1 Complexities of recapitulating polygenic effects in natural populations: replication of genetic
2 effects on wing shape in artificially selected and wild caught populations of *Drosophila*
3 *melanogaster.*
4
5

6 Katie Pelletier*[1], William R. Pitchers[2]*, Anna Mammel[2,7], Emmalee Northrop-Albrecht[2,5], Eladio
7 J. Márquez[3,6], Rosa A. Moscarella[3,4], David Houle[3], Ian Dworkin[1,2]
8 *Co-first authors
9 Corresponding authors: dworkin@mcmaster.ca, dhoule@bio.fsu.edu

10 Author affiliations: 1- Department of Biology, McMaster University

11   2 - Department of Integrative Biology, Michigan State University

12   3- Department of Biological Science, Florida State University

13   4-Current address: Department of Biology, University of Massachusetts

14   5- Current address Division of Gastroenterology and Hepatology, Mayo
15 Clinic, Rochester MN

16   6 – Current address: Dragonfly Therapeutics, Waltham MA
17
18   7-Current address: Neurocode USA. 3548 Meridian St, Bellingham, WA
19 98225

Abstract

Identifying the genetic architecture of complex traits is important to many geneticists, including those interested in human disease, plant and animal breeding, and evolutionary genetics. Advances in sequencing technology and statistical methods for genome-wide association studies (GWAS) have allowed for the identification of more variants with smaller effect sizes, however, many of these identified polymorphisms fail to be replicated in subsequent studies. In addition to sampling variation, this failure to replicate reflects the complexities introduced by factors including environmental variation, genetic background, and differences in allele frequencies among populations. Using *Drosophila melanogaster* wing shape, we ask if we can replicate allelic effects of polymorphisms first identified in a GWAS (Pitchers et al. 2019) in three genes: *dachsous (ds)*, *extra-macrochaete (emc)* and *neuralized (neur),* using artificial selection in the lab and bulk segregant mapping in natural populations. We demonstrate that multivariate wing shape changes associated with these genes are aligned with major axes of phenotypic and genetic variation in natural populations. Following seven generations of artificial selection along the *ds* shape change vector, we observe genetic differentiation of variants in *ds* and genomic regions containing other genes in the hippo signaling pathway. This suggests a shared direction of effects within a developmental network. We also performed artificial selection with the *emc* shape change vector, which is not a part of the hippo signaling network, but showed a largely shared direction of effects. The response to selection along the *emc* vector was similar to that of *ds*, suggesting that the available genetic diversity of a population, summarized by the genetic (co)variance matrix (**G**), influenced alleles captured by selection. Despite the success with artificial selection, bulk segregant analysis using natural populations did not detect these same variants, likely due to the contribution of environmental variation and low minor allele frequencies, coupled with small effect sizes of the contributing variants.

Introduction

Dissecting the genetic architecture underlying complex traits remains challenging because of the joint contributions of many alleles of small effect and genotype-by-environment interactions, among other factors. Progress in sequencing technology in conjunction with development of GWAS statistical methodologies has enabled identification of loci contributing to numerous complex traits and diseases. However, such mapping approaches identify only a subset of loci contributing to trait variation (Visscher et al., 2017). In part, this reflects the low power to detect rare alleles and those with small effects (Tam et al., 2019). For alleles that are relatively common in a population, replication rates between GWAS studies are high, even when effect sizes are small (Marigorta et al., 2018). However, GWAS studies have failed to replicate the effects observed in many candidate gene studies, in part due to the fact that many of the alleles identified in these studies are rare in a population and require very large cohorts to detect (Fritsche et al., 2016; Ioannidis et al., 2011).

In cases where an association is replicated between studies, the magnitude of the effect can vary substantially between different cohorts or populations (CONVERGE consortium, 2015; Marigorta et al., 2018). Differences can arise because of genetic background due to epistatic gene by gene (*GxG*) interactions, or due to gene-by-environment (GxE) interactions. The initial estimates of effect size will be biased upwards if statistical testing in the initial cohort is used to determine which SNPs are chosen for replication studies. It is important to understand which of these causes of differences in effect size are of practical significance when we want to generalize results to different populations or environments.

In this study, we focus on the issue of replication in a multivariate context, where the joint inheritance of multiple features are simultaneously investigated. We will refer to the suite of measured features as a 'multivariate trait' for convenience. In this case, what we want to estimate is the vector of effects of each SNP on all measured features. Each SNP may have a unique combination of effects. Univariate effects vary only in magnitude, as we can only infer effects on a single feature. For a multivariate trait, estimated genetic effects vary in magnitude, the sum of effects on all traits, and also in direction, how the total effect is allocated among different features (Melo et al., 2019). The ability to study the direction along with the

magnitude of genetic effects provides an additional and important way of assessing

repeatability. For a univariate trait, there is a 50% chance that the replicate estimate will be in

the same direction as the original estimate, even with no true effect. By contrast, the

probability of a "replicated" genetic effect sharing a similar direction by chance alone decreases

as the number of measured features increases (Marquez and Houle, 2015; Stephens, 2013).

Studying genetic effects in a multivariate context is beneficial in other ways. First, it has

been demonstrated both empirically and via simulations, that genetic mapping for multivariate

traits generally increases statistical power over trait by trait analyses (Fatumo et al., 2019;

Pitchers et al., 2019; Porter and O'Reilly, 2017; Shriner, 2012). Second, some multivariate traits

cannot be sensibly reduced to a single measurement. The wing shape we study is a great

example of such a multivariate trait. We have good reason to believe that wing shape is

important for flight (Ray et al., 2016), but we cannot yet say that any feature, such as wing

length or width, is more or less important than any other. Natural selection on wing shape may

affect any or all combinations of measurements.

Perhaps most importantly, traits are not inherited in isolation, but are the joint outcome

of an integrated developmental process that results in extensive genetic correlations that can

have important effects on evolution. The main source of such correlations are the patterns of

pleiotropic effects generated by mutational effects. Multivariate studies of inheritance allow

pleiotropic effects to be estimated in a rigorous and justifiable manner (Melo et al., 2019). The

multivariate breeder's equation, $\Delta z = G\beta$, enables short term prediction of evolutionary

responses. Key to understanding how populations respond to selection in the short term

requires an understanding of properties of the genetic (co)variance matrix (**G**), and in particular

the axis of greatest genetic variation, $\mathbf{g}_{max}$. Studies demonstrate that the direction of $\mathbf{g}_{max}$

influences evolutionary trajectories (Blows and McGuigan, 2015; McGuigan, 2006; Schluter,

1996). The degree to which genetic effects associated with particular variants align to major

axes of genetic (co)variance, expressed through **G**, may provide insights into which alleles are

most likely to be "captured" by selection (Pitchers et al., 2019).

Due to the polygenic nature of complex traits, including multivariate ones, it is

important to consider not only the direction of effect for alleles in a single gene but also

correlated effects between genes contributing to the phenotype. Interestingly, initial

comparisons of directions of genetic effects among induced mutations in two *Drosophila*

melanogaster wing development pathways showed only partially correlated effects on wing

shape within and between pathways (Dworkin and Gibson, 2006). However, recent work has

demonstrated that despite large differences in magnitude, the direction of genetic effects of

variants segregating in populations are sometimes similar to those from validation experiments

using RNAi knockdown of those same genes (Pitchers et al., 2019). Additionally, Pitchers et al.

(2019) demonstrated this shared direction of effect could also be shared between a SNP and

RNAi knockdown of other genes in the same signaling pathway, such as those involved with

hippo signaling, a key pathway involved with wing growth and morphogenesis (Pan et al.,

2018).

Pitchers et al. (2019) identified over 500 polymorphisms contributing to wing shape

variation in the *Drosophila* genetic resource panel (DGRP). Among these, the hippo pathway

was over-represented in SNPs associated with wing shape (Pitchers et al., 2019). The degree to

which identified hippo pathway variants reflect allele specific effects, differences in magnitude

of genetic effects, and even the large statistical uncertainty associated with genetic effects of

small magnitude are unclear. Given common dominance patterns, and the likely non-linear

genotype-phenotype relationships of most genetic effects, small to moderate changes in gene

function may result in modest phenotypic effects (Green et al., 2017; Melo et al., 2019; Wright,

1934). Large effect mutants and many RNAi knockdown studies have moderate to large

phenotypic effects that are not reflective of the magnitude of genetic effects of SNPs

contributing to phenotypic variance in natural populations.

The expression of genetic effects is also dependent on genetic and environmental

context, with contributions of gene by gene (GxG) and gene-by-environment (GxE) sources of

variation. The context-dependence of genetic effects for a multivariate trait has been

demonstrated for *Drosophila* wing shape. Variants in *Egfr* influencing *Drosophila* wing shape are

replicable in both lab reared and wild-caught cohorts (Dworkin et al., 2005; Palsson et al., 2005;

Palsson and Gibson, 2004). However, in replication studies, effect sizes of alleles were

diminished in both outbred populations and wild cohorts. In the latter case the same variant

explained 1/10 of the phenotypic variance explained in the initial study (Dworkin et al., 2005).

Interestingly, in a series of experimental crosses among strains, the effects of the SNP were

replicable for direction and magnitude in multiple experimental assays and crossing schemes.

Despite this, the genetic effect on wing shape from this SNP largely disappeared in one natural

population (Palsson et al., 2005). A number of reasons have been proposed for the failure to

replicate genetic effects including environmental effects, differences between controlled lab

and natural environments (Dworkin et al., 2005), and genetic background (Greene et al., 2009),

among others. Because both environment and genetic background likely affect the genotype-

phenotype map in a non-linear fashion (Wright, 1934), it is important to test observed

associations in other experimental contexts.

A promising approach to confirm the estimated effects of candidate genetic variants is

to test whether they respond to artificial selection in the direction of the inferred effect. This

approach is particularly relevant to evolutionary questions, but has rarely been used. In this

study, we use both artificial selection and bulk segregant analysis to replicate and validate SNPs

associated with three genes: *dachsous (ds), extra-macrochetae (emc)* and *neutralized (neur)*

previously identified in a GWAS of *Drosophila* wing shape (Pitchers et al., 2019). Using the

vectors of shape change based on RNAi knockdowns of each gene, we demonstrate that the

direction of shape change for these genetic effects is aligned with major axes of natural

phenotypic and genetic variation. Using artificial selection based on the direction of shape

change defined by RNAi knockdown, we were able to replicate the effects observed for *ds,* but

not *emc*, likely due to the available genetic diversity in the population. We then asked if these

effects could be replicated in a natural population using a bulk segregant approach. We discuss

our results in the context of the replicability of genetic effects and the shared direction of

genetic effects due to shared developmental processes.

Methods

Source Populations and phenotypic analysis

Drosophila strains

Phenotype data for the *Drosophila* genetic resource panel (DGRP) was collected for 184

strains as part of a GWAS study as described in Pitchers *et al* (2019). Genotype data for these

strains was obtained from freeze 2 of the DGRP (Huang et al., 2014). For replication using

artificial selection, 30 DGRP strains were used: DGRP-149, 324, 383, 486, 563, 714, 761, 787,

796, 801, 819, 821, 822, 832, 843, 849, 850, 853, 859, 861, 879, 887, 897, 900, 907, 911, 913.

These strains were selected to increase genetic variation at the *ds* locus (Supplemental Figure 1,

Table 1). Reciprocal pairwise crosses between the 30 selected DGRP strains were used to create

heterozygotes and these 30 heterozygous genotypes were successively pooled for 4

subsequent generations, allowing for recombination. After pooling, the synthetic outbred

population was maintained for approximately 47 subsequent generations (allowing for

recombination) before the start of artificial selection experiments.

For the replication in wild-caught populations using bulk segregant analysis (BSA),

individuals were collected via sweep-netting from orchards and vineyards in Michigan and after

species identification, stored in 70% ethanol. In 2013 and 2014, cohorts were collected from

Fenn Valley Winery (FVW13 and FVW14 respectively, GPS coordinates: 42.578919, -

86.144936). Additionally in 2014, cohorts were collected from Country Mill Orchard (CMO, GPS

coordinates: 42.635270, -84.796706), and Phillip's Hill Orchard (PHO, GPS coordinates:

43.117981, -84.624235). For all collected cohorts, except for the FVW14 collection, only males

were used in this study given difficulties differentiating *Drosophila melanogaster* and *D.*

simulans females. For the genomic analysis of the FVW14 wild caught population (below) we

utilized both males and females as the number of individuals was insufficient otherwise. For the

collection where females were included in the study, there is no evidence of contamination

with *D. simulans* as all dissected wings were classified as *D. melanogaster* using linear

discriminant analysis (LDA). The LDA algorithm was trained using male wings from the collected

D. melanogaster data set as well as males from a *D. simulans* data set acquired from the

Hosken lab. There was 100% agreement between the classification of females within each

species with our phenotypic classification, indicating that it is very unlikely that *D. simulans*

females were included in our samples (Supplemental Figure 2).

