

# 1 **The age of reason: Functional brain network development during childhood**

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## 1 **Abstract**

2 Human childhood is characterized by dramatic changes in the mind and brain. However, little is  
3 known about the large-scale intrinsic cortical network changes that occur during childhood due to  
4 methodological challenges in scanning young children. Here, we overcome this barrier by using  
5 sophisticated acquisition and analysis tools to investigate functional network development in  
6 children between the ages of 4 and 10 years ( $n = 92$ ). At multiple spatial scales, age is positively  
7 associated with brain network segregation. At the system level, age was associated with  
8 segregation of systems involved in attention from those involved in abstract cognition, and with  
9 integration among attentional and perceptual systems. Associations between age and functional  
10 connectivity are most pronounced in visual and medial prefrontal cortex, the two ends of a  
11 gradient from perceptual, externally oriented cortex to abstract, internally oriented cortex. These  
12 findings suggest that both ends of the sensory-association gradient may develop early, in contrast  
13 to the classical theories that cortical maturation proceeds from back to front, with sensory areas  
14 developing first and association areas developing last. More mature patterns of brain network  
15 architecture, controlling for age, were associated with better visuospatial reasoning abilities. Our  
16 results suggest that as cortical architecture becomes more specialized, children become more able  
17 to reason about the world and their place in it.

18 **Keywords:** childhood, functional network, development, graph theory, reasoning

## 19 **Significance**

20 Anthropologists have called the transition from early to middle childhood the “age of reason”,  
21 when children across cultures become more independent. We employ cutting-edge neuroimaging  
22 acquisition and analysis approaches to investigate associations between age and functional brain  
23 architecture in childhood. Age was positively associated with segregation between cortical  
24 systems that process the external world, and those that process abstract phenomena like the past,  
25 future, and minds of others. Surprisingly, we observed pronounced development at both ends of  
26 the sensory-association gradient, challenging the theory that sensory areas develop first and  
27 association areas develop last. Our results open new directions for research into how brains

1 reorganize to support rapid gains in cognitive and socioemotional skills as children reach the age  
2 of reason.

### 3 **Introduction**

4 Children's minds develop fastest during the first decade of life. Sensory and motor skills develop  
5 before complex cognitive skills: children can see and walk before they can solve abstract puzzles.  
6 Diverse skills including reasoning, executive function, emotion regulation, and social cognition all  
7 improve dramatically until 8-10 years of age, after which change slows down (see Akshoomoff et  
8 al. (2014); Kopp (1989); Wellman (2014); Whitaker et al. (2018); but see also Fortenbaugh et al.  
9 (2015)). Developmental psychologists once called these changes the "5-to-7-year shift": the  
10 transition from Piaget's preoperational stage, in which children rely on perceptual information, to  
11 the concrete operational stage, in which children are less bound by perceptual information and  
12 more able to think abstractly (Sameroff and Haith, 1996). Anthropologists have called this  
13 developmental period the "age of reason" or the "age of sense", when children become more  
14 independent from their parents, begin to build more complex social relationships with peers and  
15 other adults, and become less egocentric and more able to understand others' perspectives  
16 (Chandler and Lalonde, 1996; Lancy, 2014).

17 A core tenet of developmental cognitive neuroscience is that brain development proceeds along  
18 the sensory-association axis, with sensory areas developing first and association areas developing  
19 last (Sydnor et al., 2021; Tooley et al., 2021). This sequence is in line with data from both  
20 behavioral and cognitive development (Cole et al., 2005). The far end of the association axis is  
21 anchored by the default mode system (Smallwood et al., 2021), which is furthest from sensory  
22 input and engages primarily in abstract cognitive processes that do not rely on the current  
23 sensory environment. Examples of such processes include remembering the past, projecting the  
24 future, and taking the perspective of others (Buckner and DiNicola, 2019). Other association  
25 systems, such as the dorsal and ventral attention systems, receive and process more input from  
26 the outside world (Corbetta and Shulman, 2002). The frontoparietal system can be thought of as a  
27 toggle controlling the switch between internally and externally oriented cognition, flexibly  
28 coordinating other systems and holding sensory information online. Such processes are commonly  
29 exemplified in working memory and reasoning tasks (Owen et al., 2005; Cole et al., 2013).

1 Data on structural brain development, including cortical thinning, surface area, and white matter  
2 coherence, clearly support early development of sensory areas (Stiles and Jernigan, 2010;  
3 Raznahan et al., 2011; Whitaker et al., 2016; Gennatas et al., 2017; Reynolds et al., 2019). However,  
4 regions of the default mode system, including the medial prefrontal cortex and the precuneus, also  
5 show early structural development (Brown and Jernigan, 2012; Li et al., 2013; Wierenga et al.,  
6 2014; Li et al., 2015). Thus, another possibility is that both ends of the sensory-association axis  
7 become anchored early in life, and developmental processes differentiate and refine the  
8 boundaries of attention and executive systems along this axis later in development. Brain  
9 structure is easier to measure than function in sleeping children, so it has been better  
10 characterized in early childhood (see Houston et al. (2014) and Lenroot and Giedd (2006)).  
11 However, brain function may be more closely linked to cognition and behavior (Zimmermann et  
12 al., 2018; Dhamala et al., 2021), particularly during development when the brain is highly plastic  
13 (Chen et al., 2020).

14 Functional brain networks can be studied at multiple spatial scales: in the whole brain, across  
15 systems, and among regions or parcels. Understanding how functional networks reorganize at the  
16 whole brain level allows us to examine the extent to which segregation is an overall guiding  
17 principle of development, while studying the constituent systems (sometimes referred to as  
18 “networks” in the literature) allows for examination of relationships among specialized functional  
19 subnetworks. The parcel resolution yields more granular detail about which specific brain areas,  
20 or network nodes, might drive effects. Segregation refers to the presence of groups or  
21 subnetworks of densely interconnected nodes, and is thought to emerge partially as a result of  
22 maturing inhibitory interneurons; synchronized inhibition may be necessary for establishing  
23 segregated network function (Cardin, 2018; Kraft et al., 2020; Chini et al., 2021).

24 Functional network development has been studied predominantly in middle childhood (7-10  
25 years) or later (see Morgan et al. (2018) and Grayson and Fair (2017) for review), due to the  
26 challenges of acquiring high quality data in younger children while they are awake. From middle  
27 childhood through adolescence, at the whole brain level, networks become more modular and  
28 segregated with age, supporting improved cognition (Satterthwaite et al., 2013b; Gu et al., 2015;  
29 Grayson and Fair, 2017; Marek et al., 2019). At the system resolution, age is associated with  
30 increases in within-system connectivity, and decreases in between-system connectivity,

1 particularly between the default mode system and executive control and attention systems (Fair et  
2 al., 2008; Chai et al., 2013; Satterthwaite et al., 2013b; Gu et al., 2015; Lopez et al., 2019; Jones et  
3 al., 2021). At the regional level, effects are less consistent, perhaps because findings vary widely  
4 depending on the age range studied (Grayson and Fair, 2017; Morgan et al., 2018). Another way to  
5 examine parcel-level effects is to examine the development of the sensory-association axis across  
6 cortex. Two recent and well-powered studies found that in middle childhood, a sensory-  
7 association gradient is in place, but the most variance in patterns of connectivity is explained by  
8 separation between visual and somatomotor systems (Dong et al., 2021; Xia et al., 2022). By age  
9 12 years, however, the organization of the sensory-association gradient resembles that of adults;  
10 development of the primary sensory-association gradient may be mediated by changes in network  
11 architecture (Dong et al., 2021; Xia et al., 2022). Functional network architecture has been shown  
12 to have cognitive consequences: youth with more segregated networks, and in particular task-  
13 positive (i.e., attention and control systems) and task-negative (i.e., default mode) systems,  
14 perform better on a wide variety of cognitive tasks (Gu et al., 2015; Lopez et al., 2019; Marek et al.,  
15 2019; Jones et al., 2021; Xia et al., 2022).

16 A few studies have characterized functional network development in children younger than 6  
17 years of age, and overall suggest developmental specialization of cortex with age. *In utero*, a proto-  
18 default-mode system is detectable, and visual and motor systems show overlap with that found in  
19 adults, but attention and frontoparietal systems remain undifferentiated (Turk et al., 2019;  
20 Thomason, 2020). Infant brain networks can be studied during sleep: primary sensory systems  
21 have an adult-like architecture at birth, but default, ventral attention, and dorsal attention systems  
22 do not develop a distributed network architecture until 1-2 years of age, and executive control  
23 systems are still immature at 2 years of age (Gilmore et al., 2018). The anticorrelation between  
24 default and dorsal attention system connectivity begins to emerge around 1 year of age (see Gao et  
25 al. (2013),  $n = 147$ ). From the age of 3 months to 6 years, within-system connectivity broadly  
26 increases with age, while between-system connectivity decreases (see Bruchhage et al. (2020),  
27  $n = 196$ , natural sleep). Another way to address challenges involved in scanning young children is  
28 to have them view movies: a study of children aged 4 to 7 years showed that age was positively  
29 associated with connectivity in systems identified with an independent component analysis,  
30 including sensory, motor, default mode, and executive control systems, but not the ventral

1 attention system (see Rohr et al. (2018),  $n = 60$ ). An analysis of the same sample also found that  
2 age was negatively associated with connectivity between seeds in the dorsal attention system  
3 (intraparietal sulcus, frontal eye fields) and areas of the default mode system (Rohr et al. (2017),  
4  $n = 44$ ). In general, more mature patterns of connectivity are associated with better performance  
5 on measures of attention and cognition (Rohr et al., 2017, 2018; Bruchhage et al., 2020; Qi et al.,  
6 2021). These studies of young children have examined connectivity between specific regions or  
7 subsets of regions, but not the architecture of intrinsic cortical networks at rest. Hence, little is  
8 known about how rewiring of intrinsic functional networks supports the profound cognitive  
9 changes that take place during childhood.

