1   **Article Type: Letter**

3   **Experimental estimates of germline mutation rate in eukaryotes: a phylogenetic meta-**
4   **analysis**

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Short Running Title: mutation rates in eukaryotes

Keywords: *de novo* mutation rate, eukaryotes, indels, phylogeny

Abbreviations: MA, mutation accumulation experiment; PO, parent-offspring experiment;
SNMs, single nucleotide mutations; indels, insertions and deletion mutations; PGLMM,
phylogenetic generalised linear mixed model; Mb, Mega base pair(s); HPD, highest posterior
density.

Abstract Word Count: 280

Total Word Count: ~5200

Abstract

Mutation is the ultimate source of all genetic variation, and over the last ten years the ready

availability of whole-genome sequencing has permitted direct estimation of mutation rate

for many non-model species across the tree of life. In this meta-analysis we make a

comprehensive search of the literature for mutation rate estimates in eukaryotes,

identifying 140 mutation accumulation (MA) and parent-offspring (PO) sequencing studies

covering 134 species. Based on these data, we revisit differences in single nucleotide

mutation (SNM) rate between different phylogenetic lineages and update the known

relationships between mutation rate and generation time, genome size, and nucleotide

diversity—while accounting for phylogenetic non-independence. We do not find a significant

difference between MA and PO in estimated mutation rates, but we confirm that mammal

and plant lineages have higher mutation rates than arthropods, and that unicellular

eukaryotes have the lowest mutation rates. We find that mutation rates are higher in

species with longer generation times and larger genome sizes, even when accounting for

phylogenetic relationships. Moreover, although nucleotide diversity is positively correlated

with mutation rate, the gradient of the relationship is significantly less than one (on a

logarithmic scale), consistent with higher mutation rates in populations with smaller

effective size. For the 29 species for which data are available, we find that indel mutation

rates are positively correlated with nucleotide mutation rates, and that short deletions are

generally more common than short insertions. Nevertheless, despite recent progress, no

estimates of either SNM or indel mutation rates are available for the majority of deeply-

branching eukaryotic lineages—or even for most animal phyla. Even among charismatic

megafauna, experimental mutation rate estimates remain unknown for amphibia and scarce

for reptiles and fish.

Lay Summary

Over the past decade, the sequencing revolution has led to an ever-increasing number of

mutation-rate estimates from mutation accumulation or parent-offspring sequencing studies

in eukaryotes. However, studies rarely quantify to what extent the mutation rate varies

among these species. Also, despite strong predictions as to how mutation rate should vary

with (e.g.) generation time, there have been few recent or wide-ranging analyses of such

predictors while accounting for the inherent similarity between closely-related species. Of

particular note, there has been surprisingly little effort to robustly test the 'drift barrier'

hypothesis that mutation rates should decrease with increasing effective population size. In

this study, we used a comprehensive literature search to identify all the available

experimental estimates of mutation rate in eukaryotes and subject them to phylogenetic

mixed-model analyses. We find that per-nucleotide per-generation mutation rates differ by

orders of magnitude among species: plants and mammals tend to have higher mutation

rates than arthropods, and unicellular organisms have the lowest mutation rates. Our

analysis also shows that mutation rates increase significantly with increasing generation time

and genome size, and nucleotide diversity increases with mutation rate with a gradient less

than one—as predicted by the drift-barrier hypothesis.

Introduction

The per-nucleotide per-generation *de novo* mutation rate, μ, is a key parameter in

population and evolutionary genetics, appearing either alone or combined with effective

population size (N_e) in the compound parameter θ. However, it seems implicit that μ is of no

interest to many evolutionary studies; researchers often use μ from one species when

analysing another—as if μ were exchangeable in a way that N_e is not (e.g. Wilding, 2017).

This is partly because the historically high cost of sequencing made estimates prohibitively

expensive, but may also be because variation in μ among species is widely thought to be

negligible compared to variation in N_e. But, if the mutation rate varied as widely among

multicellular eukaryotes as many other traits, there would be much less need to invoke

variation in N_e when explaining variation in (e.g.) genetic diversity. One reason why μ may

vary relatively little is if it is minimised to a lower limit imposed by drift: the 'drift barrier'

hypothesis (Lynch, 2008, 2010). However, it might be that in reality mutation rate is

optimised to well above the drift-barrier limit by selection in favour of adaptability or the

cost of replication fidelity (Kimura, 1967; Liu et al., 2021; Peck et al., 2018). If this were the

case, then we might find that mutation rates varied dramatically, even among closely related

taxa.

