Polygenic prediction across populations is influenced by ancestry, genetic architecture, and methodology

Ying Wang[1,2,*], Masahiro Kanai[1,2,3,4], Taotao Tan[5], Mireille Kamariza[6], Kristin Tsuo[1,2], Kai Yuan[1,2], Wei Zhou[1,2], Yukinori Okada[4,7,8,9], the BioBank Japan Project, Hailiang Huang[1,2], Patrick Turley[10], Elizabeth G. Atkinson[5], Alicia R. Martin[1,2,*]

1. Analytic and Translational Genetics Unit, Massachusetts General Hospital, Boston, MA 02114, USA
2. Stanley Center for Psychiatric Research and Program in Medical and Population Genetics, Broad Institute of MIT and Harvard, Cambridge, MA 02142, USA
3. Department of Biomedical Informatics, Harvard Medical School, Boston, MA, USA
4. Department of Statistical Genetics, Osaka University Graduate School of Medicine, Suita, Japan
5. Department of Molecular and Human Genetics, Baylor College of Medicine, Houston, TX, USA
6. Society of Fellows, Harvard University, Cambridge, MA, 02138 USA
7. Laboratory for Systems Genetics, RIKEN Center for Integrative Medical Sciences, Yokohama, Japan
8. Laboratory of Statistical Immunology, Immunology Frontier Research Center (WPI-IFReC); Center for Infectious Disease Education and Research (CiDER); and Integrated Frontier Research for Medical Science Division, Institute for Open and Transdisciplinary Research Initiatives, Osaka University, Suita 565-0871, Japan
9. Department of Genome Informatics, Graduate School of Medicine, the University of Tokyo, Tokyo 113-0033, Japan.
10. Department of Economics, and Center for Economic and Social Research, University of Southern California, Los Angeles, CA, USA

Summary

Polygenic risk scores (PRS) developed from multi-ancestry genome-wide association studies (GWAS), PRS$_{multi}$, hold promise for improving PRS accuracy and generalizability across populations. To establish best practices for leveraging the increasing diversity of genomic studies, we investigated how various factors affect the performance of PRS$_{multi}$ compared to PRS constructed from single-ancestry GWAS (PRS$_{single}$). Through extensive simulations and empirical analyses, we showed that PRS$_{multi}$ overall outperformed PRS$_{single}$ in understudied populations, except when the understudied population represented a small proportion of the multi-ancestry GWAS. Notably, for traits with large-effect ancestry-enriched variants, such as mean corpuscular volume, using substantially fewer samples from Biobank Japan achieved comparable accuracies to a much larger European cohort. Furthermore, integrating PRS based on local ancestry-informed GWAS and large-scale European-based PRS improved predictive performance in understudied African populations, especially for less polygenic traits with large ancestry-enriched variants. Our work highlights the importance of diversifying genomic studies to achieve equitable PRS performance across ancestral populations and provides guidance for developing PRS from multiple studies.

Keywords: genome-wide association studies; multi-ancestry; polygenic risk scores; genetic architecture

Lead contact: Ying Wang (yiwang@broadinstitute.org)
*Correspondence: yiwang@broadinstitute.org, armartin@broadinstitute.org

Introduction

Introduction

Polygenic risk scores (PRS) have emerged as useful tools for estimating the cumulative genetic
susceptibility to complex traits and diseases. PRS are typically calculated by weighting the
number of risk alleles based on their associations in genome-wide association studies (GWAS).
PRS have shown promising potential in predicting some traits and disease risks, comparable to
monogenic variants and traditional clinical risk factors[1–5]. Achieving the most accurate and
generalizable PRS requires access to large-scale and diverse GWAS, especially with
representation that matches the specific target population. However, the current landscape of
GWAS predominantly focuses on European (EUR) ancestry populations, which have
considerably larger sample sizes compared to other populations. Although ongoing efforts are
underway to rectify these gaps, achieving global representativeness is a challenging goal.
Encouragingly, studies have shown that using GWAS data with even a small proportion of non-
European ancestry individuals has the potential to improve the predictive accuracy of PRS in
underrepresented populations[6–8]. This finding could largely be attributed to the substantial
contribution of common variants to the heritable variation of complex traits and diseases, and that
causal variants are largely shared across ancestries[9–12]. With the ever-increasing availability and
scalability of genomic data from underrepresented and ancestrally diverse populations, we are
especially interested in leveraging this greater diversity to improve PRS generalizability.
60

In particular, recently admixed populations, consisting of chromosomal segments of mosaic
ancestries, are systematically excluded in many existing genomic studies due to their
underrepresentation and complicated population structure[13–15]. However, these populations
present unique opportunities to develop more generalizable PRS as their genetic effects can be
estimated in more consistent environments, which helps reduce confounding factors compared
to estimates across different ancestry groups in different populations[16]. Furthermore, the
comprehensive characterization of phenotypes is often insufficient or inconsistently performed in
different populations. However, in the recently admixed populations, there is a greater potential
for consistency and comparability in phenotype measurements, as the genetic factors contributing
to phenotypic differences between the source populations can be decoupled in the recently
admixed populations[16,17]. The advancement of methodologies such as local ancestry inference
and association testing has further enabled ancestry-specific GWAS in admixed populations[18–20],
allowing for the construction of PRS that leverage genetic information captured by local ancestry
inference. With the ongoing accumulation of data from recently admixed populations, particularly
through initiatives like the *All of Us* Research Program[21], expanded resources will provide
unparalleled opportunities to explore the performance of PRS derived from local ancestry-
informed summary statistics within historically underrepresented populations. Furthermore, such
data will facilitate their integration with PRS derived from predominantly EUR-based cohorts.
79

Recently developed statistical methodologies leverage the increasing diversity of GWAS data to
improve PRS portability[8,22,23]. However, the effect of genetic architecture, ancestry composition of
GWAS discovery cohorts, and PRS construction methodologies on cross-ancestry predictive

accuracy remains largely unclear. For example, a recent study found no increase in accuracy
when meta-analyzing GWAS from a relatively small Ugandan cohort with larger EUR data[24].
Furthermore, theoretical frameworks for approximating expected PRS accuracy from multi-
ancestry GWAS are lacking. Current theoretical calculations for PRS accuracy rely on the
assumption of homogeneity within the ancestral discovery samples[25,26], ignoring factors that are
likely to play a role with multi-ancestry cohorts. Such factors may include differences in linkage
disequilibrium (LD), minor allele frequency (MAF), heritability, sample sizes, and genetic
correlation across different ancestries.

To provide insights into those issues, we explored the impact of ancestry compositions in
discovery GWAS on predictive accuracy of PRS constructed using different methodologies. This
exploration involved large-scale population genetic simulations as well as the utilization of real
genomic data from the BioBank Japan (BBJ)[27] and UK Biobank (UKBB)[6] across traits exhibiting
distinct genetic architectures (**Figure 1)**. In what follows, we used **single-ancestry GWAS** to
denote studies conducted exclusively within a single ancestry group (defined using genetic data),
while **multi-ancestry GWAS** refers to studies encompassing two or more distinct ancestries. In
our analyses, we performed meta-analyses of GWAS conducted in European ancestry
populations (**EUR GWAS)** and GWAS conducted in other minority populations (**Minor GWAS**) by
varying the ratios of sample sizes to mimic multi-ancestry GWAS with varying ancestry
compositions. Specifically, we focused on East-Asian (EAS) and African (AFR) minority
populations. By comparing the performance of PRS derived from single-ancestry GWAS (referred
to as **PRS$_{single}$**) and multi-ancestry GWAS (referred to as **PRS$_{multi}$**) through simulations and real
data, we consistently observed that PRS$_{multi}$ overall exhibited superior performance in comparison
to PRS$_{single}$ (primarily PRS derived from large-scale EUR GWAS, referred to as **PRS$_{EUR\_GWAS}$**). As
admixed populations remain understudied despite disproportionately yielding novel genetic
findings[28], we further conducted local ancestry inference to explore whether, how, and to what
extent PRS performance could be improved using GWAS discovery data from AFR-EUR admixed
individuals. While optimal PRS methods are trait- and context-specific, this study
comprehensively evaluates PRS accuracy across a wide range of scenarios, facilitating a set of
best practices that ultimately reduces the number of analyses necessary to optimize PRS for
specific applications.

Results

Evaluating the effects of imbalanced sample sizes across ancestries on PRS accuracy through simulations

We simulated genotypes using HapGen2 and phenotypes according to six different scenarios
with varying trait heritability (h^2 = 0.03, 0.05) and number of causal variants (M_c = 100, 500, 1000),
such that the polygenicity ranged from ~0.1% to ~1%. We assumed that the causal variants and
their effect sizes are shared across ancestries (i.e., cross-ancestry genetic correlation, r_g, is 1) in
our initial simulations. For single-ancestry GWAS, we first conducted GWAS within each bin and

then meta-analyzed GWAS across different numbers of bins (1-52 per ancestry). Each bin
represented 10,000 individuals randomly sampled from the corresponding ancestry. For multi-
ancestry GWAS, we meta-analyzed GWAS from EUR and minor populations (EAS or AFR) to
evaluate the impact of ancestry composition. We used varying numbers of bins from the EUR
GWAS (ranging from 4 to 52 with 4 increments) and varied the contribution from minority
populations (1-52 bins) from EAS or AFR GWAS. We constructed PRS using the classic pruning
and thresholding (P+T) method with varying p-value thresholds. This approach follows a greedy
heuristic algorithm wherein variants are sorted based on their p-values. The algorithm iteratively
descends in significance while retaining only those variants that do not exceed a predetermined
LD threshold with previously retained variants. We assessed the accuracy, measured by
prediction R^2, using the optimal threshold through fine-tuning in the validation cohort. Detailed
information about the simulation setup is shown in **Figure 1** and **STAR Methods**.
135

PRS predictive accuracy improved with more individuals from target populations included
in the multi-ancestry GWAS but varied with genetic architecture

When developing PRS using single-ancestry GWAS, we found that using ancestry-matched
GWAS generally outperformed using GWAS from other discovery populations (**Figure S1**).
Compared to using EUR GWAS, the benefit of using ancestry-matched GWAS was more evident
for traits with more polygenic genetic architectures and larger GWAS sample sizes. To further
evaluate the impact of ancestry composition, we compared the accuracy of PRS$_{multi}$ and PRS$_{single}$.
We constructed PRS$_{multi}$ using an LD reference panel consisting of individuals proportional to the
ancestry composition of the discovery GWAS. This reference panel yielded approximately optimal
accuracy among three different reference panels utilized in our study (**Figure S2 and**
Supplementary Note 1, 2).
147

Relative to the accuracy of PRS$_{EUR\_GWAS}$, we observed significant improvements in the
understudied target population by including more individuals from the target ancestry in multi-
ancestry GWAS. Across all simulations, a statistically significant median improvement of 0.008 in
R^2 was observed (one-sided Wilcoxon signed-rank test, p-value < 2.2e-16, **Table S1**). This trend
was more apparent in more polygenic traits. As shown in **Figure 2**, we compared accuracy
between PRS$_{multi}$ and PRS$_{EUR\_GWAS}$ derived from 320,000 EUR individuals. For traits with h^2 of
0.05, the median improvements in R^2 of PRS$_{multi}$ was 0.006, 0.014 and 0.013 with M_c of 100, 500,
and 1000, respectively, in the EAS target population. Similarly, corresponding R^2 improvements
of 0.009, 0.010 and 0.014 were shown in AFR (**Figure S3**). However, we did not consistently
observe such accuracy gains for the majority EUR population, or in scenarios where the other
understudied ancestry was not included in the multi-ancestry discovery GWAS. In our simulations
but unlike in most GWAS, populations typically understudied in current genomic studies can be
the majority in the discovery GWAS. Nevertheless, we still observed significant PRS accuracy
improvements, of median improvements in R^2 0.007 across simulations when the proportion of
understudied populations in the discovery GWAS was less than 50% (one-sided Wilcoxon signed-
rank test, p-value < 2.2e-16). We expected to observe similar relative R^2 improvements, which
measured the PRS generalizability, in the target populations using PRS$_{multi}$ compared to using
PRS$_{EUR\_GWAS}$ with the same number of bins from EUR populations (**Supplementary Note 3**).