## 10 **The present research**

11 Here we focused on functional brain network development between the ages of 4 and 10 years  
12 ( $n = 92$ ). To overcome barriers associated with resting-state data collection from young children,  
13 we applied sophisticated neuroimaging acquisition and analysis approaches to minimize motion  
14 and its impacts, including sequences optimized to reduce motion artifacts (Tisdall et al., 2012),  
15 real-time motion monitoring (Dosenbach et al., 2017), rigorous image quality assurance using  
16 open-source tools, and a preprocessing pipeline optimized to reduce the impact of head motion.  
17 We used network science tools to take a hierarchical analytical approach, asking first whether  
18 whole-brain measures of network topology are associated with age, and then which systems and  
19 parcels of cortex drive patterns of topological refinement. Finally, we asked whether network  
20 structure was associated with cognition. We focused on reasoning because it is a core skill that  
21 develops rapidly until middle childhood (Whitaker et al., 2018), is highly predictive of later  
22 academic outcomes (Fuchs et al., 2006; Ferrer et al., 2007; Pagani et al., 2017), and was assessed  
23 across the majority of our sample. If age-associated changes in network architecture support  
24 reasoning skills, then individual differences in reasoning, controlling for age, should mirror  
25 associations with age. In other words, we predict that children with more mature functional  
26 architecture, i.e., greater network segregation, should have better cognitive skills.

## 1 **Materials & Methods**

### 2 ***Participants***

3 The Institutional Review Board at the University of Pennsylvania approved this study. All parents  
4 provided informed, written consent. Children younger than age 8 provided verbal assent, and  
5 children ages 8 and older provided written assent. Participants were recruited from Philadelphia  
6 and the surrounding regions through advertisements on public transportation, partnerships with  
7 local schools, outreach programs, community family events, and social media ads. Children were  
8 between the ages of 4 and 10.59 years ( $M = 6.85$ ,  $SD = 1.38$ ). We chose to collect data from  
9 children starting at 4 years of age, as collecting functional brain imaging data from awake children  
10 younger than 4 may result in large amounts of unusable data. Parents were asked to report their  
11 child's gender and were provided four sex categories: female, male, other, and prefer not to  
12 answer. We recognize that the wording of this question conflated sex and gender, making it  
13 impossible for us to investigate the relation between brain development and the child's gender  
14 identity, whether within or outside the binary. 54% of children were reported by parents to be  
15 male and 46% were reported by parents to be female; none were reported to be other, suggesting  
16 that we might not have any intersex children in our sample. The racial and ethnic makeup of the  
17 sample was as follows: 61% Black, 36% white, 20% Asian, 8% other, and 10% Hispanic/Latino.  
18 Percentages sum to greater than 100% because parents or guardians could endorse multiple  
19 races. 49% of children had a parent with a college degree or more education and 45% had an  
20 annual family income of \$50,000 or more. For comparison, Philadelphia is 43.6 % Black, 44.8 %  
21 White, 7.8 % Asian, 3.9 % Other, and 15.2 % Hispanic or Latino, and the median household income  
22 was \$49,127 (US Census Bureau, 2020).

23 The target sample size was 123 children with usable data to detect correlations of  $r = .25$  with a  
24 power of greater than .8. However, data usability in young children can be difficult to predict, and  
25 data collection was cut short in 2020 by the COVID-19 pandemic. Resting-state scans were  
26 acquired for 138 participants. Ninety-two participants were included in the final sample.  
27 Participants were excluded for not completing the resting-state scan (e.g., due to falling asleep or  
28 wanting to end the scan early,  $n = 17$ ), or parent-reported diagnosis of Attention-

1 Deficit/Hyperactivity Disorder or developmental delay during the visit, despite not reporting a  
2 diagnosis during screening ( $n = 4$ ).

3 To mitigate the effect of image quality on our analyses, we also employed motion and quality  
4 exclusions, excluding children with average framewise displacement (FD)  $> 1$  mm ( $n = 14$ ), and  
5 censored volumes at 0.5 mm FD. We further excluded children with  $> 30\%$  of frames exceeding 0.5  
6 mm FD ( $n = 8$ , Power et al. (2012)) or artifacts ( $n = 3$ , see below for details). These criteria were  
7 selected to balance the need to include as much data as possible in a young population (Leonard et  
8 al., 2017) and the need to limit the influence of low-quality on connectivity metrics (Power et al.,  
9 2014a).

10 We conducted an additional sensitivity analysis with stricter motion cutoffs: excluding children  
11 with  $> 0.5$  mm average FD ( $n = 9$ ) and censoring volumes with  $> 0.25$  mm FD.

12 25 children were excluded for image artifacts or motion in the original sample. At the more lenient  
13 threshold, these children were younger than the included children ( $t(40.15) = -2.79, p = .008$ ),  
14 but not different on age-normed reasoning scores ( $t(37.60) = -1.47, p = .150$ ). At the stricter  
15 threshold, 34 children were excluded for image artifacts or motion: excluded children were  
16 younger than the included children ( $t(50.96) = -2.13, p = .038$ ), but not different on age-normed  
17 reasoning scores ( $t(52.15) = -1.46, p = .150$ ).

## 18 ***Data acquisition***

19 Prior to the scanning session, participants were acclimated to the scanning environment with a  
20 mock scanner that simulates typical MRI noises. Participants practiced keeping still in the mock  
21 scanner, by watching a movie that would pause each time they moved their heads more than 1  
22 mm. During the MRI session, a researcher stayed in the scanner room with the participant to  
23 reassure the child. Participants viewed a fixation cross on a gray screen throughout the resting-  
24 state scan.

25 Imaging was performed at the Center for Advanced Magnetic Resonance Imaging and  
26 Spectroscopy (CAMRIS) at the University of Pennsylvania. Scanning was conducted using a  
27 Siemens MAGNETOM Prisma 3 T MRI scanner with the vendor's 32-channel coil. 5-minute resting-  
28 state fMRI scans were acquired using a T2\*-weighted multiband gradient-echo echo-planar



1 imaging (EPI) sequence (TR = 2000 ms, TE = 30.2 ms, BW = 1860 Hz/px, flip angle = 90°, voxel size  
2 = 2 mm isotropic, matrix size = 96 × 96, 75 axial slices, FOV = 192 mm, volumes = 150–240, 5  
3 dummy scans, multiband acceleration factor = 3). We chose a multiband factor of 3 to minimize  
4 interactions between multiband and motion (Risk et al., 2021). A whole-brain, high-resolution, T1-  
5 weighted 3D-encoded multi-echo anatomical image (MEMPRAGE) was acquired (TR = 2530 ms,  
6 TEs = 1.69 ms/3.55 ms/5.41 ms/7.27 ms, BW = 650 Hz/px, 3x GRAPPA, flip angle = 7°, voxel size  
7 = 1 mm isotropic, matrix size = 256 × 256, 176 sagittal slices, FOV = 256 mm, total scan time of 4:38  
8 minutes). This anatomical sequence used interleaved volumetric navigators to prospectively track  
9 and correct for subject head motion (Tisdall et al., 2012).

10 To increase the amount of usable data, midway through data collection, we updated our  
11 acquisition strategy in two ways: (i) monitoring head motion in real-time using the Framewise  
12 Integrated Real-time MRI Monitor (FIRMM) system (Dosenbach et al., 2017), and (ii) collecting 10  
13 minutes of low-motion resting-state data (2 resting-state runs of data with FD < 1 mm) when  
14 possible. An incidental feature of this design choice is that it decouples age, and reasoning ability,  
15 from the amount of data acquired for each child. One scan was acquired for 43 children, two scans  
16 were acquired for 48 children, and three scans were acquired for one child. Participants were  
17 eligible for inclusion if they had more than 135 frames of resting-state data. Participants had an  
18 average FD of 0.3 mm ( $SD = 0.18$  mm). For participants with more than one usable resting-state  
19 run, we took an average of FD across runs, weighted by run length. All analyses controlled for  
20 average FD and total number of resting-state frames collected.

## 21 ***Image preprocessing***

22 Results included in this manuscript come from preprocessed data, where the preprocessing was  
23 performed using Freesurfer (Dale et al., 1999), fMRIPprep 1.2.6-1 (Esteban et al. (2018); Esteban  
24 et al. (2019); RRID:SCR\_016216), which is based on Nipype 1.1.7 (Gorgolewski et al. (2017);  
25 Gorgolewski et al. (2011); RRID:SCR\_002502), as well as xcpEngine 1.0 (Ciric et al., 2018). Brain  
26 surfaces were reconstructed using recon-all (Dale et al., 1999) prior to other processing, and  
27 reconstructed surfaces were used as input to fMRIprep.

28 The T1-weighted (T1w) image was corrected for intensity non-uniformity (INU) using  
29 N4BiasFieldCorrection (Tustison et al. (2010), ANTs 2.2.0), and used as T1w-reference throughout

1 the workflow. The T1w-reference was then skull-stripped using antsBrainExtraction.sh (ANTs  
2 2.2.0), using OASIS as the target template. The brain mask was refined with a custom variation of  
3 the method to reconcile ANTs-derived and FreeSurfer-derived segmentations of the cortical gray-  
4 matter of Mindboggle (RRID:SCR\_002438, Klein et al. (2017)). Spatial normalization to the ICBM  
5 152 Nonlinear Asymmetrical template version 2009c (Fonov et al. (2009), RRID:SCR\_008796) was  
6 performed through nonlinear registration with antsRegistration (ANTs 2.2.0, RRID:SCR\_004757,  
7 Avants et al. (2010)), using brain-extracted versions of both T1w volume and template. Brain  
8 tissue segmentation of cerebrospinal fluid (CSF), white matter (WM) and gray matter was  
9 performed on the brain-extracted T1w using fast (FSL 5.0.9, RRID:SCR\_002823, Zhang et al.  
10 (2001)).

11 For each of the resting-state BOLD runs found per subject, the following preprocessing was  
12 performed: A reference volume and its skull-stripped version were generated using a custom  
13 methodology of fMRIPrep. The BOLD reference was then co-registered to the T1w reference using  
14 bbregister (FreeSurfer) which implements boundary-based registration (Greve and Fischl, 2009).  
15 Co-registration was configured with nine degrees of freedom to account for distortions remaining  
16 in the BOLD reference. Head-motion parameters with respect to the BOLD reference  
17 (transformation matrices, and six corresponding rotation and translation parameters) were  
18 estimated before any spatiotemporal filtering using mcflirt (FSL 5.0.9, Jenkinson et al. (2002)).  
19 BOLD runs were slice-time corrected using 3dTshift from AFNI 20160207 (Cox and Hyde (1997),  
20 RRID:SCR\_005927). The BOLD time-series were resampled onto MNI152Nlin2009cAsym  
21 standard space by applying a single, composite transform, generating a preprocessed BOLD run in  
22 MNI152Nlin2009cAsym space.

