Opposite-sex associations are linked with annual fitness, but sociality is stable over

lifetime.

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Lay Summary

How many social connections birds have is linked to their success in the following breeding

season. However, we do not know whether this holds true across a birds' lifetime. Our

research confirms that birds with more connections to the other sex have higher annual

breeding success. However, over a lifetime we find that birds with an average amount of

social connections have better fitness than those at the extremes. This suggests that

evolution may favour an average sociality.

Abstract

Animal sociality, an individual's propensity to associate with others, has fitness

consequences through mate choice, for example, directly, by increasing the pool of

prospective partners, and indirectly through increased survival and individuals benefit

from both. Annually, fitness consequences are realised through increased mating success

and subsequent fecundity. However, it remains unknown whether these consequences

translate to life-time fitness. Here, we quantified social associations and their link to fitness

annually and over lifetimes, using a multi-generational, genetic pedigree. We used social

network analysis to calculate variables representing different aspects of an individual's

sociality. Sociality showed high within-individual repeatability. We found that birds with

more opposite-sex associates had higher annual fitness than those with fewer, but this did

not translate to lifetime fitness. Instead, for lifetime fitness, we found evidence for

stabilizing selection on opposite-sex sociality, and sociality in general, suggesting that

reported benefits are only short-lived in a wild population, and that selection favours an

average sociality.

Keywords: sociality, social network analysis, selection, social behaviour, genetic
pedigree, long-term fitness, de-lifing, lifetime reproductive success

Background

Some individuals are consistently more sociable than others, demonstrated by within-individual repeatability of social traits across vertebrate groups (Aplin et al. 2015; Thys et al. 2017; Dimitriadou, Croft and Darden 2019; Plaza et al. 2019; Beck, Valcu, and Kempenaers 2020; Proops et al. 2021; Strickland et al. 2021). This variation in individual sociality is positively linked with fitness in some taxa (Silk 2007; Silk et al. 2009) and is therefore expected to be subject to selection (Krause and Ruxton 2002). Fitness is a relative measure of an individual's genetic contribution to the population in the next generation, and thus, can only be represented comprehensively and precisely by measures of traits spanning an organism's lifetime (Endler 1986; Blankenhorn 2010; Reid et al. 2019; Moiron, Charmantier and Bouwhuis 2022). A comprehensive definition of fitness is fundamental to understand the evolutionary pressures that shape variation in sociality. In practice, however, many studies must rely on fitness correlates instead (e.g., number of broods, or survival, instead of genetic contribution). For example, in mammal societies, both variation in within- and between-sex affiliations (Archie et al. 2014) have been linked to lifetime fitness correlates, survival and longevity (Cameron, Setsaas and Linklater 2009; Silk, Alberts and Altman 2003; Silk et al. 2010; Stanton and Mann 2012, but also see Thompson and Cords 2018). Whereas, in birds, the subject of this study, the use of fitness correlates (eggs laid, chicks fledged, within-year survival etc.) are frequent against more precise fitness measures (Moiron, Charmantier and Bouwhuis 2022), that require intensive field work over a long period of time.

Although some benefits of maintaining close social associations are linked with social behaviors during the breeding period (Bebbington et al. 2017; Riehl and Strong 2018), most of these benefits tend to be short-term and contextual. Instead, benefits associated with reproduction are often linked with social behaviour during the non-breeding period (Firth and Sheldon 2016; Kohn 2017; Maldonado-Chaparro et al. 2018; Beck, Farine, and Kempenaers 2020), when group cohesion is stronger (Silk et al. 2014; Kurvers 2020). For example, communal foraging between socially associated individuals during the non-breeding period facilitates resource information transfer (Aplin et al. 2012; Firth, Sheldon

and Farine 2016; Hillemann et al. 2020) and reduces predation risk (Cresswell 1994;

Cresswell and Quinn 2011). However, these benefits may also incur costs associated with

competition for resources, and during the breeding period, competition for mates

(Birkhead and Biggins 1987; Le Galliard et al. 2005; Forstmeier et al. 2011; Mayer and

Pasinelli 2013; Grant and Grant 2019; but also see Lea et al. 2010). Likewise, individuals

that hold more central social network positions in the non-breeding period may increase

their breeding success, by maximizing their opportunity for mate choice (McDonald 2007;

Oh and Badyaev 2010; Firth et al. 2018; Beck, Farine and Kempenaers 2021). Although the

association between annual fitness correlates and non-breeding sociality has been well

described, data on the link with lifetime fitness, and selection of sociality, remains scarce.

