Synthesis reveals biotic homogenisation and differentiation are both common

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Earth's biodiversity continues to change rapidly through the Anthropocene[1], including widespread reordering of species in space[2,3] and time[4,5]. A common expectation of this reordering is that the species composition of sites is becoming increasingly similar across space, known as biotic homogenization, due to anthropogenic pressures and invasive species[6,7]. While many have argued that homogenisationis a common phenomenon (e.g., [6–10]), it is equally plausible that communities can become more different through time, known as differentiation, including through human impacts[11,12]. Here, we used a novel adaptation of Whittaker's (1960)[13] spatial-scale explicit diversity partition to assess the prevalence of biotic homogenisation and differentiation, and associated changes in species richness at smaller and larger spatial scales. We applied this approach to a compilation of species assemblages from 205 metacommunities that were surveyed for 10-64 years, and 54 'checklists' that spanned 50-500+ years. Scale-dependent changes of species richness were highly heterogeneous, with approximately equal evidence for homogenisation(i.e., lower β-diversity) and differentiation (i.e., higher β-diversity)

through time across all regions, taxa and data types. Homogenisation was most often due to increased numbers of widespread species, which tended to increase both local and regional richness through time. These results emphasise that an explicit consideration of spatial scale is needed to fully understand biodiversity change in the Anthropocene.

Humans are fundamentally altering the Earth's climate, water and nutrient cycles, and are appropriating over 50% of the Earth's terrestrial net primary production[14]. With this increasing human footprint, consequent changes in Earth's biodiversity are inevitable. There is substantial evidence that humans are accelerating the global extinction rate[15]. However, diversity declines at local scales have not been prevalent, with several syntheses indicating little directional trends amidst substantial variability[4,5,16,17], though not without controversy[18]. It has also been proposed that spatially distinct locations are becoming more similar to one another in species composition through time, commonly called biotic homogenization[6,7]. Homogenisationcan occur, for example, when widespread species replace rare species through mechanisms such as introduction of invasive species, loss of rare species, and homogenisation of landscapes via anthropogenic pressures (e.g., agricultural practices or urbanisation). While homogenisation is frequently reported in empirical studies[8–10,19], it is not clear how common it is relative to other types of scale-explicit biodiversity change[12]. For example, the opposite of biotic homogenization, known as biotic differentiation, can occur when landscapes are fragmented or otherwise made more heterogeneous via human activities, or when exotic species are introduced but do not become widespread[11,12,20]. A comprehensive empirical assessment of the relative frequency of homogenisation, differentiation, or no directional change is lacking.

Change in spatial differentiation of community composition is logically equivalent to differential rates of change in diversity at two spatial scales[11,12]. This can be clearly seen with Whittaker's (1960)[13] diversity partition where the diversity of a single site is α-diversity and the sum of the diversity of several local sites (i.e., a region) is γ-diversity. Variation in local community composition is referred to as β-diversity, and given by: $\beta = \gamma / \bar{\alpha}$ (where $\bar{\alpha}$ is the average local diversity across sites in a region). Whenever rates of change in α- and γ-diversity are not equal through time, there will be some change in β-diversity. Here, we adapt Whittaker's framework such that change in β-diversity emerges naturally when changes at α- and γ-scales are considered jointly (Figure 1). Moreover, these changes in β-diversity can be mathematically linked to changes in the number of sites species occupy. Average occupancy,

or the fraction of sites within a region occupied by species i (o_i), is related to Whittaker's

formula by $\beta = \gamma/\bar{\alpha} = \gamma/(\Sigma o_i) = \gamma/(\gamma\bar{o}) = 1/\bar{o}$ [21]. This allows us to connect α- and γ-

diversity change with changes in β-diversity and the underlying processes of colonisation and

extinction that drive changes in average occupancy.

