CORRESPONDENCE OF NEWCOMB-BENFORD NUMBER LAW WITH ECOLOGICAL PROCESSES

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ABSTRACT

The Newcomb-Benford number law has been applied in the natural sciences for decades, with little ecological attention. Empirical data reveal statistical correspondence between the discrete Benford probability distribution and physical systems in dynamic equilibrium along a continuum of stability. Analytic methods are presented to detect this mathematical representation across multiple levels of ecological organization and spatial scale. Case studies demonstrate novel application to help identify bidirectional regime changes to alternative states of dynamic equilibrium. Widespread documentation of the surprising phenomenon is anticipated as ecologists revisit historic sets of random measurement data and design future sampling protocols. Controlled experiments with measurement variables that span multiple orders of magnitude would be well suited for future empirical and theoretical inquiry.

Introduction

Organization in ecological systems results from negative and positive feedback interactions between organisms and their abiotic environment [1]. The complex mix of stochastic and non-random (natural selection) interactions generates interest to identify patterns that unify ecological processes at multiple hierarchical levels and spatial scales [2]. Pattern in empirical data has been attributed to 'law-like'

mathematical representations [3-6], a phenomenon not well investigated in ecology

[7].

The purpose of this note is to communicate statistical correspondence between a

long-known probability distribution (the Newcomb-Benford law [NBL], Newcomb,

1881 [8], Benford, 1938 [9], Pinkham, 1961 [10]) with ecological processes related to

complexity and stability. The Benford distribution is an empirical observation that the

first significant digits (FSD) from certain sets of naturally occurring data (i.e., with

minimal human intervention) follow a discrete logarithmic probability distribution

(Figure 1, Table 1). The NBL has been applied in the natural sciences for decades

[11-18], with little ecological attention. Potential ecological applications arise from

reports the dynamic equilibrium of physical and social systems (defined *sensu* Thoms

et al. [19] as long-term <u>balanced</u> fluctuations about short-term constantly changing

system conditions) can be well described by the Benford FSD sequence (Table 2

[14-18]). The recursive theme is that interactions of system components that maintain

structural stability [20] (i.e., persistence, constancy, resilience) can be represented by

the informational distance between two discrete probability distributions, one from a

FSD transformation of random measurements on system variables, the second from

the regularity of FSD probabilities predicted by the Newcomb-Benford law (see

Appendix Figures 1-4). As the distance between distributions increases in response to

external perturbations, build-up of informational uncertainty can lead to a phase

transition away from a balanced threshold of minimal entropy production [21-23].

Prigogine [24] identified 'dissipative structures' to denote self-organized dynamic

systems with a propensity of minimal entropy production over time to maintain

stability. The phenomenon is consistent with Ulanowicz's ascendency view of

ecological stability [25], defined as the ability of an ecosystem to "*remain within an*

arbitrarily-designated nominal range of behavior" in response to changing

environmental conditions.

Figure 1. The discrete Benford probability distribution of first significant digits:

$P(d)$=Log$_{10}$ [1+(1/d)], where $d=1,2,…9$ [8,9].

Given that dynamical ecological systems are open to short-term flux of matter,

energy, and information, they also should adhere to reports of stability predicted by

the first significant digit sequence of the NBL (Table 2). Anecdotal evidence comes

from observations of seasonally predictable counts of lake algae that are well

described by the Benford distribution [26,27], persistent gene sequences over

evolutionary time scales that distinguish taxonomic categories above the species level

[28], and annually consistent pollen counts from atmospheric samples [29]. Biotic

communities at equilibrium have been associated with the log-normal distribution

[30], which mathematically conforms with the Benford distribution [31]. Skewed

Benford-like probability distributions (Figure 1) predict food-web stability [32] and

have been suggested for use as an early warning signal of impending regime change

for otherwise stable ecosystems [33].

Table 1. First significant digit relative frequencies for the Benford probability

distribution.

First Digit	Relative Frequency	Fractional
1	0.30103	$Log_{10}(2/1)$
2	0.17609	$Log_{10}(3/2)$
3	0.12494	$Log_{10}(4/3)$
4	0.09691	$Log_{10}(5/4)$
5	0.07918	$Log_{10}(6/5)$
6	0.06695	$Log_{10}(7/6)$
7	0.05799	$Log_{10}(8/7)$
8	0.05115	$Log_{10}(9/8)$
9	0.04576	$Log_{10}(10/9)$

75 **Table 2. Physical and social systems in stable dynamic equilibrium that conform**

76 **to the first significant digits frequencies of the Benford distribution.**

Modes of Dynamic Equilibrium	References
Atmospheric stability	Li and Fu, 2016 [14]
Stability of cardiac processes	Seenivasan et al., 2016 [15]
Persistent economic patterns	Villas-Boas et al., 2015 [16]
Recurrence of seismic event times	Sttili, et al., 2012 [17]
Quantum phase transitions	De and Sen, 2011 [18]

Methods

Background The Newcomb-Benford law, also known as the first digit law and Benford's law, is a mathematical oddity in the probability of first significant (non-zero) digits observed from certain sets of naturally occurring data (Figure 1). A Benford distribution emerges when a set of random measurements on system components i.e., (0.061, 0.14, 9.21, 20.2, 23.0, 723.0, 3345.0,…) are transformed into a compressed vector of FSD (6, 1, 9, 2, 2, 7, 3,…). The expected probability P(d) of leading digits (d) conforms to the logarithmic equation:

$$\text{P}(d) = \text{Log}_{10}\,[1 + (1/d)], \text{ where } d = 1, 2, \ldots 9 \tag{1}$$

The leading digit #1 appears about 30% of the time, digit #9 near 5%, with a Log_{10} digit sequence (Table 1). First presented by Newcomb in 1881 [8], the phenomenon was rediscovered in 1938 by Benford [9] and empirically verified from 20,000+ measurements taken on sets of naturally occurring data compiled across multiple fields of study. The Benford distribution is uniquely both scale [10] and base [34] invariant and strongly correlated (R^2=0.998, Table 1 data) with a power-law function over the first digit sequence (X1:9): Probability (Y)=0.3135 X $^{(-0.8637)}$. Increased conformity to law predictions is expected when data from multiple variables are aggregated. Additional information from analysis of second and third significant digits can improve the predictive power for some applications [35].

