Title

Breakdown in seasonal dynamics of ant communities with land-cover change

Authors

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Abstract

Concerns about widespread human-induced declines in insect populations are mounting, yet little is known about how land-use change modifies the dynamics of insect communities, particularly in understudied biomes. Here we examine how the seasonal patterns of ant activity, key drivers of ecosystem functioning, vary with human-induced land cover change on a subtropical island landscape. Using trap captures sampled biweekly from a biodiversity monitoring network covering Okinawa Island, Japan, we processed 1.2 million individuals and reconstructed activity patterns within and across habitat types. We determined that communities inside the forest exhibited more variability than those in more developed areas. Using time-series decomposition to deconstruct this pattern, we found that ant communities at sites with greater human development exhibited diminished seasonality, reduced synchrony, and higher stochasticity compared to those at sites with greater forest cover. We demonstrate that our results cannot be explained by variation in either regional or *in situ* temperature patterns, or by differences in species richness or composition among sites. We conclude that the breakdown of natural seasonal patterns of functionally key insect communities may comprise an important and underappreciated consequence of global environmental change that must be better understood across Earth's biomes.

Keywords

ant, insect, monitoring, Japan, subtropics, synchrony, temporal, variability

Introduction

Insects comprise 95% of described terrestrial animal species on Earth and are key drivers of a multitude of ecosystem functions and services, including pollination, food provisioning, pest control, water filtration, carbon sequestration, and decomposition (Schowalter, 2013). There is also mounting concern about the possibility of long-term global insect declines due to human-induced environmental change, which can disrupt key functions and services (Goulson, 2019; Wagner, 2020). But while such trends measured over broad timescales are informative for monitoring populations (Dornelas et al., 2013), they overlook changes in natural seasonal patterns that can only be detected at fine temporal resolutions. When these seasonal patterns are broken down by human pressures, it can result in the erosion of biodiversity (Tonkin et al., 2017) and degradation of the ecosystem functions and services on which human societies depend (Ross et al., 2021; Stevenson et al., 2015). For this reason, a better understanding is needed of how fine-scale temporal patterns of insect communities are affected in this era of global change.

Land-cover modification is a major driver of biodiversity loss (Cardinale et al., 2012), and many studies have examined its consequences for insect diversity (e.g., Castro et al., 2018; Corro et al., 2019; Knop, 2016; Oliver et al., 2016; Senapathi et al., 2015). However, the corresponding effects on temporal dynamics of insect communities are not well-known. For example, human-altered land cover with higher densities of alien species can lead to boom-and-bust dynamics that cause high variability (Lester & Gruber, 2016; Simberloff & Gibbons, 2004). At the same time, urbanization can lead to biotic homogenization and reduced community variability over time (Hung et al., 2017; McKinney, 2006). Although these documented effects of land-cover change on insect temporal dynamics differ, investigations into how insect communities are affected are crucial for tropical and subtropical biomes undergoing rapid land-cover change (Janzen & Hallwachs, 2021), and also for islands that have restricted land area, frequent introductions of alien species, and exposure to increasingly extreme weather events (Russell & Kueffer, 2019).

Here, we examine the effects of anthropogenic land-cover change on the fine-scale temporal variability of ant community activity. Ants are a dominant insect group in the many environments they inhabit around the world, underpinning a broad spectrum of ecological interactions and playing a major role in regulating ecosystem function (Hölldobler & Wilson, 1990), which makes them ideal bioindicators for ecosystem change (Andersen & Majer, 2004). Community seasonality is linked to the maintenance of key ecosystem functions and services (Stevenson et al., 2015). However, even though ants can exhibit significant seasonality in tropical and subtropical regions (Basu, 1997; Samways, 1990), little is known about how land-cover change affects ant community seasonality in these diverse biomes.

To address this, we used high-resolution (biweekly) monitoring data on ant activity for subtropical Okinawa Island in southern Japan to examine whether, and how, the temporal variability of ant communities varies with land cover change. We then determined which temporal components of variability (i.e., seasonality, trend, or stochasticity) contribute the most to differences across the gradient (Fig. S1). We also tested whether other potential drivers, including differences across sites in climate seasonality, species richness, and community composition, could be responsible for observed differences in temporal community variability, and additionally examined the roles played by invasive species in moderating these effects.

Materials and Methods

Study sites and land-cover data

Our study focuses on the subtropical main island of Okinawa in the southern Ryukyu archipelago of Japan. The island has a historical land-cover gradient, spanning from the minimally developed north to the more urban south (Fig. 1). Okinawa's extensive Yambaru Forest in the north was recently recognized as part of Japan's newest Natural World Heritage Site for its high endemism and biodiversity. Although the island is only 0.3% of Japan's total land area, it provides habitat to more than one-third of the country's ant species (Terayama et al. 2009, Terayama et al. 2014). The few studies of ant communities in Okinawa include biodiversity surveys (e.g., Ito et al., 1998) and comparisons between native and alien ant populations in Yambaru (Katayama & Tsuji, 2010; Suwabe et al., 2009; Yamauchi & Ogata, 1995). This study, the first to sample island-wide ant diversity and temporal patterns in Okinawa, uses data from a biodiversity monitoring system covering the land-cover gradient of the island, administered by the Okinawa Environmental Observation Network (OKEON) Churamori Project (https://okeon.unit.oist.jp/).

We used a land-cover map for Okinawa (year 2015) to calculate representative land-cover values for the 24 OKEON monitoring sites. The sites (pairwise distances: minimum = 389 m, mean = 33 km, maximum = 102 km) are located in or near forested areas across multiple land-cover types varying from broadly contiguous forest to highly agricultural or urban areas (Fig. 1). To characterize each site, we calculated the proportion of seven main land-cover classes (forest, agriculture, urban, grass, sand, freshwater, miscellaneous) around circular buffers with radius 1 km (see Ross et al., 2018 for details). Buffers with this size were chosen to characterize representative land-cover values for areas surrounding sites without including regions that may have resident unsampled colonies. Sites near the coast included ocean area within the buffer, so we used relative proportions for all sites that considered only land area. As correlations were high between the main land-cover classes of interest (that is, forest, agriculture, and urban; Table S1), we used principal component analysis (R function *prcomp*) after applying an arcsine transformation to the proportions to improve normality (e.g., Van Buskirk, 2005) and retained the first and second axes (variance explained: PC1 = 81% and PC2 = 11%) for use as explanatory variables. PC1 represents the forested (high) to developed (low; urban and/or agriculture) gradient, while PC2 represents the rural (high; agriculture and/or grass) to urban (low) gradient (Fig. S2). All analyses were conducted in R (R Core Group, 2021), and spatial analyses were conducted with R packages *sf* (Pebesma, 2018) for vector data and *raster* (Hijmans, 2021) for gridded data.

Ant activity data

We sampled worker ant activity biweekly with Sea, Land, and Air Malaise (SLAM) traps for two years and identified samples to species level, resulting in 1,378,324 classified individuals. Specifically, we sampled one SLAM trap (Large BT1005, MegaView Science Co., Ltd.) at each of three stations per OKEON site (pairwise distances within sites: mean = 90 m, minimum = 19 m, maximum = 195 m) every two weeks from March 2016 to March 2018 (further details in Supporting Information A). The stations are predominantly deployed under forest canopy, though several are in nearby open areas. This sampling method mainly targets terrestrial browsing species, though does occasionally catch arboreal and subterranean individuals. Insect

count data sampled with passive traps are usually interpreted as a measure of activity, or the rate at which individuals intersect a point in space, rather than direct estimates of abundance, though changes in abundance are generally reflected in ant activity data (Kaspari et al., 2022a).

We additionally categorized all species by invasive status as either native, alien, or uncertain based on expert opinion and biogeographical databases (Guénard et al. 2017, Janicki et al. 2016, Terayama et al. 2009, https://www.antwiki.org/, https://www.antweb.org/). Native refers to species with enough data to establish historical occupancy in Okinawa without human-aided dispersal from other regions. Alien refers to species that definitively have native ranges in different regions and have dispersed to Okinawa via human activity. Uncertain species lack sufficient data to make such determinations.

Some prevalent species had extreme count outliers, such as *Technomyrmex brunneus* (representing counts greater than 5,000), likely caused by large colony sizes and/or increased activity following mating flights, heterospecific competition, or other biological phenomena. We avoided bias in variability estimates by limiting count values to a conservative threshold of 500 individuals for a given species at a station within a two-week period. We also examined changes in our results for a range of thresholds. For each species during each sampling period, we summed counts over the three stations per site to determine site-level activity (Fig. 2), and over all sites per land-cover group for grouped species-level activity. We then calculated total site density as the sum of all counts per site.

Sampled abundance and variability can be correlated due to sampling error alone because of mean-variance scaling (McArdle & Gaston, 1995). For example, even if species dynamics are identical across two sites, if we sample fewer individuals from site A than site B (e.g., due to effects of sampling station placement), site B may have a higher temporal variability simply because its counts are higher. While in some sense this could reflect the actual variability of observed species at specific locations (that is, at the stations), we wanted to disentangle differences in community variability from those attributable to sampling error. Thus, we rarefied the data from all sites so that the total count (sum of all samples across the two years) for each site equals the total count at the site with the least abundance (Fig. S1, Supporting Information B; Gaston & McArdle, 1994). Finally, to determine whether species richness differences could drive any observed differences in temporal variability, we calculated site-specific total species richness, and the richness of each invasive status group from the observed and rarefied data, making extrapolations to account for incomplete sampling using Hill numbers ($q = 0$) with the R package *iNEXT* (Hsieh et al., 2016).

Temporal variability and time-series decomposition

We calculated two aspects of temporal variability: functional and compositional (Supporting Information C). Functional variability refers to activity changes at the aggregate scale (e.g., Hillebrand & Kunze, 2020) and is calculated as the coefficient of variability (ratio of the standard deviation to the mean) of total ant counts across all species per site (McArdle & Gaston, 1995). Compositional variability estimates the variation in community composition over time and is calculated as temporal beta diversity (Legendre & De Cáceres, 2013) using the R package *adespatial* (Dray et al., 2022), which additionally records species' individual contributions to

beta diversity. As we use ant count data, compositional changes would thus represent changes in relative activity rates of different species in the local species pool, which in turn is a function of the abundance and patterns of behavior (Kaspari et al., 2022a).