Six qualitatively distinct scenarios emerge in the intersecting space of changing α- and γ-

diversity (Figure 1). The 1:1 line (i.e., $\Delta\gamma = \Delta\alpha$) represents equal log-proportional changes at

both scales (i.e. $\log[\gamma_{t2}/\gamma_{t1}] = \log[\alpha_{t2}/\alpha_{t1}]$), and delineates the boundary between

homogenisation and differentiation (Figure 1). When $\Delta\alpha > \Delta\gamma$ (i.e., below the 1:1 line),

homogenisation occurs. One such scenario involves α-diversity increasing proportionately

more than γ-diversity increases via increase of high occupancy species (Figure 1i). Increases

in high occupancy species can accompany anthropogenic changes that favour widespread,

generalist and/or non-native species. Another scenario involves widespread species replacing

low occupancy species causing average occupancy increases and thus homogenisation which

would be associated with α-diversity increases, but γ-diversity decreases (Figure 1ii).

However, homogenisation can also accompany simultaneously decreasing α- and γ-diversity,

as would occur when low occupancy species, such as endemic species or those restricted to

only a few sites, go regionally extinct, for example, due to habitat loss or degradation (Figure

1iii). These distinct scenarios all describe decreasing β-diversity and biotic homogenisation,

but the outcomes for α- and γ-diversity, and the implications for policy and conservation, are

very different. Parallel distinct scenarios of differentiation are also possible. If increasing

habitat heterogeneity results in fewer widespread or high occupancy species, increased β-

diversity would be associated with lower α- and γ-diversity (Figure 1iv). Increased habitat

heterogeneity could also lead to increases in γ-diversity, accompanied by either α-diversity

declines (Figure 1v) or increases (Figure 1vi).

Figure 1: **Mechanisms that underpin spatial changes in spatial differentiation in composition can be understood by examining the relationship between changes in regional- and local-scale species richness through time.** When richness changes at regional ($\Delta\gamma$) and local ($\Delta\alpha$) scales are calculated as proportional changes, assemblages below the dashed 1:1 line, i.e., $\Delta\gamma < \Delta\alpha$, are being homogenised, and β-diversity is decreasing. Conversely, assemblages above the diagonal dashed 1:1 line, i.e., $\Delta\gamma > \Delta\alpha$, are differentiating, and β-diversity is increasing. These changes can be linked to changes in species occupancy, for regions where **(i)** $\Delta\gamma < \Delta\alpha$, and $\Delta\gamma > 0$, $\Delta\alpha > 0$, the number of species with high occupancy (i.e., that occupy the majority of sites in the region) is increasing; **(ii)** $\Delta\gamma < \Delta\alpha$, $\Delta\gamma < 0$ and $\Delta\alpha > 0$, average occupancy is increasing, e.g., due to species with low occupancy being replaced by those with high occupancy; **(iii)** $\Delta\gamma < \Delta\alpha$ and $\Delta\gamma$, $\Delta\alpha < 0$, the number of species with low occupancy is decreasing; **(iv)** $\Delta\gamma > \Delta\alpha$ and $\Delta\gamma$, $\Delta\alpha < 0$, the number of species with high occupancy is decreasing; **(v)** $\Delta\gamma > \Delta\alpha$, $\Delta\gamma > 0$ and $\Delta\alpha < 0$, species with low occupancy (i.e., occupy few sites in the region) are replacing those with high occupancy; **(vi)** $\Delta\gamma > \Delta\alpha$ and $\Delta\gamma$, $\Delta\alpha > 0$, the number of species with low occupancy (i.e., occupy relatively few sites in the region) is increasing.

Here, we use the framework of Figure 1 to: i) estimate the overall change in β-diversity (above or below the 1:1 line) observed in empirical studies, and ii) classify changes in β-

diversity into six distinct scenarios. To do so, we required data that recorded species richness at a minimum of two time points across multiple locations. First, we compiled a total of 205 studies that had data on species abundances (or occurrences) through time from at least four locations and at least ten years between the first and last samples. One hundred and thirty-eight of these studies came from already published compiled databases (e.g., [22–24]) and 67 datasets had a similar structure (e.g., species abundances from samples through time), but were not included in these previous compilations and are compiled for the first time here. Second, we used 54 studies that were based on 'checklists' where species occurrence was recorded for sites at historical and contemporary time periods. We chose to include species checklist data here despite their coarse nature, as they have made key contributions to our understanding of long-term trends in introductions and extinctions[25,26], as well biotic homogenisation[27,28]. We analysed whether these distinct data types affected the results.

