Internal selective attention is delayed by competition between endogenous and exogenous factors.

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Summary

External selective attention is mediated by competition between endogenous (voluntary) and exogenous (involuntary) factors, with the balance of competition determining which stimuli are selected. Occasionally, exogenous factors "win" this competition and drive the selection of salient but task-irrelevant stimuli. Endogenous and exogenous selection mechanisms may also compete to control the selection of internal representations (e.g., those stored in working memory), but how this competition manifests and whether it is resolved in the same way as external attention is unknown. Here, we leveraged the high temporal resolution of human EEG to test three different models describing how competition between endogenous and exogenous factors might influence the selection of internal representations. Unlike external attention, competition between endogenous and exogenous factors did not prompt a selection of task-irrelevant working memory content. Instead, it simply delayed the endogenous selection of task-relevant content by several hundred milliseconds. Thus, competition between endogenous and exogenous factors affects both external and internal selective attention, but in different ways.

Efficient behavioral selection requires rapid comparisons of sensory inputs with internal representations of motor affordances and goal states. Many of these comparisons take place in working memory (WM), a capacity- and duration-limited system that forms a temporal bridge between fleeting sensory phenomena and possible actions (D'Esposito & Postle, 2015; van Ede & Nobre, 2022). Capacity limits in WM necessitate the existence of external selection mechanisms that gate access to this system (i.e., "input gating"; Vogel et al., 2005; O'Reilly & Frank, 2006; Rac-Lubashevsky & Frank, 2021), while rapidly changing environmental circumstances necessitate the existence of internal selection mechanisms that prioritize behaviorally relevant subsets of information stored in WM for action (i.e., "output gating"; Chatham et al., 2014; Ester et al., 2018). Whether similar mechanisms mediate the selection of internal and external information is hotly debated (O'Reilly & Frank, 2006; Chun et al., 2011; Chatham & Badre, 2015; Souza & Oberauer, 2016; van Ede et al., 2020; Rac-Lubashevsky & Frank, 2021; van Ede & Nobre, 2022).

External sensory inputs can be selected based on top-down goals (i.e., endogenous selection) or bottom-up salience (i.e., exogenous selection), with selection ultimately determined by the balance of competition between these factors (Desimone & Duncan, 1995; Corbetta & Shulman, 2002; Buschman & Miller, 2007; Baluch & Itti, 2011; Wolfe, 2020). For example, exogenous factors can reflexively trigger the selection of salient, unexpected, or valuable stimuli (e.g., Theeuwes, 1992; Folk et al., 1992; Bacon & Egeth, 1994; Folk et al., 2009; Fukuda & Vogel, 2009; Anderson et al., 2011), disrupting top-down searches for less salient but task-relevant stimuli (e.g., Wolfe, 2020). These disruptions – termed attention capture – are frequently accompanied by

concurrent shifts in cortical and subcortical spatial priority maps that mediate eye movements and endogenous covert spatial attention (Corbetta & Shulman, 2002; Silver et al., 2005; Bisley & Goldberg, 2010; Sprague & Serences, 2013; Cosman et al., 2018).

Endogenous and exogenous factors may also compete to control the selection of WM content (e.g., van Ede et al., 2020; van Ede & Nobre, 2022). However, less is known about how this competition is resolved. One possibility is that exogenous factors have the same effect on internal attention as they do on external attention: an involuntary and reflexive selection of salient but task-irrelevant information. For example, a salient external stimulus might trigger an exogenous selection of similar WM content. This would be the converse of WM-guided attention capture, where external attention is reflexively oriented to task-irrelevant stimuli that incidentally match attributes of stimuli stored in WM (Olivers et al., 2011). A second possibility is that exogenous factors trigger an automatic "refreshing" of matching but otherwise task-irrelevant WM content. This alternative is motivated by neurocomputational models of conjunctive coding (e.g., Schneegans & Bays, 2017; 2018) and short-term synaptic plasticity (Mongillo et al., 2008; Manohar et al., 2019; Li et al., 2021) where bottom-up inputs can reactivate spiking patterns and/or synaptic networks responsible for holding stimulus-specific information in WM. A third possibility is that internal selection is mediated by competition between endogenous and exogenous factors, with the latter delaying the selection of task-relevant information by the former. Although this possibility has been rejected in the external attention literature (e.g., Remington et al., 2001), it may explain a recent finding documenting delays in oculomotor biases to the locations of items

stored in WM when experimental factors place endogenous and exogenous selection mechanisms in conflict (van Ede et al., 2020).

Here, we tested these models. We recorded EEG while participants performed a retrospectively cued WM task typically used to study internal attention (e.g., Griffin & Nobre, 2003; Landman et al., 2003; see Souza & Oberauer, 2016 and Myers et al., 2017 for recent comprehensive reviews). In different experimental blocks, a cue presented during WM storage indicated which of two memorized positions would be probed for recall (pro-cue trials) or which position would not be probed for recall (anti-cue trials). We reasoned that the anti-cue manipulation would create a state of conflict between endogenous and exogenous selection mechanisms, a point supported by studies documenting visual search costs when participants are cued to the identity of an upcoming distractor (e.g., Moher & Egeth, 2012; van Ede et al., 2020). To discriminate between the internal selection models described in the previous paragraph, we quantified how anti-cues influenced the quality of cue-matching but task-irrelevant location-specific signals decoded from scalp EEG activity during WM storage. The capture and refreshing models predict that exogenous selection mechanisms engaged by anti-cues should prompt an inadvertent selection or refreshing of the cue-matching but task-irrelevant item on a subset of trials, yielding a transient increase in task-irrelevant position decoding performance. Conversely, the delayed selection model predicts that conflict generated by exogenous selection delays the endogenous selection of task-relevant WM content, yielding a delay in improved decoding performance during anti-cue relative to pro-cue trials.

