*Title*

**Estimating clade-specific diversification rates and palaeodiversity dynamics from**

**reconstructed phylogenies**

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*Abstract*

One of the central goals of macroevolution is to understand palaeodiversity dynamics through

time and space, and its potential drivers. Estimating palaeodiversity dynamics has been

historically addressed with the fossil record because it directly reflects the past variations of

biodiversity. Unfortunately, some groups or regions lack a good fossil record, and dated

phylogenies can be useful to estimate deep-time diversification dynamics. Recent

methodological developments have unlocked the possibility to investigate palaeodiversity

dynamics with dated phylogenies by using birth-death models with non-homogeneous rates

through time and across clades. One of them, which models rate heterogeneity as a local

clock, seems particularly promising to detect clades whose diversity has declined through

time. However, empirical applications of the method have been hampered by its lack of

automation as leading the analysis by hand was costly and error-prone.

Here we propose an automation with additional features accounting for recent developments

and we implement it in the R package RPANDA. We also test the approach with simulations

focusing on its ability to detect shifts of diversification and to infer palaeodiversity dynamics.

Finally, we illustrate the automation by investigating the palaeodiversity dynamics of Cetacea,

Vangidae, Parnassiinae, and Cycadales.

Simulations showed that we accurately detected shifts of diversification although false shift

detections were higher for complex diversification models. The median global error of

palaeodiversity dynamics estimated with the automated model is low showing that the method

can capture diversity declines. We detected shifts of diversification for three of the four

empirical examples considered (Cetacea, Parnassiinae and Cycadales). These heterogeneities

of diversification rates unveil a waxing-and-waning pattern due to a phase of negative net

diversification rate embedded in the trees after isolating recent radiations.

Our work allows to easily apply non-homogeneous models of diversification in which rates can vary through time and across clades to reconstruct palaeodiversity dynamics. By doing so, we detected palaeodiversity declines among three of the four groups tested, highlighting that such periods of negative net diversification might be common. We discuss the extent to which this approach might provide reliable estimates of extinction rates and we provide guidelines for users.

*Keywords*

Birth-death model, macroevolution, palaeodiversity, rate heterogeneity, RPANDA.

## 1 | Introduction

Estimating palaeodiversity dynamics over geological timescales is a thrilling goal of macroevolution (Alroy et al. 2008; Quental and Marshall 2010; Morlon et al. 2011). This question is generally addressed with paleontological data because fossil records are direct witnesses of past diversity dynamics (Marx and Uhen 2010; Ezard et al. 2011). Unfortunately, the fossil record is not homogeneous over time, space, and groups (Close et al. 2020). Consequently, investigating evolutionary history of poorly fossilised groups relies on diversification analyses based on dated phylogenies (time-trees) mainly composed of extant species. Methodological developments in the field have focused on estimating the speciation rate ($\lambda$) and extinction rate ($\mu$), which represent the pace at which lineages appear and go extinct, respectively, as well as the net diversification rate (denoted r) defined as the difference between them (Ricklefs 2007; Stadler 2013; Morlon 2014). A large variety of phylogenetic models allows testing evolutionary hypotheses on whether diversification rates are constant, vary over time (Rabosky and Lovette 2008; Morlon et al. 2010), across clades (Morlon et al. 2011; Rabosky et al. 2013; Maliet et al. 2019) or in association with species' traits (Maddison et al. 2007; FitzJohn 2012; Beaulieu and O'Meara 2016) or with environmental changes (Condamine et al. 2013, Condamine, Rolland & Morlon 2019). However, early studies have shown that reconstructed dated phylogenies can display heterogeneity of diversification rates across clades (Rabosky et al. 2007, Esquerré et al. 2019).

Among the methods developed to detect this heterogeneity (Alfaro et al. 2009; Stadler and Bokma 2013; Rabosky et al. 2013; Maliet et al. 2019; Barido-Sottani et al. 2020; Maliet & Morlon 2022), the model developed by Morlon et al. (2011) considers heterogeneity in diversification rate by testing for shifts in rates of diversification at specific branching points in the phylogeny. The diversification of subclades identified *a priori* based on taxonomy or

ecological traits, and the diversification of the ancestral group leading to these clades,

hereafter called the backbone, can be analysed separately. The main advantage of this method

is that each part of a phylogeny (subclades and backbone) can follow a different birth-death

model, offering the opportunity to test biological hypotheses that create heterogeneity of

diversification rate and thus reconstruct palaeodiversity dynamics for each part of the tree

(Nee et al. 1992; Morlon et al. 2011; Morlon 2014). Doing so for Cetacea, Morlon et al.

(2011) showed a good consistency in the trend of the palaeodiversity dynamics between the

molecular phylogeny and the fossil record. Recent advances have improved palaeodiversity

estimation with a probabilistic approach providing confidence intervals of palaeodiversity

values (Billaud et al. 2020). Subclades might experience recent adaptive radiations due to the

acquisition of key innovations (Jønsson et al. 2012; Drummond et al. 2012) while the

backbone tree might display a different dynamic such as a waxing-waning pattern of

diversification. The resulting heterogeneity of diversification rates cannot be detected in

clade-homogeneous birth-death models because recently radiating clades can overshadow the

estimation of extinction in other parts of the tree, and likely hamper the inference of a

negative net diversification rate ($r < 0$, $\mu > \lambda$) that may explain such a period of diversity

decline (Morlon et al. 2011). The waxing-waning pattern detected by Morlon et al. (2011) was

promising, yet the approach developed in that paper, part of which was made accessible in the

R package RPANDA (Morlon et al. 2016), has been applied to only a few empirical cases

since it was proposed (Bacon et al. 2013; Legendre and Condamine 2018; Billaud et al. 2020;

Chazot et al. 2021) because of the lack of automation.