In total, our analysis on 259 datasets with a total of 16,359 locations is the largest compilation of data sources, ecosystem types, and taxon groups used to examine the question of changes in β-diversity through time to date (Extended Data Figure 1). We estimated temporal changes in species richness for every dataset at the smaller, α-scale, where a sample was taken, and a larger γ-scale, where richness was estimated as the sum of species in all of the local samples. The grain of the α-scale and extent of γ-scales varied among datasets; ranging, for example, from quadrat samples of plant communities collected over small spatial extents ($< 1\text{km}^2$) to species checklists of birds on islands distributed across several oceans. For both the α- and γ-scales, we quantified change through time as the log-ratio of species richness in the most recent time point over the species richness in the initial sample for every location (for α-diversity) or the sum of all locations (for γ-diversity) within each dataset, and then divided that by the number of years between the two samples to get a standardised annual rate of change independent of time series length. We then fit a multilevel model for each scale that estimated the average change occurring (i.e., a non-varying intercept only), which also included a random term for variation between datasets (i.e., varying intercepts). For results presented in the main text, species richness in the initial and final time point were calculated using a single year. However, we repeated all analyses using multiple years to estimate the average richness for two periods (data permitting) so as to verify that starting or ending periods were not having undue influence, and found that results were qualitatively consistent (see Materials and Methods and Extended Data Figure 2).

Overall across datasets, we found many instances for each of the six scenarios identified in

Figure 1 (Figure 2a, 2b). Averaged over the entire data set, β-diversity change showed a weak

trend towards homogenisation but with the 90% credible interval for this trend clearly

overlapping an average β-diversity change of zero (Figure 2c). The weak trend (-0.001) is the

equivalent of the loss per decade of 1 out of 100 entirely distinct (i.e., no shared species)

communities[29]. Moreover, few individual datasets showed strong evidence for changes in β-

diversity in either direction (Figure 2f). Qualitatively, when we count the number of empirical

estimates that fall into each outcome, the commonly hypothesised homogenisation scenario[6,7]

of high occupancy species causing extinction and replacement of low occupancy species and

a concomitant decline in regional diversity is among the least frequent pathways (sector ii in

Figure 2b). Where differentiation is occurring (i.e., increases of β-diversity through time), the

prevalence of increases in both α- and γ-diversity (sector vi) were approximately balanced by

declines at both scales (sector iv in Figure 2b). A common scenario is increases in diversity at

both the local and regional scale (sectors i and vi), with small changes in average occupancy

(i.e., increasing or decreasing) tipping a given system towards homogenisation or

differentiation.

Figure 2: Patterns of homogenisation and differentiation are approximately balanced at the global scale. (a) Empirical estimates of γ-scale changes as a function of α-scale changes, both axes show log-ratios standardised by the number of years between the estimates; the black point shows the γ- and α-scale intercepts and the 90% credible interval from multi-level models fit separately to these data at each scale; colour represents categories of change from Figure 1, shape represents sample type (circle = resurvey, triangles = checklist). Dashed lines show x = 0, y = 0, and x = y. (b) Count of the number of regions in each of the different qualitative outcomes depicted on Figure 1. (c) Kernel density plot of change in β-diversity per year calculated as the distance from 1:1 line (left = homogenisation, right = differentiation) of 1000 draws of α- and γ-scale intercept posterior distributions; black point shows median, bar represents 50% (thick) and 90% (thin) credible intervals. Estimates of change for each region at the (d) α-, (e) γ-, and (f) β-scales; each point represents a single region, with the bar showing the 90% credible interval; regions are in the same order on panels d-f, arranged by the magnitude of the α-scale estimate.