Results

We recorded EEG while 40 human volunteers performed a retrospectively cued spatial recall task (Fig 1A). Participants remembered the positions of two discs over a brief delay, and a retrospective color cue presented 1.25 seconds later instructed participants to continue remembering the positions of both discs (i.e., uninformative trials) or to prioritize one of the discs for subsequent recall (i.e., informative trials). At the end of the trial, participants recalled the position of the task-relevant disc via mouse click. Behavioral performance was quantified via average response times and average absolute recall error (i.e., the average absolute difference between the correct and reported disc position). In separate experimental blocks, participants performed a pro-cue task or an anti-cue task. During the pro-cue task informative cues were assigned 100% validity; during the anti-cue task informative cues were assigned 0% validity (i.e., the cue color indicated which disc was task-irrelevant). This allowed us to disentangle the effects of endogenous and exogenous factors on the selection of WM content: during the pro-cue task the color cue indicates which of the two remembered stimuli are task relevant, and endogenous and exogenous selection mechanisms are aligned. During the anti-cue task, however, the color cue indicates which of the two stimuli are task-irrelevant, placing endogenous and exogenous selection mechanisms in competition (e.g., Moher & Egeth, 2012; van Ede et al., 2020). Task order (i.e., pro-followed by anti-cue or vice versa) was counterbalanced across participants, and participants were explicitly reminded about cue validity at the beginning of every block.

Figure 1. Memory Task and Performance. *(A) Participants remembered the locations of two discs over a blank delay. Each disc could appear at one of eight positions along the perimeter of an imaginary circle centered at fixation (upper right panel). (B) Effects of cue type (informative, uninformative) and task type (pro-cue, anti-cue) on average absolute recall errors. (C) We estimated the effects of exogenous factors on recall performance by computing the difference between informative pro-cue trials (i.e., where endogenous and exogenous factors are aligned) and informative anti-cue trials (i.e., where endogenous and exogenous factors are opposed). We estimated the effects of endogenous factors on recall performance by computing the difference between informative pro-cue trials and uninformative pro-cue trials minus the estimated effect of exogenous factors (see text for specifics). Identical analyses were also applied to participants response times (D, E). Error bars depict the 95% confidence interval of the mean.*

Endogenous and Exogenous Factors Influence the Selection of Task-Relevant WM Content, but in Different Ways.

A two-factor repeated measures analysis of variance (ANOVA) applied to participants' average absolute recall errors (Fig 1B) revealed a main effect of cue type (i.e., informative vs. uninformative; [$F(1,39) = 15.854$, $p = 0.0003$, $\eta^2 = 0.289$]), a main effect of task (i.e., pro- vs. anti-cues; [$F(1,39) = 8.168$, $p = 0.0068$, $\eta^2 = 0.1732$]), and a significant interaction between these factors [$F(1,39) = 5.35$, $p = 0.0261$]. A complementary analysis of response times (Fig 1D) revealed a main effect of cue type [$F(1,39) = 483.046$, $p < 0.0001$, $\eta^2 = 0.9253$], no main effect of task [$F(1,39) = 0.022$, $p = 0.884$, $\eta^2 = 0.060$], and a significant interaction between these factors [$F(1,39) = 30.362$, $p < 0.0001$]. In planned comparisons, we sought further clarity on how endogenous and

exogenous factors influenced participants' memory performance. We reasoned that during informative pro-cue trials endogenous and exogenous factors are aligned while during informative anti-cue trials they conflict. Thus, to isolate the effects of exogenous factors on memory performance we compared participants' recall errors and response times across informative pro- and anti-cue trials (i.e., the simple effect of task for informative cues). Conversely, to quantify the effect of endogenous selection on memory performance, we first calculated differences in participants' recall errors and response times across informative and uninformative trials in the pro-cue task (i.e., the simple effect of cue type for the pro-cue task). Next, we subtracted the effects of exogenous factors estimated in the previous step from these differences to isolate the effects of endogenous factors on memory performance (i.e., while accounting for the fact that during the pro-cue task endogenous and exogenous cues are aligned while during the anti-cue task they are opposed). Endogenous factors had a facilitatory effect on task performance, lowering recall errors (M = 1.78°; 95% CI = 0.645°-3.112°; Fig 1C) and speeding response times (M = 0.165 sec; 95% CI = 0.014-0.305 sec; Fig 1E). In contrast, exogenous factors significantly worsened participants' recall errors (M = 0.961°; 95% CI = 0.191°-1.863°; Fig 1C) but had no effect on response times (M = -0.055, 95% CI = -0.073-0.189; Fig 1E). Thus, endogenous and exogenous factors had faciliatory and deleterious effects on participants' memory performance, respectively.

Manipulation Check: The Anti-cue Task Requires a Greater Degree of Cognitive Control than the Pro-cue task. A key assumption of our experimental approach holds that the anti-cue task produces conflict between endogenous and exogenous selection

Figure 2. Frontal Theta Power is Greater During the Anti- vs. Pro-Cue Task. *(A) Difference in theta-power (4-7 Hz) scalp topography during the pro- and anti-cue tasks. Pro- and anti-cue theta power estimates were averaged over a period spanning 2.5-3.0 sec after trial start (i.e., 750-1250 ms after cue onset). Electrode-wise power estimates during the pro-cue task were subtracted from corresponding estimates during the anti-cue task, i.e., larger values indicate higher theta power during the anti- vs. pro-cue task. (B) Time-resolved differences in pro- and anti-cue theta power computed from frontal electrode sites. Theta power estimates were larger during the anti- vs. pro-cue task beginning approximately 600 ms after cue onset. Shaded regions depict the 95% confidence interval of the mean. Vertical solid lines at times 0.00 and 3.00 depict the onset of the sample and recall displays, respectively; the vertical dashed line at time 1.75 depicts the onset of an informative retrocue. The horizontal black bar at the top of the plot marks periods where the difference between anti- and pro-cue theta power was significantly greater than zero (cluster-based permutation tests; see Methods). See also Figure S1.*

mechanisms while the pro-cue task does not. Thus, we reasoned that a greater degree

of cognitive control would be required during the anti-cue task compared to the pro-cue

task. To verify this assumption, we computed estimates of theta power (4-7 Hz) over

frontal electrode sites during the pro- and anti-cue tasks. Frontal theta power has

robustly linked with the need for cognitive control (Cavanagh & Frank, 2014), scales

with WM load (Jensen & Tesche, 2002), predicts successful working memory updating

(Itthipurripat et al., 2013), and is inversely correlated with default mode network activity

(Scheeringa et al., 2008). Thus, we expected larger frontal theta power estimates during

the anti-cue vs. the pro-cue task. We tested this prediction by analyzing data from

informative trials during the pro- and anti-cue tasks. An analysis of theta power topography revealed significantly higher frontal theta power during the anti-cue relative to the pro-cue task (Fig 2A; Fig S1A, B). Time-resolved analyses of frontal theta power revealed that task-level differences began ~600 ms after retrocue onset and persisted well into the probe period (Fig 2B; S1C). Thus, the anti-cue task required a greater degree of cognitive control than the pro-cue task, supporting the general logic of our experimental approach.