Compared with using PRS_{EUR_GWAS}, we found that PRS_{multi} derived from GWAS with much smaller
sample sizes could achieve comparable or better predictive accuracy (**Table S1**). For example,
in the scenario with M_c of 1000 and h^2 of 0.03, the meta-analysis of 16 EUR and 2 AFR bins
achieved a comparable accuracy of 0.008 to that of using 32 EUR bins in the AFR population.
Overall, adding fewer individuals from the target populations saturated accuracy improvements
for less polygenic traits faster than more polygenic traits. Similarly, larger sample sizes from AFR
populations were required to achieve comparable PRS accuracy to EAS populations especially
for more polygenic traits, likely due to the larger effective population size in AFR populations and
larger genetic divergence between EUR and AFR populations. As shown in **Figure S3**, when h^2
was 0.03, the accuracy improvement of PRS_{multi} in the AFR population plateaued to ~0.005 with
11 and 20 AFR bins for M_c of 100 and 500, respectively, but continued to increase with more AFR
bins for M_c of 1000. Similarly, when h^2 was 0.03, including 2 and 12 EAS bins in PRS_{multi} yielded
an accuracy improvement of >0.005 in EAS for M_c of 100 and 500, respectively (**Figure 2**). In
comparison to PRS derived from Minor GWAS alone (PRS_{Minor_GWAS}), we found that the accuracy
improvement of PRS_{multi} gradually diminished as the sample size of Minor GWAS increased
(**Figure S4 and Table S1**). We showed that for more polygenic traits, PRS_{multi} achieved little to
no improvement when the understudied target populations accounted for more than half of the
sample size in multi-ancestry GWAS (**Supplementary Note 4**).

Because genetic correlation estimates between populations can be significantly less than 1, we
also modified our simulations by varying the r_g to be 0.6 and 0.8. We investigated two simulation
scenarios that represent the extremes in per-variant variance explained: the least polygenic
scenario 1 with $M_C = 100$ and $h^2 =0.05$, and the most polygenic scenario 2 with $M_C = 1000$ and
$h^2 =0.03$ (**STAR Method**). Consistent with our previous findings, PRS_{multi} exhibited improved
predictive accuracy in the target population when a greater number of individuals from the same
ancestry were included, as compared to relying solely on large-scale EUR GWAS (**Figure S5-A,
B**). This improvement was more pronounced for scenario 2. Moreover, we needed a larger
number of individuals from the target ancestry to saturate accuracy improvements in scenario 1
when r_g was moderately reduced. Furthermore, as the sample sizes of the Minor GWAS
increased and the values of r_g decreased, the advantage of utilizing PRS_{multi} over PRS_{Minor_GWAS}
diminished and eventually vanished (**Figure S5-C, D**). Details are shown in **Table S2** and
S**upplementary Note 5.**

200 Empirical analysis of PRS accuracy within and across ancestries using 17
201 quantitative phenotypes

202 Genetic architecture of 17 studied phenotypes

203 To understand how trait genetic architecture influences predictive accuracy of PRS across
204 ancestries, we conducted a comprehensive analysis involving 17 phenotypes in the UKBB and
205 BBJ. Specifically, we estimated key parameters influencing different aspects of genetic
206 architecture, including SNP-based heritability, polygenicity (the proportion of SNPs with nonzero

effects) and a coefficient of negative selection (S, measuring the relationship between MAF and estimated effect sizes). To obtain these estimates, we employed a Bayesian method called summary-data-based BayesS (SBayesS), which leverages GWAS summary statistics as input data[29].

The phenotypes included in this study varied widely in genetic architecture across these estimated parameters (**Figure 3**, **Table S3 and Table S4**). The polygenicity estimates spanned a broad range, from low values (0.001-0.005) for traits like mean corpuscular hemoglobin concentration (MCHC), basophil count (basophil), mean corpuscular hemoglobin (MCH), and mean corpuscular volume (MCV), to higher values (0.012-0.047) for traits such as height and body mass index (BMI). SNP-based heritability estimates similarly ranged from <0.1 for basophil and MCHC to 0.54 and 0.33 for height using UKBB and BBJ, respectively, regardless of polygenicity. The median S parameters were -0.63 and -0.47 using UKBB and BBJ, respectively. While the negative S values indicate negative selection (i.e., rarer variants have larger effects), it remains unclear to what degree population stratification could confound such estimates[30,31]. We found that the polygenicity estimates using UKBB were mostly higher than those using BBJ, which could be due to the higher statistical power with larger sample sizes in the UKBB resulting in the detection of more variants with small effects. Similarly, we observed significantly higher SNP-based heritability in the UKBB compared to BBJ with the exception of MCHC and basophil, indicating possible phenotype heterogeneity between the two cohorts. These results are expected from the biobank designs, as BBJ is a hospital-based cohort with participants recruited with certain diseases, whereas UKBB is a population-based cohort with overall healthier participants and thus a wider range of natural variation in complete blood counts. This finding is also consistent with the previous study using estimates from LD score regression (LDSC) and stratified-LDSC[32]. Moreover, as described previously[32], the estimated cross-ancestry genetic correlations between UKBB and BBJ for those traits were not statistically different from 1 (p-value > 0.05/17) except for a few including basophil (r_g = 0.5945, SE = 0.1221), height (r_g = 0.6932, SE = 0.0172), BMI, (r_g = 0.7474, SE = 0.0230), diastolic blood pressure (DBP, r_g = 0.8354, SE = 0.0509), and systolic blood pressure (SBP, r_g = 0.8469, SE = 0.0430).

Multi-ancestry GWAS-derived PRS usually improves predictive performance compared to single-ancestry GWAS-derived PRS

We constructed PRS_{single} using the P+T and PRS-CS methods with GWAS from UKBB and BBJ, respectively. The GWAS sample sizes varied based on the number of Bin_{Total}, which represented the total number of bins specific to each trait as shown in **Table S3.** Each bin consisted of 5,000 individuals randomly selected from the respective cohort. We found that employing target ancestry-matched GWAS, even with smaller sample sizes, yielded comparable accuracy to utilizing large-scale EUR GWAS but depended on PRS methodology and trait-specific genetic architecture (**Figure S6, Figure S7, Table S5** and **Supplementary Note 6**). We evaluated predictive accuracy by computing incremental R^2 using linear regression, while accounting for the potential impact of covariates (**STAR Methods**).

For comparison, we developed PRS_{multi} using both P+T and PRS-CS, where we meta-analyzed
single-ancestry GWAS from UKBB and BBJ. Similar to the simulation setup, we mimicked
proportional ancestry composition in the multi-ancestry GWAS by meta-analyzing EUR GWAS in
the UKBB with GWAS in the BBJ while varying number of bins (each bin of 5,000 individuals,
UKBB bins ranging from 8 to 64 with an increment of 8, see **STAR Methods** and **Figure 1**). The
ratio of EUR/EAS samples in the multi-ancestry GWAS varied from 64:1 to $8/Bin_{Total}$. Thus, 85%
of the multi-ancestry GWAS had a higher proportion of EUR samples (>50% EUR). Consistent
with our findings from the simulations, where we observed that the choice of LD reference panel
had a limited impact on the predictive accuracy of more polygenic traits, we observed only a slight
improvement of median R^2 of 0.002 for P+T when employing a combined LD reference panel that
was proportional to the ancestries represented in the multi-ancestry GWAS. We compared this
result with PRS developed using a reference panel that was matched with the majority population
of the discovery GWAS (**Figure S8 and Table S6**). Because the majority of PRS was constructed
from GWAS predominantly composed of EUR individuals, we hereafter reported the results using
1KG-EUR as the LD reference.

In our analysis comprising 3,160 comparisons between single-ancestry PRS derived from UKBB
GWAS (PRS_{EUR_GWAS}) and multi-ancestry PRS (PRS_{multi}), we observed encouraging results.
Specifically, in the UK Biobank East-Asian population (UKBB-EAS), PRS_{multi} showed accuracy
improvements in 99.7% and 92.4% of these comparisons when using P+T and PRS-CS,
respectively (**Table S7 and Figure S9**). Accuracy increased with more EAS samples in the multi-
ancestry GWAS (**Figure 4**). For example, when comparing PRS_{multi} with PRS_{EUR_GWAS} using P+T,
the largest relative improvements in R^2 were 80.9% (0.085 vs. 0.047) for platelet count (PLT),
152.2% (0.058 vs. 0.023) for BMI and 91.9% (0.071 vs. 0.037) for height. We observed these
improvements when using multi-ancestry GWAS including EAS bins from BBJ, which were either
concordant with or proximal to Bin_{Total}, along with 64 EUR bins from UKBB. Similarly, the
corresponding relative R^2 improvements for these same three traits were 19.8% (0.0126 vs.
0.101), 50.0% (0.075 vs. 0.050) and 15.5% (0.097 vs. 0.084) when using PRS-CS. We did not
consistently observe the upward trend for white blood cell count (WBC) with PRS-CS, which can
be attributed to the lack of accuracy improvement with larger sample sizes of BBJ (**Figure S6**).
We also found that P+T showed greater improvement compared to PRS-CS but worse accuracy
overall, regardless of the number of bins from EUR GWAS; the median improvements in R^2 across
traits were 0.014 and 0.008, respectively. However, the upward trend in PRS accuracy was not
consistently shown in the UKBB-EUR, particularly when using PRS-CS (**Figure S10 and Table
S7**). This pattern aligned with our simulation results and previous reports that PRS accuracy for
minority populations included in the multi-ancestry GWAS benefited more from adding more
ancestry-matched individuals compared to other populations, including EUR populations[33]. We
noted that the accuracy of PRS_{multi} remained largely unchanged or slightly decreased when the
number of bins from BBJ was small (e.g., 1 or 2 bins), which was consistent with previous
studies[6,33]. In contrast to PRS derived from BBJ (PRS_{Minor_GWAS}), we noted a diminishing trend in
accuracy improvements of PRS_{multi} as the sample sizes of BBJ increased, especially for traits such
as height, PLT, MCH and MCV (**Figure S11**). Furthermore, we observed greater variation in
accuracy among traits from real data compared to simulations, which could be attributed to the
smaller sample sizes and the more complicated genetic architecture.