The method developed by Morlon et al. (2011) requires manually isolating subclades

and backbones, accounting for all pruned speciation events, applying diversification models,

and testing all backbone(s)/subclade(s) combinations by hand. This is time-consuming, error-

prone, and makes the method underutilized. In addition to making this method more

automatic and user-friendly, several aspects of its performance remain to be tested. It was

shown that the method recovers unbiased rates of diversification and can detect periods of

diversity decline (Morlon et al. 2011), but its ability to detect clade-specific shifts of

diversification remains untested and can make the palaeodiversity dynamic challenging to

recover (Burin et al. 2019). More generally, several studies have questioned our ability to

estimate extinction rate from molecular phylogenies (Rabosky 2010, 2016; Louca and Pennell

2020; but see Nee et al. 1994; Morlon 2014; Beaulieu and O'Meara 2015). Historically, Nee

et al. (1994) have demonstrated how to estimate constant extinction rate from molecular-dated

phylogenies with extant species only. Some studies relaxed the assumption of constant rates

(Rabosky and Lovette 2008, Stadler 2011a), but these models are homogeneous across clades.

By being non-homogeneous across clades, the model of Morlon et al. (2011) brings better

extinction rate estimates by revealing ancient periods of palaeodiversity decline (Stadler

2011b) in line with the fossil record (Marx and Uhen 2010). This model opens the possibility

to investigate palaeodiversity dynamics for groups without good fossil records.

In this study, we automatize the approach developed by Morlon el al. (2011) and

implement it in the R package RPANDA (Morlon et al. 2016). In addition, we provide several

advances compared to the original method. In particular, we implement the possibility to test

for a shift at the crown age of subclades, whereas the original version of the method tested

shifts at the stem age of subclades (Morlon et al. 2011; Billaud et al. 2020). This rationale is

motivated by the idea that clade radiations – may it be adaptive or simply rapid compared to

the rest of the tree – are characterized by long stem branches at the crown of subclades, with

fast speciation starting at the crown. We provide functions to calculate rates and diversity

dynamics through time, including recent developments from Billaud et al. (2020) with new

functions using automation outputs. Using simulations, we assess the performance of the

approach to detect shifts and to infer palaeodiversity dynamics under various diversification

scenarios. We further illustrate the approach empirically by investigating the evolutionary

history of four groups with well-sampled molecular-dated phylogenies (Cetacea, Vangidae,

Parnassiinae and Cycadales), which have been previously analysed with other methods. We

finally discuss both simulations and empirical cases to provide guidelines and

recommendations when using this approach.

## 134    2 | Methods

### 136    2.1 | The clade-dependent time-dependent diversification model of Morlon et al. (2011)

Morlon et al. (2011) developed an exact analytical expression of a birth-death model

of cladogenesis that estimates rate variation through time while accounting for undersampling

thanks to a sampling fraction (denoted by $f$, $f$=1 if all described species are included in the

phylogeny). The expression is defined backwards in time such that values for $\lambda$ and $\mu$ for

time-dependent models correspond to rate values at present ($t$=0) and $\alpha$ and $\beta$ are their

dependency parameters, respectively. Then, the assumed changes of diversification rates

occur at specific branching points (nodes) of the reconstructed phylogeny, meaning that

potential rate shifts on unobserved (extinct or unsampled) lineages are ignored (Moore et al.

2016; Barido-Sottani et al. 2020). Under such assumptions, their analytical expression holds

and in particular their expressions of the probability that a lineage alive at time $t$ has no

descendant in the sample at $t$=0 ($\Phi(t)$), and the probability that a lineage alive at time $t$ leaves

exactly one descendant lineage at time $s < t$ ($\Psi(s, t)$, *i.e.* the probability of survival from $t$ to $s$,

see Appendix S8 and S9 from Morlon et al. 2011). These two expressions allow expressing

the likelihood of the backbone phylogeny by including the speciation event corresponding to

the subclade. The probability $1 - \Phi(t_1^{cl})$ that the subclade survives from this branch point $t_1^{cl}$ is

then included in the likelihood of the subclade expressed with the same formula. Finally, the

product of the likelihood of the subtree and the backbone phylogeny gives the global

likelihood of the phylogeny while accounting for heterogeneity of diversification rates across

clades. This global likelihood can be compared to the likelihood of a diversification model

with a homogeneous model of diversification to see if the shift of diversification rate is

supported. This can be extended to a phylogeny with multiple subclades following their own

diversification rates and even shifts of diversification rates inside the backbone creating

multiple backbones.

