Comparison between synonymous transition/transversion with non-synonymous transition/transversion reveals different purifying selection on coding sequences

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ABSTRACT

Base substitution is a major event in genome variation. Therefore, a correct estimation of the base substitution event is important to understand the process of selection. Here, we developed a method (link) to estimate $\frac{ti}{tv}$ ratio by accounting codon degeneracy as well as pretermination codons in a coding sequence. A distinct observation of the influence of codon degeneracy and pretermination codon on $\frac{ti}{tv}$ was observed in 2516 genes of *Escherichia coli* analyzed in this study. The mean Sti, Stv, Nti and Ntv values revealed the decreasing order of their variation dynamics, respectively in *E. coli* genome. While a strong positive correlation was observed

between Sti, and Stv values, the correlation between Sti, and Nti was poor in coding sequences.

In several instances coding sequences with same Sti values were found to have different Nti

values. This indicated that the varying strength of purifying selection on Nti in genes is

responsible for the poor correlation above. Our approach is convenient to visualize the base

substitution variation dynamics as well as selection in coding sequences. Also, the proposed

method is useful to estimate different $\frac{ti}{tv}$ ratios accurately in coding sequences and is insightful

from evolutionary perspectives.

INTRODUCTION

The genetic code table is a beautiful gift of nature to study molecular evolution (1,2). The

different degenerate codons influence the transition (*ti*) to transversion (*tv*) ratio $(\frac{ti}{tv})$ as follows:

synonymous variation in a *two-fold degenerate* (TFD) codon occurs *via* only one transition

whereas the same in a *four-fold degenerate* (FFD) codon occurs via one transition and two

transversions (Supplementary Figure S1A). In addition to the above, compositional variation of

pretermination codons (PTC) (Supplementary Figure S1B and S1C) (3) among the coding

sequences is also likely to influence the $\frac{ti}{tv}$ ratio because non-synonymous variations leading to

termination codons purged out rapidly in a population due to strong purifying selection (4-6).

Considering in general, a *ti* being more frequent than a *tv* (7-13) which is known as Kimura's two

parameter model (14), and a stronger purifying selection on non-synonymous variation than a

synonymous variation (15), the $\frac{ti}{tv}$ ratios of TFD codons and FFD codons are expected to be

different. Accordingly, it has been recently reported that the synonymous variation in TFD

codons is relatively higher than the FFD codons in *Escherichia coli* (16). Therefore, in any

organism, the $\frac{ti}{tv}$ ratio is likely to vary across the coding sequences having differences in

degenerate codon composition. Hence, a comparative study of $\frac{ti}{tv}$ among coding sequences in a

genome will be always insightful.

Considering the number of possibilities of the two types of base variations in genomic

DNA (17) (Figure 1). The ratio of *ti* to *tv* is theoretically expected to be 0.5. But the ratio is

generally observed more than 0.5 in the genomes (7,8). Because of the following factors such as

geometry of DNA double helix (18), modification of DNA bases such as deamination of cytosine

and adenine (19-21), secondary structure in the RNA (9) and codon degeneracy are known to

elevate *ti* phenomenon in a genome over *tv* (16,22). In a standard genetic code table out of 549

total single nucleotide variations (SNVs) involving 61 sense codons (9 SNVs per codon), only

134 SNVs are synonymous (24.4%). Among synonymous (*Sti* and *Stv*) and non-synonymous (*Nti*

and *Ntv*) variations, the theoretically estimated ($_e$) numbers of *Sti*, *Stv*, *Nti* and *Ntv* in a genetic

codon table are 62, 72, 121 and 294, respectively (see Supplementary Table S1). Yet the

observed synonymous variations are always higher than the expected proportion across genes

(23-26) due to stronger purifying selection on non-synonymous variations than the synonymous

variations (15). Hence, by accounting the codon degeneracy as well as PTC, the *Sti*, *Stv*, *Nti* and

Ntv in the genetic codon table can be estimated theoretically for each codon (see Supplementary

Table S2).

The codon substitution model suggests that sites within codons evolve at different rates

and consequently they should not be equally treated (22,27). In the codon substitution model,

using likelihood approach the $\frac{ti}{tv}$ ratio was estimated in accordance with codon degeneracy class

which was implemented in estimating *dN/dS* in coding sequences (22,28,29). Recently, by

accounting $\frac{ti}{tv}$ substitution rate difference in codons as well as nonsense variations in PTC, the

dN/dS values have been improved (16). This finding implies that factors such as codon

degeneracy and PTC composition should be considered while analyzing the evolution of protein-

coding genes. The ratio of *ti* to *tv* has been studied for reliable estimation of sequence distance

and phylogeny reconstruction (30). Using maximum likelihood approach, the $\frac{ti}{tv}$ ratio (commonly

known as kappa *κ*) has been used to mainly compare the $\frac{ti}{tv}$ values among different species (30).

The parameter *κ* has not been implemented to compare coding sequences within a genome in the

context of their compositional differences of different degenerate classes before. As the selective

pressure on synonymous and non-synonymous variations are heterogeneous (31), it is important

to analyze the $\frac{ti}{tv}$ values among different genes within an organism by accounting synonymous

sites and non-synonymous sites regarding their compositional differences of different degenerate

codons.

