Genomic prediction with whole-genome sequence data in intensely selected pig lines

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Abstract

Background

Early simulations indicated that whole-genome sequence data (WGS) could improve

prediction accuracy and its persistence across generations and breeds. However,

results in real datasets have been ambiguous so far. Large data sets that capture most

of the genome diversity in a population must be assembled so that allele substitution

effects are estimated with higher accuracy. The objectives of this study were to use a

large pig dataset to assess the benefits of using WGS for genomic prediction

compared to using commercial marker arrays, to identify scenarios in which WGS

provides the largest advantage, and to identify potential pitfalls for its effective

implementation.

Methods

We sequenced 6,931 individuals from seven commercial pig lines with different

numerical size. Genotypes of 32.8 million variants were imputed for 396,100

individuals (17,224 to 104,661 per line). We used BayesR to perform genomic

prediction for 8 real traits and 9 simulated traits with different genetic architectures.

Genomic predictions were performed using either data from a marker array or variants

preselected from WGS based on linkage disequilibrium, functional annotation, or

association tests. Both single and multi-line training sets were explored.

Results

Using WGS improved prediction accuracy relative to the marker array, provided that

training sets were sufficiently large, especially for traits with high heritability and low

number of quantitative trait nucleotides. The performance of each set of predictor

variants was not robust across traits and lines. The most robust results were obtained

when preselected variants with statistically significant associations were added to the

marker array. Under this method, average improvements of prediction accuracy of 2.5

and 4.2 percentage points were observed in within-line and multi-line scenarios,

respectively, with training sets of around 80k individuals.

Conclusions

40 Our results evidenced the potential for WGS to improve genomic prediction accuracy

41 in intensely selected pig lines. Although the prediction accuracy improvements

42 achieved so far were modest at best, we would expect that more robust improvements

43 could be attained with a combination of larger training sets and optimised pipelines.

Introduction

Whole-genome sequence data (WGS) has the potential to empower the

identification of causal variants that underlie quantitative traits or diseases [1–4],

increase the precision and scope of population genetic studies [5,6], and enhance

livestock breeding. Genomic prediction has been successfully implemented in the

main livestock species and it has increased the rate of genetic gain [7]. Genomic

prediction has provided many benefits such as greater accuracies of genetic

evaluations and the reduction of the generational interval in dairy cattle. However,

since its early implementations, genomic prediction is typically performed using

marker arrays that capture the effects of the (usually unknown) causal variants via

linkage disequilibrium. Alternatively, WGS are assumed to contain the causal variants

themselves. For this reason, it was hypothesized that such data could further improve

prediction accuracy and its persistence across generations and breeds. Early

simulations indicated that causal mutations from WGS could increase prediction

accuracy [8–13]. One simulation study indicated that the magnitude of prediction

accuracy improvement relative to dense marker arrays ranged from 2.5 to 3.7%, with

a persistence of over 10 generations [11]. Another one reached improvements of 30%

if causal variants with low minor allele frequency could be captured by the WGS [9].

However, benefits could be on the lower end of that range in standard livestock

populations due to small effective population sizes and long-term negative selection

[10].

During the last few years, there have been several attempts at improving the

accuracy of genomic prediction with the use of WGS in the main livestock species.

Results have been ambiguous so far. When predicting genomic breeding values within

breed or line, some studies found no relevant improvement of prediction accuracy for

WGS compared to marker arrays [14–18]. Other studies found small, and often

unstable, improvements (e.g., from 1 to 5% or no improvement depending on

prediction method [19–21], or trait-dependent results [21,22]). When predicting

genomic breeding values across populations, the identification of causal variants from

WGS can improve prediction accuracy [23–26], especially for small populations

where initial prediction accuracy was low or that were not included in the training

population [23,25–28].

One of the most successful strategies to exploit WGS consists in augmenting

available marker arrays with preselected variants from WGS based on their

association with the trait of interest [29–32]. In some cases, this strategy improved

prediction accuracy by up to 9% [31] and 11% [32]. However, it did not improve

prediction accuracies in other within-line scenarios [16]. Nevertheless, this shows

how identifying causal variants could enhance genomic prediction with WGS. Whole-

genome sequence data has already been applied in genome-wide association studies

(GWAS) to identify variants associated to a variety of traits in livestock [2,33–35],

including pigs [36,37]. However, the fine-mapping of causal variants remains

challenging due to the pervasive long-range linkage disequilibrium across extremely

dense variation.

High accuracy in estimating allele substitution effects and, ideally, the

identification of causal variants amongst millions of other variants are important for

the usefulness of WGS in research and breeding. This requires large data sets able to

capture most of the genome diversity in a population. Despite that low-cost

sequencing strategies have been developed, which typically involve sequencing a

subset of the individuals in a population at low coverage and then imputing WGS for

the remaining individuals [38–40], the cost of generating accurate WGS at this scale,

as well as the large computational requirements for the analyses of such datasets, have

limited the population sizes or number of populations tested in some of the previous

studies. This hinders the interpretation of results across studies, which are very

diverse in population structures, sequencing strategies and prediction methodologies

used. The largest studies on the use of WGS for genomic prediction to date have been

performed in cattle, for which large multi-breed reference panels are available from

the 1000 Bull Genomes Project [2,19,33]. This has enabled the imputation of WGS

for cattle populations. The lack of such available reference panels has been cited as an

important limiting factor for performing similar studies in other species, such as pigs

[36].

We have previously described our approach to impute WGS in large pedigreed

populations without the need for haplotype phased reference panels [41]. Following

that strategy, we generated WGS for 396,100 pigs from seven intensely selected lines

with diverse genetic backgrounds and numerical size. The objectives of this study

were to use this large pig dataset to assess the benefits of using WGS for genomic

prediction compared to using commercial marker arrays, to identify scenarios in

which WGS provides the largest advantage, and to identify potential pitfalls for its

effective implementation.

Materials and Methods

Populations and sequencing strategy

We performed whole-genome re-sequencing of 6,931 individuals from seven

commercial pig lines (Genus PIC, Hendersonville, TN) with a total coverage of

approximately 27,243x. Sequencing effort in each of the seven lines was proportional

to population size. Approximately 1.5% (0.9 to 2.1% in each line) of the pigs in each

line were sequenced. Most pigs were sequenced at low coverage, with target coverage

of 1 or 2x, but a subset of pigs were sequenced at higher coverage of 5, 15, or 30x.