23 Several confounding time-series were calculated based on the preprocessed BOLD: framewise  
24 displacement (FD), DVARS and three region-wise global signals (CSF, WM, and the whole-brain).  
25 FD and DVARS were calculated for each functional run, both using their implementations in  
26 Nipype (following the definitions by Power et al. (2014b)). The head-motion estimates calculated  
27 in the correction step were also placed within the corresponding confounds file.

28 All resamplings can be performed with a single interpolation step by composing all the pertinent  
29 transformations (i.e. head-motion transform matrices and co-registrations to anatomical and

1 template spaces). Gridded (volumetric) resamplings were performed using antsApplyTransforms  
2 (ANTs), configured with Lanczos interpolation to minimize the smoothing effects of other kernels  
3 (Lanczos, 1964).

4 Further preprocessing was performed using a confound regression procedure that has been  
5 optimized to reduce the influence of participant motion (Satterthwaite et al., 2013b; Ciric et al.,  
6 2017; Parkes et al., 2018); preprocessing was implemented in XCP Engine 1.0, a multi-modal  
7 toolkit that deploys processing instruments from frequently used software libraries, including FSL  
8 (Jenkinson et al., 2012) and AFNI (Cox, 1996). Further documentation is available at  
9 <https://xcpengine.readthedocs.io> and <https://github.com/PennBBL/xcpEngine>. Functional  
10 timeseries were demeaned, and linear and quadratic trends were removed. Confound regression  
11 was performed using a 36-parameter model; confounds included mean signal from the whole  
12 brain, WM, and CSF compartments, 6 motion parameters as well as their temporal derivatives,  
13 quadratic terms, and the temporal derivatives of the quadratic terms (Satterthwaite et al., 2013a).  
14 Motion censoring was applied, with outlier volumes exceeding FD = 0.5 mm or standardized  
15 DVARS = 1.75 flagged and removed from confound regression. Outlier volumes were interpolated  
16 over using least squares spectral analysis (Power et al., 2014a) prior to band-pass filtering to  
17 retain frequencies between 0.01 Hz and 0.08 Hz, then re-censored. Prior to confound regression,  
18 all confound parameters were band-pass filtered in a fashion identical to that applied to the  
19 original timeseries data, ensuring comparability of the signals in frequency content (Hallquist et  
20 al., 2013).

### 21 ***Image quality and exclusion criteria***

22 The quality of imaging data was assessed using fMRIPrep's visual reports and MRIQC 0.14.2  
23 (Esteban et al., 2017). Two raters manually examined all structural and functional images  
24 between preprocessing steps for image quality issues. Functional images were visually inspected  
25 for whole-brain field of view coverage, signal blurring or artifacts, and proper alignment to the  
26 anatomical image. Participants were excluded for: unusable anatomical image ( $n = 1$ ), artifact in  
27 functional data (due to hair glitter,  $n = 1$ ), incorrect registration at the scanner ( $n = 1$ ), average  
28 FD greater than 1 mm ( $n = 14$ ), more than 30% of resting-state frames exceeding FD > 0.5 mm  
29 ( $n = 8$ , Power et al. (2012)). All participants that were flagged for dropout or signal blurring were

1 ultimately excluded for not meeting motion criteria. For participants with more than one usable  
2 resting-state run, FD was averaged across runs, weighted by run length. All analyses controlled for  
3 average FD and total number of resting-state frames.

4 To ensure that our results were not driven by motion, we conducted an additional analysis with a  
5 more stringent preprocessing pipeline and motion exclusion criteria. In this pipeline, motion  
6 censoring was applied with a threshold for outlier volumes of  $FD > 0.25$  mm or standardized  
7 DVARS  $> 1.75$ . One participant was lost during preprocessing as they did not have adequate  
8 degrees of freedom. Additionally, we excluded participants who had average FD greater than 0.5  
9 mm ( $n = 8$  additional participants), for a total of  $n = 83$  total participants.

### 10 ***Functional network analysis***

11 After preprocessing and nuisance regression, we extracted residual mean BOLD time series from a  
12 400-region cortical parcellation (Schaefer et al., 2018), and represented the functional  
13 connectivity matrix as a graph or network (Bassett et al., 2018). To evaluate whether our results  
14 were dependent on specific node definitions, we also extracted residual mean BOLD time series  
15 from a 200-region cortical parcellation (Schaefer et al., 2018). Results were qualitatively similar  
16 between the two parcellations (see online at  
17 [https://github.com/utooley/Tooley\\_2022\\_childhood\\_functional\\_network\\_dev](https://github.com/utooley/Tooley_2022_childhood_functional_network_dev)).

18 We assigned regions, or nodes, to systems based on a 7-system partition (Yeo et al., 2011), or  
19 assignment of nodes to systems. Here, we use the term *system* to refer to a set of regions  
20 previously defined *a priori* (i.e. the dorsal attention system, comprising a set of regions), while we  
21 use the term *network* to refer to the representation of the functional connectivity matrix as a  
22 graph. Regions were represented by network nodes, and the functional connectivity between  
23 region  $i$  and region  $j$  was represented by the network edge between node  $i$  and node  $j$ . We used  
24 this encoding of the data as a network to produce an undirected, signed, and weighted adjacency  
25 matrix  $A$ . We estimated the functional connectivity between any two brain regions by calculating  
26 the product-moment correlation coefficient  $r$  between the mean activity time series of region  $i$   
27 and the mean activity time series of region  $j$  (Zalesky et al., 2012). Correlations were subsequently  
28  $r$ -to- $z$ -transformed.

1 Recent evidence has demonstrated that the maintenance of edge weights is critical for an accurate  
2 understanding of the underlying biology of neural systems (Cole et al., 2012; Bassett and  
3 Bullmore, 2017), and work in applied mathematics has demonstrated that graph-related  
4 calculations are markedly more robust in weighted graphs than in binary graphs (Good et al.,  
5 2010). In light of these two lines of evidence and recent work in the field developing methods  
6 sensitive to the topologies present in weak versus strong edges (Rubinov and Sporns, 2011), we  
7 maintained all edge weights without thresholding and studied the full graph including both  
8 positive and negative correlations (Bassett et al., 2012; Santarnecchi et al., 2014). Functional  
9 connectivity matrices were averaged across runs for each participant, weighted by the number of  
10 frames in each run passing the quality threshold.

11 Across the cortex, we calculated the following summary functional network measures. *System*  
12 *segregation* is a measure of segregation that quantifies the difference between mean within-  
13 system connectivity and mean between-system connectivity as a proportion of mean within-  
14 system connectivity (Chan et al., 2014; Wig, 2017), given an *a priori* partition of nodes into  
15 systems, in this case the 7-system partition referenced earlier (Yeo et al., 2011). *Modularity*,  
16 quantified by the modularity quality index ( $Q$ ), is a measure of mesoscale network segregation  
17 that estimates the extent to which a network's nodes can be subdivided into groups or "modules"  
18 characterized by strong, dense intramodular connectivity and weak, sparse intermodular  
19 connectivity. Our approach is built on the modularity quality function originally defined in  
20 Newman (2006). Unlike system segregation, the modularity quality index is independent of a  
21 mapping of nodes to functional systems. Higher modularity is indicative of a more highly  
22 segregated network at the mesoscale. The *clustering coefficient* is a measure of local segregation  
23 that quantifies the amount of connectivity between a node and its strongest neighbors (Achard et  
24 al., 2006; Bartolomei et al., 2006; Bassett et al., 2006; Xu et al., 2016a). A node has a high clustering  
25 coefficient when a high proportion of its neighbors are also strong neighbors of each other. The  
26 *participation coefficient* quantifies the diversity of a node's connections across systems (Guimer'a  
27 and Nunes Amaral, 2005; Rubinov and Sporns, 2010). A node has a high participation coefficient  
28 when it is evenly and strongly connected to many different systems. A lower participation  
29 coefficient is indicative of a more highly segregated network. We specifically chose measures of  
30 functional network topology that were suitable for weighted, signed networks, when possible.

## 1 **System segregation**

2 System segregation quantifies the difference in mean within-system connectivity and mean  
3 between-system connectivity as a proportion of within-system connectivity. Previous work has  
4 linked this measure to aging-related changes in brain networks and poorer cognitive ability across  
5 age (Chan et al., 2014). In these analyses, we define system segregation as in (Chan et al., 2014),  
6 as:

$$7 \quad \frac{\bar{a}_{within} - \bar{a}_{between}}{\bar{a}_{within}}$$

8 Where  $\bar{a}_{within}$  is the mean edge weight between nodes within the same system and  $\bar{a}_{between}$  is the  
9 mean edge weight between nodes of one system to all nodes in other systems. We assigned nodes  
10 to systems based on a 7-system partition (Yeo et al., 2011). Freely available MATLAB code from  
11 [https://github.com/mychan24/system\\_matrix\\_tools](https://github.com/mychan24/system_matrix_tools) was used to calculate system segregation.

## 12 **Modularity quality index**

13 Statistics that quantify the modular structure of a network assess the extent to which a network's  
14 nodes can be subdivided into groups or modules characterized by strong, dense intramodular  
15 connectivity and weak, sparse intermodular connectivity. We considered the most commonly  
16 studied mesoscale organization—assortative community structure—that is commonly assessed  
17 by maximizing a modularity quality function (Porter et al., 2009; Fortunato, 2010). Our approach  
18 is built on the modularity quality function originally defined by Newman (Newman, 2006) and  
19 subsequently extended to weighted and signed networks by various groups.