Given that there is a significant variation of substitution rates among protein coding

sequences within a genome (32), an appropriate methodology for the estimation of *ti* to *tv* ratio is

required. We are presenting an improved equation that builds upon the research conducted by

Yasuo Ina in 1998 (33). Considering the differential selective constraint in synonymous and non-

synonymous variations it is crucial to separately analyze the synonymous $\frac{ti}{tv}$ and non-

synonymous $\frac{ti}{tv}$ to gain a comprehensive understanding of the overall $\frac{ti}{tv}$. The improved equations

used in this study are helpful to estimate the *ti* and *tv* rate differences in coding sequences by

accounting the compositional difference in terms of codon degeneracy and PTC. Though it is

known that TFD codons only undergo synonymous *ti*, the $\frac{Sti}{Stv}$ ratio among the coding sequences

in relation to their compositional differences (TFD, FFD and PTC) has not been reported

previously. Further, we observed a correlation between TFD: FFD and $\frac{Sti}{Stv}$ values and the PTC%

and $\frac{Nti}{Ntv}$ values. We tried to establish the biological significance of codon composition and $\frac{ti}{tv}$ ratio

across 2516 coding sequences of *E. coli*. Interestingly, coding sequences with similar

synonymous variation dynamics are found to have a profound variation regarding their non-

synonymous variation dynamics. The improved method developed in the present study may be

applied to study different $\frac{ti}{tv}$ ($\frac{Sti}{Stv}$ and $\frac{Nti}{Ntv}$) across coding sequences.

MATERIALS AND METHODS

Coding sequences of *E. coli* genome

In this study, we carried out a SNVs based computational analysis of 2516 coding sequences

across 157 strains of *E. coli* (34). The information regarding the dataset selected for the present

study is provided in Supplementary Table S3. Although the public database has ~4500 coding

sequences, we excluded coding sequences showing length variation and improper

alignment/annotation. Codons having variations in more than one position or variations leading

to ambiguous nucleotides were not considered in the analyses of the variations in this study. A

detailed procedure for estimating variation is represented in (see Supplementary Table S4) using

example of a hypothetical sequence. This procedure is based on intra-species genome sequence

comparison to find out single nucleotide variations in bacterial genomes. A reference sequence

has been generated considering the most frequent nucleotide occurrence at a site in the gene

sequence. The logic is that the most abundant nucleotide in a species is the ancestral nucleotide

in the species (9,16,35,36).The observed SNVs (SNV$_o$) at different positions of each codon were

categorized into *Sti$_o$*, *Stv$_o$*, *Nti$_o$* and *Ntv$_o$*, where S for synonymous, N for non-synonymous, *ti$_o$* for

observed transition and tv_o for observed transversion, and subsequently ti_o ($Sti_o + Nti_o$) and tv_o

(Stv_{o+} Ntv_o) for each gene were calculated (see Supplementary Table S5). After finding observed

variation values for a gene, we calculated the theoretically estimated variation values Sti_e, Stv_e

Nti_e, Ntv_e, ti_e and tv_e. We derived the reference sequence for each gene based on the methodology

explained (see Supplementary Table S3). The reference sequences were further used to find the

codon counts of each gene. Then we multiplied the count of each codon by the theoretically

estimated variation values for each codon. For example, if the codon frequency of GGG is

observed to be 10 in a gene, then the total theoretically estimated value for Sti, Stv, Nti and Ntv

are 10, 20, 20 and 40 (as estimated Sti, Stv, Nti and Ntv for each GGG is 1, 2, 2 and 4,

respectively). We then summed up the individual variation values for 61 sense codons to get the

total theoretically estimated individual substitution values (Sti_e, Stv_e, Nti_e and Ntv_e) for the gene

(see Supplementary Table S6). The values of Sti_o, Stv_o, Nti_o, Ntv_o, ti_o, and tv_o and the values of

Sti_e, Stv_e Nti_e, Ntv_e, ti_e, and tv_e of a coding sequence are then applied in the modified equation to

find out different $\frac{ti}{tv}$ ratio (see Supplementary Table S7). A detailed workflow is provided in

Figure 2.

Improved equations

We proposed an improved mathematical equation for the calculation of different $\frac{ti}{tv}$ ratio as

described below.

$$\frac{ti'}{tv'} = \frac{\left(ti_o/ti_e\right)}{\left(tv_o/tv_e\right)}$$ **Equation 1**

$$\frac{Sti'}{Stv'} = \frac{\left(Sti_o/Sti_e\right)}{\left(Stv_o/Stv_e\right)}$$ **Equation 2**

$$\frac{Nti'}{Ntv'} = \frac{\left(Nti_o / Nti_e \right)}{\left(Ntv_o / Ntv_e \right)}$$ **Equation 3**

Where, $\frac{ti'}{tv'}$ is improved $\frac{ti}{tv}$ ratio, $\frac{Sti'}{Stv'}$ is improved $\frac{Sti}{Stv}$ ratio, $\frac{Nti'}{Ntv'}$ is improved $\frac{Nti}{Ntv}$ ratio, ti_o is transition

observed, ti_e is the theoretically estimated transition, tv_o is transversion observed, tv_e is

transversion estimated, Sti_o is synonymous transition observed, Sti_e is the theoretically estimated

synonymous transition, Stv_o is synonymous transversion observed, Stv_e is the theoretically

estimated synonymous transversion, Nti_o is non-synonymous transition observed, Nti_e is the

theoretically estimated non-synonymous transition, Ntv_o is non-synonymous transversion observed,

Ntv_e is the theoretically estimated non-synonymous transversion. We calculated $\frac{ti}{tv}$ ratios for all the

genes using a program written in the <mark>C language. (to be modified later)</mark>

It is noteworthy that in a genome the number of potential replacements is greater for non-

synonymous alterations, whereas the frequency of actual variations is higher for synonymous

alterations among coding sequences. The improved estimators $\frac{ti'}{tv'}$, $\frac{Sti'}{Stv'}$, and $\frac{Nti'}{Ntv'}$ account for the

observed variations out of the total possible variations, on the contrary, the conventional method

only takes an account of the ti_o and tv_o. The possible sites for *Sti* and *Stv* are also different among

genes, which is dependent upon the codon composition belonging to different degeneracy classes.