Early estimates of the mutation rate leveraged visible phenotypes, such as the occurrence of

genetic diseases in humans (Nute et al., 1984; Trimble et al., 1974). For example, using

haemoglobin M disease Stamatoyannopoulos et al. (1982) estimated the mutation rate in

the human haemoglobin beta gene at 7.4×10^{-9} per bp per generation. However, such

estimates are prone to bias, for example if the mutation is required to be autosomal

dominant and non-lethal (Nute et al., 1984), and such approaches are not easily applicable in

organisms for which disease is hard to observe. More commonly, using the equivalence

between mutation and neutral substitution rates (Kimura, 1968), many studies have used

time-calibrated sequence divergence to estimate μ. For example, comparison between

humans and chimpanzees gives rise to estimates on the order of 10^{-8} per bp per generation

(Drake et al., 1998; Kondrashov et al., 1993; Nachman et al., 2000). However, such indirect

estimates are also limited (Kondrashov, 2003). First, because generation time and calibration

dates are prone to substantial uncertainty (e.g. Obbard et al., 2012). Second, because it is

hard to distinguish between unconstrained and weakly constrained sites (Harmon et al.,

2021). Third, because mutation saturation in hotspots may lead to an underestimation,

especially for sequences with a long divergence time (Sigurðardóttir et al., 2000). In addition,

phylogenetic-calibration approaches reflect the long-term mutation rate, averaging across

biologically important factors such as generation time, sex differences, life-stage differences,

and inter-individual variation.

In contrast to phenotypic and phylogenetic approaches, mutation-accumulation (MA) and

parent-offspring (PO) sequencing can provide direct estimates of the per-generation

mutation rate. Mutation-accumulation utilises the accumulation of spontaneous mutations

in a single inbred or asexual genome over multiple generations, by minimising the

effectiveness of natural selection (Halligan et al., 2009). Given the low mutation rates in

eukaryotes, the multiple generations in an MA experiment allow sufficient mutations for

them to be detected at the end. Although the first use of MA to estimate μ phenotypically

dates back to a study by Mukai et al. (1977), the availability of whole genome sequencing

and its decreasing price have made sequencing approaches more widely applicable.

Mutation accumulation best suits easily-maintained species with short generation times, for

example nematodes (Denver et al., 2004; Konrad et al., 2019; Meier et al., 2014; Saxena et

al., 2019), fruit flies (Haag-Liautard et al., 2007; Huang et al., 2016; Keightley et al., 2009;

Schrider et al., 2013), green algae (Lopez-Cortegano et al., 2021; Ness et al., 2012), and

water fleas (Ho et al., 2020; Keith et al., 2021). However, MA may underestimate mutation

rates (Lynch et al., 1998) because natural selection will still come into play for mutations that

are very highly deleterious, such as those that cause complete sterility or are lethal (Baer et

al., 2007).

By sequencing parents and offspring, and counting the mutations arising over a single

generation, the parent-offspring approach avoids purging deleterious mutations, except

those that are dominant lethal, and so provides the most direct estimate of mutation rate. It

is not limited to model species (Keightley et al., 2014; Yang et al., 2015), and is suitable for

large animals outside of the laboratory (Harland et al., 2017; Wang et al., 2022). Parent-

offspring sequencing has been particularly widely used in humans, where MA is not

applicable. However, the challenge of robustly identifying new mutations in natural

heterozygous genomes is formidable, as the number of mutations arising within one

generation is small when compared to the impact of sequencing/mapping errors and the

genetic variation of heterozygotes. For example, Bergeron et al. (2022) compared the

estimates from five different research groups for the same family of rhesus macaques and

found that the highest estimates could be twice the lowest ones. Minimising the rates of

false positive and false negative mutation-calls is therefore critical for the PO approach

(Yoder et al., 2021).

Relaxed-clock phylogenetic studies generally indicate there is substantial among-lineage

variation in the evolutionary rate, suggestive of variation in mutation rates among taxa (Ho

et al., 2015). But such studies cannot separate the impact of generation time and mutation

rate, and may be biased by variation in the action of (weak) selection (Harmon et al., 2021).