Hill [36] derived the Benford distribution to be a phenomenon of collecting random samples from multiple random distributions: *"if distributions are selected at random (in an unbiased way) and random samples are taken from each of the distributions, then the significant digit frequencies of the combined samples converge to Benford's distribution though the individual distributions selected may not closely follow the law."* The correspondence is not intuitive. One would expect a set of

random measurements collected on organisms, perhaps the biomass of species nested

within a herbivore functional group in a grassland habitat, to yield a FSD sequence

with as many biomass measurements that begin with 1, 2, 3 as 7, 8, 9 (e.g.,~33.3% per

triplet set), however, the NBL informs the final naturally occurring frequencies will

be 60.2% and 15.5% respectively (Table 1).

Empirical support for the Benford phenomenon comes from multiple fields of

study including weight of molecular compounds, numbers on economic spreadsheets,

algae cell counts in lakes, earthquakes, and astronomic bright objects identified by the

Fermi space telescope [12]. Mathematical functions of interest to ecologists that

conform to the Benford distribution include: reciprocal density [37], exponential [38],

log-normal [31], Laplace transform [39], Markov processes [40], and Poisson [41].

The most common application is from economics where deviation from a FSD

sequence is used to detect the falsification of data on spreadsheets and tax returns

[35,42]. Mathematical reviews include Berger and Hill, [43] and Miller [44]. A

real-time online bibliography of Berger, Hill, and Rogers [45] provides 1000+ articles

and books with theoretical and applied examples of the Newcomb-Benford law:

http://www.benfordonline.net.

A variety of constraints (measurement requirements) associated with the

collection and reporting of data are prerequisites for the emergence of the Benford

distribution (Table 3). An increase in the number and magnitude of *constraint*

violations increases the likelihood the set of data under investigation will not conform

with NBL predictions. A review of literature identifies the following constraints

associated with ecological applications: (1) the collection of data follows methods of

random sampling such that each of the nine first digit categories has the same chance

of occurrence for each measurement taken, with location of sample sites chosen to

include habitat heterogeneity (i.e., such as riffles, runs, and deep pools in lotic

systems); (2) the set of data span several orders of magnitude, ideally (10^6) suggested

by Fewster [46], but as few as (10^2) reported here; (3) total sample size is sufficient to

populate all FSD categories and appropriate for methods adopted for statistical

inference; (4) the full range of data analyzed is not filtered within arbitrary limits; (5)

the sampled habitat is naturally occurring with regional background levels of global

environmental disturbance from climate change, and minimal direct anthropogenic

perturbation (i.e., habitat alteration, chemical pollution); (6) tabulation of data is free

from intentional human alteration; and (7) tabulated data are generated by

non-additive mathematical operations such as multiplication, division, or numbers

raised to integer powers [42,47], which are associated with multiple biotic processes.

The Benford FSD phenomenon has been demonstrated for a variety of ecological

variables including counts of individuals and taxa, biomass, animal movements, and

taxonomic classifications. Cell counts/colony of the cyanobacterium *Microcystis*

aeruginosa collected from five lakes in Spain were well described by NBL

predictions [25], as well as persistent communities of lake algae sampled over five

years [26] and annually consistent pollen counts from atmospheric samples [28].

Campos et al. [48] observed that naturally assigned taxonomic categories of

angiosperms conformed with NBL predictions, those artificially assigned did not. The

number/ha of Red-Listed indicator species from 2707 conservation forest plots in

Sweden visually conformed with the Benford distribution [49]. Pröger et al. [50]

applied a Benford FSD analysis to telemetry data for wild red deer freely moving in

natural habitats. Ninety one percent of 1132 weekly sets of data conformed with NBL

predictions. Disturbance from human activity was suggested as a possible causal

factor for non-conformance.

Table 3. Data collection and reporting constraints associated with emergence of the Benford distribution.

Data Collection Constraints
Measurements on organisms from random sampling
Set of data with multiple orders of magnitude
Sample size (N) sufficient to populate all digit categories
Range of data not artificially limited
Tabulated data free from human transcription error
Sampled habitat naturally occurring *
Habitat with minimal direct human disturbance
Data from non-additive mathematical operations **

* Includes *in situ* experimental protocols

** Includes multiplication, division, and numbers raised to powers

Statistical Analysis

Empirical data from the ecological literature were analyzed to test a proposition that naturally occurring ecological systems reported to be progressing toward stable dynamic equilibrium can be well represented by the Benford probability distribution. A weight-of-evidence analytic approach was adopted to quantify differences between microscopic and macroscopic attributes of observed *vs* expected FSD probability distributions.

Analytic Method #1: The following macroscopic null hypothesis (H_O) was adopted: the measured nine FSD probability sequence is well represented by the sequence predicted by the Newcomb-Benford law, with alternative hypothesis (H_A): the measured FSD sequence is not well represented. The (H_O) was tested with a nine-dimensional Euclidean Distance formula, recommended for Benford applications

170 by Coe and Gaines [51], as modified by Morrow [52] to allow for statistical inference:

$$d*_N = (\sqrt{N}) \sqrt{\sum_{d1}^{9} [(\text{Obs} - \text{Exp})]^2} \qquad (2)$$

where d = digits (1:9), N = total sample size of observations, Obs = probability a

random measurement x has a FSD = d for total sample size N, and Exp =

Newcomb-Benford law expected FSD probabilities from Table 1 as $\text{Log}_{10} (1+1/d)$.

