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What Happens After an Error?

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Author Contributions: TA, HZ, and TL conceptualized the studies, designed the experiments, and wrote the paper. TA and HZ programmed the tasks, collected the data, and analyzed the data.

Keywords: adaptive control, cognitive control, errors, post-error slowing, action slips

Abstract:

A ubiquitous finding in cognitive science across a wide variety of tasks is that humans tend to slow down after making an error. The dominant account of this post error slowing is that people engage in adaptive control and are simply more cautious after an error. However, this explanation is challenged by the fact that, although people are slower, accuracy typically does not improve following an error. Errors negatively impact cognitive processing, but characterizing the nature of error-based impairments has been a challenge in standard paradigms. Here, we adopt a recently developed experimental approach to uncover the time course of stimulus-response processing following an error by exerting tight control over the timing of responses. This method allows us to apply a computational model of response preparation that allows us to estimate the latency of cognitive processing underlying responses. In a four alternative forced-choice task with arbitrary stimulus-response mappings, we find that human participants are less accurate after an error even when given up to two seconds to make a response. Our modeling results ruled out the possibility that errors lead to a subsequent slowing of the cognitive processing underlying responses. Instead, we found that the "efficacy" of cognitive processing in producing an intended response is impaired following errors as people commit more perseverative slips of action regardless of when a response is made. These results suggest that prior observations of post-error slowing may be an adaptive response to impaired cognitive processing rather than a strategic shift in the speed-accuracy tradeoff.

Significance Statement:

What happens after we make a mistake? It has long been established that human behavior changes after committing an error, but it has been surprisingly hard to establish how our mental processing is affected. By forcing people to respond at predetermined times, we uncover the time-course of how we go from a stimulus to a response. We find that errors do not affect the timing or the variability of stimulus-response processing. Instead, errors lead people to be more likely to immediately slip up and repeat their past mistakes, even when given ample time to recover. Our results that even though people may slow down to take their time to get it right, they are fundamentally less effective after making an error.

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We're taught from a young age that we all make mistakes: "It's how you bounce back that matters!" But how do people respond after an error? You can try to take your time to get it right, but do we bounce back stronger or are there lingering costs when you know you got it wrong?

A longstanding finding in the field of psychology is that people respond more slowly immediately following errors in decision-making (P. M. Rabbitt, 1966; P. Rabbitt & Rodgers, 1977). This phenomenon is often referred to as "post-error slowing" (Fairweather, 1978). Post-error slowing sometimes coincides with increased response accuracy and has thus been widely assumed to reflect an adaptive strategic adjustment to prevent future errors. The dominant account of this phenomenon suggests that people are more cautious in responding after they make an error. Neural network models explain these post-error effects in terms of a decrease in baseline activation of a response (Botvinick et al., 2001) and evidence accumulation models explain these post-error effects in terms of an increase in the threshold of evidence required to make a response(Dutilh et al., 2012). This would all predict that we should be more accurate after making an error as we shift along the same speed-accuracy tradeoff curve. However, accuracy is

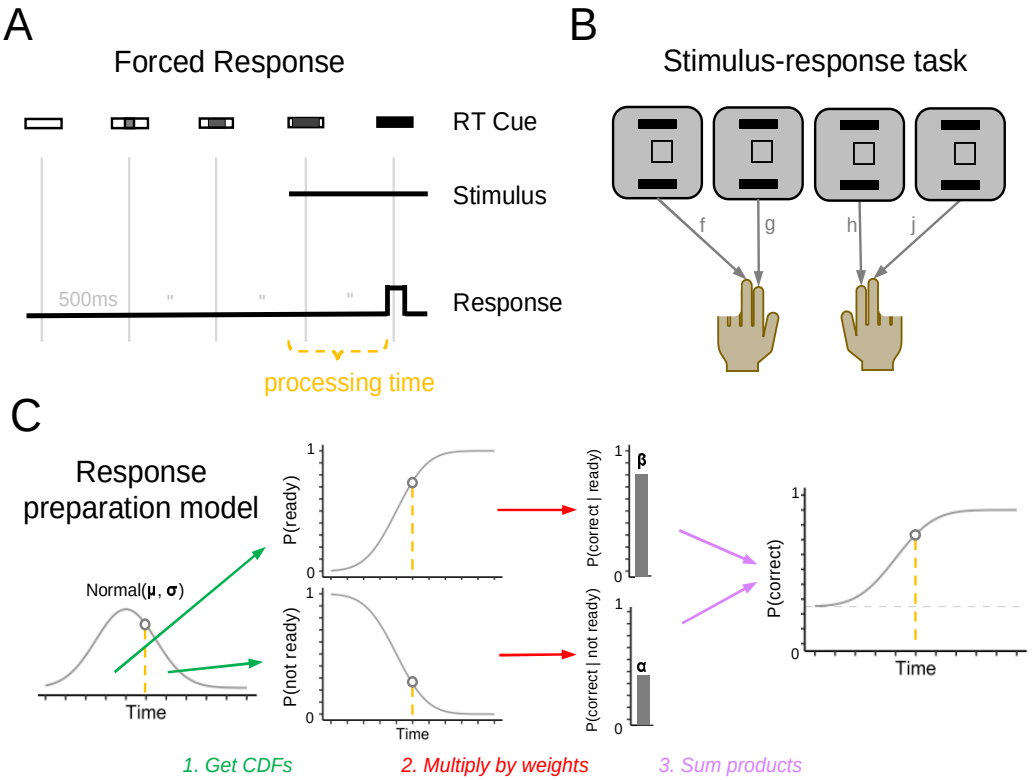

Figure 1. Forced response SR task. Participants performed a stimulus-response task (4AFC) with forced responding. **A.** Participants were trained and cued to respond when an empty rectangle was completely filled (at 2000 ms) and the target stimulus appeared at a random time between 0 and 2000 ms. **B.** Participants were instructed to press the 'f' key with their left middle finger, the 'g' key with their left index finger, the 'h' key with their right index finger, or the 'j' key with the left middle finger, depending on the symbol (Exp. 1 & 2). **C.** Response preparation model used to predict participants responses as a function of the time available for processing (orange dashed line). The model assumes that time at which participant has processed the stimulus and prepared the response is normally distributed. The probability that a response is prepared (or not) at a given time is determined by the cumulative distribution function (CDF) of the Normal(μ, σ). These probabilities are multiplied by weights representing the probability of expressing the correct response given that it is prepared (β) or not (α). Summing these products gives the probability of expressing correct response at a given time.

quite often stable or even reduced after an error(Notebaert et al., 2009; P. Rabbitt & Rodgers, 1977). This has led some researchers to instead conclude that post-error slowing is a *maladaptive* response reflecting impaired processing rather than a cognitive control adjustment aimed at improving behavior. Although some researchers have attempted to reconcile adaptive and maladaptive accounts of post-error

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slowing (Danielmeier & Ullsperger, 2011; Purcell & Kiani, 2016; Wessel & Aron, 2017), characterizing the nature of impairments in cognitive processing underlying post-error responses has remained a challenge in standard paradigms. It is currently unknown exactly *why* accuracy is often so poor following errors given how much slower people are to respond.