PRS derived from meta-analyzed multi-ancestry GWAS versus weighted PRS from single-ancestry GWAS in understudied populations

In contrast to PRS_{multi}, an alternative approach proposed in previous studies to enhance predictive
accuracy in diverse populations is the linear combination of PRS derived from GWAS conducted
on populations with different ancestries[34]. Here, we implemented this approach by developing a
weighted PRS (**$PRS_{weighted}$**) using P+T and PRS-CS. This combination involved linearly weighting
PRS derived from single-ancestry GWAS conducted in the UKBB and BBJ. Additionally, we
employed a more advanced Bayesian method called PRS-CSx[8], which jointly models GWAS and
LD information from multiple populations. Similarly, we constructed $PRS_{weighted}$ using ancestry-
specific posterior SNP effects. Furthermore, we developed PRS by integrating ancestry-specific
posterior SNP effects using the inverse-variance weighted meta-analysis strategy, also referred
to as PRS_{multi} (see **STAR Methods**).

Among the three PRS methods evaluated in the UKBB-EAS, PRS-CSx exhibited the highest
performance, followed by PRS-CS and P+T. Specifically, for PRS_{multi}, the corresponding median
R^2 values across traits were 0.051, 0.048 and 0.037, while for $PRS_{weighted}$, they were 0.051, 0.045
and 0.021, respectively (**Figure 5, Table S8 and Table S9**). Notably, we observed that PRS_{multi}
for BMI using PRS-CS yielded significantly better accuracy compared to PRS-CSx (median R^2:
0.057 vs. 0.055, p-value < 2.2e-16, one-sided Wilcoxon signed-rank test). Out of the 3,160
comparisons between PRS_{multi} and $PRS_{weighted}$ in the UKBB-EAS, 91.4% and 78.0% showed higher
accuracy of PRS_{multi} when using P+T and PRS-CS, respectively, with median improvements in
R^2 of 0.011 (p-value < 2.2e-16) and 0.003 (p-value < 2.2e-16). Although we found better
performance overall with PRS_{multi}, we found that $PRS_{weighted}$ significantly outperformed PRS_{multi} for
PLT using P+T (median R^2: 0.086 vs. 0.081, p-value < 2.2e-16) and for height using PRS-CS
(median R^2: 0.091 vs. 0.082, p-value = 2.6e-04). Contrary to trends observed with other methods,
in 59.7% of the comparisons, $PRS_{weighted}$ outperformed PRS_{multi} when using PRS-CSx, although
we observed no significant accuracy difference across traits. However, $PRS_{weighted}$ showed
superior performance compared to PRS_{multi} (p-value < 0.05/17) for several traits, including MCV
(median R^2: 079 vs. 0.072), MCH (median R^2: 0.079 vs. 0.073), Basophil (median R^2: 0.010 vs.
0.007) and hemoglobin concentration (HB, median R^2: 0.025 vs. 0.024).

Moreover, the extent of accuracy improvements using PRS_{multi}, in contrast to $PRS_{weighted}$, largely
varied across traits and ancestry compositions. For example, when evaluating accuracy within
the UKBB-EAS using P+T, we observed 3.25-fold increase in R^2 with PRS_{multi} compared to
$PRS_{weighted}$ for monocyte count (monocyte, 0.065 vs. 0.020). This improvement was achieved with
a bin ratio 56:15 for the discovery GWAS, consisting of 56 bins from UKBB and 15 bins from BBJ.
Similarly, using a bin ratio of 40:25, we achieved a 4-fold increase in R^2 for DBP (0.048 vs. 0.012)
with PRS_{multi} compared to $PRS_{weighted}$. When developing PRS_{multi} using PRS-CS, we observed
notable relative improvements in R^2 when compared to $PRS_{weighted}$, specifically a 24.7% increase
for PLT (0.091 vs. 0.073) with bin ratio of 24:1, and a 57.1% increase for lymphocyte (0.044 vs.
0.028) with a bin ratio of 16:1. Additionally, we found that PRS-CSx showed better performance
in comparison to PRS-CS, especially when EUR GWAS was smaller or Minor GWAS was larger.

However, such improvements were less pronounced with large-scale EUR GWAS or small Minor
GWAS (**Figure S12**). While sharing ancestry-specific GWAS summary statistics is highly
beneficial for determining optimal approaches, our findings highlight the value of pragmatic
approaches that directly construct PRS from large-scale meta-analyzed multi-ancestry GWAS.
Such studies are often more accessible than ancestry-specific GWAS summary statistics.
341

342 PRS derived from local ancestry-informed GWAS can improve accuracy for
343 some less polygenic traits

We next conducted a comparative analysis to evaluate the optimal PRS approaches for admixed
populations, utilizing local ancestry-informed GWAS. Specifically, we used Tractor[19] to perform
GWAS in AFR tracts within admixed AFR-EUR individuals, referred to as **AFR$_{Tractor}$**. This
approach enabled us to construct ancestry-specific PRS across 17 traits in the understudied AFR
population. We developed PRS using both P+T and PRS-CS, and subsequently compared the
accuracies of PRS derived from AFR$_{Tractor}$ with those derived from large-scale EUR GWAS
performed with standard linear regression (**EUR$_{Standard}$**). To maximize discovery sample size, we
also developed PRS$_{weighted}$ by combining EUR$_{Standard}$-derived PRS and AFR$_{Tractor}$-derived PRS
through linear weighting; we compared its performance to PRS derived from multi-ancestry meta-
analyzed GWAS (referred to as **Meta$_{Standard}$**, see **STAR Methods**).
354

Local ancestry-informed ancestry-specific GWAS had a much smaller sample size relative to the
EUR-inclusive GWAS, as is typical for GWAS of underrepresented populations. As expected, we
did not observe significant predictive accuracy of AFR$_{Tractor}$-derived PRS for most traits such as
height and BMI (**Figure 6 and Table S10**). However, we observed notable improvements for 5
traits, including WBC, neutrophil count (neutrophil), MCV, MCH and MCHC, where AFR$_{Tractor}$-
derived PRS achieved significantly higher R^2 compared to EUR$_{Standardr}$-derived PRS when using
P+T (0.040 vs. 0.007, one-sided paired t-test, p-value = 0.038), despite a much larger sample
size for EUR$_{standard}$. This improvement might be attributed to the presence of large-effect AFR-
enriched variants, particularly for MCV, MCH and MCHC, which are effectively captured by
Tractor GWAS[6,19]. Consistent with our previous findings, P+T generally outperformed PRS-CS
for these traits characterized by much sparser genetic architectures, with a mean R^2 of 0.040
compared to 0.022. Given that heritability bounds predictive accuracy, which can vary among
populations and contexts, we also compared SNP-based heritability estimates between the AFR
and EUR populations in the Pan-UK Biobank Project
(https://pan.ukbb.broadinstitute.org/docs/heritability/index.html). In line with our PRS accuracy
results, we observed higher estimates of SNP-based heritability for WBC (h^2 = 0.41, SE = 0.19
vs. h^2 = 0.17, SE = 0.01), neutrophil (h^2 = 0.44, SE = 0.26 vs. h^2 = 0.15, SE = 0.01), and MCHC
(h^2 = 0.15, SE = 0.11 vs. h^2 = 0.06, SE = 0.01) in the AFR population compared to the EUR
population. However, these differences did not reach statistical significance, which can be
attributed to the large standard errors resulting from the limited small sample size of AFR
population and the sparser genetic architectures, leading to less stable heritability estimates using
LDSC.
377

The best local ancestry-informed PRS approach that we evaluated for the 5 less polygenic traits
with large ancestry-specific effects was a weighted linear regression approach. This approach
combined PRS derived from $AFR_{Tractor}$ and $EUR_{Standard}$ using linear regression and outperformed
predictive accuracy compared to using $Meta_{Standard}$-derived PRS. This finding aligns with our
earlier observations, where $PRS_{weighted}$ outperformed PRS_{multi} for traits with large effect ancestry-
enriched variants, while PRS_{multi} exhibited superior overall performance for traits lacking such
variants. Specifically, the mean accuracies of $PRS_{weighted}$ using P+T, PRS-CS and PRS-CSx for
those 5 traits were 0.044, 0.031, and 0.028, respectively, with no significant differences observed
among the three PRS methods. The mean accuracies of $Meta_{Standard}$ were 0.016 and 0.008 using
PRS-CS and P+T, respectively. Additionally, we did not observe significant accuracy differences
between PRS derived from GWAS conducted using standard linear regression in admixed
populations and $AFR_{Tractor}$-derived PRS **(Table S10**). It is worth noting that the effective sample
size of local ancestry-informed GWAS is approximately 20% smaller due to the reduction from
deconvolving ancestral tracts. Moreover, PRS derived from traditional GWAS in admixed
populations necessitate an in-sample LD reference panel. In contrast, local ancestry-informed
GWAS-based PRS, as shown in this study, can leverage external LD reference panels,
eliminating the need for direct access to individual-level genotypes of admixed populations.
395

Discussion

In this study, we extensively evaluated PRS performance through a combination of simulation
and empirical analyses to explore the impact of various factors on PRS predictive accuracy and
generalizability across populations. We demonstrated that increasing genetic diversity of
discovery GWAS improved predictive accuracy in understudied populations. The extent of
improvement was influenced by factors such as sample size ratios between EUR GWAS and
Minor GWAS, genetic architecture, PRS methodology, and LD reference panels. Among those
factors, between-ancestry genetic architecture differences, such as ancestry-enriched variants
with large effects, affected accuracy improvement more than other factors. While leveraging large-
scale EUR GWAS continues to benefit PRS accuracy given the current scale of understudied
populations, we may not expect accuracy improvement when meta-analyzing extremely small
Minor GWAS[24].
408

Our study also revealed that directly meta-analyzing datasets from diverse ancestral groups could
yield greater accuracy improvements than linearly combining PRS through an optimized weighting
strategy, especially for P+T. Such improvements from meta-analyzed GWAS supports the
common implicit assumption that causal variants are shared between ancestries. Consistent with
this assumption, when smaller target populations lack representation, leveraging genetic
information from a different population with larger sample sizes improves PRS accuracy, even
when it is ancestrally diverged. Notably, when employing the more sophisticated genome-wide
PRS method, PRS-CSx, accuracy differences between PRS_{multi} and $PRS_{weighted}$ were marginal.
Moreover, PRS-CSx generally outperformed PRS-CS, with the exception of BMI. The
improvement was most pronounced for traits with ancestry-specific variants, such as MCV and
MCH.

