1	Ageing in a collective: The impact of ageing individuals on social network structure
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24	Abstract
25 26	Ageing affects many phenotypic traits, but its consequences for social behaviour have
27	only recently become apparent. Social networks emerge from associations between individuals.
28	The changes in sociality that occur as individuals get older are thus likely to impact network
29	structure, yet this remains unstudied. Here we use empirical data from free-ranging rhesus
30	macaques and an agent-based model to test how age-based changes in social behaviour feed up
31	to influence: (1) an individual's level of indirect connectedness in their network; and (2) overall
32	patterns of network structure. Our empirical analyses revealed that female macaques became less
33	indirectly connected as they aged for some, but not all network measures examined, suggesting
34	that indirect connectivity is affected by ageing, and that ageing animals can remain well

integrated in some social contexts. Surprisingly, we did not find evidence for a relationship
between age distribution and the structure of female macaque networks. We used an agent-based
model to gain further understanding of the link between age-based differences in sociality and
global network structure, and under which circumstances global effects may be detectable.
Overall, our results suggest a potentially important and underappreciated role of age in the
structure and function of animal collectives, which warrants further investigation.

- 41
- 42 Introduction
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The costs and benefits of an animal's behaviour toward members of their collectives can 44 45 shape lifespans, life history evolution, and the pace of senescence. For example, the support that 46 cooperative breeding species receive from non-breeding helpers may lead to longer lifespans than those of related solitary species [1]. The transfer of enhanced ecological knowledge to close 47 48 relatives during collective foraging may have contributed to the evolution of reproductive 49 senescence and prolonged post-reproductive lifespan (i.e., menopause) in female killer whales [2– 4]. The formation of strong relationships between group mates has also been linked to enhanced 50 51 longevity in a range of group-living taxa [5]. It is therefore clear that living in collectives can alter 52 the adult ageing process. However, an important question is whether ageing, in turn, influences 53 the behaviour of adults in collectives, and ultimately the structure of collectives themselves.

54 Growing evidence suggests that older adults differ in their social behaviours and social 55 relationships from young adults [6–11]. One pattern that seems to be emerging across taxa is that 56 older adults interact with fewer individuals than do younger adults, concentrating social 57 relationships on close associates and kin [6,12–15]. Given that social networks are an emergent 58 feature of association rules between individuals [16], shifts in patterns of social behaviour with 59 age might not only affect who ageing individuals associate with directly (i.e., their direct 60 connectedness), but could also affect higher-order network structure. Age-based changes in social 61 behaviour could scale up to alter an individual's connections to the partners of their social partners 62 (their indirect connectedness [17]) as well as the overall structure ("topology") of the social 63 network, both of which can have consequences for disease transmission [18–23], information 64 transfer [24–26], the cohesive movement of groups [27,28], and many other eco-evolutionary 65 dynamics [29,30]. Yet little attention has been given to understanding the impact of social ageing 66 for the polyadic social world or the structure of the collective.

67 Understanding how ageing shapes an individual's indirect connectedness may be 68 particularly relevant as such connections are tightly linked to processes that can directly influence fitness, including those described above. Declines in indirect connectedness with age 69 70 may help limit exposure to disease [20-22] which might be beneficial in aged animals 71 experiencing immunosenescence, but simultaneously could inhibit the transfer of important 72 socio-ecological information [26,31] which could exacerbate pre-existing fitness declines in old 73 age. In some cases, indirect connectivity may be an even more important predictor of fitness than 74 direct connections dyadic associations [17,32]. Therefore, changes in indirect connections may 75 be a particularly important component of the social ageing phenotype to investigate. Recent work 76 has offered some glimpses into how measures of indirect connectedness can differ between 77 young and old adults. In marmots and Barbary macaques, older adults have measures of indirect 78 connectedness that suggest they have partners who are themselves not well connected [9,33], 79 (but this is not the case in rhesus macaques, see [34,35]). Older marmots and rhesus macaques 80 are less effective at reaching disparate nodes in the network compared to younger adults [34–36]. 81 In contrast, older adults are more strongly embedded in cliques or clusters in their networks 82 compared to younger adults in marmots [9], but not in either Barbary or rhesus macaques

83 [33,35]. However, most research to date has compared differences in measures of indirect 84 connectedness among adults of different age classes (e.g., old versus young), but has lacked the 85 longitudinal data required to quantify how the social positioning of individuals changes across 86 their lifetimes (c.f., [9]). Such longitudinal analyses have the potential to reveal important 87 patterns that might otherwise be masked by differences between individuals or cohorts [37]. 88 Tracking within-individual changes in measures of indirect connectedness is essential for more 89 firmly placing changes in sociality across the lifespan within the larger ageing phenotype and 90 therefore understanding the causes and consequences of these patterns of social ageing [38]. 91 Populations composed of a greater proportion of older (or younger) adults may also be 92 structured in meaningfully different ways, and this could affect important processes such as 93 communication and cooperation. For example, the loss of old individuals through age-related 94 disease or trophy hunting can disrupt intergenerational flow of accumulated social and ecological 95 knowledge, impeding collective movement and the ability to locate critical resources [4,39–42]. 96 The age structure of a group can also regulate the behaviour of younger individuals [43,44], 97 influencing aggression rates and social cohesion [45]. The impact of diminished cohorts of 98 younger individuals on overall network structure is less well understood, but likely to have 99 repercussions for network connectivity and cohesiveness given that younger adults are more 100 socially active in many populations [8,10,12-14]. For example, the simulated removal of 101 juvenile killer whales led to networks that were more fragmented than when random individuals 102 were removed, suggesting an important role of young individuals in maintaining network 103 cohesion [46]. Despite the established ecological and evolutionary importance of network structure [29,30] the underlying drivers of variation in network structure remain understudied 104 105 [47–50]. Ageing, as an important process underlying patterns of individual-level variation in

sociality, might therefore provide a window into how simple processes can generate complexnetwork structures [51,52].