With the recent development of tools to construct and analyse social networks (Wey et al.

2008; Farine and Whitehead 2015) the study of sociality has become popular in ecology

and evolution. Yet, to describe the association between sociality and fitness any potential

study must first overcome two problems: (1) A social association must be clearly defined

relative to the behavior of the study system (Figure 1; Psorakis et al. 2012, 2015), and (2)

to study the evolution of social behaviour, precise measures of individual fitness must be

quantified to test for correlation with a social trait – selection. Although annual fitness

correlates are widely used, lifetime fitness is more precise than correlates, not subject to

the stochasticity of annual measures (Dobson, Murie and Viblanc 2020; Alif et al. 2022) and

thus, can better describe selection pressure acting on a trait (Endler 1986; Blankenhorn

2010; Reid et al. 2019). However, lifetime fitness requires wild animals to be monitored

throughout their whole lives, and all breeding attempts, and the fates of their offspring,

must be recorded to determine recruitment. All of these require a multi-generational,

genetic pedigree (Kruuk 2004; Korsten et al. 2013).

Our study system, an island population of house sparrows *Passer domesticus* (hereafter,

sparrow/s) where we monitor all individuals from birth to death, without capture bias

(Simons et al. 2015), overcomes both problems. (1) We have sociality data from birds that

are electronically registered visiting a feeder. Sociality measures are repeatable across

different timescales and contexts in this and other populations (Plaza et al. 2019). (2) We

have lifetime recruitment data available, from a multi-generational genetic pedigree that,

because our population is closed, we can use to compute precise annual and lifetime fitness

estimates (Schroeder et al. 2015; Alif et al. 2022).

Here, we aim to better understand the potential for selection on sociality, and therefore,

our two aims are: (1) We confirmed that the sociality traits we measured were meaningful

by testing for individual repeatability of sociality. (2) We tested for selection on non-

breeding sociality using annual, and lifetime fitness measures.

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111 **Methods**
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Study system

We used data from the Lundy sparrow system, a long-term study based on the island of

Lundy (51.11N, 4.40W), ≈19 km off North Devon, UK, in the Bristol Channel. The sparrows

on Lundy breed in nest boxes, sited in groups around the only village on the island. The

island is rodent-free and therefore the sparrows have no predators but for the occasional

vagrant raptor. House sparrows are a model organism in behavioral ecology and evolution,

and much is known about their biology, physiology, and life-history (Andersson 1994;

Sánchez-Tójar et al. 2018). Females are socially monogamous, but 25% of broods show

they can be genetically promiscuous (Schroeder et al. 2016). On Lundy, they have on

average 2-3 broods of 4-5 eggs per breeding season (Westneat et al. 2014). The sex ratio is

stable, and the mean lifespan of recruits is three years (Alif et al. 2022). Although sparrows

are territorial during the breeding season, during the non-breeding period they form

gregarious groups that forage communally for seeds and at supplementary bird feeders

(Summer-Smith 1963), both of which are available year-round on Lundy.

Most sparrows were first captured, and tissue sampled at their natal site during the

breeding season (April to August) and a few during the post-fledging period (Schroeder et

al. 2011; Girndt et al. 2019). We genotyped sparrow DNA at <22 microsatellite loci suitable

for parentage assignment in sparrows (Dawson et al. 2012). Using the genetic data, we

assembled a near-complete genetic pedigree (Schroeder et al. 2015, 2016), which at the

time of writing spans twenty years, 2000–2019, and 8,379 individuals. We fitted all

sparrows with a unique combination of a coded metal ring and three coloured leg rings. We

also provided each sparrow with a subcutaneous Passive Identification Transponder (PIT

tag; TROVANID100: 12 × 2 mm and 0.1 g), which we have previously shown have no

detrimental effect on subsequent fitness (Schroeder et al. 2011). These tagged individuals

were then recorded when they visited a custom made 19.8cm x 19.8cm Radio Frequency

Identification antenna (RFID; DorsetID) mounted below a seed reservoir (Sánchez-Tójar et

al. 2017; Brandl et al. 2019), positioned centrally within our study site. The feeder was

open access, and explicitly not limited to a single bird feeding at one time, as is the case at a

hanging bird feeder (Youngblood 2019; Beck, Farine and Kempenaers 2020), taking into

consideration the gregarious feeding behaviour of house sparrows.