We have comprehensively evaluated characteristics that impact PRS performance, including in
recently admixed populations. We have shown the advantage of leveraging GWAS in admixed
populations by accounting for local ancestry, which could improve PRS predictive performance in
understudied populations even without direct access to individual genotypes of admixed
populations. Specifically, we found that $PRS_{weighted}$ consistently outperformed PRS_{multi} for traits
with ancestry-enriched variants. However, the sample size of admixed individuals here was
relatively small, and we anticipate that future analyses incorporating larger datasets, such as the
All of Us Research Program, will provide further insights into optimal PRS strategies for improved
accuracy and generalizability using PRS derived from local ancestry-informed GWAS.

While previous studies have shown the advantages of leveraging increased genetic diversity to
improve PRS accuracy in global populations[7,35], most have used GWAS with primarily European
ancestry. Here, we have provided additional best practices for developing PRS for understudied
populations using diverse discovery cohorts, particularly when GWAS encompass different
ancestry compositions across various trait genetic architectures (**Figure 7**). Our
recommendations primarily revolve around general guidelines for constructing PRS_{single} and
PRS_{multi} (or $PRS_{weighted}$), depending on factors examined in this study (**Figure S13**).

First, in the development of PRS_{single}, we employed a theoretical equation[36] to enhance the
selection of input GWAS (**Supplementary Note 7),** as a function of the cross-ancestry genetic
correlation, SNP-based heritability in discovery and target populations, discovery GWAS sample
size, and the number of genome-wide independent segments in the discovery population[36]. For
traits with relatively low r_g and a sizable ancestry-matched GWAS (e.g., > 20-40% of EUR
GWAS), such as BMI and height, PRS accuracy in the target population improves when ancestry-
matched GWAS are utilized. On the other hand, for traits with high r_g and SNP-based heritability,
we expect larger-scale EUR GWAS to outperform smaller-scale ancestry-matched GWAS.
However, it is important to consider the characteristics of the target cohort and phenotype
precision. Additionally, we expect Bayesian methods tailored to trait-specific genetic architecture
to outperform classic P+T methods. However, this superior performance may not hold true for
traits that exhibit large-effect ancestry-enriched variants or with a very sparse genetic architecture,
which are attributes typically informed by prior knowledge or information gleaned from literature
and public resources[35,37–39]. To enhance accuracy in such scenarios, we recommend employing
a grid-search approach with a finer-scale adjustment of the hyper-parameters in Bayesian
methods.

Second, in comparison to PRS_{single} derived from large-scale EUR GWAS, we recommend using
PRS_{multi}, unless the target ancestry-matched GWAS is extremely small (<10,000). PRS_{multi} is
generally preferred for traits with high r_g, SNP-based heritability, and large sample sizes. We find
increasing evidence supporting the notion that the effects of most common variants are shared
between ancestries, indicating a high r_g for most traits[9,11]. However, estimates of r_g can be
affected by phenotypic and environmental heterogeneity across populations[10,40]. When
constructing PRS_{multi} using summary-level based methods such as P+T and PRS-CS,
researchers should carefully consider which LD reference panel best approximates the LD

structure between SNPs while being most readily accessible. We have shown that when EUR
remains the majority population in the discovery GWAS, using the EUR-based reference panel
effectively approximates the LD of discovery GWAS, consistent with our previous findings[7].
467
Third, our findings indicate the advantages of PRS_{multi} compared to $PRS_{weighted}$, particularly when
employing P+T and PRS-CS. However, there are some notable exceptions, such as the higher
accuracy observed when using $PRS_{weighted}$ with PRS-CS for traits with low r_g, such as height.
Furthermore, when incorporating local ancestry-informed GWAS and large-scale EUR GWAS,
$PRS_{weighted}$ outperformed PRS_{multi} for traits with AFR-enriched variants, such as WBC and MCHC,
in the UKBB-AFR. On the other hand, we note that the accuracy of PRS_{multi} could be more affected
by the choice of LD reference panel, while $PRS_{weighted}$ was not limited in this regard due to its easy
accessibility of external ancestry-matched reference panels. PRS-CSx, which accounts for
variations in allele frequencies and LD structures across ancestries, is recommended when
ancestry-specific GWAS from multiple populations are available, especially with considerable
sample sizes (e.g., > 25,000~50,000) in the Minor GWAS. These results highlight the importance
of making ancestry-specific GWAS summary statistics publicly available.
480
In summary, there is no one-size-fits all approach for constructing PRS, as the optimal approach
depends on genetic architecture, ancestry composition, statistical power, and other factors. These
factors can be complex, particularly as a deluge of methods are being developed to address the
PRS generalizability problem. To inform optimal approaches across a wide range of scenarios,
we have distilled the results of extensive simulations and empirical analyses across trait genetic
architectures, ancestries, and methods into a set of guidelines from parameters that are typically
evaluated at the outset of a genetic study.
488

Limitations of the study

We acknowledge some limitations and future directions in our study. First, we focused on common
variants in different populations, while population-enriched variants have lower frequencies and
larger effect sizes in some populations. The role of such variants in polygenic prediction are worth
exploring across phenotypes when there are sufficient sample sizes for different ancestral
populations. Second, as we used external LD reference panels for PRS construction, PRS
performance decreases with LD mismatch between the discovery population and LD reference
panel, especially using multi-ancestry GWAS. While we show that LD reference panel differences
have a relatively modest effect on PRS accuracy, they have a much larger effect on fine-
mapping[41], so future efforts are warranted to share in-sample LD without direct access to
individual-level genotypes, especially for large consortia with numerous and diverse cohorts.
Alternatively, developing more sophisticated individual-level PRS methods that preserve privacy
and are scalable to current biobank-scale data is also promising. Third, while our primary focus
pertains to quantitative phenotypes characterized by diverse genetic architectures, we expect our
findings can be broadly applied to binary traits, as we have investigated previously[7]. However,
binary phenotypes introduce additional complexities due to factors such as variable case/control
ratios, phenotype definitions, environmental differences, and smaller effective sample sizes or

lower statistical power. Fourth, while we have provided theoretical expectations of cross-ancestry prediction, the reliability of parameter estimates such as cross-ancestry genetic correlation and the effective number of independent genome-wide segments poses significant challenges, particularly in the context of multi-ancestry GWAS with highly imbalanced sample sizes. Finally, it is important to acknowledge that our study focused on selected methods, which consistently exhibit similar trends[42]. Although we anticipate that our findings are broadly applicable to alternative methods, such as XPASS[43] and XP-BLUP[42], further research is needed to explore the generalizability of our findings to other polygenic prediction approaches. Despite the limitations, our study highlights the advantages of leveraging the increasing diversity of current genomics studies to improve polygenic prediction across populations. We emphasize the necessity of diversifying not only the ancestry but also phenotypic spectrum when collecting genomic data from global populations, which will contribute to achieve a more equitable and effective use of PRS for traits with varying genetic architectures.

Acknowledgements

A.R.M. and Y.W. were supported by funding from the National Institutes of Health (K99/R00MH117229 to A.R.M.) as well as funding from European Union's Horizon 2020 research and innovation program under grant agreement 101016775. A.R.M. was also supported by funding from U01HG011719. H.H. acknowledges support from NIDDK K01DK114379, NIDDK 1R01DK129364, NIMH U01MH109539, the Zhengxu and Ying He Foundation, and the Stanley Center for Psychiatric Research. E.G.A. is supported by K01MH121659 from NIMH, the Caroline Wiess Law Fund for Research in Molecular Medicine, and the ARCO Foundation Young Teacher-Investigator Fund at Baylor College of Medicine. P.T. is supported by funding from NIH/NIA (R00 AG062787) and by the Good Ventures Foundation (010623-00001).

Author Contributions

Conceptualization: Y.W., A.R.M., E.G.A., M.Kanai.
Formal analysis: Y.W., M.Kanai., T.T., M.Kamariza., K.Y., W.Z., P.T.
Writing – Original Draft: Y.W., A.R.M., E.G.A., M.Kanai., T.T.
Writing – Review & Editing: Y.W., M.Kanai., T.T., M.Kamariza, K.T., K.Y., W.Z., Y.O., H.H., P.T., K.T., E.G.A., A.R.M.

Declaration of interests

H.H. received consultancy fees from Ono Pharmaceutical and honorarium from Xian Janssen Pharmaceutical. All other authors declare no competing interests.

Figure Legends

Figure 1. Study design in both simulations and empirical analyses

1) In the context of single-ancestry GWAS, we randomly split individuals in European (EUR) and other minority populations, including East-Asian and African populations, into equally sized bins. Simulations involved a total of 52 bins per population, each containing 10,000 individuals. For empirical analysis, bin number was dependent on the sample size of respective phenotype in that population (**Table S3**), with 5,000 individuals per bin. GWAS was conducted within each bin for each individual population, followed by meta-analysis of GWAS from various numbers of bins within each population. To construct PRS derived from single-ancestry GWAS (**PRS$_{single}$**) in the target population, we applied P+T for both simulations and empirical analyses, utilizing PRS-CS for the latter. Subsequently, we combined PRS$_{single}$ developed from EUR GWAS (**PRS$_{EUR\_GWAS}$**) and other minority population-based GWAS (**PRS$_{Minor\_GWAS}$**) through a linear weighted strategy (denoted as **PRS$_{weighted}$**, highlighted in red box) for empirical analyses. Note that PRS$_{weighted}$ was also developed using PRS-CSx, which utilizes GWAS summary statistics from multiple populations. 2) For meta-analyzed multi-ancestry GWAS (referred to as **Meta),** we ran meta-analyses on EUR GWAS and Minor GWAS with varying ancestry compositions. In simulations, we incrementally included 4 bins from EUR GWAS for the meta-analysis, while in empirical analyses, we increased the number to 8 bins. Simultaneously, we varied the number of bins in Minor GWAS from 1 to the total number. Following the meta-analysis, we constructed PRS based on Meta (referred to as **PRS$_{multi}$**), using the P+T method for simulations, and employing both P+T and PRS-CS for empirical analyses.

Figure 2. Improvement of PRS accuracy through meta-analyzed multi-ancestry GWAS compared to large-scale European GWAS across 6 simulated genetic architectures.

The multi-ancestry GWAS included populations of European (EUR) and East-Asian (EAS) ancestry, with the EAS sample size varying as indicated on the x-axis. For illustrative purposes, we present the results using 32 EUR bins, each consisting of 10,000 individuals, which were included in both EUR GWAS and multi-ancestry GWAS. PRS was separately evaluated in African (AFR), EAS and EUR populations. Full results are shown in **Table S1**. M_c indicates the number of causal variants and h^2 refers to SNP-based heritability. In each panel, the red vertical dashed line indicates the point where an equal number of bins from EUR and EAS populations were included in the multi-ancestry GWAS. The error bars represent the standard errors of predictive accuracy differences between PRS derived from multi-ancestry GWAS (PRS$_{multi}$) and PRS derived from EUR GWAS (PRS$_{EUR\_GWAS}$).

Figure 3: Genetic architecture of 17 studied traits between Biobank Japan (BBJ) and UK Biobank (UKBB).