We detected the strongest signature of biotic homogenisation in the freshwater realm, and this

was primarily driven by increased average occupancy (i.e., increases in α-diversity and a

weaker decrease in γ-diversity; Figure 3a). This is not unexpected since freshwater systems

are often regarded as among the most low-connectivity, dispersal-limited ecological

systems[30,31]. Indeed, human introductions of non-native species in lakes and rivers are among

the most well-known examples of biotic homogenisation (e.g., [32–35]). Homogenisation and

differentiation were approximately balanced among terrestrial assemblages (Figure 3a), and

associated with weak gains in species richness that did not differ from zero at α- and γ-scales.

In contrast, we found that there was a slight tendency towards differentiation in the marine

realm (Figure 3a), which was associated with gains in γ-diversity. One possible reason for

this is that marine assemblages are typically highly fluid and dynamic, with high connectivity

across locations, which might allow species favoured by global change conditions to colonise

new regions more readily than in the other realms. Across latitudinal bands (Figure 3b),

tropical, subtropical and temperate latitudes showed a trend towards homogenisation driven

by relatively large gains in α-diversity, while polar regions showed a trend to differentiation

driven by gains in γ-diversity with comparatively small changes in α-diversity, though

confidence intervals overlapped no change and sample sizes were small. Data type (resurvey

vs checklist), and spatial and temporal scale did not show substantial effects on our results

(Figure 3c, Extended Data Figure 2), although larger spatial and temporal scales, including

checklist data, were slightly more likely to show homogenisation. Lastly, across taxa,

mammals tended towards differentiation driven by declines in α-diversity that were

proportionately larger than declines in γ-diversity, while fish tended towards homogenisation

due to small γ-scale richness increases of high occupancy species (Figure 3d). As 55/80 fish

studies were from freshwater, this parallels our findings for the freshwater realm. But again,

β-diversity effect sizes by taxonomic group, latitudinal band, and realm were small and

confidence intervals all overlapped with no change with the exception of the freshwater realm

(Figure 3).

Figure 3: α- and γ-scale species richness changes, and patterns of homogenisation and differentiation varied among sample types, realms, latitudinal bands and taxon groups.
γ-scale estimates of the log-ratio as a function of α-scale estimates (left) and distance from 1:1 line (right) for comparisons between (**a**) realms, (**b**) latitudinal bands, (**c**) sample types: checklists and resurveys; and (**d**) taxon groups. Distances from the 1:1 line were calculated using 1000 draws from posterior distributions, where each x,y coordinate was a single draw

from the population-level (fixed effect) parameters of α- and γ-scale models, respectively.

Points and whiskers show median, 50 and 90% credible intervals.

Our results suggest that β-diversity trends show considerable variability across locations,

taxa, and time, with differentiation slightly less likely than homogenisation. This matches

findings on local diversity[4,5,16], and population trends[36], where variation in the direction of

change means that the strength of overall net trends up or down are weak, and most often

statistically indistinguishable from zero. This should not be surprising. While humans are

having many impacts that could lead to homogenisation including transporting species and

recreating urban or high intensity agricultural landscapes repeatedly, many other impacts

could lead to differentiation, including: substantial fragmentation of the landscape, the

creation of strong spatial gradients of human impact intensity, applying spatially varying

resource management practices and land use regulations, and causing climate change, which

induces species to shift at different rates, all leading to spatial heterogeneity[11,37]. In this

context, we stress that the lack of prevalence of homogenisation should not in any way be

taken to indicate that humans are not having a large impact on biodiversity. Changes in β-

diversity either negatively, as in homogenisation, or positively, as in differentiation, are likely

often a result of humans modifying nature.