**2.2 | Automation**

**2.2.1 | Implementation in RPANDA**

Few empirical studies have applied the approach using a manual and sequential

approach (Morlon et al. 2011; Bacon et al. 2013; Billaud et al. 2020; Chazot et al. 2021). Each

shift was selected one by one among subclades defined *a priori* until either the inclusion of a

new shift was no longer statistically supported or shifts at all nodes were supported. Our

automated approach still relies on the *a priori* definition of a limited number of subclades in

which shifts may have occurred, but contrary to the sequential approach it compares all

combinations of these potential shifts at the same time. The simplest way to make an *a priori*

selection of subclades is to follow the approach of Morlon et al. (2011) by choosing

taxonomic subclades such as families. Moreover, taxonomic databases then often provide the

information required to compute sampling fractions (i.e. clade-level taxonomic diversity) that

are required by phylogenetic diversification models. The tested subclades can also correspond

to clades of species sharing key-innovations or colonisation events to test the role of these

biological events on the clade's diversification. Once subclades have been chosen *a priori,* the

whole analysis is divided in four steps: (1) calculating sampling fractions for the subclades

from the taxonomic database, (2) calculating all potential combinations of subclade(s) and

corresponding backbone(s) from the set of preselected subclades, (3) applying a set of

diversification models (see section 2.2.2) to each part (subclades and backbone) of each

combination to calculate global likelihoods and the corrected Akaike Information Criterion

(AICc) for each combination and comparing their statistical support, and (4) computing the

palaeodiversity dynamics with deterministic or probabilistic approaches. The four steps of the

whole analysis are illustrated in **Figure 1**. By default, two constraints are applied: (1) $\lambda$ and $\mu$

must be strictly positive, and (2) r ($\lambda$ - $\mu$) must be positive (r > 0) at the beginning of the

diversification history(otherwise the clade goes extinct immediately).

Figure 1. Conceptualization of the analytical pipeline for the automation of the method presented in Morlon et al. (2011). A) Step 1 calculates sampling fractions (*f*) from a database that contains all described species of the clade of interest and monophyletic subgroups (based on taxonomy or ecology); B) Step 2 calculates all the possible combinations of shifts (*e.g.* combination 1 has two shifts Clade A and D, the rest of the tree is included into the backbone tree); C) Step 3 applies a model comparison to select the best model for each part (subclades and backbone) of each combination. Then, it computes the likelihood and AICc of each combination, when using the best models. Finally, it selects the best supported combination. The final result is a list containing four elements: the model comparison for (1) the phylogeny analysed as a whole, (2) all subclades apart, (3) all backbones apart, and (4) the

comparison of all combinations of heterogeneous diversification models based on their global

AICc (see Appendix S4 for example). D) Step 4 computes palaeodiversity dynamics either

with the deterministic (function *paleodiv*) or probabilistic approach (function

*apply_prob_dtt*).

**2.2.2 | Set of diversification models**

Several features have been implemented to perform different analyses. The minimum

clade size for applying diversification can be defined at each step and should be the same

during the application of the whole pipeline (argument clade.size, default is 5 tips). The set of

diversification models applied by default in the model comparison corresponds to the six

following birth-death models: (1) constant pure-birth model (BCST), (2) constant birth-death

model (BCST_DCST), (3) pure-birth model with exponential time-dependency such as $\lambda(t) =$

$\lambda_0 \times \exp(\alpha \times t)$ (BVAR), (4) birth-death model with exponential time-dependency for

speciation rate and constant extinction rate (BVAR_DCST), (5) birth-death model with

constant speciation rate and exponential time-dependency for extinction rate such as $\mu(t) =$

$\mu_0 \times \exp(\beta \times t)$ (BCST_DVAR), and (6) birth-death model with exponential time-dependency

for both speciation rate and extinction rate (BVAR_DVAR). A subset of these models can be

used for all parts (subclades and backbones) with the argument models or only for subclades

if, for example, extinction is not expected among subclades but only in the backbone(s) (see

help files in RPANDA).

**2.2.3 | Testing shifts at crown or stem age**

Originally, the probability of the speciation event at the stem age was included in the

likelihood of the backbone, while the probability that the stem of the subclade survives to the

crown age was included in the likelihood of the subclade. Hence, the shift of diversification is

located at the stem age. However, when phylogenies are usually analysed with diversification

models, the stem branch is rarely included. Moreover, adaptive or rapid radiations are

expected to start at the crown age of the clade. We implemented the possibility to test for

shifts at crown ages of subclades by including the probability that the stem of the subclade

survives to the crown age in the likelihood of the backbone. In other words, it means that the

change in diversification rates synchronously occurs with the cladogenesis event as it could be

expected under an adaptive radiation. Users can choose to test for a shift at the stem age

(backbone.option="stem.shift") as in Morlon et al. (2011), or at the crown age of the subclade

(backbone.option="crown.shift", see Appendix S1 for more details).

**2.2.4 | Multiple backbones**

Shifts of diversification can also be nested creating combinations of shifts with more

than one backbone (multiple backbones) in the phylogeny (Figure 1B, combination 3). We

recommend to start investigating the heterogeneity of diversification with a simple backbone

(no shift in the backbone) to first capture the main tendency of the backbone (option

multi.backbone=F in function shift.estimates). This might be interesting if there are a lot of

combinations with multiple backbones. Combinations with multiple backbones can then be

analysed using the option multi.backbone=T or all combinations can be analysed at once with

the option multi.backbone="all".

**2.2.5 | Palaeodiversity dynamic estimation and constraints**

Morlon et al. (2011) used a deterministic calculation of the palaeodiversity dynamics

corresponding to the solution of the differential equation relating the time variation in the

number of lineages to the rate values of the birth-death model (see Box 4 in Morlon 2014 for

more details). Our automation also includes recent developments on palaeodiversity dynamic

estimates (Billaud et al. 2020), a probabilistic approach to infer a confidence interval around

the palaeodiversity curve. Functions that are already included in the R package RPANDA 2.0

(Morlon et al. 2016) have been slightly modified to be used with outputs of the

shift.estimates() through a new function called apply_prob_dtt(). Palaeodiversity dynamics

are computed independently for subclades and backbones. Two optional constraints are

embedded in the optimisation process: maximum rate value and maximum number of species

estimates can be constrained with a maximum value (argument rate.max and n.max) for each

part of the combinations. The constraint applied on palaeodiversity uses the deterministic

approach to calculate maximum values of palaeodiversity dynamics and penalised the

likelihood during the optimisation process if the maximum values exceed the constraint.