Decoding Remembered Locations from Scalp EEG Activity

To understand how endogenous and exogenous factors influence the selection of WM content, we examined how pro- and anti-cues influenced our ability to decode stimulus positions from scalp EEG. Our approach builds on studies demonstrating that stimulus- and location-specific information can be decoded from scalp EEG signals (e.g., Foster et al., 2016) and that attending to an item or location stored in WM selectively boosts decoding for the attended information (e.g., Lewis-Peacock et al., 2012; LaRocque et al., 2013; Rose et al., 2016; Ester et al., 2018; Nouri & Ester, 2020).

Figure 3. Decoding Analysis. *(A) We computed the multivariate (Mahalanobis) distance between patterns of alpha-band activity measured during each trial of a training data set with position-averaged patterns measured in an independent training data set. (B) If patterns of alpha-band activity contain robust position-specific signals, then multivariate distances should be smallest when comparing activation patterns associated with memory for a specific location in the test data set with activation patterns associated with memory for the same location in the training data set, and the multivariate distance between test and training data patterns should increase with polar distance.*

We implemented a multivariate distance-based decoding analysis (e.g., Wolff et al., 2017) that was customized for our (parametric, circular) location space (e.g., Fig 1A; Fig 3). This approach is conceptually similar to image reconstruction techniques (i.e., "inverted encoding models", Sprague et al., 2014; 2016; Ester et al., 2018) but does not require the experimenter to measure or specify a specific coding model or basis set. A complete description of this analysis can be found in the accompanying STAR methods, but briefly we decoded the location of each disc by comparing trial-wise multielectrode patterns of occipitoparietal alpha-band (8-13 Hz) activity in a test data set with position-averaged patterns in an independent training data set. A key assumption of this approach is that patterns of activity associated with memory for neighboring positions in the training and test data sets should be more similar than patterns of activity associated with memory for distal positions. Thus, multivariate distances should be smallest when comparing activation patterns associated with memory for the same

location in the training and test data sets, and multivariate distance should increase with the polar distance between the compared locations (e.g., Fig 3A). We therefore computed the multivariate (Mahalanobis) distance between activity patterns measured during each trial and time point of a test data set with averaged activity patterns associated with memory for all possible stimulus positions in an independent training data set, yielding an *n* location x 1 vector of distance estimates. Trial-wise distance estimates were circularly shifted to a common center (0°, by convention) and averaged. If the measured EEG signal contains robust location information, this analysis should reveal an inverted "tuning-curve" like function over time (Fig 3B). This function was inverted and convolved with a cosine function, yielding a single decoding estimate for each test trial and timepoint. This entire analysis was performed separately for each disc (i.e., red vs. blue) for each cue condition (i.e., informative vs. uninformative) and task type (i.e., pro- vs. anti-cue), yielding a series of condition-specific decoding time series for each participant. To facilitate comparisons across cue conditions and tasks participant-level decoding time series were sorted by task relevance: during the pro-cue task decoding performance for the cue-matching disc was designated task-relevant and decoding performance for the cue-nonmatching disc was designated task-irrelevant; during the anti-cue task this mapping was reversed. Finally, we verified that our key findings (e.g., Figs 5-7) were not dependent on this specific decoding approach; an inverted encoding model analysis (STAR Methods) yielded qualitatively similar results (Fig S2).

Our pro-cue task is a direct replication of two earlier studies that reported rapid increases in task-relevant location decoding performance following the appearance of

Figure 4. Selection Models. *Each panel depicts a hypothetical retrocue-locked decoding time course for task-relevant and task-irrelevant stimuli. Task relevance is defined by the relationship between the color of the retrospective cue and the position of the color-matching disc. For example, during pro-cue blocks the cue-matching disc is task-relevant while the cue-nonmatching disc is task-irrelevant; during anti-cue blocks the cue-nonmatching disc is task-relevant while the cue-matching disc is task-irrelevant.*

an informative color cue (Ester et al., 2018; Nouri & Ester, 2020). Thus, we expected to

replicate these findings in the current study (Figure 4A). Next, we considered two

models describing how exogenous mechanisms might influence the selection of task-

relevant WM content during the anti-cue task (Fig 4). The first model was motivated by

studies reporting exogenous capture but salient but task-irrelevant stimuli in the external

attention literature (e.g., Theeuwes, 1992; Folk et al., 1992) and neurocomputational

models of WM storage predicated on conjunctive coding (Schneegans & Bays, 2017;

2018) or pattern completion (e.g., Mongillo et al., 2008; Manohar et al., 2019). A

capture-based account of internal selection holds that exogenous factors prompt

reflexive selection of WM contents that incidentally match the anti-cue; conjunctive

coding and pattern completion models predict that the anti-cue should prompt a "re-

activation" of spiking or synaptic mechanisms responsible for storing color-location

bindings in WM. Both accounts predict a transient increase in decoding performance for

the task-irrelevant but cue-nonmatching item during the anti-cue task (Fig 4B). We used

this prediction as a benchmark to evaluate an alternative model where exogenous mechanisms delay the selection of task-relevant WM content by endogenous mechanisms (Fig 4C). This model predicts that competition between endogenous and exogenous factors delays the selection of task-relevant WM content, yielding an increase in task-relevant decoding performance that begins significantly later during the anti-cue task (where endogenous and exogenous factors compete to control selection) compared to the pro-cue task (where endogenous and exogenous factors are aligned).