Next, we used temporal decomposition models to determine the relationships of temporal variability components with observed patterns. Time series can be decomposed into components describing different underlying temporal patterns in the data: "seasonal" processes that repeat cyclically over some temporal scale (i.e., phenology), "trend" that describes directional change, and "remainder" that represents any residual variation not captured by the other components (i.e., short-term stochastic fluctuations; Hyndman & Athanasopoulos, 2021). A community with high temporal variability may have strong cyclical patterns, increasing or decreasing trends, high stochasticity, or some combination of these. For temporal datasets spanning many annual cycles, models that can estimate complex seasonal responses (e.g., wavelet analysis; Tonkin et al., 2017) are often employed, but for studies spanning few annual cycles at a fine temporal resolution, simpler models with fewer assumptions and problems with overfitting are preferable. We estimated additive temporal components for the ant activity data by fitting time-series linear models with temporal predictors using the R package *fable* (O'Hara-Wild et al., 2021). As our time-series was short, we used simple predictors consisting of a linear trend and a seasonal signal approximated with an annual Fourier term with the simplest maximum order (K) of 1—this models seasonality as a sine wave (Hyndman & Athanasopoulos, 2021). We fitted this model to the rarefied data at both the site-level and species-level for land-cover groups, then decomposed the count value at each time-step. For each site, we measured the absolute variance of each temporal component, but also the relative component variance, defined here as the individual component variance divided by summed variance of all components (Supporting Information D).

Finally, as spatial autocorrelation between sites did not affect results (Table S3; Supporting Information E), we used simple linear models to estimate relationships between land cover and our focal metrics of temporal variability. Specifically, we fitted models for total ant counts (before rarefaction), functional and compositional temporal variability, and absolute and relative variances of time-series components. To explore differences between native and alien species, we also fitted models separately by invasive status for site richness and summed species' contributions to beta diversity. We performed model selection on linear combinations of the predictor variables land-cover PC1 and PC2 using the Akaike Information Criterion corrected for small sample sizes (AICc).

Assessing effects of differences in temperature and community composition on observed seasonality patterns

To determine if other differences among sites may be responsible for any observed variation in ant community activity and relationships with land cover, we additionally examined differences in 1) regional and site-level temperature and 2) tested whether seasonality differences persisted after standardizing community composition between land-cover groups. For (1), we downloaded regional climate data for the collection period from the Japan Meteorological Agency (JMA) database (http://www.jma.go.jp/en/amedas_h/map65.html, accessed 04/01/2020) and calculated temperature means and extremes from the six climate stations that collect temperature data (Miyagijima, Itokazu, Naha, Ashimine, Nago, Oku). We also collected *in situ* site-level data

throughout the collection period at one station per site to characterize local air and soil temperature patterns (WatchDog 2900ET Weather Station 1.5 m above ground and SMEC WaterScout 300 Soil Sensor 10 cm below ground, Spectrum Technologies). With the same workflow as for the ant activity data, we calculated absolute variance of seasonality for regional and site-level temperature time-series, then estimated relationships with land-cover PC1 using linear models.

For (2), we conducted analyses to test whether differences in community composition alone, without variation in seasonality within species, could be driving the observed patterns in seasonality, or alternatively if habitat-dependent dynamics of species occurring in both habitat types were responsible. We used land-cover PC1 to partition the sites into two groups with eight sites each—forested and developed (comprising sites with more agricultural and/or urban land cover)—and retained only those species shared between groups and with total counts greater than 100. We excluded the eight sites with intermediate values of PC1 to better represent the most characteristic sites for each land-cover group. To make balanced comparisons between land-cover groups, we rarefied each species' activity data to the minimum count of that species between groups (Fig. S1, Supporting Information B). To determine whether community composition affected our results, we fitted time-series linear models to each species in the land-cover groups and calculated the relative variance of their seasonality (forested versus developed sites). As differences between the land-cover groups were not normally distributed (Shapiro-Wilk test; $p = 0.011$), we used the nonparametric Wilcoxon signed-rank test to determine whether the same species between groups exhibit significant differences in relative seasonality variance. Additionally, to determine how temporal mismatch changes with invasive status, we calculated the synchrony of seasonality for each invasive status between land-cover groups and standardized them to the same scale. We used the R package *codyn* (Hallett et al., 2016) to calculate synchrony using the method described by Loreau & De Mazancourt (2008). We then compared seasonal curves to the average seasonality curve for each land-cover group to visualize the degree of temporal mismatch by invasive status.

Results

Our sampling for 2016–2018 resulted in the recovery of >1.2 million individuals (before thresholding) and 16,595 unique count records at the site-level across 91 ant species. Total count was less than 5 individuals for 21 species and greater than 10,000 individuals for 15 species, with a minimum of 1 (*Aenictus ceylonicus*, *Crematogaster* c.f. *matsumurai*, *Ectomomyrmex sp.*, *Erromyrma latinodis*, *Hypoponera* sp., *Leptogenys confucii*, *Strumigenys hirashimai*, *Strumigenys mazu*) and a maximum of 259,180 (*Tetramorium bicarinatum*). The pattern of ant activity dropping in the winter and rising in the summer qualitatively diminished as human disturbance increased (Fig. 2). Total ant activity (after thresholding to 500) had a nearly 23-fold difference between the site with the minimum count (Yona Forest, $n = 3616$) and that with maximum count (Sueyoshi Forest, $n = 82{,}781$). The natural log of total density had a strong negative relationship with land-cover PC1 ($R^2 = 0.71$, $p < 0.001$), indicating that ant activity overall was higher in developed areas than forested areas (Fig. 3). Rarefaction resampled all sites to the minimum site count before analysis.

The linear models suggested strong relationships with the land-cover PC1 for most metrics of temporal variability (Figs. 3, 4). Based on model selection via AICc, all models included PC1 as the sole predictor except the relative variances of seasonality and stochasticity, which also used PC2 (Table S3). Functional and compositional temporal variability both had strong positive relationships with PC1 (functional: $R^2 = 0.67$, $p < 0.001$; compositional: $R^2 = 0.60$, $p < 0.001$; Fig. 3), and these two variabilities also had a strong relationship to each other ($R^2 = 0.62$, Fig. S3). The species' contributions to beta diversity (Fig. S4) that explained individual species' impacts on compositional variability were also positively correlated with PC1 for native species ($R^2 = 0.40$, $p < 0.001$), but uncertain ($R^2 = 0.28$, $p < 0.01$) species had a weaker, negative relationship (alien species showed no relationship). Absolute seasonality and stochasticity variance had positive relationships with PC1, though that for seasonality ($R^2 = 0.57$, $p < 0.001$) was stronger than that for stochasticity ($R^2 = 0.15$, $p < 0.05$; Fig. 4). On the other hand, relative seasonality variance ($R^2 = 0.23$, $p < 0.05$) was positively related to PC1, while relative stochasticity variance ($R^2 = 0.36$, $p < 0.01$) had a negative relationship (both had $p > 0.05$ for PC2; Fig. 4). All linear regression results are found in Table S3.

We found no evidence that differences in species richness, temperature at different scales, or species composition are candidates as drivers of observed differences in ant community temporal variability. We found no relationship between PC1 and total species richness (either observed or extrapolated), and a weak correlation with rarefied richness ($R^2 = 0.15$, $p < 0.05$) due to the loss of rare species during the rarefaction process, though we consider this unlikely to bias our results (Fig. S5). We did find strong positive relationships with PC1 for native richness values (observed: $R^2 = 0.47$, $p < 0.001$; extrapolated: $R^2 = 0.35$, $p < 0.005$; rarefied: $R^2 = 0.61$, $p < 0.001$) indicating that more native species can be found in the more forested sites, and weaker negative relationships with alien richness values (observed: $R^2 = 0.29$, $p < 0.005$; extrapolated: $R^2 = 0.20$, $p < 0.05$; rarefied: $R^2 = 0.20$, $p < 0.05$) (Table S3, Fig. S5). Moreover, we found no clear patterns in the absolute variance of regional temperature seasonality from JMA climate stations spanning the island from north to south, nor did we find relationships between the absolute seasonality variance of *in situ* air or soil temperature and PC1 (Fig. S6). Only two sites had exposed sensors that were not located below canopy cover (Oyama Park and OIST Open [OYA and OIT, respectively, in Fig. S6]), and these had higher absolute variances than the other sites. Concerning the tests of different threshold sizes, although we observed some differences, our results remained similar (Supporting Information F, Fig. S7).

Lastly, individual species showed differing seasonality between land-cover groups (more forested and more developed). Comparisons revealed that, in general, the same species have a higher relative seasonality variance at forested sites than at sites with more human development ($p < 0.005$), though the alien species *Tetramorium lanuginosum* was a notable exception (*Tlanu* in Fig. 5). While seasonality was stronger in the forested group for all species, when separated by invasive status we observed higher synchrony in the forested group for native ($n = 11$, forested: 0.75, developed: 0.58) and uncertain species ($n = 8$, forested: 0.5, developed: 0.28), but the opposite pattern for alien species ($n = 4$, forested: 0.73, developed: 0.86) (Fig. 5).

Discussion

Our study sheds light on how ant community activity dynamics vary across a gradient of human-induced land-cover change in a subtropical island environment. Our results complement an extensive number of recent studies focusing on the temperate zone that examine change of insect populations over time (Hallmann et al., 2017; Kaspari et al., 2022a; Seibold et al., 2019; Uhler et al., 2021), including those with enough intra-annual sampling to examine changes to seasonality (Kaspari et al., 2022b). The notable difference is that because climatic seasonality is much weaker in the subtropics, differences in seasonality of insect populations across environmental gradients are likely to be more difficult to detect than for populations in the temperate zone that experience stronger seasonal contrasts (Kishimoto-Yamada & Itioka, 2015).