As more direct experimental estimates have become available, it has become possible to

draw comparisons within particular clades such as primates (Chintalapati et al., 2020), or

vertebrates (Bergeron et al., 2022; Bergeron et al., 2023; Yoder et al., 2021). However, there

has not been a wider analysis, nor one that includes both PO and MA studies.

In addition, many previous studies have not taken an explicitly phylogenetic analytical

approach (e.g. Bergeron et al., 2022; Katju et al., 2019; Yoder et al., 2021). This could in

principle lead to pseudoreplication in the analyses that is caused by the non-independence

among taxa that arises from their shared ancestry (reviewed in Freckleton, 2009). However,

this problem can be mitigated by explicitly modelling covariance among species that is

contributed by their patterns of relationship, and generalised linear mixed modelling

approaches are available to do this (Hadfield et al., 2010; Halliwell et al., 2022; Lynch, 1991).

Here we perform a comprehensive phylogenetic meta-analysis of published eukaryotic single

nucleotide mutation (SNM) and insertion-deletion (indel) mutation-rate estimates that use

direct sequencing of multiple generations, either through MA or PO comparison. We focus

on the variation in mutation rate among species, and the relationship between mutation

rate estimates and the experimental method, generation time, genome size, and genetic

diversity.

Results and Discussion

Studies included

We searched for eukaryotic mutation-rate studies in Clarivate Web of Science published prior to 21st September 2022, and then we manually examined these studies and identified those containing mutation rate estimates (see Methods). For the majority of our analyses we also included data from one further recent study of 68 vertebrate species by Bergeron et al. (2023), published after our initial submission. In total, we included 140 studies using either MA or PO approaches, covering 134 species (see Availability of data and materials). Studies on primate mutation rate accounted for over a third of this literature, and 82% (45/55) of the primate studies were on humans. Arthropods accounted for 17% of the identified literature, followed by fungi (11%), plants (8%), non-primate mammals (7%), and nematodes (5%) (**Fig. 1A**). We only identified four studies on fish, two studies on birds and one study on reptiles. Across all groups, most studies focused on model species: yeast in fungi, *Drosophila* in arthropods, mice in mammals, and *Arabidopsis* in plants. These estimates tended to derive from MA analyses rather than PO studies (61 vs. 43 among non-human studies). However, as sequencing has become cheaper, there has been a rapid increase in the number of published experimental estimates of the mutation rate, with an average of 11 studies each year in eukaryotic species between 2010 and 2022 (**Fig. 1B**).

Single nucleotide mutation rates across eukaryotes

To summarise the single nucleotide mutation rates in different lineages, we obtained a time-

scaled phylogenetic tree of all represented species from http://www.timetree.org/ (Kumar

et al., 2022), and used this to fit the covariance among species in a Phylogenetic Generalised

Linear Mixed Model (PGLMM) using the Bayesian mixed model package MCMCglmm

(Hadfield et al., 2010). Treating species as a random effect, this provides a posterior estimate

of mutation rate for each species, given the observations and relationships to other species

(see Methods). In these data, the phylogenetic effect accounts for 96.4% of the random-

effects variance, with a 95% HPD interval of 94.9% to 97.9%.

Among all the studied species, a ciliate (*Tetrahymena thermophila*) had the lowest mutation

rate, estimated at 0.01×10^{-9} per generation per base pair, while a fungus (*Marasmius*

oreades) had the highest rate, estimated at 55.58×10^{-9}—a difference of over five thousand-

fold. The estimates for unicellular species were generally more than an order of magnitude

lower than rates in multicellular species (below; **Fig. 2**), even compared to species in the

same clade, e.g., fungi. This may either be attributable to there being a single cell division

per generation, or possibly to a limit set by the drift barrier in species with larger N_e (Sung et

al., 2012). We also estimated relatively low SNM rates for nematodes at 2.44×10^{-9},

arthropods at 2.82×10^{-9}, and fish at 5.55×10^{-9}. For plants, the clade root was predicted as

3.66×10^{-9}, but a large difference was observed between duckweed and other plant species

(0.22×10^{-9} vs. 19.8×10^{-9}). The high mutation rates in most plant species might be

attributable to the lack of a segregated germline (Burian, 2021; Cruzan et al., 2022; Lanfear,

2018; Wang et al., 2019). Mammals had moderate mutation rates with a root estimate at

8.36×10^{-9} and species (excluding primates) ranging from 5.64×10^{-9} in pigs to 11.68×10^{-9} in

killer whales. Primates had higher mutation rates, with a root estimate at 9.37×10^{-9}. Of

note, as the most studied species, human mutation rate was estimated as 13.41×10^{-9} with

95% CI 11.61 to 15.28×10^{-9}, which is also the highest among mammals. Similar to mammals,

birds and reptiles had high mutation rate, with root estimates at 10.43×10^{-9} and 12.41×10^{-9}

^{9}, respectively.

