

1 **Life history adaptations to fluctuating environments: Combined effects of**
2 **demographic buffering and lability of demographic parameters**

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10 **Abstract**

11 Demographic buffering and lability have been identified as adaptive strategies to optimise fit-
12 ness in a fluctuating environment. These are not mutually exclusive, however we lack efficient
13 methods to measure their relative importance for a given life history. Here, we decompose the
14 stochastic growth rate (fitness) into components arising from nonlinear responses and variance-
15 covariance of demographic parameters to an environmental driver, which allows studying joint
16 effects of buffering and lability. We apply this decomposition for 154 animal matrix population
17 models under different scenarios, to explore how these main fitness components vary across
18 life histories. Faster-living species appear more responsive to environmental fluctuations, ei-
19 ther positively or negatively. They have the highest potential for strong adaptive demographic
20 lability, while demographic buffering is a main strategy in slow-living species. Our decomposi-
21 tion provides a comprehensive framework to study how organisms adapt to variability through
22 buffering and lability, and to predict species responses to climate change.

23

24 **Keywords:** comparative study, convexity, demographic buffering, demographic lability, envi-
25 ronmental variance, climate change, matrix population model, nonlinearity, temporal covari-
26 ance, stochasticity

27 Introduction

28 Understanding life history adaptations to fluctuating environments is increasingly important,
29 as anthropogenic climate change is altering the temporal variability of multiple climatic drivers
30 (IPCC, 2021; Laufkötter *et al.*, 2020; Pendergrass *et al.*, 2017). For instance, while an increased
31 variance in daily and seasonal temperature and precipitation is expected across much of Europe
32 in summer, a decrease is projected in other regions (Huntingford *et al.*, 2013; IPCC, 2021; Kotz
33 *et al.*, 2021; Pendergrass *et al.*, 2017). Fluctuations in abiotic and biotic environmental drivers
34 experienced by organisms may affect their relative fitness and select for specific adaptations to
35 live in variable environments.

36 Two main processes have been identified as adaptations to environmental variability, opti-
37 mizing fitness: Demographic buffering reduces the variance in demographic parameters (e.g.,
38 survival, fertility), thereby minimizing the effects of bad environments (Morris & Doak, 2004;
39 Hilde *et al.*, 2020), while demographic lability lets the organisms take advantage of good envi-
40 ronments by mounting a large increase in some demographic parameters compared to an average
41 or bad environment, and therefore increasing their mean (Koons *et al.* 2009; Jongejans *et al.*
42 2010; Barraquand & Yoccoz 2013; see Box 1 for Glossary). The two processes are not mutually
43 exclusive but can be selected simultaneously, so that different demographic parameters of a
44 given life history can show different responses to an environmental driver. Yet, these processes
45 have often been investigated separately, and we lack efficient methods to disentangle and pre-
46 dict their relative importance for a given life history and environment. To understand how
47 organisms combine lability and buffering of their demographic parameters to enhance fitness
48 in varying environments, we need a demographic model framework to predict two main fitness
49 components: i) the effects of nonlinearity in responses of all demographic parameters to an
50 environmental driver, and ii) the effects of variance-covariance of these parameters. While the
51 latter is well described in stochastic demographic theory (Lande *et al.*, 2003), we know much

52 less about the impacts of nonlinearity, representing the potential for adaptation to varying
53 environments through lability.

54 A key prediction from classical theory for evolutionary bet-hedging and stochastic popu-
55 lation growth is that the long-term fitness will be reduced if the temporal variance of fitness
56 is increased (Lewontin & Cohen, 1969). This result is assuming an unstructured population
57 with annual population growth rates that are IID (independently and identically distributed).
58 The fitness is then the logarithm of the geometric mean of these growth rates (Lewontin &
59 Cohen, 1969). In structured populations, the stochastic growth process is more complex due
60 to fluctuations in the (st)age structure that introduce autocorrelation in the annual growth
61 rates (Caswell, 2001). Still, under the assumption of small fluctuations in the demographic
62 parameters, Tuljapurkar (1982) derived an important approximation of long-term growth rate
63 in stage-structured populations, emphasising how the variance in fitness is linked to variances
64 and covariances of demographic parameters in different stages (equation 1). The key conclusion
65 from this approximation is that temporal variability in demographic parameters and/or positive
66 covariance will have a negative effect on fitness, and should be selected against, in particular
67 for demographic parameters that have a large impact on fitness in the mean environment. Ac-
68 cordingly, the demographic buffering hypothesis predicts that natural selection should favour
69 a reduction in variance of the demographic parameters with the strongest influence on popu-
70 lation growth (Boyce *et al.*, 2006; Gaillard & Yoccoz, 2003; Hilde *et al.*, 2020; Pfister, 1998;
71 Tuljapurkar & Orzack, 1980).

72 However, positive effects of environmental variability have also been demonstrated under
73 strong negative covariance among demographic parameters (Colchero *et al.*, 2019; Doak *et al.*,
74 2005; Tuljapurkar, 1990), negative environmental autocorrelation (Tuljapurkar, 1982), and con-
75 vex relationships between demographic parameters and the environment. The latter represents
76 a case of adaptive lability as described by Koons *et al.* (2009). In contrast to adaptive demo-
77 graphic buffering, which optimizes fitness by reducing the variance of most influential demo-

78 graphic parameters, lability can be adaptive if the benefit of an increase in the arithmetic mean
79 of the annual growth rates through increased demographic parameter means can overcome the
80 negative effect of increased demographic variance on fitness (Box 1). Nonlinearity in population
81 and demographic parameter responses to environmental drivers may be common in the wild
82 (Barraquand & Yoccoz, 2013; Clark & Luis, 2020; Dahlgren *et al.*, 2011; Drake, 2005; Hansen
83 *et al.*, 2021; Jenouvrier *et al.*, 2012; Lawson *et al.*, 2015; Louthan & Morris, 2021; Mysterud
84 *et al.*, 2001), highlighting the potential importance of lability as an adaptation to environmen-
85 tal variability. However, with structured life histories the combined effects of nonlinearity in
86 different demographic parameters on fitness are challenging to predict (Koons *et al.*, 2009).

87 Somewhat contrasting predictions have been made as to which demographic parameters
88 should be labile or buffered, and the relative importance of each process for a given life history.
89 Demographic lability has been suggested to affect mainly the demographic parameters with
90 least effect on fitness (Hilde *et al.*, 2020), as a consequence of selection for buffering of more
91 influential demographic parameters. Other studies suggest that lability can be equally impor-
92 tant to demographic buffering, and that it can also occur in highly influential demographic
93 parameters (Jongejans *et al.*, 2010; Koons *et al.*, 2009; McDonald *et al.*, 2017). Based on the
94 latter prediction, recent research suggests that adaptive lability and buffering can be located
95 at the opposite ends of a continuum, encompassing a wide range of demographic strategies
96 (Salguero-Gómez, 2021; Santos *et al.*, 2021). Yet, the extent to which lability among the least
97 or the most influential demographic parameters can be adaptive strategies for coping with vary-
98 ing environments, relative to buffering, remains largely unexplored (e.g., Barraquand & Yoccoz
99 2013).