Competition Between Endogenous and Exogenous Mechanisms Delays the Selection of Task-Relevant WM Content

Our experimental task (Fig 1A) was deliberately constructed so that the effects of endogenous and exogenous factors on the selection of WM contents could be measured during informative *and* uninformative trials. For example, during uninformative trials participants received an uninformative retrospective cue instructing them to remember the positions of both discs. Upon presentation of the probe display, this uninformative cue was replaced by a 100% valid (pro-) or 0% valid (anti-) cue instructing participants which disc to report. Conversely, during informative trials pro- and anti-cues were presented midway through the storage period. Since informative and uninformative trials had different response demands (i.e., pro- and anti-cues presented at the end of uninformative trials required an immediate response while pro- and anti-cues presented during the memory delay during informative trials did not), we analyzed data from these conditions separately.

Figure 5. Location Decoding Performance During Neutral Trials. *(A, B) Decoding performance for task-relevant and task-irrelevant locations during pro-cue and anti-cue blocks, respectively. (C) Overlay of task-relevant location decoding performance for pro-cue and anti-cue blocks (i.e., the blue lines in panels A and B). Solid vertical lines at time 0.00 and 3.00 depict the onset of the sample and probe displays, respectively. The dashed vertical line at time 1.75 depicts the onset of the (uninformative) retrocue. Gray shaded region spanning 0.00-0.50 marks the duration of the sample display. Horizontal bars at the top of each plot mark intervals where decoding performance was significantly greater than zero (nonparametric cluster-based randomization test; see Methods) or intervals where decoding performance for one location was significantly greater than decoding performance for the other location. Shaded regions around each line depict bootstrapped confidence intervals of the mean.*

We first considered data from uninformative trials (Fig 5). Task-relevant and task-irrelevant location decoding performance in the pro-cue (Fig 5A) and anti-cue (Fig 5B) tasks increased rapidly during the sample display but returned to chance levels by the time the (uninformative) retrocue was presented 1.75 sec later. Task-irrelevant decoding performance remained at chance levels through the retrocue and probe displays while task-relevant decoding performance increased from chance- to above-

chance levels during the probe period. Visual comparisons of probe-locked task-relevant decoding performance suggested that above-chance decoding performance was reached earlier during the pro- relative to the anti-cue task (Fig 5C). To quantify this effect, we extracted and compared probe-locked task-relevant decoding time courses during the pro- and anti-cue tasks via cross-correlation. Specifically, we computed correlations between the timeseries of task-relevant decoding performance during the pro- and anti-cue tasks while temporally shifting the former by -1.0 to +1.0 sec in 4 msec intervals relative to the latter, yielding a correlation-by-lag function. Next, we compared the observed cross-correlation function to a null distribution computed by randomizing participant-level task labels (i.e., pro vs. anti-cue) 10,000 times. Thus, correlations larger than those expected by chance at negative lag values and correlations smaller than those expected by chance at positive lag values indicate that task-relevant decoding performance reaches above-chance levels earlier during the pro- vs. anti-cue task. Observed cross-correlation coefficients (Fig 7A) exceeded those expected by chance over lags spanning -0.33 to -0.22 sec and fell below those expected by chance over a period spanning +0.25 to +0.35 sec, confirming that task-relevant decoding performance reached above chance levels earlier during the pro- vs. anti-cue task.

A complementary analysis of cue-locked decoding performance during informative trials yielded similar conclusions (Fig 6; see also Fig S2). Here, task-relevant (but not task-irrelevant) decoding performance reached above-chance levels during the retrocue period during both the pro- and anti-cue tasks (Figs 6A and 6B, respectively). Visual comparisons of cue-locked decoding performance suggested that above-chance decoding performance was reached earlier during the pro- vs. anti-cue

Figure 6. Location Decoding Performance During Informative Trials. *(A, B) Decoding performance for task-relevant and task-irrelevant locations during pro-cue and anti-cue blocks, respectively. (C) Overlay of task-relevant location decoding performance for pro-cue and anti-cue blocks (i.e., the blue lines in panels A and B). Conventions are identical to Figure 4.*

task (Fig 6C). This effect was confirmed by a cross-correlation analysis (Fig 7B): we observed cross-correlation coefficients exceeding those expected by chance over lags spanning -0.44 to -0.35 sec and falling below those expected by chance over lags spanning -0.14 to +0.13 and +0.28 to +0.70 sec.

Alternative Explanations

Figure 7. Cross-correlation Results. (A) Cross-correlation between the time series of task-relevant location decoding during the response period of neutral pro-cue and neutral anti-cue trials. Peaks over negative values depict epochs where the time series of task-relevant decoding performance during pro-cue blocks leads the time series of decoding performance during anti-cue blocks. (B) An identical analysis was applied to decoding time series measured during the cue-to-probe interval of informative pro-cue and anti-cue trials. The black line and shaded area depict the mean and 95% confidence interval of cross-correlation coefficients computed by randomizing time series condition labels (i.e., pro- vs. anti-cue block) 10,000 times. Blue horizontal bars at the top of the plot depict intervals where observed cross-correlations exceeded the 95th or 2.5th percentile of the null distribution.

The first model of internal exogenous attention we considered (Fig 4B) predicts that exogenous factors prompt either a reflexive selection of the cue-matching but task-irrelevant disc or an automatic refreshing of position information during the anti-cue task. This prediction is difficult to reconcile with the findings summarized in Figs 5-7: during the anti-cue task, we observed no evidence for above-chance task-irrelevant decoding performance during uninformative (Fig 5B) or informative trials (Fig 6B). However, evidence for retrospective capture by the cue-matching but task-irrelevant disc may have been obscured by trial averaging. For example, perhaps the retrocapture effect is small, short, lived, or intermittent (i.e., occurring on only a subset of trials). To examine these possibilities in greater detail, we recomputed task-irrelevant decoding performance during the anti-cue task after sorting participants' data by median recall error (i.e., "high" vs. "low"), reasoning that since exogenous factors have a deleterious effect on participants' recall errors (Fig 1C),

exogenous selection of the task-irrelevant disc – as indexed by higher task-irrelevant decoding performance – should be more evident during high recall error trials. However, this was not the case: we observed no evidence for above-chance task-irrelevant decoding performance during low- or high-error informative (Fig S3A) or uninformative trials (Fig S3B). Thus, it is unlikely that the pattern of exogenous-then-endogenous selection predicted by the retro-capture model was obscured by trial-averaging.