To assess a possible impact of the experimental method on rate estimates, we included the

experimental approach (MA or PO) as a fixed effect in the PGLMM. Overall, we did not find a

significant difference between the two methods (MCMC P = 0.155). We then examined

comparisons between MA and PO in fruit flies and mice more closely, as these two species

have been investigated using both methods. The rate estimated using MA in fruit flies was

5.28×10^{-9}, but 3.20×10^{-9} using PO (**Fig. S1**), and in mice the MA estimate was 5.40×10^{-9}

while the PO estimate was 4.05×10^{-9} (**Fig. S2**), and these differences were similarly not

significant. Note, however, that the mutation rate may be heterogeneous among different

populations: for *D. melanogaster*, the estimated mutation rate from African populations was

lower than that from European population and North American population (Chan et al., 2012;

Wang et al., 2023). To make a more robust analysis would require samples of the same

population but using different methods (MA or PO), which are not yet available.

Biological predictors of the SNM rate

Many biological factors have been proposed to be associated with mutation rate, including

generation time (Bailey et al., 1991; Keightley et al., 2000; Li et al., 1987; Wu et al., 1985)

and genome size (Drake, 1991; Lynch, 2010). To test for a dependence of mutation rate on

these factors, we ran three further models: the first two including these factors each

separately as an additional fixed effect in a PGLMM, and the last one including both of them

simultaneously.

We found a significant positive relationship with generation time (Bayesian PGLMM MCMC P

< 0.0001; **Fig. 3A**), such that a longer generation time predicted a higher per-site mutation

rate per generation. If mutations resulted only from replication errors, and most species had

a similar number of cell divisions per generation, then the per-generation mutation rate

would not scale with generation time (Thomas et al., 2014; Wu et al., 1985). However,

although few estimates are available, in reality the number of germline cell divisions per

generation varies within and between species. In humans, there are around 401 cell

divisions per 30-year generation in males and 31 in females (Drost et al., 1995; Ohno, 2019),

but in mice it is 62 per 9-month generation in males and 25 in females. In *Drosophila* the

sexes are much more similar, with an estimated 35.5 cell divisions in a 18-day generation for

males but 36.5 in an 25-day generation in females in *D. melanogaster* (Drost et al., 1995).

Within a species, a long generation time is likely to permit more cell divisions and thus a

higher mutation rate per generation, and many studies in humans and primates have shown

that mutations are heavily male-biased and correlated with paternal ages (Kaplanis et al.,

2022; Kong et al., 2012; Thomas et al., 2018; Wang et al., 2020).

Alternatively, mutations may not result predominantly from replication errors. Recent

studies suggest that not all germline mutations track cell divisions, and DNA damage may

contribute to the mutation rate (Wu et al., 2022; Wu et al., 2020). Independent of cell

divisions, this mutation rate would be correlated with calendar time, and thus species with a

longer generation time would have a higher rate. The relationship we observe between

generation time and per-generation mutation rate could therefore be a consequence of

either a greater number of cell divisions or of accumulating damage over time, and it will be

hard to separate these potential causes given the limited data available on the number of

germline cell divisions for different species (Gao et al., 2016; Wu et al., 2022).

We also found evidence that larger genome size predicts a higher per-generation mutation

rate (MCMC P = 0.0020; **Fig. 3B**). Mutation rate has previously been found to correlate

negatively with genome size in prokaryotes, a phenomenon termed "Drake's rule"

(Bourguignon et al., 2020; Drake, 1991; Marais et al., 2020), but positively in eukaryotes

(Lynch, 2010).