100 We thus need a more thorough understanding of how the opportunity for selection on de-
101 mographic buffering and lability depends on major axes of life history variation such as the
102 slow-fast continuum (Stearns, 1992; Gaillard *et al.*, 2016; Salguero-Gómez *et al.*, 2016b). For
103 instance, populations of fast-living species have been predicted to be more responsive to envi-

104 ronmental variability than those of slow-living species, and to be more likely to show adaptive
105 lability in their demographic parameters (Dalglish *et al.*, 2010; Iles *et al.*, 2019; Koons *et al.*,
106 2009; Morris *et al.*, 2008). According to demographic buffering hypothesis, species towards the
107 slow end of the continuum benefit most from reduced variance in annual survival of the ma-
108 ture stages, while fast-living species gain relatively more from reduction of variance in annual
109 fertility and/or survival of the immature stages (Hilde *et al.*, 2020; Gaillard & Yoccoz, 2003;
110 Rotella *et al.*, 2012). These effects can be predicted from Tuljapurkar’s small noise approxi-
111 mation (Tuljapurkar, 1982; equation 1), but we lack a similar expression to describe the net
112 impact of nonlinearity in different demographic parameters of the same life history. Here, we
113 introduce a new ‘nonlinearity index’ to predict changes in the arithmetic mean arising from non-
114 linearity in different demographic parameter responses to an explicit environmental driver. We
115 decompose the stochastic growth rate into contributions from nonlinearity effects and variance-
116 covariance effects. We then apply the decomposition to study how organisms may combine
117 adaptive buffering and lability responses depending on generation time, which closely correlates
118 with the species’ position along the slow-fast continuum (Gaillard *et al.*, 2005). We use pop-
119 ulation models from the COMADRE animal matrix database (Salguero-Gómez *et al.*, 2016a)
120 as a starting point for our calculations, representing a broad range of life histories in the mean
121 environment. We then add stochastic environmental variation and perform the decomposition
122 under different scenarios for nonlinearity and covariance among demographic parameters (Fig.
123 1). Our study provides a method to disentangle the effects of buffering and lability for any
124 given life history, and the subsequent analysis addresses two main questions: First, what is the
125 opportunity for positive effects due to adaptive lability to overcome negative impacts through
126 the variance-covariance of demographic parameters, and how does this pattern depend on gen-
127 eration time? Second, are demographic parameters that show adaptive lability typically the
128 least or most influential demographic parameters for fitness?

129 Material and methods

130 To explore fitness responses to environmental variability along the slow-fast continuum, we de-
131 composed the long-term stochastic growth rate $\ln(\lambda_s)$, a measure of fitness (Tuljapurkar, 1990;
132 Caswell, 2001; Lande *et al.*, 2003), into main components capturing effects of nonlinearity in
133 demographic parameters as a function of an environmental driver z , and effects of variance-
134 covariance among the parameters. Our approach builds on Tuljapurkar's approximation which
135 assumes linear relationships between demographic parameters and an IID environmental vari-
136 able: (Tuljapurkar, 1990):

$$\ln(\lambda_s) \approx \ln\lambda_0 - \sigma^2, \text{ with } \sigma^2 = \frac{Var(\lambda_t)}{2\lambda_0^2}. \quad (1)$$

137 Here λ_0 is the arithmetic growth rate in the mean environment, which is assumed equal to the
138 mean arithmetic growth rate $\bar{\lambda}$ (ignoring non-linear responses), while $Var(\lambda_t)$ is the variance
139 in annual population growth caused by temporal variance and covariance in the demographic
140 parameters. We show in the next section that including nonlinear effects of the environment on
141 demographic parameters mainly affects $\ln(\lambda_s)$ through the mean arithmetic growth rate $\bar{\lambda}$, but
142 also through the variance-covariance term. After defining main components of the stochastic
143 growth rate, we perform a theoretical exploration of how the different components will vary
144 across generation time, using different scenarios regarding nonlinear functions for survival and
145 fertility (Fig. 1A-C). We also confront hypotheses about demographic lability, through scenarios
146 that specifically consider effects of nonlinearity in the demographic parameters of immature or
147 mature individuals only, keeping other parameters constant ('forced buffering' scenarios, Fig.
148 1C). All simulations and calculations were performed in R, version 4.0.3 (R Core Team, 2020).
149 R code is provided in Supporting information S7.

150 **Decomposing the stochastic growth rate with nonlinear effects**

151 We assume that the environment at each time step is described by a stochastic variable z
 152 (IID), with mean 0 and variance σ_z^2 . Population growth from one time step to the next is
 153 given by $\mathbf{n}_{t+1} = \mathbf{A}(z_t)\mathbf{n}_t$, where \mathbf{n}_t is the vector containing the number of individuals in each
 154 stage at time t , and $\mathbf{A}(z)$ is the population projection matrix. Elements of $\mathbf{A}(z)$ are the
 155 demographic parameters describing survival, fertility and transitions as functions of z . To
 156 derive the stochastic growth rate, we approximate this projection matrix using $\mathbf{A}(z) \approx \mathbf{A}(0) +$
 157 $\frac{\sigma_z^2}{2}\mathbf{A}''(0) + \varepsilon$, where ε is the matrix describing the noise terms with mean elements 0, $\mathbf{A}(0)$ is
 158 the projection matrix of the mean environment (with asymptotic growth rate λ_0) and $\mathbf{A}''(0) =$
 159 $\mathbf{A}''(z)|_{z=0}$ contains the second derivatives of elements of $\mathbf{A}(z)$. Using this second derivative
 160 matrix, we define a nonlinearity index (Supporting information S3)

$$D = \mathbf{v}\mathbf{A}''(0)\mathbf{u} = \sum_{ij} \frac{d\lambda_0}{dA_{ij}(0)} A''_{ij}(0), \quad (2)$$

161 which measures the overall degree of nonlinearity in the life history defined by $\mathbf{A}(z)$. A positive
 162 D indicates adaptive lability. A matrix element (i.e., demographic parameter) with strong
 163 convex curvature may still have a low impact on D if the corresponding sensitivity of λ_0 to that
 164 element is low, and vice versa.