A (H_O) rejection range (alpha 0.10 $d*_N$ = 1.22; alpha 0.05 $d*_N$ = 1.33; alpha 0.01

$d*_N$ = 1.57) is presented by Morrow derived from Monte Carlo simulation calibrated

for sample size: N>80<500. Values of $d*_N$ >1.33 reject the (H_O). This narrow range of

sample size is a limitation in the use of the Morrow ($d*_N$) statistic, but should be

sufficient for most ecological applications. Euclidean distance (>1.22<1.33) was

adopted to serve as an early warning signal of impending regime change. Goodness of

fit statistics for Benford applications can yield false negative errors at larger sample

sizes [51,52,53]. Kossovsky [53] advocates use of squared deviations (SSD) to

quantify the magnitude of distance between measured data with NBL predictions. The

SSD is calculated as percentages of $[(\text{Obs-Exp})]^2$ without consideration of the ($\sqrt{N}$)

sample size operator (equation #2). Kossovsky proposes SSD>100 as a signal that

measurement data are not well represented by the Benford probability distribution.

Analytic Method #2:   Cohen [54] recommends calculation of *effect size* to

estimate the magnitude of distance between two probability phenomena to

complement interpretation of results from statistical inference. As relates to a Benford

FSD analysis of data, a higher effect size is expected with an increase in the number

and magnitude of constraint violations associated with collection of data (see Table 3).

One measure of effect size that is independent of sample size and compliments

information from Euclidean distance is the Cohen-W:

$$\text{Cohen-W} = \sqrt{\sum_1^9 [[\,(\text{Obs} - \text{Exp})\,]^2/\text{Exp}]} \qquad (3)$$

where Obs and Exp are the same as used to calculate Euclidean distance (d^*_N). Cohen

[54] offers thresholds to identify weak, moderate, and strong effect sizes respectively

(W=0.1, 0.3, 0.5). A Cohen-W>0.5 provides strong non-statistical inference of

increased magnitude of constraint violations leading to non-conformity of NBL

predictions. A Cohen-W range (>0.3<0.5) was adopted to serve as an early warning

signal of impending regime change resulting from effects on skewness related with

constraint violations.

Analytic Method #3: Microscopic multivariate volatility of FSD frequencies

was estimated from the Pearson Residual (PR) formula presented by Sharpe [55]:

$$\text{PR} = (\,|(\text{Obs count}) - (\text{Exp count})|) / \sqrt{\text{Exp count}} \qquad (4)$$

where Obs = measured FSD count frequency, and Exp = predicted FSD frequency

based on total sample size N (e.g., from Table 1, the Exp counts for FSD #1=

0.30103*N). A PR value >1.96 (Z-Test, p = 0.05) for any FSD category was adopted

as an early warning signal of impending regime change related to interaction effects

of organisms with each other and their abiotic environment.

Results

Analytic results for four case reports are presented in Appendix A and summarized in

Table 4. The weight of evidence supports a proposition that naturally occurring

ecological systems progressing toward a state of stable dynamic equilibrium can be

well represented by the first significant digit sequence of the Newcomb-Benford law.

Table 4. Summary of analytic test results from Appendix A case reports.

Case report number & FSD comparison	N	OM	d*$_N$	Max PR	Cohen W	SSD
1: Below ground food-web stability: Early succession stages	159 #	6	0.782	1.38	0.184	38.4
1: Below ground food-web stability: Late succession stages	208 #	6	0.664	1.59	0.171	21.2
2: Experimental diatom colonization: Small islands	105 #	6	0.759	1.87	0.250	54.9
2: Experimental diatom colonization: Large islands	148	6	1.147	**1.98** *	**0.316** *	88.9
3: *Plethodon* salamander populations: Pre-anthropogenic effects	96 #	2	0.783	1.13	0.195	63.9
3: *Plethodon* salamander populations: Post-anthropogenic effects	173	3	**1.437** **	**2.23** *	**0.328** *	**119** **
4: Regional biomass of lotic freshwater fish communities with exceptional biotic integrity	141 #	3	0.638	1.46	0.182	28.9

N=FSD sample size, OM=orders of magnitude, d*$_N$= Modified Euclidean Distance, Max PR=Maximum Pearson Residual, Cohen W=Cohen-W effect size, SSD=Kossovsky Sum of Square Deviations, *=Early warning signal of pending bidirectional regime change, **=Non-conformity of measured data with NBL predictions, #=Inclusive analytic conformity.

The phenomenon was demonstrated at multiple levels of organization and spatial scales: (1) annual functional group biomass of below ground food-webs during developmental stages of succession, (2) counts of diatom cells on glass slides progressing toward dynamic equilibrium of taxa richness during island colonization, (3) constancy in abundances of meta-populations of woodland salamanders over a

large geographic area, and (4) biomass of isolated meta-communities of lotic freshwater fish with persistent biotic integrity. Inclusive conformity with FSD predictions of the the Benford distribution, for five case report comparisons, was within analytical ranges far away from a threshold limit of minimal entropy production (Table 4): Euclidean distance (d^*_N= 0.638-0.783), Cohen-W effect size (0.171- 0.250), and Kossovsky SSD (21.2-63.9). Results from two case studies highlight the importance of including analytic test results from Pearson Residuals and Cohen-W effect size to help detect signals of impending regime change to compliment information from goodness-of-fit statistical inference.