Post-error effects are almost exclusively examined in tasks that measure free response times (RT) and error rates that are analyzed separately. A critical issue with free RTs is that they confound the cognitive processing necessary for response preparation (e.g., stimulus identification and action selection) with response initiation (i.e. emitting the motor response). Indeed, recent work has shown that responses are accurately prepared and ready to deploy much more quickly than free RT would indicate (Haith et al., 2016). This work argues that preparation and response initiation are independent motor control parameters (Haith et al., 2016; Hardwick et al., 2019). On this view, people not only decide *what* response to make, but also *when* to make it. Prior studies on post-error effects do not directly distinguish between selection and initiation in their experimental designs or in their theoretical models. This presents a problem in that slower RTs after errors could be due to strategic delays in response initiation without any change in the cognitive processes underlying response preparation per se. Although researchers have used mathematical models such as the drift diffusion model to attempt to tease apart response caution from maladaptive cognitive processing following an error (e.g., Dutilh et al., 2012; Purcell & Kiani, 2016), this approach fundamentally relies on fitting RT distributions with the assumption that RT is a reliable indicator of the total duration of cognitive processing necessary to produce a response[1]. Additionally, the drift diffusion framework makes fairly strong assumptions about the nature of the decision-making process (a single evidence accumulation process, a static evidence accumulation rate per condition, etc.).

To address these shortcomings, we examined post-error effects using a forced-response paradigm which controls response time and treats processing time as an independent variable (Haith et al., 2016; Hardwick et al., 2019). In this paradigm (Figure 1a), participants are cued to respond at the same time on each trial while the onset of the target stimulus is uniformly varied in a 4-alternative forced choice stimulus-response task (Figure 1b). This allows us to examine post-error effects on processing per se by controlling the time of response initiation to query the state of cognitive processing as time unfolds. We use these data to fit a model that makes minimal assumptions about the cognitive processes leading to a response: 1) each stimulus leads to the preparation of the appropriate response with some mean latency and normally distributed trial-to-trial variability; 2) if stimulus-based response preparation is not yet complete, participants will guess randomly; and 3) the ability of cognitive processing to produce a desired or intended response isn't perfect, and slips of action sometimes occur (Norman, 1981). These assumptions lead directly to three free parameters in this model: μ, the average speed of cognitive processing underlying a correct response; σ, the standard deviation (i.e. trial-to-trial noise) of the speed of cognitive processing; and β, the probability a correct response will be produced when this cognitive processing is complete (i.e. the "efficacy" of processing). Note that $1-\beta$ is the probability that an action slip occurs even if the correct response has been prepared. When combined, these parameters can be used to predict accuracy when the amount of time given for stimulus-response processing is known (Figure 1c; See Methods for complete modelling details). In combination with controlling the time of response initiation, this modeling framework allows us to distinguish among several distinct ways in which cognitive processing might be affected following an error. Following an error, stimulus-response processing might be slower (Figure 2a), more variable (Figure 2b), or simply less effective at producing the correct response resulting in more frequent slips of action (Figure 2c). Each of these possibilities would produce both a different pattern of behavioral results and a quantitative change in one of the parameters of our model.

Across four experiments (Figure 1b), we found that accuracy was reduced after errors, even when there was ample time to prepare a response (up to 2 seconds). Our modeling results revealed that this effect

[1] Although the Drift Diffusion Model does include a parameter called 'non-decision time' that can influence RT, this parameter is not thought to reflect the cognitive processing that leads to a decision being made. Instead, it simply shifts the entire RT distribution earlier or later in time.

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could not be explained by slowed or more variable cognitive processing (i.e., more time required to select a response or trial-to-trial variability in this processing time) after an error. Instead, the observed behavior was due to a decrease in efficacy, or the probability of executing prepared responses. In other words, even when the stimulus timing suggested that a correct response was highly likely to be prepared, participants had more perseverative slips of action following an error. These results suggest that the increased response caution observed in prior studies examining post-error slowing is not simply reflecting a shift along the speed-accuracy curve, but may rather be an adaptive response to impairments in cognitive processing following an error.

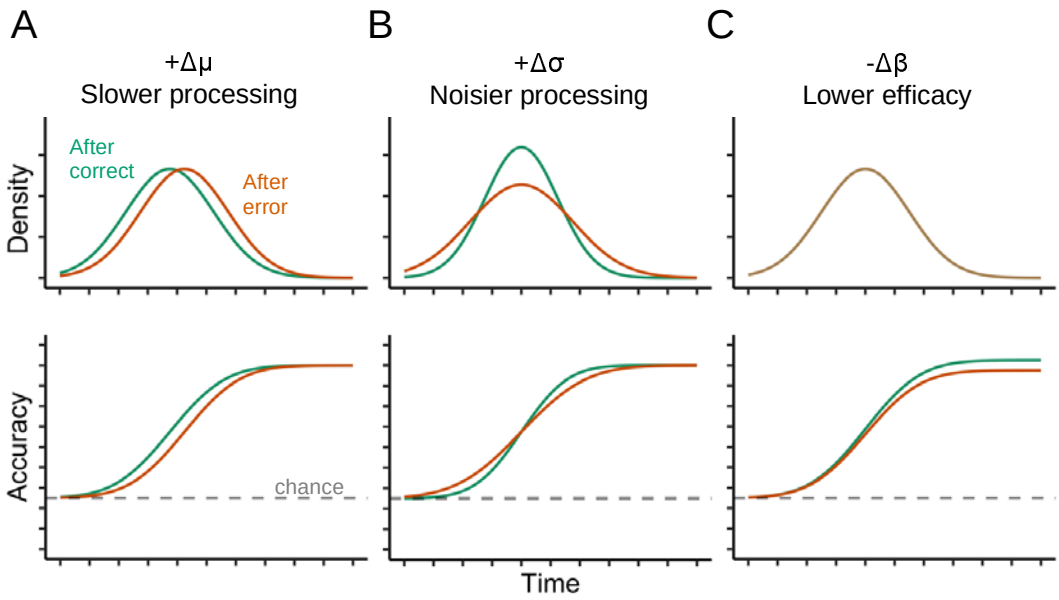

Figure 2. Hypothetical effects of previous error on response preparation. The first row depicts the effects in terms of the underlying response preparation distributions posited by the response preparation model used in the present study. The second row depicts the effects in terms of behavioral accuracy as a function of preparation time. Each column represents the effect of past error on a unique parameter in the model. Column A depicts slower response preparation after an error, Column B depicts noisier response preparation after an error, and Column C depicts reduced efficacy of prepared responses after an error. Efficacy is not a property of the underlying distributions but is instead a probability weight assigned to the prepared response—i.e., how likely is the participant to overtly express a response if it has already been prepared. A decrease in efficacy can be understood as an increase in the propensity for an action slip. Note that these parametric changes are associated with distinct predictions about the observable conditional accuracy functions (Row 2).