165 Applying a Taylor approximation to the mean change of the logarithm of the total repro-
 166 ductive value $V_t = \sum_j n_{j,t}v_j$ (where reproductive values \mathbf{v} are calculated for the matrix $\mathbf{A}(0)$),
 167 we show in Supporting information S3 that the stochastic growth rate is given by

$$\ln(\lambda_s) = \mathbb{E}[\ln V_{t+1} - \ln V_t | V_t] \approx \underbrace{\ln \lambda_0 + \frac{\sigma_z^2}{2\lambda_0} D \left(1 - \frac{\sigma_z^2}{4\lambda_0} D\right)}_{\ln \bar{\lambda} - \ln \lambda_0} - \underbrace{\frac{\sigma_z^2}{2\lambda_0^2} \left(B + \frac{\sigma_z^2}{2} C\right)}_{\sigma^2} \quad (3)$$

168 where D is the nonlinearity index defined above, $B = \sum_{ij} \sum_{kl} \frac{d\lambda_0}{dA_{ij}(0)} \frac{d\lambda_0}{dA_{kl}(0)} A'_{ij}(0) A'_{kl}(0)$ (where

169 $\mathbf{A}'(0) = \mathbf{A}'(z)|_{z=0}$ is the matrix of first derivatives), and

170 $C = \sum_{ij} \sum_{kl} \frac{d\lambda_0}{dA_{ij}(0)} \frac{d\lambda_0}{dA_{kl}(0)} A''_{ij}(0) A''_{kl}(0)$. The stochastic growth rate is thus decomposed into
171 the growth rate in the mean environment, $\ln \lambda_0$, plus two additive terms describing changes
172 mainly due to nonlinearity ($\ln \bar{\lambda} - \ln \lambda_0$), and changes mainly due to variance-covariance (σ^2) of
173 demographic parameters in a stochastic environment. The first term can be positive or negative,
174 depending on the nonlinearity index D , and can be further decomposed into effects of survival
175 and fertility coefficients (code in Supporting information S7). The second term corresponds
176 largely to the variance-covariance term in the approximation of Tuljapurkar (1982), except that
177 here there is also a small effect of nonlinearity through C . However, the effect of nonlinearity
178 on the second term is very small compared to the effect of nonlinearity on the mean, therefore
179 we refer to the first term as the nonlinearity component and second term as the variance-
180 covariance component. In the Supporting information S4 we demonstrate the accuracy of this
181 approximation using simulations.

182 **Applying the decomposition**

183 To explore life history variation in the main components of the stochastic growth rate, we used
184 age- and stage-structured matrix population models (MPMs) from the COMADRE Animal Ma-
185 trix Database (v.4.20.5; Salguero-Gómez *et al.*, 2016a) as a starting point, considering different
186 scenarios for effects of the environment z on the demographic parameters. Each MPM includes
187 a projection matrix that depends on the (st)age-specific fertilities, transitions, and survival rates
188 for a given time interval (see Fig. 1). We let this projection matrix represent the matrix in
189 the mean environment, $\mathbf{A}(0)$. We selected MPMs from unmanipulated and free-ranging pop-
190 ulations, considering only 'mean' matrices (i.e., one matrix per population) with annual time
191 steps. Before the analysis we standardized all MPMs to have $\lambda_0 = 1$ by dividing each matrix
192 element by λ_1 calculated from the original matrix (see Supporting information S1 for complete
193 description of selection criteria). One hundred fifty-four MPMs were selected, describing two

194 amphibian, 35 bird, 22 bony fish, three insect, 61 mammal, and 31 reptile populations, belong-
195 ing to 107 species. Generation time was calculated as the mean age of parents at the stable
196 (st)age distribution (Bienvenu & Legendre, 2015) and ranged from 1.1 years to 265.6 years.

197 ***Nonlinear relationships*** - We added environmental effects to the survival and fertility coeffi-
198 cients. Since some models were stage-structured, we first separated out the two matrices contain-
199 ing these coefficients: Each stage structured projection matrix can be written as $\mathbf{A} = \mathbf{GS} + \mathbf{QB}$
200 (Vindenes *et al.*, 2021). Here \mathbf{G} and \mathbf{Q} are matrices describing the stage transition rates of
201 individuals and new offspring, respectively, assumed constant in our analysis. The matrix \mathbf{B}
202 contains the stage-specific fertility coefficients $f_j(z)$ on the diagonal and zeroes elsewhere, while
203 the matrix \mathbf{S} contains stage-specific survival rates $s_j(z)$ on the diagonal and zeroes elsewhere.
204 For each MPM, we chose a link function for the survival rates $s_j(z)$ (logistic or loglog link)
205 and a link function for the fertility coefficients $f_j(z)$ (logistic, loglog, or log link), defining the
206 relationship of $\mathbf{A}(z)$ to the environmental driver z . For each scenario we defined different link
207 functions (Fig. 1B-C), where $s_j(0)$ and $f_j(0)$ corresponded to the values from the standardized
208 MPM in COMADRE. For instance, with a loglog link function, the survival rates are defined
209 as $s_j(z) = \exp(-\exp(-\beta_0 - \beta_{z_S} z))$, and the parameter β_0 is defined as $\beta_0 = -\ln(-\ln(s_j(0)))$.
210 The parameter β_{z_S} defines the strength of the environmental effect on $s_j(z)$, and affects the
211 curvature and variance of survival probability in stage j (Fig. 1A; Fig. S6 shows survival
212 and fertility coefficients for different β_{z_S} and β_{z_F} values). Fertility coefficients are defined in a
213 similar way, but here we also defined a maximum $\text{MaxF} = M * f_j(0)$ with $M = 2.5$ (results
214 for different values of M are shown in Supporting information S5), so that the fertility in the
215 mean environment was set as a proportion of the maximum fertility. The values in the mean
216 environment $s_j(0)$ and $f_j(0)$, defined by the standardized MPM, affect the second derivatives
217 of the link functions (Fig.1A and Fig.S2). A complete description including equations for all
218 link functions and their derivatives is provided in Supporting Information S2 and S7.

219 To limit the number of scenarios we made the simplifying assumption that survival rates of

220 different (st)ages all have the same value of β_{z_S} , and similarly all fertility coefficients have the
221 same β_{z_F} . This means that there is always positive covariance among survival rates of different
222 (st)ages and among fertilities of different (st)ages, while covariance between survival and fertility
223 is controlled in our scenarios by the sign of β_{z_S} and β_{z_F} . These assumptions are biologically
224 relevant for populations where individuals of different (st)ages live in the same environment,
225 and where survival of different stages and reproduction of different stages are affected similarly
226 by a key driver. A range of other scenarios are also possible but not considered here, such as
227 no covariance among demographic parameters.

228 **Scenarios** - The decomposition of the stochastic growth rate was done under 13 scenarios (Fig.
229 1C) varying 1) the type of link function defining $s_j(z)$ and $f_j(z)$, 2) the covariance between
230 survival and fertility; negative or positive (scenarios 1-8), and by applying 3) special cases of
231 forced buffering, turning off the effect of z for certain demographic parameters (thus nonlinearity,
232 variance and covariance of demographic parameters were affected; scenarios 9-13). In the first
233 eight scenarios, effects of z were added to survival and fertility of all stages as described above.
234 Four combinations of link functions were tested, including logistic functions for all parameters,
235 loglog link functions for all parameters, and two combinations of log-link function for $f_j(z)$
236 with logistic or loglog link functions for $s_j(z)$. Each of these four combinations was tested
237 using positive or negative covariance between survival and fertility (Fig. 1). In the scenarios of
238 demographic lability with forced buffering, mature stages were defined as all stages equal to or
239 larger than the stage with first non-zero fertility, and immature stages as all stages preceding this
240 stage. Either $s_{mature}(z)$ or all $s_j(z)$ (scenarios 9-11), or all $f_j(z)$ and $s_{immature}(z)$ (scenarios
241 12-13) were held constant and equal to their value in the mean environment as reported in
242 the standardized COMADRE MPM. We used logistic functions for the other demographic
243 parameters (Fig. 1C). These scenarios reflect different assumptions of demographic lability
244 and buffering within the least or the most influential demographic parameters on population
245 growth, assessed qualitatively depending on the position of the populations along the slow-fast

246 continuum of life histories (Stearns, 1989; Sæther & Bakke, 2000; Gaillard & Yoccoz, 2003).
247 Survival of immature stages and fertility coefficients are assumed to show a higher contribution
248 to population growth in fast-living species, while survival rates of the mature stages are assumed
249 to be more influential for slow life histories.