<u>Morphometric Data</u>

Landmark and semi-landmark data were captured from black and white TIFF images

using the pipeline described in Houle et al. (2003). First, two landmark locations, the humeral

break and alula notch, were digitized using tpsDig2 (version 2.16). Wings (Van der Linde 2004–

2014 , v3.72) software was used to fit nine cubic B-splines and correct errors. All shape data

was subjected to full Procrustes superimposition to scale images to the same size, translate

configurations to the same location and rotate to minimize distance between points (Rohlf and

Slice, 1990). Generalized Procrustes superimposition (registration) and extraction of 14

landmark and 34 semi landmark positions was done with the program CPR v1.11 (Márquez

2012–2014, Figure 1). Sliding of semi-landmarks utilized minimization of Procrustes Distance as

the objective function. Superimposition results in the loss of 4 possible dimensions of variation

while semi-landmarks are constrained to vary over only one axis, restraining these points to

approximately a single dimension of variation. This results in a total of ~58 available dimensions

of shape variation, that can be summarized using the first 58 Principal components (PCs).

Allometry was adjusted for in the analysis by fitting a model for landmark coordinates onto

centroid size, and using the residuals from this model (Klingenberg, 2022). By accounting for

the allometric component of shape, shape variation due to size variation is accounted for

(Supplemental Figure 3). For most analyses, 'allometry corrected' landmarks were used, with

the exception of shape models fit using the Geomorph package in R, where Procrustes

landmarks were used and centroid size was included as a predictor in the model.

<u>Generation of shape vectors for artificial selection and Bulk segregant analysis</u>

A panel of shape change vectors was estimated using the progesterone-inducible

Geneswitch GAL4 construct under the regulation of an ubiquitous *tubulin* driver, to drive the

expression of RNAi for genes of interest, as previously described in Pitchers *et al.,* 2019. GAL4

expression was induced throughout larval development by adding mifepristone, an analog of

progesterone, to the larval food.   Knockdown was varied by assaying phenotypes at

mifepristone concentrations of 0.3, 0.9, and 2.7 µM, plus a control without mifepristone. Wing

shape change associated with knockdown of the gene of interest was estimated using

multivariate regression of shape on concentration of mifepristone. Shape change vectors for

dachsous (*ds*)*, extramacrochaete* (*emc*) and *neuralized* (*neur*) were used in this experiment

(Figure 1b, Supplemental Figure 4). The magnitude ('length') of the vector measures how much

shape change occurs per unit change in mifepristone. In general, vectors of greater magnitudes

enable better estimate of direction of effect for shape change. As reported in Pitchers et al.,

(2019), the magnitude (l^2-norm) of vectors for RNAi knockdown of these genes are 5.5 for *ds*,

2.8 for *neur*, and 0.44 for *emc*.

Shape data collected as part of a previous study (Pitchers et al., 2019) was used to

assess the relationships between shape change vectors from the RNAi titration and $\mathbf{g_{max}}$, the

first eigenvector of the **G** matrix estimated from DGRP line means. The effects of sex, centroid

size and their interaction were removed using a linear model and these residuals were used to

calculate shape score by projecting the data (see Supplemental Figure 5) onto the shape change

vector estimated in each knockdown experiment. To assess major axes of genetic variation

among the DGRP strains, principal component analysis was performed on the allometry

adjusted model residuals (Supplemental Figure 5B). PCA was done in a similar manner for the

individuals from the wild caught cohorts. Correlations between the first three eigenvectors

("genetic PCs" including $\mathbf{g_{max}}$), the first three PCs from the wild caught cohorts and the shape

scores for *ds*, *emc* and *neur* were calculated (Figure 1a, Supplemental Figure 5). From this, *ds,*

emc and *neur* shape change vectors were selected for further experiments given high

correlation with directions of natural genetic variation (Figure 1, Supplemental Figure 5).

<u>Artificial selection of synthetic outbred population</u>

The synthetic outbred population resulting from pooling DGRP lines was used as the

parent population for artificial selection. Both the *ds* and *emc* artificial selection experiment

was carried out with three independent replicates of each "up" and "down" selection regimes,

along with unselected control lineages. Each generation, wings of live flies were imaged using

the 'wingmachine' system and shape data collected (Houle et al., 2003, Van der Linde 2004–

2014 ,v3.72). Shape scores were calculated by projecting the data onto the *ds* or *emc* shape

change vector as described above, and the 40 individuals each with highest or lowest shape

scores, were selected to found the next generation (Supplemental Figure 5A). For the control

lineages, 40 individuals were randomly selected for the next generation within each replicate

lineage. Following seven generations of selection, 75 individuals from each lineage were

selected for pooled sequencing, described below. The response to selection was evaluated both

by computing Procrustes distance (PD) between average shape of wings between generations

one and seven, and using shape scores (projections) with a linear mixed effect model allowing

for the fixed effect factors of treatment and sex, continuous predictors of centroid size and

generation, with third order interactions among these effects. The effect of generation was

allowed to vary by replicate lineages (lmer(ds ~ (CS + Sex + line + gen0)^3 + (1 +

gen0|line:rep)). Realized heritabilities were estimated separately for up and down selection

lineages, from the slope of the regression of cumulative selection differentials on cumulative

selection response, averaging over sex and with a random effect of replicate lineage.

Wild populations

For the Bulk segregant Analysis (BSA), wings for wild caught individuals were dissected

and mounted in 70% glycerol in PBS. Images of wings were captured using an Olympus DP30B

camera mounted on an Olympus BX51 microscope (Olympus software V.3,1,1208) at 20X

magnification. When possible, both left and right wings were dissected, imaged and averaged

to calculate an individual's mean shape. For some individuals a wing was damaged so only one

wing could be used. Shape was captured as described above. The total number of individuals

phenotyped from each cohort can be found in Supplemental table 1.

To remove allometric effects in the data, shape was regressed onto centroid size and

the model residuals were used for all subsequent morphometric analysis. Only data from males

was used to compare shape in wild populations, although, including females from the FVW14

population and regressing shape onto centroid size and sex gave equivalent results

(Supplemental Figure 6). To test for shape differences between collection cohorts, the effect of

centroid size and collection cohort on shape were modeled (procD.lm(shape ~ CS + pop_year))

using the procD.lm function in Geomorph v 3.1.3. (Adams and Otárola-Castillo, 2013) and

distances between populations were calculated using the pairwise function. To select

individuals for sequencing, a 'shape score' was calculated using the method described above.

Shape data was projected onto the vector of shape change defined by the *ds* or *neur*

knockdowns. The *emc* projection vector was not used for bulk segregant analysis due to the

high similarity with *ds* shape change (Figure 1), and the similarity of the selection response. Its

inclusion would result in selection of largely the same cohorts of individuals for sequencing for

both *ds* and *emc*. As an alternative, we utilized the *neur* shape vector as it was largely

uncorrelated with that of *emc* and *ds*, but strongly correlated with natural variation of shape.

The 75 most extreme individuals within each wild-caught cohort on the shape score distribution

were selected for pooled sequencing. Allele frequencies within each population was estimated

by sequencing 75 random individuals within each cohort. The difference vector between mean

shapes of selected pools (within each population) was used to calculate Procrustes distance

(PD) between pools and the correlation of this shape change vector with the selection vector

used. An estimate of genetic distances between populations was calculated using allele

frequencies (mapping pipeline described below) in the pools of the 75 randomly selected

individuals using Bray's distance with the vegdist() function from the vegan package (v2.6-2) in

R.

<u>Sequencing and Genomic Analysis</u>

DNA extractions from pools of selected individuals was performed using a Quiagen

DNeasy DNA extraction kit. Library preparation and Illumina sequencing was performed at the

research technology support facility at Michigan State University. All library samples were

prepared using the Rubicon ThruPLEX DNA Library Preparation kit, without a procedure for

automatic size selection of samples. Paired end libraries (150bp) were sequenced using Illumina

HiSeq 2500, with each sample (either one pool of 75 individuals in the BSA or one pooled

replicate lineage in the artificial selection) being run on two lanes.

Reads were trimmed with Trimmomatic (v0.36) to remove adapter contamination and

checked for quality using FastQC prior to alignment (Bolger et al., 2014). Trimmed reads were

aligned to the *Drosophila melanogaster* genome (v6.23) using BWA-MEM (v0.7.8)(Li and

Durbin, 2010). Sequencing replicates of the same biological samples were merged using

SAMtools (v1.11). PCR duplicates were removed using Picard with the MarkDuplicates tool (v

2.10.3) and reads with a mapping quality score less than 20 were removed using SAMtools (Li et

al., 2009). A local realignment around indels was performed using GATK using the

IndelRealigner tool (v3.4.46). For artificial selection experiments, reads were merged for all up,

down and control selection lines as replicates lineages were independent. For wild cohorts,

pools were not merged between populations. mpileup files were created using SAMtools and

used for subsequent genomic analysis. Highly repetitive regions of the *Drosophila* genome were

identified and subsequently masked in mpileup files using RepeatMasker (v4.1.1) with default

settings. INDELs and regions within 5bp of an indel were identified and masked using

popoolation2 scripts. Population genetic statistics were calculated using PoPoolation (v1.2.2)

and PoPoolation2 (v1.201) (Kofler et al., 2011b, 2011a).

For the bulk segregant analysis (BSA) in the wild-caught cohorts, a modified Cochran-

Mantel-Haenszel (CMH) test was used to measure significantly differentiated sites between

pools of individuals. Sampling effects were accounted for using the ACER package (v.1.0) in R,

assuming $N_e = 10^6$ with 0 generations of differentiation between selected pools (Spitzer et al.,

2020). To adjust for multiple testing, the p-value was corrected using a Benjamini and Hochberg

correction (Benjamini and Hochberg, 1995) with an adjusted alpha of 0.05. For each significant

site from the CMH test, using an adjusted p-value cut-off of 0.05, we identified the nearest

gene using BEDtools (v2.19.1) (Quinlan and Hall, 2010). In addition, to account for sampling

variation, we sampled genomic coverage to 75x for all samples, dropping sites that did not

meet this threshold and repeating the CMH test. We confirmed that there was no association

between genetic and shape differentiation between populations, and that the populations do

not show strong phenotypic differentiation based on either overall shape variation, or shape

scores used to identify selected individuals for BSA (Supplemental Figures 3 and 7). There was

some variation among populations in overall wing size (Supplemental Figure 8), however we

(assuming common allometry) adjusted for allometric effects on shape.

For artificial selection experiments, F_{ST} was calculated in 5000bp windows. We chose

this window size as it is expected that blocks of LD in the synthetic outbred population will be

much larger in comparison to that of the wild caught samples (King et al., 2012a; King et al.,

2012b; Marriage et al., 2014). This statistic was used to compare the "up" selected pools to the

"down" selected pools to help identify regions of differentiation between selected populations.

For the artificial selection comparisons, genes in regions of high F_{ST} were identified by

finding overlaps between outlier windows and annotated *Drosophila* genes using

GenomicRanges (v1.46.1) in Bioconductor. High F_{ST} was defined as F_{ST} values greater than three

standard deviations above the mean. GO terms associated with identified genes were

annotated using TopGO package (v2.34.0) (Alexa et al., 2006) in Bioconductor. GO enrichment

was then performed to identify those terms overrepresented in the identified list using TopGO

and a Fisher's exact test. Over representation of 2 GO terms in outlier windows (hippo

signaling, GO:0035329; negative regulation of hippo signaling GO:0035331) were tested using a

permutation test that randomly sampled genomic windows from the total windows for which

F_{ST} was calculated and the permutation was run 1000 times. The distribution of the ratio of

observed to expected genes annotated with the term of interest within randomly sampled

regions was compared to the number observed in the data.