Discussion

MacArthur [57] suggested the science of ecology is in large part a search for repeating mathematical patterns in nature. The Fibonacci pattern (1,1,2,3,5,8,13,21…), first applied to explain breeding in rabbits, is associated with the phenotypic expression of amino acids during embryonic development [58]. Margalef [59] equated the cybernetic stability of ecosystems to the 'constancy of numbers'. Data presented here support a hypothesis that ecological systems progressing toward a threshold of stable dynamic equilibrium can be well represented by constancy in the sequence of first significant digits predicted by the Newcomb-Benford number law (Figure 1). The results concur with similar observations for multiple physical and social systems (Table 2), mathematical constants of nature [56], and anecdotal biotic observations [26-29,48-50].

A Benford FSD analysis of random measurement data presents a variety of theoretical and empirical lines of inquiry for future ecological research: (1) demonstration of sample size effects on NBL predictions in the range N>80<500 at

multiple levels of organization and spatial scale, (2) expansion of the list of

measurement variables, (3) identification of alternative statistical methods for early

identification of regime shifts, (4) experiments investigating data collection

constraints associated with human perturbation of environmental conditions, (5)

analyses of second and third significant digits to improve predictive power, (6)

comparisons of random sampling methods, (7) sets of big-data created via remote

sensing [60], and (8) a FSD analysis using metabolism of detritus as a measurement

variable, suggested by Wetzel [61] to be an integral driver of stability for both aquatic

and terrestrial ecosystems. Further investigation also is warranted concerning the

mechanisms by which the Newcomb-Benford mathematical representation is

compatible with ecological concepts related to scaling laws, including Taylor's power

law [62], MacArthur's broken stick model of niche appointments [63], maximum

information entropy [64,65], effective complexity [66], ecosystem sustainability [67],

and cybernetic Markovian processes [68,40].

Skewed probability distributions have been suggested to detect ecological

systems that depart from stability thresholds [33]. Case reports suggest a hypothesis

that changes in skewness characteristics of the Benford FSD distribution can be used

as a bidirectional early warning signal of regime change. Results from a colonization

experiment (case report #2) document the natural progression of a diatom community

towards increasing dynamic equilibrium of species richness as predicted by theory of

island biogeography, whereas analysis of long-term observations of persistent

meta-populations of woodland salamanders (case report #3) reveal regime change

away from stable species abundances in response to human-influenced environmental

perturbation.

An important caveat for ecological application is that not all biotic systems in

stable dynamic equilibrium that are well described by NBL predictions will

necessarily be of beneficial use to humans. Examples include lentic habitats with

seasonally stable blooms of cyanobacteria that can reach limits of toxicity to wildlife,

and biotic communities with dominant and stable populations of exotic species that

alter community biotic integrity and well-being. Research directed toward a mix of

sample variables is suggested to partition the detection of the Benford distribution

with and without the presence of non-beneficial populations.

Consensual agreement on the measurement of ecological stability remains elusive.

The prevailing view of Landi et al. [69] is that no single system parameter, such as

species diversity, can aggregate all possible interpretations of stability. Ives and

Carpenter [70] argue for a holistic approach: "*it will be more profitable to study*

stability comprehensively, including diversity as only one of the possible factors that

affect ecosystem responses to environmental change." A Benford FSD analysis of

measurement data can serve to quantify the comprehensive stability of complex

adaptive systems. Following Ulanowicz's ascendency view of stable ecological

development [25], as measurements collected on organisms transition from one

microscopic FSD category to another (i.e., a measurement at time x of 0.19 transforms

to 0.20 at $x+1$ and populates a higher first digit category), the macroscopic sequence

of FSD probabilities remains within a threshold range of minimal entropy production.

Comprehensive system stability (*sensu* Ives and Carpenter) emerges as information

gained from the FSD probabilities of measurement *variables* (i.e., system properties

that vary over time from one individual to another), with additional data from

statistical *parameters*, such as species diversity, applied to dynamical modelling to

monitor factors that affect responses of structural stability to environmental change.

Wilson [71] proposed that future research in ecology be directed towards a better

understanding of *"transformations that reassemble components in ways that retain*

system properties holistically," with future success measured by the ability of such

transformations to identify emergent phenomena that are useful. A simple Benford

FSD transformation of arithmetic data randomly collected on organisms, reassembled

into a mathematical representation of first significant digit categories of logarithmic

information, affords one holistic approach forward to investigate this consilience goal.

Conclusion

Reviewed empirical data support a hypothesis that complex adapted systems

trending toward a threshold of minimal entropy production and stable dynamic

equilibrium can be well represented by the Benford probability distribution of first

significant digits. Analytic methods using multidimensional Euclidean distance and

Cohen-W effect size magnitude are presented to detect the surprising correspondence

at multiple levels of ecological hierarchical organization and spatial scales. Case

studies demonstrate that the Benford distribution can serve as a bidirectional early

warning signal of regime change in response to both naturally occurring and

anthropogenic modified environmental conditions. Widespread documentation of the

mathematical representation is expected as ecologists revisit historic data and design

sampling protocols with attention to constraints associated with collection of data.

Experiments using conserved variables that span multiple orders of magnitude, for

populations with a short generation time, would be well suited for future empirical

and theoretical inquiry. A simple transformation of an arithmetic set of random

measurement data into a vector of first significant digits, nested into discrete

categories of logarithmic probabilities, offers novel potential to monitor the effective

comprehensive stability and complexity of complex adaptive systems over ecological

and evolutionary lengths of time.

Data Availability

The availability of data cited in Appendix A case reports is included in the text.

Conflicts of Interest

The author declares that there is no conflict of interest regarding the publication

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                          APPENDIX A: CASE REPORTS

 Case Report #1: Food-web stability during succession

    Background: Neutel et al., [72] sampled two below ground food-webs

 representing four developmental stages of succession at the Schiermonnikoog [SCH]

 island national park and Hulschorsterzand [HUL] mainland nature reserve,

 Netherlands. Food-web stability was quantified from dynamic models related to three

 omnivorous feedback loops. Non-random predator-prey interactions within these

 loops were reported to play an important role preserving food-web stability at both

 sample locations as productivity and functional group complexity increased during

 succession.