20 Specifically, we follow Rubinov and Sporns (2011) by first letting the weight of a positive  
21 connection between nodes  $i$  and  $j$  be given by  $a_{ij}^+$ , the weight of a negative connection between  
22 nodes  $i$  and  $j$  be given by  $a_{ij}^-$ , and the strength of a node  $i$ ,  $s_i^\pm = \sum_j a_{ij}^\pm$ , be given by the sum of the  
23 positive or negative  $j$  connection weights of  $i$ . We denote the chance expected within-module  
24 connection weights as  $e_{ij}^+$  for positive weights and  $e_{ij}^-$  for negative weights, where  $e_{ij}^\pm = \frac{s_i^\pm s_j^\pm}{v^\pm}$ . We  
25 let the total weight  $v^\pm = \sum_{ij} a_{ij}^\pm$  be the sum of all positive or negative connection weights in the  
26 network. Then the asymmetric generalization of the modularity quality index is given by:

$$Q^* = \frac{1}{v^+} \sum_{ij} (a_{ij}^+ - e_{ij}^+) \delta_{M_i M_j} - \frac{1}{v^+ + v^-} \sum_{ij} (a_{ij}^- - e_{ij}^-) \delta_{M_i M_j}$$

where  $M_i$  is the community to which node  $i$  is assigned, and  $M_j$  is the community to which node  $j$  is assigned. We use a Louvain-like locally greedy algorithm as a heuristic to maximize this modularity quality index subject to a partition  $M$  of nodes into communities. We ran the Louvain algorithm 100 times per network, and detected on average 3 ( $M=3.44$ ,  $SD=0.483$ ) communities using modularity maximization in our developmental sample.

### Clustering coefficient

To assess local network segregation, we used a commonly studied graph measure of local connectivity—the clustering coefficient—that is commonly interpreted as reflecting the capacity of the system for processing within the immediate neighborhood of a given network node (Achard et al., 2006; Bartolomei et al., 2006; Bassett and Bullmore, 2006; Xu et al., 2016b). We specifically used a formulation that was recently generalized to signed weighted networks (Zhang and Horvath, 2005; Costantini and Perugini, 2014). This version is sensitive to nonredundancy in path information based on edge sign as well as edge weight and importantly distinguishes between positive triangles and negative triangles, which have distinct meanings in networks constructed from correlation matrices.

Let the functional connectivity network of a single participant be represented as the graph  $G = (V, E)$ , where  $V$  and  $E$  are the vertex and edge sets, respectively. Let  $a_{ij}$  be the weight associated with the edge  $(i, j) \in E$ , and define the weighted adjacency matrix of  $G$  as  $A = [a_{ij}]$ . The clustering coefficient of node  $i$  with neighbors  $j$  and  $q$  is given by

$$C_i = \frac{\sum_{jq} (a_{ji} a_{iq} a_{jq})}{\sum_{j \neq q} |a_{ji} a_{iq}|}$$

The clustering coefficient of the entire network was calculated as the average of the clustering coefficient across all nodes as follows:

$$C = \frac{1}{n} \sum_{i \in N} C_i$$

1 In this way, we obtained estimates of the regional and global clustering coefficient for each subject  
2 in the sample.

### 3 ***Participation coefficient***

4 The participation coefficient is a measure of network integration that quantifies the diversity of a  
5 node's connections across communities, and has been linked in older children and adolescents to  
6 developmental changes in network segregation (Marek et al., 2015; Baum et al., 2017). In these  
7 analyses, we define the participation coefficient  $P_i$  of a node  $i$  as:

$$8 \quad P_i = 1 - \sum_{k \in K} \left( \frac{a_{ik}}{s_i} \right)$$

9 where  $k$  is a system in a set  $K$  of systems, in this case defined by the a priori mapping of nodes to  
10 intrinsic functional systems (Yeo et al., 2011),  $a_{ik}$  is the positive (negative) weight of edges  
11 between node  $i$  and nodes in system  $k$ , and  $s_i$  is the positive (negative) strength of node  $i$ . The  
12 participation coefficient was calculated separately on negative and positive weights (Rubinov and  
13 Sporns, 2010).

14 As in our analyses of local segregation, the participation coefficient of the entire network was  
15 calculated as the average positive (negative) participation coefficient across all nodes as follows:

$$16 \quad P = \frac{1}{n} \sum_{i \in N} P_i$$

17 The average positive and negative participation coefficient for each participant's network were  
18 averaged to obtain a global measure of network integration.

### 19 ***System connectivity***

20 Within- and between-system connectivity were estimated as the average connectivity between  
21 nodes within a functional system or between pairs of functional systems. Results were corrected  
22 for multiple comparisons using the Benjamini-Hochberg false discovery rate (FDR).



## 1 ***Parcel-level connectivity***

2 When examining results at the parcel resolution, we applied a similar model as that at the whole-  
3 brain and system level across all 79,800 edges in each child's functional brain network. As  
4 correction for multiple comparisons in this situation raises the risk of missing true effects, we  
5 alternatively employed a stringent significance threshold for display of edge-level data ( $p <$   
6  $0.001$ ). Data are presented at  $p < .01$  and at  $p < .00001$  online at  
7 [https://github.com/utooley/Tooley\\_2022\\_childhood\\_functional\\_network\\_dev](https://github.com/utooley/Tooley_2022_childhood_functional_network_dev).

## 8 ***Statistical models***

9 All statistical analyses were conducted in MATLAB R2018a and R 3.6.1 [MATLAB:2018a; R Core  
10 Team (2013)]; code is publicly available at  
11 [https://github.com/utooley/Tooley\\_2022\\_childhood\\_functional\\_network\\_dev](https://github.com/utooley/Tooley_2022_childhood_functional_network_dev). We examined  
12 effects of age using generalized additive models with the *mgcv* package in R (Wood, 2011;  
13 Satterthwaite et al., 2014). We first tested for nonlinear effects of age. The penalty parameters for  
14 the nonlinear spline terms were fit as random effects and tested using restricted likelihood ratio  
15 tests (RLRTs) with *RLRsim* (Scheipl et al., 2008). Note that these tests of nonlinearity are  
16 constructed so as to test for nonlinear effects over and above any linear effects that may be  
17 present. We did not observe significant nonlinear relationships between age and whole-brain or  
18 system-level measures of network structure. 7.9% of edges showed significant nonlinear effects of  
19 age, as compared to 12.5% of edges that showed linear effects. Nonlinear effects at the parcel level  
20 are presented online at  
21 [https://github.com/utooley/Tooley\\_2022\\_childhood\\_functional\\_network\\_dev](https://github.com/utooley/Tooley_2022_childhood_functional_network_dev).

22 We modeled the linear effect of age while controlling for in-scanner motion (average FD), sex  
23 (male or female), total number of volumes across runs, and average functional network weight.  
24 Average network weight was included to control for global differences in connectivity strength  
25 (Van Wijk et al., 2010; Ginestet et al., 2011; Yan et al., 2013). Multiple comparisons correction was  
26 applied across models at the parcel and system resolutions using Benjamini-Hochberg false  
27 discovery rate (FDR) correction (Benjamini and Hochberg, 1995). Surfaces and partitions were  
28 shown on cortical surfaces generated by Freesurfer (Dale et al., 1999), using *fsbrain* 0.4.2 and  
29 *freesurfer-formats* 0.1.14 (Schäfer and Ecker, 2020).

## 1 ***Measurement and analyses of visuospatial reasoning ability***

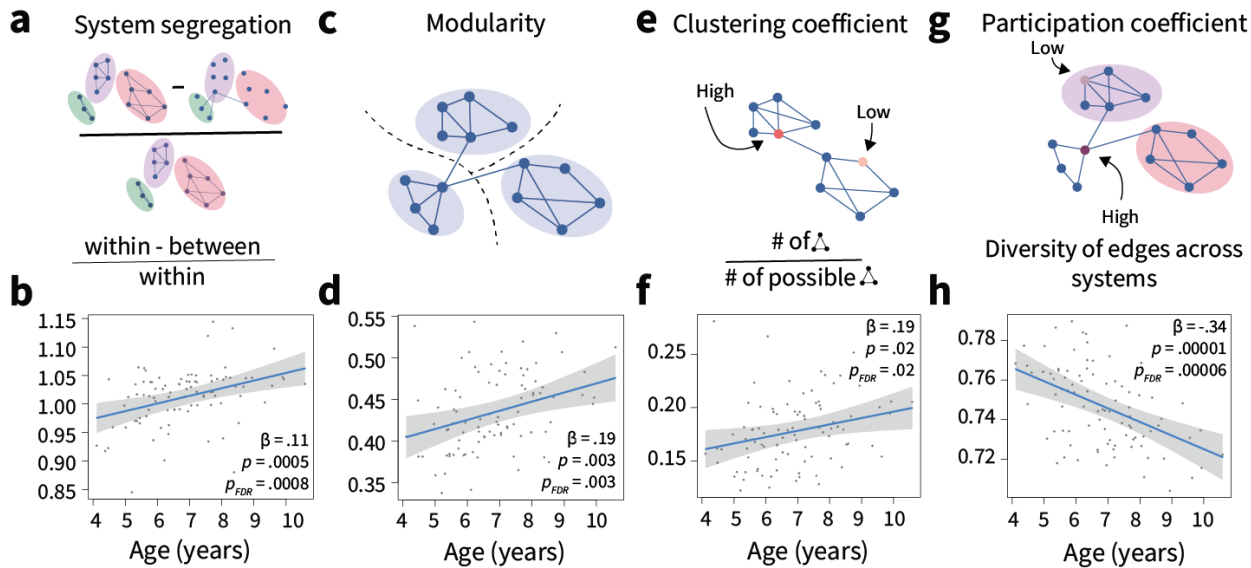
2 To assess reasoning, we administered matrix reasoning tests from Wechsler batteries. We used  
3 age-appropriate versions to avoid ceiling and floor effects. Children between the ages of 4 and 7  
4 years, 7 months completed the Matrix Reasoning subtest of the Wechsler Preschool & Primary  
5 Scale of Intelligence (WPPSI-IV, Wechsler and Corporation (2012);  $n = 63$ ). Children over age 7  
6 years, 7 months took the Matrix Reasoning subtest of the Wechsler Intelligence Scale for Children  
7 (WISC, Wechsler et al. (2014);  $n = 23$ ). Test items in both versions require identifying and  
8 integrating patterns in abstract shapes. For example, in Figure 4a, the foreground and background  
9 shapes switch across columns, and the shape type and color change across rows. To answer the  
10 question correctly, it is necessary to integrate these two relations. The WPPSI is normed down to  
11 2.5 years old so it begins with simpler items than the WISC. Therefore, raw scores on the WPPSI  
12 cannot be combined with raw scores on the WISC. Age was positively associated with raw scores  
13 on the WPPSI (mean raw score: 15.31, range 3-23, max possible score: 26,  $t(62) = 2.78, p = .007$ ).  
14 Age was not associated with raw scores on the WISC (mean raw score: 16.13, range 7-24, max  
15 possible score: 32,  $t(21) = 0.40, p = .694$ ). Scaled scores were used for all brain analyses. Models  
16 examining relationships between reasoning and system connectivity controlled for age, sex, in-  
17 scanner motion, total number of volumes across runs, and average functional network weight.  
18 Associations between system connectivity and reasoning ability were examined only for systems  
19 showing significant associations with age and the frontoparietal system (FDR-corrected for  
20 multiple comparisons across 5 systems).