108 Using both empirical data and a theoretical model, we explore how social ageing of 109 individuals relates to measures of indirect connectedness and overall network structure in a 110 group-living primate, the rhesus macaque (Macaca mulatta), which is an emerging model in 111 social ageing research [38,53]. As female rhesus macaques grow older, they show clear changes 112 in their patterns of direct connectedness: they reduce the size of their social networks and focus 113 their social effort on a few important partners including close kin [6]. Despite this, females do 114 not reduce the rate at which they engage in social interactions as they age, indicating that 115 although their networks get smaller, older females continue to invest the same amount of time 116 into fewer relationships [6].

117 Given these previously established changes in direct connections, here, we set out to test 118 if age-based changes in social behaviour relate to measures of an individual's level of indirect 119 connectedness [17,54], and if they scale up to influence network structure as a whole. Individual-120 level social network metrics, including measures of direct and indirect connectedness, relate to 121 underlying, putatively simple, social differences or processes, such as individual-level variation 122 in general sociability or reassociation tendency [51]. For example, an individual's general 123 tendency to be sociable can be intuitively quantified using a direct network metric (e.g., strength, 124 or the sum of the weights of an individual's ties to their partners), or using an indirect metric 125 (e.g., weighted eigenvector centrality, which measures how well-connected an individual is to its 126 partners and how well connected those partners are to others [51]). Drawing on this idea that 127 underlying processes can predict an individual's position in the network, and on the age-based 128 changes in direct connections we have previously documented in rhesus macaque females [6],

129 we made predictions for four common measures of indirect connectedness: eigenvector 130 centrality, betweenness, closeness, and clustering coefficient (Table 1). Adult female rhesus 131 macaques maintain their strength of grooming ties to others as they age, despite reducing their 132 number of partners [6]. We therefore predicted that weighted eigenvector centrality (a measure 133 of overall connectedness in the network) would also remain stable with age, as by retaining 134 strong connections to some of their partners, females therefore (indirectly) retain connections to 135 the partners of their partners. We also previously found that as female rhesus macaques age, they 136 increase their likelihood of interacting with certain partners, especially their kin [6]. That is, 137 ageing females increase both their tendency to reassociate with others, and their tendency to 138 interact with someone from their own (kin-based) sub-group. Individuals with higher within-139 group association tendencies mix less widely in their networks and those with greater 140 reassociation tendency are less likely to associate with new individuals and connect disparate 141 parts of the network [51]. Therefore, we predicted that as females aged, they would have lower 142 measures of betweenness (capacity for linking discrete clusters in a network) and closeness 143 (capacity to reach others or be reached), but a higher clustering coefficient (cliquishness).

144 To determine if age-based changes at the individual level can result in changes in overall 145 network structure we used empirical data from 19 networks of female rhesus macaques to test 146 how variable proportions of old individuals in a network was related to three common measures 147 of global network structure: mean degree, network diameter and transitivity (Table 1). We 148 predicted that networks with a greater proportion of old individuals would be more sparsely 149 connected due to older animals having fewer social partners (i.e., have lower mean degree and 150 network diameter). Given greater kin clustering with age, we expected networks with more old 151 individuals would be more clustered (i.e., have higher transitivity). Finally, to help inform our

152	empirical findings and better understand the link between age distribution and global structure,
153	we built an agent-based model (parameterized using information from our empirical data) to
154	simulate how different proportions of old individuals would be expected to affect network
155	structure under simplified conditions where everything else is equal. Our results provide a first
156	step to understanding how and when individual changes in social tendencies with age might scale
157	up to detectable effects on global network structure, offering important insight into the
158	consequences of demography for the structure and function of collectives.
159 160	Methods
160 161 162	Study population and data
163	Data used in this study were collected on a well-studied population of rhesus macaques
164	on the island of Cayo Santiago, off the southeastern coast of Puerto Rico. The current population
165	is maintained by the Caribbean Primate Research Center (CPRC) and is descended from 409
166	macaques that were introduced to the island from India in 1938. The animals are food
167	supplemented and provided with ad libitum access to water. There is no medical intervention,
168	and so the major causes of death at this predator-free site are illness and injury [55,56]. The
169	CPRC staff collect demographic data five days per week and thus track dates of birth and death
170	of all individuals with a high degree of accuracy.
171	Rhesus macaques are highly social cercopithecine primates that live in matrilineal kin-
172	groups and exhibit clearly differentiated social relationships with kin-biased affiliation
173	[55,57,58]. At six years old, females are deemed adults [59] and previous research on the
174	macaque population of Cayo Santiago has shown that, for females that survive to reproductive
175	age, the median lifespan is 18 years with a maximum lifespan of about 30 years [53,60]. Female
176	rhesus macaques have a strict dominance hierarchy with maternal rank inheritance and youngest

ascendency [61]. Patterns of social interactions and social attention vary between young and old
adults [6,34,62]. Female macaques show clear evidence of within-individual declines in the
number of grooming partners with age, although the amount of time spent giving and receiving
grooming remains constant across adulthood [38].

181 For this study, subjects were mature adult females 6 years and older [59] from 6 naturally 182 formed mixed-sex social groups. We used data collected between 2010-2017, a time period for 183 which we had detailed behavioural data from which to estimate social networks. We collected 184 behavioural data between 07:30 and 14:00, which are the working hours of the field station, 185 using 10-min focal animal samples and recording all behaviours continuously [63]. We balanced 186 data collection to ensure equal sampling of individuals throughout the day and over the course of 187 the year, resulting in approximately the same number of focal samples per individual per year. 188 For these analyses we used grooming interactions to build our networks, given the clear age-189 based changes in grooming associations previously demonstrated in this system [6]. Grooming 190 behaviour was collected by recording the duration of a grooming bout along with the identities of 191 the interactants and the direction of grooming. We focused only on interactions between adult 192 females (\geq 6 years old) and did not include interactions with infants (<1 year old), juveniles (2-3 193 years old), or sub-adult females (4-5 years old). We also did not include interactions with males 194 as we wanted to avoid capturing changes in socio-sexual behaviour with age. We established 195 dominance ranks for all females in a given year using observed win-loss interactions (as per 196 [64,65]). Rank was assigned as "high" ($\geq 80\%$ of other females dominated), "medium" (50%-197 79% of other females dominated) or "low" (\leq 49% of other females dominated).