250 **Decomposition** - For each population in each scenario, we calculated and decomposed the
251 stochastic growth rate $\ln(\lambda_s)$ following equation 3. Since all the MPMs from COMADRE were
252 standardized so that $\ln(\lambda_0) = 0$, the stochastic growth rate is a sum of the nonlinearity and the
253 variance-covariance component. The sign of the stochastic growth rate directly reflects whether
254 the fitness effects of environmental variance (σ_z^2) are positive or negative in that population and
255 scenario. All calculations shown in the main text use the value $\sigma_z^2 = 1$, and altering this value
256 only affects the magnitude of the effects (Supporting information S4). In our analyses, $|\beta_{z_F}|$ and
257 $|\beta_{z_S}|$ were both set to 0.4 (Fig. 1A; results for other values shown in Supporting information
258 S4).

259 Results

260 Combined effects of nonlinearity and variance-covariance among demographic pa- 261 rameters

262 In all scenarios, life histories with short to intermediate generation times (< 10 years) showed
263 consistently stronger fitness responses to environmental variability than slow life histories (Fig.
264 2-3). Whether these responses are positive or negative, strongly depends on the combined im-
265 pacts of covariance structure between the (st)age-specific survival rates and fertility coefficients
266 and their curvatures.

267 Positive effects of lability on the mean fitness $\ln(\bar{\lambda})$ were found mainly among the fast-
268 living species, and positive effects occurred through both survival and fertility (Fig. 2-3). The
269 nonlinearity index D correlated almost perfectly with this nonlinearity component (Spearman
270 coefficient > 0.999 and 0.928 in all scenarios without and with bony fish MPMs, respectively),

271 suggesting that this is a reliable indicator of adaptive lability. However, as fitness $\ln(\lambda_s)$ also
272 depends on the variance-covariance structure of the demographic parameters, this must also be
273 taken into account.

274 With positive covariance between survival and reproduction, $\ln(\lambda_s)$ was consistently reduced
275 compared to the mean environment, regardless of the type of link functions used (e.g., Fig. 2a,
276 b). In these scenarios, positive nonlinearity components still occurred, but were not sufficient
277 to overcome the negative variance-covariance component. In contrast, populations of fast-
278 living species could show an overall positive fitness $\ln(\lambda_s)$ if survival and fertility covaried
279 negatively (Fig. 2 and 3), although less frequent when $s_j(z)$ and $f_j(z)$ were defined by loglog
280 link functions (Fig. 2d). Positive effects were stronger when we used a log-link function for
281 the fertility coefficients, so that they increased exponentially with the environmental driver z
282 leading to strong convexity (Fig. 3c-d). For bony fish MPMs, the signs of the nonlinearity and
283 variance-covariance components were the same as for the other MPMs, but the magnitude was
284 stronger. Here the underlying models from COMADRE showed very high fertility coefficients
285 and low survival rates, yielding extremely high variance in demographic parameters. Under
286 scenarios using loglog link functions for $s_j(z)$ and/or $f_j(z)$, the small noise assumption behind
287 our decomposition of $\ln(\lambda_s)$ was violated to a degree where the approximation broke down for
288 these MPMs (Supporting information S6).

289 **Demographic lability with forced buffering**

290 In these scenarios, some survival probabilities and fertility coefficients were kept constant and
291 buffered, while others were allowed to vary. The identity of labile demographic parameters,
292 together with the position of the species along the slow-fast continuum, affected each fitness
293 component and their combined impact on fitness (Fig. 4). When lability in all survival rates
294 $s_j(z)$ or in only the mature stages $s_{mature}(z)$ was combined with a constant fertility (Fig. 4a-
295 b), only the fastest-living species showed a positive $\ln(\lambda_s)$. This positive fitness resulted from

296 a positive nonlinearity effect of survival rates and a low negative variance-covariance effect,
297 reflecting buffering. When lability in fertility $f_j(z)$ and survival rates of the immature stages
298 $s_{immature}(z)$ was combined with constant survival rates in all stages or mature stages, positive
299 values of $\ln(\lambda_s)$ were also detected, especially when immature survival rates and reproduction
300 covaried negatively (Fig. 4c-e).

301 In contrast, for intermediate and slow-living species, labile survival rates of the reproductive
302 stages $s_{mature}(z)$ combined with constant fertility $f_j(z)$ and constant survival of immature
303 stages $s_{immature}(z)$ (Fig. 4b) always produced negative nonlinearity components, and very small
304 negative variance-covariance components, leading to an overall negative $\ln(\lambda_s)$. The scenarios of
305 lability in fertility coefficients combined with constant (st)age-specific survival rates or in only
306 the mature stages (Fig. 4c-e) showed a weak negative variance-covariance component while the
307 nonlinearity component was zero or slightly positive, leading to overall fitness $\ln(\lambda_s)$ having
308 values close to zero. In other words, constant (st)age-specific survival rates associated with
309 labile fertility coefficients have a stabilizing effect on $\ln(\lambda_s)$ of slow life histories (generation
310 time > 10 years; Fig. 4c VS Fig. 4a-b).

311 Discussion

312 This study emphasizes the importance of considering explicit links between environmental
313 drivers and demographic parameters to understand the effects of environmental variability on
314 fitness, as these links allow effects on nonlinearity to be quantified. We extended Tuljapurkar's
315 approximation of the stochastic growth rate to incorporate effects of nonlinearity in demo-
316 graphic parameters. We also defined a nonlinearity index to measure the overall nonlinearity in
317 a given life history, reflecting the potential for positive fitness effects of environmental variabil-
318 ity. Our decomposition of the stochastic growth rate into nonlinearity and variance-covariance
319 components creates a new framework to study their joint impacts on fitness, expanding ear-
320 lier theory focusing mainly on buffering through the variance-covariance component. Applying

321 this decomposition across a range of scenarios and life histories, we identified the faster-living
322 species as the most responsive to environmental fluctuations, both through the nonlinearity
323 and variance-covariance components. Positive fitness values were only found when positive
324 nonlinearity components were combined with negative covariance between survival and fertility,
325 leading to a smaller negative variance-covariance component. In scenarios with some demo-
326 graphic parameters being constant (forced buffering), lability in both the least and the most
327 expected influential demographic parameters were found to benefit fitness to some extent, but
328 mainly for short-lived species. Our decomposition provides a step forward in our understand-
329 ing of potential adaptations to environmental variability in a wide range of life histories, and
330 stresses the importance of characterising both nonlinearity and covariance structure of demo-
331 graphic parameters with respect to key environmental drivers. Our framework is also useful for
332 predicting population responses to increased variability under global change.