    Mean functional group biomass data (gms C ha^{-1} cm depth^{-1}) are provided in

 online supplemental Table 1 of Neutel et al. [doi:10.1038/nature06154]. Soil samples

 were randomly collected in upper soil layers with four replications per development

 stage, three times, to establish yearly mean biomass statistics. Measurements span six

 orders of magnitude at both sample locations with sufficient sample size to test

 conformity with NBL predictions when data are combined for both SCH and HUL

 locations, and segregated into early (1,2) and late (3,4) developmental stages.

 Minimal anthropogenic disturbance is expected for the two natural areas.

    Results and discussion: Functional group biomass counts at combined SCH and

 HUL locations provide insufficient weight of evidence to reject the tested (H$_O$) of

 similarity between observed *vs* NBL expected FSD probability distributions

 (Appendix Table 1). Pearson Residuals (max<1.96) and Cohen-W effect sizes (max

 <0.18) agree with results from Euclidean distance (max d*n<0.782, p>0.10) that mean

 annual functional group biomass is well described by the Benford distribution, within

 a nominal range of uncertainty, at both early and late developmental stages of

succession (Appendix Table 1). The Benford FSD analysis agrees with findings of

Neutel et al. that the SCH and HUL food-webs have maintained stable dynamic

equilibrium over time (Appendix Figure 1). No evidence of pending regime change

was observed for combined SCH and HUL locations. As documented by Neutel et al.,

[72] an increase in FSD functional group diversity over time (N 159 to 208, Appendix

Table 1) did not lead to a decrease in comprehensive food-web stability as would be

predicted by purely random models of species interactions.

Altenia et al. [73] revisited data of Neutel et al. to investigate effects of weighted

connectence of biomass flux rates on food-web stability. The study was motivated in

part by reports that species interactions in ecosystems are not evenly distributed in

space and time, but instead follow skewed probability patterns, such as observed in

Appendix Figure 1. A Benford FSD analysis of biomass flux rate data provides a

novel approach to investigate Altenia et al.'s finding that food-web stability has

increased over time at the individual SCH and HUL sample locations, a trend

suggested at combined sample locations (Appendix Table 1).

Appendix Table 1. Observed first significant digit counts of annual biomass for fifteen functional groups *vs* counts expected by the Benford probability distribution. Counts from combined SCH and HUL sample locations during early and late stages of food-web development. Raw data from Neutel et al., Supp. Table 1 [72].

EARLY SUCCESSION (DEVELOPMENT STAGES 1 & 2)

First Digits	1	2	3	4	5	6	7	8	9
Observed counts	42	29	26	16	12	9	13	7	5
NBL Expected Counts	47.9	28.0	19.9	15.4	12.6	10.6	9.2	8.1	7.3
Pearson Residuals	0.85	0.19	1.38	0.15	0.17	0.50	1.24	0.40	0.84
Observed Probability %	26.4	18.2	16.4	10.1	7.5	5.7	8.2	4.4	3.1
Expected Probability %	30.1	17.6	12.5	9.7	7.9	6.7	5.8	5.1	4.6

FSD sample size N = 159. Max Pearson Residual<1.96; Euclidean distance $d*n$ = 0.782 (p>0.10); Cohen-W = 0.184 (95% CI: 0.141-0.227); Kossovsky SSD = 38.4

LATE SUCCESSION (DEVELOPMENT STAGES 3&4)

First Digits	1	2	3	4	5	6	7	8	9
Observed counts	64	36	28	22	10	19	13	7	9
NBL Expected Counts	62.6	36.6	26.0	20.2	16.5	13.9	12.1	10.6	9.5
Pearson Residuals	0.18	0.10	0.39	0.41	1.59	1.36	0.27	1.12	0.17
Observed Probability %	30.8	17.3	13.5	10.6	4.8	9.1	6.2	3.4	4.3
Expected Probability %	30.1	17.6	12.5	9.7	7.9	6.7	5.8	5.1	4.6

FSD sample size N = 208. Max Pearson Residual<1.96; Euclidean distance $d*n$ = 0.664 (p>0.10); Cohen-W = 0.171 (95% CI: 0.127-0.216); Kossovsky SSD = 21.2

Early Developmental Stages (1&2)   **Late Developmental Stages (3&4)**

Appendix Figure 1. Comparison between FSD counts expected by Newcomb-Benford law and the probability distribution of observed counts for annual biomass of fifteen functional groups at combined SCH and HUL sample locations [72]. Graphs of data from Appendix Table 1.

Case Report #2: Dynamic equilibrium of diatom community colonization

Background: MacArthur and Wilson [74], their Table 6, provide diatom cell counts for a colonization experiment conducted by Ruth Patrick at Roxborough Spring, Pennsylvania. Over a span of two weeks, diatom species colonized glass slides representing two size 'islands' (12 & 25 mm^2). Replicate samples are reported as experiments 1&3 and 2&4. At experiment end, a diverse community of 54 diatom taxa and >600,000 diatom cell counts were recorded. MacArthur and Wilson cited the Patrick data in support of their equilibrium model of island biogeography, which predicts the ecological process of diatom colonization was progressing toward a threshold of stable dynamic equilibrium of species richness in response to stabilizing selection effects in rates of species immigration and extinction.

The diatom count data of Patrick meet sampling constraints required to test

correspondence with NBL predictions (Table 3). The set of data spans six orders of

magnitude with sufficient sample size to test conformity with the Benford distribution

when data from replicated experiments for same sized islands are combined over the

two week span of colonization. The experimental spring was described elsewhere by

Patrick [75] as perennial flowing and free from human disturbance.