Results

Experiment 1

Following a brief training session to familiarize participants with stimulus-response mappings and the response timing required, participants completed 400 trials of a four alternative forced-choice stimulus response task (see Fig. 1 and Methods for complete details). Participants were required to make their responses between 1900-2100 ms following the start of each trial, but the stimulus presentation varied randomly between 0-2000 ms. This approach allowed us to investigate participants' accuracy while tightly controlling the amount of time available for the cognitive processing required to translate the stimulus information into a response.

Overall, there was strong evidence that participants were moderately less likely to perform the correct response if they made an error on the previous trial (b = -0.28, 95% CI = [-0.41, -0.15], pd = 1.0).

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However, a sliding window analysis revealed that the effect of previous error on accuracy depended on the amount of time available for preparation (Fig. 3a). For earlier processing times (PT < 500ms), there was no evidence for an effect of past error on future accuracy (b = 0.06, CI = [-0.25, 0.35], pd = .68). For later processing times (PT > 1000ms), there was strong evidence that participants were much less likely to perform the correct response following an error on the previous trial (b = -0.63, CI = [-0.98, -0.25], pd = 1.0).

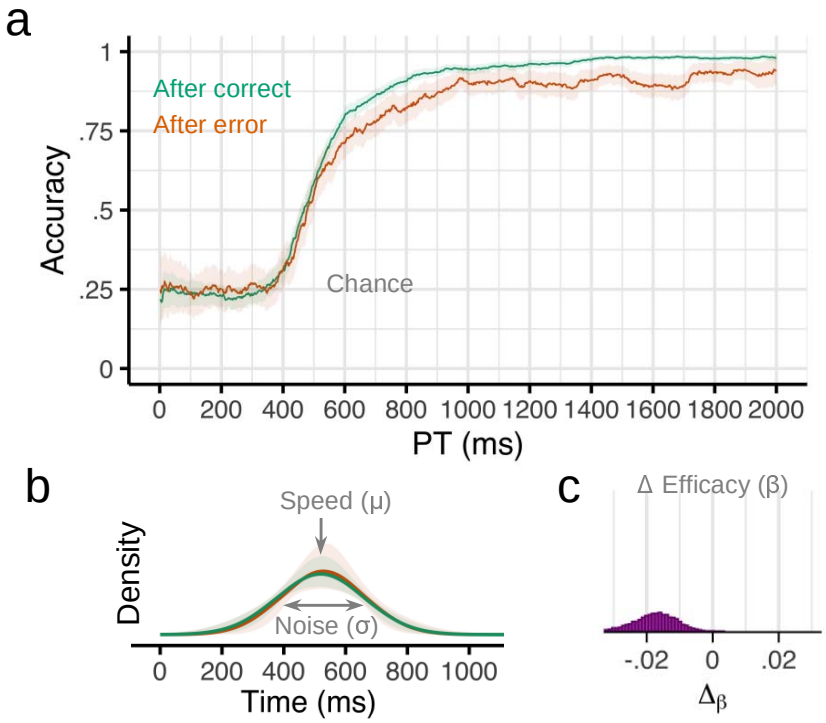

Figure 3. Results from Experiment 1. **a** Smoothed response accuracy as a function of preparation time and previous error. Bold lines represent smoothed means and ribbons represent smoothed 95% confidence intervals (i.e., standard error times 1.96). Note in the response preparation model, upper bound accuracy is controlled by the efficacy parameter β. **b** Model-estimated probability densities representing the time required to prepare responses following correct and incorrect trials. Densities were computed using group-level intercepts and slopes for μ and σ. Bold lines represent the posterior medians, and ribbons represent the 95% quantile intervals of the posterior. **c** Posterior (MCMC) distribution for group-level effects of previous error (Δ) on efficacy β.

We characterized the observed conditional accuracy functions (Fig. 3a) using a mathematical model of the underlying response preparation processes (Fig. 1c). Overall, the expected cognitive processing time required to prepare a response (μ) in this task was 523ms (CI = [502, 544]), the noise in preparation time (σ) was 140ms (CI = [109, 175]), and the efficacy of the prepared response (β) was .97 (CI = [.96, .98]). Efficacy was slightly lower after an error (M = .96, CI = [.94, .97]) compared to after no error (M = .97, CI = [.96, .98]. The Δ_β parameter was negative for nearly all posterior samples, indicating that efficacy post-error was consistently lower than efficacy following a correct trial (M_{diff} = -0.02, CI = [-0.03, -0.01], pd = .999; Fig. 3c). There was no evidence that preparation speed was affected by previous errors (M_{diff} = -0.01, CI = [-0.04, 0.05], pd = .68) or that preparation noise was affected by previous errors (M_{diff} = -0.01, CI = [-0.06, 0.05], pd = .59; Fig. 3b).

Experiment 2

We conducted an exact replication of the first experiment to test the reproducibility of the above finding that errors reduce efficacy on subsequent trials. There was some evidence that participants were slightly less likely to perform the correct response if they made an error on the previous trial (b = -0.11, 95% CI =

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[-0.24, 0.04], pd = .93). However, a sliding window analysis revealed that the effect of previous error on accuracy depended on the amount of time available for preparation (Fig. 3a). For later processing times (PT > 1000ms), there was strong evidence that participants were less likely to perform the correct response following an error on the previous trial (b = -0.50, CI = [-0.93, -0.05], pd = .99). For earlier processing times (PT < 500ms), accuracy was closer to chance levels and there was only some evidence