Verification of *ds* indel in DGRP

Sanger sequencing was performed on individuals from a cross between DGRP lines predicted to

have the polymorphism (DGRP 195, 28, 96, 48, 59, 801) and those without (DGRP 129, 301, 69,

385, 75, 83, 491, 34, 774) crossed to a line carrying a deletion in the region of interest (BDSC

24960) to account for potential residual heterozygosity in otherwise inbred strains. DNA was

prepared by incubating flies in DNA extraction buffer (1mM EDTA, 25mM NaCl, 10 mM TrisHCl

pH 7.5) for 10 minutes, followed by storage at -20 C. PCR application of the region of interest

(Forward primer ggagtacaagttgctcgaac: Reverse Primer: cagatcgtgttcccttagc) using Taq DNA

polymerase (Gene DirectX) (PCR mix: 1uL DNA, 1uL forward primer, 1uL reverse primer, 0.5uL

10mM dNTPs, 2uL 10x PCR Buffer, 0.1 uL taq, H20 to 20uL). PCR conditions were as follows: 5

minutes 95°, (30 seconds 95°, 30 seconds 55°, 30 seconds 72°)x30. Reactions were checked on a

gel and cleaned with the GenepHlow™ Gel/PCR Kit (Geneaid). Sanger sequencing reactions

were performed by the Mobix Lab at McMaster University. All alignments were created using

ClustalOmega (Madeira et al., 2022).

Results

dachsous (ds) shape change is aligned with major axes of genetic and phenotypic variation in

natural populations

To assess the relationship between shape change vectors and axes of variation in the

DGRP, mean shape vectors were calculated for each DGRP line then used in a PCA to summarize

axes of variation among strains. The mean shape vectors for each DGRP line were projected

onto vectors for *ds* and *emc* defined from the RNAi knockdown and the correlations between

them (Figure 1, Supplemental Figure 5) and with PC1 generated from DGRP line means (*ds*: r = -

0.56; *emc*: r = -0.45; Figure 1). A similar response was observed with the data projected onto

the *ds* and *emc* shape change vectors (Figure 1, r= 0.69) due to the high correlation between

shape change vectors themselves (r = 0.65). Projections of the data onto the vector defining the

neur shape change is aligned with PC1 (r = -0.69) and PC3 (r = -0.64), indicating that this is also

an important axis of shape variation in this population (Figure 1) that differs from projections

onto both *ds* (r = 0.034) and *emc* (r = 0.30) shape change vectors. Because of these observed

correlations, and previous associations observed (Pitchers et al., 2019) *ds, emc* and *neur* were

selected as focal genes for subsequent studies.

We also examined the relationship between direction of phenotypic effects with the

wild caught cohorts, where phenotypic variance for shape (and thus principal components) is

due to the joint contribution of genetic and environmental effects. To illustrate the difference

in shape variance in wild populations and the DGRP, we calculated correlations between the

first three eigenvectors for shape in the DGRP, the combined wild cohorts as well as the CMO

cohort alone, and observed low correlation between the DGRP eigenvectors and those

estimated from wild populations (Supplemental Table 2). As observed with the DGRP, there is a

substantial correlation between projections of individual shapes onto the *ds* shape change

vector and PC1 (defined by phenotypic variation among wild caught files) in most of the

sampled cohorts (PHO: r = 0.78; CMO: r = 0.87; FVW13: r = -0.22; FVW14: r = 0.95, Figure 2,

Supplemental Figure 9). In the cohorts where the *ds* shape change vector was not correlated

with PC1, specifically the FVW13 collection, this vector is correlated with PC2 (PHO: r = 0.12;

CMO r = -0.44; FVW13: r = -0.63; FVW14: r = 0.19; Figure 2; Supplemental Figure 5) from the

wild caught individuals. We also observe a correlation between *neur* shape change and PC1 in

most cohorts (PHO: r = 0.51; CMO: r = -0.051; FVW12: r = -0.95 FVW13; FVW14: r = -0.084;

Figure 2; Supplemental Figure 7). As with the *ds* shape change vector, in cases such as the CMO

cohort and PC1 there is a correlation between the *neur* shape change vector and PC2 (PHO: r =

0.22; CMO: r = -0.57; FVW13: r = -0.0.059; FVW14: r = 0.85; Figure 2; Supplemental Figure 7).

Interestingly, in the CMO natural cohort of flies, the correlations between the projection of

shape data onto the *ds* and *neur* shape change vectors is low (r = 0.11, Figure 2).

<u>Multiple loci linked to hippo signaling - including *ds*- respond to artificial selection for *ds* and</u>

<u>*emc* shape changes.</u>

To examine if variants in *ds* are contributing to shape variation, replicating the findings

of the earlier GWAS (Pitchers et al. 2019), we performed artificial selection experiment for wing

shape along the *ds* shape change vector and then examined the genomic response to selection.

By the final generation of selection, we observed a substantial shape change in both the "up"

(females: Procrustes Distance (PD) = 0.039, males: 0.044) and "down" directions (females: PD

= 0.022, males: PD = 0.022) compared to the base population at the start of the selection

experiment. In comparison, the shape change among unselected control lineages was much

smaller (females: PD = 0.005, males: PD = 0.005) (Figure 3, Supplemental Figure 10). The

direction of phenotypic shape change after seven generations of selection was in a similar

direction to the *ds* shape change vector (defined by the RNAi knockdown) for both the up

(females: r = 0.90, males: r = 0.90) and down (females: r = -0.82, males: r = -0.77) selection

lineages. Realized heritabilities, averaged over sex and replicate (Supplemental Figure 11, up =

0.38, 95% CI: 0.25 – 0.50; down = 0.28, 95% CI: 0.24 – 0.50) were moderate. Hippo signaling,

including the effects of *ds*, is often associated with changes in size (Pan, 2007). However, we do

not observe a significant change of wing size in our selection lineages in either sex

(Supplemental Figure 12). It is possible that with more generations of selection we would have

observed a clear change in size, as there is a trend indicating such divergence (Supplemental

Figure 12).

Genome-wide patterns of F_{ST} were examined between up and down *ds* selection

lineages. We observed strong genetic differentiation linked with the *ds* locus (Figure 3,

Supplemental Figure 13). We also observed striking genetic differentiation across several

regions in the genome. In addition, one of the SNPs (2L:702560) in the intron of *ds* identified in

Pitchers et al. (2019) through GWAS, showed the expected pattern of response to selection,

with opposing sign in up and down selection lineages, with the SNP going to high frequency in

all three up selection lineages (Table 1). It should be noted that this SNP represents a complex

polymorphism including an insertion of 18bp that may result in inaccurate genotyping at this

locus which we further evaluated with Sanger sequencing of this region (Supplemental Figure

14). Gene ontology analysis for genes in regions of the genome with an F_{ST} greater than 0. 345

(three standard deviations from mean F_{ST}), show enrichment for hippo signaling loci

(Supplemental Table 3). The top 20 enriched terms are all related to cell signaling and

development. Of particular note is the inclusion of the terms for 'negative regulation of hippo

signaling' (GO:0035331), and 'hippo signaling' (GO:0035329) in this list (Supplemental Table 3,

Supplemental Figure 13). Using a permutation test we confirmed this result by selecting

random sets of genomic intervals equal in size to the number of observed outlier windows, and

measuring the ratio of genes annotated to the expected number of genes in these regions. The

observed value for the terms for hippo signaling (ratio = 4.76) and negative regulation of hippo

signaling (ratio = 9.23) were compared to the distribution of values for 1000 random sets of loci

(Supplemental Figure 15). For the hippo signaling term the observed to expected ratio for

random loci was larger than the observed value 5 times out of 1000, while the value from

random loci was never larger than the observed.

For the artificial selection experiment based on the *emc* shape change vector we

observed phenotypic differentiation under artificial selection in both up (females: PD = 0.043,

males: PD = 0.040) and down directions (females: PD = 0.021, males: PD = 0.020) with little

change in control lineages (females: PD = 0.009, males: PD = 0.008) (Figure 4). The direction of

phenotypic change is correlated with the *emc* (RNAi knockdown) shape change vector in both

up (females: r = 0.75, males: r = 0.69) and down (females: r = -0.69, males: r = -0.75) directions.

Realized heritabilities, averaged over sex and replicate were calculated for both up and down

lineages (Supplemental Figure 16, up = 0.38, 95% CI: 0.29 – 0.47; down = 0.28, 95% CI: 0.21 –

0.35). Genetic differentiation linked to the *emc locus* was modest following selection, but we

again observed striking genetic differentiation linked to *ds* (Figure 4, Supplemental Figure 13).

Notably, as seen in Supplemental Figure 1, the site frequency spectrum (SFS) suggests modest

allelic variation at the *emc* locus in the synthetic outbred population. Using a three standard

deviation cut-off for F_{ST}, we did observe enrichment for various developmental GO terms, but

not of hippo signaling terms (Supplemental Table 4, Supplemental Figure 13).

<u>Bulk segregant analysis in wild caught cohorts does not recapitulate effects of the GWAS or</u>

<u>artificial selection</u>

Having demonstrated that variants in *ds* (or linked to *ds*) respond to artificial selection
for wing shape along the *ds* shape change vector, we next wanted to determine whether we
could recapitulate these findings with wild caught individuals. In addition to determining
whether we can replicate effects in wild cohorts, it provides the opportunity to identify causal
SNPs because of low LD generally observed in wild caught *Drosophila*. Wild caught populations
introduce considerably more environmental variation for shape along with a different site
frequency spectrum for variants contributing to shape variation (and *ds* like shape changes
specifically). In particular, it is known that several of the variants that the original GWAS
detected in *ds* have low minor allele frequency (MAF) (Pitchers et al., 2019) (Table 2). The SNP
at 2L:702560 does appear to be at intermediate frequency but it occurs both directly before
and after an indel, making the alignment and variant calling in this region challenging
(Supplementary Figure 14). We have included the frequencies (Table 2), but these results
should be interpreted with caution due to the technical complexities of mapping and variant
calling close to indels.

As we sampled multiple cohorts of wild-caught flies in different locations and years in
Michigan (USA), we wanted to confirm that any phenotypic differentiation among these
samples was modest and would not impact genomic analysis for the BSA. We observe modest,
statistically significant wing shape differences among cohorts from a Procrustes ANOVA,
utilizing permutations of the residuals for the relevant "null" model (Supplemental Table 5; R^2=
0.16, F = 351, Z_{RRPP} = 18.3, p = 0.001) (Collyer and Adams, 2018). This appears to be due to
differences in wing shape between the PHO population and other populations based on
pairwise Procrustes Distances (Supplemental Table 5, Supplemental Figure 3). In a joint PCA
including all populations, there is very modest separation between populations using allometry
adjusted shape (Supplemental Figure 3). Most relevant to the bulk segregant analysis approach
we used, when we project all wild caught individuals onto the *ds* and *neur* vectors, we see no
clear separation among sampling locales (Supplemental Figure 7). There is some variation in

wing size between populations (Supplemental Figure 8), but this is unlikely to influence

downstream analysis as we use size adjusted estimates. There is little evidence of genetic

differentiation between populations with the two collections from Fenn Valley Winery

separating more on a Principal Co-ordinate Analysis (PCoA)(Supplemental Figure 17) than other

sampling locales. There is also no relationship between genetic and phenotypic distances

between samples (Supplemental Figure 18). These results suggests that the multiple sampling

locales should not influence downstream genomic analysis as individuals used for generating

pools were compared within each population, and we observe little evidence for substantial

differences among populations.

Because there is a single bout of phenotypic selection distinguishing pools in the bulk

segregant analysis (BSA), shape and allele frequency changes between pools are expected to be

modest. We observe shape difference between the two pools within each population (PD =

CMO: 0.033; PHO: 0.036; FVW13: 0.040; FVW14: 0.041; Supplemental Figure 19). Correlations

of the shape difference vectors of the pools (i.e. difference between the two pools created

from the extremes along the *ds* shape change axis) and the direction of the *ds* shape change

vector used for selection is high (CMO: 0.94, PHO: 0.79, FVW13: 0.92, FVW14: 0.90).

The genome scans from the bulk segregant analysis show little evidence of genetic

differentiation linked to the *ds gene* (Figure 5). Across the genome, 15 sites were detected as

significantly differentiated between "up" and "down" selected pools based on a CMH test with

a FDR cut-off of 5% (Figure 5, Table 3). The genes nearest to these sites are not associated with

hippo signaling pathways or implicated in the development of the *Drosophila* wing (Table 3).