Thus, the average individual coverage was 3.9x, but the median coverage was 1.5x.

The number of pigs sequenced and at which coverage for each line is summarized in

Table 1.

The sequenced pigs and their coverage were selected following a three-part

sequencing strategy developed to represent the haplotype diversity in each line. First

(1), sires and dams with the highest number of genotyped progeny were sequenced at

2x and 1x, respectively. Sires were sequenced at a greater coverage because they

contributed with more progeny than dams. Then (2), the individuals with the greatest

genetic footprint on the population (i.e., those that carry more of the most common

haplotypes) and their immediate ancestors were sequenced at a coverage between 1x

and 30x (AlphaSeqOpt part 1; [42]). The sequencing coverage was allocated with an

algorithm that maximises the expected phasing accuracy of the common haplotypes

from the cumulated family information. Finally (3), pigs that carried haplotypes with

low cumulated coverage (below 10x) were sequenced at 1x (AlphaSeqOpt part 2;

[43]). Sets (2) and (3) were based on haplotypes inferred from marker array genotypes

(GGP-Porcine HD BeadChip; GeneSeek, Lincoln, NE), which were phased and

imputed using AlphaPhase [44] and AlphaImpute [45].

Most sequenced pigs and their relatives were also genotyped either at low

density (15k markers) using the GGP-Porcine LD BeadChip (GeneSeek) or at high

density (80k markers) using the GGP-Porcine HD BeadChip (GeneSeek). Quality

control of the marker array data was based on the individuals genotyped at high

density. Markers with minor allele frequency below 0.01, call rate below 0.80, or that

failed the Hardy-Weinberg equilibrium test were removed. After quality control,

38,634 to 43,966 markers remained in each line.

Sequencing and data processing

Tissue samples were collected from ear punches or tail clippings. Genomic

DNA was extracted using Qiagen DNeasy 96 Blood & Tissue kits (Qiagen Ltd.,

Mississauga, ON, Canada). Paired-end library preparation was conducted using the

TruSeq DNA PCR-free protocol (Illumina, San Diego, CA). Libraries for

resequencing at low coverage (1 to 5x) were produced with an average insert size of

350 bp and sequenced on a HiSeq 4000 instrument (Illumina). Libraries for

resequencing at high coverage (15 or 30x) were produced with an average insert size

of 550 bp and sequenced on a HiSeq X instrument (Illumina). All libraries were

sequenced at Edinburgh Genomics (Edinburgh Genomics, University of Edinburgh,

Edinburgh, UK).

DNA sequence reads were pre-processed using Trimmomatic [46] to remove

adapter sequences from the reads. The reads were then aligned to the reference

genome *Sscrofa11.1* (GenBank accession: GCA_000003025.6) using the BWA-MEM

algorithm   [47].   Duplicates   were   marked   with   Picard

(http://broadinstitute.github.io/picard). Single nucleotide polymorphisms (SNPs) and

short insertions and deletions (indels) were identified with the variant caller GATK

HaplotypeCaller (GATK 3.8.0) [48,49] using default settings. Variant discovery with

GATK HaplotypeCaller was performed separately for each individual and then a joint

variant set for all the individuals in each population was obtained by extracting the

variant positions from all the individuals.

We extracted the read counts supporting each allele directly from the aligned

reads stored in the BAM files using a pile-up function to avoid biases towards the

reference allele introduced by GATK when applied on low-coverage WGS [50]. That

pipeline uses the tool pysam (version 0.13.0; https://github.com/pysam-

developers/pysam), which is a wrapper around htslib and the samtools package [51].

We extracted the read counts for all biallelic variant positions, after filtering variants

in potential repetitive regions (defined as variants that had mean depth values 3 times

greater than the average realized coverage) with VCFtools [52]. This amounted to a

total of 55.6 million SNP (19.6 to 31.1 million within each line) and 10.2 million

indels (4.1 to 5.6 million within each line). A more complete description of the

variation across the lines is provided in [53].

Genotype imputation

Genotypes were jointly called, phased and imputed for a total of 483,353

pedigree-related individuals using the 'hybrid peeling' method implemented in

AlphaPeel [54,55]. This method used all the available marker array and WGS.

Imputation was performed separately for each line using complete multi-generational

pedigrees, which encompassed from 21,129 to 122,753 individuals each (Table 1).

We have previously published reports on the accuracy of imputation in the same

populations using this method [41]. The estimated average individual-wise dosage

correlation was 0.94 (median: 0.97). Individuals with low predicted imputation

accuracy were removed before further analyses. An individual was predicted to have

low imputation accuracy if itself or all of its grandparents were not genotyped with a

marker array or if it had a low degree of connectedness to the rest of the population.

These criteria were based on the analysis of simulated and real data on imputation

accuracy [41]. A total of 396,100 individuals remained, with each line comprising

between 17,224 and 104,661 individuals (Table 1). The expected average individual-

wise dosage correlation of the remaining individuals was 0.97 (median: 0.98)

according to our previous estimates. We also excluded from the analyses variants with

a minor allele frequency lower than 0.023, as their estimated variant-wise dosage

correlations was lower than 0.90 [41]. After imputation, 32.8 million variants (14.5 to

19.9 million within each line) remained for downstream analyses, out of which 9.9

million segregated across all seven lines.

Traits

We analysed data of 8 traits that are commonly included in selection

objectives of pig breeding programmes: average daily gain (ADG, g), backfat

thickness (BFT, mm), loin depth (LD, mm), average daily feed intake (ADFI, kg),

feed conversion ratio (FCR), total number of piglets born (TNB), litter weight at

weaning (LWW, kg), and return to oestrus 7 days after weaning (RET, binary trait).

Most pigs with records were born during the 2008–2020 period. Breeding values were

estimated by line with a linear mixed model that included polygenic and non-genetic

(as relevant for each trait) effects. Deregressed breeding values (dEBV) were obtained

following the method by VanRaden and Wiggans [56]. Only individuals in which the

trait was directly measured were retained for further analyses. The number of records

for each trait used in the analyses of each line is detailed in Table 2.