The conventional $\frac{ti}{tv}$ ratio does not adequately account for the limitations of variable selection

against synonymous and non-synonymous variations. Hence, all these parameters make the

requirement of an improved estimator instead of the conventional estimator.

Comparative representation of different $\frac{ti}{tv}$ through conventional and improved method

We have done a comparative analysis of $\frac{ti}{tv}$ calculated using the existing and the improved

method to support the validation of method proposed in this study. Also, we have calculated the

TFD: FFD and the percentage of PTC (PTC%) of each gene and performed statistical analysis

between different parameters. A summary of the different parameters observed throughout the

dataset is represented in Table 1. We also calculated the percentage (%) change in $\frac{Sti}{Stv}$ and $\frac{Nti}{Ntv}$

values obtained from conventional to improved method as follows.

(a) $\frac{Sti}{Stv}$ % change $= \dfrac{\frac{Sti}{Stv} - \frac{Sti'}{Stv'}}{\frac{Sti}{Stv}} \times 100$

(b) $\frac{Nti}{Ntv}$ % change $= \dfrac{\frac{Nti}{Ntv} - \frac{Nti'}{Ntv'}}{\frac{Nti}{Ntv}} \times 100$

We found out $\frac{ti}{tv}$ in all these 2516 genes using maximum likelihood (ML) approach

available in MEGA-X (37). Using our improved approach, we computed the normalized *Sti, Stv*,

Nti, and *Ntv* values for the dataset. These normalized values were determined by dividing the

observed values by the corresponding estimated values in the dataset and further analyzed in the

variation dynamics study.

Statistical analyses

Comparison among the conventional, ML approach and the improved $\frac{ti}{tv}$ values were done using

different statistical methods to demonstrate the performance and accuracy of the improved

method. OriginPro, Version 2022, OriginLab Corporation, Northampton, MA, USA was used to

draw the Box-plot as well as to perform the Mann-Whitney test (38) to find out the *p*-value. The

correlation plot was drawn for R^2 value and Pearson's correlation coefficient (39) was also

calculated between the $\frac{ti}{tv}$ values obtained using the conventional and improved method.

<mark>**Software and programme**</mark>

Programme written in Python-language is available at GitHub in the following <mark>link.</mark> (). Further

queries regarding computer programme may contact SSS (ssankar@tezu.ernet.in).

RESULTS

Impact of codon degeneracy and PTC on $\frac{Sti}{Stv}$ and $\frac{Nti}{Ntv}$ values in *E. coli*

We found out the total number of TFD, FFD, and PTC codons across 2516 coding sequences of *E.*

coli. The percentage of TFD, FFD, and PTC was calculated (see Supplementary Table S3). The

percentage (%) compositional differences in the 2516 coding sequences represented as TFD%,

FFD%, and PTC% is shown in a box-plot (Figure 3). A variable range of TFD: FFD ratio in the

coding sequences of *E. coli* was observed. The minimum to maximum TFD:FFD ratio was observed

as 0.229 and 3.333 in *sugE* and *mntS*, respectively. It suggests that FFD codons are present

approximately four times more frequent than TFD codons in certain coding sequences. Similarly,

TFD codons are present approximately three times more frequent than FFD codons in the selected

coding sequences. Among 2516 coding sequences, only 35 coding sequences possessed the

TFD:FFD ratio of 1.000. Hence, this observation suggests that there is a compositional asymmetry

of TFD and FFD codons in the coding sequences. The minimum to maximum PTC% was observed

in the gene *atpE* as 15.00% and in the gene *chaB* as 57.14% (Figure 3). The compositional variation

(TFD, FFD and PTC) observed in coding sequences were considered to validate the accuracy of

different $\frac{ti}{tv}$ ratio using the conventional and improved method.

We estimated $\frac{ti}{tv}$ by accounting codon degeneracy as well as PTC in coding sequences and

compared with the conventional method of $\frac{ti}{tv}$ in coding sequences. Theoretically, a coding region

entirely composed of FFD codons will result into a twice $\frac{Sti'}{Stv'}$ value than $\frac{Sti}{Stv}$ value, similarly a coding

region composed of same proportion of FFD and TFD codons will result into equal values of $\frac{Sti'}{Stv'}$

and $\frac{Sti}{Stv}$. The increase in TFD codon proportion will directly impact the difference between $\frac{Sti}{Stv}$ values

and $\frac{Sti'}{Stv'}$ values. We have observed a similar pattern as well (Supplementary Figure S2). The

complete set of observed and estimated values of different variations is enlisted (see Supplementary