Because PHO had somewhat distinct shape variation from the other populations and had a

lower correlation of the difference vector between selected pools and *ds* shape change vector,

we repeated the CMH test with this population left out. We observe significant differentiation

at 174 sites between "up" and "down" pools (Supplemental Table 6, Supplemental Figure 20).

We identified the nearest genes to these sites and GO analysis indicated enrichment for wing

development terms, in particular related to Wnt signaling, but not hippo signaling terms

(Supplemental Table 6). Importantly, we do not observe differentiation linked to *ds* or any

other hippo loci. To ensure that the results we obtained were not due to uneven coverage

between samples, we down-sampled genomic coverage to 75x for each sample, dropping sites

that did not meet this threshold. Significant differences were detected at 19 sites

(Supplemental Figure 21, Supplemental Table 7), but none of these overlapped with those

identified using all the genomic data. Two of the significant sites are located in the *dumpy* gene,

a gene known to have a role in wing morphogenesis during pupation (Etournay et al., 2015). F_{ST}

between selected and random pools within each cohort are generally low (Supplemental Figure

22).

In addition to the BSA selection based upon the *ds* shape change, we also selected pools

of individuals based on the *neur* shape change vector. We did not use *emc* shape change in this

experiment due to the high similarity between the *ds* and *emc* shape change vectors (r = 0.65).

We selected the *neur* shape change vector as it is not aligned with *ds* but does align with

directions of variation in wild populations (Figure 1, 2, Supplemental Figure 9). Additionally,

there is little relationship between the *ds* and *neur* shape change axis (r = 0.12, Supplemental

Figures 7 and 8). We observe shape changes between pools of individuals (PD = CMO: 0.027;

PHO: 0.028; FVW14: 0.041; FVW13: 0.038, Supplemental Figure 23). There is little evidence of

genetic differentiation between *neur* selected pools (Supplemental Figure 24). Only 4 sites were

identified as being significantly differentiated between pools and none of these sites are

associated with wing development (Supplemental Table 9). When population differentiation

between pools within populations is measured using F_{ST}, genetic differentiation remains low

across the genome (Supplemental Figure 25)

Discussion

The primary goal of this study was to determine whether we could recapitulate genetic

effects initially observed through a traditional GWAS using an "inverted" approach: artificially

selecting on phenotypes and observing changes in allele frequencies. We observed that shape

changes associated with the *ds*, *emc* and *neur* genes were associated with major axes of genetic

variation among a panel of wild type strains (DGRP) reared in the lab, and axes of phenotypic

variation among wild caught individuals (Figure 1, 2). After observing a strong response to

artificial selection along two shape change vectors (*ds* and *emc*), we examined patterns of

genomic differentiation and observed substantial changes in allele frequency for markers linked

with *ds* itself, and also markers linked to numerous other genes associated with hippo signaling

(Figure 3, 4).

In contrast, our bulk segregant experiments using pools of wild caught individuals chosen to

be phenotypically divergent on the same shape vectors, did not detect differences in the loci

identified in the artificial selection experiments (Figure 5, Supplemental Figure 24). As we

discuss in detail below, these seemingly contradictory results are in fact not that surprising.

Following artificial selection based on *ds* shape change we observe allele frequency changes

not only at *ds* but also linked to a number of other hippo signalling loci (Figure 3, Supplemental

Table 3). The previous GWAS study identified a number of loci associated with wing shape

variation in the DGRP, however, this approach cannot predict which alleles are causative

(Pitchers et al., 2019). In our synthetic outbred population, we maximized variation among

haplotype blocks containing many of the candidate SNPs in *ds*, increasing our ability to detect

frequency changes at and near the implicated variants. Although LD blocks in the outcrossing

population from this study remain large, *ds* variants exist on multiple distinct haplotypes,

allowing for an examination of allele frequency changes for variants. Of particular interest is

SNP 2L:702560, which was previously identified though GWAS (Pitchers et al., 2019) as

influencing wing shape variation. It was driven to near fixation in each of the artificial selection

lineages (Table 1). Although this polymorphism is annotated as a SNP, this region may contain a

complex polymorphism (Supplemental Figure 14), making it difficult to accurately assess

genotypic calls. Because of this, the predicted allele frequency in the founding population and

allele frequencies in this region may be inaccurate. Previous studies demonstrate the

importance of alleles at intermediate frequency in founding populations to those contributing

to responses to selection over short timescales (Kelly and Hughes, 2019). If this polymorphism

is at a more intermediate frequency in the founding population, it would be more likely to be

captured by selection during these experiments. Additionally, we recognize that the haplotype

blocks are large, and contain many potential functional variants. However, based on the results

of both the current and previous studies, these *ds* variants associated with 2L:702560 are good

candidates for functional validation in future work.

When selecting on the *emc* shape change vector, which is similar but not identical to *ds*

shape change, we observe only a modest allele frequency change at *emc*, and a more robust

response at *ds* (Figure 4). In hindsight, this is not particularly surprising and there are multiple

contributing factors. Given the increased genetic diversity at *ds* compared to *emc* in the

founding population, alleles in *ds* may have provided a more robust genetic target influencing

phenotypic variation for wing shape along this axis of variation. Additionally, if our estimated

direction of effects and selection for *emc* (based on RNAi knockdown) was not well aligned with

the actual direction of *emc* SNP effects, this could result in weaker selection on variants at the

emc locus. It is worthwhile pointing out the small magnitude of the *emc* shape change vector

(0.44) relative to *ds* (5.5). However, previous work has indicated that there is a relationship

between this estimated *emc* shape change vector (from RNAi) and the effect of SNPs in *emc* on

shape change (Pitchers et al., 2019).

In addition to a response on allele frequency associated with *ds*, our results suggest a

response on segregating variation at other hippo signaling loci. Earlier work has suggested that

the direction of effects within signaling pathways are inconsistent for alleles of small effect

(Dworkin and Gibson, 2006). However, allelic effect sizes in the 2006 study were heterogeneous

and may result in direction and magnitude being confounded. In contrast, in both the current

and the Pitchers et al*.* (2019) studies, we estimated the direction of genetic effects by titrating

gene knockdown. The strength of this approach is highlighted in the result that segregating

variation at multiple hippo loci was selected on (Figure 3). Our finding is consistent with models

for the architecture of complex traits that predict that many alleles of small effect will

contribute to trait variation with many genes within developmental pathways (Boyle et al.,

2017; Wray et al., 2018). This pathway response has also been demonstrated in human

adaptation to pathogen resistance (Daub et al., 2013) and adaptation to high altitude (Gouy et

al., 2017). These results are consistent with the expectation that polymorphisms in the same

developmental pathway would show correlated phenotypic effects and therefore correlated

genomic responses to selection. However, this may not be reflective of all wild caught

populations. In this study, we generated a population that had high diversity at *ds*, while these

variants are at much lower frequency in natural populations (Table 2). The amount of

selectable variation a variant causes is dependent on both effect size, *a*, and variant

frequencies, *p* as $V_A= 2p(1-p)a^2$. When allele frequencies are near 0 or 1, even variants with

large effects will have only a small contribution to short term selection response.   Therefore,

the outcrossed population we created here is an ideal situation to validate the existence of the

measured effects. It is unlikely to be typical of natural populations where functional variants

may be rare.

Given the clear and robust response observed in the artificial selection experiment, it may

seem surprising that we do not observe allele frequency changes in the bulk segregant analysis

using the wild cohorts. Indeed, previous work has demonstrated that variants in *Egfr*, could be

replicated in wild caught samples (Dworkin et al., 2005; Palsson et al., 2005) and were also

found in genome wide associations (Pitchers et al., 2019). However, there are many

explanations for why we may not have been able to detect these allele frequency changes in

our experiment. First, the addition of environmental variation to the system introduces

additional complications. In the aforementioned example of variants in *Egfr*, the genetic effect

of the SNP in wild caught cohorts was only 10% of the magnitude estimated in lab reared flies.

As discussed previously, the *ds* variants implicated in the previous GWAS study are at low

frequency in the natural cohorts (Table 2). Given that natural populations of *Drosophila* are

generally large and wing shape is likely under weak selection (Gilchrist and Partridge, 2001),

mutation-drift-selection balance may maintain most variation, resulting in low minor allele

frequencies at these sites. Because allelic contribution to wing shape are expected to be both

rare in wild populations and of small phenotypic effect, we do not expect large allele frequency

changes given only one "generation" of selection. Using the approach of ACER (Spitzer et al.,

2020) to account for sampling effects, we observe few differentiated sites, and none in the *ds*

gene, indicating that bulk segregant analysis (BSA) may not be a well-suited approach to

identify modest allele frequency changes and as such may not be particularly effective for

highly polygenic traits. Although our approach was tailored to look for variants that had

consistent direction of frequency changes across the four collection cohorts, it is possible that

different loci were contributing variation within each cohort. We attempted to address this

question by examining allele frequency changes between selected pools within each cohort

(Supplemental Figures 22 and 24) but could not identify specific loci contributing to differences

within any one population. Previous successful BSA studies identified a smaller number of

contributing loci with few polymorphisms contributing to the trait of interest. For example, in

Drosophila, a number of melanin synthesis genes contributing to variance in pigmentation

between populations were identified using a BSA (Bastide et al., 2013). Pigmentation may

represent a relatively 'simpler' genetic architecture (fewer variants of individually larger genetic

effect, smaller impact of environmental variation, smaller mutational target size) and if so, this

may have enabled the success of the BSA approach with such systems. In the case of wing

shape, we know that many alleles of small effect contribute to variation in the trait (Pitchers et

al., 2019).

Our approach for the BSA was to perform the same phenotypic selection within each of

four distinct "populations". It is important to recognize that there was heterogeneity among

our populations, not only in allele frequencies, but in environmental variance and potentially

GxE, even though all were caught in locales in lower Michigan. We detected small degrees of

phenotypic and genetic differences between the four cohorts, however these effects are

neither correlated with one another, nor related to the *ds* and *neur* shape scores we used for

selecting individuals (Supplemental Figures 3, 7, 18).The population from the Phillips Orchard

(PHO) was phenotypically distinct from the other populations. When we performed the BSA

without this population, we observed a larger set of variants associated with shape

(Supplemental Figure 20), albeit still not showing any effects at the *ds* or *neur* genes

themselves. One possibility is that the increased number of sites when the PHO sample is

removed from analysis represents an unknown statistical artefact we have not identified.

However, a more likely explanation is that there are some large unknown environmental

influences (**E**), or that the genetic effects show a degree of GxE (with a specific environment in

PHO) that contributed to shape variation along the *ds* direction in this population. Such

obfuscating effects have been observed before with the previously discussed *Egfr* example,

where the same SNP effect that was initially identified and validated in multiple contexts

(Dworkin et al., 2005; Palsson et al., 2005; Palsson and Gibson, 2004) could not be detected in

one natural population, despite being at intermediate frequencies in each sample (Palsson et

al., 2005). Importantly, we did detect differentiation at sites associated with developmental

processes in the wild cohorts, suggesting that the failure to detect variation linked to *ds* or

other hippo signaling loci (Table 3, Supplemental Table 6,7) is not due simply a lack of power.

The response to selection at *ds* and other hippo signaling loci in the artificial selection

experiment based on *ds* shape change indicates that this is an important axis of variation for

wing shape. Coupled with the alignment of phenotypic effects of perturbations in genes in this

pathway with directions of **G** and **P**, this finding may seem to suggest a developmental bias in

available variation. However, we caution against such interpretations based solely on the

findings in this study. The structure of the **G** matrix strongly influenced our findings as we

artificially created a population to maximize genetic diversity at *ds*. When another effect is

aligned with *ds* shape change, as in the case of *emc* shape change, we observed the same

response at the hippo signaling loci and not at *emc.* Only the genetic diversity in the starting

population was available to be selected on so this influenced selection towards the "spiked in"

ds variants, even if the inferred phenotypic effects of *emc* variants are very similar.