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Sociality measures

To quantify the sociality of individual sparrows we used presence data from the RFID

antenna, collected during the non-breeding periods, November–January, of 2015/16 (15

weeks recording time) and 2016/17 (13 weeks recording time, hereafter referred to as two

events). An association between birds observed from these data can either reflect two or

more individuals that choose to maintain some social cohesion with one another. We

consider this to be a non-random association. Another option is that two individuals

without pre-established social cohesion coincide in time and space. We consider this

random mixing. The data records the presence of birds without distinguishing between

these two types of association. Hence, we derived a method to infer non-random

associations from these data.   Further, dyadic interactions can also be aggressive

interactions, and our data cannot exclude these cases. Sparrows have long considered to be

a text-book model for signaling dominance hierarchies. Although, a recent meta-analysis

did not support the idea that the black bib of the house sparrows predicts dominance

status (Sánchez-Tójar et al. 2018). In sparrows' dominance hierarchies are typically linear,

while, not particularly steep, and there is no evidence for a correlation with reproductive

fitness (Sánchez-Tójar 2018).

The 'gambit of the group' is a common approach used to identify discrete groups among all

associating individuals (Whitehead and Dufault 1999; Figure 1A). However, given the

gregarious nature of sparrows and the high activity at our feeder, at which non-discrete

groups of sparrows accumulated, the gambit of the group approach overestimated

associations between individuals (Figure 1A; also see Ferreira et al. 2020), capturing both

random mixing and non-random associations. One solution to this would be to define a

non-random association as whenever two individuals overlapped by a defined period (Δ_t)

in the proximity of our bird feeder. However, in the sparrow system, again due to the near

constant visitations, this resulted in linear network structures, e.g., linking the first bird to

the second, then the second to the third, and so on for the whole observation period (Figure

1B). To account for the social behaviour, and dominance hierarchies, of sparrows, we

instead assumed that small groups and non-random social associations are established

before they attend the feeder (Summer-Smith 1963). Therefore we defined an association

simply as two individuals that arrived to feed within 150 seconds (Δ_t) of each other. Here,

an arrival is defined as the (re)appearance of the individual at the feeder after being absent

for a period of minimally >300 seconds (Δ_I). From our experience watching sparrows in

the field and on pre-recorded footage, arriving at our feeder (Plaza et al., 2019) to conclude

that Δ_t = 150 seconds was long enough to detect and link all individuals who arrive

together in a group (see Figure 1C), and the resulting data better samples non-random

associations between individuals.

From the resulting association matrices of the two events, 2015/16 and 2016/17, we built

a series of weighted, non-directional, social networks (hereafter, network/s), where the

vertices represent individual sparrows and interconnecting edges their associations. First,

we built individual networks for each of the 15 weeks in 2015/16, and 13 weeks in

2016/17, to estimate within-individual repeatability in social network metric, to validate

our sociality measures against the individual repeatability already demonstrated in this

system using both RFID and video data (Plaza et al. 2019). Then, we built a network for

each event to extract non-breeding sociality. Finally, we also built two bipartite networks

from each event (sub-graphs), which only considered association strength between

opposite-sex individuals.

From the first non-breeding networks we extracted three measures of sociality for each

individual using the 'iGraph' R package (Csardi and Nepusz 2006): We selected the sociality

measures to reflect aspects of an individual's social preference, following similar studies on

sociality (McDonald 2007; Beck, Farine and Kempenaers 2021). Degree represents the

number of associated individuals. Strength represents the quality of those relationships.

We calculated strength following (Farine 2013), using the sum of dyadic Simple Ratio

Indices (the association probability between a dyad, from 0, never associated, to 1, always

associated), which we transformed to give a measure of net association quality (Eq. 1,

where S denotes strength, $d_{(i)}$ and sum $SRI_{(i)}$ the degree and SRI of a given individual

respectively and N(V) = the number of vertices in a network)

$$(Eq.\,1) \quad S \,=\, d_{(i)}\frac{SRI_{(i)}}{N(V)}$$

We used eigenvector centrality (following McDonald 2007, hereafter centrality) to quantify

the influence of an individual to all others within the network (Newman 2004); Finally, we

extracted opposite-sex degree from the two bipartite sub-graphs, and which we used to

represent an individual's pool for mate choice as a fourth measure of sociality (following

Beck, Farine and Kempenaers 2021). We did not threshold our networks (following Farine

and Whitehead 2015) to maintain network structure.