Results and discussion: Diatom cell counts for *smaller islands* provide

insufficient analytic evidence to reject the tested (H_O) of similarity between observed

vs expected FSD probability distributions (Appendix Table 2 & Figure 2). No early

warning evidence of pending regime change is evident (max PR<1.96; Cohen-W<0.3;

d*n<1.22). Test results support an interpretation the diatom community on the smaller

islands attained a state of stable dynamic equilibrium. Supporting evidence for this

hypothesis comes from independent research at Roxborough Spring by Patrick [75]

where constancy of diatom taxa richness was recorded after a two week span of

colonization for an eight week experiment.

In contrast to results for smaller sized islands, analytic data for *larger islands*

reveal higher levels of FSD volatility (Appendix Table 2), with both Person Residual

(PR>1.96) and Cohen-W effect size (W = 0.316, 95% CI:0.250-0.381) within an early

warning range of regime change. Theory of island biogeography predicts the

increased volatility in response to effects of species interactions related to competition

and taxa extinctions, a hypothesis open to experimental investigation by extending the

colonization time for larger size islands to eight weeks. Results highlight the

importance of including measures of Pearson Residual volatility and magnitude of

Cohen effect size to complement statistical inference from calculation of Euclidean

distance to detect marginally stable systems trending toward future bidirectional

regime change (Appendix Table 2). Future research is suggested using protocols of

Patrick [75] to increase sample size and island replicates to monitor time-series trends

of FSD frequencies, as well as controlled experiments that manipulate data collection

constraints related to human bias and alteration of habitat conditions.

Appendix Table 2. Observed first digit cell counts *vs* Newcomb-Benford law

expected counts for diatom colonization experiments (MacArthur and Wilson

[74], their Table 6).

EXPERIMENTS 1&3 (12 mm^2 slides, small 'islands')
678
679   **Combined One & Two Week Exposure**

First Digits	1	2	3	4	5	6	7	8	9
Observed counts	32	15	9	11	7	12	6	7	6
NBL Expected Counts	31.6	18.5	13.1	10.2	8.3	7.0	6.1	5.4	4.8
Pearson Residuals	0.07	0.81	1.14	0.26	0.46	1.87	0.04	0.70	0.55
Observed Probability %	30.5	14.3	8.6	10.5	6.7	11.4	5.7	6.7	5.7
Expected Probability %	30.1	17.6	12.5	9.7	7.9	6.7	5.8	5.1	4.6

680
681   FSD sample size N = 105. Max Pearson Residual<1.96; Euclidean distance d*n =

682   0.759 (p>0.01); Cohen-W = 0.250 (95% CI: 0.182-0.317); Kossovsky SSD = 54.9

684   **EXPERIMENTS 2&4 (25 mm^2 slides, large 'islands')**

685   **Combined One & Two Week Exposure**

First Digits	1	2	3	4	5	6	7	8	9
Observed counts	48	26	27	19	5	7	4	4	8
NBL Expected Counts	44.6	26.1	18.5	14.3	11.7	9.9	8.6	7.6	6.8
Pearson Residuals	0.52	0.01	**1.98**	1.23	1.96	0.92	1.56	1.30	0.47
Observed Probability %	32.4	17.6	18.2	12.8	3.4	4.7	2.7	2.7	5.4
Expected Probability %	30.1	17.6	12.5	9.7	7.9	6.7	5.8	5.1	4.6

686
687   FSD sample size N = 148. Max Pearson Residual>1.96; Euclidean distance d*n =

688   1.147 (p>0.10); Cohen-W = 0.316 (95% CI: 0.250-0.381); Kossovsky SSD = 88.9.

Experiments 1&3 (small islands)   **Experiments 2&4 (large islands)**

Appendix Figure 2. Comparison between FSD counts expected by Newcomb-Benford law and the probability distribution of observed counts for diatom cells colonizing islands at small and large spatial scales [74]. Graphs of data from Appendix Table 2.

Case Study #3. Regional decline of stable amphibian populations.

Background: Published databases suitable for a Benford FSD analysis with random samples collected before and after human disturbance of natural habitats, that also meet required data collection constraints (Table 3), are not common in the ecological literature. One example known to the author is that of Highton [76] for visual encounters of woodland salamanders (genus *Plethodon*) that span five decades. From 1957-1999 Highton and colleagues recorded observation of individual salamanders from 38 species of the genus *Plethodon* from 127 sample locations in forested habitats of Eastern United States. Pre-1990 surveys revealed long-term constancy of salamander abundance for 205 *Plethodon* populations (1957-1988). Post-1990 surveys (1991-1999) revealed decline in many of previously stable

populations, with a statistically significant reduction in mean observations/site visit from 8.77 to 3.65 (Sign Test, $\chi^2 = 117$, p < 0.001). At 32 previously stable locations, *Plethodon* salamander species were extirpated. Although the causes for salamander declines were not investigated, logging was observed as a possible factor at multiple locations. Fungal infections, warming of forest temperatures from climate change, and changes in soil pH from acid rain are other suspected long-term stresses for woodland adapted salamanders.

The mean number of salamanders observed/site for Pre-1990 and Post 1990 are given in Table 8-1 of Highton [76]. The data span two and three orders of magnitude respectively. With the exception of the observed human habitat alteration for the Post-1990 surveys, the data meet sampling constraints required to test correspondence with NBL predictions (Table 3). Data were normalized to 1-6 site visits per observer, the maximum number reported for post-1990 surveys. This normalization process resulted in sample sizes N=96 (pre-1990) and N=173 (post-1990) to allow for statistical analysis.