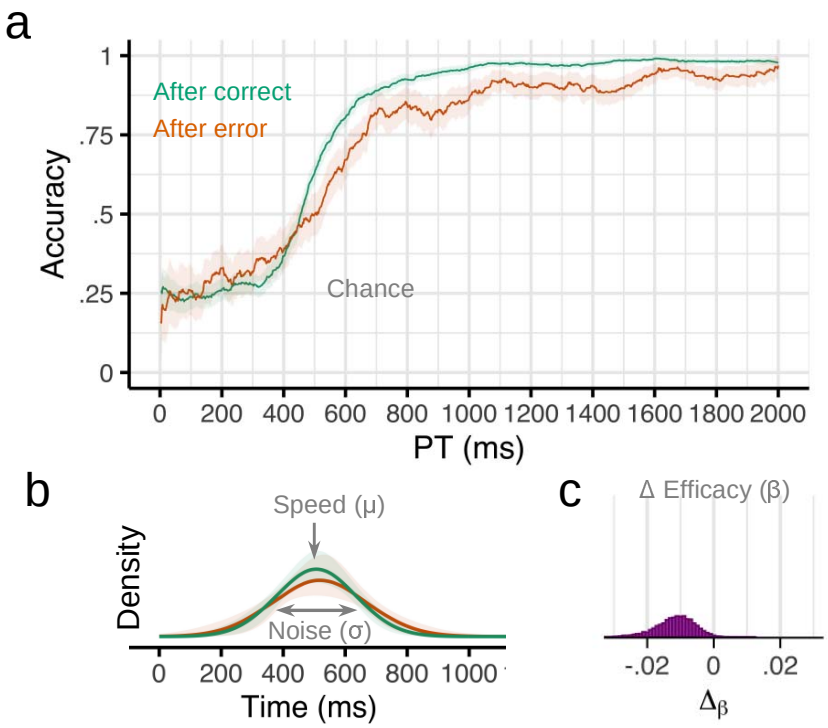

Figure 4. Results from Experiment 2. **a** Smoothed response accuracy as a function of preparation time and previous error. Bold lines represent smoothed means and ribbons represent smoothed 95% confidence intervals (i.e., standard error times 1.96). Note in the response preparation model, upper bound accuracy is controlled by the efficacy parameter β. **b** Model-estimated probability densities representing the time required to prepare responses following correct and incorrect trials. Densities were computed using group-level intercepts and slopes for μ and σ. Bold lines represent the posterior medians, and ribbons represent the 95% quantile intervals of the posterior. **c** Posterior (MCMC) distribution for group-level effects of previous error (Δ) on efficacy β.

that accuracy was higher after an error (b = 0.17, CI = [-0.08, 0.42], pd = .91).

Next, we accounted for the observed conditional accuracy functions (Fig. 4a) using a mathematical model of the underlying response preparation processes. Overall, the expected cognitive processing time required to prepare a response (speed, μ) was 512ms (CI = [481, 542]), the noise in preparation time (noise, σ) was 141ms (CI = [108, 178]), and the probability of expressing a response if it is prepared (efficacy, β) was .97 (CI = [.96, .98]). Efficacy was slightly lower after an error (M = .97, CI = [.95, .98]) compared to after no error (M = .98, CI = [.97, .99]. Although the credible intervals overlapped, the Δ_β parameter was negative for nearly all posterior (MCMC) samples, indicating that efficacy post-error was consistently lower than efficacy post-correct (M_{diff} = -0.01, CI = [-0.02, -0.002], pd = .99; Fig. 3c). There was little evidence that preparation speed was affected by previous errors (M_{diff} = 0.01, CI = [-0.02, 0.05], pd = .71) or that preparation noise was affected by previous errors (M_{diff} = 0.03, CI = [-0.02, 0.07], pd = .84; Fig. 3b). These results replicate the finding from Experiment 1 that errors impair subsequent cognitive processing by reducing the efficacy of a prepared response, thereby leading to an increase in slips of action.

Experiment 3

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One possibility for why we did not observe post-error slowing of cognitive processing in the first two experiments is because the time between trials was too long. Previous research has suggested that post-error slowing is reduced as the duration between a stimulus and a previous response grows (Danielmeier & Ullsperger, 2011; Jentzsch & Dudschig, 2009). We therefore conducted a pair of replication studies to assess whether the post-error effects observed previously depended on the duration of time between trials. Whereas in the first two experiments, the inter-trial-interval (ITI) was 1000ms, in the present Experiment 3, the ITI was 0ms. Overall, there was again strong evidence that participants were less likely to perform the correct response if they made an error on the previous trial (b = -0.42, 95% CI = [-0.62, -0.22], pd = 1.0). However, a sliding window analysis revealed that the effect of previous error on accuracy depended on the amount of time available for preparation (Fig. 5a). For earlier processing times (PT < 500ms), there was little evidence for an effect of past error on future accuracy (b = -0.14, CI = [-0.44, 0.15], pd = .82). For later processing times (PT > 1000ms), there was strong evidence that participants were much less likely to perform the correct response following an error on the previous trial (b = -0.87, CI = [-1.20, -0.49], pd = 1.0).

Again, we characterize the observed conditional accuracy functions (Fig. 5a) using a mathematical model of the underlying response preparation processes (Fig. 1c, 2). Overall, the expected cognitive processing time required to prepare a response (μ) in this task was 434 ms (CI = [385, 479]), the noise in preparation time (σ) was 202 ms (CI = [155, 265]), and the efficacy of the prepared response (β) was .94 (CI = [.92, .96]). Efficacy was slightly lower after an error (M = .91, CI = [.87, .94]) compared to after no error (M = .96, CI = [0.95, 0.97]. The Δ_β parameter was negative for all posterior samples, indicating that efficacy post-error was consistently lower than efficacy post-correct (M_{diff} = -0.05, CI = [-0.08, -0.03], pd = 1.0; Fig. 5c). There was no evidence that preparation speed was affected by previous errors (M_{diff} = 0.00, CI = [-0.05, 0.05], pd = .50; Fig. 5b) or that preparation noise was affected by previous errors (M_{diff} = -0.02, CI = [-0.09, 0.06], pd = .73; Fig. 5b).

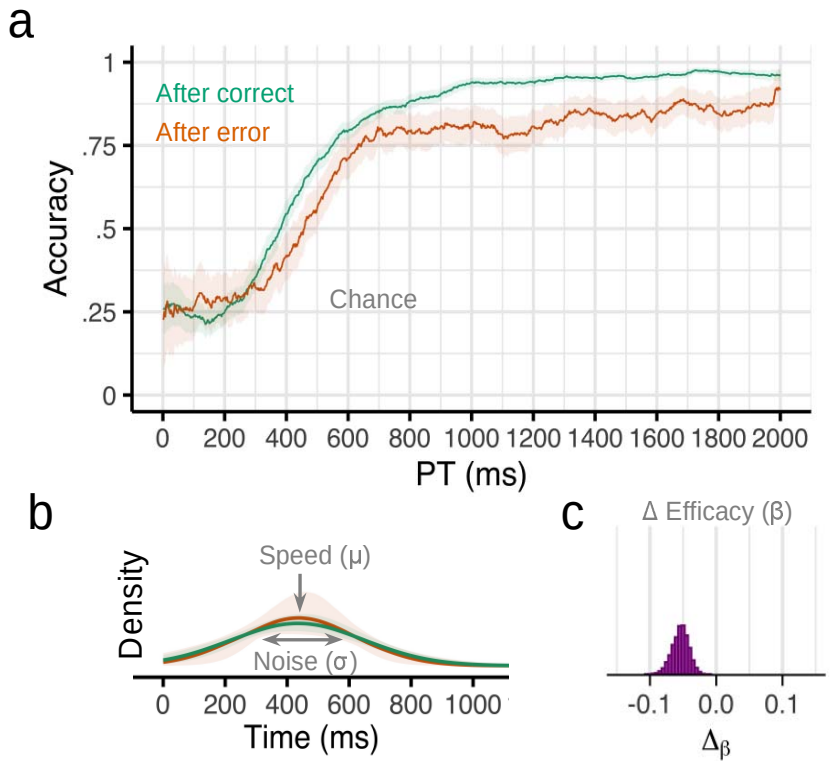

Figure 5. Results from Experiment 3. **a** Smoothed response accuracy as a function of preparation time and previous error. Bold lines represent smoothed means and ribbons represent smoothed 95% confidence intervals (i.e., standard error times 1.96). Note in the response preparation model, upper bound accuracy is controlled by the efficacy parameter β. **b** Model-estimated probability densities representing the time required to prepare responses following correct and incorrect trials. Densities were computed using group-level intercepts and slopes for μ and σ. Bold lines represent the posterior medians, and ribbons represent the 95% quantile intervals of the posterior. **c** Posterior (MCMC) distribution for group-level effects of previous error (Δ) on efficacy β.