The delayed selection model (Fig 4C) predicts that endogenous and exogenous selection compete to control internal selection during anti-cue trials, and that selection proceeds only once this competition has been resolved. The data reported here are nominally consistent with this prediction: task-relevant decoding performance reached above-chance levels during the pro-cue task (i.e., when endogenous and exogenous factors are aligned) earlier than during the anti-cue task (i.e., when endogenous and exogenous factors were opposed; Figs 5C, 6C; 7). We considered several alternative explanations for this pattern. For example, one trivial possibility is that that it simply takes participants longer to interpret the anti-cues. However, this explanation is difficult to reconcile with the fact that neither the main effect of task (i.e., pro-cue vs. anti-cue; Fig 1D) nor the simple effect of task (Fig 1E) on response times during informative cue trials reached significance. A second possibility is that delayed above-chance decoding performance during the anti-cue task was caused by carryover effects. For example, although task order was counterbalanced across observers, perhaps participants who completed the pro-cue task followed by the anti-cue task had extra difficulty interpreting anti-cues compared to participants who performed the anti-cue task followed by the pro-cue task. To test this possibility, we compared the time-courses of task-relevant

decoding performance during informative anti-cue trials in participants who performed

the pro-cue task followed by the anti-cue task (N = 17) or vice versa (N = 23). For both

groups, task-relevant decoding performance reached above chance levels shortly

before or immediately after the onset of the probe display (Fig S4). If anything, the onset

of above-chance decoding performance occurred later for participants who performed

the anti-cue task second vs. those who performed the anti-cue task first, though this

difference was not significant (p = 0.146; randomization test, see Methods). Thus, order

effects cannot account for delays in task-relevant decoding performance during the anti-

cue vs. pro-cue blocks.

Finally, we considered the hypothesis that the selection of cue-matching but task-

Figure 8. Frontal Theta Power and Decoding Performance. (A) Difference in frontal theta power during the anti-cue and pro-cue tasks. Values > 0 depict larger frontal theta estimates during the anti- vs. pro-cue task. (B) Scalp distribution of the data shown in (A). (C) Decoding performance for the cue-matching but task-irrelevant disc during low-theta power and high-theta power anti-cue trials.

irrelevant positions during the anti-cue task were obscured by successful cognitive control. Specifically, we reasoned that selection of the task-irrelevant position might be more likely during trials where participants "spaced out" compared to trials when they were actively engaged in the task. To test this hypothesis, we re-computed cue-matching but task-irrelevant location decoding performance after sorting participants' alpha-band EEG data by frontal theta power (Fig 2), reasoning that inadvertent selection of cue-matching but task-irrelevant stimuli would be more likely during trials where frontal theta power was low vs. high. However, we observed no evidence for above-chance decoding of the cue-matching but task-irrelevant position during high- or low-theta power trials (Fig 8). This analysis provides converging evidence suggesting that exogenous factors do not lead to a selection or re-activation of cue-matching but task-irrelevant WM content, but instead delay the endogenous selection of task-relevant WM content.

Discussion

Selective attention can be allocated to sensory inputs and internal representations based on voluntary, endogenous factors or involuntary, exogenous factors. An enormous literature suggests that external selection is mediated by competition between endogenous and exogenous factors, with the focus of selection determined by the balance of competition between these factors (e.g., Desimone & Duncan, 1995; Corbetta & Shulman, 2002; Connor et al., 2004). Here, we show that exogenous factors influence internal selection by a different mechanism: delayed selection. Capitalizing on recent findings (e.g., Ester et al., 2018; Nouri & Ester, 2020) we used EEG to decode the contents of spatial WM after a 100% informative or 0% informative retrospective cue (i.e., pro- and anti-cues, respectively). We found no evidence for exogenous selection or refreshing of cue-matching but task-irrelevant WM content during the anti-cue task. Instead, the selection of task-relevant WM content was simply delayed by ~250 ms during the anti-cue relative to the pro-cue task. Thus, the selection of task-relevant WM content is delayed until competition between endogenous and exogenous factors is resolved.

The current findings have major implications for models of external and internal attentional selection. External attention is directed to sensory inputs, and competition between endogenous and exogenous factors sometimes results in attention "capture" by unexpected, salient, and valuable but otherwise task-irrelevant stimuli (Theeuwes, 1992; Folk et al., 1992; Bacon & Egeth, 1994; Folk et al., 2009; Fukuda & Vogel, 2009; Anderson et al., 2011). Endogenous and exogenous factors also compete to control the selection of internal representations (e.g., van Ede et al., 2020; van Ede & Nobre,

2022), but this competition is resolved in a fundamentally different way. We observed no evidence for the selection of salient internal representations akin to capture effects seen in external attention. Instead, we found that the selection of task-relevant internal representations was delayed when endogenous and exogenous selection mechanisms were placed in conflict compared to when they were aligned. This finding supports claims that the selection of external sensory inputs and internal memory representations are mediated by different mechanisms (Chun et al., 2011; van Ede & Nobe, 2022).