Most published tests of the frequency of homogenisation use pairwise comparisons of sites

and then averages across all possible pairs, and thus only addresses the α-scale, but not

changes at the γ-scale. For example, previous tests of biotic homogenisation concluding that

homogenisation was driven by loss of rare species used pairwise averages that are only able to

capture single site extirpations, not landscape or regional scale extinctions, and thus can

exaggerate the perception of loss. We found that the hypothesised mechanism of widespread

species, like invasive species, replacing locally rare species (orange sector ii in Figure 1) is

among the rarest pathways of spatial homogenisation. Similarly, efforts to use β-diversity to

link changes at local scales (e.g., [5]) to regional or global scale changes (e.g., [38]) also cannot be

based on pairwise metrics, which have no representation of the larger scale. Although

Whittaker's β-diversity measures and pairwise β-diversity measures (e.g., 1-Jaccard's index)

are equivalent for two sites, the two metrics can give very different answers when more than

two sites are examined[39,40]. Thus, it is essential to directly estimate changes at each of smaller

and larger scales, rather than relying on pairwise similarity metrics that can give misleading

answers to how γ-diversity might change through time. For all of these reasons, we believe

that using a perspective based on the Whittaker diversity partition, combined with the new

classification we propose, provides a more accurate and constructive way of measuring β-

diversity to inform policy.

In summary, we propose moving beyond a belief in the predominance of homogenisation, and

instead working to understand the variability in how spatial β-diversity changes through time,

embracing the multiscale nature of biodiversity change. Specifically, our conceptual

classification emphasises that different long-term studies are experiencing fundamentally

different types of temporal change in spatial β-diversity (via changes in the number of species

with relatively high or low occupancy). We found that homogenisation was most often

characterised by gains in the number of widespread species. This suggests that there is still

time to protect many rare and narrow-ranged species before they ultimately contribute to

homogenisation, and that efforts to slow the spread of widespread (possibly non-native)

species are vital for preventing biotic homogenisation. Furthermore, by simultaneously

considering change in α-, β- and γ-diversity through time, we show how conservation

practice can embrace a multiscale approach. It is increasingly recognised that many species

require protection across multiple sites or at landscape (or larger) spatial scales for effective

conservation[41,42]. The framework introduced here can be used to detect the occupancy

changes that underpin β-diversity change, and provides links back to changes in species

richness across spatial scales.

Extended Data Figure 1: Counts of locations in hexagonal grid cells (cell area = 69,968

km2). The data include 259 regions (i.e., γ-scale) and 16,359 locations (i.e., α-scale).

Extended Data Figure 2: Results were qualitatively consistent when multiple years were used to determine richness in the two periods. Parameter estimates of models fit to changes calculated using first and last year only (as per main text results) as a function of those estimated when richness was averaged over multiple years in two periods. (a) Intercept estimates for changes at each scale, (b) Realm estimates for each scale, (c) Latitudinal band estimates for each scale, (d) Sample type estimates for each scale, and (e) Taxon group estimates for each scale. Diagonal dashed line is 1:1 line in all panels; points show median of the posterior distribution, and whiskers show 90% credible interval.

Extended Data Figure 3: **Changes in β-diversity as a function temporal and spatial**

scales. (a) Temporal duration and spatial extent for all datasets in our analysis; and, empirical

estimates of (b) α-scale and (c) γ-scale richness change as a function of temporal duration,

and (d) α-scale and (e) γ-scale richness change as a function of spatial extent. Points are

coloured according to whether the empirical estimates indicated differentiation or

homogenisation (i.e., above or below the 1:1 line in our conceptual figure 1); shapes denote

sample type (checklist or resurvey). Dashed vertical lines on (a) show the land area of

Germany and continental USA for reference. Two datasets where empirical estimates of α-

and γ-scale richness change were equal to zero (and hence neither homogenising or

differentiating) are not shown.