198 Social networks



We built 19 grooming networks including all adult females (≥ 6 years old) from the

200 following group-years (group F 2010-2017; group HH 2014; group KK 2015; group R 2015-201 2016; group S 2011; group V 2015-2016), with data availability based on the focus of projects 202 over time given limited person power. We used weighted network metrics, as these are more 203 robust and provide higher resolution than binary measures [66]. In these weighted networks, 204 edges represented the undirected rate of grooming between a pair of individuals (seconds of 205 grooming/total number of hours that both individuals were observed in focal animal samples). 206 We note that using undirected grooming rates may fail to capture some finer nuances of age-207 based variation in social behaviour if, for instance, older individuals were likely to give less 208 grooming than they received [61]. However, our most recent within-individual analyses suggest 209 that while the number of grooming partners does decline as females age, both the amount of 210 grooming given and received remains constant [6]. Therefore, the use of undirected grooming 211 rates to build networks should provide a comprehensive picture of how individual connectedness 212 to the wider network changes with age, without missing essential changes in the form that the 213 connectedness takes. The average size of our networks was 50.7 (\pm SE = 3.9) adult females. All 214 network metrics were calculated in R (version 4.2.0; [67]) using the igraph package (version 215 1.3.1; [68]).

216 Empirical analyses

All empirical models were fitted in a Bayesian framework with different error structures and random effects dependent on the data analysed (see below). We conducted all analyses using R (version 4.2.0; [67]) and fitted all models in the Bayesian software STAN [69] using the brms package (version 2.17.0; [70]). All fixed effects were given weakly informative priors (see Supplementary Information for more details). We ran all models for 10,000 iterations across two chains with a warm-up period of 2,000 iterations. We assessed model convergence by examining

trace plots to assess sampling mixing and by ensuring Rhat = 1. We considered estimates of
fixed effects to be significantly different from zero when the 95% credible intervals of the
posterior distribution did not overlap zero.

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227 Investigating the relationship between age and indirect connectedness

229 For these analyses we set out to test how individuals' level of indirect connectedness 230 to their social network changed from prime adulthood into later life. In our dataset, the median 231 age of adult females was 10-years and previous studies in this system have shown that 232 individuals aged 10 and beyond show clear evidence of physical [71–74], immunological [75], 233 reproductive [71], and social ageing [6]. To capture later-life changes in indirect connectedness 234 we focused our analyses on individuals aged 10 and older, in line with previous ageing studies 235 [6]. Females in this analysis therefore ranged between 10 and 28 years old, although, to be clear, 236 all measures of indirect connectedness for these females were extracted from the networks 237 including all adult females (≥ 6 years old; see above). We had 563 macaque years of data over 238 204 unique females, with an average of 2.8 years of data per individual (range: 1-8 years; Fig. 239 S1). Because there was variation in the age-ranges over which individuals were sampled, we 240 used a within-individual centering approach to capture changes in indirect connectedness across 241 individuals' lifespans [76]. Briefly, following the methodology of van de Pol & Wright (2009) 242 we split our age term into a between-individual effect (calculated as the mean age of an 243 individual across all observations; hereafter called average age) and a within-individual effect 244 (calculated as the deviation of an individual's age from their mean age; hereafter called within-245 individual age). This within-individual age term was our primary variable of interest and reflects 246 how an individual's deviation from its age affects its indirect connectedness in the network (see 247 [6] for a more detailed description of these methods).

248 We fitted four models with our four response variables of interest: eigenvector centrality, 249 betweenness, closeness, and clustering coefficient. All statistical models included average age 250 and within-individual age as continuous fixed effects. Given the strict dominance hierarchy 251 exhibited by female rhesus macaques [61], we assessed whether social status affected the change 252 in an individual's measures of indirect connectedness with age [6] by fitting an interaction 253 between rank and within-individual age in all models. We removed the interaction when not 254 significant. We included individual ID, group, and year, as random effects to account for 255 repeated observations and to capture any variation in indirect connectedness measures that might 256 be due to differences between individuals, groups, or years. We also fitted within-individual age 257 as a random slope over individual ID to capture any among-individual variation in the change in 258 indirect connectedness with age. We had no biological reason to expect non-linearities in the 259 relationship between age and measures of indirect connectedness given that we were looking at 260 changes from prime adulthood to old age and we've previously found a linear relationship 261 between direct measures of connectedness and age [6]. Nevertheless, we fitted a model with 262 quadratic terms for within-individual age and average age and compared that model to the model 263 with only linear age terms using leave-one-out cross-validation in the brms package (version 264 2.17.0) [70]. The quadratic terms never improved the model fit and so were not considered 265 further. For the eigenvector centrality model, we logit-transformed the response variable to 266 improve model fit and fitted a model with a Gaussian error distribution. For betweenness we 267 fitted a model with a zero-inflated Poisson distribution. For closeness and clustering coefficient 268 we fitted models with a Beta error distribution and a zero-one-inflated Beta error distribution, 269 respectively.

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272 Investigating the relationship between age distribution and network structure 273

274 Our analyses quantifying changes in indirect connectedness with age focused on females 275 aged 10 and older to capture individual changes in social network position from prime age and 276 beyond. Here we were interested in how the age distribution of a group was linked to network structure. For these analyses we therefore used our complete networks including all adult females 277 278 aged 6 and older. We had three response variables of interest (three measures of global network 279 structure) and fitted a separate model for each: mean degree, diameter, and transitivity. To ask if 280 age distribution relates to network structure, we included the proportion of old individuals as a 281 continuous predictor in all our models. Since 18 is the median age of death in this population and 282 maximum lifespan is about 30 years, we considered individuals above 18 to be "old" [52,61]. 283 The proportion of old females in our 19 networks ranged from 0.03 to 0.19 (Fig. 1; Fig. S2). 284 Given that rhesus macaques live in matrilineal groups with kin-biased behaviour [53,77], we 285 included the average relatedness of the network as a continuous covariate to account for 286 differences between groups in general relatedness. We also included network density (calculated 287 as the number of existing ties in the network divided by the number of possible ties) as a 288 continuous fixed effect in the models of diameter and transitivity to account for the fact that 289 density can drive variation in other global network metrics. We included year as a random effect 290 to account for yearly variation that might affect network structure. For the diameter model, we 291 log-transformed the response variable to improve model fit. We fitted all models with a Gaussian 292 error distribution. To account for a potential non-linear relationship between age distribution and 293 network structure we fitted all models with a quadratic term for proportion of old individuals and 294 compared that to the model with only linear terms using leave-one-out cross-validation in the

brms package (version 2.17.0; [70]). The quadratic terms never improved model fit and so werenot considered further.