333 **Lability and buffering in fast vs. slow life histories**

334 Several studies have shown evidence that populations located at the fast end of the slow-fast
335 continuum are more sensitive to changes in the different components of climate change. These
336 populations tend to respond more strongly to changes in climate drivers (e.g., Compagnoni
337 *et al.* 2021), to environmental variability (e.g., Dalglish *et al.* 2010; Drake 2005; Koons *et al.*
338 2009; Morris *et al.* 2008, but see Le Coeur *et al.* 2021; Santos *et al.* 2021), to shifts in temporal
339 autocorrelation in the environment (e.g., Paniw *et al.* 2018), and to shifts in the correlation
340 structure of demographic parameters (Iles *et al.*, 2019). In line with these previous studies, we
341 found that populations of faster-living species have larger absolute values of both nonlinearity
342 and variance-covariance components of fitness in a stochastic environment compared to those
343 of slow living-species. On one hand, fast-living species are more vulnerable to environmental
344 fluctuations due to higher negative variance-covariance components, as reported in previous
345 studies (e.g., Dalglish *et al.* 2010; Morris *et al.* 2008). On the other hand, they have the largest

346 potential for adaptive lability through convex demographic responses. Our results show that
347 a positive nonlinearity component can overcome the negative variance-covariance and lead to
348 increased fitness especially when there is a negative correlation between fertility and survival.
349 We found that the nonlinearity index D is a reliable predictor of the nonlinearity component of
350 the stochastic growth rate (eq. 3).

351 A majority of studies have focused on effects of the variance-covariance component alone,
352 without explicit reference to the underlying environmental drivers, even though other studies
353 (Drake, 2005; Henden *et al.*, 2008; Koons *et al.*, 2009) highlighted the potentially critical im-
354 portance of including such links. Our results support this conclusion, and show that the total
355 impact of environmental fluctuations on the fitness of structured populations may be either
356 positive or negative if nonlinear demographic responses are present (eq. 3). With explicit links,
357 where some are convex, positive fitness responses are possible, but we highlight that the net
358 effect also depends on the variance-covariance component and the type of link functions. Ev-
359 idence of convex relationships between demographic parameters or underlying vital rates and
360 key environmental drivers is still limited for natural populations, due to data limitation or *a*
361 *priori* linear assumptions in the statistical models. Our study highlights the need for empirical
362 research to determine more systematically the shape and curvature of demographic parameter
363 responses to accurately predict fitness responses to environmental variance. Quantifying the
364 relationships between environmental drivers and all demographic parameters remains, however,
365 a statistical challenge for wild populations (e.g., separating link functions; Gill, 2001) and re-
366 quires long-term monitoring data (see Lee, 2017 for an alternative method to study nonlinearity
367 in the growth rate response to an environmental driver with discrete levels). This highlights
368 the need to continue and increase the ongoing collection of demographic data.

369 The decomposition of the stochastic growth rate considers nonlinearity and variance-covariance
370 of demographic parameters, which in turn are functions of underlying vital rates. For instance,
371 fertility depends on both fecundity and survival of offspring or parents, depending on the census

372 of the matrix model. Studies applying the method for specific empirical systems should care-
373 fully consider how the demographic parameters depend on lower-level parameters as functions
374 of environmental drivers. Our qualitative conclusions on demographic buffering and lability
375 across generation time are general, but quantitative differences are likely present for instance
376 for models based on pre- vs. post-reproductive census, when environmental effects arise through
377 lower-level parameters. This presents an interesting area for future research using the decom-
378 position.

379 **Role of temporal covariance between (st)age-specific demographic parameters**

380 While negative covariance between demographic parameters could arise from life history trade-
381 offs (Stearns, 1989) or opposite responses to the same environmental driver, positive covariances
382 between these parameters are just as likely to occur in a population. Previous theoretical work
383 has shown that positive covariance enhances the variance in population growth while negative
384 covariance reduces it (Tuljapurkar, 1982, 1990). Our results are in line with this result, showing
385 reduced negative variance-covariance component when survival and fertility covaried negatively
386 compared to positively.

387 Interestingly, there is no general consensus on the degree to which positive or negative
388 covariance in demographic parameters are more common in the wild, nor if the sign, magnitude
389 or type of (st)age-specific demographic parameters involved correlate with the position of a
390 species along the fast–slow continuum (but see a recent comparative study, Fay *et al.*, 2022).
391 From empirical studies, positive covariances have been reported predominantly in long-lived
392 species (e.g., Dahlgren *et al.*, 2016; Rotella *et al.*, 2012; van de Pol *et al.*, 2010) with substantial
393 (e.g., Coulson *et al.*, 2005) or weak (e.g., Altwegg *et al.*, 2007; Compagnoni *et al.*, 2016; Johnson
394 *et al.*, 2010) effects on fitness. In contrast, negative covariances were less often detected (Fay
395 *et al.*, 2022), with often small effects on $\ln(\lambda_s)$. To our knowledge, relatively few studies have
396 specifically addressed this question among species towards the fast-end of the continuum.

397 In our scenarios, we assumed a perfect, positive temporal covariance between (st)age-specific
398 survival rates and between (st)age-specific fertilities, respectively, but positive or negative co-
399 variances between survival and fertility. While these assumptions on the direction of covariance
400 between stages and type of demographic parameters are plausible, they are strong in terms
401 of magnitude, and a main environmental driver is unlikely to explain all of the (co)variance in
402 demographic parameters. Our results may therefore overestimate the magnitude of the variance-
403 covariance component in the decomposition, compared to wild populations where correlations
404 are likely not perfect. Even though we assumed perfect correlation, we found that variance-
405 covariance had negligible effects on fitness of slow-living populations, reflecting a large degree
406 of buffering in these species. For fast-living species, covariance had contrasting effects on the
407 fitness components. These effects were strengthened in scenarios where link functions implied
408 more asymmetric relationships between demographic parameters and environmental driver.