We found that temporal variability diminishes for the subtropical ant communities of Okinawa Island as land cover transitions from forested to more developed, and that this is related to a breakdown in seasonality. Forested sites were more variable overall, but their resident ant communities were more seasonal and synchronous, while community variability for developed sites (i.e., urban and agricultural) was characterized by stochastic variation. Although native ant richness was higher in forested sites, we determined the observed differences in seasonality across the land-cover gradient were not due to differences in regional or local temperature patterns, nor to species composition among sites. Compositional differences between communities in different land-cover types, such as the prevalence of one or more highly seasonal species, could drive observed relationships between community seasonality and land cover. However, we found that populations of the same species occupying forested sites were more seasonal than those with more human development. Further, when we separated species by invasive status, we found that the strong seasonal patterns we observed in the forest were more synchronous for native than alien species, though both synchrony values are reasonably high and the considerably smaller sample size for alien species should be taken into consideration.

Our results suggest that increasing agricultural and urban development can dampen seasonal community variability and thus disrupt phenological patterns for insects. In a number of studies, loss of seasonality has been explained by biotic homogenization, whereby communities become more similar and generalist, and this has been linked with urbanization for a variety of taxa, including bees (Hung et al., 2017), bugs and leafhoppers (Knop, 2016), flowering plants and butterflies (Uchida et al., 2018), and birds (La Sorte et al., 2014; Leveau, 2018). However, the majority of studies on the temporal variability of natural communities compare measurements between years and rarely within them (e.g., Cottingham et al., 2001; de Mazancourt et al., 2013; Olivier et al., 2020; Tilman et al., 2006), meaning that loss of seasonality due to environmental or anthropogenic factors may be overlooked. Yearly samples taken for highly seasonal communities at peak productivity over many annual cycles should show higher stability than for a more stochastic community, yet more frequent sampling would reveal considerable variation due to higher synchrony and seasonality. As other studies have demonstrated, temporal scale is a crucial factor that determines how stability should be interpreted (Clark et al., 2021; McArdle et al., 1990), and as community variability and its drivers vary across temporal scales (Hatosy et al., 2013), it is important to measure community responses to the environment and anthropogenic disturbance across scales to improve future predictions.

We found a positive relationship between functional and compositional temporal variability, meaning that those sites with communities that varied more as a whole tended also to have higher total variation of the individual species in the local pool. The relationship between functional and compositional variability can range from positive to negative depending on the system and exposure to disturbance (Hillebrand et al., 2018; Ross et al., 2022; White et al., 2020). Two characteristics of more forested sites may help explain this positive relationship. The first is that more forested sites had higher variability in general at the biweekly timescale. The second is that, at more forested sites, native species had higher richness and the highest contributions to beta diversity. Sites with high richness that include substantial numbers of alien species can have community dynamics that differ considerably from similar sites in which native species dominate (Krushelnycky & Gillespie, 2008; Sanders et al., 2003), thus including alien species in richness estimates can affect how diversity-stability relationships are interpreted (e.g., Moore & Olden, 2017).

Our results raise the question of what mechanisms drive these patterns. We outline several hypotheses for why seasonal variation in activity could vary across the land-cover gradient. First, there is a possibility that ants are simply responding to the temperature they experience on a physiological level, reflecting temperature and microclimatic variation across habitats. Ants are well-known to be sensitive to temperature variation, with foraging limited to certain temperature ranges (Bernstein 1979, Stuble et al., 2013). Regional climatic variation across the island is very modest, and with no consistent differences in seasonality. Further, we found no relationships between the seasonal variance of *in situ* air and soil temperature and the main land-cover gradient PC1. However, differences in canopy openness, vegetation, and urbanization (via heat-island effects, etc.) may create conditions that are relevant for promoting and depressing ant foraging, and the prevalence of such conditions might be higher in more degraded habitats. Comprehensive measurements with small sensors at various habitat strata for representative sites could help elucidate whether this phenomenon could differ across land-cover types. Another possibility is that ants are responding to resource availability patterns. In this scenario, it is possible that resources are more seasonal inside the forest rather than more human-impacted areas. This could occur, for example, because of anthropogenic food sources that are more static throughout the year, or because of differences in phenology of vegetation inside or outside the forest (Penick et al., 2015). A related question is whether these differences in seasonality are limited to ants or are observed across the arthropod community. Were arthropod seasonality responding more generally to land cover, this would provide a resource pattern that could affect ant activity. Related to this, it is unclear why there seems to be little seasonal niche differentiation for ant species in forested areas, which have activity patterns that follow temperature quite closely in synchrony. One possibility is that species become most active at different times during the same warm season to avoid competition—we saw some evidence of this in the standardized seasonal curves (Fig. 5). This could be due to the inability to achieve high activity levels during cooler months due to physiological limitations, which may be absent in more developed environments with anthropogenic temperature refugia. Other potential explanations relating to differences in the ecology or physical environments across the two habitats likely exist, but more targeted research is needed to address this question.

We measured variability over two years, but there may be other types of variability that only manifest on longer timescales and are also correlated with land cover. For example, introduced

ants are known to have boom-bust dynamics when they reach new localities (Lester & Gruber, 2016). Over longer timescales, communities in more human-dominated areas may prove to be more dynamic with successive shifts in the community. Indeed, we witnessed such an event in our two-year period at one of our sites (Sefautaki Forest), where *Pheidole megacephala*, a notoriously impactful invasive ant, arrived and quickly reached high abundance, raising variability of this site beyond what was typical of others with similar land-cover characteristics. Recent research in Yambaru Forest has documented resilience of native ant communities to disturbance based on surveys of invasions at developed areas after regrowth (Shimoji et al., 2022), so this site too may eventually recover to its previous structure. Thus, successive regime shifts in these more disturbed communities could occur over longer timescales, but only monitoring over more years would reveal such dynamics.

Conclusions

This study provides novel insight into the seasonal dynamics of subtropical insect communities across differing levels of human development. Importantly, it is the first study to our knowledge to measure seasonality of ant activity with high-frequency sampling across a human disturbance gradient. The results of our study, which link anthropogenic disturbance in the form of agricultural and urban development to a reduction in intra-annual temporal variability due to weakening seasonality, support the growing evidence that human development reduces natural seasonal patterns (Hung et al., 2017; La Sorte et al., 2014; Leveau, 2018; Uchida et al., 2018). Although this study focuses on correlative relationships and thus cannot fundamentally demonstrate causal mechanisms, it does provide compelling evidence that continuing habitat loss and fragmentation can lead to increasing homogenization of insect communities in parts of the globe with the highest insect diversity, the subtropics and tropics. As loss of seasonal patterns has been linked to a subsequent reduction in key ecosystem functions and services (Ross et al., 2021; Stevenson et al., 2015), increasing development could result in disruptions to nutrient cycling and food security (Bommarco et al., 2013). To better understand the generality of these patterns, future studies using high-frequency sampling should focus on monitoring a greater diversity of insect communities across different habitat strata (i.e., arboreal, subterranean; Gotelli et al., 2011) in a variety of biomes. Greater focus on the impacts of land-cover change on insect community seasonality should aid ongoing management and conservation efforts to help preserve the important ecosystem functions and services insects provide.

Data Availability

Raw species' count data before processing and R code used for analysis will be found on Data Dryad after article acceptance. Please contact <jamie.m.kass (at) gmail.com> with data requests.

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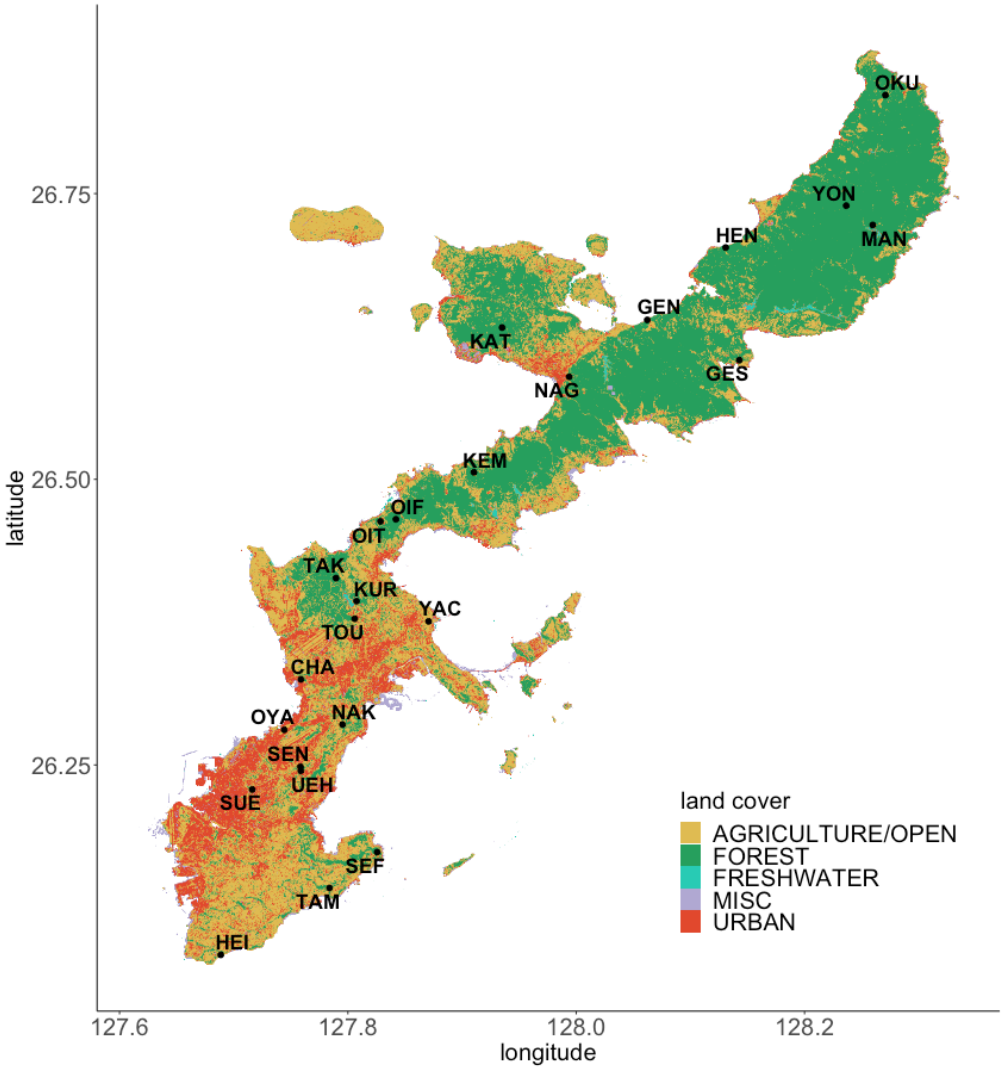

Figure 1. Major land-cover classes of Okinawa Island with locations of the OKEON monitoring stations, which cover the full gradient of human development across the island. Ants were sampled from three stations at each sampling site biweekly from 2016–2018 (total of 52 sampling periods).