Table S3). These values were used to calculate the $\frac{ti}{tv}, \frac{Sti}{Stv}$ and $\frac{Nti}{Ntv}$. Further the improved equation

was used to calculate the $\frac{ti'}{tv'}, \frac{Sti'}{Stv'}$ and $\frac{Nti'}{Ntv'}$. We compared the values derived using the conventional

method ($\frac{ti}{tv}, \frac{Sti}{Stv}$ and $\frac{Nti}{Ntv}$) and improved method ($\frac{ti'}{tv'}, \frac{Sti'}{Stv'}$ and $\frac{Nti'}{Ntv'}$) (see Supplementary Table S3). A

box-plot showing the comparison between the values of different $\frac{ti}{tv}$ ratio obtained through the

conventional and improved method is represented in Figure 4. The minimum % change in $\frac{Sti}{Stv}$

between conventional and improved method was observed in the gene *garP* as -30.99% and

maximum in the gene *yfhL* as 30.76%. Similarly, the minimum % change in $\frac{Nti}{Ntv}$ between

conventional and improved method was observed in the gene *thiS* as -194.675% and maximum in

the gene *bhsA* as -129.600%. The detailed % change in $\frac{Sti}{Stv}$ and $\frac{Nti}{Ntv}$ is represented (see

Supplementary Table S3). While it is evident that, TFD codons and FFD codons have different

probabilities of obtaining synonymous variations, the composition of TFD: FFD was cross-checked

against the % change in $\frac{Sti}{Stv}$ for each gene. Likewise, the impact of PTC% in estimating the $\frac{Nti}{Ntv}$

values have been considered in this study.

The synonymous variation in the genes is likely to be different considering the differential

composition of TFD codons and FFD codons, because when FFD codon proportion is higher in a

gene, $\frac{ti}{tv}$ value is going to be lower. Similarly, when the FFD codon proportion is lower, $\frac{ti}{tv}$ value is

going to be higher. So, this observation is distinct, and the improved estimator is taking care of the

compositional differences (estimated numbers) of different degenerate codons. The Pearson

correlation coefficient value between TFD: FFD and % change in $\frac{Sti}{Stv}$ (conventional to improved)

was obtained to be 0.891, signifying a strong correlation. Similarly, the Pearson correlation

coefficient value between PTC% and % change in $\frac{Nti}{Ntv}$ was obtained to be -0.488, a moderate

negative correlation between the two variables. The results were found to be statistically significant

at $p<0.01$, for $\frac{ti}{tv}$ & $\frac{ti'}{tv'}$, $\frac{Sti}{Stv}$ & $\frac{Sti'}{Stv'}$ and $\frac{Nti}{Ntv}$ & $\frac{Nti'}{Ntv'}$. Further the correlation graph between % change in

$\frac{Sti}{Stv}$ and TFD: FFD and % change in $\frac{Nti}{Ntv}$ and PTC% also substantiate the role of codon composition

and $\frac{ti}{tv}$ ratio (Figure 5A & 5B). In the trendline analyses the % change in $\frac{Sti}{Stv}$ was perceived to be

directly influenced by the TFD: FFD with a R^2 value of 0.795. A weak negative association was

obtained between % change in $\frac{Nti}{Ntv}$ and PTC% with a R^2 value of 0.216. Hence, it is pertinent to note

that the overall % change in $\frac{Sti}{Stv}$ and $\frac{Nti}{Ntv}$ were directly influenced by the TFD and PTC codon

composition across the coding sequences. We performed a simulation study to substantiate the

impact of the compositional difference of degeneracy class in estimating $\frac{Sti}{Stv}$ by using our improved

method in coding sequences. For detailed methodology and results (see Supplementary Table S8,

Supplementary Figure S3). The results of comparative analysis between ML based method and

improved method are presented in supplementary files for 2516 genes (see Supplementary Table S9,

Supplementary Figure S4).

***Sti*, *Stv*, *Nti* and *Ntv* variation dynamics in coding sequences through improved approach**

A total of 2516 coding sequences were subjected to analysis for finding out the different variation

dynamics such as *Sti*, *Stv*, *Nti*, and *Ntv*. The mean *Sti*, *Stv*, *Nti*, and *Ntv* values were 0.113, 0.034,

0.014 and 0.004 respectively (see Supplementary Table S11, Figure 6). The mean *Sti*, *Stv*, *Nti*, and

Ntv values revealed the decreasing order of their variation dynamics, respectively. This comparison

of the values suggests that *Sti* is 3.3 times more frequent than *Stv*, while *Nti* is 3.7 times more

frequent than *Ntv*. Moreover, the ratio of *Sti* to *Nti* was 8.2, and the ratio of *Stv* to *Ntv* was 9.3. This

indicates a similar purifying selection on *Nti* and *Ntv* and this difference was found to be statistically

significant ($p<0.01$) between *ti* and *tv* of both synonymous and non-synonymous variations. Table 2

presents the Pearson correlation coefficient values for the pairs *Sti*, *Stv*, *Nti*, and *Ntv* in the 2516

coding sequences, obtained using both conventional and improved methods. It shows the

correlations between these variables under both approaches. By using the conventional approach, the

Pearson correlation between *Sti* and *Nti* was calculated to be 0.52, while the correlation between *Stv*

and *Ntv* was found to be 0.54. However, with the improved method, the correlation values decreased

substantially to 0.19 for *Sti* and *Nti*, and 0.25 for *Stv* and *Ntv*, respectively. Our method is more

appropriate in estimating the distinctive selection patterns between *Sti* and *Nti* even though the

pattern of substitution is similar in both. As part of our comparative study, we investigated the

coding sequences of similar *Sti* and different *Nti* values. In a set of 20 coding sequences with similar

Sti values such as 0.131, while the *Nti* values varied significantly, ranging from 0.000 to 0.030. This

wide range of *Nti* values, with a maximum 30-fold variation, highlights the presence of a variable

Nti pattern within the respective coding sequences. Additionally, these results suggest that different

substitution types undergo distinct selection processes.