Results and Discussion: Pre-1990 observations of salamander encounters provide insufficient weight of evidence to reject the tested (H_O) of similarity between observed *vs* expected FSD probability distributions (Appendix Table 3 & Figure 3). Pearson Residuals (max<1.96) and Cohen-W effect size magnitudes (max<0.2) agree with results from statistical inference (d*n=0.783, p>0.10) that the regional abundances of 38 *Plethodon* salamander species were at stable dynamic equilibrium prior to 1988 surveys. This interpretation from the Benford FSD analysis matches pre-1988 observations of Highton [76] of "*consistent patterns of salamander surface activity from year to year*" at the 127 sample locations.

In contrast with Pre-1990 observations, Post-1990 test results reject the tested

(H_O). Signals of the regime change are presented for all analytical test methods

(Pearson Residuals>1.96, Euclidean distance $d*n = 1.437$ ($p<0.05$), Cohen-W = 0.328

(95%CI 0.261-0.394). The Kossovsky SSD (119.4) adds support to an interpretation

that the Post-1990 data are not well described by the Benford distribution (Appendix

Table 3 & Figure 3). Results from the Benford FSD analysis concur with observation

of Highton that widespread populations of *Plethodon* salamander species in the

eastern United States from 1990 to 1999 were in various local modalities of unstable

dynamic equilibrium compared to pre-1988 observations.

Post-1999 surveys on *Plethodon* species within the geographic range sampled by

Highton by Caruso and Lips [77] and Caruso et al. [78] confirm the continuation of

salamander declines, in particular at sample locations dominated by larger-bodied

Plethodon species. Future investigation of measurement variables associated with

gene flow among geographically isolated populations and allocation of energy to

growth is suggested to investigate Benford FSD predictions of stability for those

remaining large-bodied *Plethodon* populations under greatest selective pressure for

extirpation.

759 **Appendix Table 3. Observed first digit counts of mean number of *Plethodon***

760 **salamanders per site visit *vs* Newcomb-Benford law expected counts for stable**

761 **(pre-1990) and declining (post-1990) populations from Eastern United States.**

762 **(Highton [76], his Table 8-1). Analyzed data normalized to 1-6 site visits.**

763
764 **Stable Pre-1990 Salamander Counts**
765

First Digits	1	2	3	4	5	6	7	8	9
Observed counts	35	15	9	11	7	5	4	6	4
NBL Expected Counts	28.9	16.9	12.0	9.3	7.6	6.4	5.6	4.9	4.4
Pearson Residuals	1.13	0.46	0.86	0.56	0.22	0.56	0.66	0.49	0.19
Observed Probability %	36.5	15.6	9.4	11.5	7.3	5.2	4.2	6.3	4.2
Expected Probability %	30.1	17.6	12.5	9.7	7.9	6.7	5.8	5.1	4.6

766
767
768 FSD sample size N = 96. Max Pearson Residuals<1.96; Euclidean distance d*n =

769 0.783 (p>0.10); Cohen-W = 0.195 (95% CI: 0.153- 0.237); Kossovsky SSD = 63.9.

771 **Declining Post-1990 Salamander Counts**

First Digits	1	2	3	4	5	6	7	8	9
Observed counts	65	32	21	10	21	4	11	5	4
NBL Expected Counts	52.1	30.5	21.6	16.8	13.7	11.6	10	8.8	7.9
Pearson Residuals	1.79	0.28	0.13	1.65	**1.97**	**2.23**	0.31	1.29	1.39
Observed Probability %	37.6	18.5	12.1	5.8	12.1	2.3	6.4	2.9	2.3
Expected Probability %	30.1	17.6	12.5	9.7	7.9	6.7	5.8	5.1	4.6

773
774
775 FSD sample size N = 173. Max Pearson Residual >1.96; Euclidean distance d*n =

776 1.437 (p>0.05); Cohen-W = 0.328 (95% CI: 0.261-0.394); Kossovsky SSD = 119.4

Stable Populations (Pre-1990)   **Declining Populations (Post-1990)**

Appendix Figure 3. Comparison between FSD counts expected by Newcomb-Benford law and the probability distribution of observed counts for mean observations/site visit of multiple *Plethodon* salamander populations from Eastern United States [76]. Graphs of data from Appendix Table 3.

Case Report 4: Regional fish communities with stable biotic integrity

Background: For the past six decades the State of Ohio Environmental Protection Agency (Ohio EPA) has conducted surveys of fish communities from streams and rivers to determine attainment of Federal Clean Water Act goals. Fish are sampled using standardized electrofishing protocols [79]. The overall condition and health of fish communities is evaluated using a modified Index of Biotic Integrity (IBI) [80], and agency developed Modified Index of Well-Being (MIwb). These indices are combined with information for benthic macroinvertebrate communities to assign Clean Water Act designated uses for protection of aquatic life. Stream segments designated Exceptional Warmwater Habitat (EWH) have unique species of fish with persistent levels of biotic integrity over time, and are associated with watersheds that

have minimal background levels of anthropogenic disturbance. Chemical and

biological survey results are available from Ohio EPA online technical documents:

[https://epa.ohio.gov/divisions-and-offices/surface-water/reports-data/biological-

and-water-quality-reports]**.**

Five EWH stream segments from various ecoregions of Ohio known to the

author were selected to investigate correspondence of biomass measurements

(reported as relative weight) with FSD predictions of the Benford distribution. Details

of sample locations are as follows (sample date, river mile, # fish spp., # individual

fish captured, length of stream sampled): Grand River (2004, 8.5, 26, 301, 0.42 km);

Chagrin River (2003, 36.4, 21, 361, 0.2 km); Big Darby Creek (2014, 30.2, 39, 1012,

0.5 km); Scioto River (2011, 9.1, 33, 519, 0.5 km); Mohican River (2007, 6.5, 29, 360,

0.5 km).