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298 Agent-based model of the relationship between age distribution and network structure

299 We expected that age-based changes in patterns of direct association would scale up to 300 affect overall network structure. However, our empirical data revealed no effect of the proportion 301 of old individuals in the network on global network metrics (see Results section below). To 302 better understand these results, we built an individual-based model where we could manipulate 303 age-based differences in sociality in isolation of other variables to better explore how different 304 age distributions would be expected to affect overall network structure, and why such scaling up 305 of age-based differences in social behaviour may not be detectable in our data. Our general 306 approach was to simulate artificial populations that exhibited the social age-dependencies 307 observed in the rhesus macaque system on Cayo Santiago and use these to investigate the 308 relationship between age distribution and global network metrics. As mentioned above, adult 309 female rhesus macaques change two aspects of their sociality with age: their number of social 310 partners, and the proportion of partners that are kin [6]. These individual changes in sociality 311 across adulthood lead to differences between young and old individuals in both the number of 312 social connections and probability of connecting with kin [6].

The model therefore simulates social networks of varying age distributions, in which the probability that two nodes (i.e., individuals) have an edge (i.e., a grooming link) depends on the age of the individuals and their and kinship to each other. Individuals belonged to two age categories (as in the empirical analysis for global structure; old adults and young adults, hereafter "old" and "young") and two kin categories (kin and non-kin). This gives 6 dyad types, which we denote by the age categories of the two individuals and their kinship status (e.g., old/old kin). We

319 established the linking probability (i.e., the chance of having an edge in the social network) for 320 each dyad type based on the mean proportion of dyads of that type that had an edge across the 19 321 empirical macaque grooming networks. The linking probability for dyads that were old 322 individuals who were related to each other (i.e., old/old kin) was 0.33; for old/old non-kin dyads 323 it was 0.02; for old/young kin dyads it was 0.37; for old/young non-kin dyads it was 0.05; for 324 young/young kin dyads it was 0.27 and for young/young non-kin dyads it was 0.08. We fixed 325 group size at 50 individuals, which approximates the mean number of adult females in real 326 groups on Cayo Santiago (mean \pm SE = 50.7 \pm 3.9). Each simulated network had 10 clusters of 327 individuals who were related to each other (i.e., kin groups) with 5 individuals in each. The 328 number of individuals within each kin group mirrors the mean number of close adult female kin 329 (relatedness coefficient ≥ 0.125) that adult female rhesus macaques have on Cayo Santiago 330 (mean \pm SE = 5.2 \pm 0.87). In the model, kinship between pairs of individuals was determined by 331 their kin group membership: individuals from the same kin group were classed as kin and 332 individuals from different kin groups were classed as non-kin. 333 Each simulation round (i.e., construction and quantification of one network) proceeded as 334 follows: We first randomly drew the number of old individuals in the group (n_{old}) from a uniform 335 distribution with a set range [0, 50]. We then randomly assigned all 50 group members to age-336 groups (n_{old} old individuals and n_{voung} = 50- n_{old} young individuals) and kin groups. This allowed 337 the age structure of the groups to vary across the range from only young individuals (0% old) to only old individuals (100% old). We then constructed the social network by drawing links, where 338

the chance of each dyad getting a link depended on their dyad type (linking probabilities given

above). To determine whether a dyad was given a link, we extracted a random value from a

341 binomial distribution with a sample of 1 and probability equal to the linking probability of the

type of dyad. If the extracted value was 1, the dyad was given a link, if it was 0, the dyad was not
given a link. A schematic representation of the various steps of the model-building process can
be found in Fig S3.

345 To confirm that the model was working as intended, we ran 10000 simulations (i.e., we 346 generated 10000 networks), from which we calculated the proportion of dyads of each type that 347 had a link. We confirmed that the means of these proportions (old/old kin = 0.33; old/old non-kin 348 = 0.02; old/young kin = 0.37; old/young non-kin = 0.05; young/young kin = 0.27; young/young 349 non-kin = 0.08) corresponded well to their respective empirical linking probability (see above, 350 and Fig. S4). We also confirmed that our model was not highly sensitive to our input linking 351 probabilities (see Supplementary Methods, Table S1, Fig. S5 for more information). Given that 352 our input values are only estimates of the real values, robustness to these values is important for 353 drawing general conclusions about the potential effects in the real system.

354 To investigate the general relationship between age distribution and network structure, 355 we ran 100000 replicates of the simulation (i.e., generated 100000 networks) where we allowed 356 the proportion of old individuals in the network to range from 0.00 to 1.00 and visualised the 357 relationship between the proportion of old individuals and network mean degree, diameter, and 358 transitivity. Additionally, we used the agent-based model to investigate if we should expect to 359 find relationships between the proportion of old individuals and network metrics in the empirical 360 analysis. To do this, we limited the proportion of old individuals in each network to between 361 0.04 and 0.20 (mirroring the empirical variation in age distributions), and we ran 19 simulations 362 (mirroring the 19 empirical networks). We repeated this 50 times (i.e., 50 sets of 19 networks) to gauge consistency of the results. 363

364

365 **Results**

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367 Relationship between age and indirect measures of network connectedness