DISCUSSION

Codon degeneracy influences the transition to transversion ratio in coding sequences because a

TFD codon undergoes synonymous variation only by transition while a FFD codon undergoes

synonymous variation by transition as well as transversion. In addition, there are only 18

pretermination codons in the genetic code table where non-sense substitution is rare to observe.

Coding sequences in a genome are with varying with composition of codons belonging to

different degeneracy as well as PTC. Therefore, it is inappropriate to compare coding sequences

regarding the $\frac{ti}{tv}$ ratio without accounting degeneracy and pretermination nature of codons. In this

study we have developed a methodology to estimate $\frac{ti}{tv}$ ratio by accounting degeneracy and

pretermination nature of codons, which has been implemented to analyze 2516 coding sequences

in *E. coli*. Our result suggested that coding sequences are different regarding their $\frac{ti}{tv}$ ratio, which

is independent of the composition of degenerate codons and PTC. In fact, the $\frac{Sti}{Stv}$ values are

observed to be variable among the coding sequences, which is intriguing. We did not observe any

significant correlation of $\frac{Sti}{Stv}$ values with the gene expression represented as CAI values. The role

of context dependent mutation (40-43), which has been described recently, might be a

contributing factor that needs to be investigated in future.

Our study on *ti* and *tv* has revealed that the base variation dynamics among the coding

sequences is in the decreasing order such as *Sti > Stv > Nti > Ntv*. The lower value of *Nti* and *Ntv*

indicates the role purifying selection (5) on non-synonymous variations. The strength of

purifying selection on *Nti* and *Ntv* seems comparable, which is unlike the report by Zou and

Zhang (44). Our study is on *E. coli* genes, whereas their research is on 90 clades representing all

domains of life. In our approach we have considered degeneracy and PTC, which has not been

implemented in their study.   As transitions are more frequent events than transversion, transition

can be considered alone to understand the selection on coding sequences. Considering *Sti*

representing the variation dynamics in a coding sequence, the *Nti* representing the strength of

purifying selection acting on the coding sequence, a comparison between *Sti* and *Nti* in a coding

sequence might represent the strength of purifying selection acting on it. In several instances, we

have observed that coding sequences with equal *Sti* dynamics are highly different regarding their

Nti dynamics, which indicates disproportionate purifying selection on coding sequences.

Various methods have been implemented in the estimation of the *ti* to *tv* ratio by

molecular evolutionary biologists to understand the mutation as well as selection bias. The

phylogeny-based method and Bayesian methods have been implemented by researchers in the

inter-species study of $\frac{ti}{tv}$ (45,46). Usually the *dN/dS* approach has been used to visualize the

selection in coding sequences (16) that considers both *ti* and *tv*. However, our method

implemented in this study is simple, succinct, easy to follow for researchers in the field with

limitations in mathematics and provides an insight to understand the evolution of protein coding

sequences. We also did a comparison with the ML method. There are several genes where we

could observe the $\frac{ti}{tv}$ values estimated by the improved method is more in concordance with the

observed *ti* and *tv* values than the ML method. It has been said that the likelihood methods are

computer intensive and difficult to apply to large datasets and can fall into local traps (47).

Furthermore, the extrapolation of our method in various inter-species studies, as well as

phylogenetic tools can enhance the horizon of our understanding regarding mutational bias in

coding sequences.

DATA AVAILABILITY

294 The authors confirm that the data supporting the findings of this study are available within the

295 article and Supplementary Data.

ACKNOWLEDGMENTS

PKB is grateful to Tezpur University for the institutional fellowship. RA is thankful for the

JRF fellowship from the DBT grant (BT/511/NE/TBP/2013) and (BT/403/NE/U-Excel/2013). PS is

grateful to UGC, GoI New Delhi, for the JRF. SD is thankful to DBT for MSc. Scholarship. SSS and

SKR are thankful to DBT, GoI for the twinning grant (BT/PR16361/NER/95/192/2015 date 18-10-

2016) to them. NDN, SSS & SKR thankfully acknowledge the DBT funded Bioinformatics and

Computational Biology Centre at Tezpur University.

Authors contributions

304 P.K.B, S.S.S. and S.K.R: Idea, design, executed, data analysis, writing, and discussion; P.S

305 and R.A.: executed, data analysis, and writing; S.D: writing and discussion. N.D.N, E.F:

306 Writing, Discussion, and critical analysis.

Author Disclosure Statement

308 The authors declare no potential conflict of interest. The authors also declare no competing

309 financial interests.

FUNDING

311 The authors received no specific funding for this work.

REFERENCE

1. Crick, F.H.C., Barnett, L., Brenner, S. and Watts-Tobin, R.J. (1961) General nature of the genetic code for proteins. *Nature*, **30**;192:1227-32. https://doi.org/ 10.1038/1921227a0.

2. Khorana, H.G., Büchi, H., Ghosh, H.,Gupta, N., Jacob, T. M., Kössel, H., Morgan, R., Narang, S.A., Ohtsuka, E. and Wells, R.D. (1966) Polynucleotide synthesis and the genetic code. Cold Spring Harbor Symp Quant Biol. **31**:39-49. https://doi.org/10.1101/sqb.1966.031.01.010.