Fish biomass measurements meet data collection constraints required to test NBL

predictions. Fish were randomly sampled from all available habitats (riffles, run,

pools) and the set of data spans three orders of magnitude, with total sample size

sufficient for statistical inference when data from five EWH stream distributions are

combined. Because raw data are scattered in multiple online technical support

documents, biomass measures for 141 fish species from the five stream locations are

given in Appendix Table 4, sorted by Benford FSD categories. This format of data

presentation allows for calculation of biomass mean and variance for each of the nine

FSD categories. As a point of reference, the 80.82 value in Appendix Table 4

represents the relative weight (kg/km) of 156 Golden Redhorse (*Moxostoma*

erythrurum) collected from the Big Darby Creek, which represents ~27.5% of total

fish community biomass at that stream location.

825 *Results and discussion:* FSD biomass counts for 141 fish species collected from

826 five EWH stream segments provide insufficient analytic weight of evidence to reject

827 the tested (H_O) of similarity between observed *vs* expected FSD probability

828 distributions (Appendix Table 4 & Figure 4). Maximum Person Residuals (<1.96) and

829 Cohen-W effect size magnitude (W=0.182) agree with results from statistical

830 inference (d*n=0.759) that distributions of EWH fish communities over a wide

831 geographic area of Ohio persist within a threshold range of stable dynamic

832 equilibrium.

833 Odum [81] defined an ecological system as a set of interacting components

834 forming a unified whole at the population, community, and ecosystem levels of

835 organization. The geographically isolated distributions of EWH fish communities

836 from this case report, combined with observations of widespread stable populations of

837 salamanders in the eastern United States (case report #3), are consistent with the

838 meta-community and meta-population concepts of organization proposed by Leibold

839 et al. [82]. Benford FSD predictions of stability of meta-distributions are maximized

840 at sample locations where constraint violations related to collection of data are

841 minimized.

842 Taylor's power law ($V = aM^b$), where V = variance, M = mean, a = constant, and

843 b=correlation slope, is an empirical observation relating the variance of measurements

844 on organisms to mean unit area of sampling effort [62]. Fish species community

845 biomass measurements from EWH stream segments reveal high coefficient of

846 determination ($Y = 3.1X^{2.1}$, $R^2 = 0.91$) between Taylor's power law and the mean (X)

847 and variance (Y) of data for the nine Benford FSD categories (Appendix Table 4).

848 This correlation leads to a hypothesis that the FSD sequence of the

849 Newcomb-Benford law provides an underlying mathematical explanation of the

correlation slope metric of Taylor's Law for different levels of ecological

organization and spatial scale [62].

Appendix Table 4. Biomass (kg/km) of 141 species of fish collected from five streams in Ohio designated Exceptional Warmwater Habitat (EWH) for protection of aquatic life. Measurements sorted by Benford first digit categories.

First Digits	1	1	2	3	4	5	6	7	8	9
	0.01	1.06	0.02	0.03	0.04	0.05	0.06	0.07	0.08	0.09
	0.01	1.08	0.02	0.03	0.04	0.05	0.06	0.73	0.08	0.09
	0.01	1.10	0.02	0.03	0.04	0.05	0.60	0.76	0.87	0.90
	0.01	1.20	0.02	0.03	0.04	0.56	0.63	70.18	80.82	0.99
	0.01	1.20	0.02	0.03	0.04	0.56	0.64			9.0
	0.01	1.52	0.02	0.03	0.04	5.10	0.64			9.31
	0.01	1.68	0.02	0.03	0.40	5.20	0.69			
	0.01	1.71	0.02	0.32	0.46	5.51	6.30			
	0.01	1.98	0.02	0.34	0.48	5.67	6.70			
	0.01	10.52	0.02	3.20	4.0	53.85	62.25			
	0.01	11.03	0.20	3.2	4.01		67.5			
	0.01	12.0	0.20	3.6	4.8					
	0.11	14.11	0.21	3.68	44.90					
	0.11	14.72	0.29	3.70	48.17					
	0.11	15.8	2.0	32.99						
	0.12	16.94	2.0	33.60						
	0.14	18.22	2.23	38.03						
	0.14	19.50	2.25							
	0.15	114.5	2.41							
	0.16		2.49							
	0.16		2.80							
	0.16		2.80							
	0.17		2.85							
	0.17		21.01							
	0.18		21.30							
	0.19		21.35							

858 **Appendix Table 4, continued.**

First Digits	1	2	3	4	5	6	7	8	9
Observed Counts	47.0	27.0	18.0	14.0	10.0	11.0	4.0	4.0	6.0
NBL Expected Counts	42.4	24.8	17.6	13.7	11.1	9.4	8.2	7.2	6.5
Pearson Residuals	0.71	0.44	0.09	0.08	0.33	0.52	1.47	1.19	0.20
Observed % Probability	33.3	19.1	12.8	9.9	7.1	7.8	2.8	2.8	4.3
Expected % Probability	30.1	17.6	12.5	9.7	7.9	6.7	5.8	5.1	4.6
Digit Mean (X)	5.94	3.28	8.74	7.68	7.67	13.27	17.93	20.46	3.40
Digit Variance (Y)	301	43	208	274	270	658	1213	1619	20

FSD sample size N = 141. Max Pearson Residual<1.96; Euclidean distance d*n =

0.638 (p>0.10); Cohen-W = 0.182 (95% CI: 0.134- 0.230); Kossovsky SSD = 28.9.

Taylor's Power Law: $Y = 3.1 X^{2.1}$, $R^2 = 0.91$.

Appendix Figure 4. Comparison between FSD biomass counts expected by

Newcomb-Benford law and the probability distribution of observed counts for

141 species of fish collected from stream segments with persistent exceptional

biotic integrity. Graph of data from Appendix Table 4.