestmannøy (Araya-Ajoy 355 et al., 2019; Pärn, Ringsby, Jensen, & Sæther, 2012). Inbreeding depression is expected to 356 have more severe consequences under environmental stress (Armbruster & Reed, 2005; Reed 357 et al., 2002), such as harsh weather or competition (de Boer et al., 2018a; Fox & Reed, 2011; 358 Marr, Arcese, Hochachka, Reid, & Keller, 2006). Telomeres shorten due to environmental 359 stressors such as harsh abiotic conditions (Chatelain et al., 2020). We speculate that 360 environmental differences between the habitats of the two sparrow populations may explain 361 the exacerbated effects of inbreeding on TL in the Træna population. For instance, in juvenile 362 Seychelles warblers a negative relationship between homozygosity and TL was found only in 363 poor seasons, i.e. when food availability was low (Bebbington et al., 2016). In adult 364 Seychelles warblers, the effect of homozygosity on TL was consistently negative across 365 seasons, suggesting that the negative effects of inbreeding accumulate through life and are 366 reflected in telomere erosion (Bebbington et al., 2016). Here, we showed that inbreeding 367 manifests in TL already at the nestling stage in a similar wild passerine.

We measured TL in blood, thus it is possible that inbreeding or heterosis only affected telomeres in erythrocytes (Manning et al., 2002; Olsson, Geraghty, Wapstra, & Wilson,

370 2020). However, this is unlikely because TLs often correlate well across tissues within the 371 organism (Daniali et al., 2013; Demanelis et al., 2020; Reichert, Criscuolo, Verinaud, Zahn, 372 & Massemin, 2013), especially in early-life (Prowse & Greider, 1995). Although genomic 373 inbreeding estimates were only available for first-year survivors, we may have avoided 374 confounding effects of selective mortality of inbred individuals at much older ages by 375 measuring TL already at the nestling stage (Hemmings, Slate, & Birkhead, 2012; Sánchez-376 Montes et al., 2020). Furthermore, since the mutation accumulation theory of senescence 377 predicts that deleterious effects of inbreeding increase with age (Charlesworth & Hughes, 378 1996; Keller, Reid, & Arcese, 2008), we may expect that the effect on TL is persistent and 379 potentially stronger in adult sparrows. Thus, future studies are required to investigate if 380 inbreeding leads to persistently eroded TL throughout life, and if there are combined fitness 381 consequences of any interaction between TL and inbreeding in wild populations. Even in the 382 absence of a mechanism directly linking inbreeding and TL via the effects of oxidative stress 383 (cf. the introduction), we may find inbred individuals to have short telomeres, because 384 inbreeding impairs other physiological processes that affects both fitness and TL (Bebbington 385 et al., 2016). Thus, the conflicting evidence in the literature of an effect of inbreeding on TL 386 (reviewed in the introduction) suggests that an experimental procedure is needed to further 387 elucidate the mechanisms underlying the correlation reported here (Manning et al., 2002), 388 especially in wild populations.

In conclusion, the negative associations between inbreeding levels and TL found in this study suggest that TL may reveal subtle somatic costs of inbreeding in wild populations, and thereby demonstrates a potential route by which inbreeding negatively impacts the physiological state of an organism in early life. The observation of a potential heterosis effect on TL suggests that maintenance of dispersal within this metapopulation is important for mitigating the negative effects of inbreeding.