**2.3 | Simulations**

We used simulations to test whether the approach successfully detects diversification shifts

and correctly estimates palaeodiversity dynamics (**Figure 2**). We simulated four

diversification scenarios following Billaud et al. (2020). All trees were simulated forward in

time over 100 million years (Myrs, $T_{max}=100$) with the R package TESS 2.1.0 (Höhna et al.

2016). For both waxing-waning models, $\alpha$ and $\beta$ values were defined such that the diversity

decline starts between the last third and the last quarter of the phylogenetic tree (*i.e.* between

33 and 25 Myrs ago in our case). Rate values were selected such as they respect the following

conditions:

1  Expanding diversity: constant pure-birth model (BCST) with $\lambda \in [0.05; 0.25]$ and $\mu=0$

(**Figure 2A**).

2  Saturing diversity: pure-birth model with exponential time-dependency (BVAR) with

$\lambda_{(t)}=\lambda_0 \exp(-\alpha t)$ with $\lambda_0 \in [0.05; 0.25]$ and $\alpha$ such that $T_{eq}$ satisfies $\exp(-\alpha T_{eq})=0.5$ is in

$[T_{max}2/3; T_{max}3/4]$ and $\mu=0$ (**Figure 2B**).

3 Waxing-waning (with decreasing speciation): birth-death model with exponential

time-dependency for speciation rate and constant extinction rate (BVAR_DCST) with

$\lambda_{(t)}=\lambda_0\exp(-\alpha t)$ with $\lambda_0 \in [0.2; 0.4]$ and $\mu=\lambda_0/2$ and $\alpha$ such that $T_{eq}$ satisfies $\exp(\alpha T_{eq})=\mu/\lambda$

in $[T_{max}2/3; T_{max}3/4]$ (**Figure 2C**).

4 Waxing-waning (with increasing extinction): birth-death model with constant

extinction rate and exponential time-dependency for speciation rate (BCST_DVAR) $\lambda$

$\in [0.2; 0.4]$ and $\mu(t)=\mu_0\exp(\beta t)$ with $\mu_0=\lambda/2$ and $\beta$ such that $T_{eq}$ satisfies $\exp(\beta T_{eq})=\lambda/\mu$ in

$[T_{max} 2/3; T_{max}3/4]$ (**Figure 2D**).

We retained the simulated trees containing at least 40 species. We simulated the

diversification shift by grafting a new simulated tree at a specific node defined as the more

recent node among nodes between 33 and 25 Myrs ago (**Figure 2F**). This node must have a

minimum of five extant descendant species to avoid applying diversification models on a

species-poor clade when testing for the detection of a shift on raw phylogenies (*i.e.* when

there is no shift). For trees with a shift of diversification, the grafted tree follows a pure-birth

model (BCST) with $\lambda=0.1$ (**Figure 2E**). This tree was grafted at its crown age and must

contain at least 15 species. Because the shift was simulated at the crown age, the search for

shift was also conducted at crown ages.

We quantified whether the model detected a shift of diversification while no shift was

simulated (*i.e.* false detection of shifts). Following Billaud et al. (2020), we also calculated the

global error $D$ between the estimated and the expected diversity curves, calculated with the

deterministic approach for each simulation (see Appendix S2 for details). Simulations were

set up with Snakemake 5.2.4 and R 3.6.3 using the R packages ape 5.5 (Paradis et al. 2019),

phytools 0.7-80 (Revell 2012), TESS 2.1.0 (Höhna et al. 2016), phangorn 2.7.0 (Schliep

2011), and RPANDA 2.0 (Morlon et al. 2016, see Appendix S2 for more details).

**Figure 2. Simulation pipeline.** A) to D) Diversification rates through time for the models used to simulate the backbone (grey). E) Diversification model used for the subclade (orange). F) Example of a simulated phylogeny with a rate shift shown in orange (the backbone is simulated with the Model 1).

## 2.4 | Applications to empirical phylogenies

We tested the automated pipeline on four clades representing a large diversity of taxa

to illustrate the potential of the approach. First, we applied the approach to the original

example of Morlon et al. (2011), the cetacean phylogeny published by Steeman et al. (2009).

To compare with previous results, we used the same taxonomy as in Morlon et al. (2011) with

a global sampling fraction of 98% (87/89 species). We looked for diversification shifts for the

four main families (Balaenopteridae, Delphinidae, Phocoenidae and Ziphiidae) and for the

two parvorders (Odontoceti and Mysticeti). We tested shifts at the crown age

(backbone.option="crown.shift")      of      subclades      but      also      at      the      stem      age

(backbone.option="stem.shift" see Appendix S2) as in Morlon et al. (2011) to compare both

approaches.

Second, we tested a group without expected diversity decline, and studied the

phylogeny of Vangidae (Jønsson et al. 2012), which is a small but fully sampled group (32/32

species). In their study, the authors detected two diversification shifts and concluded that this

group has experienced adaptive radiation involving changes in beak size and body mass. We

thus tested these hypotheses.

Third, we applied the approach to Parnassiinae, a subfamily of swallowtail butterflies

(Condamine et al. 2018) with a good sampling fraction (96.6%, 85/88 species). Shifts of

diversification can also come from ecological changes such as colonisation of new

environments (Yoder et al. 2010) or host-plant shifts (Ehrlich and Raven 1964). We tested

hypotheses related to host-plant shifts, Himalayan Mountain colonisations, and taxonomic

groups.