Simulated traits

To assist in the interpretation of results, we also created 9 simulated traits with

different numbers of quantitative trait nucleotides (QTN; 100, 1,000 or 10,000 QTN)

and heritability levels (h^2; 0.10, 0.25 or 0.50). Positions of the QTN were sampled

randomly amongst all variants called across all lines. Because QTN were sampled

from all variants, some QTN were fixed in some of the lines while segregating in

others. There were only negligible differences in the number of segregating QTN per

line (53 to 61, 531 to 583, or 5375 to 6058, respectively). Marker effects of the QTN

were sampled from a gamma distribution with shape=2 and scale=5. After a polygenic

term was calculated for each individual using these marker effects, residual terms

were sampled from a normal distribution with a variance parameter adjusted to

produce the desired heritability level. The number of records for the simulated traits is

detailed in Table 2. In these simulations, we used the imputed genotypes as real

genotypes and, therefore, implicitly cancelled any errors that might arise from the

processing of the sequencing reads and genotype imputation.

Training and testing sets

We split the individuals in each population into training and testing sets. The

testing sets were defined as those individuals from full-sib families from the last

generation of the pedigree (i.e., individuals that did not have any progeny of their

own). Only families with a minimum of 5 full-sibs were considered. The training set

was defined as all those individuals that had a pedigree coefficient of relationship

lower than 0.5 with any individual of the testing set. This design was chosen to mimic

a realistic situation in which breeding companies evaluate the selection candidates

available in the selection nucleus at any given time.

Genome-wide association study

234   To assess whether variants from the WGS could provide a finer mapping of

235   causal variants than marker array data, and to provide an association-based criterion

236   to preselect variants for the genomic prediction tests, we performed a GWAS for each

237   trait and line. This step included only the individuals in the training set. We fitted a

238   univariate linear mixed model that accounted for the genomic relationship matrix as:

239
$$\mathbf{y} = \mathbf{x}_i\beta_i + \mathbf{u} + \mathbf{e},$$

where $\mathbf{y}$ is the vector of dEBV, $\mathbf{x}_i$ is the vector of genotypes for the ith SNP coded as

0 and 2 if homozygous for either allele or 1 if heterozygous, β_i is the additive effect

of the ith SNP on the trait, $\mathbf{u} \sim N(0, \sigma_u^2 \mathbf{K})$ is the vector of polygenic effects with the

covariance matrix equal to the product of the polygenic additive variance σ_u^2 and a

genomic relationship matrix $\mathbf{K}$, and $\mathbf{e}$ is a vector of uncorrelated residuals. Due to

computational limitations, the genomic relationship matrix $\mathbf{K}$ was calculated using

only imputed SNP genotypes in the marker array regardless of whether the association

study involves the SNPs in the marker array or the variants in WGS. We used the

FastLMM software [57,58] to fit the model.

We used the same p-value threshold ($p<10^{-6}$) for both marker array and for

sequence associations, because while the WGS contains many more variants, they are

also expected to be in higher linkage disequilibrium. This threshold was based on

Bonferroni's multiple test correction assuming that the markers from the marker array

were independent. For the simulated traits, we defined genomic regions that contained

significant associations and assessed whether or not they contained a QTN. These

regions were defined by overlapping 500-kb segments centered on the significant

markers.

Genomic prediction in within-line scenarios

258    To test whether variants from the WGS could provide greater prediction

259 accuracy than the marker array, we tested genomic prediction using variants from the

260 marker array, from the WGS, or combining them. The marker array data (referred to

261 as 'Chip') was set as the benchmark for prediction accuracy. It contained all ~40k

262 variants in the marker array. For the sequence-based predictors, we preselected sets of

263 variants because currently available methods for genomic prediction are not yet

264 capable of handling datasets as large as the complete WGS. We tested different

265 alternative strategies for preselecting the predictor variants:

266 • *LDTags*. Tag variants retained after pruning based on linkage disequilibrium.

267   Variants were removed so that no pairs of SNPs with $r^2>0.1$ remained in any 10-

268   Mb window (windows slid by 2,000 variants) using Plink 1.9 [59]. The number

269   of predictor variants preselected by this method was on average of 30k variants

270   (range: 5k to 80k).

271 • *Top40k*. Variants preselected based on GWAS analyses. To mimic the number of

272   variants in Chip, we preselected the variants with the lowest p-value (not

273   necessarily below the significance threshold) in each of consecutive non-

274   overlapping 55-kb windows along the genome. In addition, to test the impact of

275   variant density on prediction accuracy, we preselected 10k, 25k, 75k, or 100k

276   predictor variants following the same criterion.

277 • *ChipPlusSign*. Variants preselected based on GWAS analyses as in Top40k, but

278   only significant variants ($p \leq 10^{-6}$) were preselected and merged with those in

279   Chip. When a 55-kb window contained more than one significant variant, only

280   that with the lowest p-value was selected as a proxy, in order to reduce the

preselection of multiple significant SNPs tagging the same causal variant. On

average, 309 significant variants were identified per trait and line (range: 23 to

1083; Table 3). These significant variants were merged with those in Chip.

• *Functional.* Variants that were annotated as loss-of-function or missense

according to Ensembl Variant Effect Predictor (Ensembl VEP; version 97, July

2019) [60]. The most severe predicted consequence type for each variant was

retrieved. The number of predictor variants preselected by this method was on

average of 35k variants (range: 27k to 40k).

• *Rand40k.* The same number of predictor variants as in Chip, chosen randomly.

Genomic prediction was performed by fitting a univariate model with BayesR

[61,62], with a mixture of normal distributions as the prior for variant effects,

including one distribution that sets the variant effects to zero. The model was:

$$\mathbf{y} = \mathbf{1}\mu + \mathbf{X}\boldsymbol{\beta} + \mathbf{e},$$

where $\mathbf{y}$ is the vector of dEBV, $\mathbf{1}$ is a vector of ones, μ is the general mean, $\mathbf{X}$ is a

matrix of genotypes, $\boldsymbol{\beta}$ is a vector of variant effects, and $\mathbf{e}$ is a vector of uncorrelated

residuals. The prior variance of the variant effects in $\boldsymbol{\beta}$ had four components with

variances $\sigma_1^2 = 0$, $\sigma_2^2 = 0.0001\sigma_g^2$, $\sigma_3^2 = 0.001\sigma_g^2$, or $\sigma_4^2 = 0.01\sigma_g^2$, where σ_g^2 is the

total genetic variance. We used a uniform and almost uninformative prior for the

mixture distribution. We used a publicly available implementation of BayesR

(https://github.com/syntheke/bayesR; accessed on 30 April 2021), with default

settings. Prediction accuracy was calculated in the testing set as the correlation

between the genomic estimated breeding value and the dEBV. Bias of the prediction

accuracy was calculated as the regression coefficient of the dEBV on the genomic

estimated breeding values.