The error bar is the standard deviation of the corresponding estimate. The vertical dashed line was the median estimate. Full results are shown in **Table S4**. The phenotypes were ranked according to their polygenicity estimates using GWAS from UKBB, including: BMI (body mass index), Height, DBP (diastolic blood pressure), SBP (systolic blood pressure), WBC (white blood cell count), Lymphocyte (lymphocyte count), RBC (red blood cell count), Neutrophil (neutrophil count), HB (hemoglobin concentration), HT (hematocrit percentage), Eosinophil (eosinophil count), PLT (platelet count), Monocyte (monocyte count), MCV (mean corpuscular volume), MCH (mean corpuscular hemoglobin), Basophil (basophil count), MCHC (mean corpuscular hemoglobin concentration).

Figure 4. Accuracy improvement of PRS in the UK Biobank East-Asian population (UKBB-EAS) using multi-ancestry GWAS comprare to using European (EUR) GWAS for P+T and PRS-CS.

The multi-ancestry GWAS were obtained by meta-analyzing EUR GWAS and EAS GWAS, with the EAS sample size from the Biobank Japan (BBJ) varying as indicated on the x-axis. For illustrative purposes, we present the results using 64 EUR bins, each containing 5,000 individuals, which were included in both EUR GWAS and multi-ancestry GWAS. PRS were constructed using P+T and PRS-CS and evaluated in the UKBB-EAS. The y-axis is the accuracy difference of PRS when using multi-ancestry GWAS (PRS_{multi}) compared to using EUR GWAS (PRS_{EUR_GWAS}). The error bars indicate the standard error of accuracy improvement. The red dashed line is y=0. We showed the results for 7 traits with SNP-based heritability > 0.1 in both BBJ and UKBB, and they were ranked by polygenicity estimates using UKBB (**Figure 3**). Full results are shown in **Table S7**.

Figure 5. Predictive accuracy using different PRS methods in the UK Biobank East-Asian population (UKBB-EAS).

PRS_{multi} represents PRS derived from multi-ancestry GWAS, while $PRS_{weighted}$ denotes PRS constructed from a weighted linear combination (see **STAR Methods** for details). PRS were constructed with three methods, including P+T, PRS-CS and PRS-CSx. We showed the results for 7 traits with SNP-based heritability > 0.1 in both Biobank Japan (BBJ) and UKBB. Traits were ranked by polygenicity estimates using UKBB (**Figure 3**). Boxes represent the first and third quartiles, with the whiskers extending to 1.5-fold the interquartile range. Full results are shown in **Table S8 and Table S9**.

Figure 6. Accuracy of PRS derived from local-ancestry informed GWAS versus other discovery GWAS in the UK Biobank African population (UKBB-AFR)

We evaluated PRS performance in the UKBB-AFR by utilizing various methods on different discovery GWAS. Specifically, $AFR_{Tractor}$ denotes the AFR-specific GWAS performed using Tractor on the UKBB admixed African-European individuals. $EUR_{Standard}$ refers to standard GWAS performed on the European (EUR) population in the UKBB. $Meta_{Standard}$ is the meta-analysis performed on $AFR_{Tractor}$ and $EUR_{Standard}$. Furthermore, we constructed a weighted PRS by combining PRS generated from $AFR_{Tractor}$ and $EUR_{Standard}$ through a linear weighted approach. The figure shows the results for traits with SNP-based heritability > 0.1 in the UKBB-AFR. Full results are shown in **Table S10**.

Figure 7. General practices for developing PRS using different discovery GWAS.

We summarized the general practice for developing PRS A) using single-ancestry GWAS (PRS_{single}); and B) using GWAS from multiple ancestries (PRS_{multi} or $PRS_{weighted}$). Abbreviations: Cross-ancestry genetic correlation (r_g), SNP-based heritability in discovery (h_d^2) and target populations (h_t^2), discovery GWAS sample size (N_d) and the number of genome-wide independent segments in the discovery population (M_d).

STAR Methods

Resources Availability

Lead Contact

Further information and requests for resources and reagents should be directed to and will be fulfilled by the lead contact, Ying Wang (yiwang@broadinstitute.org).

Materials Availability

This study did not generate new unique reagents.

Data and code availability

1000 Genome Phase 3 data can be accessed at ftp://ftp.1000genomes.ebi.ac.uk/vol1/ftp/data_collections/1000_genomes_project/data. We used UK Biobank data via application 31063. The software used in this study can be found at: Plink (https://www.cog-genomics.org/plink/), PRS-CS (https://github.com/getian107/PRScs), PRS-CSx (https://github.com/getian107/PRScsx), Tractor (https://github.com/Atkinson-Lab/Tractor), HapGen2 (https://mathgen.stats.ox.ac.uk/genetics_software/hapgen/hapgen2.html) and SBayesS/GCTB (https://cnsgenomics.com/software/gctb/). The Pan UK Biobank Project can be

Methods Details

Simulations

Simulated genotypes in three populations

To explore the potential improvement of predictive accuracy within an underrepresented target
ancestry through the inclusion of additional samples included in the multi-ancestry discovery
GWAS, we simulated genotypes of chromosome 22 for 560,000 individuals in each population
including European ancestry (EUR), East Asian ancestry (EAS) and African ancestry (AFR) using
the software HapGen2 v2.1.2[44]. We used the haplotypes from 1000 Genome Project (1KG, Phase
3)[45] as the sample pool. We excluded Americans of African Ancestry in SW USA and African
Caribbeans in Barbados from the AFR samples due to their high degree of recent admixture. We
used default parameters in HapGen2 with effective sample sizes of 11,375, 12,239 and 17,380
for EUR, EAS and AFR, respectively[44]. After simulating the genotypes on chromosome 22, we
ran analyses with a total of 87,938 overlapping SNPs across the three ancestries which passed
quality control filters: minor allele frequency (MAF) > 0.01, Hardy-Weinberg Equilibrium (HWE) *p*-
value > 10^{-6} and genotype missingness rates across individuals < 0.05. We then removed 2nd-
degree related individuals using the software KING[46], resulting in 534,352, 533,996 and 537,498
unrelated individuals from EUR, EAS and AFR, separately. We randomly sampled 10,000 and
520,000 individuals from each ancestry as the withheld target population and discovery
population, respectively.

Simulated phenotypes with varying trait genetic architecture

For the sake of simplicity, we assumed that causal variants are shared across populations and
their effect sizes are perfectly correlated (cross-ancestry genetic correlation, $r_g = 1$) in our initial
simulations. The pairwise r_g among K populations is represented by a $K * K$ matrix, denoted as
R, where the off-diagonal elements of **R** had the value of r_g and diagonal elements of **R** were set
to 1. In our study, K was equal to 3, indicating the number of populations considered. We
simulated phenotypes based on the simple additive model: $y = g + e$, where $g = \sum_{j=1}^{M_c} x_{ij} \beta_j$.
M_c is the number of causal variants, x_{ij} is the genotype coded as 0, 1, or 2 for the jth SNP in the
ith population. The effect size of jth SNP across K populations is drawn from a multivariate normal
distribution, $\beta \sim MVN(0, \Sigma)$, where for the $K * K$ variance-covariance matrix, Σ, the diagonal and
off-diagonal elements were $\frac{h^2}{2 f_{ij}(1-f_{ij})M_c}$ and $\mathbf{R} \bullet \frac{h^2}{2 f_{ij}(1-f_{ij})M_c}$, respectively. We denoted f_{ij} as the
MAF of jth SNP in the ith population and h^2 as the trait heritability. We simulated the
environmental effects to follow a normal distribution with 0 mean and $1 - h^2$ variance, $e \sim N(0, 1 -$

h^2). We simulated different levels of heritability for chromosome 22 ($h^2 = 0.03$ and 0.05).
Additionally, we randomly sampled various numbers of causal variants (M_c = 100, 500, and 1000)
from all the 87,938 SNPs. As a result, we defined a total of 6 distinct simulation scenarios that
encompass a realistic spectrum of polygenicity, ranging from ~0.1% to ~1% of causal variants.
To assess the impact of r_g on PRS performance, we expanded our simulation study by
considering two scenarios. These scenarios aimed to capture different levels of per-variant
variance explained. In scenario 1 characterized by M_c = 100 and $h^2 = 0.05$, the per-variant
variance explained was higher. Conversely, scenario 2 involved M_c = 1000 and $h^2 = 0.03$,
resulting in a lower per-variant variance explained. For each scenario, we varied the values of r_g
to 0.6 and 0.8, respectively.

Downsampling and meta-analyzed GWAS in simulations

To provide the requisite discovery data for constructing PRS, we proceeded to perform GWAS on
the simulated phenotypes. Specifically, we split the discovery population, which consisted of
520,000 unrelated individuals, into 52 evenly distributed bins, each comprising 10,000 individuals
(denoted as Bin_1, Bin_2, …, Bin_{total}). Subsequently, we ran GWAS on each of those 52 bins
independently within the three populations, using simple linear regression implemented in PLINK
v2.0[47]. We excluded the causal variants when running GWAS to mimic the phenomenon of
imperfect tagging. We then employed an iterative process of meta-analysis, employing the
inverse-variance weighted method using METAL[48], gradually incorporating a varying number of
bins. Specifically, we commenced the meta-analysis with Bin_1+Bin_2, subsequently progressing to
Bin_1+Bin_2+Bin_3, and so forth, until we encompassed the complete set of bins
(Bin_1+Bin_2+Bin_3+...+Bin_{tota}) for each population.
709
To simulate a scenario resembling a meta-analysis involving multiple ancestries with varying
proportions, we opted for an arbitrary selection of subsets from EUR GWAS. Specifically, we
chose a range of bins, from 4 to 52 bins, with increments of 4. Subsequently, we systematically
incorporated different numbers of bins, spanning from 1 to 52, from EAS and AFR populations
into the EUR GWAS dataset via meta-analysis. The meta-analysis was conducted utilizing the
inverse-variance weighted fixed effects model implemented in the METAL software. This iterative
process allowed us to achieve a range of sample size ratios between EUR and EAS as well as
EUR and AFR, encompassing ratios from 52:1 to 4:52, in the meta-analyzed multi-ancestry
GWAS (referred to as **Meta**). The simulation configuration is visually depicted in **Figure 1**.
719

Pruning and Thresholding (P+T) in simulations

We employed PLINK v1.90 to clump quasi-independent SNPs within 500Kb windows, utilizing a
LD threshold of $r^2 < 0.1$. To explore the impact of various LD reference panels on predictive
accuracy of PRS, we used a total of four different LD reference panels: one for single-ancestry
and three for multi-ancestry GWAS, with consideration to the ancestry composition of the
discovery GWAS and the target population.
726

For the single-ancestry GWAS, we used a LD reference panel consisting of 10,000 individuals
from the target population that were matched to the ancestry of the discovery GWAS. In the case
of multi-ancestry GWAS, we used three LD reference panels. These panels included two
composed of a single ancestry that did not mirror the ancestral makeup of the discovery GWAS.
Specifically, one panel comprised 10,000 withheld EUR individuals, while the other panel
encompassed individuals from understudied populations, either 10,000 EAS or 10,000 AFR
individuals, consistent with the minority population represented in the discovery GWAS. The third
LD reference panel consisted of individuals from different ancestries in proportions proportional
to the discovery GWAS, amounting to a total of 10,000 samples.
736
We calculated PRS in the target population using 8 different p-value thresholds: 5×10^{-8}, 1×10^{-6},
1×10^{-4}, 1×10^{-3}, 0.01, 0.05, 0.1, and 1. We denoted PRS constructed from single-ancestry
GWAS as single-ancestry PRS (PRS_{single}) and those from meta-analyzed multi-ancestry GWAS
as multi-ancestry PRS (PRS_{multi}). We calculated the predictive accuracy as the variance explained
by the PRS (R^2) through linear regression: $y \sim PRS$ and computed corresponding 95%
confidence intervals (CIs) through bootstrap. To identify the optimal p-value threshold associated
with the highest predictive accuracy, we evenly divided the target population into a test cohort
and a validation cohort. The p-value threshold was optimized through a process of
hyperparameter tuning in the validation cohort, and subsequently, the accuracy of the model was
assessed using the test cohort.
747

Empirical analysis of 17 quantitative traits in the UK Biobank (UKBB) and Biobank Japan (BBJ)

750 We further explored how the findings from simulations generalized in real data using 17
751 quantitative traits shared between UKBB and BBJ, including anthropometric traits (BMI and
752 height) and blood panel traits studied previously (**Table S3**)[32]. The selection of these traits was
753 motivated by their widespread availability within biobanks and their substantial statistical power,
754 attributable to their quantitative properties.