Finally, we investigated the diversification history of Cycadales (Nagalingum et al.

2011). The sampling fraction of this group is equal to 72.5% (231/327 species, IUCN 2021).

This group originated 275 Myr ago and almost all extant species emerge from recent

radiations during the Miocene (Condamine et al. 2015). We tested shifts for all genera with

more than five species described except *Dioon*, which is the least sampled (8/13 species). To

account for uncertainty in divergence time estimates, we applied the approach to posterior

trees and calculated the corresponding range of the palaeodiversity dynamics with the

deterministic approach. This allows measuring variation of palaeodiversity estimates on a

large panel of posterior trees and therefore accounting for uncertainties on divergence time

estimates. We also calculated the palaeodiversity dynamics with the probabilistic approach to

account for uncertainty of diversity estimates (Billaud et al. 2020).

For all empirical cases, we did not use the BVAR_DVAR model because it often

produced unrealistic estimates and has already been discussed as problematic (Burin et al.

2019). We also detailed an example with the BVAR_DVAR model in Appendix S4 to

illustrate our choice.

**2.5 | Posterior predictive analyses**

To support the results on empirical phylogenies, we implemented posterior predictive

analyses to test whether data simulated from empirical results are similar to empiricial data.

For each empirical case where shifts of diversification are identified, we simulated 500

phylogenies with the same combination of shifts that the best result by using the crown age

and the diversification rates of subclade(s) and backbone(s). We developed the

simul_posteriors() function to simulate those phylogenies, which are then analysed by looking

at the proportion of points of the empirical LTT contained in the 95% confidence interval of

the LTT from the simulated trees (explained in detail in Appendix S4 for each case).

**3 | Results**

**3.1 | Simulations**

The simulation plan was set up to test both shift detections and palaeodiversity dynamic

estimates. The analyses of the simulations initially revealed high proportions of false

detections of shift (>11%, **Figure 3**, light orange). These false detections of shift led us to

look at what they would be without the model selection procedure, *i.e.* if we fitted only the

true (simulated) models rather than selecting the best model from the set of models applied to

subclades, backbones and the whole phylogeny. By doing this, proportions of false detections

of shift dropped below 10% (**Figure 3**, light blue). This finding sheds light on two problems:

(1) the model selection, and (2) even with the simulated models, proportions of false

detections of shift are higher than 5%. This latter problem shows that when a homogeneous

model is simulated, splitting the optimisation of this model into a subclade and a backbone

provides a gain of likelihood at the subclade level that is not penalised enough to avoid

selecting a shift. We solved this by arbitrarily adding an extra penalisation with an additional

parameter for the shift in the AICc calculation. This additional parameter can be thought as

corresponding to the shift location as in previous methods such as MEDUSA (Alfaro et al.

2009); although the locations of the shifts are less free to vary in our approach that in

MEDUSA, where all nodes are tested, they still need to be selected among a set. With this

additional parameter and when fitting the models that were simulated, proportions of false

detections of shift decreased below 5% (**Figure 3**, dark blue). If we use this additional

parameter but used models that were selected rather than those that were simulated,

proportions of false detections of shift are equal to 3.3% for BCST, 2.4% for BVAR, 13.8%

for BVAR_DCST, and 22.7% for BCST_BVAR (**Figure 3**, dark orange). These results are

satisfactory for simple models (BCST and BVAR) but false detections of shift are still

important for more complex models.

**Figure 3. Proportions of false detection of shifts by models used to simulate the backbone depending on the method.** Values for best-fitting models (in shades of orange) include the step of model selection for each part (subclade and backbone) while values for simulated models (in shades of blue) are calculated using the simulated models (no model selection). "Old version" corresponds to the AICc calculation from Morlon et al. (2011) and "+1 parameter" corresponds to the addition of one parameter for the location of the shift.

Concerning palaeodiversity dynamics, the main tendency (decline or not) is well recovered (BCST=97.35%, BVAR=84.6%, BVAR_DCST=89.5%, BCST_DVAR=92.65%).

More precisely, we compared palaeodiversity curves from simulated rate values with

palaeodiversity curves from estimated rate values. Figure 4 represents these comparisons for

simulations that get a global median error $D$ equal to the three quartiles of $D$ for each model.

We found rather low global median errors $D$ for simpler models (BCST and BVAR models)

(**Figure 4**). More complex models have slightly higher median values of $D$ (BVAR_DCST

and BCST_DVAR models), **Figure 4**). First and third quartiles follow the same tendency as

the median across models, with higher values of $D$ for more complex models. The mean of

median values of $D$ across the different models is equal to 0.233. There is no specific bias

neither towards overestimation nor underestimation of palaeodiversity (Appendix S3).

Figure 4. Comparison of inferred versus simulated palaeodiversity dynamics by model, with increasing global error *D*. For each model (in row) inferred palaeodiversity (purple) and simulated palaeodiversity (in green) are compared for the three quartiles of the global error *D* distributions by model (in column); see Appendix S3 for details. Note that the inferred palaeodiversity with the BVAR model (median) corresponds to a false detection of shift with the selection of a BVAR DCST model for the backbone creating this decline.

## 3.2 | Empirical evidence of extinction from phylogenies

### 3.2.1 | Cetacea

As in Morlon et al. (2011), we tested the presence of shifts in diversification at the

base of the four main and recently diverged families (Balaenopteridae, Delphinidae,

Phocoenidae and Ziphiidae) and the two parvorders (Mysteceti and Odotonceti) of Cetacea.