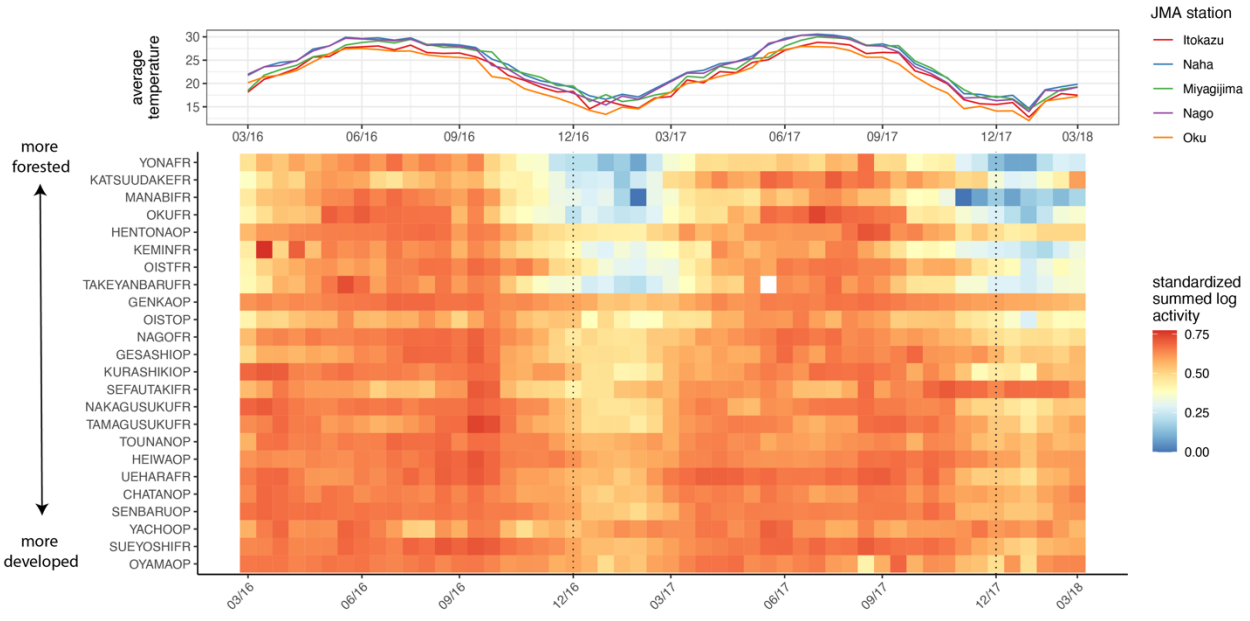

Figure 2. Heat map showing the standardized ant activity (i.e., logs of summed site counts divided by the log of the grand total count) over two years in Okinawa. Mean temperature measurements from the Japan Meteorological Agency across the island, ordered from south to north, are included for reference. Heat map sites are ordered by values of land-cover PC1, explaining the anthropogenic stress gradient (top: more forested, bottom: more urban). Dotted lines show the last sampling period before the new year.

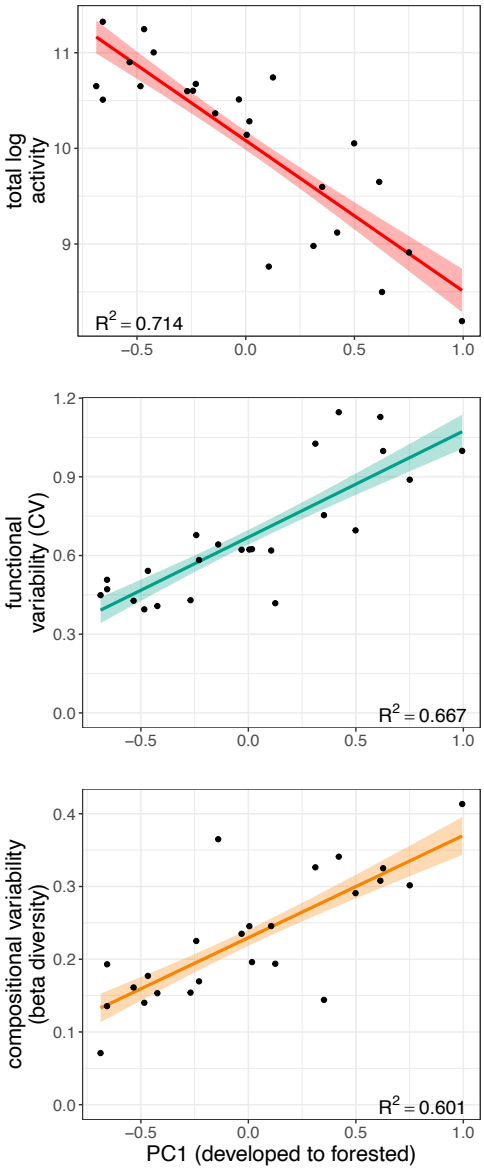

Figure 3. Trends in total log activity, functional variability, and compositional variability for ants across a gradient of anthropogenic impact. Relationships with PC1, explaining the anthropogenic stress gradient (low: more developed, high: more forested). Total log ant activity (before rarefaction) correlates negatively with PC1 while indices of temporal variability (functional: coefficient of variation, compositional: beta diversity) correlate positively.

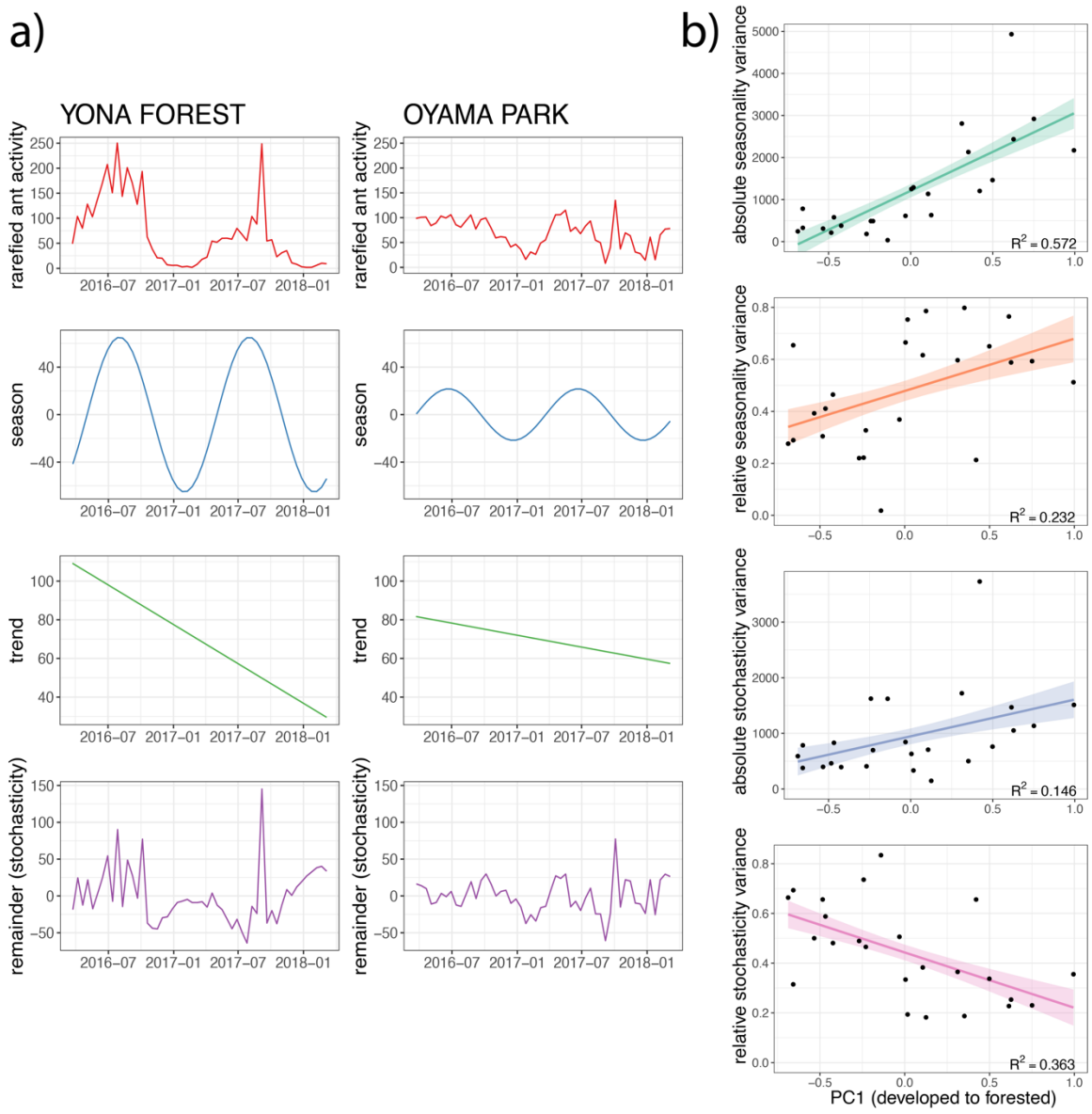

Figure 4. a) Time-series decomposition of rarefied ant activity for the sites with the most forest (Yona Forest) and most human development (Oyama Park) within a 1 km buffer. b) Relationships between absolute and relative variance of seasonality and stochastic temporal components and PC1 across all 24 sites, explaining the anthropogenic stress gradient (low: more developed, high: more forested). Relative variance here is calculated as individual component variance divided by summed variance of all components.