Datasets and Quality Control (QC)

UK Biobank (UKBB): The details of assigning ancestry for each individual in the UKBB are
described in the Pan-UK Biobank Project (Pan UKBB: https://pan.ukbb.broadinstitute.org/).
Briefly, a random forest classifier trained on reference data from 1KG and Human Genome
Diversity Project (HGDP)[49] was used to classify cohort individuals under continental population
labels based on the top 6 principal components (PCs). In this study, we used a total of 361,144
and 2,684 unrelated EUR and EAS participants, respectively. We obtained unrelated individuals
through running hl.maximal_independent_set using Hail (https://hail.is/). Specifically, within each
population, we ran PC-Relate[50] with k=10 and min_individual_maf=0.05. We used the individuals
assigned EAS ancestry as the target dataset. For EUR samples, we first randomly retained 5,000
individuals with complete phenotype information for all 17 studied phenotypes as the target
population. Subsequently, we split the remaining individuals into evenly distributed bins, each
containing 5,000 individuals, for each phenotype. The number of total bins for each studied

phenotype ranged from 68 to 71, depending on phenotype missingness (**Table S3**). The bins
were labeled sequentially from 1 to the total number of bins, following the same procedure as
described in our simulations.
771
BioBank Japan (BBJ): BBJ is a multi-institutional hospital-based biobank which has recruited
approximately 200,000 participants from 12 medical institutions in Japan between fiscal years
2003 and 2007[27]. Written informed consents were obtained from all the participants, as approved
by the ethics committees of the RIKEN Center for Integrative Medical Sciences, and the Institute
of Medical Sciences, the University of Tokyo. The participants were genotyped using either (i) the
Illumina HumanOmniExpressExome BeadChip or (ii) a combination of the Illumina
HumanOmniExpress and HumanExome BeadChips. The genotypes were then prephased using
Eagle[51] and imputed using Minimac3[52] with a reference panel that consists of 1KG samples (N =
2,504) and whole-genome sequencing (WGS) data of Japanese individuals (N = 1,037)[53].
Standard quality controls of participants and genotypes were applied as described elsewhere[53].
Briefly, we excluded samples with low call rates (< 98%), closely related individuals (PLINK
PI_HAT > 0.175), or non-Japanese outliers based on the principal component analysis (PCA).
We then excluded genotyped variants with call rate < 98%, HWE P-value < 1.0×10^{-6}, number of
heterozygotes < 5, or low concordance rate (< 99.5%) with WGS for a subset of individuals (N =
939). Phenotypes were retrieved from medical records and prepared as described previously[54].
787
788 **1000 Genomes Project Phase 3 (1KG):** We used 1KG phase 3 data as LD reference panels in
789 this study. Specifically, we kept 495 unrelated EUR, 498 unrelated EAS, and 484 unrelated AFR
790 individuals from 1KG. The AFR individuals were solely utilized for analyses pertaining to recently
791 admixed populations.
792
Quality Controls: The imputation strategies for UKBB and BBJ have been described in detail
elsewhere[55,56]. After imputation, we first excluded ambiguous variants (e.g., A/T and C/G) and
further filtered to keep those variants with imputation INFO score > 0.3, MAF > 0.01, HWE p-
value > 10^{-6}, and genotyping missing rates across individuals < 0.05. Consequently,
approximately 8.6 million and 6.6 million SNPs were retained for the UKBB and BBJ, respectively.
For our analyses, we exclusively utilized SNPs that passed these quality control measures,
resulting in approximately 3.6 million SNPs that were shared among both biobanks and 1KG.
800

PRS construction for 17 traits in empirical analysis

Discovery GWAS: All phenotypes were curated and transformed to be normally distributed as
described previously[32]. Subsequently, we performed GWAS on the rank normalized phenotypes
using simple linear regression implemented in PLINK v2.0. We included age, sex, age^2, age ×
sex, age^2 × sex, and the first 20 PCs as the covariates. In line with the GWAS strategy outlined
in the *Simulations* section, we initially performed GWAS within individual bins and then engaged
in an iterative meta-analysis, employing inverse-variance weighted meta-analysis in METAL,
separately for UKBB and BBJ cohorts. For the meta-analysis of GWAS results derived from
single-ancestry analyses in the UKBB and BBJ (referred to as "Meta"), we incorporated a variable

number of EUR bins from UKBB, ranging from 8 to 64 with an increment of 8. Subsequently, we
systematically integrated additional EAS bins from BBJ.

PRS construction methods: We used different methods to construct PRS in the target
populations, specifically UKBB-EAS and UKBB-EUR. In accordance with *Simulations*, we also
explored the impact of LD reference panels on PRS performance by utilizing multiple panels from
1KG, while taking into account the ancestry composition of discovery GWAS for P+T. Additionally,
we implemented PRS-CS[39], a Bayesian regression framework that integrates a continuous
shrinkage prior to infer the posterior mean effects of SNPs. To ensure computational efficiency,
we employed the auto model in the PRS-CS framework, which automatically estimates the hyper-
parameter *phi* (the proportion of SNPs with non-zero effects) based on the input GWAS (see
Supplementary Note 8). For both UKBB and Meta, we used 1KG-EUR as the LD reference
panel, while for BBJ, we utilized 1KG-EAS reference panel.

To further explore the performance of PRS incorporating GWAS from multiple ancestries, we
constructed a weighted PRS by linearly combining PRS derived from single-ancestry GWAS[34].
Specifically, the weighted PRS was calculated as **PRS$_{weighted}$** = w_1* PRS$_{EUR\_GWAS}$ + w_2 *
PRS$_{Minor\_GWAS}$, where w_1 and w_2 were weights attached to individual PRS. Furthermore, we used
a more sophisticated method, PRS-CSx[8], to generate ancestry-specific posterior SNP effects
using multiple GWAS summary statistics. PRS-CSx, an extension of PRS-CS, can model
ancestry-specific allele frequencies and LD patterns. Similar to PRS-CS, we used the ancestry-
matched LD reference panel from 1KG and performed the auto model implemented in PRS-CSx.
We also incorporated the *--meta* flag, which enables inverse-variance weighted meta-analysis in
the Gibbs sampler. Consequently, we developed two types of PRS from PRS-CSx, one was
based on the meta-analyzed effects (referred to as PRS$_{multi}$) and the other, PRS$_{weighted}$, was
dependent on the ancestry-specific posterior SNP effects.

PRS performance evaluation: We assessed the predictive accuracy of PRS by measuring the
incremental R^2 using linear regression, where we accounted for the influence of covariates. Two
models were compared: 1) $H_0: Phenotype \sim covariates$, representing the baseline model, and 2)
$H_1: Phenotype \sim PRS + covariates$, incorporating PRS as the full model. The incremental R^2 was
utilized to quantify the improvement in model accuracy resulting from the inclusion of PRS, thus
providing a measure of the specific contribution made by PRS to the predictive power of the
model. We computed the corresponding 95% confidence intervals (CIs) through bootstrap. To
maximize the predictive accuracy of P+T and PRS$_{weighted}$, we employed an optimization strategy
to identify the optimal *p*-value thresholds for P+T and the weights (w_1 and w_2) assigned to various
PRS components for PRS$_{weighted}$. This optimization process entailed a random partitioning of the
target population into two equally sized subsets, namely the validation dataset and the test
dataset. The hyperparameter was identified in the validation dataset, and subsequently, the
accuracy of the model was assessed using the test dataset. We replicated the process 100 times
and calculated the standard error of predictive accuracy across 100 replicates. This approach
allowed us to maximize the performance of P+T and PRS$_{weighted}$ by iteratively refining the *p*-value
thresholds and weight parameters, thereby enhancing their predictive capabilities.

Measures of genetic architecture using summary-data-based BayesS (SBayesS)[29]

To better understand the impact of trait genetic architecture on PRS predictive performance, we
evaluated three parameters including the polygenicity (proportion of SNPs with nonzero effects),
SNP-based heritability and S (the relationship between MAF and effect sizes) for 17 studied
phenotypes (**Table S3**). These parameters were estimated using SBayesS implemented in the
GCTB software (https://cnsgenomics.com/software/gctb/). For the analysis, we employed meta-
analyzed GWAS data obtained from the comprehensive UKBB and BBJ datasets. Specifically,
the number of bins included in the GWAS was equal to the total number of bins associated with
the respective phenotype (**Table S3**). We used the LD reference panel provided by GCTB for
UKBB GWAS. We constructed a shrunk LD matrix using 50,000 unrelated individuals from BBJ
as the LD reference panel for BBJ GWAS. We used 4 chains for the Markov Chain Monte Carlo
process, which calculated the Gelman-Rubin convergence diagnostic (also known as potential
scale reduction factor) for these three parameters. We performed the analyses using other default
settings for SBayesS. Given the potential convergence issues associated with Bayesian models,
we deemed a threshold value of less than 1.2 for the Gelman-Rubin convergence diagnostic as
indicative of good convergence for the estimated parameters.
870

UK Biobank recent admixture ancestry analysis

To investigate one explanation for poor transferability of PRS across populations – genetic
divergence between the discovery and target cohorts – we further explored whether PRS
constructed from ancestry-specific summary statistics generated with local ancestry-informed
GWAS in admixed populations improves predictive performance in underrepresented
populations. Specifically, we used the Tractor method[19], accounting for both local ancestry and
risk allele information, to run GWAS in two-way admixed AFR-EUR individuals from the UKBB (N
= 4,576). The average AFR proportion was 62.9%. We used 4,022 unrelated relatively
homogeneous AFR individuals, which are independent from the admixed individuals, as the target
cohort.
881

We followed the same criteria for QC and individual selection as described in Atkinson et al.[19].
For sample QC, we excluded individuals that had <95% call rate, withdrew from the study, had
closer than 2nd degree relatives present in the sample, or that had sex chromosome aneuploidies.
For variant QC we restricted to biallelic SNPs with >90% call rate, HWE p-value $> 10^{-6}$, and MAF
of at least 0.5%. We selected two-way admixed AFR-EUR individuals from the UKBB by first using
the PC loadings from the reference dataset described previously for ancestry inference (1KG +
HGDP) to project UKBB individuals into the same PC space. We applied the same random forest
ancestry classifier described previously to the projected UK Biobank PCA data and assigned AFR
ancestry if the probability was >50%. We restricted to only two-way admixed AFR-EUR ancestry
individuals by selecting those individuals assigned the 'AFR' population label, then filtering to
those with at least 12.5% European ancestry, at least 10% African ancestry, and who did not
deviate more than 1 standard deviation from the AFR-EUR cline based on their PC loadings. This
process resulted in 4,576 individuals.
895

We ran local ancestry deconvolution on this set of admixed individuals using RFmix v2[18] with 1
EM iteration and a window size of 0.2 cM with the HapMap combined recombination map[57] to
inform switch locations. The -n 5 flag (terminal node size for random forest trees) was included to
account for an unequal number of reference individuals per reference population. We used the --
reanalyze-reference flag, which recalculates admixture in the reference samples for improved
ability to distinguish ancestries. As a reference panel, we used continental AFR and EUR
individuals from the 1KG.