Methods

Data compilation: Our conceptual framework requires estimates of species richness changes at two scales. We refer to them as local (α) and regional (γ), the exact definition of which varies among data sources. To make our data search and synthesis as comprehensive as possible, we searched broadly for data that met these criteria, where regions had at least four plots or locations, and where richness changes were estimated over a period of at least ten years. We started by identifying 80 relevant datasets within the BioTIME database[22] that monitored patterns of species abundances within assemblages. To this, we added: (1) similar assemblage-level time series of studies not (yet) included in BioTIME (e.g., [23,24]); (2) data from studies using 'resurveys', where sites associated with a historical dataset were revisited and re-surveyed using similar methodology in more recent times; (3) data from 'checklist' studies where species known to be present in a given locality (and region) at a 'historical' point in time were indicated together with species present in that locality at a later point in time (minus those that went extinct from a site plus those that newly colonised that site); and, (4) data from studies that reported changes in species richness at two spatial scales, but for which the underlying raw data were not available. Because of the relatively specific data requirements, literature searches were conducted in an ad-hoc fashion, rather than using a formal literature search. In all, we compiled a total of 259 regions and a total of 16,359 locations that met our criteria (Extended Data Figure 1); 205 regions documented repeated samples of species assemblages through time; 54 regions were compiled from checklist studies.

Exploratory variables: In addition to differentiating sample types as either a resurvey or a checklist, we retained metadata to use in subsequent analyses that included:

- realm (freshwater, marine and terrestrial),
- geographic coordinates of all locations,
- spatial extent (km^2). Extent was calculated in two ways: most often, as the area of a convex hull (or bounding box) around all locations within a region, or for some checklist data as the sum of the area of each location within a region (e.g., sum of island areas for birds on islands distributed across the Pacific, Indian, Atlantic Oceans and the Caribbean Sea).
- the following taxon groups: mammals, herpetofauna (reptiles, amphibians), plants, birds, fish, invertebrates. For studies in BioTIME labelled as having multiple taxa, we identified the dominant taxon group (using the same groups listed above).

Data standardisation: To quantify changes in *β*-diversity that emerged from combined changes occurring at the local- and regional-scale, we required that the starting and end years for all locations within a given region were the same. This ensured that change estimated across the different locations within a region covered the same period of time, and meant that regional changes estimated by aggregating all species across all locations within regions also covered the same time period. Additionally, to ensure that our analyses did not quantify changes in species richness due to variation in sampling effort, we needed to standardise sampling effort (e.g., the number of plots or transects) across all locations for each time point within regions. The heterogeneous nature of the data that we compiled meant that we needed slightly different procedures to identify combinations of locations and years for different data sources. For clarity, we delineate broad categories of data structures, and describe separately how locations and years were selected and sample-effort standardised for the different structures.

Checklist data: Checklist data typically consisted of species lists for locations within regions, compiled for two time periods, historical and contemporary. These lists were compiled either from samples and/or observations collected during the two periods, or more frequently, by counting native species only to determine the richness of the historical period, with the contemporary species richness calculated as the sum of native and introduced species (minus any species that went extinct). For our analyses, we selected regions with at least four locations, removed locations that documented species lists for only one period, and finally, ensured that all locations within each region had the same year for both the historical and contemporary species lists.

Resurvey data: We distinguish three different data structures that we refer collectively to as resurvey data:

(i) data that document repeated samples of assemblages, e.g., BioTIME [22], RivFishTime[23], and InsectChange[24], and similar data that we compiled for this study. We first filtered data to ensure that samples from all locations within regions had a temporal duration of at least ten years, and that at least four locations were sampled per year. Locations within regions were identified using geographic coordinates in the data, although we also regions with only one geographic coordinate where discrete, unique samples could be identified, e.g., plots within a site.