The current findings also have implications for neurocomputational models of WM based on conjunctive coding (e.g., Schneegans & Bays, 2017; 2018) or synaptic storage (e.g., Manohar at al. 2019; Li et al., 2021). On the one hand, conjunctive coding models predict that WM representations are maintained by spiking activity in feature- and/or location-specific neural populations. While the exact mechanisms vary by implementation (see Scheegans & Bays, 2017 and Oberauer & Lin, 2017 for representative examples), these models generally predict that a feature probe in one dimension (e.g., orientation) activates spiking patterns in neural populations that code this feature and those that code other features of the same object (e.g., color) and/or its location. This, in turn, enables robust read-out of the probed and non-probed stimulus dimensions by downstream neural populations. While these models were not developed to describe the anti-cue task contemplated here, one could reasonably predict an increase in task-irrelevant decoding performance after presentation of an anti-cue based on their general architecture. We observed no evidence for such an effect, and it remains to be seen whether these models can be modified to predict behavioral and neural data during pro- and anti-cue tasks.

On the other hand, pattern completion models predict that the contents of WM reside in different neural states – an "active" state mediated by sustained spiking activity and a "latent" state mediated by short-term synaptic plasticity. Presentation of a feature probe that matches a stimulus stored in a latent format reinstates activity patterns evoked when that stimulus was encoding, prompting and/or "refreshing" of the neural representation of the probe-matching item through pattern completion (Mongillo et al., 2008; Manohar et al., 2019). This prediction enjoys some support: a representation stored in WM item can be "re-activated" (as indexed by above-chance EEG decoding performance) by a task-irrelevant sensory input (Wolff et al., 2017) or a transcranial magnetic stimulation (TMS) pulse applied over sensory cortex (Rose et al., 2016). Conversely, in the current study we found no evidence for a reactivation of cue-matching but task-irrelevant WM content following presentation of an anti-cue. However, one salient difference between the current study and prior work is that in the latter, an informative retrospective cue instructed presented prior to the "impulse" stimulus instructed participants which of two remembered stimuli should be prioritized for report. Thus, re-activation of information stored in synaptic traces may be contingent on the network responsible for storing information to be selected or otherwise primed for decision making and action.

A recent study reported that the onset of small-but-robust biases in gaze position (referred to hereafter as microsaccades) towards the memorized location of a task-relevant stimulus were delayed following the appearance of an anti- vs. pro-cue (van Ede et al., 2020). The authors of this study speculated that delays in selecting task-relevant information during an anti-cue task were caused by an automatic "refreshing" of

cue-matching but task-irrelevant WM content. However, as the authors of this study acknowledged, this hypothesis cannot be substantiated based on microsaccades alone. First, although attention-related modulations of cortical and subcortical processing are larger during trials containing microsaccades towards the location of a (covertly) attended stimulus, clear attention-related modulations are also observed in the absence of microsaccades (Yu et al., 2022; Liu et al., 2022; though see Lowet et al. 2018 for a contradictory result). Second, any of the three models examined in the current study (i.e., retro-capture, automatic refreshing, or delayed selection) could have produced the delays in microsaccade biases reported by van Ede et al. (2020). Although we did not (and could not, lacking a functional eyetracker at the time this study was conducted) record eye position with sufficient spatiotemporal resolution needed to replicate the findings reported by van Ede et al., we observed no evidence for an enhancement of cue-matching but task-irrelevant position information during the anti-cue task that would be predicted by the account endorsed by van Ede and colleagues (2020). Thus, we speculate that the delays in microsaccade biases reported by these authors were generated by delayed selection of the appropriate WM content. Further research could compare the time-courses of oculomotor and neural signals associated with internal selection to test this possibility.

To summarize, the current findings support recent suggestions that endogenous and exogenous selection mechanisms compete to control access to internal WM representations. However, this competition is resolved in a fundamentally different way than that seen during external attention. Specifically, endogenous and exogenous competition does not produce an errant selection or refreshing of salient but task-

irrelevant WM content. Instead, this competition delays the selection of task-relevant memory content by endogenous mechanisms. This result reveals a fundamental difference between the operation of external and internal attentional selection, and suggests that they are mediated by different mechanisms.

Methods

Participants. 42 human volunteers (both sexes) participated in a single 2.5-hour testing session. Participants were recruited from the Florida Atlantic University community via campus advertisements and remunerated at $15/h in Amazon.com gift cards. All participants gave both written and oral informed consent in compliance with procedures established by the local institutional review board, and all participants self-reported normal or corrected-to-normal visual acuity. Two participants voluntarily withdrew from the study prior to completing both cue conditions (i.e., pro-cue vs. anti-cue); data from these participants were excluded from final analyses. Thus, the data reported here reflect the remaining 40 participants.

Testing Environment. Participants were seated in a dimly-lit and acoustically shielded recording chamber for the duration of the experiment. Stimuli were generated in MATLAB and rendered on a 17'' Dell CRT monitor cycling at 85 Hz (1024 x 768 pixel resolution) via PsychToolbox3 software extensions (Kleiner et al., 2007). Participants were seated approximately 65 cm from the display (head position was unconstrained). To combat fatigue and/or boredom, participants were offered short breaks at the end of each experimental block.

Spatial Retrocue Task. A task schematic is shown in Fig 1A. Each trial began with the presentation of an encoding display lasting 500 ms. The encoding display contained two colored circles (blue and red; subtending 1.75° visual angle from a viewing distance of 65 cm) rendered at of eight polar locations (22.5° to 337.5° in 45° increments) along the perimeter of an imaginary circle (radius 7.5° visual angle) centered on a circular fixation point (subtending 0.25°) rendered in the middle of the display. The locations of the two

discs were counterbalanced across each task (i.e., pro-cue vs. anti-cue), though not necessarily within an experimental block. Participants were instructed to maintain fixation and refrain from blinking for the duration of each trial.

The sample display was followed by a 1.25 sec blank display and a 1.25 sec retrocue display. Retrocues were defined by a change in the color of the fixation point; during informative trials the fixation point changed colors from black to either blue or red (i.e., matching the color of a remembered disc). During neutral trials, the fixation point changed colors from black to purple (the "average" of blue and red). At the end of the trial, a response display containing a mouse cursor, "?" symbol, and a probe circle appeared. Participants were required to report the location of a remembered disc by clicking along the perimeter of the probe circle (during neutral trials, the fixation point changed colors from purple to either blue or red during the response display, indicating which item should be reported). Participants were instructed to prioritize accuracy over speed, and no response deadline was imposed. The trial terminated as soon as the participant clicked on a location. Sequential trials were followed by a 1.5-2.5 sec blank period (randomly and independently selected from a uniform distribution after each trial).