Alternatively, the inferred *emc* direction of effects (via RNAi knockdown) may be sufficiently

"distant" from true effects of *emc* variants. If this was the case, we were ineffectively selecting

for *emc* shape changes. In other cases where single genes are implicated in divergence between

multiple populations, such as *mc1r* in mice (Steiner et al., 2007) or *pitx1* in stickleback (Chan et

al., 2010), other factors such as low pleiotropy, developmental and mutational constraints and

history of selection in the population are used to explain why these genes are so often

implicated in evolutionary change (Gompel and Prud'homme, 2009; Martin and Orgogozo,

2013; Stern and Orgogozo, 2008). In our case, it is not *ds* itself that is special but rather the

orientation of the **G** matrix to align g_{max} with the direction of effect for *ds* that shapes our

results. Selection acts on variants aligned with the vector of selection (Reddiex and Chenoweth,

2021). By varying the orientation of g_{max} in the parental population, we would be able to

address questions about the repeatability of hippo overrepresentation and if this can be

explained by more than just the orientation of **G.**

Despite the need for skepticism about the potential for developmental bias influencing

directions of variation, the correlated response of sites linked to multiple other hippo signaling

genes is intriguing. Coupling of more traditional mapping approaches like GWAS with short

term artificial selection provides an additional route to validation and replication of genetic

effects. It also suggests that using multivariate data to address the distribution of genetic

effects will pay long-term dividends to our understanding of both inheritance and the evolution

of multivariate traits.

Data Availability:

All code is available on GitHub at: https://github.com/KatiePelletier/WingShapeBSA

All images and phenotypic data are available on Dryad at:

All sequence data has been submitted to NCBI

Acknowledgements:

We would like to thank the Hosken lab for *D. simulans* wing shape data. Yun Bo Xi aided

with the sequencing of *ds* haplotype.

References:
Adams, D.C., Otárola-Castillo, E., 2013. geomorph: an R package for the collection and analysis
of geometric morphometric shape data. Methods Ecol. Evol. 4, 393–399.
https://doi.org/10.1111/2041-210X.12035
Alexa, A., Rahnenfuhrer, J., Lengauer, T., 2006. Improved scoring of functional groups from
gene expression data by decorrelating GO graph structure. Bioinformatics 22, 1600–
1607. https://doi.org/10.1093/bioinformatics/btl140
Bastide, H., Betancourt, A., Nolte, V., Tobler, R., Stöbe, P., Futschik, A., Schlötterer, C., 2013. A
Genome-Wide, Fine-Scale Map of Natural Pigmentation Variation in Drosophila
melanogaster. PLoS Genet. 9, e1003534. https://doi.org/10.1371/journal.pgen.1003534
Benjamini, Y., Hochberg, Y., 1995. Controlling the False Discovery Rate: A Practical and Powerful
Approach to Multiple Testing. J. R. Stat. Soc. Ser. B Methodol. 57, 289–300.
https://doi.org/10.1111/j.2517-6161.1995.tb02031.x
Blows, M.W., McGuigan, K., 2015. The distribution of genetic variance across phenotypic space
and the response to selection. Mol. Ecol. 24, 2056–2072.
https://doi.org/10.1111/mec.13023
Bolger, A.M., Lohse, M., Usadel, B., 2014. Trimmomatic: a flexible trimmer for Illumina
sequence data. Bioinformatics 30, 2114–2120.
https://doi.org/10.1093/bioinformatics/btu170
Boyle, E.A., Li, Y.I., Pritchard, J.K., 2017. An Expanded View of Complex Traits: From Polygenic to
Omnigenic. Cell 169, 1177–1186. https://doi.org/10.1016/j.cell.2017.05.038

Chan, Y.F., Marks, M.E., Jones, F.C., Villarreal, G., Shapiro, M.D., Brady, S.D., Southwick, A.M.,
Absher, D.M., Grimwood, J., Schmutz, J., Myers, R.M., Petrov, D., Jónsson, B., Schluter,
D., Bell, M.A., Kingsley, D.M., 2010. Adaptive Evolution of Pelvic Reduction in
Sticklebacks by Recurrent Deletion of a Pitx1 Enhancer. Science 327, 302–305.
https://doi.org/10.1126/science.1182213
Collyer, M.L., Adams, D.C., 2018. RRPP: An ʀ package for fitting linear models to high-
dimensional data using residual randomization. Methods Ecol. Evol. 9, 1772–1779.
https://doi.org/10.1111/2041-210X.13029
CONVERGE consortium, 2015. Sparse whole-genome sequencing identifies two loci for major
depressive disorder. Nature 523, 588–591. https://doi.org/10.1038/nature14659
Daub, J.T., Hofer, T., Cutivet, E., Dupanloup, I., Quintana-Murci, L., Robinson-Rechavi, M.,
Excoffier, L., 2013. Evidence for Polygenic Adaptation to Pathogens in the Human
Genome. Mol. Biol. Evol. 30, 1544–1558. https://doi.org/10.1093/molbev/mst080
Dworkin, I., Gibson, G., 2006. Epidermal Growth Factor Receptor and Transforming Growth
Factor-β Signaling Contributes to Variation for Wing Shape in *Drosophila melanogaster*.
Genetics 173, 1417–1431. https://doi.org/10.1534/genetics.105.053868
Dworkin, I., Palsson, A., Gibson, G., 2005. Replication of an *Egfr* -Wing Shape Association in a
Wild-Caught Cohort of *Drosophila melanogaster*. Genetics 169, 2115–2125.
https://doi.org/10.1534/genetics.104.035766
Etournay, R., Popović, M., Merkel, M., Nandi, A., Blasse, C., Aigouy, B., Brandl, H., Myers, G.,
Salbreux, G., Jülicher, F., Eaton, S., 2015. Interplay of cell dynamics and epithelial tension
during morphogenesis of the Drosophila pupal wing. eLife 4, e07090.
https://doi.org/10.7554/eLife.07090
Fatumo, S., Carstensen, T., Nashiru, O., Gurdasani, D., Sandhu, M., Kaleebu, P., 2019.
Complimentary Methods for Multivariate Genome-Wide Association Study Identify New
Susceptibility Genes for Blood Cell Traits. Front. Genet. 10, 334.
https://doi.org/10.3389/fgene.2019.00334
Fritsche, L.G., Igl, W., Bailey, J.N.C., Grassmann, F., Sengupta, S., Bragg-Gresham, J.L., Burdon,
K.P., Hebbring, S.J., Wen, C., Gorski, M., Kim, I.K., Cho, D., Zack, D., Souied, E., Scholl,
H.P.N., Bala, E., Lee, K.E., Hunter, D.J., Sardell, R.J., Mitchell, P., Merriam, J.E., Cipriani,
V., Hoffman, J.D., Schick, T., Lechanteur, Y.T.E., Guymer, R.H., Johnson, M.P., Jiang, Y.,
Stanton, C.M., Buitendijk, G.H.S., Zhan, X., Kwong, A.M., Boleda, A., Brooks, M., Gieser,
L., Ratnapriya, R., Branham, K.E., Foerster, J.R., Heckenlively, J.R., Othman, M.I., Vote,
B.J., Liang, H.H., Souzeau, E., McAllister, I.L., Isaacs, T., Hall, J., Lake, S., Mackey, D.A.,
Constable, I.J., Craig, J.E., Kitchner, T.E., Yang, Z., Su, Z., Luo, H., Chen, D., Ouyang, H.,
Flagg, K., Lin, D., Mao, G., Ferreyra, H., Stark, K., von Strachwitz, C.N., Wolf, A., Brandl,
C., Rudolph, G., Olden, M., Morrison, M.A., Morgan, D.J., Schu, M., Ahn, J., Silvestri, G.,
Tsironi, E.E., Park, K.H., Farrer, L.A., Orlin, A., Brucker, A., Li, M., Curcio, C.A., Mohand-
Saïd, S., Sahel, J.-A., Audo, I., Benchaboune, M., Cree, A.J., Rennie, C.A., Goverdhan, S.V.,
Grunin, M., Hagbi-Levi, S., Campochiaro, P., Katsanis, N., Holz, F.G., Blond, F., Blanché,
H., Deleuze, J.-F., Igo, R.P., Truitt, B., Peachey, N.S., Meuer, S.M., Myers, C.E., Moore,
E.L., Klein, R., Hauser, M.A., Postel, E.A., Courtenay, M.D., Schwartz, S.G., Kovach, J.L.,
Scott, W.K., Liew, G., Tan, A.G., Gopinath, B., Merriam, J.C., Smith, R.T., Khan, J.C.,
Shahid, H., Moore, A.T., McGrath, J.A., Laux, R., Brantley, M.A., Agarwal, A., Ersoy, L.,

Caramoy, A., Langmann, T., Saksens, N.T.M., de Jong, E.K., Hoyng, C.B., Cain, M.S.,
Richardson, A.J., Martin, T.M., Blangero, J., Weeks, D.E., Dhillon, B., van Duijn, C.M.,
Doheny, K.F., Romm, J., Klaver, C.C.W., Hayward, C., Gorin, M.B., Klein, M.L., Baird, P.N.,
den Hollander, A.I., Fauser, S., Yates, J.R.W., Allikmets, R., Wang, J.J., Schaumberg, D.A.,
Klein, B.E.K., Hagstrom, S.A., Chowers, I., Lotery, A.J., Léveillard, T., Zhang, K., Brilliant,
M.H., Hewitt, A.W., Swaroop, A., Chew, E.Y., Pericak-Vance, M.A., DeAngelis, M.,
Stambolian, D., Haines, J.L., Iyengar, S.K., Weber, B.H.F., Abecasis, G.R., Heid, I.M., 2016.
A large genome-wide association study of age-related macular degeneration highlights
contributions of rare and common variants. Nat. Genet. 48, 134–143.
https://doi.org/10.1038/ng.3448

Gilchrist, A.S., Partridge, L., 2001. The contrasting genetic architecture of wing size and shape in
Drosophila melanogaster. Heredity 86, 144–152. https://doi.org/10.1046/j.1365-
2540.2001.00779.x

Gompel, N., Prud'homme, B., 2009. The causes of repeated genetic evolution. Dev. Biol. 332,
36–47. https://doi.org/10.1016/j.ydbio.2009.04.040

Gouy, A., Daub, J.T., Excoffier, L., 2017. Detecting gene subnetworks under selection in
biological pathways. Nucleic Acids Res. 45, e149–e149.
https://doi.org/10.1093/nar/gkx626

Green, R.M., Fish, J.L., Young, N.M., Smith, F.J., Roberts, B., Dolan, K., Choi, I., Leach, C.L.,
Gordon, P., Cheverud, J.M., Roseman, C.C., Williams, T.J., Marcucio, R.S., Hallgrímsson,
B., 2017. Developmental nonlinearity drives phenotypic robustness. Nat. Commun. 8,
1970. https://doi.org/10.1038/s41467-017-02037-7

Greene, C.S., Penrod, N.M., Williams, S.M., Moore, J.H., 2009. Failure to Replicate a Genetic
Association May Provide Important Clues About Genetic Architecture. PLoS ONE 4,
e5639. https://doi.org/10.1371/journal.pone.0005639

Houle, D., Mezey, J., Galpern, P., Carter, A., 2003. Automated measurement of Drosophila
wings. BMC Evol. Biol. 3, 25. https://doi.org/10.1186/1471-2148-3-25

Huang, W., Massouras, A., Inoue, Y., Peiffer, J., Ramia, M., Tarone, A.M., Turlapati, L., Zichner,
T., Zhu, D., Lyman, R.F., Magwire, M.M., Blankenburg, K., Carbone, M.A., Chang, K., Ellis,
L.L., Fernandez, S., Han, Y., Highnam, G., Hjelmen, C.E., Jack, J.R., Javaid, M., Jayaseelan,
J., Kalra, D., Lee, S., Lewis, L., Munidasa, M., Ongeri, F., Patel, S., Perales, L., Perez, A., Pu,
L., Rollmann, S.M., Ruth, R., Saada, N., Warner, C., Williams, A., Wu, Y.-Q., Yamamoto,
A., Zhang, Y., Zhu, Y., Anholt, R.R.H., Korbel, J.O., Mittelman, D., Muzny, D.M., Gibbs,
R.A., Barbadilla, A., Johnston, J.S., Stone, E.A., Richards, S., Deplancke, B., Mackay, T.F.C.,
2014. Natural variation in genome architecture among 205 Drosophila melanogaster
Genetic Reference Panel lines. Genome Res. 24, 1193–1208.
https://doi.org/10.1101/gr.171546.113

Ioannidis, J.P.A., Tarone, R., McLaughlin, J.K., 2011. The False-positive to False-negative Ratio in
Epidemiologic Studies: Epidemiology 22, 450–456.
https://doi.org/10.1097/EDE.0b013e31821b506e

Kelly, J.K., Hughes, K.A., 2019. Pervasive Linked Selection and Intermediate-Frequency Alleles
Are Implicated in an Evolve-and-Resequencing Experiment of *Drosophila simulans*.
Genetics 211, 943–961. https://doi.org/10.1534/genetics.118.301824