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367	After applying these filters, the number of locations sampled per year often
368	varied considerably within regions, and we sought to identify locations, as well
369	as start and end years that balanced a trade-off between the number of
370	locations and the duration of the sampling period for each region. To do this
371	we first identified all year-pairs -- combinations of start and end year with at
372	least ten years separating them -- for all locations within a given region. We
373	then determined different thresholds for what proportion of the total number of
374	locations we wanted to retain, using a combination of the total number of
375	locations in a region, and visual inspection of locations sampled in each year.
376	For example, for resurvey data newly collated for this study, we selected
377	starting and end years where the proportion of the maximum locations was at
378	least 90% for regions with fewer than 20 locations, 50% for regions with more
379	than 20 locations, and 25% for the NERC Countryside survey data[43–47], which
380	had between 60 and 300 locations across the UK (and where the lower
381	threshold meant that the duration of the region increased by more than ten
382	years). For regions in the BioTIME and RivFishTime databases, we identified
383	year-pairs with at least 75% and 90% of the maximum number of locations,
384	respectively. Multiple year-pairs often remained following this, and we
385	selected the pair of years with the longest duration, and finally, broke any
386	remaining ties by selecting the pair of years with the most locations. For other
387	data, specifically mosquito data sourced from Vectorbase
388	(https://vectorbase.org/vectorbase/app), this process of selecting locations and
389	the start and end years for each region was done visually.
390	
391	Next, we ensured that sampling effort was consistent across all years and
392	locations within regions, using sample-based rarefaction[48] where required to
393	standardise effort. Note that for many data (e.g., InsectChange and other
394	invertebrate data) where sampling took place across multiple months within
395	years, we used sample-based rarefaction to resample equal numbers of samples
396	across the same months for all locations within a region, which were then
397	compiled to provide one sample per year for each location. Additionally, for
398	data collected using multiple sampling methodologies (e.g., mosquitoes
399	sampled using different attractants, or freshwater fishes collected with different

400		techniques), we identified the methodology that ensured the maximum number
401		of time series, and standardised sampling effort using data collected with one
402		methodology only.
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404	(ii)	We collated data from studies where sites associated with a historical dataset
405		were revisited and re-surveyed using the same methodology in more recent
406		times, sometimes referred to as "legacy" studies (e.g., [49]). Again, we required
407		each region to have at least four locations and ten years or more between the
408		historical and contemporary samples.
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410	(iii)	Finally, we collated studies that estimated species richness changes at two
411		scales, where there were at least four sites at the smaller scale and ten years
412		between the first and last sample. For these studies raw data were not available
413		($n = 15$), and we extracted an estimate of the average local richness at two time
414		points, and a single value for regional richness at two time points.
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Estimating richness and its change: For the majority of the data, we calculated species

richness from the effort-standardised locations and years as the number of distinct species,

though higher classifications, such as genera, were sometimes used where studies only

classified organisms to genus. We calculated species richness for each location within each

region for every available year to document changes in local-scale species richness. Regional-

scale richness was calculated as the number of species in all sites combined for each region

and each year. However, because this method of calculating regional richness yields a single

number for each region at each time point, which limited our ability to fit statistical models to

these data, we additionally calculated two types of resamples of regional richness: jackknife

or bootstrap. Jackknife resamples were calculated by systematically leaving each location out

of the regional richness calculation once, and we retained all ($n_{locations}$) resamples for our

regional scale analyses. For some datasets where effort-standardisation was more complex

and required the use of sample-based rarefaction, we used 200 bootstrap resamples (i.e.,

richness was estimated using all locations, not $n_{locations}-1$); to prevent these resampled data

dominating the data to which models were fit, we subsampled the bootstrap resamples down

to the same size as a jackknife would have been (i.e., we used a random subset of the

bootstrap resamples equal to $n_{locations}$ for the given dataset).

Many data sources, e.g., the fifty-four regions documented by checklist data, had only two

years of data available. So, to maximise the number of regions in our analysis, we calculated

richness change using the log-ratio of species richness in the most recent time point and

species richness in the initial sample, divided by the number of years between the two

samples (i.e., $log\left(\frac{S_{t2}}{S_{t1}}\right) . t^{-1}$, where S_{t2} is species richness in the most recent sample (t2 =

year), S_{t1} is species richness in the first sample (t1 = year), and $t = t2 - t1 + 1$ is the number of

years between the samples. This was done separately for each location in each region. These

same data were aggregated and used to calculate concomitant changes in regional diversity

through time, quantified as the log-ratio of jackknife resamples of species richness at the

regional scale (i.e., the species richness of (n-1) locations within a region) in the most recent

sample and jackknife resamples of species richness in the initial sample, divided by the

number of years between the two samples.