This selection of subclades creates 25 combinations of shifts with a single backbone and 47

combinations of shifts if we add combinations with multiple backbones. When testing the

shifts at the crown of subclades, we detected a single best combination with shifts for the four

families tested (Table S1, $\Delta$AICc=8.50). All the subclades follow a model with a decreasing

speciation rate through time (see Appendix S4 for details). The simple backbone of this

combination follows a model of diversification with a constant speciation rate and an

increasing extinction rate leading to a phase of decline starting at 10 Ma before present, with a

peak of palaeodiversity at 596 species (**Figure 5**). The same analysis with shifts tested at stem

ages provides a best combination with a lower global AICc (Appendix S1, Table S5).

**Figure 5. Diversification shifts and palaeodiversity dynamics estimated for Cetacea with shifts tested at crown ages.** A) The phylogeny of Cetacea with shifts highlighted in colours and best models in parenthesis. Red dots correspond to all tested nodes. B) The evolution of diversification rates through time. C) The palaeodiversity dynamic estimated with the probabilistic approach. Dotted line represents the confidence interval of diversity estimates for the backbone at 95% calculated with the probabilistic approach. For the sake of clarity, confidence intervals of diversity estimates for subclades are not represented.

### 3.2.2 | Vangidae

We tested shifts at three nodes in the Vangidae phylogeny corresponding to the Malagasy Vangidae clade and two nodes where shifts of morphological evolution have been detected in a previous study (Jønsson et al. 2012). Because these nodes are nested and there are only 32 species, only five combinations result from this initial selection of subclades. We did not detect any shift of diversification, and a model for the phylogeny without shift remains better (Table S2, ΔAICc=3.38) than other combinations. The whole phylogeny follows a model with a decreasing speciation rate resulting in a palaeodiversity increase through time, which slows down to present and reaches the current diversity of 32 species (**Figure 6**).

**Figure 6. Diversification shifts and palaeodiversity dynamics estimated for Vangidae**

**with stems included in backbone analysis.** A) The phylogeny of Vangidae without shift

detected (crown shifts). B) Diversification rates through time. C) The palaeodiversity dynamic

estimated with the probabilistic approach. Dotted line represents the confidence interval of

diversity estimates for the backbone at 95% calculated with the probabilistic approach.

**3.2.3 | Parnassiinae**

We tested 11 shifts corresponding to tribes (Luehdorfiini, Parnassiini, Zerynthiini), genera

(*Allancastria, Parnassius*), subgenera (*Driopa, Tadumia, Kailasius, Koralius, Parnassius*), of

which some match with host-plant shifts (genus *Parnassius*) or colonisation of Himalayan

mountains (*Kailasius, Koramius*). This selection creates 392 combinations to compare with a

single backbone and 1079 by adding combinations with multiple backbones. First results

showed unrealistically high rates of diversification leading to unlikely palaeodiversity curves

highlighting limits of some models (see Appendix S4 for details). We then repeated the

analysis by constraining the rates to a maximum value of 2 and found three best combinations

among the 1079 combinations (Table S3, ΔAICc=2.13). Among these three best

combinations, the first and the third combination of shifts showed highly unrealistic

palaeodiversity dynamics despite the applied constraint on rate values. Given that the three

combinations are statistically equivalent (ΔAICc < 2), we only kept the second combination

that was more biologically realistic (ΔAICc=1.11 with the first combination, see Appendix S4

for details). Among the tested shifts, seven shifts compose the best combination with tribes

(Luehdorfiini, Zerynthiini) and subgenera (*Driopa, Tadumia, Koramius, Kailasius* and

*Parnassius,* **Figure 7**). The tribes Luehdorfiini, Zerynthiini and both subgenera *Parnassius*

and *Tadumia* follow a model with a constant speciation rate while subgenera *Driopa,*

*Kaliasius* and *Koramius* follow a model with a decreasing speciation rate through time. The

backbone follows a model with a constant speciation rate and an increasing extinction rate through time (**Figure 7b**), leading to a period of negative net diversification rate from the beginning of the Miocene (19 Myrs ago) from the peak of the palaeodiversity of 324 species to the diversity of 88 extant species. This creates a palaeodiversity decline at the beginning of the Miocene.

**Figure 7. Diversification shifts and palaeodiversity dynamics estimated for Parnassiinae with stems included in backbone analysis.** A) The phylogeny of Parnassiinae with shifts highlighted (crown shifts) in colours and best models in parenthesis. Red dots correspond to all tested nodes. B) Diversification rates through time. C) The palaeodiversity dynamic estimated with the probabilistic approach. Dotted line represents the confidence interval of

diversity estimates for the backbone at 95% calculated with the probabilistic approach. For

the sake of clarity, confidence intervals of diversity estimates for subclades are not

represented.

**3.2.4 | Cycadales**

We tested five shifts with the genera *Cycas, Ceratozamia, Zamia, Macrozamia*, and

*Encephalartos*. This selection creates 32 combinations with a simple backbone and none with

multiple backbones. The best combination is composed of all the tested subclades (Table S4,

ΔAICc=11.96). The genera *Cycas, Ceratozamia* and *Macrozamia* follow pure-birth models

with a constant speciation rate, while *Zamia* and *Encephalartos* follow models with a

decreasing speciation rate (**Figure 8a,** Appendix S4). The diversification model of the

backbone has a constant speciation rate and an increasing extinction rate, creating a negative

net diversification rate starting 130 Myrs ago during the Early Cretaceous (**Figure 8b**). The

global palaeodiversity increased again and reached the extant diversity thanks to the

radiations of extant genera during the Miocene. The peak of palaeodiversity calculated with

the deterministic approach reaches 200 species and ranges on posterior trees from 500 to 100

species (**Figure 8c**). The probabilistic approach estimates 150 species at the palaeodiversity

with a confidence interval ranging from 400 to 0 species (**Figure 8d**).