409 **Demographic lability and buffering of different demographic parameters**

410 The set of scenarios combining lability in some demographic parameters with forced buffering
411 in others, yielded insights into possible demographic strategies along the slow-fast continuum.
412 While different predictions have been made as to which demographic parameters should be
413 selected for lability (Hilde *et al.*, 2020; McDonald *et al.*, 2017), we found that demographic
414 lability in the demographic parameters assumed to be the least ($f_j(z)$ and/or $s_{immature}(z)$)
415 or most ($s_{mature}(z)$) important to fitness, could both lead to enhanced fitness in many fast-
416 living life histories due to positive nonlinearity components and reduced variance-covariance
417 components. However, such positive effects on fitness were stronger and more prevalent with
418 lability in both fertility and survival of the immature stages (the most influential in fast life
419 histories). In contrast, for slow-living life histories, lability in the survival rates of mature stages,
420 believed to have the highest impact on fitness led to negative effects on fitness due to negative
421 nonlinearity components. Selection for a reduction in variance in (and in positive covariance

422 between) the demographic parameters that contribute the most to fitness, combined with other
423 parameters varying more freely, as stated by the demographic buffering hypothesis, seems likely
424 for slow-living species, at least those with a similar animal life history as in our analysis.

425 Labile and buffered demographic parameters in our scenarios were qualitatively assigned
426 based on expectations from the demographic buffering and life history theories (Stearns, 1989;
427 Sæther & Bakke, 2000; Gaillard & Yoccoz, 2003). This simple categorization, while accurate
428 for some life histories, may be different for other populations with the same generation time.
429 Further insights would require differentiating labile and buffered (st)age-specific demographic
430 parameters and underlying vital rates in a population based on elasticities of the growth rate
431 in the mean environment.

432

433 In conclusion, this study provides a comprehensive framework for assessing the contribu-
434 tions of demographic lability and buffering on fitness of any given population. Positive effects
435 of environmental fluctuations on fitness are only possible to detect if we account for the im-
436 pacts of nonlinear relationships between demographic parameters and environmental drivers.
437 Our decomposition of the stochastic growth rate into components of nonlinearity and variance-
438 covariance provides a tool to quantify their relative impacts in different life histories and scenar-
439 ios, and is easily applicable for other study systems and scenarios not considered here. Across
440 the slow-fast continuum of animal life histories, faster-living species have the largest potential for
441 using demographic lability as an adaptive response to variability, while demographic buffering
442 is a main adaptive response in slow-living species. These findings have important implications
443 for predicting population and species responses to changes in environmental fluctuations under
444 climate change and other anthropogenic impacts.

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Text box 1: Glossary

Stochastic population growth rate - $\ln(\lambda_s)$: The long-term rate of population growth on a logarithmic scale, a measure of fitness in a stochastic density-independent environment.

Growth rate in the mean environment - $\ln(\lambda_0)$: dominant eigenvalue of the projection matrix in the mean environment ($z=0$) $\mathbf{A}(\mathbf{0})$ on a logarithmic scale.

Mean growth rate - $\ln(\bar{\lambda})$: dominant eigenvalue of the mean projection matrix across variable environments $\bar{\mathbf{A}}$ on a logarithmic scale.

Demographic lability / labile demographic parameter: A labile demographic parameter fluctuates with temporal variation in environmental conditions. The relationship between a labile demographic parameter and the environment (e.g., a key environmental driver) can be convex, concave or linear, so that the average value of this demographic parameter in a variable environment becomes $>$, $<$, or $=$ to the demographic parameter estimated in the mean environment ($z=0$), respectively. The same definition applies to labile vital rates (e.g., survival, fecundity, transition).

Adaptive demographic lability (demographic lability hypothesis): selection for demographic parameters to track environmental fluctuations that leads to an overall increased fitness, $\ln(\lambda_s)$. Increase in $\ln(\lambda_s)$ occurs when an increase in the demographic parameter means due to convexity in their responses leads to a shift in the arithmetic mean of annual population growth rates $\ln(\bar{\lambda})$, that overcomes the negative effect of temporal variance in the annual population growth rates (variance-covariance component σ^2). This hypothesis relies on the assumption that the nonlinearity index D (defined below) is positive.

Nonlinearity index (D): D measures the total effect of nonlinearity of demographic parameters in a life history, and is a key component to describe the nonlinearity component of the fitness decomposition (equation 3). This index corresponds to the sum over all (st)ages of the second derivatives of the demographic parameters (depending on vital rates) in the mean environment ($z=0$), weighted by the sensitivities of λ_0 to the corresponding demographic parameters (matrix elements). When positive (/negative), D is an indicator of adaptive (/non-adaptive) lability through overall positive (/negative) contributions from convexity (/concavity) of the demographic parameters. Adaptive lability can create a positive overall effect of environmental variability if D is positive and the negative effects of increased variance-covariance of the demographic parameters are not too large (see equation 3).

Demographic buffering / buffered demographic parameters: Low variance of a demographic parameter in response to temporal variation in the environmental variable z . A more flat relationship between the demographic parameter and the environment z leads to such low parameter variance, and to the mean demographic parameter in the variable environment remaining approximately equal to demographic parameter value in the mean environment ($z=0$). The same definition applies to buffered vital rates (e.g., survival, fecundity, transition).

Adaptive demographic buffering (demographic buffering hypothesis): The prediction that natural selection should favour a reduction in variance of the demographic parameters with the strongest influence on fitness in the mean environment, reducing the variance-covariance component σ^2 and leading to an overall stable or increased fitness in variable environments. The assumption that $\ln(\bar{\lambda})$ is not affected by environmental variance ($\ln(\lambda_0) \approx \ln(\bar{\lambda})$), is often made for this hypothesis.

Figure captions

Figure 1. Framework used to study the effects of environmental variability on fitness (stochastic growth rate $\ln(\lambda_s)$). **A.** Our calculations define demographic parameters as non-linear functions of the environmental driver z (see methods), where $\mathbf{A}(0)$ (from our selected, standardized COMADRE models, $N_{tot}=154$) defines the values of (st)age-specific survival rates $s_j(0)$ and fertilities $f_j(0)$ in the mean environment ($z=0$). Different levels of environmental variance levels σ_z^2 and environmental strength $|\beta_z|$ of z on the demographic parameters were considered. In the analytical approach, $\ln(\lambda_s)$ was calculated and decomposed into main components capturing nonlinearity and variance-covariance effects following equation 3. The accuracy of this decomposition was tested using simulations (Supporting information S4). **B.** Two or three different link-functions were considered for survival $s_j(z)$ and fertility $f_j(z)$, respectively. **C. Scenarios 1-8:** Four combinations were examined including logistic functions for all parameters, loglog link functions for all parameters and two combinations of exponential fertilities $f_j(z)$ (log link) with logistic or loglog link function for $s_j(z)$. Positive or negative covariance between survival and fertility was tested for each combination, assuming positive covariance between $s_j(z)$, and between $f_j(z)$. **Scenarios 9-11:** Scenarios of forced buffering considering demographic lability in the fertility coefficients and survival rates of the immature stages ($S_{immature}$). **Scenarios 12-13:** Scenarios of forced buffering assuming demographic lability in all survival rates $s_j(z)$ or in only the mature stages (S_{mature}). Logistic functions were used to define lability while the other rates were held constant and fixed to the values reported in the standardized COMADRE projection matrix.