3. Modiano, G., Battistuzzi, G. and Motulsky, A.G. (1981) Nonrandom patterns of codon usage and of nucleotide substitutions in human alpha-and beta-globin genes: an evolutionary strategy reducing the rate of mutations with drastic effects? *Proc Natl Acad Sci,* **78**:1110–1114. https://doi.org/ 10.1073/pnas.78.2.1110

4. van den Eynden, J., Basu, S. and Larsson, E. (2016) Somatic mutation patterns in hemizygous genomic regions unveil purifying selection during tumor evolution. *PLoS Genet,* **12**: e1006506. https://doi.org/10.1371/journal.pgen.1006506

5. Morales, A.C., Rice, A.M., Ho, A.T.l., Mordstein, C., Mühlhausen, S., Watson, S., Cano, L., Young, B., Kudla, G. and Hurst, L.D. (2021) Causes and consequences of purifying selection on SARS-CoV-2. *Genome Biol Evol,* **13**: evab196. https://doi.org/10.1093/gbe/evab196

6. Li, W-H., Gojobori, T. and Nei, M. (1981) Pseudogenes as a paradigm of neutral evolution. *Nature,* **292**:237–239. https://doi.org/10.1038/292237a0

7. Duchêne, S., Ho, S.Y.W. and Holmes, E.C. (2015) Declining transition/transversion ratios

through time reveal limitations to the accuracy of nucleotide substitution models. *BMC*

Evol Biol, **15**:1–10. https://doi.org/10.1186/s12862-015-0312-6

8. Stoltzfus, A. and Norris, R.W. (2016) On the causes of evolutionary transition: transversion

bias. *Mol Biol Evol,* **33**:595–602. https://doi.org/10.1093/molbev/msv274

9. Sen, P., Aziz, R., Deka, R.C., Feil. E. J., Ray, S.K. and Satapathy S.S. (2022) Stem region

of tRNA genes favors transition substitution towards keto bases in bacteria. *J Mol Evol,*

90:114–123. https://doi.org/10.1007/s00239-021-10045-x

10. Seplyarskiy, V.B., Kharchenko, P., Kondrashov, A.S. and Bazykin, G.A. (2012)

Heterogeneity of the transition/transversion ratio in Drosophila and Hominidae genomes.

Mol Biol Evol, **29**:1943–1955. https://doi.org/10.1093/molbev/mss071

11. Lewis Jr, C.A., Crayle, J., Zhou, S. Swanstrom, R. and Wolfenden, R. (2016) Cytosine

deamination and the precipitous decline of spontaneous mutation during Earth's history.

Proc Natl Acad Sci, **113**:8194–8199. https://doi.org/ 10.1073/pnas.1607580113.

12. Lyons, D.M. and Lauring, A.S. (2017) Evidence for the selective basis of transition-to-

transversion substitution bias in two RNA viruses. *Mol Biol Evol,* **34**:3205–3215.

https://doi.org/ 10.1093/molbev/msx251

13. Schroeder, J.W., Randall, J.R., Hirst, W.G., O'Donnel, M.E. and Simmons, L.A. (2017)

Mutagenic cost of ribonucleotides in bacterial DNA. *Proc Natl Acad Sci,* **114**:11733–

11738. https://doi.org/10.1073/pnas.1710995114

14. Kimura, M. (1980) A simple method for estimating evolutionary rates of base substitutions

through comparative studies of nucleotide sequences. *J Mol Evol,* **16**:111–120.

https://doi.org/ 10.1007/BF01731581

15. Hurst, L.D. and Pál, C. (2001) Evidence for purifying selection acting on silent sites in
BRCA1. *TRENDS in Genetics,* **17**:62–65. https://doi.org/ 10.1016/s0168-9525(00)02173-9

16. Aziz, R., Sen, P., Beura, P.K., Das, S., Tula, D., Dash, M., Namsa, N.D.N., Deka, R.C.,
Feil, E.J. Satapathy, S.S. *et al* (2022) Incorporation of transition to transversion ratio and
nonsense mutations, improves the estimation of the number of synonymous and non-
synonymous sites in codons. *DNA Research,* **29**: dsac023.
https://doi.org/10.1093/dnares/dsac023

17. Fersht, A.R. and Knill-Jones, J.W. (1981) DNA polymerase accuracy and spontaneous
mutation rates: frequencies of purine. purine, purine. pyrimidine, and pyrimidine.
pyrimidine mismatches during DNA replication. *Proc Natl Acad Sci,* **78**(7):4251–
4255.https://doi.org/10.1073/pnas.78.7.4251

18. Topal, M.D. and Fresco, J.R. (1976) Complementary base pairing and the origin of
substitution mutations. *Nature,* **263**(5575):285-289. https://doi.org/10.1038/263285a0.

19. Kino K, Sugiyama H (2001) Possible cause of G. C> C. G transversion mutation by
guanine oxidation product, imidazolone. *Chem. Biol*, 8, 369–378.
https://doi.org/10.1016/S1074-5521(01)00019-9

20. Rocha, E.P.C., Touchon, M. and Feil. E. J. (2006) Similar compositional biases are caused
by very different mutational effects. *Genome Res,* **16**, 1537–1547. https://doi.org/
10.1101/gr.5525106

21. Bhagwat, A. S., Hao, W., Townes, J.P., Lee, H., Tang, H. and Foster, P.L. (2016) Strand-
biased cytosine deamination at the replication fork causes cytosine to thymine mutations in
Escherichia coli. *Proc Natl Acad Sci,* **113**, 2176–2181.
https://doi.org/10.1073/pnas.1522325113

22. Muse, S.V. and Gaut, B.S. (1994) A Likelihood Approach for Comparing Synonymous and
Nonsynonymous Nucleotide Substitution Rates, with Application to the Chloroplast
Genome. *Mol Bio Evol,* **1 1(5)**:7 15-724. https://doi.org/
10.1093/oxfordjournals.molbev.a040152