Fitness measures

For each of the sparrows that survived to the following breeding period we used the

genetic pedigree to calculate two fitness measures, and for each of these, we calculated one

annual measure, and one across the lifetime of the sparrow. Both fitness measures are

based on the number of recruits, and we have shown that they correlate well with

reproductive value, thus represent fitness in an evolutionarily meaningful way (Alif et al.

2022). We defined recruits as offspring that survived and produced genetic offspring

themselves.

(1) *Recruits*. For the number of annual recruits, we summed individual recruits within the

breeding year following the social events. We then again summed individual recruits across

a lifetime, or up to 2020 as a measure of lifetime recruits. Note that the latter category only

contained five sparrows that were still alive at the point of census, and as such, our

recruitment data can be considered near complete.

(2) **De-lifed fitness**. As a second fitness measure, we used de-lifed fitness (p_{ti}; Eq. 2), which

estimates an individual's genetic contribution to the population (Coulson et al. 2006). De-

lifed fitness is a retrospective measure of realized fitness, relative to the population each

year, calculated by removing (de-lifing) an individual and its offspring from the pedigree

and recalculating the resulting change in population growth.

$$(Eq.\,2) \quad p_{ti} = \frac{\xi_{t(i)} - w_{t(i)}}{N_t - 1}$$

Here, p_{ti} is the contribution of individual (i) to population growth during a specific period t.

Further, $\xi_{t(i)}$ is a measure of individual performance, here the number of surviving

offspring of individual i at the end of the breeding period t. We added a value of one if the

individual i itself survived to the next breeding period $t + 1$. The population size at time t is

N_t at the beginning of each breeding cycle (here April). To estimate the individual's

contribution to population growth, we use w_t, which represents the ratio of the population

size at $t + 1$ to the population size at t. This de-lifed fitness p_{ti} is an annual value per

individual, and we calculated it for all birds which produced at least one recruit. We then

also summed p_{ti}, within individuals as a lifetime de-lifed fitness measure, p_i.

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Individual repeatability in sociality

First, we validated our sociality measures by confirming that they were repeatable within-

individuals between weeks, but within season. We modelled degree, strength, centrality,

and opposite-sex degree respectively, as response variables each in four Gaussian linear

mixed models against the intercept, and with bird identity modelled as a random effect. We

then divided the variance explained by bird identity by the total phenotypic variance of the

trait to quantify repeatability (Nakagawa and Schielzeth 2010). We ran repeatability

models using package MCMCglmm default parameters and priors – the models converged

robustly and reliably.

Selection on sociality – annual fitness

We quantified the association between sociality measures and fitness. As all four sociality measures are inherently correlated, we modeled each separately to avoid collinearity (Webster, Schneider and Vander Wal 2020). For each sociality measure, we ran two models, with annual recruits and annual de-lifed fitness as the response variables. In the models with annual numbers of recruits as response, we assumed a Poisson error distribution with a log link function, and in the models explaining de-lifed fitness we assumed a Gaussian error distribution. We mean-centered all social measures within each year, eliminating between-year differences, and modelled them as fixed covariates. To test for stabilizing or disruptive selection (Wolf et al. 2007), where averages are favored over the extremes, we added each sociality variable also as quadratic effect. Bird identity was modelled as a random effect on the intercept to account for pseudo-replication, and cohort to account for environmental stochasticity. We modelled fixed effects for sex (male, 1 or female, 0) and age (in calendar years) and age as a quadratic effect, to account for variation in fitness as explained by demography (Schroeder et al. 2012). We added sex as an interaction term with age to account for the extra-pair behaviour of older males (Girndt et al. 2018).

Selection on sociality – lifetime fitness

We modelled lifetime recruits and lifetime de-lifed fitness in the same way as the annual ones, but instead of age we used lifespan, or maximum age at year 2020. Because each bird was only represented once in this dataset, we only modelled cohort as a random effect.