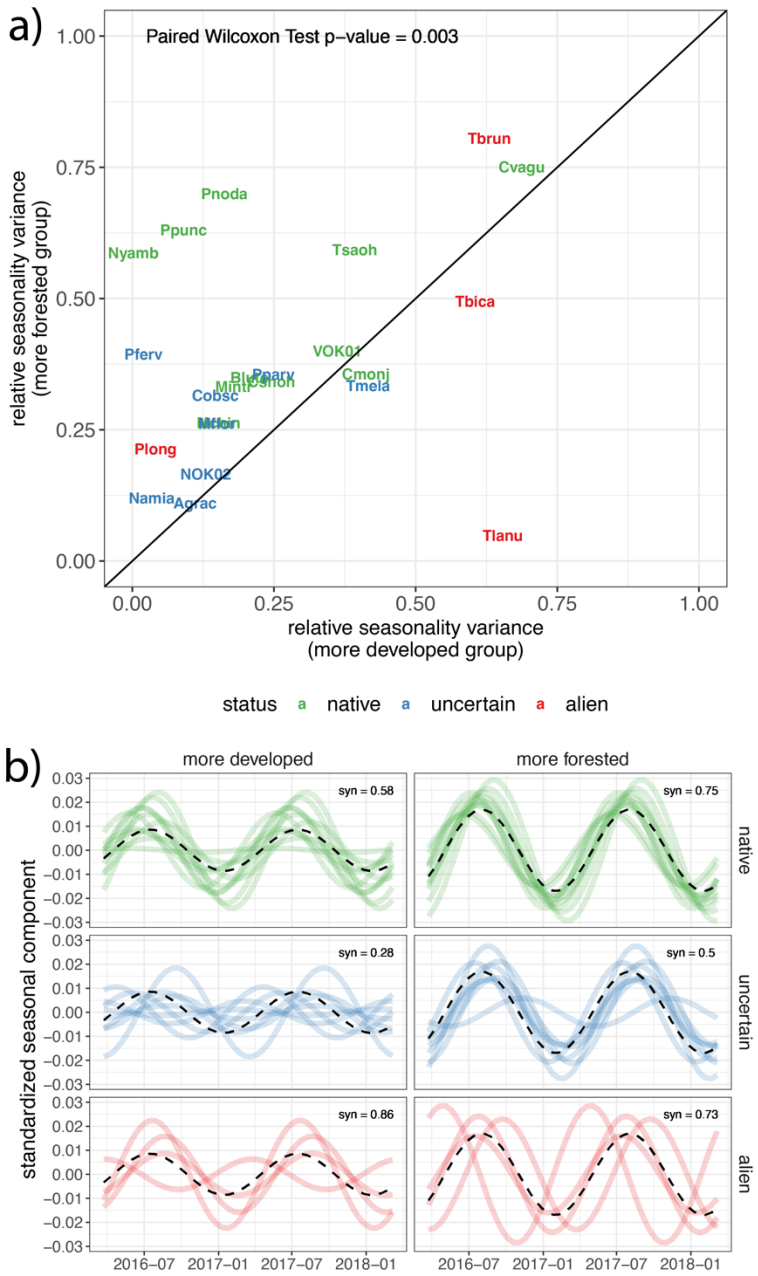

Figure 5. a) Relative seasonality variance calculated per ant species in two representative land-cover groups each containing eight sites: more forested and more developed. Only species shared between groups (abbreviations of those found in Table S2) and with total counts ≥100 were retained, whereupon species counts were rarefied to the minimum between groups. On average, species have a higher relative seasonality variance in the forested group (paired Wilcoxon Test; p = 0.003), and no apparent patterns exist regarding alien status. b) Seasonal component time-series for each species (lines) found in the more forested and more developed groups, standardized by total species activity and separated by invasive species status. The dashed black line shows the mean standardized seasonality across all species per land-cover group, showing much greater temporal mismatch in ant communities at sites with greater human development.

692 Supporting Information

693 A. Ant activity data details

694 As sites across the island were sampled in sequential groups rather than simultaneously for each
695 biweekly period (differences are on the order of 7 days), to allow for appropriate site
696 comparisons, we assigned each sample a time step (52 in total) based on the relative sampling
697 order for that site. To this activity dataset, we filled in absence data when species that were
698 observed at least once at a site were not observed for other time steps for that site in order to
699 calculate temporal variability metrics. Values of NA (n = 8) were assigned to samples that
700 experienced trap malfunctions or data loss via extreme weather events and were removed from
701 analyses.
702

703 B. Rarefaction permutations

704 Community samples collected over time are typically biased by sampling error, which results in
705 fluctuations in activity due to site placement or other external factors, but this bias can be
706 reduced via data rarefaction (Gaston & McArdle, 1994). As natural differences in total
707 community count across sites were high (from n = 3,616 [Yona Forest] to n = 82,781 [Sueyoshi
708 Forest]), we rarefied site- and group-level count while retaining original activity patterns to
709 reduce the effects of sampling error, allowing us to make more even comparisons of temporal
710 variability. For site-level activity, we randomly resampled the individual counts of species i at
711 time t for each site to equal the minimum total site count (n = 3,616). For group-level activity,
712 we first limited the species considered to those with total count greater than 100 for the two-year
713 study period, then randomly resampled (with replacement) the individual counts of species i at
714 time t for each group to equal that species' minimum total count between groups. We performed
715 1000 iterations of each rarefaction and conducted all proceeding analyses on these iterations.
716

717 C. Community and compositional variability calculations

718 We measured community temporal variability by calculating the coefficient of variation (CV) of
719 summed ant count by site over the 2-year sampling period. We measured compositional temporal
720 variability by calculating temporal beta diversity for each site community as the total variability
721 of the species composition matrix (here, species × time step) after a Hellinger transformation (to
722 ensure purely relative count data; Legendre & De Cáceres, 2013) using the function *beta.div*
723 from the R package *adespatial* (Dray et al., 2022). As an additional product of this function, we
724 derived species contributions to beta diversity per site, which is calculated as species variance
725 divided by the total community variance (summing to 1). We then summed these contributions
726 for each invasive status category (native, uncertain, alien) to determine the contribution of each
727 category. We calculated both indices of temporal variability on each rarefied site activity dataset,
728 then found the mean values across datasets. These mean rarefied site values were used in
729 subsequent models.
730

731 D. Formulas for temporal decomposition and associated variance

732 We used the TSLM() function in the R package *fable* (O'Hara-Wild et al., 2021) to fit time-
733 series linear models with the following formula:

$$y_t = \beta_0 + \beta_1 t + \beta_2 x_1$$

where y_t is the total count of ants at a site, with t equal to the range 1 to T, the number of time intervals; and where x_1 is the 1st order Fourier term $\sin(\frac{2\pi t}{m})$, with m equal to the number of seasonal periods (in this case, $m = 2$ for 2 years).

We calculated the temporal components in the following ways:

$$trend = \beta_0 + \beta_1 t$$

$$season = y_t - trend - resid$$

$$remainder = resid$$

where *resid* is the model residuals.

In addition to calculating absolute variance of temporal components, we calculated relative variance with the following formula:

$$relative\ var_z = \frac{var(z)}{var(trend) + var(season) + var(remainder)}$$

where z is one of the temporal components: trend, season, or remainder.

E. Assessing effects of spatial autocorrelation on models

We first built models using ordinary least squares (OLS) regression to estimate the relationships between land cover and the variables of interest. As some OKEON sites are relatively close to others, we assessed the spatial autocorrelation present in the OLS model residuals by calculating Moran's I and testing with random permutations using the function *moran.randtest*() from the R package *adespatial* (999 repetitions; Dray et al., 2022). But as this test was not significant for any of the models ($p > 0.05$), we fitted OLS models instead of more complex generalized least squares (GLS) models with spatial structure. We determined goodness-of-fit with the adjusted coefficient of determination (R-squared) and compared these between different count threshold choices.

F. Effects of count threshold choice

As we avoided bias in temporal variability estimates by limiting all count values to an assigned threshold, we examined how robust our results are over a range of different thresholds. We report results in the main text for the relatively conservative threshold of 500 set at the station-level, but additionally ran the same analyses with no threshold, and with thresholds 100, 200, 1000, and 2000, applied at both the station- and site-levels. We observed some differences across

the different thresholding schema, but no threshold fundamentally changes our results (which we report for threshold 500 set at the station-level). The adjusted R-squared values for the metrics calculated were sensitive in varying degrees to threshold (Fig. S7), with the exception of compositional temporal variability—this is because our beta diversity calculation includes a Hellinger transformation that mollifies the effects of outliers in the data. Further, thresholding at the station- and site-levels showed the same general trends. The highest threshold considered (2000) for the station-level was most often associated with the lowest R-squared values. The absence of a threshold had the most dramatic differences. This reduced the R-squared value for functional temporal variability to 0.11 from a thresholded minimum of 0.40, and seasonality variance to 0.16 from 0.39. R-squared values for other metrics were similarly reduced, but more closely followed decreasing trends as thresholds became larger. These reductions in correlations with PC1 occurred because, without thresholding, sites with extreme count values for one or several species spike in functional variability (calculated with CV) and lose seasonality, and the existence of such spikes has little to do with land cover.

Table S1. Pearson correlations between proportions of land-cover classes within 1 km buffers of sampling sites. “MISC” stands for miscellaneous and includes the classes for unclassified pixels and those classified as bare rock or soil. As pairwise correlations were high, we conducted a principal component analysis (PCA) to derive the main orthogonal land-cover gradients.