## 21 **Results**

### 22 ***Functional network segregation increases with age***

23 We first investigated age effects on measures of whole-brain functional network segregation (Fig.  
24 1). Measures of functional network segregation were consistently positively associated with age,  
25 including average within-system connectivity ( $\beta = 0.3, t(86) = 3.75, p < .001, p_{FDR} = 0.0006$ ),  
26 average between-system connectivity ( $\beta = -0.06, t(86) = -3.75, p < .001, p_{FDR} = 0.0006$ ),  
27 overall system segregation ( $\beta = 0.11, t(86) = 3.60, p = .001, p_{FDR} = 0.0008$ ), the modularity  
28 quality index ( $\beta = 0.19, t(86) = 3.06, p = .003, p_{FDR} = 0.004$ ), and the clustering coefficient ( $\beta =$   
29  $0.19, t(86) = 2.35, p = .021, p_{FDR} = 0.021$ ). Consistent with these associations, we found that the

1 average participation coefficient, a measure that inversely tracks network segregation, was  
 2 negatively correlated with age ( $\beta = -0.35$ ,  $t(86) = -4.68$ ,  $p < .001$ ,  $p_{FDR} = 0.00006$ ).



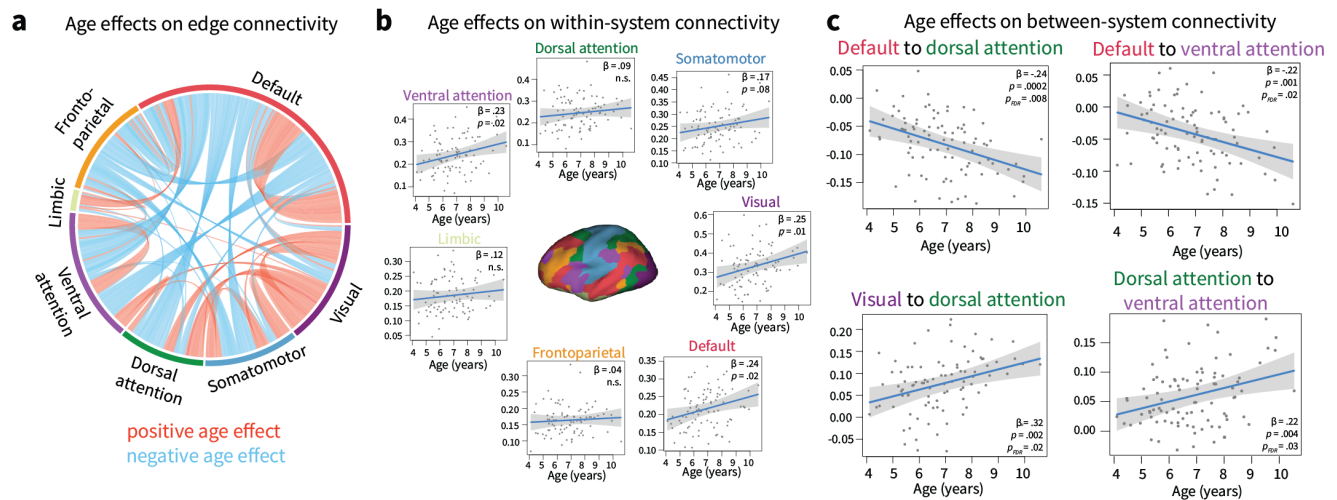
3

4 **Figure 1. Functional network segregation is positively associated with age.** **a.** System segregation is a whole-  
 5 brain measure of functional network segregation that quantifies the difference between mean within-system  
 6 connectivity and mean between-system connectivity as a proportion of mean within-system connectivity. **b.** System  
 7 segregation is positively associated with age. **c.** Modularity is a measure of mesoscale network segregation that  
 8 estimates the extent to which a network's nodes, or in this case brain regions, can be subdivided into modules  
 9 characterized by strong, dense intramodular connectivity and weak, sparse intermodular connectivity. Note that the  
 10 modules are data-driven, not a priori defined as functional systems. **d.** Modularity is positively associated with age. **e.**  
 11 The clustering coefficient is a measure of local segregation that quantifies the amount of connectivity between a node  
 12 and its neighbors. A node has a high clustering coefficient when a high proportion of its neighbors are also strongly  
 13 connected to one another. In a weighted network, the clustering coefficient measures the strength of triangles around  
 14 a node. **f.** The average clustering coefficient is positively associated with age. **g.** The participation coefficient quantifies  
 15 the diversity of a node's connections across systems. A node has a high participation coefficient when it is evenly  
 16 connected to many different systems. A lower participation coefficient is indicative of a more segregated network. **h.**  
 17 The average participation coefficient is negatively associated with age.

### 18 ***Systems specializing in perceptual processing segregate from systems for abstract thought***

19 We next tested for age effects at the system level by dividing the cortex into seven systems (Yeo et  
 20 al., 2011). We first visualized the balance of significant positive and negative age effects within and  
 21 between systems (Fig. 2a). Within systems, 94.6% of significant age effects were positive and 5.3%  
 22 were negative. Between systems, 27.7% of significant age effects were positive and 72.2% were

1 negative. Age was positively, but weakly, associated with within-system connectivity in the visual  
 2 (Fig. 2b,  $\beta = 0.25$ ,  $t(86) = 2.51$ ,  $p = .014$ ) and default mode systems ( $\beta = 0.24$ ,  $t(86) = 2.42$ ,  $p =$   
 3  $.017$ ), as well as in the ventral attention system ( $\beta = 0.24$ ,  $t(86) = 2.41$ ,  $p = .018$ ). The  
 4 significance of these associations did not survive correction for multiple comparisons. In contrast,  
 5 age was strongly associated with between-system connectivity (Fig. 2c). Age was negatively  
 6 associated with connectivity between the default mode and dorsal attention systems ( $\beta = -0.24$ ,  
 7  $t(86) = -3.79$ ,  $p < .001$ ,  $p_{FDR} = 0.01$ ), and connectivity between the default mode and ventral  
 8 attention systems ( $\beta = -0.22$ ,  $t(86) = -3.36$ ,  $p = .001$ ,  $p_{FDR} = 0.02$ ). Additionally, age was  
 9 positively correlated with connectivity between the visual and dorsal attention systems ( $\beta = 0.32$ ,  
 10  $t(86) = 3.15$ ,  $p = .002$ ,  $p_{FDR} = 0.02$ ) and with connectivity between the dorsal attention and  
 11 ventral attention systems ( $\beta = 0.22$ ,  $t(86) = 2.92$ ,  $p = .004$ ,  $p_{FDR} = 0.03$ ).



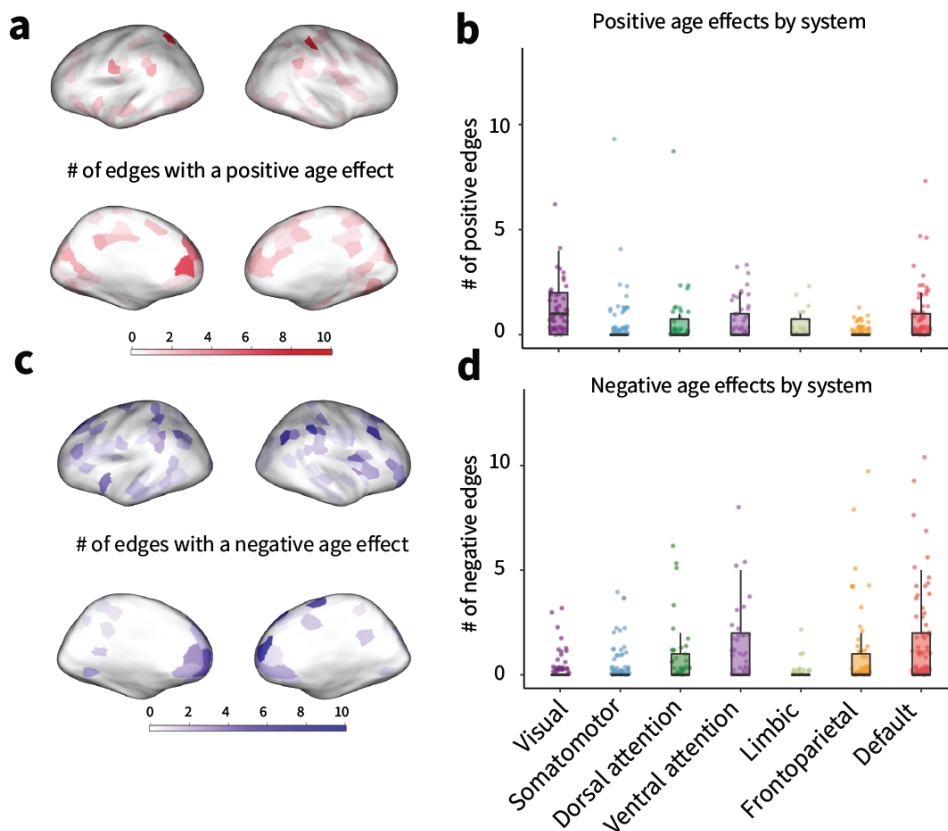
12

13 **Figure 2. System-level effects of age on system connectivity.** **a.** Age effects on edge connectivity. Note that only  
 14 edges with significant age effects at  $p_{unc} < 0.001$  are shown. **b.** Age effects on within-system connectivity. No  
 15 relationships survive FDR correction across systems. **c.** Age effects on between-system connectivity. All effects shown  
 16 survive FDR correction across systems.