King, Elizabeth G., Macdonald, S.J., Long, A.D., 2012. Properties and Power of the Drosophila
Synthetic Population Resource for the Routine Dissection of Complex Traits. Genetics
191, 935–949. https://doi.org/10.1534/genetics.112.138537
King, E. G., Merkes, C.M., McNeil, C.L., Hoofer, S.R., Sen, S., Broman, K.W., Long, A.D.,
Macdonald, S.J., 2012. Genetic dissection of a model complex trait using the Drosophila
Synthetic Population Resource. Genome Res. 22, 1558–1566.
https://doi.org/10.1101/gr.134031.111
Klingenberg, C.P., 2022. Methods for studying allometry in geometric morphometrics: a
comparison of performance. Evol. Ecol. 36, 439–470. https://doi.org/10.1007/s10682-
022-10170-z
Kofler, Orozco-terWengel, P., De Maio, N., Pandey, R.V., Nolte, V., Futschik, A., Kosiol, C.,
Schlötterer, C., 2011a. PoPoolation: A Toolbox for Population Genetic Analysis of Next
Generation Sequencing Data from Pooled Individuals. PLoS ONE 6, e15925.
https://doi.org/10.1371/journal.pone.0015925
Kofler, Pandey, R.V., Schlotterer, C., 2011b. PoPoolation2: identifying differentiation between
populations using sequencing of pooled DNA samples (Pool-Seq). Bioinformatics 27,
3435–3436. https://doi.org/10.1093/bioinformatics/btr589
Li, H., Durbin, R., 2010. Fast and accurate long-read alignment with Burrows–Wheeler
transform. Bioinformatics 26, 589–595. https://doi.org/10.1093/bioinformatics/btp698
Li, H., Handsaker, B., Wysoker, A., Fennell, T., Ruan, J., Homer, N., Marth, G., Abecasis, G.,
Durbin, R., 1000 Genome Project Data Processing Subgroup, 2009. The Sequence
Alignment/Map format and SAMtools. Bioinformatics 25, 2078–2079.
https://doi.org/10.1093/bioinformatics/btp352
Madeira, F., Pearce, M., Tivey, A.R.N., Basutkar, P., Lee, J., Edbali, O., Madhusoodanan, N.,
Kolesnikov, A., Lopez, R., 2022. Search and sequence analysis tools services from EMBL-
EBI in 2022. Nucleic Acids Res. 50, W276–W279. https://doi.org/10.1093/nar/gkac240
Marigorta, U.M., Rodríguez, J.A., Gibson, G., Navarro, A., 2018. Replicability and Prediction:
Lessons and Challenges from GWAS. Trends Genet. 34, 504–517.
https://doi.org/10.1016/j.tig.2018.03.005
Marquez, E.J., Houle, D., 2015. Dimensionality and the statistical power of multivariate
genome-wide association studies (preprint). Genomics. https://doi.org/10.1101/016592
Marriage, T.N., King, E.G., Long, A.D., Macdonald, S.J., 2014. Fine-Mapping Nicotine Resistance
Loci in *Drosophila* Using a Multiparent Advanced Generation Inter-Cross Population.
Genetics 198, 45–57. https://doi.org/10.1534/genetics.114.162107
Martin, A., Orgogozo, V., 2013. The Loci of Repeated Evolution: A Catalog of Genetic Hotspots
of Phenotypic Variation. Evolution 67, 1235–1250. https://doi.org/10.1111/evo.12081
McGuigan, K., 2006. Studying phenotypic evolution using multivariate quantitative genetics:
EVOLUTIONARY QUANTITATIVE GENETICS. Mol. Ecol. 15, 883–896.
https://doi.org/10.1111/j.1365-294X.2006.02809.x
Melo, D., Marroig, G., Wolf, J.B., 2019. Genomic Perspective on Multivariate Variation,
Pleiotropy, and Evolution. J. Hered. 110, 479–493.
https://doi.org/10.1093/jhered/esz011

Palsson, A., Dodgson, J., Dworkin, I., Gibson, G., 2005. Tests for the replication of an association
852   between Egfr and natural variation in Drosophila melanogaster wing morphology. BMC
853   Genet. 6, 44. https://doi.org/10.1186/1471-2156-6-44
854   Palsson, A., Gibson, G., 2004. Association Between Nucleotide Variation in Egfr and Wing Shape
855   in Drosophila melanogaster. Genetics 167, 1187–1198.
856   https://doi.org/10.1534/genetics.103.021766
857   Pan, D., 2007. Hippo signaling in organ size control. Genes Dev. 21, 886–897.
858   https://doi.org/10.1101/gad.1536007
859   Pan, Y., Alégot, H., Rauskolb, C., Irvine, K.D., 2018. The dynamics of Hippo signaling during
860   *Drosophila* wing development. Development 145, dev165712.
861   https://doi.org/10.1242/dev.165712
862   Pitchers, W., Nye, J., Márquez, E.J., Kowalski, A., Dworkin, I., Houle, D., 2019. A Multivariate
863   Genome-Wide Association Study of Wing Shape in *Drosophila melanogaster*. Genetics
864   211, 1429–1447. https://doi.org/10.1534/genetics.118.301342
865   Porter, H.F., O'Reilly, P.F., 2017. Multivariate simulation framework reveals performance of
866   multi-trait GWAS methods. Sci. Rep. 7, 38837. https://doi.org/10.1038/srep38837
867   Quinlan, A.R., Hall, I.M., 2010. BEDTools: a flexible suite of utilities for comparing genomic
868   features. Bioinformatics 26, 841–842. https://doi.org/10.1093/bioinformatics/btq033
869   Ray, R.P., Nakata, T., Henningsson, P., Bomphrey, R.J., 2016. Enhanced flight performance by
870   genetic manipulation of wing shape in Drosophila. Nat. Commun. 7, 10851.
871   https://doi.org/10.1038/ncomms10851
872   Reddiex, A.J., Chenoweth, S.F., 2021. Integrating genomics and multivariate evolutionary
873   quantitative genetics: a case study of constraints on sexual selection in *Drosophila*
874   *serrata*. Proc. R. Soc. B Biol. Sci. 288, 20211785. https://doi.org/10.1098/rspb.2021.1785
875   Rohlf, F.J., Slice, D., 1990. Extensions of the Procrustes Method for the Optimal Superimposition
876   of Landmarks. Syst. Zool. 39, 40. https://doi.org/10.2307/2992207
877   Schluter, D., 1996. Adaptive Radiation Along Genetic Lines of Least Resistance. Evolution 50,
878   1766. https://doi.org/10.2307/2410734
879   Shriner, D., 2012. Moving toward System Genetics through Multiple Trait Analysis in Genome-
880   Wide Association Studies. Front. Genet. 3. https://doi.org/10.3389/fgene.2012.00001
881   Spitzer, K., Pelizzola, M., Futschik, A., 2020. Modifying the Chi-square and the CMH test for
882   population genetic inference: Adapting to overdispersion. Ann Appl Stat 14, 202–220.
883   https://doi.org/10.1214/19-AOAS1301
884   Steiner, C.C., Weber, J.N., Hoekstra, H.E., 2007. Adaptive Variation in Beach Mice Produced by
885   Two Interacting Pigmentation Genes. PLoS Biol. 5, e219.
886   https://doi.org/10.1371/journal.pbio.0050219
887   Stephens, M., 2013. A Unified Framework for Association Analysis with Multiple Related
888   Phenotypes. PLoS ONE 8, e65245. https://doi.org/10.1371/journal.pone.0065245
889   Stern, D.L., Orgogozo, V., 2008. THE LOCI OF EVOLUTION: HOW PREDICTABLE IS GENETIC
890   EVOLUTION? Evolution 62, 2155–2177. https://doi.org/10.1111/j.1558-
891   5646.2008.00450.x
892   Tam, V., Patel, N., Turcotte, M., Bossé, Y., Paré, G., Meyre, D., 2019. Benefits and limitations of
893   genome-wide association studies. Nat. Rev. Genet. 1. https://doi.org/10.1038/s41576-
894   019-0127-1

Visscher, P.M., Wray, N.R., Zhang, Q., Sklar, P., McCarthy, M.I., Brown, M.A., Yang, J., 2017. 10
Years of GWAS Discovery: Biology, Function, and Translation. Am. J. Hum. Genet. 101,
5–22. https://doi.org/10.1016/j.ajhg.2017.06.005
Wray, N.R., Wijmenga, C., Sullivan, P.F., Yang, J., Visscher, P.M., 2018. Common Disease Is More
Complex Than Implied by the Core Gene Omnigenic Model. Cell 173, 1573–1580.
https://doi.org/10.1016/j.cell.2018.05.051
Wright, S., 1934. Physiological and Evolutionary Theories of Dominance. Am. Nat. 68, 24–53.
https://doi.org/10.1086/280521
903
904

Figure 1. Projections of data onto RNAi shape change vectors are correlated with major axes of shape variation among DGRP strains. (A) Shape change vectors from RNAi titration experiments for *ds, emc* and *neur* were used, and DGRP line means were projected onto these vectors to calculate shape scores. Eigenvectors for the PCA were estimated based on the same DGRP line means. Vector correlations between shape change vectors from RNAi knockdown: *ds − emc*: 0.65, *ds − neur*: 0.03, *emc − neur*: 0.30. (B) Effect of *ds* shape change estimated from RNAi knockdown, effects not magnified. (C) Landmarks (red) and semi-landmarks (blue) used in geomorphic morphometric analysis on a *Drosophila* wing. PCs 1−3 account for 22%, 20% and 9% of the overall, among DGRP shape variance.

Figure 2. Projections of data onto RNAi shape change vectors are correlated with major axes of shape variation in wild-caught *Drosophila*. Correlations between projection of shape data from CMO population onto *ds, emc* and *neur* RNAi shape change vectors, and the first three eigenvectors from the PCA, calculated from shape data from all samples in the CMO population. PC1−3 account for 24%, 18% and 9% of overall shape variance in the CMO population.

Figure 3. Artificial selection along *ds* shape change vector influences allele frequencies of variants at *ds*, and other hippo signalling genes. (A) Phenotypic response to selection based on *ds* shape change vector. Only data from females is plotted for ease of visualization. Each replicate of up (squares), control (dots) and down (triangles) selection lineages are plotted (greys). Estimated response to selection shown along red lines. Wing plots represent the effect of selection on shape change between generation one and seven (red, effects not magnified). (B) Genomic differentiation (F_{ST}) between up and down selection treatments measured in 5000bp windows. Red line represents the location of the *ds* locus. Grey line represents 3sd from genome wide mean F_{ST}.

Figure 4. Artificial selection along *emc* shape change vector has modest influence on allele frequencies at *emc*, but a greater impact at the *ds locus*. (A) Phenotypic response to selection based on the *emc* shape change vector. Only data from females is plotted for ease of visualization. Each replicate of up (squares), control (dots) and down (triangles) selection lineages are plotted in greys. Estimated response to selection shown along red lines. Shape change between generation 1 and 7 is indicated on the right. Shape effects have been magnified 5x. (B) Genomic differentiation between up and down selection lineages (F_{ST}) measured in 5000bp sliding windows. Red and purple vertical lines represents genomic locations of *ds* and *emc* respectively. Grey line represents 3sd from genome wide mean F_{ST}.

.

Figure 5. Genetic Differentiation between pools selected based on *ds* shape change among the wild-caught cohorts. (A) Genome-wide scan for differentiated loci between pools selected based on *ds* shape change vector using the CMH test implemented in ACER. Points in red indicate sites with significant differentiation. (B) Genomic differentiation at *ds* between pools selected based on *ds* shape change vector. No sites are significantly differentiated in *ds*. The large gap in sites is due to a masked region in the genome due to repetitive sequence and poor (syntenic) mapping scores. (C) Shape difference between selected pools of individuals from one representative (CMO) population, with the mean shape of pools represented in black and red.

Table 1. Variants from Pitchers et al. (2019) in *ds* artificial selection experiment. Estimated effect sizes for SNPs are estimated from a GWAS in the DGRP using LASSO regularized coefficients. Average frequency is given with replicate lineage frequencies in brackets. Estimated effect is the l^2-*norm* of shape differences associated with the variant. MAF = minor allele frequency.