Figure 2. Mid panels: Stochastic growth rate (fitness) $\ln(\lambda_s)$ across generation time, under four scenarios of covariance and link-functions of the demographic parameters. Left panels: Illustration of scenarios, with grey and black lines corresponding to the (st)age-specific survival rates $s_j(z)$ and fertility coefficients $f_j(z)$, respectively (functions varied for each stage depending on $s_j(0)$ and $f_j(0)$; only one function is shown for survival and fertility here). We assumed positive covariance between survival rates of different (st)ages and between the fertilities of different (st)ages. For each scenario and for each population, positive (panels a,b) or negative (panels c,d) covariance between $f_j(z)$ and $s_j(z)$ were considered, treating $f_j(z)$ and $s_j(z)$ as logistic functions (panels a,c) or loglog link functions (panels b,d) of the environment z . Right panels: Decomposition of $\ln(\lambda_s)$ into main components capturing variance-covariance effects (blue triangles) and lability effects generated by nonlinear responses of $f_j(z)$ (red circles) and $s_j(z)$ (orange circles). Results for bony fish populations and populations with generation time

> 62 years are not shown ($N_{MPMs} = 129$; see Fig. S14 for all MPMs).

Figure 3. Mid panels: Stochastic growth rate (fitness) $\ln(\lambda_s)$ across generation time, considering positive (panels a,b) or negative (panels c,d) covariance between (st)age-specific survival rates $s_j(z)$ and fertilities $f_j(z)$, treating $s_j(z)$ as logistic (panels a,c) or loglog (panels b,d) link functions of the environment z and $f_j(z)$ as log link functions. We assumed positive covariance between survival rates of different (st)ages and between the fertilities of different (st)ages. See Fig. 2 for explanation of left and right panels. Results for bony fish populations and populations with generation time > 62 years are not presented ($N_{MPMs} = 129$; see Fig. S15 for all MPMs).

Figure 4. Results from scenarios of forced buffering assuming demographic lability only in (a) (st)age-specific survival rates, (b) survival rates of the reproductive stages only, (c) (st)age-specific fertilities and (d-e) fertilities and survival rates of the immature stages. For each scenario, the long term fitness $\ln(\lambda_s)$ and its main components reflecting variance-covariance effects (blue triangles) and lability effects due to nonlinearity of $f_j(z)$ (red circles) and $s_j(z)$ (orange circles) are plotted against generation time (mid and right panels; see Fig. S16 for all MPMs). See Fig. 2 for explanation of left panel.

Figure 1.

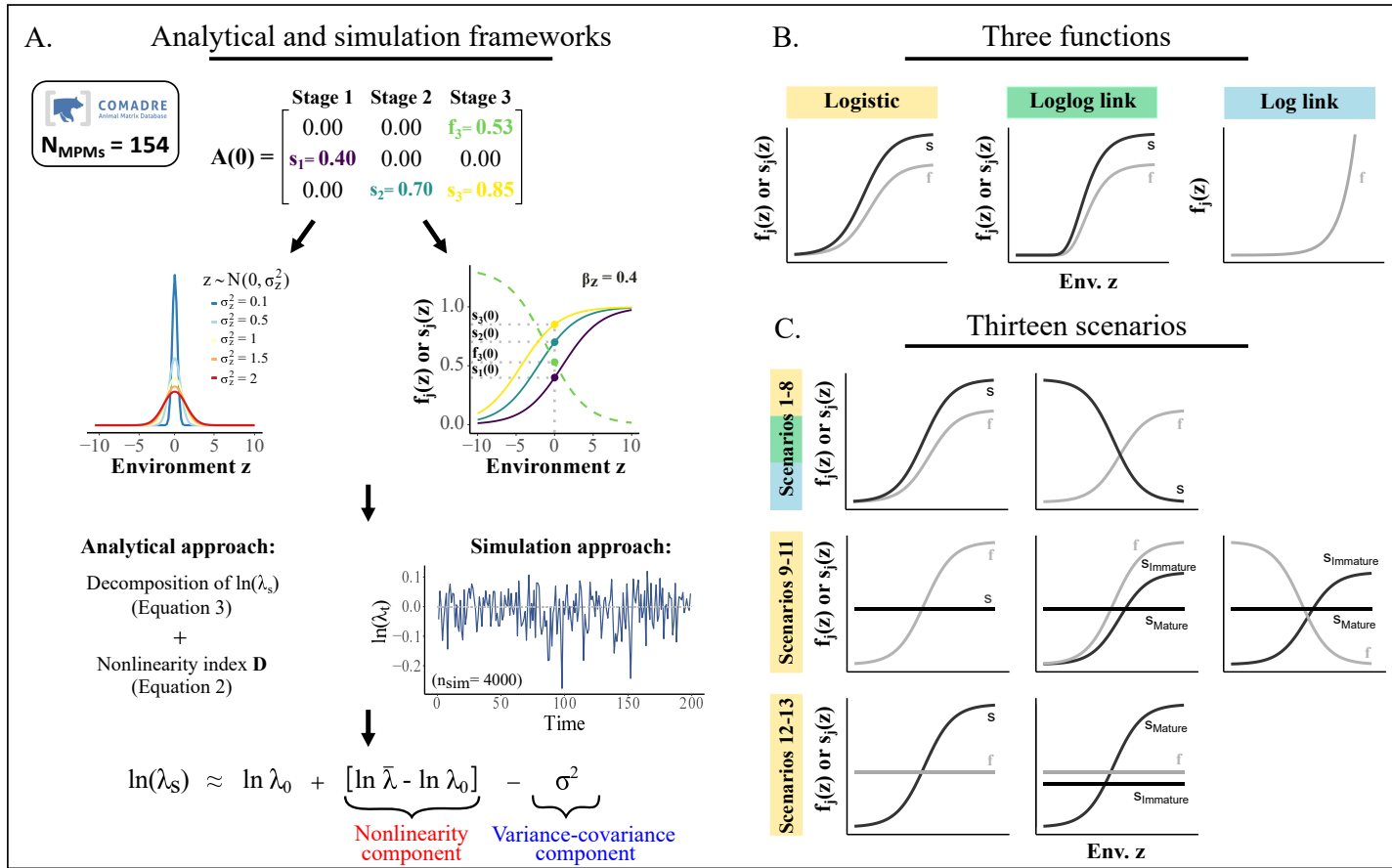


Figure 2.