17 ***Age effects are concentrated at both ends of the sensory-association gradient***

18 We next examined age effects at the parcel level to characterize regional specificity. In particular,  
 19 we determined which parcels had the most edges with significant age effects. Parcels with the  
 20 highest number of positive edge-level age effects were observed in the intraparietal sulcus (2  
 21 parcels with 9 significant edges), the medial prefrontal cortex (7 edges), and the occipital cortex (6

1 edges; see Fig. 3a). When parcels were grouped by system (Yeo et al., 2011), positive associations  
2 with age were most common in the visual system, followed by the default mode system, and the  
3 ventral attention system (Fig. 3b). Parcels with negative edge-level age effects were also  
4 concentrated in the medial prefrontal cortex and the intraparietal sulcus, but not in lower-level  
5 sensory or motor areas (Fig. 3c). Edge-level age effects were most pronounced in a medial  
6 prefrontal cortex parcel in the default mode system (top parcel: medial prefrontal cortex, 13  
7 negative age-associated edges). The top five non-anatomical meta-analytic associations on  
8 Neurosynth for the medial prefrontal cortex region (MNI coordinates of centroid:  $x = 8, y = 54, z =$   
9  $12$ ) were “mind,” “theory mind,” “autobiographical,” “mentalizing,” and “mental states”. Negative  
10 associations with age were most common in the default mode system and the ventral attention  
11 system, followed by the dorsal attention and frontoparietal systems. Very few negative  
12 associations were found in the visual, somatomotor, or limbic systems (Fig. 3d).



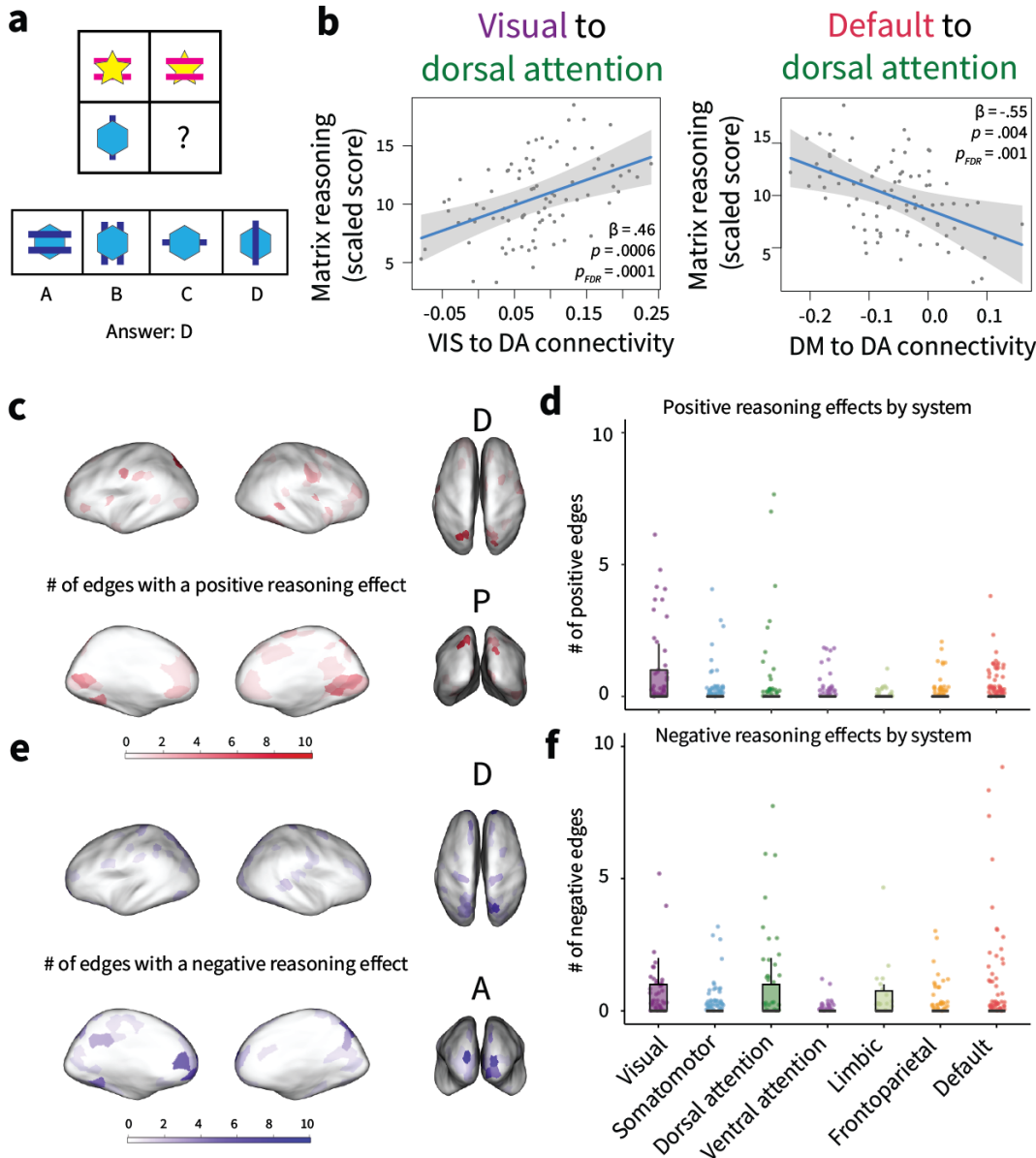
13

14 **Figure 3. Parcel-level effects of age on network connectivity.** **a.** Number of edges from each parcel showing a  
15 significant positive age association; significance was defined as  $p_{unc} < 0.001$ . **b.** Number of edges with positive effects  
16 of age, grouped by system. Each datapoint represents a parcel. **c.** Number of edges from each parcel showing a

1 significant negative age association; significance was defined as  $p_{unc} < 0.001$ . **d.** Negative edge effects for each parcel  
2 grouped by system. Each datapoint represents a parcel.

### 3 ***Functional network structure is associated with cognition***

4 Finally, we explored the cognitive consequences of age-associated network segregation by  
5 examining relationships between functional architecture and visuospatial reasoning (matrix  
6 reasoning from Wechsler tests (Wechsler and Corporation, 2012; Wechsler et al., 2014); see Fig.  
7 4a). Controlling for age, reasoning was positively associated with average within-system  
8 connectivity ( $\beta = 0.34$ ,  $F(1,79) = 4.78$ ,  $p = .032$ ,  $p_{FDR} = 0.08$ ), and negatively associated with  
9 average between-system connectivity ( $\beta = -1.57$ ,  $F(1,79) = 4.78$ ,  $p = .032$ ,  $p_{FDR} = 0.08$ ) and  
10 average participation coefficient ( $\beta = -0.35$ ,  $F(1,79) = 4.39$ ,  $p = .039$ ,  $p_{FDR} = 0.08$ ). However,  
11 these associations did not pass correction for multiple comparisons, and reasoning was not  
12 associated with other measures of whole-brain network architecture ( $p$ -values  $> .05$ ). At the  
13 system level, we focused on pairs of cognitive systems that show significant associations with age  
14 (see Fig. 2c), and found that connectivity between the visual and dorsal attention systems was  
15 positively associated with reasoning ability (Fig. 4b,  $\beta = 0.46$ ,  $t(79) = 4.06$ ,  $p < .001$ ,  $p_{FDR} =$   
16  $0.0006$ ). We also found that connectivity between the default and dorsal attention systems was  
17 negatively associated with reasoning ability (Fig. 4b,  $\beta = -0.55$ ,  $t(79) = -2.93$ ,  $p = .004$ ,  $p_{FDR} =$   
18  $0.01$ ). Further, motivated by prior studies linking the frontoparietal system to reasoning (Prado et  
19 al., 2011; Wertheim and Ragni, 2018), we tested whether reasoning was associated with within-  
20 system frontoparietal connectivity; we found no effect ( $\beta = 0.07$ ,  $t(79) = 0.51$ ,  $p = .613$ ,  $p_{FDR} =$   
21  $0.68$ ). At the parcel level, connections with the intraparietal sulcus (top parcel: 8 edges), as well as  
22 the medial prefrontal and occipital areas, showed positive relationships with reasoning (Fig. 4c).  
23 Parcels with positive reasoning effects were most numerous in the visual system (Fig. 4d).  
24 Connections with the frontal pole (top parcel: 8 edges), the intraparietal sulcus, the medial  
25 prefrontal cortex, and visual areas showed negative associations with reasoning (Fig. 4e). Parcels  
26 with negative reasoning effects were most numerous in the visual, default mode, and dorsal  
27 attention systems (Fig. 4f).



1

2 **Figure 4. Associations between functional network structure and visuospatial reasoning.** a. Example reasoning  
 3 item. Reasoning was assessed with the Matrix Reasoning subscale of the Weschler assessments. b. System-level  
 4 associations with reasoning, controlling for age. Reasoning is associated with connectivity between the visual and  
 5 dorsal attention systems, and with connectivity between the default mode and dorsal attention systems. c. The  
 6 number of edges from each parcel showing a significant positive reasoning association; significance was defined as  
 7  $p_{unc} < 0.001$ . d. The number of edges with positive effects of reasoning, grouped by system. Each datapoint  
 8 represents a parcel. e. The number of edges from each parcel showing a significant negative reasoning association;  
 9 significance was defined as  $p_{unc} < 0.001$ . f. The number of edges with negative effects of reasoning, grouped by  
 10 system. Each datapoint represents a parcel.