	AGRICULTURE	FOREST	WATER	GRASS	SAND	URBAN	MISC
AGRICULTURE	1						
FOREST	-0.89	1					
WATER	-0.05	0.02	1				
GRASS	0.53	-0.44	0.02	1			
SAND	0.54	-0.48	-0.04	0.23	1		
URBAN	0.56	-0.84	-0.08	0.01	0.20	1	
MISC	0.51	-0.73	0.13	0.18	0.53	0.74	1

797 Table S2. Invasive status for all observed ant species based on existing literature and expert

798 opinion.

799

species	**status**
Aenictus ceylonicus	native
Aenictus lifuiae	native
Anoplolepis gracilipes	uncertain
Aphaenogaster concolor	native
Aphaenogaster irrigua	native
Brachyponera chinensis	native
Brachyponera luteipes	native
Camponotus bishamon	native
Camponotus devestivus	native
Camponotus monju	native
Camponotus OK01	native
Camponotus yambaru	native
Cardiocondyla kagutsuchi	uncertain
Cardiocondyla minutior	uncertain
Cardiocondyla obscurior	uncertain
Cardiocondyla wroughtonii	native
Carebara hannya	native
Carebara oni	native
Carebara yamatonis	native
Colobopsis shohki	native
Crematogaster cf. matsumurai	native

Crematogaster nawai	native
Crematogaster vagula	native
Cryptopone tengu	native
Diacamma OK01	native
Discothyrea kamiteta	native
Ectomomyrmex OK01	native
Erromyrma latinodis	native
Euponera pilosior	native
Hypoponera nippona	native
Hypoponera OK01	native
Hypoponera punctatissima	alien
Hypoponera sauteri	native
Leptogenys confucii	native
Lioponera daikoku	native
Monomorium chinense	native
Monomorium floricola	uncertain
Monomorium hiten	native
Monomorium intrudens	native
Monomorium pharaonis	alien
Myrmecina ryukyuensis	native
Nylanderia amia	uncertain
Nylanderia OK02	uncertain
Nylanderia OK03	uncertain
Nylanderia ryukyuensis	native
Nylanderia yambaru	native
Ochetellus glaber	native

Odontomachus kuroiwae	native
Ooceraea biroi	uncertain
Paratrechina longicornis	alien
Pheidole fervens	uncertain
Pheidole megacephala	alien
Pheidole noda	native
Pheidole parva	uncertain
Pheidole pieli	native
Plagiolepis alluaudi	alien
Polyrhachis dives	native
Polyrhachis moesta	native
Ponera takaminei	native
Ponera tamon	native
Pristomyrmex punctatus	native
Proceratium japonicum	native
Protanilla lini	native
Rhopalomastix OK01	native
Solenopsis tipuna	native
Stigmatomma sakaii	native
Stigmatomma silvestrii	native
Strumigenys circothrix	native
Strumigenys emmae	uncertain
Strumigenys exilirhina	native
Strumigenys hexamera	native
Strumigenys hirashimai	native
Strumigenys lewisi	native

Strumigenys mazu	native
Strumigenys membranifera	uncertain
Strumigenys minutula	native
Strumigenys OK01	uncertain
Strumigenys strigatella	native
Tapinoma melanocephalum	uncertain
Tapinoma saohime	native
Technomyrmex brunneus	alien
Temnothorax indra	native
Temnothorax OK01	native
Tetramorium bicarinatum	alien
Tetramorium kraepelini	uncertain
Tetramorium lanuginosum	alien
Tetramorium nipponense	native
Tetramorium simillimum	alien
Tetramorium smithi	alien
Trichomyrmex destructor	uncertain
Vollenhovia OK01	native

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Table S3. Results from linear regression models built with two land-cover PCA axes: PC1, explaining the forested-developed gradient, and PC2, explaining the rural-urban gradient. All model residuals were tested for spatial autocorrelation by calculating Moran's I and testing with random permutations (*moran.randtest*() from the R package *adespatial*; 999 repetitions), yet none had significant results.

		PC1			PC2			Adj. R^2	Moran *p*
Response variable	Intercept	Estimate	Std. Error	*p*	Estimate	Std. Error	*p*		
Total activity (natural log)	10.082	-1.575	0.206	0	NA	NA	NA	0.714	0.269
Total richness (rarefied)	22.742	3.453	1.557	0.037	NA	NA	NA	0.145	0.727
Native richness (rarefied)	13.662	8.351	1.367	0	NA	NA	NA	0.612	0.789
Uncertain richness (rarefied)	5.863	-3.585	1.019	0.002	NA	NA	NA	0.331	0.179
Alien richness (rarefied)	3.217	-1.314	0.509	0.017	NA	NA	NA	0.198	0.59
Total richness (observed)	30.708	NA	NA	NA	NA	NA	NA	0	0.887
Native richness (observed)	18.708	7.483	1.686	0	-7.594	4.611	0.114	0.47	0.8
Uncertain richness (observed)	7.875	-5.188	1.516	0.002	NA	NA	NA	0.318	0.312
Alien richness (observed)	4.125	-2.052	0.643	0.004	NA	NA	NA	0.285	0.271
Total richness (estimated)	38.923	NA	NA	NA	NA	NA	NA	0	0.862
Native richness (estimated)	23.781	13.457	3.646	0.001	NA	NA	NA	0.354	0.767
Uncertain richness (estimated)	8.427	-5.717	1.636	0.002	NA	NA	NA	0.328	0.487
Alien richness (estimated)	4.292	-2.092	0.802	0.016	NA	NA	NA	0.202	0.3

Functional temporal variability	0.67	0.405	0.059	0	NA	NA	NA	0.667	0.269
Compositional temporal variability	0.23	0.141	0.024	0	NA	NA	NA	0.601	0.703
Native SCBD	0.53	0.237	0.059	0.001	NA	NA	NA	0.397	0.776
Uncertain SCBD	0.229	-0.16	0.05	0.004	NA	NA	NA	0.283	0.524
Alien SCBD	0.257	-0.087	0.05	0.094	NA	NA	NA	0.082	0.465
Seasonal variability (absolute)	1209.034	1850.927	328.626	0	NA	NA	NA	0.572	0.918
Seasonal variability (relative)	0.478	0.201	0.081	0.022	-0.373	0.222	0.108	0.232	0.39
Trend variability (absolute)	194.354	311.24	90.185	0.002	NA	NA	NA	0.322	0.567
Trend variability (relative)	0.079	NA	NA	NA	NA	NA	NA	0	0.722
Stochastic variability (absolute)	946.566	660.74	297.568	0.037	NA	NA	NA	0.146	0.591
Stochastic variability (relative)	0.443	-0.223	0.066	0.003	0.345	0.18	0.069	0.363	0.403

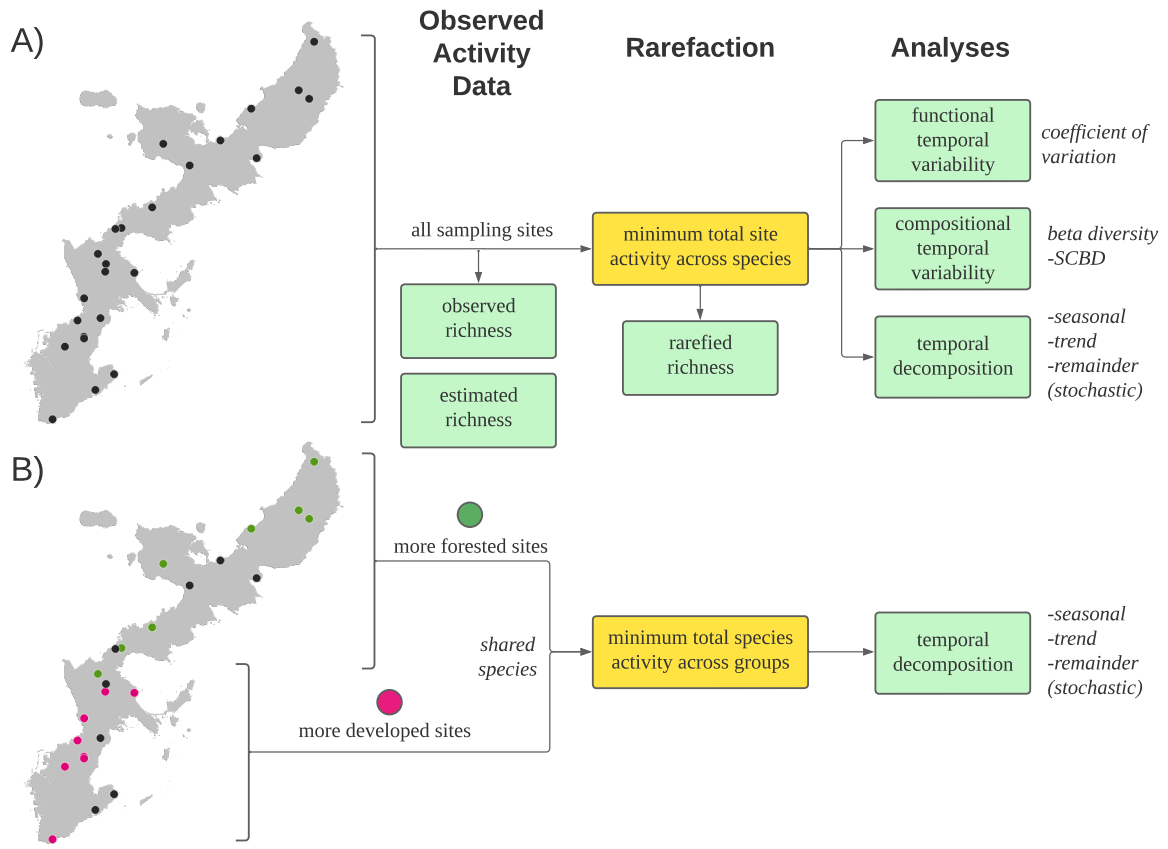

Figure S1. Conceptual workflow schematic describing the analysis steps. A) The main analysis uses ant activity data from all sites. Richness is calculated for observed data and estimated via Hill numbers, then the dataset is rarefied by randomly resampling individual workers per site to match the minimum total site activity; rarefied richness is calculated from this rarefied dataset. Next, the following analyses are conducted: functional temporal variability, compositional temporal variability (including species contribution to beta diversity [SCBD]), and temporal decompositions to derive seasonal, trend, and remainder (stochastic) components. B) The group-level analysis first splits sites into two groups (green and pink points) based on land-cover PC1, explaining the forested-developed gradient. Both groups share the same species, and those with mid-level PC1 values are excluded (black points). These grouped data are then rarefied by randomly resampling individual workers per species within each group to match the minimum total species activity between groups. The same temporal decomposition is then conducted on these grouped data.