Variant	Estimated Effect	DGRP MAF	Estimated MAF in synthetic outcross	Average allele frequency "up" selection	Average allele frequency "down" selection	Average allele frequency "control" selection
2L:655894	0.072	0.44	0.067	0 (0, 0, 0)	0 (0, 0, 0)	0.003 (0, 0, 0.0105)
2L:702560*	**0.159**	**0.056**	**0.06**	**0.995 (1, 0.98, 1)**	**0.446 (0.32, 0.35, 0.67)**	**0.705 (0.69, 0.56, 0.87)**
2L:702798	0.101	0.089	0.1	0.007 (0, 0.0217, 0)	0 (0, 0, 0)	0.005 (0, 0, 0.139)
2L:718623	0.225	0.033	0	0 (0, 0, 0)	0 (0, 0, 0)	0 (0, 0, 0)
2L:718627	0.11	0.033	0	0 (0, 0, 0)	0 (0, 0, 0)	0 (0, 0, 0)

* This is a complex polymorphism with linked SNPs and INDELs, in Pitchers *et al* (2019) a SNP in this region was found to be linked. However, the variant calling pipeline used in this work recognized an INDEL in this region which was used for counting.

Table 2. *ds* Variants from Pitchers et al. (2019) in wild-caught cohorts used in the present study. Estimated effect sizes for SNPs are estimated from the DGRP GWAS with LASSO regularized coefficients. MAF in wild cohorts was estimated from sequenced pools of 75 random individuals.

Variant	Estimated Effect (Pitchers 2019)	DGRP MAF	Estimated MAF CMO	Estimated MAF FVW13	Estimated MAF FVW14	Estimated MAF PHO
2L:655894	0.072	0.445	0	0	0	0
2L:702560*	0.159	0.056	0.375	0.473	0.485	0.336
2L:702798	0.101	0.089	0.077	0.101	0.044	0.034
2L:718623	0.225	0.033	0.051	0.021	0.044	0.100
2L:718627	0.11	0.033	0.055	0.020	0.046	0.099

* This is a complex polymorphism with linked SNPs and INDELs, in Pitchers *et al* (2019) a SNP in this region was found to be linked. However, the variant calling pipeline used in this work recognized an INDEL in this region which was used for counting.

Table 3. Significantly differentiated variants for *ds* shape change from the wild-caught cohorts (BSA).

Location	CMH p-value (FDR corrected)	Gene	FlyBase ID	Distance from ORF (bp)
2R:17491270	0.026	*NT5E-2*	FBgn0050104	0
2R17498059	0.034	*CG30103*	FBgn0050103	2061
2R:17515133	0.022	*CG4853*	FBgn0034230	0
2R: 20537878	0.013	*CG13423*	FBgn0034513	0
2R:23601278	0.005	*CG10332*	FBgn0260455	0
2R:23601278	0.005	*IM18*	FBgn0067903	0
2R:23613785	0.013	*Eglp4*	FBgn0034885	0
2R:23613785	0.013	*Eglp2*	FBgn0034883	0
2R: 23646252	0.016	*retn*	FBgn0004795	0
3L:12831924	0.005	*CG10960*	FBgn0036316	0
3L: 20999119	0.022	*skd*	FBgn0003415	0
3R: 21523866	0.013	*CG7956*	FBgn0038890	0
3R: 2559549	0.011	*Pzl*	FBgn0267430	0
X: 14891220	0.013	*Flo2*	FBgn0264078	0
X: 14891220	0.013	*CG9514*	FBgn0030592	0
X:16039731	0.017	*Muc14a*	FBgn0052580	0
X: 793052	0.011	*CG16989*	FBgn0025621	95
X: 9448676	0.034	*mgl*	FBgn0261260	0

Supplemental Figure 1. Allele Frequency spectra demonstrate more variation at *ds* compared with *emc* in the synthetic outbred population used for artificial selection. Figures show estimated alternate allele frequencies at *ds* and *emc*. Alternate allele frequencies are estimated using parental strain genotype data and assuming an equal contribution from each parent to the founding population. Note the different y axis scale between *ds* and *emc*.

Supplemental Figure 2. Projection of FVW14 wings onto *ds* shape change vector shows clear distinction between *D. melanogaster* and *D. simulans.* Clear separation between the FVW14 samples (black) and *D. simulans* data (gold) indicates that *D. melanogaster* females were accurately identified.

Supplemental Figure 3. Principal component analysis of shape variation within and among populations for wild collected *Drosophila*. First four axes from the PCA for shape variation are shown. (A) and (B) PCA includes all shape variation. (C) and (D) use the 'allometry corrected' landmarks, (residuals from a model regressing landmarks onto centroid size of wings).

Supplemental Figure 4. Shape change effects due to RNAi knockdown of *ds*, *emc* and *neur*. Scaling of effects provided to aid in visualization of shape changes. Different magnifications are provided to account for the disparate magnitudes for estimated shape change vectors: *ds* = 5.5, *emc* = 0.44, *neur* = 2.8. The vectors from these analyses were used for projections in this study.

Supplemental Figure 5. Illustration of projections onto shape vectors to generate shape scores used in this study. The *ds* shape change vector is used for demonstration. (A) Calculating shape score using projection. y_1 and y_2 represent vectors of landmarks for two representative individuals. Dotted arrows represent projection of y_1 and y_2 onto the shape change vector defined by *ds* RNAi to generate the *ds* shape score. (B) Similarly, y_1 and y_2 are projected onto the g_{max} (PC1) vector, representing the direction of maximum genetic variation in the genetic variance-covariance matrix, **G**. (C) Comparing g_{max} and shape change vectors, using the correlation (*r*) of vectors directly, or via the angle θ, between vectors. (D) Hypothetical relationship between *ds* shape scores and g_{max}, indicating a relationship between directions of *ds* induced shape change and the direction of genetic variation in shape variation.

A   B

Supplemental Figure 6. Including wild-caught females in the analyses does not change interpretation for FVW14 data. Correlation between projection of shape data from FVW14 population onto *ds, emc* and *neur* shape change vectors and the first three PCs are calculated from shape data from all wings in the FVW14 population. (A) male only data, (B) females and male data. Inclusion of female data does not change the conclusions drawn from the relationships between directions shape change vectors and PCs. Note that the "flipped" directions of the PCs between (A) and (B) represents the arbitrary sign of the set of eigenvectors for the PCA (i.e. the set of eigenvectors can equivalently be multiplied by -1).

Supplemental Figure 7. Shape variation in field collected samples by projecting individual shape data onto *ds* and *neur* shape change vectors. Projections were performed used size "adjusted" landmark data onto *ds* and *neur* shape change vectors. The correlation between the ds and neur shape scores is r = 0.12. In linear models with shape scores (*ds* and *neur* respectively) regressed onto population and wing size, the partial R^2 for population effects is 0.040 (*ds*) and 0.18 (*neur*).

Supplemental Figure 8. Distribution of wing sizes (males only) in wild caught cohorts. When size is linearly regressed onto population, model $R^2 = 0.27$.

Supplemental Figure 9. **Projections of data onto RNAi shape change vectors are correlated with major axes of shape variation in wild-caught *Drosophila* from each population**. Correlation between projection of shape data from wild cohorts onto *ds, emc* and *neur* shape change vectors and the first three PCs are calculated from shape data from all wings in each cohort independently. Note that the "flipped" directions of the PCs for FVW12 represents the arbitrary sign of the set of eigenvectors for the PCA (i.e. the set of eigenvectors can equivalently be multiplied by -1). Note that the eigenvectors representing PC1 and PC2 in FVW12 have "swapped" because of similarities in variance accounted for.

A

B

Supplemental Figure 10. Variation in wing shape among individuals from artificial selection along the *ds* shape change vector. All wings from generation 7 of up (red) and down (blue) selection lineages plotted, with females in (A) and males in (B). Black wireframe is mean shape from the experiment.

Supplemental Figure 11. Response to artificial selection along the *ds* shape change vector. Regression of cumulative selection differential onto cumulative response was used to estimate realized heritability in both treatments ("up" & "down") independently.

Supplemental Figure 12. Modest (and not significant) changes in wing size following artificial selection based on *ds* shape change. Mean size estimated for up and down *ds* selection lineages estimated from a linear mixed model with replicate lineage fitted as a random effect. Shaded regions represent 95% confidence bands for the correlated response to selection.

Supplemental Figure 13. Genetic differentiation between artificial selection pools (Figure 3B), with core hippo signaling loci (red) and *emc* (purple) marked. Genomic differentiation between up and down selection lineages (F_{ST}) measured in 5000bp windows for the artificial selection along *ds* (A) and *emc* shape change vectors (B). Horizonal grey lines represents 3sd from mean F_{ST}.

```
195   GGATTATATGGGGTTTACTAGAAGTGTATATGACCCATCTTAGTTGGCATATTTGTGCTT   108
28    GGATNATATGGGGTTTACTAGAAGTGTATATGACCCATCTTAGTTGGCATATTTGTGCTT   106
96    GGATNATATGGGGTTTACTAGAAGTGTATATGACCCATCTTAGTTGGCATATTTGTGCTT   106
48    GGATNATATGGGGTTTACTAGAAGTGTATATGACCCATCTTAGTTGGCATATTTGTGCTT   106
59    GGATTATATGGGGTTTACTAGAAGTGTATATGACCCATCTTAGTTGGCATATTTGTGCTT   109
801   GGATTATATGGGGTTTACTAGAAGTGTATATGACCCATCTTAGTTGGCATATTTGTGCTT   108
ref   ggattatatggggtgtacaagaagtgtatatgactcatcttag-----------------   163
129   GGATTATATGGGGTGTACTAGAAGTGTATATGACTTATCTTAG-----------------   102
301   GGATTATATGGGGTGTACTAGAAGTGTATATGACTTATCTTAG-----------------   100
69    GGATTATATGGGGTGTACTAGAAGTGTATATGACTTATCTTAG-----------------   102
385   GGATTATATGGGGTGTACTAGAAGTGTATATGACTTATCTTAG-----------------   100
75    GGATTATATGGGGTGTACTAGAAGTGTATATGACTCATCTTAG-----------------   101
83    GGATTATATGGGGTGTACTAGAAGTGTATATGACTTATCTTAG-----------------   100
491   GGATTATATGGGGTGTACTAGAAGTGTATATGACTCATCTTAG-----------------   101
34    GGATTATATGGGGTGTACAAGAAGTGTATATGACTCATCTTAG-----------------   99
774   GGATTATATGGGGTGTACAAGAAGTGTATATGACTCATCTTAG-----------------   102
      ****  *********  ***  ***************   *******

195   AGGTGCATATTTCTAACAAAAATCCGTAACTATTTTGAGTGGCCAAGAAAATTTCAAATT   168
28    AGGTGCATATTTCTAACAAAAATCCGTAACTATTTTGAGTGGCCAAGAAAATTTCAAATT   166
96    AGGTGCATATTTCTAACAAAAATCCGTAACTATTTTGAGTGGCCAAGAAAATTTCAAATT   166
48    AGGTGCATATTTCTAACAAAAATCCGTAACTATTTTGAGTGGCCAAGAAAATTTCAAATT   166
59    AGGTGCATATTTCTAACAAAAATCCGTAACTATTTTGAGTGGCCAAGAAAATTTCAAATT   169
801   AGGTGCATATTTCTAACAAAAATCCGTAACTATTTTGAGTGGCCAAGAAAATTTCAAATT   168
ref   -ttggcatatttctaacaaaaatccgtaactattttgagtgtccaagaaaatttcaaa-t   221
129   -TTGGCATATTTCTAACAAAAATCCGTAACTATTTTGAGTGGTAGAAA-ATTTCAAAT-T   159
301   -TTGGCATATTTCTAACAAAAATCCGTAACTATTTTGAGTGGTAGAAA-ATTTCAAAT-T   157
69    -TTGGCATATTTCTAACAAAAATCCGTAACTATTTTGAGTGGTAGAAA-ATTTCAAAT-T   159
385   -TTGGCATATTTCTAACAAAAATCCGTAACTATTTTGAGTGGTAGAAA-ATTTCAATT-T   157
75    -TTGGCATATTTCTAACAAAAATCCGTAACTATTTTGAGTGGTAGAAA-ATTTCAAAT-T   158
83    -TTGGCATATTTCTAACAAAAATCCGTAACTATTTTGAGTGGTAGAAA-ATTTCAAAT-T   157
491   -TTGGCATATTTCTAACAAAAATCCGTAACTATTTTGAGTGGTAGAAA-ATTTCAAAT-T   158
34    -TTGGCATATTTCTAACAAAAATCCGTAACTATTTTGAGTGTCCAAGAAAATTTCAAA-T   157
774   -TTGGCATATTTCTAACAAAAATCCGTAACTATTTTGAGTGTCCAAGAAAATTTCAAA-T   160
       ***********************************        *  *  * **   *    *

195   TTTTTTTGCGTGTGCTTGGATGTTGTAATTTTGCATTGTTTGGCTAATCTACGTAAATCT   228
28    TTTTTTTGCGTGTGCTTGGATGTTGTAATTTTGCATTGTTTGGCTAATCTACGTAAATCT   226
96    TTTTTTTGCGTGTGCTTGGATGTTGTAATTTTGCATTGTTTGGCTAATCTACGTAAATCT   226
48    TTTTTTTGCGTGTGCTTGGATGTTGTAATTTTGCATTGTTTGGCTAATCTACGTAAATCT   226
59    TTTTTTTGCGTGTGCTTGGATGTTGTAATTTTGCATTGTTTGGCTAATCTACGTAAATCT   229
801   TTTTTTTGCGTGTGCTTGGATGTTGTAATTTTGCATTGTTTGGCTAATCTACGTAAATCT   228
ref   tttttttgcgtgtgcttggatgttgtaatttcgcattgtttggctaatctacgtaaatct   281
129   TTTTTTTGCGTGTGCTTGGATGTTGTAATTTCGCATTGTTTGGCTAATCTACGTAAATCT   219
301   TTTTTTTGCGTGTGCTTGGATGTTGTAATTTCGCATTGTTTGGCTAATCTACGTAAATCT   217
69    TTTTTTTGCGTGTGCTTGGATGTTGTAATTTCGCATTGTTTGGCTAATCTACGTAAATCT   219
385   TTTTTTTGCGTGTGCTTGGATGTTGTAATTTCGCATTGTTTGGCTAATCTACGTAAATCT   217
75    TTTTTTTGCGTGTGCTTGGATGTTGTAATTTCGCATTGTTTGGCTAATCTACGTAAATCT   218
83    TTTTTTTGCGTGTGCTTGGATGTTGTAATTTCGCATTGTTTGGCTAATCTACGTAAATCT   217
491   TTTTTTTGCGTGTGCTTGGATGTTGTAATTTCGCATTGTTTGGCTAATCTACGTAAATCT   218
34    TTTTTTTGCGTGTGCTTGGATGTTGTAATTTCGCATTGTTTGGCTAATCTACGTAAATCT   217
774   TTTTTTTGCGTGTGCTTGGATGTTGTAATTTCGCATTGTTTGGCTAATCTACGTAAATCT   220
      ******************************** ***************************
```