## 1 ***Sensitivity analyses***

2 We conducted a set of sensitivity analyses to ensure that our results were not dependent on  
3 specific analytical choices. Specifically, we conducted our main analyses with a more stringent  
4 preprocessing pipeline and motion exclusion criteria. In this pipeline, measures of functional  
5 network segregation were consistently positively associated with age, including average within-  
6 system connectivity ( $\beta = 0.22, t(77) = 3.02, p = .003, p_{FDR} = 0.007$ ), average between-system  
7 connectivity ( $\beta = -0.04, t(77) = -3.02, p = .003, p_{FDR} = 0.007$ ), overall system segregation ( $\beta =$   
8  $0.1, t(77) = 2.53, p = .013, p_{FDR} = 0.01$ ), the modularity quality index ( $\beta = 0.23, t(77) = 3.74,$   
9  $p < .001, p_{FDR} = 0.002$ ), and the clustering coefficient ( $\beta = 0.21, t(77) = 2.67, p = .009, p_{FDR} =$   
10  $0.01$ ). Consistent with these associations, we found that the average participation coefficient, a  
11 measure that inversely tracks network segregation, was negatively correlated with age ( $\beta = -0.23,$   
12  $t(77) = -2.66, p = .009, p_{FDR} = 0.01$ , see Fig. 5a).

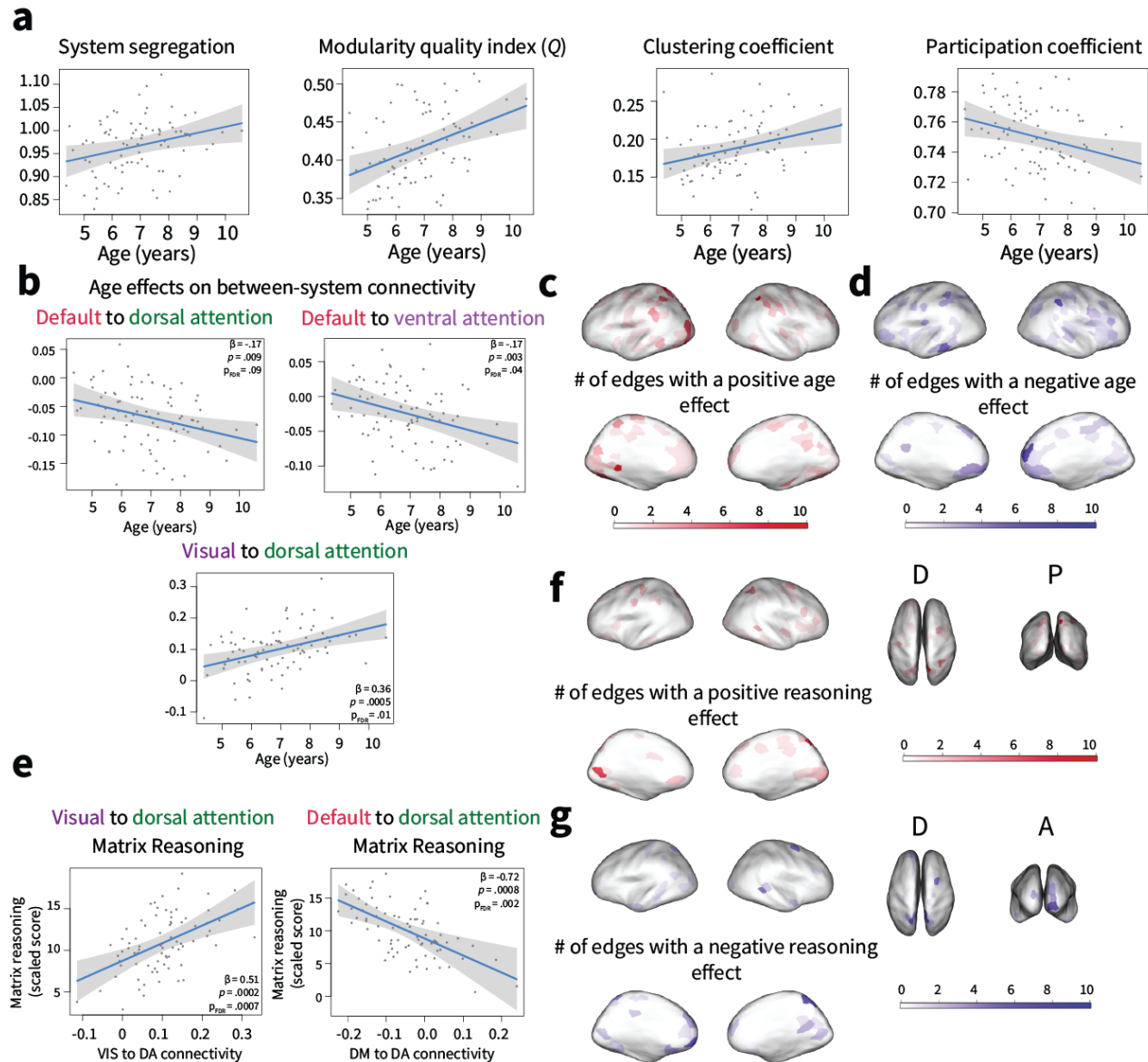
13 At the system resolution, age was positively, but weakly, associated with within-system  
14 connectivity in the visual ( $\beta = 0.25, t(77) = 2.21, p = .030$ ) and limbic ( $\beta = 0.2, t(77) = 2.05, p =$   
15  $.044$ ) systems, and was marginally positively associated with within-system connectivity in the  
16 default mode ( $\beta = 0.17, t(77) = 1.83, p = .071$ ) and somatomotor systems ( $\beta = 0.18, t(77) =$   
17  $1.93, p = .057$ ). None of these associations survived correction for multiple comparisons. In  
18 contrast, age was strongly associated with between-system connectivity (Fig. 5b). Age was  
19 negatively associated with connectivity between the default mode and ventral attention system  
20 ( $\beta = -0.17, t(77) = -3.09, p = .003, p_{FDR} = 0.04$ ). Age was also negatively associated with  
21 connectivity between the default mode and dorsal attention system ( $\beta = -0.17, t(77) = -2.66,$   
22  $p = .009, p_{FDR} = 0.09$ ), but this association was marginal after FDR correction. Additionally, age  
23 was positively correlated with connectivity between the visual and dorsal attention systems ( $\beta =$   
24  $0.36, t(77) = 3.64, p < .001, p_{FDR} = 0.01$ ).

25 We next examined age effects at the parcel level to characterize regional specificity. Parcels with  
26 the highest number of positive edge-level age effects were observed in the superior parietal  
27 lobule/intraparietal sulcus (2 parcels with 11 and 9 significant edges) and the occipital cortex (2  
28 parcels with 10 and 9 significant edges; see Fig. 5d). Parcels with the highest number of negative



1 edge-level age effects occurred in medial prefrontal cortex (8 edges) and intraparietal sulcus (7  
2 edges, Fig. 5e).

3 Finally, we examined relationships between functional architecture and visuospatial reasoning.  
4 Controlling for age, reasoning was marginally positively associated with average within-system  
5 connectivity ( $\beta = 0.32$ ,  $F(1,72) = 2.96$ ,  $p = .090$ ,  $p_{FDR} = 0.18$ ), and marginally negatively  
6 associated with average between-system connectivity ( $\beta = -1.62$ ,  $F(1,72) = 2.96$ ,  $p = .090$ ,  
7  $p_{FDR} = 0.18$ ). However, these associations did not pass correction for multiple comparisons, and  
8 reasoning was not associated with other measures of whole-brain network architecture ( $p$ -values  
9  $> .05$ ). At the system level, we focused on pairs of cognitive systems that showed associations with  
10 age in the main analyses, and found that connectivity between the visual and dorsal attention  
11 systems was positively associated with reasoning ability (Fig. 5e,  $\beta = 0.51$ ,  $t(72) = 4.00$ ,  $p < .001$ ,  
12  $p_{FDR} = 0.0008$ ). We also found that connectivity between the default and dorsal attention systems  
13 was negatively associated with reasoning ability ( $\beta = -0.72$ ,  $t(72) = -3.51$ ,  $p = .001$ ,  $p_{FDR} =$   
14  $0.0019$ ). Reasoning was not associated with within-system frontoparietal connectivity ( $\beta = 0.17$ ,  
15  $t(72) = 1.01$ ,  $p = .318$ ,  $p_{FDR} = 0.53$ ). At the parcel level, connections with superior parietal cortex  
16 (8 edges) and visual areas (7 edges) showed positive relationships with reasoning (Fig. 5f).  
17 Connections with the frontal pole (7 edges), superior parietal cortex (2 parcels with 7 and 6  
18 edges), medial prefrontal cortex (6 edges), and visual areas showed negative associations with  
19 reasoning (Fig. 5g).



1

2 **Figure 5. Replication with stricter motion exclusions.** In this pipeline, we censored volumes with FD > 0.25mm  
 3 and excluded participants with average FD > 0.5 mm. **a.** Whole-brain measures of functional network segregation  
 4 (system segregation, modularity, and the clustering coefficient) are positively associated with age. The participation  
 5 coefficient is a measure of functional network integration and is negatively associated with age. **b.** Age effects on  
 6 between-system connectivity. **c.** Number of edges from each parcel showing a significant positive age association;  
 7 significance was defined as  $p_{unc} < 0.001$ . **d.** Number of edges from each parcel showing a significant negative age  
 8 association; significance was defined as  $p_{unc} < 0.001$ . **e.** System-level associations with reasoning, controlling for age.  
 9 Reasoning is associated with connectivity between the visual and dorsal attention systems, and with connectivity  
 10 between the default mode and dorsal attention systems. **f.** The number of edges from each parcel showing a

1 significant positive reasoning association; significance was defined as  $p_{unc} < 0.001$ . **g.** The number of edges from each  
2 parcel showing a significant negative reasoning association; significance was defined as  $p_{unc} < 0.001$ .

### 3 **Discussion**

4 We investigated the development of cortical functional network architecture during childhood. At  
5 the whole-brain level, age was positively associated with multiple measures of functional network  
6 segregation, consistent with prior work on development later in childhood and adolescence (Fair  
7 et al., 2009; Marek et al., 2015; Lopez et al., 2019). At the system level, age was associated with a  
8 segregation of systems involved in attention from those involved in abstract, internally oriented  
9 cognition, as well as an integration among attentional and perceptual systems. At the parcel level,  
10 age effects on functional connectivity were strongest in medial prefrontal areas of the default  
11 mode system, and in areas of the visual system. Classically, brain development is thought to move  
12 from back to front, from sensory areas to association areas. Our results suggest another  
13 possibility: both ends of the sensory-association gradient are anchored early, perhaps by the  
14 presence or absence of sensory input, and then boundaries along the gradient are gradually  
15 solidified. This possibility is consistent with the very early emergence of the default mode network  
16 in utero and in infancy (Gao et al., 2009; Thomason et al., 2014; Gilmore et al., 2018; Hodel, 2018),  
17 and with work showing that medial prefrontal cortex, like primary sensory areas, is already highly  
18 segregated in adolescence (Baum et al., 2020). These findings fill a critical gap in our  
19 understanding of how intrinsic functional network remodeling supports the profound cognitive  
20 development that takes place during early and middle childhood.