During the first half of the experiment (e.g., experimental blocks 1-8), each participant was assigned to the pro-cue or anti-cue task. In the pro-cue task, participants were informed that during informative would be required to click on the location of the disc matching the retrocue color. Conversely, in the anti-cue task participants were informed that during informative trials they would be required to click on the location of the disc that did not match the retrocue color. Participants completed

eight blocks of 56 trials in the both the pro-cue and anti-cue tasks. Task order (i.e., eight blocks of the pro-cue task followed by eight blocks of the anti-cue task or vice versa) was counterbalanced across participants.

EEG Acquisition and Preprocessing. Continuous EEG was recorded from 63 uniformly distributed scalp electrodes using a BrainProducts "actiCHamp" system. The horizontal and vertical electrooculogram (EOG) were recorded from bipolar electrode montages placed over the left and right canthi and above and below the right eye, respectively. Live EEG and EOG recordings were referenced to a 64th electrode placed over the right mastoid and digitized at 1 kHz. All data were later re-referenced to the algebraic mean of the left- and right mastoids, with 10-20 site TP9 serving as the left mastoid reference.

Data preprocessing was carried out via EEGLAB software extensions (Delorme & Makeig, 2004) and custom software. Data preprocessing steps included the following, in order: (1) resampling (from 1 kHz to 250 Hz), (2) bandpass filtering (1 to 50 Hz; zero-phase forward- and reverse finite impulse response filters as implemented by EEGLAB), (3) epoching from -1.0 to +5.0 relative to the start of each trial, (4) identification, removal, and interpolation of noisy electrodes via EEGLAB software extensions, and (5) identification and removal of oculomotor artifacts via independent components analysis as implemented by EEGLAB. After preprocessing, location decoding analyses focused exclusively on the following 10-20 occipitoparietal electrodes: P7, P5, P3, Pz, P2, P4, P6, P8, PO7, PO3, POz, PO2, PO4, PO8, O1, O2, Oz.

Data Cleanup. Prior to analyzing participants' behavioral or EEG data, we excluded all trials where the participant responded with a latency of < 0.4 sec (we attributed these trials to accidental mouse clicks following the onset of the probe display rather than a

deliberative recall of a specific stimulus position) and more than 3 standard deviations above the average response time across all experimental conditions. This resulted in an average (±1 S.E.M.) loss of 14.43 ±0.93 trials (or 1.67% ± 0.11% of trials) across participants but had no qualitative effect on any of the findings reported here.

Decoding Spatial Positions from Posterior Alpha-Band EEG Signals. Location decoding was based on the multivariate distance between EEG activity patterns associated with memory for specific positions. This approach is an extension of earlier parametric decoding methods (Wolff et al., 2017) designed for use in circular feature spaces (e.g., Fig 2A). Following earlier work (e.g., Ester et al., 2018), we extracted spatiotemporal patterns of alpha-band activity (8-13 Hz) from 17 occipitoparietal electrode sites (see *EEG Acquisition and Preprocessing* above). The raw timeseries at each electrode was bandpass filtered from 8-13 Hz (zero-phase forward-and-reverse filters as implemented by EEGLAB software), yielding a real-valued signal f(t). The analytic representation of f(t) was obtained via Hilbert transformation:

$$z(t) = f(t) + if(t)$$

where *i* is the imaginary operator and *if(t)* = $A(t)e^{i\varphi(t)}$. Alpha power was computed by extracting and squaring the instantaneous amplitude A(t) of the analytic signal z(t).

Location decoding performance was computed separately for each disc (i.e., blue vs. red), trial type (i.e., informative vs. neutral) and each task (i.e., pro-cue vs. anti-cue) on a timepoint-by-timepoint basis. In the first phase of the analysis, we sorted data from each condition into 5 unique training and test data sets using stratified sampling while

ensuring that each training set was balanced across remembered positions (i.e., we ensured that each training data set contained an equal number of observations where the location of the remembered stimulus was at 22.5°, 67.5°, etc.). We circularly shifted the data in each training and test data set to a common center (0°, by convention) and computed trial-averaged patterns of responses associated with memory for each disc position in each training data set. Next, we computed the Mahalanobis distance between trial-wise activation patterns in each test data set with position-specific activation patterns in the corresponding test data set, yielding a location-wise set of distance estimates. If scalp activation patterns contain information about remembered positions then distance estimates should be smallest when comparing patterns associated with memory for identical positions in the training and test data set and largest when comparing opposite positions (i.e., those ±180° apart), yielding an inverted gaussian-shaped function (e.g., Fig 2B). Trial-wise distance functions were averaged and sign-reversed for interpretability. Decoding performance was estimated by convolving timepoint-wise distance functions with a cosine function, yielding a metric where chance decoding performance is equal to 0. Decoding results from each training- and test-data set pair were averaged (thus ensuring the internal reliability of our approach), yielding a single decoding estimate per participant, timepoint, and task condition.

Cross-correlation Analyses. Temporal differences in task-relevant location decoding performance were estimated via cross-correlation analyses. For neutral trials, we extracted task-relevant decoding performance during pro- and anti-cue blocks over a period spanning 0.0 to 1.0 seconds following the onset of the probe display (i.e., when

an informative cue instructed participants which disc to recall). We computed correlation coefficients between pro- and anti-cue decoding time series while systematically shifting the pro-cue time series from -1.0 to +1.0 sec relative to the anti-cue decoding time series (Fig 6A, blue line). We compared these correlation coefficients to a distribution of correlation coefficients computed under the null hypothesis (i.e., no systematic difference in pro- and anti-cue decoding time series) by repeating the same analysis 10,000 times while randomizing the decoding condition labels (i.e., pro- vs. anti-cue) for each participant (Fig 6A, black line and shaded area). An identical analysis was performed on task-relevant pro- and anti-cue decoding task performance from 0.0 to 1.75 during informative trials (Fig 6B). We deliberately selected a longer temporal interval for analysis during informative trials as we expected increases in pro- and anti-cue decoding performance to begin during the retrocue period and persist into the ensuing response period.