Statistical analyses

We used Bayesian Markov Monte-Carlo simulations, using MCMCglmm (Hadfield 2010), to run all models. For all models we used inverse Wishart priors for random effects, and ran each over 343,000 iterations, with a burn-in of 3,000 and a thinning interval of 200. We visually checked the posterior trace plots for all model outputs and ensured that autocorrelation was below 0.1 and effective sample sizes between 1,000 and 2,000. The fixed effects were considered statistically significant when the 95% credible interval (CI) of its posterior distribution did not span zero.

Null models and dominance interactions

We ran a node-permutation null-model by randomizing bird identity against arrival time,

without replacement, before constructing new networks and calculating mode network

centrality measures from 1000 randomized permutations (Farine, 2017). We used these

mean random measures to re-run all fitness models. Finally, to exclude the possibility that

dominance was interacting with sociality measures, we tested for correlations between the

sociality measures and between ELO ratings calculated from video observations taken

within our events (Sánchez-Tójar 2018).

Results

Our data consisted of 150 individual birds making 410,114 visits to the RFID feeder within

our study period (mean = 2,734 visits per bird, SD = 8,116), across both events. After

constructing the networks, we identified 3,783 associations between 118 PIT tagged

individuals during 2015/2016, and 874 associations between 69 individuals in 2016/2017.

These networks contained 66.3% of 122 and 26.3% of 205 breeding birds in 2016 and

2017 respectively. Combined, we had 130 records for annual and lifetime fitness from 102

individuals, and where 33 were recorded in both years (for summary statistics see Table

1). The correlation between degree and opposite-sex degree implying that those with more

opposite-sex associates also tend to also have more associates of either sex.

Individual repeatability in sociality

299 We confirmed individual repeatability by week in all four sociality measures between 15

300 weeks in 2015/16 and 13 weeks in 2016/17 (Table 1).

Selection on sociality – annual fitness

Opposite-sex degree had a statistically significant, negative linear relationship with the

number of annual recruits and annual de-lifed fitness. Strength and centrality had a

negative quadratic association with annual recruitment (Figure 2; Table 2; Table 3). Age

and sex both also predicted annual recruits, with younger individuals and females

recruiting more offspring. Age also positively predicted annual de-lifed fitness (Table2;

Table3). Older males had greater recruitment annually.

Selection on sociality – lifetime fitness

None of our sociality measures statistically significantly predicted lifetime recruitment

(Supplementary material; Table 4), however, all four had a statistically significant negative

quadratic relationship with lifetime de-lifed fitness (Figure 2).

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Null models and dominance interactions

We found no link between fitness and sociality, nor any evidence of selection from our null

models (supplementary files, Figure 3). Likewise, dominance was not strongly correlated

with sociality, implying that our method of assigning associations based on arrival time,

rather than shared space at a bird feeder is not influenced by dominance (supplementary

files).

Table 1. Summary statistics for fitness and summary statistics and repeatability for social measures for individual house
sparrows on Lundy Island during two non-breeding events (November–January 2015/2016 and 2016/2017). Each measure is
given as mean and standard deviation, as well as an R value which measures repeatability and a 95% CI.

Fitness Variable	Scale	2015/16 Mean, SD		2016/17 Mean, SD	
Recruits	Annual	2.65, 1.98		0.78, 1.07	
	Lifetime	5.34, 3.76		4.15, 3.86	
Sociality measure		Mean, SD	R, 95CI	Mean, SD	R, 95CI
Degree		64.11, 32.03	0.29, 0.15–0.39	25.33, 17.37	0.35, 0.16–0.49
Strength		0.06, 0.05	0.22, 0.10–0.32	0.09, 0.13	0.33, 0.16–0.48
Centrality		0.35, 0.30	0.15, 0.03–0.27	0.25, 0.31	0.23, 0.10–0.36
Opposite-sex degree		32.00, 15.86	0.27, 0.13–0.4	12.18, 9.19	0.45, 0.24–0.61

326 Table 2. Model outputs, one for each of our four sociality measures (degree, strength, centrality, and opposite-sex degree),
327 against annual recruits of house sparrows on Lundy Island, with a Poisson distribution. We inferred significance where the
328 95% CI did not span zero, positive effects on the response variable are highlighted in red, and negative in blue.