**Figure 8. Diversification shifts and palaeodiversity dynamics estimated for Cycadales with stems included in backbone analysis.** A) The phylogeny of cycads with shifts highlighted (crown shifts) in colours and best models in parenthesis. Red dots correspond to all tested nodes. B) The evolution of diversification rates through time. C) The palaeodiversity estimated with the deterministic approach. The red area represents the range of palaeodiversity estimates calculated on posterior trees. D) The palaeodiversity dynamic estimated with the probabilistic approach. Dotted line represents the confidence interval of diversity estimates for the backbone at 95% calculated with the probabilistic approach. For the sake of clarity, confidence intervals of diversity estimates for subclades and species names are not represented.

## 3.3 | Posterior predictive analyses

The posterior predictive analyses allowed us to check how empirical data were close to a set

of trees simulated from the parameters of the best combinations of shifts for each empirical

case. Overall, 98.7% of the empirical LTT values of the Cetacea falls into the 95% confidence

interval defined on the LTT from the posterior predictive trees (Figure S5A). This statistic is a

bit lower for Parnassiinae with 82.5% and slightly higher for Cycadales with 99.6% (Figure

S5B and C; see Appendix S4 for details).

**4 | Discussion**

508            In this study, we automated and implemented in RPANDA, tested with simulations,

and illustrated with empirical cases the approach developed by Morlon et al. (2011) that

allows detecting shifts of diversification and estimating palaeodiversity dynamics through

time from dated phylogenies. Several options and functions have been added to make this

approach largely accessible to address an array of questions about heterogeneity of

diversification rates and palaeodiversity dynamics by consolidating the challenging field of

macroevolution in adopting hypothesis-driven approaches (Morlon et al. 2022). This

emphasises the importance of additional information from biological knowledge on the focal

groups to compare scenarios motivated by biological hypotheses (*e.g.* Legendre and

Condamine 2018). In this sense, integrative taxonomic work is fundamental as it summarises

broad morphological, ecological, and geographical divergences without neglecting historical

divergence of clades thanks to phylogenetic analyses. Beyond taxonomy, ecological data, or

other analyses such as ancestral state estimations can ground hypotheses for selecting

subclades if the sampling fraction is known.

522            From a biological point of view, diversification histories are more complex than birth-

523    death models. Accordingly, one should keep in mind that underlying assumptions of models

524    influence the results. Among time- and clade-heterogeneous approaches, the most popular

525    model, BAMM, investigates shifts along branches and does not allow periods of decline of

diversity (Rabosky et al. 2013). The Multi-Type Birth-Death model developed by Barido-Sottani et al. (2020) provides a Bayesian framework that investigates shifts thanks to type-specific birth and death rate without defining location of the type changes. ClaDS models rate heterogeneity with small but frequent shifts (Maliet et al. 2019). MiSSE is an extension of HiSSE models that focuses on trait-free shifts and investigates heterogeneity of diversification rates at the tips of phylogenies (Vasconcelos et al. 2022). These alternative approaches can be used in a complementary way to study heterogeneity of diversification but will not necessarily provide similar results because of their underlying assumptions on diversification rates. Globally, each model is better at finding what it is made for: BAMM performs better to find large discrete shifts and will underestimate finer heterogeneity while ClaDS and, to a lesser degree MiSSE, will perform relatively worst in looking for large shifts but will be more accurate for small and more frequent changes of diversification rates (see Vasconcelos et al. 2022 for a comparison between BAMM, ClaDS and MiSSE). In our model, rates of birth-death models are either constant or exponentially variable over time and shifts occur at specific branch points. Despite the uncertainty being higher deep in the past, the main advantage of our approach is to provide inference of palaeodiversity dynamics by better estimating palaeodiversity values during deep periods, with the probabilistic approach to calculate palaeodiversity dynamics (see Billaud et al. 2020).

Previous studies on Vangidae (Jønsson et al. 2012), on Cetacea (Rabosky 2014), on Parnassiinae (Condamine et al. 2018), and on Cycadales (Condamine et al. 2015) have investigated heterogeneity of diversification with other methods. The cladogenesis test used for the Vangidae found one significant and one marginal shift among the nodes we tested. Contrary to the cladogenesis test that looks at the number of descendants from a node relatively to what is expected under a null birth-death model, our method applies and compares various birth-death models. Finding no shift can then come from the small size of

the phylogeny but might also highlight that our approach is more conservative in detecting

shifts in small phylogenies. Interestingly, BAMM analyses on the three other empirical cases

estimated less diversification shifts (one in Cetacea vs. four here, one in Parnassiinae vs.

seven here, and four in Cycadales vs. five here). This confirms that BAMM might be

conservative in rate shift detection, although it has been shown to concern shifts with small

clade sizes (Mitchell et al. 2019). Moreover, BAMM did not detect palaeodiversity decline

contrary to our inferences. For instance in Parnassiinae, we found a decline starting at 19 Ma

from a peak at 324 species to the current diversity of the backbone (8 species). This is

consistent with Condamine et al. (2018) who applied models assuming instantaneous drops of

diversity through time at specific time points (Stadler 2011a; May et al. 2016). This study

found a sudden extinction around 15 Ma, 4 Myrs later, with a loss of 89% of the diversity

(from 90 species to 10). Both methods recover a decline of diversity, although different

assumptions behind each model (mass extinction vs. exponential dependency to time)

influence the inferred timing and the scale of the decline.