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Experiment 4

In this experiment, we imposed an inter-trial-interval (ITI) of 2000ms. If the post-error effects reported above were due to an orienting response to an unexpected event (Notebaert et al., 2009) or transient distraction (i.e., flustering), then these effects might disappear during a sufficiently long ITI—however, this was not the case. Overall, there was strong evidence that participants were less likely to perform the correct response if they made an error on the previous trial (b = -0.23, 95% CI = [-0.43, -0.04], pd = .99). A sliding window analysis revealed that the effect of previous error on accuracy was relatively constant across processing times (Fig. 6a). There was evidence that participants were less likely to perform the correct response following an error on the previous trial for earlier processing times (PT < 500ms; b = -0.42, CI = [-0.78, -0.10], pd = .995) as well as for later processing times (PT > 1000ms; b = -0.34, CI = [-0.67, 0.04], pd = .97).

Again, we characterize the observed conditional accuracy functions (Fig. 6a) using a mathematical model of the underlying response preparation processes (Fig. 1c, 2). Overall, the expected cognitive processing time required to prepare a response (μ) in this task was 434 ms (CI = [396, 471]), the noise in preparation time (σ) was 214 ms (CI = [175, 257]), and the efficacy of the prepared response (β) was .96 (CI = [.94, .97]). Efficacy was slightly lower after an error (M = .95, CI = [.93, .97]) compared to after no error (M = .97, CI = [.95, .98]. The Δ_β parameter was negative for nearly all posterior samples, indicating that efficacy post-error was consistently lower than efficacy post-correct (M_{diff} = -0.01, CI = [-0.03, -0.00], pd = .98; Fig. 6c). There was some evidence that preparation speed was slower after an error (M_{diff} = 0.03, CI = [-0.01, 0.08], pd = .93; Fig. 6b) but there was no evidence that preparation noise was affected by previous errors (M_{diff} = 0.00, CI = [-0.06, 0.07], pd = .54; Fig. 6b).

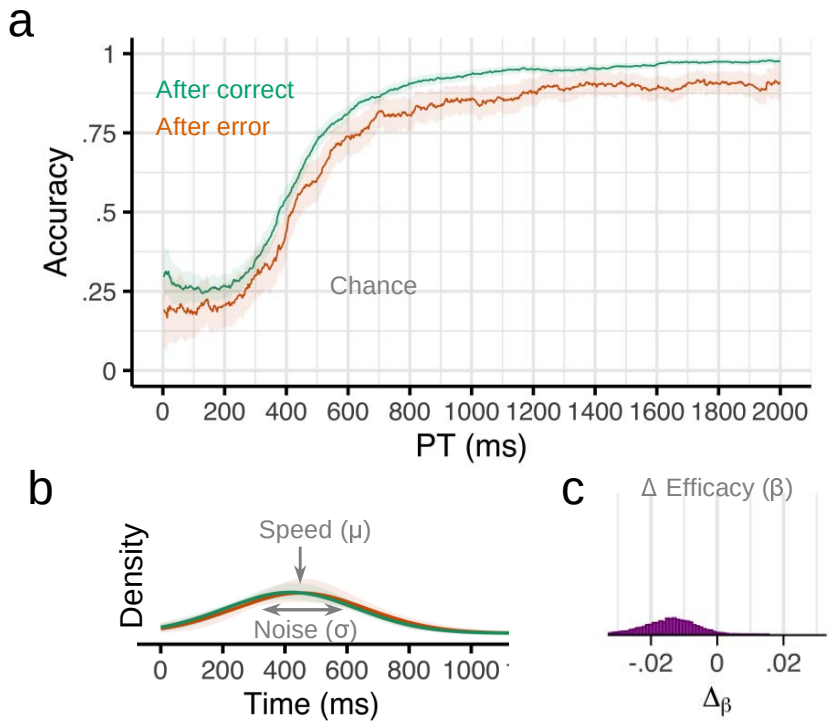

Figure 6. Results from Experiment 4. **a** Smoothed response accuracy as a function of preparation time and previous error. Bold lines represent smoothed means and ribbons represent smoothed 95% confidence intervals (i.e., standard error times 1.96). Note in the response preparation model, upper bound accuracy is controlled by the efficacy parameter β. **b** Model-estimated probability densities representing the time required to prepare responses following correct and incorrect trials. Densities were computed using group-level intercepts and slopes for μ and σ. Bold lines represent the posterior medians, and ribbons represent the 95% quantile intervals of the posterior. **c** Posterior (MCMC) distribution for group-level effects of previous error (Δ) on efficacy β.

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Next, we explored the persistence of post-error effects over time. In particular, we examined the effect of inter-trial-interval on the post-error effect by comparing the results from Experiment 3 (0ms ITI) with that of Experiment 4 (2000ms ITI) in terms of the size of the post-error effect on response efficacy (β) (Fig. 7b). We observe that the post-error effect was greater in Experiment 3 (Δ_β = -0.0535, CI = [-0.0802, -0.0314]) compared to Experiment 4 (Δ_β = -0.0141, CI = [-0.0312, -0.00097]). This result suggests that although providing more time for stimulus processing and response preparation during a trial does not eliminate post-error deficits, these post-error effects may dissipate during the time between trials.

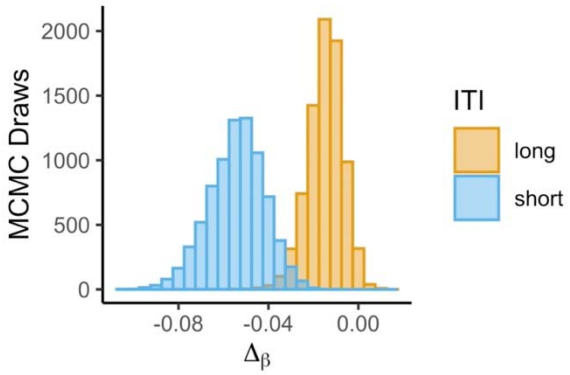

Figure 7. ITI modulation of post-error effects. Posterior distributions of post-error effects on response efficacy (β) for Experiment 3 (0 ms ITI) and Experiment 4 (2000ms ITI).