Annual recruits	Posterior mode 95% credible intervals			
Variable	Degree	Strength	Centrality	Opposite-sex degree
Fixed terms				
(Intercept)	3.49 1.46—6.11	3.69 1.58—6.46	4.19 1.60—6.31	3.78 1.58—6.50
Social measure	0.17 -0.01—0.34	0.22 -0.03—0.44	0.26 -0.01—0.45	0.25 0.05—0.41
Social measure^2	-0.1 -0.26—0.06	-0.12 -0.32—0.07	-0.16 -0.34—0.08	-0.04 -0.16—0.10
Sex (male)	-0.67 -1.41—0.05	-0.7 -1.46—-0	-0.74 -1.48—-0	-0.57 -1.47—0
Age	-1.14 -2.03—-0.30	-1.15 -2.04—-0.28	-1.27 -1.99—-0.26	-1.07 -2.00—-0.26
Age^2	0.05 -0.07—0.17	0.04 -0.08—0.17	0.06 -0.08—0.16	0.03 -0.07—0.17
Age*Sex (male)	0.21 -0.01—0.48	0.22 -0.02—0.46	0.28 -0.02—0.49	0.21 0—0.49
Random terms				
*Bird ID	0 0—0.13	0 0—0.12	0 0—0.12	0 0—0.13
*Cohort	1.88 0.25—15.69	1.5 0.19—15.57	2.09 0—15.89	2.29 0.11—16.90
Residuals	0 0—0.27	0 0—0.28	0 0—0.29	0 0—0.24

330 Table 3. Model outputs from GLMMs for each of our four sociality measures (Degree, Strength, Centrality and Opposite-sex
331 degree), against lifetime recruits in the Lundy house sparrow system. We inferred significance where the 95%CI do not span
332 zero, positive effects on the response variable are highlighted in red, and negative in blue. († Age in lifetime models denotes
333 either lifespan, or age in 2020, whichever is greatest)

Lifetime recruits	Posterior mode, 95% credible intervals (lower – upper)			
Variable	Degree	Strength	Centrality	Opposite-sex degree
Fixed term				
(Intercept)	0.85 0—1.93	0.85 0.02—1.95	0.99 -0.06—1.87	0.8 -0.03—1.83
Social measure	0.1 -0—0.25	0.1 -0.09—0.3	0.12 -0.07—0.31	0.14 -0.05—0.25
Social measure^2	-0.05 -0.19—0.11	-0.01 -0.24—0.11	-0.07 -0.24—0.11	-0.03 -0.16—0.12
Sex (male)	0.33 -0.35—0.99	0.28 -0.36—0.91	0.28 -0.33—0.98	0.48 -0.24—1.03
†Age	0.1 0—0.23	0.11 0—0.24	0.09 0.01—0.25	0.12 0—0.24
†Age*Sex (male)	-0.06 -0.20—0.07	-0.07 -0.19—0.07	-0.07 -0.2—0.06	-0.06 -0.21—0.04
Random effects				
*Cohort	NA 0.07—2.78	0.47 0.07—2.76	0.45 0.06—3	0.39 0.10—2.77
Residuals	NA 0—0.27	0.12 0.01—0.29	0.11 0.02—0.28	0.14 0—0.26

Discussion

We found evidence for annual fitness benefits of sociality, where individuals with more opposite-sex associates had higher fitness in the following breeding period, than those with fewer opposite-sex associates, but that this did not translate to lifetime fitness. For lifetime fitness, we found evidence for stabilizing selection on sociality measures, including opposite-sex degree, suggesting that such benefits are only short-lived, or contextual, in a wild population.

We constructed our networks by linking dyads of birds that arrived together to a bird feeder, but ignored the time that they spent there, to eliminate random associations. Other studies have also adapted similar approaches in high-density and open feeder systems, or have considered the same implicit problems (Gomes, Boogert and Cardoso 2021). Ferreira et al. (2020), for example, identified flocks arriving, but then defined associations by proximity at the feeder. Further research could optimize our approach for other systems, either by refining the time after which an individual is determined to have left the feeder ($\triangle_i$), which we defined by watching a feeder with binoculars and thus, is likely to be specific to our system, or similarly, the time it takes all members of a group to interact with the feeder upon arrival ($\triangle_t$).