21 Age effects were pronounced in areas of medial prefrontal cortex that are activated by self-  
22 referential thought and social perception tasks in adults (de la Vega et al., 2016; Meyer and  
23 Lieberman, 2018; Parelman et al., 2021). This result is consistent with evidence for major changes  
24 in social cognition between the ages of 3 and 10 years, supported by changes in the structure and  
25 function of the medial prefrontal cortex, the precuneus, and the temporoparietal junction (Weimer  
26 et al., 2021). Though we did not collect a behavioral or imaging measure of social cognition in this  
27 sample, we speculate that the medial prefrontal regions that show age effects may support  
28 improvements in social cognition in this age range. In this context, it is notable that medial  
29 prefrontal cortex continues developing after 10 years of age, and shows a protracted course of

1 age-associated change through adolescence and into adulthood (Baum et al., 2020). The age-  
2 associated remodeling we observe in medial prefrontal cortex may be simply an early  
3 manifestation of the ongoing anchoring of the far end of the sensorimotor-association gradient  
4 that continues into adulthood. It is also possible that changes in medial prefrontal connectivity  
5 more broadly support self-regulation processes that are required for efficiently completing most  
6 types of tasks (Akshoomoff et al., 2014; de la Vega et al., 2016; Meyer and Lieberman, 2018).

7 Age effects were also pronounced in the visual system. The visual network showed increased  
8 integration with the dorsal attention network, particularly along the dorsal stream. The majority  
9 of inputs into primary visual cortex come from higher-order visual areas and attention areas  
10 (Muckli and Petro, 2013), so it is possible that inputs from attention systems are reflected in the  
11 structure and function of perceptual areas. Indeed, attention improves substantially in early  
12 childhood (Amso and Scerif, 2015). Further, connectivity within regions of the dorsal attention  
13 and visual systems is positively associated with attention skills in 4-7-year-old children (Rohr et  
14 al., 2017, 2018), suggesting that the age effects we observe in regions of the visual system may  
15 also support developing attention skills.

16 Better reasoning abilities were associated with more mature patterns of brain network  
17 architecture, after controlling for age. At the parcel level, reasoning was associated with the  
18 connectivity of medial prefrontal and visual areas, as well as the intraparietal sulcus and the  
19 frontal pole. At the systems level, reasoning was associated with integration between the visual  
20 and dorsal attention systems, and with segregation between the default mode and dorsal attention  
21 systems. Prior work in older children and adults has linked structure and function of the  
22 frontoparietal system to reasoning skills, with a specific focus on rostral lateral prefrontal cortex  
23 and parietal areas (Prado et al., 2011; Vendetti and Bunge, 2014; Wertheim and Ragni, 2018).  
24 Interestingly, one study found that neural correlates of reasoning depended on age: after age 8  
25 years, stronger reasoning skills were associated with stronger functional connectivity between  
26 rostral lateral prefrontal cortices and the inferior parietal lobe, whereas before age 8 years, there  
27 were no such associations (Wendelken et al., 2016). Similarly, we found no association between  
28 frontoparietal system connectivity and reasoning ability in our age range. By taking a whole brain  
29 approach rather than focusing on the frontoparietal network, we found that visuospatial  
30 reasoning is associated with integration between perceptual and attentional systems in children.

1 We also found that reasoning was associated with segregation between task-positive and task-  
2 negative systems, consistent with other work across multiple age ranges and cognitive domains  
3 (Chan et al., 2014; Keller et al., 2015; Marek et al., 2019; Bruchhage et al., 2020). The involvement  
4 of perceptual systems such as the visual and dorsal attention systems in reasoning may not be as  
5 surprising as it initially seems: in children and adults, reasoning tasks engage visual areas more  
6 than non-reasoning control tasks (Soulières et al., 2009; Mackey et al., 2015; Whitaker et al.,  
7 2018). There is also evidence that reasoning performance relies more on lower-level skills like  
8 processing speed and visuospatial attention than on higher-level skills like working memory and  
9 relational integration early in childhood (Fry and Hale, 1996; Kail and Hulme, 2016). Broadly, our  
10 results suggest that maturation of brain network architecture, in particular in areas at two ends of  
11 the sensory-association gradient, supports the development of reasoning abilities.

12 Making decisions about motion criteria is difficult because of tradeoffs between data quality and  
13 generalizability, as motion is often highly correlated with other sample characteristics of interest  
14 (Hodgson et al., 2017; Leonard et al., 2017; Bolton et al., 2019). Our approach here was to analyze  
15 the data at two motion thresholds, a more lenient threshold that included more children and more  
16 data, and a more conservative threshold that minimized motion concerns. At both thresholds, the  
17 general pattern of findings was the same. At the whole-brain level, age was positively associated  
18 with measures of segregation. At the system level, age was positively associated with segregation  
19 between external and internal attention systems, and integration between attentional and  
20 perceptual systems. Although the specific parcel-level results were not identical, the broad pattern  
21 of results is similar, with age effects on functional connectivity strongest in medial prefrontal  
22 cortex, superior parietal cortex, and visual areas. Better reasoning abilities were associated with  
23 more mature patterns of brain network architecture. This suggests that our findings are robust  
24 and are not driven by motion.

25 Several potential limitations should be noted. First, our dataset is cross-sectional and of a  
26 relatively small sample size. Future work with longitudinal data will be necessary to establish the  
27 temporal sequence of the relationships we report, as well as to better evaluate nonlinearities and  
28 ideally, developmental trajectories in children younger than age 4 years. Longitudinal data would  
29 also make it possible to test whether changes in network structure mediate age-related  
30 improvements in reasoning. Fortunately, such a study — the HEALthy Brain and Cognitive

1 Development Study (Volkow et al., 2020) — is about to begin. Second, by carefully excluding data  
2 with motion artifacts, we may have limited the generalizability of our findings. Most young  
3 children move in the scanner, so it is essential to develop more motion-resilient sequences to  
4 allow investigators to acquire data in a more representative sample of young kids. Third, our  
5 cognitive measures were limited. Future work is necessary to link changes in functional  
6 organization to changes across a broader set of cognitive and social skills, including abilities that  
7 might diminish with age, such as creativity and imagination (Thompson-Schill et al., 2009; Gopnik,  
8 2020). Fourth, major cognitive and social changes during middle childhood (called the “age of  
9 reason” by anthropologists Sameroff and Haith (1996)) have been observed across many cultures  
10 all over the world, but our sample only captures development in our specific geographic and  
11 cultural context. Finally, we could not determine the causes of the developmental patterns we  
12 uncovered. More work is needed to understand whether these patterns were associated with  
13 specific experiences, for example formal schooling (Brod et al., 2017; Nolden et al., 2021), or  
14 simply reflect biological experience-independent maturation.

15 In sum, age effects on functional cortical architecture during childhood parallel long-known age  
16 effects on behavior. As children learn to resist the lure of perceptual information, and begin to  
17 reason abstractly, cortical systems for perception and abstraction separate, while connections that  
18 facilitate attention tend to strengthen. As children’s concept of self matures, the connectivity of the  
19 medial prefrontal cortex changes. Our results provide new insights into how changes in cortical  
20 organization give rise to changes in the mind as children reach the age of reason.

## 21 **Citation diversity statement**

22 Recent work in several fields of science—including neuroscience, where our work here is  
23 situated—has identified a bias in citation practices such that papers from women and other  
24 marginalized scholars are under-cited relative to the number of such papers in the field (Maliniak  
25 et al., 2013; Mitchell et al., 2013; Caplar et al., 2017; Dion et al., 2018; Dworkin et al., 2020;  
26 Chatterjee and Werner, 2021; Teich et al., 2021; Wang et al., 2021). Here we sought to proactively  
27 consider choosing references that reflect the diversity of the field in thought, form of contribution,  
28 gender, race, ethnicity, and other factors. First, we obtained the predicted gender of the first and  
29 last author of each reference by using databases that store the probability of a first name being

1 carried by a woman (Dworkin et al., 2020; Zhou et al., 2020). By this measure (and excluding self-  
2 citations to the first and last authors of our current paper), our references contain 26%  
3 woman(first)/woman(last), 16% man/woman, 20% woman/man, and 38% man/man  
4 categorization. This method is limited in that a) names, pronouns, and social media profiles used  
5 to construct the databases may not, in every case, be indicative of gender identity and b) it cannot  
6 account for intersex, non-binary, or transgender people. Second, we obtained the predicted  
7 racial/ethnic category of the first and last author of each reference by databases that store the  
8 probability of a first and last name being carried by an author of color (Ambekar et al., 2009; Sood  
9 and Laohaprapanon, 2018). By this measure (and excluding self-citations), our references contain  
10 9.83% author of color (first)/author of color(last), 16.45% white author/author of color, 18.51%  
11 author of color/white author, and 55.21% white author/white author. This method is limited in  
12 that a) names and Florida Voter Data to make the predictions may not be indicative of  
13 racial/ethnic identity, and b) it cannot account for Indigenous and mixed-race authors, or those  
14 who may face differential biases due to the ambiguous racialization or ethnicization of their  
15 names. We look forward to future work that could help us to better understand how to support  
16 equitable practices in science.

## 17 **Author contributions**

18 U.A. Tooley, D.S. Bassett, and A.P. Mackey developed the study concept. All authors contributed to  
19 the study design. Testing and data collection were performed by U.A. Tooley, A.T. Park, J.A.  
20 Leonard, A.L. Boroshok, and C.L. McDermott. U.A. Tooley performed the data analysis and  
21 interpretation under the supervision of A.P. Mackey and D.S. Bassett. U.A. Tooley and A.P. Mackey  
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