- 369 In line with our predictions, female macaques did not show any change in eigenvector centrality with age (within-individual age: $\beta = -0.04$; 95% CI = -0.30, 0.24; Fig. 2A-B, Table S2). 370 371 That is, as females aged, the strength of their relationships to their partners, and to their partners' 372 partners, was stable. We did find evidence of a within-individual decline in betweenness with age. 373 However, this effect was rank dependent and seems to be driven primarily by a decline in 374 betweenness as high-ranking individuals got older (within-individual age:rankH: $\beta = -0.35$; 95% 375 CI = -0.41, -0.29; Fig. 2C-D, Table S3), although mid-ranking individuals also showed greater 376 declines in betweenness with age than did low-ranking females (within-individual age:rankM: β = 377 -0.14; 95% CI = -0.18, -0.09; Fig. 2C-D, Table S3). Therefore, some, but not all individuals 378 became less effective at reaching disparate nodes in the network as they got older. Individuals showed a reduction in their closeness with age (within-individual age: $\beta = -0.18$; 95% CI = -0.35, 379 380 -0.02; Fig. 2E-F, Table S4), meaning they were harder to reach and be reached in the network, as 381 we predicted. Contrary to our predictions, there was no within-individual effect of age on 382 clustering coefficient (within-individual age: $\beta = 0.04$; 95% CI = -0.06, 0.14; Fig. 2G-H, Table 383 S5). That is, individual cliquishness was stable as individuals aged.
- 384

385 Relationship between age distribution and network structure

386

387 Empirical results

388 Contrary to our predictions, we found no evidence that groups with a greater number of 389 old individuals were structured differently from groups with fewer old individuals. There was no 390 overall effect of the proportion of old individuals in the group on mean degree ($\beta = -2.76$; 95%

391 CI = -10.24, 5.25; n = 19; Fig. 3A, Table S6), or network diameter ($\beta = -0.12; 95\%$ CI = -3.93,392 3.67; n = 19; Fig. 3B, Table S7). Therefore, networks were not more sparsely connected as the 393 proportion of older animals increased. Older networks were also not more clustered or cliquish, 394 as measured by transitivity ($\beta = 0.20; 95\%$ CI = -0.29, 0.70; n = 19; Fig. 3C, Table S8).

395

396 Agent-based model results

397 Different group demographics in age appeared to have important consequences for 398 network structure in our simulations (Fig. 4; Fig. S6). Specifically, our 100,000 simulations 399 where we allowed the proportion of old individuals in the network to range from 0.00 to 1.00 400 showed that, as the proportion of old individuals in the population increased, mean degree (i.e., 401 the mean number of partners that each individual associated with) decreased (Fig. 4D), and the 402 diameter of the network (i.e., the longest path length in the network) increased (Fig. 4E), both as 403 expected. Network transitivity (i.e., the degree of clustering in the network; Fig. 4F) showed 404 little variation in relation to age distribution.

405 Importantly, these global metrics did not always exhibit linear relationships with network 406 age (as might have initially been intuitively expected). The potential for non-linearity is 407 important for understanding when the effects of age on network structure might be the most 408 pronounced and most detectable in the real world. In particular, the increase in network diameter 409 with an increasing proportion of old individuals showed an accelerating trend (Fig. 4E). That is, 410 the relationship between network diameter and proportion of old individual was stronger (i.e., the 411 slope was steeper) when the proportion of old individuals in the network was high, suggesting 412 that effects of age on network structure might be most detectable when, for example, greater than 413 50% of the individuals in the network are old. While we did not find any clear change in 414 transitivity with the proportion of old individuals in the network, there was also some evidence

415	of an accelerating trend at the higher end of the age distribution. These metrics (transitivity,
416	diameter) are driven to a greater extent by modularity in networks and the non-linear relationship
417	could arise from the kin-biased nature of the relationships in our (modelled) study system. We
418	also observed that transitivity and network diameter were more variable when there was a greater
419	proportion of old individuals in the network. This suggests that changes in network structure
420	with changing age distribution might be unpredictable, for example, networks might become
421	more clustered or less clustered as the population gets older. Overall, these results suggest that
422	compared to networks with more young individuals, networks with more old individuals are
423	sparser and less cohesive, and can take on more varied structures.
424	When we limited the sample size and variation in age distribution to those of the
425	empirical data (i.e., 19 networks where the proportion of old individuals was only allowed to
426	vary between 0.04 and 0.20) we observed no clear change in network structure as the proportion
427	of old individuals in the network increased (Fig. 4G-I).
428	
429	Discussion
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	Discussion Age has begun to emerge as an important attribute shaping the social decisions of
430	
430 431	Age has begun to emerge as an important attribute shaping the social decisions of
430 431 432	Age has begun to emerge as an important attribute shaping the social decisions of individuals [6,7,14]. This implies that age is a potentially significant feature underlying the
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430 431 432 433 434 435 436	Age has begun to emerge as an important attribute shaping the social decisions of individuals [6,7,14]. This implies that age is a potentially significant feature underlying the behaviour of animals in collectives. Here we have shown that age-based changes in social behaviour are likely to have important consequences for an individual's position within the wider social network and may scale up to influence network topology. As female rhesus macaques aged, they showed declines in some, but not all of their indirect measures of connectedness. But

440 based differences in social behaviour do necessarily, as expected, scale up to affect network 441 structure, but that these effects may be nonlinear and so may not always be easily detectable. In line with our predictions, we found no evidence for changes in eigenvector centrality 442 443 with age. This was expected because, despite reducing their number of social partners, female 444 rhesus macaques maintain the amount of time they engage in social interactions as they get older 445 [6], allowing them to continue to have strong connections to some of their partners, and 446 (indirectly) to the partners of their partners. A similar pattern has been observed in a population 447 of wild rhesus macaques in China whereby older individuals had fewer partners but exhibited 448 similar weighted eigenvector centrality to those of younger individuals [34]. Unexpectedly, we 449 did not find any evidence that clustering coefficient increased as individuals aged in this 450 population. Females increase their preference for grooming kin as they age [6], suggesting that 451 within-group (e.g., homophilic) tendencies increase with age, which should lead to greater 452 clustering as females from the same matriline would cluster together in the network. However, 453 although kin-biases increase with age, they exist for females of all ages [58]. It is therefore 454 possible that ties with kin are the main driver of clustering coefficient scores regardless of age. 455 The loss or removal of a few non-kin relationships as an individual ages could thus have a 456 relatively minor impact on their clustering coefficient, leading to stable values for this metric 457 across the adult lifespan. In general, the stability of weighted eigenvector centrality and 458 clustering coefficient across the adult lifespan suggest that older animals can remain socially 459 central and well-integrated in some respects, despite other aspects of their social life changing 460 (e.g., fewer partners overall, and the patterns observed below). Old individuals may therefore continue to reap some of the advantages of social relationships [5]. 461