Exploring the nature of post-error processing deficits

Experiments 1-4 established that post-error deficits in performance are not due to the slowing (Δμ) of cognitive processing underlying response preparation or an increase in trial-to-trial variability (Δσ) in the speed of cognitive processing. Instead, we observed more slips of action (Δβ) at all time points where a correct response was likely to be prepared. What could be driving this static impairment in executing the correct response? One possibility is that people are simply biased away from repeating a response that just resulted in an error. We explored whether incorrect responses after an error were driven by a response bias—that is, a tendency to avoid/repeat the same key as on the previous trial. This analysis focused on incorrect trials for which the preparation time was greater than 1000ms because this was the locus of the post-error effect. We fit hierarchical Bernoulli regression models to these data with repetition as the outcome and previous error as the covariate. Across all four datasets, we found that people were *more* likely to repeat the previous keypress after an error compared to after a correct trial (Fig. 8). We found strong evidence for increased perseveration after an error in Experiment 1 (b = 1.03, CI = [0.03, 1.96]), Experiment 2 (b = 0.87, CI = [0.19, 1.53], Experiment 3 (b = 1.33, CI = [0.76, 1.90]), and Experiment 4 (b = 1.72, CI = [1.03, 2.43]). These results suggest that an increase in perseverative action slips following an error could at least partly explain post-error deficits in performance.

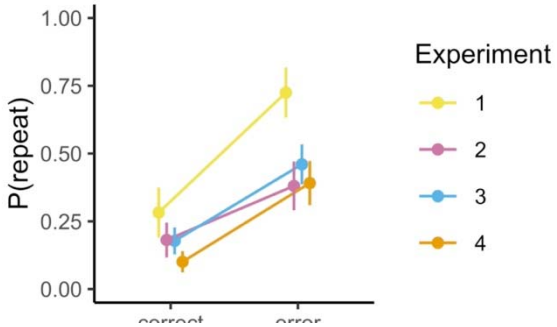

Figure 8. Post-error perseveration. Proportion of incorrect trials (PT > 1000ms) in which the participant pressed the same key as on the previous trial, grouped by previous trial accuracy and experiment. Points are means and error bars are 95% confidence intervals (+/- SEM * 1.96)

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Experiments 3 and 4 established that performance recovers somewhat when there is more time to recover from an error before the start of the next trial. However, our participants were still more prone to slips of action after an error when given a full two seconds between trials. When does performance fully recover? To address this question, we explored the persistence of post-error effects across trials. In particular, we examined the lingering effect of errors on the subsequent two trials. We were especially interested in whether post-error effects persisted when there was an intervening correct trial. We re-fit the response preparation model with N-2 error as the covariate, separately for N-1 correct and N-1 error trials. For trials preceded by a correct response (N-1), we found little evidence for N-2 error effects (Exp. 1: Δ_β = -0.008, CI = [-0.021, 0.001]; Exp. 2: Δ_β = -0.002, CI = [-0.013, 0.009]; Exp. 3: Δ_β = -0.017, CI = [-0.036, 0.0001]; Exp. 4: Δ_β = -0.017, CI = [-0.033, -0.004]). However, for trials preceded by an error (N-1), we found evidence for N-2 error effects on current trial response efficacy (β) in Experiment 1 (Δ_β = -0.024, CI = [-0.053, -0.005) and Experiment 4 (Δ_β = -0.055, CI = [-0.095, -0.024]), but not in Experiment 2 (Δ_β = -0.015, CI = [-0.040, 0.004) or Experiment 3 (Δ_β = -0.033, CI = [-0.075, 0.002]). These results suggest that post-error effects are 'reset' if there is an intervening correct response, but they may persist (or even compound) if there is an intervening incorrect response.

Discussion

In the present study, we examined post-error effects in a stimulus-response task in which the time available for response preparation was manipulated. Across four experiments, we found that participants' response accuracy was lower if they made an error on the previous trial. This deficit was observed even when participants were given ample time to prepare their responses. A model-based analysis revealed that the post-error effect on accuracy was driven by a decrease in the probability of expressing a response given that it was highly likely to be prepared (β). These erroneous slips of action were predominantly perseverative repeats of the button pressed on the previous trial. Our analyses ruled out that the post-error effect on accuracy are due to a slowing of the cognitive processing required to prepare a response (μ) or variability in the latency of cognitive processing (σ). These results suggest that previous findings of post-error slowing are unlikely to be due to a decrease in the speed of cognitive processing underlying the preparation of responses to stimuli. Furthermore, our results indicate that delaying initiation of a prepared response could be an adaptive response to impaired efficacy of cognitive processing after an error rather than a strategic shift along the speed-accuracy curve.

Our conclusions rely on the view that response preparation and response initiation are independent motor control parameters(Haith et al., 2016; Hardwick et al., 2019; Wong et al., 2017). People not only decide *what* response to make, they also decide *when* to make their response. Prior studies on post-error effects do not directly distinguish between selection and initiation, in their experimental designs or in their theoretical models. One issue is that these studies use free RT, which is some combination of the duration of time spent selecting a response and the duration of delay after selection before initiation. Typical models used to explain post-error effects, such as the drift diffusion model, are models of how people decide *what* to do—i.e., response selection. However, recent work in motor control shows that people initiate responses (in a free RT task) long after they have prepared those same responses as revealed in a forced-RT task (Haith et al., 2016). Thus, free RTs could be used to disentangle selection and initiation only with a more complete model that also describes the processes underlying the decision about *when* to initiate a selected response. Absent such a model, an alternative approach is to control response initiation using a forced-RT task, as in the present study.

Several studies in the past have shown that people respond more slowly after an error (Danielmeier & Ullsperger, 2011; Laming, 1979; P. Rabbitt & Rodgers, 1977). A prominent account of these effects is that participants are more cautious after an error(Danielmeier & Ullsperger, 2011). For example, evidence accumulation models have been used to argue that participants alter their decision thresholds to accumulate more evidence before deciding on a response(Dutilh et al., 2012). On this view, responses might take longer to prepare after an error. In the present study, we found that participants responded *less* accurately after an error, even if they were given up to two seconds to prepare their response.

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Contrary to some previous accounts, we found no evidence that the cognitive processing required to prepare responses occurred more slowly after an error. That is, our estimate of the latency at which a response is prepared (μ) was unaffected by errors. Our data provide evidence against the view that post-error effects on performance are only due to an increase in the evidence required to select a particular response. Instead, post-error slowing effects might be better explained by adaptive delays in response initiation to compensate for impairments stimulus-response processing that lead to more frequent action slips.