395

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- 400 Availability of data and material: Data will be available on Dryad or another open data401 repository.
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- 403 Authors' contributions: MLP measured telomeres, analyzed data, and wrote the manuscript
- 404 with contributions from all authors. WB supervised telomere measurements. HJ, AKN, and
- 405 TK contributed to the genotype data processing, pedigree construction, and in designing
- 406 statistical analyses. THR, BE-S, and HJ initiated the study system. THR, HJ, and TK
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- 408 Ethics approval: Fieldwork was carried out in accordance with permits from the Ringing
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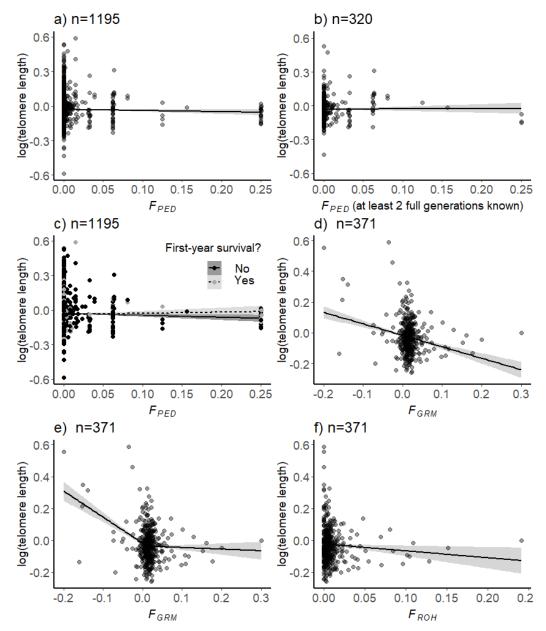
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865 **TABLES AND FIGURES**

866 Table 1: Number of nestling house sparrows of each sex and in total with early-life telomere 867 length and inbreeding coefficient measurements within each population (Hestmannøy and 868 Træna) for each measure of inbreeding (microsatellite pedigree-based inbreeding coefficient 869 $[F_{PED}]$, genomic inbreeding coefficient $[F_{GRM}]$, and runs-of-homozygosity $[F_{ROH}]$). Number of individuals with at least two known full ancestral generations (gen.) are shown. Number of 870 871 individuals with F_{GRM} values above and below the mean F_{GRM} , which is used as a break point 872 to differentiate individuals that were more and less inbred than average, respectively, are also 873 shown.

Island population:	Hestmannøy						
	Males	Females	Sum:	Males	Females	Sum:	Sum:
<i>F</i> _{PED} (≥1.5 gen.)	511	546	1057	78	60	138	1195
F_{PED} (≥ 2 full gen.)	148	165	313	4	3	7	320
F _{GRM}	140	135	275	49	47	96	371
$F_{GRM} > 0.016$	43	63	106	26	32	58	164
$F_{GRM} < 0.016$	97	72	169	23	15	38	207
F _{ROH}	140	135	275	49	47	96	371



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876 Fig. 1: Associations between early-life telomere length (log₁₀-transformed) and various 877 individual measures of inbreeding in wild house sparrows: a) microsatellite pedigree-based 878 inbreeding coefficient (F_{PED}), b) F_{PED} for individuals with at least two full ancestral generations known, c) testing for an interaction effect between F_{PED} and first-year survival 879 880 (survivors: n=206 in grey, dotted regression line; non-survivors: n=989 in black, solid 881 regression line), d) genomic inbreeding coefficient F_{GRM} , e) regression with a break point at the mean F_{GRM} (0.016), and f) runs-of-homozygosity F_{ROH} . Black lines show the predicted 882 effect of the inbreeding coefficient on TL from LMMs described in the text and the grey area 883 884 shows 95% confidence intervals. Note that the y-axis is not scaled equally across panels and 885 that color of points are graduated for visibility.

Table 2: Estimates, standard errors (SE), lower and upper 95% confidence intervals (CI) from the highest ranked model of F_{PED} predicting variation in early-life TL (*n*=1195, see Table S2 and Fig. 1a). The model included random intercepts for brood identity (ID) and year. Estimates with CIs not overlapping 0 are shown in italics.

891					
091	Response variable: log ₁₀ (TL)	Estimate	SE	Lower CI	Upper CI
892	intercept	-3.1E-4	0.037	-0.072	0.071
002	inbreeding coefficient (F_{PED})	-0.169	0.101	-0.366	0.028
893	tarsus length	-0.003	0.002	-0.008	0.001
894	sex [female]	-0.006	0.006	-0.017	0.005
	island identity [Hestmannøy]	0.025	0.012	0.001	0.049
895	age	-0.003	0.002	-0.007	0.001
896	hatch day	-1.4E-4	1.5E-4	-4.4E-4	1.7E-4
000	$\sigma^2_{\text{brood ID}}$ (<i>n</i> =500)	0.002		0.001	0.003
897	$\sigma_{\text{year}}^2(n=20)$	0.003		0.001	0.006
898	Marginal \mathbb{R}^2 / Conditional \mathbb{R}^2 : 0	.014 / 0.39	3		

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Table 3: Estimates, standard errors (SE), lower and upper 95% confidence intervals (CI) from the highest ranked model of F_{GRM} predicting variation in early-life TL (*n*=371, see Table S6 and Fig. 1d).