Supplemental Figure 14. Alignment of sanger sequencing of the region containing the *ds* polymorphism from several DGRP lines (line numbers indicated at line starts) to reference sequence (from *Drosophila melanogaster* genome). DGRP lines are indicated with DGRP 195, 28, 96, 48, 59 and 801 predicted to have the insertion, and, DGRP 129, 301, 69, 385, 75, 83, 491, 34 and 774 without. In addition to the insertion in these lines, several other associated SNPs are found linked to the insertion.

Supplemental Figure 15: Permutation test for over representation of hippo signaling terms in outlier regions. The red line represents the observed ratio of significant to expected genes in outlier windows (F_{ST} greater than 3 standard deviations from the mean). The permutation test selected random windows from the genome, equal in number to those identified as outliers. For each random draw of loci, the ratio of significant to expected genes in the term of interest was calculated for 1000 permutations.

Supplemental Figure 16. Response to selection based on projections onto *emc* shape change vector. Regression of cumulative selection differential onto cumulative response was used to estimate realized heritability in both "up" and "down" selection regimes independently.

Supplemental Figure 17. Minimal genetic structure among wild populations used in this study. Principal coordinate plot calculated from Bray's distances estimated from allele frequency data (altF), between the four wild cohorts used in this analysis. Axis 1-3 explains 45%, 30% and 25% of the variance in genetic distances respectively.

Supplemental Figure 18. Genetic distance is not correlated with phenotypic distance among wild cohorts used in this study. Phenotypic distances are the Euclidian distances from model estimated shape vectors for each population from a multivariate regression of shape onto for population and wing centroid size.

Supplemental Figure 19. Shape variation within *ds* selected pools for bulk segregant analysis by population. Wings within selected pools (one red and one blue, representing the "up" and "down" pools) are plotted to show the phenotypic extremes for the shape scores used for selecting individuals within and between pools. Black line indicates mean shape between pools.

Supplemental Figure 20. Removing PHO from the BSA analysis increases the number of differentiated sites between *ds* shape change pools. Genome-wide scan for differentiated loci between pools selected based on *ds* shape change vector using the CMH test using ACER. Variants in *ds* are still not implicated in this analysis. Points in red indicate sites with significant differentiation based with a FDR of 0.05.

Supplemental Figure 21. Down sampling genome coverage does not impact results for CMH tests substantially. When all pools are sampled to a coverage depth of 75 reads, while preserving allele frequency, we do not find an increase in number of differentiated sites. Although the sites identified as significantly differentiated do change somewhat, in part due to sampling procedure as any sites without a minimum depth of 75 reads in each of the 4 populations was dropped from this analysis. Variants in *ds* are still not implicated in this analysis.

Supplemental Figure 22. F_{ST} between pools of individuals selected along the *ds* shape change axis within each population. Calculated in 100 bp windows using PoPoolation2 program. Elevated F_{ST} on the X chromosome is due to sampling of fewer X chromosomes, relative to autosomes as most pools, with the exception of FVW14, consist of only males. Red line indicates the mean F_{ST} for the chromosome, which as expected is very low.

Supplemental Figure 23. Shape variation within *neur* selected pools for bulk segregant analysis by population. Wings within selected pools (one red and one blue, representing "up" and "down" extreme pools respectively) are plotted.

Supplemental Figure 24. Genome-wide scan for differentiated loci between pools selected based on *neur* shape change vector using the CMH test implemented with ACER. (A) Whole genome scan for differentiation. Points in red indicate sites with significant differentiation based with a FDR of 0.05. (B) No significantly differentiated sites within *neur.* (C) shape change between selected pools based on *neur* shape change vector, effect size is multiplied by 2 for visualization.

Supplemental Figure 25. F_{ST} between pools of individuals selected along the *neur* shape change axis within each population. Calculated in 100 bp windows using PoPoolation2 program. Elevated F_{ST} on the X chromosome is due to sampling of fewer X chromosomes, relative to autosomes as most pools, with the exception of FVW14, consist of only males. Red line indicates the mean F_{ST} for the chromosome.

Supplemental Table 1. Number of individuals used for BSA from each wild-caught cohort

Collection	Females Phenotyped	Males Phenotyped	Females – "up" pool	Females – "down" pool	Males – "up" pool	Males – "down" pool
FVW13	0	2184	0	0	75	75
FVW14	403	797	59	0	16	75
PHO14	0	1232	0	0	75	75
CMO14	0	1001	0	0	75	75

Supplemental Table 2. Directions of major axes of variation differ somewhat between the among line (DGRP) covariance matrix and the wild caught cohorts, likely due to a combination of differences in allele frequencies and the major contribution of environmental variation (and potentially GxE) in the wild caught individuals. Correlations for the first three eigenvectors computed from among the DGRP, wild-cohorts (all populations) and CMO shape variation. |Estimates| < 10^{-15} are treated as 0.

	dgrp.PC1	dgrp.PC2	dgrp.PC3	wild.PC1	wild.PC2	wild.PC3	cmo.PC1	cmo.PC2	cmo.PC3
dgrp.PC1	1	0	0	-0.17	-0.46	0.26	-0.35	-0.54	0.066
dgrp.PC2		1	0	-0.18	-0.21	0.61	-0.10	-0.60	-0.27
dgrp.PC3			1	0.56	-0.12	0.40	-0.32	0.27	-0.71
wild.PC1				1	0	0	-0.39	0.68	-0.29
wild.PC2					1	0	0.90	0.31	-0.061
wild.PC3						1	0.056	-0.55	-0.74
cmo.PC1							1	0	0
cmo.PC2								1	0
cmo.PC3									1

Supplemental Table 3. Top 50 enriched GO terms of linked differentiated sites following artificial selection based on *ds* shape change
See: ds_artSelGOtop50.csv

Supplemental Table 4. Top 50 enriched GO terms of linked differentiated sites following artificial selection based on *emc* shape change
See: emc_artSelGOtop50_3sd.csv

Supplemental Table 5. Pairwise Procrustes (Euclidian approximation) distances between mean shapes across wild cohorts. Pairwise comparisons between population means estimated from a model with fixed effects for centroid size and population. Z-score and p-value are calculated using permutation of residuals as implemented in the RRPP package using a null model with only the effect of centroid size. Comparisons to the PHO samples, which show the largest pairwise differences are in bold.

Population Comparison	Distance	95% permutation threshold	Z-score	p
CMO-FVW13	0.0070	0.0011	9.70	0.001
CMO-FVW14	0.0084	0.0014	11.0	0.001
CMO-PHO	**0.021**	**0.0013**	**15.2**	**0.001**
FVW13-FVW14	0.0065	0.0012	7.25	0.001
FVW13-PHO	**0.022**	**0.0012**	**13.0**	**0.001**
FVW13-PHO	**0.022**	**0.0014**	**13.5**	**0.001**

Supplemental table 6. Significantly differentiated sites when PHO population is left out of CMH test.
See: ds_wild_significantSites_zeroGen_genes_fdr_noPho.csv

Supplemental table 7. GO analysis of significantly differentiated sites found in CMH test when PHO population is left out of analysis.
See: GOanalysisTop50_cmhSites_fdr_noPHO_ds_wild.csv

Supplemental Table 8. Significantly differentiated variants for *ds* shape change (CMH test) from the wild-caught cohorts (BSA) when all data is down-sampled to 75X coverage per site.

Location	CMH p-value (FDR corrected)	Gene	FlyBase ID	Distance from ORF (bp)
X:14833408	5.42 x10^{-5}	*CG14411*	FBgn0030582	0
X:14833412	3.29 x10^{-5}	*CG14411*	FBgn0030582	0
X:15381676	7.54 x10^{-5}	*CG9164*	FBgn0030634	0
X:16037404	2.21 x10^{-4}	*Muc14A*	FBgn0052580	0
X:22956575	4.58 x10^{-2}	*lncRNA:CR45502*	FBgn0267058	1501
X:4118816	9.49 x10^{-3}	*tyf*	FBgn0026083	0
X:4130993	9.84 x10^{-3}	*GlcAT-I*	FBgn0066114	218
2L:14793860	1.15 x10^{-2}	*CG18420*	FBgn0028866	656
2L:1480390	3.03 x10^{-2}	*CG31928*	FBgn0051928	11905
2L:20626829	7.29 x10^{-4}	*lncRNA:CR44909*	FBgn0266214	623
2L:4481772	1.54 x10^{-2}	*dpy*	FBgn0053196	0
2L:4481788	3.70 x10^{-2}	*dpy*	FBgn0053196	0
2L:4843588	4.62 x10^{-2}	*mxt*	FBgn0031637	0
2R:11876761	2.73 x10^{-2}	*CG13185*	FBgn0033661	0
2R:16405045	3.70 x10^{-6}	*Sema2b*	FBgn0264273	0
2R:9943212	5.15 x10^{-3}	*Mef2*	FBgn0011656	0
3L:19082031	4.49 x10^{-3}	*Mkp3*	FBgn0036844	0
3L:19082032	2.73 x10^{-3}	*Mkp3*	FBgn0036844	0
3L:23237904	3.05 x10^{-2}	*nAChRalpha4*	FBgn0266347	0

Supplemental table 9. Significantly differentiated variants for *neur* shape change from the wild-caught cohorts (BSA).

Location	CMH p-value (FDR corrected)	Gene	FB ID	Distance from ORF (bp)
2L:15775767	0.026	*CG43760*	FBgn0264260	675
2L:15967220	0.017	*Beat-Ib*	FBgn0028645	4785
2L:15967222	0.017	*Beat-Ib*	FBgn0028645	4787
2L:8651597	0.015	*Sema1a*	FBgn0011259	0
3R: 21898160	0.024	*CG6678*	FBgn0038917	0