The approach we automated has already been recognized for its ability to better

estimate extinction rates (Stadler 2011b). Our simulations confirm the capacity of the model

to recover non-zero extinction rate in the backbone that can exceed speciation rate, resulting

in a diversity decline. Despite the high values of false detection of shifts for complex models,

the difference between observed and simulated palaeodiversity was reasonable with a mean of

median global error $D$=0.233 on average over models and close to the value found in Billaud

et al. (2020) with a $D$=0.22. This result shows that even with an error on the detection of shift,

the model estimates a trustworthy palaeodiversity dynamic (**Figure 4**). Moreover, our

empirical analyses, confirmed by posterior predictive analyses, suggest diversity declines are

more widespread than previously considered. Cetacea, Parnassiinae, and Cycadales show a

common pattern with a single period of decline, but we can address the question to what

extent patterns with many diversity declines might be detected. Indeed, the main limit of our

approach is that important levels of extinction are likely to erode both diversity and statistical

power. Splitting a phylogeny into backbones and subclades requires enough species to be

shared in different parts that will be analysed separately. Dated phylogenies with only a few

extant species cannot be investigated by our approach although we sometimes know from the

fossil record that they experienced extinction. The clade of Ginkgophyta is a caricatural

example of this problem as it only contains one lineage (*Ginkgo biloba*) despite 275 Ma of

evolution, and probably experienced a lot of extinction events (Royer et al. 2003). An

interesting perspective will be to simulate more complex patterns of extinction with multiple

backbones and multiple periods of decline to see whether we can properly recover these

patterns.

Concerning subclades, we recovered only models without extinction, although some

subclades are quite ancient (*e.g.* in Parnassiinae). We ran some simulations that tend to

confirm the difficulty to select models with extinction in subclades (see Appendix S3 for

details). This is not surprising given the small size of subclades (*e.g.* six species of

Phocoenidae in Cetacea), which do not provide sufficient sample size to select parameter-rich

models (i.e. with extinction). Low extinction rates in subclades can also be a real pattern in

some cases. For the Cetacea, the phylogenetic placement of fossils indicates that extinct taxa

are concentrated in the backbone (in particular along stems of the four main families (Lloyd

and Slater 2021), which is consistent with our results indicating that the backbone bears most

of the extinction signal in the tree.

Given the difficulty to estimate extinction rates (Morlon 2014; Beaulieu and O'Meara

2015; Rabosky 2016), studies have repeatedly raised debates about models in macroevolution

(Kubo and Iwasa 1995; Rabosky and Lovette 2008; Crisp and Cook 2009; Burin et al. 2019;

Pannetier et al. 2021). Recently, Louca and Pennell (2020) argued that diversification rates

should not be estimated from birth-death models with clade-homogeneous rates because any

reconstructed phylogeny can be equally well explained with an infinite number of alternative

models. Here we adopt a hypothesis-testing approach, in which we consider a limited set of

alternative models chosen *a priori*, some of which will explain the phylogeny better than

others (Helmstetter et al. 2022, Morlon et al. 2022). Even when allowing for potential rate

shifts in only few clades, our approach sometimes compared thousands of combinations of

shifts; still, cases with many statistically equivalent combinations were rare. Methodological

developments and collecting more data (*e.g.* to update taxonomy, to increase sampling

fractions) are required to investigate questions in such a challenging field while keeping in

mind that rates inferred from extant time-trees are tentative estimates of a long history and

should be interpreted cautiously (Helmstetter et al. 2022).

In conclusion, the automated implementation of the approach originally developed by

Morlon et al. (2011) provides a user-friendly tool to investigate heterogeneity of

diversification. We improved this approach by (1) adding the possibility to test shifts at crown

age that seems to drastically improve the global likelihood of the approach (see Cetacea for

example), (2) select a set of constant or time-dependent diversification models, (3) easily set

up multiple backbone investigation, (4) make it possible to constrain rate estimates, (5) infer

palaeodiversity dynamics while considering uncertainty both within rate estimation

(probabilistic approach) and across many trees that can differ in terms of topology or

divergence times, and (6) confirm empirical results with posterior predictive analyses.

Nonetheless, we encourage users to adopt a hypothesis-driven approach and to be

parsimonious concerning the number of nodes to test. We also recommend being critical on

results of inferred palaeodiversity of the backbone in terms of raw number of species that

results from rate values. We also caution that the most complex model where both speciation

and extinction rates vary exponentially (BVAR_DVAR) is prone to provide highly unrealistic

patterns of palaeodiversity dynamics as shown in an example with Cetacea (Appendix S4).

Echoing the results from Burin et al. (2019), the BVAR_DVAR model should be avoided or

constrained. Keeping in mind these limitations, our study highlights that extinction signature

and palaeodiversity declines can be investigated in groups without species-level fossil

records.

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**Conflict of interest**

None declared.

**Authors' contributions**

N.M. and F.L.C. designed the research with advice from P.-H.F.; N.M. performed the

research and analysed the data with advice from H.M. and F.L.C.; and N.M. wrote the

manuscript, and all authors edited the text.

**Data availability statement**

This open-source software is written entirely in the R language and is freely available through

the Comprehensive R Archive Network (CRAN) at:

https://cran.r-project.org/web/packages/RPANDA/index.html. The scripts used, simulation

results, model fits and case study data are available at:

https://figshare.com/s/6e8924774165b5ed81ec

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