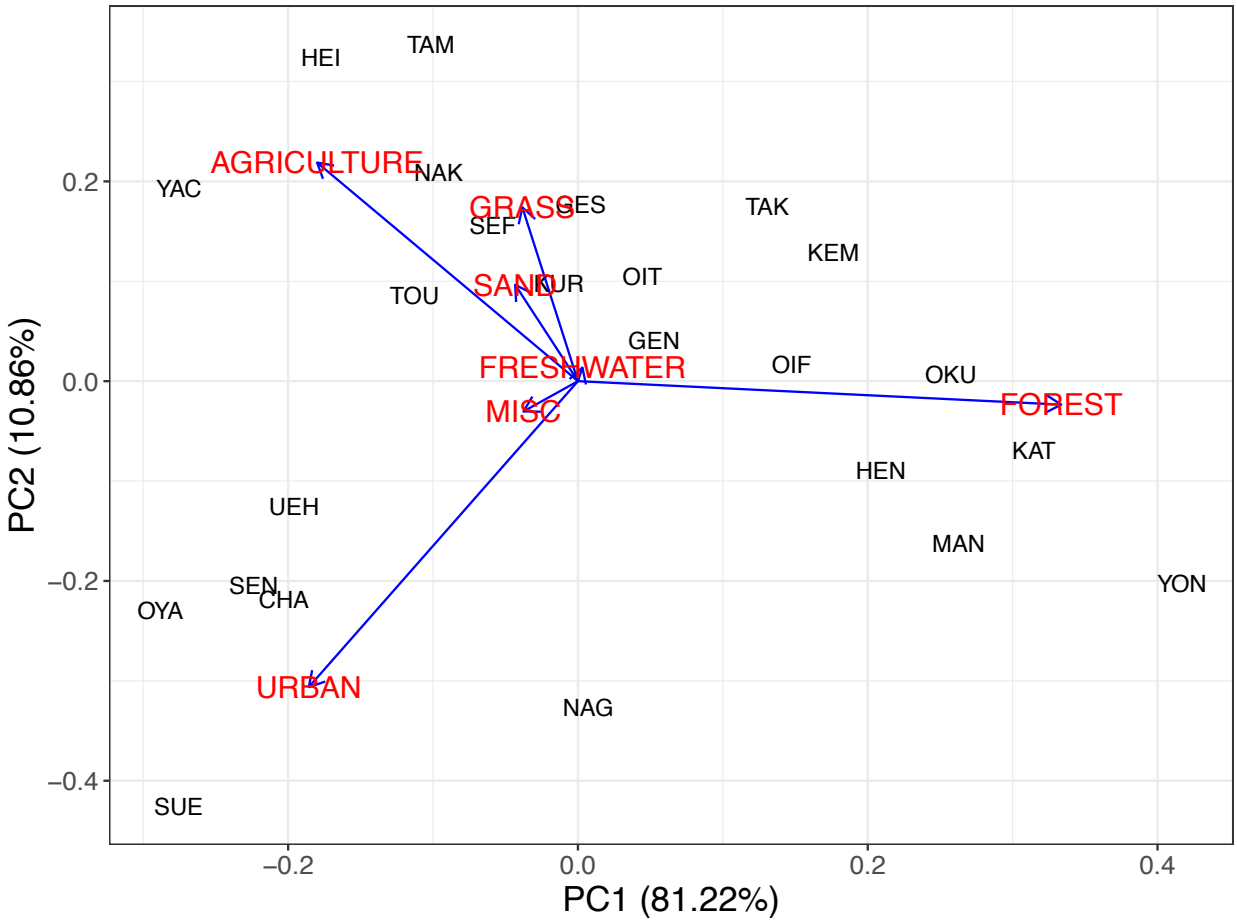

Figure S2. Biplot of the principal component analysis for the land-cover proportion dataset. PC1 explains the main anthropogenic stress gradient (low: more forested, high: more developed), while PC2 explains the rural (high; agriculture and/or grass) to urban (low) gradient.

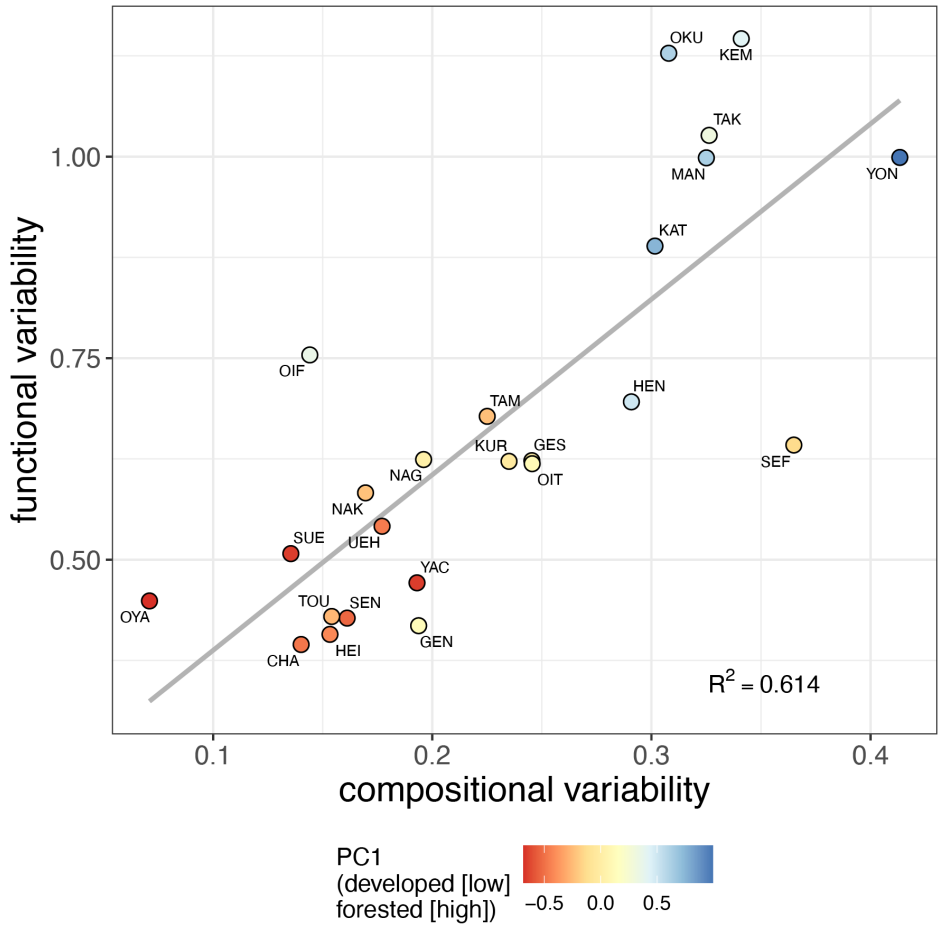

Figure S3. Plot of compositional variability by functional variability, showing the sampling sites with colors corresponding to the forested-developed gradient (PC1), where higher values (blue) are more forested and lower values (red) have more human development. Site abbreviations correspond to those used in Figure 1. The gray line shows the relationship between two variabilities modeled with linear regression.

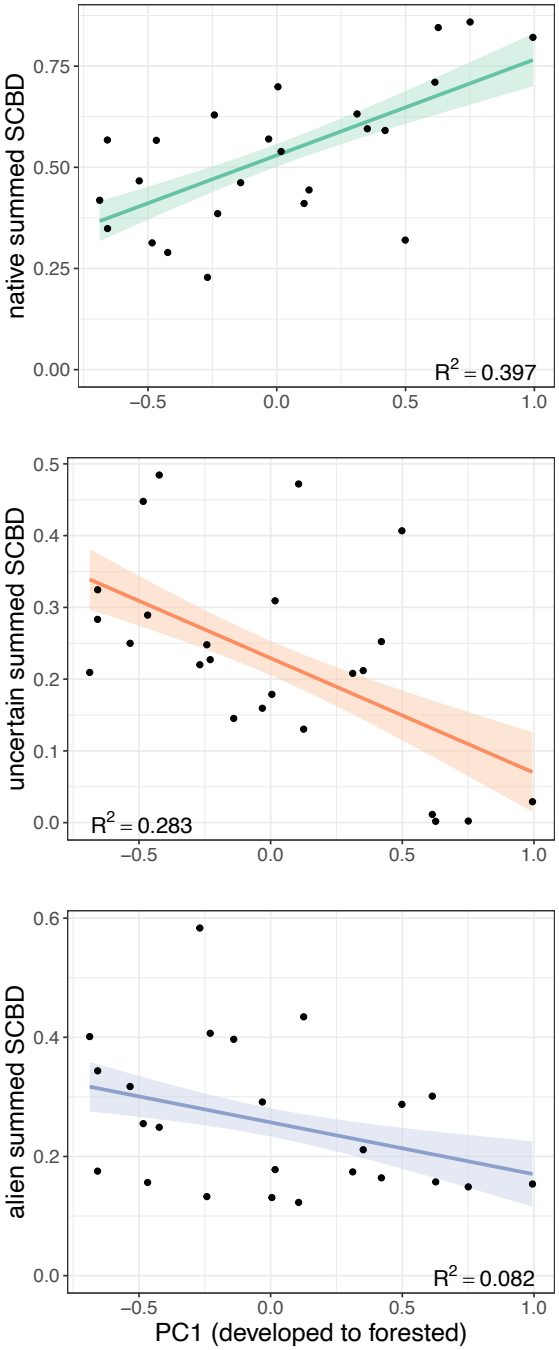

Figure S4. Relationships between PC1 and summed species contributions to beta diversity (SCBD) for native, uncertain, and alien species for the rarefied activity data. PC1 explains the anthropogenic stress gradient (low: more developed, high: more forested).