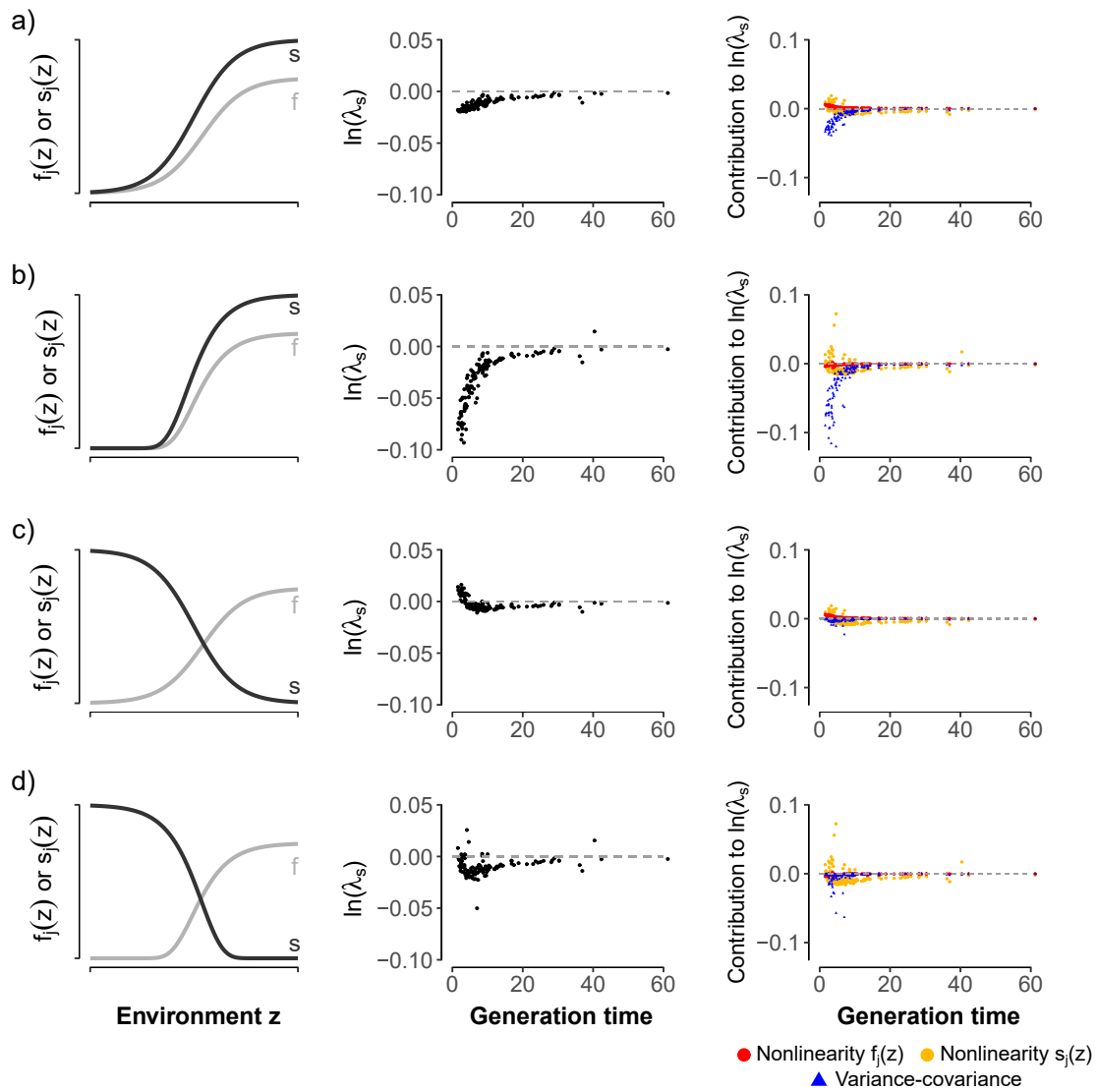


Figure 3.

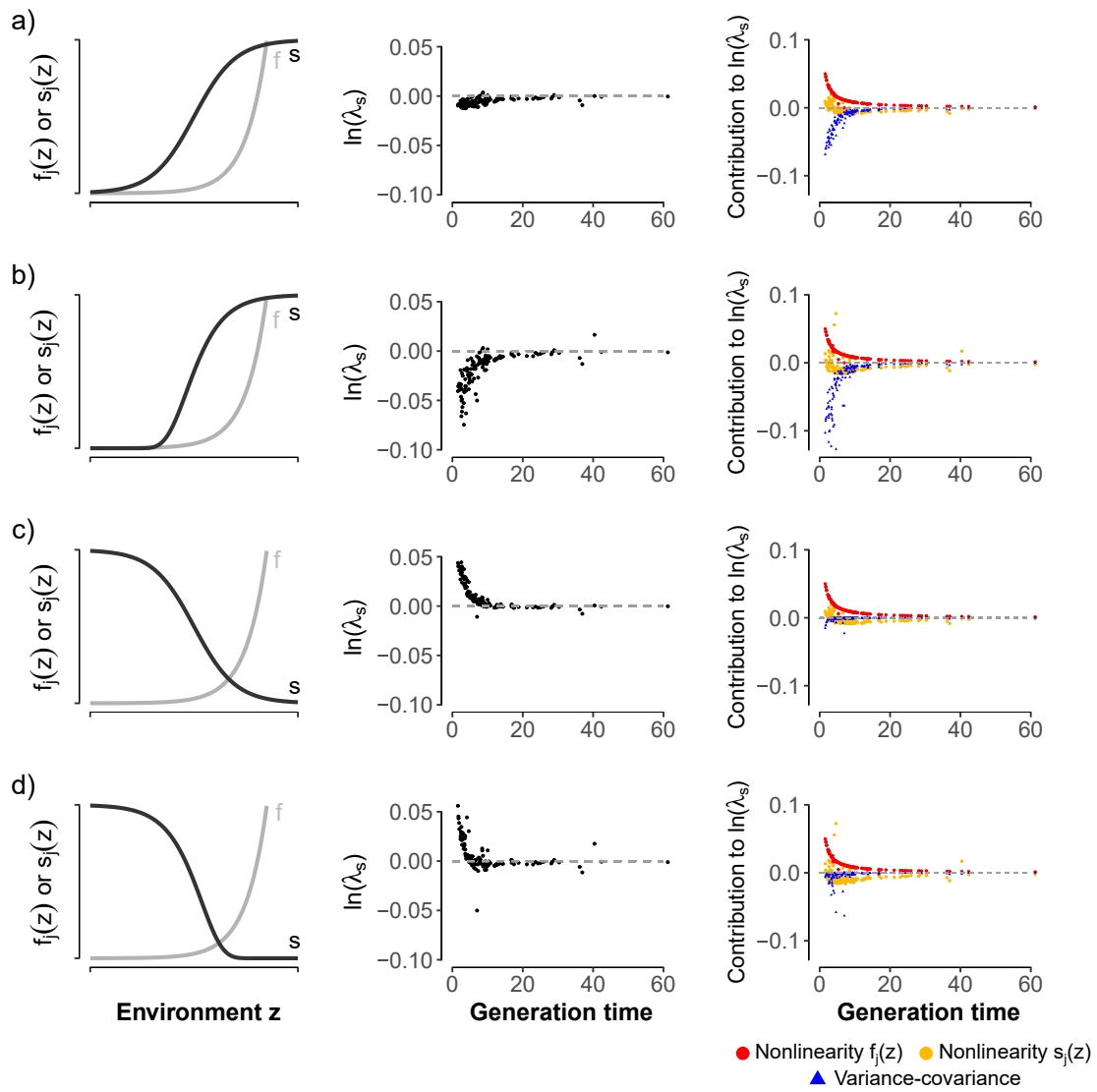


Figure 4.