It has been noted that using the same reference individuals for preselecting

variants through GWAS and for training the predictive equation can reduce prediction

accuracy and bias the predicted breeding values [16,63]. To account for that, we

reanalysed some of the scenarios after splitting the training set into two exclusive

subsets, one for GWAS to preselect the predictor variants and one for training the

predictive equation. The GWAS subset was defined by randomly selecting either 10%

or 50% of the individuals in the original training set. Those individuals were excluded

from the subset used for training the predictive equation afterwards.

Genomic prediction in multi-line scenarios

We considered multi-line scenarios in which the training set consisted of

merging the training sets that had been defined for each line. All analyses were

performed as for the within-line scenarios but with a line effect. In the multi-line

scenarios, all SNPs from the marker array that passed quality control and were

imputed for at least one line were included in the baseline (referred to as 'ML-Chip').

For ease of computation, the strategies for preselection of predictor variants from

WGS were applied only to the subset of 9.9 million variants that had been called and

imputed in all seven lines. Thus, we defined the predictor sets 'ML-Top40k' and

'ML-ChipPlusSign' by preselecting variants following the same criteria as in within-

line scenarios, but using a multi-line GWAS analyses with line effect instead. For

ML-ChipPlusSign, 60 to 7247 significant variants were identified per trait (Table 3)

and merged with those in ML-Chip. For comparison purposes, prediction accuracy

was calculated for the testing set of each individual line.

Results

Prediction accuracy within line

Whole-genome sequence data can improve prediction accuracy of marker

array data when there is a sufficiently large training set and if an appropriate set of

predictor variants is preselected. Figure 1 shows the prediction accuracy for the case

with the largest training set using different sets of predictor variants. In this case, all

tested sets of variants from the WGS, except for LDTags, yielded increases of

prediction accuracy that ranged from +2.0% to +9.2%. Using WGS also reduced bias

relative to Chip in some scenarios. However, the performance across predictor

variants set was not robust for the most part, and differed for each trait and line

(Additional File 1), often leading to no improvements of prediction accuracy or even

reduced prediction accuracy relative to Chip. One stable feature of the results was

LDTags showing a noticeable decrease in prediction accuracy in most traits and lines.

The size of the training set was one of the main factors that determined the

capacity of predictor variants from the WGS to improve the baseline prediction

accuracy of Chip. Figures 2 and 3 show the difference in prediction accuracy of

Top40k and ChipPlusSign with respect to the baseline of Chip against the number of

phenotypic records available in the training set. We observed large variability for the

difference in prediction accuracy, especially when the training set was small. This

variability was larger in Top40k than in ChipPlusSign, in a way that shrinkage of

variation as the training set was larger was more noticeable in ChipPlusSign. Gains in

prediction accuracy were low-to-moderate in the most favourable cases. In the most

unfavourable ones we observed large losses in prediction accuracy for Top40k but

more restrained losses for ChipPlusSign with moderate training set sizes. For both

sets of predictor variants, there was a positive trend that supported the need for large

training sets. This trend was clearer in ChipPlusSign than in Top40k, because of the

apparent lower robustness of the latter. Results for the other sets of predictor variants

are provided in Additional File 2.

The genetic architecture of the traits was also related to the success of WGS

for improving prediction accuracy. As the true genetic architecture of real complex

traits is mostly unknown, we used simulated traits to show that traits with high

heritability and low number of QTN were more likely to show larger improvements in

predictive performances. With Top40k (Figure 4), heritability seemed to be the main

factor that affected the expected improvement with large training sets (from null

improvements when $h^2=0.1$ to improvements of approximately 0.05 when $h^2=0.5$,

regardless of number of QTN, with a training set of 92k individuals). With

ChipPlusSign (Figure 5), the expected improvements with the same training set (92k

individuals) were not only greater in magnitude but depended on both heritability and

number of QTN (from null improvements when $h^2=0.1$ to improvements of

approximately 0.03 to 0.10 when $h^2=0.5$ with a number of 100 to 10k QTN,

respectively). Results confirmed the trends observed for the real traits (Figures 4 and

5); for instance, the higher robustness of ChipPlusSign compared to Top40k.

We observed diminishing returns when we increased the density of the

predictor variants. Increasing the number of predictor variants from the 40k in

Top40k to 75k selected in the same way yielded small improvements in prediction

accuracy compared to Top40k, but increases up to 100k variants provided smaller or

null additional gains (Additional File 3).

Splitting the original training set into two exclusive subsets, one for the

GWAS-based preselection of the variants and one for the training of the predictive

equation did not improve the prediction accuracy (Additional File 4). For

ChipPlusSign, this strategy reduced the bias but prediction accuracy decreased too,

probably because of the smaller subset available for training the predictive equation.

Prediction accuracy in multi-line scenarios

The performance of genomic predictors trained with multi-line datasets was

systematically lower than in the within-line scenarios (Additional File 5).

Nonetheless, the ML-ChipPlusSign predictor variants in general increased prediction

accuracy relative to ML-Chip (Figure 6). The increase in genomic prediction accuracy

for each line was largely dependent on the number of individuals of each line in the

training set. Therefore, the greatest improvements were achieved for the largest lines.

However, in the multi-line scenarios we observed increases of prediction accuracy for

some traits and lines for which no improvements were observed in the within-line

scenarios (Figure 7). In contrast, results for ML-Top40k were not robust (Additional

File 6).