Where previous studies on populations of wild birds have suggested links between aspects of sociality and annual fitness correlates (for examples see Firth and Sheldon 2016; Kohn 2017; Beck, Farine, and Kempenaers 2020), we also used lifetime measures, which better reflect the genetic contribution of the individual to population growth, and thus, fitness. This also allows us to describe how selection acts upon sociality across the population. We found that sociality had little influence on fitness at the annual scale, bar for opposite-sex association, which was linked to increased recruitment. Although not tested explicitly, our study corroborates that annual fitness benefits described elsewhere, particularly regarding mate choice (Oh and Badyaev 2010; Beck, Farine, and Kempenaers 2020, Beck, Farine, and Kempenaers 2021), directly translate into increased annual fitness. At the lifetime scale

however, our study also provides some insight into the evolution of social behaviors, which

we found to be maintained at the population average through stabilizing selection.

Stabilizing selection in this case may be driven by factors such as high mate fidelity or

changing sociality with age (Oh and Badyaev 2010; Albery et al. 2022), removing the need

to constantly maintain opposite-sex associations over lifetime while maintaining individual

fitness. However, those opposite-sex associations may be beneficial in an extra-pair

context from the male perspective (Beck, Farine, and Kempenaers 2020), but support for

sociality as a mechanism for 'good-genes' hypotheses are scarce (Hsu et al 2015) and

require further research.

Our sociality measures were associated with de-lifed fitness and less so with recruitment at

both annual and lifetime scales. This is probably because, although sociality may enhance

mate choice, recruitment is also dependent on parental effects and relationships within the

breeding season, which were not quantified here, although they have been suggested

elsewhere (Bebbington et al. 2017; Riehl and Strong 2018), whereas de-lifed fitness also

captures long-term survival. We found that older males recruited more offspring, likely by

virtue of older males siring more extra-pair offspring (Girndt et al. 2018). Likewise,

younger birds had lower annual de-lifed fitness, because younger birds had not recruited

any offspring in previous years that would contribute to their current de-lifed fitness.

Overall, de-lifed fitness better represents fitness as it is a relative measure of the

contribution to population growth (Alif et al. 2022). The number of recruits, while

intuitively appealing, is not relative, and in good years, more birds may have a higher

number of recruits, while in poor years, having one recruit may be an achievement. As

such, this measure is not always comparable between years, and this might explain the

discrepancy in our findings.

In conclusion, this study is the first to demonstrate a correlation between lifetime fitness

measures and sociality, and present evidence for stabilizing selection on sociality in a wild

population of passerine birds.

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Figure 1. Three versions of a simulated event (A, B and C) show the interval over which five individuals (1–5, black/grey bars)

spent at a resource over time (t), and the derived social networks from each approach: A = gambit of the group, which links all

individuals in a discrete group equally; B = time-window overlap (by Δ_t), which links individuals who overlap at a resource;

and C = arrival time (developed for this study), which links individuals who arrive together to a resource. Red bars denote the

time period during which individuals are considered to be associated: A, all individuals within a group; B, where they are

physically present at the same time (thick red bar), or shortly after they depart to account for birds which were present, but

not currently being recorded by the antenna, in that case, over-lapping by Δ_t (thin red bar, typically a few seconds); or, C,

where they arrive within Δ_t of each other, but the subsequent time spent at a feeder is irrelevant. However, note that the

function of Δ_t differs between B and C; Where in B, Δ_t functions to detect when birds are in the same place but where one (or

more) are not currently being detected by the antenna, in C the function is to link all individuals which arrive together, while

ignoring those already present at the resource, which has the potential to link two separate groups in A and B. In the case of C,

an additional interval (Δ_i) is required to define when birds have left the resource, after which they can be recorded as arriving

again.

Figure 2. De-lifed fitness as response variables against social measures from eight linear mixed models, at two scales, from the

Lundy Island house sparrows: Explanatory variables for Annual de-lifed fitness (A); and Lifetime de-lifed fitness (B), where [2]

denotes a quadratic function, also shown in the four adjacent panels for A and B, and their 95% credible intervals. Credible

intervals are given as solid bars for each explanatory variable, where a solid point denotes the posterior mode. Black bars

denote no effect on the response variable; red denote a positive and blue, a negative, relationship with the response. In

adjacent panels, quadratic functions of each response variable presented in A and B (on the Y axis: A Centrality, Degree, Opp.

Degree, Strength, and B Centrality, Degree, Opp. Degree, Strength). Blue curves represent a negative interaction with fitness

measures (on the X axis). Measures with no effect are not shown in figure. We found no link to sex, and age was also subject to

stabilizing selection, (given in supplementary files, table 4).