In each of the two models above, we fitted generation time and genome size individually as

predictors of mutation rate in a regression model. However, these factors might be highly

correlated: large organisms tend to have longer generation times and often larger genomes

(Buffalo, 2021; Martin et al., 1993), and is therefore hard to distinguish between casual

relationships and confounding correlations. Including generation time and genome size

together as fixed terms in a single model resulted in only generation time being significant

with MCMC P < 0.001 (versus MCMC P = 0.301 for genome size).

The drift-barrier hypothesis

The most direct test of the drift-barrier hypothesis would be a regression of SNM rate on N_e.

However, long-term N_e can only be estimated from genetic diversity (i.e. π) by making use of

the relationship $\theta_\pi = 4N_e\mu$, and any uncertainty in μ could induce a spurious correlation in a

regression of μ on $\pi/4\mu$. Instead, error in both π and μ can be accounted for by fitting a

bivariate linear mixed model, parameterised in terms of a regression of π on μ (see

Methods). Then, on a logarithmic scale, the gradient of this relationship would be expected

to be 1 if μ and N_e are independent of each other, and lower than 1 if species with large N_e

tended to have lower μ. Nevertheless, such a relationship could be driven by other factors.

For example, if species with small N_e tend to have a longer generation time, and a longer

generation time causes higher mutation rates (i.e. **Fig. 3A**), then a higher μ in species with

low N_e could be driven by a mechanistic generation-time effect, rather than the drift barrier.

We attempted to account for this by including generation time in our analysis as a predictor

of both μ and π, thereby accounting for the direct impact of generation time on the

relationship between them. In this bivariate phylogenetic model, μ increased with

generation time (MCMC $P < 0.001$) whereas π did not (MCMC $P = 0.501$), and the gradient of

the regression of π on μ was significantly less than 1 (gradient = 0.535, 95% CI [0.259, 0.832];

MCMC $P = 0.004$). This suggests that populations with larger N_e tend to have a lower

mutation rate even after accounting for their shorter generation times—as predicted by the

drift-barrier hypothesis (Lynch, 2010). Note that excluding generation time as a predictor

resulted in an even stronger effect (gradient = 0.300 [0.098, 0.486]; MCMC $P < 0.001$; **Fig.**

3C), whereas removing unicellular species from the analysis (because of their low mutation

rates) resulted in a gradient that was still much less than 1 (gradient = 0.439 [0.155, 0.746];

MCMC $P < 0.001$).

The mutation rates of short indels across eukaryotes

Although the rate of indel mutation is much less commonly estimated than the SNM rate,

our literature search identified estimates for 29 species. We found that the indel mutation

rate across eukaryotes displayed a similar pattern to that of the SNM rate (**Fig. 4**), with a

correlation coefficient between the two of 0.75 (95% HPD CI: 0.44, 0.98; MCMC *P* < 0.001)

using PGLMM, with the SNM rate being generally higher than the indel mutation rate,

except in two lineages (nematodes and amoebae) for which the indel mutation rate was

apparently higher (**Fig. 5A**). However, the mutations in the nematode study (Denver et al.,

2004) were specifically collected from mononucleotide repeats, which may artificially inflate

the indel mutation rate estimate in that study (Konrad et al., 2019; Rajaei et al., 2021). The

unicellular species had the lowest indel mutation rate, from 0.02×10^{-9} in yeast to 0.26×10^{-9}

in algae, followed by arthropods with $0.55 (0.08, 1.24) \times 10^{-9}$. Mammals had a relatively

higher indel mutation rate with $1.13 (0.43, 2.05) \times 10^{-9}$, with humans at the higher end of

this range at $1.74 (1.16, 2.50) \times 10^{-9}$. Of all the studied species, plants had the highest indel

mutation rate, which was as high as $2.38 (0.51, 4.96) \times 10^{-9}$.

Short deletions generally outnumbered short insertions across eukaryotes (**Fig. 5B, Fig.S3**).

The ratio of deletions over insertions was as high as 15.0 in the water flea, 6.0 in mouse, 3.4

in *D. melanogaster*, 2.6 in humans, 1.8 in *C. elegans*, and 1.6 in *Arabidopsis*. Honeybees,

diatoms and two plants (rice and peach) exhibited more insertions than deletions, but the

difference wasn't statistically significant (*P* > 0.05, binomial test). Fission yeast showed

significantly more insertions overall.

A bias toward deletion events has been known for at least 40 years, since de Jong et al.