Association tests

First, we assessed the performance of GWAS using the simulated traits. Table

4 shows the number of regions with significant associations that were detected using

either Chip or WGS, and whether they contained zero, one or multiple true QTN. The

WGS allowed the detection of a much larger proportion of true QTN than the Chip,

especially for the traits with high heritability and with large population sizes. The

most favourable scenarios for identifying regions that contained unequivocally a

single QTN with WGS were those in which the trait was controlled by a low number

of true QTN. However, even though the genetic architecture was very simple and

consisted of additive effects alone, the regions with significant associations only

captured a small fraction of the QTN that segregated within each line. Moreover,

using WGS also increased the number of regions with significant associations that

contained no QTN, which could therefore be considered as false positives. Some of

the selected regions contained multiple QTN, which could indicate either a 'hit by

chance' or an inability to disentangle multiple causal variants. While false positives

also occur with Chip, their incidence was more severe with the WGS, especially for

traits with a large number of QTN. Large population sizes further aggravated the

inflation of genome-wide p-values.

Despite this, with the real traits we found that GWAS using WGS can

contribute to a better understanding of the genetic mechanisms that underlie the traits

of interest. To illustrate this, we examined the GWAS results for BFT in line A, for

which a large number of phenotypic records were available. Figure 8 shows the

results for chromosome 1 as an example, while Additional File 7 shows the results for

six genomic regions of interest. The main genomic regions and candidate genes

associated to BFT detected with Chip in the same genetic lines studied here were

reported elsewhere [64]. We will use the candidate genes reported there to refer to the

genomic regions with significant associations. Using Chip, we identified 6 genomic

regions ($p<10^{-6}$). Using WGS (with a more stringent significance threshold of $p<10^{-9}$

to focus on the most significant associations), we confirmed 3 of these genomic

regions that co-located to candidate genes *MC4R*, *DOLK*, and *DGKI* or *PTN*.

However, the most associated variants in each of these genomic regions located

outside the coding region of these putative causal genes. These signals sometimes had

very strong evidence of association for some variants that were relatively distant from

our candidate functional gene, which could cast doubts about the fine-mapping of the

causal mutation. The region at SSC18, 9–13 Mb, contained two candidate genes

DGKI and *PTN,* but the WGS revealed significantly associated variants within *DGKI*

and none within *PTN*, despite that the strongest associations were away from both

genes at 10.5-11 Mb. Using the WGS we also detected 24 additional genomic regions

that contained candidate genes such as *CYB5R4, IGF2,* and *LEPR*. These genes were

previously detected in other lines using the Chip but not in this one [64], sometimes

because there were no markers for the associated region in Chip (SSC2, 0–4 Mb). The

region at SSC1, 52.5–53.5 Mb, showed many significant variants that encompassed

not only the previously identified candidate gene *CYB5R4*, but also *MRAP2*

(annotated with functions on feeding behavior and energy homeostasis). In contrast,

candidate gene *LEPR* was located within the region at SSC6, 146.5–147.0 Mb, where

many significant variants were located, although the most significant variants were

not in the coding regions of the gene. Using the WGS we also identified additional

candidate genes that had not been previously detected in any of the lines, such as

CYP24A1 (annotated with functions on fatty acid omega-oxidation and vitamin D

metabolism; not shown). For many of the other genomic regions, it was difficult to

pinpoint a candidate gene with the available information or there were no annotated

genes.

Discussion

Our results evidenced the potential for WGS to improve genomic prediction

accuracy in intensely selected pig lines, provided that the training sets are large

enough. Improvements achieved so far were modest at best. On one hand, these

modest improvements indicated that the strategies that we tested were likely

suboptimal. On the other hand, the positive trend for the largest training sets indicated

that we might have not reached the critical mass of data that is needed to leverage the

potential of WGS, especially in scenarios where genomic prediction with marker

arrays is already yielding high accuracy. The results from several traits and lines with

different training set sizes and the use of simulated phenotypes allowed us to identify

the most favourable scenarios for genomic prediction with WGS. We will discuss (1)

the prediction accuracy that we achieved with WGS compared to commercial marker

array data and the scenarios in which WGS may become beneficial, and (2) the

potential pitfalls for its effective implementation and the need for an optimised

strategy.

Prediction accuracy with whole-genome sequence data

We compared the genomic prediction accuracy of the current marker array

(Chip) with sets of preselected sequence variants in a way that the number of variants

remained similar across sets. Improvements of prediction accuracy can be limited if

current marker arrays are already sufficiently dense to capture a large proportion of

the genetic variance in intensely selected livestock populations. These populations

typically have small effective population size [10,19]. Nevertheless, modest

improvements have been achieved under certain scenarios. In our study, the most

robust results were obtained for the ChipPlusSign set, where variants that showed

statistically significant associations to the trait were preselected and added to the

information from the marker array. This is consistent with previous reports that

showed an improvement of prediction accuracy under similar approaches [29–32].

We added 23 to 1083 significant variants to those in Chip in different scenarios. In the

most successful ones, at least around 200 significant variants were added and average

improvements of prediction accuracy of 2.5 percentage points were observed with

training sets of around 80k individuals. In other instances, however, additions of a

larger number of variants have been proposed. The addition of 1623 variants

(preselected as the combination of 3-5 variants for each of the top QTL per trait and

breed) to a 50k array increased prediction reliability (accuracy squared) by up to 5

percentage points in Nordic cattle [29]. Adding the 16k SNPs with largest estimated

effects to a 60k array increased prediction reliability on average by 2.7 (up to 4.8)

percentage points in Holstein cattle [30]. For the custom 50k array for Hanwoo cattle,

it has been reported that adding at least around 12k SNPs (3k for each of four traits)

improved prediction accuracy by up to ~6 percentage points [32]. The addition of

~400 variants preselected by GWAS with regional heritability mapping to a 50k array

increased prediction accuracy by 9 percentage points in sheep [31]. In other cases in

Nordic cattle, however, the addition of ~1500 variants preselected by GWAS to a 54k

panel produced negligible improvements in the prediction of traits with low

heritability [65].

Preselecting an entirely new set of predictor variants from WGS, as in

Top40k, proved more challenging than ChipPlusSign. In Top40k, we preselected the

variants with the lowest p-value in each of consecutive non-overlapping 55-kb

windows along the genome. This strategy did not perform much differently from just

taking random variants from these windows, as in Rand40k. One possible reason for

these results is that at this variant density, random variants effectively tag the same

associations as Top40k thanks to linkage disequilibrium. Denser sets of predictor

variants provided only small further improvements of prediction accuracy with

diminishing returns.