Quantifying Frontal Theta Power. Analyses of frontal theta power focused on informative trials from the pro- and anti-cue tasks. The raw timeseries at each scalp electrode was bandpass filtered from 4-7 Hz (zero-phase forward-and-reverse filters as implemented by EEGLAB software), yielding a real-valued signal f(t). The analytic representation of this signal was obtained via Hilbert transformation, and theta power was computed by extracting and squaring the instantaneous amplitude A(t) of the analytic signal z(t). Topographic maps of theta power during the pro- and anti-cue tasks (Figs 4A, 4B, 7C) were obtained by averaging power estimates over trials and a temporal window spanning 2.5 to 3.0 sec following the start of each informative trial (i.e., 750 to 1250 ms after cue onset). Based on these maps, we limited further analyses

to power estimates measured at four frontal electrode sites: AFz, Fz, F1, and F2. Data from these electrodes were used to compute time-resolved estimates of theta power during the pro- and anti-cue tasks (Fig S4C) and task differences in theta power (Fig 7C). In a final analysis, we extracted and computed trial-wise estimates of theta power during the anti-cue task (using the same electrodes and temporal window described in the previous paragraph). We sorted participants' anti-cue EEG data into low- and high-theta power groups after applying a media split to power estimates, then decoded the location of the cue-matching but task-irrelevant item within each group (Fig 7C). This allowed us to test whether evidence for exogenous capture and/or "refreshing" of cue-matching memory traces was more likely to occur on trials with low- vs. high theta power.

Inverted Encoding Model. To verify the generality of our findings across analytic approaches, we reconstructed position-specific WM representations from spatiotemporal patterns of alpha-band activity using an inverted encoding model (Fig S1). Our approach was conceptually and quantitatively identical to that used in earlier studies (Foster et al., 2016; Ester et al., 2018). We modeled alpha power at each scalp electrode as a weighted sum of eight location-selective channels, each with an idealized tuning curve (a half-wave rectified cosine function raised to the 8th power, with the maximum response of each function normalized to 1). The predicted responses of each channel during each trial were arranged in a *k* channels by *n* trials design matrix C. The relationship between the EEG data and the predicted responses in *C* is given by a general linear model of the form:

$$B = WC + N$$

where *B* is an *m* electrode by *n* trial training data matrix, W is an *m* electrode by *k* channel weight matrix, and *N* is a matrix of residuals (i.e., noise).

   To estimate *W*, we constructed a training data set containing an equal number of trials for each stimulus position (i.e., 22.5-337.5° in 45° increments). We identified the location φ with the fewest *r* repetitions and constructed a training dataset B_{trn} (*m* electrodes by *n* trials) and weight matrix C_{trn} (*n* trials by *k* channels) by randomly selecting (without replacement) 1 to *r* trials for each of the eight possible stimulus positions. The training dataset was used to compute a weight for each channel C_i using ordinary least-squares estimation:

$$V_i = \frac{\Sigma_i^{-1} W_i}{W_i^T \Sigma_i^{-1} W_i}$$

where *T* and -1 are the matrix transpose and inversion operations, respectively. Σ_i is the regularized noise covariance matrix for each channel *i*, estimated as:

$$\sum_i \frac{1}{n-1} \varepsilon_i \varepsilon_i^T$$

where *n* is the number of training trials and ε_i is a matrix of residuals:

$$\varepsilon_i = B_{trn} - W_i C_{trn,i}$$

Estimates of ε_i were obtained by regularization-based shrinkage using an analytically determined shrinkage observation. In this way, an optimal spatial filter v_i was estimated for each channel C_i, yielding an *m* electrode by *k* filter matrix *V*.

   Next, we constructed a test dataset B_{tst} (*m* electrodes by *n* trials) containing data from all trials not included in the training data set and estimated trial-by-trial channel responses C_{tst} (*k* channels by *n* trials):

$$C_{tst} = V^T B_{tst}$$

Trial-wise channel response estimates were interpolated to 360°, circularly shifted to a common center (0°, by convention), and averaged, yielding a single reconstruction per participant, time point, cue condition (i.e., informative vs. uninformative) and task (i.e., pro vs. anti-cue). Condition-wise channel response functions were averaged, converted to polar form, and projected onto a vector with angle 0°:

$$r = |z| \cos(\arg(z)), \qquad z = ce^{2i\rho}$$

Where c is a vector of estimate channel responses and ρ is a vector of angles at which the channels peak. To ensure internal reliability, this entire analysis was repeated 100 times, and unique (randomly chosen) subsets of trials were used to define the training and test data sets during each iteration. The results were then averaged across permutations to obtain a single reconstruction strength estimate for each participant, task condition, and timepoint.

Statistical Comparisons – Behavioral Data. Participants' behavioral data (i.e., absolute average recall error and average response time; Fig 1B-E) were analyzed using standard repeated-measures parametric statistics (e.g., t-test, ANOVA); for these comparisons we report test statistics, p-values, and effect size estimates.

Statistical Comparisons – Decoding Performance & Inverted Encoding Model. The decoding analysis and inverted encoding model we used assume chance-level performance of 0. Likewise, direct comparisons of decoding performance or reconstruction strength across conditions (e.g., pro-cue vs. anti-cue) assume null statistics of 0. Thus, we evaluated task-relevant and task-irrelevant decoding performance (Figs 4-5, 7, S1-S3) by generating null distributions of decoding performance (or differences in decoding performance across conditions) by randomly inverting the sign of each participant's data with 50% probability and averaging the data

across participants. This procedure was repeated 10,000 times, yielding a 10,000-element null distribution for each time point. Finally, we implemented a cluster-based permutation test (Maris & Oostenveld, 2007; Wolff et al. 2017) with cluster-forming and cluster-size thresholds of $p < 0.05$ to correct for multiple comparisons across time points.

Data Availability

Stimulus presentation software, raw and preprocessed data, and analytic software associated with this project are freely available at https://osf.io/2fvt6/

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