(1981) found a four-fold excess of deletions over insertions in protein sequences. Petrov

(2002) argued that the deletion bias was likely due to a thermodynamic disparity between

short deletions and insertions that makes deletion events easier. However, studies in plants

have suggested that unequal homologous recombination and illegitimate recombination

may be another important cause of small deletions (Bennetzen et al., 2005). Whatever the

cause, the dominance of deletions over insertions seems convincing, and many early studies

argued that these short deletions resulted in genome contraction (Gregory, 2003; Petrov,

2002; Thomas et al., 2003). However, the role of deletions in shaping genome size remains

controversial (Gregory, 2003, 2004). A recent study in birds and mammals indicated that

short deletions alone cannot explain the DNA content loss (Kapusta et al., 2017), and we fail

to find any evidence of a relationship between deletion rate and genome size here (**Fig. S4;**

but note that power is likely to be low). It seems probable that other structural mutations,

such as large deletions, gene duplication and transposable elements, may have greater

impacts on genome size. For example, in *Drosophila* Wang et al. (2023) found transposable

element insertions to be more frequent than SNMs, and an order of magnitude more

frequent than short indels. Together, these studies suggest that the impact of short

deletions on genome size is likely to be small, despite their higher frequency relative to short

insertions.

Concluding remarks

All genetic variation arises from *de novo* mutations. To estimate the rate at which these arise,

studies have moved from investigating a few genes that result in phenotypic change, to

whole genome sequencing of MA lines or families. Ten years ago, Campbell et al. (2013)

called for more efforts to sequence genomes from non-human primate families to

understand 'how the rate has changed in different lineages'. And, in the present data-rich

and tool-rich era we have now seen the sequencing of not just multiple non-human primates,

but many taxa representing major clades across the tree of life. Nevertheless, the 'tree of

mutations' is far from complete, at either the large or small scale. Even within animals, the

vast majority of phyla have never been examined, and with the exception of primates, no

large groups of very close relatives have been examined. Even for vertebrates—and despite

of the recent large study by Bergeron et al. (2023) who examined 68 species–data are still

scarce for reptiles and entirely lacking for amphibians.

However, as progress is being made toward a tree of mutation rates, attention still needs to

focus on methodological heterogeneity. Different data filtering strategies may result in non-

negligible differences in mutation rate estimates (Bergeron et al., 2022). Filters that are too

conservative may discard some true mutations, while relaxed filters may inflate the false

positive rate. Although many strategies have been proposed (Bergeron et al., 2022), no

standardised pipeline has yet been agreed upon or widely adopted. All studies must

therefore carefully address the issues of false positives by manual curation and/or PCR

follow-up, and false negatives by simulation. Here, we did not examine the impact of the

false positive rate or false negative rate, but we remind the readers of the potential impact

from methodological difference, and caution that attention should be paid to the methods

or parameters used in each study when comparisons are made.

Another potential source of bias may also come from the specific choice of samples used in

each study. Beside the biological factors that affect mutation rates (e.g. age, sex),

environmental factors can also influence mutation rates, for example temperature (Belfield

et al., 2021; Waldvogel et al., 2021), osmotic stress (Hasan et al., 2022), or ultraviolet light

(Xu et al., 2019). Samples collected from the wild rather than reared in laboratories may

better reflect the 'real' mutation rate, but the variation that this introduces may also make it

harder to draw broad conclusions. Despite these challenges, the data to come are exciting.

In particular, as sequencing technology advances, we expect improved accuracy in long-read

sequencing facilitating the phasing of mutations and permitting the easy detection of

complex mutations and larger-scale structural changes.

Methods

We made an exhaustive literature search for publications on the experimental estimation of

mutation rate. We first searched Clarivate Web of Science up to the 21st September 2022 for

'Title/Keywords/Abstract' fields containing ("mutation rate" | "mutation rates" |

"mutational rate" | "mutational rates"). As much of the resulting literature was related to

somatic mutations (e.g., cancers) or virus mutations, we excluded '("tumor" | "cancer" |

"clinical" | "virus")' in the search field of 'Title/Keywords/Abstract'. This search strategy

resulted in 9462 studies with the earliest one dated 1928. We also searched for references

to mutation rate estimates in other papers not captured by our search. As bacteria are

outside of our study, we manually removed studies related to bacterial mutations. These

filters resulted in 174 studies. Of these studies, we further excluded 20 studies that used

phylogenetic approaches to infer mutation rates, 14 studies that used phenotypic markers,

and two other studies that used recombination landscape or the site frequency spectrum to

infer mutation rates statistically. In addition, we included one more recent study published

after our initial submission (Bergeron et al., 2023), which estimated the mutation rates for

68 vertebrates using PO sequencing. This resulted in a total of 140 studies in our final

dataset.