The modest performance of ChipPlusSign and Top40k could also be a

consequence of the difficulty for fine-mapping causal variants through GWAS with

WGS. Theoretically, the identification of all causal variants associated with a trait

should enable the improvement of prediction accuracy [12]. Even though WGS allows

the detection of a very large number of associations, problems such as false positives

or p-value inflation also become more severe in a way that added noise might offset

the detected signal. For instance, results in cattle showed that GWAS with WGS did

not detect clearer associated regions relative to marker arrays and failed to capture

QTL for genomic prediction [14], as the effect of potential QTL were spread across

multiple variants. Therefore, WGS performed better with simple genetic architectures

(i.e., traits with low number of QTN). This is consistent with expectations and

simulation results [8] that indicated that the benefit of using WGS for genomic

prediction would be limited by the number and size of QTN. When there are many

QTN with small effects it becomes much more difficult to properly estimate their

effects accurately. Therefore, for largely polygenic traits (as most traits of interest in

livestock production), training sets need to be very large before WGS can increase

prediction accuracy [8].

The advantage of using WGS might be limited by the current training set

sizes, especially in scenarios where marker arrays are already yielding high prediction

accuracy [14,20]. Multi-line training sets could be particularly beneficial with the use

of WGS because they allow a larger training set with low pairwise relationship degree

among individuals. Previous simulations suggested that WGS might be the most

beneficial with multi-breed reference panels [66], especially for numerically small

populations. Our results with a multi-line training set indicated that WGS can improve

prediction accuracy in scenarios that are less optimised than within-line genomic

prediction. The average improvements of prediction accuracy of 4.2 percentage points

were observed for the populations that contributed around 80k individuals to the

training set. However, in general those predictions were still less accurate than using

variants preselected under within-line training sets. In our multi-line scenarios we

only used variation that segregated across all seven lines. We observed that

population-specific variation accounted only for small fractions of genetic variance

[53] and it seems unlikely that they would contribute much to prediction accuracy

across breeds. Another possible obstacle is the differences in the allele substitution

effects of the causal mutations across breeds. This can be caused by differences in

allele frequency, contributions of non-additive effects and different genetic

backgrounds, or even gene-by-environment interactions among others [24,67].

We observed low robustness of genomic prediction with WGS across traits

and lines, and drops in prediction accuracy in those scenarios where genomic

prediction with WGS failed. Regarding bias, we did not observe a systematic increase

for ChipPlusSign despite using the same individuals for variant preselection and for

training the predictors [16,63]. When we split the training set into two subsets, one for

GWAS-based variant preselection and the other for training of the predictive

equations, we did not observe any improvement in accuracy or bias. One hypothesis is

that both subsets belong to the same population and therefore retained similar inter-

relationship degrees (i.e., they are not strictly independent sets of individuals).

Moreover, the reduction in individuals available for training the predictors negatively

affected prediction accuracy.

We did not directly test persistence of prediction accuracy, but previous

studies with real data found no higher persistence of prediction accuracy for WGS,

not even with low degree of relationship between training and testing sets [14]. We

would expect such obstacles to persistence of accuracy until causal variants can be

successfully identified.

Suboptimal strategy and pitfalls

The use of WGS for genomic prediction can only be reached after many other

steps are completed to produce the genotypes at whole-genome level. Each of these

steps has its pitfalls. It is unavoidable that the success of using of WGS is sensitive

not only to the prediction methodology itself but also to the strategy followed until

genotyping. This strategy includes the choice of which individuals to sequence, the

bioinformatics pipeline to call variants, the imputation of the WGS and choice of

variant filters. When combined with the multiplicity of prediction methods and the

preselection of predictor variants (which is unavoidable with current datasets,

predictive methodologies and computational capacities), there are many options and

variables in the whole process that can affect the final result and that are not yet well

understood. Therefore, a much greater effort for optimising such strategies is required.

Here we tested relatively simple approaches to see how they performed with large

WGS datasets. We have discussed what in our opinion are the main pitfalls of our

approach for selection of the individuals to sequence [55] and the biases that may

appear during processing of sequencing reads [50] elsewhere, and therefore here we

will focus on imputation of WGS and its use for genomic prediction.

Imputation accuracy

It is widely recognized that imputation from marker arrays to WGS from very

few sequenced individuals can introduce genotyping errors and that genotype

uncertainty can be high [17,21,68,69]. The accuracy of the imputed WGS is one of the

main factors that may limit its performance for genomic prediction. In a simulation

study, van den Berg et al. [17] quantified the impact of imputation errors on

prediction accuracy and showed that prediction accuracy decreases as errors

accumulate, especially in the testing set.

Imputation of WGS is particularly challenging because typically we have to

impute a very large number of variants for a very large number of individuals from

very few sequenced individuals. We assessed the imputation accuracy of our approach

elsewhere [41,55] and recommended that ~2% of the population should be sequenced.

In our study, line D was the line where prediction accuracy with Top40k performed

the worst, mostly performing below Chip predictors. In this line, only 0.9% of the

individuals in the population had been sequenced and therefore lower imputation

accuracy could be expected. Although there was not enough evidence for establishing

a link between these two features (sequencing effort and prediction accuracy), we

recommend cautious design of a sequencing strategy that is suited to the intended

imputation method [55].

Prediction accuracy could be improved by accounting for genotype uncertainty

of the imputed WGS. For that, it could be advantageous to use allele dosages rather

than best-guess genotypes [69], although most current implementations cannot handle

such information.

Preselection of predictor variants

Simply using WGS to increase the number of markers does not improve

prediction accuracy [18,21,24]. Due to the large dimensionality of WGS, there is a

need to remove uninformative variants [24,31,66,68,70]. Predictor variants must be

causal or at least informative of the causal variants, which depends on the distance

between the markers and the causal variants [13]. For this reason, variants that are in

weak linkage disequilibrium with causal mutations have a 'dilution' effect, i.e., they

add noise and limit prediction accuracy [24,31,70]. However, if too stringent filters

are applied during preselection of predictor variants, there is a risk of removing true

causal variants, and that would debilitate persistence of accuracy across generations

and across populations [66,71]. For instance, the impact of removing predictor

variants with low minor allele frequency can vary depending on the minor allele

frequency of the causal variants as well as the distance between predictor and causal

variants [13]. Losing causal or informative variants would negatively affect multi-line

or multi-breed prediction.