23. Tamura, K. (1992) The rate and pattern of nucleotide substitution in Drosophila
mitochondrial DNA. *Mol Biol Evol,* **9**:814–825.
https://doi.org/10.1093/oxfordjournals.molbev.a040763

24. Wolfe, K.H., Sharp, P.M. and Li, W-H. (1989) Rates of synonymous substitution in plant
nuclear genes. *J Mol Evol,* **29**:208–211. https://doi.org/10.1007/BF02100204

25. Moriyama, E.N. and Powell, J.R. (1997) Synonymous substitution rates in Drosophila:
mitochondrial versus nuclear genes. *J Mol Evol,* **45**:378–391. https://doi.org/
10.1007/pl00006243

26. Ngandu, N.K., Scheffler, K., Moore, P., Woodman, Z., Martin, D. and Seoighe, C. (2008)
Extensive purifying selection acting on synonymous sites in HIV-1 Group M sequences.
Virol J, **5**:1–11. https://doi.org/ 10.1186/1743-422X-5-160

27. Shapiro, B., Rambaut, A. and Drummond, A.J. (2006) Choosing appropriate substitution
models for the phylogenetic analysis of protein-coding sequences. *Mol Biol Evol,* **23**, 7–9.
https://doi.org/10.1093/molbev/msj021

28. Goldman, N. and Yang, Z. (1994) A codon-based model of nucleotide substitution for
protein-coding DNA sequences. *Mol Biol Evol,* **11**(5), 725-736. https://doi.org/
10.1093/oxfordjournals.molbev.a040153.

29. Arenas, M. (2015) Trends in substitution models of molecular evolution. *Front Genet,*
6:319. https://doi.org/10.3389/fgene.2015.00319

30. Yang, Z. and Yoder, A.D. (1999) Estimation of the Transition/Transversion Rate Bias and Species Sampling. *J Mol Evol* **48**:274–283. https://doi.org/10.1007/PL00006470

31. Bartolomé. C., Maside. X., Yi. S., Grant. A.L. and Charlesworth, B. (2005) Patterns of selection on synonymous and nonsynonymous variants in Drosophila miranda. *Genetics,* **169**(3):1495-507. https://doi.org/10.1534/genetics.104.033068

32. Wu, C-I. and Li, W.H. (1985) Evidence for higher rate of nucleotide substitution in rodents than in man. *Proc Natl Acad Sci,* **82**(6):1741-1745. https://doi.org/ 10.1073/pnas.82.6.1741

33. Ina, Y. (1998) Estimation of the transition/transversion ratio. *J Mol Evol,* **46**:521–533. https://doi.org/10.1007/PL00006333

34. Thorpe, H.A., Bayliss, S.C. and Hurst, L.D. and Feil, E.J. (2017) Comparative analyses of selection operating on nontranslated intergenic regions of diverse bacterial species. *Genetics,* **206**:363–376. https://doi.org/10.1534/genetics.116.195784

35. Sen, P. (2022) Computational Analysis of Codon Usage Bias Single Nucleotide Polymorphism and RNA Secondary Structures in Microbial Genome Sequences. Ph.D. Thesis, Tezpur University http://hdl.handle.net/10603/473607

36. Beura, P.K., Sen, P., Aziz, R., Satapathy, S.S. and Ray, S.K. (2023) Transcribed intergenic regions exhibit a lower frequency of nucleotide polymorphism than the untranscribed intergenic regions in the genomes of Escherichia coli and Salmonella enterica. *J Genet,* **102**:22. https://doi.org/10.1007/s12041-023-01418-w

37. Kumar, S., Stecher, G., and Li, M., Knyaz, C. and Tamura, K. (2018) MEGAX: Molecular evolutionary genetics analysis across computing platforms. *Mol Biol Evol,* **35**(6):1547–1549 https://doi.org/ 10.1093/molbev/msy096

38. Mann, H.B. and Whitney, D.R. (1947) On a test of whether one of two random variables is

stochastically larger than the other. *Ann Math Statist,* **18**(1), 50-60. https://doi.org/

10.1214/aoms/1177730491

39. Pearson, K. (1896) Mathematical contributions to the theory of evolution. On a form of

spurious correlation which may arise when indices are used in the measurement of organs.

Proc R Soc Lond, **60**(359-367), 489-498. https://doi.org/10.1098/rspl.1896.0076

40. Sung, W., Ackerman, M.S., Gout, J.F., Miller, S.F., Williams, E., Foster, P.L. and Lynch,

M. (2015) Asymmetric context-dependent mutation patterns revealed through mutation-

accumulation experiments. *Mol Biol Evol.* 32:1672–1683. doi: 10.1093/molbev/msv055.

41. Zhu, Y., Neeman, T., Yap, V.B. and Huttley. G.A. (2017) Statistical methods for

identifying sequence motifs affecting point mutations. *Genetics.*

doi: 10.1534/genetics.116.195677.

42. Aikens. R.C., Johnson. K.E. and Voight. B.F. (2019) Signals of variation in human

mutation rate at multiple levels of sequence context. *Mol Biol Evol.* 36:955–965.

doi: 10.1093/molbev/msz023.

43. Ling. G., Miller, D., Nielsen, R. and Stern. A. A. (2020) Bayesian framework for inferring

the influence of sequence context on point mutations. *Mol Biol Evol.* 37:893–903.

doi: 10.1093/molbev/msz248.