462 As predicted, we found declines in betweenness and closeness with age, although some 463 of these results were rank dependent. As individuals got older, particularly those that were higher 464 ranking had less influence on and became less well-connected to the wider network. 465 Betweenness and closeness are associated with an individual's level of reassociation tendency 466 and within-group association [51]. In other words, the more likely individuals are to reassociate 467 with the same partners, and to select partners of similar characteristics to themselves, the less 468 they mix with the wider network and the less likely they are to connect distinct subgroups (i.e., 469 have high betweenness) or be easily reached by all others in the network (i.e., have high 470 closeness). We've previously shown that ageing rhesus macaque females have fewer partners, 471 but engage with those partners more often, and increase their preference for close relatives. In 472 other words, while mean dyadic association strength to associates increases with age, mixing 473 with the wider network should decline [6]. Our betweenness and closeness results seem to reflect 474 these underlying changes in behavioural patterns. The declines in betweenness with age were 475 rank dependent, with mid-ranking and high-ranking rhesus macaque females showing a greater 476 decline in betweenness with age than low-ranking individuals. This may simply be due to a floor 477 effect – because low ranking individuals have such low betweenness to begin with, there is little 478 capacity for further decline with age.

Generally, our findings demonstrate that how indirectly connected an individual is to their social world can change across their lifespan. How an individual is positioned in the wider social environment can modulate their exposure to information [26,31], parasites [18,77], and pathogens [20–22]. Individuals with high betweenness and closeness occupy a critical position in the acquisition and transfer of 'goods' within a network [17,54,78]. Decreases in both of these measures of indirect connectedness with age may benefit aged individuals who may experience

485	greater susceptibility to disease or illness as a result of immunosenescence [79]. Such benefits
486	need not imply that individuals actively change complex network positions with age. By
487	changing simple behavioural rules or processes with age, changes in polyadic ties with age are
488	likely to emerge [51]. Measures of indirect connectedness have been linked to fitness proxies
489	including future social status [80,81], survival [82,83], and reproductive success [32,64,81].
490	Indirect network metrics can be an even stronger predictor of fitness proxies [17,32] than
491	measures of direct connections and may more strongly reflect the underlying behavioural rules
492	that give rise to individual differences in sociality [51]. As such, documenting how ageing shapes
493	the polyadic social world may be particularly relevant for understanding how changes in
494	sociality across the lifespan influence patterns of senescence and fitness in later life.
495	While age is clearly associated with changes in the behaviour and social connectedness of
496	animals living in groups [10,13,14], including in this population of rhesus macaques [6,60], we
497	did not find any empirical evidence that age impacts the overall structure of those groups. We
498	found no relationship between age distribution and network mean degree, diameter, or
499	transitivity in the observed macaque data. This was surprising to us given that we expected that
500	
	age-based changes at the individual level would scale up to the network level. That is, it should
501	be self-evident that a network with more old individuals who each interact with fewer partners
501 502	
	be self-evident that a network with more old individuals who each interact with fewer partners

505 expected effects on network structure.

504

506 By modelling age-based differences in two interaction patterns (number of social partners507 and the tendency to link with kin) based on the age-based changes we observed in the female

emerge and whether under a scenario where everything else is equal, we could recreate these

508 macaques [6], we found that the age composition of a group can have important consequences 509 for its cohesiveness and connectedness. As predicted, mean degree declined and diameter 510 increased as the proportion of old individuals increased, while transitivity exhibited no strong 511 relationship to age-distribution. Interestingly, these effects did not necessarily scale in the linear 512 manner that might be expected in response to a linear increase in proportion of old individuals in 513 the network. Network diameter showed a steeper increase as networks became older, suggesting 514 that the strength of the effect, and thereby the ability to statistically detect an effect in a real-515 world case, might depend on where along the continuum of age distributions one's data lie. The 516 combination of this and a limited sample size could potentially have led to the null result in our empirical analysis. This interpretation is supported by the fact that when we limited the age 517 518 variation and sample size (number of networks) of our model to match that of the empirical data 519 (i.e., only ran the model for 19 simulations and allowed the proportion of old individuals in the 520 network to range between 0.4-0.20) there was no clear relationship between the proportion of old 521 individuals and network structure.

522 The results of our model suggest that populations made up of very large numbers of 523 young, or old, individuals may have detectable levels of divergence in network structure, 524 whereas populations within a smaller range of age distributions may not. Sample size also 525 appears to have played an important role in our ability to detect an effect of network age on 526 network structure. Each "bout" of 19 networks (simulations) produced a slightly different result. 527 Sometimes slopes were positive, sometimes negative, sometimes flat, suggesting there is a fair 528 amount of stochasticity in our ability to detect the true effect when the sample size is small, even 529 in the model world where there are no other confounding factors. While nineteen rhesus 530 macaque networks from 6 groups collected across 8 years represents a considerable investment

531 of research time and effort, it may not have been enough to detect a clear pattern. However, it is 532 also possible that our empirical results reflect a true null result if there are processes occurring that our model has not accounted for. Our model is a simplistic version of the real world and 533 534 does not fully replicate the complete suite of changes that might occur as individuals age. For 535 example, our model does not allow for potential rewiring between nodes which might occur if 536 younger individuals choose to build new ties in response to the loss of ties with older individuals. 537 Such rewiring could explain why some measures of network structure do not change despite 538 individual declines in sociality with age.

539 While the general result from the agent-based model was expected, these simulations 540 help to inform our understanding of the empirical data and offer us a glimpse into how age-based 541 differences in sociality may affect network structure and when these effects might be detected. 542 The fact that, even in our simplified model where there is nothing to obscure individual linking 543 probabilities from scaling up to affect global network metrics, changes in some measures of 544 network structure appear most likely to be detected in very old populations raises the question of 545 when such effects are likely to be detectable in wild systems. The macaques of Cayo Santiago 546 are a managed population that are food supplemented and have no predators. As such the 547 proportion of old individuals seems likely to be higher than in many wild populations and aligns 548 with other natural but non-predated systems (e.g., red deer [84]) where the proportion of old 549 individuals in the population appears to top out at about 20%. Our model suggests that the effects 550 of age on network structure are also present at this lower end of the age-distribution, but are less 551 strong, or negligible, for some network measures. Hence, detecting an effect in real systems 552 could necessitate large amounts of data. We note, however, that the effect on the network 553 measures (the shapes of the curves) will depend both on how sociality changes with age, and the

social organisation of the system (such as relatedness patterns), and while other systems may be
similar to the one here investigated in these regards, making predictions for detectability of
social ageing effects in general would require more comprehensive modelling.