Subsequently, we performed GWAS for the 17 quantitative traits utilizing the Tractor method on
the 4,576 individuals with mixed AFR-EUR ancestry from the UKBB. This analysis yielded the
generation of ancestry-specific summary statistics for the AFR ($AFR_{Tractor}$) and EUR ($EUR_{Tractor}$)
ancestry components. To evaluate the performance of PRS in the UKBB-AFR, we developed
PRS using Tractor GWAS. Furthermore, we compared these local-ancestry informed PRS with
those derived from GWAS conducted using standard methodologies. Specifically, we constructed
PRS using GWAS performed on the same set of admixed individuals utilizing the simple linear
regression model ($ADM_{Standard}$). Additionally, GWAS summary statistics obtained from UKBB
($EUR_{standard}$, N = 320,000) from the previous section were utilized, and a meta-analysis was
conducted to combine the $AFR_{Tractor}$ with $EUR_{standard}$ ($Meta_{standard}$, N = 324,576). We constructed
PRS based on HapMap3 SNPs, as previous studies have shown comparable performance
between using reliable HapMap3 SNPs exclusively and the use of genome-wide SNPs[7,58].
Additionally, we constructed weighted PRS by incorporating GWAS of $AFR_{Tractor}$ and $EUR_{Standard}$,
for P+T, PRS-CS and PRS-CSx, respectively. Considering the ancestry composition of the
discovery GWAS, we used different sets of reference panels for each respective GWAS.
Specifically, we used 1KG-EUR as the LD reference panel for $EUR_{Tractor,}$ $EUR_{standard}$ and
$Meta_{standard}$, while using 1KG-AFR for $AFR_{Tractor}$. We used an in-sample LD panel for $ADM_{Standard}$.
We calculated the predictive accuracy in the UKBB-AFR using incremental R^2 as described
above. We repeated the process 100 times and reported the standard error of predictive accuracy
across 100 estimates.

Excel Table Title and Legends

Table S1. The comparison of using different LD reference panels across various simulation scenarios. Related to Figure S2, Figure 2, Figure S3 and Figure S4.

Table S2. Impact of cross-ancestry genetic correlation on predictive performance. Related to Figure S5.

Table S5. Predictive accuracy for P+T and PRS-CS across phenotypes using single-ancestry discovery GWAS from UKBB and BBJ. Related to Figure S6 and Figure S11.

Table S6. Impact of LD reference panel on P+T performance using multi-ancestry GWAS for 17 traits. Related to Figure S7, Figure S8 and Figure S11.

Table S7. Accuracy differences between using PRS derived from multi-ancestry GWAS and using PRS derived from EUR GWAS. Related to Figure S9 and Figure S10.

Table S8. Accuracy differences between PRS derived from multi-ancestry GWAS and using PRS from weighted linear combination. Related to Figure 5 and Figure S12.

Table S9. Predictive accuracy in the UKBB using PRS-CSx for 17 traits. Related to Figure 5 and Figure S12.

Table S10. Predictive accuracy in the UKBB-AFR using various discovery GWAS. Related to Figure 6.

References

1. Inouye, M. *et al.* Genomic Risk Prediction of Coronary Artery Disease in 480,000 Adults: Implications for Primary Prevention. *J. Am. Coll. Cardiol.* **72**, 1883–1893 (2018).

2. Khera, A. V. *et al.* Genome-wide polygenic scores for common diseases identify individuals with risk equivalent to monogenic mutations. *Nat. Genet.* **50**, 1219–1224 (2018).

3. Mars, N. *et al.* The role of polygenic risk and susceptibility genes in breast cancer over the course of life. *Nat. Commun.* **11**, 6383 (2020).

4. Maas, P. *et al.* Breast Cancer Risk From Modifiable and Nonmodifiable Risk Factors Among White Women in the United States. *JAMA Oncol* **2**, 1295–1302 (2016).

5. Craig, J. E. *et al.* Multitrait analysis of glaucoma identifies new risk loci and enables

polygenic prediction of disease susceptibility and progression. *Nat. Genet.* **52**, 160–166

(2020).

6. Majara, L. *et al.* Low generalizability of polygenic scores in African populations due to

genetic and environmental diversity. *Cold Spring Harbor Laboratory* 2021.01.12.426453

(2021) doi:10.1101/2021.01.12.426453.

7. Wang, Y. *et al.* Global biobank analyses provide lessons for computing polygenic risk

scores across diverse cohorts. *bioRxiv* (2021) doi:10.1101/2021.11.18.21266545.

8. Ruan, Y. *et al.* Improving polygenic prediction in ancestrally diverse populations. *Nat.*

Genet. **54**, 573–580 (2022).

9. Wang, Y. *et al.* Theoretical and empirical quantification of the accuracy of polygenic scores

in ancestry divergent populations. *Nat. Commun.* **11**, 3865 (2020).

10. Guo, J. *et al.* Quantifying genetic heterogeneity between continental populations for human

height and body mass index. *Sci. Rep.* **11**, 5240 (2021).

11. Shi, H. *et al.* Localizing Components of Shared Transethnic Genetic Architecture of

Complex Traits from GWAS Summary Data. *Am. J. Hum. Genet.* **106**, 805–817 (2020).

12. Ding, Y. *et al.* Polygenic scoring accuracy varies across the genetic ancestry continuum in

all human populations. *bioRxiv* 2022.09.28.509988 (2022).

13. Pfaff, C. L. *et al.* Population structure in admixed populations: effect of admixture dynamics

on the pattern of linkage disequilibrium. *Am. J. Hum. Genet.* **68**, 198–207 (2001).

14. Pritchard, J. K. & Przeworski, M. Linkage disequilibrium in humans: models and data. *Am.*

J. Hum. Genet. **69**, 1–14 (2001).

15. Fatumo, S. *et al.* A roadmap to increase diversity in genomic studies. *Nat. Med.* **28**, 243–

250 (2022).

16. Hou, K. *et al.* Causal effects on complex traits are similar for common variants across

segments of different continental ancestries within admixed individuals. *Nat. Genet.* **55**,

549–558 (2023).

17. Kim, J., Edge, M. D., Goldberg, A. & Rosenberg, N. A. Skin deep: The decoupling of

genetic admixture levels from phenotypes that differed between source populations. *Am. J.*

Phys. Anthropol. **175**, 406–421 (2021).

18. Maples, B. K., Gravel, S., Kenny, E. E. & Bustamante, C. D. RFMix: a discriminative

modeling approach for rapid and robust local-ancestry inference. *Am. J. Hum. Genet.* **93**,

278–288 (2013).

19. Atkinson, E. G. *et al.* Tractor uses local ancestry to enable the inclusion of admixed

individuals in GWAS and to boost power. *Nat. Genet.* **53**, 195–204 (2021).

20. Pasaniuc, B. *et al.* Enhanced statistical tests for GWAS in admixed populations:

assessment using African Americans from CARe and a Breast Cancer Consortium. *PLoS*

Genet. **7**, e1001371 (2011).

21. Ramirez, A. H. *et al.* The All of Us Research Program: Data quality, utility, and diversity.

Patterns (N Y) **3**, 100570 (2022).

22. Weissbrod, O. *et al.* Leveraging fine-mapping and multipopulation training data to improve

cross-population polygenic risk scores. *Nat. Genet.* **54**, 450–458 (2022).

23. Zhang, H. *et al.* Novel Methods for Multi-ancestry Polygenic Prediction and their

Evaluations in 3.7 Million Individuals of Diverse Ancestry. Preprint at

https://doi.org/10.1101/2022.03.24.485519.

24. Majara, L. *et al.* Low and differential polygenic score generalizability among African

populations due largely to genetic diversity. *HGG Adv* **4**, 100184 (2023).

25. Wray, N. R. *et al.* Pitfalls of predicting complex traits from SNPs. *Nat. Rev. Genet.* **14**, 507–

515 (2013).

26. Daetwyler, H. D., Villanueva, B. & Woolliams, J. A. Accuracy of predicting the genetic risk

of disease using a genome-wide approach. *PLoS One* **3**, e3395 (2008).

27. Nagai, A. *et al.* Overview of the BioBank Japan Project: Study design and profile. *J.*

Epidemiol. **27**, S2–S8 (2017).

28. Morales, J. *et al.* A standardized framework for representation of ancestry data in genomics

studies, with application to the NHGRI-EBI GWAS Catalog. *Genome Biol.* **19**, 21 (2018).

29. Zeng, J. *et al.* Widespread signatures of natural selection across human complex traits and

functional genomic categories. *Nat. Commun.* **12**, 1164 (2021).

30. Berg, J. J. *et al.* Reduced signal for polygenic adaptation of height in UK Biobank. *Elife* **8**,

(2019).

31. Sohail, M. *et al.* Polygenic adaptation on height is overestimated due to uncorrected

stratification in genome-wide association studies. *Elife* **8**, (2019).

32. Martin, A. R. *et al.* Clinical use of current polygenic risk scores may exacerbate health

disparities. *Nat. Genet.* **51**, 584–591 (2019).

33. Lehmann, B. C. L., Mackintosh, M., McVean, G. & Holmes, C. C. Optimal strategies for

learning multi-ancestry polygenic scores vary across traits. Preprint at

https://doi.org/10.1101/2021.01.15.426781.

34. Márquez-Luna, C., Loh, P.-R., South Asian Type 2 Diabetes (SAT2D) Consortium, SIGMA

Type 2 Diabetes Consortium & Price, A. L. Multiethnic polygenic risk scores improve risk

prediction in diverse populations. *Genet. Epidemiol.* **41**, 811–823 (2017).