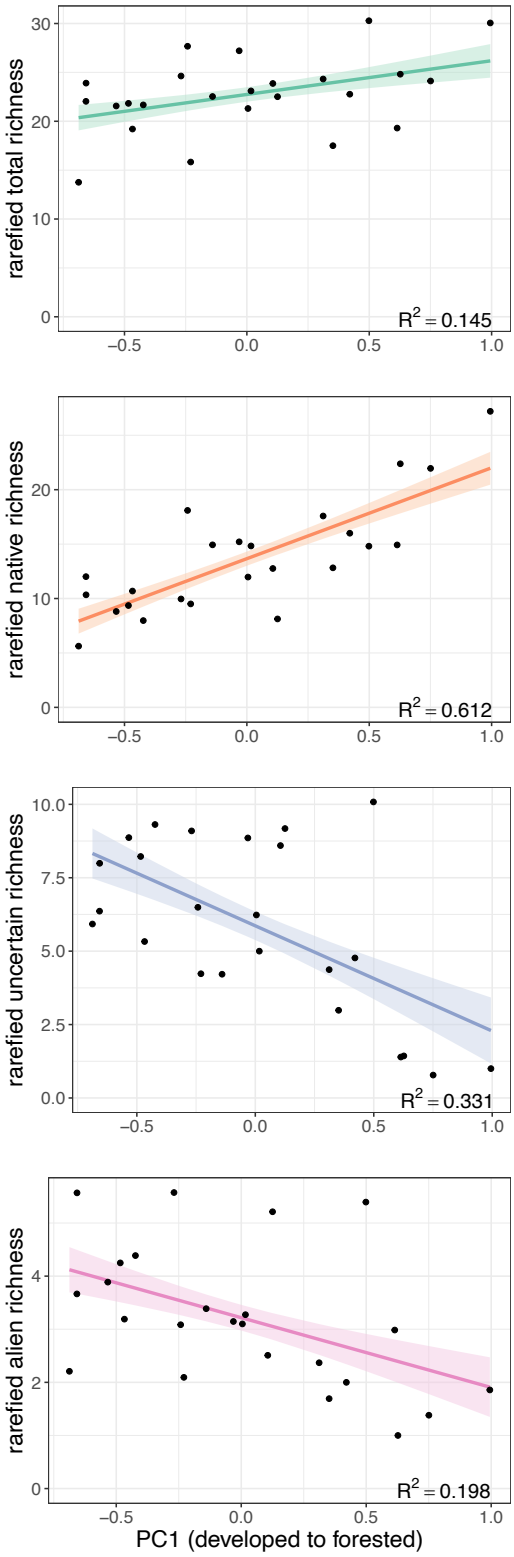

Figure S5. Relationships between PC1 and total richness, as well as richness of native, uncertain, and alien species for the rarefied activity data. PC1 explains the anthropogenic stress gradient (low: more developed, high: more forested).

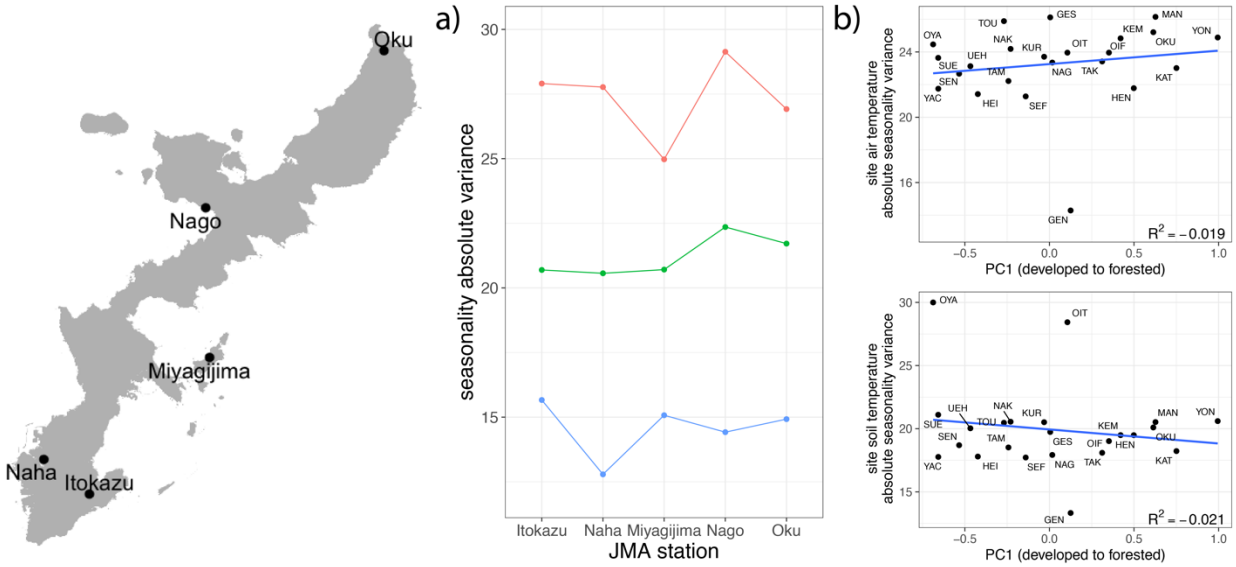

Figure S6. Absolute seasonality variance of temperature at the a) regional and b) site-level scales over the two-year study period (2016 – 2018). Regional temperature measurements were taken from Japanese Meteorological Association weather stations, and station names correspond to those on the map. Site-level temperature measurements were taken for air and soil at one station per sample site. Oyama Park (OYA) and OIST Open (OIT) were the only sites with sensors not located below forest canopy, and thus have higher variance for soil temperature seasonality.

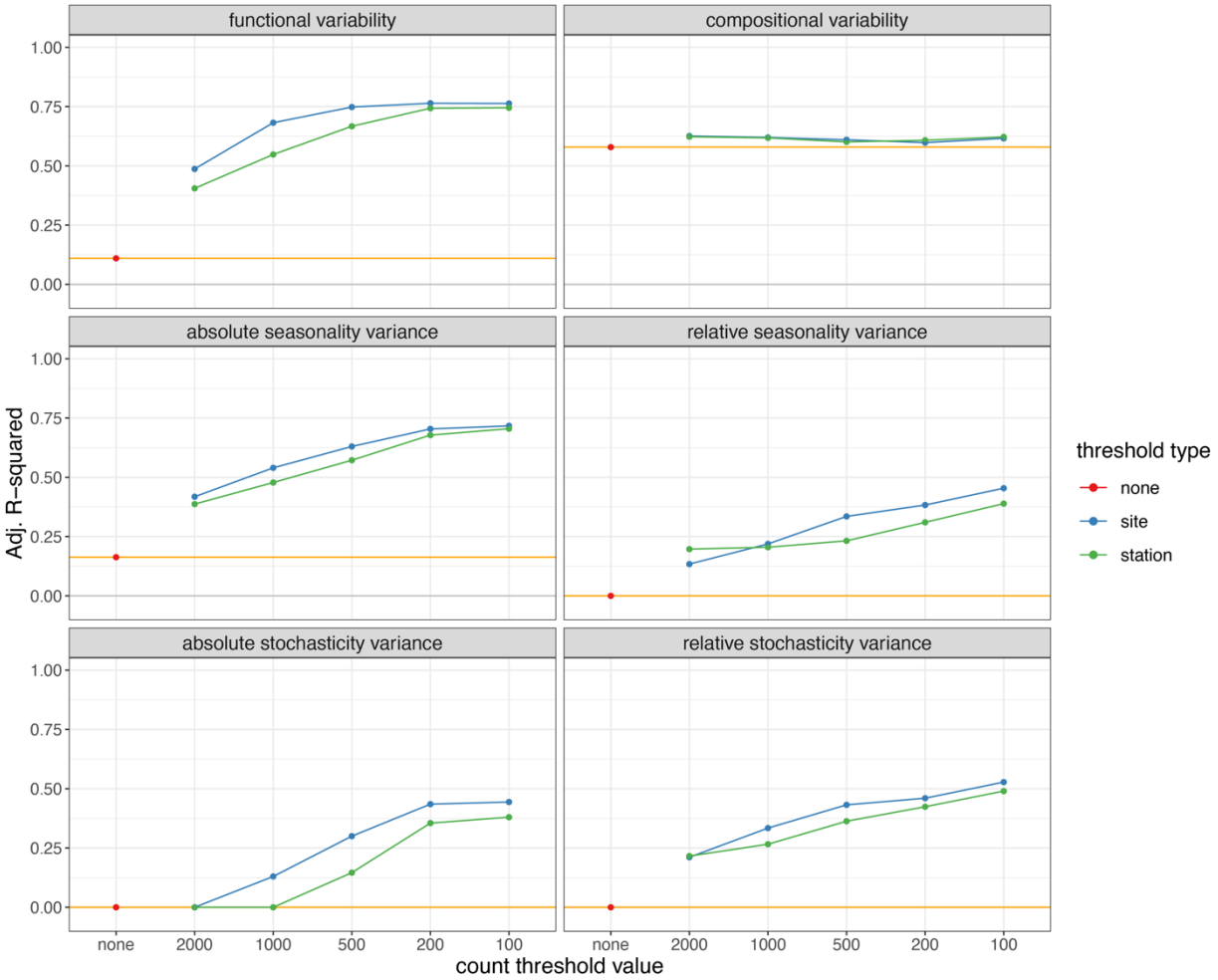

Figure S7. The adjusted R-squared value for the metrics measured in this study at different count thresholds and spatial levels for thresholding (station, site). The threshold value "none" is used as a reference and is symbolized by the red point and orange line. For reference, the analysis described in this paper is a threshold value of 500 at the station level.