44. Zou, Z. and Zhang, J. (2021) Are Nonsynonymous Transversions Generally More

Deleterious than Nonsynonymous Transitions? *Mol Biol Evol,* **38**(1):181–191.

https://doi.org/ 10.1093/molbev/msaa200

45. Purvis, A. and Bromham, L. (1997) Estimating the transition/transversion ratio from independent comparisons with an assumed phylogeny. *J Mol Evol,* **44**, 112–119. https://doi.org/ 10.1007/pl00006117

46. Huelsenbeck, J. and Ronquist, F. (2001) MrBayes: Bayesian inference of phylogenetic trees. *Bioinform* **17**, 754–755. https://doi.org/ 10.1093/bioinformatics/17.8.754.

47. Golding, B. and Felsenstein, J. (1990) A maximum likelihood approach to the detection of selection from a phylogeny. *J Mol Evol,* **31**:511-523. https://doi.org/10.1007/bf02102078

Figure_legend_Table

Figure 1. Schematic representation of the different substitutions occurring in the DNA. The transition *(ti)* is marked with blue arrows and *(tv)* is marked with orange arrows. A and G are categorized under purine (R) and C and T are categorized under pyrimidine (Y). Out of the twelve possible base substitutions, four are *ti* in which a purine (or pyrimidine) nucleotide is replaced by

another purine (or pyrimidine) nucleotide, and eight are *tv* in which a purine (or pyrimidine)

nucleotide is replaced by another pyrimidine (or purine) nucleotide.

Figure 2. Schematic representation of the workflow showing the steps to execute our improved

method. By considering each codon as j, ti_e^j and tv_e^j can be calculated for each codon as presented

in Supplementary Table S2. The reference sequence needs to be calculated from the coding region

(gene/query). From the reference sequence the observed ($_o$) substitutions are calculated for the gene.

The other usage of the reference sequence is to find out the codon count of the gene. Then, the

respective codon count is multiplied by estimated transition and transversion for each codon. The

next step is to find out the sum of *ti* and sum of *tv* in all codons to represent the estimated ($_e$) *ti* and *tv*

for the gene. After getting the observed and estimated values for *ti* and *tv*, now the improved

equation is applied to find out the $\frac{ti'}{tv'}$ of the gene. This scenario is only showing the $\frac{ti'}{tv'}$ ratio, whereas

in similar way the $\frac{Sti'}{Stv'}$ & $\frac{Nti'}{Ntv'}$ can be calculated.

Figure 3. The box-plot presents the percentage (%) of TFD, FFD and PTC codons in the coding

sequences considered to validate the improved method proposed in this study. The minimum and

maximum TFD% in the dataset was observed as 10.38% and 53.25%, similarly the minimum and

maximum FFD% in the dataset was observed as 13.95% and 52.05%. The minimum and maximum

PTC% in the dataset was observed as 15.00% and 57.14%.

Figure 4. The box-plot presents the different $\frac{ti}{tv}$ ratio values. Here we have compared the different $\frac{ti}{tv}$ ratios of the entire set of genes through conventional and improved method. $\frac{ti}{tv}$ shows the conventional method and $\frac{ti'}{tv'}$ shows the result of improved method. Similarly, $\frac{Sti}{Stv}$ and $\frac{Nti}{Ntv}$ represent the conventional values and $\frac{Sti'}{Stv'}$ and $\frac{Nti'}{Ntv'}$ represent the improved values. This illustration shows that the median values between $\frac{ti}{tv}$ & $\frac{ti'}{tv'}$, $\frac{Sti}{Stv}$ & $\frac{Sti'}{Stv'}$ and $\frac{Nti}{Ntv}$ & $\frac{Nti'}{Ntv'}$ are statistically different (*p*-value < 0.01).

A

B

Figure 5. Regression plot of coding sequence considered to validate the improved method. The x-axis shows the codon composition (TFD: FFD and PTC%) and y-axis shows the % change in $\frac{Sti}{Stv}$

and $\frac{Nti}{Ntv}$. (A) % change in $\frac{Sti}{Stv}$ is showing a strong correlation with the TFD: FFD with a R^2 value of

0.795. (B) % change in $\frac{Nti}{Ntv}$ is showing a weak negative correlation with the PTC% with a a R^2 value

of 0.216.

Figure 6. The figure illustrates the box-plot of the normalized *Sti*, *Stv*, *Nti*, and *Ntv* values. The mean

Sti and *Stv* were observed as 0.113 and 0.031 respectively, similarly the mean *Nti* and *Ntv* values

were observed as 0.014 and 0.004. *Sti* was observed to be highly frequent than *Stv* and *Nti* was

observed to be highly frequent than *Ntv*. The pairs such as (*Sti* and *Stv)* and (*Nti* and *Ntv)* values

show a statistically significant difference ($p<0.01$).

514   **Table 1.** Summary of the 2516 coding sequences selected for the present study. It represents the
515   minimum to maximum range of different parameters like GC%, TFD/FFD, PTC% and the mean and
516   standard deviation (S.D.) of the TFD%, FFD% and PTC%

Genes considered	Size bp (Min-Max)	GC% (Min-Max)	TFD/FFD (Min-Max)	PTC% (Min-Max)	TFD%	FFD%	PTC%	
					33.384	34.084	36.350	Mean
2516	93-4554	32.95-63.80	0.22-3.33	15.00-57.14	5.726	4.761	5.625	S. D

519   **Table 2.** Comparative Pearson correlation coefficient study among Sti, Sti, Nti and Ntv through
520   conventional and improved estimator

Pearson correlation coefficient	Conventional method	Improved method
	Observed values	Normalized values (Obs/Est)
Sti & Stv	0.920	0.795
Nti & Ntv	0.789	0.612
Sti & Nti	0.527	0.192
Stv & Ntv	0.544	0.257