557 Future work in other systems, conditions, and contexts could reveal consequences of age-558 based changes in sociality for network structure that are even more stark. For example, greater 559 differences between age classes in the overall probability of forming a social tie may be expected 560 in societies that experience selective disappearance of more social individuals or where declines 561 in sociality with age are exacerbated by partners dying and not being replaced (which we do not 562 observe in this system [6]). The kin structure of populations is also likely to be important if the 563 probability of connecting decreases with age but there is no compensatory increase in the 564 probability of forming ties with kin. This could lead to stronger absolute effects on global 565 network metrics, including substantial changes in transitivity, and thus would likely have further 566 consequences for network connectivity and spreading processes. While our analyses bin 567 individuals into age categories of "old" and "young", there may be much more nuanced effects 568 of age on network structure if the full age distribution of a population or group is taken into 569 account.

It is also important to recognize that the results presented here only reflect how changes in female sociality across adulthood are expected to influence adult female social networks. Understanding how social networks of entire groups are affected by age-based changes in behaviour, more broadly, will necessitate incorporating an understanding of the patterns of social ageing in both sexes, as well as an understanding of how interactions with juveniles shift as adults get older. While much less is known about the social ageing trajectories of males, some studies suggest that they show lower levels of connectedness at older ages, similar to those

observed in females [13,33]. But even in cases where social ageing trajectories are similar across
the sexes, changes in socio-sexual interactions with age in addition to shifting interactions
between age classes are likely to produce much more complex network dynamics than observed
here, necessitating further research.

581 Despite these limitations, our study provides a first glimpse into the effects that age-based 582 changes in behaviour might have on network structure, offering a deeper understanding of the 583 potential importance of demographic changes for the structure and function of collectives. Much 584 attention has been given to the potential for the social structure of populations to be affected by 585 the removal of older individuals, for example through practices like trophy hunting or fishing 586 [85,86]. This is relevant given that older individuals can play a critical role in their social groups 587 through leadership [4,39], intergenerational transfers [40], and the stabilisation of social 588 relationships [87]. Despite the importance of the loss of older individuals from networks, our 589 results suggest that asking what happens when populations contain large numbers of older 590 individuals may be an equally salient question.

591 In humans, population ageing is poised to be one of the most significant social 592 transformations of the 21^{st} century, with the global number of older persons projected to more 593 than double over the next three decades [88], facilitated by falling fertility rates and increasing 594 longevity [89,90]. Meanwhile, in other group-living species, anthropogenic challenges, including 595 climate change, poaching, disease outbreaks, and pollution, among others, are likely to cause 596 major population declines and demographic shifts [91,92]. The consequences of declining 597 fertility and/or increasing mortality rates may be further exacerbated by the concomitant shifts in 598 age structure and resulting implications for network structure and function. For instance, declines 599 in cohesiveness and structural connectedness in older networks may limit or slow the transfer of

information [24,93], restrict the potential for cooperation [94] or reduce the stability of
populations [30], feeding back to further limit population growth. Alternatively, such changes
could reduce the vulnerability of populations to infectious diseases or parasites [30,95] or
enhance cooperation if network modularity is increased [96]. Consequently, through its effects
on network structure, shifting age demography may have broader implications than previously
appreciated for group dynamics and persistence and warrants further research.

606

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622 **Competing Interest Statement:** The authors declare no competing interests.

- **Table 1.** Predictions for how indirect measures of connectedness are expected to change with
- 624 age in female rhesus macaques from Cayo Santiago and how global network metrics are
- 625 expected to change with increasing proportions of old individuals in the population.

Network Metric	Definition	Prediction	Rationale
Indirect network metrics			
Eigenvector centrality	Measures how well connected an individual is in the network to individuals who are themselves well connected.	Remain stable with increasing age	Females retain strength of connections to partners as they age.
Betweenness	Calculates the number of shortest paths between others in a network that pass through an individual. Individuals with low betweenness have low capacity for linking discrete clusters in the network.	Decrease with increasing age	Females increase their probability of interacting with strong, stable partners
Closeness	Measures the inverse distance to all other individuals in the network. The lower an individual's closeness score, the more difficult it is for them to reach others or be reached.	Decrease with increasing age	and kin. Therefore, females increase their tendency to reassociate with existing partners and have higher within-group association tendencies as they age.
Clustering	Measures cliquishness or subgrouping. A high clustering coefficient indicates that an individual's partners are highly connected to each other.	Increase with increasing age	
Global network metrics			
Mean degree	The average number of ties that each individual has to others in the network.	Decrease with increasing proportion of old individuals	Older females have fewer social partners.
Diameter	Measures the overall connectedness of the network. Networks with a larger diameter are less connected.	Increase with increasing proportion of old individuals	
Transitivity	The degree of clustering in the network.	Increase with increasing proportion of old individuals	Older females cluster more with kin.

632 Figure 1. (A) Distribution of ages in each of the 6 empirical macaque groups for each year that 633 those groups were observed (19 networks total observed between 2010-2017). The data are 634 presented as a stacked bar chart, so to compare the proportion of x-year-olds in a single group 635 across years one should compare the height of coloured bars within a single age. The dotted 636 black line at age 18 indicates the median age of death in this population and the cut off at which we considered individuals "old" for the sake of calculating the proportion of old individuals in 637 638 the group. See also Fig. S2. (B) Proportion of old individuals in each group for each year that 639 group was observed. Years with no bars indicate years in which the group was not observed.