A popular strategy to preselect the predictor variants is based on association

tests. Genome-wide association studies on WGS are expected to confirm associations

that were already detected with marker arrays and identify novel associations (e.g.,

[36,72]). However, preliminary inspection of our GWAS results for the real traits

showed that the added noise could easily offset the added information and fine-

mapping remains challenging. Multi-breed GWAS [4] and meta-analyses [73] are

suitable alternatives for GWAS to accommodate much larger population sizes and for

combining results of populations with diverse genetic backgrounds. Multi-breed

GWAS can be more efficient to identify informative variants than single-breed

GWAS, which may benefit even prediction within lines [74]. Because the signal of

some variants may go undetected for some traits but not for other correlated traits,

combining GWAS information of several traits can also help identifying weak or

moderate associations [25]. We did not test whether combining the significant

markers from the different single-trait GWAS yielded greater improvements in

prediction accuracy [29,32]. Multi-trait GWAS models could be more suited for that

purpose [72,75]. To improve fine-mapping, other GWAS models that incorporate

biological information have been proposed (e.g., functional annotation [76] or

metabolomics [77]).

There have been other suggested methods that may improve variant

preselection for genomic prediction. VanRaden et al. [30] suggested that preselecting

variants based on the genetic variance that they contribute rather than the significance

of the association could be more advantageous, as the former would indirectly

preselect variance with higher minor allele frequency. Other authors proposed

preselection of variants using statistics that do not depend on GWAS, such as the

fixation index (F_{ST}) score between groups of individuals with high and low phenotype

values [70], as an alternative to avoid the negative impact of spurious associations.

Preselecting predictor variants based on functional annotation was not useful,

as it reduced prediction accuracy in several traits and lines. Previous studies showed

that subsets of variants based on functionality either did not improve or reduced

prediction accuracy [20] and that adding preselected variants from coding regions to

marker arrays produced lower prediction accuracy than just adding the same number

of variants without considering functional classification [32]. A plausible explanation

is that functional variants are enriched for lower minor allele frequency, which can be

less informative for prediction [13]. Furthermore, functional annotation does not

necessarily capture true effects, and the method we used is biased towards protein-

coding variants, which may lead to an underrepresentation of functional non-coding

variants that may explain a large fraction of quantitative trait variance. Xiang et al.

[78] found that expression QTL and non-coding variants explained more variation in

quantitative traits in cattle than protein-coding functional variants. When functional

annotation is not considered, intergenic variants are more likely to be preselected by

chance. Such variants tend to be more common and widespread across populations,

and therefore can act as tag variants and capture much larger fractions of trait variance

[53].

Another popular strategy to reduce the number of variants is to prune variants

based on linkage disequilibrium (LDTags). This strategy performed very poorly in our

populations. Other studies reported different outcomes, where pruning for $r^2 > 0.9$

provided positive results [18,21]. It is possible that this was in part due to the stringent

threshold ($r^2 > 0.1$) that we used in order to retain only a small number of variants.

New models and methods

It is also likely that models, methods, and their implementations need to be

improved to handle the complexity of WGS and to efficiently estimate marker effects

of so many variants with high accuracy, among other features. This is a very active

area of research and multiple novel methodologies have been proposed over the last

years. Some examples are a combination of subsampling and Gibbs sampling [79],

and a model that simultaneously fits a GBLUP term for a polygenic effect and a

BayesC term for variants with large effects selected by the model (BayesGC) [26].

Testing alternative models and methods for genomic prediction was out of the scope

of this report. However, together with refinements in the preselection of predictor

variants, it remains an interesting avenue for further optimisation of the analysis

pipeline.

Some of the most promising methods are designed to incorporate prior

biological information into the models. One of such methods is BayesRC [23], which

extends BayesR by assigning flatter prior distributions to classes of variants that are

more likely to be causal [19,22]. Similarly, GFBLUP [80] could be used to

incorporate prior biological information from either QTL databases or GWAS as

genomic features [21,35,68]. The model MBMG [27], which fits two genomic

relationship matrices according to prior biological information, has also been

proposed for multi-breed scenarios to improve genomic prediction in small

populations. Haplotype-based prediction methods could provide greater prediction

accuracy with WGS than SNP-based methods in pigs [81] and cattle [82]. These

methods reduce the number of model dimensions. However, the uptake of such

methods has been limited so far due to their greater complexity, for example, to define

haplotype blocks.

Conclusion

Our results evidenced the potential for WGS to improve genomic prediction

accuracy in intensely selected pig lines. The performance of each set of predictor

variants was not robust across traits and lines and the improvements that we achieved

so far were modest at best. The most robust results were obtained when variants that

showed statistically significant associations to the trait were preselected and added to

the marker array. With this method, average improvements of prediction accuracy of

2.5 and 4.2 percentage points were observed in within-line and multi-line scenarios,

respectively, with training sets of around 80k individuals. We would expect that a

combination of larger training sets and improved pipelines could help achieve greater

improvements of prediction accuracy. The robustness of the whole strategy for

generating WGS at the population level must be carefully stress-tested and further

optimised.

Ethics approval and consent to participate

690 The samples used in this study were derived from the routine breeding activities of
691 PIC.

Consent for publication

692 Not applicable.

Availability of data and material

693 The software packages AlphaSeqOpt, AlphaPhase, AlphaImpute and AlphaPeel are
694 available from the AlphaGenes website (http://www.alphagenes.roslin.ed.ac.uk). The
695 datasets generated and analysed in this study are derived from the PIC breeding
696 programme and not publicly available.

Competing interests

697 The authors declare that they have no competing interests. BDV, CYC, and WOH are
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Authors' contributions

704 RRF, GG and JMH designed the study; CYC assisted in preparing the datasets; RRF,
705 AW and MJ performed the analyses; RRF wrote the first draft; AW, CYC, BDV,

706 WHO, GG and JMH assisted in the interpretation of the results and provided

707 comments on the manuscript. All authors read and approved the final manuscript.

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Based Haplotypes Inferred From High-Density Chip and Imputed Sequence Variants
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Figures

961
962 **Figure 1.** Prediction accuracy for each set of predictor variants for trait ADG in line

963 A. Left: Correlation (left). Dashed line at value of Chip as a reference. Values indicate

964 relative difference to Chip. Right: Bias. Dashed line at the ideal value.