Response variable: log ₁₀ (TL)	Estimate	SE	Lower CI	Upper CI
intercept	0.069	0.038	-0.004	0.145
inbreeding coefficient (F_{GRM})	-1.517	0.293	-2.150	-0.920
tarsus length	-0.001	0.005	-0.011	0.009
sex [female]	-0.016	0.011	-0.039	0.006
island identity [Hestmannøy]	-0.036	0.016	-0.068	-0.004
age	-0.006	0.003	-0.012	4.6E-4
hatch day	-3.3E-4	2.8E-4	-0.001	2.1E-4
F _{GRM} * island [Hestmannøy]	0.824	0.339	0.142	1.529
F _{GRM} * sex [female]	0.644	0.314	0.034	1.262
$\sigma^2_{\text{brood ID}}$ (n=273)	0.004		0.002	0.006
$\sigma_{\text{year}}^2(n=17)$	0.001		0.000	0.003

- 911 Table 4: Estimates, standard errors (SE), lower and upper 95% confidence intervals (CI)
- from the highest ranked model from Table S7 including a break point at F_{GRM} =0.016 (n=371,

913 see also Table S1). These effects of F_{GRM} are shown in Fig. 1e.

Estimate	SE	Lower CI	Upper CI
0.021	0.037	-0.051	0.095
-2.177	0.372	-3.051	-1.379
0.189	0.498	-0.780	1.153
-0.001	0.005	-0.011	0.008
-0.006	0.010	-0.027	0.014
-0.009	0.016	-0.041	0.024
-0.005	0.003	-0.011	0.001
-3.7E-4	2.7E-4	-0.001	1.5E-4
1.562	0.465	0.610	2.576
-0.026	0.561	-1.114	1.061
0.003		0.001	0.005
0.001		0.000	0.003
	0.021 -2.177 0.189 -0.001 -0.006 -0.009 -0.005 -3.7E-4 1.562 -0.026 0.003	0.021 0.037 -2.177 0.372 0.189 0.498 -0.001 0.005 -0.006 0.010 -0.009 0.016 -0.005 0.003 -3.7E-4 2.7E-4 1.562 0.465 -0.026 0.561 0.003	$\begin{array}{cccccccccccccccccccccccccccccccccccc$

Marginal R^2 / Conditional R^2 : 0.106 / 0.458

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916 Table 5: Estimates, standard errors (SE), lower and upper 95% confidence intervals (CI)

917	from the highest ranked model from of F_{ROH} predicting variation in early-life TL ($n=371$, see
918	Table S8 and Fig. 1f).

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Response variable: log ₁₀ (TL)	Estimate	SE	Lower CI	Upper CI
intercept	0.051	0.040	-0.027	0.130
inbreeding coefficient (F_{ROH})	-1.148	0.512	-2.144	-0.153
tarsus length	-0.001	0.005	-0.011	0.010
sex [female]	-0.018	0.012	-0.041	0.005
island identity [Hestmannøy]	-0.020	0.016	-0.052	0.012
age	-0.005	0.003	-0.012	0.001
hatch day	-2.9E-4	3.0E-4	-0.001	2.9E-4
F_{ROH} * sex [female]	0.915	0.610	-0.270	2.102
$\sigma^{2}_{\text{brood ID}}$ (<i>n</i> =273)	0.006		0.004	0.008
$\sigma_{\text{year}}^2(n=17)$	0.002		4.6E-4	0.004
Marginal \mathbf{R}^2 / Conditional \mathbf{R}^2 .	0.029 / 0.5	79		

Marginal R^2 / Conditional R^2 : 0.029 / 0.579