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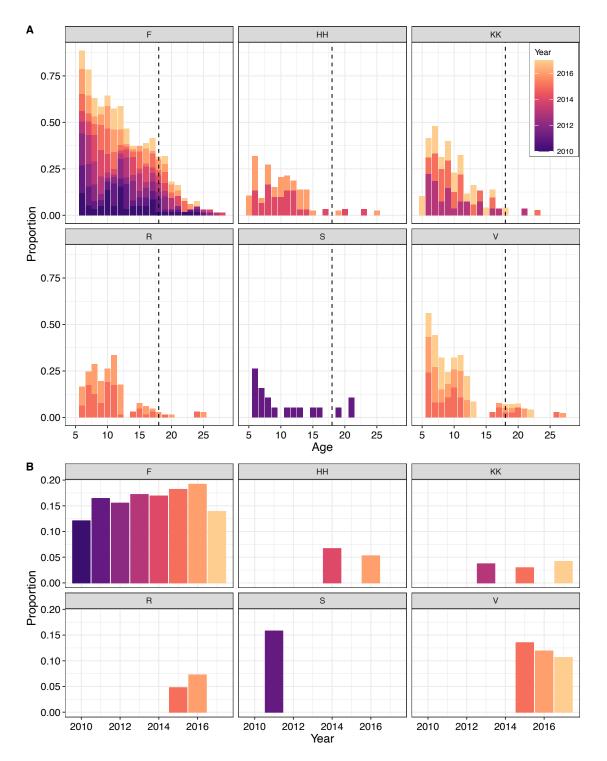


Figure 2. Relationship between within-individual changes in age and indirect measures of
connectedness including (A & B) eigenvector centrality, (C & D) betweenness, (E & F)
closeness, and (G & H) clustering coefficient in female rhesus macaques. In A, C, E, G the
points represent raw data. Shaded grey bars indicate 95% confidence intervals around the
predicted values. In B, D, F, H effect sizes and 95% credible intervals for all fixed effects and
interaction terms are shown. Instances where the 95% CI overlaps zero are coloured in purple.

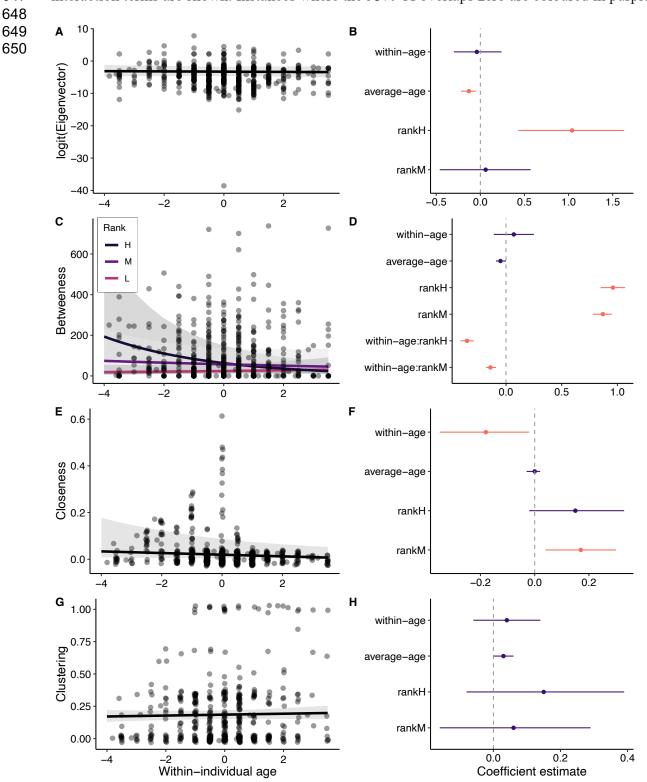
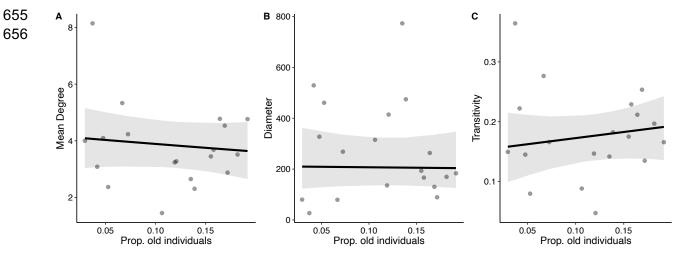
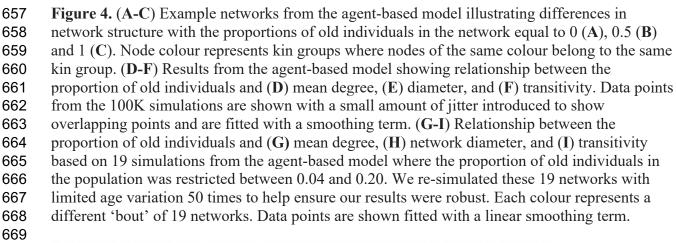
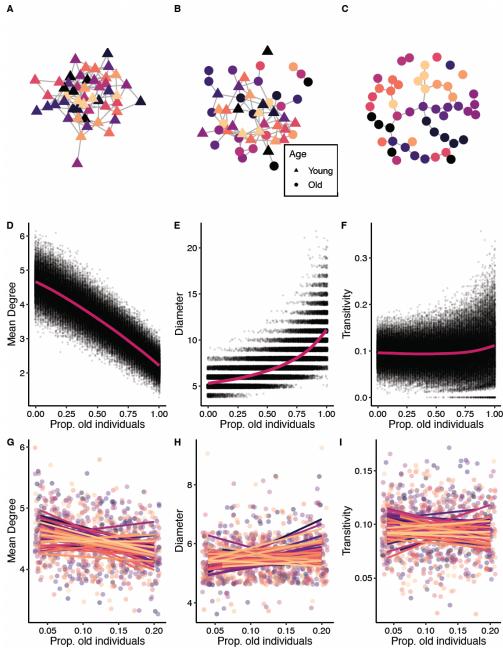


Figure 3. The relationship between the proportion of old individuals in female rhesus macaque
networks and: (A) mean degree, (B) network diameter, and (C) network transitivity. Points
represent raw data and shaded grey bars indicate 95% confidence intervals around the predicted
values.







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