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Figure 2. Genomic prediction accuracy with the Top40k predictor variants for the

real traits. The difference of prediction accuracy between Top40k and Chip is shown,

for all traits and lines (left) or by trait (right). Red dashed line at 'no difference'.

Figure 3. Genomic prediction accuracy with the ChipPlusSign predictor variants for

the real traits. The difference of prediction accuracy between ChipPlusSign and Chip

is shown, for all traits and lines (left) or by trait (right). Red dashed line at 'no

difference'.

979
980 **Figure 4.** Genomic prediction accuracy with the Top40k predictor variants for the

981 simulated traits. The difference of prediction accuracy between Top40k and Chip is

982 shown by heritability (h^2) and number of quantitative trait nucleotides (nQTN) of the

983 simulated traits. Red dashed line at 'no difference'.

986
987 **Figure 5.** Genomic prediction accuracy with the ChipPlusSign predictor variants for

988 the simulated traits. The difference of prediction accuracy between ChipPlusSign and

989 Chip is shown by heritability (h^2) and number of quantitative trait nucleotides (nQTN)

990 of the simulated traits. Red dashed line at 'no difference'.

993
994 **Figure 6.** Genomic prediction accuracy with the ML-ChipPlusSign predictor variants

995 for the real traits. The difference of prediction accuracy between ML-ChipPlusSign

996 and ML-Chip is shown, for all traits and lines (left) or by trait (right). Red dashed line

997 at 'no difference'.

1000
1001 **Figure 7.** Comparison of the difference in genomic prediction accuracy in the multi-

1002 line scenarios (between ML-ChipPlusSign and ML-Chip) and in the within-line

1003 scenarios (between ChipPlusSign and Chip). Red dashed line at 'no difference'. Blue

1004 dashed line is the bisector.

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Figure 8. Genome-wide association study results for trait BFT in line A. Only chromosome 1 is displayed as an example. In red, results for the variants in the marker array (Chip); in black, results for the whole-genome sequence data (WGS). The blue dashed line indicates significance threshold with Bonferroni's multiple test correction assuming that the markers from the marker arrays were independent (p-value $\leq 10^{-6}$).

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Tables

1015 **Table 1.** Number of sequenced pigs and pigs with imputed data.

Line	Individuals sequenced	Individuals sequenced by coverage				Individuals used in analyses	
		1x	2x	5x	15–30x	Pedigree	Imputed
A	1,856	1,044	649	73	90	122,753	104,661
B	1,366	685	545	44	92	88,964	76,230
C	1,491	628	728	54	81	84,420	66,608
D	731	362	311	16	42	79,981	60,474
E	760	394	274	27	65	50,797	41,573
F	381	193	137	16	35	35,309	29,330
G	445	217	176	15	37	21,129	17,224

1016

1017   **Table 2.** Number of phenotypic records per trait and line.

Trait	A	B	C	D	E	F	G
ADG	88,342	64,285	56,173	51,061	35,423	26,335	15,452
BFT	80,146	62,027	55,233	47,509	34,527	23,872	15,268
LD	85,233	64,141	56,026	48,509	35,495	26,453	15,274
ADFI	21,960	9,525	9,062	12,256	12,444	4,105*	4,851
FCR	21,200	9,217	8,654	12,044	12,316	4,016*	4,754
TNB	13,581	10,721	9,626	7,729*	6,506*	-	3,230*
LWW	-	9,112	7,251	-	-	-	2,813*
RET	-	6,978	6,327	-	-	-	1,669*
Simulated	104,661	76,230	66,608	60,474	41,573	29,330	17,224

1018   *ADG* average daily gain, *BFT* backfat thickness, *LD* loin depth, *ADFI* average daily

1019   feed intake, *FCR* feed conversion ratio, *TNB* total number of piglets born, *LWW* litter

1020   weight at weaning, *RET* return to oestrus 7 days after weaning.

1021   *Included in multi-line scenarios, but excluded in within-line scenarios because of the

1022   limited size of the testing set.

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1024 **Table 3.** Number of significant variants from the whole-genome sequence data that

1025 were added to the marker array in ChipPlusSign.

Trait	A	B	C	D	E	F	G	Multi-line
ADG	646	581	424	498	279	219	143	4731
BFT	1083	758	664	518	1030	218	237	6149
LD	633	579	458	518	222	215	43	7247
ADFI	145	224	169	23	183	-	119	767
FCR	198	224	162	95	56	-	134	1369
TNB	71	117	161	-	-	-	-	248
LWW	-	32	73	-	-	-	-	480
RET	-	184	31	-	-	-	-	60

1026 *ADG* average daily gain, *BFT* backfat thickness, *LD* loin depth, *ADFI* average daily

1027 feed intake, *FCR* feed conversion ratio, *TNB* total number of piglets born, *LWW* litter

1028 weight at weaning, *RET* return to oestrus 7 days after weaning.

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1030 **Table 4.** Number of significantly associated genomic regions in the genome-wide

1031 association study for the simulated phenotypes that contained 0, 1 or 2 or more

1032 quantitative trait nucleotides (QTN).

h^2	nQTN	Line size	Chip		Whole-genome sequence		
			0 QTN	1 QTN	0 QTN	1 QTN	≥2 QTN
0.10	100	27k	4	1	8	6	0
		56k	11	3	19	19	0
		92k	10	7	44	19	0
	1k	27k	1	0	4	0	1
		56k	1	0	16	3	1
		92k	1	0	283	9	0
	10k	27k	1	0	1	0	0
		56k	0	0	16	2	1
		92k	2	0	186	17	12
0.25	100	27k	11	6	26	15	1
		56k	22	8	44	28	3
		92k	20	7	90	34	1
	1k	27k	0	0	8	1	3
		56k	3	0	34	15	6
		92k	6	0	692	49	16
	10k	27k	0	0	2	0	0
		56k	0	0	90	9	22
		92k	4	0	564	56	164
0.50	100	27k	18	9	24	24	1
		56k	30	13	116	41	3
		92k	17	9	425	44	1
	1k	27k	6	0	22	9	6
		56k	5	1	238	59	32
		92k	11	1	903	169	120
	10k	27k	0	0	4	0	0
		56k	0	0	360	77	172
		92k	10	0	379	116	508

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