From each of the potential studies we extracted the year of the publication, the species, the

population, reproductive mechanism (sexual, asexual), mutation type (SNM, short insertion,

short deletion), the units of mutation rate (per bp per year, per bp per generation), the

mean and 95% CI of mutation rate, the number of identified mutation events, the number of

callable sites, the methods used to infer mutation rate (PO, MA or others), the sample size,

the number of generations (if MA), the sequencing technology, the sequencing depth and

other pertinent observations. Note that for multi-generation studies, such as MA, the

number of 'callable sites' incorporates the possibility that a mutation may have arisen in any

of several generations. If a 95% confidence or credible interval was not provided in the

original study, we calculated a 95% confidence interval using 'binconf()' from the Hmisc R

package (version 4.7-0), based on the number of mutation events and callable sites and

assuming a binomial distribution (Girard et al., 2011). If a study did not provide the mutation

count, we inferred an effective number of mutation events based on the reported mutation

rate and its confidence interval. For studies in which the number of callable sites was not

reported, an estimate was made from the number of mutation events and the reported

mutation rate. If a study *only* reported the mutation rate, we assumed one mutation event

and calculated the corresponding callable sites based on mutation rates. This latter approach

will necessarily underestimate the precision with which rates are estimated (minimising their

weight in the analysis), while keeping the estimate in the meta-analysis with unchanged

mean values.

To compare the SNM rates across different species, we conducted a phylogenetic meta-

analysis (i.e., a combined analysis of published data) using the mutation data from MA or PO

studies. We obtained the phylogenetic relationships and relative divergence dates from

http://www.timetree.org/ (Kumar et al., 2022) and visualised using ggtree (Yu, 2020). We

performed the analyses using a phylogenetic generalized linear mixed model with a unit-

scaled phylogeny as random effect, assuming a Poisson distribution for the occurrence of *de*

novo mutations, off-set by the number of callable sites (i.e. enforcing a directly proportional

relationship between the expected number of mutations and the number of sites observed).

The analysis was conducted using the Bayesian mixed-model R-package MCMCglmm

(version 2.32) (Hadfield, 2010) (see **Supporting Materials**). To investigate the effects of

methods (MA or PO) used to estimate mutation rate, we additionally included 'Method' as

an additional fixed effect in the above model.

We also investigated the relationships between mutation rate and generation time and

genome size using univariate phylogenetic mixed models that included these factors

separately as fixed effects, and a phylogenetic species term as a random effect. A log-

transformation was made on these predictors before fitting the model (see **Supporting**

Materials), and model predictions were made conditioning on each of the different species

as random effect levels. The estimated generation time for each species was gathered from

the literature (see Availability of data and materials) and we obtained estimates for 116

species. Genome sizes were taken to be the genome assembly size and gathered from the

NCBI database, which resulted in 108 species with genome size available.

To test for a relationship between nucleotide diversity (π) and the per generation per-site

SNM rate (μ) while accounting for uncertainty in both, we used MCMCglmm to fit a bivariate

PGLMM with mutation number and π as responses, again including the number of callable

sites as an offset on the number of mutations and generation time as a fixed effect (see

Supporting Materials). We collected nucleotide diversity data for 102 species from Buffalo

(2021), Bergeron et al. (2023) or literature (see Availability of data and materials).

AVAILABILITY OF DATA AND MATERIALS

The datasets and the code to analyse the data are available on GitHub:

https://github.com/Yiguan/mutation_literature

ACKNOWLEDGEMENTS

We thank the authors of the many studies we analyse for making their data available, and

particularly Lucie Bergeron and co-authors who generously shared recently published

nucleotide diversity data with us. We thank Jarrod Hadfield for suggesting the overall

statistical approach, helping with MCMCglmm syntax and interpretation, and drafting much

of the supporting material. This work was funded by the UK Biotechnology and Biological

Sciences Research Council through grant BB/T007516/1. For the purpose of open access, the

authors have applied a Creative Commons Attribution (CC BY) licence to any Author

Accepted Manuscript version arising from this submission.