Other researchers have similarly suggested that there are unavoidable negative consequences to errors and unexpected events (Notebaert et al., 2009; Purcell & Kiani, 2016; Wessel & Aron, 2017). Notebaert and colleagues (Notebaert et al., 2009) proposed that infrequent, surprising events might cause an 'orienting response' that distracts participants from the processing of a subsequent stimulus. Another account advanced the idea that errors trigger a transient global inhibitory response which affects both motor and cognitive function (Wessel & Aron, 2017). Although similar in providing evidence for a maladaptive response following errors, our results do not fully support either of these accounts. From this prior work, one would expect that when ample time is given to make a response we should observe identical levels of accuracy following both correct and incorrect responses. An orienting response and a transient inhibitory response should resolve relatively quickly and performance should recover. Although we did find some evidence that a longer inter-trial interval lessened the deleterious impact of errors on subsequent performance (Experiments 3 and 4), participants' accuracy did not improve when they were given up to two seconds to respond following stimulus presentation. Purcell and Kiani observed more slowed responses following errors at low stimulus strength in a motion discrimination task (Purcell & Kiani, 2016). Using drift diffusion modeling of response time distributions, they described this result as a combination of an increase in the response threshold and a decrease in sensory signal-to-noise ratio (SNR) following errors. This would make the preparation of response less accurate, but also either slower or more variable. However, here we did not observe any slowing in the estimated latency (μ) of the cognitive processing underlying response preparation following an error. We also did not find evidence of any increase in variability in the time it takes to complete this cognitive processing (σ). Instead, regardless of the amount of time given to prepare, participants were less accurate and were more prone to perseverative slips of action. Although the stimulus remained on the screen until a response was made in our task, it is possible that errors cause an initial impairment in subsequent stimulus-response processing that cannot be corrected online.

The experiments used in the present study differ in one critical way from previous experiments examining post-error effects: the time at which participants initiate their responses was controlled. This so-called interrogation method disables participants from strategically delaying the initiation of their responses, including after an error(Bogacz et al., 2006). We found that, under this constraint, participants were less accurate after an error. A response preparation model explained this effect in terms of a decrease in the efficacy of the cognitive processing underlying prepared responses. Here efficacy was defined as the probability that a participant will express a response after it has been prepared. A natural psychological interpretation of efficacy is participants' confidence in their selected response. It may be that if people are led astray by decisions in the past, they become less confident about their decisions in the future. This reduced confidence would make them less likely to act on their decisions and potentially more prone to make random responses. Note that this is distinct from requiring more processing time to prepare a response. As noted above, we found little evidence to support post-error slowing of processing demands (μ).

It is well-known that free RTs are not normally distributed and instead tend to follow distributions incorporating some skewness such as the an ex-gaussian distribution (Heathcote et al., 1991). Although allowing us to easily interpret modeling outcomes, one potential limitation of our study is that we assume that the amount of time it takes to complete the cognitive processing necessary to select a response can be approximated well by a normal distribution. However, we do not believe this limitation greatly affects the conclusions drawn here. The fit of our model to the data is quite good, perhaps specifically because the time deadlines do not allow for the long tail sometimes observed in RT distributions. Additionally, if it were the case that errors caused an increase in the skewness of the latency of cognitive processing, this

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would not come out as a reduction in the efficacy of cognitive processing (β) as we observed here. Instead, we would have observed a shift in the mean (μ) and/or standard deviation (σ) of our estimates of the response preparation distributions to try to account for a longer tail. We found no evidence of this across our four experiments.

Our interpretation of the results relies on the assumption that cognitive processing in free RT tasks is similar to that seen in the forced response paradigm used here. We believe this assumption is reasonable. In the context of conflict tasks (e.g. the Simon task), we have previously observed similar effects on accuracy in a forced response paradigm as is observed in response time in free RT tasks (Adkins & Lee, 2021). However, it is possible that forcing participants to respond at a predetermined time changes the nature of the task and the cognitive processes underlying responses. For example, a forced response paradigm might lead to more task engagement and leave individuals less prone to inattentiveness. Nevertheless, there is no particular reason to privilege data from free RT tasks when attempting to understand post-error effects on performance. The results presented here are not readily explained by prevailing theories of post-error effects that have been developed from free RT tasks. The forced response paradigm provides a window into post-error impairments in cognitive processing that have been difficult to examine with standard techniques.

In sum, we provide evidence against the view that cognitive processing is slower after an error. Instead, our data suggest that decisions about what to do are less likely to be translated into the appropriate motor responses after an error and lead to more perseverative slips of action. Previous observations of post-error slowing may reflect strategic delays in response initiation to compensate for this impaired efficacy in cognitive processing underlying response preparation.

Methods

Participants

Participants were recruited online using the online platform Prolific with the following inclusion criteria: US or Canadian nationality, fluent English speaker, approval rate >95%. In Experiment 1 there were 46 participants (24 female) with a mean age of 33 years old. In Experiment 2 there were 37 participants (33 female) with a mean age of 27 years old. In experiment 3 there were 47 participants (24 female, 1 declined to answer) with a mean age of 29 years old. In Experiment 4, there were 46 participants (22 female, 3 declined to answer) with a mean age of 32 years old.

Experiments

First, participants trained for 60 trials in a stimulus-response task. For this task, participants were instructed to press ‘f’, ‘g’, ‘h’, or ‘j’ depending on the identity of the stimulus (see below). At the start of each trial (t = 0ms) a stimulus was presented, and the identity of the stimulus was randomly sampled from a uniform categorical distribution over four unique stimuli (colors or symbols). During each trial, the four possible stimulus-response mappings were presented at the bottom of the screen to help participants learn these mappings. After each trial, participants were given feedback for 500ms about the accuracy of their response. Response times were unconstrained during this phase of the experiment.

Next, participants trained for 20 trials in a fixed response timing task. For this task, participants were instructed to press a key exactly when two empty white rectangles were filled h color, exactly two seconds after the trial began. At the beginning of each trial, two empty rectangles (PsychoPy height unit: .35*0.03 % of screen) were shown above and below where the stimuli had appeared in the previous training block. Every 500ms, the rectangle was filled in by an additional 25%. After 2100ms, all stimuli were removed from the display. The purpose of this cuing was to guide participants to respond at the same time on every trial. Participants were encouraged beforehand to alternate between ‘f’, ‘g’, ‘h’, and ‘j’ to practice timing with all keys. After each trial, participants were given feedback for 500ms about whether they responded too quickly (RT < 1900ms), too slowly (RT > 2100ms), or with perfect timing.

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Finally, we turned to the main experimental task, illustrated in Figure 1. Participants performed 10 blocks of 40 trials in the stimulus-response task a fixed response timing and a stimulus presentation time that was parametrically varied. As in the first phase of training, participants were instructed to press one of four keys depending on which one of the four stimuli was shown. The onset and identity of the stimulus was varied randomly as in the first phase of training. However, there was no prompt showing the S-R mappings during the trial. As in the second phase of training, participants were instructed to respond only when the rectangle timing cue was filled. This approach allows us to measure the accuracy of responses when the exact amount of time allowed for stimulus processing and response preparation is known. After each trial, participants were given feedback, but the exact specifications of this feedback varied across experiments.

In Experiment (Exp.) 1, the target stimuli were letters in the Armenian alphabet (□, □, □, □; PsychoPy height unit: 0.2% of screen) and participants were given feedback for 1000ms about whether they were too slow, too fast, or had perfect timing, followed by a 1000ms inter-trial-interval (ITI). Exp. 2 was an exact replication of Exp. 1.