Statistical models: To estimate local- and regional-scale richness changes, we fit multilevel

(also called mixed effects or hierarchical) models to data from each scale separately. These

models took the form:

$$ES_{ij} \sim N(\mu, \sigma),$$

$$\mu = \alpha + \alpha_i,$$

where ES_{ij} is assumed to have a Gaussian error distribution and is either the jth local-scale or

jackknife regional-scale estimate of species richness change in region i, α is the overall

intercept and average rate of change estimated for each scale, and α_i is the departure from the

overall intercept for each region (i.e., the varying intercept for regions). Models were fit using

Bayesian methods and we assumed the following, weakly regularising priors:

$$\alpha \sim N(0,1),$$

$$\alpha_i \sim N(0,1),$$

$$\sigma \sim N(0,1).$$

In addition to these models to estimate the overall rate of change at the local- and regional-

scales, we also used models of a similar structure to examine variation between the different

sample types (i.e., resurveys and checklists), realms (freshwater, marine and terrestrial),

latitudinal bands (polar: |latitude| > 60º, temperate: 35º < |latitude| < 60º, subtropical: 23.5º <

|latitude| < 35º, tropical: |latitude| < 23.5º), and taxon groups (birds, fish, herpetofauna,

invertebrates, mammals, multiple taxa, and plants). These models replaced the single overall

intercept estimated in the initial model with a parameter for each category. All models were

fit using the Hamiltonian Monte Carlo (HMC) sampler Stan[50], and coded using the brms

package[51]. Models were fit with four chains and 20,000 iterations, with 10,000 used a

warmup and further thinned by 10, resulting in 4000 samples of the posterior distribution.

Visual inspection of the HMC chains showed excellent convergence.

To quantify changes in β-diversity ($\Delta\beta$) using these models, we combined overall estimates

of changes at the local- ($\Delta\alpha$) and regional-scale ($\Delta\gamma$) as x- and y-coordinates, respectively,

and calculated the distance of these points from the 1:1 line. Specifically, 1000 draws from

the posterior distribution of the local-scale estimate were designated as the x-coordinate, and

combined with 1000 draws from the corresponding regional-scale estimate as the y-

coordinate, and the distance from the 1:1 line calculated. Accordingly, changes in β-diversity

are in units of effective numbers of communities[29].

To visualise counts of the different scenarios of change, we counted the empirical effect sizes

(i.e., the log-ratio standardised by duration [ES] defined above) that fell into each category;

empirical estimates that fell on the border between different scenarios (e.g., when $\Delta\gamma = 0$), the

count was divided between the different scenarios (i.e., we added fractions to counts when

estimates fell on the border between scenarios). To visualise scale-dependence of our

estimates of local- and regional-scale richness changes, we plotted the empirical effect sizes

as a function of spatial and temporal extent (Extended Data Figure 2). Note that to show

local-scale estimates as a single point on these figures, we used the mean of the local scale *ES*

for each region.

Sensitivity analysis: To examine whether our results were sensitive to our use of two time

points only in the main analyses, we repeated all analyses using multiple years to estimate the

average richness for two periods where the data permitted. Specifically, for time series where

more than three years were sampled, we split each time series into two halves, standardised

sampling effort in each of the two periods, and calculated the average species richness in each

period. Then, similar to our main analyses, we used the log-ratio of richness in the second

period divided by richness in the first period, standardised by the duration of sampling in the

region as our estimate of diversity changes occurring at each scale. We fit the same statistical

models to these data, and our results were qualitatively consistent (Extended Data Figure 3).

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