CONFLICT OF INTEREST

The authors declare that they have no competing interests.

AUTHOR CONTRIBUTIONS

YW participated in data collection, performed analyses, and wrote the first draft. DJO wrote

a part of the first draft and made extensive contributions to the development, drafting, and

editing of the manuscript and supporting material, as well as model explanation. All authors

gave final approval for publication.

Figure legends

Figure 1: A summary of *de novo* mutation rate studies published up to 21st September 2022.
(A) The proportion of studies in different taxonomic groups. (B) The number of studies by
year. MA: mutation accumulation studies, PO: parent-offspring studies.

Figure 2. A phylogenetic meta-analysis of SNM rates for different species identified from the
literature. Mutation rates and 95% CIs were estimated using a PGLMM. For each annotated
clade, we estimated the rate for their most recent common ancestor (clades marked in
colour; estimates in dashed lines and below taxon icons to the right). The vertical black
dashed line indicates the mutation rate of common ancestor of the tree.

Figure 3. Phylogenetic linear regressions of SNM rate on (A) generation time for 116 species,
(B) on genome size for 108 species. The solid black lines represent simple linear regressions;
the coloured lines represent different PGLMM regressions conditional on each of the species
random effects. (C) a regression of pairwise nucleotide diversity (π) on SNM rate (μ), fitted
as a bivariate model in MCMCglmm for 102 species. The solid grey line indicates the model

regression line; the dashed grey line indicates the expected slope for $\theta_\pi = 4N_e\mu$ on the log

scale. All axes are plotted on a log scale.

Figure 4. An analysis of short indel mutation rates for different species. The estimated

mutation rates are presented on a log scale for clarity.

Figure 5. The correlation between indel mutation rates and SNM rates plotted on a log scale

with a 1:1 ratio shown as a dashed line (A). The aggregated log ratios of deletion/insertion in

eukaryotes (B).

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Species	Mutations	Sites (1e9)	μ (Log scale)	μ (1e-9) 95%CI
Drosophila melanogaster	277	1089.71		0.61 [0.31,0.95]
Drosophila simulans	4	5.68		0.62 [0.30,0.99]
Chironomus riparius	35	23.94		1.39 [0.34,2.86]
Bombus terrestris	2	6.45		0.76 [0.11,1.77]
Apis mellifera	8	7.27		0.96 [0.20,2.11]
Daphnia pulex	36	814.08		0.11 [0.03,0.25]
Caenorhabditis elegans	971	1201.97		1.78 [0.69,3.09]
Mus musculus	13	47.58		0.79 [0.21,1.58]
Homo sapiens	14130	15035.70		1.74 [1.16,2.50]
Bos taurus	17	18.81		1.22 [0.30,2.45]
Hanseniaspora valbyensis	148	3361.86		0.05 [0.02,0.09]
Hanseniaspora uvarum	96	2318.84		0.05 [0.02,0.10]
Hanseniaspora osmophila	117	3136.73		0.05 [0.01,0.10]
Saccharomyces cerevisiae	152	11974.26		0.02 [0.01,0.04]
Saccharomyces paradoxus	2	426.44		0.02 [0.01,0.04]
Schizosaccharomyces pombe	454	3903.10		0.13 [0.02,0.29]
Rhodotorula toruloides	83	3772.73		0.05 [0.00,0.12]
Picochlorum costavermella	2	20.75		0.20 [0.01,0.56]
Chlamydomonas reinhardtii	2644	13565.97		0.25 [0.10,0.44]
Ostreococcus tauri	24	327.17		0.13 [0.03,0.30]
Micromonas pusilla	14	86.96		0.26 [0.02,0.66]
Prunus persica	46	30.07		2.49 [0.59,5.32]
Arabidopsis thaliana	429	289.44		3.05 [1.22,5.37]
Oryza sativa	14	6.04		2.75 [0.69,5.81]
Phaeodactylum tricornutum	1	63.29		0.05 [0.00,0.17]
Dictyostelium discoideum	72	1403.50		0.06 [0.01,0.14]

Mya   −1000   0
−28   −26   −24   −22   −20