In Experiments 3 and 4, the target stimuli were color-filled circles (orange, blue, red, or purple; PsychoPy size: 0.2 x 0.2). Here the response keys were ‘d’, ‘f’, ‘j’, and ‘k’. Participants were given 400 ms feedback if and only if they responded too quickly or too slowly; no feedback if their timing was perfect. Feedback was followed by an ITI of 0ms in Exp. 3 and an ITI of 2000ms in Exp. 4.

All experiments were built using PsychoPy/JS and run online using Pavlovia.

Pre-processing

Our analyses focused exclusively on behavior in the test phase of the experiment. We excluded trials in which participants responded too quickly (RT < 1900 ms) or slowly (RT > 2100 ms) because we were interested in how people behaved when their response times were fixed to the imperative cue. Since we were interested in the effects of immediately preceding errors, we also excluded trials for which there was no immediately preceding trial (i.e., the first trial of each block for each participant). Response accuracy (y) was set to 1 if the response was correct and 0 if the response was incorrect. Preparation time (PT) was defined as the duration between the stimulus onset and the response time. PT was re-scaled range from zero to one by dividing by 2000 to facilitate prior specifications. Each trial was labeled with the outcome on the previous trial. Previous-error (err_{n-1}) was set to 0.5 if the previous response was incorrect and -0.5 if the previous response was correct, so that the intercepts in models with post-error slopes can be interpreted as the average across levels of the condition.

Analysis

We used a sliding window technique to visualize the mean of response accuracy as a function of PT. The sliding window was performed separately for each level of err_{n-1}. The width of the sliding window was 100ms, the step-size was 1ms, and the window moved from 0 to 2000ms.

We used Bernoulli regression models to assess the credibility of apparent effects in the smoothed conditional accuracy functions. The models had hierarchical slopes and intercepts (i.e., participant-level variables sampled from group-level distributions) and focused on specific contiguous intervals over PT (e.g., 0ms < PT < 500ms). We used the R-package brms (Bürkner, 2017) to specify the models and to approximate posterior distributions over unobserved variables, given the observed data.

We used a response preparation model to explain the observed time-courses of performance (Hardwick et al., 2019). This model enables inferences about the speed with which responses were prepared (μ), the noise in this preparation (σ), the probability that a prepared response will expressed if it is prepared (efficacy, β), as well as effects of covariates (Δ) on these variables. We specified the model using Stan (Carpenter et al., 2017). In this model, the unobserved variables mentioned above were computed as linear functions of err_{n-1} and the effects of err_{n-1} were captured by a set of delta parameters. All intercepts

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and slopes in the model (μ_0, Δ_μ, σ_0, Δ_σ, β_0, Δ_β) were hierarchical (i.e., participant-level variables sampled from group-level distributions) and we assigned weakly informative priors to the group-level variables. The model specification is reported below, where y is the observed response accuracy, p is the probability of correct response, f is the response preparation function which computes p given the observed preparation time t and the unobserved variables $\theta = \{\mu, \sigma, \beta, \alpha\}$, and θ_{jk} is the unobserved variable θ for participant j and condition k (i.e., post-error level).

$$
\begin{aligned}
&y \sim \mathrm{Bernoulli}(p)\\
&p = f\left(t, \mu_{jk}, \sigma_{jk}, \beta_{jk}, \alpha_j\right)\\
&f \leftarrow \varphi \cdot \psi\\
&\varphi = \left[1 - \mathrm{Normal}_{\mathrm{cdf}}\left(t, \mu_{jk}, \sigma_{jk}\right), \mathrm{Normal}_{\mathrm{cdf}}\left(t, \mu_{jk}, \sigma_{jk}\right)\right]\\
&\psi = \begin{bmatrix} \alpha_j \\ \beta_{jk} \end{bmatrix}
\end{aligned}
$$

As mentioned above, unobserved variables (except for α) were computed as linear functions of err_{n-1}. Below is the specification of these variables, where θ_j^0 is the intercept for the variable θ and participant j, x is the value of err_{n-1}, Δ_j^θ is the effect of x on θ for participant j. Inverse-logit link functions were used to constrain the values of the variables between 0 and 1 (seconds for μ and σ, probability for α and β).

$$
\begin{aligned}
\mu_{jk} &= \mathrm{logit}^{-1}\left(\mu_j^0 + \Delta_j^\mu * x\right)\\
\sigma_{jk} &= \mathrm{logit}^{-1}\left(\sigma_j^0 + \Delta_j^\sigma * x\right)\\
\beta_{jk} &= \mathrm{logit}^{-1}\left(\beta_j^0 + \Delta_j^\beta * x\right)\\
\alpha_j &= \mathrm{logit}^{-1}\left(\alpha_j^0\right)
\end{aligned}
$$

The slopes and intercepts of the variables above were defined hierarchically. Below is the specification of these hierarchical variables, where θ_{loc}^0 is the group-level mean of the intercepts for a parameter θ, $\theta_{\mathrm{scale}}^0$ is the group-level standard deviation of the intercepts, $\Delta_{\mathrm{loc}}^\theta$ is the group-level mean of the err_{n-1} effects on θ, and $\Delta_{\mathrm{scale}}^\theta$ is the group-level standard deviation of the err_{n-1} effects. Note that in the stan code we used a non-centered parameterization, despite presenting the centered parameterization below for ease of comprehension.

$$
\begin{aligned}
\mu_j^0 &\sim \mathrm{Normal}(\mu_{\mathrm{loc}}^0, \mu_{\mathrm{scale}}^0)\\
\Delta_j^\mu &\sim \mathrm{Normal}(\Delta_{\mathrm{loc}}^\mu, \Delta_{\mathrm{scale}}^\mu)\\
\sigma_j^0 &\sim \mathrm{Normal}(\sigma_{\mathrm{loc}}^0, \sigma_{\mathrm{scale}}^0)\\
\Delta_j^\sigma &\sim \mathrm{Normal}(\Delta_{\mathrm{loc}}^\sigma, \Delta_{\mathrm{scale}}^\sigma)\\
\beta_j^0 &\sim \mathrm{Normal}(\beta_{\mathrm{loc}}^0, \beta_{\mathrm{scale}}^0)\\
\Delta_j^\beta &\sim \mathrm{Normal}(\Delta_{\mathrm{loc}}^\beta, \Delta_{\mathrm{scale}}^\beta)\\
\alpha_j^0 &\sim \mathrm{Normal}(\alpha_{\mathrm{loc}}^0, \alpha_{\mathrm{scale}}^0)
\end{aligned}
$$

Finally, we assigned weakly informative priors to group-level location and scale variables. Scale variables were constrained to be positive and assigned Normal (0, .5) priors. Location variables were left unconstrained. The prior was Normal (-.5, .5) for the location of the μ intercept, Normal (-2, .5) for the σ intercept, Normal (2, .5) for the β intercept and Normal (-1, .5) for the α intercept. Delta location variables were assigned Normal (0, .5) priors. Note that these normal distributions are akin to beta distributions after being transformed by the inverse logit function.

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