35. Graham, S. E. *et al.* The power of genetic diversity in genome-wide association studies of

lipids. *Nature* **600**, 675–679 (2021).

36. de Vlaming, R. *et al.* Meta-GWAS Accuracy and Power (MetaGAP) Calculator Shows that

Hiding Heritability Is Partially Due to Imperfect Genetic Correlations across Studies. *PLoS*

Genet. **13**, e1006495 (2017).

37. Ni, G. *et al.* A Comparison of Ten Polygenic Score Methods for Psychiatric Disorders

Applied Across Multiple Cohorts. *Biol. Psychiatry* **90**, 611–620 (2021).

38. Lloyd-Jones, L. R. *et al.* Improved polygenic prediction by Bayesian multiple regression on

summary statistics. *Nat. Commun.* **10**, 5086 (2019).

39. Ge, T., Chen, C.-Y., Ni, Y., Feng, Y.-C. A. & Smoller, J. W. Polygenic prediction via

Bayesian regression and continuous shrinkage priors. *Nat. Commun.* **10**, 1776 (2019).

40. Hou, K. *et al.* Causal effects on complex traits are similar across segments of different

continental ancestries within admixed individuals. *bioRxiv* (2022)

doi:10.1101/2022.08.16.22278868.

41. Kanai, M. *et al.* Meta-analysis fine-mapping is often miscalibrated at single-variant

resolution. *Cell Genomics* 100210 (2022).

42. Coram, M. A., Fang, H., Candille, S. I., Assimes, T. L. & Tang, H. Leveraging Multi-ethnic

Evidence for Risk Assessment of Quantitative Traits in Minority Populations. *Am. J. Hum.*

Genet. **101**, 218–226 (2017).

43. Cai, M. *et al.* A unified framework for cross-population trait prediction by leveraging the

genetic correlation of polygenic traits. *Am. J. Hum. Genet.* **108**, 632–655 (2021).

44. Su, Z., Marchini, J. & Donnelly, P. HAPGEN2: simulation of multiple disease SNPs.

Bioinformatics **27**, 2304–2305 (2011).

45. 1000 Genomes Project Consortium *et al.* A global reference for human genetic variation.

Nature **526**, 68–74 (2015).

46. Manichaikul, A. *et al.* Robust relationship inference in genome-wide association studies.

Bioinformatics **26**, 2867–2873 (2010).

47. Chang, C. C. *et al.* Second-generation PLINK: rising to the challenge of larger and richer

datasets. *Gigascience* **4**, 7 (2015).

48. Willer, C. J., Li, Y. & Abecasis, G. R. METAL: fast and efficient meta-analysis of

genomewide association scans. *Bioinformatics* **26**, 2190–2191 (2010).

49. Li, J. Z. *et al.* Worldwide human relationships inferred from genome-wide patterns of

variation. *Science* **319**, 1100–1104 (2008).

50. Conomos, M. P., Reiner, A. P., Weir, B. S. & Thornton, T. A. Model-free Estimation of

Recent Genetic Relatedness. *Am. J. Hum. Genet.* **98**, 127–148 (2016).

51. Loh, P.-R., Palamara, P. F. & Price, A. L. Fast and accurate long-range phasing in a UK

Biobank cohort. *Nat. Genet.* **48**, 811–816 (2016).

52. Das, S. *et al.* Next-generation genotype imputation service and methods. *Nat. Genet.* **48**,

1284–1287 (2016).

53. Akiyama, M. *et al.* Characterizing rare and low-frequency height-associated variants in the

Japanese population. *Nat. Commun.* **10**, 4393 (2019).

54. Sakaue, S. *et al.* A cross-population atlas of genetic associations for 220 human

phenotypes. *Nat. Genet.* **53**, 1415–1424 (2021).

55. Bycroft, C. *et al.* The UK Biobank resource with deep phenotyping and genomic data.

Nature **562**, 203–209 (2018).

56. Kanai, M. *et al.* Genetic analysis of quantitative traits in the Japanese population links cell

types to complex human diseases. *Nat. Genet.* **50**, 390–400 (2018).

57. International HapMap 3 Consortium *et al.* Integrating common and rare genetic variation in

diverse human populations. *Nature* **467**, 52–58 (2010).

58. Privé, F., Vilhjálmsson, B. J., Aschard, H. & Blum, M. G. B. Making the Most of Clumping

and Thresholding for Polygenic Scores. *Am. J. Hum. Genet.* **105**, 1213–1221 (2019).

1072

1073

Figure 1. Study design in both simulations and empirical analyses

1) In the context of single-ancestry GWAS, we randomly split individuals in European (EUR) and other minority populations, including East-Asian and African populations, into equally sized bins. Simulations involved a total of 52 bins per population, each containing 10,000 individuals. For empirical analysis, bin number was dependent on the sample size of respective phenotype in that population (**Table S3**), with 5,000 individuals per bin. GWAS was conducted within each bin for each individual population, followed by meta-analysis of GWAS from various numbers of bins within each population. To construct PRS derived from single-ancestry GWAS (**PRS$_{single}$**) in the target population, we applied P+T for both simulations and empirical analyses, utilizing PRS-CS for the latter. Subsequently, we combined PRS$_{single}$ developed from EUR GWAS (**PRS$_{EUR\_GWAS}$**) and other minority population-based GWAS (**PRS$_{Minor\_GWAS}$**) through a linear weighted strategy (denoted as **PRS$_{weighted}$**, highlighted in red box) for empirical analyses. Note that PRS$_{weighted}$ was also developed using PRS-CSx, which utilizes GWAS summary statistics from multiple populations. 2) For meta-analyzed multi-ancestry GWAS (referred to as **Meta),** we ran meta-analyses on EUR GWAS and Minor GWAS with varying ancestry compositions. In simulations, we incrementally included 4 bins from EUR GWAS for the meta-analysis, while in empirical analyses, we increased the number to 8 bins. Simultaneously, we varied the number of bins in Minor GWAS from 1 to the total number. Following the meta-analysis, we constructed PRS based on Meta (referred to as **PRS$_{multi}$**), using the P+T method for simulations, and employing both P+T and PRS-CS for empirical analyses.

Figure 2. Improvement of PRS accuracy through meta-analyzed multi-ancestry GWAS compared to large-scale European GWAS across 6 simulated genetic architectures.

The multi-ancestry GWAS included populations of European (EUR) and East-Asian (EAS) ancestry, with the EAS sample size varying as indicated on the x-axis. For illustrative purposes, we present the results using 32 EUR bins, each consisting of 10,000 individuals, which were included in both EUR GWAS and multi-ancestry GWAS. PRS was separately evaluated in African (AFR), EAS and EUR populations. Full results are shown in **Table S1**. M_c indicates the number of causal variants and h^2 refers to SNP-based heritability. In each panel, the red vertical dashed line indicates the point where an equal number of bins from EUR and EAS populations were included in the multi-ancestry GWAS. The error bars represent the standard errors of predictive accuracy differences between PRS derived from multi-ancestry GWAS (PRS$_{multi}$) and PRS derived from EUR GWAS (PRS$_{EUR\_GWAS}$).

Figure 3: Genetic architecture of 17 studied traits between Biobank Japan (BBJ) and UK Biobank (UKBB).

The error bar is the standard deviation of the corresponding estimate. The vertical dashed line was the median estimate. Full results are shown in **Table S4**. The phenotypes were ranked according to their polygenicity estimates using GWAS from UKBB, including: BMI (body mass index), Height, DBP (diastolic blood pressure), SBP (systolic blood pressure), WBC (white blood cell count), Lymphocyte (lymphocyte count), RBC (red blood cell count), Neutrophil (neutrophil count), HB (hemoglobin concentration), HT (hematocrit percentage), Eosinophil (eosinophil count), PLT (platelet count), Monocyte (monocyte count), MCV (mean corpuscular volume), MCH (mean corpuscular hemoglobin), Basophil (basophil count), MCHC (mean corpuscular hemoglobin concentration).

Figure 4. Accuracy improvement of PRS in the UK Biobank East-Asian population (UKBB-EAS) using multi-ancestry GWAS comprare to using European (EUR) GWAS for P+T and PRS-CS.

The multi-ancestry GWAS were obtained by meta-analyzing EUR GWAS and EAS GWAS, with the EAS sample size from the Biobank Japan (BBJ) varying as indicated on the x-axis. For illustrative purposes, we present the results using 64 EUR bins, each containing 5,000 individuals, which were included in both EUR GWAS and multi-ancestry GWAS. PRS were constructed using P+T and PRS-CS and evaluated in the UKBB-EAS. The y-axis is the accuracy difference of PRS when using multi-ancestry GWAS (PRS$_{multi}$) compared to using EUR GWAS (PRS$_{EUR\_GWAS}$). The error bars indicate the standard error of accuracy improvement. The red dashed line is y=0. We showed the results for 7 traits with SNP-based heritability > 0.1 in both BBJ and UKBB, and they were ranked by polygenicity estimates using UKBB (**Figure 3**). Full results are shown in **Table S7**.

Figure 5. Predictive accuracy using different PRS methods in the UK Biobank East-Asian population (UKBB-EAS).

PRS$_{mutli}$ represents PRS derived from multi-ancestry GWAS, while PRS$_{weighted}$ denotes PRS constructed from a weighted linear combination (see **STAR Methods** for details). PRS were constructed with three methods, including P+T, PRS-CS and PRS-CSx. We showed the results for 7 traits with SNP-based heritability > 0.1 in both Biobank Japan (BBJ) and UKBB. Traits were ranked by polygenicity estimates using UKBB (**Figure 3**). Boxes represent the first and third quartiles, with the whiskers extending to 1.5-fold the interquartile range. Full results are shown in **Table S8 and Table S9**.

Figure 6. Accuracy of PRS derived from local-ancestry informed GWAS versus other discovery GWAS in the UK Biobank African population (UKBB-AFR)

We evaluated PRS performance in the UKBB-AFR by utilizing various methods on different discovery GWAS. Specifically, $AFR_{Tractor}$ denotes the AFR-specific GWAS performed using Tractor on the UKBB admixed African-European individuals. $EUR_{Standard}$ refers to standard GWAS performed on the European (EUR) population in the UKBB. $Meta_{Standard}$ is the meta-analysis performed on $AFR_{Tractor}$ and $EUR_{Standard}$. Furthermore, we constructed a weighted PRS by combining PRS generated from $AFR_{Tractor}$ and $EUR_{Standard}$ through a linear weighted approach. The figure shows the results for traits with SNP-based heritability > 0.1 in the UKBB-AFR. Full results are shown in **Table S10**.

Figure 7. General practices for developing PRS using different discovery GWAS.

We summarized the general practice for developing PRS A) using single-ancestry GWAS (PRS$_{single}$); and B) using GWAS from multiple ancestries (PRS$_{multi}$ or PRS$_{weighted}$). Abbreviations: Cross-ancestry genetic correlation (r_g), SNP-based heritability in discovery (h_d^2) and target populations (h_t^2), discovery GWAS sample size (N_d) and the number of genome-wide independent segments in the discovery population (M_d).